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# AN ALTERNATIVE PROOF FOR THE IDENTIFIABILITY OF INDEPENDENT VECTOR ANALYSIS USING SECOND ORDER STATISTICS

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## ABSTRACT

In this paper, we present an alternative proof for characterizing the (non-) identifiability conditions of independent vector analysis (IVA). IVA extends blind source separation (BSS) to several mixtures by taking into account statistical dependencies across mixtures. We focus on IVA in the presence of real Gaussian data with temporally independent and identically distributed samples. This model is always non-identifiable when each mixture is considered separately. However, it can be shown to be generically identifiable within the IVA framework. Our proof differs from previous ones by being based on direct factorization of a closed-form expression for the Fisher information matrix (FIM). Our analysis is based on a rigorous linear algebraic formulation, and leads to a new type of factorization of a structured matrix. Therefore, the proposed approach is of potential interest for a broader range of problems.

**Index Terms**— Blind source separation, independent vector analysis, uniqueness, matrix factorization, data fusion

## 1. INTRODUCTION

In this theoretical paper, we present an alternative proof for the identifiability of independent vector analysis (IVA) [1]. By identifiability we mean the minimal set of constraints such that the source estimates are subject only to the unavoidable scale and permutation indeterminacies. Characterizing the uniqueness and identifiability properties of a model is necessary in order to achieve interpretability, i.e., attach physical meaning to its output (e.g., [2, 3]).

Consider  $T$  observations of  $K$  vectors  $\mathbf{x}^{[k]}(t)$ , modelled as

$$\mathbf{x}^{[k]}(t) = \mathbf{A}^{[k]} \mathbf{s}^{[k]}(t) \quad 1 \leq t \leq T, 1 \leq k \leq K, \quad (1)$$

where  $\mathbf{A}^{[k]}$  are  $M \times M$  invertible matrices that may be different  $\forall k$ , and  $\mathbf{x}^{[k]}(t)$  and  $\mathbf{s}^{[k]}(t)$  are  $M \times 1$  vectors. For fixed  $k$ , each mixture (1) corresponds to classical blind source separation (BSS). In IVA, the elements of the  $K \times 1$  vector  $\mathbf{s}_i(t) = [s_i^{[1]}(t), \dots, s_i^{[K]}(t)]^\top$ ,  $i = 1, \dots, M$ , are statistically *dependent* whereas the pairs  $(\mathbf{s}_i(t), \mathbf{s}_j(t))$  are statistically *independent* for all  $i \neq j \in \{1, \dots, M\}$ . Therefore, IVA aims at extracting  $M$  mutually independent vector elements (whence its name) from  $K$  sets of measurements by exploiting not only the statistical *independ-*

*ence within each set of measurements* but also the *dependence across sets of measurements*. In fact, some of these assumptions can be relaxed while maintaining the identifiability of the model [4, 5, 6].

The original motivation for IVA was dealing with the arbitrary permutation in frequency domain BSS of convolutive mixtures, because the dependence across mixtures imposes the estimates of the corresponding sources in all mixtures to be aligned, i.e., same permutation [1]. Recently, IVA has shown useful for multimodal data analysis [7], temporal dynamics [8], and more.

Among the many possible variants of IVA, in this paper, we deal only with the identifiability of instantaneous invertible mixtures of real Gaussian random variables with temporally independent and identically distributed (i.i.d.) samples. Our motivation to study this specific setup, proposed by [9], is that each individual BSS is always non-identifiable (e.g., [10, 11]). Hence, the identifiability results reflect only the added value of the multi-set diversity (i.e., the link between datasets) [12, 13, 14]. Identifiability in the presence of additional types of diversity is discussed in Section 4. A motivation to focus on second-order statistics (SOS) is that non-identifiability of IVA is associated only with the presence of source samples with Gaussian distribution [5, 6]. Furthermore, the relative simplicity and mathematical tractability of this model allow us to reformulate our derivations in form of a new type of factorization of a structured matrix, which may be useful beyond this specific task. We thus define the cross-covariance matrix  $\mathbf{S}^{[k,l]} \triangleq E \{ \mathbf{s}^{[k]}(t) \mathbf{s}^{[l]H}(t) \} \in \mathbb{C}^{M \times M}$ ,  $k, l = 1, \dots, K$ , which is diagonal due to the statistical independence, and  $\mathbf{S}_{ii} \triangleq E \{ \mathbf{s}_i(t) \mathbf{s}_i^H(t) \} \in \mathbb{C}^{K \times K}$ . In what follows, we assume that  $\mathbf{S}_{ii}$  are invertible  $\forall i$ .

