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# ON DISTRIBUTIONS DETERMINED BY THEIR UPWARD, SPACE-TIME WIENER-HOPF FACTOR

LOÏC CHAUMONT AND RON DONEY

ABSTRACT. According to the Wiener-Hopf factorization, the characteristic function  $\varphi$  of any probability distribution  $\mu$  on  $\mathbb{R}$  can be decomposed in a unique way as

$$1 - s\varphi(t) = [1 - \chi_-(s, it)][1 - \chi_+(s, it)], \quad |s| \leq 1, t \in \mathbb{R},$$

where  $\chi_-(e^{iu}, it)$  and  $\chi_+(e^{iu}, it)$  are the characteristic functions of possibly defective distributions in  $\mathbb{Z}_+ \times (-\infty, 0)$  and  $\mathbb{Z}_+ \times [0, \infty)$ , respectively.

We prove that  $\mu$  can be characterized by the sole data of the upward factor  $\chi_+(s, it)$ ,  $s \in [0, 1)$ ,  $t \in \mathbb{R}$  in many cases including the cases where:

- 1)  $\mu$  has some exponential moments;
- 2) the function  $t \mapsto \mu(t, \infty)$  is completely monotone on  $(0, \infty)$ ;
- 3) the density of  $\mu$  on  $[0, \infty)$  admits an analytic continuation on  $\mathbb{R}$ .

We conjecture that any probability distribution is actually characterized by its upward factor. This conjecture is equivalent to the following: *Any probability measure  $\mu$  on  $\mathbb{R}$  whose support is not included in  $(-\infty, 0)$  is determined by its convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$ .* We show that in many instances, the sole knowledge of  $\mu$  and  $\mu^{*2}$  restricted to  $[0, \infty)$  is actually sufficient to determine  $\mu$ . Then we investigate the analogous problem in the framework of infinitely divisible distributions.

## 1. INTRODUCTION

Let  $\mu$  be any probability measure on  $\mathbb{R}$ . Denote by  $(S_n)$  a random walk with step distribution  $\mu$ , such that  $S_0 = 0$ , a.s. Define the first ladder times associated to  $(S_n)$  by

$$\tau_- = \inf\{n \geq 1 : S_n < 0\}, \quad \tau_+ = \inf\{n \geq 1 : S_n \geq 0\}.$$

Then the Wiener-Hopf factorization of the characteristic function  $\varphi(t) = \int e^{itx} \mu(dx)$  of  $\mu$  can be written as,

$$(1.1) \quad 1 - s\varphi(t) = [1 - \chi_-(s, it)][1 - \chi_+(s, it)], \quad |s| \leq 1, t \in \mathbb{R},$$

where  $\chi_-$  and  $\chi_+$  are the downward and upward space-time Wiener-Hopf factors,

$$\chi_-(s, it) = \mathbb{E}(s^{\tau_-} e^{itS_{\tau_-}} 1_{\{\tau_- < \infty\}}) \quad \text{and} \quad \chi_+(s, it) = \mathbb{E}(s^{\tau_+} e^{itS_{\tau_+}} 1_{\{\tau_+ < \infty\}}).$$

To paraphrase W. Feller [3], XVIII.3, *the remarkable feature of the factorization (1.1) is that it represents an arbitrary characteristic function  $\varphi$  in terms of two (possibly defective) distributions, one being concentrated on the half line  $(-\infty, 0)$  and the other one on the half line  $[0, \infty)$ .* However, this feature only exploits identity (1.1) for fixed  $s \neq 0$  and reflects the fact that  $\mu$  is determined by the knowledge of the distributions of both  $S_{\tau_-}$  and  $S_{\tau_+}$ . But one may wonder about the extra information brought by the joint distributions  $(\tau_-, S_{\tau_-})$  and  $(\tau_+, S_{\tau_+})$ . In particular, is it true in general that  $\mu$  is determined by only one of these joint distributions? or equivalently, is it true that  $\varphi$  is determined by only

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one of its space-time Wiener-Hopf factors?

The aim of this paper is an attempt to answer the latter question. We will actually show that  $\mu$  is determined by  $\chi_+(s, it)$  in some quite large classes of distributions including the case where  $\mu$  has some positive exponential moments, or when  $t \mapsto \mu(t, \infty)$  is completely monotone on  $(a, \infty)$ , for some  $a \geq 0$ , or satisfies some property which is slightly stronger than analyticity. Obviously all these assumptions can be verified from the sole data of  $\chi_+(s, it)$ . These different cases cover a sufficiently large range of distributions for us to allow ourselves to raise the following conjecture. Let  $\mathcal{M}_1$  be the set of probability measures on  $\mathbb{R}$ .

**Conjecture C.** *Any distribution  $\mu \in \mathcal{M}_1$  whose support is not included in  $(-\infty, 0)$  is determined by its upward space-time Wiener-Hopf factor  $\chi_+(s, it)$ ,  $|s| < 1$ ,  $t \in \mathbb{R}$ .*

A crucial step in the proof of (1.1) is the following development of the factor  $\chi_+(s, t)$ , for  $|s| < 1$  and  $t \in \mathbb{R}$ ,

$$(1.2) \quad \log \frac{1}{1 - \chi_+(s, it)} = \sum_{n=1}^{\infty} \frac{s^n}{n} \int_{[0, \infty)} e^{itx} \mu^{*n}(dx),$$

see [3], XVIII.3, where  $\mu^{*n}$  is the  $n$ th fold convolution product of  $\mu$  by itself. We will actually refer to  $\mu^{*n}$ ,  $n \geq 0$  as the convolution powers of  $\mu$ . This identity shows that the data of  $\chi_+$  is equivalent to the knowledge of the measures  $\mu^{*n}$ ,  $n \geq 1$  on  $[0, \infty)$  and leads to the following equivalent conjecture.

**Conjecture C'.** *Any distribution  $\mu \in \mathcal{M}_1$  whose support is not included in  $(-\infty, 0)$  is determined by its convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$ .*

Each of the next sections corresponds to a class of probability distributions for which Conjecture C holds. For the first one in Section 2, we prove that distributions having some particular exponential moments satisfy conjecture C. Then in Section 3 we consider three other classes for which a much stronger result than Conjectures C and C' is true. We will see that there are actually many distributions which are determined by the sole data of  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ . This is the case when the function  $t \mapsto \mu(t, \infty)$  is smooth enough. In Subsection 3.2 we will consider the case where the function  $t \mapsto \mu(t, \infty)$  is completely monotone on  $(a, \infty)$ , for some  $a \geq 0$  and in Subsection 3.3 we will make a slightly stronger assumption than analyticity on this function. We will also present the discrete counterpart of the completely monotone case in Subsection 3.4. Finally in Section 4 we will consider Conjecture C in the restricted class of infinitely divisible distributions and show that if the upper tail of the Lévy measure is completely monotone, then  $\mu$  is determined by its upper Wiener-Hopf factor. Then we end this paper in Section 5 with some important remarks on the possibility of extending the classes of distributions studied.

In all this paper, we will denote by  $\mathcal{C}$  the set of distributions satisfying Conjecture C. Let us give a proper definition of this set.

**Definition 1.1.** *Let  $\mathcal{C}$  be the set of distributions  $\mu \in \mathcal{M}_1$  which are determined by their upward Wiener-Hopf factor  $\chi_+(s, it)$ , for  $|s| \leq 1$  and  $t \in \mathbb{R}$  or equivalently by the data of their convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$ . More formally,*

$$\mathcal{C} = \{\mu \in \mathcal{M}_1 : \text{if } \mu_1 \in \mathcal{M}_1 \text{ satisfies } \mu^{*n} = \mu_1^{*n}, \text{ on } [0, \infty), \text{ for all } n \geq 0, \text{ then } \mu = \mu_1\}.$$

The problem we investigate here originates from a result in V. Vigon's PhD thesis [9], see Section 4.5 therein, where a question equivalent to Conjecture C is raised in the setting of Lévy processes. Our question is actually more general since it concerns distributions which are not necessarily infinitely divisible. In particular, a positive answer to Conjecture C would imply that the law of any Lévy process  $(X_t, t \geq 0)$  is determined by one of its space-time Wiener-Hopf factors or equivalently by the marginals of the process  $(X_t^+, t \geq 0)$ , see Section 4.

## 2. WHEN $\mu$ ADMITS EXPONENTIAL MOMENTS.

**2.1. Recovering characteristic function and moment generating function.** In this paper, we will always assume that the support of the measure  $\mu$  is not included in  $(-\infty, 0)$ . Let us observe that from the data of the upward Wiener-Hopf factor  $\chi_+(s, it)$ , for  $|s| \leq 1$  and  $t \in \mathbb{R}$  or equivalently from the data of the measures  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$ , we know the sequences,  $\mathbb{P}(S_n < 0)$  and  $\mathbb{P}(S_n \geq 0)$ ,  $n \geq 0$ , as well as the distributions of both  $\tau_-$  and  $\tau_+$ . In particular we know whether  $(S_n)$  oscillates, drifts to  $-\infty$ , or drifts to  $\infty$ . The next result shows that provided  $n \mapsto \mathbb{P}(S_n < 0)$  tends to 0 sufficiently fast along some subsequence, it is possible to recover the characteristic function  $\varphi$  of  $\mu$  on some interval containing 0, from the measures  $\mu^{*n}$  restricted to  $[0, \infty)$ .

