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# **A Multi-State Physics Modeling Approach for the Reliability Assessment of Nuclear Power Plants Piping Systems**

Francesco Di Maio<sup>1</sup>, Davide Colli<sup>1</sup>, Enrico Zio<sup>1,2</sup>, Liu Tao<sup>3</sup>, Jiejuan Tong<sup>3</sup>

<sup>1</sup> *Energy Department, Politecnico di Milano*

*Via Ponzio 34/3, 20133 Milano, Italy*

[francesco.dimaio@polimi.it](mailto:francesco.dimaio@polimi.it)

<sup>2</sup> *Chair on System Science and Energetic Challenge*

*European Foundation for New Energy – Electricite de France*

*Ecole Centrale, Paris, and Supelec, Paris, France*

<sup>3</sup> *Institute of Nuclear and New Energy Technology, Beijing, China, 100084*

## **ABSTRACT**

A Multi-State Physics Modeling (MSPM) approach is here proposed for degradation modeling and failure probability quantification of Nuclear Power Plants (NPPs) piping systems. This approach integrates multi-state modeling to describe the degradation process by transitions among discrete states (e.g., no damage, micro-crack, flaw, rupture, etc.), with physics modeling by (physic) equations to describe the continuous degradation process within the states. We propose a Monte Carlo (MC) simulation method for the evaluation of the time-dependent transition rates between the states of the MSPM. Accountancy is given for the uncertainty in the parameters and external factors influencing the degradation process. The proposed modeling approach is applied to a benchmark problem of a piping system of a Pressurized Water Reactor (PWR) undergoing thermal fatigue. The results are compared with those obtained by a continuous-time homogeneous Markov Chain Model.

Keywords: Piping reliability; Multi-State Physics Modeling; Degradation process; Markov Chain Model, Monte Carlo; time-dependant transition rates; Nuclear Power Plant.

## **1 INTRODUCTION**

Piping systems are one of the most risk-sensitive structural elements of Nuclear Power Plants (NPPs). Then, the analysis of these systems for the quantification of their failure probability is of utmost importance [Gopika et. al., 2003]. Different approaches have been applied for the estimation of piping systems failure probabilities. The most straightforward approach is to obtain estimates of piping component failure rates by statistical estimation based on field data collected from piping

systems service experience, e.g. by Maximum Likelihood Estimation (MLE) approaches [Gosselin et. al., 1997]. Alternatively, Probabilistic Fracture Mechanics (PFM) models can be used to predict crack initiation and growth from existing flaws [Verma et. al., 2011]. To apply PFM models, the quantity and quality of information needed to perform the necessary computations are not negligible and the model evaluation can be time consuming, and even in cases where PFM approaches are appropriate, it is highly desirable to be able to benchmark and validate the results with some field data [Simola et. al., 2002]. In fact, for example, one limitation of this approach is that historical data reflect the influence of previous piping systems inspection programs, and, if changes to these programs are proposed, such changes may render the previous failure rate estimates no longer relevant [Fleming, 2004]. To overcome these problems, a Markov Chain Model (MCM) has been proposed in [Fleming et. al., 2008], in which the interaction between damage mechanisms and inspection, detection and repair strategies are explicitly defined and treated. However, transition rates between the degradation states and the holding times in any of the states of the MCM are assumed to be constant and to follow exponential distributions, respectively. This assumption is no longer acceptable when compared with field data of service experience (or when considering a new design of components with different geometry, material properties, degradation mechanisms and thermal-hydraulic behaviors) [Chatterjee et. al., 2008].

This paper presents the development and application of a novel framework for evaluating piping system reliability, specifically considering the situation of a new design of, or an already employed, piping system for which field data are not available or difficult to collect. The underlying model is non-Markovian because the transition rates are time-dependent, and include the uncertainties due to lack of knowledge of physical phenomena and parameters related to and influencing the degradation process [Li et. al., 2012; Unwin et. al., 2011]. The approach undertaken is based on a Multi-State Physics Model (MSPM) and Monte Carlo (MC) simulation, which i) accounts for the uncertain parameters and external factors influencing the process of transitions between degradation states and ii) relaxes the constraints of time-independent transition rates of continuous-time homogeneous MCM. The MSPM is, conceptually, a MCM in which the degradation processes (and thus, the transition rates) are described by physics model equations. In this work, the evaluation of the time-dependent transition rates is based on the outcomes of the Monte Carlo (MC) simulations of the degradation processes, directly described by physical models; this allows determining the distributions of the holding times of the components in the different states, from which the transition rates are derived. The knowledge added by the physics modeling and the inclusion of the effects of the uncertain parameters and the external factors which influence the degradation process allow a more realistic description of the piping system behavior and a more accurate estimation of

its reliability (which, in the numerical case study considered avoids underestimation with respect to MCM, and the risk of exposing the system to catastrophic consequences in case of piping rupture). The paper organization is as follows. Section 2 presents the general definition of the Multi-State Physics Model (MSPM) approach, also in consideration of the MCM approach, and the estimation procedure for the transition rates. Section 3 presents the MSPM application to a Pressurized Water Reactor (PWR) piping system. Section 4 contains the conclusions of the work.

## 2 Multi-State modeling for piping systems degradation

### 2.1 The continuous-time homogeneous Markov Chain Model (MCM)

Under the framework of Multi-State modeling, when the dynamics of component degradation is described by a continuous-time homogeneous Markov Chain Model (MCM), the transitions among a finite number  $M$  of discrete states  $\bar{T}$ ,  $\bar{T} = \{T_0, T_1, \dots, T_M\}$ , are modeled by constant rates  $\lambda_{i,j}$  of transition from state  $i$  to state  $j$ , from which the state probability vector  $\bar{P}(t)$  can be derived at any time instant  $t$ ,  $\bar{P}(t) = \{p_0(t), p_1(t), \dots, p_M(t)\}$  [Li et. al., 2012].

A general MCM to describe the piping systems degradation mechanisms is illustrated in Figure 1, where  $\bar{T} = \{S, F, L, R\}$  are the binary states healthy  $S$  (i.e., no detectable damage), degraded  $F$  and  $L$  (i.e., detectable flaw, detectable leak) and rupture  $R$ , respectively. The transition rates between states  $\bar{T}$  are denoted as  $\lambda_{S,F}$ ,  $\lambda_{S,L}$ ,  $\lambda_{S,R}$ ,  $\lambda_{F,L}$ ,  $\lambda_{F,R}$ ,  $\lambda_{L,R}$ ,  $\mu$  and  $\omega$ . Transitions among states  $\bar{T}$  can occur due to damage mechanisms at the pipe base metal (e.g., flow accelerated corrosion), on welds or in the heat-affected zone near welds (e.g. thermal fatigue), wall thinning, crack propagation, severe loading (e.g., water hammer and overpressure), and their various combinations [Fleming, 2004; Bush et. al., 1996].

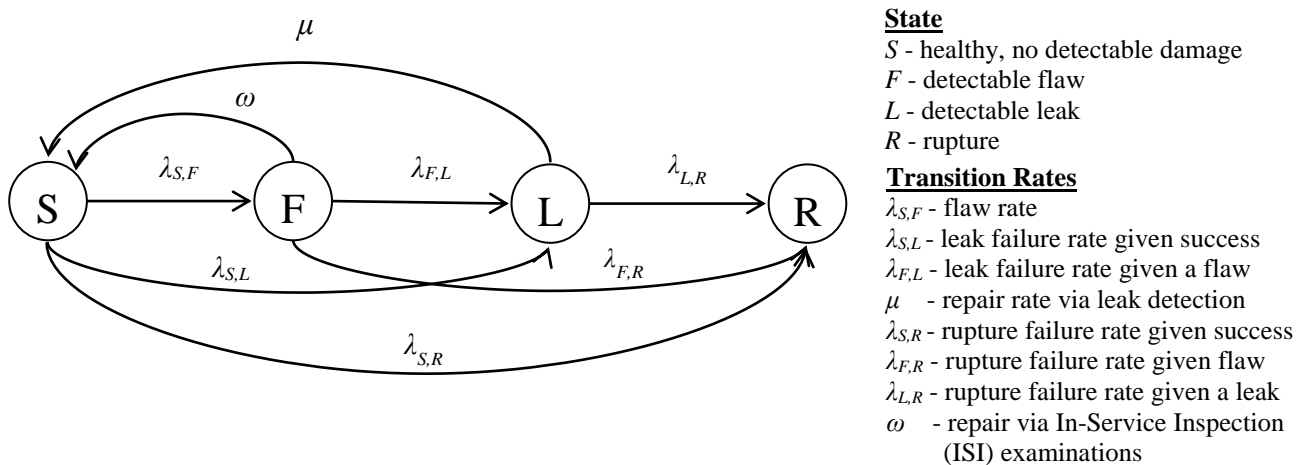


Fig. 1. Four-state MCM for degradation mechanisms in piping systems [Fleming, 2004]

Under the assumption that all the transition rates are constant, the MCM equations consist of a set of linear differential equations with constant coefficients and the state probability vector  $\bar{P}(t) = \{p_S(t), p_F(t), p_L(t), p_R(t)\}$  at any time  $t$  is determined by solving Eq. (1) (analytically or numerically) [Fleming, 2004]:

$$\begin{cases} \frac{dp_S(t)}{dt} = -(\lambda_{S,L} + \lambda_{S,R} + \lambda_{S,F})p_S(t) + \omega p_F(t) + \mu p_L(t) \\ \frac{dp_F(t)}{dt} = \lambda_{S,F}p_S(t) - (\lambda_{F,L} + \lambda_{F,R} + \omega)p_F(t) \\ \frac{dp_L(t)}{dt} = \lambda_{S,L}p_S(t) + \lambda_{F,L}p_F(t) - (\lambda_{L,R} + \mu)p_L(t) \\ \frac{dp_R(t)}{dt} = \lambda_{S,R}p_S(t) + \lambda_{F,R}p_F(t) + \lambda_{L,R}p_L(t) \end{cases} \quad (1)$$

## 2.2 The Multi-State Physics Model (MSPM)

In this work, the MSPM transition rates,  $\lambda_{i,j}(\tau_{i,j}, \bar{\delta})$ , are assumed to be functions of the influencing factors  $\bar{\delta}$  (i.e., the physical parameters used to model the degradation transition phenomena) and of  $\tau_{i,j}$  (i.e., the holding time of the system in state  $i$ , provided that the arrival state will be  $j$ ). On this premise, the objective of the MSPM framework is to solve for the state probability vector  $\bar{P}(t) = \{p_S(t, \bar{\delta}), p_F(t, \bar{\delta}), p_L(t, \bar{\delta}), p_R(t, \bar{\delta})\}$  where, differently from the homogeneous MCM of Section 2.1, the transitions among states are described by  $\tau$ - and  $\bar{\delta}$ -dependent transition rates. To obtain the state probability vector  $\bar{P}(t, \bar{\delta})$  at each time  $t$ , the differential equations of Eq. (2) below need to be jointly solved:

$$\begin{cases} \frac{dp_S(t, \bar{\delta})}{dt} = -(\lambda_{S,L}(\tau_{S,L}, \bar{\delta}) + \lambda_{S,R}(\tau_{S,R}, \bar{\delta}) + \lambda_{S,F}(\tau_{S,F}, \bar{\delta}))p_S(t, \bar{\delta}) + \omega p_F(t, \bar{\delta}) + \mu p_L(t, \bar{\delta}) \\ \frac{dp_F(t, \bar{\delta})}{dt} = \lambda_{S,F}(\tau_{S,F}, \bar{\delta})p_S(t, \bar{\delta}) - (\lambda_{F,L}(\tau_{F,L}, \bar{\delta}) + \lambda_{F,R}(\tau_{F,R}, \bar{\delta}) + \omega)p_F(t, \bar{\delta}) \\ \frac{dp_L(t, \bar{\delta})}{dt} = \lambda_{S,L}(\tau_{S,L}, \bar{\delta})p_S(t, \bar{\delta}) + \lambda_{F,L}(\tau_{F,L}, \bar{\delta})p_F(t, \bar{\delta}) - (\lambda_{L,R}(\tau_{L,R}, \bar{\delta}) + \mu)p_L(t, \bar{\delta}) \\ \frac{dp_R(t, \bar{\delta})}{dt} = \lambda_{S,R}(\tau_{S,R}, \bar{\delta})p_S(t, \bar{\delta}) + \lambda_{F,R}(\tau_{F,R}, \bar{\delta})p_F(t, \bar{\delta}) + \lambda_{L,R}(\tau_{L,R}, \bar{\delta})p_L(t, \bar{\delta}) \end{cases} \quad (2)$$

Notice that the four considered states  $\bar{T}$  are mutually exclusive and form a complete set: thus,  $p_S(t, \bar{\delta}) + p_F(t, \bar{\delta}) + p_L(t, \bar{\delta}) + p_R(t, \bar{\delta}) = 1$  at any time  $t = 1, 2, \dots, T_{miss}$ , where  $T_{miss}$  is the mission time of the piping system. The calculation of the analytical solution  $\bar{P}(t, \bar{\delta})$  of Eq. (2) with  $\tau$ - and  $\bar{\delta}$ -dependent transition rates is a difficult (or even impossible in certain cases) task [Li et. al., 2012; Lisnianski et. al., 2008]. To overcome this problem, a MC simulation framework is here proposed.

