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## A bound of the $\beta$ -mixing coefficient for point processes in terms of their intensity functions

Arnaud Poinas<sup>1</sup>

IRMAR, Campus de Beaulieu, Bat. 22/23, 263 avenue du Général Leclerc, 35042 Rennes, France

#### Abstract

We prove a general inequality on  $\beta$ -mixing coefficients of point processes depending uniquely on their *n*-th order intensity functions. We apply this inequality in the case of determinantal point processes and show that the rate of decay of the  $\beta$ -mixing coefficients of a wide class of DPPs is optimal.

Keywords: lower sum transform, determinantal point processes

#### 1. Introduction

In asymptotic inference for dependent random variables, it is necessary to quantify the dependence between  $\sigma$ -algebras. Some of the first measures of dependence that have been introduced are the alpha-mixing coefficients [28] and the beta-mixing coefficients [29]. They have been used to establish moment inequalities, exponential inequalities and central limit theorems for stochastic processes (see [7, 24, 27] for more details about mixing) with various applications in statistics, see for instance [6, 10]. In this paper, we focus on spatial point processes. As detailed below, for these models, alpha-mixing has been widely studied and exploited in the literature, but not beta-mixing in spite of its stronger properties. In a lesser extent, some alternative measures of dependence have also been used for spatial point processes, namely Brillinger mixing [4, 16] (which only applies to stationary point processes but has been established in [16] under suitable conditions on the  $\beta$ -mixing coefficients) and association [20, 25].

Email address: arnaud.poinas@univ-rennes1.fr (Arnaud Poinas)

The main models used in spatial point processes are Gibbs point processes, Cox processes and determinantal point processes, see [23] for a recent review. An  $\alpha$ -mixing inequality is established for Gibbs point processes in the Dobrushin uniqueness region in [11]. It has been used to show asymptotic normality of maximum likelihood and pseudo-likelihood estimates [18]. Similarly, some inhomogeneous Cox processes like the Neyman-Scott process have also been showed to satisfy  $\alpha$ -mixing inequalities in [31]. These inequalities are at the core of asymptotic inference results in [8, 26, 31]. Finally, an  $\alpha$ -mixing inequality has also been showed for determinantal point processes in [25] and used to get the asymptotic normality of a wide class of estimators of these models.

On the other hand,  $\beta$ -mixing is a stronger property than  $\alpha$ -mixing. It implies stronger covariance inequalities [27] as well as a coupling theorem known as Berbee's Lemma [3] used in various limit theorems (for example in [2, 30]). Nevertheless, it rarely appears in the literature in comparison to  $\alpha$ -mixing. This is especially true for point processes where there has been no  $\beta$ -mixing property established for any of the above examples. Nethertheless,  $\beta$ -mixing coefficients have still been used several times in random geometry and point process statistics [12, 13, 14, 15]. In particular, it is argued in [14] that the  $\beta$ mixing coefficient cannot be replaced by the  $\alpha$ -mixing coefficient when used to obtain bounds for point process characteristics related with the Palm distribution. Our goal is to establish a general inequality for the  $\beta$ -mixing coefficients of a point process in terms of its intensity functions.

We begin in Section 2 by recalling the basic definitions and properties of the  $\alpha$ -mixing and  $\beta$ -mixing coefficients and we introduce the lower sum transform which is the main technical tool that we use throughout the paper. Then, a general inequality for the  $\beta$ -mixing coefficients of a point process that depends only on its *n*-th order intensity functions is proved in Section 3. As an example, we deduce a  $\beta$ -mixing inequality in the special case of determinantal point processes (DPPs) in Section 4 whose rate of decay is optimal for a wide class of DPPs.

#### 2. Preliminaries

#### 2.1. Intensities of point processes

In this paper, we consider simple point processes on  $(\mathbb{R}^d, \mathcal{B}(\mathbb{R}^d), \mu)$  equipped with the euclidean norm  $\|.\|$  where d is a fixed integer,  $\mathcal{B}(\mathbb{R}^d)$  the Borel- $\sigma$ -algebra and  $\mu$  the Lebesgue measure (more information on spatial point processes can be found in [9, 22]). We denote by  $\Omega$  (resp.  $\Omega_F$ ) the set of locally finite (resp. finite) point configurations in  $\mathbb{R}^d$ . For all functions  $f: \Omega_F \to \mathbb{R}, n \in \mathbb{N}$  and  $x = (x_1, \dots, x_n) \in (\mathbb{R}^d)^n$ , we write f(x) for  $f(\{x_1, \dots, x_n\})$  by an abuse of notation. Finally, we write |A| for the cardinal of a finite set A and  $||f||_{\infty}$  for the uniform norm of a function f.

