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Mechanics Industry

Spectral methods plasticity modeling for fatigue damage estimation: experimental application – uniaxial case

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Abstract – In this paper, we present techniques for fatigue damage evaluation using spectral methods and a model taking into account confined elasto-plastic behavior. The model is associated with a local fatigue approach, covering the whole endurance domain (low cycle and high cycle fatigue). It uses Neuber's method and is valid for limited plasticity. To validate this modeling, we perform a correlation between spectral methods, modified spectral methods and experimental tests. Results presented here are focused on the uniaxial loading case.

Key words: Spectral methods / plasticity / random process / Rainflow Count / fatigue damage

1 1 Introduction

The phenomenon of fatigue has become a major design 2 3 criterion in the engineering industry, requiring fatigue methods that better model real physical phenomena. Con-4 ventional fatigue methods are defined in time domain 5 because their loading input data is usually a stress or 6 strain time history. Algorithms such as Rainflow [1, 2]7 extract stress cycles from the time history to obtain a 8 9 stress distribution. To determine the lifetime generated by the stress distribution, we can use a fatigue law such 10 as Basquin equation and a damage accumulation rule such 11 as Palmgren-Miner law. Nevertheless, these methods re-12 quire knowing the stress time history or the envelope of 13 this time history. 14

15 In the case of random vibrations, the mechanical structure may have a dynamic response. It is difficult to 16 17 know the stress time history and to make the classical 18 fatigue analysis. In time domain, it is only calculated the 19 damage associated with one realization of the random vibrations process. If one wants the damage caused by 20 random processes (in this paper, only stationary ergodic 21 random processes are considered), it is necessary to calcu-22 late damages on several different realizations of the same 23 stochastic process. This approach is called Monte Carlo 24 25 simulation. The main drawback of Monte Carlo simula-26 tions is the computation time [3].

To overcome these problems it is suitable to work in 27 the frequency domain. In addition, we model the random 28 vibrations as a gaussian stochastic process, characterized 29 by its PSD. Fatigue spectral methods are entirely defined 30 in the frequency domain. The average damage is then cal-31 culated from the PSD of stress. The stress cycles distribu-32 tion is not calculated but is estimated with a probability 33 density function. This probability density depends on pa-34 rameters of the PSD and the selected spectral method. 35 Spectral methods have shown their reliability in fatigue 36 life predictions, however, their theory requires several as-37 sumptions about the structure response (Fig. 1a) that 38 limite their applications. 39

Our proposal extends the range of SM to low cycle 40 fatigue, by modeling confined plastic behavior of materi-41 als and we use a fatigue law that take into account this 42 plasticity. Figure 1b defines the framework in which we 43 will expand the use of spectral methods. In the presence 44 of non-linearity, the response of the structure is a station-45 ary ergodic non-gaussian stochastic process. Although the 46 assumption of zero mean is no longer true, we make the 47 assumption that the response of the structure with a non-48 linearity is zero mean because the non-linearity (plastic-49 ity) is confined. 50

To allow comparison between the proposed method 51 and the methods of bibliography, we conducted vibration 52 fatigue tests which will serve as a reference. Tests will be 53 performed in the case of a nearly uniaxial response to be 54 consistent with the definition of the methods studied. 55

