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To cite this version:

HAL Id: hal-01068996
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Submitted on 26 Sep 2014

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Comparison of two parameterizations of a turbulence-induced flocculation model through global sensitivity analysis

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Abstract

This paper describes a global sensitivity analysis of a fractal-based turbulence-induced flocculation model. The quantities of interest in this analysis are related to the floc diameters in two different configurations. The input parameters with which the sensitivity analyses are performed are the floc aggregation and breakup parameters, the fractal dimension and the diameter of the primary particles. Two related versions of the flocculation model are considered, evenly encountered in the literature: (i) using a dimensional floc breakup parameter, and (ii) using a non-dimensional floc breakup parameter. The main results of the sensitivity analyses are that only two parameters of model (ii) are significant (aggregation and breakup parameters) and that the relationships between parameter and quantity of interest remain simple. Contrarily, with model (i), all parameters have to be considered. When

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identifying model parameters based on measures of floc diameters, this analysis hence suggests the use of model (ii) rather than (i). Further, improved models of the fractal dimension do not seem to be required when using the non-dimensional model (ii).

Keywords: Flocculation, Sensitivity Analysis, Analysis of Variance, Gravity Currents, Turbidty Currents.

1. Introduction

Flocculation is the property of small cohesive particles to form so-called flocs under the action of electrostatic forces. The flocs are larger groups of particles whose influence can be dramatic in many physical problems, such as for turbidity currents (Thomas et al., 1999; Geyer et al., 2004; Jarvis et al., 2005; Meiburg and Kneller, 2010). The study and modeling of these turbidity currents (Cantero et al., 2006; Elias et al., 2008; Hiester et al., 2011; Nasr-Azadani et al., 2013; Guerra et al., 2013; Espath et al., 2014) is particularly difficult because the particle concentration, and hence the driving buoyancy force, fluctuates with time and position along the current, as well as in the vertical direction. The influence of flocculation is large in that case because the settling velocity of mud flocs can be a few orders of magnitude larger than that of the primary particles, so that its effect on particle concentration is very important.

The flocculation process is influenced by sediment size, sediment concentration, turbulence intensity, temperature, and the presence of organic matter. It is governed by three processes that drive particles to collide: (i) brownian motion, (ii) differential settling of fast particles overtaking slower
particles, and (iii) turbulent motion. In estuaries and coastal regions, the first is arguably negligible (Winterwerp, 1998; Cuthbertson et al., 2010). Concerning differential settling, which is likewise often disregarded, a model is described in Lick et al. (1993) and experimental results are reported in Cuthbertson et al. (2010). About the effect of turbulent motion, more widely studied, aggregation theory suggests that aggregation rates increase with the turbulence intensity, due to the increased number of inter-particle encounters. However, the same turbulent motions may generate shear stresses that can limit the growth rate through floc disruption and break-up (disaggregation) processes.

Following the pioneering work of von Smoluchowski (1917), who derived the time variation of the concentration of a mono-disperse suspension due to Brownian diffusion, most studies of flocculation in turbidity currents have considered population balance equations along with empirical models of creation and disruption of flocs of different sizes (see for instance Thomas et al. (1999); Son and Hsu (2008); Chao and Jia (2011) and references therein). Many different population balance equations based on Lagrangian or Eulerian formulations have been proposed (Friedlander, 1957; O’Melia, 1980; Spicer and Pratsinis, 1996; Serra and Casamitjana, 1998; McAnally and Mehta, 2002; Rahmani et al., 2004), each taking into account different processes. Winterwerp (1998) developed a model for the turbulence-induced flocculation, a priori aimed at modeling flocculation in estuary and coastal areas. Although initially considering the evolution of flocs of a single characteristic diameter, this model was later extended to the poly-disperse case (Xu et al., 2008). Some authors also extended the model to fractal dimensions
dependent on the floc size (Maggi and Winterwerp, 2004; Khelifa and Hill, 2006; Vahedi and Gorczyca, 2011). All these models rely on several parameters that have to be identified in experiments, and several parameterizations sometimes exist for the same model.

The choice of parameterization influences the identifiability of models. For instance, if a given quantity of interest (QoI) does not depend on one of the model parameters, it means that this parameter can be chosen at will without degrading the estimation of that QoI. Conversely, the sensitivity of a QoI with respect to the parameters indicates the parameters that will be polluted by noise and those that will be identified with more confidence. The reason why a given QoI might be more or less sensitive to one or another parameter is due, among other things, (i) to the particular form of the model (the relation between parameters and QoI might be very non-linear), and (ii) to differences of orders of magnitudes between the parameters. Sensitivity analysis (Saltelli et al., 2000) is the study of these dependencies between model parameters and QoIs. It hence allows to identify parameters that should be the focus of more attention.

In the context of flocculation models, such sensitivity analyses have been performed for two QoIs (floc diameter and settling velocity) and with respect to some parameters in Khelifa and Hill (2006); Son and Hsu (2008); Maerz et al. (2011). In particular, a major conclusion of the sensitivity study in Maerz et al. (2011) points at the fractal dimension as the most influential parameter on the floc diameter, while its experimental identification is generally difficult (Bushell et al., 2002). Among other things, this conclusion motivates the use of refined models of that fractal dimension. In Mietta
et al. (2008), the authors presented a sensitivity study with respect to the breakup functions, that is to say with respect to different functional forms of the breakup model. Note that all these analyses remained local, in the sense that the influence of variations of the parameters (one at the time) was observed around one reference model. Conclusions drawn from a local sensitivity analysis (SA) around a given reference might be modified by the consideration of another reference point.

Contrarily to local SA, global SA aims at investigating the functional form of a model rather than the behavior around a particular design point. The objectives of global SA are the same as in local SA, but the influence of the parameters is considered jointly over the entire range of all the parameters. In particular, this means that non-linearities between parameters and QoIs are explicitly considered, while local SA explore linearized models. To the best of our knowledge, this paper presents the first global SA of flocculation models. We concentrate here on Winterwerp’s model (Winterwerp, 1998), for a unique characteristic diameter and constant fractal dimension, and perform a Monte-Carlo based global SA (Helton et al., 2006; Saltelli et al., 2008) with respect to four model parameters. Two of the parameters we consider are common to all flocculation models (although sometimes presented in the form of efficiency parameters): the flocculation parameter \(k_A\) and the breakup parameter \(k_B\). Two parameters are particular to fractal-based models: the fractal dimension \(n_f\) and the primary particle diameter \(D_P\). The QoIs that we consider are all related to floc diameters (average, minima and maxima). This SA will be performed for two versions of the model, equally encountered in the literature: (i) using a dimensional floc
breakup parameter, and (ii) using a non-dimensional floc breakup parameter. The analysis allows to propose a clear recommendation on the most appropriate parameterization. In particular, it will allow us to reconsider the need for refined models of the fractal dimension parameter.