**Lemma 2.1.** *Assume that there is  $\alpha > 0$  such that, at least for a subsequence,*

$$(2.3) \quad \mathbb{P}(S_n < 0) \leq e^{-\alpha n}.$$

*Then for all  $t$  such that  $|\varphi(t)| > e^{-\alpha}$ , along this subsequence,*

$$\lim_{n \rightarrow +\infty} \mathbb{E}(e^{itS_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} = \varphi(t).$$

*In particular, if (2.3) holds then  $\varphi$  can be determined on some neighborhood of 0.*

*Proof.* It suffices to write,

$$(2.4) \quad \begin{aligned} \mathbb{E}(e^{itS_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} &= [\varphi(t)^n - \mathbb{E}(e^{itS_n} \mathbb{I}_{\{S_n < 0\}})]^{1/n} \\ &= \varphi(t)[1 - \varphi(t)^{-n} \mathbb{E}(e^{itS_n} \mathbb{I}_{\{S_n < 0\}})]^{1/n}. \end{aligned}$$

Then from the assumption, for all  $t$  such that  $|\varphi(t)| > e^{-\alpha}$  and for all  $n$  such that  $\mathbb{P}(S_n < 0) < e^{-\alpha n}$ ,

$$\begin{aligned} |\varphi(t)^{-n} \mathbb{E}(e^{itS_n} \mathbb{I}_{\{S_n < 0\}})| &\leq |\varphi(t)^{-n} \mathbb{P}(S_n < 0)|, \\ &\leq (|\varphi(t)|e^\alpha)^{-n}. \end{aligned}$$

Therefore, the left hand side of the above inequality tends to 0 along a subsequence and this yields the result thanks to equation (2.4).  $\square$

Since, for all  $n \geq 0$ ,

$$\mathbb{P}(S_1 < 0, S_2 - S_1 < 0, \dots, S_n - S_{n-1} < 0) = \mathbb{P}(S_1 < 0)^n \leq \mathbb{P}(S_n < 0),$$

(2.3) cannot hold for all  $\alpha > 0$ , unless  $\mathbb{P}(S_1 \geq 0) = 1$ . Note also that if (2.3) holds then the random walk  $(S_n)$  cannot drift to  $-\infty$ . Moreover, if it holds along all subsequences, then  $(S_n)$  necessarily drifts to  $\infty$  thanks to Spitzer's criterion which asserts that this happens if and only if  $\sum n^{-1} \mathbb{P}(S_n < 0) < \infty$ .

Let us denote the moment generating function of  $\mu$  by  $\phi(\lambda) := \int_{\mathbb{R}} e^{\lambda x} \mu(dx)$ ,  $\lambda \in \mathbb{R}$ .

**Lemma 2.2.** *For  $\lambda < 0$ ,*

1.  $\phi(\lambda) < \infty$  *if and only if*  $\mathbb{E}(e^{\lambda S_{\tau_-}} \mathbb{I}_{\{\tau_- < \infty\}}) < \infty$ .
2. *If*  $\phi(\lambda) < 1$ , *then (2.3) holds.*

*Proof.* Assume that  $\mathbb{E}(e^{\lambda S_{\tau_-}} \mathbb{I}_{\{\tau_- < \infty\}}) < \infty$ . Then we derive from the following inequality,

$$\begin{aligned} \phi(\lambda) &= \mathbb{E}(e^{\lambda S_1} \mathbb{I}_{\{\tau_- = 1\}}) + \mathbb{E}(e^{\lambda S_1} \mathbb{I}_{\{\tau_- > 1\}}) \\ &\leq \mathbb{E}(e^{\lambda S_{\tau_-}} \mathbb{I}_{\{\tau_- < \infty\}}) + \mathbb{P}(\tau_- > 1), \end{aligned}$$

that  $\phi(\lambda) < \infty$ . Conversely, assume that  $\phi(\lambda) < \infty$ , and write for  $x > 0$ ,

$$\begin{aligned} \mathbb{P}(-S_{\tau_-} > x, \tau_- < \infty) &= \sum_{k \geq 1} \mathbb{P}(S_1 \geq 0, \dots, S_{k-1} \geq 0, -X_k > x + S_{k-1}) \\ &= \sum_{k \geq 1} \int_{[0, \infty)} \mathbb{P}(S_1 \geq 0, \dots, S_{k-2} \geq 0, S_{k-1} \in dy) \mathbb{P}(-S_1 > x + y) \\ (2.5) \quad &= \int_{[0, \infty)} v(dy) \mathbb{P}(-S_1 > x + y), \end{aligned}$$

where  $v(dy) = \sum_{k \geq 0} \mathbb{P}(S_1 \geq 0, \dots, S_{k-1} \geq 0, S_k \in dy)$  is the renewal measure on  $[0, \infty)$  of the (weak) upward ladder height process of  $(S_n)$ , see Chap. XII.2 in [3]. In particular, this measure satisfies  $\int_{[0, \infty)} e^{\alpha y} v(dy) < \infty$  for all  $\alpha < 0$ . Then from (2.5) and the assumption, we obtain

$$\begin{aligned} \mathbb{E}(e^{\lambda S_{\tau_-}} \mathbb{I}_{\{\tau_- < \infty\}}) &= \int_{[0, \infty)} e^{\lambda y} v(dy) \int_y^\infty e^{\lambda z} \mathbb{P}(S_1 \in dz) \\ &\leq \int_{[0, \infty)} e^{\lambda y} v(dy) \phi(\lambda) < \infty. \end{aligned}$$

Finally in order to prove assertion 2, it suffices to write,

$$\mathbb{P}(S_n < 0) \leq \mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n < 0\}}) \leq \phi(\lambda)^n,$$

so that if  $\phi(\lambda) < 1$ , then (2.3) is satisfied. □

**Theorem 2.1.** *Assume that we know the measures  $\mu^{*n}$  restricted to  $[0, \infty)$ . Then we can determine if (2.3) holds. Moreover if (2.3) holds, then we can determine if*

$$(2.6) \quad \text{there is } \lambda < 0 \text{ such that } \phi(\lambda) < \infty.$$

*When both (2.3) and (2.6) hold, the measure  $\mu$  belongs to the class  $\mathcal{C}$ .*

*Proof.* The first assertion is trivial since  $1 - \mu^{*n}[0, \infty) = \mathbb{P}(S_n < 0)$ . Assume that (2.3) holds. Then from Lemma 2.1, there is  $\varepsilon > 0$  such that the characteristic function  $\varphi(t)$  is known for  $t \in [-\varepsilon, \varepsilon]$ . Recall from Section 1 that the knowledge of the measures  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  is equivalent to the knowledge of  $\chi_+(s, it)$ , for all  $|s| \leq 1$  and  $t \in \mathbb{R}$ .

Therefore, from (1.1), we know  $\chi_-(1, it)$ , for  $t \in [-\varepsilon, \varepsilon]$ . Note that  $t \mapsto \chi_-(1, it)/\chi_-(1, 0)$  is the characteristic function of the random variable  $S_{\tau_-}$  under  $\mathbb{P}(\cdot | \tau_- < \infty)$ . Since this random variable is negative, the knowledge of its characteristic function on  $[-\varepsilon, \varepsilon]$  allows us to determine if the moments  $m_n = \mathbb{E}((-S_{\tau_-})^n | \tau_- < \infty)$  are finite for all  $n$  and if there is  $\lambda < 0$  such that  $\mathbb{E}(e^{\lambda S_{\tau_-}} | \tau_- < \infty) = \sum_{n \geq 0} \frac{m_n}{n!} (-\lambda)^n < \infty$ .

We have proved that, the knowledge of the characteristic function  $t \mapsto \chi_-(1, it)/\chi_-(1, 0)$  on  $[-\varepsilon, \varepsilon]$  allows us to determine if  $\mathbb{E}(e^{\lambda S_{\tau_-}} \mathbb{I}_{\{\tau_- < \infty\}}) < \infty$  for some  $\lambda < 0$ , which is equivalent to  $\phi(\lambda) < \infty$ , from Lemma 2.2. When this is the case, the knowledge of the characteristic function  $t \mapsto \chi_-(1, it)/\chi_-(1, 0)$  on  $[-\varepsilon, \varepsilon]$  is enough to determine the law of the random variable  $S_{\tau_-}$  under  $\mathbb{P}(\cdot | \tau_- < \infty)$ . Then from (1.1) we can determine the law  $\mu$ , so that it belongs to class  $\mathcal{C}$ .  $\square$

**Remark 2.1.** *Actually, Lemma 2.1 might be exploited in a better way. Indeed, assume that  $\mu$  satisfies (2.3). Then the characteristic function  $\varphi$  is known on some neighborhood of 0. If for instance  $\varphi$  is infinitely differentiable around 0, then all the moments of  $\mu$  can be recovered. If moreover,  $\mu$  is determined by these moments, then it belongs to class  $\mathcal{C}$ . In particular  $\mu$  does not need to have exponential moments as in Theorem 2.1. The question of finding such an example is still open.*

Recall that  $\phi$  is a convex function on the interval  $\{\alpha : \phi(\alpha) < \infty\}$ . Moreover, since the support of  $\mu$  is not included in  $(-\infty, 0)$ ,  $\phi$  is nondecreasing. If  $\lambda \in \mathbb{R}$  is such that  $\lambda = \inf\{\alpha : \phi(\alpha) < \infty\}$ , then  $\phi'(\lambda)$  will be understood as the right derivative of  $\phi$  at  $\lambda$ . Similarly, if  $\lambda = \sup\{\alpha : \phi(\alpha) < \infty\}$ , then  $\phi'(\lambda)$  will be the left derivative of  $\phi$  at  $\lambda$ .

