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ON THE EXIT TIME FROM AN ORTHANT FOR BADLY ORIENTED RANDOM WALKS

RODOLPHE GARBIT

ABSTRACT. In a recent paper, K. Raschel and R. Garbit proved that the exponential decreasing rate of the probability that a random walk (with all exponential moments) stays in a d -dimensional orthant is given by the minimum on this orthant of the Laplace transform of the random walk increments, as soon as this minimum exists. In other cases, the random walk is “badly oriented”, and the exponential rate may depend on the starting point x . We prove here that this rate is nevertheless asymptotically equal to the infimum of the Laplace transform, as some selected coordinates of x tend to infinity.

1. INTRODUCTION

1.1. **Context.** This work is a continuation of the paper [6] in which the authors studied the exponential decreasing rate of the probability that a random walk (with all exponential moments) stays in a d -dimensional orthant, and found that this rate is equal to the minimum on this orthant of the Laplace transform of the random walk increments, as soon as this minimum exists. Here, we extend this result so as to cover also the remaining cases where the random walk is “badly oriented” – a terminology that will be explained later – with respect to the orthant. In order to be more specific, let us introduce some notations.

For any fixed probability distribution μ on \mathbb{R}^d , let \mathbb{P}_μ^x denote the probability measure on \mathbb{R}^∞ under which the canonical process $(S_0, S_1, \dots, S_n, \dots)$ is a random walk started at x (meaning that $S_0 = x$ a.s.) whose independent increments $S_{n+1} - S_n$ have distribution μ . We are interested in the tail distribution of the exit time

$$\tau_Q = \inf\{n \geq 1 : S_n \notin Q\}$$

of the random walk from the positive orthant

$$Q = \{x \in \mathbb{R}^d : x_i \geq 0, i = 1 \dots d\},$$

where x_i denotes the i th coordinate of x in the standard basis (e_1, e_2, \dots, e_d) .

For random walks without drift, the precise asymptotics

$$\mathbb{P}_\mu^x(\tau_Q > n) = c\rho^n n^{-\alpha}(1 + o(1)), \quad n \rightarrow \infty$$

was derived (for more general cones) by Denisov and Wachtel in [3] from the corresponding tail distribution for Brownian motion by using a strong approximation theorem of random walks by Brownian motion. They also obtained a local limit theorem from which Duraj

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could derive in [4] the precise asymptotics for random walks with drift in the special case where the global minimum on Q of the Laplace transform of μ is reached at an interior point of Q .

Our goal in [6] was to understand how the drift vector $m = \mathbb{E}(\mu)$ and the variance-covariance matrix $\Gamma(\mu)$ would influence the exponential decreasing rate

$$\rho_x = \liminf_{n \rightarrow \infty} \mathbb{P}_\mu^x(\tau_Q > n)^{1/n}.$$

Not surprisingly, for random walks with all exponential moments, it turned out that ρ_x also depends on other moments of μ , and the way this dependence occurs is via its Laplace transform

$$L_\mu(z) = \int_{\mathbb{R}^d} e^{\langle z, y \rangle} \mu(dy).$$

Indeed, the main result in [6] is that

$$(1) \quad \rho_x = \min_Q L_\mu,$$

for all x far enough from the boundary of Q , as soon as this minimum exists. Note that in this case, there is essentially no dependence in the starting point x . It was also proved there that, for a truly d -dimensional distribution μ , i.e. a distribution whose support is not included in any linear hyperplane, the condition that μ reaches a global minimum on Q is equivalent to the following geometric condition:

(H) The support of μ is not included in any half-space $u^- = \{x \in \mathbb{R}^d : \langle x, u \rangle \leq 0\}$ with $u \in Q \setminus \{0\}$.

Random walks with a distribution μ that does not fulfill condition (H) are called *badly oriented*. In this case, the exponential rate ρ_x may depend on the starting point x .

Example 1. Consider the 2-dimensional lattice distribution μ defined by

$$\mu(1, -1) = \mu(-1, 1) = q, \quad \mu(-1, -1) = p, \quad p + 2q = 1, \quad p, q > 0.$$

The corresponding random walk is badly oriented since the support of μ is included in $(1, 1)^-$. Its Laplace transform satisfies the relations

$$L_\mu(i, j) = 2q \cosh(i - j) + p e^{-(i+j)} > 2q,$$

and

$$\lim_{i \rightarrow +\infty} L_\mu(i, i) = 2q.$$

Therefore, the infimum $2q$ of L_μ on Q is not a minimum. It is proved in [6] that, for $x = (i, j) \in \mathbb{N}_+^2$ with $i + j = 2N$,

$$\rho_x = 2q \cos\left(\frac{\pi}{2N + 2}\right).$$

Thus ρ_x depends on x , but we nevertheless observe that

$$\lim_{\|x\| \rightarrow \infty} \rho_x = 2q = \inf_Q L_\mu.$$

The aim of this paper is to explain this phenomenon by giving a “universal” result that applies to both well and badly oriented random walks. More precisely, we shall prove that the equality (1) is in fact a particular case of the following equality

$$\lim \rho_x = \inf_Q L_\mu,$$

where the interpretation of the limit symbol depends on a linear subspace associated with μ that we call the *reduced support* of μ .

The central idea in [6] was to perform a standard Cramér transformation (i.e. an exponential change of measure) based on the point x_0 where L_μ reaches its minimum on Q . The main novelty here is that we provide a way to achieve this transformation when x_0 is “at infinity”.

1.2. Reduction. The assumption (H) was used in [6] to ensure the existence of a global minimum of L_μ on Q . The implication follows easily from the following property of the Laplace transform (see [6, Lemma 3]):

$$\lim_{t \rightarrow +\infty} L_\mu(z + tu) = \begin{cases} +\infty & \text{if } \mu(u^-) < 1 \\ \int_{u^\perp} e^{\langle z, y \rangle} \mu(dy) & \text{if } \mu(u^-) = 1, \end{cases}$$

where u^\perp denotes the hyperplane orthogonal to u . Indeed, for a distribution satisfying assumption (H), the Laplace transform tends to infinity in each direction $u \in Q \cap \mathbb{S}^{d-1}$. Since this function is convex, this implies by a compactness argument that L_μ is coercive in Q , and the existence of a global minimum follows.

