



HAL
open science

Support indices: Measuring the effect of input variables over their supports

J Fruth, O Roustant, S Kuhnt

► **To cite this version:**

J Fruth, O Roustant, S Kuhnt. Support indices: Measuring the effect of input variables over their supports. 2016. hal-01113555v3

HAL Id: hal-01113555

<https://hal.science/hal-01113555v3>

Preprint submitted on 13 Nov 2016 (v3), last revised 26 Sep 2018 (v4)

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

SUPPORT INDICES: MEASURING THE EFFECT OF INPUT VARIABLES OVER THEIR SUPPORTS

Jana Fruth⁽¹⁾ and Olivier Roustant⁽²⁾ and Sonja Kuhnt⁽³⁾

(1) TU Dortmund University, Germany

(2) Mines Saint-Etienne, UMR CNRS 6158, LIMOS, F42023 Saint-Etienne, France

(3) Dortmund University of Applied Sciences and Arts, Germany

Abstract: Two new sensitivity indices are presented which give an original solution to the question in sensitivity analysis of how to determine regions within the input space for which the model variation is high. The indices, as functions over the input domain, give insight into the local influence of input variables over the whole domain when the other variables lie in the global domain. They can serve as an informative extension to a standard analysis and in addition are especially helpful in the specification of the input domain, a critical, but often vaguely handled issue in sensitivity analysis. In the usual framework of independent continuous input variables, we present theoretical results that show an asymptotic connection between the presented indices and Sobol' indices, valid for general probability distribution functions. Finally, we show how the indices can be successfully applied on analytical examples and on a real application.

Key words and phrases: Sensitivity analysis, support analysis, Sobol' indices, DGSM

1. Introduction

Global sensitivity analysis is an essential tool in many of the steps in the analysis of computer experiments, like screening, investigation, model building, or optimization (Saltelli et al., 2000). Variance-based indices have become the method of choice among practitioners, as they offer a clear interpretable quantification of the influence of variables on the output variance. Often a surrogate or meta-model of the computer experiment is build first and indices are derived from this model. In general, different types of variance-based indices have been proposed in the literature. The single influence of a variable is e.g. measured by the first-order Sobol' index (Sobol', 1993). The total sensitivity index (Homma and Saltelli, 1996) quantifies the joint influence of a variable and all its interactions and thus can be used to detect unessential variables. A screening of interactions can be obtained by the total interaction index (Muehlenstaedt et al., 2012; Fruth et al., 2014), which captures the total impact of the second-order interaction of a pair of variables.

An alternative concept to global sensitivity are derivative-based global indices (DGSM) (Sobol' and Gershman, 1995; Kucherenko et al., 2009). The indices use the integrated square derivative of the model output over the domain of the inputs as indicator of the input influence over the output

and, like total sensitivity indices, can be used to detect unessential variables.

For the estimation of both, variance-based indices and DGSM, input variables are usually treated as random variables following a distribution defined by the user. The minimum and maximum value of the support of this distribution is specified from the region of interest of the parameter in question. They are often chosen vaguely, but can have an important impact on the sensitivity of the response.

As an example, let us consider the Ishigami function by Ishigami and Homma (1990),

$$f(x_1, x_2, x_3) = \sin(x_1) + 7 \sin^2(x_2) + 0.1x_3^4 \sin(x_1).$$

This function is often taken as a benchmark in sensitivity analysis due to its strong non-linearities and interactions. The usual setting considers independent variables X_1, X_2, X_3 distributed according to uniform law on $[-\pi, \pi]$, and investigates the Sobol' indices of $f(X_1, X_2, X_3)$. A direct computation (see e.g. Saltelli et al., 2000) shows that the unnormalized total sensitivity indices relative to X_i , denoted by D_i^T , of the Ishigami function are equal to $D_1^T = 7.72$, $D_2^T = 6.13$, and $D_3^T = 3.37$. Now, if we reduce the support of the distribution of X_3 by 10 percent, $X_3 \sim U[-\pi + \frac{\pi}{10}, \pi - \frac{\pi}{10}]$, we get $D_1^T = 4.05$, $D_2^T = 6.13$, and $D_3^T = 1.45$. The first and the last variable have rapidly lost influence in comparison to the second one, the ranking of

variables has changed. This can be explained by the interaction between X_1 and X_3 , which is much stronger at the borders than in the rest of the function. The change in the output when changing X_2 stays the same as X_2 does not interact with X_3 .

In their book on sensitivity analysis, Saltelli et al. (2000) mention as one goal of sensitivity analysis “to determine the optimal regions within the space of the factors for which the model variation is maximum”. We present a method which performs this task by returning the local influences over the whole domain. We extend two scalar indices, namely Sobol’ and total sensitivity index. They explore the local sensitivity behavior over the support of the input variables and thus show the impact of the choice of the input distribution.

Our new method is related to the idea of *local sensitivities* see e.g. Turanyi (1990) for general review on different types of sensitivities. There, local sensitivities are defined as partial derivatives at fixed base-line points. Our method shares with them that partial derivatives are used to obtain local influences. The main difference is that in local sensitivity analysis, the influence of a variable is studied in a local space of the other inputs whereas our method studies the local influence of a variable in a global space of the other variables. The local influences are studied over the whole

input space, so no space is ignored. Welch et al. (1992) suggest to plot curves of the *main effects*, i.e. the expectation of the model conditional on each input, in a framework with a Gaussian process metamodel. They argue that these plots give a visual indication of the effects and provide insight into their relative magnitude. Our idea shares these aims, but gives two different local sensitivities corresponding to first-order Sobol' indices and total indices, which has to our knowledge not been done before.

We next discuss some recent developments towards measuring local or regional sensitivity. Rakovec et al. (2014) suggest a hybrid method which looks at global as well as local sensitivity. They base their approach on gradients at multiple points in the input space. Li et al. (2014) and Wei et al. (2014) consider measures of regional importance which evaluate the impact of inputs with respect to reduced ranges of the individual input space while keeping the input space of the remaining variables global. Similar to our method, this results in sensitivity functions over the input space. However, as their measures are not functions directly on the input space but on either a quantile value or the boundaries of the considered interval, we find our new sensitivity indices and plots more intuitive and informative. The approach of Pannier and Graf (2015) follows the same general ideas as our method. They subdivide the input space into a finite number of

subdomains and calculate derivative based global sensitivity measures on these subdomains. Our method goes a step further as it can be seen as reducing the discrete subdomains to points of the input space.

With our method, the effects of the choice of the input support do not only become visible, they are also quantified. It gives insight into the function by presenting a *function of sensitivities* as extension to Sobol' and total sensitivity indices. It can further be a helpful tool in sequential optimization, where the input behavior of certain optimal regions is of interest, which might differ extensively from the global behavior.

After an overview on global sensitivity measures in Section 2, we define the support index functions and show that they are connected to classical indices when the variable in question is restricted to an interval whose size goes to zero (Section 3). Section 4 deals with the estimation of the new support indices. To illustrate the method, it is applied to two toy functions and a real life application in sheet metal forming in Section 5. Conclusions and perspective are given in Section 6 followed by a Section 7 on proofs.

2. Overview of global sensitivity measures

We consider a black-box function $Y = f(X)$ of independent random variables $\mathbf{X} = (X_1, \dots, X_d)$ with distribution $\mu = \mu_1 \otimes \dots \otimes \mu_d$. The function f is defined on $\Delta \subseteq \mathbb{R}^d \rightarrow \mathbb{R}$ such that $f(\mathbf{X}) \in L^2(\mu)$.