We begin by recalling that the *n*-th order intensity functions (also called *n*-th order product density) are defined the following way (see [22]).

**Definition 2.1.** Let X be a simple point process on  $\mathbb{R}^d$  and  $n \ge 1$  be an integer. If there exists a non negative function  $\rho_n : (\mathbb{R}^d)^n \to \mathbb{R}$  such that

$$\mathbb{E}\left[\sum_{x_1,\cdots,x_n\in X}^{\neq} f(x_1,\cdots,x_n)\right] = \int_{(\mathbb{R}^d)^n} f(x)\rho_n(x)\mathrm{d}\mu^n(x).$$
 (1)

for all locally integrable functions  $f : (\mathbb{R}^d)^n \to \mathbb{R}$  then  $\rho_n$  is called the nth order intensity function of X.

In the rest of the paper, all point processes will be considered to admit bounded *n*-th order intensity function for all  $n \ge 1$ .

#### 2.2. Mixing

Consider a probability space  $(\mathcal{X}, \mathcal{F}, \mathbb{P})$  and  $\mathscr{A}, \mathscr{B}$  two sub  $\sigma$ -algebras of  $\mathcal{F}$ . Let  $\mathbb{P}_{\mathscr{A}}$  and  $\mathbb{P}_{\mathscr{B}}$  be the respective restrictions of  $\mathbb{P}$  to  $\mathscr{A}$  and  $\mathscr{B}$  and define the probability  $\mathbb{P}_{\mathscr{A}\otimes\mathscr{B}}$  on the product  $\sigma$ -algebra by  $\mathbb{P}_{\mathscr{A}\otimes\mathscr{B}}(A \times B) = \mathbb{P}(A \cap B)$  for all  $A \in \mathscr{A}$  and  $B \in \mathscr{B}$ . The  $\alpha$ -mixing and  $\beta$ -mixing coefficients (also called strong-mixing and absolute regularity coefficients) are defined as the following measures of dependence between  $\mathscr{A}$  and  $\mathscr{B}$  [24, 27]:

$$\alpha(\mathscr{A},\mathscr{B}) := \sup\{|\mathbb{P}(A \cap B) - \mathbb{P}(A)\mathbb{P}(B)| : A \in \mathscr{A}, B \in \mathscr{B}\},\tag{2}$$

$$\beta(\mathscr{A},\mathscr{B}) := \|\mathbb{P}_{\mathscr{A}\otimes\mathscr{B}} - \mathbb{P}_{\mathscr{A}}\otimes\mathbb{P}_{\mathscr{B}}\|_{TV},\tag{3}$$

where  $\|.\|_{TV}$  is the total variation of a signed norm.

For a given point process X and a bounded set  $A \subset \mathbb{R}^d$ , we denote by  $\mu(A) := \int_A d\mu(x)$  the volume of A and  $\mathcal{E}(A)$  the  $\sigma$ -algebra generated by  $X \cap A$ . Finally, for all  $A, B \subset \mathbb{R}^d$ , we write dist(A, B) for the infimum of ||y - x|| where  $(x, y) \in A \times B$ . The  $\beta$ -mixing coefficients of the point process X are then defined by

$$\beta_{p,q}(r) := \sup\{\beta(\mathcal{E}(A), \mathcal{E}(B)) : \mu(A) \leq p, \mu(B) \leq q, \operatorname{dist}(A, B) > r\},\$$

and we say that the point process X is beta-mixing if  $\beta_{p,q}(r)$  vanishes when  $r \to +\infty$  for all p, q > 0. The  $\alpha$ -mixing coefficients can be defined in a similar way.

Our goal is to prove that under appropriate assumptions over the intensity functions  $\rho_n$ of X we have a  $\beta$ -mixing property.