### 2.2.1 Monte Carlo estimation of $\tau$ - and $\bar{\delta}$ - dependent transition rates

The differential equations in Eq. (2) can be written in a general form as

$$\frac{d}{dt} p_i(t|\bar{\delta}) = \sum_{k \neq i}^M p_k(t|\bar{\delta}) q_{k,i}(t|\bar{\delta}) \lambda_k(t, \bar{\delta}) - p_i(t|\bar{\delta}) \lambda_i(t, \bar{\delta}) \quad (3)$$

where  $i=1, 2, \dots, M$ ,  $\lambda_i(t, \bar{\delta})$  is the total transition rate of departure from state  $i$  and

$$q_{k,i}(t|\bar{\delta}) = \frac{\lambda_{i,j}(t, \bar{\delta})}{\lambda_i(t, \bar{\delta})} \quad (4)$$

is the conditional probability that, given the transition out of any other state  $k$ , the arrival state will be  $i$ . The quantification of  $\bar{P}(t, \bar{\delta})$  calls for the solution of Eq. (3), that can be obtained introducing the integrating factor  $M_i(t, \bar{\delta}) = \exp \left[ \int_0^t \lambda_i(t', \bar{\delta}) dt' \right]$  [Li et. al., 2012]:

$$p_i(t|\bar{\delta}) = p_i(0) \cdot \exp \left[ - \int_0^t \lambda_i(t', \bar{\delta}) dt' \right] + \int_0^t \exp \left[ - \int_t^{t'} \lambda_i(t'', \bar{\delta}) dt'' \right] \sum_{k \neq i}^M p_k(t'|\bar{\delta}) q_{k,i}(t'|\bar{\delta}) \lambda_k(t', \bar{\delta}) dt' \quad (5)$$

Realistically,  $\lambda_i(t|\bar{\delta})$ ,  $i=1, \dots, M$ , is unknown and, thus, cannot be used directly for calculating  $p_i(t|\bar{\delta})$ . Instead, based on the physical knowledge of the degradation mechanisms that determine the transitions among the states, the holding times  $\tau_i$  can be estimated (by simulating the degradation mechanisms  $N_c$  times) and, then, the transition rate from state  $i$  to another state  $j$  can be indirectly determined. In this setting, the transition rates can be expressed as functions of  $\tau_{i,j}$  as:

$$\lambda_{i,j}(\tau_{i,j}, \bar{\delta}) = \frac{f(\tau_{i,j}|\bar{\delta})}{R(\tau_{i,j}|\bar{\delta})} \cong \lim_{\Delta\tau \rightarrow 0} \frac{F(\tau_{i,j} + \Delta\tau|\bar{\delta}) - F(\tau_{i,j}|\bar{\delta})}{(1 - F(\tau_{i,j}|\bar{\delta})) \times (\Delta\tau)} \quad (6)$$

where  $\tau_{i,j}$  is the holding time in state  $i$ , provided that the arrival state will be  $j$ ,  $R(\tau_{i,j}|\bar{\delta})$  is the reliability of the component at time  $\tau_{i,j}$ ,  $f(\tau_{i,j}|\bar{\delta})$  and  $F(\tau_{i,j}|\bar{\delta})$  are the probability density function and cumulative distribution function of the holding time between states  $i$  and  $j$ , respectively.

The total transition rate  $\lambda_i$  of leaving state  $i$  towards any arrival state  $j$ ,  $j=1, \dots, M$ ,  $j \neq i$  at time  $\tau_i$  is, therefore, equal to:

$$\lambda_i(\tau_i, \bar{\delta}) = \sum_{\substack{j=0 \\ j \neq i}}^M \lambda_{i,j}(\tau_i, \bar{\delta}) \quad (7)$$

The total transition rate  $\lambda_i(\tau_i, \bar{\delta})$ , can, thus, be expressed as the convolution of all the transition rates that have determined any possible transition before time  $t$ , each with its holding time  $\tau$ :

$$\lambda_i(t, \bar{\delta}) = \begin{cases} \int_0^t \lambda_1(t_1 = \tau_1, \bar{\delta}) dt_1 & \text{for } i = 1 \\ \int_0^t \lambda_1(t_2 = \tau_1, \bar{\delta}) \lambda_2(t - t_2 = \tau_2, \bar{\delta}) dt_2 & \text{for } i = 2 \\ \int_0^t \int_0^{t_1} \dots \int_0^{t_{n-1}} \lambda_1(t_2 = \tau_1, \bar{\delta}) \lambda_2(t_3 - t_2 = \tau_2, \bar{\delta}) \dots \lambda_{n-1}(t_n - t_{n-1} = \tau_{n-1}, \bar{\delta}) \lambda_n(t - t_n = \tau_n, \bar{\delta}) dt_2 dt_3 \dots dt_n & \text{for } i = n \end{cases} \quad (8)$$

The lifetime  $t$  at which the system will be in state  $j^*$  is, therefore, equal to  $t = \sum_{k=0}^i \tau_k$ .

In the end, the procedure for calculating  $\lambda_i(t, \bar{\delta})$  reduces to solving Eq. (6) that entails evaluating the cumulative distribution function  $F(\tau_{i,j}|\bar{\delta})$  and the transition rates  $\lambda_{i,j}(\tau_{i,j}, \bar{\delta})$  as follows:

- 1) *Build the physical models that describe the degradation process (e.g., fatigue, thermal fatigue and stress corrosion cracking (SCC)).*
- 2) *Select a characteristic variable  $x$  (e.g. crack depth, crack length, etc.), that represents the degradation process state and its threshold value  $X_{cr}$ , that defines the transition from one state to another: the time  $\tau_{i,j}$  at which the system moves from state  $i$  to state  $j$  is that at which  $x = X_{cr}$ .*
- 3) *Sample the values of the parameters  $\bar{\delta}$  of the physical models, treated as random variables whose values follow given distributions representing their uncertainties.*
- 4) *Simulate the degradation process  $N_c$  times for estimating the state holding time  $\tau_{i,j}$  distributions: the algorithm for the estimation of the probability density function  $f(\tau_{i,j}|\bar{\delta})$  and of the cumulative density function  $F(\tau_{i,j}|\bar{\delta})$  is sketched in the following pseudo-code, where  $N_{succ}$  is the number of MC simulations in which  $x \geq X_{cr}$ , at time  $\tau_{i,j}$  and  $N_c$  is the total number of trials. The time space is discretized by choosing a discrete timeline with  $\Delta\tau$  as interval size.*

Set the threshold dimension  $X_{cr}$ , the number of MC repeated trials  $N_c$ , the interval time size  $\Delta\tau$  and the mission time  $T_{miss}$ . Define  $N_{succ}$  as a vector of  $T_{miss}/\Delta\tau$  elements, each one representing a discrete step on the timeline equal to  $\Delta\tau$ .

Consider a physics equation  $x = g(\tau, \bar{\delta})$  that models  $x$  as a function of  $\tau$  and  $\bar{\delta}$

**For**  $N = 1:N_c$

$\tau = 0$

Sample physics parameters  $\bar{\delta}$  from their distributions

$x = g(\tau, \bar{\delta})$

**While**  $x \leq X_{cr}$

$\tau = \tau + \Delta\tau$

$x = g(\tau, \bar{\delta})$

**End While**

$N_{succ}(\tau/\Delta\tau + 1) = N_{succ}(\tau/\Delta\tau + 1) + 1$

**End For**

$f(\tau|\bar{\delta}) = N_{succ}/N_c$

$F(1) = f(1)$

**For**  $N = 2:T_{miss}/\Delta\tau$

$F(N) = F(N - 1) + f(N)$

**End For**

5) Estimate the transition rates by applying Eq. (6) with the selected  $\Delta\tau$ .

### 2.2.2 Estimation of the State Probability Vector $\bar{P}(t, \bar{\delta})$

Once the transition rates  $\lambda_{i,j}(\tau_{i,j}, \bar{\delta})$  are estimated using Eq. (6), the state probability vector  $\bar{P}(t, \bar{\delta})$  can be obtained by performing  $N_{max}$  direct MC simulations of the random walks through the states of the MCM. The time  $\tau_{i,j}$  is sampled from  $F(\tau_{i,j}|\bar{\delta})$  for each  $j = 1, \dots, M$ ; the holding time  $\tau_i$  is  $\tau_i = \min(\tau_{i,j})$ , whereas the arrival state  $j$  is the one that corresponds to  $\arg[\min(\tau_{i,j})]$ : the system enters state  $j$  at the respective time  $\tau_{i,j}$ . The state probability vector  $\bar{P}(t, \bar{\delta})$  is, then, estimated by counting the number of visits  $n(t)$  at time  $t$  to each state  $i = 1, 2, 3, 4$  ( $S, F, L, R$ , respectively) and dividing by the total number of random walk simulations  $N_{max}$  performed. The algorithm for the simulation of the process of component degradation on the time horizon  $[0, T_{miss}]$  is sketched in the following pseudo-code.

Initialize the system at time  $t_0 = 0$  and healthy state  $i = 1(S)$ . Set the total number of simulations  $N_{max}$ , the mission time  $T_{miss}$  and the state visit counters  $n_i(t) = 0, i = 1, 2, 3, 4$ . Let  $\tau_i$  be the time after which the system leaves state  $i$ .

**While**  $N < N_{max}$

$i = 1$

$t = t_0$

**While**  $t \leq T_{miss}$

Sample the state holding times  $\tau_{i,j}, j = 1, \dots, M$ , for  $j \neq i$

The system moves to state  $j$ , at which corresponds the minimum  $\tau_{i,j}$

$\tau_i = \min(\tau_{i,j})$

$n_i(t: \tau_i) = n_i(t: \tau_i) + 1$

$t = t + \tau_i$

**End While**

$N = N + 1$

**End While**

$$\bar{P}(t, \bar{\delta}) = \frac{1}{N_{max}} [n_1(t), n_2(t), n_3(t), n_4(t), ]$$

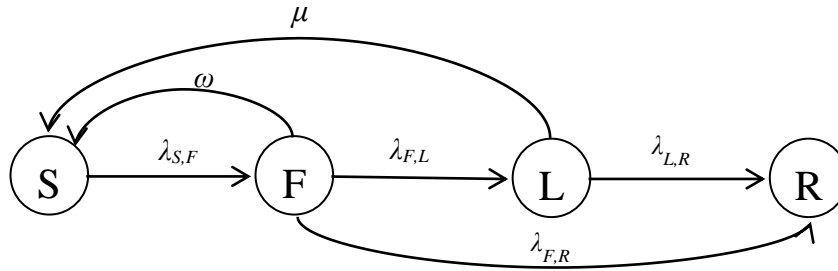


### 3 Application to a PWR piping system

#### 3.1 System description

The modelling and simulation framework proposed is applied for the evaluation of the probability of rupture due to thermal fatigue of a mixing tee between the hot and cold legs during a Loss of Coolant Accident (LOCA) in the Reactor Cooling System (RCS) of a Pressurized Water Reactor (PWR) [Fleming, 2004; Radu et. al., 2007 a]. The MCM model that describes this degradation process is sketched in Figure 2. The simplification of the model with respect to Figure 1 is based on the assumption that for fatigue damage mechanisms:

- Crack is initiated when the component shows a detectable Flaw ( $F$ ).
- Crack propagates leading to a Leak ( $L$ ) in case of circumferential crack that propagates to a through-wall circumferential crack and, then, to Rupture ( $R$ ), or directly to Rupture ( $R$ ) in case of fully-circumferential crack.
- The considered piping system is not subject to severe loading conditions: transitions between no damage state ( $S$ ) to Rupture ( $R$ ), or Leak ( $L$ ) are not considered realistic.



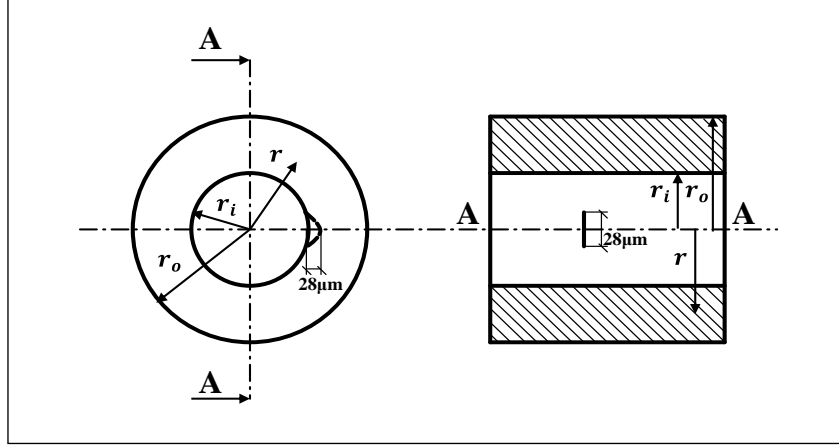
**Fig. 2 MCM for a PWR component subjected to thermal fatigue [Fleming, 2004]**

The operating conditions of the downstream mixing tee are: pressure of 36 bar, hot leg water temperature at 180°C and cold leg water temperature at 20°C, pipe inner radius in the damage zone equal to  $r_i = 120$  mm and outer radius to  $r_o = 129$  mm [Radu et. al., 2007 b]. The piping material is austenitic stainless steel 304L and the maximum temperature fluctuation ( $\Delta\theta_{max}$ ), due to turbulent mixing or vortices on the inner surface of the pipe is estimated to be 120°C [Radu et. al., 2007 b].

