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# A bound of the $\beta$-mixing coefficient for point processes in terms of their intensity functions 

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#### Abstract

We prove a general inequality on $\beta$-mixing coefficients of point processes depending uniquely on their $n$-th order intensity functions. We apply this inequality in the case of determinantal point processes and show that the rate of decay of the $\beta$-mixing coefficients of a wide class of DPPs is optimal.


Keywords: lower sum transform, determinantal point processes

## 1. Introduction

In asymptotic inference for dependent random variables, it is necessary to quantify the dependence between $\sigma$-algebras. Some of the first measures of dependence that have been introduced are the alpha-mixing coefficients [28] and the beta-mixing coefficients [29]. They have been used to establish moment inequalities, exponential inequalities and central limit theorems for stochastic processes (see [7, 24, 27] for more details about mixing) with various applications in statistics, see for instance [6, 10]. In this paper, we focus on spatial point processes. As detailed below, for these models, alpha-mixing has been widely studied and exploited in the literature, but not beta-mixing in spite of its stronger properties. In a lesser extent, some alternative measures of dependence have also been used for spatial point processes, namely Brillinger mixing [4, 16] (which only applies to stationary point processes but has been established in [16] under suitable conditions on the $\beta$-mixing coefficients) and association [20, 25].

[^0]The main models used in spatial point processes are Gibbs point processes, Cox processes and determinantal point processes, see [23] for a recent review. An $\alpha$-mixing inequality is established for Gibbs point processes in the Dobrushin uniqueness region in [11]. It has been used to show asymptotic normality of maximum likelihood and pseudo-likelihood estimates [18]. Similarly, some inhomogeneous Cox processes like the Neyman-Scott process have also been showed to satisfy $\alpha$-mixing inequalities in [31]. These inequalities are at the core of asymptotic inference results in [8, 26, 31]. Finally, an $\alpha$-mixing inequality has also been showed for determinantal point processes in [25] and used to get the asymptotic normality of a wide class of estimators of these models.

On the other hand, $\beta$-mixing is a stronger property than $\alpha$-mixing. It implies stronger covariance inequalities [27] as well as a coupling theorem known as Berbee's Lemma [3] used in various limit theorems (for example in [2, 30]). Nevertheless, it rarely appears in the literature in comparison to $\alpha$-mixing. This is especially true for point processes where there has been no $\beta$-mixing property established for any of the above examples. Nethertheless, $\beta$-mixing coefficients have still been used several times in random geometry and point process statistics [12, 13, 14, 15]. In particular, it is argued in [14] that the $\beta$ mixing coefficient cannot be replaced by the $\alpha$-mixing coefficient when used to obtain bounds for point process characteristics related with the Palm distribution. Our goal is to establish a general inequality for the $\beta$-mixing coefficients of a point process in terms of its intensity functions.

We begin in Section 2 by recalling the basic definitions and properties of the $\alpha$-mixing and $\beta$-mixing coefficients and we introduce the lower sum transform which is the main technical tool that we use throughout the paper. Then, a general inequality for the $\beta$-mixing coefficients of a point process that depends only on its $n$-th order intensity functions is proved in Section 3. As an example, we deduce a $\beta$-mixing inequality in the special case of determinantal point processes (DPPs) in Section 4 whose rate of decay is optimal for a wide class of DPPs.

## 2. Preliminaries

### 2.1. Intensities of point processes

In this paper, we consider simple point processes on $\left(\mathbb{R}^{d}, \mathcal{B}\left(\mathbb{R}^{d}\right), \mu\right)$ equipped with the euclidean norm $\|$.$\| where d$ is a fixed integer, $\mathcal{B}\left(\mathbb{R}^{d}\right)$ the Borel- $\sigma$-algebra and $\mu$ the Lebesgue measure (more information on spatial point processes can be found in [9, 22]). We denote by $\Omega$ (resp. $\Omega_{F}$ ) the set of locally finite (resp. finite) point configurations in $\mathbb{R}^{d}$. For all functions $f: \Omega_{F} \rightarrow \mathbb{R}, n \in \mathbb{N}$ and $x=\left(x_{1}, \cdots, x_{n}\right) \in\left(\mathbb{R}^{d}\right)^{n}$, we write $f(x)$ for $f\left(\left\{x_{1}, \cdots, x_{n}\right\}\right)$ by an abuse of notation. Finally, we write $|A|$ for the cardinal of a finite set $A$ and $\|f\|_{\infty}$ for the uniform norm of a function $f$.

