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### An effective streamflow process model for optimal reservoir operation using Stochastic Dual Dynamic

### <sup>3</sup> Programming

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**Abstract.** We present an innovative streamflow process model to be used

<sup>5</sup> for reservoir operational rule design in Stochastic Dual Dynamic Program-

<sup>6</sup> ming (SDDP). Model features, which can be applied independently, are: i)

<sup>7</sup> a non-linear multiplicative process model for the forward phase, which pro-

<sup>8</sup> duces positive streamflow values only, and its linearized version for the back-

<sup>9</sup> ward phase, and ii) a non-uniform time-step, which divides the hydrologi-

 $_{10}$  cal period in time-steps of different length in order to have a process with

<sup>11</sup> approximately constant variance. Model identification is straightforward as

 $_{12}$   $\,$  for additive periodic autoregressive model generally used in SDDP. We ap-

<sup>13</sup> plied this model on the Senegal River for the optimal operation of Manan-

14 tali reservoir, and evaluated the proposed solutions against streamflow pro-

<sup>15</sup> cess model currently used in the water management literature.

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#### 1. Introduction

In reservoir operation, present benefits must be balanced with future, uncertain ones 16 [Soncini-Sessa et al., 2007; Castelletti et al., 2008]. After each release decision, new infor-17 mation becomes available and partially reduces uncertainty. Optimal reservoir operation 18 can be framed as a Multistage Stochastic Programming problem [Birge and Louveaux, 19 1997; Shapiro and Andrzej, 2003, which, for long horizon, is conveniently solved by 20 Stochastic Dynamic Programming [Bellman and Dreyfus, 1966]. SDP, notwithstanding 21 its elegance and potential, is affected by the so-called "curse of dimensionality", limiting 22 its application to systems made of few variables [Stedinger et al., 1984; Trezos and Yeh, 23 1987]. In the literature, some alternatives try to circumvent these limitations, for example 24 by defining an optimal trajectory [Turgeon, 1980] or fixing the release policy family and find the parameters by evolutionary algorithm [Nicklow et al., 2009; Reed et al., 2013]. 26 These solutions, even if advantageous for some aspects, have rarely been tested over large 27 systems, i.e. made of a large number of reservoirs. Stochastic Dual Dynamic Programming [Pereira and Pinto, 1991] (SDDP) is an approximation of SDP that largely attenuates the 29 curse of dimensionality. SDDP, however, requires the optimization problem to be modeled 30 as linear, since problem linearity ensures cost-to-go function convexity. 31

SDDP requires identifying a linear stochastic inflow model which reproduces the streamflow process and its uncertainty. Streamflow process model identification is a critical step in dynamic programming problem setting, sometimes referred to as "curse of modelling" [*Tsitsiklis and Van Roy*, 1996; *Bertsekas and Tsitsiklis*, 1995], to stress that model identification can be problematic, hence a limitation for the methodology. SDDP applications

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generally use a standardized Periodic Autoregressive (PAR) model of lag 1 [*Tilmant et al.*, 37 2008; Tilmant and Kelman, 2007; Tilmant et al., 2007, 2010, 2012; Tilmant and Kinzel-38 bach, 2012; Tilmant et al., 2009; Goor et al., 2010; Arjoon et al., 2014; Margues and 30 Tilmant, 2013; Gjelsvik et al., 2010], also known as Thomas-Fiering model [Loucks, 1992]. 40 The drawback of this additive model is the non-neglegible probability of negative dis-41 charge values. This is to be avoided, because negative discharges have no physical sense. 42 Existing solutions dealing with negative discharges [Stedinger and Taylor, 1982; Pereira 43 et al., 1984; Bezerra et al., 2012] use non-linear transformations that make these models 44 not usable in SDDP. 45

The monthly time-step preserves the process as markov, but it risks to underestimate 46 the system adaptivity to changing conditions. Such a large time-step, in fact, may not 47 take into account the adaptivity at a smaller time-step, and it can be a limitation to the 48 analysis of system response to fast processes, such as flood, resulting in an underestimation 49 of system capacity to react to this type of events. 50

Short-term system adaptation can be taken into account by a time decomposition ap-51 proach [Karamouz et al., 2003]. In time decomposition, long-term policies are refined by 52 optimizations at shorter-term windows, using results from the long-term optimization as 53 boundary conditions. Even if time decomposition increases the accuracy of performance 54 estimation, short-term optimization can use more information than the long-term opti-55 mization supposes, leading to an underestimation of the performance value in the long 56 term planning [Weijs, 2011]. Therefore, aggregating discharges at large time steps is 57 an approximation with negative impact on performance, and time-step length must be a 58 trade-off between i) calculation time and capacity to represent the process as markoy, and 59

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ii) accurate representation of the relevant processes. The former requires a long time-step,
 the latter a short one.

Time-step length depends on both the system characteristics and the hydrological process that we intend to model. For some hydrological systems, variability is not uniform along the year, depending on local climate. For example, where a rainy season is separated from a dry one, the latter has generally less variability than the first, for drought is a relatively slow process compared to flood.

In this paper we present an innovative stochastic streamflow model, to be used within 67 SDDP, that avoids some important limitations of existing models. This paper is struc-68 tured as in the following. In Section 2 we introduce the methodology, from the original 69 optimal control problem, until the SDDP as a way to solve a Multistage Stochastic Pro-70 gramming problem; Section 2.1 presents a procedure to estimate a streamflow model with 71 a multiplicative stochastic component, and its linearized version. This model guarantees 72 a negligible probability of negative discharge values, and its identification is straightfor-73 ward. Section 2.2 exposes a procedure to identify non-uniform time-step lengths, to take 74 advantage of this hydrological variability to have better distributed decision instants. The 75 proposed features are independent from each other, then each of them can be applied sep-76 arately. In Section 3 we test the proposed solutions for modeling the streamflow process 77 on the Senegal river, West Africa. In Section 4 we draw the conclusions. 78

#### 2. Methodology

<sup>79</sup> Consider a water system composed of  $N_{\rm res}$  reservoirs that is operated by  $N_{\rm dec}$  discharge <sup>80</sup> decisions. Discharge decisions are diversions from rivers and releases from reservoirs. A <sup>81</sup> reservoir may have multiple releases (by different structures or for different users). The

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system is influenced by  $N_{\rm scen}$  scenarios, such as future inflows. Equations (1) define the

<sup>83</sup> control problem.