In what follows, we perform the reliability assessment of the piping system considering  $\tau$ - and  $\bar{\delta}$ -dependent transition rates on a mission time  $T_{miss} = 100$  years.  $T_{miss}$  has been chosen reasonably longer than a typical NPP lifetime of 40 years, to give account to possible life extension. We resort to the MSPM approach introduced in Section 2.2; the results will be compared with the MCM solutions obtained in [Fleming, 2004] for a typical PWR RCS piping system.

### 3.2 Monte Carlo estimation of $\tau$ - and $\bar{\delta}$ - dependent transition rates

#### 3.2.1 Transition rate $\lambda_{S,F}(\tau_{S,F}, \bar{\delta})$



**Fig. 3 Circumferential crack onset**

The crack onset (i.e., the transition between states Safe ( $S$ ) and Flaw ( $F$ )) on the tee piping component connecting the hot and cold legs of a RCS of a PWR, due to thermal fatigue, can be represented as in Figure 3 and modeled referring to the total equivalent strain rate  $\dot{\epsilon} = \dot{\epsilon}_{eq}^{tot}$ , that can be calculated as:

$$\dot{\epsilon}_{eq}^{tot} = K_v \cdot \dot{\epsilon}_{eq}^{elastic} \quad (10)$$

where  $K_v$  and  $\dot{\epsilon}_{eq}^{elastic}$  are functions of the effective equivalent stress intensity range  $\Delta\sigma_{eq}$ . Thus, [Anacllet et. al., 2007]:

$$\dot{\epsilon}_{eq}^{tot} = K_v(\Delta\sigma_{eq}) \cdot \frac{2 \cdot (1+\nu) \Delta\sigma_{eq}}{3 E} \quad (11)$$

where  $\nu$  is the Poisson coefficient,  $K_v(\Delta\sigma_{eq})$  can be found by empirical correlation (plotted in Figure 4), whereas  $\Delta\sigma_{eq}$  is equal to [Radu et. al., 2007 b]:

$$\Delta\sigma_{eq} = \sqrt{\frac{(\Delta\sigma_r - \Delta\sigma_\vartheta)^2 + (\Delta\sigma_z - \Delta\sigma_\vartheta)^2 + (\Delta\sigma_r - \Delta\sigma_z)^2}{2}} \quad (12)$$

where  $\Delta\sigma_i$ ,  $i = r, z, \vartheta$ , are the maximum stress intensities ranges due to radial ( $\sigma_r$ ), axial ( $\sigma_z$ ) and hoop ( $\sigma_\vartheta$ ) thermal stresses, respectively. To evaluate the thermal stresses  $\sigma_r$ ,  $\sigma_z$  and  $\sigma_\vartheta$ , because of the simple geometry the pipe can be represented as a hollow cylinder and it is possible to use the analytical solutions proposed in [Radu et. al., 2007 a], in which the piping system is assumed to be subjected to sinusoidal transients of thermal loads  $\theta(\omega, \theta_0)$  that well approximate the cold and hot leg mixing phenomenon occurring at the tee piping component.

Therefore, the radial, hoop and axial stresses are [Radu et. al., 2007 a]:

$$\sigma_r(r, \omega, \theta_0, k) = \frac{\alpha \cdot E}{1-\nu} \cdot \left[ -\frac{1}{r^2} \cdot I_1(r, \omega, \theta_0, k) + \frac{r^2 - r_i^2}{r^2 \cdot (r_o^2 - r_i^2)} \cdot I_2(r, \omega, \theta_0, k) \right] \quad (13)$$

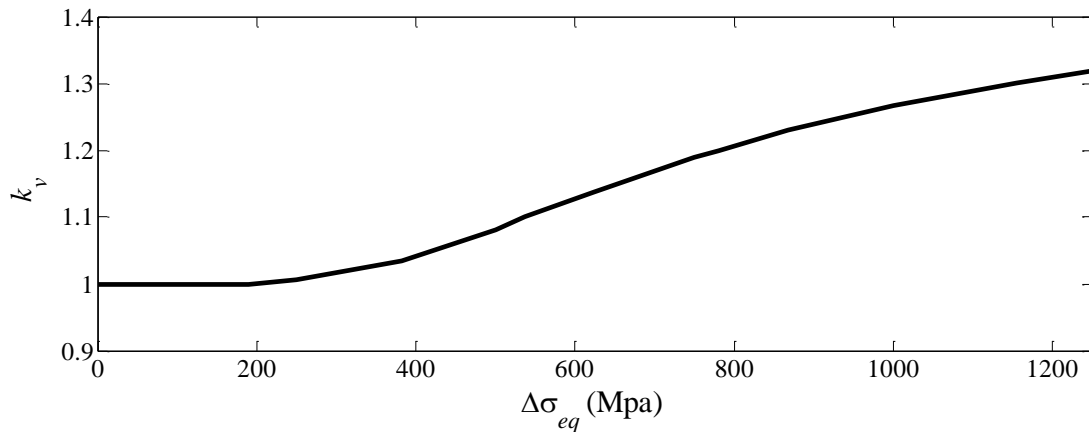
$$\sigma_\theta(r, \omega, \theta_0, k) = \frac{\alpha \cdot E}{1-\nu} \cdot \left[ \frac{1}{r^2} \cdot I_1(r, \omega, \theta_0, k) + \frac{r^2 - r_i^2}{r^2 \cdot (r_o^2 - r_i^2)} \cdot I_2(r, \omega, \theta_0, k) - \theta(r, \omega, \theta_0, k) \right] \quad (14)$$

$$\sigma_z(r, \omega, \theta_0, k) = \frac{\alpha \cdot E}{1-\nu} \cdot \left( \frac{2\nu}{r_o^2 - r_i^2} \cdot I_2(r, \omega, \theta_0, k) - \theta(r, \omega, \theta_0, k) \right) \quad (15)$$

where  $r$  is the radial distance from the center of the pipe,  $\omega = 2\pi f$  and  $\theta_0$  are the frequency and amplitude of the temperature sinusoidal wave, respectively,  $k$  is the thermal diffusivity coefficient,  $\alpha$  is the thermal expansion coefficient,  $E$  is the modulus of elasticity. The mathematical relationships for  $I_1(r, \omega, \theta_0)$ ,  $I_2(r, \omega, \theta_0)$  and the temperature distribution across the wall thickness  $\theta(r, \omega, \theta_0)$  are given in Appendix 1, with the theoretical details. Table 1 lists the parameters  $\bar{\delta} = (f, \theta_0, E, \alpha, \nu, k)$ , together with their distributions, affecting the model behaviour and, thus, the crack onset.

Parameter	Description	Unit	Type of distribution	Lower value	Upper value
$f$	Frequency of temperature wave	Cycle/sec	uniform	$10^{-2}$	100
$\theta_0$	Amplitude of temperature wave	°C	uniform	0	60
Parameter	Description	Unit	Value		
$E$	Modulus of elasticity	N/m <sup>2</sup>	$177 \times 10^9$		
$\alpha$	Thermal expansion coefficient	1/°C	$16.4 \times 10^{-6}$		
$\nu$	Poisson coefficient	/	0.3		
$k$	Thermal diffusivity coefficient	m <sup>2</sup> /s	$3.93 \times 10^{-6}$		

**Table 1. Parameters values and uncertainties**

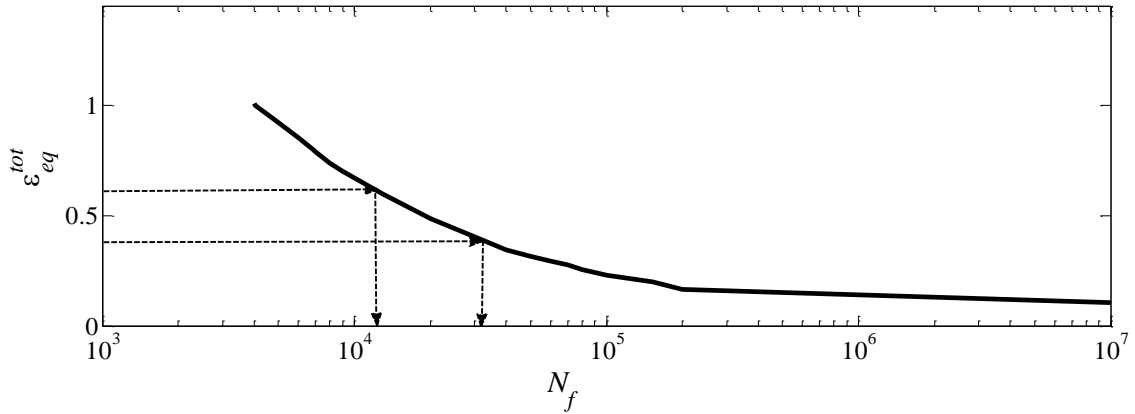


**Fig. 4 Evolution of  $k_v(\Delta\sigma_{eq})$**

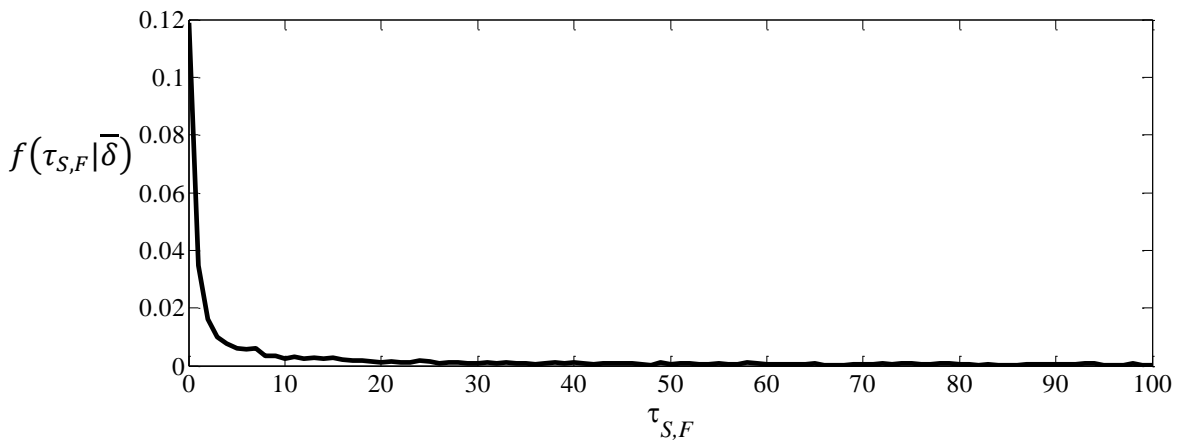
For a given value of  $x = \varepsilon_{eq}^{tot}$ , the piping system experiences a circumferential crack onset (with initial crack depth and length equal to  $x_1^{in} = x_2^{in} = 28\mu m$  [Varfolomeyev, 2006; Sester et. al., 2000]), after a number of thermal cycles  $N_f$  that is given by the empirical correlation shown in Figure 5, relating  $x$  to  $N_f$  [Anaclet et. al., 2007]. The state holding time  $\tau_{S,F}$  for a specific set of values of  $\bar{\delta}$  affecting the duration of the crack onset is, thus, calculated from  $N_f$  as:

$$\tau_{S,F}(years) = \frac{N_f}{f \cdot 3600 \cdot 24 \cdot 365} \quad (16)$$

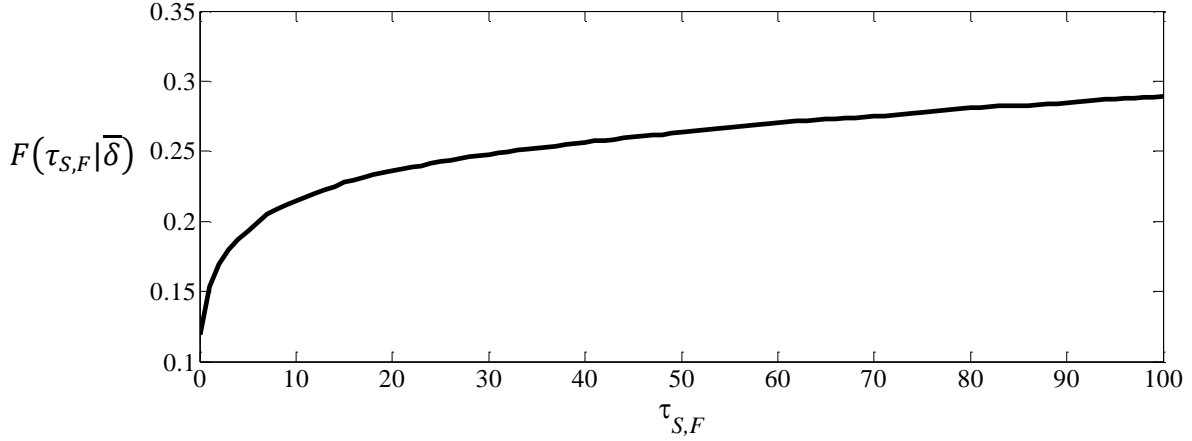
where we consider 365 days per year, 24 hours per day and 3600 seconds per hour. The probability distribution of  $\tau_{S,F}$  is estimated by applying the MC simulation proposed at item (4) of sub-Section 2.2.1, with a number of simulations  $N_c = 10000$ . For each one of these  $N_c$  trials,  $\varepsilon_{eq}^{tot}$  is evaluated applying Eqs. (10)-(15) given a sampled batch of values of the parameters  $\bar{\delta}$  (Table 1):  $N_c$  values of  $N_f$  are collected and the probability density function  $f(\tau_{S,F}|\bar{\delta})$  and cumulative distribution function  $F(\tau_{S,F}|\bar{\delta})$  can be built (shown in Figures 6 and 7, respectively). By applying Eq. (6) with a time step of one year ( $\Delta\tau = 1$ ), the transition rate  $\lambda_{S,F}(\tau_{S,F}, \bar{\delta})$  is computed (shown in Figure 8).