In this paper,  $\cdot^\top$ ,  $\cdot^{-\top}$ ,  $\cdot^H$  denote transpose, inverse transpose, and Hermitian, respectively. Scalars, vectors and matrices are denoted as normal ( $\lambda$ ,  $M$ ), bold lower ( $\boldsymbol{\lambda}$ ,  $\mathbf{x}$ ) or upper case ( $\mathcal{H}$ ,  $\mathbf{A}$ ) symbols, respectively. The  $(k, l)$ th entry of  $\mathbf{P}$  is denoted  $p^{[k,l]}$  and the  $k$ th entry of  $\boldsymbol{\lambda}$  is  $\lambda^{[k]}$ .  $\otimes$ ,  $\otimes$ ,  $\odot$ , denote the Hadamard, Kronecker and column-wise Khatri-Rao products, respectively.  $\text{diag}\{\boldsymbol{\lambda}\}$  is a matrix whose diagonal is  $\boldsymbol{\lambda}$ , whereas  $\text{diag}\{\mathbf{X}\}$  is a diagonal matrix by setting to zero the off-diagonal terms of a square matrix  $\mathbf{X}$ .  $\mathbf{I}_M$  and  $\mathbf{0}_{M \times N}$  denote an identity or all-zero matrix, respectively.

The rest of this paper is organized as follows. In Section 2 we present the principal concepts behind the previous proofs, and then explain the underlying idea of ours. Section 3 is dedicated to the actual mathematical analysis, and presents our proof in detail. It is concluded with a

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short discussion of the meaning of these results. Section 4 concludes our work with a broader view of uniqueness and identifiability, especially in the context of data fusion.

## 2. BACKGROUND: PREVIOUS PROOFS VS. OUR APPROACH

A first proof was presented by Vía et al. [4], who study SOS-based IVA in the complex domain. Their proof is based directly on the definition of identifiability, by characterizing the admissible set of uniquely-defined solutions of the IVA model, up to the unavoidable indeterminacies. The derivation is based on the observation that for normalized sources, i.e.,  $\mathbf{S}^{[k,k]} = \mathbf{I}_M$ , the trivial ambiguities boil down to  $\mathbf{S}^{[k,l]} = \mathbf{Q}^{[k]} \widehat{\mathbf{S}}^{[k,l]} \mathbf{Q}^{[l]}$ , where  $\mathbf{Q}^{[k]} \in \mathbb{C}^{M \times M}$  are unitary matrices, and both  $\mathbf{S}^{[k,l]}$  and  $\widehat{\mathbf{S}}^{[k,l]}$  are elements of the admissible set. By reformulating this relationship as  $\Sigma^{[k,l]} = \mathbf{Q}^{[k]} \widehat{\Sigma}^{[k,l]} \mathbf{Q}^{[l]}$ , where  $\Sigma^{[k,l]} \triangleq \mathbf{S}^{[k,l]} (\mathbf{S}^{[k,l]})^H$  and  $\widehat{\Sigma}^{[k,l]} \triangleq \widehat{\mathbf{S}}^{[k,l]} (\widehat{\mathbf{S}}^{[k,l]})^H$ , the authors show that each column of  $\mathbf{Q}^{[k]}$  contains exactly one non-null entry. Eventually, the theorem that they provide corresponds to Scenario 1 (Section 3.2), generalized to the complex case, but only with  $D = K$ . They mention that additional identifiability conditions can be found in the case of vanishing correlation values.

