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How big is the comeback? Japanese exchange rate pass-through assessed by Time-Varying FAVAR

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How big is the comeback? Japanese exchange rate pass-through assessed by Time-Varying FAVAR

ZAKARIA MOUSSA*

March 2, 2016

Abstract

We examine the evolution and the magnitude of exchange rate pass-through (ERPT) to Japanese prices. We employ the Time-Varying-Parameters Factor-Augmented Vector autoregression model (TVP-FAVAR), which enables us to include a large enough of data to better control for variables impacting prices and exchange rate. Our results confirm the decline in ERPT rates until the late 2000s and their resurgence in last years. Our findings provide additional support to the notion that exchange rates can impact import and domestic prices, possibly helping avoid deflation. We also find that the ERPT into aggregate prices can hide considerable variation in price sensitivity to exchange rates across stages of demand and industries. ERPT decreases along the pricing chain, from imported raw materials and intermediate goods prices to domestic prices. Finally, we find price sensitivity consistent throughout, indicating that incorporating extra information leads to more robust estimates.

JEL classification: E31, F31, F41.

Keywords: exchange rate pass-through; Japan; import prices; domestic prices; TVP-FAVAR model.

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1 Introduction

The exchange rate pass-through (ERPT) in domestic prices has recently been attracting attention again. Exchange rates may offer a plausible alternative monetary policy instrument in the current zero interest rate environment. The global financial crisis forced most major central banks to employ unconventional monetary policy tools. These included quantitative and qualitative easing, alone or in combination, as well as an explicit commitment to maintain low nominal short-term interest rates. (Bernanke and Reinhart, 2004; Walsh, 2009; Woodford, 2012) There was skepticism regarding the employment of the exchange rate by central banks (McCallum, 2006), even though this was part of the unconventional tools already proposed to the BoJ (McCallum, 2000, 2006; Svensson, 2001; Coenen and Wieland, 2003). Currently, exchange rate depreciation is increasingly used to fight deflation (Borio and Disyatat, 2010; Ishi et al., 2011; Franta et al., 2014; Shioji, 2012, 2014, 2015), which raises the question of how high ERPT needs to be before this tactic is effective.

An extensive literature revealed the decline of the exchange rate pass-through into domestic prices across many countries from the early 1990s (Taylor, 2000; Campa and Goldberg, 2005; Goldberg and Campa, 2010; Sekine, 2006). Declining ERPT into Japanese domestic prices is explained in two main ways by empirical studies. Otani et al. (2003, 2006); Sekine (2006) blame it chiefly on changing price-setting behavior of Japanese firms and pricing to market by foreign suppliers. In **contrast**, Ito et al. (2015) claim that Japanese firms' choices of invoicing currency are the major driver of ERPT decline. They show that the increase in intra-firm trades results in more Local-Currency Pricing (LCP), which likely reduces exchange rate pass-through.

However, recent papers by Shioji (2012, 2014, 2015); Hara et al. (2015) provide evidence of the resurgence of exchange rate pass-through into Japanese domestic prices in recent years. Shioji (2012) explores how yen movements are transmitted to domestic prices by using disaggregate data either on import or on domestic prices. A rise in ERPT was detected again in the late 2000s, suggesting that the Bank of Japan's aggressive monetary easing, launched to support the Abenomics in 2012, might be more successful at defeating deflation

through the impact of yen depreciation. [Shioji \(2014, 2015\)](#) attribute the resurgence of ERPT to changing production structure and an increasing share of imported intermediate goods in production. Moreover, [Shioji \(2015\)](#) shows that ERPT to consumer prices is more pronounced for frequently purchased goods. No investigation is made of a possible price setting behavior effect, as behavior is assumed to be unchanged over the sample period. [Hara et al. \(2015\)](#) provide additional evidence of recent comeback of ERPT. They put this down not only to a changing production structure, but also to price setting behavior, which may be a major factor in this resurgence.

One strand of the literature examines the impact of exchange rate shocks to import and domestic prices, using vector autoregression (VAR) models ([Hahn \(2003\)](#), [Choudhri and Hakura \(2006\)](#) and [Ito and Sato \(2008\)](#), among others). The notable advantage of VAR methodology, as explained in [Ito and Sato \(2008\)](#), is that it avoids the endogeneity problem inherent in single-equation pass-through estimation ([Campa and Goldberg \(2005\)](#), [Goldberg and Campa \(2010\)](#), [Hara et al. \(2015\)](#), among many others) and thus accounts for the interrelation between exchange rate and prices. More recently, a growing literature has been examining ERPT using Time-Varying Parameters (TVP)-VAR models. Coefficients and the variance-covariance matrix of error terms are allowed to vary over time, thus making it possible to identify changes in ERPT at each point in time. [Shioji \(2012\)](#) estimates different Time-Varying Parameters (TVP)-VAR models separately for different stages in the transmission of exchange rate movements along the Japanese pricing chain. Results show significant ERPT variability over time.

This paper reexamines the evolution of the exchange rate pass-through, extending [Shioji \(2012\)](#) in two principal respects using the TVP-FAVAR model developed by [Koop and Korobilis \(2010\)](#) over a data set spanning the period 1985Q1-2015Q2. First, as clearly stated in ([Bernanke et al., 2005](#)) and generally recognized, the limited information included in a simple VAR system can lead to bias estimation results. Adding more macroeconomic and financial information, which may affect either exchange rate or prices, has the potential to resolve many inconsistencies observed empirically. Biased estimates of the pass-through coefficient could arise if, for instance, aggregate demand, the size of the country, the pene-

tration ratio, price-setting behavior of foreign exporters and monetary policy environment are excluded from the analysis (Taylor, 2000; Smets and Wouters, 2002; Devereux et al., 2004; Sekine, 2006).¹ Moreover, relatedly, instead of estimating (TVP)-VAR models separately for different stages of the ERPT along the pricing chain, the (TVP)-FAVAR model allows us to analyze the effect of the exchange rate variations at different stages from import to domestic prices in a unified framework. Second, as pointed out by Sato (2014) and Shioji (2015), analyzing ERPT through aggregate prices masks a large variation in price elasticity to exchange rate movements across different stages of demand, as well as across industries or categories. Therefore, to explore and understand how exchange rate movements affect different prices throughout the Japanese economy, our study includes disaggregate prices both by stage of demand and by industry-breakdown on import and domestic price data series.

We report three main findings: First, our results clearly show that sensitivity to change in the exchange rate varies across disaggregate prices in terms of both magnitude and time variation. However, in most cases, price sensitivity follows a downward trend starting from 1985Q1 to the late 2000s, prices subsequently becoming more sensitive to exchange rate changes. This confirms the resurgence of ERPT detected in recent studies. Although sensitivity was still not particularly high, what is important is the trend shift starting from the late 2000s, which may become persistent in future. Second, the incomplete short- and long-run ERPT detected for different prices is decreasing along the pricing chain. In addition, the higher ERPT into import final goods, compared with raw and intermediate goods, suggests that the growing Japanese production network and invoicing choices may impact intra-firm trade in terms of price sensitivity to exchange rate movements. Overall, our results confirm the resurgence of ERPT to Japanese prices detected in the recent empirical literature, and provide additional support for the notion that exchange rates can impact import and domestic prices, thereby helping avoid deflation. Although not common to all disaggregate export prices, overall export price sensitivity therefore moves in the opposite direction from import and domestic producer price sensitivity, showing a

¹Some anomalies are identified in Shioji (2012) when, for instance, the ERPT to import price significantly exceeds the unit or the export reaction puzzle.

flatter upward trend throughout the sample period from almost zero percent in 1985Q1 to about 20% in the long-run. Japanese exporters increasingly pass on exchange rate movements to foreign importers. Finally, price sensitivity is consistent throughout our results, indicating that incorporating extra information leads to more robust estimates.

The paper is organized as follows. The next section presents the econometric methodology and the identification strategy and describes the data employed in the empirical investigation. Section 3 presents the empirical findings and section 4 concludes.

2 Econometric Methodology

2.1 The model

We adopt the TVP-FAVAR model proposed by [Koop and Korobilis \(2010\)](#) and [Mumtaz \(2010\)](#):

$$Y_t = \alpha_t + \sum_{p=1}^P \beta_{t,p} Y_{t-p} + v_t \quad (1)$$

where $Y_t = [F_t \ E_t]'$. Y_t contains unobserved factors F_t , and an observable factor E_t of dimension $(\nu \times 1)$, which represents the effective exchange rate. The $((K + \nu) \times 1)$ vector of error terms v_t is mean 0 with covariance matrix Ω_t of dimension $((K + \nu) \times (K + \nu))$. The unobservable factors, F_t are estimated in a first step using a singular value decomposition of data and then they are included in the equation in a second step. These factors are extracted from X_t is $(N \times 1)$, which in addition to import and domestic prices, contains financial data and variables related to economic activity and foreign variables. X_t are assumed to be related to the unobservable factors F_t and the observable factors E_t as follows:

$$X_t = \Lambda_t^f F_t + \Lambda_t^E E_t + e_t \quad (2)$$

where e_t are errors with mean zero and $\Psi = \text{diag}(\exp(\psi_{1,t}), \dots, \exp(\psi_{n,t}))$ as variance-covariance matrix. The error terms e_t are assumed to be either weakly correlated or uncorrelated; these can be interpreted as idiosyncratic components. Λ^f and Λ^E are the

