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A mixed-effects model with time reparametrization for longitudinal univariate manifold-valued data

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Abstract. Mixed-effects models provide a rich theoretical framework for the analysis of longitudinal data. However, when used to analyze or predict the progression of a neurodegenerative disease such as Alzheimer's disease, these models usually do not take into account the fact that subjects may be at different stages of disease progression and the interpretation of the model may depend on some implicit reference time. In this paper, we propose a generative statistical model for longitudinal data, described in a univariate Riemannian manifold setting, which estimates an average disease progression model, subject-specific *time shifts* and *acceleration factors*. The *time shifts* account for variability in age at disease-onset time. The *acceleration factors* account for variability in speed of disease progression. For a given individual, the estimated time shift and acceleration factor define an affine reparametrization of the average disease progression model. This statistical model has been used to analyze neuropsychological assessments scores and cortical thickness measurements from the Alzheimer's Disease Neuroimaging Initiative database. The numerical results showed that we can distinguish between slow versus fast progressing and early versus late-onset individuals.

1 Introduction

The Alzheimer's disease (AD), and most neurodegenerative diseases alike, is a slowly progressive neurodegenerative disorder and a growing public health issue as the number of people diagnosed with AD is steadily increasing over the years, partly due to an increasing ageing population. The progressive loss of cognitive functions, such as memory, language and reasoning, are the observable consequences of pathological processes which started to affect the brain several years before. In [1], Clifford R. Jack et al. proposed a widely accepted hypothetical model which describes the temporal evolution of the AD biomarkers in which the decline in cognitive functions and memory loss occur in the late stages of

the disease. The impairment of the cognitive functions is preceded by decreasing cerebrospinal fluid (CSF) amyloid- β_{42} ($A\beta_{1-42}$) peptide, increasing CSF tau protein and neurodegeneration, characterized by cerebral atrophy and neuron loss. In order to increase the chances of success of a given treatment, the need for an early diagnosis and the need to understand the progression of the disease are crucial.

Longitudinal data, which consists in repeated observations of a given individual or group of individuals over time, allows to study the progression of a disease and the statistical analysis of longitudinal data may provide useful informations to help diagnose AD in its early phase. Mixed effects models provide a rich framework to analyze longitudinal measurements. The fixed (and respectively random) effects allow to describe the model at the group (respectively subject) level and the distribution of the random effects accounts for inter-subject variability. A particular case of the linear mixed-effects models introduced by Laird and Ware in [5] can be used to model a linear relationship between univariate longitudinal measurement and the times of observation : the random slope and intercept model. If $y_{i,j}$ denotes the j -th observation of the i -th subject associated to time $t_{i,j}$, the random slope and intercept model would write : $y_{i,j} = (\bar{a} + a_i)(t_{i,j} - t_0) + (\bar{b} + b_i) + \varepsilon_{i,j}$, where $\varepsilon_{i,j}$ denotes a Gaussian noise with zero mean. The fixed effects \bar{a} and \bar{b} respectively correspond to the slope and intercept of an average trajectory $\bar{D}(t) = \bar{a}(t - t_0) + \bar{b}$ whereas t_0 plays the role of a reference time. The subject-specific slope a_i and b_i correspond to the slope and intercept of the trajectory of the i -th individual.

If the longitudinal observations arise, for example, from developmental studies, breeding studies or pharmacological studies, there is a natural choice for the baseline time t_0 , which may represent the time of birth, beginning of the study or time at which a certain drug has been administered. In those situations, the baseline time is known and is therefore not estimated along with the parameters of the model. However, in studies on neurodegenerative diseases, such as Alzheimer's disease, there may be no natural choice for the reference time. As pointed out by Yang et al. in [4] and Delor et al. in [3], the disease-onset time is probably different for every individual and, at a given age, two individuals may be at very different stages of disease progression. When the baseline (or disease-onset) time is unknown, one may consider it as a fixed effect of the given model. In [2], the authors proposed an interesting mixed-effect model for scalar measurements but the model did not account for the difference in stages progression among individuals. Yang et al. [4] and Delor et al. [3] addressed this issue by including time shifts in their models but Yang et al. did not estimate the time shifts in a statistical framework. In Delor et al., the observations of a given individual were shifted in time before the estimation of a disease-onset time. In both situations, the reference time was not estimated. In [8], Singh et al. proposed an interesting model for longitudinal manifold-valued observations using parallel transport. However, if the model is written for univariate longitudinal data, the model does include a random effect to account for the variability in speed of disease progression.