Suppose now that the random walk is badly oriented, i.e. that there exists a direction $u_1 \in Q \cap \mathbb{S}^{d-1}$ such that $\mu(u_1^-) = 1$. Then, the “best” way to stay in Q is certainly not to perform a step outside of u_1^\perp . Thus, it seems natural to compare our original random walk with a conditioned version of it, namely the random walk with distribution $\nu(*) = \mu(*|u_1^\perp)$. Since this new random walk may still be badly oriented, we shall construct by induction a *reduced support* V in the following way:

Definition 2. An r -tuple $(u_1, u_2, \dots, u_r) \in Q^r$ is *admissible* if

- (i) the vectors u_1, u_2, \dots, u_r are linearly independent, and
- (ii) they satisfy the following relations:

$$\begin{cases} \mu(u_1^-) = 1 \\ \mu(u_2^- \cap u_1^\perp) = \mu(u_1^\perp) \\ \vdots \\ \mu(u_r^- \cap (u_{r-1}^\perp \cap \dots \cap u_1^\perp)) = \mu(u_{r-1}^\perp \cap \dots \cap u_1^\perp). \end{cases}$$

An admissible r -tuple (u_1, u_2, \dots, u_r) is *maximal* if there is no $u \in Q$ such that $(u_1, u_2, \dots, u_r, u)$ be admissible.

If there exists an admissible tuple (i.e. if the random walk is badly oriented), then there exist maximal tuples, and the linear subspace $[u_1, u_2, \dots, u_r]$ generated by a maximal tuple does not depend on the specific maximal tuple chosen (Lemma 8).

Definition 3. The orthogonal complement

$$V = [u_1, u_2, \dots, u_r]^\perp$$

of the subspace generated by any maximal tuple is called the *reduced support* of μ . By convention, if there is no admissible tuple, we set $V = \mathbb{R}^d$.

With the reduced support V , we associate the set I of indices $i \in \llbracket 1, d \rrbracket$ such that $e_i \in V$, and the cone

$$V^+ = \{x \in V : \langle x, e_i \rangle \geq 0, \forall i \in I\}.$$

Then (Lemma 9) the infimum on Q of the Laplace transform of μ is given by the relation

$$(2) \quad \inf_Q L_\mu = \inf_{v \in V^+} \int_V e^{\langle v, y \rangle} \mu(dy).$$

If $\inf_Q L_\mu = 0$, then

$$\rho_x = \liminf_{n \rightarrow \infty} \mathbb{P}_\mu^x(\tau_Q > n)^{1/n} = 0$$

for all $x \in Q$, since $\inf_Q L_\mu$ is a universal upper bound for ρ_x (see equation (5)). Thus, from now on, we will exclude this degenerate case by assuming

$$\inf_Q L_\mu > 0.$$

Therefore $\mu(V) > 0$ and equality (2) can be rewritten as

$$(3) \quad \inf_Q L_\mu = \mu(V) \inf_{V^+} L_{\mu|V},$$

where $\mu|V$ denotes the conditional distribution $\mu|V(*) = \mu(*|V)$. Now, the ‘‘maximality’’ of V (in the sense of Definition 2) ensures that the conditioned random walk with distribution $\mu|V$ is well oriented with respect to V^+ , i.e. that the infimum on V^+ is in fact a minimum (Lemma 10), so that a Cramér transformation can be applied to this conditioned random walk. We interpret this as a Cramér transformation at infinity.

1.3. Main result. In what follows, μ is any probability distribution on \mathbb{R}^d with all exponential moments. To avoid trivialities, we assume

$$\inf_Q L_\mu > 0.$$

We denote by V the reduced support of μ (see Definitions 2 and 3), and define

$$I = \{i \in \llbracket 1, d \rrbracket : e_i \in V\}, \quad I^\perp = \{i \in \llbracket 1, d \rrbracket : e_i \in V^\perp\},$$

and, for x in Q ,

$$d(x) = \min_{i \notin I \cup I^\perp} x_i,$$

with the convention that $d(x) = \infty$ when $I \cup I^\perp = \llbracket 1, d \rrbracket$. Finally, we set

$$Q_\delta = Q + \delta(1, 1, \dots, 1) = \{x \in \mathbb{R}^d : x_i \geq \delta, \forall i \in \llbracket 1, d \rrbracket\}.$$

We are now in position to state our main result.

Theorem 4. *There exists $\delta \geq 0$ such that*

$$\lim_{\substack{d(x) \rightarrow \infty \\ x \in Q_\delta}} \rho_x = \inf_Q L_\mu.$$

Let us illustrate this theorem with some basic examples.

Example 5. If μ satisfies assumption (H), then $V = \mathbb{R}^d$, $I = \llbracket 1, d \rrbracket$, $V^+ = Q$, $I^\perp = \emptyset$, and $d(x) = \infty$. Thus we recover the non-asymptotic theorem of [6]: There exists $\delta \geq 0$ such that

$$\rho_x = \inf_Q L_\mu,$$

for all $x \in Q_\delta$.

Example 6. Consider the 2-dimensional lattice distribution μ defined by

$$\mu(-1, 0) = \alpha, \quad \mu(0, 1) = \beta, \quad \mu(0, -1) = \gamma, \quad \alpha + \beta + \gamma = 1, \quad \alpha, \beta, \gamma > 0.$$

The associated random walk is badly oriented since $\mu(e_1^-) = 1$. Now, (e_1) is maximal since $\beta, \gamma > 0$, therefore $V = e_1^\perp$, $I = \{2\}$, $V^+ = \{(0, j) : j \geq 0\}$, $I^\perp = \{1\}$ and $d(x) = \infty$, meaning that

$$\rho_x = \inf_Q L_\mu,$$

for all $x \in Q_\delta$, for some constant $\delta \geq 0$. Let us compute the value of ρ_x . The Laplace transform of μ is given by

$$L_\mu(i, j) = \alpha e^{-i} + \beta e^j + \gamma e^{-j}.$$

Minimizing first on $i \geq 0$ leads to the relation

$$\inf_Q L_\mu = \inf_{j \geq 0} (\beta e^j + \gamma e^{-j}) = (\beta + \gamma) \inf_{j \geq 0} (\beta' e^j + \gamma' e^{-j}),$$

where $\beta' = \beta/(\beta + \gamma)$ and $\gamma' = \gamma/(\beta + \gamma)$ sum up to 1. Notice that the above formula corresponds exactly to equality (3) with $V = e_1^\perp$. An easy computation shows that

$$\inf_{j \geq 0} (\beta' e^j + \gamma' e^{-j}) = \begin{cases} 2\sqrt{\beta'\gamma'} & \text{if } \gamma' \geq \beta' \\ 1 & \text{else.} \end{cases}$$