$(N \times K)$, $(N \times \nu)$ matrices of factor loadings.²

The structural shocks are recovered from the reduced form innovations using the Cholesky decomposition of the variance-covariance matrix $\Omega_t = A_t^{-1} H_t (A_t^{-1})'$. The time-varying matrices H_t and A_t are defined as:

$$H_t \equiv \begin{bmatrix} h_{1,t} & 0 & \cdots & 0 \\ 0 & h_{2,t} & \cdots & 0 \\ \vdots & \cdots & \ddots & \vdots \\ 0 & 0 & \cdots & h_{(K+\nu),t} \end{bmatrix} \quad A_t \equiv \begin{bmatrix} 1 & 0 & \cdots & 0 \\ a_{21,t} & 1 & \ddots & \vdots \\ \vdots & \cdots & \ddots & 0 \\ a_{(K+\nu)1,t} & \ddots & a_{(K+\nu)k,t} & 1 \end{bmatrix} \quad (3)$$

Following [Primiceri \(2005\)](#) and [Koop and Korobilis \(2010\)](#), all the parameters in equation 1 are assumed to evolve as random walks as follows:³

$$\begin{cases} B_t & = & \phi_{t-1} + \eta^{B_t} \\ a_t & = & a_{t-1} + \eta^{a_t} \\ \ln h_{i,t} & = & \ln h_{i,t-1} + \eta^{h_t} \end{cases} \quad (4)$$

where $B_t = [\alpha_t \ \beta_{t,p}]$ and $h_{i,t}$ evolve as geometric random walks. We assume that the innovation vectors $\eta_t^\phi \sim N(0, Q_\phi)$ and are independent from each other and from e_t and v_t and, where $\phi = [B_t \ a_t \ \ln h_{i,t}]$.

2.2 Estimation

Following [Koop and Korobilis \(2010\)](#), we summarize the estimation method in this section.⁴ Before summarizing the basic algorithm, we need to explain our choice of estimation method.⁵ As in [Koop and Korobilis \(2010\)](#) and [Korobilis \(2013\)](#), we estimate the model using a two-step procedure. In the first step, the unobservable factors are estimated by

²Factors are estimated using the principal component method. For the model to be identified, we use the standard normalization implicit in the principal components, as in [Bernanke et al. \(2005\)](#), and take $C'C/T = I$, where $C(\cdot)$ represents the common space spanned by the factors.

³Following [Korobilis \(2013\)](#) and [Mumtaz \(2010\)](#), in this model factor loadings and variance-covariance matrix in equation 2 are not allowed to vary over time.

⁴Detailed description of the prior distribution and the sampling method is given in [Koop and Korobilis \(2010\)](#).

⁵Our estimations are based on the latest version of TVP-FAVAR code provided by [Koop and Korobilis \(2010\)](#), where the corrigendum of [Negro and Primiceri \(2013\)](#) is taken into account

the principal component technique described in [Bernanke et al. \(2005\)](#). In the second step, we estimate the VAR system consisting of the estimated factors and the Japanese effective exchange rate.⁶

Four extracted factors are selected for inclusion in the VAR system. This number of factors is chosen not only by checking for sensitivity of the results to the inclusion of different numbers of factors, but also by the test of [Alessi et al. \(2010\)](#).⁷

Bayesian estimation of the model is performed through the Gibbs sampler, the most appropriate method of dealing with the problem of estimating a highly dimensional parameter model. Here we summarize the basic algorithm, which involves the following steps:

1. Initialize factors, model parameters and hyperparameters.⁸
2. Simulate the TVP-FAVAR coefficients B_t and the off-diagonal elements of the matrix A_t using [Carter and Kohn \(1994\)](#)'s algorithm,⁹ except for h which is simulated using [Kim et al. \(1998\)\(1998\)](#)'s algorithm.
3. Draw the hyperparameters $Q_{\eta^{ht}}^T$ using the inverse gamma distribution and draw the remaining Q_{η}^T from an inverse Wishart distribution.
4. Simulate the factors conditional on all other parameters using the method described by [Kim and Nelson \(2001\)](#).
5. Go to step 2.

We take 100,000 Gibbs sampling replications and discard the first 80,000.

⁶An alternative estimation method is a one-step method through a state-space representation. However, and as well explained in [Bernanke et al. \(2005\)](#), [Korobilis \(2013\)](#) and [Mumtaz and Surico \(2009\)](#), this method is very computationally demanding, especially to estimate high-dimensional models like the TVP-FAVAR model.

⁷As in [Bernanke et al. \(2005\)](#), we took the ad-hoc approach to choose the optimal number of factors. This is justified by the fact that, while the statistical identification determines the number of factors present in the dataset, it does not determine the number of factors to include in the VAR system. However, when we apply the refinement version of [Bai and Ng \(2002\)](#)'s test of [Alessi et al. \(2010\)](#), this also select the same number of factors. Results are provided in Figure 4 in Appendix B.

⁸For this step, we follow [Primiceri \(2005\)](#) and [Koop and Korobilis \(2010\)](#) and use an informative prior based on a training sample, corresponding to the first forty observations.

⁹We do check for stationarity of the estimated VAR coefficient and reject draws if nonstationary.

2.3 Data treatment and identification

For our empirical application, we use 78 quarterly time series data spanning the period from 1975Q1 to 2015-Q2.¹⁰ The database contains Japanese disaggregate price indices: import, corporate goods, export and consumer price indices. As in [Shioji \(2012\)](#), and as [Sato \(2014\)](#) suggests, we use a disaggregate Domestic Corporate Good Price Index (DCGPI) for different stages, namely, raw, intermediate and final goods. We also introduce industry-breakdown CGPI series, as the transmission of exchange rate variations may also vary across industries. This may, for instance, indicate how the growing regional production network consisting of Japan and other Asian countries affects the degree of ERPT in different sectors. Moreover, as shown by [Shioji \(2014\)](#), ERPT could also differ with degree of frequency of purchase of consumer goods. We therefore include in the analysis CPI series for different goods representing frequently-purchased and less frequently-purchased goods.

In addition, in order to obtain accurate ERPT estimation, we include macroeconomic variables that affect ERPT as identified by ([Taylor, 2000](#); [Devereux et al., 2004](#)). We add the output gap as a proxy for domestic demand shocks. The monetary aggregate and short term interest rates are included in order to take into account monetary policy effects on prices and on the ERPT ([Devereux et al. \(2004\)](#)). The crude oil price is used to control its effect on prices which could interfere with price reaction to exchange rate movements. Foreign variables related to activity for the United States and United Kingdom are proxies for world-wide demand. In addition, in order to take into account foreign price-setting behavior, we use two different proxies for the marginal cost of foreign exporters¹¹. In the empirical literature, this is constructed either as the weighted average of trade partners' CPI, as in [Marazzi et al. \(2005\)](#) and [Sekine \(2006\)](#), or as unit labor cost, as in [Campa and Goldberg \(2005\)](#). In this paper, we take both measures. These variables are also necessary controls in estimating the exchange rate pass-through as argued by [Campa and Goldberg \(2005\)](#). We include the penetration ratio, calculated on the basis of the ratio between

¹⁰Since we use the first forty observations as a training sample in the estimation procedure, our sample actually started from 1985Q1 to 2015Q2.

¹¹Foreign prices (unit labor cost) = (nominal effective exchange rate/ CPI(ULC)-based real effective exchange rate)* consumer price index excluding food and energy.

imports and real GDP, as an indicator of competitive pressures and changes in import composition. Finally, inflation and exchange rate volatilities¹² are included to control for the effects of the inflation environment (Taylor, 2000) and exchange volatility (Devereux et al., 2004) on ERPT.

Data treatment and transformation are performed as follows. First, the CPI series are adjusted for the consumption tax: each is regressed on dummies corresponding to the introduction of the tax in April 1989, and the subsequent two hikes in April 1997 and April 2014. Then, the series are de-seasonalized when necessary, using the X-12 procedure. Finally, data are transformed in order to make them stationary¹³, when necessary, using the first log-difference, and then demeaned and standardized.

Finally, as the FAVAR model is intended to measure monetary policy and its macroeconomic effects, some adaptations in terms of specification are required for our study on ERPT. Therefore, like Bernanke et al. (2005), we choose to distinguish between slow-moving variables, which are not affected by other shocks in the current period, and fast-moving variables which react contemporaneously to the exchange rate shocks. In this paper, we adopt the identification scheme of Bernanke et al. (2005) for monetary policy analysis, adapting it to ERPT analysis. Therefore, the first block, consisting of the slow-moving group, comprises variables such as oil price, output gap and foreign variables; the second block contains the nominal effective exchange rate; and the third block is the fast-moving group comprising domestic price variables.

3 Empirical results

This section reports and discusses the estimated time varying ERPT into different prices. The convention in the literature (see for instance Shioji (2012)), is to measure the elasticity

¹²Volatility is measured by the annual unconditional standard deviation.