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Nomenclature

$E[\bullet]$	Mathematical expectation
f	Frequency (Hz)
$G_{xx}(f)$	Power spectral density of $x(t)$
$R_{xx}(\tau)$	Autocorrelation function of $x(t)$
λ_m	Spectral moment of order m
$ u_0$	Number of zero-upcrossings
ν_a	Average number of maxima
γ	Irregularity factor
N	number of cycles (lifetime)
σ	Stress amplitude
σ_m	mean stress
ε	Strain amplitude
C , k	Material parameters of Basquin model
\mathbf{E}	Young's modulus
ε_f	Fatigue ductility coefficient
σ_{f}	Fatigue strength coefficient
b, c	Material parameters of Manson-Coffin-Basquin model
h(P, V)	Joint probability density function of peak and valley
$P_a(\sigma)$	Amplitude probability density function of $\sigma(t)$
d	Damage at a given amplitude σ
E[d]	Average damage in the interval σ and $\sigma + d\sigma$
E[D]	Average damage
Τf	Life time (seconds, hours or days)
X(t)	Gaussian stochastic process
Z(t)	Non-Gaussian stochastic process
$G(\bullet)$	Direct transformation
$G^{-1}(\bullet)$	Inverse transformation
MOY_{\bullet}	Average of stationary ergodic stochastic process $\bullet(t)$
STD_{\bullet}	Standard deviation of the stationary ergodic stochastic process $\bullet(t)$
S	Skewness
K	Kurtosis
h_3, h_4	Parameters of Hermite model
k'	Cyclic strength coefficient
n'	Cyclic strain hardening exponent
$\varepsilon_{e_i} \sigma_{e_i}$	Strain and stress obtained by elastic behavior for class i
$\varepsilon_{r_i} \sigma_{r_i}$	Strain and stress obtained by Neuber's model for class i
H(ullet)	Cycles distribution with linear behavior
$H_{ep}(\bullet)$	Cycles distribution with elasto-plastic behavior

Acronyms

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PSD	Power Spectral Density
PDF	Probality Density Function
RMS	Root Mean Square
SM	Spectral Method
SMM	Spectral Method Modified
HM	Hermite Model + spectral method

Excitation $\longrightarrow S$	Structure \longrightarrow	Response	Excitation -	\longrightarrow Structure \longrightarrow J	Response
Stationary Ergodic Gaussian Zero mean	Linear	Stationary Ergodic Gaussian Zero mean	Stationary Ergodic Gaussian Zero mean	Nonlinear - Elasto-plastic behavior	Stationary Ergodic Non Gaussian

(a) Assumptions necessary for using spectral methods

(b) Assumptions of our case study

Fig. 1. Definition of required assumptions in spectral methods and positionning of our framework.

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1 2 General purpose

2 2.1 Properties of stochastic processes

3 Power spectral density

In the case of a stationary Gaussian stochastic process, the PSD contains information defining this process [4]. Moreover, if the process is ergodic, the PSD can be calculated from the autocorrelation function of a realization of the random process x(t) as follows, equation (1):

$$G_{xx}(f) = \int_{-\infty}^{+\infty} R_{xx}(\tau) \mathrm{e}^{-i2\pi f\tau} \mathrm{d}\tau \qquad (1)$$

9 where $R(\tau) = E[x(t)x(t+\tau)]$. This equation (1) is known 10 as the Wiener-Khintchine theorem.

11 Spectral moments and signal parameters

12 The spectral moments are important characteristics of 13 a random process, they are calculated by equation (2).

$$\lambda_m = \int_{-\infty}^{+\infty} G_{xx}(f) \mid f \mid^m \mathrm{d}f \tag{2}$$

14 They are related to the amount of information contained 15 in a random process in the time domain. If for exam-16 ple x(t) is a displacement and a random process with 17 zero mean, λ_0 is the variance of the displacement, λ_2 is 18 the variance of the velocity and λ_4 is the variance of the 19 acceleration.

20 Spectral moments determine signal parameters of 21 random processes. These results were demonstrated by 22 Rice [5]. The number of zero-upcrossings is defined in 23 equation (3).

$$\nu_0 = \sqrt{\frac{\lambda_2}{\lambda_0}} \tag{3}$$

24 We define in equation (4) the number of maxima as:

$$\nu_a = \sqrt{\frac{\lambda_4}{\lambda_2}} \tag{4}$$

Another useful parameter is the irregularity factor. It
represents the bandwidth of the PSD. It is calculated
from the number of maxima and the number of zeroupcrossings, equation (5).

$$\gamma = \frac{\nu_0}{\nu_a} = \sqrt{\frac{\lambda_2}{\lambda_0 \lambda_4}} \tag{5}$$

29 2.2 Fatigue law

Spectral methods commonly use the Basquin fatigue
model [6]. This gives a power law relation between stress
amplitude and life. This model leads to a linear StressLife curve in a log-log scale. It is defined in equation (6)
.

$$N\sigma^{\kappa} = C \tag{6}$$

In the case of the proposed approach, the fatigue model used is Basquin-Manson-Coffin's formulation [6]. This is a non linear model that determines total strain amplitude as a function of life. It is governed by equation (7): 38

$$\varepsilon = \frac{\sigma'_{f}}{\mathbf{E}} (2N)^{b} + \varepsilon'_{f} (2N)^{c} \tag{7}$$