We begin by recalling that the $n$-th order intensity functions (also called $n$-th order product density) are defined the following way (see [22]).

Definition 2.1. Let $X$ be a simple point process on $\mathbb{R}^{d}$ and $n \geqslant 1$ be an integer. If there exists a non negative function $\rho_{n}:\left(\mathbb{R}^{d}\right)^{n} \rightarrow \mathbb{R}$ such that

$$
\begin{equation*}
\mathbb{E}\left[\sum_{x_{1}, \cdots, x_{n} \in X}^{\neq} f\left(x_{1}, \cdots, x_{n}\right)\right]=\int_{\left(\mathbb{R}^{d}\right)^{n}} f(x) \rho_{n}(x) \mathrm{d} \mu^{n}(x) . \tag{1}
\end{equation*}
$$

for all locally integrable functions $f:\left(\mathbb{R}^{d}\right)^{n} \rightarrow \mathbb{R}$ then $\rho_{n}$ is called the $n$th order intensity function of $X$.

In the rest of the paper, all point processes will be considered to admit bounded $n$-th order intensity function for all $n \geqslant 1$.

### 2.2. Mixing

Consider a probability space $(\mathcal{X}, \mathcal{F}, \mathbb{P})$ and $\mathscr{A}, \mathscr{B}$ two sub $\sigma$-algebras of $\mathcal{F}$. Let $\mathbb{P}_{\mathscr{A}}$ and $\mathbb{P}_{\mathscr{B}}$ be the respective restrictions of $\mathbb{P}$ to $\mathscr{A}$ and $\mathscr{B}$ and define the probability $\mathbb{P}_{\mathscr{A} \otimes \mathscr{B}}$ on the product $\sigma$-algebra by $\mathbb{P}_{\mathscr{A} \otimes \mathscr{B}}(A \times B)=\mathbb{P}(A \cap B)$ for all $A \in \mathscr{A}$ and $B \in \mathscr{B}$. The $\alpha$-mixing and $\beta$-mixing coefficients (also called strong-mixing and absolute regularity coefficients) are defined as the following measures of dependence between $\mathscr{A}$ and $\mathscr{B}$ [24, 27]:

$$
\begin{align*}
& \alpha(\mathscr{A}, \mathscr{B}):=\sup \{|\mathbb{P}(A \cap B)-\mathbb{P}(A) \mathbb{P}(B)|: A \in \mathscr{A}, B \in \mathscr{B}\}  \tag{2}\\
& \beta(\mathscr{A}, \mathscr{B}):=\left\|\mathbb{P}_{\mathscr{A} \otimes \mathscr{B}}-\mathbb{P}_{\mathscr{A}} \otimes \mathbb{P}_{\mathscr{B}}\right\|_{T V}, \tag{3}
\end{align*}
$$

where $\|.\|_{T V}$ is the total variation of a signed norm.
For a given point process $X$ and a bounded set $A \subset \mathbb{R}^{d}$, we denote by $\mu(A):=\int_{A} \mathrm{~d} \mu(x)$ the volume of $A$ and $\mathcal{E}(A)$ the $\sigma$-algebra generated by $X \cap A$. Finally, for all $A, B \subset \mathbb{R}^{d}$, we write $\operatorname{dist}(A, B)$ for the infimum of $\|y-x\|$ where $(x, y) \in A \times B$. The $\beta$-mixing coefficients of the point process $X$ are then defined by

$$
\beta_{p, q}(r):=\sup \{\beta(\mathcal{E}(A), \mathcal{E}(B)): \mu(A) \leqslant p, \mu(B) \leqslant q, \operatorname{dist}(A, B)>r\},
$$

and we say that the point process $X$ is beta-mixing if $\beta_{p, q}(r)$ vanishes when $r \rightarrow+\infty$ for all $p, q>0$. The $\alpha$-mixing coefficients can be defined in a similar way.

