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► **To cite this version:**

Wenjie Li, Francesca Bassi, Davide Dardari, Michel Kieffer, Gianni Pasolini. Low-complexity distributed fault detection for wireless sensor networks. IEEE International Conference on Communication, Jun 2015, London, United Kingdom. pp.6712 - 6718, 10.1109/ICC.2015.7249395 . hal-01260552

HAL Id: hal-01260552

<https://hal.science/hal-01260552>

Submitted on 22 Jan 2016

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Low-Complexity Distributed Fault Detection for Wireless Sensor Networks

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Abstract—To guarantee its integrity, a wireless sensor network needs to efficiently detect faulty nodes producing erroneous measurements. This paper proposes a fully distributed fault detection algorithm. A node first collects the measurements of its neighborhood, processes them to decide whether they contain outliers, and broadcasts the result. Then, it decides autonomously about its functioning status. The detection algorithm is proposed in two variants, depending on the proportion of faulty nodes in the network. A theoretical analysis of the probability of error and of the convergence of the algorithm is provided. The trade-off between false alarm probability and detection probability is characterized using simulation.¹

I. INTRODUCTION

A wireless sensor network (WSN) consists of many randomly deployed, autonomous devices, capable of acquiring physical measurements, of processing them, and of communicating over wireless channels. The role of a WSN is to evaluate some function of the whole set of measurements to accomplish a predefined task (*e.g.*, monitoring object detection and tracking) [1].

Outliers are defined as abnormal measurements which cannot be justified by the mere effects of sensing noise. They may be informative because associated to the insurgence of some critical event (*e.g.*, a forest fire), or caused by some sensor impairment. Outliers exhibit spatial and/or temporal correlation when indicating some event of interest (when the forest fire occurs, an entire region should have significantly higher temperature) [2], while can be assumed as independent if originated by defective sensors. Since the presence of uninformative outliers in the set of measurements may seriously disrupt the functionality of the system, the detection of the defective sensors is very important.

Many distributed fault detection (DFD) algorithms have already been proposed in the literature. The basic assumption is that all the nodes take noisy measurements of the same physical quantity. The algorithm devised in [3] allows each node to estimate its own functioning status by comparing the local measurement with the ones at neighboring nodes. In a first phase only a tendency status (*good, faulty, likely good, or likely faulty*) is determined. In the second phase the tendency status of neighbors are collected and associated to their measurements to obtain a more reliable assessment. In

[4] the local test is based on the comparison between the local measurement and the median of the measurements of the neighbors. Iterative algorithms are proposed in [5], [6], where the weighted-median and the weighted average criterion are considered, respectively. In both cases the local test weights the measurements of the neighbors by the confidence level obtained from the previous detection round, under the assumption of permanent node failure. Intermittent faults in sensing and in communication are considered in [7], which proposes an adaptive DFD algorithm with a decision threshold adjusted at each round.

In this work we propose a low complexity DFD algorithm for large WSNs. Our approach differs with respect to classical solutions in the fact that the comparison of the local and neighboring measurements in the first phase of the algorithm only determines whether an outlier is present in the measurement set. This is usually an easier task than attempting to identify it. Consider, for example, three sensors measuring some constant temperature, *e.g.*, with actual value $t = 20^\circ\text{C}$. Suppose that a non-defective sensor has a bounded measurement error, *e.g.*, ± 1 degree. Assume that the local measurement of the first sensor is $t_1 = 19.5^\circ\text{C}$, and that two other sensors provide $t_2 = 20.8^\circ\text{C}$ and $t_3 = 18.2^\circ\text{C}$, respectively. Supposing that there is no defective sensor, and taking into account the bounded measurement noise, one deduces that $t \in \mathbf{t}_1 = [t_1 - 1, t_1 + 1] = [18.5, 20.5]$, $t \in \mathbf{t}_2 = [19.8, 21.8]$, and $t \in \mathbf{t}_3 = [17.2, 19.2]$. Since $\mathbf{t}_1 \cap \mathbf{t}_2 \cap \mathbf{t}_3 = \emptyset$, there exists at least one outlier. However, determining which sensor produces the outlier is difficult as $\mathbf{t}_1 \cap \mathbf{t}_2 \neq \emptyset$ and $\mathbf{t}_1 \cap \mathbf{t}_3 \neq \emptyset$. This type of result is reminiscent of group testing (GT) [8], a statistical technique which allows to identify a small number of defective items within a large population, with only a limited number of tests. A DFD algorithm based on centralized GT is proposed in [9], under the assumption of vanishing ratio of defective sensors. The centralized approach suffers a significant drawback in the communication overhead necessary for node indexing and for the dissemination of the test results from the cluster heads to the rest of the WSN.

In this paper we propose an algorithm able to overcome the mentioned problems. The algorithm works in two stages: measurement testing and fault detection. In the first stage each node collects measurements from its neighborhood, performs a local test to determine whether outliers are present among them, and broadcasts the result. Such tests are easily available,

¹This work has been partly supported by the NEWCOM# NoE

e.g., in the context of bounded-error estimation [10], [11]. Fault detection is performed in the second stage, on the basis of the outcomes of the measurement tests performed in the neighborhood. The robustness of the decision strategy is measured in terms of detection probability (P_D) and of false alarm probability (P_{FA}), defined, respectively, as the probability of defective sensors to be correctly detected, and the probability of non-defective sensors to be wrongly diagnosed as defective. The algorithm is proposed in two variants, non-adaptive and adaptive, the second being better suited to WSNs with a large proportion of defective sensors.

II. NOTATIONS AND SYSTEM MODEL

A. Network model

Consider an infinite plane where sensor nodes are uniformly and independently deployed, with spatial density ρ , according to a 2D Poisson point process (PPP) [12]. Let \mathcal{S} denote the set of sensor nodes. A sensor is defective if it produces measurement outliers, i.e., measurement corrupted by noise samples whose characteristics differ significantly from those of non-defective sensors. An example of outlier models are introduced in Section IV-A. The sets of defective and non-defective (good) sensors are denoted by \mathcal{D} and \mathcal{G} respectively, with $\mathcal{D} \cup \mathcal{G} = \mathcal{S}$. The densities of defective and non-defective sensors are ρ_d and ρ_g , with $\rho = \rho_d + \rho_g$. Let θ_i denote the status (defective or non-defective) of an arbitrary sensor $i \in \mathcal{S}$, i.e., $\theta_i = 1$ if $i \in \mathcal{D}$ and $\theta_i = 0$ if $i \in \mathcal{G}$. In this work, it is assumed that θ_i remains constant during the time interval in which the DFD is performed.