2.1 Variance-based sensitivity measures

The FANOVA decomposition (Hoeffding, 1948; Sobol', 1993) gives a decomposition of the function f into additive terms which are centered and orthogonal:

$$f(\mathbf{X}) = f_0 + \sum_{i=1}^d f_i(X_i) + \sum_{1 \leq i < j \leq d} f_{i,j}(X_i, X_j) + \cdots + f_{1,\dots,d}(X_1, \dots, X_d). \quad (2.1)$$

The terms represent first-order effects ($f_i(X_i)$), second-order interactions ($f_{i,j}(X_i, X_j)$), and all higher combinations of input variables. From the orthogonality in Eq. (2.1), the variance of f can be decomposed accordingly,

$$\text{var}(f(\mathbf{X})) = \sum_{I \subseteq \{1, \dots, d\}} \text{var}(f_I(\mathbf{X}_I)), \quad (2.2)$$

where the overall variance is denoted by $D = \text{var}(f(\mathbf{X}))$, and the variance terms by $D_I = \text{var}(f_I(\mathbf{X}_I))$. These terms quantify the influence of the set of variables X_I on the output variance. The Sobol' sensitivity indices S_I are defined as the ratio of variance explained by X_I divided by the overall variance, ensuring the indices to fall in $[0, 1]$ and sum to one

$$S_I = \frac{D_I}{D}. \quad (2.3)$$

In a sensitivity analysis procedure, the first-order Sobol' indices $S_i, i = 1, \dots, d$ are usually estimated to quantify the single influences of the input variables. In addition, often the unnormalized and normalized total

sensitivity indices

$$D_i^T = \sum_{J \ni \{i\}} D_J \quad \text{and} \quad S_i^T = \frac{D_i^T}{D}. \quad (2.4)$$

are estimated, which give the influence of each variable including all interactions and thus are suitable for variable screening.

2.2 Derivative-based measures

A further sensitivity measure, the *derivative-based global sensitivity measures (DGSM)*, was introduced by Sobol' and Gershman (1995) and studied in Kucherenko et al. (2009). Here, the notion of sensitivity is based on the derivative of the function, rather than on the output variance. It thus summarizes the local variation instead of the global one. The DGSM are defined by

$$\nu_i = \int \left(\frac{\partial f(\mathbf{X})}{\partial x_i} \right)^2 d\mu(\mathbf{X}). \quad (2.5)$$

It is often convenient to consider the corresponding weights. The normalized DGSM are thus defined by

$$\nu_i^* = \frac{\nu_i}{\sum_{j=1}^d \nu_j}$$

Contrarily to variance-based indices, DGSM are not associated to a functional decomposition. However they can be connected to total sensitivity indices through

$$D_i^T \leq C(\mu_i)\nu_i, \quad (2.6)$$

provided that μ belongs to a class of distributions that satisfy a Poincaré inequality,

$$\int g(\mathbf{X})^2 d\mu(\mathbf{X}) \leq C(\mu) \int \|\nabla g(\mathbf{X})\|^2 d\mu(\mathbf{X}),$$

for all functions g in $L^2(\mu)$ such that $\int g(\mathbf{X}) d\mu(\mathbf{X}) = 0$, and $\|\nabla g\| \in L^2(\mu)$, where ∇g denotes the gradient of g and $C(\mu_i) > 0$ constant.

The physical interpretation is that the energy of a centered function, measured by the $L^2(\mu)$ norm, is controlled by the energy of its derivative. This explains the Poincaré-type inequality between total Sobol' index and DGSM. Indeed, denote by $g_i^T(\mathbf{x}) := \sum_{J \supseteq \{i\}} g_J(\mathbf{x}_J)$ the part of $g(\mathbf{x})$ involving x_i . Then one can check that the energy of $g_i^T(\mathbf{x})$ is equal to the total Sobol' index D_i^T and the energy of its derivative to the DGSM ν_i . See Lamboni et al. (2013) or Roustant et al. (2014) for a comprehensive summary. Log concave probability measures fall under this class, thereby most well-known distributions as e.g. uniform, normal and gamma distribution are covered (Lamboni et al., 2013). Equation (2.6) implies that the DGSM can be used as upper boundaries for the total sensitivity indices. It can thus serve as a cheaper method for the screening of unessential input variables, especially when the computer experiment additionally provides partial derivatives.

3. Support index functions

We now derive new sensitivity indices to quantify the changing effect of the input variables over their support. In addition to the previous notation, we assume that $\Delta = [0, 1]^d$, f is of class C^1 , and that X_1, \dots, X_d have continuous density functions with support $[0, 1]$. We will also denote $\mathbf{X}_{-i} = (X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d)$, and for any function g and any symbol t , the notation $g(t, \mathbf{X}_{-i})$ will assume that t is at position i . We are searching for sensitivity measures that quantify the local influence of one variable, including or not its interactions, in the global space of the others. Therefore we restrict X_i to vary only over a small interval around t and let the size of the interval go to zero. Proposition 1 summarizes the asymptotic behaviour of the overall variance, the Sobol' index and the total sensitivity index of one variable, when the interval size of this variable goes to zero.

Proposition 1. Let X_1, \dots, X_d be independent random variables with continuous density functions with support $[0, 1]$. Let f be a function defined on $\Delta = [0, 1]^d$ compact and assume that f is of class C^2 , ensuring that $f(\mathbf{X})$, $\frac{\partial f}{\partial x_\bullet}(\mathbf{X})$, $\frac{\partial^2 f}{\partial x_\bullet \partial x_\bullet}(\mathbf{X})$ are in L^2 . Let one of the random variables, say X_i , be restricted to $X_i^h \sim X_i | X_i \in [t - h/2, t + h/2]$ for $t \in]0, 1[$ and $h > 0$. Then we have that

$$\frac{\text{var}(X_i^h - t)}{h^2} \xrightarrow{h \rightarrow 0} \frac{1}{12},$$

and for the limits of the overall variance, the Sobol' index of X_i^h , denoted by $D_{X_i^h}$, and the total sensitivity index of X_i^h , denoted by $D_{X_i^h}^T$, it holds that

1. $\text{var}(f(X_i^h, \mathbf{X}_{-i})) \xrightarrow{h \rightarrow 0} \text{var}[f(\mathbf{X})|X_i = t]$,
2. $\frac{D_{X_i^h}}{h^2} = \frac{\text{var}(\mathbb{E}[f(X_i^h, \mathbf{X}_{-i})|X_i^h])}{h^2} \xrightarrow{h \rightarrow 0} \frac{1}{12} \left(\mathbb{E} \left[\left. \frac{\partial f}{\partial x_i}(\mathbf{X}) \right| X_i = t \right] \right)^2$,
3. $\frac{D_{X_i^h}^T}{h^2} = \frac{\mathbb{E}(\text{var}[f(X_i^h, \mathbf{X}_{-i})|\mathbf{X}_{-i}])}{h^2} \xrightarrow{h \rightarrow 0} \frac{1}{12} \mathbb{E} \left[\left. \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)^2 \right| X_i = t \right]$.

The first part of the proposition says that, locally, i.e. restricted to $[t - h/2, t + h/2]$, X_i has asymptotically the same variance than the uniform distribution, whatever its global distribution μ_i . Since that variance equals $h^2/12$, it makes a link to the second part of the proposition, where the term $h^2/12$ appears in the speed of convergence of $D_{X_i^h}, D_{X_i^h}^T$ when h tends to zero. Notice that the expression of the limits confirm that $D_{X_i^h}^T \geq D_{X_i^h}$, since the expectation of a square is always greater than the square of an expectation. We base the definition of our new support function on these findings.

Definition 1. Support index functions.

The *first-order support index* $\zeta_i(t)$ of $f(\mathbf{X})$ relative to X_i is the squared

conditional expectation of its partial derivative with respect to X_i :

$$\zeta_i(t) = \left(\mathbb{E} \left[\frac{\partial f}{\partial x_i}(\mathbf{X}) \middle| X_i = t \right] \right)^2.$$

The *total support index* $\zeta_i^T(t)$ of $f(\mathbf{X})$ relative to X_i is the conditional expectation of its squared partial derivative with respect to X_i :

$$\zeta_i^T(t) = \mathbb{E} \left[\left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)^2 \middle| X_i = t \right].$$

Taking the expected value over the support index functions reveals interesting connections to other indices, which give a little more insight into the interpretation of the functions. For the total support index, we immediately get equality to the DGSM index. On the other hand, the expected value over the first-order support index function can be connected to variance-based indices when we replace f as underlying function by the partial derivative of f with respect to x_i .