Find 
$$\pi_t, \forall t \in \{1, \dots, H\}$$
: (1a)

$$\max_{\pi_t} \sum_{t=1}^{H} \mathbb{E} \Big[ g_t(\mathbf{v}_t, \mathbf{r}_t, \mathbf{q}_t) \Big]$$
(1b)

Subject to:

$$\mathbf{v}_t = \mathbf{v}_{t-1} + \Delta t \cdot (I \cdot [\mathbf{r}_t, \mathbf{q}_t] - O \cdot [\mathbf{r}_t, \mathbf{q}_t])$$
(1c)

$$\mathbf{c}_t(\mathbf{v}_t, \mathbf{r}_t, \mathbf{q}_t) \le 0 \tag{1d}$$

$$\mathbf{q}_t \sim f_{\mathbf{Q}_t} \tag{1e}$$

In Problem (1), vectors  $\mathbf{v}_t \in \mathbb{R}^{N_{\text{res}}}, \mathbf{r}_t \in \mathbb{R}^{N_{\text{dec}}}, \mathbf{q}_t \in \mathbb{R}^{N_{\text{scen}}}$  represent reservoir volumes, 84 discharge decisions, and scenarios, at instant t for stocks and in the period  $[t - \Delta t, t)$  for 85 flows;  $\pi_t$  is the optimal release rule, that suggests the optimal release decision  $\mathbf{r}_t$  in function 86 of the occurring scenario, i.e. the realization of  $\mathbf{q}_t$ . In Expression (1b),  $g_t(\cdot)$  is a  $\mathbb{R}^N$  to  $\mathbb{R}$ 87 function, representing the system objective at t, where  $N = N_{\text{res}} + N_{\text{dec}} + N_{\text{scen}}$ . Equation 88 (1c) is the continuity equation, represented by the reservoirs mass balance. In Equation 89 (1c),  $\Delta t$  is time-step length, I and O are the input and output matrix, of dimension  $N_{\rm res} \times$ 90  $(N_{\text{dec}} + N_{\text{scen}})$ , associating at each inflow and discharge decision to its reservoir. O(i, j)91 and I(i, j) is 1 if the *i* variable is input or output of reservoir *j*, 0 elsewhere. Scenarios 92  $q_t$ , in Expressions (1e), are either stochastic or deterministic scenarios. Deterministic 93 scenarios are a vector of values, stochastic scenarios are random variables distributed as  $f_{\mathbf{Q}_t}(\mathbf{q}_t)$ . Future inflows to the reservoirs are described as stochastic scenarios, while other

variables, such as evaporation, for which uncertainty can be neglected, are considered as
 deterministic scenarios.

 $\mathbf{c}_t$ , in Inequality (1d), defines other constraints that apply to the system, such as physical constraints, or other legal or environmental requirements treated as constraints. For example, discharge decisions can have a physical upper limit and be limited by water availability.

<sup>102</sup> Problem (1) is to be solved for an optimization horizon of H time-steps, from t = 1, <sup>103</sup> where the initial condition,  $\mathbf{v}_0$ , is given. Decisions and realizations of stochastic variables <sup>104</sup> come in recursive mode, therefore, at each decision step, release can be adjusted thanks to <sup>105</sup> the new information on the occurring scenario. In this case the optimization problem is <sup>106</sup> set as Multistage Stochastic Programming [*Shapiro and Andrzej*, 2003], as in Expression <sup>107</sup> (2):

$$\max_{\mathbf{r}_{1}} g_{1}(\mathbf{v}_{1}, \mathbf{r}_{1}, \mathbf{q}_{1}) + \mathbb{E}_{q_{2,i}} \left[ \max_{\mathbf{r}_{2,i}} g_{2}(\mathbf{v}_{2,i}, \mathbf{r}_{2,i}, \mathbf{q}_{2,i}) + \mathbb{E}_{q_{3,i}} \left[ \dots \right] \right]$$

$$\dots + \mathbb{E}_{q_{H,i}} \left[ \max_{\mathbf{r}_{H,i}} g(\mathbf{v}_{H,i}, \mathbf{r}_{H,i}, \mathbf{q}_{H,i}) \right] \dots \right]$$

$$(2)$$

<sup>108</sup> under conditions given by Equations (1c,??,1d), and with an initial condition  $\mathbf{v}_0$ . In <sup>109</sup> some case, a condition on final time-step  $\mathbf{c}_H(\mathbf{v}_H) \leq \mathbf{b}_H$  may be present, in the form of <sup>110</sup> hard or soft constraint [van Overloop et al., 2008].

<sup>111</sup> By solving Problem (2), the optimization procedure finds an optimal release rule for a <sup>112</sup> future horizon H. In expression (2), the release rule  $\{\pi_t\}_{t=1}^H$  is a decision tree,  $\mathbf{r}_{t,i} \forall t, \forall i$ , <sup>113</sup> made of multiple bifurcations at each time-steps,  $t \in [1 : H]$ , representing the optimal <sup>114</sup> decision strategy adapted to the *i* realisation of the stochastic variable,  $\mathbf{q}_{t,i}$ . The main <sup>115</sup> drawback of Multistage Stochastic Programming is its computational complexity, which <sup>116</sup> increases exponentially with H. Multistage Stochastic Programming can be applied when

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H is small, for example in short term management [Raso et al., 2014]. In same case the 117 problem has been aggregated and reduced to seasonal decisions [Seifi and Hipel, 2001]. 118 For long-term optimization, however, MSP can be considered as a theoretical, rather than 119 practical, method [Mayne et al., 2000]. 120

Stochastic Dynamic Programming (SDP) decomposes the MSP problem in step-by-121 step optimal decision problems. Then, the optimization problem in Expression (2) can 122 be written as Bellman Chain, as in Equation (3). SDP can solve problems with a much 123 longer horizon, because problem complexity increases only linearly with H. 124

$$F_t(\mathbf{v}_t, \mathbf{q}_t) = \max_{\mathbf{r}_t} g_t(\mathbf{v}_t, \mathbf{r}_t, \mathbf{q}_t) + \mathbb{E}_{\mathbf{q}_{t+1}} \Big[ F_{t+1}(\mathbf{v}_{t+1}, \mathbf{q}_{t+1}) \Big]$$
(3)

In Equation (3),  $F_t$  is the cost-to-go function, the average cost for leaving the system in 125 the state  $[\mathbf{v}_t, \mathbf{q}_t]$ , which is the compromise of present and future benefits,  $g_t(\cdot)$  and  $F_{t+1}$ . 126 In SDP, Equation (3) is the release rule  $\pi_t$ , which maps the system state to the optimal 127 release decisions. 128

Equation (3) can be solved backwards, from t = H to the initial time-step. Condition 129 (1e) is substituted by Condition (4). The probability transition from  $\mathbf{q}_{t-1}$  to  $\mathbf{q}_t$ , together 130 with the continuity Equation (1c), makes up the transition equations, which describes the 131 system dynamic from one state to the next one. 132

SDP requires the stochastic transition to be expressed as a Markov process [Rabiner 133 and Juang, 1986], i.e. the probability of each event,  $f(\mathbf{q}_t)$ , depends only on the state at 134 previous instant,  $\mathbf{q}_{t-1}$ , as in Equation (4). 135

$$\mathbf{q}_t \sim f_{Q_t}(\mathbf{q}_t | \mathbf{q}_{t-1}) \tag{4}$$

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When an autoregressive lag 1 process is not sufficient,  $\mathbf{q}_{t-1}$  can be enlarged to contain 136 all the informative variables [Turgeon, 1980]. Then, the process 'memory"  $\mathbf{q}_{t-1}$  includes 137  $[\mathbf{q}_{t-1},\ldots,\mathbf{q}_{t-p},\mathbf{e}_{t-1},\ldots,\mathbf{e}_{t-q}]$ , where  $\mathbf{e}_{t-j}$  is the difference between the observed value and 138 the expected value of the model forecast at  $\mathbf{q}_{t-j}$ . In SDP, this state variables augmentation 139 allows to represent the stochastic process by an Periodic ARMA(p,q) model. However, 140 hydrological processes representation with more than one variable has rarely been applied, 141 and SDP applications are often limited to strategic reservoir operation. In fact, SDP 142 suffers from the so-called "curse of dimensionality,, i.e. complexity increases exponentially 143 with the number of system variables. 144