¹³It would have been better to take account of Uhlig (2009) critique, and use *heavy* transformations especially on CPI series. However, as we need to calculate cumulative impulse response functions for price variables, cumulating twice the series transformed using second log-difference gives uninterpretable impulse response functions. We therefore adopt in this paper *light* transformations differencing price series only once. Note, however, that we did estimate the model with *heavy* transformations; when we compared cumulative impulse response functions of CPI series with those from *soft* transformations, the results were fairly similar.

or ERPT rate based on (i) the cumulative impulse response functions of different prices to a positive exchange rate shock (indicating depreciation), and (ii) the impulse responses of the exchange rate to an exchange rate shock.¹⁴

3.1 The first stage of ERPT: from exchange rate to import prices

Figure 1 displays the rate of ERPT into the aggregate overall import price (IPI) as well as disaggregate import prices by stages of demand namely the import price index of raw materials (IP-RAW), intermediate (IP-INTERM) and final goods (IP-FG). The first common feature for all import prices stages is a decrease in both short- and long-run ERPT over time, from 1985 until the late 2000s. In the short run, ie. until the fourth quarter, ERPT into overall import price is incomplete (between 40% and 60%), suggesting that the exchange rate movements are only partially passed instantaneously to import prices. This reflects mixed LCP and PCP behavior by firms, and is consistent with [Betts and Devereux \(2000\)](#)' theoretical findings: under imperfect competition and nominal rigidity, only a fraction of firms set their prices in the home currency (PCP), the remainder setting prices in the foreign currency (LCP or PTM). As predicted by the theory, when the full effect of exchange rate movements is manifested, long-run ERPT increases gradually over horizons at each point of time. However, only at the start of the sample, in 1985, is the long-run ERPT into import prices virtually complete, except for raw materials (about 70%). And once again, the long-run ERPT is decreases quite sharply over the 1990s and up to the late 2000s. Moreover, this result is common to all product categories, as shown in Figure 5. This well documented downward trend in ERPT is common to developed countries,¹⁵ and can be explained by a changing world trade structure, inflation environment, or monetary policy framework.

However, another likely explanation for the evolving ERPT into imported, especially intermediate, good prices is the growing network of Japanese producer. As pointed out

¹⁴More exactly, pass-through rate is calculated according to: $ERPT_{s,t} = \frac{P_{s,t}}{E_{s,t}}$, where $P_{s,t}$ is the cumulative impulse response of a price variable at horizon s in period t and $E_{s,t}$ is the cumulative change of the nominal effective exchange rate at horizon s in period t .

¹⁵For example, [Campa and Goldberg \(2005\)](#) for the United states, [Otani et al. \(2003\)](#) and [Otani et al. \(2006\)](#) for Japan, [Mumtaz et al. \(2006\)](#) for UK and [Stulz \(2007\)](#) for Switzerland.

in [Ito et al. \(2015\)](#), intra-firm trade between subsidiaries and the head office in Japan is often invoiced in yen. Specifically, subsidiaries selling intermediate inputs to Japan tend to choose yen invoicing when there are low fluctuations in the exchange rate. This can explain both declining ERPT to intermediate goods prices and thus to total import prices over time, and its resurgence during the Asian crisis. When faced with large fluctuations in their local currency, subsidiaries reduce yen invoicing transactions, thereby increasing ERPT to Japanese import prices. As we can see from different ERPT estimations, the Asian crisis significantly altered ERPT to all import prices. Exporters from neighboring economies, subsidiaries or not, transmitted about 60% to 75% of exchange rate changes into different import prices.

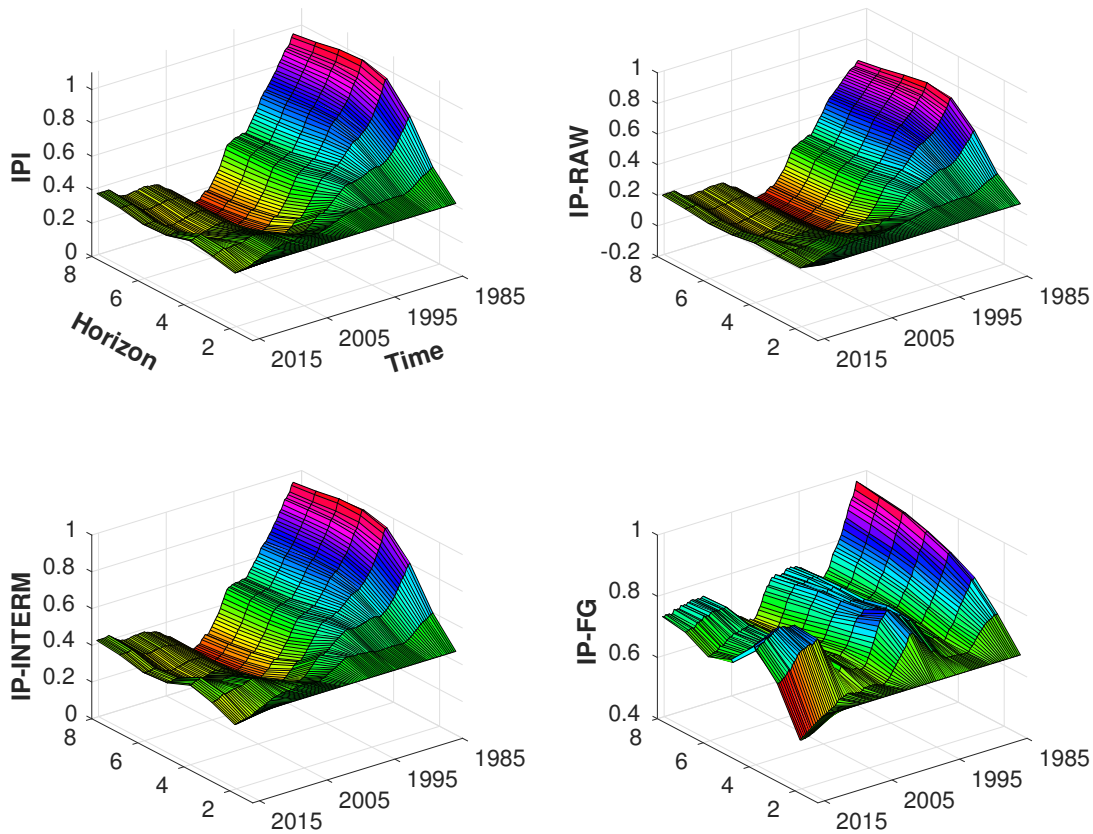
An additional interesting finding by [Ito et al. \(2015\)](#) is that the subsidiaries' choice of invoice currency depends on the type of export goods. Unlike intermediate imported goods, subsidiaries tend to reduce yen invoicing for finished goods exported to Japan. This may therefore explain the higher ERPT into imported final goods prices compared with raw and intermediate goods prices.¹⁶ Our result also confirms that of [Fujii \(2004\)](#), indicating that for imported final goods, the PCP theory is likely to apply. This suggests that, despite international competition firms exporting final goods to Japanese importers and consumers maintain a certain pricing power leading to nearly perfect long-run ERPT. A high proportion of imported final goods belong to sectors subject to high product differentiation and/or limited Japanese domestic competition, which may confer or strengthen the market and pricing power of foreign exporters, allowing them to pass on more of the exchange rate movements into Japanese imported final goods prices. [Figure 10](#) shows that foreign prices¹⁷ become more sensitive to exchange rate movements starting from the late 2000s. This confirms the view that foreign exporters have become more likely to change their prices frequently in recent years, passing **more** exchange rate changes on to prices.

The second common feature is the reversing upward trend in import prices at all stages since the late 2000s. While the long-run ERPT to IPI of 40%, observed at the end of the

¹⁶Particularly non-durable consumer and capital final goods. As [Figure 5](#) shows, imported final goods' reaction to exchange rate variations is mostly governed by these type of goods.

¹⁷[Figure 10](#) shows elasticities of the two measures of the theoretical notion of foreign prices used in [Marazzi et al. \(2005\)](#) and [Campa and Goldberg \(2005\)](#). Results are almost identical.

Figure 1: Responses of import prices



Note: three dimensional plots display the evolution of rates of exchange rate pass-through (ERPT) into import prices, over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. These price elasticities are calculated based on impulse response functions to a one-standard-deviation positive exchange rate shock (equivalent to a depreciation). More precisely, $ERPT_{s,t} = \frac{P_{s,t}}{E_{s,t}}$, where $P_{s,t}$ is the cumulative impulse response of a price variable at horizon s in period t and $E_{s,t}$ is the cumulative change of the nominal effective exchange rate at horizon s in period t . In the first row, we report ERPT into overall imported price (IPI) and import raw materials price (IP-RAW), and in the second row, into import intermediate materials price (IP-INTERM) and imported final goods price (IP-FG).

sample, may not seem particularly large, what is important is the trend shift starting in the late 2000s and which may be persistent in the future. This is consistent with the ERPT comeback view of [Shioji \(2012\)](#) and [Hara et al. \(2015\)](#) on Japanese import prices. This suggests that the yen's depreciation in recent years could regain its inflationary effect. The increased ERPT to prices for imported intermediate goods and raw materials is expected subsequently to affect domestic intermediate and final producer prices, in turn likely to impact retail prices. Moreover, very recent studies show that Japan's experience of ERPT resurgence after the financial crisis is not unique. [Kristin Forbes and Nenova \(2015\)](#) show that ERPT in UK increased between 2010 and 2013.

Still, our finding that ERPT to import prices is greatest for final goods (about 70%

at the end of the sample) suggests that it is final goods that can most strongly impact domestic consumption prices by passing on exchange rate variations. Yen depreciation is likely to have more impact on Japanese consumer price inflation directly via imported final goods than indirectly via the production chain using imported raw materials and intermediate goods.

The reaction of real import elasticity (figure 10) to yen depreciation is in line with results from import prices. The ERPT into real imports is negative for all the period, namely between 1985Q1 and 2015Q2, and evolves in virtually the same way as the ERPT into import prices. As with import prices, the elasticity of real imports to exchange rate movements followed a downward trend until the late 2000s, subsequently reversing due to rising ERPT into import prices for the same period.

