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A Robust Statistical Framework for Instantaneous Electroencephalogram Phase and Frequency Analysis

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Abstract

In recent decades, instantaneous phase (IP) and instantaneous frequency (IF) of the electroencephalogram (EEG) has received great attention as a notable complement for conventional EEG spectral analysis. The calculation of these parameters commonly includes narrow-bandpass filtering followed by the calculation of the analytical form of the signal using the Hilbert transform. In this research, using this widely accepted phase calculation approach and well-established methods from statistical signal processing, a stochastic model is proposed for the superposition of narrow-band foreground and background EEG activities. Using this model, the probability density functions of the instantaneous envelope (IE) and IP of EEG signals are derived analytically. It is rigorously shown that the IP estimation quality highly depends on the IE and any phase of frequency interpretations in low IE are unreliable for phase analysis. Based on these findings, a Monte Carlo estimation scheme is proposed for the accurate estimation and smoothing of phase related parameters. The impact of this approach on previous studies including time-domain phase synchrony, phase resetting, phase locking value and phase amplitude coupling are studied with examples.

Keywords: Electroencephalogram phase calculation; time-domain phase synchrony; phase resetting; phase locking value; phase amplitude coupling; phase slipping

1. Introduction

In recent decades, instantaneous phase and frequency of the electroencephalogram (EEG) has received great attention within the neuroscience and brain research community as a notable complement for conventional EEG spectral analysis. The concept has found broad applications in brain computer interface problems [1, 2], brain cognitive studies [3, 4] and brain connectivity assessment [5, 6]. Phase-related quantities have also been widely used in the analysis of clinical and pathological cerebral conditions, including epilepsy [7, 8], dementia [9, 10], autism [11] and many more.

The reliable estimation of the EEG phase and frequency is an essential prerequisite for all these studies. Therefore, various techniques such as the Hilbert transform (HT), wavelet transform (WT), and other time/frequency representations have been utilized for reliable and unambiguous phase calculation (see Section 2). Among all, the HT has been considered as the most reliable tool in EEG phase estimation, as a method which does not violate the physical and theoretical conditions essential for instantaneous phase extraction [12]. However, a major difficulty in using the HT (and other instantaneous EEG phase extraction methods) is to distinguish between physiological state transitions due to brain activity and spurious variations, spikes and phase jumps (known as “phase slipping” [13]). It has been previously reported that unambiguous

measurements of the EEG phase cannot be made around phase slipping epochs, nor near the ends of signal segments [14, 13]. Other aspects of EEG phase analysis, such as temporal and band-pass filtering, have been investigated in [13, 15]. The authors reported notable discontinuities (phase slipping events) in the EEG phase calculated from the human scalp and attempted to minimize these effects. A comprehensive study was performed on various phase extraction tools (particularly HT and WT) and their impacts and interpretations in neural synchrony [16]. In [12], physical and theoretical conditions and steps required for obtaining a meaningful phase sequence were extensively studied, besides investigating the most reliable approaches for this purpose.

Although the problem of reliable EEG phase and frequency estimation has been extensively studied in the neuroscience and brain communities, to the best of our knowledge, some of the fundamental statistical aspects of the problem have not been studied. The objective of the current study is to present a rigorous statistical framework for studying the problem of instantaneous EEG phase and frequency estimation. Using the most widely accepted data model, and profound theories from signal detection and estimation [17, 18, 19, 14], it is shown that issues such as EEG phase slipping and phase/frequency deviations can be fully explained and justified, and fake (unreliable) phase variations can be accurately discriminated from physiolog-

ically relevant EEG phase/frequency variations. The EEG signal's instantaneous analytical form envelope (which has been empirically reported in previous studies) is shown to have a fundamental impact on the accuracy of the EEG phase contents.

In Section 2, the basic steps required for EEG phase extraction are reviewed. The proposed data model for the EEG and the statistical properties of this model are described in Section 3. These statistical measures are used in Sections 4-6 to derive a Monte Carlo method for robust phase-related index extraction. Various applications of the proposed framework and its impact on previous EEG phase studies are presented in Section 7.

2. Instantaneous Envelope, Phase and Frequency Calculation

The classical procedure of calculating the instantaneous envelope (IE), instantaneous phase (IP) and instantaneous frequency (IF) of a signal is to use its analytical form [19]. These measures are only uniquely defined for narrow-band signals. Therefore, considering a discrete-time signal y_n (such as an EEG), it is first bandpass (BP) filtered around the desired center frequency $\omega_0 = 2\pi f_0/f_s$ with a narrow bandwidth bw .

$$x_n = \mathcal{BP}(y_n; p) = y_n * h_n \quad (1)$$

where h_n is the equivalent BP filter's impulse response, f_s is the sampling frequency in Hz, and

$$p = \{type, method, \omega_0, bw, \delta_s, \delta_p, \dots\} \quad (2)$$

is the set of design parameters describing the BP filter, including the filter *type*, design *method*, center frequency ω_0 , bandwidth bw , stop-band attenuation δ_s , pass-band ripple δ_p , etc. The significance of the parameter set p in the current study is studied in Section 4.

The BP filter is a linear transform. Accordingly, the phase of x_n over the entire Nyquist band, is the summation of the input signal's phase and the filter's phase. Therefore, for phase analysis applications, linear-phase finite impulse response (FIR) filters are used to prevent any phase distortion due to the filtering procedure. However, the order of narrow-band FIR filters can become very high. Moreover, the addition of even a linear phase to the input signal can cause fake phase jumps due to the phase wrapping of the filter's phase. To avoid this, we herein use zero-phase forward-backward FIR (or IIR) filtering¹, which does not cause any phase distortions and guarantees the phase synchrony between the input and output of the BP filter.

Next, the *analytical form* of the BP filtered signal is calculated using the discrete-time Hilbert transform $\mathcal{H}\{\cdot\}$.

$$\tilde{x}_n = x_n + j\mathcal{H}\{x_n\} \quad (3)$$

¹Zero-phase forward-backward filtering can be implemented using the *filtfilt* function in Matlab, Octave, or R.

The IE, IP, and IF are calculated from the analytical form as follows.

$$A_n \triangleq \sqrt{\mathcal{Re}(\tilde{x}_n)^2 + \mathcal{Im}(\tilde{x}_n)^2} \quad (4)$$

$$\phi_n \triangleq \text{atan2}[\mathcal{Im}(\tilde{x}_n), \mathcal{Re}(\tilde{x}_n)] \quad (5)$$

$$f_n \triangleq \frac{f_s}{2\pi} [\phi_n - \phi_{n-1}] \pmod{2\pi} \quad (6)$$

where atan2 represents the four-quadrant inverse tangent and the difference operator is used as an approximation for differentiation in (6) (cf. [18] for alternative approximations). In order to avoid, instantaneous jumps in the IF, the IP sequence is commonly *unwrapped* before calculating the difference in (6).