A second proof was proposed by Anderson et al. [5, 6]. Their identifiability analysis is quite comprehensive, as it admits non-Gaussian, non-stationary and statistically dependent samples, all in the real domain. Their approach is to determine the necessary and sufficient conditions under which the Fisher information matrix (FIM) is positive definite. Their proof is based on a sequence of statements, starting from the claim that non-positive-definiteness of the FIM for each pair of sources  $(i, j)$ , formulated as

$$\mathbf{F}_{i,j} = \text{Cov} \left\{ \begin{bmatrix} \text{diag}\{\Psi_i \Omega_j^\top\} \\ \text{diag}\{\Psi_j \Omega_i^\top\} \end{bmatrix} \right\} \in \mathbb{R}^{K \times K} \quad (2)$$

amounts to non-positive-definiteness of its sample estimate, that is, if and only if (iff)  $\exists(\boldsymbol{\mu}, \boldsymbol{\nu}) \neq (\mathbf{0}, \mathbf{0})$  such that

$$\boldsymbol{\mu}^\top \text{diag}\{\Psi_i \Omega_j^\top\} - \boldsymbol{\nu}^\top \text{diag}\{\Psi_j \Omega_i^\top\} = 0 \quad (3)$$

where  $\Omega_i$  and  $\widehat{\Omega}_i$  are  $K \times T$  matrices whose  $(k, t)$ th element is the source sample  $s_i^{[k]}(t)$  or its estimate, respectively, and  $\Psi_i \triangleq \Psi_i(\widehat{\Omega}_i) = \frac{\partial \log(p_i(\widehat{\Omega}_i))}{\partial \widehat{\Omega}_i} \in \mathbb{R}^{K \times T}$  is the multivariate score function. This analysis eventually leads to a general expression for IVA identifiability [6, Theorem 1], from which, as explained in [5, 6], various special cases can be derived, including all those in Section 3.2, which correspond to our model.

In this paper, we propose a different approach, which is based on finding the necessary and sufficient conditions for non-positive-definiteness of the FIM by *direct factorization of the closed-form expression* of the pairwise FIM. In [15], it is shown that the inverse of the  $2K \times 2K$  sym-

metric matrix<sup>1</sup>

$$\mathcal{H} = \begin{bmatrix} \mathbf{S}_{jj} \circledast \mathbf{S}_{ii}^{-1} & \mathbf{I}_K \\ \mathbf{I}_K & \mathbf{S}_{jj}^{-1} \circledast \mathbf{S}_{ii} \end{bmatrix} \quad (4)$$

is the key ingredient in the Cramér-Rao lower bound (CRLB) on the estimation error of the oblique projection matrices, which extract the sources from their mixtures. Hence, (4) corresponds to the (pairwise) FIM. Therefore, the identifiability of the model consists in characterizing the sufficient and necessary conditions for the invertibility of  $\mathcal{H}$ . Although this idea may be reminiscent of the approach of [5, 6], our analytical derivation is in fact more similar to that in [16, Appendix B]: we factorize the explicit closed-form FIM directly, using algebraic identities.

We mention that an expression similar to (4), but with sample estimates of the covariance matrices, appears in [17, Equation (9)]. However, neither [17] nor [4] mention any explicit link of this matrix to the performance of the model or its identifiability.

## 3. IDENTIFIABILITY: ALTERNATIVE PROOF

### 3.1. Analyzing $\mathcal{H}$

For  $\mathcal{H}$  to be positive-definite, we require that for any non-zero vector  $\mathbf{x} \in \mathbb{R}^{2K \times 1}$ ,  $\mathbf{x}^\top \mathcal{H} \mathbf{x} > 0$ . Conversely, for  $\mathcal{H}$  to be non-positive-definite, there must exist some non-zero  $\mathbf{x} \in \mathbb{R}^{2K \times 1}$  such that

$$0 = \mathbf{x}^\top \mathcal{H} \mathbf{x} = \mathbf{x}^\top \mathbf{V}^\top \mathbf{V} \mathbf{x} \Leftrightarrow \mathbf{V} \mathbf{x} = \mathbf{0} \quad (5)$$