Hence

$$\rho_x = \inf_Q L_\mu = \begin{cases} 2\sqrt{\beta\gamma} & \text{if } \gamma \geq \beta \\ \beta + \gamma & \text{else.} \end{cases}$$

Example 7. Consider the 2-dimensional lattice distribution μ defined by

$$\mu(-1, -1) = \alpha, \quad \mu(-1, 1) = \beta, \quad \mu(1, -1) = \gamma, \quad \alpha + \beta + \gamma = 1, \quad \alpha, \beta, \gamma > 0.$$

The associated random walk is badly oriented since the support of μ is included in $(1, 1)^-$. Here, $V = (1, 1)^\perp$, $I = I^\perp = \emptyset$, $V^+ = V$ and $d(x) = \min\{x_1, x_2\}$. Therefore, we obtain

$$\lim_{x_1, x_2 \rightarrow \infty} \rho_x = \inf_Q L_\mu = 2\sqrt{\beta\gamma},$$

as the reader may check.

The rest of the paper is organized as follows: In Section 2, we present the proof of Theorem 4. The construction of the reduced support enables us to perform a Cramér's transformation at infinity, and then compare our initial exponential decreasing rate with that of a conditioned random walk whose distribution is more "favorable". The exponential rate of this conditioned random walk is then analysed in Section 3. The Appendix at the

end of the paper provides some material on polyhedra that is needed in a technical lemma and for which no reference were found.

2. PROOF OF THEOREM 4

The first subsection recalls the strategy used in [6] in order to obtain the exponential decreasing rate in the case where the random walk is well oriented. Though not necessary, its reading is recommended in order to become acquainted with the basic ideas of the proof of Theorem 4. We will then present the construction and the main properties of the reduced support, and finally use it to perform our Cramér's transformation at infinity in order to conclude the proof of Theorem 4.

2.1. Sketch of proof for well oriented random walks. The reader may know, or at least should let his intuition convince himself that $\rho_x = 1$ when the drift vector $m = \mathbb{E}(\mu)$ belongs to the orthant Q . In other cases, the basic idea is to carry out an exponential change of measure in order to transform the drift vector m into a new one, m_0 , that belongs to Q . To do this, fix any $x_0 \in \mathbb{R}^d$ and define

$$\mu_0(dy) = \frac{e^{\langle x_0, y \rangle}}{L_\mu(x_0)} \mu(dy).$$

Clearly, μ_0 is a probability distribution with Laplace transform

$$L_{\mu_0}(z) = \frac{L_\mu(z + x_0)}{L_\mu(x_0)}.$$

Furthermore, it is well-known that the expectation of a probability distribution is equal to the gradient of its Laplace transform evaluated at 0. Thus

$$m_0 = \mathbb{E}(\mu_0) = \nabla L_\mu(x_0) / L_\mu(x_0).$$

If L_μ reaches a global minimum on Q at x_0 , then its partial derivatives are non-negative at x_0 and, therefore, the drift m_0 belongs to Q . So, the distribution μ_0 is "nice" in the sense that the value of the exponential rate for a random walk with that distribution is known. The link between the distribution of our canonical random walk under \mathbb{P}_μ^x and $\mathbb{P}_{\mu_0}^x$ is expressed via Cramér's formula (see [6, Lemma 6] for example) which leads in our case to the relation

$$(4) \quad \mathbb{P}_\mu^x(\tau_Q > n) = L_\mu(x_0)^n e^{\langle x_0, x \rangle} \mathbb{E}_{\mu_0}^x \left(e^{-\langle x_0, S_n \rangle}, \tau_Q > n \right).$$

From this, we already notice that for all $x_0 \in Q$ (being a minimum point or not),

$$\limsup_{n \rightarrow \infty} \mathbb{P}_\mu^x(\tau_Q > n)^{1/n} \leq L_\mu(x_0),$$

since $\langle x_0, S_n \rangle \geq 0$ as soon as $S_n \in Q$. Thus, the upper bound

$$(5) \quad \limsup_{n \rightarrow \infty} \mathbb{P}_\mu^x(\tau_Q > n)^{1/n} \leq \inf_Q L_\mu$$

holds for any probability distribution μ (with all exponential moments).

In order to obtain the corresponding lower bound in the case where L_μ reaches a global minimum on Q at x_0 , we use formula (4) with our nice distribution μ_0 . What is not so nice here is the $\exp(-\langle x_0, S_n \rangle)$ term, which could become very small since (S_n) has a drift

$m_0 \in Q$. But, in fact, the growth of $\langle x_0, S_n \rangle$ can be controled thanks to the following observation: Let K be the set of indices i such that $x_0^{(i)} > 0$. Since x_0 belongs to Q , the other coordinates are equal to zero, and

$$\langle x_0, S_n \rangle = \sum_{i \in K} x_0^{(i)} S_n^{(i)}.$$

Furthermore, if $x_0^{(i)} > 0$ for some index i , then 0 is a local minimum of the partial function $t \in [-x_0^{(i)}, +\infty) \mapsto L_\mu(x_0 + te_i)$ and therefore $m_0^{(i)} = 0$, because it is proportional to $\partial_{x_i} L_\mu(x_0) = 0$. So, the coordinates of the random walk that we need to control have zero mean. Since \sqrt{n} is a natural bound for the norm of a square integrable centered random walk, one can expect that adding the constraint

$$(6) \quad \max_{i \in K} |S_n^{(i)}| \leq \sqrt{n}$$

will not decrease too much the probability on the right-hand side of equation (4). As proved in [6] this happens to be true: For any distribution μ satisfying assumption (H),

$$\liminf_{n \rightarrow \infty} \mathbb{P}_{\mu_0}^x \left(\max_{i \in K} |S_n^{(i)}| \leq \sqrt{n}, \tau_Q > n \right)^{1/n} = 1.$$

Under the constraint (6), the exponential term $\exp(-\langle x_0, S_n \rangle)$ in (4) is bounded from below by $\exp(-\|x_0\|\sqrt{n})$, a term that disappears in the n th root limit. Thus, we obtain the lower bound

$$\rho_x \geq L_\mu(x_0) \liminf_{n \rightarrow \infty} \mathbb{P}_{\mu_0}^x \left(\max_{i \in K} |S_n^{(i)}| \leq \sqrt{n}, \tau_Q > n \right)^{1/n} = L_\mu(x_0),$$

and this concludes the analysis.