Our goal is to prove that under appropriate assumptions over the intensity functions $\rho_{n}$ of $X$ we have a $\beta$-mixing property.

### 2.3. Lower sum transform

The main tool we use throughout this paper is the so-called lower sum operator (see [1]). Notice that when $f$ is a symmetric function the term in the expectation in (1) can be written as $n!\sum_{Z \subset X} f(Z) \mathbb{1}_{|Z|=n}$. This motivates the following definition:

Definition 2.2. Let $f$ be a real function defined over $\Omega_{F}$. The lower sum of $f$ is the linear operator defined by

$$
\begin{equation*}
\hat{f}: X \mapsto \sum_{Z \subset X} f(Z) . \tag{4}
\end{equation*}
$$

As shown in Example 4.19 in [1], this operator admits the following inverse transform.
Proposition 2.3 ([1, Theorem 4.18]). The operator (4) admits an inverse transform $\check{f}$, called the lower difference of $f$, defined by

$$
\begin{equation*}
\check{f}: X \mapsto \sum_{Z \subset X}(-1)^{|X \backslash Z|} f(Z) . \tag{5}
\end{equation*}
$$

These definitions extend to functions over $\Omega_{F}^{2}$ by defining

$$
\hat{f}:\left(X_{1}, X_{2}\right) \mapsto \sum_{Z_{1} \subset X_{1}, Z_{2} \subset X_{2}} f\left(Z_{1}, Z_{2}\right) \text { and } \check{f}:\left(X_{1}, X_{2}\right) \mapsto \sum_{Z_{1} \subset X_{1}, Z_{2} \subset X_{2}}(-1)^{\left|X_{1} \backslash Z_{1}\right|+\left|X_{2} \backslash Z_{2}\right|} f\left(Z_{1}, Z_{2}\right) .
$$

In a similar way, we could also extend these definitions to $\Omega_{F}^{n}$ for any $n$ but we will only need the case $n \leqslant 2$ for the remaining of the paper. These operators allow us to give an explicit expression for the expectation of a functional of a point process with respect to its intensity functions.

Proposition 2.4. If $X$ is an almost surely finite point process such that $\mathbb{E}\left[4^{|X|}\right]<+\infty$, then

$$
\begin{equation*}
\mathbb{E}[f(X)]=\sum_{n=0}^{+\infty} \frac{1}{n!} \int_{\left(\mathbb{R}^{d}\right)^{n}} \check{f}(x) \rho_{n}(x) \mathrm{d} \mu^{n}(y) \tag{6}
\end{equation*}
$$

for all bounded functions $f: \Omega_{F} \rightarrow \mathbb{R}$. Moreover, if $X^{\prime}$ is a point process independent from $X$ satisfying the same assumptions than $X$ and with $n$-th order intensity functions $\rho_{n}^{\prime}$, then

$$
\begin{equation*}
\mathbb{E}\left[f\left(X, X^{\prime}\right)\right]=\sum_{m, n=0}^{+\infty} \frac{1}{m!n!} \int_{\left(\mathbb{R}^{d}\right)^{m+n}} \check{f}(x, y) \rho_{m}(x) \rho_{n}^{\prime}(y) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(x) \tag{7}
\end{equation*}
$$

for all bounded functions $f: \Omega_{F}^{2} \rightarrow \mathbb{R}$.
Proof. Using the bound $|\check{f}(x)| \leqslant\|f\|_{\infty} \operatorname{card}\{Z, Z \subset X\}=\|f\|_{\infty} 2^{|x|}$ we get

$$
\begin{equation*}
\sum_{n \geqslant 0} \mathbb{E}\left[\left|\sum_{Z \subset X}^{|Z|=n}\right| ~ \check{f}(Z) \mid\right] \leqslant \sum_{n \geqslant 0} \mathbb{E}\left[2^{|X|}\binom{|X|}{n}\right]\|f\|_{\infty}=\|f\|_{\infty} \mathbb{E}\left[4^{|X|}\right]<+\infty \tag{8}
\end{equation*}
$$