Proposition 2. 1. For the total support index $\zeta_i^T(\cdot)$ and the DGSM ν_i (see Eq. (2.5)) of a variable X_i , we have

$$\mathbb{E} [\zeta_i^T(X_i)] = \nu_i.$$

2. Consider the derivative function $\frac{\partial f}{\partial x_i}(\cdot)$ with global mean $E \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)$ and denote its first-order Sobol' index with respect to variable X_i by

$D_i \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)$, then:

$$E[\zeta_i(X_i)] = D_i \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right) + \left(E \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right) \right)^2.$$

Remark 1. Notice that the second point can be connected to Iooss et al. (2012), who also apply Sobol' indices of partial derivatives of f . They use the index of X_i of the derivative w.r.t to X_j as indicator for the interaction effect between X_i and X_j and the index of X_i of the derivative w.r.t to X_i as indicator of the nonlinear effect of X_i .

From the first point, it follows that the expected value of the total support index serves as an upper bound for the total sensitivity index,

$$D_i^T \leq C(\mu_i) E[\zeta_i^T(X_i)],$$

with $C(\mu_i)$ the Poincaré constant of μ_i as before.

Prop. 2 suggests a simple normalization of support indices. Indeed, it may be useful to compare the support indices both to the global value on the support and to other input variables. Furthermore, a same quantity should be used for first-order and total support indices. Hence, since $\zeta_i(t) \leq \zeta_i^T(t)$, we suggest to normalize $\zeta_i(t), \zeta_i^T(t)$ by the quantity $\sum_{j=1}^d \int \zeta_j^T(s) \mu_j(ds)$, which is equal to the sum of DGSMs (Prop. 2): $\sum_{j=1}^d \nu_j$.

Definition 2. Normalized support index functions.

Let *normalized first-order and total support index functions* relative to X_i be defined as

$$\zeta_i^*(t) = \frac{\zeta_i(t)}{\sum_{j=1}^d \nu_j}, \quad \zeta_i^{*T}(t) = \frac{\zeta_i^T(t)}{\sum_{j=1}^d \nu_j}$$

This normalization extends in our continuous setting the normalization proposed in Pannier and Graf (2015) for the sectional global sensitivity measure (Eq. 14), where distribution supports are divided in a finite number of intervals. However, contrarily to the discrete setting, the normalized support indices do not lie in $[0, 1]$, but only their global means do:

$$0 \leq \mathbb{E}(\zeta_i^*(X_i)) \leq \mathbb{E}(\zeta_i^{*T}(X_i)) \leq 1$$

Similarly, the expectation of normalized total support indices sum to 1:

$$\mathbb{E} \left(\sum_{i=1}^d \zeta_i^{*T}(X_i) \right) = 1$$

4. Estimation of support indices

In practice, we do not know f and have only access to a sample of input variables $X = (\mathbf{X}^{(n)})_{1 \leq n \leq N}$ and corresponding function evaluations $Y = (f(\mathbf{X}^{(n)}))_{1 \leq n \leq N}$. An important question is then to estimate the whole set of support indices $\zeta_i(\cdot), \zeta_i^T(\cdot)$ for $1 \leq i \leq d$, at a reasonable cost.

Up to square and square-root transformations, support indices can be viewed as regression curves of partial derivatives of $\frac{\partial f}{\partial x_i}(\mathbf{X})$ with respect to input variables X_i . Notice that the same input variable X_i serves as a reference both for differentiation and conditioning. Estimators can then be defined on the diagonal scatterplots $\left(X_i, \frac{\partial f(\mathbf{X})}{\partial x_i}\right)$ of the gradient of f , relying on the usual estimators for conditional estimations, based on smoothing or non-parametric estimation (see e.g. Hastie and Tibshirani, 1990). In the sequel we denote by $\widehat{m}_{X,Y}(t)$ such an estimator of $E(Y|X = t)$. We now give an algorithm for support indices estimation.

For $i = 1, \dots, d$, do:

1. Compute finite differences of f at sample points: Given a small $h > 0$, compute the vector $\mathbf{Z}_i = \left(\frac{f(X_1^{(n)}, \dots, X_i^{(n)} + h, \dots, X_d^{(n)}) - f(\mathbf{X}^{(n)})}{h} \right)_{1 \leq n \leq N}$
2. Compute the estimators for support indices from conditional expectation estimators: $\widehat{\zeta}_i(t) = (\widehat{m}_{\mathbf{X}_i, \mathbf{Z}_i}(t))^2$, $\widehat{\zeta}_i^T(t) = \widehat{m}_{\mathbf{X}_i, (\mathbf{Z}_i)^2}(t)$.
3. Compute the empirical estimate of DGSMs: $\widehat{\nu}_i = \overline{\mathbf{Z}_i^2}$

Finally, deduce the estimates for normalized support indices by dividing $\widehat{\zeta}_i(t), \widehat{\zeta}_i^T(t)$ by $\sum_{j=1}^d \widehat{\nu}_j$, for $i = 1, \dots, d$.

The first step can be skipped if the gradient is supplied at the same time as the function is evaluated. Assuming that the most important cost is function evaluation (possibly including gradient evaluation if supplied), we then have the following result:

Proposition 3. The whole set of first-order and total support indices, with and without normalization, can be estimated at the cost of $N(d+1)$ function evaluations, and only N function evaluations if the gradient of f is supplied.

Notice that the computational cost for estimating the support indices is comparable to the cost of $N(d+2)$ given in (Saltelli, 2002, Theorem 1) for estimating the whole set of usual first-order and total Sobol' indices.

Remark 2. An alternative for estimating the (unnormalized) support indices functions $\zeta_i(\cdot)$'s is to use the fact that if f is C^1 on $[0, 1]^d$ then

$$\zeta_i(t) = \left(\frac{\partial}{\partial x_i} \mathbb{E}[f(\mathbf{X}) | X_i = t] \right)^2.$$

Thus an estimate is obtained by numerical derivation from the regression curve of $f(\mathbf{X})$ with respect to X_i , avoiding to compute the gradient of f . However, such simplification seems possible neither for estimating the total support indices $\zeta_i^T(\cdot)$'s nor for estimating the DGSM ν_i 's.

5. Examples and application

We now look at a polynomial function and the Ishigami function from the introduction to show how the new indices capture the changing impact of the variable over the support. Then we give a complete sensitivity analysis of a real-life application in the field of sheet metal forming.

5.1 A polynomial function

Let us consider the second order polynomial

$$f(x_1, x_2, x_3) = x_2 - x_3^2.$$

For sensitivity analysis, we consider independent random variables X_1, X_2, X_3 drawn uniformly on $[-1, 1]$. Clearly X_1 is not influential, X_2 has a linear effect whereas X_3 has a quadratic linear effect. This behaviour is easily discovered on scatterplots of f (Figure 1).

The results of the support analysis are then shown in Figure 2. A 5 000-point sample was used, although analytical computations are easy here. More precisely, normalized support curves $\zeta_i^*(\cdot)$ and total support curves $\zeta_i^{*T}(\cdot)$ are represented for $i = 1, 2, 3$. For each x_i , the overall mean of the latter is also represented, as it is equal to the normalized DGSM ν_i^* (Prop. 2) and can be easily interpreted. In particular $\sum_{i=1}^d \nu_i^* = 1$.