#### 2.3. Lower sum transform

The main tool we use throughout this paper is the so-called lower sum operator (see [1]). Notice that when f is a symmetric function the term in the expectation in (1) can be written as  $n! \sum_{Z \subset X} f(Z) \mathbb{1}_{|Z|=n}$ . This motivates the following definition:

**Definition 2.2.** Let f be a real function defined over  $\Omega_F$ . The lower sum of f is the linear operator defined by

$$\hat{f}: X \mapsto \sum_{Z \subset X} f(Z). \tag{4}$$

As shown in Example 4.19 in [1], this operator admits the following inverse transform.

**Proposition 2.3** ([1, Theorem 4.18]). The operator (4) admits an inverse transform  $\mathring{f}$ , called the lower difference of f, defined by

$$\check{f}: X \mapsto \sum_{Z \subset X} (-1)^{|X \setminus Z|} f(Z).$$
(5)

These definitions extend to functions over  $\Omega_F^2$  by defining

$$\hat{f}: (X_1, X_2) \mapsto \sum_{Z_1 \subset X_1, Z_2 \subset X_2} f(Z_1, Z_2) \text{ and } \check{f}: (X_1, X_2) \mapsto \sum_{Z_1 \subset X_1, Z_2 \subset X_2} (-1)^{|X_1 \setminus Z_1| + |X_2 \setminus Z_2|} f(Z_1, Z_2).$$

In a similar way, we could also extend these definitions to  $\Omega_F^n$  for any n but we will only need the case  $n \leq 2$  for the remaining of the paper. These operators allow us to give an explicit expression for the expectation of a functional of a point process with respect to its intensity functions.

**Proposition 2.4.** If X is an almost surely finite point process such that  $\mathbb{E}[4^{|X|}] < +\infty$ , then

$$\mathbb{E}[f(X)] = \sum_{n=0}^{+\infty} \frac{1}{n!} \int_{(\mathbb{R}^d)^n} \check{f}(x) \rho_n(x) \mathrm{d}\mu^n(y)$$
(6)

for all bounded functions  $f: \Omega_F \to \mathbb{R}$ . Moreover, if X' is a point process independent from X satisfying the same assumptions than X and with n-th order intensity functions  $\rho'_n$ , then

$$\mathbb{E}[f(X,X')] = \sum_{m,n=0}^{+\infty} \frac{1}{m!n!} \int_{(\mathbb{R}^d)^{m+n}} \check{f}(x,y) \rho_m(x) \rho'_n(y) \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(x)$$
(7)

for all bounded functions  $f: \Omega_F^2 \to \mathbb{R}$ .

*Proof.* Using the bound  $|\check{f}(x)| \leq ||f||_{\infty} \operatorname{card} \{Z, Z \subset X\} = ||f||_{\infty} 2^{|x|}$  we get

$$\sum_{n \ge 0} \mathbb{E} \left[ \left| \sum_{\substack{Z \subset X \\ |Z|=n}} \check{f}(Z) \right| \right] \leqslant \sum_{n \ge 0} \mathbb{E} \left[ 2^{|X|} \binom{|X|}{n} \right] \|f\|_{\infty} = \|f\|_{\infty} \mathbb{E} \left[ 4^{|X|} \right] < +\infty.$$
(8)

Since we can write

$$f(X) = \hat{\check{f}}(X) = \sum_{Z \subset X} \check{f}(Z) = \sum_{n \ge 0} \sum_{\substack{Z \subset X \\ |Z|=n}} \check{f}(Z) \text{ a.s.}$$

then

$$\mathbb{E}[f(X)] = \sum_{n \ge 0} \mathbb{E}\left[\sum_{\substack{Z \subset X \\ |Z|=n}} \check{f}(Z)\right] = \sum_{n=0}^{+\infty} \frac{1}{n!} \int_{(\mathbb{R}^d)^n} \check{f}(x) \rho_n(x) \mathrm{d}\mu^n(x)$$

where the inversion of the first sum and the expectation is a consequence of (8). Similarly, for all functions  $f: \Omega_F^2 \to \mathbb{R}$  we have

$$\begin{split} \mathbb{E}[f(X,X')] &= \mathbb{E}[\mathbb{E}[f(X,X')|X']] \\ &= \mathbb{E}\left[\sum_{m=0}^{+\infty} \frac{1}{m!} \int_{(\mathbb{R}^d)^m} \left(\sum_{z \subset x} (-1)^{m-|z|} f(z,X')\right) \rho_m(x) \mathrm{d}\mu^m(x)\right] \\ &= \sum_{m,n=0}^{+\infty} \frac{1}{m!n!} \int_{(\mathbb{R}^d)^{m+n}} \check{f}(x,y) \rho_m(x) \rho'_n(y) \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(y), \end{split}$$

where all inversions of expectation with sum and integrals can be justified in a similar way than (8).