Since we can write

$$
f(X)=\hat{\tilde{f}}(X)=\sum_{Z \subset X} \check{f}(Z)=\sum_{n \geqslant 0} \sum_{\substack{Z|X\\| Z \mid=n}} \check{f}(Z) \text { a.s. }
$$

then

$$
\mathbb{E}[f(X)]=\sum_{n \geqslant 0} \mathbb{E}\left[\sum_{\substack{Z \subset X \\|Z|=n}} \check{f}(Z)\right]=\sum_{n=0}^{+\infty} \frac{1}{n!} \int_{\left(\mathbb{R}^{d}\right)^{n}} \check{f}(x) \rho_{n}(x) \mathrm{d} \mu^{n}(x)
$$

where the inversion of the first sum and the expectation is a consequence of (8). Similarly, for all functions $f: \Omega_{F}^{2} \rightarrow \mathbb{R}$ we have

$$
\begin{aligned}
\mathbb{E}\left[f\left(X, X^{\prime}\right)\right] & =\mathbb{E}\left[\mathbb{E}\left[f\left(X, X^{\prime}\right) \mid X^{\prime}\right]\right] \\
& =\mathbb{E}\left[\sum_{m=0}^{+\infty} \frac{1}{m!} \int_{\left(\mathbb{R}^{d}\right)^{m}}\left(\sum_{z \subset x}(-1)^{m-|z|} f\left(z, X^{\prime}\right)\right) \rho_{m}(x) \mathrm{d} \mu^{m}(x)\right] \\
& =\sum_{m, n=0}^{+\infty} \frac{1}{m!n!} \int_{\left(\mathbb{R}^{d}\right)^{m+n}} \check{f}(x, y) \rho_{m}(x) \rho_{n}^{\prime}(y) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y),
\end{aligned}
$$

where all inversions of expectation with sum and integrals can be justified in a similar way than (8).

## 3. $\beta$-mixing of point processes with known intensity functions

Our main result is the following inequality showing that if all $\rho_{m}(x) \rho_{n}(y)-\rho_{m+n}(x, y)$ vanish fast enough when $\|y-x\| \rightarrow+\infty$ for all $m, n \in \mathbb{N}$, then the underlying point process is $\beta$-mixing.

Theorem 3.1. Let $X$ be a simple point process on $\left(\mathbb{R}^{d}, \mu\right)$ such that $\mathbb{E}\left[4^{|X \cap A|}\right]<+\infty$ for all bounded subsets $A \subset \mathbb{R}^{d}$. Then, for all $p, q, r \in \mathbb{R}_{+}$,

$$
\begin{equation*}
\left.\beta_{p, q}(r) \underset{\substack{\mu(A)<p, \mu(B)<q \\ \text { dist }(A, B)>\mathrm{r}}}{\leqslant} \sup _{m, n=0}^{+\infty} \frac{2^{n+m-1}}{m!n!} \int_{A^{m} \times B^{n}}\left|\rho_{m}(x) \rho_{n}(y)-\rho_{m+n}(x, y)\right| \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y)\right) \tag{9}
\end{equation*}
$$

Before giving the proof of Theorem 3.1, we need the following lemmas showing the behaviour of $f(X \cap A, X \cap B)$ and $f\left(X \cap A, X^{\prime} \cap B\right)$ under the lower difference operator.