3.2 The domestic side

In this section, we investigate the effect that exchange rate movements have on both by-stage-breakdown and industry-breakdown CGPI. This helps us identify the transmission mechanism which may differ both across stages of production and across industries. Retail prices, or consumer price index (CPI) data, are also included in the analysis, completing the domestic price chain. As shown in [Shioji \(2015\)](#), ERPT may differ between frequently and less frequently purchased items. This aspect is therefore taken into account in our analysis.

Figure 2 displays the degree of ERPT along the pricing chain for domestically produced goods, namely the ERPT to the CGPI for raw materials (raw), intermediate goods (INTERM) and final goods (FINALG) that are domestically produced. As with import prices, ERPT for domestic prices declined and followed a downward trend from the late 1980s until the late 2000s, with the exception of the 1997 Asian crisis period. ERPT rates for different types of goods still vary in terms of starting values. The long-run ERPT value was around 25% for raw materials at the beginning of the sample in 1985, about 30% for final goods and almost 70% for intermediate goods. From the late 2000s, however, ERPT rates began a new upward trend, although they remained low; 10% for raw and

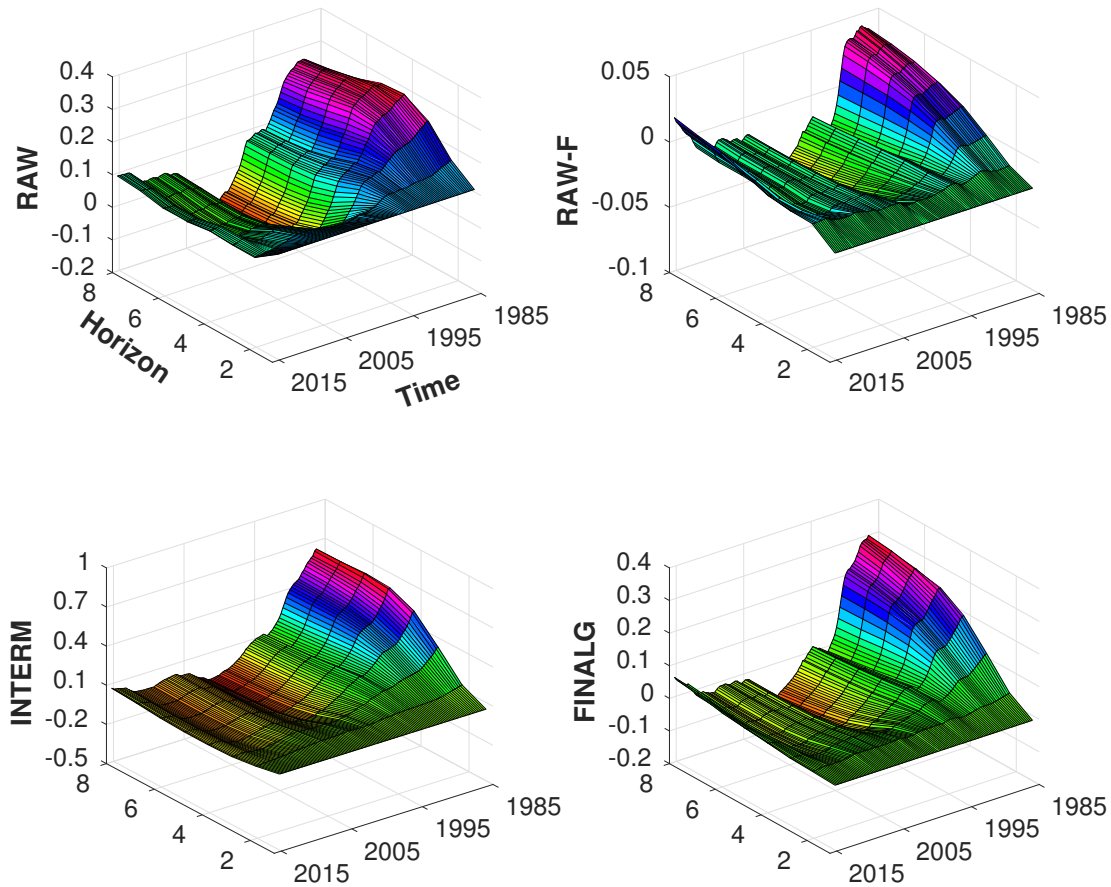
intermediate materials and even less than 10% for final goods in the long-run.

A close look at Figure 6 shows that ERPT rates also vary depending on the category of final goods, both in terms of level and of evolution over time. In the medium and long-run, while the ERPT rates for both consumer and consumer non-durable (CON-ND) goods have almost the same starting (30%) and final values (5%), the starting ERPT rate for capital goods (Capital) is only around 20%. However, capital goods have recently shown almost the same long-run ERPT rates (between 5% and 10%). The ERPT for consumer non-durable goods differs completely: except at the very beginning of the sample, it never exceeds 10%, remaining almost unchanged over time. The relatively high ERPT rate for consumer non-durables, which seems to govern the total final goods price, patterns, was predictable, since this category includes gasoline. Nevertheless, the ERPT rate is only about half that of ERPT to CGPI for intermediate goods. This is exactly in line with the results of Shioji (2012), indicating that domestic companies absorb in their profit margin (in their mark-ups) a large part of exchange rate through-put to intermediate goods domestically produced.

Overall, the analysis of the ERPT to different CGPI indicates that the pass-through of the yen depreciation quickly vanishes through the distribution chain. First, prices of domestically produced goods respond less to yen depreciation than imported goods, as the share of imported goods decreases along the distribution chain and also because domestic firms may choose not to fully transmit price changes of imported goods. Second, the relatively low ERPT to domestic final goods compared with domestic intermediate goods indicates that Japanese firms absorb the difference in their own mark-ups.

It is also worth noting that the industry-breakdown CGPI data analysis (Figure 7) identifies significant heterogeneity in ERPT rates across industries. These results support our decision to include industry-breakdown CGPI data, as suggested by Sato (2014). In so doing, we are able to take account of the growing Japanese production network that may impact price response to exchange rate movements of domestically produced goods, depending on the proportion of intra-firm imported goods in their production processes. For instance, at the beginning of the sample while the long-run ERPT rates (horizon 8)

Figure 2: Responses of disaggregate CGPI prices



Note: three dimensional plots display the evolution of exchange rate pass-through (ERPT) rates into corporate goods price indices (CGPI), over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the first row, we report ERPT rates into CGPI for raw materials (RAW) and for raw material of fuel (RAW-F), and in the second row, into CGPI for intermediate goods (INTERM) and for final goods (FINALG).

for plastic and chemical items are about 60%, rates for other categories range from about 28% (FBTF and P-Machinery) to 15% (ELECTRONIC). Clearly, industries which are more likely to use intra-firm imported intermediate goods, like machinery, and particularly the electronic and information and communication equipment sectors, experience weak ERPT to their prices. However, there is no difference in patterns of the ERPT evolution: the ERPT rate declined over time until the late 2000s and since then has followed a new upward trend, as previously found. The one exception is Information and Communications Equipment (IC-EQUIP), where ERPT rate behaves quite differently.

Figures 8 and 9 in the Appendix D show how different CPIs are influenced by exchange rate movements. As with import and CGPI prices, ERPT rates vary considerably across CPI series. A common feature of all ERPT rates for CPIs is their resurgence since the late

2000s, showing that prices of frequently-purchased goods do not systemically appear to be more sensitive to yen changes. Also, Figures 8 and 9 show a variety of shapes for ERPT rates, which remain quite low for all prices, even negative for some, in both the frequently- and the less frequently-purchased item groups. Among frequently-purchased items (Figure 9), the CPI for gasoline stands out as the highest long-run ERPT.