The upper mentioned stages of IE, IP, and IF calculation (BP filtering plus analytical signal calculation) are the main stages of the majority of EEG phase analysis studies. This procedure is fully deterministic and does not account for the stochastic properties of the EEG.

In order to show the motivation of the current study, let us consider a synthetic signal diluted by additive noise, as shown in Fig. 1. The signal model is as follows:

$$y(t) = A(t) \cos(2\pi \int_0^t f(\tau) d\tau) + n(t) \quad (7)$$

where $A(t) = \sin(2\pi \times 1.1t) + 0.7 \sin(2\pi \times 3.3t + \pi/3)$ is a slowly varying amplitude, $f(t) = \sin(2\pi \times 0.8t) + 20$ Hz is the instantaneous frequency and $n(t)$ is white Gaussian noise with a variance adjusted to set the total signal-to-noise ratio (SNR) of $y(t)$ to 25 dB. The data is sampled at $f_s = 400$ Hz for illustration. According to Fig. 1, it is clearly seen that although the amplitude variations are rather slow and the initial phase of $y(t)$ is constant, abrupt random phase jumps and instantaneous frequency errors are inevitable in low analytical signal envelopes (or in low SNR); since in low signal amplitudes, a small additive noise can cause large phase variations (even beyond $\pm\pi$) and randomize the IF. This suggests that the probability of phase jumps and the variance of the IF, somehow increase in low analytical signal envelopes.

This example shows the necessity of a statistical formulation for the problem of EEG IP and IF analysis, which systematically considers the effect of low/high instantaneous envelope and background noise variance.

3. Statistical Properties of Narrow-Band EEG

For a formal description of the problem, a data model is required. Herein, we use the most widely accepted model in the literature.

An implicit assumption in all EEG phase analysis studies is that a desired (narrow-band) brain activity, referred to as *foreground EEG* in the sequel, coexists in a pool of *background EEG*, considered as noise. The background

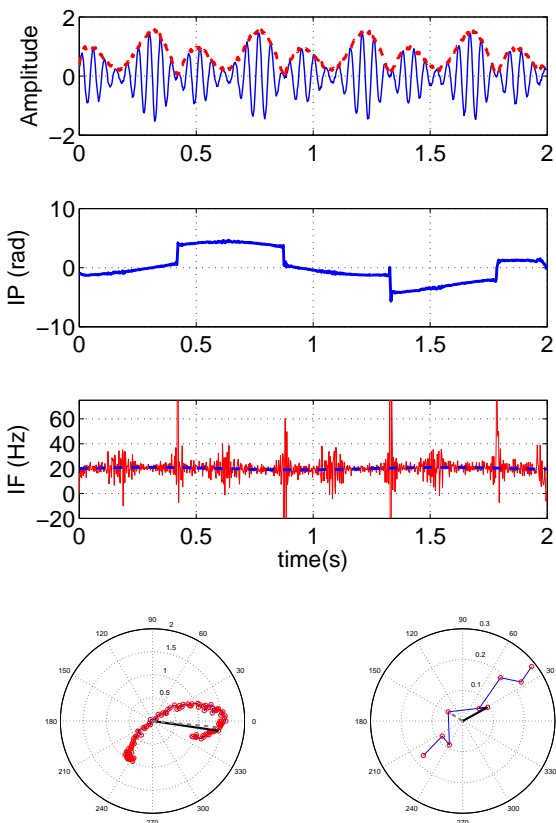


Figure 1: (Top row) a synthetic signal and its analytical form envelope in dashed lines; (second row) the instantaneous phase; (third row) the estimated instantaneous frequency and the true instantaneous frequency in dashed lines; (last row) the phasor representation of a segment of the analytical form in low envelope (left) and a zoom-in during a low envelope segment (right). The two lines starting from the origin, indicate two successive phasors in both cases. Notice the random phase jumps and IF errors in low analytical signal envelopes.

EEG is considered to be spontaneous (as compared with the foreground EEG). It is the superposition of numerous simultaneous and non-coherent cortical activities, which according to the *central limit theorem* tends to a Gaussian distribution and is spread over the entire frequency band of study. On the other hand, the foreground EEG can be the result of some mental activity, external stimuli (audio, visual, etc.), or any other evoked activity, which discriminates it from the spontaneous EEG. Based on these assumptions, the following additive model can be considered as a generic model for the EEG signal at the BP filter output.

$$x_n = X_n \cos(\omega_0 n + \theta_n) + v_n = s_n + v_n \quad (8)$$

where the first part (s_n) models the foreground EEG, X_n is the non-negative foreground envelope with slow variations (lowpass with respect to the center frequency ω_0), θ_n is the instantaneous foreground EEG phase (again with rather slow variations), and v_n is the narrow-band counterpart of the background EEG, which generally overlaps

with the foreground EEG spectrum. In this model, the rather slow variations of X_n and θ_n is the main assumption that distinguishes the foreground and background EEGs.

Due to the linearity of the BP filter, the background EEG remains Gaussian through the filtering process, i.e., $v_n \sim \mathcal{N}(0, \sigma_n^2)$. The objective of EEG phase analysis is to estimate and utilize the phase signal θ_n .

If the BP filter is sufficiently narrow-band (as required for a canonical definition of the IE, IP, and IF [17, 19]), its analytical form can be written as follows.

$$\tilde{x}_n = x_n + j\mathcal{H}\{x_n\} = X_n e^{j(\omega_0 n + \theta_n)} + \eta_n = A_n e^{j(\omega_0 n + \phi_n)} \quad (9)$$

where $\eta_n = r_n e^{j(\omega_0 n + \psi_n)}$ ($r_n \geq 0$) is the analytical form of v_n — a complex valued Gaussian random process with zero-mean independent real and imaginary parts and σ_n^2 variance [20, Ch. 8]. This additive model is perhaps the most generic model for narrow-band EEG and in accordance with the most widely accepted data models in EEG phase analysis studies (cf. [21] for a detailed discussion about the most common EEG phase models). Previous studies have mainly studied the narrow-band foreground part of model (8) within a deterministic framework; neglecting the statistical properties of the background EEG. The basic idea that is followed in this work is that the problem should be studied in a stochastic framework and the signal detectability and reliability of the desired EEG phase totally depend on the instantaneous envelopes and variances of the foreground and background EEG.

3.1. Probability Density Functions of the EEG envelope and phase

The statistical properties of the signal model in (8) have been extensively studied in the signal processing literature [22, 20]. Herein, we use some of these well-known properties to find a statistically justifiable (reliable) method for EEG phase analysis. For simplicity, we assume the sine wave envelope X_n and the noise variance σ_n^2 to be known (or estimated as discussed in Section 3.3).