Assume that any pair of nodes (i, j) can communicate if and only if $r_{i,j} \leq R_0$, where $r_{i,j}$ represents their physical distance and R_0 is the communication range, dependent on the transmission power, transmitter and receiver characteristics, and on the propagation scenario. Define $\mathcal{N}_i = \{j \in \mathcal{S} \mid 0 < r_{i,j} \leq R_0\}$ as the set of the neighbors of the node i . Let $\mathcal{U}_i = \mathcal{N}_i \cup \{i\}$. The probability of having n sensors in \mathcal{N}_i is [12]

$$\Pr\{|\mathcal{N}_i| = n\} = \frac{\mu^n}{n!} \exp(-\mu), \quad (1)$$

where $\mu = \rho\pi R_0^2$ is the average number of sensors in the neighborhood of sensor i . Similarly, one has $\mu_g = \rho_g\pi R_0^2$ and $\mu_d = \rho_d\pi R_0^2$. The probability of having n_d defective sensors and $n_g = n - n_d$ non-defective sensors in \mathcal{N}_i is [12]

$$\Pr\{|\mathcal{N}_i \cap \mathcal{G}| = n_g, |\mathcal{N}_i \cap \mathcal{D}| = n_d\} = \frac{\mu_d^{n_d} \mu_g^{n_g}}{n_d! n_g!} \exp(-\mu), \quad (2)$$

notice that $\mu = \mu_g + \mu_d$.

B. Outlier detection test

Consider a generic set of sensors $\mathcal{A} \subseteq \mathcal{S}$ and define

$$\varphi(\mathcal{A}) = \begin{cases} 1, & \text{if } \mathcal{A} \cap \mathcal{D} \neq \emptyset, \\ 0, & \text{otherwise,} \end{cases} \quad (3)$$

as the indicator function for the presence of defective sensors in \mathcal{A} . At a given time instant, let the random variable $M_i \in \mathbb{R}$ denote the measurement provided by sensor $i \in \mathcal{A}$. The measurements of all sensors in \mathcal{A} are gathered in $\mathbf{M}_{\mathcal{A}} =$

$[M_i]_{i \in \mathcal{A}} \in \mathbb{R}^{|\mathcal{A}|}$. The physical quantities measured by the sensors may be spatially correlated. The measurement noises produced by different sensors are assumed as uncorrelated. For a given realization $\mathbf{m}_{\mathcal{A}} \in \mathbb{R}^{|\mathcal{A}|}$ of $\mathbf{M}_{\mathcal{A}}$, denote $T(\mathbf{m}_{\mathcal{A}})$ the outcome of some *local outlier detection test*

$$T(\mathbf{m}_{\mathcal{A}}) = \begin{cases} 1, & \text{if at least an outlier is detected in } \mathbf{m}_{\mathcal{A}}, \\ 0, & \text{otherwise.} \end{cases} \quad (4)$$

Even if $\varphi(\mathcal{A}) = 1$, i.e., at least one defective sensor belongs to \mathcal{A} , the noise characteristics of the defective sensors may produce some realizations $\mathbf{m}_{\mathcal{A}}$ that do not allow the detection of the outliers. Thus, one introduces the probability of outlier detection

$$q_D(\mathbf{M}_{\mathcal{A}}) = \Pr\{T(\mathbf{M}_{\mathcal{A}}) = 1 \mid \varphi(\mathcal{A}) = 1\}. \quad (5)$$

Similarly, one considers the probability of false alarm

$$q_{FA}(\mathbf{M}_{\mathcal{A}}) = \Pr\{T(\mathbf{M}_{\mathcal{A}}) = 1 \mid \varphi(\mathcal{A}) = 0\}. \quad (6)$$

Specific properties of q_{FA} and q_D are provided in Section IV. They depend on the considered fault model.

III. DISTRIBUTED FAULT DETECTION ALGORITHM

This section proposes two variants of a DFD algorithm (non-adaptive and adaptive) to estimate the status of each sensor using some local outlier detection test T . The variant to be chosen depends on the defective sensor ratio μ_d/μ .

A. Non-adaptive DFD

The non-adaptive version of the proposed DFD algorithm consists of two successive phases, repeated for L rounds, followed by a final phase, whose outcome is, at each node i , an estimate $\hat{\theta}_i$ of θ_i .

Algorithm 1 Non-adaptive DFD

- 1) Initialize $\hat{\theta}_i^{(0)} = 0$ and $z_i^{(0)} = 0$ for all $i \in \mathcal{S}$.
- 2) For each round $1 \leq \ell \leq L$:

- Phase I: each sensor i broadcasts a packet containing its local measurement $m_i^{(\ell)}$, receives the measurements produced by the sensors in \mathcal{N}_i and performs the test (4) with outcome

$$y_i^{(\ell)} = T(\mathbf{m}_{\mathcal{U}_i}^{(\ell)}); \quad (7)$$

- Phase II: each sensor broadcasts $y_i^{(\ell)}$ generated in Phase I and updates $z_i^{(\ell)}$ as follows

$$z_i^{(\ell)} = z_i^{(\ell-1)} + \sum_{j \in \mathcal{U}_i} y_j^{(\ell)}, \quad (8)$$

to accumulate the number of neighbors that have a positive test result at Phase I up to round ℓ .

- 3) After L rounds:

- Phase III: each sensor i estimates its status θ_i

$$\hat{\theta}_i = \begin{cases} 1 & \text{(defective)} & \text{if } \frac{z_i^{(L)}}{L|\mathcal{U}_i|} > \gamma, \\ 0 & \text{(non defective)} & \text{otherwise,} \end{cases} \quad (9)$$

where γ is some threshold.

