

# Recent climate hiatus revealed dual control by temperature and drought on the stem growth of Mediterranean Quercus ilex

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1	RECENT CLIMATE HIATUS REVEALED DUAL CONTROL BY TEMPERATURE AND
2	DROUGHT ON THE STEM GROWTH OF MEDITERRANEAN $Q$ UERCUS ILEX
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5	Running head: STEM GROWTH AND CLIMATE HIATUS
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# 35 Abstract

36 A better understanding of stem growth phenology and its climate drivers would improve 37 projections of the impact of climate change on forest productivity. Under a Mediterranean 38 climate, tree growth is primarily limited by soil water availability during summer, but cold 39 temperatures in winter also prevent tree growth in evergreen forests. In the widespread 40 Mediterranean evergreen tree species *Quercus ilex*, the duration of stem growth has been shown 41 to predict annual stem increment, and to be limited by winter temperatures on the one hand, and by the summer drought onset on the other. We tested how these climatic controls of Q. ilex 42 43 growth varied with recent climate change by correlating a 40-year tree ring record and a 30year annual diameter inventory against winter temperature, spring precipitation, and simulated 44 45 growth duration. Our results showed that growth duration was the best predictor of annual tree 46 growth. We predicted that recent climate changes have resulted in earlier growth onset (-10 47 days) due to winter warming and earlier growth cessation (-26 days) due to earlier drought 48 onset. These climatic trends partly offset one another, as we observed no significant trend of 49 change in tree growth between 1968 and 2008. A moving-window correlation analysis revealed 50 that in the past, Q. ilex growth was only correlated with water availability, but that since the 51 2000s, growth suddenly became correlated with winter temperature in addition to spring 52 drought. This change in the climate-growth correlations matches the start of the recent 53 atmospheric warming pause also known as the 'climate hiatus'. The duration of growth of Q. 54 *ilex* is thus shortened because winter warming has stopped compensating for increasing drought 55 in the last decade. Decoupled trends in precipitation and temperature, a neglected aspect of climate change, might reduce forest productivity through phenological constraints and have 56 57 more consequences than climate warming alone.

58

# 59 **LIST OF ABBREVIATIONS USE IN PAPER**

- 60 **FS**: Long-term field survey of diameter at breast height (DBH), measured from 1986 to 2013.
- 61 **RW**: Ring-width series, measured from 1942 to 2008.
- 62 **AD**: Automatic dendrometer series, measured from 2004 to 2013.
- 63 **Puéchabon station**: weather station located in the study site since 1984.
- 64 SML station: weather station located in St Martin-de-Londres, 12 km away from the study site,
- data available from 1966 to 2013 (Meteo France).
- 66 **T**JFM: Mean of daily temperature from January to March (°C).
- 67 **P**<sub>AMJ</sub>: Sum of precipitation from April to June (mm).
- 68 **WSI**: Water stress integral, a time-cumulated drought severity index (MPa day)
- 69 **t**<sub>0</sub>: Day of year when stem growth starts

- t1: Day of year when stem growth stops in early summer  $\Delta$ tt0-t1: Duration of the period between t0 and t1 computed for each set of simulated phenological
- thresholds.
- **BAI**: Basal area increment of the stems, expressed in mm<sup>2</sup> year<sup>-1</sup> **PET**: Potential evapotranspiration (mm)

## 76 **INTRODUCTION**

Stem growth, recorded in annual tree rings, is a synthetic surrogate of carbon input in standing biomass (Babst *et al.*, 2014) and an index for tree vitality and fitness (Bigler & Bugmann 2003; Benito-Garzon *et al.*, 2013). In return, it is the record of the yearly climatic conditions that have enabled stem growth over a period lasting from decades to centuries (Fritts 1976; Briffa *et al.*, 2002). Understanding the factors that control stem growth is therefore crucial to assess the impact of past and future climate change on forests, but remains a major scientific issue due to a series of currently unresolved uncertainties.

84 First, the climate sensitivity of tree growth is generally considered to depend on photosynthesis and respiration fluxes because most process-based models assume that tree 85 86 growth is carbon limited (Davi et al., 2006; Gaucherel et al., 2008; Keenan et al., 2011). Carbon 87 allocation to sapwood is therefore the key process driving the carbon sink, and is at the heart of 88 tree growth simulations in most terrestrial biosphere models (e.g. Sitch et al., 2003; Friedlingstein et al., 2006; Fisher et al., 2010). Several allocation schemes exist (e.g. fixed 89 90 allocation, the pipe model, and hierarchical allocation between plant organs, Schippers et al., 91 2015), but there is accumulating experimental evidence that cambial activity (sink limitation) 92 is more sensitive to environmental stressors than carbon assimilation (source limitation; Körner 93 2003; Fatichi et al., 2014; Guillemot et al., 2015; Körner 2015; Delpierre et al., 2016a). Taking 94 sink limitation into account in terrestrial biosphere models is thus likely to improve their ability 95 to simulate past and future forest productivity (Leuzinger et al., 2013; Fatichi et al., 2014), but 96 a better characterization of the phenology and climate sensitivity of wood formation is a 97 prerequisite (Rossi et al., 2011; Rossi et al., 2014; Lempereur et al., 2015; Delpierre et al., 98 2016a; Delpierre et al., 2016b).

99 Second, global climate change may simultaneously exert opposing influences on forest 100 functioning, and these impacts are difficult to disentangle. For example, increased water use 101 efficiency due to atmospheric CO<sub>2</sub> enrichment is sometime reported to increase tree growth in 102 spite of concurrently increasing aridity (Martinez-Vilalta et al., 2008; Koutouvas 2013), but the 103 opposite finding is far more common (Peñuelas et al., 2011). Warming also exerts opposing 104 impacts on tree growth: on one hand it increases the length of the growing season (Keeling et 105 al., 1996; Dragoni et al., 2011) while on the other hand it exacerbates atmospheric evaporative 106 demand and hence water limitation (Angert et al., 2005; Zhao & Running, 2010). Warming 107 may therefore affect tree growth positively (e.g. Rossi et al., 2014) or negatively (e.g. Brzostek 108 et al., 2014) depending on the ecosystem under consideration.

109 Third, relationships between tree growth and climate may change under the influence of 110 climate change (e.g. Briffa et al., 1998; Büntgen et al., 2006; Carrer & Urbinati 2006; D'Arrigo et al., 2008). A classic example of this phenomenon is the so-called 'divergence problem' in 111 112 northern forests, which is a weakening of the positive temperature response of tree growth in strongly temperature limited ecosystems since the middle of the 20th century (D'Arrigo et al., 113 114 2008). The consequence of such temporal variations in tree growth sensitivity to climate is that 115 empirical dendrochronological models based on statistical links between tree growth and climate cannot be extrapolated to long term historical and future climate conditions with 116 117 confidence (e.g. Gea-Izquierdo et al., 2013; Subedi & Sharma 2013).

118 Fourth, climate change itself is a complex phenomenon that may result in changes in the mean climate (IPCC 2014), in the frequency of extreme climatic events (Meehl & Tebaldi 2004) 119 120 and in seasonality (Giorgi et al., 2011; Ruffault et al., 2013), all of which may have contrasted 121 impacts on tree growth and function. In addition, climate change is not monotonous and the 122 recent slowing down of the warming trend, early identified by Easterling & Wehner (2009) and 123 later called the 'climate hiatus' (Trenberth & Fasullo, 2013), is currently a topic of increasing 124 scientific interest. But if most scientific studies have focused on explaining the causes and 125 processes behind this 'climate hiatus' or pause in atmospheric warming (Meehl et al., 2011; 126 Balmaseda et al., 2013; England et al., 2014; Steinman et al., 2015), few have attempted to 127 assess its ecological consequences, especially for processes that are highly sensitive to climate, 128 such as tree growth.