2.2. Construction and properties of the reduced support. In this section, we prove the uniqueness of the reduced support of μ (Definitions 2 and 3) and establish some of its properties. When μ satisfy assumption (H), the reduced support V is equal to \mathbb{R}^d – by definition – and all properties listed below are trivial, as the reader may check. Therefore, in what follows, we assume the existence of a direction $u_1 \in Q \cap \mathbb{S}^{d-1}$ such that $\mu(u_1^-) = 1$. Thus, there exists at least an admissible tuple, namely (u_1) , and the existence of a maximal tuple follows immediately.

First, we prove the uniqueness¹ of the linear space generated by maximal admissible tuples.

Lemma 8. *Any two maximal admissible tuples generate the same linear space.*

Proof. Let (u_1, u_2, \dots, u_r) and (v_1, v_2, \dots, v_s) be two maximal admissible tuples and let

$$A = [u_1, u_2, \dots, u_r] \quad \text{and} \quad B = [v_1, v_2, \dots, v_s]$$

denote the linear spaces generated by this two tuples, respectively. Since (u_1, u_2, \dots, u_r) is maximal, $(u_1, u_2, \dots, u_r, v_1)$ is not admissible. On the other hand, $\mu(v_1^-) = 1$ so that

$$\mu(v_1^- \cap A^\perp) = \mu(A^\perp),$$

¹This property will not be used, but it is a very natural question since the conclusion of Theorem 4 depends on this subspace.

and the tuple $(u_1, u_2, \dots, u_r, v_1)$ satisfies condition (ii) of Definition 2. Thus, the reason why this tuple fails to be admissible is that condition (i) is not satisfied, i.e. v_1 belongs to A . Now, suppose that v_1, v_2, \dots, v_{k-1} all belong to A for some $k \in \llbracket 2, s \rrbracket$, and denote $B_k = [v_1, v_2, \dots, v_{k-1}]$ the linear space they generate. By hypothesis, we have

$$A^\perp \subset B_k^\perp \quad \text{and} \quad \mu(v_k^- \cap B_k^\perp) = \mu(B_k^\perp).$$

This implies $\mu(v_k^- \cap A^\perp) = \mu(A^\perp)$ and therefore v_k must belong to A , otherwise $(u_1, u_2, \dots, u_r, v_k)$ would be admissible. By induction, we obtain the inclusion $B \subset A$, and the equality $A = B$ follows by interchanging the role of A and B . \square

Thanks to Lemma 8 we can define the *reduced support* V of μ as the orthogonal complement of the linear space generated by any maximal admissible tuple (Definition 3). In what follows, we fix a maximal admissible tuple (u_1, u_2, \dots, u_r) and set

$$V = [u_1, u_2, \dots, u_r]^\perp \quad \text{and} \quad V^+ = V \cap \left(\bigcup_{h \in V^\perp} Q - h \right).$$

The reduced support V and the cone V^+ play a fundamental role in our analysis. First of all, they provide a useful expression for the infimum of the Laplace transform.

Lemma 9. *The following equality holds:*

$$\inf_Q L_\mu = \inf_{v \in V^+} \int_V e^{\langle v, y \rangle} \mu(dy).$$

Proof. It follows from the orthogonal decomposition $\mathbb{R}^d = V \oplus V^\perp$ that

$$\begin{aligned} \inf\{L_\mu(x) : x \in Q\} &= \inf\{L_\mu(v+h) : (v, h) \in V \times V^\perp, v+h \in Q\} \\ &= \inf_{v \in V^+} \inf_{h \in V^\perp \cap (Q-v)} L_\mu(v+h). \end{aligned}$$

Therefore, we have to prove that for all $v \in V^+$

$$(7) \quad \inf_{h \in V^\perp \cap (Q-v)} L_\mu(v+h) = \int_V e^{\langle v, y \rangle} \mu(dy).$$

Let $v \in V^+$. First, for all $h \in V^\perp$, we notice that

$$L_\mu(v+h) \geq \int_V e^{\langle v+h, y \rangle} \mu(dy) = \int_V e^{\langle v, y \rangle} \mu(dy),$$

since $\langle h, y \rangle = 0$ for all $y \in V$. Now, pick $h \in V^\perp \cap (Q-v)$ (such a h exists since $v \in V^+$), and let $\lambda_1, \lambda_2, \dots, \lambda_r$ denote its coordinates with respect to the basis (u_1, u_2, \dots, u_r) of V^\perp . It is clear that $h = \sum_{k=1}^r \lambda_k u_k$ will still belong to $V^\perp \cap (Q-v)$ if one increases the value of any λ_i . Hence, the equality (7) will follow from

$$(8) \quad \lim_{\lambda_r \rightarrow \infty} \dots \lim_{\lambda_2 \rightarrow \infty} \lim_{\lambda_1 \rightarrow \infty} L_\mu \left(v + \sum_{k=1}^r \lambda_k u_k \right) = \int_V e^{\langle v, y \rangle} \mu(dy).$$

Since (u_1, u_2, \dots, u_r) is admissible, $\mu(u_1^-) = 1$ and consequently

$$L_\mu \left(v + \sum_{k=1}^r \lambda_r u_r \right) = \int_{u_1^\perp} e^{\langle v, y \rangle} e^{\sum_{k=2}^r \lambda_k \langle u_k, y \rangle} \mu(dy) + \int_{\{\langle u_1, y \rangle < 0\}} e^{\langle v, y \rangle} e^{\sum_{k=1}^r \lambda_k \langle u_k, y \rangle} \mu(dy).$$