Based on (5), we look for a meaningful decomposition  $\mathcal{H} = \mathbf{V}^\top \mathbf{V}$ . The desired factorization

$$\mathcal{H} = \underbrace{\begin{bmatrix} (\mathbf{S}_{jj}^{\frac{1}{2}\top} \circledast \mathbf{S}_{ii}^{-\frac{1}{2}})^\top \\ (\mathbf{S}_{jj}^{-\frac{1}{2}} \circledast \mathbf{S}_{ii}^{\frac{1}{2}\top})^\top \end{bmatrix}}_{\mathbf{V}^\top} \underbrace{\begin{bmatrix} \mathbf{S}_{jj}^{\frac{1}{2}\top} \circledast \mathbf{S}_{ii}^{-\frac{1}{2}} & \mathbf{S}_{jj}^{-\frac{1}{2}} \circledast \mathbf{S}_{ii}^{\frac{1}{2}\top} \end{bmatrix}}_{\mathbf{V} \in \mathbb{R}^{K \times 2K}}$$

follows from introducing the square root factorization  $\mathbf{S}_{jj} = \mathbf{S}_{jj}^{\frac{1}{2}} \mathbf{S}_{jj}^{\frac{1}{2}\top}$ , and then applying Identity 1 (Appendix 5.1).

Next, we find  $\mathbf{x} \neq \mathbf{0}$  such that  $\mathbf{V} \mathbf{x} = \mathbf{0}$ . Without loss of generality, we look for  $\mathbf{x}$  in the general form  $\mathbf{x} \triangleq [\boldsymbol{\mu}^\top \quad -\boldsymbol{\nu}^\top]^\top$  where  $\boldsymbol{\mu}$  and  $\boldsymbol{\nu}$  are  $K \times 1$  vectors.  $\mathbf{V} \mathbf{x} = \mathbf{0}$  implies that

$$(\mathbf{S}_{jj}^{\frac{1}{2}\top} \circledast \mathbf{S}_{ii}^{-\frac{1}{2}}) \boldsymbol{\mu} - (\mathbf{S}_{jj}^{-\frac{1}{2}} \circledast \mathbf{S}_{ii}^{\frac{1}{2}\top}) \boldsymbol{\nu} = \mathbf{0} \quad (6)$$

for some non-zero  $\boldsymbol{\mu}$  and/or  $\boldsymbol{\nu}$ <sup>2</sup>. We now turn to finding these  $\boldsymbol{\mu}$  and  $\boldsymbol{\nu}$ . Using Identity 2 (Appendix 5.1), equality (6) rewrites as

$$\text{vec}\{\mathbf{S}_{ii}^{-\frac{1}{2}} \text{diag}\{\boldsymbol{\mu}\} \mathbf{S}_{jj}^{\frac{1}{2}}\} = \text{vec}\{\mathbf{S}_{ii}^{\frac{1}{2}\top} \text{diag}\{\boldsymbol{\nu}\} \mathbf{S}_{jj}^{-\frac{1}{2}\top}\}.$$

Removing the “vec” notation and using the invertibility of  $\mathbf{S}_{ii}^{\frac{1}{2}}$  and  $\mathbf{S}_{jj}^{\frac{1}{2}}$ , we obtain

$$\text{diag}\{\boldsymbol{\mu}\} \mathbf{S}_{jj} = \mathbf{S}_{ii} \text{diag}\{\boldsymbol{\nu}\}. \quad (7)$$

<sup>1</sup>Expression (4) is a special case of the more general result in [15], in which the sources within each mixture are random vectors instead of scalar random variables.

<sup>2</sup>Note the difference between (6) and (3)

Hence, non-identifiability exists for non-zero  $(\boldsymbol{\mu}, \boldsymbol{\nu})$  for which equality (7) holds.

The identifiability problem (7) can be further simplified into characterizing all the solutions to

$$\text{diag}\{\boldsymbol{\lambda}\}\mathbf{R} = \mathbf{P} \text{diag}\{\boldsymbol{\lambda}\} \quad (8)$$

where  $\boldsymbol{\lambda}$  is an arbitrary  $K \times 1$  vector (instead of the original problem with a  $2K \times 1$  vector  $\mathbf{x}$ ). The  $K \times K$  symmetric positive definite matrices  $\mathbf{P}$  and  $\mathbf{R}$  are normalized versions of  $\mathbf{S}_{ii}$  and  $\mathbf{S}_{jj}$  such that their diagonals are the identity,  $p^{[k,k]} = 1 = r^{[k,k]} \forall k$ . The normalization scheme is