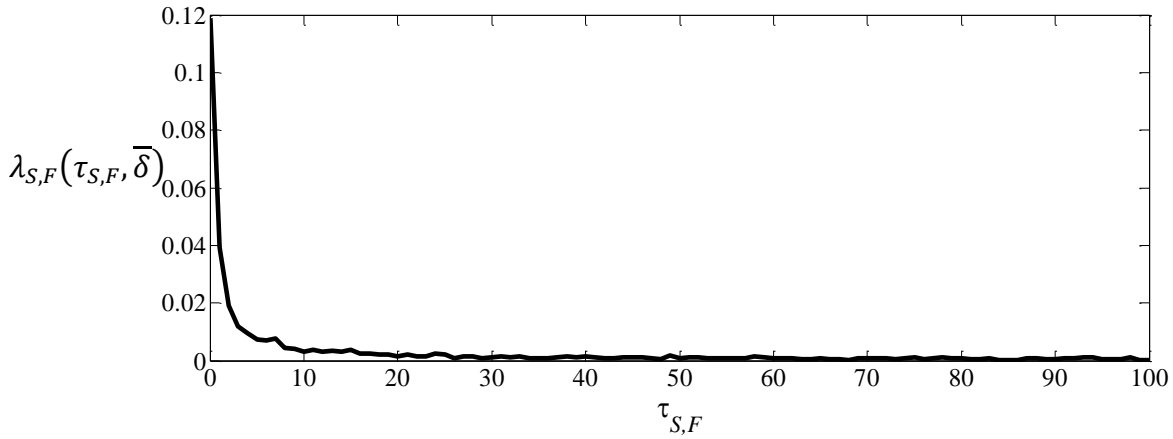
**Fig. 5 Total equivalent strain rate  $\varepsilon_{eq}^{tot}$  vs number of cycles to failure  $N_f$  [Anaclet et. al., 2007]**



**Fig. 6 Probability density function of the state holding time  $\tau_{S,F}$**



**Fig.7 Cumulative distribution function of the state holding time  $\tau_{S,F}$**

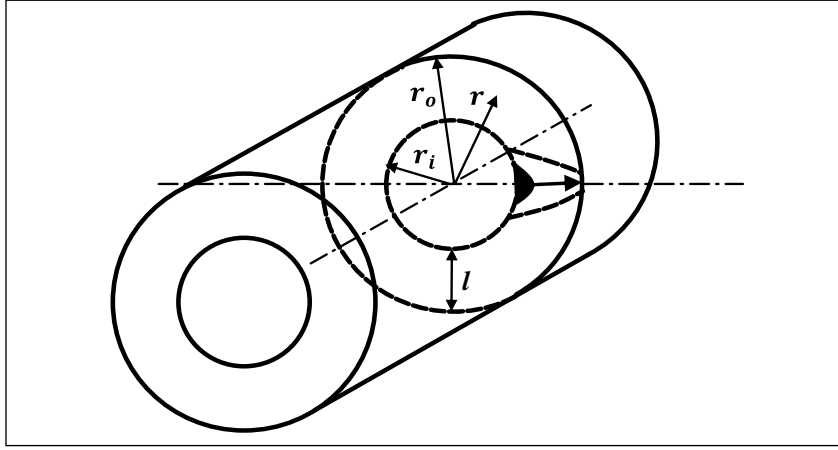


**Fig. 8 Transition rate between states  $S$  and  $F$**

The cumulative distribution function  $F(\tau_{S,F}|\bar{\delta})$  describing the (uncertain) timing of the onset of the degradation process, reaches a value of 0.29 at  $\tau_{S,F} = 100$  years: this confirms that i) the mixing tee between the hot and cold legs is a resistant component capable of withstanding the onset of a flaw due to thermal fatigue for a very long period of time, and that ii) even if unlikely during the life of an NPP, if a crack onset occurs, it appears in the early stage of the component life (as shown in Figures 6, 8).

### 3.2.2 Transition rate $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$

The transition between states Flaw ( $F$ ) (i.e., the onset of a circumferential crack on the inner pipe surface) and Leak ( $L$ ) (i.e., a through-wall circumferential crack) is determined when the crack propagating in the radial direction, reaches the piping wall dimension  $l = r_o - r_i$  (shown in Figure 9). Thus, the critical value of the characteristic variable (i.e., the crack depth  $x_1$ ) is  $X_{cr,1} = l = 9mm$ .



**Fig. 9 Circumferential crack onset**

We assume that the crack growth rate  $dx_1/dN$  follows a generalized Paris law equation (17) [Kozin et. al., 1989; Di Maio et. al., 2013]:

$$\frac{dx_1}{dN} = C \cdot (\Delta K_{eff}(x_1))^n \quad (17)$$

where  $N$  is the number of thermal cycles,  $C$  and  $n$  are material coefficients [Zio et. al., 2012] and  $\Delta K_{eff}$  is the effective stress intensity factor range, that can be expressed as a function of the maximum stress intensity factor range  $\Delta K_I$  and of a parameter  $q(R)$  [Radu et al., 2007 b]:

$$\Delta K_{eff}(\sigma_z, x_1) = q(R) \cdot \Delta K_I(\sigma_z, x_1) \quad (18)$$

where  $\sigma_z$  is the axial stress of Eq. (15) and  $\Delta K_I(\sigma_z, x_1)$  and  $q(R)$  are defined as in Eqs. (19)-(22):

$$\Delta K_I(\sigma_z, x_1) = K_I^{max}(\sigma_z, x_1) - K_I^{min}(\sigma_z, x_1) \quad (19)$$

$$R(\sigma_z, x_1) = \frac{K_I^{min}(\sigma_z, x_1)}{K_I^{max}(\sigma_z, x_1)} \quad (20)$$

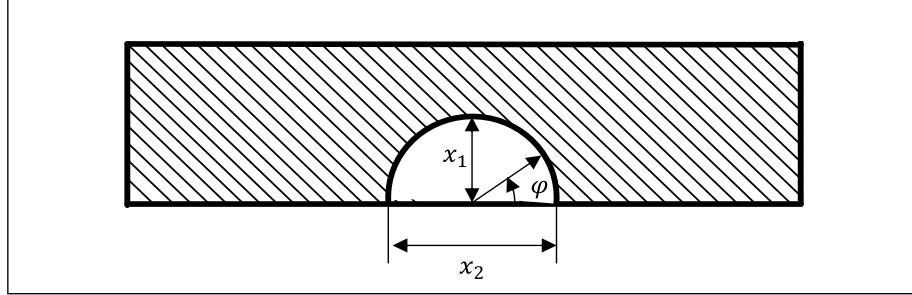
$$q(R) = \frac{1-0.5R}{1-R} \quad \text{if } R < 0 \quad (21)$$

$$q(R) = \frac{1}{1-0.5R} \quad \text{if } R > 0 \quad (22)$$

To calculate  $K_I(\sigma_z, x_1)$ , and its maximum ( $K_I^{max}(\sigma_z, x_1)$ ) and minimum ( $K_I^{min}(\sigma_z, x_1)$ ), we suppose that:

- the crack propagates radially such that the ratio between crack depth ( $x_1$ ) and crack length ( $x_2$ ) is equal to 1 [Anacleto et. al., 2007] (i.e., when the crack depth reaches the wall thickness dimension, at the inner surface of the pipe, the crack length is equal to the wall thickness  $l$ )

- $l/r_i$  is small enough to approximate the cylinder to a flat plate, in fact, for  $l/r_i < 0.1$  the specific model for circumferential crack propagation in a hollow cylinder cannot be used, and the  $K_I$  solution is approximated using the flat plate model.
- the propagation phenomenon is studied at the deepest point of the crack ( $\varphi = \pi/2$ , see Figure 10).



**Fig. 10** Circumferential crack for a cylinder that approximates a flat plate

Theoretical details are given in Appendix 2.

Tables 1 and 2 list the parameter  $\bar{\delta} = (f, \theta_0, E, \alpha, \nu, k, n, C)$  and the distributions used in the evaluation of  $\Delta K_{eff}(\sigma_z, x_1)$  and  $N$  [Chapuliot et. Al., 2005].

Parameter	Description	Unit	Value
$n$	material coefficient		4
$C$	material coefficient	m/cycle	$7.5 \times 10^{-13}$

**Table 2.** Paris law equation parameter values

The number of cycles ( $N$ ) needed to propagate a circumferential through-wall crack, starting from a circumferential crack onset, is given by:

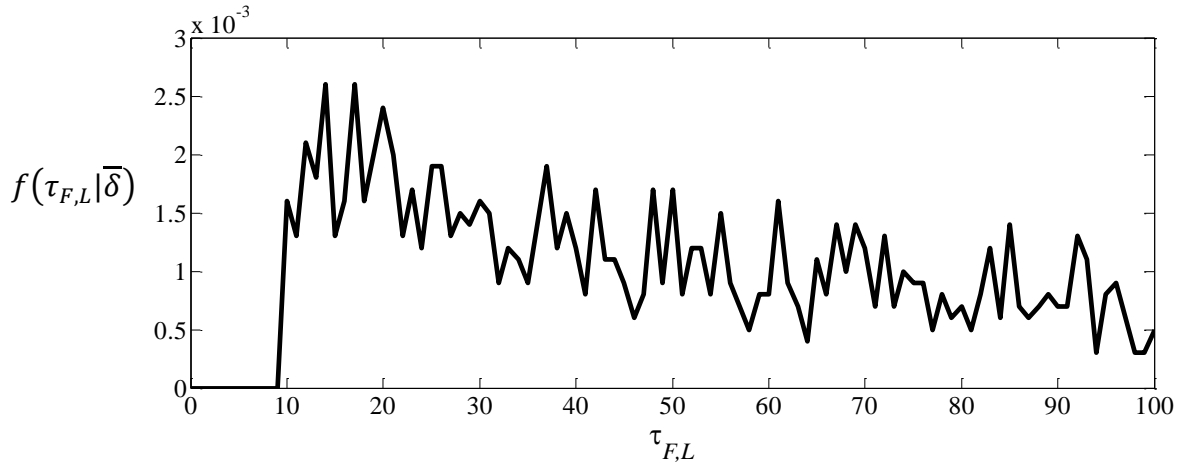
$$N = \int_{x_1^{in}}^{X_{cr,1}} \frac{dx_1}{C \cdot (\Delta K_{eff}(\sigma_z, x_1))^n} \quad (23)$$

The equation has been solved using a MATLAB solver. The state holding time  $\tau_{F,L}$  for a specific set of  $\bar{\delta}$  affecting the radial propagation of the crack is, thus, calculated from  $N$  as:

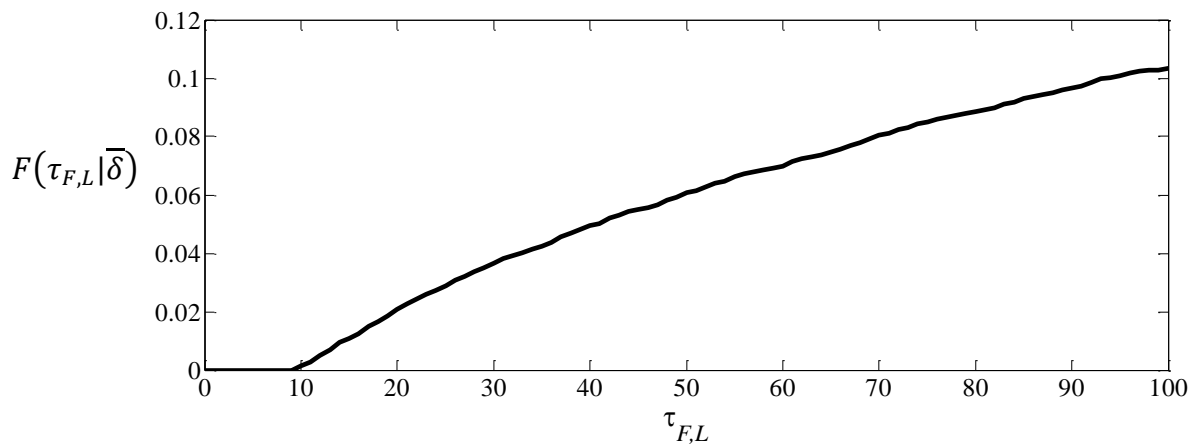
$$\tau_{F,L}(\text{years}) = \frac{N}{f \cdot 3600 \cdot 24 \cdot 365} \quad (24)$$

The probability distribution of  $\tau_{F,L}$  is estimated by applying the MC simulation proposed at item (4) of sub-Section 2.2.1, with a number of trials  $N_c = 10000$ . For each one of these  $N_c$  trials,  $\Delta K_{eff}(\sigma_z, x_1)$  is evaluated given a sampled batch of values of the parameters  $\bar{\delta}$  (Tables 1, 2). Thus,

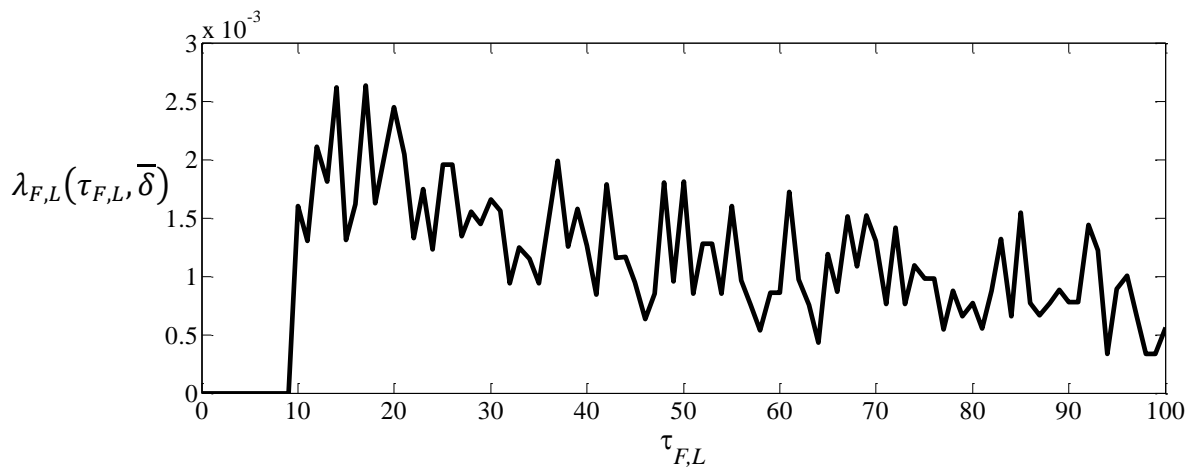
the distributions  $f(\tau_{F,L}|\bar{\delta})$  and  $F(\tau_{F,L}|\bar{\delta})$  can be built (shown in Figures 11 and 12, respectively). By applying Eq. (6) with a time step of one year ( $\Delta\tau = 1$ ), the transition rate  $\lambda_{F,L}(\tau_{F,L},\bar{\delta})$  is computed (shown in Figure 14).