Stochastic Dual Dynamic Programming (SDDP) [Shapiro, 2011] is an approximation 145 of the original SDP problem. SDDP attenuates the curse of dimensionality, and can be 146 applied to larger systems. SDDP does not require variables discretization, but the time-147 step optimization problem must be linear: transition equations (1c, 4) and objective (1b) 148 must be linear, and equations defining the constraints space (1d) must be affine. SDDP is 149 solved iteratively forward and backward. In the backward stage, at each  $t = H, \ldots, 2$ , the 150 optimization finds the minimum average cost to pass from  $[\mathbf{v}_{t-1}, \mathbf{q}_{t-1}]$  to  $[\mathbf{v}_t, \mathbf{q}_t]$ , adding 151 an extra cut  $l_k(\mathbf{v}_t, \mathbf{q}_t)$  to the approximation of the cost-to-go function  $\mathcal{F}_t(\mathbf{v}_t, \mathbf{q}_t)$ , such 152 that  $\mathcal{F}_t(\cdot) := \max\{\mathcal{F}_t(\cdot), l_k(\cdot)\}$ . In the forward stage, the approximate problem is solved 153 from t = 1 to H to find the optimal trajectories to be used in the next backward phase. 154 By successive iterations,  $\mathcal{F}_t$  converges to the real cost-to-go function  $F_t$ , as demonstrated, 155 under mild conditions, by *Philpott and Guan* [2008] and *Linowsky and Philpott* [2005]. 156

#### 2.1. Multiplicative error model identification

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In this section we propose a multiplicative error model to deal with the problem of 157 generating negative discharges. 158

In SDDP, the stochastic hydrological model must be linear, but no conditions are im-159 posed on distribution  $f_t(\mathbf{q}_t)$ . There is not specific reason to prefer separate additive terms 160 for the predictive and the uncertain parts of the model [Koutsoyiannis, 2009]. In the 161 following, we show a procedure to estimate a linear model in  $\mathbf{q}_t$ , with a multiplicative 162 stochastic component. 163

We start from a multivariate signal of observed discharge, which we want to reproduce, 164  $\mathbf{q}_t$ , made of T samples for  $N_y$  years of data, where T is the number of time-steps per year. 165

The original signal is transformed into  $\mathbf{y}_t$  according to Equation (5). 166

$$\mathbf{y}_t = \log(\mathbf{q}_t) - \log(\overline{\mathbf{q}}_{\tau}) \tag{5}$$

In Equation (5),  $\tau = 1, \ldots, T$  is the periodic time index. Equation (6) defines  $\overline{\mathbf{q}}_{\tau}$ , i.e. 167 the periodic geometric average of  $\mathbf{q}_t$  for all  $\mathbf{t} = \tau$ , where  $\mathbf{t} = (t-1) \mod T+1$ , and mod 168 is the modulus (or remainder) operator. 169

$$\overline{\mathbf{q}}_{\tau} = \left(\prod_{\mathbf{t}}^{N_y} \mathbf{q}_t\right)^{1/N_y} = \exp\left(\frac{1}{N_y} \sum_{\mathbf{t}}^{N_y} \log \mathbf{q}_t\right), \ \forall \mathbf{t} = \tau \tag{6}$$

Logarithm smooths extremes and long tails, making model identification easier. In 170 fact, logarithm is a particular case of the Box-Cox transformation, suggested in Box et al. 171 [1970] to deal with non-normal residuals. We use the  $\mathbf{y}_t$  signal to identify an ARMA 172 model on the  $N_{\text{stoch}}$  dimensional signal. Contemporaneous ARMA (CARMA) [Camacho 173 et al., 1987] are an effective sub-class of multivariate ARMA model that proved effective 174 in hydrological applications. A CARMA model is identified as a set of  $N_{\rm stoch}$  univariate 175

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ARMA models, with correlated residuals. For simplicity of explanation and notation, we describe the model identification procedure for an univariate ARMA model, recalling that possible correlation among sub-catchments can be included in the residuals covariance matrix [*Salas et al.*, 1985].

In hydrological systems, parameters are generally time-variant [Salas, 1980; Hipel and McLeod, 1994], because of climatic effect and because different hydrological processes are dominant at different periods in the years. Time-variant parameters can effectively dealt with Periodic ARMA model (PARMA). Equation (7) defines the  $PARMA_{\tau}(p,q)$  model.

$$y_t = \sum_{i=1}^p \phi_{\tau,i} y_{t-i} + \sum_{j=1}^q \psi_{\tau,i} \varepsilon_{t-i} + \varepsilon_t \tag{7}$$

In Equation (7),  $\varepsilon_t$  is the stochastic process, extracted from  $\mathcal{N}(0, \sigma_\tau^2)$ , independent on previous  $\varepsilon_{t-i}$ , where  $\varepsilon_{t-i} = y_{t-i} - \hat{y}_{t-i}$ ;  $y_t$  is observed value, and  $\hat{y}_t$  is expected value of Model (7), i.e. when  $\varepsilon_t = 0$ . Identifying a  $PARMA_\tau(p,q)$  model is defining parameters  $\phi_{\tau,i}, \psi_{\tau,i}, \sigma_\tau$  of Equation (7)  $\forall i, \forall j, \forall \tau$ .

Equation (7) can be written as transition from  $q_{t-i}$  to  $q_t$ . By inversion of Equation (5) and some rearrangement, one can obtain Equation (8).

$$q_t = \alpha_\tau \cdot \prod_{i=1}^p q_{t-i}{}^{\phi_{\tau,i}} \cdot \prod_{j=1}^q \xi_{t-i}{}^{\psi_{\tau,j}} \cdot \xi_t \tag{8}$$

where  $\xi_t \sim \ln \mathcal{N}(0, \sigma_{\tau}), \ \alpha_{\tau} = \overline{q}_{\tau} / \prod_{i=1}^p \overline{q}_{\tau-i}^{\phi_{\tau,i}}$ , with  $\overline{q}_{\tau}$  defined as in Equation (6), and  $\xi_{t-j} = \frac{q_{t-j}}{\hat{q}_{t-j}}.$ 

If in model (7) residuals are normal and additive, in model (8) residuals become lognormal and multiplicative. For positive initial condition,  $q_0 > 0$ , multiplicative random

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process ensures non-negative values of inflow process, offering a better representation of 194 the hydrological process. 195

In Model (8), dependencies on  $q_{t-i}$  are non-linear. Model (8) can be directly employed 196 in the forward phase of SDDP, also using a less parsimonious model than in backwards, 197 which can produce, in some cases, considerably better results [Bartolini and Salas, 1993]. 198 SDDP forward phase requires in fact a streamflow time series, regardless of the method 199 to generate it. In backward phase, instead, SDDP must solve a linear optimization, 200 therefore transition from  $q_{t-1}$  to  $q_t$  must be linear. Model (8), to be applied in backward 201 phase of SDDP, must be written as in Equation (9), which is Model (8) linearized by 202 taylor expansion on the argument median, i.e.  $q_{t-i} = \overline{q}_{\tau-i}$  and  $\xi_{\tau-j} = 1, \forall \tau, \forall i, \forall j$ . 203

$$q_t = \left[\sum_{i=1}^p \rho_{\tau,i} q_{t-i} + \sum_{j=1}^q \omega_{\tau,i} \xi_{t-i} + \kappa_\tau\right] \cdot \xi_\tau \tag{9}$$

where parameters are defined in Equations (10) and derived in Appendix A.