$$\mathbf{P} = \boldsymbol{\Omega}_{ii} \mathbf{S}_{ii} \boldsymbol{\Omega}_{ii}^\top \quad \text{and} \quad \mathbf{R} = \boldsymbol{\Omega}_{jj} \mathbf{S}_{jj} \boldsymbol{\Omega}_{jj}^\top \quad (9)$$

where  $\boldsymbol{\Omega}_{ii} \triangleq \text{diag}\{(\mathbf{S}_{ii})^{-\frac{1}{2}}\} \forall i$ . The proof of (8) is given in Section 5.2.

In order to extract further information from (8), we multiply (8) either on the left or on the right by  $\text{diag}\{\boldsymbol{\lambda}\}$ ,

$$\text{diag}\{\boldsymbol{\lambda}\}\mathbf{R} \text{diag}\{\boldsymbol{\lambda}\} = \mathbf{P} \text{diag}^2\{\boldsymbol{\lambda}\} \quad (10a)$$

$$\text{diag}\{\boldsymbol{\lambda}\}\mathbf{P} \text{diag}\{\boldsymbol{\lambda}\} = \text{diag}^2\{\boldsymbol{\lambda}\}\mathbf{R} \quad (10b)$$

Since the LHS of (10) is symmetric, so must be the RHS:

$$\text{diag}^2\{\boldsymbol{\lambda}\}\mathbf{P} = \mathbf{P} \text{diag}^2\{\boldsymbol{\lambda}\} \quad (11a)$$

$$\text{diag}^2\{\boldsymbol{\lambda}\}\mathbf{R} = \mathbf{R} \text{diag}^2\{\boldsymbol{\lambda}\} \quad (11b)$$

which can be reformulated as the set of equations  $\forall k, l$

$$(\lambda^{2[k]} - \lambda^{2[l]})p^{[k,l]} = 0 \quad (12a)$$

$$(\lambda^{2[k]} - \lambda^{2[l]})r^{[k,l]} = 0 \quad (12b)$$

where  $\lambda^{2[k]} \triangleq (\lambda^{[k]})^2$ . Note that  $k = l$  nor  $\lambda^{2[k]} = \lambda^{2[l]}$  do not yield any constraints.

### 3.2. Necessary and Sufficient Conditions for Identifiability

It remains to characterize the non-identifiability scenarios associated with  $\boldsymbol{\lambda} \neq \mathbf{0}^3$ . Non-zero  $\boldsymbol{\lambda}$  is associated with two different types of non-identifiability, as we now explain.

**Scenario 1.** The first type is associated with

$$2 \leq D \leq K : \underbrace{\lambda^{[1]} = \dots = \lambda^{[D]}}_{\neq 0} \neq \underbrace{\lambda^{[k>D]}}_{\text{whatever}} \quad (13)$$

i.e., the first  $D \geq 2$  elements of  $\boldsymbol{\lambda}$  are equal, non-zero and different from the rest. Such  $\boldsymbol{\lambda}$  are associated with non-identifiability where there exists a pair  $(i, j)$  of sources whose covariance matrices have the structure

$$\mathbf{S}_{ii} = \begin{bmatrix} \mathbf{S}_{ii}^{[1:D,1:D]} & \mathbf{0}_{D \times (K-D)} \\ \mathbf{0}_{(K-D) \times D} & \mathbf{S}_{ii}^{[D+1:K, D+1:K]} \end{bmatrix} \quad (14a)$$

$$\mathbf{S}_{jj} = \begin{bmatrix} \mathbf{S}_{jj}^{[1:D,1:D]} & \mathbf{0}_{D \times (K-D)} \\ \mathbf{0}_{(K-D) \times D} & \mathbf{S}_j^{[D+1:K, D+1:K]} \end{bmatrix} \quad (14b)$$

and the blocks on the top left obey

$$\mathbf{S}_{ii}^{[1:D,1:D]} = \text{diag}\{\boldsymbol{\omega}\}\mathbf{S}_{jj}^{[1:D,1:D]} \text{diag}\{\boldsymbol{\omega}\} \quad (14c)$$

<sup>3</sup>A similar type of analysis is used, e.g., in [18, Chapter 4.1].

where  $\boldsymbol{\omega}$  is a  $D \times 1$  vector with strictly positive values, and  $\mathbf{S}_{ii}^{[1:D,1:D]}$  is the  $D \times D$  sub-matrix of  $\mathbf{S}_{ii}$  with row and column indices running from 1 to  $D$  (similar to Matlab notation). The proof of Scenario 1 is given in Section 5.3.