When the time reference time t_0 is considered unknown, the linear model given above is not identifiable and there is a relation between the distribution of the intercept random effect b_i and the fixed effect t_0 , which makes the interpretation of the distribution of the intercept random effect b_i difficult.

To address the issues we pointed out, we propose to replace the intercept random effect b_i , which characterizes the distribution of the measurement values at time t_0 , by a time shift τ_i which measures the delay (or advance) in disease progression. We therefore consider the model : $y_{i,j} = v_0\alpha_i(t_{i,j} - t_0 - \tau_i) + \varepsilon_{i,j}$ where α_i is a random effect log-normally distributed with mean equal to 1 and τ_i is the time shift random effect normally distributed with mean 0. The fixed effect v_0 correspond to the slope of the mean trajectory and is the fixed effect associated to α_i whereas t_0 is now the fixed effect associated to the random effect τ_i . If D denotes the straight line $D(t) = v_0(t - t_0)$, the model writes : $y_{i,j} = D(\alpha_i(t_{i,j} - t_0 - \tau_i) + t_0) + \varepsilon_{i,j}$. As a consequence, the random effects α_i and τ_i allow to define a subject-specific affine reparametrization of the straight line D , which can be thought of as an average trajectory. The *acceleration factor* α_i and the time shift τ_i of the affine subject-specific reparametrization provide informations on the dynamical evolution of a given subject with respect to the average model: the *acceleration factor* will determine whether an individual is progressing faster than the individual trajectory. The *time shift* determines whether the individual is lagging behind or ahead of the average model. Therefore, the random effects of the model will allow to determine, among a group of individuals, who are fast or slow progressing individuals and whether a given individual might develop the disease earlier (early-onset) or later (late-onset) than the estimated reference time t_0 .

Moreover, since the subject-specific trajectories are affine reparametrizations of an average trajectory, this model can easily be generalized to longitudinal measurements on a Riemannian manifold. The average trajectory may be replaced with a geodesic of the manifold and the subject-specific trajectories will remain geodesics on the manifold. As a matter of fact, longitudinal measurements arising from neuroimaging, neuropsychological tests or clinical examinations may belong to nonlinear spaces. The scores to a neuropsychological test often belong to a bounded interval. Riemannian manifolds offer a very flexible framework which includes the previous examples. The generalization of this statistical model to Riemannian manifold-valued observations could be used either to analyze shapes, cortical thickness measurements or neuropsychological assessment scores. After a proper renormalization, the ADAS-Cog 13 scores belong to the open unit interval $(0, 1)$ which, equipped with a specific metric, can be seen as a univariate Riemannian manifold. For this metric, the geodesics are logistic curves. The models presented in [2] and [3] do not generalize to manifold-valued data. In [9], Lorenzi et al. used Riemannian manifold techniques to estimate a model of the brain's normal ageing from healthy individuals T1 MR scans. The model was used to compute a time shift, called *morphological age shift*, which corresponds to the actual anatomical age of the subject with respect to an estimated average age for healthy subjects. However, the subject-specific time shifts

were not estimated as parameters of a statistical model. In [7], Datar et al. generalized a particular linear mixed-effects model to longitudinal shape analysis but the proposed model did not include time shifts to account for the variability in stages of disease progression among the population. Even though the model is used to analyze shapes (given by a set of points), the model is not described in a Riemannian manifold setting.