By the dominated convergence theorem, the second integral on the right-hand side of the above equation goes to zero as λ_1 goes to infinity. Hence

$$\lim_{\lambda_1 \rightarrow \infty} L_\mu \left(v + \sum_{k=1}^r \lambda_r u_r \right) = \int_{u_1^\perp} e^{\langle v, y \rangle} e^{\sum_{k=2}^r \lambda_k \langle u_k, y \rangle} \mu(dy).$$

Now, by hypothesis, we have $\mu(u_2^- \cap u_1^\perp) = \mu(u_1^\perp)$, so that the same argument as above leads to

$$\lim_{\lambda_2 \rightarrow \infty} \lim_{\lambda_1 \rightarrow \infty} L_\mu \left(v + \sum_{k=1}^r \lambda_r u_r \right) = \int_{u_1^\perp \cap u_2^\perp} e^{\langle v, y \rangle} e^{\sum_{k=3}^r \lambda_k \langle u_k, y \rangle} \mu(dy).$$

The equality in (8) is obtained by using repeatedly the same argument. \square

Lemma 10. *Assume $\inf_Q L_\mu > 0$. Then:*

- (1) $\mu(V) > 0$.
- (2) For all $v \in V^+ \cap \mathbb{S}^{d-1}$, $\mu(v^- | V) < 1$.
- (3) The Laplace transform $L_{\mu|V}$ has a global minimum on V^+ .
- (4) $\inf_Q L_\mu = \mu(V) \min_{V^+} L_{\mu|V}$.

Proof. The formula in Lemma 9 shows that $\inf_Q L_\mu = 0$ as soon as $\mu(V) = 0$. The first item follows by contraposition.

Let $v \in V^+ \cap \mathbb{S}^{d-1}$. By definition of V^+ there exists $h \in V^\perp$ such that $u = v + h \in Q$. For any $y \in V$, we have $\langle v, y \rangle = \langle u, y \rangle$ and consequently

$$v^- \cap V = u^- \cap V.$$

So, it suffices to show that $\mu(u^- | V) < 1$ and the second assertion of the lemma will follow. By definition, $u = v + h$ (with $v \in V$, $v \neq 0$ and $h \in V^\perp$) is linearly independent of $V^\perp = [u_1, u_2, \dots, u_r]$ and belongs to Q . But by maximality, the tuple $(u_1, u_2, \dots, u_r, u)$ is not admissible. Thus, we must have $\mu(u^- \cap V) < \mu(V)$. This proves the second assertion of the lemma.

The third assertion follows from the second one since V^+ is a closed cone (see [6, Lemma 4] – note that the hypothesis (H1) is not used for the part of the lemma that we need here).

Finally, the last item is just a reformulation of the formula in Lemma 9. \square

We shall now give a very simple description of the cone V^+ associated with the reduced support V . To this end, we define

$$I = \{i \in \llbracket 1, d \rrbracket : e_i \in V\}.$$

Lemma 11. *The cone V^+ has the following expression:*

$$V^+ = \{x \in V : \langle x, e_i \rangle \geq 0, \forall i \in I\}.$$

Proof. We first note that $x \in \bigcup_{h \in V^\perp} Q - h$ if and only if there exists $h \in V^\perp$ such that $\langle x + h, e_i \rangle \geq 0$ for all i . But, for $i \in I$, $\langle x + h, e_i \rangle = \langle x, e_i \rangle$ since $e_i \in V$ and $h \in V^\perp$. Hence, the condition splits into:

- (1) $\langle x, e_i \rangle \geq 0$ for all $i \in I$, and
- (2) there exists $h \in V^\perp$ such that $\langle x + h, e_i \rangle \geq 0$ for all $i \notin I$.

Therefore, it remains to prove that the last condition holds for all x . To this end, recall that $V^\perp = [u_1, u_2, \dots, u_r]$ with $u_k \in Q$, and set $u_0 = \sum_{k=1}^r u_k$. Since $\langle u_0, e_i \rangle = \sum_{k=1}^r \langle u_k, e_i \rangle$ and $\langle u_k, e_i \rangle \geq 0$ for all k , we see that $\langle u_0, e_i \rangle \geq 0$ and equality occurs if and only if $\langle u_k, e_i \rangle = 0$ for all k , which means exactly that $e_i \in [u_1, u_2, \dots, u_r]^\perp = V$. Thus, by definition of the set I , we have $\langle u_0, e_i \rangle > 0$ for all $i \notin I$. For any fixed x , this property allows to find $\lambda > 0$ such that $\langle x, e_i \rangle + \lambda \langle u_0, e_i \rangle \geq 0$ for all $i \notin I$, so that condition (2) holds with $h = \lambda u_0$. This proves the lemma. \square

2.3. Comparison with the conditioned random walk with increments restricted to the reduced support. As mentioned earlier (see (5)), the infimum of L_μ on Q is always an upper bound for the exponential decreasing rate

$$\rho_x = \liminf_{n \rightarrow \infty} \mathbb{P}_\mu^x(\tau_Q > n)^{1/n}.$$

Therefore, our task is to show that it is also a lower bound for ρ_x (at least as $x \rightarrow \infty$ in the sense of Theorem 4).