**Scenario 2.** The second type is associated with

$$1 \leq D \leq K : \underbrace{\lambda^{[1]} \neq \dots \neq \lambda^{[D]}}_{\text{one may be zero}} \neq \underbrace{\lambda^{[k>D]}}_{\text{whatever}} \quad (15)$$

i.e., the first  $D \geq 1$  elements of  $\boldsymbol{\lambda}$  are different from each other and also from the rest. Such  $\boldsymbol{\lambda}$  are associated with non-identifiability where there exists a pair  $(i, j)$  of sources whose covariance matrices have the structure

$$\mathbf{S}_{ii} = \begin{bmatrix} \text{diag}\{\mathbf{S}_{ii}^{[1:D,1:D]}\} & \mathbf{0}_{D \times (K-D)} \\ \mathbf{0}_{(K-D) \times D} & \mathbf{S}_{ii}^{[D+1:K, D+1:K]} \end{bmatrix} \quad (16a)$$

$$\mathbf{S}_{jj} = \begin{bmatrix} \text{diag}\{\mathbf{S}_{jj}^{[1:D,1:D]}\} & \mathbf{0}_{D \times (K-D)} \\ \mathbf{0}_{(K-D) \times D} & \mathbf{S}_j^{[D+1:K, D+1:K]} \end{bmatrix} \quad (16b)$$

The proof of (16) is given in Section 5.4.

All other types of  $\boldsymbol{\lambda}$  are either not associated with non-identifiability, or subsumed within the previous two scenarios by reordering the source or mixture indices.

One can readily verify that there is one-to-one correspondence between the identifiability conditions in Scenarios 1–2 and those in [4, 5, 6], for our model. We have thus completed our (alternative) proof.

### 3.3. Discussion

We conclude this section with a brief discussion of these identifiability results.

In terms of source separation, Scenario 1 implies that the IVA model is non-identifiable if in  $D \geq 2$  mixtures there exist two sources  $(i, j)$  which (i) are not correlated with any of the other  $(K - D)$  mixtures, as imply the zeros on the off-block-diagonal, and (ii) (14c) holds. Scenario 1 deals with non-identifiability when there is a specific type of similarity between the covariance and cross-covariance profiles of two sources, across mixtures. Therefore, it can be regarded as the multi-set analogue of the ‘‘spatial diversity’’ constraint for BSS, when non-stationarity or colour diversity is concerned; see, e.g., [12].

For  $D = K$ , Scenario 1 corresponds to [4, Theorem 1] and [6, Theorem 3]. For  $D \geq 2$ , this is the case of factorizable probability density function (pdf) in [6, Section VII]. For  $K = 2$ , SOS IVA amounts to generalized eigenvalue decomposition (GEVD) [19]. One can readily verify that the GEVD is indeed not unique when Scenario 1 holds.

In terms of source separation, Scenario 2 implies that the IVA model is non-identifiable if in each of the mixtures indexed  $k = 1, \dots, D$ , there exist two sources  $(i, j)$  that do not have any statistical correlation with any of the other  $(K - 1)$  mixtures, as imply the zeros on the off-diagonal. The non-identifiability of this scenario finds its counterpart in the well-known notion that in classical BSS, one cannot separate more than one Gaussian source (e.g., [20]). Indeed, for  $K = 1$ , this non-identifiable scenario amounts to a single BSS problem with (at least two) Gaussian

sources [6, Theorem 5]. If the structure (16) holds for all pairs of sources  $(i, j)$ ,  $i, j = 1, \dots, N$ , then this scenario amounts to the presence of  $D$  individual BSS mixtures; scenario 2 validates the known fact that this setup is not identifiable using SOS alone.