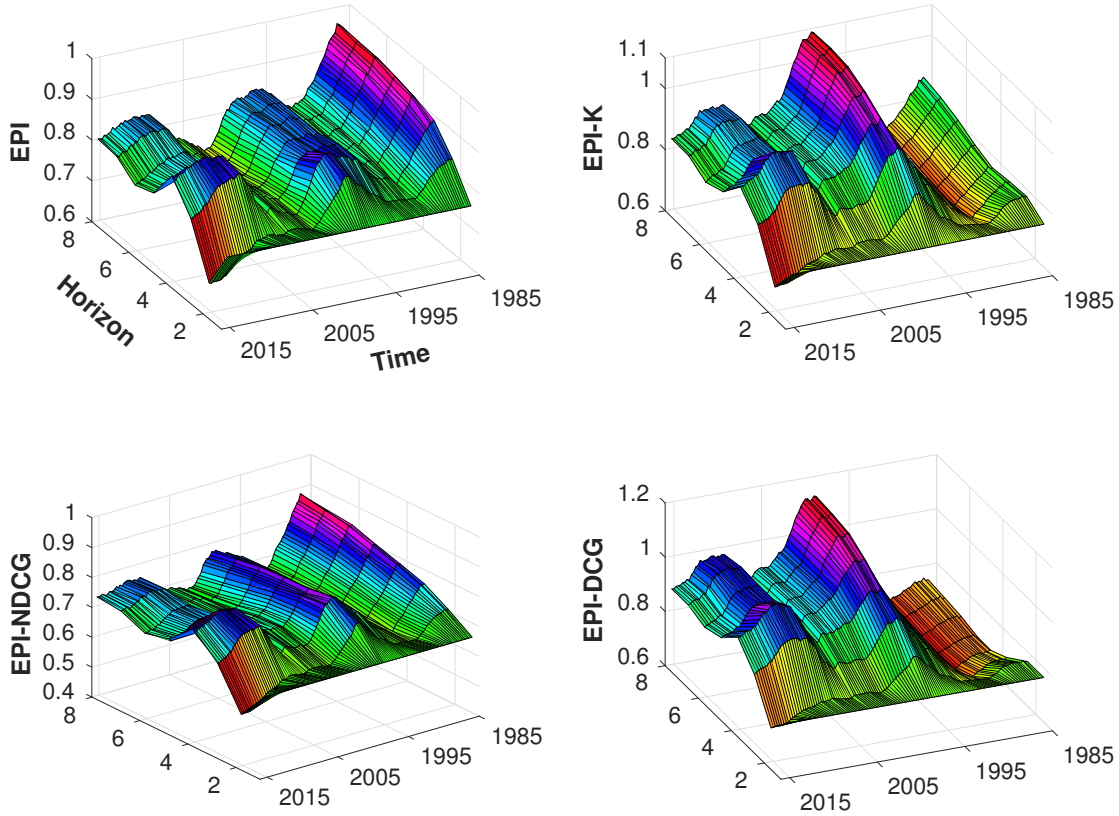
3.3 The export side

As export prices are expressed in JPY units, the rate of pass-through into export prices should be read as one minus given the elasticity in Figure 3. The second expected benefit from the yen depreciation is that of generating an export competitiveness boosting Japanese GDP. Assuming, of course, that the yen depreciation is strongly transmitted to Japanese export prices. Results shown in Figure 3 indicate a gradual rise in ERPT into overall export (EPI) and non-durable consumer goods prices over time, whereas the opposite picture emerges for capital and durable consumer good prices. Throughout the horizon, however, a common downward trend is found for ERPT into different export prices: fairly high short-run rates of ERPT of roughly 30% to 40% (slightly more for non-durable goods prices, between 40% and 50%), which then rapidly diminish to reach a long-run level of some 20% within about four quarters.

Japanese intra-firm export has grown rapidly in recent years. [Bernard et al. \(2010\)](#), using U.S. trade transaction data, observe that in 2000, 74% of U.S. imports from Japan are intra-firm and traded between Japanese parents and U.S. subsidiaries. Moreover, the authors show that products traded predominantly via intra-firm transactions are either capital intensive or durable goods.¹⁸ On the other hand, [Ito et al. \(2013\)](#) focus on the invoicing choice of Japanese exporters and find that they tend to choose yen invoicing for intra-firm trade. This should therefore lower the ERPT into export prices, especially for capital and durable good prices. This intuition is supported by our results: capital and durable good price elasticities to yen depreciation decrease over time. However, the overall export price is increasing and seems to be governed by non-durable consumer good prices.

¹⁸page 5, by industry

Figure 3: Responses of disaggregate export prices



Note: three dimensional plots display the evolution of rates of exchange rate pass-through (ERPT) into export prices, over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the first row, we report ERPT into overall export price (EPI) and export price of capital goods (EPI-K), and in the second row, into export price of non-durable (EPI-NDCG) and durable (EPI-DCG) consumer goods.

Although the ERPT into overall export prices may partly hide divergence in elasticity among export price categories, the very short-run ERPT remains on average relatively strong, indicating that export prices respond quite strongly and immediately to yen depreciations. However, contrary to expectations, Figure 10 shows that real export elasticity to yen depreciation moves over time in the opposite direction from that of the export price index. That is to say, when ERPT into export price is weak at the beginning of the sample, real exports respond relatively strongly, and conversely at the end of the sample. This means that exports were far more responsive to export prices than at present. Still, at the end of the sample, the yen depreciation resulted in a very short-lived increase in real exports (Figure 10), which rapidly returned to the initial zero level.

4 Conclusion

Contributing to the current debate about the ability of exchange rate movements to impact domestic inflation and help escape deflation, this paper adopts the TVP-FAVAR model to re-examine the magnitude of the exchange rate pass-through (ERPT) into Japanese import, domestic and export prices. Compared with VAR and TVP-VAR models, widely used in previous studies on Japan, the TVP-FAVAR model offers two advantages. First, it allows us to incorporate a large amount of information, thus considering the macroeconomic variables affecting ERPT to prices such as output gap, oil price, openness of country, monetary policy instrument. At the same time, we can take into account different indirect transmissions of exchange rates to prices through other variables like mark-ups and marginal costs of trading partners. Second, the model allows us to incorporate many disaggregate prices series corresponding to different stages of demand and different categories and industries. It thereby becomes possible to evaluate the ERPT to different prices along the pricing chain and for different categories and industries in a unifying framework. Both advantages lead to more accurate estimation of the ERPT.

Our empirical results show that the evolution of ERPT assumes different patterns and magnitudes across disaggregate prices, justifying the incorporation of these variables in the analysis. The ERPT follows a downward trend starting from 1985Q1 to the late 2000s; subsequently, prices become more sensitive to exchange rate changes, confirming the comeback of the ERPT detected in recent empirical studies. These findings provide additional support to the notion that exchange rates can impact import and domestic prices, thereby possibly helping avoid deflation. Our results display only a little or no evidence of inconsistent price sensitivity, indicating that incorporating extra information leads to more accurate estimates.

Two interesting extensions to our analysis would be, first, to examine the effect of the yen appreciation in order to check whether price sensitivity is asymmetric ([Delatte and López-Villavicencio, 2012](#)). A second avenue worth exploring would be the notion that the underlying cause of exchange rate changes can impact the exchange rate through-put to

prices, as put forward in [Kristin Forbes and Nenova \(2015\)](#) and [Shambaugh \(2008\)](#).

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A Dataset

Table 1: **Variable list**

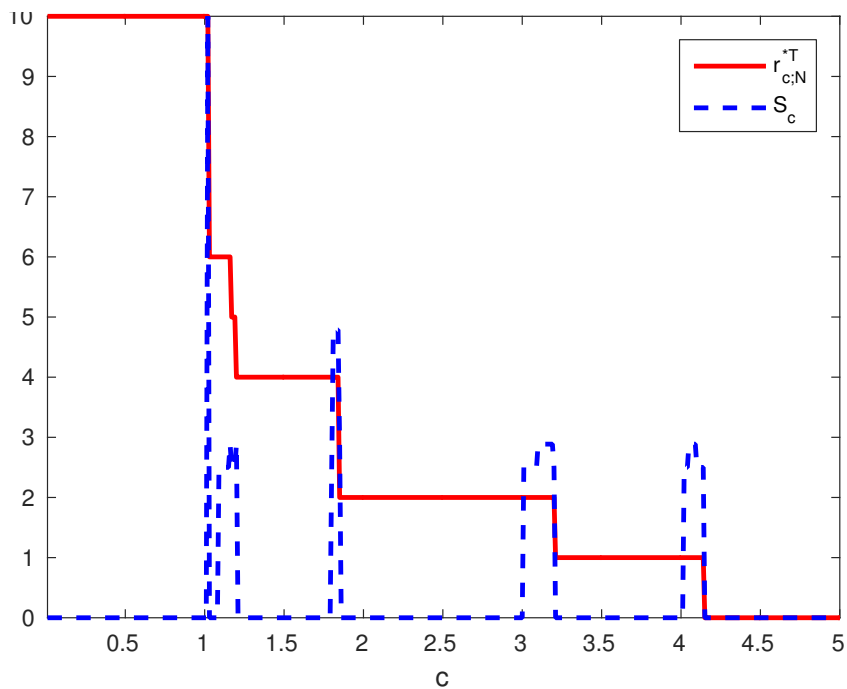
Data are extracted from Reuters EcoWin database. The transformation codes (T) are: 1: no transformation; 5: first difference of logarithm.

#	T	Mnemonic	Description
Slow moving			
1	5	WORLDOIL	World Commodity Prices: Crude Oil NADJ
2	5	USGDP	US GDP (AR) CONA
3	5	UKGDP	UK GDP At Market Prices (CVM) CONA
4	5	FFR	US Federal Funds Rate NADJ
5	5	UKrate	UK short term rate
7	5	M2CD	JP Money Supply: M2 (METHO-BREAK, APR. 2003) CURA
8	1	TOPIX	JP Share Prices: TSE TOPIX ALL SHARES INDEX
9	1	OUTPUTGAP	Output gap provided by the BoJ
10	5	RGFCFPNR	JP Private Non-Residential GFCF (Real) (AR) CONA
11	5	RFDEXP	JP Final Domestic Expenditure (Real) (AR) CONA
12	5	RTDEXP	JP Total Domestic Expenditure (Real) (AR) CONA
13	5	NEER	Nominal/Effective Exchange Rates (2010=100) NEER
Fast moving			
14	5	FOREIGNP	Foreign prices
15	5	FOREIGNULC	Foreign Unit Labor Cost
16	5	IMPORTALL	Import Price Index/yen basis. All commodities
17	5	IMPORTPCNG	Import Price Index/Group/Petroleum, coal & natural gas
18	5	IMPORTFF	Import Price Index/Group/Foodstuffs & feedstuffs
19	5	IMPORTRAW	Import Price Index/Raw materials
20	5	IMPORTRAWP	Import Price Index/Raw materials/Raw materials for processing
21	5	IMPORTINTERM	Import Price Index/Intermediate materials
22	5	IMPORTINTERSFG	Import Price Index/Intermediate materials/Semi-finished goods
23	5	IMPORTINTERFE	Import Price Index/Intermediate materials/Fuel & energy
24	5	IMPORTFG	Import Price Index/Final goods
25	5	IMPORTFGCG	Import Price Index/Final goods/Capital goods
26	5	IMPORTFGCONSG	Import Price Index/Final goods/Consumer goods
27	5	IMPORTFGCGD	Import Price Index/Final goods/Durable consumer goods
28	5	IMPORTFGCGND	Import Price Index/Final goods/Nondurable consumer goods
29	5	CGPIRM	CGPI (2010 Base)/Raw materials (Domestic goods)
30	5	CGPIRMP	CGPI (2010 Base)/Raw materials/Raw materials for processing
31	5	CGPIRMF	CGPI (2010 Base)/Raw materials/Fuel
32	5	CGPIRMOTH	CGPI (2010 Base)/Raw materials/Other raw materials
33	5	CGPIINTM	CGPI (2010 Base)/Intermediate materials (Domestic goods)
34	5	CGPIINTMSFG	CGPI (2010 Base)/Intermediate materials/Semi-finished goods
35	5	CGPIINTMFE	CGPI (2010 Base)/Intermediate materials/Fuel & energy
36	5	CGPIINTOTH	CGPI (2010 Base)/Other intermediate materials