129 In water limited ecosystems, where forest productivity is primarily limited by drought 130 (Churkina & Running 1998; Zhao & Running 2010), the supply of water from rainfall appears 131 to be the key variable to explain tree growth (Babst et al., 2013). In the case of the 132 Mediterranean species *Quercus ilex*, the conclusions of dendrochronological studies converge 133 more specifically to identify winter and spring rainfall amounts as the key drivers of annual 134 stem growth (Gutiérrez et al., 2011). In contrast, Q. ilex stem growth is generally poorly 135 correlated with winter and spring temperatures. Daily measurements of basal area increment 136 over an eight year period (2004-2011) with automatic dendrometers in a Q. ilex forest, 137 demonstrated that the duration of spring growth was an accurate predictor of annual stem 138 growth (Lempereur et al., 2015). After observing that the duration of spring growth was 139 determined by winter temperature for growth start and a lower threshold of predawn plant water 140 potential of -1.1 MPa for growth cessation, Lempereur et al. (2015) proposed a new hypothesis 141 wherein the annual growth of Q. ilex is under the dual control of winter temperature and spring 142 water limitation.

143 In this contextual framework, we used both tree ring records over a 40-year period and 144 an annual diameter inventory over a 30-year period to (i) test whether the stem growth 145 phenology approach proposed by Lempereur et al. (2015) can explain Q. ilex growth over a 146 long retrospective period with different growth estimation methods, (ii) assess the relative 147 importance of drought and temperature in limiting Q. ilex growth under a Mediterranean 148 climate, and (iii) evaluate the stationarity of these limitations in recent decades. In particular, 149 we aimed to identify whether any temporal shift in the climate controls of *Q*. *ilex* growth 150 (drought or temperature plus drought) are concomitant to, and the functional consequence of, 151 the recent pause in atmospheric warming.

152

## 153 MATERIALS AND METHODS

#### 154 *Site description*

155 The study site is located 35 km north-west of Montpellier (southern France), on a flat plateau, in the Puéchabon state forest (43°44'29"N, 3°35'45"E, 270 m a.s.l.). This forest has 156 been managed as a coppice for centuries and the last clearcut took place in 1942. Vegetation is 157 158 largely dominated by a dense overstory of the evergreen oak *Quercus ilex*. In 2014, the average 159 top canopy height was 5.5 m and the stem density was 4700 ( $\pm$ 700) stems ha<sup>-1</sup>. Understory 160 evergreen species, Buxus sempervirens, Phyllirea latifolia, Pistacia lentiscus and Juniperus 161 *oxycedrus*, compose a sparse shrubby layer (height < 2 m) with less than 25% cover. The 162 climate is Mediterranean, with 80% of the rainfall occurring between September and April. Average annual precipitation for the 1984-2013 period was 916 mm (550 to 1549 mm). Mean 163 164 annual temperature over the same period was 13.2 °C with a minimum in January (5.5 °C) and 165 a maximum in July (22.9 °C). The very shallow bedrock has a hard Jurassic limestone origin. 166 The volumetric fraction of stones and rocks averages 0.75 in the top 0-50 cm and 0.90 below. 167 The stone free fine fraction of the soil in the 0-50 cm layer is a homogeneous silty clay loam 168 (USDA texture triangle) comprising 38.8% clay, 35.2% silt and 26% sand.

169

#### 170 Monitoring annual stem growth

171 Stem diameter at breast height (DBH) was measured annually from 1986 to 2013 on 319 172 trees distributed in eight circular plots (diameter 20 m) within an area of 2 ha. The 319 trees 173 used in this field survey (hereafter FS, Fig. 1) were selected so as to include all the diameter 174 size classes in 1986. DBH was measured every winter under dry conditions using a diameter 175 tape at a height identified by a paint mark on the stem. In January 2008, the stem diameters 176 were distributed as follows: 58% < 8 cm DBH, 24% between 8 and 10 cm and 18% >10 cm. 177 Because the smaller size classes exhibited a very low growth and a high mortality rate, we 178 restricted our analysis to the 125 trees with a DBH > 8 cm in 2008 to obtain a stronger growth 179 signal (Table 1). From 2004 to 2013, stem growth was measured more frequently and with 180 greater accuracy using automatic band dendrometers on a subset of trees in two neighboring plots (for more details, see Lempereur et al., 2015). Automatic dendrometers (ELPA-98, 181 182 University of Oulu, Finland; hereafter AD) were set up at a height of 1.3 m above the ground on 6 to 12 trees with DBH > 7 cm. The two datasets (FS and AD) were strongly correlated 183 184 during their overlapping period (Table 1), thus demonstrating the relative accuracy of FS in 185 estimating the annual growth of the largest trees.

186



187

Fig. 1 Chronological timeline of the main data sources used in this study. Stem growth series: automatic
 dendrometers (AD), field survey (FS) of stem diameter at breast height (DBH) and ring width series (RW).
 Climatic data: Puéchabon meteorological station and St Martin-de-Londres (SML) meteorological station.

191

# 192 Dendrochronological procedures

193 In 2005 and 2008, 15 and 12 stems, respectively, were selected among the largest size 194 classes of the field survey (FS) sample (10 cm < DBH < 16 cm) and cut down to measure ring-195 width chronology. Cross sections were collected at stump height, air dried, sanded and polished 196 (40 to 400 grit). The longest growth radii in each cross section were selected and compared 197 pairwise under a binocular magnifier. In old coppiced oaks, rings are narrow and sometimes 198 not clearly visible after 40 years of cambial age; consequently only 12 of the sampled stems 199 could be cross dated with confidence, seven stems in 2005 and five stems in 2008. The first 200 ring was formed in 1942 after the clearcut in winter 1941-1942, so data in the ring width series

201	(hereafter RW) were obtained from 1942 to 2008 (Table 1, Fig. 1). Cross dating was facilitated
202	by the presence of frost rings corresponding to the severe winters of 1963, 1985 and 1987.
203	These rings, typical for years with severe freezing episodes, were abnormally wide because the
204	cells of their initial area were crushed and dislocated, a typical constraint on tree ring
205	interpretation in woody species under a Mediterranean climate (Cherubini et al., 2003). This
206	particular feature made them useful as markers for cross dating but prevented the reliable
207	measurement of annual stem growth, so these years were excluded from subsequent analyses.
208	Cross correlation coefficients on annual tree ring indices calculated among the 12 individual
209	series were all higher than 0.6 (P-value<0.05). The ring width measured at stump height and
210	along the longest growth radius was then rescaled to the tree DBH at the time of the cut. This
211	required correction for the tapering between stump height (20 cm) and breast height (130 cm)
212	as well as correction for the bark thickness (see Fig. S1 for details). The average tree ring width
213	after rescaling, which was used to calculate the basal area increment (BAI), was 954 $\mu m$
214	(CV = 48%) between 1942 and 2008, and 792 $\mu m$ (CV = 38%) over the period 1968-2008 used
215	in our study.