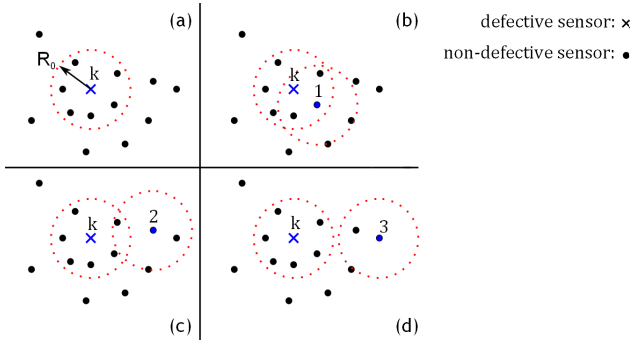


Fig. 1. Example where k is the only defective sensor. (a) shows that $z_k^1 = |\mathcal{U}_k| = 7$. In (b), sensor 1 belongs to \mathcal{U}_k and $z_1^1 = 4$. In (c), sensor 2 is not in \mathcal{U}_k but $r_{2,k} \leq 2R_0$, which results in $z_2^1 = 1$. In (d), sensor 3 has a distance $r_{3,k} > 2R_0$, so $z_3^1 = 0$.

The decision (9) in Phase III can result in both false alarm (with probability P_{FA}) and non-detection (with probability $1 - P_D$). The values of γ and L determine the trade-off between P_D and P_{FA} . They have to be adjusted to meet the targeted performance. When γ is too large, P_D may be low, with a significant number of defective sensors diagnosed as good. On the other hand, P_{FA} may be high for small values of γ . Increasing L provides a better averaging effect in (9), which reduces the variance of $\hat{\theta}_i$. However, L cannot be taken too large, to preserve the assumption of constant θ_i s during L rounds.

Example 1. Figure 1 depicts a toy example. Let k be the only defective sensor. Assume that both q_{FA} defined in (6) and q_{ND} defined in (5) are negligible: all the sensors in \mathcal{U}_k successfully detect outliers. Moreover consider $L = 1$ (unique round) and take $\gamma = 0.7$. At the end of Phase II, $z_k^{(1)} = |\mathcal{U}_k| = 7$, $z_1^{(1)} = |\mathcal{U}_k \cap \mathcal{U}_1| = 4$, $z_2^{(1)} = |\mathcal{U}_k \cap \mathcal{U}_2| = 1$, and $z_3^{(1)} = |\mathcal{U}_k \cap \mathcal{U}_3| = 0$. Since $z_k^{(1)} / |\mathcal{U}_k| = 1 > \gamma$, $z_1^{(1)} / |\mathcal{U}_1| \approx 0.67 < \gamma$, $z_2^{(1)} / |\mathcal{U}_2| = 0.25 < \gamma$, and $z_3^{(1)} / |\mathcal{U}_3| = 0 < \gamma$, only sensor k determines itself as defective, according to (9), while sensors 1, 2, and 3 diagnose themselves as non-defective.

As it will be discussed in Section VI, the non-adaptive DFD performs well when $\mu_d \ll \mu_g$. However, P_{FA} rapidly increases with μ_d / μ . In the critical situation, where every non-defective sensor has at least one defective sensor in its neighborhood, for all sensors $i \in \mathcal{S}$ the test outcome is $y_i^{(\ell)} = 1$ with a high probability. As a consequence, non-defective sensors are frequently diagnosed as defective.

B. Adaptive DFD

To improve the non-adaptive DFD algorithm for increasing μ_d / μ , an adaptive version is described in Algorithm 2. In the adaptive algorithm, sensors are allowed to take temporary decisions about their status. At each of the first L_1 rounds of the algorithm, measurements are exchanged with neighbors and a temporary decision $\hat{\theta}_i^{(\ell)}$ is taken. The decision at round $\ell - 1$ affects the set of sensors from which sensor i gets information at round ℓ , i.e., $\tilde{\mathcal{U}}_i^{(\ell)} = \{j \in \mathcal{U}_i \text{ s.t. } \hat{\theta}_j^{(\ell-1)} = 0\}$. This first part of the algorithm allows to discover most of the defective sensors, while keeping P_{FA} small.

Algorithm 2 Adaptive DFD

- 1) Set $\ell = 1$; $\hat{\theta}_i^{(0)} = 0$ for all $i \in \mathcal{S}$.
- 2) If $\hat{\theta}_i^{(\ell-1)} = 0$, sensor i broadcasts its measurement, calculates the sum $z_i^{(\ell)} = \sum_{j \in \tilde{\mathcal{U}}_i^{(\ell)}} y_j^{(\ell)}$, and performs the decision

$$\hat{\theta}_i^{(\ell)} = \begin{cases} 1, & \text{if } z_i^{(\ell)} / |\tilde{\mathcal{U}}_i^{(\ell)}| > \gamma_1, \\ 0, & \text{otherwise.} \end{cases} \quad (10)$$

- 3) If $\hat{\theta}_i^{(\ell-1)} = 1$, sensor i is silent, i.e., it does not broadcast its measurement and performs the decision $\hat{\theta}_i^{(\ell)} = y_i^{(\ell)}$.
- 4) $\ell = \ell + 1$. If $\ell \leq L_1$, go to 2.
- 5) After round L_1 :

- each sensor i such that $\hat{\theta}_i^{(L_1)} = 0$ sets $z_i^{(L_1)} = 0$.
- each sensor i such that $\hat{\theta}_i^{(L_1)} = 1$ is determined as defective. It stops broadcasting its measurements. It does not participate in the non-adaptive DFD during the following rounds.

- 6) Non-adaptive DFD is performed during L_2 rounds with threshold γ_2 .

At the beginning of the algorithm, a large amount of non-defective sensors may be diagnosed as defective. In Step 3, sensors become silent and stop broadcasting their measurements if they are temporarily determined as defective. Step 4 gives a chance to revive to the non-defective sensors erroneously deemed defective (the P_{FA} is high at the beginning): if $\hat{\theta}_i^{(\ell-1)} = 1$, the decision is made only based on the local test of sensor i . In few rounds, P_{FA} can thus be reduced to an acceptable level. The following L_2 rounds help to increase P_D , as shown in Section VI.

IV. OUTLIER DETECTION TEST

This section presents an example of outlier model and a possible detection test.