37	5	CGPIFG	CGPI (2010 Base)/Final goods
38	5	CGPIFGCAPG	CGPI (2010 Base)/Final goods/Capital goods
39	5	CGPIFGCONSG	CGPI (2010 Base)/Final goods/Consumer goods
40	5	CGPIFGDCONSG	CGPI (2010 Base)/Final goods/Durable consumer goods
41	5	CGPIFGNDCONSG	CGPI (2010 Base)/Final goods/Nondurable consumer goods
42	5	FBTF	CGPI: Food, beverages, tobacco & feedstuffs
43	5	TEXTILE	CGPI: Textile products
44	5	CHEMICAL	CGPI: Chemicals & related products
45	5	PLASTIC	CGPI: Plastic products
46	5	CERSTCL	CGPI: Ceramic, stone & clay products
47	5	BOMACH	CGPI: Business oriented machinery
48	5	INFCMEQUIP	CGPI: Information & communications equipment
49	5	TPTEQUIP	CGPI: General purpose machinery
50	5	PRDMACH	CGPI: Production machinery
51	5	ELECCOMP	CGPI: Electronic components & devices
52	5	ECMACHEQUIP	CGPI: Electrical machinery & equipment
53	5	JPCPIEFAP	CPI: All Items Less Fresh Food NADJ
54	5	JPCPIEIAF	CPI: All Items Less Imputed Rent NADJ
55	5	JPCPIGLAF	CPI: All items NADJ
56	5	JPCPIEFFF	CPI: Food Less Fresh Food NADJ
57	5	JPCPIFDAF	CPI: Food NADJ
58	5	JPCPIRXIF	CPI: Rent Less Imputed Rent NADJ
59	5	JPCPPVSWF	CPI: Processed Vegetables & Seaweeds NADJ
60	5	JPCPIUEAF	CPI: Electricity NADJ
61	5	JPCPIACAF	CPI: Clothes NADJ
62	5	JPCPIUGAF	CPI: Gas NADJ
63	5	JPCPIRTAF	CPI: Rent NADJ
64	5	JPCPGASAF	CPI: Gasoline NADJ
65	5	JPCPAUTAF	CPI: Automobiles NADJ
66	5	JPCPNWPAF	CPI: Newspapers NADJ
67	5	JPCPIHHAF	CPI: Domestic Services NADJ
68	5	JPCPIDRAF	CPI: Durable Goods NADJ
69	5	EXPPALL	CGPI (2010 Base)/ Export Price Index (yen basis)
70	5	EXPPCAPG	CGPI (2010 Base)/ Exports/Capital goods
71	5	EXPPCONSG	CGPI (2010 Base)/Exports/Consumer goods
72	5	EXPPDCG	CGPI (2010 Base)/ Exports/Durable consumer goods
73	5	EXPPNDCG	CGPI (2010 Base)/ Exports/Nondurable consumer goods
74	5	WAGE	JP Labour Markets: Wage Rates NADJ
75	5	REXPORT	Real Export, Unit CY2010=100
76	5	RIMPORT	Real Import, Unit CY2010=100
77	5	TOT	JP Terms of Trade Index NADJ
78	5	PRATIO	Import Penetration Ratio (the share of imports of goods and services in real GDP)

B Number of factor test

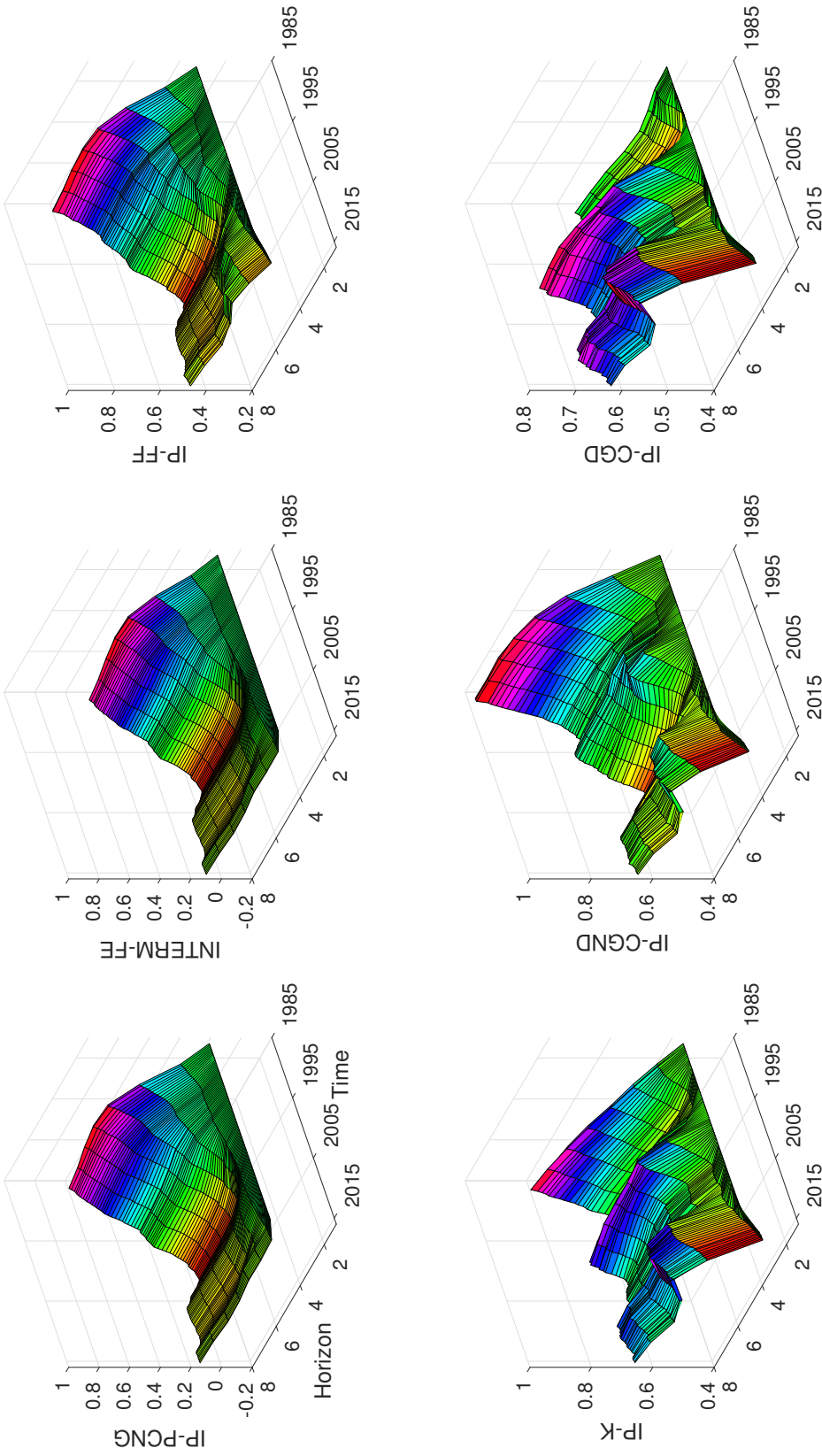
Figure 4: Estimated number of factors



Note: Dashed line represents the suggested number of factors r_c and the solid line is a measure of the instability of r_c . The number must correspond to the smallest value of c for which both a plateau of the dashed line and a zero of the solid line occur. In this sense, [Bai and Ng \(2002\)](#) criteria constitute only a special case when $c = 1$. This refined test suggests therefore 2 factors for our dataset. Fuller detail on this approach is given in [Alessi et al. \(2010\)](#).