In terms of data fusion, Scenario 1 implies that if a pair of sources is linked across at least two data sets, then these sources are generically identifiable. Both scenarios indicate that non-identifiability is associated with anomalous behaviour that is shared by a pair of sources, across one or more data sets. This pairwise characteristic is not so surprising, since non-identifiability essentially means that two (or more) sources cannot be distinguished. A possible remedy to Scenario 1 is adding data sets, which may reduce the risk of (14c) to occur. A possible remedy to both non-identifiability scenarios is to exploit additional types of diversity, at least in the mixtures with problematic pairs of sources.

#### 4. CONCLUSION

In this paper, we presented an alternative proof for the identifiability of a special case of IVA. The proposed approach is based on a rigorous linear algebraic formulation. This type of analysis, as well as the new matrix factorization associated with it, may turn out useful in a broader range of problems.

In this work, we discussed only SOS-based IVA when each individual BSS problem was not identifiable. We mention that analogous results, for coupled tensor decompositions, when each tensor decomposition is not unique individually, can be found in [21]. As noted, e.g., by [4, 5, 6], additional types of diversity such as higher-order statistics (HOS), complex values, and non-stationarity, potentially further improve IVA identifiability (with some concrete mathematical proofs therein). Therefore, rigorous analysis of uniqueness and identifiability of various coupled models, not limited to IVA, with different types and combinations of diversity, is an ongoing challenge. For data fusion, there is particular interest in such results, as a large number of questions regarding the combination of various modalities and data sets is still largely open; see, e.g., [14, Section V.E] and references therein.

#### 5. APPENDIX

##### 5.1. Some Algebraic Properties

**Identity 1.** Consider four matrices  $\mathbf{A} \in \mathbb{R}^{\mu \times \xi}$ ,  $\mathbf{B} \in \mathbb{R}^{\nu \times \xi}$ ,  $\mathbf{C} \in \mathbb{R}^{\mu \times \eta}$ ,  $\mathbf{D} \in \mathbb{R}^{\nu \times \eta}$ . Then,

$$\mathbf{A}^\top \mathbf{C} \circledast \mathbf{B}^\top \mathbf{D} = (\mathbf{A} \circledast \mathbf{B})^\top (\mathbf{C} \circledast \mathbf{D}). \quad (17)$$

*Proof.* Can easily be proved by evaluating the  $(i, j)$ th element at each side.  $\square$

The special case  $\mathbf{A} = \mathbf{C}$ ,  $\mathbf{B} = \mathbf{D}$  can be found, e.g., in [22, Chapter 3.2, Equation (12)].

**Identity 2.** For any matrices  $\mathbf{A}$ ,  $\mathbf{B}$  and vector  $\mathbf{x}$  with appropriate dimensions,

$$(\mathbf{B} \circledast \mathbf{A})\mathbf{x} = \text{vec}\{\mathbf{A} \text{diag}\{\mathbf{x}\}\mathbf{B}^\top\} \quad (18)$$

Identity 2 can be found, e.g., in [23, Table III, T3.13].

##### 5.2. Proof of $\text{diag}\{\boldsymbol{\lambda}\}\mathbf{R} = \mathbf{P} \text{diag}\{\boldsymbol{\lambda}\}$

Applying the normalization scheme (9) to (7) and changing sides,

$$\underbrace{\boldsymbol{\Omega}_{ii} \text{diag}\{\boldsymbol{\mu}\}\boldsymbol{\Omega}_{jj}^{-1}}_{\triangleq \text{diag}\{\boldsymbol{\lambda}\}} \mathbf{R} = \mathbf{P} \underbrace{\boldsymbol{\Omega}_{ii}^{-\top} \text{diag}\{\boldsymbol{\nu}\}\boldsymbol{\Omega}_{jj}^\top}_{\triangleq \text{diag}\{\boldsymbol{\varphi}\}} \quad (19)$$

Since  $\mathbf{P}$  and  $\mathbf{R}$  are normalized, the  $k$ th diagonal term of (19) is  $\lambda^{[k]} = \varphi^{[k]} \forall k$ . Hence,  $\text{diag}\{\boldsymbol{\lambda}\} = \text{diag}\{\boldsymbol{\varphi}\}$ . Applying this insight to (19) concludes our proof.