Lemma 3.2. Let $A \subset \mathbb{R}^{d}, f: \Omega_{F} \rightarrow \mathbb{R}$ and define $f_{A}: X \mapsto f(X \cap A)$. Then,

$$
\widetilde{f_{A}}(X)=\check{f}(X) \mathbb{1}_{X \subset A}
$$

Proof. If $X \subset A$ then the result is trivial. Otherwise, there exists $x \in X \backslash A$ and we can write

$$
\begin{aligned}
\widetilde{f_{A}}(X) & =\sum_{Z \subset X, Z \ni x}(-1)^{|X \backslash Z|} f(Z \cap A)+\sum_{Z \subset X, Z \not \supset x}(-1)^{|X \backslash Z|} f(Z \cap A) \\
& =\sum_{Z \subset X, Z \not \supset x}(-1)^{|X \backslash Z|-1} f((Z \cup\{x\}) \cap A)+\sum_{Z \subset X, Z \not \supset x}(-1)^{|X \backslash Z|} f(Z \cap A) \\
& =\sum_{Z \subset X, Z \not \supset x}(-1)^{|X \backslash Z|-1} f(Z \cap A)+\sum_{Z \subset X, Z \not \supset x}(-1)^{|X \backslash Z|} f(Z \cap A)=0 .
\end{aligned}
$$

This result can be extended to multivariate functions: The lower difference of $\left(X_{1}, X_{2}\right) \rightarrow$ $f\left(X_{1} \cap A_{1}, X_{2} \cap A_{2}\right)$ is $\check{f}\left(X_{1}, X_{2}\right) \mathbb{1}_{\left\{X_{1} \subset A_{1}\right\}} \mathbb{1}_{\left\{X_{2} \subset A_{2}\right\}}$.

Lemma 3.3. For all $f: \Omega_{F}^{2} \rightarrow \mathbb{R}$ and $A, B$ disjoint subsets of $\mathbb{R}^{d}$, let us define the function $g: X \mapsto f(X \cap A, X \cap B)$. The lower difference of $g$ satisfies

$$
\check{g}(X)=\check{f}(X \cap A, X \cap B) \mathbb{1}_{\{X \subset A \cup B\}}
$$

Proof. Using Lemma 3.2 we get that $\check{g}(X)=0$ whenever $X$ is not a subset of $A \cup B$. Otherwise, since $A$ and $B$ are disjoint sets,

$$
\check{g}(X)=\sum_{Z \subset X}(-1)^{|X \backslash Z|} f(Z \cap A, Z \cap B)=\sum_{\substack{U \subset X \cap A \\ V \subset X \cap B}}(-1)^{|(X \cap A) \backslash U|+|(X \cap B) \backslash V|} f(U, V)
$$

which, by definition, is equal to $\check{f}(X \cap A, X \cap B)$.
We now have the necessary tools required for the proof of Theorem 3.1.
Proof of Theorem 3.1. Let $p, q>0$ and $A, B$ be two disjoint subsets of $\mathbb{R}^{d}$ such that $\mu(A) \leqslant p$ and $\mu(B) \leqslant q$. Using one of the characterizations of the total variation distance, the $\beta$-mixing coefficient between $\mathcal{E}(A)$ and $\mathcal{E}(B)$ can be expressed as

$$
\beta(\mathcal{E}(A), \mathcal{E}(B))=\frac{1}{2} \sup _{\|f\|_{\infty}=1}\left|E[f(X \cap A, X \cap B)]-E\left[f\left(X \cap A, X^{\prime} \cap B\right)\right]\right|
$$

where $X^{\prime}$ is an independent copy of $X$. Since $X \cap A, X^{\prime} \cap B$ and $X \cap B$ are finite a.s. we can apply (7) which, combined with Lemma 3.2, gives us

$$
\begin{equation*}
\mathbb{E}\left[f\left(X \cap A, X^{\prime} \cap B\right)\right]=\sum_{m, n=0}^{+\infty} \frac{1}{m!n!} \int_{A^{m} \times B^{n}} \check{f}(x, y) \rho_{m}(x) \rho_{n}(y) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y) . \tag{10}
\end{equation*}
$$

On the other hand, by combining (6) with Lemma 3.3, we get

$$
\mathbb{E}[f(X \cap A, X \cap B)]=\sum_{n=0}^{+\infty} \frac{1}{n!} \int_{(A \cup B)^{n}} \check{f}(x \cap A, x \cap B) \rho_{n}(x) \mathrm{d} \mu^{n}(x) .
$$