C Disaggregate Import Prices

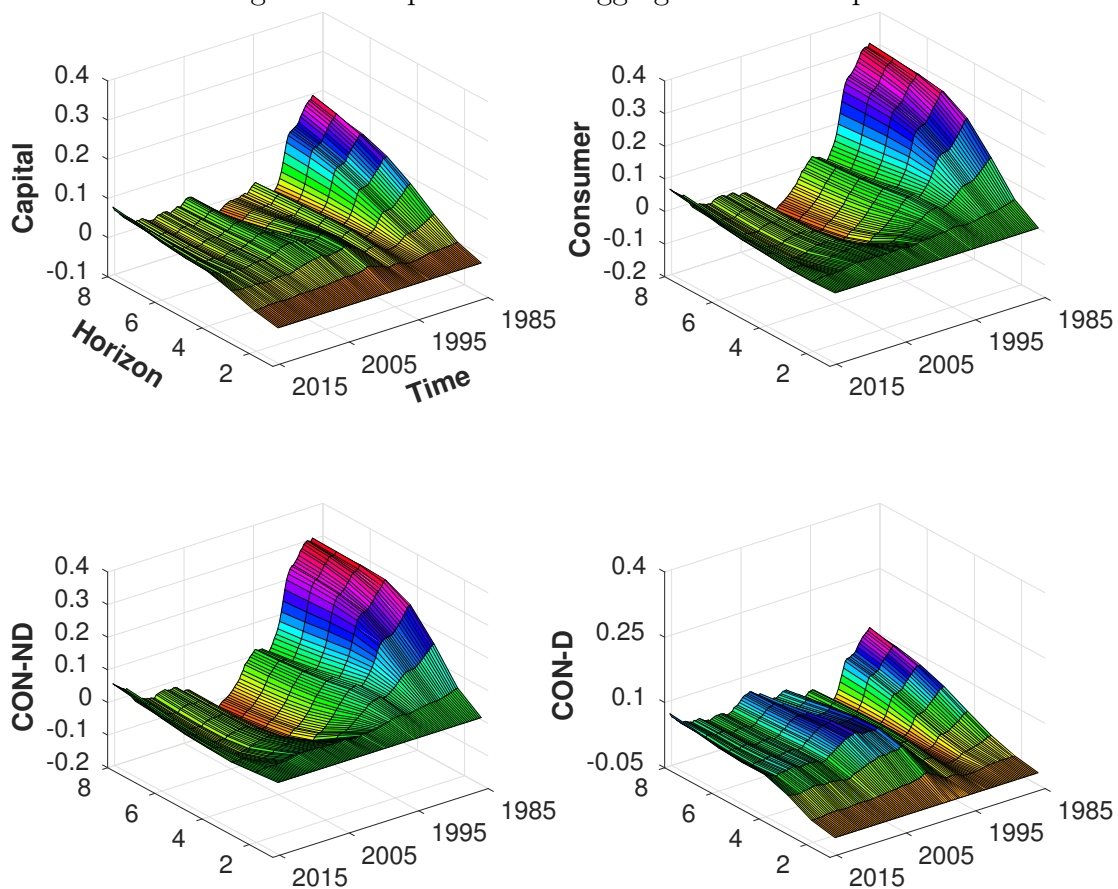
Figure 5: Responses of import prices



Note: three dimensional plots display the evolution of rates of exchange rate pass-through (ERPT) into import prices, over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the top panel, we report ERPT into import price of petroleum, coal and natural gas (IP-PCNG), intermediate fuel and energy (INTERM-FE) and foodstuffs and feedstuffs (IP-TF), and in the bottom panel into import price of capital goods (IP-K) and import price of non-durable (IP-NDCG) and durable (IP-DCG) consumer goods.

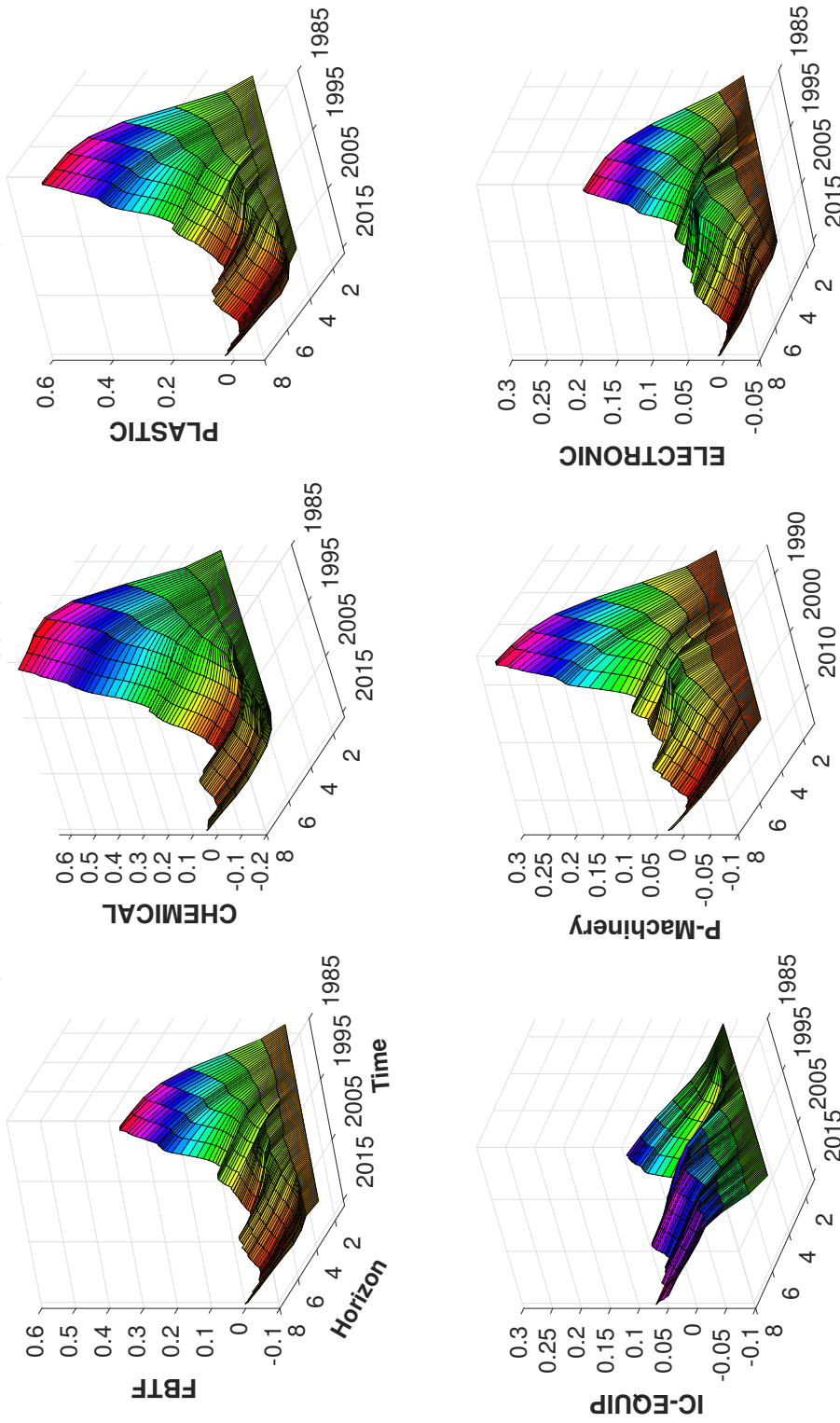
D Domestic side

Figure 6: Responses of disaggregate domestic prices



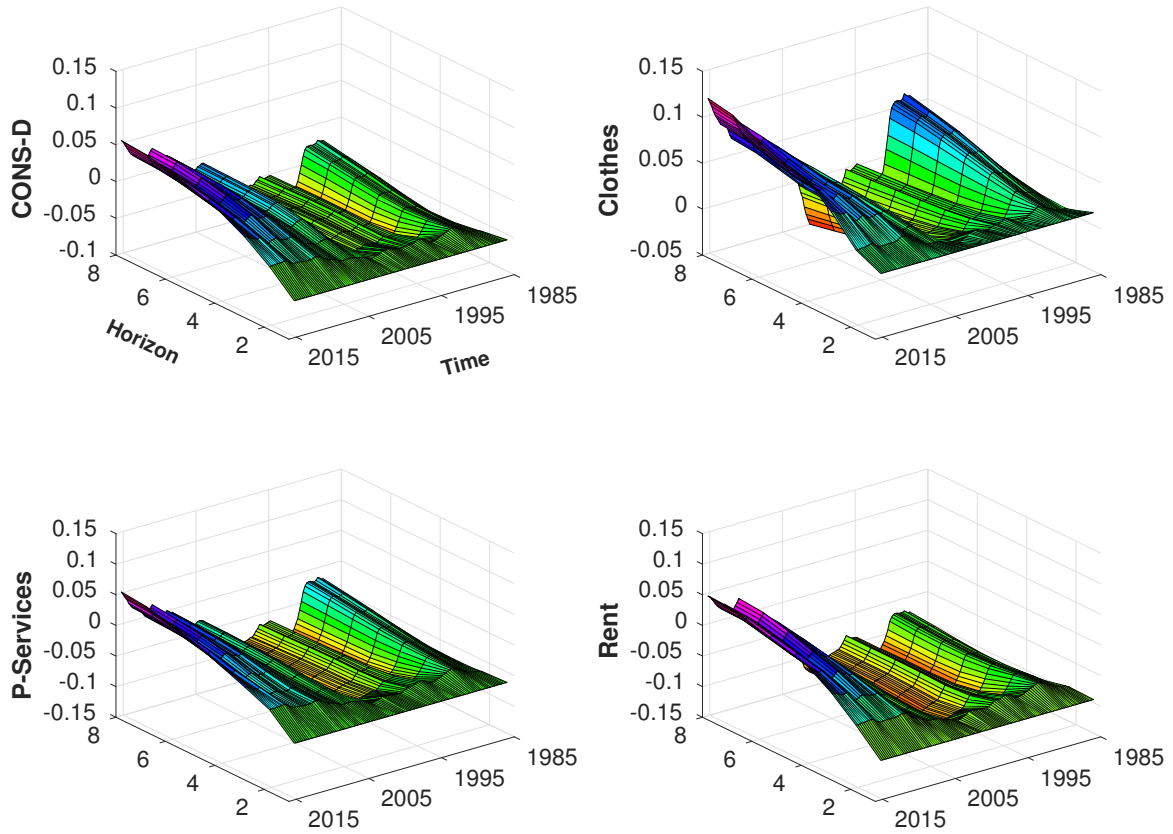
Note: three dimensional plots display the evolution rates of exchange rate pass-through (ERPT) into corporate goods price indices (CGPI), over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the first row, we report ERPT into CGPI for capital goods (Capital) and for consumer goods (Consumer), and in the second row, into CGPI for consumer non-durable (CON-ND) and durable goods (CON-D).