Let V be the reduced support of μ . From now on, we assume that $\inf_Q L_\mu > 0$, so that $\mu(V) > 0$ (Lemma 10). Remember that we have introduced the reduced support with the idea that the best way to stay in Q was to never perform any step outside of V . Hence, denoting $\xi_1, \xi_2, \dots, \xi_n$ the increments of the random walk, it is natural to use the lower bound

$$(9) \quad \mathbb{P}_\mu^x(\tau_Q > n)^{1/n} \geq \mathbb{P}_\mu^x(\tau_Q > n, \xi_1, \xi_2, \dots, \xi_n \in V)^{1/n} = \mu(V) \mathbb{P}_{\mu|V}^x(\tau_Q > n)^{1/n}.$$

A look at the last formula of Lemma 10,

$$(10) \quad \inf_Q L_\mu = \mu(V) \min_{V^+} L_{\mu|V},$$

then explains our strategy: Theorem 4 will follow from a comparison between

$$\liminf_{n \rightarrow \infty} \mathbb{P}_{\mu|V}^x(\tau_Q > n)^{1/n} \quad \text{and} \quad \min_{V^+} L_{\mu|V}.$$

To simplify notations, set $\nu = \mu|V$. For all x in Q , write $x = v + w$ the orthogonal decomposition with respect to V and V^\perp . Then

$$\mathbb{P}_\nu^x(\tau_Q > n) = \mathbb{P}_\nu^v(S_1, S_2, \dots, S_n \in Q - w).$$

Under \mathbb{P}_ν^v , the random walk S_1, S_2, \dots, S_n almost surely belongs to V (since $v \in V$, and $\nu(V) = 1$). Thus, we have to focus our attention on the geometry of $(Q - w) \cap V$.

To this end, we recall that

$$I = \{i \in \llbracket 1, d \rrbracket : e_i \in V\}.$$

We also define

$$I^\perp = \{i \in \llbracket 1, d \rrbracket : e_i \in V^\perp\},$$

and, for $x \in Q$,

$$d(x) = \min_{i \notin I \cup I^\perp} x_i.$$

Notice that $\|a - x\| \leq d(x)$ implies that $a_i \geq 0$ for all $i \notin I \cup I^\perp$.

Let $V_1 = [e_i, i \in I]$ and write $V = V_1 \oplus V_2$ the orthogonal decomposition of V . Define the positive orthant of V_1 as

$$V_1^+ = \{y \in V_1 : \langle y, e_i \rangle \geq 0, \forall i \in I\},$$

and notice that Lemma 11 asserts that

$$V^+ = V_1^+ \oplus V_2.$$

For any $y \in V$, let $y^{(1)}$ and $y^{(2)}$ be the projections of y onto V_1 and V_2 , respectively.

Lemma 12. *For all $x = v + w \in Q$, holds the inclusion:*

$$\{y \in V : y^{(1)} \in V_1^+ \text{ and } \|y^{(2)} - v^{(2)}\| \leq d(x)\} \subset (Q - w) \cap V.$$

Proof. Let $y \in V$ be such that $y^{(1)} \in V_1^+$ and $\|y^{(2)} - v^{(2)}\| \leq d(x)$. We have to show that $y + w$ belongs to Q .

First of all, for any $i \in I^\perp$, we have $e_i \in V^\perp$ so that

$$\langle y + w, e_i \rangle = \langle v + w, e_i \rangle = x_i \geq 0.$$

Similarly, since $\langle y + w, e_i \rangle = \langle y^{(1)}, e_i \rangle$ for all $i \in I$, the condition $y^{(1)} \in V_1^+$ rewrites $\langle y + w, e_i \rangle \geq 0$ for all $i \in I$.

It remains to check that the conclusion also holds when $i \notin I \cup I^\perp$. To this end, we notice that

$$\|(v^{(1)} + y^{(2)} + w) - x\| = \|y^{(2)} - v^{(2)}\| \leq d(x).$$

Therefore, for all $i \notin I \cup I^\perp$,

$$\langle v^{(1)} + y^{(2)} + w, e_i \rangle \geq 0.$$

But for those indices i , we have $e_i \perp V_1$ and consequently

$$\langle y + w, e_i \rangle = \langle v^{(1)} + y^{(2)} + w, e_i \rangle \geq 0.$$

This concludes the proof of the lemma. \square

This lemma provides the convenient lower bound:

$$(11) \quad \mathbb{P}_\nu^x(\tau_Q > n) \geq \mathbb{P}_\nu^v(\tau_{V^+} > n, \max_{k \leq n} \|S_k^{(2)} - v^{(2)}\| \leq d(x)).$$

We now analyse this lower bound with the help of Cramér's transformation. We know by Lemma 10 that there exists $v_0 \in V^+$ such that

$$\lambda := L_\nu(v_0) = \min_{V^+} L_\nu > 0.$$

Let ν_0 be the probability measure on V defined by

$$\lambda \nu_0(dy) = e^{\langle v_0, y \rangle} \nu(dy).$$

Thanks to Cramér's formula (see [6, Lemma 6]), the lower bound in equation (11) can be written as

$$(12) \quad \lambda^n \mathbb{E}_{\nu_0}^v \left(e^{-\langle v_0, S_n - v \rangle}, \tau_{V^+} > n, \max_{k \leq n} \|S_k^{(2)} - v^{(2)}\| \leq d(x) \right).$$

Since $v_0 \in V^+$, we have $\langle v_0, e_i \rangle \geq 0$ for all $i \in I$. Define

$$K = \{i \in I : \langle v_0, e_i \rangle > 0\}.$$

Then $\langle v_0, e_i \rangle = 0$ for all $i \in I \setminus K$, so that

$$\begin{aligned} |\langle v_0, S_n - v \rangle| &= \left| \sum_{i \in K} \langle v_0, e_i \rangle \langle S_n^{(1)} - v^{(1)}, e_i \rangle + \langle v_0, S_n^{(2)} - v^{(2)} \rangle \right| \\ &\leq \left(\sum_{i \in K} \langle v_0, e_i \rangle + \|v_0\| \right) d(x), \end{aligned}$$

as soon as $|\langle S_n^{(1)} - v^{(1)}, e_i \rangle| \leq d(x)$ for all $i \in K$ and $\|S_n^{(2)} - v^{(2)}\| \leq d(x)$. Under this additional constraint, the term $\exp(-\langle v_0, S_n \rangle)$ inside the expectation in (12) is bounded from below by some positive constant that will disappear in the n th root limit.