$$\rho_{\tau,i} = \phi_{\tau,i} \cdot \frac{\overline{q}_{\tau}}{\overline{q}_{\tau-i}} \tag{10a}$$

$$\omega_{\tau,i} = \psi_{\tau,i} \cdot \overline{q}_{\tau} \tag{10b}$$

$$\kappa_{\tau} = \overline{q}_{\tau} \cdot \left( 1 - \sum_{i=1}^{p} \phi_{\tau,i} - \sum_{j=1}^{q} \psi_{\tau,i} \right)$$
(10c)

Linearisation introduces an approximation error that must be quantified. Equation (11)204 defines  $e_t$ , the error due to linearisation. 205

$$e_t = q_t^{\rm lin} - q_t^{\rm nl} \tag{11}$$

where  $q_t^{\text{lin}}$  and  $q_t^{\text{nl}}$  are the output of Model (9) and Model (8) for  $\xi_t = 1$ . Considering, 206 for simplicity of notation, an univariate Multiplicative Periodic Autoregressive Model, 207

<sup>208</sup>  $PAR\tau(1)$ , then  $e_t$  can be written as function of  $q_{t-1}$  only. Knowing the distribution <sup>209</sup>  $f_{\tau}(q_{t-1})$  we can estimate that of  $e_t$ . Specifically, we are interested in the average and <sup>210</sup> extreme quantiles of  $e_t$ , defined as in Equations (12).

$$\mathbb{E}(e_{\tau}) = \int_0^{+\infty} e_t(q_{t-1}) \cdot f_{\tau}(q_{t-1}) \cdot dq_{t-1}$$
(12a)

$$q_p(e_\tau) = e_t \left( F_{q_{\tau-1}}^{-1}(p) \right)$$
 (12b)

<sup>211</sup> Where  $q_p(e_{\tau})$ , is the *p* quantile of  $e_{\tau}$ , and  $F_{q_{\tau-1}}(\cdot)$  is the Cumulative Density Function <sup>212</sup> of  $q_{\tau-1}$ . Distribution of  $q_{\tau-1}$  can be obtained from  $f(y_{\tau-1})$  through Equation (5). From <sup>213</sup> Vecchia [1985] Bartolini et al. [1988], we know that  $f(y_{\tau-1})$  is normal with known average <sup>214</sup> and variance. Therefore  $f(q_{\tau-1})$  is lognormal, with parameters derived from  $f(y_{\tau-1})$ <sup>215</sup> average and standard deviation.

 $\mathbb{E}(e_{\tau})$ , from Equation (12a), can be used to correct results of linear Model (9) by shifting the output value in order to have zero bias. In this case,  $-\mathbb{E}(e_{\tau})$  is to be added in Equation (10c).

#### 2.2. Non-uniform time-step length

In this section we propose a non-uniform discharge aggregation, which modulates the time-step length to have a fine discharge representation only when needed.

Hydrological models in SDDP have so far always used a fixed time-step and periodically variable parameters. Predictive uncertainty, changing along the period, is included in the process model by considering heteroscedastic residuals, i.e. residuals with different variability. We consider here, instead, a variable time-step that divides the hydrological

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period in time-steps of different length in order to maintain an approximately homoge-225 neous variance. 226

In the following, we describe the proposed procedure to select the non-uniform time-step 227 length,  $\Delta k(\tau)$ , to aggregate data from the finest time-step,  $\tau_d = 1, \ldots, T_d$  to the desired 228 level of aggregation,  $\tau = 1, \ldots, T$ , such that the aggregated time-step is  $\Delta \tau = \Delta \tau_d \cdot \Delta k$ . 229 Starting from the data at the finest time aggregation of period  $T_d$  and time index 230  $\tau_d = 1, \ldots, T_d$ , we identify a  $\text{PAR}_{T_d}(1)$  model on  $q_t$ , having parameters  $\phi_{\tau_d}, \sigma_{\tau_d}^2, \forall \tau_d \in$ 231  $\{1, \ldots, T_d\}$ , defined as in Equation (7). Even if such simple model may be not accurate 232 enough for prediction purposes, it is generally sufficient to catch the dominant dynamics. 233 Time-steps of homogeneous variability,  $\Delta k(\tau)$ , must be such that  $VAR(q_{\tau_0+\Delta k(\tau)}|q_{\tau_0})$  is 234 approximately homogeneous for all  $\tau$ , and  $\sum \Delta k(\tau) = T_d$ . Equation (13) defines variance 235 of  $q_{\tau_0+\Delta k(\tau)}$  conditional to  $q_{\tau_0}$  as function of time-step residual variance,  $\sigma_{\tau_d}^2$ . 236

$$VAR(q_{\tau_0+\Delta k(\tau)}|q_{\tau_0}) = \sum_{\tau_d=\tau_0+1}^{\tau_0+\Delta k(\tau)} \prod_{i=\tau_d+1}^{\tau_0+\Delta k(\tau)} \phi_i^2 \cdot \sigma_{\tau_d}^2$$
(13)

where, by convention,  $\phi_i^2 = 1$  if  $i > \tau_0 + \Delta k(\tau)$ . 237

In Equation (13),  $\phi_{\tau_d}$  is generally close to one, especially for  $\tau_d$  where variability is 238 small, and time-steps  $\Delta k$  can be larger. Considering  $\prod \phi_{\tau_d}^2$  in Equation (13) as equal to 239 one,  $VAR(q_{\tau_0+\Delta k_i}|q_{\tau_0})$  can be written as proportional to the sum of variances only, which 240 we take as indicator of variability. This allows us to define a cumulated variability that 241 depends on  $\tau_d$  only. 242

The residual variance is used to define a Cumulative Variability in function of  $\tau$ , as in 243 Equation (14). 244

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$$CV(\tau_d) = \frac{\sum_{t=1}^{\tau_d} \underline{\sigma}_{\tau_d}^2}{\sum_{t=1}^{T_d} \underline{\sigma}_{\tau_d}^2}$$
(14)

In Equation (14), the numerator is the cumulative variance until  $\tau_d$ , the denominator is the cumulated variability for the entire period,  $T_d$ , to standardize CV between zero and one. The non-uniform time-step  $\Delta k(\tau)$  is then chosen by splitting the hydrological period in T time-steps having approximately homogenous variability, as in Equation (15).

$$\Delta k(\tau) = \left\lfloor CV^{-1}\left(\frac{\tau}{T}\right) - CV^{-1}\left(\frac{\tau-1}{T}\right) \right\rceil; \ \forall \tau = \{1, \cdots, T\}$$
(15)

where operator  $\lfloor \cdot \rfloor$  returns the nearest integer. By convention,  $CV^{-1}(0) = 0$ .

The discharge signal is aggregated using the variable time-step, and a model is identified on the aggregated signal according to procedure in Section 2.1, or others.

#### 3. Application to the Senegal river

The Senegal River, West Africa, is a 1790 km long river. Its drainage basin extension is 270.000  $km^2$ , over Guinea, Mali, Senegal and Mauritania. The river inflow is extremely variable, following the tropical raining seasonality with a marked difference between the dry season, in January-June, and the raining one, in July-October, when most of the water falls in the upper part of the basin [*Albergel et al.*, 1997]. Figure 2 displays the discharge at Soukoutali, inflow to Manantali for 64 years.