Figure 7: Responses of disaggregate CGPI prices by industry



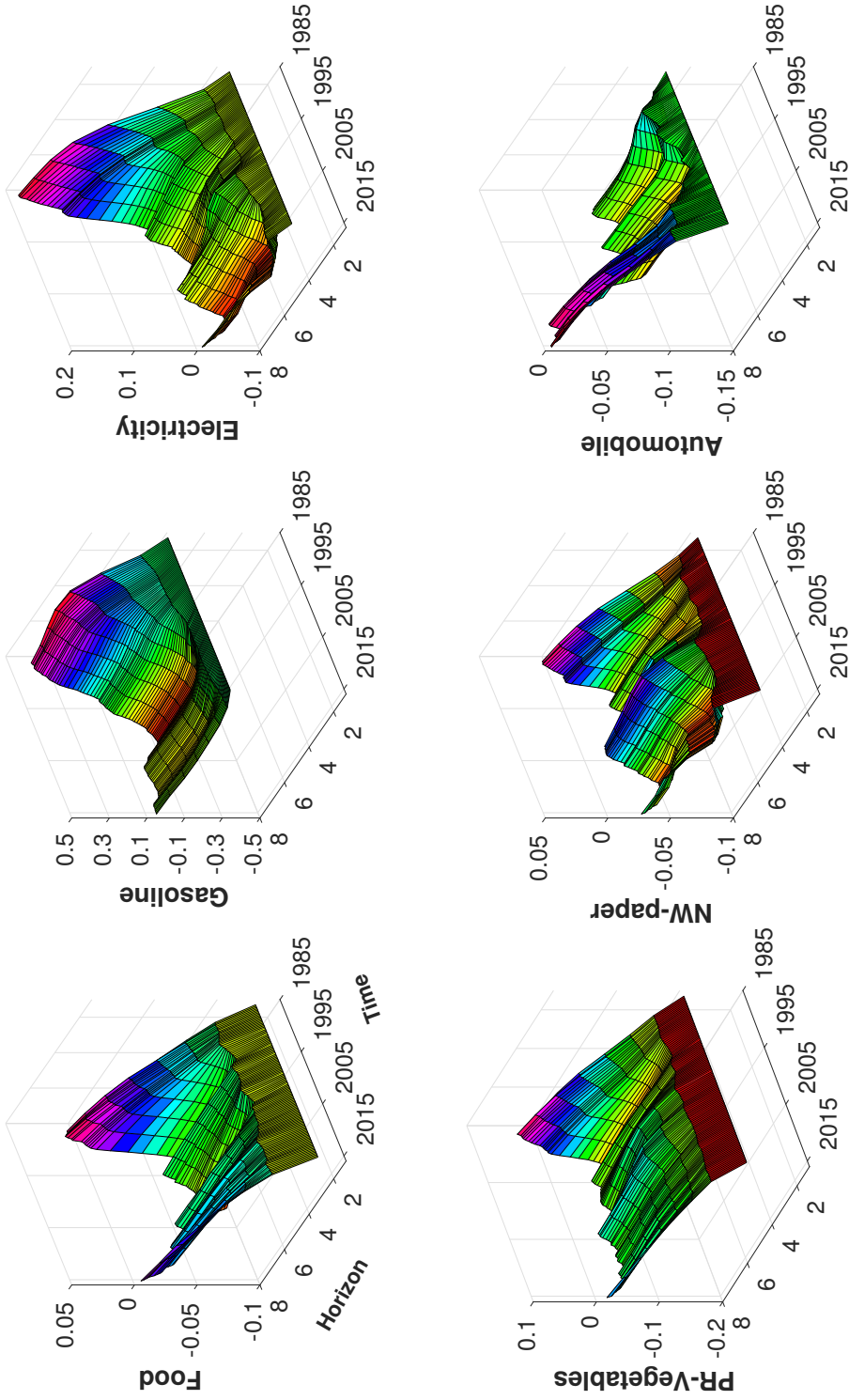
Note: three dimensional plots display the evolution of rates of the exchange rate pass-through (ERPT) into industry-breakdown corporate goods price indices (CGPI), over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the top panel, we report ERPT into CGPI for food, beverage, tobacco and feedstuffs (FBTF), for chemical products (CHEMICAL), and plastic products (PLASTIC), and in the bottom panel, into CGPI for information and communication equipment (IC-EQUIP), production machinery (P-Machinery) and electronic components and devices (ELECTRONIC).

Figure 8: Responses of CPI for less frequent purchased items



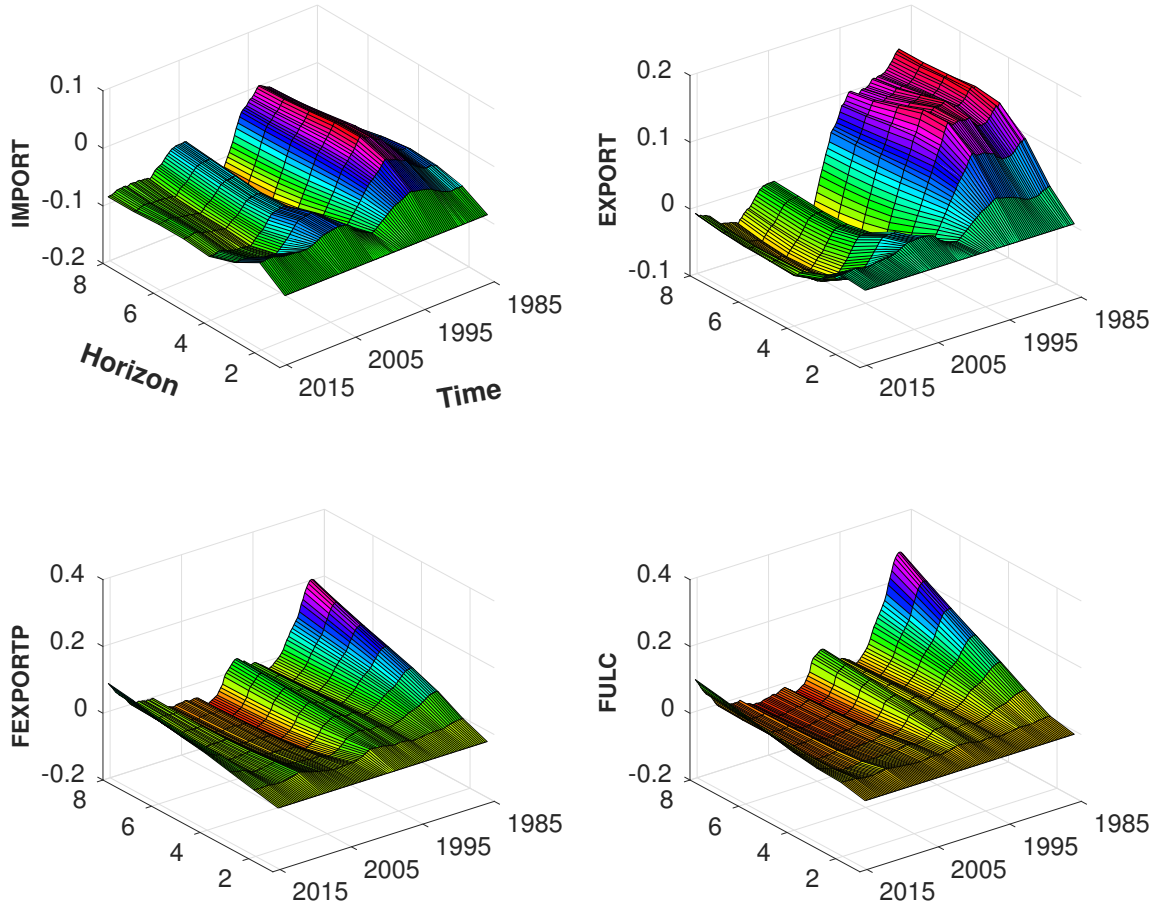
Note: three dimensional plots display the evolution of rates of the exchange rate pass-through (ERPT) into consumer price indices (CPI) for less frequent purchased items, over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the first row, we report ERPT into CPI for durable consumer goods (CONS-D) and for clothes, and in the second row, into CPI for domestic services (P-services) and for rent.

Figure 9: Responses of CPI for frequent purchased items



Note: three dimensional plots display the evolution of rates of the exchange rate pass-through (ERPT) into consumer price indices (CPI) for frequent purchased items, over horizon (8 quarters) and over time from 1985Q1 to 2015Q2. In the top panel, we report ERPT into CPI for food, gasoline and electricity, and in the bottom panel, into CPI for processed vegetables and seaweeds (PR-Vegetables), news-papers and automobiles.

Figure 10: Responses of activity variables



Note: three dimensional plots display the evolution of rates of the exchange rate pass-through (ERPT) into real import and real export and into the two measures of foreign prices, namely the CPI-based real effective exchange rate (FEXPORTP) and the ULC-based REER (FULC), over horizon (8 quarters) and over time from 1985Q1 to 2015Q2.