We rediscover on Figure 2 that X_1 has neither a global nor a local influence as well as the intuition that the location of X_2 has no impact on its influence to the output. This is explained by its pure linear effect on the

output and the uniform distribution over the support. Similarly, X_3 is all the more influent than closer to the support boundary, explained by its pure quadratic effect and the symmetry with respect to the support. This information seems not easily seen on the usual visualization tools such as scatterplots of Figure 1. Actually, as proved in Prop.1, this requires an estimate of the derivatives of f , which was used to create the new figure.

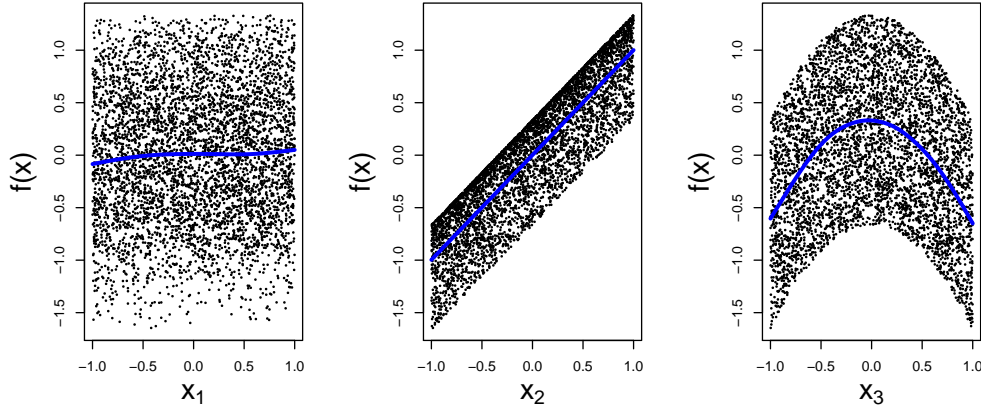


Figure 1: Scatterplots of the polynomial function and main effects estimated by smoothing. A 5 000-point sample was used. The global mean has been removed.

5.2 Ishigami function

We continue the introductory example about Ishigami function:

$$f(x_1, x_2, x_3) = \sin(x_1) + 7 \sin^2(x_2) + 0.1x_3^4 \sin(x_1).$$

For sensitivity analysis, we consider independent random variables X_1, X_2, X_3

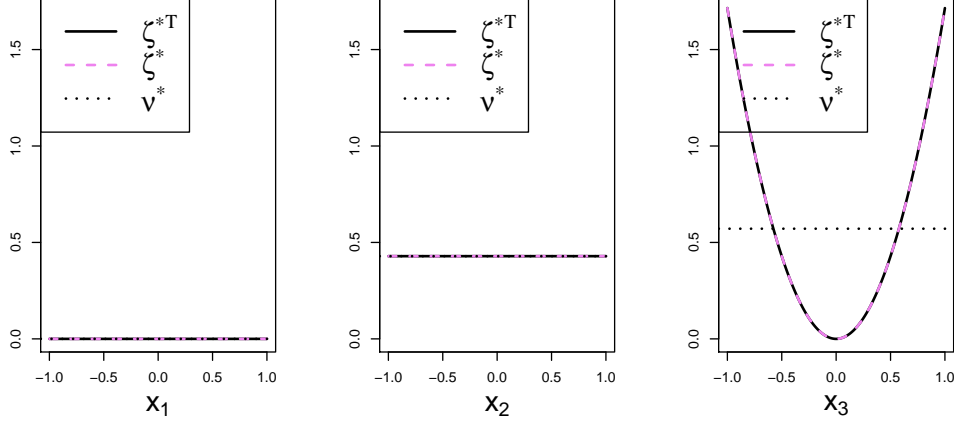


Figure 2: Support analysis of the polynomial function. Estimated curves are normalized by the sum of DGSM values.

drawn uniformly on $[-\pi, \pi]$. First, some elements of a standard global sensitivity analysis are shown in Table 1 and Figure 3, including Sobol' indices and DGSM values, as well as a plot of main effects. We then show the new support index functions in Figure 4.

The results of support analysis confirm the effect observed in the Introduction, that the location of X_3 on its support has a large influence on its total Sobol' index. Indeed, looking at the (normalized) total support curve $\zeta_3^{*T}(\cdot)$, we can see that this influence is neglectible at the middle of the support and increasing when X_3 gets closer to the boundary. We also see that this behaviour is entirely due to interactions, since the support index

values $\zeta_3^*(t)$ are null. Beside this the sinusoidal behavior of X_1 and X_2 can be observed. As the calculation of Sobol' indices over subdomains in Panier and Graf (2015) is along the same line as our approach, their Figure 11 for the Ishigami function shows the same general behaviour. However, it consists of a piecewise constant function corresponding to fixed size intervals. In our Figure, we directly obtain a smooth function corresponding to infinitesimal intervals, which gives a more local information.

Surprisingly here, similar insight can be viewed by looking at the variability of the projected points on the scatterplots of f (Figure 3). This is due to the fact that, for Ishigami function, the scatterplots of f have common shapes with diagonal scatterplots $\left(X_i, \frac{\partial f(\mathbf{X})}{\partial x_i}\right)$ of the gradient of f (Figure 5), on which the support indices are estimated. Actually the conditional variance, visible in the scatterplots of f , is the limit of the overall variance when the support length tends to zero, as proved in Theorem 1. This is much different from the asymptotic behaviour of Sobol' indices, measured by support index functions. In general, scatterplots of f cannot be used to deduce all the information concerning support analysis.

5.3 Application to sheet metal forming

The industrial process of deep drawing is a fundamental procedure for

	Sobol' index	total Sobol' index	DGSM	normalized DGSM
X_1	0.314	0.558	7.720	0.179
X_2	0.442	0.442	24.5	0.567
X_3	0	0.244	10.987	0.254

Table 1: Global sensitivity values for Ishigami function.

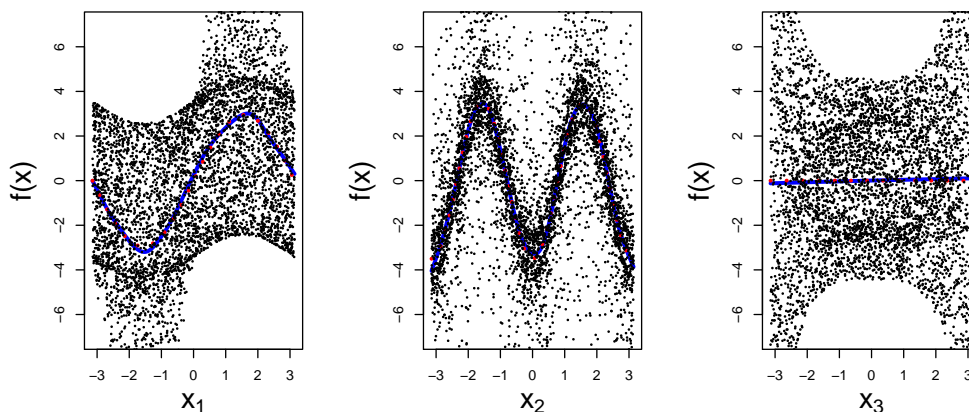


Figure 3: Scatterplots and main effects of the Ishigami function, estimated by smoothing from a 5000-point sample. The global mean has been removed. Theoretical curves are superimposed in dotted lines.

forming sheet metal into desired shapes. It is for example extensively applied in the automotive industry in the production of car bodies. In the deep drawing process, a flat sheet metal is pressed with a punch into a die while so-called blankholders keep the metal fixed at the metal borders. The application aims at analyzing the influence of several input parameters

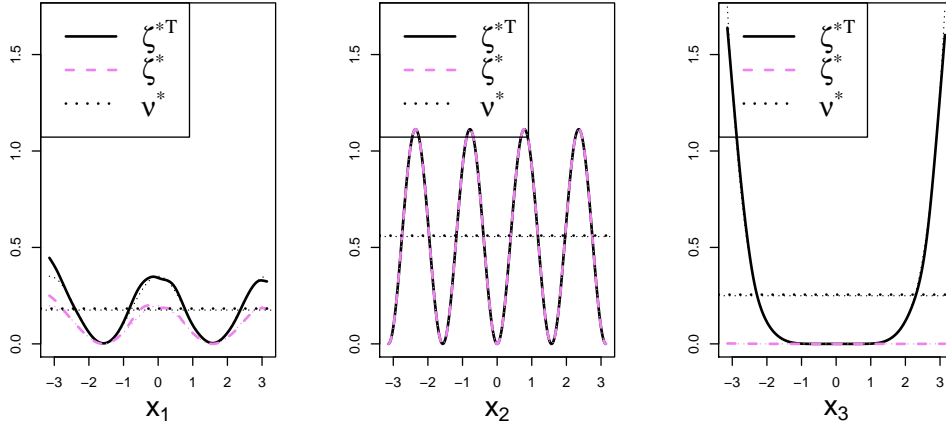


Figure 4: Support analysis of the Ishigami function. Estimated curves are normalized by the sum of DGSM values. Theoretical curves are superimposed in thinner lines.