216

Table 1 Main characteristics of the three growth series used in the study, automatic dendrometer (AD), field survey
 (FS) and ring width (RW), over their respective complete temporal series and their overlap periods. Coefficients
 of correlation (r) between FS and AD over the period 2004-2013 and between RW and FS over the period 1987 2008 are given (P-value < 0.05: \*; P-value < 0.001: \*\*\*).</li>

Characteristics of the growth series						
	Method	Period	Mean BAI (mm <sup>2</sup> )	DBH 2008 (cm)	Sample size	
AD	Auto Dendro	2004-2013	185.8(±70,00)	11.1 (±2.45)	12	
FS	Diameter tape	1987-2013	154.4 (±44,47)	10.5 (±1.85)	125	
RW	Ring width	1968-2008	204.6 (±55,39) 12.7 (±1.75)		12	
Correlations on overlap periods						
	Ovelap period	Growth series	Mean BAI (mm <sup>2</sup> )	r		
	2004-2013	AD	185.8(±70.00)	0.00 (***)		
AD <i>vs.</i> F5		FS	142.3 (±55.37)	0.90 (****)		
ES un DW	1987-2008	FS	158.2 (±44.68)	0 55 (*)		
г <i>э vs</i> . кw		RW	196.2(±45.00)	0.55 (*)		

221

### 222 Climate variables

The Puéchabon meteorological station (hereafter Puéchabon) is located in a clearing 200 m away from the study plot and has provided daily on-site climate data since 1984 (Fig. 1). 225 Precipitation was measured with a tipping bucket rain gauge (ARG100; Environmental 226 Measurements, Sunderland, UK) calibrated to 0.2 mm per tip and placed 1 m above the ground, 227 air temperature was recorded with a MP100 sensor (Rotronic, Bassersdorf, Switzerland) at a height of 2 m, and net radiation was measured with a pyranometer (SKS1110; Skye Instruments,
UK) at a height of 2 m above the ground.

To extend our analysis before 1984, we used climate variables (daily rainfall, daily 230 231 minimum and maximum temperature) from the St Martin-de-Londres meteorological station 232 (hereafter SML; 43°47'06''N; 3°43'48''E, altitude 194 m a.s.l., located about 12 km away from 233 the study site; source Meteo-France) over the period 1966-2013 (Fig. 1). The mean daily 234 temperature was calculated as the average of the minimum and maximum daily temperature. 235 The climate data from SML station were compared to on-site measurements in the overlap 236 period 1984-2013 (Table S1 and Fig. S2). The comparison showed close agreement for 237 temperature but a higher rainfall amount in SML than in Puéchabon. No significant biases 238 (slope not different from 1, and distance to origin not different from 0) were observed between 239 the two meteorological stations in the two crucial variables related to stem growth: the mean 240 temperature from January to March (T<sub>JFM</sub>) and the sum of precipitation from April to June 241 (P<sub>AMJ:</sub> Fig. S2).

242 Daily solar radiation at the SML station was calculated from the processing chain 243 described in Kumar et al. (1997). In the first step, theoretical clear sky solar radiation (R<sub>a</sub>) was 244 calculated from the daily timing of sunrise and sunset, the bi-hourly sun azimuthal angle and 245 the atmospheric transmittance according to longitude using R Cran Packages 'RAtmosphere' 246 and 'oce'. No topographical effects were taken into consideration. Actual solar radiation (Rs, in MJ m<sup>-2</sup> day<sup>-1</sup>) was then calculated from maximum and minimum daily temperature (T<sub>max</sub> and 247  $T_{min}$ , respectively) and the clear sky theoretical radiation following the Hargreaves equation: 248  $R_s = k_{Rs} \sqrt{(T_{max} - T_{min})} R_a$ , where the adjustment coefficient k<sub>Rs</sub> was set to 0.16 for interior 249 location (Allen et al., 1998). Results of the Hargreaves equation were validated against solar 250 251 radiation measured in Puéchabon since 1998.

252

# 253 Modeling predawn leaf water potential and water stress integral

Soil water storage integrated over the rooting depth, c.a. 4.5 m, was measured during the vegetative periods of 1984-1986 and from July 1998 to August 2009 at approximately monthly intervals, using a neutron moisture gauge (503DR Hydroprobe, CPN, Concord, CA, USA). Discrete measurements were interpolated at a daily time step using the soil water balance model described by Rambal (1993) and further used by Rambal *et al.* (2014). The model was driven by daily values of incoming solar radiation, minimum and maximum temperatures, and rainfall amount. Potential evapotranspiration was computed using the Priestley-Taylor equation 261 (Priestley & Taylor, 1972). The reduced major axis (RMA) regression between neutron moisture gauge measurements and model simulations yielded an  $R^2$  of 0.93, the slope was 262  $0.94 \pm 0.05$  (P < 0.0001, n = 91) and the intercept not significantly different from 0. Soil water 263 264 storage and soil water potential were linked by a Campbell-type retention curve (Campbell, 265 1985) whose parameters are strongly dependent on soil texture (Saxton et al., 1986; Rambal et 266 al., 2003). Predawn leaf water potential used for model validation was measured about eight 267 times a year between April and October from 2003 to 2009 on a subsample of four trees among 268 those equipped with automatic dendrometers (see Limousin et al., 2012). RMA regressions between measured and simulated values of predawn leaf water potential yielded an R<sup>2</sup> of 0.84, 269 270 the slope was  $0.93 \pm 0.05$  (P < 0.0001, n = 54), and the intercept was not significantly different 271 from 0. We used the simulations of predawn water potential rather than soil water content, as 272 the former is more closely linked with plant functioning (Rambal et al., 2003). The daily 273 simulations of predawn water potential were performed with the climate data from the Puéchabon and SML stations for the periods 1984-2012 and 1966-2012, respectively. The water 274 275 stress integral (WSI), defined by Myers (1988) as the seasonally or yearly sum of predawn 276 water potential, was used as a drought severity index to quantify annual or seasonal water stress.

277

# 278 Duration of spring basal area increment: calculations of dates $t_0$ and $t_1$

279 The dates of stem growth phenology that bounded the spring growth period, the DOY 280 (day of the year) when stem growth starts (hereafter t<sub>0</sub>) and the DOY when stem growth stops 281 in early summer (hereafter t<sub>1</sub>, Fig. S3), were estimated for each year. The relationships between 282  $t_0$  or  $t_1$  and climate variables were calibrated by Lempereur *et al.* (2015) on observations of  $t_0$ 283 and t<sub>1</sub> obtained with automatic dendrometers from 2004 to 2011. Briefly, t<sub>0</sub> was defined as the 284 first day at which basal area exceeded the culmination of the previous year, and t<sub>1</sub> as the first 285 day when BAI became null or negative (see Lempereur et al., 2015 for methodological details). 286 t<sub>0</sub> and t<sub>1</sub> were estimated using climate data from both Puéchabon and the nearby SML station 287 over their corresponding timeframes. t<sub>0</sub> was predicted by a nonlinear relationship with the mean 288 temperature from January to March ( $T_{JFM}$ ). The relationships were fitted between t<sub>0</sub> and  $T_{JFM}$ 289 measured in Puéchabon ( $t_0 = 849.2 \exp(-0.6436 T_{JFM}) + 121$ ;  $R^2 = 0.95$ ; RMSE = 2.6 days; 290 Lempereur *et al.*, 2015) and between  $t_0$  and  $T_{JFM}$  measured in SML ( $t_0 = 4300 \exp(-0.943)$ 291  $T_{JFM}$ )+124.4;  $R^2 = 0.95$ ; RMSE = 2.3 days; Fig. S4). t<sub>1</sub> was predicted by the DOY when the plant water potential simulated using climate data from Puéchabon reached a threshold of -292 293 1.1 MPa ( $R^2 = 0.75$ ; RMSE = 7 days, Lempereur *et al.*, 2015), corresponding to a DOY when

the plant water potential simulated using climate data from SML reached a threshold of -1.1 MPa +  $\varepsilon$  (with  $\varepsilon$  = -0.1; R<sup>2</sup> = 0.73; RMSE = 13.5 days) due to the slight differences in temperature and precipitation between the two stations (see Fig. S5). The duration  $\Delta t_{t0-t1}$ corresponds to the duration of the period between t<sub>0</sub> and t<sub>1</sub> and was computed for each pair of simulated phenological thresholds.