3.1.1. pdf of narrow-band background EEG

Considering a Gaussian probability density function (pdf) for the background EEG, the magnitude and phase of its analytical form ($\eta_n = r_n e^{j(\omega_0 n + \psi_n)}$) are independent [20]. The magnitude has a Rayleigh distribution and the phase is uniform over $[-\pi, \pi]$.

$$f(r_n | \sigma_n^2) = \begin{cases} \frac{r_n}{\sigma_n^2} \exp\left(-\frac{r_n^2}{2\sigma_n^2}\right), & \text{for } r_n \geq 0 \\ 0, & \text{elsewhere} \end{cases} \quad (10)$$

$$f(\psi_n | \sigma_n^2) = \begin{cases} \frac{1}{2\pi}, & \text{for } -\pi \leq \psi_n \leq \pi \\ 0, & \text{elsewhere} \end{cases} \quad (11)$$

3.1.2. pdf of narrow-band foreground plus background EEG

Due to the Gaussian assumption on the background EEG, it can be shown that the envelope and phase of the analytical form (9) have the following joint pdf:

$$f(A_n, \phi_n, \theta_n | X_n, \sigma_n^2) = f(A_n, \Delta\phi_n | X_n, \sigma_n^2) = \frac{A_n}{2\pi\sigma_n^2} \exp\left(-\frac{A_n^2 + X_n^2 - 2X_n A_n \cos(\Delta\phi_n)}{2\sigma_n^2}\right) \quad (12)$$

where $\Delta\phi_n = \phi_n - \theta_n$ is the instantaneous random phase difference (error) between the noisy and noiseless sinusoidal part of the model (observed versus foreground EEG phase in our case).

In this case, the envelope and phase are no longer independent (a key point in our study). The marginal distribution of the analytical envelope has a Rician distribution.

$$f(A_n | X_n, \sigma_n^2) = \frac{A_n}{\sigma_n^2} \exp\left(-\frac{A_n^2 + X_n^2}{2\sigma_n^2}\right) I_0\left(\frac{X_n A_n}{\sigma_n^2}\right) \quad (13)$$

where $I_0(\cdot)$ is the *modified Bessel function of the first kind* [20, Ch. 8]. Using the Bayes' rule and combining (12) and (13), the conditional pdf of the phase error $\Delta\phi_n$ is found.

$$f(\Delta\phi_n | A_n, X_n, \sigma_n^2) = f(\Delta\phi_n | \kappa_n) = \frac{\exp(\kappa_n \cos(\Delta\phi_n))}{2\pi I_0(\kappa_n)} \quad (14)$$

where $\kappa_n \triangleq A_n X_n / \sigma_n^2$. Apparently, this distribution is symmetric ($E\{\Delta\phi_n | A_n, X_n, \sigma_n^2\} = 0$) and $E\{\Delta\phi_n^2 | A_n, X_n, \sigma_n^2\}$ is an envelope-dependent phase-error variance. These properties are later used in the smoothing stage of the proposed algorithm.

Finally, integrating (12) over A_n , the marginal distribution of the IP error is found.

$$f(\Delta\phi_n | \rho_n) = \frac{1}{2\pi} \exp(-\rho_n^2) \left[1 + \frac{\sqrt{\pi} \rho_n \cos(\Delta\phi_n) \operatorname{erfcx}(-\rho_n \cos(\Delta\phi_n))}{\sqrt{\pi} \rho_n \cos(\Delta\phi_n) \operatorname{erfcx}(-\rho_n \cos(\Delta\phi_n))} \right] \quad (15)$$

where $\rho_n \triangleq \sqrt{\operatorname{SNR}_n} = X_n / \sqrt{2\sigma_n^2}$ denotes the root square of the instantaneous SNR and $\operatorname{erfcx}(\cdot)$ is the *scaled complementary error function*.

The envelope, phase error, and conditional phase error distributions of narrow-band Gaussian background EEG and foreground plus background EEG are shown in Figs. 2, 3, and 4, respectively.

The calculation of the IF pdf additionally requires the joint pdf of θ_n and θ_{n-1} . Due to the lack of ground truth regarding the temporal dependence of the EEG phase, previous results on the IF pdf of sinusoidal signal plus noise [22, 23, 24], which either assume a constant or fully random phase are not directly applicable in this case. A rigorous approach is to model the temporal dynamics of the IP, e.g., by assuming a Markov model for the IP, which is beyond the scope of the current study. However, as a rule of thumb, previous research have shown that the IF pdf

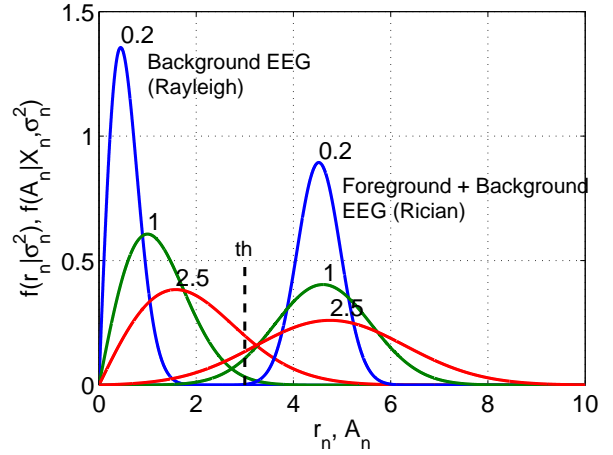


Figure 2: Envelope distributions of Gaussian background EEG (left three curves) and foreground EEG oscillations plus Gaussian background EEG (left right curves) for $X_n = 4.5$ and $\sigma_n^2 = 0.2, 1, 2.5$, equivalent to $\operatorname{SNR}_n = 17, 10, 6$ dB.

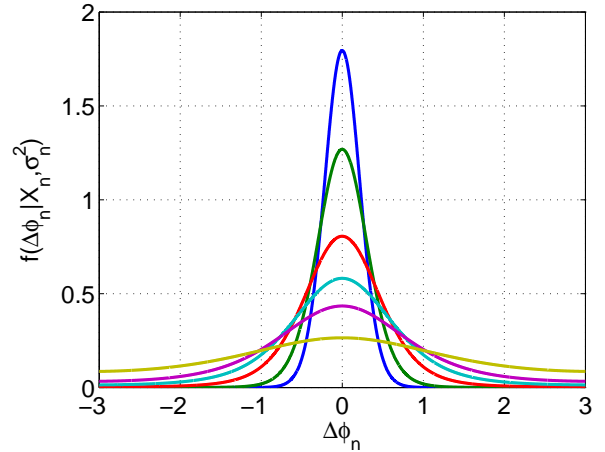


Figure 3: Phase distributions of foreground EEG oscillations plus Gaussian background EEG for $X_n = 4.5$ and from top to bottom $\sigma_n^2 = 1, 2, 5, 10, 20, \text{ and } 100$, equivalent to $\operatorname{SNR}_n = 10, 7, 3, 0, -3, \text{ and } -10$ dB.

becomes sharper around ω_0 as the IP becomes more correlated in time (has slow temporal variations) [23]. Moreover, the IF pdf is peaked around ω_0 in high SNR and it tends to a uniform distribution over $[-\pi, \pi]$ (the entire Nyquist band) in low SNR.