A. Outlier model

Assume that each sensor i gets a noisy observation of some scalar physical quantity ϕ

$$m_i = \phi + w_i, \quad \forall i \in \mathcal{S}. \quad (11)$$

The components w_i of the measurement noise in (11) are assumed to be realizations of independent Gaussian random variables W_i with the same standard deviation σ . For good sensors (i.e., $i \in \mathcal{G}$), the measurement noise is assumed zero-mean. For the defective sensors, W_i has a random bias E_i uniformly distributed over $[-3\eta\sigma, 3\eta\sigma]$, where $\eta > 1$. With this setting, each defective sensor has a different value of bias.

B. Outlier detection test

Consider the interval $[m_i] = [m_i - \Delta, m_i + \Delta]$ of width 2Δ centered around the measurement m_i . If $i \in \mathcal{G}$, one has $\Pr\{\phi \in [m_i]\} = 1 - \text{erfc}\left(\frac{\Delta}{\sqrt{2}\sigma}\right)$. If $i \in \mathcal{D}$, similarly, one evaluates the probability that $\phi \notin [m_i]$ as a function of σ , Δ , and η , when the fault model is that of Section IV-A.

Consider now the intersection of all $[m_i], \forall i \in \mathcal{A}$:

$$\left[\widehat{\phi}(\mathbf{m}_{\mathcal{A}}) \right] = \bigcap_{i \in \mathcal{A}} [m_i]. \quad (12)$$

With and without presence of outliers, one is able to evaluate $\Pr \left\{ \left[\widehat{\phi} \right] = \emptyset \right\}$ as a function of σ , η , and Δ . These two quantities may be used to define a low-complexity outlier detection test

$$T(\mathbf{m}_{\mathcal{A}}) = \begin{cases} 1, & \text{if } \left[\widehat{\phi}(\mathbf{m}_{\mathcal{A}}) \right] = \emptyset, \\ 0, & \text{else.} \end{cases} \quad (13)$$

With this test, q_{FA} and q_{D} defined in (6) and (5) only depend on n_{g} and n_{d} , where $n_{\text{g}} = |\mathcal{A} \cap \mathcal{G}|$ and $n_{\text{d}} = |\mathcal{A} \cap \mathcal{D}|$. As $\left[\widehat{\phi}(\mathbf{m}_{\mathcal{A}} \cup m_j) \right] \subseteq \left[\widehat{\phi}(\mathbf{m}_{\mathcal{A}}) \right]$, $T(\mathbf{m}_{\mathcal{A}})$ is more probable to get close to 1 as the number of measurements increase. Therefore, $q_{\text{FA}}(n_{\text{g}})$ is an increasing function of n_{g} , and $q_{\text{D}}(n_{\text{g}}, n_{\text{d}})$ is an increasing function of both n_{g} and n_{d} .

V. THEORETICAL ANALYSIS

In this section we characterize analytically the probabilities of fault detection P_{D} and of false alarm P_{FA} of the proposed DFD algorithm, when the decision is taken after one single round of the non-adaptive algorithm, *i.e.*, after a single collection of measurements from neighbors. Conditions for the convergence of the adaptive algorithm are also evaluated.

A. P_{D} and P_{FA} for a single round of the non-adaptive DFD algorithm

Consider the event

$$\mathcal{E}_{i,\text{D}}(n) = \left\{ \sum_{j \in \mathcal{U}_i} Y_j > \gamma(n+1) \mid \theta_i = 1, |\mathcal{N}_i| = n \right\}, \quad (14)$$

representing, according to (9), the situation where a defective sensor i is detected as defective considering data coming from n neighbors. The detection probability of a defective sensor can be hence expressed as

$$P_{\text{D}} = \sum_{n=0}^{\infty} \Pr \{ \mathcal{E}_{i,\text{D}}(n) \} \Pr \{ |\mathcal{N}_i| = n \} \quad (15)$$

As sensors cannot perform outlier detection when they have no neighbors, $\Pr \{ \mathcal{E}_{i,\text{D}}(0) \} = 1$.

Similarly, consider a non-defective sensor i , the conditional false alarm event is

$$\mathcal{E}_{i,\text{FA}}(n) = \left\{ \sum_{j \in \mathcal{U}_i} Y_j > \gamma(n+1) \mid \theta_i = 0, |\mathcal{N}_i| = n \right\}. \quad (16)$$

The probability of false alarm is given by

$$P_{\text{FA}} = \sum_{n=0}^{\infty} \Pr \{ \mathcal{E}_{i,\text{FA}}(n) \} \Pr \{ |\mathcal{N}_i| = n \}. \quad (17)$$

Lemma 1. Both P_{D} and P_{FA} are decreasing functions of γ .

Proof: The monotonicity of P_{D} and P_{FA} with respect to γ is obvious, since if $\gamma_1 > \gamma_2$, $\sum_{j \in \mathcal{U}_i} Y_j > \gamma_1(n+1)$ implies $\sum_{j \in \mathcal{U}_i} Y_j > \gamma_2(n+1)$. ■

Consider now the event

$$\mathcal{Y}_{j,\text{D}}(n) = \left\{ Y_j = 1 \mid j \in \mathcal{U}_i, \theta_i = 1, |\mathcal{N}_i| = n \right\}, \quad (18)$$

with $n > 0$ and $j \in \mathcal{U}_i$. For any $j' \in \mathcal{U}_i$ such that $j \neq j'$, $\mathcal{Y}_{j,\text{D}}(n)$ and $\mathcal{Y}_{j',\text{D}}(n)$ are dependent. Their dependence comes from the fact that in general $\mathcal{N}_j \cap \mathcal{N}_{j'} \neq \emptyset$. The probability mass function (pmf) of $\sum_{j \in \mathcal{U}_i} Y_j$, necessary to evaluate P_{D} , is thus quite difficult to evaluate, since the dependency between the Y_j s is not explicit. The exact form of P_{D} then becomes quite complicated, but we can make the following conjecture.

Conjecture 1. P_{D} is an increasing function of μ_{d} . Moreover, $\mu_{\text{d}} P_{\text{D}}$ is a convex function of μ_{d} .