299

# 300 Data processing and statistical analyses

301 Annual stem growth (DBH and ring widths) was converted into annual basal area 302 increments (BAI, expressed in mm<sup>2</sup> year<sup>-1</sup>) for the two data series. Analyses were performed 303 with averaged BAI values of 125 and 5-12 individual trees for FS and RW, respectively. The 304 year 1992 presented two periods when predawn plant water potential fell below the threshold 305 of -1.1 MPa, in spring and autumn (Fig. S6). This unusual bi-modal drought prevented us from 306 determining the duration  $\Delta t_{t0t1}$  so 1992 was excluded from subsequent analyses. The growth series BAIFS included measurements that appeared to be outliers compared to BAIRW or BAIAD 307 308 for the years 2002 and 2010 (see Fig. S7). In 2002, BAI<sub>FS</sub> was particularly low when BAI<sub>RW</sub> 309 was close to the expected value, and in 2010, BAI<sub>FS</sub> was clearly higher than the BAI<sub>AD</sub> value. 310 Because BAI<sub>FS</sub> is less accurate than the other two methods, and was calculated as the difference 311 between two subsequent measurements of DBH, these two values were considered to be 312 unreliable and excluded from the analyses performed on BAI<sub>FS</sub>.

313 To investigate the links between tree growth and climate, we computed the correlations 314 between our two different growth datasets (FS and RW) and a set of relevant climate predictors. 315 For each correlation tested, we reported the Pearson's correlation coefficient and its 316 significance based on the standard bootstrap method with 1000 samples taken from the original 317 distribution of climate and tree ring data. Within each sample, the number of observations was 318 random and followed a geometric distribution. We used the 'treeclim' R package built upon the 319 DENDROCLIM2002 statistical tool dedicated to tree ring analysis (Biondi & Waikul, 2004; 320 Zang & Biondi, 2015). Climate predictors were derived from temperature and precipitation at 321 monthly, seasonal and annual time scales, and from our functional index of spring growth 322 duration ( $\Delta t_{t0t1}$ ). To test for the stationarity of the drought and temperature controls over the 323 studied period, we also performed the same bootstrap sampling Pearson correlation analysis on 324 a 10-year moving window along the whole series.

The temporal trends in stem growth and climate variables were characterized using both trend tests and breakpoint analyses. Temporal trends in time series were estimated using the Theil-Sen test (Sen 1968) from the 'openair' R cran package, applied on a 1000 block bootstrap simulations to account for auto-correlated variables (Kunsch 1989), with the block length set to n/3, n being the length of the time series. Temporal breakpoints in the time series were assessed by computing the yearly F-statistic (sequential F-test) and the *supF* statistic was then used to test for their significance.

332

# 333 **RESULTS**

### 334 Sensitivity of stem growth to the dual temperature-drought drivers



335

Fig. 2 Time series of annual stem basal area increment (BAI) for the ring width series (RW; 1942-2008), the long-term field survey (FS; 1986-2013) and the automatic dendrometer series (AD; 2004-2013). Error bars are standard errors among the sampled trees (see Table 1 for sample size).

339

340 Annual stem growth (BAI<sub>RW</sub>) from 1942 to 2008 exhibited two distinct phases, with a 341 significant breakpoint around 1967 (Fstat = 81.96; P-value<0.0001; Fig. 2). The first phase (1942-1967) lasted 25 years after the clearcut and showed a linear increase of 5.5 (±0.62) mm<sup>2</sup> 342 y<sup>-1</sup> in annual BAI. This phase, which corresponds approximately to the cut frequency formerly 343 used in the traditional management of this Q. ilex coppice (Floret et al., 1992), was 344 345 characterized by incomplete canopy cover and notable self-thinning. The second phase (1968-2008) showed a stabilization of BAI<sub>RW</sub> around a mean value of 203.8 ( $\pm$ 54.84) mm<sup>2</sup> y<sup>-1</sup>. The 346 influence of climate on stem growth was consequently only considered from 1968 on, to avoid 347 348 the confounding effects caused by changes in competition, space-filling and self-thinning 349 during the first phase.

350 The two long-term growth series  $BAI_{RW}$  and  $BAI_{FS}$  were significantly correlated with each

other during their overlap period 1986-2008 (r = 0.55, P-value < 0.05, Table 1). Both series of

BAI exhibited large between year variations (coefficient of variation, CV = 27% and 29% for

BAI<sub>RW</sub> and BAI<sub>FS</sub>, respectively). The minimum value of BAI was recorded in 2006 for both series:  $BAI_{RW} = 114 \text{ mm}^2 \text{ y}^{-1}$  (SE = 36.5 mm<sup>2</sup> y<sup>-1</sup>) and  $BAI_{FS} = 58 \text{ mm}^2 \text{ y}^{-1}$  (SE = 13.2 mm<sup>2</sup> y<sup>-1</sup>). The maximum BAI was recorded in 1977 for RW (BAI<sub>RW</sub> = 348 mm<sup>2</sup> y<sup>-1</sup>, SE = 22.1 mm<sup>2</sup> y<sup>-1</sup>) and in 2001 for FS (BAI<sub>FS</sub> = 246 mm<sup>2</sup> y<sup>-1</sup>; SE = 20.8 mm<sup>2</sup> y<sup>-1</sup>, *vs.* BAI<sub>RW</sub> = 213 mm<sup>2</sup> y<sup>-1</sup>; SE = 24 mm<sup>2</sup> y<sup>-1</sup>.2).

358



359

**Fig. 3** Relationship between yearly basal area increment (BAI) and the duration of the spring growth period ( $\Delta t_{t0}$ . 1). Ring width (RW, dark grey circle) is shown for the period 1984-2008 and field survey (FS, light grey square) for the period 1986-2013. The linear relationships between  $\Delta t_{t0-t1}$  and  $BAI_{RW}$  (dark grey line;  $BAI_{RW} = 1.34*\Delta t_{t0}$ . 1).  $t_1 + 122.9$ ;  $R^2 = 0.56$ ; P-value < 0.001) or  $BAI_{FS}$  (light grey line;  $BAI_{FS} = 1.09*\Delta t_{t0-t1} + 95.5$ ;  $R^2 = 0.35$ ; P-value < 0.01) are represented. Error bars are standard errors.