**Fig. 11** Probability density function of the state holding time  $\tau_{F,L}$



**Fig. 12** Cumulative distribution function of the state holding time  $\tau_{F,L}$



**Fig. 13** Transition rate from states  $F$  to state  $L$

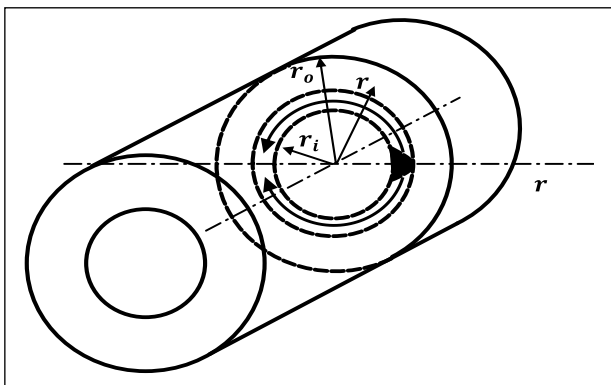


The cumulative distribution function  $F(\tau_{F,L}|\bar{\delta})$  of Figure 12, describing the (uncertain) timing of the transition between states  $F$  and  $L$ , reaches a value of 0.10 at  $\tau_{F,L} = 100$  years, showing that, (hopefully) the laps of time considered is not enough to guarantee that the crack onset propagates radially till a through-wall crack and generates a leakage phenomenon. The transition rate distribution  $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$ , shown in Figure 13 and describing the variation of the probability of the system to leave states  $F$  to enter state  $L$ , shows a discontinuity from  $\lambda_{F,L}(9, \bar{\delta}) = 0$  to  $\lambda_{F,L}(10, \bar{\delta}) = 0.0016$  and a larger value of  $\tau_{F,L}$  around  $10 \leq \tau_{F,L} \leq 20$  years. This leads us to conclude that the circumferential crack that propagates across the piping wall needs at least 10 years to reach a through-wall circumferential characteristic. Moreover, Figures 11 and 13 show that the estimation of  $f(\tau_{F,L}|\bar{\delta})$  and  $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$  provided by the crude MC simulation of Section 2.2.1 is more irregular than  $f(\tau_{S,F}|\bar{\delta})$  and  $\lambda_{S,F}(\tau_{S,F}, \bar{\delta})$  of Figures 6 and 8. This can be improved by resorting to more efficient MC techniques able to dealing with low probability estimation [Zio et. al., 2011].

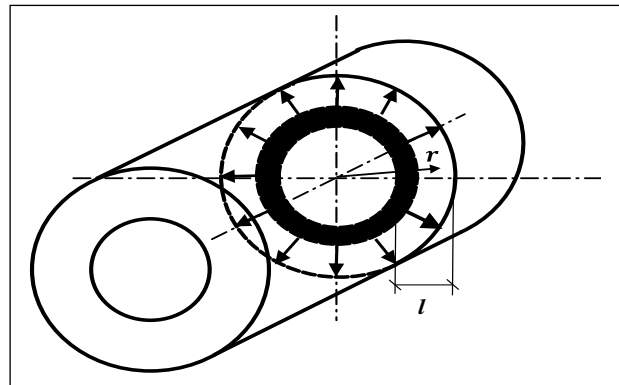
### 3.2.3 Transition rate $\lambda_{F,R}(\tau_{F,R}, \bar{\delta})$

A piping system subjected to thermal fatigue may also break without showing any leakage phenomenon (i.e., the transition between states Flaw ( $F$ ) and Rupture ( $R$ )), This event can be modeled considering two stages:

1. crack propagation along the circumference of the pipe (as shown in Figure 14), that generates a fully-circumferential crack. This phenomenon can be modelled referring to the crack length ( $x_2$ ), whose threshold value  $X_{cr,2}$  is equal to  $2\pi r_i = 754 \text{ mm}$ .
2. radial propagation of the fully-circumferential crack (as shown in Figure 15); this phenomenon can be modelled referring to the crack depth ( $x_1$ ), whose threshold  $X_{cr,1}$  value is equal to  $l = 9 \text{ mm}$ .



**Fig. 14 Circumferential crack propagating along the circumference**



**Fig. 15 Fully-circumferential crack propagating radially**

To obtain the crack growth rate  $dx_i/dN$ , the same procedure explained in sub-Section 3.2.2 is followed and Eqs. (19)-(22) are used. Concerning the propagation of the circumferential crack,  $\sigma_z$  of Eq. (15) is related to the thermal fatigue loading to the fully-circumferential crack [Radu et. al., 2007 b]. The stresses intensity factors  $(K_I^i(\sigma_z, x_i))$   $i = 1, 2$  are computed by applying the BS 7910 procedure for a circumferential crack propagating along the circumference and the API 579 procedure for a fully-circumferential crack propagating radially [FITNET FFS – MK 7].

We suppose that:

- in the first stage of crack propagation, the crack depth ( $x_1 = x_1^{in}$ ) is constant and we approximates the hollow cylinder to a flat plate (see sub-section 3.2.2 hypothesis for the evaluation of  $K_I$ )
- the propagation phenomenon is studied at the inner surface of the pipes ( $\varphi = 0$ , see Figure 10).

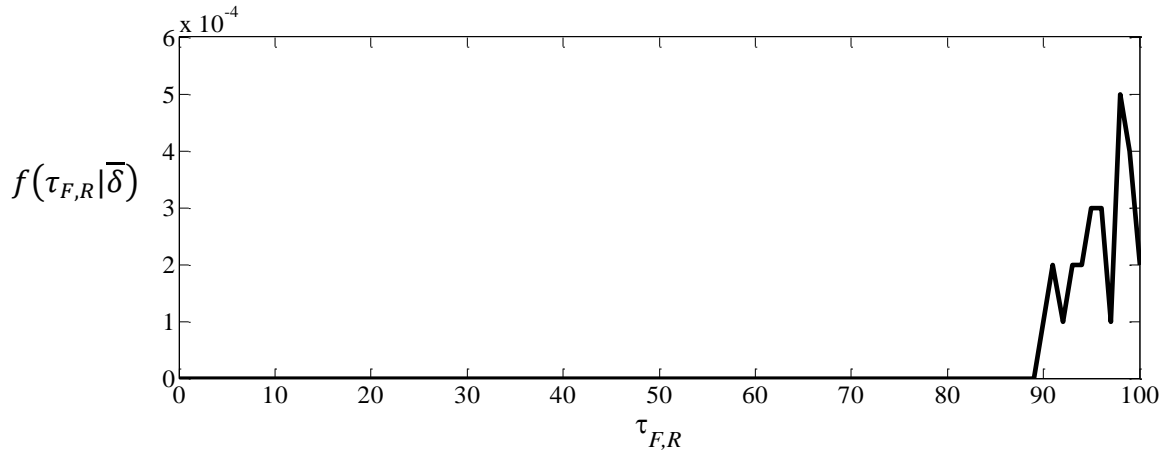
Theoretical details are given in Appendix 2. Tables 1 and 2 list the parameters  $\bar{\delta} = (f, \theta_0, E, \alpha, \nu, k, n, C)$  and the distributions used to evaluate  $\Delta K_{eff}^i(\sigma_z, x_i)$  and  $N$  [Chapuliot et. al., 2005]. The number of cycles ( $N$ ) needed to propagate a fully-circumferential crack, starting from a circumferential crack onset, is given by:

$$N = \int_{x_1^{in}}^{X_{cr,1}} \frac{dx_1}{C \cdot (\Delta K_{eff}^1(\sigma_z, x_1))^n} + \int_{x_2^{in}}^{X_{cr,2}} \frac{dx_2}{C \cdot (\Delta K_{eff}^2(\sigma_z, x_2))^n} \quad (25)$$

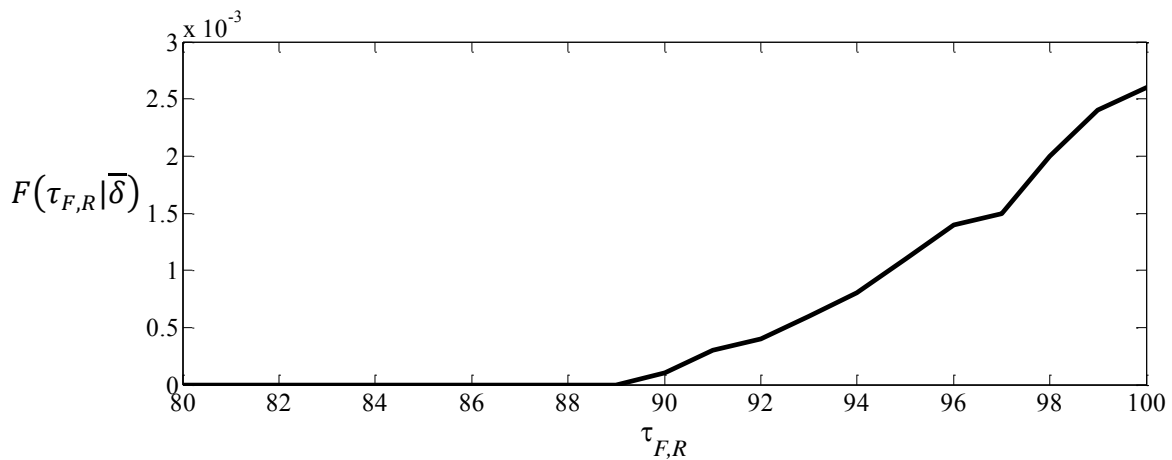
The equation has been solved using a MATLAB solver. The state holding time  $\tau_{F,R}$  for a specific set of  $\bar{\delta}$  affecting the radial propagation of the crack is, thus, calculated from  $N$  as:

$$\tau_{F,R}(years) = \frac{N}{f \cdot 3600 \cdot 24 \cdot 365} \quad (26)$$

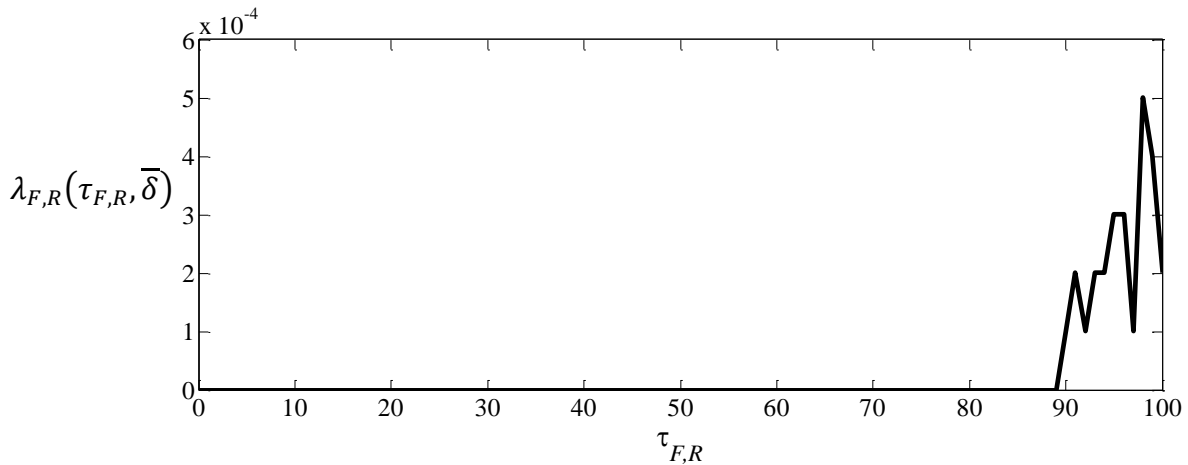
The probability distribution of  $\tau_{F,R}$  is estimated by applying the MC simulation proposed at item (4) of sub-Section 2.2.1, with a number of trials  $N_c = 10000$ . For each one of these  $N_c$  trials,  $\Delta K_{eff}^i(\sigma_z, x_i)$  are evaluated given a sampled batch of values of the parameters  $\bar{\delta}$  (Tables 1, 2). Then, the distributions  $f(\tau_{F,R}|\bar{\delta})$  and  $F(\tau_{F,R}|\bar{\delta})$  can be built (shown in Figures 16 and 17, respectively). By applying Eq. (6) with a time step of one year ( $\Delta\tau = 1$ ), the transition rate  $\lambda_{F,R}(\tau_{F,R}, \bar{\delta})$  is computed (shown in Figure 18).