The work presented herein will generalize the model written above to a Riemannian manifold included in \mathbb{R} . In Section 2, we will introduce our model and we will present two particular cases of this model : the *straight lines model* and the *logistic curves model*. We also explain how the parameters of the model are estimated. In Section 3, the straight lines model and logistic curves model were used to analyze longitudinal ADAS-Cog and cortical thickness measurements from the ADNI database. We show that the estimated random effects of the model allowed to distinguish between slow versus fast progressing individuals and early versus late-onset individuals.

2 A mixed-effects model with time reparametrization for manifold-valued observations

2.1 Model description

Let us assume that we observe p different individuals. For each individual, we have n_i observations, obtained at times $t_{i,1} < \dots < t_{i,n_i}$ and the observations of the i -th individual are denoted by $y_{i,1}, \dots, y_{i,n_i}$. We assume that each observation $y_{i,j}$ is a point on a one-dimensional Riemannian manifold (M, g) included in \mathbb{R} . In addition to this, we assume that the geodesics of M are defined on the entire real line.

Let p_0 be a point in M , t_0 in \mathbb{R} and $v_0 \in T_{p_0}M$, a tangent vector to M at the point p_0 . The triplet (p_0, t_0, v_0) allows to define an *average trajectory* γ_{p_0, t_0, v_0} , the geodesic of M which passes through the point $p_0 \in M$ at time t_0 and with the velocity $v_0 \in T_{p_0}M \simeq \mathbb{R}$. This average trajectory summarises the progression of all the individuals and could be interpreted as the evolution of an “average subject”. While the average subject passes through the point p_0 at time t_0 with the velocity v_0 , we assume that the trajectory of the i -th subject is the geodesic γ_i which passes through the point p_0 at time $t_0 + \tau_i$ with velocity $\alpha_i v_0$. The parameters (p_0, t_0, v_0) will be the fixed effects of the model whereas α_i and τ_i will be the random effects of the model. The point p_0 in M can be understood as an *observation point*. A simplistic interpretation of the role played by α_i and τ_i is the following : an observer standing at the point p_0 in M and aware that the average subject passes through the point at the time t_0 , with velocity v_0 , will be able to tell whether the i -th subject passes through this point earlier or later than the average subject and whether he passes faster or slower. More precisely, τ_i is a time shift which will provide an information on whether the progression of the i -th trajectory is ahead of the average trajectory or not

and α_i provides the information on whether the i -th subject is evolving faster or slower than the average subject. As a consequence, the random effects α_i and τ_i take into account the fact that two different subjects may be at very different stages of disease progression and allow for inter-subjects comparisons. The description above makes the assumption that every subject-specific trajectory of disease progression will eventually reach a common value p_0 . This assumption does not make sense in a higher-dimensional Riemannian manifold. In order to generalize the proposed model to a higher-dimensional setting, one would need to add another random effect, associated to the fixed effect p_0 .

For a point $p_0 \in M$ and $t_0 \in \mathbb{R}$, $v_0 \in T_{p_0}M$, we define $\gamma_{p_0, t_0, v_0} := \text{Exp}_{t_0, p_0}(\alpha_i v_0)(\cdot)$ as the geodesic which passes through the point p_0 at time t_0 with velocity v_0 : γ_{p_0, v_0, t_0} is the unique curve, drawn on M and defined on \mathbb{R} , which satisfies $\nabla_{\dot{\gamma}_{p_0, t_0, v_0}} \dot{\gamma}_{p_0, t_0, v_0} = 0$ and $\gamma_{p_0, t_0, v_0}(t_0) = p_0$, $\dot{\gamma}_{p_0, t_0, v_0}(t_0) = v_0$. For the dataset $(t_{i,j}, y_{i,j})$ ($1 \leq i \leq p$ and $1 \leq j \leq n_i$), we assume the following model :

$$y_{i,j} = \text{Exp}_{t_0 + \tau_i, p_0}(\alpha_i v_0)(t_{i,j}) + \varepsilon_{i,j} \quad (1)$$

where $\alpha_i = \exp(\eta_i)$ and :

$$\begin{cases} \eta_i \sim \otimes_{i=1}^p \mathcal{N}(0, \sigma_\eta^2) \\ \tau_i \sim \otimes_{i=1}^p \mathcal{N}(0, \sigma_\tau^2) \\ \varepsilon_{i,j} \sim \otimes_{i,j} \mathcal{N}(0, \sigma^2). \end{cases}$$