Section 2 recalls the main components of the two versions of Winterwerp's model, along with their implementation in a coupled system of transport equations and the definitions of equilibrium diameter and time. Section 3 describes a literature survey of ranges of variations of the flocculation parameters and proposes probabilistic descriptions of these parameters for use in the SA. Section 4 describes the SA performed for the equilibrium diameter with the two different parameterizations of Winterwerp's model. Its conclusion is that the non-dimensional model should be used rather than the dimensional one. Section 5 describes the SA performed for a 0D tidal example taken from Maerz et al. (2011) that confirms the previous result. Finally, some conclusions are presented in Section 6.

2. Turbulent flocculation model

In this section, which is mostly based on Winterwerp’s original paper (Winterwerp, 1998), we recall the flocculation model that will be considered throughout. It is based on the hypothesis of fractal behavior of the floc size. It proposes a flocculation source term in a population balance equation for the number of flocs per unit volume \( N \).

2.1. Self-similar flocculation model

In this model, the aggregates are treated as self-similar entities (Kranenburg, 1994). The model therefore assumes that the flocs are composed of an
ensemble of so-called primary particles, with diameter $D_P$, which bond in a very specific way. This assembly process induces a relation between the mass concentration $c$, the number concentration $N$ and the mass of one primary particle $m_P$

$$c = m_P N \left( \frac{D}{D_P} \right)^{n_f}$$

$$N = \frac{c}{m_P} \left( \frac{D_P}{D} \right)^{n_f}$$

$$D = D_P \left( \frac{c}{m_P N} \right)^{1/n_f}, \quad (1)$$

where $n_f$ is called the fractal dimension. Fractal dimension and primary particle diameter are two parameters that have to be identified through experiments. Typical values will be discussed in Sec. 3.2.

The flocculation model proposed in Winterwerp (1998) combines linearly a literature model for the aggregation (Levich, 1962) and an original model for the breakup of flocs. It describes the evolution of the number of flocs as:

$$\delta N = -k_A \frac{m_P}{\rho_s} G \left( \frac{D}{D_P} \right)^3 N^2 + k_B \sqrt{\frac{\mu}{F_y}} G^{3/2} D \left( \frac{D}{D_P} - 1 \right)^{3-n_f} N, \quad (2)$$

where $\rho_s$ is the density of particles, $k_A$ is a dimensionless flocculation coefficient, $k_B$ is a dimensionless floc breakup parameter, $F_y$ is the yield strength of flocs, and $\mu$ is the dynamic viscosity of the fluid. Note that the model presented here is slightly simplified from the original article (Winterwerp, 1998). In particular, processes of deposition and erosion, which are taken into account through additional parameters, are here disregarded.

2.2. Coupled system of transport equation

The previous equation describes only the creation and destruction of flocs through flocculation. The space and time evolution of flocs is controlled by two coupled balance equations, for which the above model acts as a source term. Assuming that the flocs are both convected by the flow of carrying
fluid with a velocity \( \mathbf{w} \), and also naturally settle with a (vertical) velocity \( \mathbf{w}_s \), this transport equation is:

\[
\frac{\partial N}{\partial t} + \nabla \cdot ((\mathbf{w} - \mathbf{w}_s)N - K_N \nabla N) = \delta N, \tag{3}
\]

where \( K_N \) is a diffusion tensor (assumed given and diagonal). The settling velocity of the flocs depends on their size. Assuming that the flocs are spherical, and at low Reynolds number, the settling velocity can be modeled as:

\[
\mathbf{w}_s(D) = -\frac{(\rho_s - \rho_w)gD_P^2}{18\mu} \left( \frac{D}{D_P} \right)^{n_f-1} \mathbf{e}_z \tag{4}
\]

\[
= -\frac{(\rho_s - \rho_w)gD_P^2}{18\mu} \left( \frac{c}{m_P N} \right)^{1-1/n_f} \mathbf{e}_z, \tag{5}
\]

where \( \rho_w \) is the density of water, \( g \) is the acceleration of gravity, and \( \mathbf{e}_z \) is the vertical unit vector pointing upward.

Finally the model is closed with a more classical transport equation for the mass concentration \( c \):

\[
\frac{\partial c}{\partial t} + \nabla \cdot ((\mathbf{w} - \mathbf{w}_s)c - K_c \nabla c) = 0, \tag{6}
\]

where \( K_c \) is a mixing tensor (assumed given and diagonal). Winterwerp’s model therefore consists of a set of two transport equations (3)-(6) for the two unknowns \( c \) and \( N \). These two transport equations are coupled through both the flocculation source term \( \delta N \) (2) and the settling velocity \( \mathbf{w}_s \) (4).

Rewriting the two equations as functions of \( c \) and \( N \) only, and with the coupling terms on the right-hand sides, yields:

\[
\frac{\partial c}{\partial t} + \nabla \cdot (\mathbf{w}c - K_c \nabla c) = w_P \frac{\partial}{\partial z} \left( c \left( \frac{c}{m_P N} \right)^{1-1/n_f} \right), \tag{7}
\]
\[
\frac{\partial N}{\partial t} + \nabla \cdot (wN - KN \nabla N) = \\
w_P \frac{\partial}{\partial z} \left( \frac{c}{m_P} \left( \frac{m_P N}{c} \right)^{1/n_f} \right) - q_A N^2 \left( \frac{c}{m_P N} \right)^{3/n_f} \\
+ q_B N \left( \frac{c}{m_P N} \right)^{1/n_f} \left( \left( \frac{c}{m_P N} \right)^{1/n_f} - 1 \right)^{3-n_f}, \tag{8}
\]

where we have introduced a volume flux of floc aggregation \(q_A = k_A m_P G/\rho_s\), a number flux of floc breakup \(q_B = k_B D_P G \sqrt{\mu G/F_y}\) and the settling velocity of a primary particle (assumed spherical) \(w_P = (\rho_s - \rho_w) g D_P^2/(18 \mu)\).

Boundary conditions for that transport model can be taken (for example) as \((w c + K c) \cdot e_z = 0\) and \((w N + K N \nabla N) \cdot e_z = 0\) at the free surface and on the bottom (assuming there is no erosion or deposition). The above system of equations can be approximated by numerical techniques (see Section 5).