366 The DOY of the start  $(t_0)$  and stop  $(t_1)$  of the spring stem growth, simulated over the period 367 1984-2013 using climate data from the Puéchabon station, occurred on average in mid-May (DOY 133, SD = 9.9 days) and early July (DOY 184, SD = 23.5 days), respectively. The growth 368 369 duration  $\Delta t_{t0-t1}$  varied considerably among years (CV = 50%) with values ranging from 2 days 370 in 1995 to 95 days in 2008 (Fig. 3). BAI was linearly correlated with  $\Delta t_{t0-t1}$  for both stem growth series ( $R^2 = 0.56$ ; P-value < 0.0001 and  $R^2 = 0.35$ ; P-value < 0.01 for RW and FS respectively; 371 372 Fig. 3). Moreover,  $\Delta t_{t0-t1}$  was the best explanatory variable for the inter-annual variations in 373  $BAI_{RW}$  compared to other climate variables over the period 1984-2008 (Table 2; 3). We thus 374 concluded that the dual control of annual stem growth by temperature and precipitation, stated 375 by Lempereur et al. (2015) for the period 2004-2011, remained valid for a longer retrospective 376 period and when growth is measured with less accuracy than with automatic dendrometers. The 377 intercepts of the linear relationships between BAI and  $\Delta t_{t0-t1}$ , which represent the residual 378 autumnal growth (Lempereur et al., 2015), differed significantly between the two BAI series (F = 17.97; P-value < 0.001), with 122.9 ( $\pm$ 34.42) mm<sup>2</sup> y<sup>-1</sup> and 95.5 ( $\pm$ 39.09) mm<sup>2</sup> y<sup>-1</sup> for RW 379 380 and FS, respectively (Fig. 3). By contrast, the slopes of these relationships did not differ

significantly between RW and FS (F = 0.46; P-value > 0.05), and the common slope equals 1.23 mm<sup>2</sup> day<sup>-1</sup>. The sensitivity of BAI<sub>RW</sub> to  $\Delta t_{t0-t1}$  was thus identical to the sensitivity of BAI<sub>FS</sub>.

383 Furthermore, neither the extension of the study period to 1968-2008 nor the switch in

384 meteorological stations from Puéchabon to SML, removed the significant correlation between

- 385 BAI<sub>RW</sub> and  $\Delta t_{t0-t1}$  (R<sup>2</sup> = 0.31; P-value < 0.001; Table 2).
- 386

387 Table 2 Mean Pearson's correlation coefficients (and significance derived from a 1000 classical bootstrap 388 sampling) between chronologies of annual stem growth (BAI) and the main phenology explanatory variables. The 389 correlations are given using climate variables measured in Puéchabon field surveys (FS) over the period 1986-390 2013 and ring width (RW) over the period 1984-2008, and climate data from the SML meteorological station over 391 the period 1968-2008 for RW (see Fig. 1). The explanatory variables tested were the start and stop growth dates 392 (t<sub>0</sub> and t<sub>1</sub>, respectively) and the duration of stem growth ( $\Delta t_{10-1}$ ). The coefficient of correlation (r) and the level of 393 statistical significance (\*P-value < 0.05; \*\*P-value < 0.01; \*\*\*P-value < 0.001) are given. Significant correlations 394 are in bold.

	FS	RW	RW
Period	1986-2013	1984-2008	1968-2008
$t_0$	-0.43**	-0.51***	0.18
$t_1$	0.56***	0.68***	0.66***
$\Delta t_{t0-t1}$	0.60***	0.73***	0.56***

395

#### 396 Stem growth response to climate trends from 1968 to 2013

397 Over the longer period 1968-2008,  $\Delta t_{t0-t1}$  was not a better predictor of BAI than the sum 398 of spring precipitation (P<sub>AMJ</sub>) or t<sub>1</sub> alone (r = 0.75; P-value < 0.001 and r = 0.66; P-value < 0.001 399 respectively; Table 2, 3), and we also observed a loss of correlation between BAI and t<sub>0</sub> 400 (r = 0.18; P-value = 0.31; Table 2).

401  $P_{AMJ}$  was the main explanatory climate variable for  $t_1$  (over the period 1968-2008: 402 r = 0.80; P-value<0.001) and it exhibited no temporal trend between 1968 and 2013 (P-403 value = 0.31; Fig. 4b). However, spring water limitation increased (WSI in spring: -0.48 MPa day y<sup>-1</sup>; P-value<0.001; Fig. 4d, Table S3) as a result of increasing spring potential 404 evapotranspiration (PET; +0.028 mm day<sup>-1</sup> y<sup>-1</sup>; P-value<0.001; Fig. 4c, Table S3). This trend 405 406 in PET was mainly due to the significant warming trend in spring temperatures throughout the period 1968-2013 (+0.07°C y<sup>-1</sup>; P-value<0.001; Fig. 4a, Table S3). As a result, drought onset 407  $t_1$  exhibited a temporal trend toward earlier dates (-0.57 day y<sup>-1</sup>; P-value<0.01; Fig. 5). The 408 409 correlation between BAIRW and PAMJ remained significant, however, throughout most of the 410 1968-2008 period (Fig. 6a).

411



412 Fig. 4 Temporal trends in (a) spring (AMJ) temperature (grey circles) and winter (JFM) temperature (light grey 413 inverted triangles), (b) mean annual precipitation (dark grey triangles) and AMJ precipitation (grey circles), (c) 414 mean annual potential evapotranspiration (ETP; dark grey triangles) and AMJ ETP (grey circles) and (d) annual 415 water stress integral (WSI, dark grey triangles) and AMJ three months WSI (grey circles). The solid lines represent 416 significant temporal trends (P-value<0.05). The dashed lines represent the extended trend for temperature after 417 1998, or before 1978 for PET. The effect of the warming pause on the JFM temperature (panel a) is shown by the 418 hatched area between the points and the regression line, the vertical dotted lines indicate the start of the winter 419 warming pause in 2002.

420

421 In parallel, T<sub>JFM</sub> increased significantly throughout the period (+0.04 °C v<sup>-1</sup>; P-422 value<0.05; Fig. 4a) which led to a significant trend of t<sub>0</sub> toward earlier dates (-0.21 day y<sup>-1</sup>; P-423 value<0.05; Fig. 5). However, a pause in the atmospheric winter warming occurred in the last 424 decade. This warming pause is apparent in the discrepancy between actual and expected T<sub>JFM</sub> 425 obtained from the temporal trend built over the period 1978-1998. The year 1978 was chosen 426 because it corresponds to the onset of a rapid warming phase in the Northern Hemisphere (Mann 427 et al., 1999). Similarly, the year 1998 was considered as the onset of the climate hiatus 428 observations (IPCC 2014). We used the period 1978-1998 to build the reference trend line and 429 extrapolate expected T<sub>JFM</sub> values over the 1999-2013 period. We then tested for a breakpoint in 430 the anomalies between observed and expected T<sub>JFM</sub> over the whole period. Such a breakpoint 431 was identified in 2002, indicating a significant slowing down in winter warming after this date 432 (Fstat = 21.52; P<0.001; Fig. 4a). This temporal pattern was also observed at seven other 433 meteorological stations in the region surrounding our study site (Table S4 and Fig. S8), thus 434 confirming the regional occurrence of the globally observed climate hiatus (Trenberth & 435 Fasullo, 2013). The date of growth onset ( $t_0$ ) was very sensitive to the warming pause in  $T_{JFM}$ 436 and a breakpoint in the  $t_0$  trend was observed in 2002 (Fstat = 21.52; P<0.001; Fig. 5). We also 437 observed that the start of the winter warming pause coincided with a change in the climate 438 controls on annual stem growth. Indeed, the correlation between BAI and T<sub>JFM</sub>, which was only 439 significant over one 10-year window between 1968 and 1994, became constantly significant 440 from the time window 1995-2004 on (Fig. 6b).