**Fig. 16 Probability density function of the state holding time  $\tau_{F,R}$**



**Fig. 17 Cumulative distribution function of the state holding time  $\tau_{F,R}$**



**Fig. 18 Transition rate from state  $F$  to state  $R$**

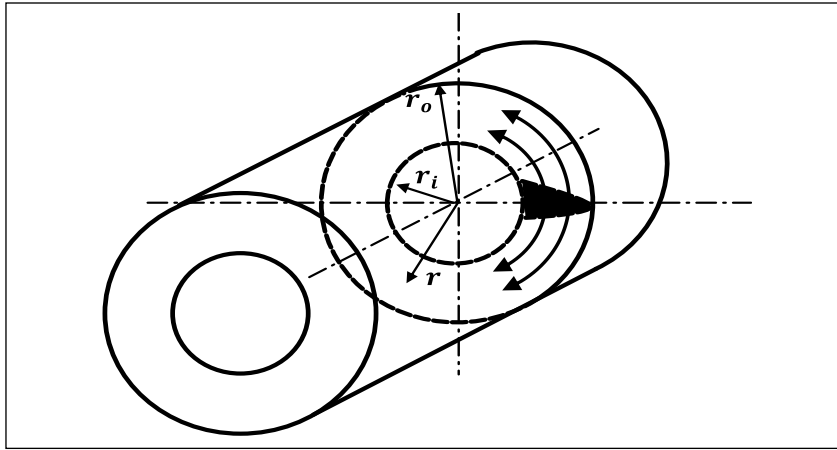
The cumulative distribution function  $F(\tau_{F,R}|\bar{\delta})$  describing the (uncertain) timing of one of the possible second steps of the degradation process (i.e., transition between states  $F$  and  $R$ ), and shown in Figure 17, shows that the transition between these two states can be considered very rare during the usual 40 years of a NPP life (once the component enter state  $F$  it has to wait at least  $\tau_{F,R} = 89$

years to leave state  $F$  and enter state  $R$  with negligible  $f(\tau_{F,R}|\bar{\delta})$  and  $F(\tau_{F,R}|\bar{\delta})$ ). Again, the discontinuity and the irregularity of the curve representing the transition rate  $\lambda_{F,R}(\tau_{F,R}, \bar{\delta})$  can be explained as for the estimation of  $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$  of Section 3.2.2.

### 3.2.4 Transition rate $\lambda_{L,R}(\tau_{L,R})$

The transition between states Leak ( $L$ ) (i.e., through-wall crack that presents leakage phenomena) and rupture ( $R$ ) (i.e., the pipe is completely broken), shown in Figure 19, occurs when the crack length ( $x_2$ ) has reached the circumference dimension in the outside surface of the pipe, that is equal to:

$$X_{cr,2} = 2\pi r_o = 819 \text{ mm} \quad (27)$$



**Fig. 19 Through-wall circumferential crack propagating along the circumference**

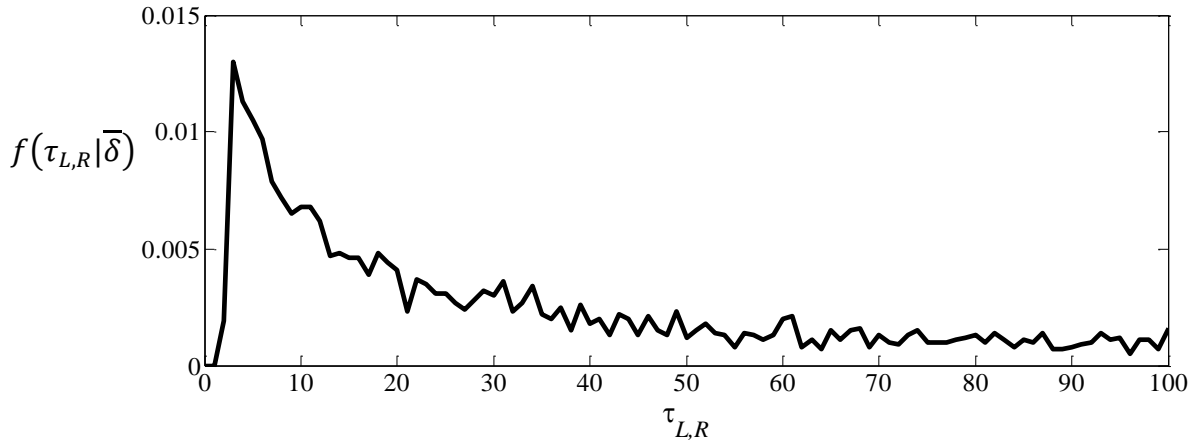
Responsible for the crack propagation is the axial stress ( $\sigma_z$ ), given in Eq. (15) [Radu et. al., 2007 b]. To determine the number of thermal cycles  $N$  after which the system will experience a rupture, the same procedure explained in Section 3.2.2 is followed and Eqs. (19)-(22) are used. The stress intensity factor ( $K_I(\sigma_z, x_2)$ ) is computed by applying the BS 7910 procedure for a through-wall circumferential crack propagating along the pipe circumference [FITNET FFS – MK 7]. Theoretical details are given in Appendix 2. Tables 1 and 2 list the parameters  $\bar{\delta} = (f, \theta_0, E, \alpha, \nu, k, n, C)$  and their distributions used to evaluate  $\Delta K_{eff}(\sigma_z, x_2)$  and  $N$  [Chapuliot et. al., 2005]. The number of cycles ( $N$ ) needed to reach  $X_{cr}$ , starting from a through-wall crack, is given by:

$$N = \int_{x_2^{in}}^{X_{cr,2}} \frac{dx_2}{C \cdot (\Delta K_{eff}(\sigma_z, x_2))^n} \quad (28)$$

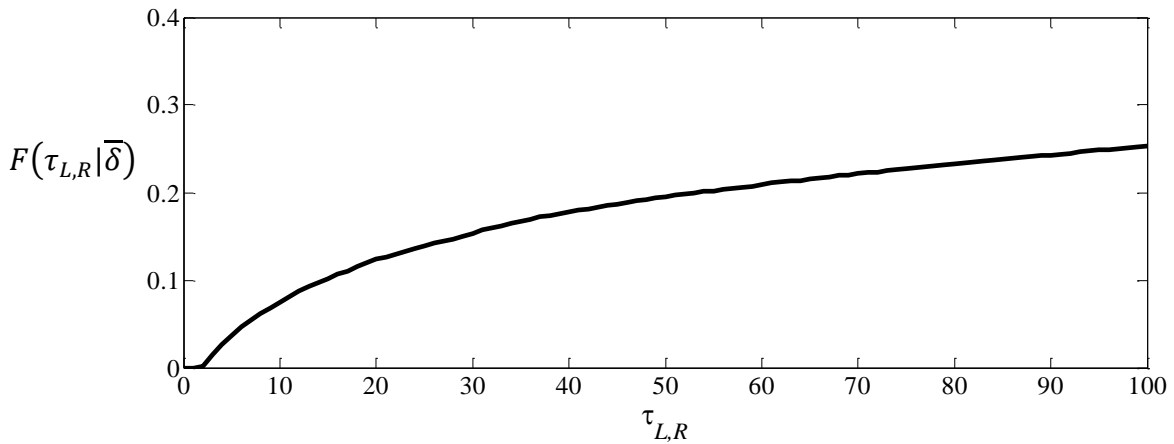
The initial crack length  $x_{in}$  is considered equal to  $x_2^{in} = 28 \mu m$ . The equation has been solved using a MATLAB solver. The state holding time  $\tau_{L,R}$  for a specific set of  $\bar{\delta}$  affecting the propagation of the crack is, then, calculated from  $N$  as:

$$\tau_{L,R}(years) = \frac{N}{f \cdot 3600 \cdot 24 \cdot 365} \quad (29)$$

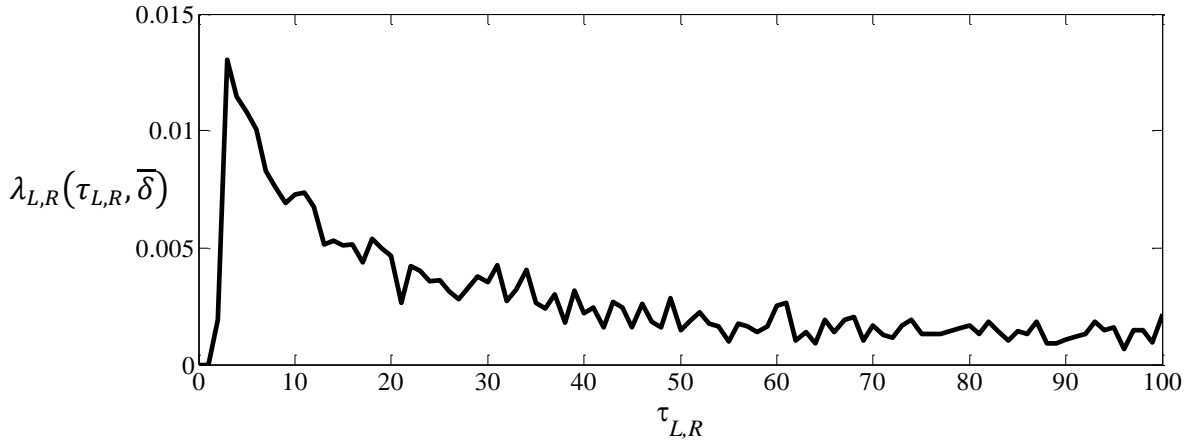
The probability distribution of  $\tau_{L,R}$  is estimated by applying the MC simulation proposed at item (4) of Section 2.2.1, with a number of trials  $N_c = 10000$ . For each one of these  $N_c$  trials,  $\Delta K_{eff}(\sigma_z, x)$  is evaluated given a sampled batch of values of the parameters  $\bar{\delta}$  (Tables 1, 2). Then, the distributions  $f(\tau_{L,R}|\bar{\delta})$  and  $F(\tau_{L,R}|\bar{\delta})$  can be built (shown in Figures 20 and 21, respectively). By applying Eq. (6) with a time step of one year ( $\Delta\tau = 1$ ), the transition rate  $\lambda_{L,R}(\tau_{L,R}, \bar{\delta})$  is computed (shown in Figure 22).



**Fig. 20 Probability density function of the state holding time  $\tau_{L,R}$**



**Fig. 21 Cumulative distribution function of the state holding time  $\tau_{L,R}$**



**Fig. 22 Transition rate from states  $L$  to state  $R$**

Figures 20 and 22 show that once the crack reaches a circumferential through-wall characteristic, it rapidly propagates (peak value at the 3<sup>rd</sup> year). Moreover, the Figure shows that provided that the component is in in state  $L$ , the transition out of this state into state  $R$  can occur with high probability.

### 3.2.5 Repair transition rates $\omega$ and $\mu$

The repair rates  $\omega$  and  $\mu$  are estimated by means of two simple models described as follows [Fleming, 2004]:

$\omega = \frac{P_I P_{FD}}{(T_{FI} + T_R)} = 2 \times 10^{-2}/yr$  components are assumed to have a 25% chance ( $P_I$ ) of being inspected for flaws detection every 10 years ( $T_{FI}$ ) with a 90% detection probability ( $P_{FD}$ ); detected Flaws will be repaired in 200 h ( $T_R=200 \text{ h}/8760 \text{ h/ year}$ ).

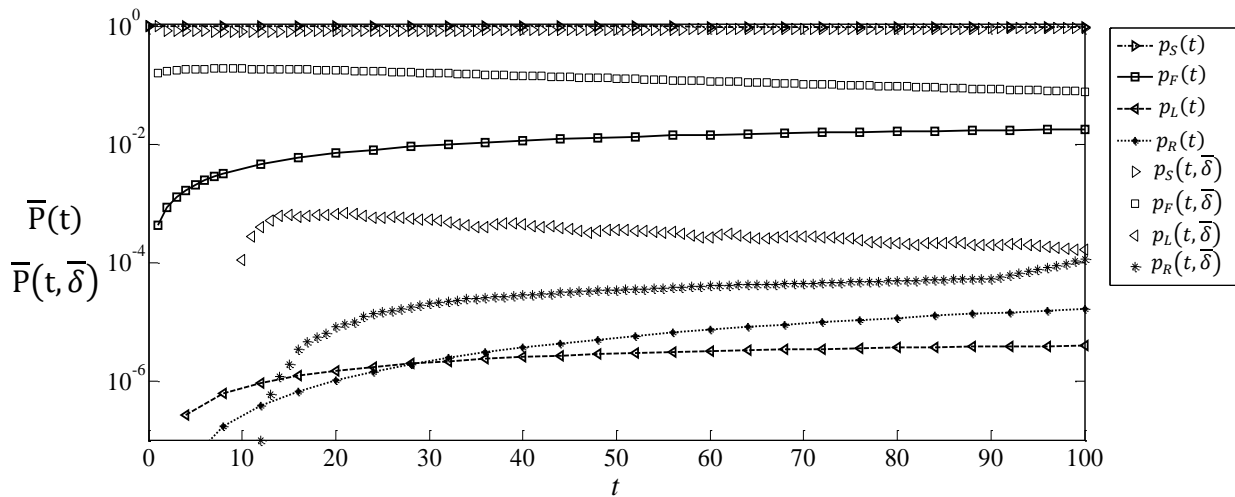
$\mu = \frac{P_I P_{LD}}{(T_{LI} + T_R)} = 7.92 \times 10^{-1}/yr$  the components are assumed to have a 90% chance ( $P_I$ ) of being inspected for leak detection every 1 years ( $T_{LI}$ ) with a 90% detection probability ( $P_{LD}$ ); detected Leaks will be repaired in 200 h ( $T_R=200 \text{ h}/8760 \text{ h/ year}$ ).