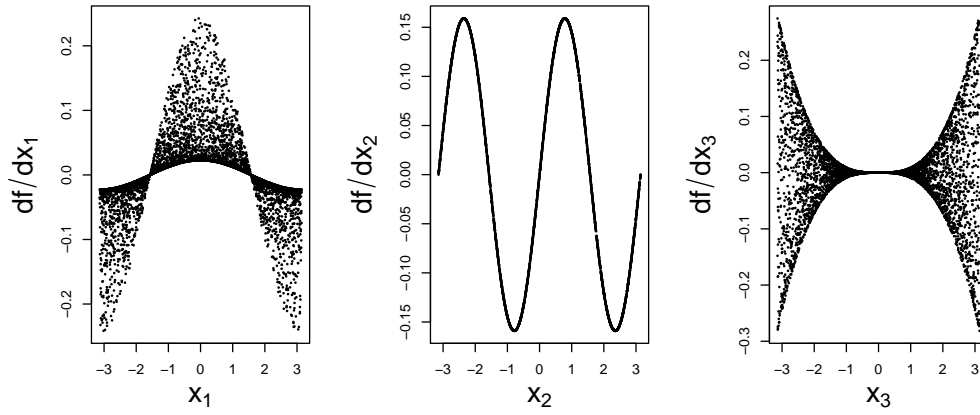


Figure 5: Diagonal scatterplots $\left(X_i, \frac{\partial f(\mathbf{X})}{\partial x_i}\right)$ of the gradient of Ishigami function, for a 5000-point sample.

on the thickness reduction of the formed part. A high thickness reduction at a region of the sheet metal means a strong local thinning, which can lead to structural deformations or actual tearing. It is thus of interest to know, which parameters influence the thinning and in which way as well as modeling and optimization.

Eight input variables are varied in this study, listed in Table 2 together with ranges assumed by the practitioners. The four variables flow stress, initial sheet thickness, hardening exponent and sheet layout concern the material. On the other hand, the two variables blankholder force and friction can be varied during the process. In practice, friction can be changed during the process by adding or removing lubricant with high pressure air and oil removing agents. To roughly integrate this change of friction over the process time, friction is modeled as three independent variables corresponding to three stages of the process time over which friction is kept constant. The output value is the maximal thickness reduction, that is the thickness reduction of the point of the formed part with the strongest thinning. It is given as the ratio between thickness reduction and initial thickness which leads to a scalar value between 0 and 1.

As the duration of the simulation lies between two and six hours per run, it is out of the scope of a direct analysis. A common approach to reduce

Table 2: Input variables of the thickness reduction application.

	input variable	feasible region
X_1	flow stress	100-200 MPa
X_2	initial sheet thickness	0.5-1.7 mm
X_3	blankholder force	50-200 kN
X_4	friction; 1 st third of process time	0-0.14
X_5	friction: 2 nd third of process time	0-0.14
X_6	friction: 3 rd third of process time	0-0.14
X_7	hardening exponent	0.1-0.3
X_8	sheet layout	100-150%

the number of evaluations is to replace the original exact model by a fast-to-evaluate metamodel, which is fitted on a comparatively small amount of evaluations (see e.g. Fang et al., 2006). For the design of these evaluations, space-filling designs like Latin hypercube design (McKay et al., 1979) are well suited. In this application, a maximin Latin hypercube design with 50 runs is applied. The input variables are scaled to lie between 0 and 1. While this does not affect the sensitivity analysis, it simplifies computations and is necessary for the comparisons in the support analysis. On the results, a Gaussian process (GP) metamodel with a standard product kernel, a single constant trend, and Matérn 5/2 covariance functions is fitted (see e.g.

Santner et al., 2003; Roustant et al., 2012). The accuracy of the metamodel is validated by leave-one-out cross-validation. The root mean square error of the cross-validated predictions equals 0.11, which is regarded sufficient by the mechanical engineers.

Using the obtained GP metamodel, first-order Sobol' and total sensitivity indices, are estimated by applying independent uniform distributions to the input variables. For the computation and visualization the R package `sensitivity` (Pujol et al., 2014) is used. The results are shown in Figure 6. According to this, variable X_8 , the sheet layout, has the strongest influence followed by the two last parts of the friction X_5 and X_6 . The total sensitivity indices show strong interactions for those variables. Additional information is provided in Table 3. Notice that the superiority of X_8 is accentuated when measured by DGSM. This is due to the larger fluctuations of that variable, as can be seen on the scatterplots of main effects (Figure 7).

	Sobol' index	total Sobol' index	DGSM	normalized DGSM
X_5	0.136	0.301	0.087	0.090
X_6	0.095	0.205	0.065	0.068
X_8	0.470	0.692	0.776	0.805

Table 3: Global sensitivity values of the GP metamodel for the sheet metal application, limited to three most influential variables.

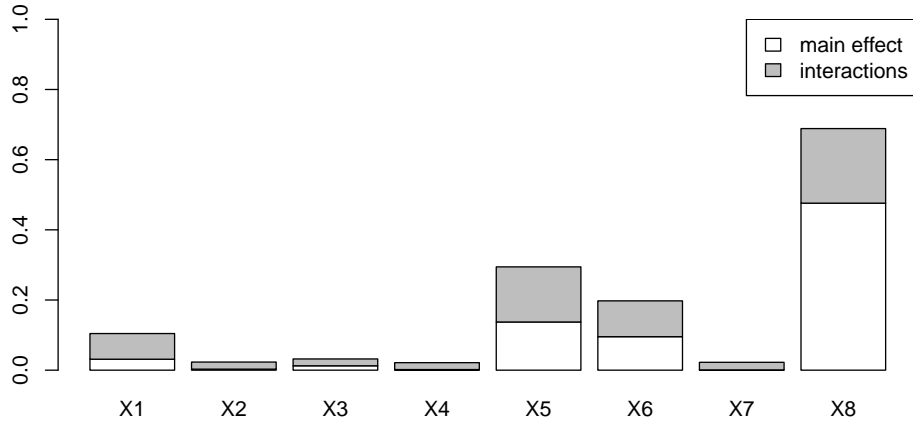


Figure 6: Sheet metal application, first-order Sobol' and total sensitivity indices.