The decision concerning the status of sensor i at the end of the DFD algorithm is based on the results of the local tests of all the sensors in \mathcal{U}_i . If the local test results are more likely to be 1, the final decision is reasonably more probable to be 1. Thus, one conjectures that the monotonicity of P_{D} with respect to μ_{d} is inherited by the monotonicity of $\Pr \{ Y_i = 1 \mid \theta_i = 1 \}$ and of $\Pr \{ Y_i = 1 \mid \theta_i = 0 \}$ with respect to μ_{d} . Define the function

$$h(n_{\text{g}}, n_{\text{d}}) = \begin{cases} q_{\text{FA}}(n_{\text{g}}), & \text{if } n_{\text{d}} = 0, \\ 1 - q_{\text{ND}}(n_{\text{g}}, n_{\text{d}}), & \text{if } n_{\text{d}} \neq 0, \end{cases} \quad (19)$$

representing the probability of having $Y_i = 1$ in a local test, knowing that $|\mathcal{G} \cap \mathcal{U}_i| = n_{\text{g}}$, $|\mathcal{D} \cap \mathcal{U}_i| = n_{\text{d}}$. Thus, one may introduce

$$\begin{aligned} f_1(\mu_{\text{g}}, \mu_{\text{d}}) &= \Pr \{ Y_i = 1 \mid \theta_i = 0 \} \\ &= \sum_{n_{\text{g}}=0}^{\infty} \sum_{n_{\text{d}}=0}^{\infty} h(n_{\text{g}}+1, n_{\text{d}}) \frac{\mu_{\text{d}}^{n_{\text{d}}} \mu_{\text{g}}^{n_{\text{g}}}}{n_{\text{d}}! n_{\text{g}}!} \exp(-\mu_{\text{d}} - \mu_{\text{g}}), \end{aligned} \quad (20)$$

and

$$\begin{aligned} f_2(\mu_{\text{g}}, \mu_{\text{d}}) &= \Pr \{ Y_i = 1 \mid \theta_i = 1 \} \\ &= \sum_{n_{\text{g}}=0}^{\infty} \sum_{n_{\text{d}}=0}^{\infty} h(n_{\text{g}}, n_{\text{d}}+1) \frac{\mu_{\text{d}}^{n_{\text{d}}} \mu_{\text{g}}^{n_{\text{g}}}}{n_{\text{d}}! n_{\text{g}}!} \exp(-\mu_{\text{d}} - \mu_{\text{g}}). \end{aligned} \quad (21)$$

One has

$$\begin{aligned} \frac{\partial f_1}{\partial \mu_{\text{g}}} &= \sum_{n_{\text{g}}=0}^{\infty} \sum_{n_{\text{d}}=0}^{\infty} \frac{\mu_{\text{d}}^{n_{\text{d}}} \mu_{\text{g}}^{n_{\text{g}}}}{n_{\text{d}}! n_{\text{g}}!} \exp(-\mu_{\text{d}} - \mu_{\text{g}}) \\ &\quad \cdot (h(n_{\text{g}}+2, n_{\text{d}}) - h(n_{\text{g}}+1, n_{\text{d}})) > 0, \end{aligned} \quad (22)$$

as $h(n_{\text{g}}+2, n_{\text{d}}) > h(n_{\text{g}}+1, n_{\text{d}})$. Similarly, one has $\frac{\partial f_1}{\partial \mu_{\text{d}}} > 0$, $\frac{\partial f_2}{\partial \mu_{\text{g}}} > 0$, and $\frac{\partial f_2}{\partial \mu_{\text{d}}} > 0$. We can conjecture that P_{D} is an increasing function of μ_{d} and $\mu_{\text{d}} P_{\text{D}}$ is a convex function of μ_{d} . These conjectures are verified experimentally in Section VI.

The monotonicity of P_{FA} with respect to μ_{g} is more complicated to analyze qualitatively. As μ_{g} increases, a good sensor that has a defective sensor as its neighbor is more probable to be diagnosed as defective. On the other hand, the probability of having one defective sensor as neighbor decreases with μ_{g} .

Conjecture 2. P_{FA} is a decreasing function of μ_{g} , while $\mu_{\text{g}} P_{\text{FA}}$ is an increasing and concave function of μ_{g} .

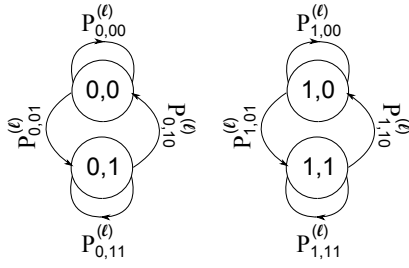


Fig. 2. Evolution with the rounds of the state of a sensor when applying the adaptive DFD, notice that both automatras do not evolve in parallel independently.

Again, Conjecture 2 is verified experimentally in Section VI.

B. Convergence of the adaptive algorithm

Let the pair $(\theta_i, \hat{\theta}_i)$ denote the *state* of sensor i , where θ_i is its actual status and $\hat{\theta}_i$ is its estimated status. Among the four possible states, (0, 0) and (1, 1) are states resulting from a correct decision, (0, 1) corresponds to a false alarm, and (1, 0) corresponds to a non-detection. Let $\mu_{\theta_i, \hat{\theta}_i}^{(\ell)}$ denote the density of sensors in the state $(\theta_i^{(\ell)}, \hat{\theta}_i^{(\ell)})$ at round ℓ . The aim of this section is to characterize the evolution of $\mu_{\theta_i, \hat{\theta}_i}^{(\ell)}$ to determine whether the adaptive algorithm converges to a steady state. The transition between states is described in Figure 2.

Before the first round of the adaptive algorithm, one has $\mu_{00}^{(0)} = \mu_g$, $\mu_{10}^{(0)} = \mu_d$, and $\mu_{01}^{(0)} = \mu_{11}^{(0)} = 0$. Note that at any round ℓ , $\mu_{00}^{(\ell)} + \mu_{01}^{(\ell)} = \mu_g$ and $\mu_{10}^{(\ell)} + \mu_{11}^{(\ell)} = \mu_d$. The transitions from round ℓ to round $\ell + 1$ are illustrated in Figure 2. Since the actual status of a sensor is assumed constant during the rounds of the DFD algorithm, the only possible transitions are between states (0, 0) and (0, 1) and between states (1, 0) and (1, 1). The evolution of the densities between rounds ℓ and $\ell + 1$ is given by

$$\begin{bmatrix} \mu_{00}^{(\ell+1)} \\ \mu_{01}^{(\ell+1)} \\ \mu_{10}^{(\ell+1)} \\ \mu_{11}^{(\ell+1)} \end{bmatrix} = \begin{bmatrix} P_{0,00}^{(\ell)} & P_{0,10}^{(\ell)} & 0 & 0 \\ P_{0,01}^{(\ell)} & P_{0,11}^{(\ell)} & 0 & 0 \\ 0 & 0 & P_{1,00}^{(\ell)} & P_{1,10}^{(\ell)} \\ 0 & 0 & P_{1,01}^{(\ell)} & P_{1,11}^{(\ell)} \end{bmatrix} \begin{bmatrix} \mu_{00}^{(\ell)} \\ \mu_{01}^{(\ell)} \\ \mu_{10}^{(\ell)} \\ \mu_{11}^{(\ell)} \end{bmatrix}, \quad (23)$$

with:

$P_{0,01}^{(\ell)}$ is the transition probability from (0, 0) to (0, 1). It is equal to P_{FA} in the first round. The sensors considered as defective in the previous round are silent, so P_{FA} is a function of $\mu_{00}^{(\ell)}$ and $\mu_{10}^{(\ell)}$, denoted as $P_{\text{FA}}(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$. Obviously, $P_{0,00}^{(\ell)} = 1 - P_{0,01}^{(\ell)}$.

$P_{0,10}^{(\ell)}$ is the transition probability from (0, 1) to (0, 0). Suppose that sensor i is in the state (0, 1). Recall that in the proposed algorithm, sensor i does not broadcast, but still collects measurements from its neighbors in state (0, 0) or (1, 0). The measurement of sensor i is only considered in its own local test. From (20), one deduces that $P_{0,10}^{(\ell)} = 1 - f_1(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$ and $P_{0,11}^{(\ell)} = f_1(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$.

$P_{1,00}^{(\ell)}$ is the transition probability from (1, 0) to (1, 0). This is the probability of mis-detection of a defective sensor, *i.e.*, $1 - P_{\text{D}}(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$. Moreover $P_{1,01}^{(\ell)} = P_{\text{D}}(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$.

$P_{1,10}^{(\ell)}$ is the transition probability from (1, 1) to (1, 0), which is similar to $P_{0,10}^{(\ell)}$. One has $P_{1,10}^{(\ell)} = 1 - f_2(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$ and $P_{1,11}^{(\ell)} = f_2(\mu_{00}^{(\ell)}, \mu_{10}^{(\ell)})$.

From (23), one gets

$$\begin{aligned} \mu_{00}^{(\ell+1)} &= \mu_{00}^{(\ell)} P_{0,00}^{(\ell)} + (\mu_g - \mu_{00}^{(\ell)}) P_{0,10}^{(\ell)} \\ &= \mu_{00}^{(\ell)} (f_1 - P_{\text{FA}}) + \mu_g (1 - f_1) \end{aligned} \quad (24)$$

and

$$\begin{aligned} \mu_{10}^{(\ell+1)} &= \mu_{10}^{(\ell)} P_{1,00}^{(\ell)} + (\mu_d - \mu_{10}^{(\ell)}) P_{1,10}^{(\ell)} \\ &= \mu_{10}^{(\ell)} (f_2 - P_{\text{D}}) + \mu_d (1 - f_2). \end{aligned} \quad (25)$$

To lighten notations, the arguments $\mu_{00}^{(\ell)}$ and $\mu_{10}^{(\ell)}$ of P_{FA} , P_{D} , f_1 , and f_2 have been omitted in (24), (25), and in what follows when there is no confusion.

Let μ_{00}^* and μ_{10}^* be the values at equilibrium of $\mu_{00}^{(\ell)}$ and $\mu_{10}^{(\ell)}$, respectively. From (24) and (25), one deduces that μ_{00}^* and μ_{10}^* should satisfy

$$\mu_g - \mu_{00}^* = \frac{\mu_{00}^* P_{\text{FA}}(\mu_{00}^*, \mu_{10}^*)}{1 - f_1(\mu_{00}^*, \mu_{10}^*)}, \quad (26)$$

and

$$\mu_d - \mu_{10}^* = \frac{\mu_{10}^* P_{\text{D}}(\mu_{00}^*, \mu_{10}^*)}{1 - f_2(\mu_{00}^*, \mu_{10}^*)}. \quad (27)$$

Lemma 2. Assume that Conjectures 1 and 2 hold, then (26) and (27) admit a unique solution.

Proof: In the case of (27), when we study the left-hand side, $\mu_g - \mu_{00}^*$ is a continuous and monotonically decreasing function of μ_{00}^* over $[0, \mu_g]$, with

$$\lim_{\mu_{00}^* \rightarrow 0} \mu_g - \mu_{00}^* = \mu_g \quad \text{and} \quad \lim_{\mu_{00}^* \rightarrow \mu_g} \mu_g - \mu_{00}^* = 0.$$

We need now to study the behavior of the right hand side for a fixed value of μ_{10}^* . As f_1 and $\mu_{00}^* P_{\text{FA}}$ are increasing functions of μ_{00}^* , obviously, $\mu_{00}^* P_{\text{FA}} / (1 - f_1)$ is an increasing function of μ_{00}^* . As $0 < f_1 < 1$, one has

$$\lim_{\mu_{00}^* \rightarrow 0} \frac{\mu_{00}^* P_{\text{FA}}(\mu_{00}^*, \mu_{10}^*)}{1 - f_1(\mu_{00}^*, \mu_{10}^*)} = 0.$$

Thus for any fixed μ_{10}^* , and for any decision threshold $0 < \gamma < 1$, there exists exactly one μ_{00}^* satisfying (26).

Similarly, one may show that for any fixed μ_{00}^* , and for any decision threshold $0 < \gamma < 1$, there exists only one μ_{10}^* satisfying (27). \blacksquare

Lemma 3. Let $G_1(\mu_{00}, \mu_{10}) = (\mu_{00} - \mu_g) \cdot f_1(\mu_{00}, \mu_{10})$ and $G_2(\mu_{00}, \mu_{10}) = (\mu_{10} - \mu_d) \cdot f_2(\mu_{00}, \mu_{10})$. Assume that

$$\frac{\partial G_1}{\partial \mu_{00}} > 0, \quad \forall \mu_{00} \in [0, \mu_g], \mu_{10} \in [0, \mu_d], \quad (28)$$

$$\frac{\partial G_2}{\partial \mu_{10}} > 0, \quad \forall \mu_{00} \in [0, \mu_g], \mu_{10} \in [0, \mu_d], \quad (29)$$

and Conjectures 1 and 2 hold, then the equilibrium point (μ_{00}^*, μ_{10}^*) is locally stable.