2.3. Stationary regime

Before considering the solution of this coupled system of transport equations, it is interesting to discuss the stationary situation in which flocs are created and destroyed at the same rate, so that the floc population and size remain constant. Considering \(\delta N = 0\) in Eq. (2) yields the following implicit definition for the floc equilibrium diameter \(D_\infty\):

\[
D_\infty \left( 1 - \frac{D_P}{D_\infty} \right)^{3-n_f} = \frac{k_A c}{k_B \rho_s} \sqrt{\frac{F_y}{\mu G}}. \tag{9}
\]

In the limit \(D_\infty >> D_P\), which is often realized in practice, one obtains the approximation

\[
D_\infty \approx (3 - n_f) D_P + \frac{k_A c}{k_B \rho_s} \sqrt{\frac{F_y}{\mu G}}. \tag{10}
\]
We will check in Sec. 4 the accuracy of that approximation. One can also
define a characteristic time scale when \( D_\infty >> D_P \) as:

\[
T_\infty = \frac{n_f}{k_B G} \sqrt{\frac{F_y}{\mu G D_\infty}} \left( \frac{D_P}{D_\infty} \right)^{3-n_f} \tag{11}
\]

This characteristic flocculation time represents the order of magnitude of the
time required to reach the equilibrium diameter \( D_\infty \) when the system is out
of equilibrium. These two quantities \( D_\infty \) and \( T_\infty \) are useful to characterize
grossly the flocculation regimes, depending on the driving phenomena (see
Section 5 for an example in a tidal-controlled cyclic flocculation).

### 2.4. Alternative parameterization

Finally, it is important to note that another parameterization is used as
often in the literature as the one that was already presented. Instead of
defining a dimensionless floc breakup parameter, this alternative parameter-
ization is based on a floc breakup parameter \( k_B' = k_B \sqrt{\mu/F_y}/(n_f D_P^{3-n_f}) \).
Although it might appear as an unimportant change of variables, it may have
a very important impact on the identifiability of the different parameters of
the model. This will be discussed in Section 4.

When considering this parameterization, the equilibrium diameter is then
given by

\[
D_\infty \left( 1 - \frac{D_P}{D_\infty} \right)^{3-n_f} = \frac{k_A c}{k'_B n_f \rho_s D_P^{3-n_f} \sqrt{G}}, \tag{12}
\]

and, in the limit \( D_\infty >> D_P \),

\[
D_\infty \approx (3 - n_f) D_P + \frac{k_A c}{k'_B n_f \rho_s D_P^{3-n_f} \sqrt{G}}, \tag{13}
\]

The characteristic time for the alternative parameterization is

\[
T_\infty = \frac{1}{k'_B G^{3/2} D_\infty^{4-n_f}}, \tag{14}
\]
It should be noted that the dimensional and non-dimensional parameters are perfectly equivalent. Given the set of parameters including the non-dimensional \( k_B \), one can define in a unique and equivalent way the dimensional parameter \( k'_B \). Likewise, given the set of parameters including the dimensional \( k'_B \), one can define in a unique and equivalent way the non-dimensional parameter \( k_B \). However, this choice of parameterization modifies drastically the relationship between the parameters and the quantities of interest (the equilibrium diameter for instance). This is particularly true here, where the relationship between \( k'_B \) and \( k_B \) involves two other parameters \( n_f \) and \( D_P \). Among other things this means that the sensitivity to errors in the parameters might be modified for the two cases.

3. Description of the flocculation parameters

SA refers to the determination of the contributions of individual uncertain inputs to the uncertainty in analysis results. A number of approaches have been developed, including differential analysis, response surface methodologies, Monte Carlo analyses, and variance decomposition methods. We will concentrate here on Monte Carlo analyses (Helton et al., 2006). The objective of the paper is to study the sensitivity of some QoIs with respect to the parameters \( k_A, k_B, n_f \) and \( D_P \). The first step of such a SA consists in choosing appropriate probability density functions for the input parameters. The approach that is chosen here consists in considering a database of experimentally-identified values of these parameters reported in the literature. From that database, average and variance are selected, and the distribution is chosen based on the Maximum Entropy Principle (Shannon,
1948; Jaynes, 1957), the known bounds on the parameters and the given average and variance. Examples of construction of stochastic models using this method can be found in Udwadia (1989); Soize (2000); Cottereau et al. (2007); Adhikari and Friswell (2007).

3.1. Literature survey of flocculation parameters

Values of the four parameters of interest found in the literature are reported in Table 1. In most of the corresponding papers, the authors do not experimentally identify the value of $n_f$, but rather choose it equal to 2 before concentrating on identifying the values of $k_A$ and $k_B$. The choice of $n_f = 2$ has the advantage that the formulas of the flocculation model simplify so that the identification becomes simpler.

Table 1: Literature values of flocculation parameters.

<table>
<thead>
<tr>
<th>Reference</th>
<th>$k_A$</th>
<th>$k_B$</th>
<th>$n_f$</th>
<th>$D_P$</th>
<th>Data from</th>
</tr>
</thead>
<tbody>
<tr>
<td>Winterwerp (1998)</td>
<td>0.31</td>
<td>$3.5 \times 10^{-5}$</td>
<td>2</td>
<td>4</td>
<td></td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>0.33</td>
<td>$1.1 \times 10^{-5}$</td>
<td>2</td>
<td>15</td>
<td>Manning and Dyer (1999)</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>0.3</td>
<td>$1.4 \times 10^{-5}$</td>
<td>2</td>
<td>15</td>
<td>Manning and Dyer (1999)</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>0.008</td>
<td>$4.3 \times 10^{-5}$</td>
<td>2</td>
<td>15</td>
<td>Biggs and Lant (2000)</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>1.02</td>
<td>$0.4 \times 10^{-5}$</td>
<td>2</td>
<td>5</td>
<td>Bouyer et al. (2004)</td>
</tr>
<tr>
<td>Mietta et al. (2008)</td>
<td>0.46</td>
<td>$24.0 \times 10^{-5}$</td>
<td>2</td>
<td>6</td>
<td>Maggi et al. (2002)</td>
</tr>
<tr>
<td>Maerz et al. (2011)</td>
<td>0.18</td>
<td>$2.9 \times 10^{-5}$</td>
<td>1.9</td>
<td>4</td>
<td></td>
</tr>
<tr>
<td>Maerz et al. (2011)</td>
<td>0.4</td>
<td>$11.5 \times 10^{-5}$</td>
<td>1.9</td>
<td>4</td>
<td></td>
</tr>
</tbody>
</table>
However, this (almost) constant value of $n_f = 2$ throughout the reported experiments does not reflect the variability that seems to arise in other sets of experiments. We therefore performed a larger study of the literature, concentrating on obtaining a relevant database for the fractal dimension. The values are reported in Table 2. Note that it is possible to derive values of the fractal dimension based on theoretical considerations of aggregation mechanisms at the molecular scale (see for instance Meakin (1984); Thouy and Jullien (1994); Vahedi and Gorczyca (2012)). However, these results were not used here, because of a lack of knowledge about these mechanisms for the situation we are considering.