These pdfs have also been used to extract lower bounds on the estimation variance of the IP and IF using different assumptions on the temporal structure of the IP [25, 26]. Again, as a general rule of thumb, the Cramér-Rao lower bound (CRLB) of the IP error variance is commonly inversely proportional to the SNR [27].

In summery, all the probability density functions derived in this section depend on the background EEG variance and the instantaneous analytical signal envelope of the foreground EEG. Specifically, in low analytical signal envelopes, the probability of phase error tends to a uni-

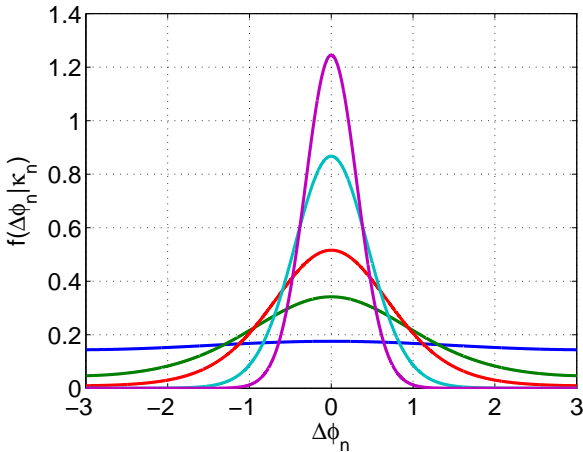


Figure 4: Conditional phase error distributions of foreground EEG oscillations plus Gaussian background EEG for $\kappa_n = 10, 5, 2, 1, 0.1$ (from top to bottom).

form distribution over $[-\pi, \pi]$. Based on these results, a statistical framework is proposed in Section 4, for reliable EEG phase analysis.

3.2. Signal Detectability

Regardless of the presence or absence of a foreground EEG activity, the procedure of IP calculation (narrow-band BP filtering, analytical signal calculation followed by IP estimation) always produces an output; even from pure random background EEG (noise). But how can one guarantee that there has been some foreground activity? Here, we either require physiological ground truth obtained from other modalities (e.g., simultaneous MEG, fMRI, etc.), or complementary information from multiple EEG leads. In absence of reliable ground truth, we face a statistical *hypothesis test*. At each time instant, we have two possible hypothesis:

$$\begin{aligned} H_0 : x_n &= v_n && \text{(background EEG alone)} \\ H_1 : x_n &= X_n \cos(\omega_0 n + \theta_n) + v_n && \text{(background plus foreground EEG)} \end{aligned} \quad (16)$$

The decision is apparently probabilistic. The objective is to determine the most likely hypothesis, which meets some predefined *probability of detection* (p_d) and *probability of false alarm* (p_f).

The detection of a sinusoidal in noise is a classical problem in detection theory [28, Ch. 6]. The hypothesis test is performed by setting an appropriate threshold th on the IE to meet the desired p_d and p_f . By definition and according to Fig. 2

$$p_f = \int_{th}^{\infty} f(r_n | \sigma_n^2), \quad p_d = \int_{th}^{\infty} f(A_n, \phi_n, \theta_n | X_n, \sigma_n^2) \quad (17)$$

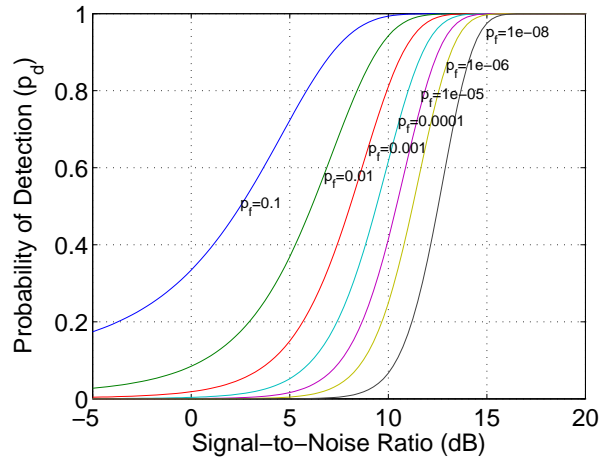


Figure 5: Detection probability (p_d) as a function of the required SNR with false alarm probability (p_f) as a parameter [28, Ch. 6]

which both depend on the instantaneous SNR. The curves for p_d as a function of the required SNR with p_f as parameter (known as the ROC curves) are shown in Fig. 5, for a fixed instantaneous phase $\theta_n = cte$. The same results are approximately applicable for sinusoids with slowly varying phase (as compared with the center frequency ω_0), a condition that is satisfied for narrow-band BP filtered EEG. Some useful rules can be derived from Fig. 5. For example, a 10dB SNR is required to detect a foreground signal with a probability of 90% and $p_f=1\%$. For the same p_f , the probability of detection drops below 0.1 in zeros or negative SNR. In practice, one may fix the desired p_d and p_f and find the corresponding SNR value from Fig. 5. By intersecting this value with the instantaneous SNR, one can determine the time instants for which the existence of foreground EEG activity are most (least) probable.

The utilization of the ROC curves requires the instantaneous SNR (or its estimate), which is not directly available for the EEG. A heuristic approach for estimating the EEG SNR is proposed in Section 3.3.

3.3. Foreground/Background EEG SNR Estimation

The estimation of the instantaneous SNR requires some prior assumptions regarding the foreground and background EEG. Considering the additive data model proposed in Section 3, and considering the foreground EEG (s_n) and background EEG (v_n) to be uncorrelated, superposition holds between their power spectral densities (PSD).

$$S_x(f) = S_s(f) + S_v(f) \quad (18)$$

$S_x(f)$ can be directly estimated from the raw EEG. If we further assume that the background EEG is flat over the narrow-band filter's passband, the background EEG PSD $S_v(f)$ can be estimated from the neighborhood bands of the frequency band of interest (as shown in Fig. 6), and $S_s(f)$ is found by subtracting $S_v(f)$ from $S_x(f)$. The area under these spectral estimates is equal to their total power.

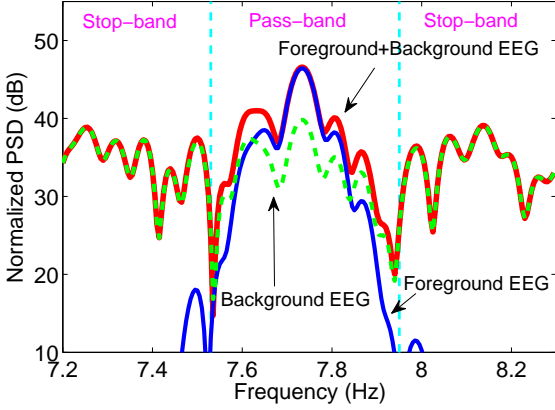


Figure 6: Normalized power spectral density (PSD) of a sample EEG segment

Therefore, the SNR in the filter's passband is obtained. The same method can be performed online, by estimating the PSD and SNR_n over a sliding window of the input signal.