Proof: Consider a linearization of (24) around equilibrium with $\mu_{00}^{(\ell)} = \mu_{00}^* + \delta_0^{(\ell)}$, one gets

$$\delta_0^{(\ell+1)} = \delta_0^{(\ell)} \cdot \frac{\partial F_1}{\partial \mu_{00}}(\mu_{00}^*, \mu_{10}^*), \quad (30)$$

where $F_1(\mu_{00}, \mu_{10}) = \mu_{00} \cdot (f_1 - P_{FA}) + \mu_g \cdot (1 - f_1)$, with $0 \leq \mu_{00} \leq \mu_g$ and $0 \leq \mu_{10} \leq \mu_d$. Since $\frac{\partial F_1}{\partial \mu_{00}}(\mu_{00}^*, \mu_{10}^*)$ does not depend on ℓ , the system (30) is locally stable around μ_{00}^* if

$$\left| \frac{\partial F_1}{\partial \mu_{00}}(\mu_{00}^*, \mu_{10}^*) \right| < 1. \quad (31)$$

Notice that

$$\frac{\partial F_1}{\partial \mu_{00}} = \underbrace{f_1}_{<1} - \underbrace{(\mu_g - \mu_{00}) \frac{\partial f_1}{\partial \mu_{00}}}_{>0} - \underbrace{\frac{\partial(\mu_{00} P_{FA})}{\partial \mu_{00}}}_{>0} < 1 \quad (32)$$

holds since $\mu_{00} P_{FA}$ has been assumed to be an increasing function of μ_{00} . On the other hand, one needs $\frac{\partial F_1}{\partial \mu_{00}} > -1$, which leads to

$$\begin{aligned} & P_{FA}(\mu_{00}, \mu_{10}) + \mu_{00} \frac{\partial P_{FA}}{\partial \mu_{00}} + (\mu_g - \mu_{00}) \frac{\partial f_1}{\partial \mu_{00}} - f_1 \\ &= \frac{\partial(\mu_{00} \cdot P_{FA})}{\partial \mu_{00}} + (\mu_g - \mu_{00}) \frac{\partial f_1}{\partial \mu_{00}} - f_1 \\ &= \frac{\partial(\mu_{00} \cdot P_{FA})}{\partial \mu_{00}} - \frac{\partial G_1}{\partial \mu_{00}} < 1 \end{aligned} \quad (33)$$

One has assumed that $\mu_{00} P_{FA}$ is concave, then

$$\frac{\partial(\mu_{00} P_{FA})}{\partial \mu_{00}} < \lim_{\mu_{00} \rightarrow 0} \frac{\partial(\mu_{00} P_{FA})}{\partial \mu_{00}} \leq 1. \quad (34)$$

From (34) and (28), one obtains (33), thus (31) holds.

We can also prove that (27) is stable at μ_{10}^* in the similar way if Conjecture 1 and (29) hold. We conclude that the equilibrium point (μ_{00}^*, μ_{10}^*) is locally stable if the conditions mentioned in Lemma 3 are satisfied. ■

VI. SIMULATION RESULTS

Consider a WSN of 1000 sensors randomly deployed according to a 2D PPP over a square of size 10×10 units, with $\mu = 7$. To avoid boundary effects, only the sensors in the square of size $(10 - 2R_0) \times (10 - 2R_0)$ units are considered in the evaluations of P_D and P_{FA} . In the outlier model presented in Section IV-A, $\eta = 10$. All results have been averaged over 2000 realizations of the WSN.

First, Conjecture 1 and 2 are verified experimentally in Figure 3 and 4 respectively, for different values of parameters. The conditions in (28) and (29) are also verified numerically in Figure 5.

The performance of a single round of the non-adaptive DFD algorithm described in Section III-A is evaluated. Setting $\mu = 7$, Figure 6 shows P_D as a function of P_{FA} for different values of the threshold $\gamma \in [0, 1]$ and for various $\mu_d/\mu \in \{0.02, 0.05, 0.1, 0.2\}$.

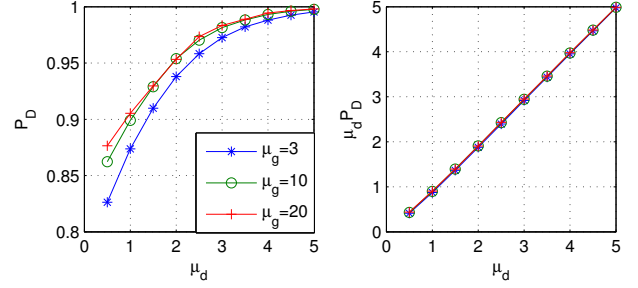


Fig. 3. Verification of Conjecture 1, with $\mu_g \in \{3, 10, 20\}$, and $\gamma = 0.8$.

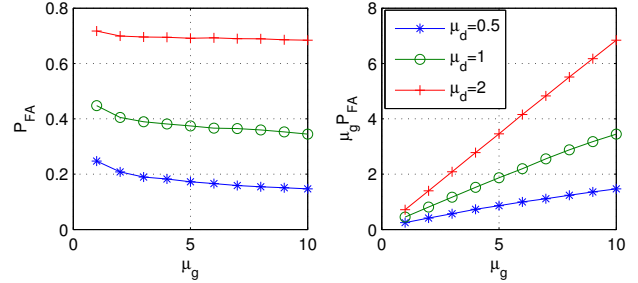


Fig. 4. Verification of Conjecture 2, with fixed $\mu_d \in \{0.5, 1, 2\}$, and $\gamma = 0.8$.

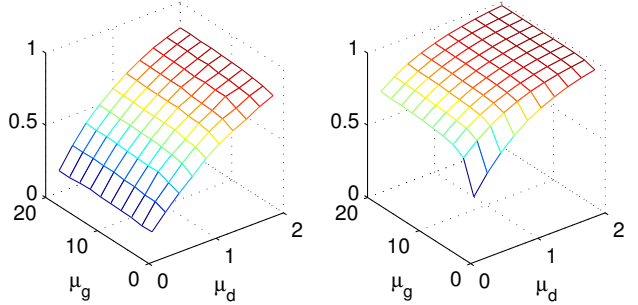


Fig. 5. Verification of $\frac{\partial G_1}{\partial \mu_{00}} > 0$ (left) and $\frac{\partial G_2}{\partial \mu_{10}} > 0$ (right) with $\mu_g = 20$ and $\mu_d = 2$.