441

442**Table 3** Mean Pearson's correlation coefficients and significance (derived from a 1000 classical bootstrap443sampling) between annual stem growth (BAI) or spring growth duration ( $\Delta t_{t0t1}$ ) and monthly or seasonal444precipitation (P, in mm) and temperature (T, in °C) for RW data series. The climate data from Puéchabon and445Saint Martin-de-Londres meteorological stations were used for the periods 1984-2013 and 1966-2013,446respectively. The coefficients of correlation (r) and the level of statistical significance (\*P-value < 0.05; \*\*P-</td>447value < 0.01; \*\*\*P-value < 0.001) are given. Significant correlations are in bold.</td>

	BAI			$\Delta t_{t0t1}$				
	1984-2008		1968-2008		1984-2013		1966-2013	
	Т	Р	Т	Р	Т	Р	Т	Р
November (t-1)	0.06	-0.08	-0.23*	-0.03	0.07	0.24	-0.31*	-0.03
December (t-1)	0.30**	0.08	-0.07	-0.04	0.25	0.04	-0.06	-0.03
January	0.50***	-0.06	0.02	0.15	0.51***	0.03	0.34***	0.14
February	0.37*	-0.27*	-0.02	0.06	0.20	-0.16	0.22	0.02
March	0.34**	-0.17	-0.10	0.21	0.41***	-0.10	0.11	0.19*
April	0.05	0.23	0.00	0.33*	-0.11	0.48***	-0.07	0.24*
May	0.12	0.47***	-0.23	0.57***	0.08	0.37	-0.25*	0.48**
June	-0.46**	0.21	-0.65***	0.44***	-0.01	0.00	-0.24*	0.22
July	-0.56***	0.32**	-0.54***	0.36**	-0.26	-0.13	-0.42***	0.01
August	-0.08	-0.04	-0.30**	0.07	0.03	0.24*	-0.37***	0.08
September	-0.39*	-0.39*	-0.42***	-0.25*	-0.18	-0.35	-0.34***	-0.10
October	-0.13	-0.13	-0.14	-0.04	-0.18	-0.34**	-0.22	-0.15
Jan. Feb. March	0.48***	-0.20	-0.05	0.18	0.47*	-0.11	0.32**	0.15
Apr. May June	-0.15	0.53***	-0.38***	0.75***	-0.02	0.49*	-0.23	0.55***
June Aug. Sept.	-0.42**	-0.34	-0.51*	-0.14	-0.18	-0.30	-0.45***	-0.05
from Nov(t-1) to Oct	0.04	-0.11	-0.37*	0.25*	0.15	0.01	-0.21*	0.17

<sup>448</sup> 

The temporal trends in  $t_0$  and  $t_1$ , both toward earlier dates (-10 and -26 days, respectively) offset each other to some extent until 2002 and resulted in a non-significant trend in  $\Delta t_{t0-t1}$ . After 2002, the prolonged trend in  $t_1$  continued to decrease while  $t_0$  stabilized, thereby leading to a sudden shortening of the growing season (Fig. 5). The timing of this phenological reduction was concomitant with the appearance of a significant relationship between T<sub>JFM</sub> and 454 BAI (Fig. 6b). The decline in  $\Delta t_{t0-t1}$  induced by the winter warming pause resulted in a decrease 455 in BAI<sub>RW</sub> after 2002. This decreasing trend over time become significant after the 10-year time 456 window 1997-2006 (Fig. 7a,b). BAI<sub>FS</sub> exhibited a similar decreasing trend in the same period, 457 but of lower amplitude and not significant.

458



459 460 461 462 463

**Fig. 5** Onset of phenological stem growth ( $t_0$ , pale grey inverted triangles) and end of stem growth ( $t_1$ , grey circles) from 1966 to 2013, expressed as day of the year (DOY). The effect of the warming pause on  $t_0$  is shown by the hatched area between the points and the regression line over the 1978-1998 period.



464

465

466 **Fig. 6** Temporal variations in the bootstrapped Pearson's correlation coefficients between (a)  $BAI_{RW}$  and  $P_{AMJ}$ , 467 and (b)  $BAI_{RW}$  and  $T_{JFM}$  over 10-year moving windows. The black curve represents the correlation coefficient and 468 the grey curves the upper and lower confidence intervals at a significance level  $\alpha = 0.05$ . Asterisks identify the 469 time windows with significant correlations.

470



471Fig. 7 Temporal trend over 10-year moving windows of (a) mean annual stem growth from 1968 to 2008 and from4721987 to 2013 for ring width (RW, dark grey circles) and field surveys (FS, grey squares), respectively. Error bars473are standard errors. (b) Slopes of the linear trends of BAI for the two series. The asterisks indicate time windows474with a significant trend (p<0.05). The dotted vertical lines in both panels indicate the first time window with a</td>475significant relationship between BAI<sub>RW</sub> and T<sub>JFM</sub> after the warming pause.

476

# 477 **DISCUSSION**

#### 478 Growth duration as a predictor of annual stem growth

479 The duration of spring stem growth ( $\Delta t_{t0-t1}$ ) was the best explanatory variable of yearly basal area increment for both BAI<sub>FS</sub> and BAI<sub>RW</sub> after 1984. This confirms earlier results 480 481 obtained with automatic dendrometers over a shorter period and a smaller sample of trees 482 (Lempereur *et al.*, 2015), and therefore validates  $\Delta t_{t0-t1}$  as a robust estimator of Q. *ilex* annual 483 stem growth over long time scales and large samples. The two independent samplings we used 484 here differed markedly in the number of trees measured, the measurement method and its 485 precision. FS corresponds to a large number of trees (125 trees) representing a broader tree size 486 distribution but measured with a less precise resolution (0.3 mm of DBH), while RW was 487 measured on a small sample of large dominant trees (12 trees) but with a better resolution 488 (0.01 mm of radius). Average stem growth was lower for  $BAI_{FS}$  than for  $BAI_{RW}$ , but both 489 responded linearly and with the same sensitivity to  $\Delta t_{t0-t1}$  (Fig. 3).

490 The spring stem growth duration  $\Delta t_{t0-t1}$  is defined as the length of the period between the 491 onset of spring stem growth  $(t_0)$  and the date of growth cessation caused by summer drought 492  $(t_1)$ . Lempereur *et al.* (2015) observed that  $t_0$  was well predicted by winter temperature, and that 493  $t_1$  occurred when the leaf predawn water potential dropped below a threshold of -1.1 MPa. A 494 similar approach was used by Rossi et al. (2011, 2014) who linked the stem growth of boreal 495 black spruce with the duration of xylem growth delimited by cold temperatures in spring and 496 autumn, and by Delpierre et al. (2016a) who linked the growth of temperate oak to the timing 497 of the summer growth cessation controlled by water limitation. By identifying the climatic 498 limits for cell division and elongation, the temporal delimitation of tree growth duration is an 499 innovative way of improving simulation of forest carbon sink response to climate change 500 (Fatichi et al., 2014). Tree growth has generally been simulated using two types of models: statistical dendroclimatological models, which are empirical and may lead to marked 501 502 uncertainties when extrapolated to future climate scenarios (e.g. Gea-Izquierdo et al., 2013; 503 Babst et al., 2014); and process based models that link tree growth to carbon assimilation (e.g. 504 Sitch et al., 2003; Friedlingstein et al., 2006; Gaucherel et al., 2008; Keenan et al., 2011), 505 ignoring experimental evidence showing a direct, and generally more limiting, climate control 506 of plant growth (Körner, 2003; Muller et al., 2011). Projecting tree growth based on growth 507 phenology as in Rossi et al. (2011) and Lempereur et al. (2015) merges the biological realism 508 of the first approach, i.e. a direct link between climate and growth, with the mechanistic 509 understanding of tree physiology used in process based models, and be a significant step 510 forward in projecting the response of forest ecosystems to climate change (Delpierre et al., 511 2016b). Showing that growth duration is an adequate predictor of tree growth for long time 512 series datasets (40 years in our study) and coarse temporal resolution (annual diameter 513 measurements) is thus an important preliminary.