An alternative approach for SNR estimation is to use oscillatory signal tracking schemes (similar to the one developed in [29] for electrocardiogram denoising applications) to estimate the foreground EEG from x_n . Having an estimate of s_n , the background EEG and the instantaneous SNR can be found.

4. Robust EEG Phase Analysis

To this point, we have shown that the instantaneous EEG phase is a random variable, with an uncertainty that depends on the instantaneous signal envelope. In this section a statistical method is proposed for finding robust estimates and confidence intervals for the IP and IF. Our basic assumption is that the phase and envelope characteristics of the foreground EEG should be rather consistent and robust to minor perturbations of the bandpass filter's design parameter set p (bandwidth, center frequency, etc.). This assumption is supported by the fact that the sub-frequency bands defined in the EEG phase analysis literature are rather subjective and minor deviations in the center frequencies and bandwidths are implicitly assumed to be irrelevant to the phase study. To the best of the authors' knowledge, only some general guidelines have been proposed for the bandpass filter specifications [13, 30], and researchers have been basically using arbitrary and un-unified bandpass filtering schemes in their studies. A similar effect is expected if the input signal is *dithered* by minor additive noise, at the order of the background EEG level. The foreground EEG is again expected to remain consistent under such minor dithers.

Based on these assumptions, for a given EEG signal y_n , the following procedure is proposed for finding robust values and statistical confidence intervals for IP and IF esti-

mates. The idea is inspired from Monte Carlo estimation methods in statistics.

1. *Random Ensemble Generation:* For $k = 1, \dots, K$ repeat the following stages:

- (a) Add minor Gaussian noise to the input EEG, to obtain dithered randomized ensembles of the EEG $d_n^{(k)} = y_n + \nu_n^{(k)}$; where $\nu_n^{(k)} \sim \mathcal{N}(0, \epsilon_n^2)$. The dither variance ϵ_n^2 should be chosen at the order of the background EEG power, which can be estimated from the neighborhood spectral bands (Fig. 6), or left as a fine-tuning parameter.
- (b) Bandpass filter $d_n^{(k)}$ using filters with infinitesimal deviations in their parameter set $p^{(k)}$ (bandwidth, center frequency, ripples, etc.)

$$x_n^{(k)} = \mathcal{BP}(d_n^{(k)}; p^{(k)}) \quad (19)$$

As shown in the results, the filter parameter deviations can simply be confined to the center frequency and bandwidth and as low as 0.001 Hz to 0.01 Hz, which are expected to be irrelevant for any EEG phase analysis application.

- (c) Calculate the analytical form of each ensemble $\tilde{x}_n^{(k)} = x_n^{(k)} + j\mathcal{H}\{x_n^{(k)}\}$.

2. *Parameter Estimation:*

- (a) Calculate the IE, IP, and IF of each ensemble

$$X_n^{(k)} = |\tilde{x}_n^{(k)}| \quad (20)$$

$$\theta_n^{(k)} = \text{atan2}[\mathcal{I}m(\tilde{x}_n^{(k)}), \mathcal{R}e(\tilde{x}_n^{(k)})] \quad (21)$$

$$f_n^{(k)} = \frac{f_s}{2\pi} [\theta_n^{(k)} - \theta_{n-1}^{(k)}] \text{ mod } 2\pi \quad (22)$$

- (b) Find the ensemble averages and variances of the parameters

$$\bar{x}_n = E\{\tilde{x}_n^{(k)}\}, \quad \lambda_n^x = E\{(\tilde{x}_n^{(k)} - \bar{x}_n)^2\} \quad (23)$$

$$\bar{X}_n = E\{X_n^{(k)}\}, \quad \lambda_n^X = E\{(X_n^{(k)} - \bar{X}_n)^2\} \quad (24)$$

$$\bar{\theta}_n = E\{\theta_n^{(k)}\}, \quad \lambda_n^\theta = E\{(\theta_n^{(k)} - \bar{\theta}_n)^2\} \quad (25)$$

$$\bar{f}_n = E\{f_n^{(k)}\}, \quad \lambda_n^f = E\{(f_n^{(k)} - \bar{f}_n)^2\} \quad (26)$$

where $E\{\cdot\}$ represents expectation over the K ensembles. The calculated variances are used to find $\sigma_n^X = \sqrt{\lambda_n^X}$, $\sigma_n^\theta = \sqrt{\lambda_n^\theta}$, and $\sigma_n^f = \sqrt{\lambda_n^f}$

as confidence intervals (percentiles) for the estimated parameters. For asymmetric densities (as the IF can be), the right and left μ -percentiles, denoted by $\sigma_n^{\mu+}$ and $\sigma_n^{\mu-}$, respectively, can be calculated and monitored².

For illustration, Fig. 7 shows 30 s of a typical EEG, recorded in a brain computer interface (BCI) experiment; together with its IE, IF and unwrapped IP minus $2\pi f_0 t$ (to permit the monitoring of θ_n in the data model (8)), calculated according to the above detailed procedure for $K = 50$ iterations. The sampling frequency of the signal was $f_s = 160$ Hz. The BP filter was a zero-phase forward-backward filter obtained from an order six moving average lowpass filter prototype. The center frequency of the filter was $f_0 = 10$ Hz and the effective bandwidth was 1 Hz. In each of the K trials, this filter was perturbed with random deviations of the bandwidth ranging from zero to 0.5 Hz, the center frequency f_0 was perturbed with a uniform random value in the range of $f_0 \pm 10^{-6}$ Hz, and the dither $\nu_n^{(k)}$ in each iteration was a zero mean Gaussian random variable with a standard deviation $\epsilon_n = 10^{-4}$. The gray shades show the results of K trials overlaid. The average instantaneous frequency \bar{f}_n and the confidence intervals bounded by $\bar{f}_n \pm \sigma_n^f$ are also shown.

Accordingly to Fig. 7, the deviations of the IP and IF significantly increase in low analytical signal envelopes. In fact, in low IE, the confidence interval is wide, which means that the IP can take any value in the range of $[-\pi, \pi]$ (and the IF in the range of $[-f_s/2, f_s/2]$), without any physiological significance and merely due to a low IE.

4.1. Parameter Selection

The proposed framework is generic and its parameters should be selected for each application. Specifically, the BP filter specifications such as the center frequency, bandwidth, and design method are all subjective and may be selected according to physiological ground truth (as considered in all previous studies). For each application, the parameter deviations considered during the Monte Carlo simulation are selected such that the deviated filter specifications would be physiologically irrelevant for the application of interest. For example, consider an EEG with a dominant 10-12 Hz (alpha-band) spectral density. Physiologically, 0.1 Hz of filter bandwidth deviation is considered irrelevant for such experiments (at least according to the current literature). Therefore, randomization of the BP filter bandwidth at this level will help the identification of reliable and unreliable EEG parameters.