#### 3. $\beta$ -mixing of point processes with known intensity functions

Our main result is the following inequality showing that if all  $\rho_m(x)\rho_n(y) - \rho_{m+n}(x,y)$ vanish fast enough when  $||y - x|| \to +\infty$  for all  $m, n \in \mathbb{N}$ , then the underlying point process is  $\beta$ -mixing.

**Theorem 3.1.** Let X be a simple point process on  $(\mathbb{R}^d, \mu)$  such that  $\mathbb{E}[4^{|X \cap A|}] < +\infty$  for all bounded subsets  $A \subset \mathbb{R}^d$ . Then, for all  $p, q, r \in \mathbb{R}_+$ ,

$$\beta_{p,q}(r) \underset{\substack{\mu(A) < p,\mu(B) < q \\ \text{dist}(A,B) > r}}{\sup} \left( \sum_{m,n=0}^{+\infty} \frac{2^{n+m-1}}{m!n!} \int_{A^m \times B^n} |\rho_m(x)\rho_n(y) - \rho_{m+n}(x,y)| \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(y) \right).$$
(9)

Before giving the proof of Theorem 3.1, we need the following lemmas showing the behaviour of  $f(X \cap A, X \cap B)$  and  $f(X \cap A, X' \cap B)$  under the lower difference operator.

**Lemma 3.2.** Let  $A \subset \mathbb{R}^d$ ,  $f : \Omega_F \to \mathbb{R}$  and define  $f_A : X \mapsto f(X \cap A)$ . Then,

$$\check{f}_A(X) = \check{f}(X) \mathbb{1}_{X \subset A}.$$

*Proof.* If  $X \subset A$  then the result is trivial. Otherwise, there exists  $x \in X \setminus A$  and we can write

$$\begin{split} \widecheck{f_A}(X) &= \sum_{Z \subset X, Z \ni x} (-1)^{|X \setminus Z|} f(Z \cap A) + \sum_{Z \subset X, Z \not\ni x} (-1)^{|X \setminus Z|} f(Z \cap A) \\ &= \sum_{Z \subset X, Z \not\ni x} (-1)^{|X \setminus Z| - 1} f\left((Z \cup \{x\}) \cap A\right) + \sum_{Z \subset X, Z \not\ni x} (-1)^{|X \setminus Z|} f(Z \cap A) \\ &= \sum_{Z \subset X, Z \not\ni x} (-1)^{|X \setminus Z| - 1} f(Z \cap A) + \sum_{Z \subset X, Z \not\ni x} (-1)^{|X \setminus Z|} f(Z \cap A) = 0. \end{split}$$

This result can be extended to multivariate functions: The lower difference of  $(X_1, X_2) \rightarrow f(X_1 \cap A_1, X_2 \cap A_2)$  is  $\check{f}(X_1, X_2) \mathbb{1}_{\{X_1 \subset A_1\}} \mathbb{1}_{\{X_2 \subset A_2\}}$ .

**Lemma 3.3.** For all  $f: \Omega_F^2 \to \mathbb{R}$  and A, B disjoint subsets of  $\mathbb{R}^d$ , let us define the function  $g: X \mapsto f(X \cap A, X \cap B)$ . The lower difference of g satisfies

$$\check{g}(X) = \check{f}(X \cap A, X \cap B) \mathbb{1}_{\{X \subset A \cup B\}}.$$

*Proof.* Using Lemma 3.2 we get that  $\check{g}(X) = 0$  whenever X is not a subset of  $A \cup B$ . Otherwise, since A and B are disjoint sets,

$$\check{g}(X) = \sum_{Z \subset X} (-1)^{|X \setminus Z|} f(Z \cap A, Z \cap B) = \sum_{\substack{U \subset X \cap A \\ V \subset X \cap B}} (-1)^{|(X \cap A) \setminus U| + |(X \cap B) \setminus V|} f(U, V)$$

which, by definition, is equal to  $\check{f}(X \cap A, X \cap B)$ .

We now have the necessary tools required for the proof of Theorem 3.1.