The random variables $(\eta_i)_{1 \leq i \leq p}$ and $(\tau_i)_{1 \leq i \leq p}$ are assumed independent. The vector $\theta = (p_0, t_0, v_0, \sigma_\eta, \sigma_\tau, \sigma)^\top$ is the vector of the parameters of the model. We will discuss later how this vector is estimated from the data.

An important feature of the model (1) is that it can be written as follows :

$$y_{i,j} = \gamma_{p_0, t_0, v_0}(\alpha_i v_0(t_{i,j} - t_0 - \tau_i)) + \varepsilon_{i,j} \quad (2)$$

$$= \gamma_i(t_{i,j}) + \varepsilon_{i,j}. \quad (3)$$

The subject-specific trajectory of the i -th subject, the geodesic γ_i , therefore appears as an affine reparametrization of the average trajectory γ_{p_0, t_0, v_0} . As a consequence, in order to compare the progression of a given individual with respect to the average trajectory (respectively compare two given subjects together), one only need to compare the affine reparametrization $t \mapsto \alpha_i v_0(t - t_0 - \tau_i)$ with the identity map $t \mapsto t$ (respectively, compare the slopes and intercepts of the two subject-specific affine reparametrization). We can observe that these reparametrization play the same role as the *time warps* which were introduced by Durrleman et al. in [11] and [10], in the context of longitudinal shape analysis. The time warps considered in those papers were diffeomorphisms of the real line which accounted for the different rates of shape changes and were constructed using the LDDMM framework : the diffeomorphism was obtained by integration of a time-varying vector field which belongs to a Gaussian reproducing kernel Hilbert space. The affine reparametrizations (with non-zero intercept) are also diffeomorphisms but only model a constant rate of progression. However, with

only two random effects, we are able to describe the progression of a given subject in a realistic and easily interpretable manner.

Considering p_0 as a fixed effect of the model and estimating p_0 along with the other fixed effects ensures that the estimated p_0 will be the best observation point possible. It should be noted that the model does not depend on a reference time because if the observations $(t_{i,j}, y_{i,j})$ are transformed into $(t_{i,j} + t'_0, y_{i,j})$ then, only the estimated value of t_0 will change whereas the variance parameters will remain the same. Therefore, the interpretation of the random effects (and between-subjects or subjects-average comparisons) do not depend on the timeline.

2.2 The straight lines model

In this section, we will illustrate the flexibility of the model we introduced in the previous section. The choice of the Riemannian metric on M determines the shape of the geodesics of M . By taking the canonical metric on \mathbb{R} , the geodesics will be straight lines.

When $M = \mathbb{R}$ is equipped with the canonical metric, M is a geodesically complete Riemannian manifold and the geodesics of M are of the form $t \mapsto p + tv$. In this case, (1) writes :

$$y_{i,j} = p_0 + \alpha_i v_0 (t_{i,j} - t_0 - \tau_i) + \varepsilon_{i,j}. \quad (4)$$

Even though the model is called “linear”, (4) does not belong to the class of linear mixed-effects models introduced in [5].