Therefore, using the notation

$$\tilde{\rho}_x = \liminf_{n \rightarrow \infty} \mathbb{P}_{\nu}^x(\tau_Q > n)^{1/n},$$

we have

$$(13) \quad \tilde{\rho}_x \geq \lambda \liminf_{n \rightarrow \infty} \mathbb{P}_{\nu_0}^v \left(\tau_{V^+} > n, \max_{k \leq n} \chi(S_k - v) \leq d(x) \right)^{1/n}$$

where

$$\chi(S_k - v) = \max \left\{ \max_{i \in K} |\langle S_k^{(1)} - v^{(1)}, e_i \rangle|, \|S_k^{(2)} - v^{(2)}\| \right\}.$$

The behavior of this last limit is analysed in Theorem 13. Indeed, the space V where the probability distribution ν lives has the cartesian product structure $V = V_1 \oplus V_2$ and the cone writes $V^+ = V_1^+ \oplus V_2$, where V_1^+ is the positive orthant of V_1 . Furthermore, Lemma 10 asserts that $\nu = \mu|V$ satisfies assumption (H') of Theorem 13 with respect to V^+ ; so does ν_0 since ν and ν_0 are absolutely continuous with respect to each other. Thus, it remains to take a look at its expectation m_0 which is given by

$$m_0 = \nabla L_{\nu_0}(0) = \lambda^{-1} \nabla L_{\nu}(v_0).$$

Let us write $m_0 = m_0^{(1)} + m_0^{(2)}$ with $m_0^{(1)} \in V_1$ and $m_0^{(2)} \in V_2$. Since v_0 is a global minimum point of L_{ν} on $V^+ = V_1^+ \oplus V_2$, it is easily seen that:

- $m_0^{(1)} \in V_1^+$,
- $\langle m_0^{(1)}, e_i \rangle = 0$ for all $i \in K$, and
- $m_0^{(2)} = 0$.

Indeed, for all $i \in I$, we have $v_i = \langle v_0, e_i \rangle \geq 0$, and the half-line $\{v_0 + te_i : t \geq -v_i\}$ is included in V^+ . Therefore, the function

$$t \in [-v_i, \infty) \mapsto f_i(t) = L_{\nu}(v_0 + te_i)$$

reaches its minimum at $t = 0$. This implies that $\langle m_0^{(1)}, e_i \rangle = \lambda^{-1} f'_i(0) \geq 0$ with equality if $v_i > 0$, i.e. if $i \in K$. On the other hand, if we take any $h \in V_2$ then the whole line $\{v_0 + th : t \in \mathbb{R}\}$ is included in V^+ , hence the function $t \in \mathbb{R} \mapsto L_\nu(v_0 + th)$ reaches a local minimum at $t = 0$. By consequence, its derivative $\langle \nabla L_\nu(0), h \rangle = \lambda \langle m_0, h \rangle$ at $t = 0$ is always equal to 0. Thus $m_0^{(2)} = 0$.

Thanks to those properties of ν_0 and its expectation m_0 , we can apply Theorem 13 which ensures the existence of some $\delta \geq 0$ such that

$$\lim_{R \rightarrow \infty} \liminf_{n \rightarrow \infty} \mathbb{P}_{\nu_0}^v \left(\tau_{V^+} > n, \max_{k \leq n} \chi(S_k - v) \leq R \right)^{1/n} = 1,$$

for all $v \in V_\delta^+ = \{v \in V : v_i \geq \delta, \forall i \in I\}$. Notice that the probability under consideration reaches its minimum on V_δ^+ when v is the ‘‘corner’’ point $v_* = \delta \sum_{i \in I} e_i$ (this follows by inclusion of events). Hence it follows from (13) that for any $x = v + w \in Q$ with $v \in V_\delta^+$,

$$\tilde{\rho}_x \geq \lambda \liminf_{n \rightarrow \infty} \mathbb{P}_{\nu_0}^{v_*} \left(\tau_{V^+} > n, \max_{k \leq n} \chi(S_k - v_*) \leq d(x) \right)^{1/n}.$$

Since $d(x)$ is now disconnected from the starting point v_* we can let $d(x) \rightarrow \infty$, thus proving that

$$\lim_{\substack{d(x) \rightarrow \infty \\ x \in Q_\delta}} \tilde{\rho}_x \geq \lambda = \min_{V^+} L_{\mu|V}.$$

Theorem 4 then follows from the combination of this inequality with (9) and (10).

3. THE FAVORABLE CASE

In this section, we consider a square integrable random walk (S_n) in $\mathbb{R}^d = \mathbb{R}^p \times \mathbb{R}^q$ with distribution μ , mean m and variance-covariance matrix Γ . For all $x \in \mathbb{R}^d$, we denote by $x_i, i = 1 \dots d$, its coordinates in the standard basis, and $x^{(1)} \in \mathbb{R}^p$ and $x^{(2)} \in \mathbb{R}^q$ its ‘‘coordinates’’ with respect to the cartesian product $\mathbb{R}^p \times \mathbb{R}^q$. Let Q denote the positive orthant of \mathbb{R}^p , i.e.

$$Q = \{x \in \mathbb{R}^p : x_i \geq 0, \forall i = 1 \dots p\}.$$

We are interested in the tail distribution of the exit time

$$\tau_K = \inf\{n \geq 1 : S_n \notin K\}$$

of the random walk from the cartesian product

$$K = Q \times \mathbb{R}^q,$$

in the *favorable* case where $m \in Q \times \{0\}^q$, but with the additional constraint that some zero-mean coordinates of the walk stay in a bounded domain. In what follows, we assume the random walk is well oriented (with respect to K), i.e. the probability distribution satisfies the following condition:

- (H') The support of μ is not included in any half-space $u^- = \{x \in \mathbb{R}^d : \langle x, u \rangle \leq 0\}$ with $u \in Q \times \{0\}^q \setminus \{0\}$.

Let J be the set of indices j such that $m_j = 0$. Denote by $\mathbb{R}^{(J)}$ the subspace of J -coordinates, that is

$$\mathbb{R}^{(J)} = \{x \in \mathbb{R}^d : x_i = 0, \forall i \notin J\},$$

and let $x^{(J)}$ be the projection of $x \in \mathbb{R}^d$ on $\mathbb{R}^{(J)}$. Let also

$$(14) \quad \|x\|_J = \|x^{(J)}\| = \left(\sum_{i \in J} |x_i|^2 \right)^{1/2}$$

be the norm of the projection of x on the subspace of J -coordinates.

We shall prove in this setting the following result that extends Theorem 13 of [6].