**Lemma 2.3.** *For all  $\lambda \in \mathbb{R}$  such that  $\phi(\lambda) < \infty$  and  $\phi'(\lambda) > 0$ ,*

$$\lim_{n \rightarrow +\infty} \mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} = \phi(\lambda).$$

*Proof.* Let  $(S_n^{(\lambda)})$  be a random walk with step distribution  $\mu_\lambda(dx) := \frac{e^{\lambda x}}{\phi(\lambda)} \mu(dx)$ . Since

$$\mathbb{E}(S_1^{(\lambda)}) = \int_{\mathbb{R}} \frac{x e^{\lambda x}}{\phi(\lambda)} \mu(dx) = \frac{\phi'(\lambda)}{\phi(\lambda)} > 0,$$

the random walk  $(S_n^{(\lambda)})$  drifts to  $\infty$ , so that  $\lim_{n \rightarrow \infty} \mathbb{P}(S_n^{(\lambda)} \geq 0) = 1$ . Then the result follows from the identity

$$\mathbb{P}(S_n^{(\lambda)} \geq 0) = \frac{\mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})}{\phi(\lambda)^n}.$$

$\square$

The following theorem shows that distributions having some negative exponential moments less than 1 or some positive exponential moments bigger than 1 belong to class  $\mathcal{C}$ .

**Theorem 2.2.** *The knowledge of the measures  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  allows us to determine if  $\phi$  satisfies one of the two following conditions:*

- (a) *There exists  $\lambda < 0$  such that  $\phi(\lambda) < 1$ .*
- (b) *There exists  $\lambda > 0$  such that  $\phi(\lambda) \in (1, \infty)$ .*

*When at least one of these two conditions holds, the measure  $\mu$  belongs to the class  $\mathcal{C}$ .*

*Proof.* From Lemma 2.2, if (2.3) does not hold, then (a) is not satisfied. Assume that (2.3) holds. From Theorem 2.1, we can determine if there is  $\lambda < 0$  such that  $\phi(\lambda) < \infty$ . If this is not the case, then clearly (a) does not hold. Assume that there is  $\lambda < 0$  such that  $\phi(\lambda) < \infty$ . From Theorem 2.1, the measure  $\mu$  belongs to class  $\mathcal{C}$  so that we can determine if (a) holds.

From our data, for all  $\lambda > 0$  and  $n \geq 1$ , the expression  $\mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})$  is known. Assume that there is  $\lambda > 0$  such that  $\lim_{n \rightarrow +\infty} \mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} > 1$ . Since  $\mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} \leq \phi(\lambda)$ , we have actually  $\phi(\lambda) > 1$ . Moreover, our data clearly allows us to know if  $\phi(\lambda) <$

$\infty$ . Then since  $\phi(0) = 1$ , by convexity of the function  $\lambda \mapsto \phi(\lambda)$ , it is clear that  $\phi'(\lambda) > 0$ , so that from Lemma 2.3,  $\lim_{n \rightarrow +\infty} \mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} = \phi(\lambda)$ . This means that from our data, we can determine if there is  $\lambda > 0$  such that  $\phi(\lambda) \in (1, \infty)$  and in this case  $\lim_{n \rightarrow +\infty} \mathbb{E}(e^{\lambda S_n} \mathbb{I}_{\{S_n \geq 0\}})^{1/n} = \phi(\lambda)$ , so that  $\phi(\lambda)$  is known. Moreover, from the continuity of  $\phi$  on the set  $\{x : \phi(x) < \infty\}$ , there is an interval  $I$  containing  $\lambda$  such that for all  $x \in I$ ,  $\phi(x) \in (1, \infty)$  and from the same reasoning as above,  $\phi(x)$  is known for all  $x \in I$ , so that the measure  $\mu$  is determined and we conclude that it belongs to the class  $\mathcal{C}$ .  $\square$

We can easily check that condition (b) of Theorem 2.2 is satisfied in the two following situations:

(b<sub>1</sub>)  $\phi(\lambda) < \infty$  for all  $\lambda > 0$ .

(b<sub>2</sub>)  $\mu$  is absolutely continuous in  $[0, \infty)$  and its density  $f$  satisfies  $\ln(f(x)) \sim -\lambda_0 x$ , as  $x \rightarrow \infty$ , for some  $\lambda_0 \in (0, \infty)$ .

Indeed, in case (b<sub>1</sub>), if  $\mu \neq \delta_0$  then since  $\phi$  is a nondecreasing convex function such that  $\phi(0) = 1$  and since the support of  $\mu$  is not included in  $(-\infty, 0)$ ,  $\lim_{\lambda \rightarrow \infty} \phi(\lambda) = \infty$ . In case (b<sub>2</sub>) it is clear that  $\lim_{\lambda \rightarrow \lambda_0^-} \phi(\lambda) = \infty$ . Checking condition (a) in Theorem 2.2, and in particular checking if  $\mu$  has negative exponential moments under condition (2.3) requires more work on the behavior of the moments of  $S_{\tau_-}$ , see the proof of Theorem 2.1.

**2.2. Skip free distributions.** A distribution  $\mu$  whose support is included in  $\mathbb{Z}$  is said to be downward (resp. upward) skip free if  $\mu(n) = 0$  for all  $n \leq -2$  (resp. for all  $n \geq 2$ ). Clearly skip free distributions possess exponential moments. Moreover, upward skip free distributions belong to class  $\mathcal{C}$  from Theorem 2.2 (b) and the note following its proof. Then in this subsection we shall see that the case of downward skip free distributions allows us to go a little beyond the cases encompassed by Theorems 2.1 and 2.2. We first need to make sure that our data allows us determine if the support of a distribution is included in  $\mathbb{Z}$ .

**Lemma 2.4.** *The support of the measure  $\mu$  is included in  $\mathbb{Z}$  if and only if the support of the measures  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  is included in  $\mathbb{Z}_+$ .*

*Proof.* The direct implication is obvious. Then assume that the support of  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  is included in  $\mathbb{Z}_+$ , whereas the support of  $\mu$  restricted to  $(-\infty, 0]$  is not included in  $\mathbb{Z}_-$ . Then there is an interval  $I \subset (-\infty, 0] \setminus \mathbb{Z}_-$  such that  $\mu(I) > 0$ . Let  $n \in \mathbb{Z}_+ \setminus \{0\}$  such that  $\mu(\{n\}) > 0$  and  $h \in \mathbb{Z}_+ \setminus \{0\}$  such that  $hn + \inf I > 0$ . Then

$$0 < \mu(I)\mu(\{n\})^h = \mathbb{P}(S_1 \in I, S_{i+1} - S_i = n, i = 1, \dots, h) \leq \mathbb{P}(S_{n+1} \in hn + I).$$

This implies that  $\mu^{*(n+1)}(hn + I) > 0$ , where  $hn + I \subset [0, \infty) \setminus \mathbb{Z}_+$ , which contradicts the assumption.  $\square$

**Theorem 2.3.** *Downward skip free distributions belong to class  $\mathcal{C}$ .*

*Proof.* Let  $\mu \in \mathcal{M}_1$  whose convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  are known. Then from Lemma 2.4 we can determine if the support of  $\mu$  is included in  $\mathbb{Z}$  or not. Let us assume that it is the case.

As already noticed at the beginning of this section, we can determine if  $(S_n)$  drifts to  $\infty$  or not. Assume first that  $(S_n)$  drifts to  $\infty$ . From Theorem 2.2 (a), we can determine if there exists  $\lambda < 0$  such that  $\phi(\lambda) < 1$ . If this is not the case, then  $\mu$  cannot be downward skip free. To see this, consider for instance the analytic extension of the Wiener-Hopf factorization (1.1) which is given in (9.3.1), p.99 of [2], for nonnegative arguments and in the case of upward skip free random walks. On the contrary, if there exists  $\lambda < 0$  such

that  $\phi(\lambda) < 1$ , then from Theorem 2.2,  $\mu$  can be determined. In particular, we know if  $\mu$  is downward skip free or not.

Now assume that  $(S_n)$  does not drift to  $\infty$ . From the same arguments as those we used to prove identity (2.5), we obtain for all  $n \geq 1$ ,

$$\mathbb{P}(S_{\tau_+} > n) = \sum_{r \geq 1} v(r) \mathbb{P}(S_1 > n + r),$$

where  $v$  is the renewal measure of the (strict) downward ladder height process of  $(S_n)$ . In particular, this renewal measure satisfies  $v(r) \leq 1$  for all  $r \geq 1$ . Moreover,  $(S_n)$  is downward skip free and does not drift to  $\infty$ , if and only if  $v(r) = 1$ , for all  $r \geq 1$ . Since it is the only unknown in the above equation, we can determine if it is the case or not. Finally, knowing that  $\mu$  is downward skip free, we immediately determine this distribution on  $\mathbb{R}$  from its knowledge on  $[0, \infty)$ .  $\square$

It appears in the proof of Theorem 2.3 that downward skip free distributions which drift to  $\infty$  actually satisfy condition (a) of Theorem 2.2. Hence the only additional case in this subsection is this of downward skip free distributions which does not drift to  $\infty$ .

We will denote by  $\mathcal{E}$  the set of measures  $\mu$  satisfying assumptions of Theorem 2.1 or those of Theorem 2.2 or those of Theorem 2.3, that is the set of measures satisfying both (2.3) and (2.6) or (a) or (b) or downward skip free distributions. It will be called the *exponential* class. From Theorems 2.1, 2.2 and 2.3, we have  $\mathcal{E} \subset \mathcal{C}$ .