2.3 Spectral methods

In a classical approach of damage, counting methods 40 define each stress cycle extracted by a peak of level P and 41 valley of level V or amplitude σ and mean value σ_m . Each 42 extracted cycle is a random event that has its own proba-43 bility of occurrence. Consider h(P, V) the joint probabil-44 ity density of cycles extracted functions of peaks P and 45 valley V. By simple variable change, the joint distribution 46 of cycles is obtained in terms of amplitude σ and mean 47 value σ_m , equation (8): 48

$$P_{a,m}(\sigma,\sigma_m) = 2h(\sigma_m + \sigma,\sigma_m - \sigma) \tag{8}$$

The amplitude probability density function (PDF) is 49 then, equation (9): 50

$$P_a(\sigma) = \int_{-\infty}^{+\infty} P_{a,m}(\sigma, \sigma_m) d\sigma_m \tag{9}$$

Damage due to each stress amplitude σ , according to 51 Basquin equation is equation (10) : 52

$$d = \frac{1}{N} = C^{-1} \sigma^k \tag{10}$$

The expectation of damage contribution of cycles whose 53 amplitude is between σ and $\sigma + d\sigma$, equation (11) : 54

$$E[d] = C^{-1}\nu_a \sigma^k P_a(\sigma) \mathrm{d}\sigma \tag{11}$$

where $\nu_a P_a(\sigma) d\sigma$ is the number of maxima per unit time 55 in the interval $[\sigma; \sigma + d\sigma]$. 56

The damage expectation per unit time for the law of 57 Palmgren-Miner is equation (12). 58

$$E[D] = C^{-1}\nu_a \int_0^{+\infty} \sigma^k P_a(\sigma) \mathrm{d}\sigma \qquad (12)$$

 $P_a(\sigma)$ is the amplitude probability density function and depends on the formulation made by each author. In our case, we use the probability density function (PDF) defined by Lalanne (Rice theory) [7] and Dirlik (Rainflow approach) [8].

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Fig. 2. Methodology diagram for calculating fatigue for non-gaussian process with a reversible transformation of the hermite model type.

1 Lalanne's method

Lalanne [7] defines the amplitude probability density
function from the work of Rice [5], given in equation (13).

$$P_a^{Lal}(\sigma) = \sqrt{\frac{1-\gamma^2}{2\pi\lambda_0}} \exp\left(\frac{-\sigma^2}{2\lambda_0(1-\gamma^2)}\right) + \frac{\gamma\sigma}{2\lambda_0} \exp\left(-\frac{\sigma^2}{2\lambda_0}\right) \times \left[1 + \operatorname{Erf}\left(\frac{\gamma\sigma}{\sqrt{2\lambda_0(1-\gamma^2)}}\right)\right]$$
(13)

4 where $\operatorname{Erf}(x)$ is the error function, defined by $\operatorname{Erf}(x) = \frac{2}{\sqrt{\pi}} \int_0^x e^{-\alpha^2} d\alpha$. The average damage can be calculated 6 from the equation (14).

$$E[D^{\text{Lal}}] = C^{-1} \nu_a \int_0^{+\infty} \sigma^k P_a^{Lal}(\sigma) \mathrm{d}\sigma \qquad (14)$$

7 Dirlik's method

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8 In 1985, Dirlik [9] proposed to determine directly from 9 the PSD the amplitude density probability based on the 10 definition of a rainflow cycle. He determines the rainflow 11 cycles probability density empirically, from a large num-12 ber of Monte Carlo simulations with different forms of 13 PSD and irregularity factors. The probability density by 14 Dirlik is given in equation (15).