<sup>258</sup> Manantali is an annual reservoir for hydropower, located in Mali and controlling about <sup>259</sup> 50% of the total water flow of the Senegal River. Manantali was completed in 1987 and <sup>260</sup> started to produce electricity in 2002. Its benefits are shared among Mali, Mauritania, <sup>261</sup> and Senegal. These three countries participate with Guinea to the Organization for the

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Valorization of Senegal River (Organisation pour la Mise en Valeur du fleuve Senegal, 262 OMVS). Manantali reservoir volume is  $12 \times 10^9 m^3$ , its installed capacity is 205 MW, the 263 average inflow is 270  $m^3/s$ , the average residence time is about one year. 264

The OMVS have gradually embraced the Integrated Water Resources Management 265 paradigm, in which water allocation decisions are based on economic, social, technical 266 and political factors, in accordance with stakeholders' interests. This led the OMVS 267 to prepare a reservoir management optimization program [Fraval et al., 2002]. Manantali 268 operation was originally designed to satisfy different rival uses: energy production and low-269 flow augmentation on one hand; flood support on the other. Operating the reservoir for 270 hydroelectric production reduces the annual streamflow variability, with negative effects 271 on the ecological equilibrium and on some traditional activities in the valley, but with 272 positive effects on water availability for irrigation and navigation, which will become a 273 more important objective in the near future [Bader et al., 2003]. This analysis, however, 274 focuses on methodological aspects, so we consider energy production objective only. 275

In the following, at section 3.1, we define the non-uniform time-step, then we identify 276 a streamflow process model as defined in Equations (8) and (9), and a classic Thomas-277 Fiering Model, for comparison. At section 3.2, we test the added value of non-uniform 278 time-step in terms of reservoir operation. 279

#### 3.1. Streamflow process model identification

The hydrological process on the Senegal river is characterised by a strong periodical 280 component. Following the procedure described in section 2.2, we select  $\Delta k(\tau)$  using 281 T = 12 time-steps, for comparison with monthly time-step. Figure 3 shows the  $CV(\tau_d)$ 282 function, defined in Equation (14), and the  $\Delta k(\tau)$  for T = 12, for the entire periodic year, 283

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on plot (a), and a zoom on the period of high variability, on plot (b), for  $\tau_d$  between 135 (May 15) and 285 (October 12).

The non-uniform time-step defined by  $\Delta k(\tau)$  is used to aggregate the daily inflow signal. The same inflow is also aggregated at the monthly time-step to be compared to the nonuniform aggregation. Figure 4 shows the aggregated observed inflow with non-uniform time-step and monthly time-step.

Non-uniform aggregation allocates only two time-steps for the entire recession curve, going from October to May, using the ten remaining time-steps for the raising part of the hydrography, against the seven and five time-steps used by monthly aggregation. A finer time-step during the raising part of the hydrograph should allow to better adapt to the incoming information on the inflow value.

Figure 5 shows the autocorrelation lag 1 on the de-trended logarithmic of discharge signal, equivalent to  $\phi_{\tau}$  of Model (7) for a PAR(1), being  $\phi_{\tau} = \frac{COV(y_t, y_{t-1})}{VAR(y_{t-1})}$  [Box et al., 1970]. In monthly aggregation,  $\phi_{\tau}$  stays close to 1 for different months during the dry period, reaching 0.5 at  $\tau = 8$ . The non-uniform aggregation autocorrelation is more regular: most of the cases lay between 0.6 and 0.8. This results suggest that the nonuniform aggregation can provide a more effective distribution of decision instants. This is further investigated in section 3.2.

From the non-uniform aggregated inflow signals, following the procedure described in 2.1, we identify a multiplicative model as in Equations (9) and (8) and an additive Thomas-Fiering model. We refer to these models as multiplicative and additive model. A PAR(1)represents the process sufficiently well. The residuals autocorrelation is approximately zero for both the multiplicative and the additive model, being within the 95% confidence

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<sup>307</sup> band, for most of the lags larger than zero. This confirms the validity of the non-uniform <sup>308</sup> aggregation, for it preserves the process as markov.

We test whether the residual sampling distribution adhere to the prescribed one, log-309 normal for the multiplicative case, normal for the additive. Both residuals must be stan-310 dardised by the time variant variance,  $\sigma_{\tau}$ . For the multiplicative residuals, we test their 311 normality on  $\log(\Xi_t)$ . Figure 6 presents the normality plot for both the logarithm of resid-312 uals for the multiplicative model and the residuals for the additive model, showing that 313 the multiplicative model follows the prescribed distribution, whereas the additive model 314 suffers from a fat tail on higher values. A Kolmogorov-Smirnov test [Dekking, 2005] 315 on the standardized logarithms of residuals does not reject the hypothesis of a normal 316 distribution, with a p-value of 0.07. The same test on the additive model standardized 317 residual gives a p-value of  $10^{-6}$ , leading to rejecting the assumption of residuals being nor-318 mally distributed with sufficient confidence. The positive results on the  $(\xi_t)$  distribution 319 correctness gives us further confidence in the validity of the multiplicative model. 320

Figure 7 shows the observed discharge data and their match with the 95% confidence 321 bands for both the multiplicative and the additive model. The confidence bands are 322  $\bar{q}_{\tau} \cdot \exp\left(\pm 2 \cdot VAR(y_t)\right)$  for the multiplicative model (bold continuous lines) and  $\mathbb{E}(q_t) \pm 1$ 323  $2 \cdot VAR(q_t)$  for the additive model (bold dashed lines). The signal variance  $VAR(q_t)$ 324 or  $VAR(y_t)$ , is derived from  $\{\sigma_{\tau}\}_{\tau=1}^T$  as described in *Bartolini et al.* [1988].From Figure 325 7 we see how the additive model has a non-negligible probability of producing negative 326 inflow. Moreover, lower confidence band for the additive model is, for some time-steps, 327 much lower than the lowest observed discharge. whereas the multiplicative model follows 328 closely the observed signal variance for the entire period. 329

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We quantify the approximation due to linearization using Equations (12) for error av-330 erage, and 5% and 95% quantile of  $q_{\tau-1}$ , examining both the average for all  $\tau$  and the 331 largest value in  $\tau$ . The average error is 0.6  $m^3/s$ , or 0.3% of median discharge, which 332 we consider a relatively small value. Even for the largest error average, at  $\tau = 7$ , is 23 333  $m^3/s$ , which is less than 3% of median discharge. Error at 5% quantile is, on average, 4% 334 of median discharge, with a peak of 7% at  $\tau = 4$  (15  $m^3/s$  in absolute value). Error at 335 95% quantile is 11% of median discharge, with a maximum absolute value of 88  $m^3/s$  at 336  $\tau = 7$ , which is 11% of median discharge. As a summary, we can state that error due to 337 linearization is adequatly small; its average is generally negligible, growing to about 5%338 and 10% of median discharge at 5% and 95% quantiles of  $q_{\tau-1}$ . 339

Figure 8 shows the detail of the error due to linearization for  $\tau = 7$ . Plot (a) compares output of nonlinear and linearized multiplicative model, i.e. model (8) and model (9) for  $\xi_t = 1$ , in function of  $q_{t-1}$ . The linear model is tangent to the nonlinear one at the linearisation point, i.e. at  $\bar{q}_{\tau-1}$ . Going further from the median the models diverge, even if the difference stays small. Plot (b) presents the error magnitude in function of  $q_{t-1}$  next to its probability of occurrence,  $f(q_{\tau-1})$ .