### 3. WHEN $\mu$ IS CHARACTERIZED BY $\mu$ AND $\mu^{*2}$ ON $[0, \infty)$ .

**3.1. Preliminary results.** We will show that in many cases, the sole data of  $\mu$  and  $\mu * \mu$  on  $[0, \infty)$  actually suffices to determine  $\mu$ . In this subsection, we give a theoretical condition for this to hold.

We first observe that in Conjectures C and C', there is no loss of generality in assuming that  $\mu$  is absolutely continuous on  $[0, \infty)$ .

**Conjecture C'.** *Any absolutely continuous distribution  $\mu \in \mathcal{M}_1$  whose support is not included in  $(-\infty, 0)$  is determined by its convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$ .*

**Lemma 3.1.** *Conjectures C, C' and C'' are equivalent.*

*Proof.* We already know from Section 1 that Conjectures C and C' are equivalent. Then clearly, it suffices to prove that if Conjecture C'' is true, then Conjecture C' is true.

Let  $\mu, \mu_1 \in \mathcal{M}_1$  be any two distributions such that the measures  $\mu^{*n}$  and  $\mu_1^{*n}$  agree on  $[0, \infty)$ , for all  $n \geq 1$ . Let  $g$  be any probability density function on  $[0, \infty)$ , i.e.  $\int_0^\infty g(x) dx = 1$  and let  $\bar{\mu}, \bar{\mu}_1 \in \mathcal{M}_1$  be the absolutely continuous measures whose respective densities are  $h(x) = \int_{\mathbb{R}} g(y - x) \mu(dy)$  and  $h_1(x) = \int_{\mathbb{R}} g(y - x) \mu_1(dy)$ ,  $x \in \mathbb{R}$ .

Denoting by  $g^{*n}$  the  $n$ -th convolution product of the function  $g$  by itself, it is plain that for all  $x \geq 0$ ,

$$\bar{\mu}^{*n}[x, \infty) = \int_0^\infty \mu^{*n}[x + y, \infty) g^{*n}(y) dy \quad \text{and} \quad \bar{\mu}_1^{*n}[x, \infty) = \int_0^\infty \mu_1^{*n}[x + y, \infty) g^{*n}(y) dy.$$

Therefore since the measures  $\mu^{*n}$  and  $\mu_1^{*n}$  agree on  $[0, \infty)$ , for all  $n \geq 1$ , the measures  $\bar{\mu}^{*n}$  and  $\bar{\mu}_1^{*n}$  also agree on  $[0, \infty)$ , for all  $n \geq 1$  and from the assumption that conjecture C''



is true, we conclude that the measures  $\bar{\mu}$  and  $\bar{\mu}_1$  agree on  $\mathbb{R}$ . Then we can identify both characteristic functions:

$$\begin{aligned}\int_{\mathbb{R}} e^{itx} \bar{\mu}(dx) &= \int_{\mathbb{R}} e^{itx} \mu(dx) \int_0^\infty e^{-itx} g(x) dx \quad \text{and} \\ \int_{\mathbb{R}} e^{itx} \bar{\mu}_1(dx) &= \int_{\mathbb{R}} e^{itx} \mu_1(dx) \int_0^\infty e^{-itx} g(x) dx.\end{aligned}$$

This implies that  $\int_{\mathbb{R}} e^{itx} \mu(dx) = \int_{\mathbb{R}} e^{itx} \mu_1(dx)$ , for all  $t \in \mathbb{R}$  thanks to the fact that the characteristic function of  $g$  cannot vanish identically in an interval, see Theorem 2 in [8] and by continuity of characteristic functions. Then we conclude that  $\mu = \mu_1$  on  $\mathbb{R}$ , from injectivity of the Fourier transform.  $\square$

We derive from Lemma 3.1 that there is no loss of generality in assuming that  $\mu$  is absolutely continuous on  $\mathbb{R}$ . We will sometimes make this assumption and denote the density of  $\mu$  by  $f$ .

**Lemma 3.2.** *For any probability density function,  $f$  on  $\mathbb{R}$  and for all  $t \geq 0$ ,*

$$(3.7) \quad \int_0^\infty f(t+s)\bar{f}(s) ds = \frac{1}{2} \left( f * f(t) - \int_0^t f(t-s)f(s) ds \right),$$

where  $\bar{f}(s) = f(-s)$ .

*Proof.* It suffices to decompose  $f * f$  as

$$\begin{aligned}f * f(t) &= \int_{\mathbb{R}} f(t-s)f(s) ds \\ &= \int_{-\infty}^0 f(t-s)f(s) ds + \int_0^t f(t-s)f(s) ds + \int_t^\infty f(t-s)f(s) ds.\end{aligned}$$

Then from a change of variables, we obtain,  $\int_t^\infty f(t-s)f(s) ds = \int_{-\infty}^0 f(t-s)f(s) ds = \int_0^\infty f(t+s)\bar{f}(s) ds$ , which proves our identity.  $\square$

The main idea of this section is to exploit identity (3.7) in order to characterize the function  $\bar{f}$  on  $[0, \infty)$  (or equivalently  $f$  on  $(-\infty, 0]$ ) from the sole data of  $f$  and  $f * f$  on  $[0, \infty)$ . More specifically, assume that  $f$  restricted to  $[0, \infty)$  fulfills the following property: for any two nonnegative Borel functions  $g_1$  and  $g_2$  defined on  $[0, \infty)$  such that  $\int_0^\infty f(t+s)g_1(s) ds < \infty$ ,  $\int_0^\infty f(t+s)g_2(s) ds < \infty$ , for all  $t \geq 0$ , the following implication is satisfied,

$$(3.8) \quad \int_{[0, \infty)} f(t+s)g_1(s) ds = \int_{[0, \infty)} f(t+s)g_2(s) ds, \quad \text{for all } t \geq 0 \Rightarrow g_1 \equiv g_2, \quad \text{a.e.}$$

Then clearly, the map  $t \mapsto \int_{[0, \infty)} f(t+s)\bar{f}(s) ds$  characterizes  $\bar{f}$  on  $[0, \infty)$  and therefore from (3.7),  $\mu$  is determined on  $\mathbb{R}$  by the sole data of  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ .

**Remark 3.1.** *There are density functions  $f$  which do not satisfy (3.8). For instance with  $f(s) = \frac{1}{2}e^{-s}$ , the operator  $t \mapsto \int_0^\infty f(t+s)\bar{f}(s) ds = \frac{1}{2}e^{-t} \int_0^\infty e^{-s}\bar{f}(s) ds$  provides a very poor information on  $\bar{f}$  and certainly cannot characterize this function on  $[0, \infty)$ . Also if  $\mu$  has a bounded support in  $[0, \infty)$ , then clearly the operator  $t \mapsto \int_0^\infty f(t+s)\bar{f}(s) ds$  cannot characterize  $\bar{f}$  outside this support.*

*Note also that from (3.7), the knowledge of  $t \mapsto \int_0^\infty f(t+s)\bar{f}(s) ds$  and  $f(t)$ , for  $t \geq 0$  is equivalent to this of the functions  $f$  and  $f^{*2}$  on  $[0, \infty)$ . Therefore, in the above examples,  $f$  is not even determined by the data of  $f$  and  $f^{*2}$  on  $[0, \infty)$ .*

The following proposition gives a sufficient condition for (3.8) to hold.

**Proposition 3.1.** *Assume that  $\mu$  is absolutely continuous on  $\mathbb{R}$  with density  $f$ . Let us introduce the following set of functions defined on  $[0, \infty)$ ,*

$$\mathcal{H} := \left\{ \sum_{k=1}^n \alpha_k f(t_k + \cdot) : n \geq 1, \alpha_k \in \mathbb{R}, t_k \geq 0 \right\}.$$

*If the restriction of  $f$  to  $[0, \infty)$  belongs to  $L^\infty([0, \infty))$  and if  $\mathcal{H}$  is dense in  $L^\infty([0, \infty))$ , then for any  $g \in L^1([0, \infty))$ , the following implication is satisfied,*

$$(3.9) \quad \int_0^\infty f(t+s)g(s)ds = 0, \text{ for all } t \geq 0 \Rightarrow g \equiv 0, \text{ a.e.}$$

*When (3.9) holds, the measure  $\mu$  is determined on  $\mathbb{R}$  by  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ .*

*Proof.* If  $\int_0^\infty f(t+s)g(s)ds = 0$ , for all  $t \geq 0$ , then clearly, since  $\mathcal{H}$  is dense in  $L^\infty([0, \infty))$ ,  $\int_0^\infty h(s)g(s)ds = 0$  for all  $h \in L^\infty([0, \infty))$  and this implies that  $g \equiv 0$ , a.e.

Assume now that the restriction of the measures  $\mu$  and  $\mu^{*2}$  are known on  $[0, \infty)$ . Recall the notation  $\bar{f}$  from Lemma 3.2 and observe that the right hand side of identity (3.7) is known for all  $t \geq 0$ . From (3.9) this determines  $\bar{f}$  on  $[0, \infty)$  and the measure  $\mu$  is determined.  $\square$

Unfortunately we do not know any example of function satisfying the condition of Proposition 3.1 and finding a simple criterion on  $f$  for it to satisfy this condition remains an open problem. More specifically, we may wonder if the converse of Proposition 3.1 holds, that is if assertion (3.9) implies that  $\mathcal{H}$  is dense in  $L^\infty([0, \infty))$ . The latter problem can be compared with Wiener's approximation theorem which asserts that for a function  $f$  in  $L^1(\mathbb{R})$  the set  $\mathcal{H}$  (thought as a set of functions defined on  $\mathbb{R}$ ) is dense in  $L^1(\mathbb{R})$  if and only if the Fourier transform of  $f$  does not vanish, see [6].