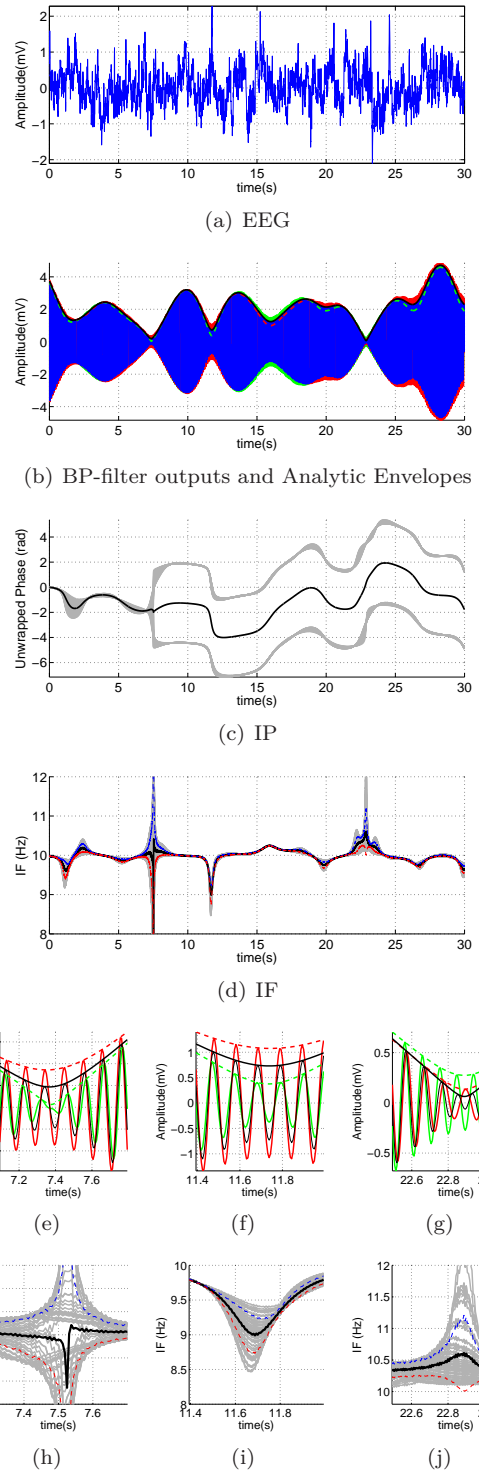


Figure 7: (a) A sample EEG segment; (b) the bandpass filtered signals using three BP filters with $f_0 = 9.99, 10, 10.01$ Hz and the corresponding analytical signal envelopes; (c) the unwrapped IP of fifty perturbed ensembles with their highlighted mean; (d) the IF of fifty perturbed ensembles with \bar{f}_n , and $\bar{f}_n \pm \sigma_n^f$ confidence intervals; (e)-(g) a zoom-in of three IE segments where the black, red and green traces correspond with the output of BP filters centered at $f_0, f_0 - \delta f$ and $f_0 + \delta f$, respectively, and envelopes shown in thick black, dashed red and dashed green; (h)-(j) zoom-in of three IF segments.

²In our notation, the right and left percentiles represent $\Pr\{X \leq \sigma^{\mu-}\} = (1 - \mu)/2$ and $\Pr\{X \geq \sigma^{\mu+}\} = (1 - \mu)/2$. For symmetric distributions $\sigma^{\mu-} = \sigma^{\mu+}$.

5. Interpretation of Instantaneous Envelope, Phase and Frequency Estimates

The results of previous sections has some major implications for a systematic interpretation of the EEG phase, including:

1. The confidence in signal detection, IP and IF estimations is directly related to the instantaneous SNR and envelope (X_n). In fact, considering a stationary background EEG variance ($\sigma_n = cte.$), in low foreground EEG envelopes (X_n), the accuracy of signal detection drops, the probability of phase slipping increases, and the IP and IF estimation qualities degrade. This suggests that the IP and IF are only reliable in high SNR, and it is not possible to validate any phase or frequency activity without considering the instantaneous envelope of the analytical form of the EEG simultaneously.
2. The EEG IP, and IF are stochastic parameters with SNR-dependent pdfs. As random variables, they can fluctuate within the range of their pdf. The confidence intervals of these parameters imply that as far as the estimated IP and IF are within the $\bar{\theta}_n \pm \sigma_n^\theta$ and $\bar{f}_n \pm \sigma_n^f$ ranges, the event is considered normal. An odd event— having perhaps a physiological origin, such as phase resetting, etc.— may only be reported when the IP or IF exceed the normal range. This shows that any phase or frequency jumps that occur in low analytical signal envelopes are statistically irrelevant and may not necessarily be associated to any biological origin. More generally, using the pdfs derived in Section 3.1, percentiles can be estimated for each of the estimated parameters to find the probability of a given IP or IF.
3. The randomized ensembles generated from the above procedure can be considered as samples driven from the density functions derived in Section 3.1. Considering the symmetry of the IP distribution, $\bar{\theta}_n$ yields the maximum a posteriori (MAP) estimate of the IP. For the IE (due to the right skew of its distribution) and IF (due to its possible pdf asymmetry, which depends on the temporal dynamics of the EEG phase), instead of taking the expectation of the randomized ensembles, the sample *mode* can be used to obtain the MAP estimate (which are no longer equal to \bar{X}_n or \bar{f}_n).
4. Considering the fundamental role of the IE in phase analysis, a hypothesis that requires future studies is that many of the phase inferred parameters of the EEG, may in fact be associated to the analytical signal envelope, rather than the EEG phase. In other words, phase-related parameters such as phase resetting may be side effects of the signal's envelope drop, rather than being an independent measure.

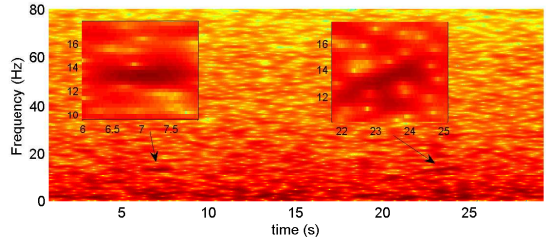


Figure 8: The short-time frequency transform (STFT) of a sample EEG. The left highlighted segment shows a vanishing component at a fixed IF; the right highlighted segment shows a component with varying IF, corresponding to scenarios (A) and (B) in Section 5.1, respectively.