Proof of Theorem 3.1. Let p, q > 0 and A, B be two disjoint subsets of  $\mathbb{R}^d$  such that  $\mu(A) \leq p$ and  $\mu(B) \leq q$ . Using one of the characterizations of the total variation distance, the  $\beta$ -mixing coefficient between  $\mathcal{E}(A)$  and  $\mathcal{E}(B)$  can be expressed as

$$\beta(\mathcal{E}(A), \mathcal{E}(B)) = \frac{1}{2} \sup_{\|f\|_{\infty} = 1} |E[f(X \cap A, X \cap B)] - E[f(X \cap A, X' \cap B)]|$$

where X' is an independent copy of X. Since  $X \cap A$ ,  $X' \cap B$  and  $X \cap B$  are finite a.s. we can apply (7) which, combined with Lemma 3.2, gives us

$$\mathbb{E}[f(X \cap A, X' \cap B)] = \sum_{m,n=0}^{+\infty} \frac{1}{m!n!} \int_{A^m \times B^n} \check{f}(x, y) \rho_m(x) \rho_n(y) \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(y).$$
(10)

On the other hand, by combining (6) with Lemma 3.3, we get

$$\mathbb{E}[f(X \cap A, X \cap B)] = \sum_{n=0}^{+\infty} \frac{1}{n!} \int_{(A \cup B)^n} \check{f}(x \cap A, x \cap B) \rho_n(x) \mathrm{d}\mu^n(x).$$

Since A and B are disjoint sets and by symmetry of  $\check{f}(x \cap A, x \cap B)\rho_n(x)$ , we can simplify the above expression into

$$\mathbb{E}[f(X \cap A, X \cap B)] = \sum_{n=0}^{+\infty} \sum_{m=0}^{n} \frac{1}{n!} \binom{n}{m} \int_{A^m \times B^{n-m}} \check{f}(x, y) \rho_n(x, y) d\mu^m(x) d\mu^{n-m}(y) = \sum_{m,n=0}^{+\infty} \frac{1}{m!n!} \int_{A^m \times B^n} \check{f}(x, y) \rho_{m+n}(x, y) d\mu^m(x) d\mu^n(y).$$
(11)

Combining (10) and (11) yields that  $|\mathbb{E}[f(X \cap A, X \cap B)] - \mathbb{E}[f(X \cap A, X' \cap B)]|$  is equal to

$$\left|\sum_{m,n=0}^{+\infty} \frac{1}{m!n!} \int_{A^m \times B^n} \check{f}(x,y) (\rho_m(x)\rho_n(y) - \rho_{m+n}(x,y)) \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(y)\right|$$

which is bounded by

$$\sum_{m,n=0}^{+\infty} \frac{2^{n+m}}{m!n!} \int_{A^m \times B^n} |\rho_m(x)\rho_n(y) - \rho_{m+n}(x,y)| \mathrm{d}\mu^m(x) \mathrm{d}\mu^n(y)$$

when  $||f||_{\infty} = 1$  and where we used the bound  $|\check{f}(x,y)| \leq 2^{|x|+|y|}$ .

#### 4. Application to determinantal point processes

We can directly apply Theorem 3.1 to determinantal point processes. First introduced in [21] under its current form to model fermion systems, DPPs are a broad class of repulsive point processes. We recall that a DPP X with kernel  $K : (\mathbb{R}^d)^2 \to \mathbb{R}$  is defined by its intensity functions

$$\rho_n(x_1, \cdots, x_n) = \det(K[x]) \quad \forall x \in (\mathbb{R}^d)^n, \ \forall n \in \mathbb{N}$$

where we denote by K[x] the matrix  $(K(x_i, x_j))_{1 \le i,j \le n}$ . Existence and uniqueness conditions as well as general information on DPPs can be found in [17]. The application of Theorem 3.1 to DPPs gives us the following  $\beta$ -mixing condition:

**Theorem 4.1.** Let X be a DPP with kernel K and define

$$\omega(r) := \sup_{\|y-x\| \ge r} |K(x,y)|.$$

If K is bounded and  $\omega(r) \xrightarrow[r \to +\infty]{} 0$  then X is  $\beta$ -mixing. In particular,

$$\beta_{p,q}(r) \leq 2pq(1+2p\|K\|_{\infty})(1+2q\|K\|_{\infty})e^{2\|K\|_{\infty}(p+q)}\omega(r)^{2}.$$

Unfortunately, this result does not give a bound for  $\beta_{p,\infty}(r)$  which yet is necessary in almost all limit theorems based on beta-mixing.