514 Nevertheless, the spring growth duration approach remains limited by both its conditions of 515 applicability and the large proportion of the growth variability not explained by  $\Delta t_{t0-t1}$ . The 516 particular case of the year 1992, with its bi-modal drought in spring and autumn and its well-517 watered summer (Fig. S6), suggests that  $\Delta t_{t0-t1}$  can only be used as a descriptor of stem growth 518 phenology in years with typical Mediterranean seasonality and a drought period occurring in 519 summer. While the lowest threshold of predawn water potential below which stem growth is 520 prevented certainly remains valid for any drought seasonality (Hsiao & Xu, 2000; Muller et al., 521 2011; Lempereur et al., 2015), we hypothesize that the relationship between stem growth and

522 the duration of periods with adequate growth conditions (the rate of growth) varies in years 523 with peculiar phenology. Variations in the rate of growth may also account for the unexplained 524 variability in the BAI against  $\Delta t_{t0-t1}$  relationships (44% and 65% for RW and FS, respectively; 525 Fig. 3). Finally, the relative extent of autumn growth compared to spring growth would deserve 526 more detailed investigation. Lempereur et al. (2015) observed that autumn growth explained 527 only approximately 30% of annual stem growth and was strongly determined by spring 528 conditions, but its magnitude over long time periods and under climate change conditions 529 remains to be described.

530

## 531 Dual limitation of growth by winter cold and spring-summer drought

532 Summers in the Mediterranean region are typically characterized by the concomitance of 533 low rainfall, high temperature and high solar radiation, while winters are cold and humid. 534 Consequently, vegetation functioning is limited by water deficit in summer and by low 535 temperatures in winter (Mitrakos, 1980; Terradas & Savé, 1992). A cessation of growth during 536 cold and dry periods is frequently observed in Mediterranean trees which display a bi-phasic 537 growth pattern over the year (Cherubini *et al.*, 2002; Campelo *et al.*, 2007; Montserrat-Marti *et 538 al.*, 2009; Camarero *et al.*, 2009, Gutiérrez *et al.*, 2011).

The sensitivity of growth to water deficit is well established in the literature (Lockhart 539 540 1965; Hsiao & Xu, 2000; Muller et al., 2011), as is the negative impact of drought on ring width (Fritts, 1976). Previous dendroclimatological studies generally stated that spring to early-541 542 summer precipitation was the main driver of annual stem growth of Q. ilex (Gutiérrez et al., 543 2011). More precisely, it is the start of the dry season, defined by a threshold of soil water 544 deficit that was identified as the main determinant of variations in inter-annual growth (Maselli 545 et al., 2014; Lempereur et al., 2015). The predominant effect of the timing of drought induced 546 growth cessation on annual stem growth is not limited to Mediterranean ecosystems, as a similar 547 effect has been observed in mesic temperate forests (Brzostek et al., 2014; Delpierre et al., 548 2016a). In our study, we defined the start of the dry season as a critical threshold of predawn 549 plant water potential of -1.1 MPa. This threshold is biologically meaningful because water 550 potential affects the cell turgor pressure necessary for cell growth (Hsiao & Xu, 2000), and it 551 has the other advantage of comprehensively accounting for climate and local soil conditions in 552 a single metric (Ruffault et al., 2013). It is, however, a more complex index to calculate than 553 climatic indices of drought as it requires substantial knowledge of soil and vegetation 554 characteristics, and accurate meteorological data at the daily time scale. Consequently, the

uncertainty on  $t_1$  was higher than the uncertainty on  $t_0$ , especially when meteorological data from SML (located at a distance of 12 km from the study site) was used instead of on-site data, but  $t_1$  nonetheless remained a good predictor of annual stem growth for whichever period (Table 2).

559 Depending on the species and on the local bio-climate, a minimum temperature 560 threshold ranging from +4 °C to +7 °C is required for stem growth to occur (Körner, 2003; 561 Rossi et al., 2007; Deslauriers et al., 2008; Gruber et al., 2010; Swidrak et al., 2011; Lempereur 562 et al., 2015). The onset of stem growth can be assessed from winter temperature (Delpierre et 563 al., 2016b), thus the mean winter temperature directly impacts the duration of cambium activity 564 and wood formation (Rossi et al., 2011). However, there is no consensus in 565 dendrochronological studies of the Mediterranean Q. ilex that winter temperature is a good 566 predictor of annual BAI, and both negative (Zhang & Romane 1990; Paton et al., 2009; Gea-567 Izquierdo et al., 2009) and positive correlations (Campelo et al., 2009; Nijland et al., 2011; 568 Gea-Izquierdo et al., 2011) between annual stem growth and winter temperature have been 569 reported. In our study, a positive correlation between growth and winter temperature was 570 observed over the period 1984-2008 (r = 0.48, P<0.001; Table 3), in accordance with 571 Lempereur et al. (2015), but the relationship was not significant for the period 1968-2008, either 572 with t<sub>0</sub> or with T<sub>JFM</sub>.

Taken together, these results suggest that summer drought is the main limiting factor for *Q. ilex* growth, which is to be expected under the dry Mediterranean climate. Nevertheless, the contrasted growth responses to winter temperature depending on the study concerned, or on the period of interest, led us to investigate the impact of recent climate change on *Q. ilex* growth in more detail.

578

# 579 Can warmer winter temperature compensate for earlier summer drought under climate 580 change?

From 1968 to 2008, the annual and spring amounts of precipitation were stable (Fig. 4b), but water limitation increased (Fig. 4d) due to increasing potential evapotranspiration with rising temperature in spring and summer (Fig. 4a and c), in accordance with regional observations (Ruffault *et al.*, 2013). Consequently, we simulated an earlier occurrence of drought onset ( $t_1$ , -26 days on average) along with an earlier growth onset, although to a lesser extent ( $t_0$ , -10 days; Fig. 5), which, taken together resulted in a non-significant decrease in growth duration. The positive effect of warming on ecosystem functioning and tree growth 588 through a longer growing season has been widely observed in temperature limited boreal and 589 temperate forests (Keeling et al., 1996; Menzel et al., 2006; D'Arrigo et al., 2008; Dragoni et 590 al., 2011). However, in water limited regions, like the Mediterranean, warming is generally 591 considered to be an aggravating factor for drought, mainly because of increased evaporative 592 demand (Angert et al., 2005; Zhao & Running, 2010; Park et al., 2012). Consequently, climate 593 warming in the Mediterranean usually reduces tree growth (Jump et al., 2006; Sarris et al., 594 2007; Peñuelas et al., 2008; Piovesan et al., 2008; Martin-Benito et al., 2010). Our results thus 595 mitigate this widely accepted conclusion and illustrate the peculiar parallel controls driven by 596 temperature, such that the benefit of an earlier stem growth is cancelled out by earlier drought 597 onset mostly caused by increasing evaporative demand.