5.1. Simultaneous Envelope, Phase and Frequency Variations

The implication of the previous study is that the instantaneous analytical signal envelope is highly fundamental for phase or frequency variations and phase and frequency measures become unreliable in low IE. However, further considerations are required for reliable interpretation of the IE. Consider the following scenarios, which both lead to a low analytical signal envelope:

- A) Stationary IF (fixed phase shift) with a dropping IE
- B) Stationary envelope with a varying IF (variable phase shift)

The two scenarios are shown in the short-time frequency transforms of a sample EEG segment in Fig. 8. In the first scenario, the IE is explicitly dropping; while in the second scenario, due to the variation of the IF, the signal is moving out of the narrow BP filter's pass-band, which indirectly results in an IE drop. In both cases, the estimated IP and IF become unreliable. However, only the first scenario can be considered as a statistical burden (without any physiological source); while the second scenario can frequently happen for an EEG to have minor fluctuations in the IF mode (considering the narrow-band of the BP filter). In order to discriminate these scenarios, the analytical signal envelope can be tracked at the output of a filter bank with three filters: a BP filter centered at f_0 and two BP filters with minor center frequency deviations, centered at $f_0 - \delta f$ and $f_0 + \delta f$. In the first scenario the envelope at all three filter outputs will drop; while in the second scenario the envelope of the main BP filter (centered at f_0) will drop and the envelope of either the right or left vicinity frequency bands will increase. This simple frequency tracking scheme is a computationally efficient alternative for time-frequency analysis of the EEG.

Fig. 7(b) illustrates the idea. A sample EEG has been passed through three zero-phase BP filters with the same bandwidth (1 Hz) and center frequencies 9.99 Hz, 10 Hz, and 10.01 Hz ($\delta f = 0.01$ Hz). The resulting signals have been overlaid for comparison, with three highlighted segments shown in Figs. 7(e), 7(f) and 7(g). Due to the

zero-phase (zero group delay) property, the input and output of these filters are phase synchronous. Accordingly, Fig. 7(e) shows a significant IE drops in all three BP filter outputs. Considering the IF in the same segment in Fig. 7(h) we notice the totally random behavior of the IF (positive and negative fluctuations around f_0), due to the very low IE of this segment. Fig. 7(f) shows a minor IE drops in all three BP filter outputs. Considering the IF in the same segment in Fig. 7(i) we notice that although the IF has a high variance in this segment, but all the randomized ensembles report an instantaneous frequency drop in this segment. The same phenomenon has occurred in Fig. 7(g) and Fig. 7(j); but this time with IF tendency towards frequencies above f_0 .

6. Post-Processing: Temporal Filtering of Instantaneous EEG Parameters

The ensemble averaging technique proposed in Section 4 does not use the temporal correlations of the IE, IP, or IF. By assuming a temporal dynamic model for these parameters, the estimation quality may be improved by conventional filtering and smoothing schemes such as the Kalman filter or Particle filter (also known as Sequential Monte Carlo). Herein, for proof of concept, we assume a first order auto-regressive dynamic model for the IP and apply a classical Kalman filter to improve the IP and IF estimates.

$$\begin{aligned}\theta_n &= \theta_{n-1} + w_n \\ \phi_n &= \theta_n + \psi_n\end{aligned}\quad (27)$$

where $w_n \sim \mathcal{N}(0, \gamma_n^2)$ is considered as process noise, and $\psi_n \sim \mathcal{N}(0, \sigma_n^2)$ is the observation noise. For this model, it can be shown that the Kalman filter equations have a single tunable parameter σ_n^2/γ_n^2 , which can be adjusted proportional to the phase and frequency variances obtained in (25) and (26). In Fig. 9, the IF of the sample EEG segment from Fig. 7 is shown after applying the Kalman smoother for $\sigma_n^2 = \lambda_n^f$ (as defined in (26), and two values $\gamma_n = 10^{-3}$ and $\gamma_n = 10^{-5}$. Apparently, the results are smoother for smaller values of γ_n (the filter relies on the assumed dynamics) and become closer to the observed noisy IF for larger values of γ_n . This example is only shown as proof of concept and the rich literature on Kalman filtering and its extensions can be used in future studies for robust IP and IF estimation.

7. Applications

The current study has significant implications in various EEG phase analysis applications. In this section, some of the major applications are studied with examples. In order to make the results reproducible, all source codes related to this study are online available in the open-source electrophysiological toolbox (OSET) [31]³.

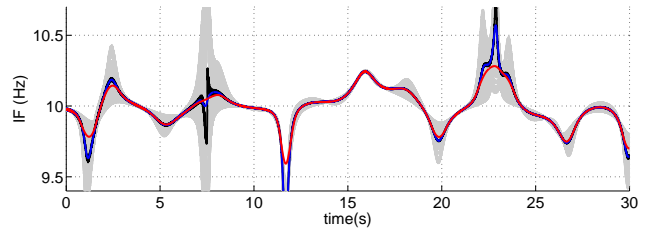


Figure 9: The estimated IF of 50 ensembles overlaid (gray shades); the average IF (black); the KF-smoothed IF with $\gamma_n = 10^{-3}$ (blue) and $\gamma_n = 10^{-5}$ (red).

7.1. Time Domain Synchrony and PLV

The synchronous firing of neurons within different brain regions has been considered as a potential source of EEG-level brain responses. Synchrony estimation consists of first calculating the phase sequences using electric or magnetic brain records and then quantifying the local stability of phase lock through computing the difference between these sequences [32]. In this context, phase-locking value (PLV) is one of the most common indexes used for calculating the coherency and synchrony between phase signals [32, 33]. The index is based on the phase difference (PD) between two EEG signals.

To show the importance of the proposed statistical approach in phase synchrony studies, the IP, IF and IE of five 20 s segments of simultaneous EEG channels are shown in Fig.10.

As illustrated in Fig.10, the initial observation shows a great synchrony between the IP of O1, P3, C3 and partially F3 (middle panel). Moreover, considering the FP1 as a reference signal, the inter-channel phase differences (PLVs), are showing a sort of *phase lock* or significant *phase shift* in the highlighted time epochs (indicated by boxes in Fig.10). Nevertheless, when comparing these plots with their corresponding IE (in the left panel of Fig.10), it is noticed that the spikes in the phase indexes are concurrent with low IE. Therefore, the phase effects are side-effects of a low analytical signal envelope during the same epochs. More specifically, the boxes indicated as (a) and (b) in the left panel have been selected such that the IE captured from electrode FP1 (the reference lead) includes two notches. During the same period, two big spikes can be observed in its corresponding phase derivatives in the middle panel of FP1, which result in notable displacement (shifts) in the calculated phase differences (PLVs). On the other hand, box (c) is selected such that a low-value analytic envelope in four other electrodes is concurrent with a high envelope in FP1. Again, the phase differences are significantly affected by unreliable spikes in phase sequences due to low IE in these channels. Apparently, as discussed before, measurements of instantaneous phase in time instants with low IE is unreliable and any phase-related quantities such as phase lock, phase difference, or PLV are affected by this issue and any physiological and pathological interpretations based on these quantities require further consideration.

³All source codes related to this paper shall be provided online after the publication of the current study.