<table>
<thead>
<tr>
<th>Ref.</th>
<th>$n_f$ [-]</th>
<th>Data from</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hsu et al. (2007)</td>
<td>2.15</td>
<td></td>
</tr>
<tr>
<td>Vahedi and Gorczyca (2012)</td>
<td>[2.55–2.99]</td>
<td></td>
</tr>
<tr>
<td>Maerz et al. (2011)</td>
<td>[1.5–2.4]</td>
<td>several databases</td>
</tr>
<tr>
<td>Khelifa and Hill (2006)</td>
<td>[2–3]</td>
<td>several databases</td>
</tr>
</tbody>
</table>

Finally, we consider the alternative parameterization described at the end of Section 2 and present literature values for the dimensional floc breakup coefficient $k'_B$ in Table 3. In this table, the values of $k'_B$ are reported with the corresponding values of $n_f$, $D_p$, $F_y$ and $\mu$ because these values are required to go from $k_B$ to and from $k'_B$. Note that, in the papers reported in the left column of Tables 1 and 3, it was either $k_B$ or $k'_B$ that was originally used.
When \( k_B \) was given in the original paper, we computed the corresponding \( k'_B \) and reported it in Table 3. When \( k'_B \) was given in the original paper, we computed the corresponding \( k_B \) and reported it in Table 1.

Table 3: Literature values for the alternative definition of the floc breakup parameter. See Table 1 for the origin of the data.

<table>
<thead>
<tr>
<th>Reference</th>
<th>( k'_B ) ([s^{1/2}/m^2])</th>
<th>( n_f )</th>
<th>( D_P ) ([\mu m])</th>
<th>( F_y ) ([kg.m/s^2])</th>
<th>( \mu ) ([kg/m/s])</th>
</tr>
</thead>
<tbody>
<tr>
<td>Winterwerp (1998)</td>
<td>14×10^3</td>
<td>2</td>
<td>4</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>1.2×10^3</td>
<td>2</td>
<td>15</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>1.5×10^3</td>
<td>2</td>
<td>15</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>4.6×10^3</td>
<td>2</td>
<td>15</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Son and Hsu (2008)</td>
<td>1.2×10^3</td>
<td>2</td>
<td>5</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Mietta et al. (2008)</td>
<td>63.2×10^3</td>
<td>2</td>
<td>6</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Maerz et al. (2011)</td>
<td>12×10^3</td>
<td>1.9</td>
<td>4</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
<tr>
<td>Maerz et al. (2011)</td>
<td>48×10^3</td>
<td>1.9</td>
<td>4</td>
<td>10^{-10}</td>
<td>10^{-3}</td>
</tr>
</tbody>
</table>

3.2. Statistical description of flocculation parameters

The number of identification experiments reported in the literature is rather limited. It is hence not reasonable to expect identifying full probability density functions from the databases reported above (see Berry et al. (2012) for a more precise discussion). Rather, we will proceed in three steps: (i) identify only the expectation and variance of the parameters \( k_A, k_B, n_f, D_P \) and \( k'_B \) from the literature values, (ii) identify the form of the first-order marginal distribution using the Maximum Entropy Principle and the
known bounds of the parameters, and (iii) identify the cross-correlations between the parameters. The first-order marginal densities are the probability density functions of each random variable considered alone, discarding the potential correlation or higher-order relations between these variables. In this approach, and because of the limited information available in the literature, we only aim at constraining (i) the first-order marginal densities and (ii) cross-correlations.

Concerning the first step, and based on the discussion above, the parameters $k_A$, $k_B$ and $D_P$ are identified from the values in Table 1. The fractal dimension $n_f$ is identified from the values in both Tables 1 and 2, considering only the middle value for the intervals. Finally, the dimensional floc breakup parameter $k'_{B}$ is identified from the values in Table 3. The statistical expectation and variance obtained for these five parameters are reported in Table 4. The unbiased estimates for the expectation and variance are used 

$$
\mu_X = \frac{\sum_{i=1}^{N} x_i}{N} \quad \text{and} \quad \sigma^2_X = \frac{\sum_{i=1}^{N} (x_i - \mu_X)^2}{(N - 1)},
$$

where the $x_i$ are the $N$ realizations of the random variable $X$. The correlation coefficient is computed as $

$$
\xi_X = \frac{\sigma_X}{\mu_X}.
$$

<table>
<thead>
<tr>
<th></th>
<th>$k_A$ [-]</th>
<th>$k_B$ [-]</th>
<th>$n_f$ [-]</th>
<th>$D_P$ [µm]</th>
<th>$k'_{B}$ [$s^{1/2}/m^2$]</th>
</tr>
</thead>
<tbody>
<tr>
<td>Expectation $\mu$</td>
<td>0.38</td>
<td>6.2×10^{-5}</td>
<td>2.2</td>
<td>8.5</td>
<td>18.2×10^{3}</td>
</tr>
<tr>
<td>Variance $\sigma^2$</td>
<td>0.09</td>
<td>64.1×10^{-10}</td>
<td>0.15</td>
<td>29.4</td>
<td>573×10^{6}</td>
</tr>
<tr>
<td>Correlation coefficient $\xi$</td>
<td>0.8</td>
<td>1.7</td>
<td>0.2</td>
<td>0.6</td>
<td>1.3</td>
</tr>
</tbody>
</table>

Table 4: Statistical estimates for the flocculation parameters (based on Tables 1 and 3).
We then turn to the choice of the first-order marginal density functions for each of the parameters. The Maximum Entropy Principle (Shannon, 1948; Jaynes, 1957) allows us to select the least informative distribution given a series of constraints. In particular, in our case, we can consider the average and variance identified above as constraints. Further, the parameters have different supports that should be considered as constraints also. By nature, the parameters $k_A, k_B, D_P,$ and $k'_B$ are positive, while the fractal dimension has bounded support $0 < n_f < 3$. Based on these constraints, the Maximum Entropy Principle proposes to use truncated Gaussian distributions for all the parameters (Udwadia, 1989). The first-order marginal distributions for $k_A, k_B, n_f$ and $D_P$ are plotted in Fig. 1, with the appropriate choice of expectation and variance.