2.3 The logistic curves model

The situation when $M = (0, 1)$ may arise when the observations are naturally bounded. For example, neuropsychological assessments such as ADAS-Cog or MMSE produce positive scores which are bounded above by a maximum score. The 13 questions ADAS-Cog assessment test is marked out of 85 points. Therefore, a typical ADAS-Cog score is a number in $[0, 85]$. By a proper renormalization, the score values can be considered as points in the interval $(0, 1)$. Fitting a linear model to such observations might not be such a good idea since the fitted trajectories will not always remain in the domain $(0, 1)$. On the other hand, $(0, 1)$ corresponds to the range of logistic curves with asymptote equal to 1. In order to equip the open interval $(0, 1)$ with a Riemannian manifold structure, we must give a Riemannian metric. At a point $p_0 \in (0, 1)$, the tangent space $T_{p_0}M$ is isomorphic to \mathbb{R} . For all p_0 in M , we define g_{p_0} to be the following inner product : $\forall (u, v) \in T_{p_0}M, g_{p_0}(u, v) = uM(p_0)v$ where $M(p_0) = \frac{1}{p_0^2(1-p_0)^2}$. The mapping $p_0 \mapsto g_{p_0}$ is smooth on M and defines a Riemannian metric on M . In this particular setting, the geodesic equation $\nabla_{\dot{\gamma}}\dot{\gamma} = 0$ writes : $\ddot{\gamma}(t) + \Gamma_{1,1}^1(\gamma(t))(\dot{\gamma}(t))^2 = 0$ where $\Gamma_{1,1}^1$ is the Christoffel symbol determined by the only coefficient $g_{1,1}(t) = M(t)$ of the metric. The

geodesic equation writes : $\ddot{\gamma}(t) - \left(\frac{1}{\gamma(t)} + \frac{1}{\gamma(t)-1}\right)(\dot{\gamma}(t))^2 = 0$, or equivalently : $\gamma(\gamma-1)\ddot{\gamma} - (2\gamma-1)(\dot{\gamma})^2 = 0$. Solving for this differential equation allows to determine that the geodesics of M are logistic curves of the form $t \mapsto \frac{1}{1+a\exp(-rt)}$ with $a > 0$ and $r \in \mathbb{R}$. The geodesics are defined on \mathbb{R} and therefore ensure that M is a geodesically complete Riemannian manifold. In this case, (1) writes :

$$y_{i,j} = \frac{1}{1 + \left(\frac{1}{p_0} - 1\right) \exp\left(-\frac{\alpha_i v_0}{p_0(1-p_0)}(t_{i,j} - t_0 - \tau_i)\right)} + \varepsilon_{i,j}. \quad (5)$$

2.4 Estimation of the parameters of the model

The statistical model described in (1) is a nonlinear mixed-effects model with gaussian random effects. Such models have been extensively studied in the literature and several methods exist for maximum likelihood estimation. The estimation of the parameters $(p_0, t_0, v_0, \sigma_\eta, \sigma_\tau, \sigma)$ was coded into the SAS software (SAS Institute) using the Gauss-Hermite quadrature approximation with one quadrature point. This quadrature approximation method is equivalent to the Laplace approximation of the observed likelihood. During the estimation procedure, the objective function resulting from the Laplace approximation was minimized using the Newton-Raphson algorithm or the Nelder-Mead simplex algorithm. It should be mentioned that these estimation procedures were quite sensitive to the choice of starting values. However we observed that this method was more efficient and more robust than the Alternating Algorithm, proposed by Lindstrom and Bates in [13] or Pinheiro and Bates in [12], and implemented in other softwares.

3 Results

3.1 Data

We used data from the Alzheimer's Disease Neuroimaging Initiative (ADNI) database [15]. The ADAS-Cog scores we used were collected for 1391 individuals enrolled in the ADNI1, ADNIGO and ADNI2 phases and cortical thickness measurements for 725 individuals from the ADNI1 database. The longitudinal follow-up of these individuals ranges from 18 months to 4 years. Diagnoses were recorded for every individual and at each visit. These subject-specific sequences of diagnoses allowed to classify the individuals into 4 groups of interest : *stable controls* (329 individuals), *stable mild cognitive impairment*, also denoted as *stable MCI* (471 individuals), *stable Alzheimer patients* (248 individuals) and *converters MCI* (282 subjects). Cortical thickness measurements were available for 194 stable controls, 182 stable MCI, 170 converters MCI and 162 Alzheimer patients. The individuals who reverted to control or MCI were not included in these groups.