Since $A$ and $B$ are disjoint sets and by symmetry of $\check{f}(x \cap A, x \cap B) \rho_{n}(x)$, we can simplify the above expression into

$$
\begin{align*}
\mathbb{E}[f(X \cap A, X \cap B)] & =\sum_{n=0}^{+\infty} \sum_{m=0}^{n} \frac{1}{n!}\binom{n}{m} \int_{A^{m} \times B^{n-m}} \check{f}(x, y) \rho_{n}(x, y) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n-m}(y) \\
& =\sum_{m, n=0}^{+\infty} \frac{1}{m!n!} \int_{A^{m} \times B^{n}} \check{f}(x, y) \rho_{m+n}(x, y) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y) . \tag{11}
\end{align*}
$$

Combining (10) and (11) yields that $\left|\mathbb{E}[f(X \cap A, X \cap B)]-\mathbb{E}\left[f\left(X \cap A, X^{\prime} \cap B\right)\right]\right|$ is equal to

$$
\left|\sum_{m, n=0}^{+\infty} \frac{1}{m!n!} \int_{A^{m} \times B^{n}} \check{f}(x, y)\left(\rho_{m}(x) \rho_{n}(y)-\rho_{m+n}(x, y)\right) \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y)\right|
$$

which is bounded by

$$
\sum_{m, n=0}^{+\infty} \frac{2^{n+m}}{m!n!} \int_{A^{m} \times B^{n}}\left|\rho_{m}(x) \rho_{n}(y)-\rho_{m+n}(x, y)\right| \mathrm{d} \mu^{m}(x) \mathrm{d} \mu^{n}(y)
$$

when $\|f\|_{\infty}=1$ and where we used the bound $|\check{f}(x, y)| \leqslant 2^{|x|+|y|}$.

## 4. Application to determinantal point processes

We can directly apply Theorem 3.1 to determinantal point processes. First introduced in [21] under its current form to model fermion systems, DPPs are a broad class of repulsive point processes. We recall that a DPP $X$ with kernel $K:\left(\mathbb{R}^{d}\right)^{2} \rightarrow \mathbb{R}$ is defined by its intensity functions

$$
\rho_{n}\left(x_{1}, \cdots, x_{n}\right)=\operatorname{det}(K[x]) \quad \forall x \in\left(\mathbb{R}^{d}\right)^{n}, \forall n \in \mathbb{N}
$$

where we denote by $K[x]$ the matrix $\left(K\left(x_{i}, x_{j}\right)\right)_{1 \leqslant i, j \leqslant n}$. Existence and uniqueness conditions as well as general information on DPPs can be found in [17]. The application of Theorem 3.1 to DPPs gives us the following $\beta$-mixing condition:

Theorem 4.1. Let $X$ be a DPP with kernel $K$ and define

$$
\omega(r):=\sup _{\|y-x\| \geqslant r}|K(x, y)| .
$$

If $K$ is bounded and $\omega(r) \underset{r \rightarrow+\infty}{\longrightarrow} 0$ then $X$ is $\beta$-mixing. In particular,

$$
\beta_{p, q}(r) \leqslant 2 p q\left(1+2 p\|K\|_{\infty}\right)\left(1+2 q\|K\|_{\infty}\right) e^{2\|K\|_{\infty}(p+q)} \omega(r)^{2} .
$$

Unfortunately, this result does not give a bound for $\beta_{p, \infty}(r)$ which yet is necessary in almost all limit theorems based on beta-mixing.

Proof. Since $\mathbb{E}\left[4^{|X \cap A|}\right]<+\infty$ for all bounded sets $A$ (see [25, Lemma B.5]) then the $\beta$-mixing coefficients of $X$ satisfy (9) by Theorem 3.1. Let $x=\left(x_{1}, \cdots, x_{n}\right)$ and $y=\left(y_{1}, \cdots, y_{m}\right)$, we need to control $|\operatorname{det}(K[x]) \operatorname{det}(K[y])-\operatorname{det}(K[x, y])|$ where $\|x-y\| \geqslant r$. By [25, Lemma B.4], we get the bound

$$
0 \leqslant \operatorname{det}(K[x]) \operatorname{det}(K[y])-\operatorname{det}(K[x, y]) \leqslant n m\|K\|_{\infty}^{n+m-2} \sum_{i=1}^{n} \sum_{j=1}^{m} K\left(x_{i}, y_{j}\right)^{2} .
$$