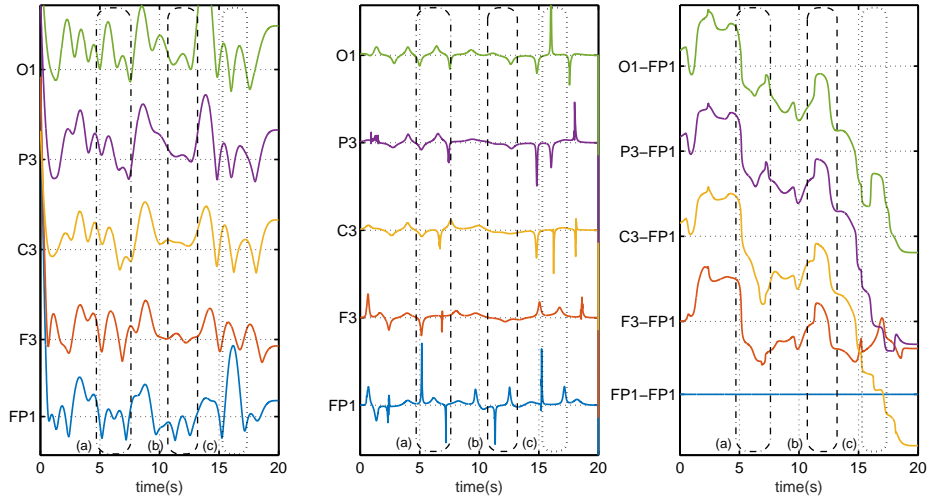


Figure 10: Time domain synchrony (middle panel), the corresponding PLV index (right-side panel), and the instantaneous analytical signal envelope (left-side panel). Notice the estimations affected by low values of analytic envelope .

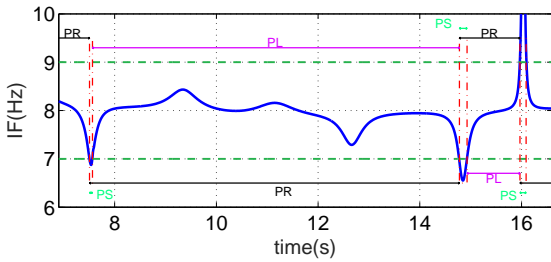


Figure 11: Illustration of Phase Resetting (PR), Phase Lock (PL) and Phase Shift (PS) events using the IF and IP of an EEG segment. The dashed green lines indicate the tunable threshold used for measuring PS [34].

7.2. Phase Resetting

Phase resetting has been defined as a phase shift (PS) followed by a phase difference stability, i.e., phase lock (PL) [14]. Each pair of PS and PL, starting from the beginning of a PS and finishing by the end of a PL (beginning of the next PS), is called a phase reset (PR). Fig. 11 shows the PR, PL and PS events calculated for a typical EEG.

Using the concepts presented in Fig. 11, in Fig. 12 phase shifts are estimated for segments of two EEG channels (FP1 and O1) together with their IE, IF, PD and the first order time difference of the PD. Accordingly, among all the detected phase shift events (curves crossing the green dashed *threshold* in second panel), only two of them have occurred in significantly high IE (may be considered relevant). Furthermore, according to the third and fourth panels, the phase shifts calculated between two channels are only relevant only during the epochs where the IE of both channels have been above the threshold (blue boxes) and the rest of the detected phase shift events (red boxes) are unreliable, due to the low-envelope analytic signal.

Phase resetting in various frequency bands of cere-

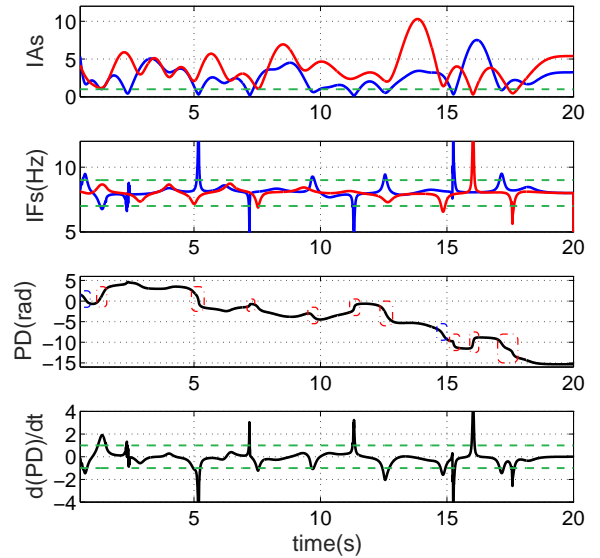


Figure 12: Phase resetting (phase shift) measurements affected by the low values in analytic envelope.

bral signals have been previously correlated to different cognitive responses such as working memory [35, 36], brain development [37], intelligence [38, 39], consciousness [40, 41, 42, 43], sensory-motor interactions [44] and many more [45, 46]. While many of the phase shifts and resettings have been during unreliable IE magnitudes. This raises some reservations regarding physiological and pathological interpretations which have been based on phase resetting of brain signals and highlights the necessity of simultaneous IE analysis for such applications, either as a complement or replacement for phase analysis.

7.3. Phase-Amplitude Coupling (PAC)

Phase amplitude coupling (PAC) is a means of investigating the coupling between the phase of lower frequency oscillations and the power of higher frequency oscillations, i.e., synchronization of IE of faster rhythms with IP of slower rhythms [47]. Due to the algorithmic details of the PAC, its study is beyond the scope of the current study. However, many aspects of PAC, including the calculation of the IP and IE in different frequency bands can be studied within the hereby proposed scheme. Specifically, the current study emphasizes the necessity of a Monte Carlo estimation of the IE, IP and IF for PAC studies to assure the statistical relevance of the calculated index.

8. Discussion and Conclusion

In this study the classical procedure of instantaneous EEG phase and frequency analysis was studied in a stochastic framework, using the most widely accepted data model representing foreground and background EEG activity. The probability density functions of the instantaneous phase and envelope and their dependence on the instantaneous SNR of the EEG were derived. By using minor perturbations in the BP filtering scheme, it was shown that the EEG phase parameters are highly dependent on the IE and are statistically unreliable in low analytical signal envelopes. The impact of this framework was shown for EEG IP and IF calculation, and well-known phase based parameters such as phase synchrony, PLV and phase resetting. The study raises some major reservations on the interpretation of previously reported physiological factors, which have been derived from the EEG phase alone (neglecting the envelope information). Considering the high impact of the IE on IP and IF, a fundamental question is whether phase related indexes can be considered as independent cerebral factors, or they are merely side effects of the IE variations. The answer to this question requires a statistical setup on a large dataset recorded under well-defined brain experiments.

In future studies, the hereby proposed randomization procedure and post-processing proposed in Section 6 can also be unified using a particle filtering or sequential Monte Carlo filtering scheme, which performs randomization and smoothing at the same time and provides a sample-based MAP estimate for EEG phase and frequency estimates.

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