$$P_a^{Dir}(\sigma) = \frac{1}{\sqrt{\lambda_0}} \left[\frac{D_1}{Q_D} \exp\left(-\frac{\varphi}{Q_D}\right) + \frac{D_2\varphi}{R^2} \exp\left(-\frac{\varphi^2}{2R^2}\right) + D_3\varphi \exp\left(-\frac{\varphi^2}{2}\right) \right]$$
(15)

where

$$\varphi = \frac{\sigma}{\sqrt{\lambda_0}}; \quad x_m = \frac{\lambda_1}{\lambda_0} \sqrt{\frac{\lambda_2}{\lambda_4}}; \quad D_1 = \frac{2(x_m - \gamma^2)}{1 + \gamma^2}$$
$$R = \frac{\gamma - x_m - D_1^2}{1 - \gamma - D_1 - D_1^2}; \quad D_2 = \frac{1 - \gamma - D_1 - D_1^2}{1 - R} \quad (16)$$
$$D_3 = 1 - D_1 - D_2; \quad Q_D = \frac{1.25(\gamma - D_3 - (D_2R))}{D_1}$$

Equation (17) indicates the average damage with the cycles probability density by Dirlik. 17

$$E[D^{Dir}] = C^{-1}\nu_a \int_0^{+\infty} \sigma^k P_a^{Dir}(\sigma) \mathrm{d}\sigma \qquad (17)$$

2.4 Hermite's model

In case the structure response is a stationary ergodic 19 non-Gaussian stochastic process, there are ways to per-20 form the calculation of the fatigue damage. Hermite's 21 model [10, 11] is a reversible transformation that can 22 transform a non-Gaussian processes in Gaussian process. 23 Combined with spectral methods, this allows fatigue life 24 calculation of a non-Gaussian stochastic process [12, 13]. 25 The global methodology is described in Figure 2. In our 26 case, the non-Gaussian processes can be generated by a 27 structure with a non-linear behavior. Here we use the Her-28 mite model because it is easy to implement but there are 29 other models such as Ochi and Ahn [14,15] and Power-law 30 model [16]. 31

Equation (18) represents the $G(\bullet)$ transformation [17] 32 applied to our problem where X_0 is the reduced centered 33 variable $(X_0 = \frac{X - MOY_x}{STD_x})$. 34

$$X_{0} = \left[\sqrt{\xi^{2}(Z) + \eta_{3}} + \xi(Z)\right]^{\frac{1}{3}} - \left[\sqrt{\xi^{2}(Z) + \eta_{3}} - \xi(Z)\right]^{\frac{1}{3}} - \eta_{1}$$
(18)

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Fig. 3. Proposed fatigue process in the spectral methods.

1 Where

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$$\xi(Z) = 1.5\eta_2 \left(\eta_1 + \frac{Z - MOY_z}{STD_z} \right) - \eta_1^3, \quad \eta_1 = \frac{h_3}{3h_4},$$

$$\eta_2 = \frac{1}{3h_4}, \quad \eta_3 = (\eta_2 - 1 - \eta_1^2)^3, \quad h_3 = \frac{S}{6},$$

$$h_4 = \frac{(K-3)}{24}$$
(19)

2 Equation (20) represents the inverse transformation of G 3 (G^{-1}) :

$$Z = MOY_z + STD_z[X_0 + h_3((X_0^2 - 1) + h_4((X_0^3 - 3X_0))]$$
(20)

4 3 Modeling of plasticity in spectral methods

5 3.1 Modelization in spectral methods

To take into acount elasto-plastic behavior, the spec-6 tral methods need two essential elements to provide a 7 good estimation of the damage. The first element is a 8 fatigue law that takes into account the plastic behavior. g The second element is a model giving the elasto-plastic 10 local responses due to the loading. These two elements are 11 taken into account by using the Basquin-Manson-Coffin 12 13 model and the rule of Neuber [18, 19].

Manson-Coffin-Basquin's law can take into account 14 the plastic behavior of the material during the accumu-15 lation of damage. Neuber's method defines a distribution 16 that takes into account the elasto-plastic behavior from a 17 linear calculation respecting the assumptions of spectral 18 methods (Fig. 1). In spectral domain, it is difficult to ac-19 20 cess to the real strain time history, but for each strain level it is possible to obtain the cycles number corresponding, 21 this is the distribution of cycles number (H function). The 22 23 H function is generally used for the discretized calculus Figure 4 shows how we apply for each stress increment 24 25 the Neuber's method to obtain the real strain. To represent the present phenomenon, we use to apply the correc-26

27 tion of Ramberg-Osgood equation and Neuber equation,

see equation (21).