3.2 Experimental results

The ADAS-Cog scores were normalized by 85, which corresponds to the maximum possible score, and analyzed using the logistic curves model. For each individual, a couple of random variables (α_i, τ_i) where $\alpha_i = \exp(\eta_i)$ were estimated along with the parameters of the model. We recall that α_i (and respectively τ_i) account for the change in speed of disease progression (and respectively delay in disease progression) of a given individual with respect to the average disease progression trajectory. The couples (η_i, τ_i) have been plotted in figure 1.

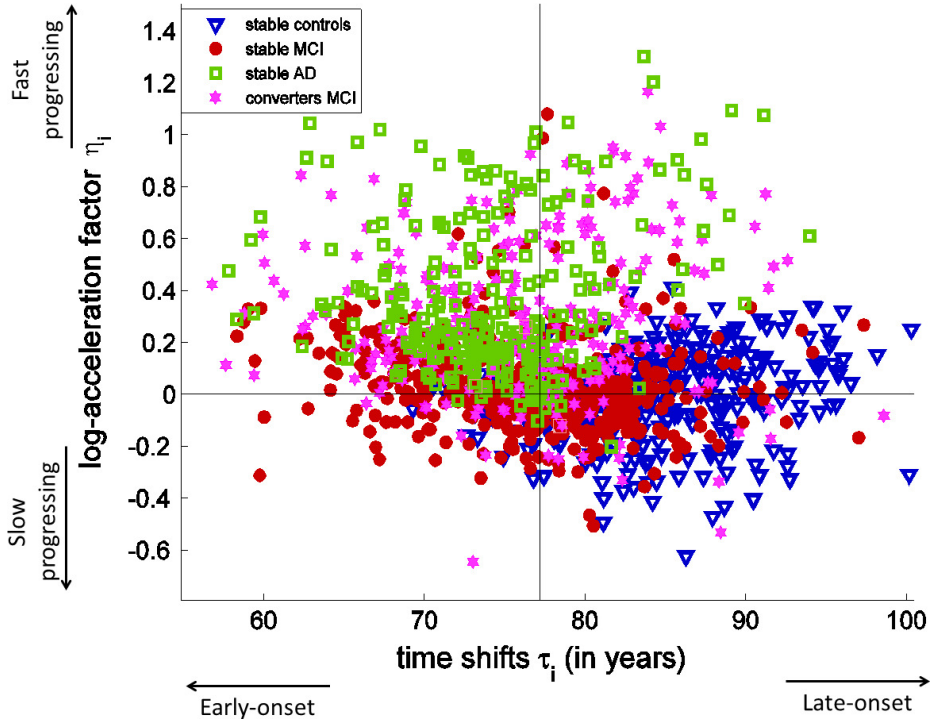


Fig. 1. Log-acceleration factors η_i plotted against the time shifts τ_i for the 1391 individuals with ADAS-Cog measurements. An horizontal line was plotted at the level $\eta_i = 0$ (no change in speed with respect to the average trajectory) and at $\tau_i = t_0 = 77.17$ (the estimated reference time t_0).

The results presented in figure 1 allow to compare the informations provided by (η_i, τ_i) across the groups of interest we previously mentioned. We can observe that stable controls have larger time shifts than other groups. The stable controls can be considered as late-onset individuals who are not, on average, evolving faster than the average disease progression trajectory. On the other hand, stable Alzheimer patients and MCI individuals tend to have smaller time

shifts than stable controls. A portion of the stable MCI and most of the stable Alzheimer patients can be considered as early-onset individuals. It appears clearly that stable Alzheimer patients and converters MCI are fast progressing individuals even though a small number of converters MCI are slow-progressers. On average, stable MCI and stable controls are not progressing faster than the average disease progression scenario. These observations are coherent with the diagnoses of the individuals and the subject-specific random effects allow to distinguish groups of individuals which are coherent with their diagnoses. The disease progression for Alzheimer patients and converters MCI is faster and started earlier than for stable controls. The Alzheimer patients are quite clearly separated from the stable controls. In the model, the random effects η_i and τ_i are assumed independent, therefore uncorrelated. For this experiment, we computed the estimated covariance matrix of the random effects. For stable controls, we find that $\text{Corr}(\eta_i, \tau_i) = 0.20$, for stable MCI : $\text{Corr}(\eta_i, \tau_i) = -0.17$, for stable AD : $\text{Corr}(\eta_i, \tau_i) = 0.12$ and for converters MCI : $\text{Corr}(\eta_i, \tau_i) = 0.18$. These results seem coherent with the shape of the point clouds in the log-acceleration-time-shift plane, especially for stable controls and stable MCI.