These two transition rates are considered constant and the state transition time will follow an exponential distribution.

### 3.3 State Probability $\bar{P}(t)$

To evaluate the probability vector solution  $\bar{P}(t)$ , the procedure explained in sub-Section 2.2.2 is followed. Figure 23 shows the solution of  $\bar{P}(t, \deltā)$  obtained with a MSPM whose transition rates values have been defined in the previous Sections and the solution obtained by solving an MCM as in [Fleming, 2004], in which the transition rates are considered constant so that the state holding

times are exponentially distributed. Figures 23 shows that the probabilities  $p_L(t, \bar{\delta})$  and  $p_R(t, \bar{\delta})$ , obtained with the MSPM, are larger than  $p_L(t)$  and  $p_R(t)$  obtained by MCM [Fleming, 2004] due to  $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$  and  $\lambda_{L,R}(\tau_{L,R}, \bar{\delta})$  of Figures 13 and 22. Moreover, it is worth noticing the decreasing trends of the  $p_F(t, \bar{\delta})$  and  $p_L(t, \bar{\delta})$ : those phenomena can be explained looking at the distributions of the transition rates  $\lambda_{F,L}(\tau_{F,L}, \bar{\delta})$  and  $\lambda_{L,R}(\tau_{L,R}, \bar{\delta})$ , shown in Figures 13 and 22, that have decreasing trends as  $p_F(t, \bar{\delta})$  and  $p_L(t, \bar{\delta})$ . This means that once the system enters states  $F$  or  $L$ , it has a larger probability to leave these states and enter states  $L$  or  $R$ , respectively, after a short time. With respect to the transition from state  $F$  to  $R$ , it can be seen that the distribution of  $\lambda_{F,R}(\tau_{F,R}, \bar{\delta})$  of Figure 18 influences the  $p_R(t, \bar{\delta})$  to be in state  $R$  at the very end of its life (90-100 years).



**Fig. 23 State probability vector solution. Comparison of MCM and MSPM approaches**

Indeed, the estimates provided by MCM (dotted line with diamonds) and MSPM (stars) differ from the early stage of the piping system operation. This is due to the fact that the integration in the MSPM of more physical information (data and models) than in MCM has allowed a more realistic degradation process modeling. As a result, below 15 years the probability of rupture is not credible for MCM (e.g., leading to a relaxation of maintenance/repair efforts, with cost savings when relying on the MSPM results), whereas at larger times the probability of rupture is underestimated by MCM (~1 order of magnitude), with the risk associated to this.

## 4 Conclusions

A Multi-State Physics Modeling (MSPM) framework for degradation modeling and failure probability quantification of Nuclear Power Plants piping systems has been developed.

The approach has been applied to a benchmark problem of a piping system of a Pressurized Water Reactor undergoing thermal fatigue. The results are compared with a Continuous-time homogeneous Markov Chain Model (MCM).

The transition rates describing the degradation phenomenon in the MSPM have been determined by simulating the degradation (physics) models that describe the different stages of the thermal fatigue degradation process of a piping component and evaluating by Monte Carlo (MC) simulation the time-dependent transition rates between the states of the MSPM.

The comparison of the MCM with the MSPM results shows that with more realistic assumptions and consistent exploitation of the available knowledge (data and models), the latter method gives larger probabilities of occurrence of a leakage/rupture in the piping system, than the MCM. This difference in the estimates can be significant from the risk point of view, as this could be underestimated with all associated consequences. This shows the importance of finding “modeling ways” to include all the knowledge and information available (in the form of data, models, expert judgments, etc.) for an informed-as-possible, faithful-as-possible description of the real degradation and failure mechanism. Finally, another advantage of the MSPM for piping systems failure probability quantification is its applicability to assess the reliability of newly designed NPPs components when lacking of field data.

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## APPENDIX 1

### 1 Thermal fatigue Stresses solution

Thermal stresses due to thermal fatigue are dependent on the temperature distribution  $\theta(r, \omega, \theta_0, k, t)$  across the wall thickness. For the simple geometry of a pipe that can be represented as a hollow cylinder, analytical solutions for temperature fields and associated elastic thermal

stresses  $\sigma_r$ ,  $\sigma_z$ , and  $\sigma_\theta$  distributions for a pipe subject to sinusoidal transient thermal loading, have been developed in [Radu et al., 2007 a].

The one-dimensional heat diffusion equation in cylindrical coordinates and with axisymmetric thermal variations is [Radu et al., 2007 a]:

$$\begin{cases} \frac{\partial^2 \theta}{\partial r^2} + \frac{1}{r} \frac{\partial \theta}{\partial r} = \frac{1}{k} \frac{\partial \theta}{\partial t} \\ \theta(r_i, t) = \theta_0 \cdot \sin(\omega t) \\ \theta(r_o, t) = 0 \\ \theta(r, 0) = 0 \end{cases} \quad (30)$$

where  $\theta = T(r, t) - T_0$  and  $T_0$ , is the unstrained temperature.

The solution for the temperature distribution during a thermal transient can be written as follows [Radu et al., 2007 a]:

$$\theta(r, \omega, \theta_0, k, t) = k \cdot \pi \cdot \sum_{n=1}^{\infty} \theta_1(r_i, r_o, s_n) \cdot \theta_2(r_i, r, s_n) \cdot \theta_3(\omega, t, s_n) \quad (31)$$

where

$$\theta_1(r_i, r_o, s_n) = \frac{s_n^2 \cdot J_0^2(s_n r_o)}{J_0^2(s_n r_o) - J_0^2(s_n r_i)} \quad (32)$$

$$\theta_2(r_i, r, s_n) = Y_0(s_n r_i) \cdot J_0(s_n r) - J_0(s_n r_i) \cdot Y_0(s_n r) \quad (33)$$

$$\theta_3(\omega, t, s_n) = \theta_0 \cdot \frac{\omega \cdot e^{-k \cdot s_n^2 \cdot t} + (k \cdot s_n^2) \cdot \sin(\omega t) - \omega \cdot \cos(\omega t)}{(k \cdot s_n^2)^2 + \omega^2} \quad (34)$$

and  $s_n$  are the positive roots of

$$Y_0(s_n \cdot r_i) \cdot J_0(s_n \cdot r_o) - J_0(s_n \cdot r_i) \cdot Y_0(s_n \cdot r_o) = 0 \quad (35)$$

where  $J_\nu(z)$ ,  $Y_\nu(z)$  are the Bessel functions of first and second kind of order ( $\nu$ ).

The temperature distribution  $\theta(r, \omega, \theta_0, k, t)$  has been used to calculate the thermal stress components. The one-dimensional equilibrium equation in the radial direction is [Radu et al., 2007 a]:

$$\frac{d\sigma_r}{dr} + \frac{\sigma_r - \sigma_\theta}{r} = 0 \quad (36)$$

The displacement technique has been used to solve the axisymmetric problems of hollow cylinders. When all the strains and stresses are only functions of the radial distance  $r$ , the strain displacement relations are [Radu et al., 2007 a]:

$$\varepsilon_r = \frac{du}{dr} \quad \varepsilon_\theta = \frac{u}{r} \quad \varepsilon_{r\theta} = 0 \quad (37)$$

where  $u$  is the radial displacement.

The components of stress in cylindrical coordinates can be expressed as [Radu et al., 2007 a]:

$$\sigma_r = \frac{E'}{1-\nu'^2} \left[ \frac{du}{dr} + \nu' \cdot \frac{u}{r} - (1 + \nu') \cdot \alpha' \cdot \theta + (1 + \nu') \cdot c' \right] \quad (38)$$

$$\sigma_\theta = \frac{E'}{1-\nu'^2} \left[ \nu' \cdot \frac{du}{dr} + \frac{u}{r} - (1 + \nu') \cdot \alpha' \cdot \theta + (1 + \nu') \cdot c' \right] \quad (39)$$

$$\sigma_{r\theta} = 0 \quad (40)$$

substituting equation (38,39) into equation (36), we get:

$$\frac{d}{dr} \left[ \frac{1}{r} \cdot \frac{d(r \cdot u)}{dr} \right] = (1 + \nu') \cdot \alpha' \cdot \frac{d\theta(r,t)}{dr} \quad (41)$$

and the general solution of equation (42) is:

$$u = (1 + \nu') \cdot \alpha' \cdot \frac{1}{r} \cdot \int_r \theta(r, t) \cdot r \cdot dr + C_1 \cdot r + \frac{C_2}{r} \quad (42)$$

where the constants  $E'$ ,  $\nu'$ ,  $\alpha'$ ,  $c'$  in Eqs. (38)-(40) are:

$$E' = \frac{E}{1-\nu^2} \quad \nu' = \frac{\nu}{1-\nu} \quad \alpha' = (1 + \nu) \cdot \alpha \quad c' = \nu \cdot \varepsilon_0 \quad (43)$$

The radial, hoop and axial stresses for an hollow cylinder made of an homogeneous isotropic material are [Radu et al., 2007 a]:

$$\sigma_r(r, \omega, \theta_0, k, t) = \frac{\alpha \cdot E}{1-\nu} \cdot \left[ -\frac{1}{r^2} \cdot I_1(r, \omega, \theta_0, k, t) + \frac{r^2 - r_i^2}{r^2 \cdot (r_o^2 - r_i^2)} \cdot I_2(r, \omega, \theta_0, k, t) \right] \quad (44)$$

$$\sigma_\theta(r, \omega, \theta_0, k, t) = \frac{\alpha \cdot E}{1-\nu} \cdot \left[ \frac{1}{r^2} \cdot I_1(r, \omega, \theta_0, k, t) + \frac{r^2 - r_i^2}{r^2 \cdot (r_o^2 - r_i^2)} \cdot I_2(r, \omega, \theta_0, k, t) - \theta(r, \omega, \theta_0, k, t) \right] \quad (45)$$

$$\sigma_z(r, \omega, \theta_0, k, t) = \frac{\alpha \cdot E}{1-\nu} \cdot \left( \frac{2\nu}{r_o^2 - r_i^2} \cdot I_2(r, \omega, \theta_0, k, t) - \theta(r, \omega, \theta_0, k, t) \right) \quad (46)$$

where  $I_1(r, \omega, \theta_0, k, t)$  and  $I_2(r, \omega, \theta_0, k, t)$  are expressed as:

$$I_1(r, \omega, \theta_0, k, t) = \int_{r_i}^r \theta(r, \omega, \theta_0, k, t) \cdot r \cdot dr = k \cdot \pi \cdot \sum_{n=1}^{\infty} \frac{s_n^2 \cdot J_0^2(s_n \cdot r_o)}{J_0^2(s_n \cdot r_o) - J_0^2(s_n \cdot r_i)} \times \\ \times \left[ \frac{1}{s_n} \{ Y_0(s_n \cdot r_i) \cdot [r \cdot J_1(s_n \cdot r) - r_i \cdot J_1(s_n \cdot r_i)] - J_0(s_n \cdot r_i) \cdot [r \cdot Y_1(s_n \cdot r) - r_i \cdot Y_1(s_n \cdot r_i)] \} \right] \times \\ \times \left[ \theta_0 \cdot \frac{\omega \cdot e^{-ks_n^2 \cdot t} + (k \cdot s_n^2) \cdot \sin(\omega \cdot t) - \omega \cdot \cos(\omega \cdot t)}{(k \cdot s_n^2)^2 + \omega^2} \right] \quad (47)$$

$$I_2(r, \omega, \theta_0, k, t) = \int_{r_i}^{r_o} \theta(r, \omega, t) \cdot r \cdot dr = k \cdot \pi \cdot \sum_{n=1}^{\infty} \frac{s_n^2 \cdot J_0^2(s_n \cdot r_o)}{J_0^2(s_n \cdot r_o) - J_0^2(s_n \cdot r_i)} \times \\ \times \left[ \frac{1}{s_n} \{ Y_0(s_n \cdot r_i) \cdot [r_o \cdot J_1(s_n \cdot r) - r_i \cdot J_1(s_n \cdot r_i)] - J_0(s_n \cdot r_i) \cdot [r_o \cdot Y_1(s_n \cdot r_o) - r_i \cdot Y_1(s_n \cdot r_i)] \} \right] \\ \times \left[ \theta_0 \cdot \frac{\omega \cdot e^{-ks_n^2 \cdot t} + (k \cdot s_n^2) \cdot \sin(\omega \cdot t) - \omega \cdot \cos(\omega \cdot t)}{(k \cdot s_n^2)^2 + \omega^2} \right] \quad (48)$$

For clarity sake, in Eqs. (13)-(15)  $t$  has not been reported.