In Subsection 3.2, we give a class of density functions such that (3.8) holds and in Subsection 3.3, we give a class of density functions which are bounded on  $[0, \infty)$  and such that (3.9) holds.

**3.2. The completely monotone class.** In this subsection, we assume that  $\mu$  is absolutely continuous with respect to the Lebesgue measure on  $\mathbb{R}$  and we denote by  $f$  its density.

We will show that if  $f$  restricted to  $(a, \infty)$ , for some  $a \geq 0$  is a completely monotone function satisfying some mild additional assumption, then  $\mu$  is characterized from  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ . Let us first recall that from Bernstein Theorem, the function  $f$  is completely monotone on  $(a, \infty)$ , for  $a \geq 0$ , if and only if there is a positive Borel measure  $\nu$  on  $(0, \infty)$  such that for all  $t > a$ ,

$$(3.10) \quad f(t) = \int_0^\infty e^{-ut} \nu(du).$$

**Theorem 3.1.** *Assume that there is  $a \geq 0$  such that the restriction of  $f$  to  $(a, \infty)$  is completely monotone and that the support of the measure  $\nu$  in (3.10) contains a sequence  $(a_n)$  such that  $\sum_n a_n^{-1} = +\infty$ . Then (3.8) holds and the measure  $\mu$  is characterized by the restriction of  $\mu$  and  $\mu^{*2}$  to  $[0, \infty)$ . In particular,  $\mu$  belongs to class  $\mathcal{C}$ .*

*Proof.* Let  $g$  be any nonnegative Borel function defined on  $[0, \infty)$  such that  $\int_0^\infty f(t+s)g(s)ds < \infty$ , for all  $t > a$ . Then from Fubini's Theorem, for all  $t > a$ ,

$$\begin{aligned} \int_0^\infty f(t+s)g(s)ds &= \int_0^\infty \int_0^\infty e^{-u(t+s)}\nu(du)g(s)ds \\ &= \int_0^\infty e^{-ut} \int_0^\infty e^{-us}g(s)ds\nu(du). \end{aligned}$$

This expression is the Laplace transform of the measure  $\theta(du) := \int_0^\infty e^{-us}g(s)ds\nu(du)$ . The knowledge of this Laplace transform for all  $t > a$  characterizes the measure  $\theta(du)$  so that a version of the density function  $u \mapsto \int_0^\infty e^{-us}g(s)ds$  is known on a Borel set  $B \subset (0, \infty)$  such that  $\nu(B) = 1$ . Since this density function is continuous, it is known on  $\overline{B}$  and hence it is known everywhere on the support of  $\nu$ . Therefore, from the assumption on  $\nu$ , we can find a sequence  $(a_n)$  such that  $\sum_n a_n^{-1} = +\infty$  and such that the Laplace transform  $\int_0^\infty e^{-a_n s}g(s)ds$  of the function  $g$  at  $a_n$  is known for each  $n$ . From a result in [4], this is enough to determine the function  $g$  and (3.8) holds, see also [3], p.430.

Then we derive from (3.8) and Lemma 3.2 that the function  $\bar{f}$  is determined on  $[0, \infty)$  from the restriction on  $[0, \infty)$  of  $\mu$  and  $\mu^{*2}$ . Therefore the measure  $\mu$  is determined.  $\square$

We will denote by  $\mathcal{M}$  the set of absolutely continuous measures  $\mu$  whose density  $f$  satisfies the assumption of Theorem 3.1. This class will be called the completely monotone class. Theorem 3.1 shows that  $\mathcal{M} \subset \mathcal{C}$ .

**Remark 3.2.** Note that since  $f$  is a density function, the measure  $\nu$  in (3.10) should also satisfy

$$(3.11) \quad \int_0^\infty f(t)dt = \int_0^\infty \frac{e^{-au}}{u}\nu(du) \leq 1.$$

**Remark 3.3.** Clearly class  $\mathcal{E}$  is not included in class  $\mathcal{M}$ . Moreover, it is easy to find an example of a measure in class  $\mathcal{M}$  which does not belong to class  $\mathcal{E}$ . First we readily check that whenever the support  $S$  of  $\nu$  is such that  $S \cap (0, \varepsilon) \neq \emptyset$ , for all  $\varepsilon > 0$ , then  $\mu$  has no positive exponential moments, so that it cannot belong to class  $\mathcal{E}$ . Then let us take for instance the measure  $\nu(du)$  on  $[0, \infty)$  with density  $u^2 \mathbb{I}_{\{u \in [0, 1]\}}$ . This measure satisfies the assumption of Theorem 3.1, for  $a = 0$ .

**Remark 3.4.** Theorem 3.1 excludes completely monotone functions of the type  $f(t) = \sum_{k=1}^n \alpha_k e^{-\beta_k t}$ ,  $t > a$ , for some  $\alpha_k, \beta_k > 0$  and some finite  $n$  since in this case the measure  $\nu(du) = \sum_{k=1}^n \alpha_k \delta_{\beta_k}(du)$  does not satisfy the condition required by this theorem. It also excludes functions  $f$  whose support is bounded since completely monotone functions on  $(a, \infty)$  are analytic on  $(a, \infty)$ . This remark is consistent with Remark 3.1.

**3.3. The analytic class.** Let us assume again that  $\mu$  has density  $f$  on  $\mathbb{R}$ . We will now exploit the same kind of arguments as in the previous subsection by assuming that  $f$  is the Fourier transform of some complex valued function.

**Theorem 3.2.** *Assume that there is a complex valued function  $k$  such that for all  $t \geq 0$ ,*

$$(3.12) \quad f(t) = \int_{\mathbb{R}} e^{iut} k(u) du.$$

*Assume moreover that*

1. *the moments  $M_n = \int_{\mathbb{R}} |u|^n |k(u)| du$ ,  $n \geq 0$  are finite and satisfy*

$$\limsup_n \frac{M_n}{n!} < \infty,$$

2. *the function  $k$  does not vanish on any interval of  $\mathbb{R}$ .*

*Then  $f$  is bounded on  $[0, \infty)$  and (3.9) holds. In particular, the measure  $\mu$  is characterized by the restrictions of  $\mu$  and  $\mu^{*2}$  to  $[0, \infty)$  and  $\mu$  belongs to class  $\mathcal{C}$ . Moreover  $f$  admits an analytic continuation on  $\mathbb{R}$ .*

*Proof.* The fact that  $f$  is bounded follows directly from (3.12) and 1. Let  $g \in L^1([0, \infty))$ , then from (3.12) and Fubini's theorem, we can write for all  $t \geq 0$ ,

$$\begin{aligned} \int_0^\infty f(t+s)g(s) ds &= \int_0^\infty \int_{\mathbb{R}} e^{iu(t+s)} k(u) du g(s) ds \\ &= \int_{\mathbb{R}} e^{iut} \int_0^\infty e^{ius} g(s) ds k(u) du. \end{aligned}$$

Assume that this expression vanishes for all  $t \geq 0$  and set  $\varphi(u) = \int_0^\infty e^{ius} g(s) ds$ . This means that the Fourier transform

$$\Psi(t) := \int_{\mathbb{R}} e^{iut} \varphi(u) k(u) du,$$

of the function  $u \mapsto \varphi(u)k(u)$ ,  $u \in \mathbb{R}$  vanishes for all  $t \geq 0$ . Then let us show that under our assumptions, the function  $\Psi$  is analytic on the whole real axis. First note that since  $|\varphi(u)| \leq 1$ ,  $u \in \mathbb{R}$  and since all the moments  $M_n$  are finite, then  $\Psi$  is infinitely differentiable on  $\mathbb{R}$  and

$$(3.13) \quad \Psi^{(n)}(t) = \int_{\mathbb{R}} (iu)^n e^{iut} \varphi(u) k(u) du, \quad t \in \mathbb{R}.$$

Then notice that for all  $t, u, x \in \mathbb{R}$ ,

$$\left| e^{iux} \left( e^{itu} - 1 + \frac{itu}{1!} - \dots - \frac{(itu)^{n-1}}{(n-1)!} \right) \right| \leq \frac{|tu|^n}{n!}.$$

We derive from this inequality and (3.13) that

$$(3.14) \quad \left| \Psi(x+t) - \Psi(x) - \frac{t}{1!} \Psi'(x) - \dots - \frac{t^{n-1}}{(n-1)!} \Psi^{(n-1)}(x) \right| \leq \frac{M_n}{n!} |t|^n.$$

Set  $c = \limsup_n M_n/n!$ , then from Stirling's formula, for  $|t| < 1/(3c)$ , the righthand side of (3.14) tends to 0, as  $n \rightarrow +\infty$ , hence the Taylor series of  $\Psi$  converges in some interval around  $x$ , for all  $x \in \mathbb{R}$ . It follows that  $\Psi$  is analytic on  $\mathbb{R}$ . As a consequence,  $\Psi$  is determined by its expression on the positive half line. Hence the Fourier transform of the continuous function  $u \mapsto \varphi(u)k(u)$  vanishes on  $\mathbb{R}$ , which means that this function vanishes a.e. on  $\mathbb{R}$ . Since  $k$  does not vanish on any interval of  $\mathbb{R}$  and  $\varphi$  is continuous, it implies that  $\varphi(u) = 0$ , for all  $u \in \mathbb{R}$  and we conclude that  $g(t) = 0$ , for almost every  $t \in [0, \infty)$ . We have proved that (3.9) holds and from the second part of Proposition 3.1,  $\mu$  is characterized by the restriction of  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ .

We have proved above that  $\Psi(t) = \int_{\mathbb{R}} e^{iut} \varphi(u) k(u) du$  is analytic on  $\mathbb{R}$ . It follows from the same arguments that the continuation of  $f$  on  $\mathbb{R}$  which is defined in a natural way by  $f(t) = \int_{\mathbb{R}} e^{iut} k(u) du$ ,  $t \in \mathbb{R}$  is analytic, which proves the last assertion of the theorem.  $\square$

We will denote by  $\mathcal{A}$  the class of distributions which satisfy the assumptions of Theorem 3.2. It will be called the analytic class. From Theorem 3.2,  $\mathcal{A} \subset \mathcal{C}$ .

It is very easy to construct examples of distributions in class  $\mathcal{A}$ , simply by choosing any symmetric function  $k$  which satisfies assumptions 1. and 2. in Theorem 3.2. Let us consider for instance  $k(u) = e^{-|u|}$ , for  $u \in \mathbb{R}$ . Then

$$f(t) = \int_{\mathbb{R}} e^{iut} e^{-|u|} du = \frac{1}{2(1+t^2)}, \quad t \geq 0.$$

Any extension on  $\mathbb{R}$  of this function into a density function determines a distribution of  $\mathcal{A}$  which does not belong to classes  $\mathcal{E}$  and  $\mathcal{M}$ . Conversely, none of the classes  $\mathcal{E}$  and  $\mathcal{M}$  is included in  $\mathcal{A}$ . It is straightforward for  $\mathcal{E}$ . Then let us consider

$$f(t) = \frac{1}{4} \left( \int_0^1 e^{-ut} u^{1/2} du + \int_1^\infty e^{-ut} u^{-1} du \right), \quad t > 0.$$

The measure  $\nu(du) = (u^{1/2} \mathbb{I}_{[0,1]}(u) + u^{-1} \mathbb{I}_{[1,\infty)}(u)) du$  satisfies the condition of Theorem 3.1 so that  $\mu \in \mathcal{M}$  (here we choose  $a = 0$ ). However,  $f$  is not bounded on  $(0, \infty)$ , hence it does not admit an analytic continuation on  $\mathbb{R}$ , so that  $f$  does not belong to class  $\mathcal{A}$ , from Theorem 3.2. (Note also that since the support of  $\nu$  intersects any interval  $(0, \varepsilon)$ ,  $\varepsilon > 0$ , the measure  $\mu$  has no positive exponential moments, see Remark 3.3.)

**3.4. A class of discrete distributions.** In this section, we present the discrete counterpart of class  $\mathcal{M}$ . More specifically, we will consider distributions whose support is included in  $\mathbb{Z}$ .

First we need the following equivalent of Lemma 3.2 for discrete distributions. Its proof is straightforward so we omit it.

**Lemma 3.3.** *Let  $(q_n)_{n \in \mathbb{Z}}$  be any probability on  $\mathbb{Z}$ . Define  $q_n^{*2} = \sum_{k \in \mathbb{Z}} q_{n-k} q_k$ . Then for all  $n \geq 1$*

$$(3.15) \quad \sum_{k=-\infty}^0 q_{n-k} q_k = \frac{1}{2} \left( q_n^{*2} - \sum_{k=1}^{n-1} q_{n-k} q_k \right),$$

where we set  $\sum_{k=1}^{n-1} q_{n-k} q_k = 0$  if  $n = 1$ .

A sequence  $(a_k)_{k \geq 0}$  of nonnegative real numbers is called completely monotone if for all  $k \geq 0$  and  $n \geq 1$ ,

$$\Delta^n a_k := \Delta^{n-1} a_k - \Delta^{n-1} a_{k+1} \geq 0,$$

where  $\Delta^0 a = a$ . A result from Hausdorff asserts that  $(a_k)_{k \geq 0}$  is completely monotone if and only if there is a probability measure  $\nu$  on  $[0, 1]$  such that for all  $k \geq 0$ ,

$$(3.16) \quad a_k = \int_0^1 t^k \nu(dt).$$

Let us set  $\mu(\{n\}) = \mu_n$ , for  $n \in \mathbb{Z}$ . Then by assuming that  $(\mu_n)_{n \geq 0}$  is completely monotone, we obtain a new class of distributions satisfying conjecture  $\mathcal{C}$  as shows the following theorem.

**Theorem 3.3.** *Assume that  $(\mu_n)_{n \geq 0}$  is completely monotone and that the support of the measure  $\nu$  in the representation (3.16) contains a sequence  $(c_n)$  such that  $\sum (-\ln c_n)^{-1} = \infty$ . Then the data of  $\mu$  and  $\mu^{*2}$  restricted to  $[0, \infty)$  allows us to determine if the support of  $\mu$  is included in  $\mathbb{Z}$ . Assume that this is the case, then the measure  $\mu$  is determined by the restriction of  $\mu$  and  $\mu^{*2}$  to  $[0, \infty)$ . In particular  $\mu$  belongs to class  $\mathcal{C}$ .*

*Proof.* First let us observe that we can derive from the data of  $\mu$  and  $\mu^{*2}$  restricted to  $[0, \infty)$  that the support of  $\mu$  is included in  $\mathbb{Z}$ . Indeed, assume that the support of  $\mu$  restricted to  $(-\infty, 0]$  is not included in  $\mathbb{Z}_-$ . Then there is an interval  $I \subset (-\infty, 0] \setminus \mathbb{Z}_-$  such that  $\mu(I) > 0$ . From (3.16),  $\mu(\{n\}) > 0$  for all  $n \geq 1$ . Let  $n \in \mathbb{Z}_+ \setminus \{0\}$  such that  $n + \inf I > 0$ , then

$$0 < \mu(I)\mu(\{n\}) = \mathbb{P}(S_1 \in I, S_2 - S_1 = n) \leq \mathbb{P}(S_2 \in n + I).$$

This implies that  $\mu^{*2}(n + I) > 0$ , where  $n + I \subset [0, \infty) \setminus \mathbb{Z}_+$ , which contradicts the assumption.

From the Haussdorff representation recalled above, there is a unique finite measure  $\nu$  on  $[0, 1]$  such that  $\mu_k = \int_0^1 t^k \nu(dt)$ . Using this representation and Fubini's theorem, we can write for all  $n \geq 0$ ,

$$\sum_{k=0}^{\infty} \mu_{-k} \mu_{n+k} = \int_0^1 t^n \left( \sum_{k=0}^{\infty} \mu_{-k} t^k \right) \nu(dt).$$

From (3.15) in Lemma 3.3 applied to  $(\mu_k)$ , we derive that this expression is determined from the knowledge of  $\mu$  and  $\mu^{*2}$  on  $[0, \infty)$ . This means that we know the moments of the measure  $(\sum_{k=0}^{\infty} \mu_{-k} t^k) \cdot \nu(dt)$ . This measure is finite and its support is included in  $[0, 1]$ , hence it is determined by its moments. Then we know the generating function  $t \mapsto \sum_{k=0}^{\infty} \mu_{-k} t^k$  of the sequence  $(\mu_{-k})_{k \geq 0}$  on the support of  $\nu$ , since this function is continuous. From the assumption, we know this generating function on a sequence  $(c_n)$  such that  $\sum (-\ln c_n)^{-1} = \infty$ . This is enough to determine the sequence  $(\mu_{-k})_{k \geq 0}$ , from [4].

We conclude that the measures  $\mu$  and  $\mu^{*2}$  restricted to  $[0, \infty)$  allow us to determine  $\mu$  on  $\mathbb{Z}$ .  $\square$

The set of measures satisfying the assumptions of Theorem 2.4 will be called the discrete monotone class and will be denoted by  $\mathcal{M}_d$ . Theorem 2.4 shows that  $\mathcal{M}_d \subset \mathcal{C}$ . Moreover, it is clear that none of the classes  $\mathcal{E}$ ,  $\mathcal{M}$  and  $\mathcal{A}$  is included in  $\mathcal{M}_d$  and that these classes do not contain  $\mathcal{M}_d$ .

#### 4. WHEN $\mu$ IS INFINITELY DIVISIBLE

The aim of this section is to present a problem equivalent to conjecture  $\mathcal{C}$  in the framework of infinitely divisible distributions. When  $\mu$  is infinitely divisible, the Wiener-Hopf factorization can be understood in two different ways: we can either factorize the characteristic function  $\varphi$  as in (1.1), or we can factorize the characteristic exponent  $\psi$ , which is defined by

$$\varphi(t) = e^{\psi(t)}, \quad t \in \mathbb{R}.$$

Then let us recall the Wiener-Hopf factorization in the latter context. Let  $(X_t, t \geq 0)$  be a real Lévy process issued from 0 under the probability  $\mathbb{P}$  and such that  $X_1$  has law  $\mu$  under this probability, that is  $\mathbb{E}(e^{iuX_t}) = e^{-t\psi(u)}$ , for all  $t \geq 0$ . The characteristic exponent of  $\mu$

is given explicitly according to the Lévy-Khintchine formula by

$$\psi(u) = iau + \frac{\sigma^2}{2}u^2 + \int_{\mathbb{R} \setminus \{0\}} (1 - e^{iux} + iux1_{\{|x| \leq 1\}}) \Pi(dx),$$

where  $a \in \mathbb{R}$ ,  $\sigma \geq 0$  and  $\Pi$  is a measure on  $\mathbb{R} \setminus \{0\}$ , such that  $\int (x^2 \wedge 1) \Pi(dx) < \infty$ . Then the Wiener-Hopf factorization of  $\psi$  has the following form:

$$(4.17) \quad s + \psi(u) = \kappa_+(s, -iu)\kappa_-(s, iu), \quad u \in \mathbb{R}, \quad s \geq 0,$$

where  $\kappa_+$  and  $\kappa_-$  are the Laplace exponents of the upward and downward ladder processes  $(\tau^+, H^+)$  and  $(\tau^-, H^-)$  of  $X$ , that is  $\mathbb{E}(e^{-\alpha\tau_t^{+/-} - i\beta H_t^{+/-}}) = e^{-t\kappa_{+/-}(\alpha, \beta)}$ . These exponents are given explicitly for  $\alpha, \beta \geq 0$  by the identities,

$$(4.18) \quad \kappa_-(\alpha, \beta) = k_- \exp \left( \int_0^\infty \int_{(-\infty, 0)} (e^{-t} - e^{-\alpha t - \beta x}) \frac{1}{t} \mathbb{P}(X_t \in dx) dt \right)$$

$$(4.19) \quad \kappa_+(\alpha, \beta) = k_+ \exp \left( \int_0^\infty \int_{[0, \infty)} (e^{-t} - e^{-\alpha t - \beta x}) \frac{1}{t} \mathbb{P}(X_t \in dx) dt \right),$$

where  $k_-$  and  $k_+$  are positive constants depending on the normalization of the local times at the infimum and at the supremum of  $X$ . The joint law of  $(\tau_1^+, H_1^+)$  is the continuous time counterpart of the joint law  $(\tau_+, S_{\tau_+})$  defined in Section 1, in the setup of random walks. We refer to Chap. VI of [1], Chap. IV of [5] or Chap. of [2] for complete definitions of these notions. Note that our formulation of the Wiener-Hopf factorization (4.17) includes compound Poisson processes since expression (4.19) takes account of a possible mass at 0 for the measure  $\mathbb{P}(X_t \in dx)$ . This slight extension can be derived from p. 24 and 25 of [9], see also the end of Section 6.4, p.183 in [5].

Set  $\bar{\Pi}(t) = \Pi(t, \infty)$ ,  $t > 0$  and let  $\mu_t$  be the law of  $X_t$ .

**Lemma 4.1.** *The data of  $\kappa_+$  is equivalent to this of  $\mu_t$  on  $[0, \infty)$ , for all  $t \geq 0$ . Moreover, the knowledge of  $\kappa_+$  allows us to determine the function  $\bar{\Pi}(t)$ ,  $t > 0$ .*

*Proof.* From identity

$$(4.20) \quad \frac{1}{t} \mathbb{P}(X_t \in dx) dt = \int_0^\infty \mathbb{P}(\tau_u \in dt, H_u \in dx) \frac{du}{u}, \quad x \geq 0, t > 0,$$

which can be found in Section 5.2 of [2], we see that the law of  $X_t$  in  $[0, \infty)$ , for all  $t \geq 0$  is determined by the law of  $(\tau, H)$  and hence by  $\kappa^+$ . (Note that equation (4.20) is also valid for compound Poisson processes.) Conversely, it follows directly from formula (4.19) that  $\kappa^+$  is determined by the data of the measure  $\mu_t$  on  $[0, \infty)$ , for all  $t \geq 0$ .

The second assertion is a consequence of the first one and Exercise 1 of chap.I in [1], which asserts that the family of measures  $\frac{1}{t} \mathbb{P}(X_t \in dx)$  converges vaguely toward  $\Pi$ , as  $t \rightarrow 0$ .  $\square$

The above lemma enables us to make the connection between the two Wiener-Hopf factorizations (1.1) and (4.17). Let us state it more specifically in the following proposition.

**Proposition 4.1.** *The Wiener-Hopf factor  $\kappa_+$  allows us to determine the Wiener-Hopf factor  $\chi_+$ .*

*Proof.* The result is straightforward from Lemma 4.1. Indeed, knowing  $\kappa_+$  we can determine  $\mu_n = \mu^{*n}$  restricted to  $[0, \infty)$ , for all  $n \geq 1$  and from Section 1 that this data is equivalent to that of  $\chi_+$ .  $\square$

**Definition 4.1.** We will denote by  $\mathcal{C}_i$  the class of infinitely divisible distributions  $\mu$  which are determined by the data of their upward Wiener-Hopf factor  $\kappa_+(s, t)$ , for  $s, t \geq 0$  or equivalently by the data of the measures  $\mu_t$ ,  $t > 0$  restricted to  $[0, \infty)$ .

Let us denote by  $\mathcal{J}$  the set of infinitely divisible distributions. Then it is straightforward that  $\mathcal{J} \cap \mathcal{C} \subset \mathcal{C}_i$ . In particular if Conjecture C is true, then  $\mathcal{C}_i = \mathcal{J}$ . It was proved in Chapter 4 of [9] that infinitely divisible distributions having some exponential moments belong to class  $\mathcal{C}_i$ , which is a consequence of our results. The latter work uses a different technique based on the analytical continuation of the Wiener-Hopf factors  $\kappa_+$  and  $\kappa_-$ .

Let  $k_-, \delta_-, \gamma_-$  and  $k_+, \delta_+, \gamma_+$  be the killing rate, the drift and the Lévy measure of the subordinators  $H_-$  and  $H_+$ , respectively and let us set  $\bar{\gamma}_+(x) = \gamma_+(x, \infty)$  and  $\bar{\gamma}_-(x) = \gamma_-(x, \infty)$ . Let also  $U_-$  be the renewal measure of the downward ladder height process  $H^-$ , that is  $U_-(dx) = \int_0^\infty \mathbb{P}(H_t^- \in dx)$ .

**Theorem 4.1.** Assume that the function  $t \in (a, \infty) \mapsto \bar{\Pi}(t)$  is completely monotone, for some  $a \geq 0$ , that is there exists a Borel measure  $\nu$  on  $(0, \infty)$  such that for all  $t > a$ ,

$$(4.21) \quad \bar{\Pi}(t) = \int_0^\infty e^{-ut} \nu(du).$$

Assume moreover that the support of  $\nu$  contains a sequence  $(a_n)$  such that  $\sum_n a_n^{-1} = +\infty$ . Then the measure  $\mu$  belongs to the class  $\mathcal{C}_i$ .

*Proof.* The proof relies on Vigon's équation amicale inversée, see [9], p.71, or (5.3.4) p.44 in [2] which can be written as

$$(4.22) \quad \bar{\gamma}_+(x) = \int_{[0, \infty)} U_-(dy) \bar{\Pi}(x+y), \quad x > 0.$$

From Lemma 4.1, given  $\kappa_+$ , we know both  $\bar{\gamma}_+(x)$ , for  $x > 0$  and  $\bar{\Pi}(t)$ , for  $t > 0$ . Then we will show that under our assumption, equation (4.22) allows us to determine the renewal measure  $U_-(dy)$ , so that the law of  $X$  will be entirely determined, thanks to the relation:

$$(4.23) \quad \begin{aligned} \hat{U}_-(z) &= \int_{\mathbb{R}_+} e^{-yz} U_-(dy) \\ &= \frac{1}{\kappa_-(0, z)}, \quad z > 0, \end{aligned}$$

and the Wiener-Hopf factorization (4.17).

From (4.21), (4.22) and Fubini's Theorem, we can write for all  $x > 0$ ,

$$(4.24) \quad \begin{aligned} \bar{\gamma}_+(x) &= \int_{[0, \infty)} U_-(dy) \int_0^\infty e^{-(x+y)z} \nu(dz) \\ &= \int_0^\infty e^{-xz} \hat{U}_-(z) \nu(dz). \end{aligned}$$

Then the left hand side of equation (4.24) determines the measure  $\hat{U}_-(z) \nu(dz)$ . Since  $z \mapsto \hat{U}_-(z)$  is a continuous function, then it is determined on the support of  $\nu$ . From our assumption on this support and [4] we derive that  $\hat{U}_-$  (and hence  $\kappa_-(0, z) = -\log \mathbb{E}(e^{-zH_1^-})$ ) for  $z > 0$ , is determined.  $\square$

Note that an analogous result to Theorem 4.1 holds when  $\mu$  has support in  $\mathbb{Z}$  and the sequence  $\Pi(n)$ ,  $n \geq 1$  satisfies the same assumptions as  $(\mu_n)_{n \geq 0}$  in Theorem 3.3. One



may also wonder if an assumption such as (3.12) for  $\bar{\Pi}(t)$  would lead to a similar result to Theorem 3.2. However, in order to use the same argument as in the proof of this theorem together with equation (4.22), we need  $\hat{U}_-(z)$  to be bounded, which is not the case in general.

**Remark 4.1.** *As already observed above, the class  $\mathcal{C}_i$  contains at least all probability measures in the set  $\mathcal{I} \cap (\mathcal{E} \cup \mathcal{M} \cup \mathcal{A} \cup \mathcal{M}_d)$  but Theorem 4.1 shows that there are other distributions in  $\mathcal{C}_i$ . Indeed, it is easy to construct an example of a compound Poisson process  $(X_t, t \geq 0)$  with intensity 1, whose Lévy measure  $\Pi$  satisfies conditions of Theorem 4.1 but such that the law  $\mu(dx) = e^{-1} \sum_{n \geq 0} \Pi^{*n}(dx)/n!$  of  $X_1$  does not belong to any of the classes  $\mathcal{E}$ ,  $\mathcal{M}$ ,  $\mathcal{A}$  and  $\mathcal{M}_d$ .*

An infinitely divisible distribution is said to be downward skip free (respectively upward skip free) if the support of the measure  $\Pi$  is included in  $(-\infty, 0]$  (respectively in  $[0, \infty)$ ). Here is a counterpart of Theorem 2.3.