The cortical thickness measurements (computed with Freesurfer) were averaged within 34 regions of interest defined with the Desikan-Killiany [19] cortical atlas. The averaged cortical thickness measurements were analyzed using the straight lines model. For each subject, time shifts and acceleration factors were estimated and the corresponding values were displayed on the cortical surface. The differences in the estimated time shifts and acceleration factors were compared between Alzheimer patients and controls (figure 2), and between converters MCI and stable MCI (figure 3). Significance level was set at 0.05, corrected for multiple comparisons using Bonferroni correction. Alzheimer patients present accelerated gray matter loss compared to stable controls in a large number of regions, with highest speed in temporal (including entorhinal cortex, parahippocampal gyrus, superior and middle temporal gyri), parietal associative (including precuneus) and frontal regions. A similar topographical pattern was found for converters MCI compared to stable MCI but with smaller accelerations than in AD patients. On the contrary, primary motor and sensitive as well as visual cortices were spared. These results are highly consistent with the spatial-temporal progression patterns of neurodegeneration evidenced in histopathological studies [16], [17]. Furthermore, accelerated atrophy has also been recently shown to coincide with disease-onset [18]. On the other hand, the estimated time shifts were not significantly different for the vast majority of regions. In the few significant regions, the magnitude of the time-shifts was small. This is in contrast with the large time-shift values found for the cognitive variables. This can be attributed to the slow but consistent age-related atrophy that is present in control subjects.

4 Discussion and perspectives

We proposed a statistical model which allows to model the progression of a neurodegenerative disease for a group of individuals. The power of our approach