$$\varepsilon_{ri} \begin{cases} \varepsilon_{ri} \sigma_{ri} = \frac{\sigma_{ei} \sigma_{ei}}{\mathbf{E}} & \text{Neuber} \\ \varepsilon_{ri} = \frac{\sigma_{ri}}{\mathbf{E}} + \left(\frac{\sigma_{ri}}{k'}\right)^{1/n'} & \text{Ramberg-Osgood} \end{cases}$$
(21)

We must add the assumption that the total number of 29 cycles from distribution remains unchanged regardless of 30 the materials behavior [20]. The distribution is defined 31 for a given time interval. In this time interval, the cycles 32 number is given by the system frequency and not by the 33 strain or stress amplitudes. We work only on the alloca-34 tion of the cycles number on the space of possible real 35 strain (Fig. 4). This means that the area under the curve 36 remains the same before and after correction, only the cy-37 cles modification on the high strain modifies the damage 38 [21]. This process is integrated in the process of spectral 39 methods as shown in Figure 3. The assumptions of spec-40 tral methods are respected since it applies in the linear 41 part of the process. 42

4 Tests and results

4.1 Specimen description

The specimen in Figure 5 is designed to meet the 45 assumption of the method described previously (con-46 fined plasticity). It is made from quenched & tempered 47 30NiCrMo8 steel, whose general properties are summa-48 rized in Table 1. The specimen consists of two parts (cylin-49 drical portion and inertia block), crimped together by a 50 shrink fitting process to allow transfer of forces. In Fig-51 ure 5c, we show in blue the clamping area of the specimen. 52 The flat part prevents any rotation of the specimen dur-53 ing the test. The cylindrical portion of the clamping area 54 allows for a recessed well defined and reproducible on all 55 specimens. Table 2 gives the experimental fatigue data 56 for the laws of Basquin and Coffin-Manson-Basquin. The 57 experimental value for the Basquin slope k is 12.03. To 58 avoid dispersion due to choice of damage laws, we chose 59

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Table 1. Mechanical properties of 30NICRMO8.

Young's modulus	Poisson's coefficient	Yield strength	Tensile strength
195 GPa	0.33	945 MPa	1135 MPa



Fig. 4. Correction method for the distribution of cycles number in the spectral methods.

1 to adjust it to the slope of the elastic line from Basquin-2 Manson-Coffin (1/0.085 = 11.764).

3 4.2 Test bench description

Each test is composed of three specimens (Fig. 6) to 4 save time and set repetitive. The specimens are encased in 5 6 a clamping system. The excitation system is a electrodynamic vibrator which is controlled by several parameters 7 8 (RMS value, leveling the PSD). The tests are performed on the second eigenmode of the specimen in order to have 9 a uniaxial loading. The PSD excitation is a PSD band 10 90 Hz-120 Hz, centered around the eigenmode. Table 5 11 in first column presents the different levels of PSD selected 12 for our tests. We have 9 specimens for each level, requir-13 ing 3 tests. Accelerometers are placed on the specimen in 14 order to detect crack initiation. 15

The lifetime is defined as the time between the start of the test and the initiation of the crack. The crack initiation is detected by the variation of the frequency and amplitude of the eigenmode. Table 2. Fatigue properties of 30NiCrMo8 steel.

S N	parametrer		E N p	aramet	er
k	C	b	c	ε'_{f}	σ'_f (MPa)
11.764	$3.659 \times 10e36$	-0.085	-0.752	1.741	1750.000

4.3 Numerical model and modal analysis

To compare the experimental lifetimes for different numerical methods, we modeled the specimen using a finite element code (ABAQUS). This allows retrieving the stress frequency response function which is the input data for calculation methods in fatigue. 25

Numerical model was refined thanks to correlations 26 with experimental data (density, Young modulus and 27 damping). Table 3 shows the experimental eigenfrequen-28 cies from average of measurements performed on 10 test 29 pieces (shock hammer and shaking table), we can see that 30 the gap with the frequency of the numerical model is 31 less than 1% in the case of our study. Figure 7 shows 32 each modal shape with excitation along \hat{Z} , we note that 33 the second eigenmode is orthogonal to the first and third 34 eigenmodes. 35