Proof. Since  $\mathbb{E}[4^{|X \cap A|}] < +\infty$  for all bounded sets A (see [25, Lemma B.5]) then the  $\beta$ -mixing coefficients of X satisfy (9) by Theorem 3.1. Let  $x = (x_1, \dots, x_n)$  and  $y = (y_1, \dots, y_m)$ , we need to control  $|\det(K[x]) \det(K[y]) - \det(K[x, y])|$  where  $||x - y|| \ge r$ . By [25, Lemma B.4], we get the bound

$$0 \leq \det(K[x]) \det(K[y]) - \det(K[x,y]) \leq nm \|K\|_{\infty}^{n+m-2} \sum_{i=1}^{n} \sum_{j=1}^{m} K(x_i, y_j)^2.$$

Injecting this bound into (9) gives us

$$\beta_{p,q}(r) \leqslant \sum_{n,m=0}^{+\infty} \frac{n^2 m^2 2^{n+m-1} p^{n-1} q^{m-1} \|K\|_{\infty}^{n+m-2}}{n!m!} \sup_{\substack{|A| < p, |B| < q \\ \operatorname{dist}(A,B) > r}} \int_{A \times B} |K(x,y)|^2 \mathrm{d}\mu(x) \mathrm{d}\mu(y) \quad (12)$$
$$\leqslant \sum_{n,m=0}^{+\infty} \frac{n^2 m^2 2^{n+m-1} p^n q^m \|K\|_{\infty}^{n+m-2}}{n!m!} \omega(r)^2$$
$$= 2pq(1+2p\|K\|_{\infty})(1+2q\|K\|_{\infty})e^{2(p+q)\|K\|_{\infty}}\omega(r)^2.$$

In particular, if  $\omega(r)$  vanishes when  $r \to +\infty$  then X is  $\beta$ -mixing.

In conclusion, the  $\beta$ -mixing coefficients of DPPs decay at the same rate than  $|K(x,y)|^2$ does when x and y deviate from each other. For example, kernels of the Ginibre ensemble or the Gaussian unitary ensemble have an exponential decay (see [17]). Moreover, among translation-invariant kernels used in spatial statistics (see [5, 19]), all kernels of the Laguerre-Gaussian family also have an exponential decay while kernels of the Whittle-Matérn and Cauchy family satisfy  $\omega(r) = o(r^{-d})$  and kernels of the Bessel family satisfy  $\omega(r) = o(r^{-(d+1)/2})$ .

It is also worth noticing that Theorem 3.1 is optimal in the sense that for a wide class of DPPs, the  $\beta$ -mixing coefficients  $\beta_{p,q}(r)$  do not decay faster, when r goes to infinity, than the supremum of  $\int_{A\times B} |K(x,y)|^2 d\mu(x) d\mu(y)$  for all A, B such that  $\mu(A) \leq p, \mu(B) \leq q$  and  $dist(A, B) \geq r$  as stated in the following proposition.

**Proposition 4.2.** Let X be a DPP with a non-negative bounded kernel K such that the eigenvalues of its associated integral operator are all in [0, M] where M < 1. Then, for all

p,q,r > 0,

$$2(1-M)^{\frac{(p+q)\|K\|_{\infty}}{M}} \sup_{\substack{\mu(A) < p, \mu(B) < q \\ \operatorname{dist}(A,B) > r}} \int_{A \times B} |K(x,y)|^2 \mathrm{d}\mu(x) \mathrm{d}\mu(y) \leqslant \beta_{p,q}(r)$$
$$\leqslant 2(1+2p\|K\|_{\infty})(1+2q\|K\|_{\infty})e^{2(p+q)\|K\|_{\infty}} \sup_{\substack{\mu(A) < p, \mu(B) < q \\ \operatorname{dist}(A,B) > r}} \int_{A \times B} |K(x,y)|^2 \mathrm{d}\mu(x) \mathrm{d}\mu(y).$$

*Proof.* The first inequality is a consequence of the fact that  $\beta_{p,q}(r) \ge 2\alpha_{p,q}(r)$  and [25, Proposition 4.3]. The second inequality is equivalent to (12) once the sum has been developed.  $\Box$ 

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