![Figure 1: Probability distributions of the flocculation parameters.](image)

Finally, the consideration of appropriate correlation between the input parameters is particularly important in sensitivity analyses. We therefore compute, again using the values from Table 1, the correlation between the four parameters. The unbiased estimate of the correlation is $\sigma^2_{XX'} = \sum_{i=1}^{N} (x_i - \mu_X)(x'_i - \mu'_X)/(N - 1)$, where the $x_i$ and $x'_i$ are the $N$ realizations of the random variables $X$ and $X'$, respectively. The correlation between the fractal dimension and the other parameters cannot be estimated, because its distribution is determined from a
different database. In that case, the Maximum Entropy Principle states that the fractal dimension should be considered as independent from the other parameters. The matrix of correlation coefficients $R$ between the different parameters (in the order $k_A$, $k_B$, $n_f$, $D_P$) is therefore:

$$R = \begin{bmatrix}
1 & 0 & 0 & -0.4 \\
0 & 1 & 0 & -0.3 \\
0 & 0 & 1 & 0 \\
-0.4 & -0.3 & 0 & 1
\end{bmatrix}$$  \hspace{1cm} (15)

Note that the lack of correlation between the aggregation and breakup parameters $k_A$ and $k_B$ arises from the data analyzed and has not been assumed as vanishing. We denote, for later use, the non-zero extra-diagonal values of the matrix as $R_{AD} = -0.4$ and $R_{BD} = -0.3$. When considering $k_B'$ instead of $k_B$ (and using the values in Table 3), one obtains $R_{B'D} = -0.46$, and the rest of the matrix remains unchanged.

The generation of realizations of the flocculation parameters following the identified first-order marginal laws (Fig. 1), and correlation matrix $R$ is discussed in Appendix A. This concludes the construction of appropriate probability distributions for the flocculation parameters. The next two sections report two sensitivity analyses for different QoIs.

4. Sensitivity analysis for equilibrium diameter

We now turn to the first SA. We want to analyze the influence of the flocculation parameters on the equilibrium diameter defined by Eq. (9), and approximated by Eq. (10). At least two reasons make this quantity important in applications: (i) it is directly measurable in identification experiments
(through direct image analysis, or through diffraction patterns (Bushell et al.,
2002)), and (ii) it has a strong influence on the settling velocity, which is an-
other measurable quantity, and one of the most important QoIs for industrial
applications.

All the computations in this section are done with $G = 10 \, \text{s}^{-1}$ and $c =
1 \, \text{g/L}$. They were repeated with different pairs of values with no qualitative
influence on the results (except for very large values of $G$ or very small
values of $c$ that induce $D_\infty \approx D_P$). The other numerical parameters are
$F_y = 10^{-10} \, \text{kg.m/s}^2$, $\mu = 10^{-3} \, \text{kg/m/s}$ and $\rho_s = 2650 \, \text{kg/m}^3$.

4.1. 1D and 2D histograms

We first propagate the probability laws of $k_A$, $k_B$, $n_f$ and $D_P$ through
the formula of Eq. (9). We use the Monte Carlo approach, with $10^7$ samples
and approximate the solution of Eq. (9) for each sample using the bisection
method. We obtain an approximation of the first-order marginal density of
the equilibrium diameter, that is represented in Figure 2(a). Note that this
histogram is obtained considering homogeneous intervals in log-scale. With
that choice, log-normal variables have densities that look gaussian. This
choice makes visualization easier for distributions considered here. It will be
followed in this paper for all variables, except the fractal dimension $n_f$, for
which homogeneous intervals will be considered in true scale. The probability
density obtained for the approximation of the equilibrium diameter given by
Eq. (10) is also shown in Figure 2(a). As expected, the true value and
its approximation only differ when the equilibrium diameter $D_\infty$ is of the
same order of magnitude as the primary diameter $D_P$, in which case the
approximation formula is not valid.
We plot second-order marginal densities between the QoI and each of the input parameters in turn on Figure 3. As before, densities are estimated for $k_A$, $k_B$, $D_P$ and $D_\infty$ using intervals homogeneous in log-scale, while the density with respect to $n_f$ is estimated using homogeneous intervals in true scale. Visually, it seems that $D_\infty$ is uncorrelated with $n_f$ as the second-order marginal density appears as a tensorization of the first-order marginal densities. The particle diameter seems only slightly correlated while the two flocculation parameters $k_A$ and $k_B$ seem to be more correlated.

4.2. Sensitivity indices

We now turn to quantitative measures of sensitivity. Correlation provides a measure of the strength of the linear relationship between the output $D_\infty$ and one of the input parameters, one after the other. The Pearson correlation
Figure 3: Second-order marginal densities of equilibrium diameter $D_\infty$ (considering the parameterization with non-dimensional $k_B$) with each of the input parameters. The contour lines correspond to the successive 10% quantiles.
coefficient between $D_\infty$ and a parameter $X$ is defined by:

$$CC_X = \frac{1}{\sigma_\infty \sigma_X} \sum_{i=1}^{N} (D_{\infty,i} - \mu_\infty)(x_i - \mu_X),$$

(16)

where $\mu_\infty = \frac{\sum_{i=1}^{N} D_{\infty,i}}{N}$ and $\sigma^2_\infty = \frac{\sum_{i=1}^{N} (D_{\infty,i} - \mu_\infty)^2}{N - 1}$, $D_{\infty,i}$ are the samples of the equilibrium diameter and $N = 10^7$ in that case. A vanishing correlation only indicates the absence of a linear relation between $D_\infty$ and the considered parameter. It does not preclude the existence of a well-defined nonlinear relationship between them. Values of the Pearson correlation coefficient for our SA are reported in Table 5 (first line). The correlation coefficients do not seem to provide any particular insight in the example we consider here.

When the relationship between the input and the output is non-linear, but monotonic, rank transformations may provide interesting insights into input-output relationship. Rank transformation consists in replacing the values of the parameters and output by their respective ranks in the data set of values. Hence, any non-linear monotonic relationship becomes a straight line. The computation of correlation coefficients for these rank-transformed parameters and outputs yields the so-called Rank Correlation Coefficient (RCC), that is reported in Table 5 (second line). Those parameters are more interesting here as there is clear evidence of correlation between $k_A$ (positive correlation) and $k_B$ (negative correlation) and the equilibrium diameter. On the other hand, these coefficients confirm the lack of correlation between $n_f$ or $D_P$ and the equilibrium diameter.