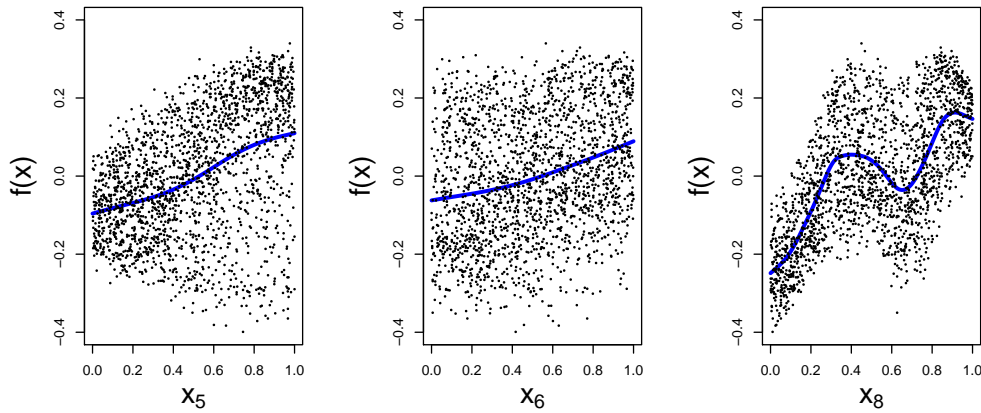


Figure 7: Scatterplots and estimated main effects for the GP metamodel in the sheet metal application. A 2000-point sample was used. The global mean has been removed.

As we are working with the very fast-to-evaluate GP metamodel, we

can explore this main effect analysis further by looking at the support indices of the three most influential variables X_5 , X_6 and X_8 . This analysis shows us the local influence of the variables over the domain and can in particular check if the chosen input ranges influence the results critically. Support indices are estimated with a 2000-point sample. The resulting support functions are drawn in Figure 8. The plots show that the chosen ranges indeed seem appropriate, as the output variation is not strong at the borders of the input supports. Clearly X_8 has a much more local influence. This is not surprising since the mean of the total support curve is equal to the normalized DGSM value, which is much more important for X_8 . The wavy shape of the support index function of that variable reveals strong nonlinearities in the relation between this variable and the thickness reduction. For instance, the local influence of X_8 — when X_5, X_6 lie in $[0, 1]$ — is strong around 0.2 and 0.8, but vanishes around 0.4, 0.6, and 0.9. The two friction parameters X_5 and X_6 show rather similar support functions, underlining that both describe the same type of parameter. The interactions seem to be rather equally spread over the support.

6. Conclusion and perspectives

Two new sensitivity indices, the first-order support index and the total support index, have been presented that extend the Sobol' index and

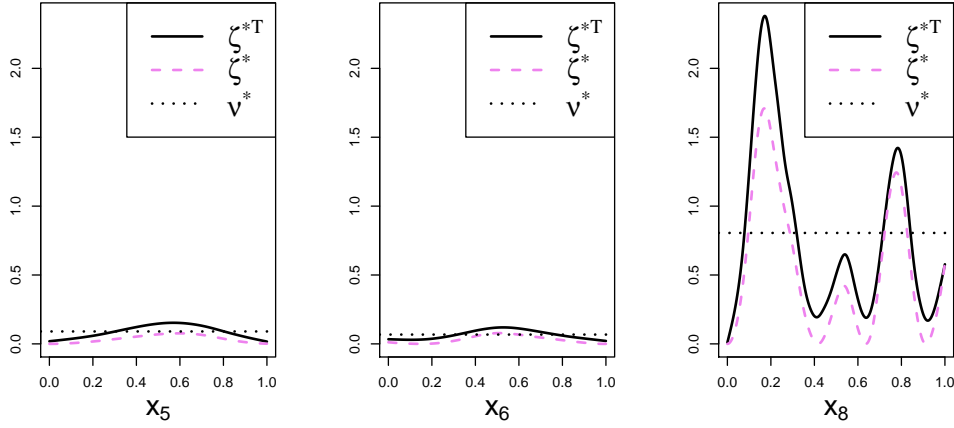


Figure 8: Support analysis of the sheet metal application. Estimated curves are normalized by the sum of DGSM values.

the total sensitivity index, respectively. They give an original and global solution to one of the challenges of sensitivity analysis named in Saltelli et al. (2000): to determine regions within the input space for which the model variation is maximum. By returning a function of sensitivities over the input space, the indices give insight into the first-order and the total local behavior of input variables.

On the theoretical level, we showed that the two support index functions are asymptotically connected to their corresponding scalar indices when the support approaches zero. Furthermore, we showed that the expected value over the support of the functions is linked to derivative-based indices and

variance-based indices, respectively. The results are general as they only require independence and work for every probability density functions of the input variables.

The functions give insight into the local influence of variables and thus can be used as an extension to a sensitivity analysis via standard scalar indices. They can be especially helpful in the specification of the input distribution, a possibly highly influential, but often vaguely determined point in sensitivity analysis. They can further be useful in optimization to explore potential regions in the input domain.

The estimation of the whole set of first-order and total support indices can be done at a reasonable cost with non-parametric smoothing techniques applied to scatterplots of the output gradient. Indeed, using a Monte Carlo sample of size N , the computational cost is equal to $N(d + 1)$ where d is the number of input variables, which is comparable to the cost for estimating Sobol' indices. Furthermore the cost can be reduced to N when the gradient is known. The accuracy of the estimators directly depends on the accuracy of the smoothers used to estimate a regression curve. This could be enhanced in future works by adapting specific smoothing techniques involving derivatives, known as Average Derivative Estimation (see e.g. Lu, 1996). Finally notice that the normalization proposed in Definition 2 was

used by Touzani and Busby (2014) in the framework of DGSMs, who empirically showed its benefit in estimation accuracy. One may wonder whether this remains true for support indices.

A limitation of support indices corresponds to the one observed for the Sobol' indices and derivative-based Sobol' measures. As the support indices represent functions of partial derivatives, strongly varying models will lead to higher support indices than smooth ones, an effect which is not reflected in the scalar indices which just summarize the output variance. This is due to the focus on local sensitivities and only important in the case of strongly varying models, in which careful interpretation is necessary.

Among future works, it would be useful to extend the approach to the case of dependent inputs, and to consider developments for Gaussian process metamodeling linked to derivatives and model error.

7. Proofs

We provide the proofs for Proposition 1 and 2.

Proof of Proposition 1

Let us first show that

$$\frac{\text{var}(X_i^h - t)}{h^2} \xrightarrow{h \rightarrow 0} \frac{1}{12}.$$

Let g be the density function of $X_i - t$, $g(x) = f_{X_i}(x + t)$. By assumption, g is continuous in 0, $g(0) > 0$. Consider the density of the truncated variable

$$X_i^h - t, g_h(x) = \frac{g(x)}{\int_{-\frac{h}{2}}^{\frac{h}{2}} g(t) dt} \mathbf{1}_{[-\frac{h}{2}, \frac{h}{2}]}(x).$$

Let $\ell_h = \inf_{x \in [-\frac{h}{2}, \frac{h}{2}]} g(x)$ and $u_h = \sup_{x \in [-\frac{h}{2}, \frac{h}{2}]} g(x)$. With continuity of f in 0, we have $u_h \xrightarrow{h \rightarrow 0} g(0)$ and $\ell_h \xrightarrow{h \rightarrow 0} g(0)$. By standard integral computations we get

$$h\ell_h \leq \int_{-\frac{h}{2}}^{\frac{h}{2}} g(x) dx \leq hu_h \quad \text{and} \quad \frac{h^3}{12}\ell_h \leq \int_{-\frac{h}{2}}^{\frac{h}{2}} g(x)x^2 dx \leq \frac{h^3}{12}u_h$$

so that for $E((X_i^h - t)^2) = \frac{\int_{-\frac{h}{2}}^{\frac{h}{2}} g(x)x^2 dx}{\int_{-\frac{h}{2}}^{\frac{h}{2}} g(x) dx}$ it holds that

$$\frac{\ell_h}{u_h} \frac{h^3/12}{h} \leq E((X_i^h - t)^2) \leq \frac{u_h}{\ell_h} \frac{h^3/12}{h}.$$