Injecting this bound into (9) gives us

$$
\begin{align*}
\beta_{p, q}(r) & \leqslant \sum_{n, m=0}^{+\infty} \frac{n^{2} m^{2} 2^{n+m-1} p^{n-1} q^{m-1}\|K\|_{\infty}^{n+m-2}}{n!m!} \sup _{\substack{|A|<p,|B|<q \\
\text { dist } \mathrm{A}, \mathrm{~B})>\mathrm{r}}} \int_{A \times B}|K(x, y)|^{2} \mathrm{~d} \mu(x) \mathrm{d} \mu(y)  \tag{12}\\
& \leqslant \sum_{n, m=0}^{+\infty} \frac{n^{2} m^{2} 2^{n+m-1} p^{n} q^{m}\|K\|_{\infty}^{n+m-2}}{n!m!} \omega(r)^{2} \\
& =2 p q\left(1+2 p\|K\|_{\infty}\right)\left(1+2 q\|K\|_{\infty}\right) e^{2(p+q)\|K\|_{\infty}} \omega(r)^{2} .
\end{align*}
$$

In particular, if $\omega(r)$ vanishes when $r \rightarrow+\infty$ then $X$ is $\beta$-mixing.
In conclusion, the $\beta$-mixing coefficients of DPPs decay at the same rate than $|K(x, y)|^{2}$ does when $x$ and $y$ deviate from each other. For example, kernels of the Ginibre ensemble or the Gaussian unitary ensemble have an exponential decay (see [17]). Moreover, among translation-invariant kernels used in spatial statistics (see [5, 19]), all kernels of the Laguerre-Gaussian family also have an exponential decay while kernels of the WhittleMatérn and Cauchy family satisfy $\omega(r)=o\left(r^{-d}\right)$ and kernels of the Bessel family satisfy $\omega(r)=o\left(r^{-(d+1) / 2}\right)$.

It is also worth noticing that Theorem 3.1 is optimal in the sense that for a wide class of DPPs, the $\beta$-mixing coefficients $\beta_{p, q}(r)$ do not decay faster, when $r$ goes to infinity, than the supremum of $\int_{A \times B}|K(x, y)|^{2} \mathrm{~d} \mu(x) \mathrm{d} \mu(y)$ for all $A, B$ such that $\mu(A) \leqslant p, \mu(B) \leqslant q$ and $\operatorname{dist}(A, B) \geqslant r$ as stated in the following proposition.

Proposition 4.2. Let $X$ be a DPP with a non-negative bounded kernel $K$ such that the eigenvalues of its associated integral operator are all in $[0, M]$ where $M<1$. Then, for all
$p, q, r>0$,

$$
\begin{aligned}
& 2(1-M)^{\frac{(p+q)\|K\|_{\infty}}{M}} \sup _{\substack{\mu(A)<p, \mu(B)<q \\
\text { dist }(\mathrm{A}, \mathrm{~B})>\mathrm{r}}} \int_{A \times B}|K(x, y)|^{2} \mathrm{~d} \mu(x) \mathrm{d} \mu(y) \leqslant \beta_{p, q}(r) \\
& \quad \leqslant 2\left(1+2 p\|K\|_{\infty}\right)\left(1+2 q\|K\|_{\infty}\right) e^{2(p+q)\|K\|_{\infty}} \sup _{\substack{\mu(A)<p, \mu(B)<q \\
\text { dist }(\mathrm{A}, \mathrm{~B})>\mathrm{r}}} \int_{A \times B}|K(x, y)|^{2} \mathrm{~d} \mu(x) \mathrm{d} \mu(y) .
\end{aligned}
$$

Proof. The first inequality is a consequence of the fact that $\beta_{p, q}(r) \geqslant 2 \alpha_{p, q}(r)$ and [25, Proposition 4.3]. The second inequality is equivalent to (12) once the sum has been developed.

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