Finally, we perform an analysis of variance on this model, through the computation of Sobol Indices (Sobol, 2001). These indices predict the influ-
Table 5: Correlations, rank-correlations, and sensitivity indices of the equilibrium diameter $D_{\infty}$ with respect to flocculation parameters, considering the parameterization with non-dimensional $k_B$ (upper half of the table) or with dimensional $k'_B$ (lower half of the table). The second-order Sobol indices that are not indicated are close to zero.

<table>
<thead>
<tr>
<th></th>
<th>$k_A$</th>
<th>$k_B$</th>
<th>$n_f$</th>
<th>$D_P$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Pearson Correlation Coefficients (CC)</td>
<td>0.02</td>
<td>-0.02</td>
<td>-0.00</td>
<td>0.02</td>
</tr>
<tr>
<td>Rank Correlation Coefficient (RCC)</td>
<td>0.61</td>
<td>-0.74</td>
<td>-0.03</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol First-Order Sensitivity Indices (SI)</td>
<td>0.00</td>
<td>0.60</td>
<td>0.00</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol Second-Order Sensitivity Indices</td>
<td>$k_A-k_B$ : 0.36</td>
<td>-</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Sobol Third-Order Sensitivity Indices</td>
<td>0.60</td>
<td>0.00</td>
<td>0.97</td>
<td>0.97</td>
</tr>
<tr>
<td>Correlation Coefficients (CC)</td>
<td>0.00</td>
<td>0.00</td>
<td>-0.00</td>
<td>-0.00</td>
</tr>
<tr>
<td>Rank Correlation Coefficient (RCC)</td>
<td>0.12</td>
<td>-0.19</td>
<td>-0.93</td>
<td>0.08</td>
</tr>
<tr>
<td>Sobol First-Order Sensitivity Indices</td>
<td>0.00</td>
<td>0.01</td>
<td>0.00</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol Second-Order Sensitivity Indices</td>
<td>$k'_B-n_f$ : 0.45</td>
<td>-</td>
<td>-</td>
<td>-</td>
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<tr>
<td>Sobol Third-Order Sensitivity Indices</td>
<td>0.03</td>
<td>0.00</td>
<td>0.03</td>
<td>0.64</td>
</tr>
</tbody>
</table>
ence of the variance of each of the input parameters onto the variance of the output, either alone (first-order Sobol indices) or in combinations (higher-order Sobol indices). These indices are listed in Table 5. They indicate that the most influential parameter for the variance of the output is $k_B$ (highest first-order Sobol index and lowest third-order Sobol index). The parameter $k_A$ is also influential, but only through its influence coupled together with $k_B$. Note that, contrarily to the case of independent parameters, the Sobol indices do not sum to one when input parameters are correlated (as is the case here). Note also that third-order indices in Table 5 are reported in the columns of the one parameter (out of the four) which is not considered for the computation of that particular index. Hence, vanishing values in a column indicate the most influential parameters.

4.3. Choice of parameterization

As discussed in the introduction, one of the main interests of SA lies in its ability to potentially help in the choice of a parameterization in the view of experimental identification. We therefore complete the above sensitivity study with the corresponding analysis based on the alternative parameterization based on the dimensional breakup function $k_B'$. The first-order marginal density of $D_P$ is presented in Figure 2(b), the sensitivity indices and correlation coefficients are presented in the lower part of Table 5, and the second-order marginal densities are presented in Figure 4. Both the first- and second-order marginal densities appear as more complex than the counterparts obtained with the parameterization with the non-dimensional $k_B$. This suggests that the influence is more widely spread over the different parameters. It is confirmed by the values of the coefficients in Table 5. No
Figure 4: Second-order marginal densities of equilibrium diameter $D_\infty$ (considering the parameterization with dimensional $k'_B$) with each of the input parameters. The contour lines correspond to the successive 10% quantiles.
parameter is notably correlated with $D_\infty$, while the rank correlation coefficients are non-zero for all the parameters. In terms of variance analysis, it is noticeable that three of them have third-order parameters close to zero.

4.4. Discussion

The differences observed between the dimensional and non-dimensional approaches are the following: (i) the general structure of the second-order marginal densities to the tensorization of first-order marginal densities in the non-dimensional case, (ii) fewer parameters are correlated significantly with the equilibrium diameter in the non-dimensional case, and (iii) fewer parameters have influence on the variance of the equilibrium diameter in the non-dimensional case. All these (in particular the last point) mean that the identification of $k_B$ (or $k_A$ to a lesser extent) could be well constrained using equilibrium diameter measures and the non-dimensional model. For this identification procedure, and as they are very little influential, it also means that the values of $n_f$ and $D_P$ could be chosen quite freely. As the situation is not so simple with the dimensional model, the identification process would have to involve more variables (although $D_P$ could probably be left aside) and hence become more difficult. Finally, it should be noted that in the case of non-dimensional models, the recent improved model of the fractal dimension (see Maerz et al. (2011)) seem unnecessary as $n_f$ has little influence on this particular QoI. This will be discussed further in Section 5.

4.5. Additional correlation between fractal dimension and flocculation parameters
The Maximum Entropy approach used in Section 3.2 indicates that the rational choice when there is no evidence of dependency between random variables is independence. In order to give more weight to the conclusions of this paper, we nevertheless investigate in this section the possibility that $n_f$ be correlated to the other flocculation parameters. To do this, we modify the correlation matrix in Eq. (15) by adding a correlation coefficient $R_{An} = -0.5$ between $n_f$ and $k_A$ (both in the dimensional and non-dimensional cases) and correlations coefficient $R_{Bn} = R_{B'n} = 0.5$ between $n_f$ and $k_B$ and $n_f$ and $k_B'$, respectively. We have no argument for choosing these values rather than others, except that they verify the expected sign. Their amplitude is also large enough to test our hypothesis that the conclusions of the paper still hold when correlation with the fractal dimension is considered. The correlation and sensitivity indices in both these cases are reported in Table 6. Although the indices in the dimensional case seem harder to analyze than without correlation, this is not so in the non-dimensional case, where results are not significantly modified. This suggests that the conclusions we put forward in this paper are really due to the simplicity of the parameter-QoI relationship in the non-dimensional case rather than some particular case of independence of parameters.

5. Sensitivity analysis in a tidal forcing experiment

We now turn to a more advanced example of tidal flocculation, already described in Maerz et al. (2011), with comparison with experiments. In this example, the evolution of the diameter of flocs in a turbulence-controlled 0D system under cyclic loading is observed. The loading parameter $G$ follows
Table 6: Correlations, rank-correlations, and sensitivity indices of the equilibrium diameter $D_\infty$ with respect to flocculation parameters considering dependent fractal dimension $n_f$, and considering the parameterization with non-dimensional $k_B$ (upper half of the table) or with dimensional $k'_B$ (lower half of the table). The second-order Sobol indices that are not indicated are close to zero.