##### 5.3. Proof of Scenario 1

We begin by the special case  $D = K$  which is the simplest. Let  $\boldsymbol{\lambda} = \lambda \mathbf{1}_K$ ,  $\lambda \neq 0$  such that  $\text{diag}\{\boldsymbol{\lambda}\} = \lambda \mathbf{I}_K$ . Then, (8) reduces to  $\lambda \mathbf{I} \mathbf{R} = \mathbf{P} \mathbf{I} \lambda$ , which implies  $\mathbf{R} = \mathbf{P}$ . Next, it follows from (9) that

$$\boldsymbol{\Omega}_{jj}^{-1} \boldsymbol{\Omega}_{ii} \mathbf{S}_{ii} \boldsymbol{\Omega}_{ii} \boldsymbol{\Omega}_{jj}^{-1} = \mathbf{S}_{jj} \quad (20)$$

The latter can be reformulated as

$$\text{diag}\{\boldsymbol{\omega}\} \mathbf{S}_{ii} \text{diag}\{\boldsymbol{\omega}\} = \mathbf{S}_{jj} \quad (21)$$

where  $\text{diag}\{\boldsymbol{\omega}\} \triangleq \boldsymbol{\Omega}_{jj}^{-1} \boldsymbol{\Omega}_{ii}$  such that  $\boldsymbol{\omega}$  is a  $K \times 1$  vector with arbitrary strictly positive values. We now prove the general case. When (13) holds, (12) implies that

$$p^{[k \leq D, l \leq D]} \underbrace{(\lambda^{2[k \leq D]} - \lambda^{2[l \leq D]})}_0 = 0 \quad (22a)$$

$$p^{[k \leq D, l > D]} \underbrace{(\lambda^{2[k \leq D]} - \lambda^{2[l > D]})}_{\neq 0} = 0 \quad (22b)$$

$$p^{[k > D, l \leq D]} \underbrace{(\lambda^{2[k > D]} - \lambda^{2[l \leq D]})}_{\neq 0} = 0 \quad (22c)$$

$$p^{[k > D, l > D]} \underbrace{(\lambda^{2[k > D]} - \lambda^{2[l > D]})}_{\text{depends on } \boldsymbol{\lambda}} = 0 \quad (22d)$$

which induces a structure

$$\mathbf{P} = \left[ \begin{array}{c|c} \mathbf{P}^{[1:D, 1:D]} & \mathbf{0}_{D \times (K-D)} \\ \hline \mathbf{0}_{(K-D) \times D} & \mathbf{P}^{[D+1:K, D+1:K]} \end{array} \right] \quad (23)$$

and similarly for  $\mathbf{R}$ . That is, a set of equal values within  $\boldsymbol{\lambda}$  induces a diagonal sub-block of entries in  $\mathbf{P}$  and  $\mathbf{R}$ , which are uncorrelated with the rest. In a second step, within this  $D \times D$  block, denoting  $\lambda^{[k \leq D]} = \lambda$ , (8) implies that  $\lambda \mathbf{I}_D \mathbf{R}^{[1:D, 1:D]} = \lambda \mathbf{I}_D \mathbf{P}^{[1:D, 1:D]}$ , which implies  $\mathbf{R}^{[1:D, 1:D]} = \mathbf{P}^{[1:D, 1:D]}$ . Undoing the normalization of (23) via (9), similarly to (20), we obtain the desired result.

##### 5.4. Proof of Scenario 2

When (15) holds, using a procedure similar to (22), equations (12) induce a structure

$$\mathbf{P} = \left[ \begin{array}{c|c} \mathbf{I}_D & \mathbf{0}_{D \times (K-D)} \\ \hline \mathbf{0}_{(K-D) \times D} & \mathbf{P}^{[D+1:K, D+1:K]} \end{array} \right] \quad (24)$$

i.e., the top left block is the identity, the top right and bottom left blocks are zero. Only the bottom right block is unaffected; and similarly for  $\mathbf{R}$ . Undoing the normalization of (24) via (9) yields (16), which concludes our proof.

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