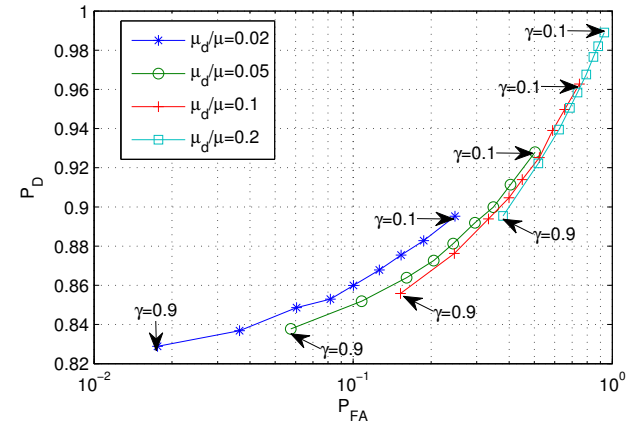


Fig. 6. P_D as a function of P_{FA} for the non-adaptive DFD with different γ and different densities of defective sensors; here $\mu = 7$.

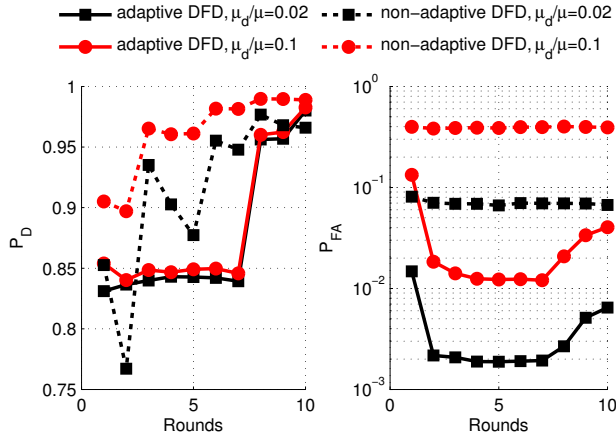


Fig. 7. P_D (left) and P_{FA} (right) as a function of the number of rounds for the DFD algorithms, with $\mu = 7$, and $\mu_d/\mu \in \{0.02, 0.1\}$

The roles played by the number of rounds on the achievable performance of both non-adaptive DFD and adaptive DFD are then investigated in Figure 7, which shows P_D and P_{FA} as a function of the round number $1 \leq \ell \leq L$, for $L = 10$ and $\mu = 7$. We set $\gamma = 0.6$ in the non-adaptive DFD. As can be observed, P_D rapidly increases, whereas P_{FA} almost remains constant. In the case of adaptive DFD, we consider $L_1 = 7$, $L_2 = 3$, $\gamma_1 = 0.95$, and $\gamma_2 = 0.6$. Here γ_1 is chosen to be large to make the adaptive DFD easier to converge and to result in small P_{FA} . P_{FA} can be reduced significantly during the first seven adaptive rounds. During the three following non-adaptive rounds, P_D improves rapidly, while P_{FA} remains small. This is due to the fact that most of the defective sensors have been detected and turned off during the previous adaptive phase, hence at the beginning of the non-adaptive rounds the residual μ_d/μ is much lower than the initial μ_d/μ . Thus, the adaptive DFD algorithm performs better than the non-adaptive DFD algorithm: to achieve the similar P_D , the resulting P_{FA} is much smaller by using the adaptive algorithm.

In Figure 8, the comparison between the proposed DFD and a reference DFD in [3] is reported in terms of P_D and P_{FA} , after 10 rounds of the algorithm, in the case where $\mu = 7$. A fair comparison between the two algorithms is quite difficult because they are based on different tests with non homogeneous parameters. In fact, in the reference DFD, the only parameter one can control is the threshold of a local test (similar to Δ in our case), whereas the performance of the proposed DFD depends on Δ , γ_1 , γ_2 , L_1 , and L_2 . For example, we fix $\Delta = 3\sigma$, $L_1 = 7$, $L_2 = 3$, $\gamma_1 = 0.95$ and vary the value of γ_2 . As can be seen in Figure 8, one may have better performance (with respect to P_{FA} and P_D) if the parameters are properly chosen in the proposed DFD.

VII. CONCLUSIONS

This work proposed two variants of a two-staged DFD algorithm, which allows each node to decide whether its sensor is producing outliers. The performance of the non-adaptive algorithm has been theoretically characterized. An analytical framework for the analysis of the adaptive variant has been sketched. The performance of both variants have been characterized by simulations, which enables to draw insights

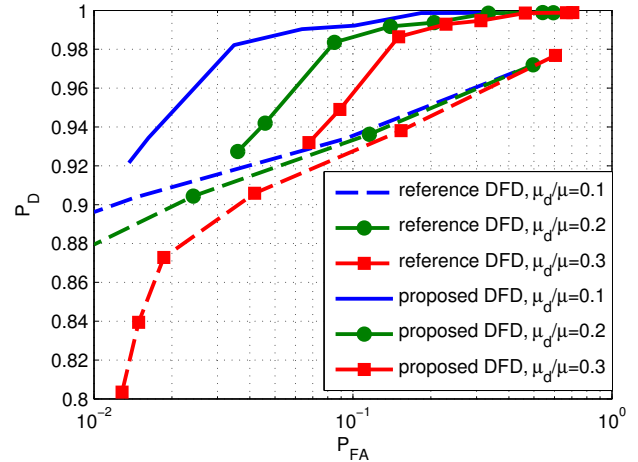


Fig. 8. P_{FA} as a function of P_D for the reference DFD and the proposed DFD after 10 rounds, with $\mu_d/\mu \in \{0.1, 0.2, 0.3\}$.

on the impact of the algorithm parameters (number of rounds, local test threshold) and of the network topology (density of faulty sensors, size of the neighborhood) on the trade-off between P_D and P_{FA} . The future work will be devoted to the study of other noise and outlier models and to a more formal verification of the conjectures.

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