<table>
<thead>
<tr>
<th></th>
<th>$k_A$</th>
<th>$k_B$</th>
<th>$n_f$</th>
<th>$D_P$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Pearson Correlation Coefficients (CC)</td>
<td>0.02</td>
<td>-0.02</td>
<td>-0.04</td>
<td>0.02</td>
</tr>
<tr>
<td>Rank Correlation Coefficient (RCC)</td>
<td>0.62</td>
<td>-0.74</td>
<td>-0.74</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol First-Order Sensitivity Indices (SI)</td>
<td>0.00</td>
<td>0.63</td>
<td>0.00</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol Second-Order Sensitivity Indices</td>
<td>$k_A-k_B$ : 0.38</td>
<td>-</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td>Sobol Third-Order Sensitivity Indices</td>
<td>0.63</td>
<td>0.00</td>
<td>1.00</td>
<td>1.00</td>
</tr>
<tr>
<td>Correlation Coefficients (CC)</td>
<td>0.00</td>
<td>-0.00</td>
<td>-0.00</td>
<td>-0.00</td>
</tr>
<tr>
<td>Rank Correlation Coefficient (RCC)</td>
<td>0.50</td>
<td>-0.60</td>
<td>-0.96</td>
<td>0.12</td>
</tr>
<tr>
<td>Sobol First-Order Sensitivity Indices</td>
<td>0.00</td>
<td>0.00</td>
<td>0.00</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol Second-Order Sensitivity Indices</td>
<td>$k'_B-n_f$ : 5.84</td>
<td>-</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sobol Third-Order Sensitivity Indices</td>
<td>0.03</td>
<td>0.00</td>
<td>0.03</td>
<td>0.64</td>
</tr>
</tbody>
</table>
a cycle that reproduces what is observed in coastal waters during a tidal cycle. It is varied from $0 \text{ s}^{-1}$ to $12 \text{ s}^{-1}$ on a six-hour time-scale, and is shown in Figure 5. A 0D version (with $|\mathbf{w}| = |\mathbf{w}_s| = 0$) of Eq. (8) is solved, with a constant mass concentration $c = 93 \text{ mg/L}$, to obtain the evolution of the volume concentration $N$ as a function of time. Eq. (1) then provides the evolution of the diameter of the flocs as a function of time. The other numerical parameters are $F_y = 10^{-10} \text{ kg.m/s}^2$, $\mu = 10^{-3} \text{ kg/m/s}$ and $\rho_s = 2650 \text{ kg/m}^3$. The initial diameter is equal to the equilibrium diameter $D_\infty$ corresponding to the initial turbulence parameter $G = 1 \text{ s}^{-1}$.

Figure 5: Tidal forcing experiment: (upper figure) evolution of the forcing turbulence parameter; and (lower figure) evolution of the floc diameters (solid lines) and equilibrium diameters (dash-dotted lines) as a function of time, for three realizations of the input parameters.

A previous (local) SA performed on this example (Maerz et al., 2011) stated the fractal dimension as the most influential parameter of the floculation model for the average diameter. The authors used a model with a dimensional floc-breakup parameter $k'_B$. This statement is therefore coherent with the results of the (global) analysis performed in the previous section.
The objective of this section is to check whether the statement is conserved when we consider the flocculation model with a non-dimensional parameter $k_B$. This statement has important implications because it is one of the reasons why several authors proposed to refine the modeling of the fractal dimension and consider it to depend on diameter.

We therefore consider in the following the flocculation model with non-dimensional breakup parameter $k_B$. We then perform a SA with respect to the amplitude of the fluctuations of the diameter, the difference between the minimum and maximum diameters over one tidal cycle, namely over the period $T = [9 – 15]$ h. We use the Monte Carlo approach, with $10^5$ samples and approximate the solution of Eq. (8) for each sample using a Crank-Nicholson scheme for the time discretization. Note that, because of the inertia of the system, these extremal diameters are not equal to the equilibrium diameters corresponding to $G = 12$ s$^{-1}$ and $G = 1$ s$^{-1}$, respectively. Indeed, the differences in the dynamic behavior of three different samples of the input parameters can be observed on Figure 5. For some samples, the evolution of the diameter (solid lines) follows very closely the evolution of the equilibrium diameter (dashed-dotted lines), while for other samples, the diameter is almost constant throughout. This is related to the value of the equilibrium time $T_\infty$, defined in Eq (11).

We plot on Figure 6 the first-order marginal densities of the minimum and maximum densities, as well as that of the amplitude of fluctuations. On Figure 7, the second-order marginal densities for the amplitude of fluctuations with respect to the input parameters are plotted. Although they are not as simple as in the previous example (in particular, one appears bi-modal), the
general shapes of the second-order marginal densities do not seem to show any kind of prominence of the fractal dimension with respect to the other parameters. This is confirmed by the correlation and sensitivity indices reported in Table 7. The fractal dimension $n_f$ does appear in the sensitivity indices, but only in conjunction with the parameter $k_B$. The QoI is exactly as sensitive to $n_f$ as it is to $k_A$, and certainly less than $k_B$. The diameter of primary particles is still unimportant for the variance of the output.

Figure 6: Tidal forcing experiment: first-order marginal distributions of minimum and maximum diameters (left figure, solid and dashed lines, respectively), and of the amplitude of fluctuations $D_{\text{max}} - D_{\text{min}}$ (right figure).

We now turn back to the conclusions of the (local) SA in Maerz et al. (2011). When using the flocculation model with dimensional floc-breakup parameter (as Maerz et al. (2011) do), it seems reasonable to refine the modeling of the fractal dimension parameter. The conclusions of our previous global SA (although for a different QoI) seem to confirm the importance of the fractal dimension observed by Maerz et al. (2011). However, the analysis also
Figure 7: Tidal forcing experiment: second-order marginal densities of the amplitude of fluctuations $D_{\text{max}} - D_{\text{min}}$. The contour lines correspond to the successive 10% quantiles.
Table 7: Correlations, rank-correlations, and sensitivity indices of the amplitude of fluctuations $D_{\text{max}} - D_{\text{min}}$. The second-order Sobol indices that are not indicated are close to zero.