**Theorem 4.2.** *Skip free infinitely divisible distributions belong to the class  $\mathcal{C}_i$ .*

*Proof.* The proof relies on Vigon's équation amicale, p.71 in [9]. See also equation (5.3.3), p.44 in [2]. If  $\delta_- > 0$ , then from [9], the Lévy measure  $\gamma_+$  is absolutely continuous and we will denote by  $\gamma_+(x)$  its density. Then Vigon's équation amicale can be written as

$$\bar{\Pi}(x) = \int_0^\infty \gamma_+(x+du)\bar{\gamma}_-(u) + \delta_- \gamma_+(x) + k_- \bar{\gamma}_+(x), \quad x > 0.$$

It is plain that in the right hand side, the term  $\int_0^\infty \gamma_+(x+du)\bar{\gamma}_-(u)$  is identically 0 if and only if the Lévy process  $X$  is spectrally positive, that is  $\mu$  is downward skip free. Moreover, (4.17) for  $u = 0$  entails that the knowledge of  $\kappa_+$  implies that of  $\kappa_-(s, 0)$ , for all  $s \geq 0$ . In particular, we know the killing rate of the subordinator  $(\tau_t^-, t \geq 0)$ , and this killing rate is the same as this of  $(H_t^-, t \geq 0)$ , that is  $k_-$ . Then we conclude that  $X$  is spectrally positive if and only if there is a constant  $\delta_-$  such that

$$\bar{\Pi}(x) = \delta_- \gamma_+(x) + k_- \bar{\gamma}_+(x), \quad x > 0,$$

and this can be determined, since from our data, we know  $k_-$ ,  $\bar{\Pi}(x)$  and  $\bar{\gamma}_+(x)$ , for  $x > 0$ .  $\square$

## 5. MORE CLASSES OF DISTRIBUTIONS

In the previous sections, we have highlighted the subclasses  $\mathcal{E}$ ,  $\mathcal{M}$ ,  $\mathcal{M}_d$  and  $\mathcal{A}$  of  $\mathcal{C}$  and proved that these sets of distributions are distinct from each other. More specifically, none of them is included into another one. Then the aim of this section is to show that some of these classes can be substantially enlarged through simple arguments.

Actually for most of the subclasses investigated in this paper, we imposed conditions bearing only on  $\mu$  restricted to  $[0, \infty)$ , but one is also allowed to make assumptions on  $\mu^{*n}$  restricted to  $[0, \infty)$ . In order to move in this direction, let us mention the following straightforward extension of results of Section 3.

**Proposition 5.1.** *Let  $\mu \in \mathcal{M}_1$  be absolutely continuous with density  $f$ . If there is  $n \geq 1$  such that the density function  $f^{*n}$  satisfies the same conditions as  $f$  in Theorems 3.1 or in Theorem 3.2, then  $\mu$  is determined by  $\mu^{*n}$  and  $\mu^{*2n}$  restricted to  $[0, \infty)$ . In particular  $\mu$  belongs to class  $\mathcal{C}$ .*

It is plain that an analogous extension of Theorem 3.3 is satisfied. Then here is a more powerful result allowing us to extend our classes of distributions.

**Theorem 5.1.** *Let  $\mu \in \mathcal{M}_1$ . If there is  $\nu \in \mathcal{M}_1$  whose support is included in  $(-\infty, 0]$  and such that  $\mu * \nu \in \mathcal{C}$ , then  $\mu \in \mathcal{C}$ .*

*Proof.* Let  $\mu, \mu_1 \in \mathcal{M}_1$  such that for each  $n \geq 1$ , the measures  $\mu^{*n}$  and  $\mu_1^{*n}$  restricted to  $[0, \infty)$  coincide. Set  $\bar{\mu} = \mu * \nu$  and  $\bar{\mu}_1 = \mu_1 * \nu$ . Then from commutativity of the convolution product,  $\bar{\mu}^{*n} = \mu^{*n} * \nu^{*n}$  and  $\bar{\mu}_1^{*n} = \mu_1^{*n} * \nu^{*n}$ . Since the support of  $\nu^{*n}$  is included in  $(-\infty, 0]$  and  $\mu^{*n}$  and  $\mu_1^{*n}$  restricted to  $[0, \infty)$  are known and coincide, the measures  $\bar{\mu}^{*n}$  and  $\bar{\mu}_1^{*n}$  restricted to  $[0, \infty)$  are known and coincide. Since  $\bar{\mu} \in \mathcal{C}$ , the measures  $\bar{\mu}$  and  $\bar{\mu}_1$  are equal. Finally, since the characteristic function of  $\nu$  does not vanish on any interval of  $\mathbb{R}$ , see Theorem 2 in [8], we obtain the identity  $\mu = \mu_1$  from injectivity of the Fourier transform.  $\square$

Theorem 5.1 entails in particular that Conjecture C' is equivalent to the following one: *Any distribution  $\mu \in \mathcal{M}_1$  whose support is not included in  $(-\infty, 0)$  is determined by its convolution powers  $\mu^{*n}$ ,  $n \geq 1$  restricted to  $[a, \infty)$ , for some  $a \geq 0$ .* Indeed, it suffices to choose  $\nu = \delta_{-a}$  in Theorem 5.1. Finding more general examples illustrating this result is an open problem. In order to do so, one needs for instance to find the characteristic function  $\varphi$  of a random variable which belongs to class  $\mathcal{C}$  and the characteristic function  $\varphi_Y$  of a nonnegative random variable  $Y$  such that the ratio  $\varphi(t)/\varphi_Y(-t)$  is the characteristic function of some random variable  $X$ . Then since the law of  $X - Y$  belongs to class  $\mathcal{C}$ , so does the law of  $X$  from Theorem 5.1.

Note that neither Proposition 5.1 nor Theorem 5.1 allows us to enlarge class  $\mathcal{C}$ . In order to do so in the same spirit as in Theorem 5.1, one needs to find an invertible transformation  $T(\mu) \in \mathcal{M}_1$  of a distribution  $\mu \in \mathcal{M}_1 \setminus \mathcal{C}$ , such that  $T(\mu)^{*n}$ ,  $n \geq 1$  restricted to  $[0, \infty)$  would be known and such that  $T(\mu)$  belongs to class  $\mathcal{C}$ .

Let us end this paper with an example of a distribution which satisfies conjecture C, although it does not belong to any of the classes studied here. Assume that the support of  $\mu$  is included in  $\mathbb{Z}$  and recall that according to Lemma 2.4, this assumption can be checked from the data of the measures  $\mu^{*n}$ ,  $n \geq 0$  restricted to  $[0, \infty)$ . Assume moreover that there are positive integers  $a$  and  $b$  such that

$$(5.25) \quad \begin{cases} \mu(n) > 0, \text{ for all } n \geq a + b \text{ and } \mu(n) = 0, \text{ for all } n = 0, \dots, a + b - 1, \\ \mu^{*2}(n) = 0, \text{ for all } n = 0, \dots, a. \end{cases}$$

Then we can determine  $\mu$  on  $\mathbb{Z}_-$ , so that  $\mu \in \mathcal{C}$ . Let us first show that  $\mu(n) = 0$ , for all  $n \leq -b$ . Assume that there is  $n \leq -b$  such that  $\mu(n) > 0$ . Then let  $k = 0, \dots, a$  such that  $k - n \geq a + b$ . By definition of the convolution product  $0 \leq \mu(k - n)\mu(n) \leq \mu^{*2}(k)$ , but from our assumptions  $\mu(k - n)\mu(n) > 0$  and  $\mu^{*2}(k) = 0$ , which is contradictory, hence  $\mu(n) = 0$ , for all  $n \leq -b$ . On the other hand, assumptions (5.25) entail that for all  $k = a + 1, \dots, a + b - 1$ ,

$$\mu^{*2}(k) = \sum_{i=a+b}^{k+b-1} \mu(k-i)\mu(i),$$

that is  $\mu^{*2}(a+1) = \mu(-b+1)\mu(a+b)$ ,  $\mu^{*2}(a+2) = \mu(-b+2)\mu(a+b) + \mu(-b+1)\mu(a+b+1)$ , ... Therefore, this system allows us to determine  $\mu(n)$ , for  $n = -b + 1, -b + 2, \dots, -1$  and

the conclusion follows.

Let us consider for instance  $a = 1$ ,  $b = 3$  and

$$\mu(-2) = \mu(-1) = \frac{1-c}{2} \quad \text{and} \quad \mu(n) = \frac{1}{n^3}, \quad n \geq 4,$$

where  $c = \sum_{n \geq 4} n^{-3}$ . Clearly, such a distribution does not belong to any of the classes  $\mathcal{A}$ ,  $\mathcal{M}$  or  $\mathcal{M}_d$ . Then let us check that it does not belong to class  $\mathcal{E}$ . The mean of  $\mu$  satisfies

$$\sum_{k \geq -2} k\mu(k) = -\frac{3}{2}(1-c) + \sum_{n \geq 4} \frac{1}{n^2} < 0,$$

so that (2.3) does not hold. Moreover  $\mu$  has no positive exponential moments. Therefore conditions of Theorems 2.1 and 2.2 are not satisfied and since  $\mu$  is not downward skip free, we obtain the conclusion. Finally it cannot be proved that  $\mu$  belongs class  $\mathcal{C}$  by applying Proposition 5.1 or Theorem 5.1. However, from the above arguments,  $\mu$  does belong to class  $\mathcal{C}$ .

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