On a multiplicative  $PAR_{\tau}(1)$ ,  $\phi_{\tau} < 1$  for most of  $\tau$ , because stationarity condition requires that  $\prod_{\tau=1}^{T} \phi_{\tau} < 1$ . This implies that model (8) is a concave function in  $q_{t-1}$  and  $e_t > 0, \forall q_{t-1}$ , i.e. Model (9) systematically overestimate its nonlinear version. A large number of time-steps per period results in a shorter time-steps length. As a consequence,  $\phi_{\tau}$  values will be closer to one, for  $\phi_{\tau}$  is the autocorrelation lag 1 in a AR model [*Box et al.*, 1970], resulting in a smaller error due to linearization.

#### 3.2. Effects of non-uniform aggregation on reservoir operation

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We compare the system performances and behaviour of reservoir operation between
 monthly and non-uniform discharge aggregation to assess the advantage of the latter.
 The system model is made of a reservoir and a hydrological component, as in Figure 9.
 Streamflow at Soukoutali, output of the hydrological model, is the input to the reservoir.
 The objective is energy production. Each element of the optimization problem is detailed
 in the following.

The reservoir is modelled as in Equation (1c). The reservoir input is the discharge from the hydrological component  $q_t$ . Reservoir outputs are: discharge through turbines  $r_{\text{turb},t}$ , and discharge through spillways,  $r_{\text{spill},t}$ . Evaporation is not considered in this analysis. Inequalities (16) are constraints on discharge decisions and reservoir volume.

$$v_{\min} \le v_t \le v_{\max} \tag{16a}$$

$$0 \le r_{\text{turb},t} \le r_{\text{turb,max}} \tag{16b}$$

$$r_{\text{spil},t} \ge m_{\text{safety}} \cdot (v_t - v_{\text{safety}})$$
 (16c)

Inequalities (16a) and (16b) are physical constraints, derived from the system characteristics. Inequality (16c) is a legal condition that forces to draw down the reservoir when its volume exceeds the safety threshold,  $v_{\text{safety}}$ . Constraint on release through spillages is large enough for never being active during the simulation period, as verified a posteriori, and therefore it is not included. Inequalities (16b) and (16c) are implemented as hard constraint, inequality (16a) is implemented as soft constraint to avoid non-feasibility.

The system objective is the yearly average energy production  $J_E = 1/N_{years} \sum_{t=1}^{T \times N_{years}} E_t$ , composed of the sum of the daily energy production,  $E_t$ , as defined in Equation (17), for the entire simulation horizon  $T \times N_{years}$ , where  $N_{years} = 43$ , from 1970 to 2012.

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$$E_t = \eta \cdot \Delta k(\tau) \cdot \Delta h_t \cdot r_{\text{turb},t} \tag{17}$$

In Equation (17),  $\Delta h_t$  is the hydraulic head [m],  $r_{\text{turb},t}$  is discharge trough turbines  $[m^3/s]$ .  $\eta$  is a multiplicative factor, such that  $\eta = \rho \cdot g \cdot \hat{\eta}(r_t, \Delta h) \cdot 24 \cdot 10^{-6}$ , where  $\rho$  is the water density, 1000 [kg/m3], g is the gravity acceleration, 9.8  $[m/s^2]$ ,  $\hat{\eta}$  is the efficiency coefficient, considered equal to 0.9, 24 and  $10^{-6}$  are unit transformation coefficients, [h/d]and [MW/W].  $E_t$  is expressed in [MWh].

Equation (17), to be employed in linear optimization within SDDP, is approximated by expressing it as linear function of  $r_t$  and  $v_t$ , linearized at an operational point, under the hypothesis of cylindric reservoir. Equation (18) defines the operational time-step objective indicator used in SDDP, being the weighted sum of releases and volume.

$$E_t^{\text{op}} = E_{0,t} +$$

$$+\eta \cdot \Delta k(\tau) \cdot \left\{ + \left[ \frac{r_{\text{turb},0}}{A_0} \right] \cdot v_t + \left[ h_0 - h_0^v - m_v (1 + r_{\text{turb},0}) \right] \cdot r_{\text{turb},t} + \left[ -m_v \right] \cdot r_{\text{spill},t} \right\}$$

$$(18)$$

In Equation (18),  $E_{0,t} = \eta \cdot \Delta k(\tau) \cdot \left\{ \left[ h_0 - h_0^v \right] \cdot r_{\text{turb},0} - \left[ \frac{r_{\text{turb},0}}{A_0} \right] \cdot v_0 - \left[ h_0 - h_0^v - m_v \cdot r_{\text{spill},0} \right] \right\}$ . Parameters of Equation (18) and their derivation are described in Appendix B.

Equation (17) is linearized at a normal operational point, that is the reservoir state at which the reservoir is mostly operated, either historically observed or deduced from system characteristics. We infer the operational point from the reservoir characteristics,

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# considering $v_0 = v_{\text{safety}}$ , i.e. the safety limit, $r_{turb,0} = \mathbb{E}(q_t)$ , i.e. the average inflow to the reservoir, and $r_{spill,0} = 0 m^3/s$ , which considers no release through the spillages.

The optimization is performed using 25 extractions for the forward phase and 25 for the 388 backward one until convergence, attained at an accuracy level of  $2 \times 10^7$ . This accuracy 389 level lays within the  $\pm 2\sigma$  of forward simulation results. We can consider with sufficient 390 confidence that the algorithm has converged to the optimum. 391

Performance results show a moderate improvement for non-uniform aggregation. Op-392 timal solution gives  $J_E$  equal to 930 GWh/year for the monthly aggregation, and 945 393 GWh/year in the non-uniform case, equivalent to an improvement of one week of average 394 energy production. 395

We analyse the reservoir operation behaviour. Figure 10 shows the reservoir volume and 396 the release trough turbines for the monthly and the non-uniform aggregation in response to 397 the 2005 inflow scenario. Inflow peak is larger than maximum discharge trough turbines; 398 therefore, to avoid spillage, the reservoir must be drawn down before the high flow period, 399 in order to create a buffer that stores part of the incoming water. SDDP optimal operating 400 rules are the results of an optimal compromise between the objectives of keeping a high 401 water level and avoiding spillages. 402

Figure 10 plot (b) shows how the reservoir operation using non-uniform steps adjusts 403 decisions at higher frequency during high uncertainty periods, adapting more rapidly to 404 the new observed discharge. Thanks to this rapid adaptation during the high discharge 405 period, reservoir operation using non-uniform aggregation can draw down the reservoir 406 less, as shown in Figure 10 plot (a). 407

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#### 4. Conclusions and Discussion