<table>
<thead>
<tr>
<th></th>
<th>$k_A$</th>
<th>$k_B$</th>
<th>$n_f$</th>
<th>$D_P$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Correlation Coefficients (CC)</td>
<td>0.01</td>
<td>-0.01</td>
<td>0.01</td>
<td>0.00</td>
</tr>
<tr>
<td>Rank Correlation Coefficient (RCC)</td>
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<td>-0.54</td>
<td>-0.07</td>
<td>-0.10</td>
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<td>Sobol First-Order Sensitivity Indices (SI)</td>
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<td>0.18</td>
<td>0.00</td>
<td>0.00</td>
</tr>
<tr>
<td>Sobol Second-Order Sensitivity Indices</td>
<td>$k_A$-$k_B$: 0.32</td>
<td>-</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td></td>
<td>$k_B$-$n_f$: 0.34</td>
<td>-</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>$k_B$-$D_P$: 0.01</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sobol Third-Order Sensitivity Indices</td>
<td>0.45</td>
<td>0.00</td>
<td>0.57</td>
<td>1.25</td>
</tr>
</tbody>
</table>

seems to suggest that using the flocculation model with a non-dimensional parameter attenuates strongly the influence of the fractal dimension. In that case, the fractal dimension would not need to be modeled more precisely, as Maerz et al. (2011) and others do, for instance by considering it to be a function of floc diameter. As it is much simpler to use the model based on a non-dimensional parameter rather than to make the fractal dimension dependent on the floc diameter, it is advisable to follow the former approach.

6. Conclusion

We have reported in this paper a global SA for two different parameterizations of a fractal-based turbulence-induced flocculation model. Using the equilibrium diameter as first QoI, we have shown that the two parameterizations yielded very different results in terms of parameter-QoI correlations and
sensitivities. Namely, the parameterization in terms of a non-dimensional floc breakup parameter showed strong sensitivity with respect to two parameters only (floc breakup and aggregation parameters), while the parameterization in terms of the dimensional floc breakup parameter presented a more widely distributed sensitivity and more complex correlation patterns. The conclusion of this analysis is that the former parameterization is probably much better suited for an identification experiment based on the measure of floc diameter, provided the fractal dimension and diameter of the primary particle can be measured through other experiments. It also indicates that the equilibrium diameter can be approximated very well by a function of only the corresponding two parameters.

A second SA considered a 0D experiment reproducing a tidal flocculation process, taken from Maerz et al. (2011). A previous SA by these authors on the characteristic diameter suggested that the fractal dimension was extremely important for this experiment. Our analysis observes that this is only true for the parameterization with dimensional floc breakup parameter. Using the non-dimensional parameterization, as recommended by our previous analysis, the diameter fluctuations appears more sensitive to the floc breakup parameters. The results of Maerz et al. (2011) motivated in particular the introduction of more complex models of the fractal dimension, varying with the diameter. Our analysis seems to indicate that this might be unnecessary, provided that the appropriate parameterization of the model is chosen.

Finally, it should be noted that other QoIs could be considered, and might lead to different conclusions. In particular, Verney et al. (2011) report local
sensitivity analyses focused on the settling velocity that seem to indicate that
the primary diameter $D_P$ might be more important for this QoI than in the
SA reported here.

Appendix A. Generation of samples of flocculation parameters

We describe here a method for generating samples of the flocculation
parameters with the chosen first-order marginal distributions and correlation.
We consider a generic random vector of dimension 4, with given first-order
marginal distributions for each of its coordinates, and a given correlation
matrix. The cumulative distributions are denoted $\Phi_i$, for $1 \leq i \leq 4$, and the
correlation matrix is denoted $R$.

We start with 4 independent unit centered gaussian random variables $G_i$
grouped in a vector $G$, and a correlation matrix $R^*$ (the value of which will
be discussed below). We then construct the 4-dimensional correlated random
vector $U = G\sqrt{R^*}$, where $\sqrt{R^*}$ comes from the Choleski decomposition of
$R^*$. Finally, we transform each of the random variables in the vector $U$ into
the appropriate first-order marginal distribution using the classical transfor-
mation $K_i = \Phi_i^{-1}(\Phi_0(U_i))$, where $\Phi_0$ is the cumulative density function of
a unit centered gaussian random variable. By doing this last transformation, we unfortunately modify the correlation structure of $U$ (which was by
definition $R^*$) into something else.

The proposed approach is therefore iterative. We start from an initial
guess of $R^*$ (the obvious choice being $R^* = R$), compute numerically the
resulting correlation for $K$ and iterate until that correlation is close to $R$.

Using this approach and the statistical parameters and laws discussed in
Section 3.2, we converged to the following correlation matrices:

- for the non-dimensional model (and $n_f$ independent of the other parameters)
  \[
  R^* = \begin{bmatrix}
  1 & 0 & 0 & -0.44 \\
  0 & 1.08 & 0 & -0.35 \\
  0 & 0 & 1 & 0 \\
  -0.44 & -0.55 & 0 & 1
  \end{bmatrix}
  \] (A.1)

- for the dimensional model (and $n_f$ independent of the other parameters)
  \[
  R^* = \begin{bmatrix}
  1 & 0 & 0 & -0.44 \\
  0 & 1.08 & 0 & -0.55 \\
  0 & 0 & 1 & 0 \\
  -0.44 & -0.55 & 0 & 1
  \end{bmatrix}
  \] (A.2)

- for the non-dimensional model (and $n_f$ dependent on the other parameters)
  \[
  R^* = \begin{bmatrix}
  1 & 0 & -0.52 & -0.44 \\
  0 & 1.08 & 0.57 & -0.35 \\
  -0.52 & 0.57 & 1 & 0 \\
  -0.44 & -0.35 & 0 & 1
  \end{bmatrix}
  \] (A.3)

- for the dimensional model (and $n_f$ dependent on the other parameters)
  \[
  R^* = \begin{bmatrix}
  1 & 0 & -0.53 & -0.44 \\
  0 & 1.08 & 0.55 & -0.53 \\
  -0.53 & 0.55 & 1 & 0 \\
  -0.44 & -0.53 & 0 & 1
  \end{bmatrix}
  \] (A.4)
Using these matrices to correlate the gaussian germs allows to generate samples of the parameters following the chosen first-order marginal laws as well as having the right correlation structure.

Acknowledgements

The authors would like to thank the support of PETROBRAS Technological Program on Basin Modeling in the name of its general coordinator, Dr. Marco Moraes. We also acknowledge the fruitful discussions within the program with Dr. Paulo Paraizo and Prof. Ben Kneller. Partial support was also provided by MCT/CNPq and FAPERJ.


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