## APPENDIX 2

### 2 BS 7910 procedure for flat plate $K_I(\sigma_m, \sigma_b, x_1, x_2)$ solution

In the BS 7910 procedure, the stress intensity factor  $K_I(\sigma_m, \sigma_b, x_1, x_2)$  is expressed as a function of the stress intensity magnification factor for membrane ( $M_m$ ) and bending ( $M_b$ ) stresses, the finite

width factor ( $f_w$ ) and the linearized membrane ( $\sigma_m(x_1)$ ) and bending ( $\sigma_b(x_1)$ ) stresses over the position of the crack depth as [FITNET FFS – MK 7]:

$$K_I(\sigma_m, \sigma_b, x_1, x_2) = f_w(x_1, x_2, l, w) \cdot [M_m(x_1, x_2, l) \cdot \sigma_m + M_b(x_1, x_2, l) \cdot \sigma_b] \sqrt{\pi x_1} \quad (49)$$

$M_m(x_1, x_2, l)$  and  $M_b(x_1, x_2, l)$ , functions of the crack depth ( $x_1$ ) and crack length ( $x_2$ ) and of the wall thin ( $l$ ), are equal to [FITNET FFS – MK 7]:

$$M_m = \left[ M_1 + M_2 \cdot \left( \frac{x_1}{l} \right)^2 + M_3 \cdot \left( \frac{x_1}{l} \right)^4 \right] \frac{gf_\vartheta}{\Phi} \quad (50)$$

$$M_b = M_m [H_1 + (H_2 - H_1)(\sin \varphi)^q] \quad (51)$$

where  $M_1, M_2, M_3, \Phi, H_1, H_2, q$ , are given in the following Eqs. (52)-(58) and Eqs. (59)-(65) for a ratio of  $x_1/x_2 \leq 0.5$  and  $0.5 < x_1/x_2 \leq 1$ , respectively [FITNET FFS – MK 7].

$$M_1 = 1.13 - 0.09 \cdot \left( \frac{x_1}{x_2/2} \right) \quad (52)$$

$$M_2 = \left[ \frac{0.89}{0.2 + \left( \frac{x_1}{x_2/2} \right)} \right] - 0.54 \quad (53)$$

$$M_3 = 0.5 - \frac{1}{0.65 + \frac{x_1}{x_2/2}} + 14 \left( 1 - \frac{x_1}{x_2/2} \right)^{24} \quad (54)$$

$$\Phi = \sqrt{1 - 1.464 \cdot \left( \frac{x_1}{x_2/2} \right)^{1.65}} \quad (55)$$

$$H_1 = 1 - 0.34 \cdot \left( \frac{x_1}{l} \right) - 0.11 \cdot \left( \frac{x_1}{l} \right) \cdot \left( \frac{x_1}{x_2/2} \right) \quad (56)$$

$$H_2 = 1 - \left[ -1.22 + 0.12 \cdot \left( \frac{x_1}{x_2/2} \right) \right] \cdot \left( \frac{x_1}{l} \right) - \left[ 0.55 - 1.05 \cdot \left( \frac{x_1}{x_2/2} \right)^{0.75} + 0.47 \cdot \left( \frac{x_1}{x_2/2} \right)^{1.5} \right] \cdot \left( \frac{x_1}{l} \right)^2 \quad (57)$$

$$q = 0.2 + \left( \frac{x_1}{x_2/2} \right) + 0.6 \cdot \left( \frac{x_1}{l} \right) \quad (58)$$

$$M_1 = \left( \frac{x_2/2}{x_1} \right)^{0.5} \left[ 1 - 0.04 \cdot \left( \frac{x_2/2}{x_1} \right) \right] \quad (59)$$

$$M_2 = 0.2 \cdot \left( \frac{x_2/2}{x_1} \right)^4 \quad (60)$$

$$M_3 = -0.11 \cdot \left( \frac{x_2/2}{x_1} \right)^4 \quad (61)$$

$$\Phi = \sqrt{1 - 1.464 \cdot \left( \frac{x_2/2}{x_1} \right)^{1.65}} \quad (62)$$

$$H_1 = 1 - \left[ 0.04 + 0.41 \cdot \left( \frac{x_2/2}{x_1} \right) \right] \cdot \left( \frac{x_1}{l} \right) + \left[ 0.55 - 1.93 \cdot \left( \frac{x_2/2}{x_1} \right)^{0.75} + 1.38 \cdot \left( \frac{x_2/2}{x_1} \right)^{1.5} \right] \cdot \left( \frac{x_1}{l} \right)^2 \quad (63)$$

$$H_2 = 1 - \left[ -2.11 + 0.77 \cdot \left( \frac{x_2/2}{x_1} \right) \right] \cdot \left( \frac{x_1}{l} \right) - \left[ 0.55 - 0.72 \cdot \left( \frac{x_2/2}{x_1} \right)^{0.75} + 0.14 \cdot \left( \frac{x_2/2}{x_1} \right)^{1.5} \right] \cdot \left( \frac{x_1}{l} \right)^2 \quad (64)$$

$$q = 0.2 + \left( \frac{x_2/2}{x_1} \right) + 0.6 \cdot \left( \frac{x_1}{l} \right) \quad (65)$$

For the parameters  $g$  and  $f_\vartheta$ , we used the simplified values listed in Table 3 [FITNET FFS – MK 7].

	$\varphi = \pi/2$		$\varphi = 0$	
	$x_1/x_2 \leq 0.5$	$0.5 < x_1/x_2 \leq 1$	$x_1/x_2 \leq 0.5$	$0.5 < x_1/x_2 \leq 1$
$g$	1	1	$1.1 + 0.35 \cdot (x_1/l)^2$	$1.1 + 0.35 \cdot \left(\frac{x_2/2}{x_1}\right) \cdot (x_1/l)^2$
$f_\vartheta$	1	$\left(\frac{x_2/2}{x_1}\right)^{0.5}$	$(x_1/c)^{0.5}$	1

**Table 3. Simplified values for  $g$  and  $f_\vartheta$  parameters**

The finite width factor ( $f_w$ ) is expressed as Eq (66), where the surface width  $W = 2\pi r_m$  and  $r_m$  is the mean radius of the pipe [FITNET FFS – MK 7]:

$$f_w = \left[ \sec \left[ \left( \frac{\pi x_2/2}{W} \right) \sqrt{\left( \frac{x_1}{l} \right)} \right] \right]^{0.5} \quad (66)$$

The linearized stresses  $\sigma_m(x_1)$  and  $\sigma_b(x_1)$  (MPa), functions of the crack depth, take values from the analytical stresses distributions computed as in Appendix 1.  $\sigma_m(x_1)$  and  $\sigma_b(x_1)$  are expressed as [FITNET FFS – MK 7]:

$$\sigma_m(x_1) = \frac{4Nx_1^2 - 6Mx_1 - 3Nx_1 + 6Ml}{x_1^3} \quad (67)$$

$$\sigma_b(x_1) = \frac{3Nx_1 - 6Ml}{x_1^3} \quad (68)$$

where the parameters  $N$  and  $M$  are:

$$N = \int_0^{x_1} \sigma(r_n) dr_n \quad (69)$$

$$M = \int_0^{x_1} x\sigma(r_n) dr_n \quad (70)$$

In this work,  $\sigma(r_n)$  is the 4<sup>th</sup> order polynomial expression fitted to the through-wall stress profile  $\sigma_z$  in the radial direction (that corresponds to the crack propagation direction) of the hollow cylinder, for an instant of time  $t_i^s$  into the sinusoidal period, and  $r_n$  is the radial coordinate starting at the inner surface of the pipe.

To evaluate the stress intensity factor, the following steps must be followed:

1. Evaluate  $i$  4<sup>th</sup> order polynomial expressions, fitted to the stress profiles in radial direction of  $\sigma_z$ , given in Eqs. (14),(15), for different times  $t_i^s$   $i = 1, 2, \dots, n$  in the entire sinusoidal period, as a function of the normalized radial distance  $\frac{r_n}{l}$ , for  $t_i^s$   $i = 1, 2, \dots, n$  [Radu et. al., 2007 b]:

$$\sigma_i(r_n/l) = \sigma_0 + \sigma_1 \cdot \left(\frac{r_n}{l}\right) + \sigma_2 \cdot \left(\frac{r_n}{l}\right)^2 + \sigma_3 \cdot \left(\frac{r_n}{l}\right)^3 + \sigma_4 \cdot \left(\frac{r_n}{l}\right)^4 \quad (71)$$

2. Evaluate  $\sigma_m(x_1)$  and  $\sigma_b(x_1)$  for  $n$  steps of crack dimension, until it reaches a selected threshold value ( $X_{cr}$ ), for each of the  $i$   $\sigma_i(r_n/l)$ .

3. Evaluate  $K_I(\sigma_m, \sigma_b, x_1, x_2)$  applying Eq.(50).
4. For each crack growth step  $n$ , select the maximum and minimum values the  $K_I(\sigma_m, \sigma_b, x_1, x_2)$  among the entire sinusoidal period.
5. Evaluate  $\Delta K_{eff}$ , function of crack depth  $x_1$  as explained in section 3.2.2., by fitting with a cubic spline interpolation (e.g. using the MATLAB function) to use in the Paris law equation.

## 1 API 579 procedure for a fully circumferential crack $K_I(\sigma(r_n/l), x_1)$ solution

$K_I(x_1)$  is evaluated following API 579 procedure, as [Radu et al., 2007 b]:

$$K_I\left(\frac{x_1}{l}\right) = \left[ G_0 \cdot \sigma_0 + G_1 \cdot \sigma_1 \cdot \left(\frac{x_1}{l}\right) + G_2 \cdot \sigma_2 \cdot \left(\frac{x_1}{l}\right)^2 + G_3 \cdot \sigma_3 \cdot \left(\frac{x_1}{l}\right)^3 + G_4 \cdot \sigma_4 \cdot \left(\frac{x_1}{l}\right)^4 \right] \cdot \sqrt{\frac{\pi x_1}{Q}} \quad (72)$$

where  $\sigma_j, j = 0, 1, 2, 3, 4$ , are the coefficients of the polynomial stress distribution (MPa),  $Q = 1$  and the coefficients  $G_0, G_1, G_2, G_3, G_4$ , are expressed in Eq. (76)-(80), respectively, [Radu et al., 2007 b] for a fully-circumferential crack, and a ratio  $r_i/l = 13$ .

$$G_0\left(\frac{x_1}{l}\right) = 1.1198 + 0.1938 \cdot \left(\frac{x_1}{l}\right) + 2.9663 \cdot \left(\frac{x_1}{l}\right)^2 - 0.5521 \cdot \left(\frac{x_1}{l}\right)^3 \quad (73)$$

$$G_1\left(\frac{x_1}{l}\right) = 0.6812 + 0.1654 \cdot \left(\frac{x_1}{l}\right) + 0.7604 \cdot \left(\frac{x_1}{l}\right)^2 + 0.1385 \cdot \left(\frac{x_1}{l}\right)^3 \quad (74)$$

$$G_2\left(\frac{x_1}{l}\right) = 0.5234 + 0.1608 \cdot \left(\frac{x_1}{l}\right) + 0.1388 \cdot \left(\frac{x_1}{l}\right)^2 + 0.3354 \cdot \left(\frac{x_1}{l}\right)^3 \quad (75)$$

$$G_3\left(\frac{x_1}{l}\right) = 0.4391 + 0.1557 \cdot \left(\frac{x_1}{l}\right) - 0.1345 \cdot \left(\frac{x_1}{l}\right)^2 + 0.4271 \cdot \left(\frac{x_1}{l}\right)^3 \quad (76)$$

$$G_4\left(\frac{x_1}{l}\right) = 0.3785 + 0.0937 \cdot \left(\frac{x_1}{l}\right) + 0.0151 \cdot \left(\frac{x_1}{l}\right)^2 + 0.2211 \cdot \left(\frac{x_1}{l}\right)^3 \quad (77)$$

Following the points 1.-4. in Appendix 2 Section 1, the stress intensity factor, function of crack depth, is evaluated.

## 2 BS 7910 procedure for a through-wall circumferential crack $K_I(\sigma(r_n/l), x_2)$ solution

The stress intensity factor is evaluated considering the following hypothesis:

1. The crack length is evaluated on the outside surface of the pipe.
2. Membrane ( $\sigma_m$ ) and bending ( $\sigma_b$ ) stresses are obtained applying Eqs. (67)-(70) with  $x_1 = 9 \text{ mm}$ .

The stress intensity factor ( $K_I(x_2)$ ) can be expressed as [FITNET FFS – MK 7]:

$$K_I(x_2) = (M_m \sigma_m + M_b \sigma_b) \sqrt{\pi x_2} \quad (78)$$

where  $M_m$  and  $M_b$  are expressed as in Table 4 as a function of  $\lambda_M$  that is equal to [FITNET FFS – MK 7]:

$$\lambda_M = [12 \cdot (1 - \nu^2)]^{0.25} \frac{x_2/2}{\sqrt{r_m}} \quad (79)$$

Following the points 1.- 4. in Appendix 2 Section 1, the stress intensity factor, function of crack depth, is evaluated.

$\lambda_M$	<b>Mm</b>	<b>Mb</b>
0	1	1
1	1.214	0.624
2	1.612	0.404
3	2.020	0.319
4	1.382	0.278
5	1.711	0.265
6	2.214	0.490
7	1.509	0.352
8	1.844	0.295
9	2.467	0.273

**Table 4.  $M_m$  function of  $\lambda_M$  for  $l/r_m=0.0723$**