A convergence study of the mesh was performed in order to ensure good estimation of stresses and strains in the critical area. We refined the mesh until the values of stress and strain does not vary between two refined meshing. We obtained that 41% of the degrees of freedom are located in the notch so that it represents only 5% of the total volume of the specimen. 42

We also performed a damping factor analysis. This pa-43 rameter must be correctly quantified because it strongly 44 affects the values of stress and strain. In our case, the 45 damping is low, the slightest error in estimating it in-46 volves lifetimes completely wrong. To determine a damp-47 ing value consistent across all tests, we calculate the 48 damping of each test during the first two hours of testing 49 where the specimen is healthy. We calculate the damp-50 ing means on all tests. We chose the damping mean as 51 damping value for our calculations. This value is 0.0016 52 or 0.16% damping. 53

4.4 Results and observations

The presented results are in the context of an exci-55 tation of the second eigenmode which generates a nearly 56 uniaxial response. Table 4 summarizes the mean and stan-57 dard deviation of the experimental life times by level of 58 excitation. Means and standard deviations were calcu-59 lated on a set of 36 specimens (9 specimens per level). 60 Data are calculated in seconds and hours. The last col-61 umn of the table indicates the ratio between mean and 62

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(a) Specimen global view

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(b) View of the critical area of (c) CATIA model view with clamping the specimen area





Fig. 6. Schema of test bench with the different elements that compose it and a picture of the test bench with specimens.

Eigenmode	Experimental eigenfrequency (Hz)	Numerical eigenfrequency (Hz)	Damping (%)	Loading type	Loading state (in critical area)
1st	94.5	$94.2 \ (-0.31\%)$	0.11	Bending and Torsion	multiaxial
2nd	102.7	$102.1 \ (-0.58\%)$	0.16	Bending	uniaxial
3rd	248.7	256.5 (+3.14%)	0.48	Torsion and Bending	multiaxial

Table 3. Dynamic parameters (experimental and numerical data).

Table 4. Experimental average lifetime by excitation level well as the associated stantard deviation.

Level of PSD acceleration (g^2/Hz)	Lifetime (initiation)	Average	Standard deviation	Standard deviation/Average
	Hours	4.4	0.8	
0.0040	Seconds	1.6 + E04	2.8 + E03	0.175
	Hours	9.6	1.2	
0.0034	Seconds	3.5E + 04	4.2E + 03	0.120
	Heures	17.6	2.2	
0.0030	Secondes	6.3E + 04	7.8E + 03	0.124
	Heures	29.7	5.0	
0.0026	Secondes	1.1E + 05	1.8E + 04	0.164

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Fig. 7. The deformations of each eigenmode with in hard, the specimen at rest and in transparent, the specimen deformed.

Table 5. Comparaison between the experimental lifetimes and the lifetimes of the different methods. in parenthesis, the relative error from the experimental lifetime.

$\begin{array}{c} Excitation \\ level \\ (g^2/Hz) \\ \end{array} \begin{array}{c} Ez \\ -m \\ lif \end{array}$	Expori				Numerical data				
	montal	montal	RMS	SMM (s)		SM(s)		Hermite (HM) (s)
	life (s)	(MPa)	Dirlik	Lalanne	Dirlik	Lalanne	Dirlik	Lalanne	
0.0040	1.6E + 04	229	1.2E+04 (-25%)	1.4E + 04	1.6E+02 (-99.1%)	1.1E + 02	3.1E+03 (-80.6%)	2.9E + 03	
0.0034	3.5E + 04	211	3.4E + 04 (-2.9%)	3.2E + 04	1.6E+03 (-95.9%)	1.5E + 03	1.4E+04 (-60.0%)	3.5E + 04	
0.0030	6.3E + 04	197	5.7E+04 (-9.5%)	5.4E + 04	9.2E+03 (-85.3%)	8.7E + 03	3.8E+04 (-39.6%)	6.3E + 04	
0.0026	$1.1E{+}05$	175	1.1E+05 (+3.6%)	$1.0\mathrm{E}{+}05$	7.6E+04 (-30.9%)	7.2E + 04	9.5E+05 (-13.6%)	$1.1\mathrm{E}{+}05$	

1 standard deviation. In our case, the ratio is quite low re-

2 gardless of the level of excitation. It appears that there is3 no relationship between this ratio and the level of excita-

4 $\,$ tion. This gives a good level of confidence in the experi-

5 mental results.