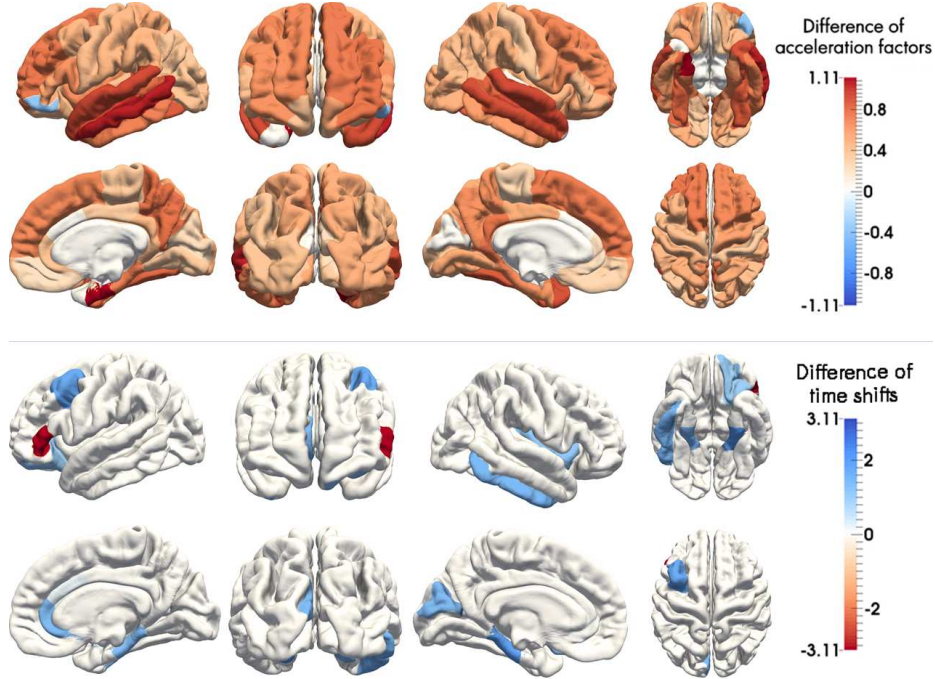


Fig. 2. At the top (respectively bottom) : the difference in averaged acceleration factors (respectively time shifts) between AD patients and stable controls is displayed on the cortex. Acceleration factors (and respectively time shifts) were averaged per regions of interest. Only regions where the difference was statistically significant ($p < 0.05$, corrected for multiple comparisons) were colored.

lies in the use of subject-specific time shifts and acceleration factors which appear as random effects of our statistical model. The acceleration factor and time shift associated to a given individual allow to define the subject-specific trajectory of disease progression as an affine reparametrization of the average disease-progression trajectory. As a consequence, the progression of a given individual can be easily compared to the estimated average scenario and two individuals can be compared together through the comparisons of their corresponding affine reparametrizations. These subject-specific time shift account for delay (or advance) in disease progression. Using the logistic curves model and the straight lines model, we showed that we were able to distinguish between fast versus slow progressing individuals and early versus late onset individuals. However, our experimental results were based on the a priori knowledge of clinical diagnostic informations and the statistical model we introduced does not allow to determine the sequence in which biomarkers reach pathological levels. Fonteijn et al. proposed such an event-based model in [14]. Finally, our model was described in a Riemannian manifold setting which ensures that the model can be applied to

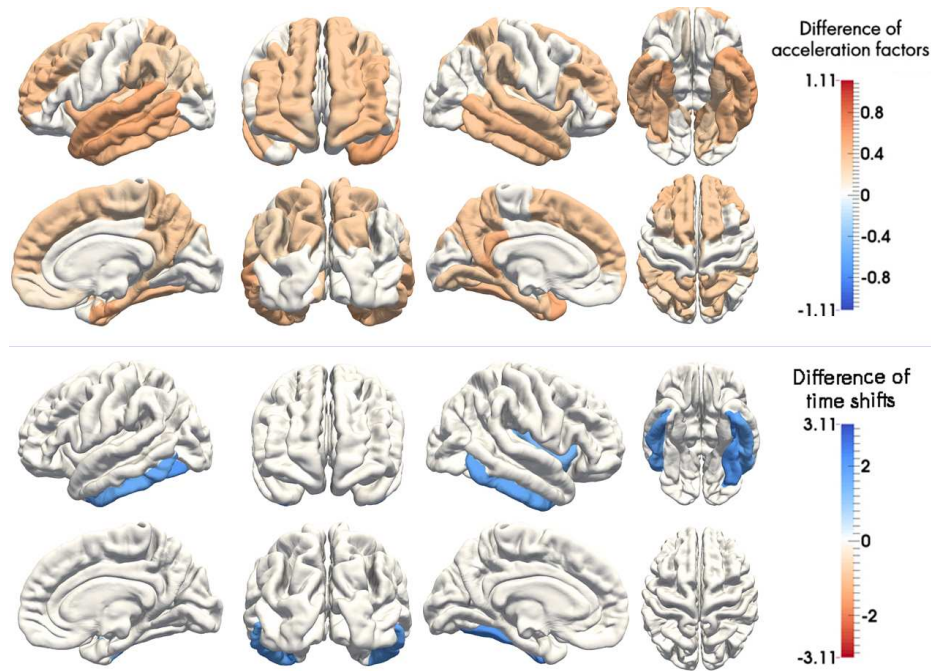


Fig. 3. At the top (respectively bottom) : the difference in averaged acceleration factors (respectively time shifts) between converters MCI and stable MCI is displayed on the cortex. Acceleration factors (and respectively time shifts) were averaged per regions of interest. Only regions where the difference was statistically significant ($p < 0.05$, corrected for multiple comparisons) were colored.

data of varying complexity. This also paves the way for a generalization of this model to a multivariate framework.

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