SDDP, to be employed for reservoir operational rules design, requires the identification 408 of a linear streamflow process model. Presently, models from the literature use almost 409 always a periodic autoregressive model with monthly time-steps. In this study we proposed 410 an innovative streamflow process model to be used in SDDP. Model features are i) a log-411 normal multiplicative stochastic component, which guarantees positive discharge values, 412 and ii) non-uniform time-steps, which makes the process approximately homoscedastic 413 i.e. having constant variability. The multiplicative non-linear model can be employed 414 in the SDDP forward phase directly, whereas a linearized version must be used in the 415 backward phase. We showed how to identify the multiplicative streamflow process model 416 and its linearized version, and how to derive the non-uniform time-steps lengths from 417 discharge data. Model identification for the proposed model is not more complex than 418 for classic periodic autoregressive models with monthly time-step. The proposed features 419 are independent from each other, then each of them can be applied separately. This work 420 address specific problems encountered in SDDP, but some results may have a broader 421 (potential) validity in time series modeling for synthetic streamflow sequences generation. 422 We applied the model to the streamflow process at Soukoutali, on the Senegal River, 423 for the operational rules design of a single reservoir system. Model identification would 424 not be different if the reservoirs were many. The proposed multiplicative model offers 425 a better representation of the streamflow process both in the forward phase, where it 426 correctly represents the streamflow dynamics and the discharge distribution, and in the 427 backward phase, where it correctly represents the residual distribution, avoiding the fat-428 tail phenomenon, otherwise present in the classic Thomas-Fiering model. The model 429

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using non-uniform time-steps has a relatively homogeneous variance. This brings in a 430 practical advantage: the non-uniform time-steps follow closer the changing hydrological 431 variability along the year, adapting the decision more frequently during high variability 432 periods, resulting in enhanced system performance evaluation. If time-step aggregation is 433 sufficiently fine, a non-uniform aggregation may even make time decomposition needless. 434 Model linearization, used in the backward phase, introduces an error. The analysis 435 on error due to linearisation show that the error average is negligible, growing to about 436 5%-10% of median discharge at 5%-95% quantiles of  $q_{\tau-1}$ , which we consider satisfactory; 437 this depends, however, on the specific test-case and, a priori, we cannot exclude it to be 438 a limitation for this model. 439

#### Appendix A: Linear model parameters derivation

Parameters  $\rho_{\tau,i}, \omega_{\tau,i}, \kappa_{\tau}$  are derived by linerization of Model (8) on the median of its deterministic inputs, at  $q_{t-i} = \bar{q}_{\tau-i}$ , with  $\bar{q}_{\tau}$  defined in Equation (6), and  $\xi_{t-j} = 1$ .

Equation A1 is model (8) for  $\xi = 1$  written as Taylor expansion on its deterministic inputs.

$$q_t^{\rm nl} \approx \sum_{i=1}^p \frac{\partial q_t^{\rm nl}}{\partial q_{t-i}} \cdot (q_{t-i}^{\rm nl} - \bar{q}_{t-i}) + \sum_{\substack{j=1\\j=1}}^q \frac{\partial q_t^{\rm nl}}{\partial \xi_{t-j}} (\xi_{t-j} - \bar{\xi}_{t-j}) + q_t^{\rm nl} (q_{t-i}, \xi_{t-j})$$
(A1)

Equations (A2) are the derivatives of  $q_t$  on inputs  $q_{t-i}$  and  $\xi_{t-j}$ .

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$$\frac{\partial q_t^{\text{nl}}}{\partial q_{t-i}} = \alpha_\tau \cdot \left[ \phi_{\tau,i} \cdot q_{t-i}^{(\phi_{\tau,i}-1)} \cdot \prod_{k \in \{1,\dots,p\} \setminus i} q_{t-k}^{\phi_{\tau,k}} \cdot \prod_{j=1}^q \xi_{t-j}^{\psi_{\tau,i}} \right]$$
(A2a)

$$\frac{\partial q_t^{\text{nl}}}{\partial \xi_{t-j}} = \alpha_\tau \cdot \left[ \psi \cdot \xi_{t-j}^{(\psi_{\tau,j}-1)} \cdot \prod_{k \in \{1,\dots,q\} \setminus j} \xi_{t-k}^{\psi_{\tau,k}} \right]$$
(A2b)

Separating members of Equation (A1) by their inputs, and considering that  $\alpha_{\tau}$  = 445  $\bar{q}_{\tau}/\prod_{i=1}^{p} \bar{q}_{\tau-i}^{\phi_{\tau,i}}$  we derive, in Equations (A3), parameters (10) of Model (9). 446

$$\begin{split} \rho_{\tau,i} &= \frac{\partial q_t^{\mathrm{nl}}}{\partial q_{t-i}} (\bar{q}_{t-i}, \bar{\xi}_{t-j}) = \\ &= \frac{\bar{q}_\tau}{\prod_{i=1}^p \bar{q}_{\tau-i}^{\phi_{\tau,i}}} \cdot \left[ \phi_{\tau,i} \cdot \frac{\bar{q}_{t-i}^{\phi_{\tau,i}}}{\bar{q}_{t-i}} \cdot \prod_{k \in \{1, \dots, p\} \setminus i} \bar{q}_{t-k}^{\phi_{\tau,k}} \right] = \end{split} \tag{A3a} \\ &= \phi_{\tau,i} \cdot \frac{\bar{q}_\tau}{\bar{q}_{\tau-i}} \\ \omega_{\tau,j} &= \frac{\partial q_t^{\mathrm{nl}}}{\partial \xi_{t-j}} (\bar{q}_{t-i}, \bar{\xi}_{t-j}) = \\ &= \frac{\bar{q}_\tau}{\prod_{i=1}^p \bar{q}_{\tau-i}^{\phi_{\tau,i}}} \cdot \left[ \psi \cdot \prod_i^p \bar{q}_{t-i}^{\phi_{\tau,i}} \right] = \end{aligned} \tag{A3b} \\ &= \psi_{\tau,i} \cdot \bar{q}_\tau \\ \kappa_\tau &= q_t^{\mathrm{nl}} (\bar{q}_{t-i}, \bar{\xi}_{t-j}) - \sum_{i=1}^p \frac{\partial q_t^{\mathrm{nl}}}{\partial q_{t-i}} (\bar{q}_{t-i}, \bar{\xi}_{t-j}) \cdot \bar{q}_{t-i} - \sum_{j=1}^q \frac{\partial q_t^{\mathrm{nl}}}{\partial \xi_{t-j}} (\bar{q}_{t-i}, \bar{\xi}_{t-j}) \cdot \bar{\xi}_{t-j} = \\ &= \bar{q}_\tau \cdot \left( 1 - \sum_{i=1}^p \phi_{\tau,i} - \sum_{j=1}^q \psi_{\tau,i} \right) \end{aligned} \tag{A3c}$$

#### Appendix B: Energy objective linearization

Energy function, from Equation 17, is written in Equation (B1) as function of problem 447 variables. 448

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 $E_t = \eta \cdot \Delta k(\tau) \cdot \left[ R_1(v_t) - R_2(r_{\text{turb},t} + r_{\text{spill},t}) \right] \cdot r_{\text{turb},t}$ (B1)

In Equation (B1), the hydraulic head  $\Delta h$ , is written as function of problem variables;  $h_t = R_1(v_t)$  is the stage-storage curve and  $h_t^v = R_2(r_{turb,t} + r_{spill,t})$  the rating-curve, where  $h_t$  is the water level in the reservoir and  $h_t^v$  the tailwater elevation, downstream of the reservoir.