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In Table 5, we compare experimental and numerical 6 lifetimes for Lalanne and Dirlik models. In the second col-7 umn, we reported the RMS value obtained by calculating 8 9 the PSD stress. This value is the same regardless of the 10 method selected because this is an input data for fatigue calculation. Third column gives the experimental lifetime 11 average per excitation level. The lifetime average per level 12 is obtained from nine specimens of three tests. The other 13 columns are the lifetimes obtained from the numerical 14 methods. The average lifetimes are calculated from the 15 probability density function (PDF) of Lalanne and Dirlik. 16 17 In case of Dirlik PDF, in parentheses, the values of rela-18 tive errors compared to experimental lifetimes. The relative error is defined in equation (22). 19

$$\frac{\text{Relative error}(\%) =}{\frac{\text{calculated lifetime} - \text{experimental lifetime}}{\text{experimental lifetime}} \times 100 \quad (22)$$

For the proposed model (SMM), the relative errors com-20 pared to experimental lifetimes are low; it shows a good 21 accuracy of the method. In contrast for classical spectral 22 methods (SM), the maximum relative error is 99.1%, it is 23 obtained between the experimental lifetime and the life-24 time with classical spectral method at the level 1. The re-25 sult is consistent because the lower the calculated lifetime, 26 27 the more elastoplastic the material behavior is. Since the spectral methods ignore the elasto-plastic behavior, this 28

29 implies skewed results. Conversely, in the case of level 4,

the loads are small and the material works global in elastic behavior. In reference [22], for the steel 30NiCrMo8, 31 an RMS value of less than 180 MPa involves elastic behavior during the fatigue test. This explains good correlation between classical spectral methods predictions and experimental results. 35

Table 5 results are illustrated in Figure 8. As X-36 axis there are the lifetimes obtained with the different 37 numerical methods. As Y-axis, there are the lifetimes ob-38 tained experimentally. A point on the graph form the 39 combination of a life of a numerical method associated 40 with the experimental life corresponding thereto. A point 41 on the median line means a simulation fitting perfectly 42 experimental data. This allows to see that the main in-43 fluence is due to material nonlinearity; the probability 44 density function selected (Lalanne or Dirlik) has little ef-45 fect. This highlights that it is the material nonlinearity 46 that are in default in classical spectral methods. 47

We also observe that the Hermite model associated 48 with spectral methods tend to give better results at any 49 level than classical spectral methods. Nevertheless, this 50 model diverges when the nonlinearity increases and moves 51 away from the experimental results. The proposed model 52 gives a good estimate of the lifetime of the specimen at 53 any level. This shows the influence of the elasto-plastic 54 behavior on the fatigue life. 55

5 Conculsion and perspectives

The proposed method allows a better estimation of 57 the lifetime by including the influence of the plastic behavior of the material on the fatigue phenomenon. To 59



Fig. 8. Synthesis of different experimental and numerical liftimes for the specimen excited on the second eigenmode.

consolidate the results obtained, other tests will be per-1 fomed on different levels and different type of solicitation. 2 We found that the Hermite model gives good re-3 sults but tend to diverge when the nonlinearity increased. 4 There may be two reasons for this. The first is that the 5 6 method does not model directly the nonlinearity. The second is that the fatigue law used is a S-N law that does 7 8 not take into account the plasticity.

9 We have shown that if the classical spectral methods 10 give incorrect results; this was due to the non-respect of 11 assumptions, in our case, the linearity of the material be-12 havior.

13 For the following, we have several axes of 14 improvement:

- We study the impact of experimental and numerical scatter on the different lifetimes results. This involves both the variability of experimental lifetimes and of numerical lifetimes. We also study the impact of the damping coefficiency the acquired lifetimes.
- damping scatter on the calculated lifetimes.

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- $_{\rm 21}$ We will study and model the inclusion of non re-
- versed strains (non-zero mean of the process studied).
- And, we will focus on the modeling of hysteresis phenomenon of material in the fatigue calculation.
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