By dividing by h^2 this leads to

$$\frac{E((X_i^h - t)^2)}{h^2} \xrightarrow{h \rightarrow 0} \frac{1}{12}.$$

It remains to show that $\frac{E(X_i^h - t)}{h} \xrightarrow{h \rightarrow 0} 0$. As

$$\int_{-\frac{h}{2}}^{\frac{h}{2}} g(x)x dx = \int_0^{\frac{h}{2}} (g(x) - g(-x))x dx$$

it holds that

$$\begin{aligned} |E(X_i^h - t)| &\leq \frac{1}{\ell_h h} \int_0^{\frac{h}{2}} |g(x) - g(-x)|x dx \leq \frac{1}{\ell_h h} \int_0^{\frac{h}{2}} (u_h - \ell_h)x dx \\ &= \frac{u_h - \ell_h}{\ell_h h} \int_0^{\frac{h}{2}} x dx = \frac{u_h - \ell_h}{\ell_h h} \frac{h^2}{8}. \end{aligned}$$

Thus

$$\frac{|E(X_i^h - t)|}{h} \leq \frac{u_h - \ell_h}{8\ell_h} \xrightarrow{h \rightarrow 0} 0.$$

In the following, denote $B_i = \frac{X_i^h - t}{h}$. Remark that B_i is bounded by $|B_i| \leq \frac{1}{2}$ and that $\text{var}(B_i) \xrightarrow{h \rightarrow 0} \frac{1}{12}$.

Proof of the first statement Write the mean value theorem between a real number $x_i \in [t - h/2, t + h/2] \subseteq [0, 1]$ and t

$$f(\mathbf{x}) = f(x_i, \mathbf{x}_{-i}) = f(t, \mathbf{x}_{-i}) + (x_i - t) \frac{\partial f}{\partial x_i}(c, \mathbf{x}_{-i})$$

for a c between x_i and t , thus $c \in [t - h/2, t + h/2]$. Replacing by random variables, we get

$$f(X_i^h, \mathbf{X}_{-i}) = f(t, \mathbf{X}_{-i}) + (X_i^h - t) \frac{\partial f}{\partial x_i}(C, \mathbf{X}_{-i}),$$

where C is a random variable such that $C \in [t - h/2, t + h/2]$. Denoting

$Q = B_i \frac{\partial f}{\partial x_i}(C, \mathbf{X}_{-i})$, we then have

$$f(X_i^h, \mathbf{X}_{-i}) = f(t, \mathbf{X}_{-i}) + hQ.$$

Now remark that Q is bounded since both B_i and $\frac{\partial f}{\partial x_i}$ are bounded (by continuity of $\frac{\partial f}{\partial x_i}$ on the compact set $[0, 1]^d$). This implies that $E(f(X_i^h, \mathbf{X}_{-i})) \xrightarrow{h \rightarrow 0}$

$E(f(t, \mathbf{X}_{-i}))$ by Lebesgue theorem as well as $E(f(X_i^h, \mathbf{X}_{-i})^2) \xrightarrow{h \rightarrow 0} E(f(t, \mathbf{X}_{-i})^2)$.

The result follows.

Proof of the second statement The proof is similar to the one of 1. Write the Taylor-Lagrange expansion of f between a real number $x_i \in [t - h/2, t + h/2] \subseteq [0, 1]$ and t

$$f(\mathbf{x}) = f(x_i, \mathbf{x}_{-i}) = f(t, \mathbf{x}_{-i}) + (x_i - t) \frac{\partial f}{\partial x_i}(t, \mathbf{x}_{-i}) + \frac{1}{2}(x_i - t)^2 \frac{\partial^2 f}{\partial x_i^2}(c, \mathbf{x}_{-i})$$

for a c between x_i and t , thus $c \in [t - h/2, t + h/2]$. Replacing by random variables, we have

$$f(X_i^h, \mathbf{X}_{-i}) = f(t, \mathbf{X}_{-i}) + hB_i \frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}) + h^2 \tilde{R},$$

with $\tilde{R} = \frac{1}{2} B_i^2 \frac{\partial^2 f}{\partial x_i^2}(C, \mathbf{X}_{-i})$, where C is a random variable, $C \in [t - h/2, t + h/2]$. Then, using the independence between X_i^h and \mathbf{X}_{-i} , we have

$$\mathbb{E} [f(X_i^h, \mathbf{X}_{-i}) | X_i^h] = \beta_0 + h\beta_1 B_i + h^2 R,$$

with $\beta_0 = \mathbb{E} (f(t, \mathbf{X}_{-i}))$, $\beta_1 = \mathbb{E} \left(\frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}) \right)$, and $R = \mathbb{E} [\tilde{R} | X_i^h]$. Notice that $\beta_1 = \mathbb{E} \left[\frac{\partial f}{\partial x_i}(\mathbf{X}) \middle| X_i = t \right]$, by using that X_i and \mathbf{X}_{-i} are independent.

Finally

$$\text{var} (\mathbb{E} [f(X_i^h, \mathbf{X}_{-i}) | X_i^h]) = h^2 \beta_1^2 \text{var}(B_i) + 2h^3 \beta_1 \text{Cov}(B_i, R) + h^4 \text{var}(R).$$

Now remark that \tilde{R} is a bounded random variable by continuity of $\frac{\partial^2 f}{\partial x_i^2}$ on the compact set $[0, 1]^d$, and thus R is bounded as well. This implies that $\text{var}(R) = O(1)$ and $\text{Cov}(B_i, R) = O(1)$, and the result follows.

Proof of the third statement With the same notations as in the proof of 2.,

we have

$$\mathbb{E} [f(X_i^h, \mathbf{X}_{-i}) | \mathbf{X}_{-i}] = f(t, \mathbf{X}_{-i}) + h\mathbb{E}(B_i) \frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}) + h^2 \mathbb{E} [\tilde{R} | \mathbf{X}_{-i}]$$

and thus

$$f(X_i^h, \mathbf{X}_{-i}) - \mathbb{E} [f(X_i^h, \mathbf{X}_{-i}) | \mathbf{X}_{-i}] = h(B_i - \mathbb{E}(B_i)) \frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}) + h^2 S,$$

with $S = \tilde{R} - \mathbb{E} [\tilde{R} | \mathbf{X}_{-i}]$. Since S is centered, we have by independence of B_i and $\frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i})$

$$\begin{aligned} \mathbb{E} \left((f(X_i^h, \mathbf{X}_{-i}) - \mathbb{E} [f(X_i^h, \mathbf{X}_{-i}) | \mathbf{X}_{-i}])^2 \right) = \\ \beta_2 \text{var}(B_i) h^2 + 2h^3 \text{Cov} \left((B_i - \mathbb{E}(B_i)) \frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}), S \right) + h^4 \text{var}(S), \end{aligned}$$

with $\beta_2 = \mathbb{E} \left(\left(\frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i}) \right)^2 \right)$. Notice that $\beta_2 = \mathbb{E} \left[\left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)^2 \middle| X_i = t \right]$, by using the independence between X_i and \mathbf{X}_{-i} . The result follows in the same way as in the proof of 2. by using that $\frac{\partial f}{\partial x_i}(t, \mathbf{X}_{-i})$ and S are bounded random variables.