<sup>453</sup> Equations (B2) are the partial derivatives of Equation (B1).

$$\frac{\partial E_t}{\partial v_t} = \eta \cdot \Delta k(\tau) \cdot R_1'(v_t) \cdot r_{\text{turb},t}$$

$$\frac{\partial E_t}{\partial r_{\text{turb},t}} = \eta \cdot \Delta k(\tau) \cdot \left[ R_1(v_t) - R_2'(r_{\text{turb},t} + r_{\text{spill},t}) \cdot r_{\text{turb},t} - R_2(r_{\text{turb},t} + r_{\text{spill},t}) \right]$$

$$\frac{\partial E_t}{\partial r_{\text{spill},t}} = -\eta \cdot \Delta k(\tau) \cdot R_2'(r_{\text{turb},t} + r_{\text{spill},t})$$
(B2)

In Equations (B3), we consider a cylindrical reservoir in proximity of the operational reservoir water level,  $h_0$ , and a linear rating curve in proximity of discharges  $r_{\text{turb},0} + r_{\text{spill},0}$ .

$$R_1(v_t) \approx h_0 + \frac{1}{A_0} \cdot (v_t - v_0)$$
 (B3a)

$$R_2(r_{\text{turb},t} + r_{\text{spill},t}) \approx h_0^v + m_v(r_{\text{turb},t} - r_{\text{turb},0} + r_{\text{spill},t} - r_{\text{spill},0})$$
(B3b)

where  $A_0$  is the reservoir surface corresponding to  $h_0$ ,  $h_0^v = R_2(r_{turb,0} + r_{spill,0})$ , and  $m_v = R_2'(r_{turb,0} + r_{spill,0})$ .

<sup>458</sup> Considering Equations (B2) and (B3), we get, in Equation (B4), the linear approxima-<sup>459</sup> tion of Equation (B1).

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$$E_{t} \approx E_{t}(v_{0}, r_{\text{turb},0}, r_{\text{spill},0}) +$$

$$+ \frac{\partial E_{t}}{\partial v_{t}} \cdot (v_{t} - v_{0}) +$$

$$+ \frac{\partial E_{t}}{\partial r_{\text{turb},t}} \cdot (r_{\text{turb},t} - r_{\text{turb},0}) +$$

$$+ \frac{\partial E_{t}}{\partial r_{\text{spill},t}} (r_{\text{spill},t} - r_{\text{spill},0}) =$$

$$= \eta \cdot \Delta k(\tau) \cdot \left\{ [h_{0} - h_{0}^{v}] \cdot r_{\text{turb},0} +$$

$$+ \left[ \frac{r_{\text{turb},0}}{A_{0}} \right] \cdot (v_{t} - v_{0}) +$$

$$+ [h_{0} - h_{0}^{v} - m_{v}(1 + r_{\text{turb},0})] (r_{\text{turb},t} - r_{\text{turb},0})$$

$$+ [-m_{v}](r_{\text{spill},t} - r_{\text{spill},0}) \right\}$$
(B4a)
(B4b)

Equation (B4b) gives the weights for  $v_t, r_{turb,t}$ , and  $r_{spill,t}$  that maximise energy, as in Equation (18). For constant tailwater elevation, then  $m_v = 0$ , and Equations (B4b) and (18) can be simplified.

#### Appendix C: List of main variables

- 463 We use here the classic convention of representing vectors in bold.
- 464  $\mathbf{v}_t$  Reservoir volumes  $[m^3]$
- 465  $\mathbf{r}_t$  Discharge decision  $[m^3/s]$
- 466  $\mathbf{q}_t$  (Flow) scenarios  $[m^3/s]$
- $_{467}$   $N_{\rm res}$  Number of reservoirs [-]
- 468  $N_{\text{dec}}$  Number of discharge decisions [-]
- $_{469}$  N<sub>scen</sub> Number of scenarios [-]
- 470 I Input matrix [-]
- $_{471}$  O Output matrix [-]

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- 472  $\mathbf{c}_t(\cdot)$  Inequality constraints [-]
- $_{473}$   $g(\cdot)$  Time-step objective function [-]
- $F(\cdot)$  Cost-to-go function [-]
- 475  $\mathcal{F}(\cdot)$  cost-to-go function approximation by Bender's cuts [-]
- 476  $\mathbb{E}$  Expected value [-]
- $f_{\mathbf{Q}_{t}}(\mathbf{q}_{t})$ ; Probability density function of  $q_{t}$  [-]
- 478  $\alpha_{\tau,i}, \phi_{\tau,i}, \psi_{\tau,i}, \sigma_{\tau}^2$  Periodic ARMA parameters
- 479  $\rho_{\tau,i}, \omega_{\tau,i}, \kappa_{\tau}$  Linearized multiplicative model parameters
- $_{480}$   $\bar{q}_{\tau}$  Climatic average of  $Q_{\tau} \ m^3/s$
- 481 t Time-step index [-]
- 482 au Periodic time-step index [-]
- $_{483}$   $\Delta t$  Daily time-step length [86400s]
- <sup>484</sup>  $\Delta k$  Number of daily time-steps [days]
- 485 T Period length [-]
- <sup>486</sup> *H* Optimization horizon [-]
- $e_t$  Error due to model linearization  $[m^3/s]$
- $_{488}$  CV Cumulative variance [-]
- $v_{\min}$  Minimum reservoir volume  $[3.9 \times 10^9 m^3]$
- $v_{\text{max}}$  Maximum reservoir volume  $[1.5 \times 10^{10} m^3]$
- $v_{\text{safety}}$  Reservoir volume safety limit  $[1.18 \times 10^{10} m^3]$
- $r_{\rm spill,max}$  Maximum discharge through turbines  $[500m^3/s]$
- $_{493}$   $J_E$  Annual average energy production [GWh]

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- 494 E Daily energy production [MWh]
- $_{495}$   $E^{\text{op}}$  Linearized energy production objective [MWh]
- $v_0$  Operational volume  $[1.18 \times 10^{10} m^3]$
- $r_{\rm turb,0}$  Operational release through turbines  $[270m^3/s]$
- $r_{\rm spill,0}$  Operational release through spillages  $[0m^3/s]$
- <sup>499</sup> Acknowledgments. Luciano Raso's research is funded by the AXA Research Fund.

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Figure 1. Map of the Senegal Basin.



Figure 2. Inflow at Soukoutali from 1 January 1950 to 31 December 2013, daily time-step.

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(b)  $\tau_d$  between 135 and 285

**Figure 3.** Cumulative Variance on  $\tau_d$ .



(b) Monthly aggregation

Figure 4. Inflow at Soukoutali from 1 January 1950 to 31 December 2013, aggregated observed discharge.

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**Figure 5.**  $\phi_{\tau}$  for non-uniform and monthly time-step



Figure 6. Normality plot for logarithms of model residuals of Model (9), Multiplicative Model (a), and residuals of Thomas-Fiering, Additive Model (b).

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Figure 7. Observed discharge data and 95% confidence bands for Model (8), or multiplicative model (continuous bold lines) and Thomas-Fiering, or additive model (dashed bold lines).



(b) error due to linearisation vs  $q_{t-1}$  distribution.

Figure 8. Error due to linearisation at  $\tau = 7$ .

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Figure 9. Schema of the Senegal River System at Manantali, including: Manantali reservoir volume,  $v_t$ , inflows at Soukoutali,  $q_t$ , and discharge decisions,  $\mathbf{r}_t$ .



**Figure 10.** Reservoir volume (a) and discharge trough turbines (b) for monthly (dashed line) and non-uniform (continuous line) aggregation, year 2005.

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