598 The effect of temperature on t<sub>1</sub>, mediated by PET, may also explain why the correlation 599 with growth was better for  $t_1$  than for  $P_{AMJ}$  over the period 1984-2008, when temperatures 600 increased significantly, but not over the longer period 1968-2008. Actually, spring PET, like 601 temperature, did not increase significantly until the early 1980s (Fig. 4a, c), suggesting that, in 602 the past, inter-annual variability in precipitation may have been a stronger driver of drought 603 onset. Alternatively, the lack of on-site precipitation measurements before 1984 may mean that 604 t<sub>1</sub> estimates based on SML are less closely linked to on-site conditions than three-month 605 cumulated precipitation. The correlation in precipitation amounts between Puéchabon and SML 606 actually increases with longer temporal resolutions.

Winter warming may also have a positive impact on stem growth by delaying growth cessation in autumn, thereby partly compensating for earlier drought onset. However, the phenology of autumn growth cessation appears to be less variable than that of  $t_0$  (Lempereur *et al.*, 2015), possibly because it is concurrently driven by photoperiod. Moreover, warm winters may even impact *Q. ilex* growth negatively if the species requires winter chilling, as is the case of the deciduous *Quercus* species (Fu *et al.*, 2015).

613

# 614 A Mediterranean "divergence problem"

When looking at the temporal variations in the correlations between stem growth and climate variables, we observed an abrupt and significant increase in the sensitivity of stem growth to temperature in the early 2000s, while at the same time, its response to precipitation weakened (Fig. 6a, b). Temporal changes in the response of tree growth to climate have been observed in a wide range of climates and tree species in recent decades (e.g. Briffa *et al.*, 1998; Büntgen *et al.*, 2006; Carrer & Urbinati 2006; Jump *et al.*, 2007; D'Arrigo *et al.*, 2008). The 621 'divergence problem' in northern forests has been defined as the tendency for tree growth at 622 previously temperature limited sites to undergo a weakening of their temperature response 623 concurrent with an increasing sensitivity to drought (D'Arrigo et al., 2008). Our results suggest 624 a Mediterranean 'divergence problem' according to which tree growth in water limited 625 Mediterranean ecosystems undergo a weakening of their response to spring-summer 626 precipitation and an increasing sensitivity to winter temperatures. Similar reports of temporal 627 changes from water driven to increasingly temperature driven tree growth in water limited 628 ecosystems have already been reported for beech forests in northeast Spain (Jump et al., 2007), 629 black pine forests in Spain (Martin-Benito et al., 2010), and Scots pine, European larch and 630 black pine in Switzerland (Feichtinger et al., 2014). These observations differ from ours by 631 reporting an increase of the overall negative effect of warming on tree growth. Nevertheless, 632 they collectively point to the increasing influence of temperature on the growth of previously 633 water limited trees.

634

# 635 Concurrent increase in drought and warming: a keystone aspect of climate change revealed 636 during the recent warming hiatus

637 A keystone result of our study is the sudden significance of the growth-temperature relationship occurring from the 1995-2004 time window on (Fig. 6b). In parallel, we observed 638 639 lower winter temperatures than would have been expected on the basis of continuous climate 640 warming (Fig. 4a). This regional pattern is therefore comparable to the globally observed 641 "warming hiatus" (Easterling & Wehner 2009; Trenberth & Fasullo 2013). Our results suggest 642 that these recent cooler winter temperatures resulted in later growth onset and led to an 643 increased temperature constraint on the duration of stem growth. In the meantime, the constraint 644 exerted by the water deficit increased constantly from 1968 onward with neither breakpoints 645 nor changes in the trend line (Fig. 4d, 5). As a result, the stem growth of *Q*. *ilex* was significantly 646 correlated only with water deficit in the past, but the shorter growth period after the 2000s 647 revealed the dual control by winter temperatures and spring-summer water deficit. The pause 648 in climate warming in turn disrupted the precarious balance between increasing winter 649 temperatures and increasing spring-summer water deficit, which temporarily sustained stem 650 growth until the end of the 1990s (Fig. 7). This is, to our knowledge, the first example of a 651 'divergence problem' in the tree growth-climate relationship triggered not by a continuous 652 climate change but instead by the warming pause. This recent warming pause, a still 653 controversial aspect of climate change (Lewandowsky et al., 2015; Wehner & Easterling 2015),

654 is mainly caused by the variations in Atlantic and Pacific multidecadal oscillations (Steinman 655 et al., 2015), and produced globally heterogeneous patterns of breakpoint in the recent warming trend (Ying et al., 2015), with an enhanced effect on winter temperature in Eurasia (Li et al., 656 657 2015). It is however, likely to be reversed in the coming decades (Steinman et al., 2015). 658 Whether a future re-acceleration of climate warming in the Mediterranean would compensate 659 for increasing water deficit by phenological stimulation of earlier tree growth again, or on the 660 contrary, further exacerbate drought stress remains an open and important question. Studying trends in targeted and concurrent climate variables such as temperature and drought should fully 661 662 capture the complexity of climate change impacts (Mazdiyani & Aghakouchak 2015). Climate 663 projections for the Mediterranean region forecast an increase in potential evapotranspiration and a decrease in summer precipitation by the end of the 21<sup>th</sup> century (Gao & Giorgi 2008; 664 Ruffault *et al.*, 2014), which are likely to move the onset of drought forward more strongly than 665 666 the onset of spring growth (Lempereur et al., 2015). Together with the current growth limiting 667 winter temperatures, these future trends in drought features could lead to a sharp reduction in 668 forest productivity and an increase in tree mortality in Mediterranean Q. ilex forests.

669

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# 678 **REFERENCES**

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# 951 SUPPORTING INFORMATION

- 952 Fig. S1 Relationships used to rescale the ring width series to the tree DBH at the time of the
- 953 cut: tree tapering (base diameter minus DBH) against DBH, and bark thickness against DBH.
- **Table S1** Summary of climate data for Puéchabon and St Martin-de-Londres meteorological
  stations.
- 956 Fig. S2 Relationships between the meteorological data of Puéchabon and St Martin-de-Londres
- 957 for the overlap period 1984-2013.
- Fig. S3 Characteristic example of annual stem growth showing the main phenological eventsfor stem growth.
- 960 Fig. S4 Relationship between the spring onset of growth  $(t_0)$  and the January to March
- 961 temperature (T<sub>JFM</sub>) for the meteorological station of St Martin-de-Londres.
- 962 Fig. S5 Relationship between observed and predicted day of the year of growth cessation due
- 963 to drought  $(t_1)$  when simulating plant water potential with the meteorological data from St
- 964 Martin-de-Londres.
- 965 Fig. S6 Seasonal variations of simulated predawn leaf water in Puéchabon in 1992.
- 966 Fig. S7 Relationship between yearly basal area increment (BAI) and the duration of the spring
- basal growth period ( $\Delta t_{t0-t1}$ ) showing the years 2002 and 2010 that are outliers for the series BAI<sub>FS</sub>.
- Table S3 Results of the Theil-Sen tests for the trends of temperature, precipitation,
  evapotranspiration and drought index.

- 971 Fig. S8 Temporal trends in winter temperature for seven meteorological stations in the region
- 972 Languedoc-Roussillon.
- 973 Table S4 Breakpoints and changes in mean winter temperature for seven meteorological
- 974 stations in the region Languedoc-Roussillon.