Proof of Proposition 2

The first point is obtained directly by writing

$$\mathbb{E} [\zeta_i^T(X_i)] = \mathbb{E} \left[\mathbb{E} \left(\left(\frac{\partial f}{\partial x_i}(X_i, \mathbf{X}_{-i}) \right)^2 \middle| X_i \right) \right] = \nu_i.$$

For the second point, compute the first-order Sobol' index of $\frac{\partial f}{\partial x_i}(\cdot)$ corresponding to the input variable X_i by using the pick-freeze formula (Sobol',

1993) with \mathbf{Z}_{-i} an independent copy of \mathbf{X}_i ,

$$D_i \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right) = \mathbb{E} \left[\frac{\partial f}{\partial x_i}(X_i, \mathbf{X}_{-i}) \frac{\partial f}{\partial x_i}(X_i, \mathbf{Z}_{-i}) \right] - \left(\left(\frac{\partial f}{\partial x_i} \right)_0 \right)^2.$$

where $\left(\frac{\partial f}{\partial x_i} \right)_0 = E \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right)$. As both factors in the expectation are independent conditionally on X_i , we can write

$$\begin{aligned} D_i \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right) &= \mathbb{E} \left[\mathbb{E} \left(\frac{\partial f}{\partial x_i}(X_i, \mathbf{X}_{-i}) \frac{\partial f}{\partial x_i}(X_i, \mathbf{Z}_{-i}) \mid X_i \right) \right] - \left(\left(\frac{\partial f}{\partial x_i} \right)_0 \right)^2 \\ &= \mathbb{E} \left[\left(\mathbb{E} \left[\frac{\partial f}{\partial x_i}(X_i, \mathbf{X}_{-i}) \mid X_i \right] \right) \left(\mathbb{E} \left[\frac{\partial f}{\partial x_i}(X_i, \mathbf{Z}_{-i}) \mid X_i \right] \right) \right] - \left(\left(\frac{\partial f}{\partial x_i} \right)_0 \right)^2, \end{aligned}$$

and due to the same distribution of \mathbf{X}_{-i} and \mathbf{Z}_{-i} , we obtain

$$D_i \left(\frac{\partial f}{\partial x_i}(\mathbf{X}) \right) = \mathbb{E} \left[\left(\mathbb{E} \left[\frac{\partial f}{\partial x_i}(X_i, \mathbf{X}_{-i}) \mid X_i \right] \right)^2 \right] - \left(\left(\frac{\partial f}{\partial x_i} \right)_0 \right)^2.$$

Acknowledgements This paper is based on investigations of the collaborative research centre SFB 708, project C3. The authors would like to thank the Deutsche Forschungsgemeinschaft (DFG) for funding. We are grateful to an anonymous associate editor and two referees for their relevant comments and suggestions, which led to substantial improvements.

References

- Fang, K., Li, R. and Sudjianto, A. (2006). *Design and modeling for computer experiments*, Chapman & Hall/CRC, London.
- Fruth, J., Roustant, O. and Kuhnt, S. (2014). Total Interaction Index: A Variance-based Sensitivity Index for Second-order Interaction Screening. *Journal of Statistical Planning and Inference* **147**, 212–223.
- Hastie, T. J. and Tibshirani, R. J. (1990). *Generalized additive models*. Chapman & Hall/CRC, London.
- Hoeffding, W. (1948). A Class of Statistics with Asymptotically Normal Distribution. *The Annals of Mathematical Statistics* **19(3)**, 293–325.
- Homma, T. and Saltelli, A. (1996). Importance measures in global sensitivity analysis of non-linear models. *Reliability Engineering & System Safety* **52(1)**, 1-17.
- Iooss, B., Popelin, A. L., Blatman, G., Ciric, C., Gamboa, F., Lacaze, S. and Lam, M. (2012). Some new insights in derivative-based global sensitivity measures. *PSAM 11 & ESREL*

2012 Conference, Helsinki, Finland.

Ishigami, T. and Homma, T. (1990). *An importance quantification technique in uncertainty analysis for computer models*. In: Proceedings of the ISUMA '90, Editor Ayyub, B.M., 398–403.

Kucherenko, S., Rodriguez-Fernandez, M., Pantelides, C. and Shah, N. (2009). Monte Carlo evaluation of derivative-based global sensitivity measures. *Reliability Engineering & System Safety* **94(7)**, 1135–1148.

Lamboni, M., Iooss, B., Popelin, A.-L. and Gamboa, F. (2013). Derivative-based global sensitivity measures: general links with Sobol' indices and numerical tests. *Mathematics and Computers in Simulation* **87**, 45–54.

Li, L., Lu, Z. and Hu, J. (2014). A new kind of regional importance measure of the input variable and its state dependent parameter solution. *Reliability Engineering & System Safety* **128**, 1–16.

Lu Z.Q. (1996). Multivariate Locally Weighted Polynomial Fitting and Partial Derivative Estimation. *Journal of Multivariate Analysis* **59**, 187–205.

McKay, M. D., Beckman, R. J. and Conover, W. J. (1979). A comparison of three methods for selecting values of input variables in the analysis of output from a computer code. *Communications in Statistics - Theory and Methods* **21(2)**, 239–245.

Mühlenstädt, T., Roustant, O., Carraro, L. and Kuhnt, S. (2012). Data-driven Kriging models based on FANOVA-decomposition. *Statistics and Computing* **22**, 723–738.

- Oakley, J. E. and O'Hagan, A. (2004). Probabilistic sensitivity analysis of complex models: a Bayesian approach. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* **66(3)**, 751–769.
- Pannier, S. and Graf, W. (2015). Sectional global sensitivity measures. *Reliability Engineering and System Safety* **134**, 110–117.
- Pujol, G., Iooss, B. and Janon, A. (2014) *sensitivity: Sensitivity Analysis*, R package version 1.10.1. <http://CRAN.R-project.org/package=sensitivity>
- Rakovec, O., Hill, M. C., Clark, M. P., Weerts, A. H., Teuling, A. J. and Uijlenhoet, R. (2014). Distributed evaluation of local sensitivity analysis (DELSA), with application to hydrologic models. *Water Resources Research* **50(1)**, 409–426.
- Roustant, O., Ginsbourger, D. and Yves Deville (2012). DiceKriging, DiceOptim: Two R Packages for the Analysis of Computer Experiments by Kriging-Based Metamodeling and Optimization. *Journal of Statistical Software* **51(1)**, 1–55.
- Roustant, O., Fruth, J., Iooss, B. and Kuhnt, S. (2014). Crossed-derivative based sensitivity measures for interaction screening. *Mathematics and Computers in Simulation* **105**, 105–118.
- Saltelli, A., Chan, K. and Scott, E.M. (2000). *Sensitivity analysis*. Wiley, Chichester.
- Saltelli, A. (2002). Making best use of model evaluations to compute sensitivity indices. *Computer Physics Communications* **145(2)**, 280–297.
- Santner, T.J., Williams, B.J. and Notz, W.I. (2003). *The design and analysis of computer*

experiments, Springer, New York.

Sobol', I.M. (1993). Sensitivity estimates for non linear mathematical models. *Mathematical Modelling and Computational Experiments* **1**, 407-414.

Sobol', I. and Gershman, A. (1995) On an alternative global sensitivity estimators. *In: Proceedings of SAMO 1995, Belgirate*, 40–42.

Turanyi, T. (1990) Sensitivity analysis of complex kinetic systems. Tools and applications. *J. Math. Chem.* **5**, 203-248.

Touzani, S. and Busby, D. (2014). Screening method using the derivative-based global sensitivity indices with application to reservoir simulator, *Oil & Gas Science and Technology – Revue d'IFP Energies nouvelles* **69(4)**, 619–632

van der Vaart, A. W. (1998) *Asymptotic statistics*. Cambridge University Press, Cambridge.

Welch, W. J., Buck, R. J., Sacks, J., Wynn, H. P., Mitchell, T. J. and Morris, M. D. (1992). Screening, Predicting, and Computer Experiments. *Communications in Statistics - Theory and Methods* **34**, 15–25.

Wei, P., Lu, Z., Ruan, W. and Song, J. (2014). Regional sensitivity analysis using revised mean and variance ratio functions. *Reliability Engineering & System Safety* **121** 121–135.

TU Dortmund University, 44221 Dortmund, Germany

E-mail: fruth@statistik.tu-dortmund.de

Mines Saint-Etienne, F-42023 Saint-Etienne, UMR CNRS 6158, LIMOS, France

E-mail: roustant@emse.fr

Dortmund University of Applied Sciences and Arts, 44227 Dortmund, Germany

E-mail: sonja.kuhnt@fh-dortmund.de