Theorem 13. *Assume μ satisfies (H'), and that $m \in Q \times \{0\}^q$. Let J be the set of indices j such that $m_j = 0$. There exists $\delta \geq 0$ such that*

$$\lim_{R \rightarrow \infty} \liminf_{n \rightarrow \infty} \mathbb{P}_\mu^x \left(\tau_K > n, \max_{k \leq n} \|S_k - x\|_J \leq R \right)^{1/n} = 1$$

for all $x \in K_\delta = Q_\delta \times \mathbb{R}^q$.

Theorem 13 will follow from the two propositions below. Roughly speaking, the first one will enable us to push the random walk as far as we want from the boundary of the cone K (with a positive probability):

Proposition 14. *Under the hypotheses of Theorem 13, there exist $\gamma > 0$, $b \geq 1$ and $\delta, R > 0$ such that*

$$\mathbb{P}_\mu^x \left(\tau_K > b\ell, S_{b\ell} \in K_\ell, \max_{k \leq b\ell} \|S_k - x\|_J \leq R \right) \geq \gamma^\ell,$$

for all $\ell \geq 1$ and $x \in K_\delta$.

Now, as soon as the random walk has reached a point y at a distance $\geq R$ from the boundary, it suffices that the walk stays in $B(y, R)$ so as to be sure that it will not leave the cone. This simple observation will enable us to derive the theorem from the second proposition:

Proposition 15. *Assume (\tilde{S}_n) is square integrable random walk with mean $m = 0$ and any variance-covariance matrix. Then*

$$\lim_{R \rightarrow \infty} \liminf_{n \rightarrow \infty} \mathbb{P}_\mu^0 \left(\max_{k \leq n} \|\tilde{S}_k\| \leq R \right)^{1/n} = 1.$$

The proofs of Proposition 14 and 15 are deferred to section 3.2 and section 3.3, respectively. First of all, let us explain precisely how the combination of those two propositions leads to Theorem 13.

3.1. Proof of Theorem 13. Proposition 14 ensures the existence of $\gamma > 0$, $b \geq 1$ and $\delta, R_0 > 0$ such that

$$\mathbb{P}_\mu^x \left(\tau_K > b\ell, S_{b\ell} \in K_\ell, \max_{k \leq b\ell} \|S_k - x\|_J \leq R_0 \right) \geq \gamma^\ell,$$

for all $\ell \geq 1$ and $x \in K_\delta$.

Let $\epsilon > 0$ be given. Applying Proposition 15 to the centered random walk $\tilde{S}_n = S_n - nm$, we obtain the existence of a number $R \geq R_0$ such that

$$\mathbb{P}_\mu^0 \left(\max_{k \leq n} \|\tilde{S}_k\| \leq R - R_0 \right) \geq (1 - \epsilon)^n,$$

for all n large enough. Now fix $\ell \geq R - R_0$ and suppose that $y \in K_\ell \cap \overline{B_J(x, R_0)}$. If

$$\max_{k \leq n} \|\tilde{S}_k - y\| \leq R - R_0 (\leq \ell),$$

then:

- (1) Clearly, \tilde{S}_k belongs to K for all $k \leq n$. Since $S_k = \tilde{S}_k + km$ and m belongs to K , the same is true for S_k , thus $\tau_K > n$.
- (2) For all $k \leq n$, we have

$$\|S_k - x\|_J = \|\tilde{S}_k - x\|_J \leq \|\tilde{S}_k - y\| + \|y - x\|_J \leq R.$$

Therefore, if we consider only trajectories such that $S_{b\ell} \in K_\ell$ and $\|S_{b\ell} - x\|_J \leq R_0$, and then use the Markov property at time $b\ell$, we obtain the lower bound

$$\begin{aligned} \mathbb{P}_\mu^x \left(\tau_K > n, \max_{k \leq n} \|S_k - x\|_J \leq R \right) &\geq \gamma^\ell \times \inf_y \mathbb{P}_\mu^y \left(\tau_K > n - b\ell, \max_{k \leq n - b\ell} \|S_k - x\|_J \leq R \right) \\ &\geq \gamma^\ell \times \inf_y \mathbb{P}_\mu^y \left(\max_{k \leq n - b\ell} \|\tilde{S}_k - y\| \leq R - R_0 \right) \\ &\geq \gamma^\ell \times (1 - \epsilon)^{n - b\ell}, \end{aligned}$$

where the infimum is taken over all $y \in K_\ell \cap \overline{B_J(x, R_0)}$. Consequently

$$\liminf_{n \rightarrow \infty} \mathbb{P}_\mu^x \left(\tau_K > n, \max_{k \leq n} \|S_k - x\|_J \leq R \right)^{1/n} \geq 1 - \epsilon,$$

and the theorem is proved.

3.2. Pushing the walk deep inside the cone. This section is devoted to the proof of Proposition 14. In what follows, the distribution μ of the random walk increments is assumed to satisfy assumption (H'). Let m be the expectation of μ and

$$F = (\ker \Gamma)^\perp,$$

where Γ is the variance-covariance matrix of μ . It is well-known that the smallest affine subspace of \mathbb{R}^d with full μ -probability is

$$m + F.$$

Therefore, assumption (H') ensures that there exists no $u \in Q \times \{0\}^q \setminus \{0\}$ such that $m + F \subset u^-$.

We define the *smoothed support* \mathcal{G} of the random walk as

$$\mathcal{G} = \mathbb{R}_+ m + F,$$

and notice that, if started at any point in \mathcal{G} , the random walk stays in \mathcal{G} forever.

In addition, we assume that

$$m \in Q \times \{0\}^q,$$

and define the set

$$J = \{j \in \llbracket 1, d \rrbracket : m_j = 0\} \supset \llbracket p+1, q \rrbracket.$$

Finally, let

$$B_J(0, R) = \{x \in \mathbb{R}^d : \|x\|_J < R\},$$

where $\|*\|_J$ is defined as in (14).

3.2.1. Some geometry. We collect here two technical lemmas related to the geometry of the problem. The first one asserts that the affine support of μ meets the interior of the cone $K = Q \times \mathbb{R}^q$. This is crucial since otherwise we couldn't expect the walk to go deep inside the cone.

Lemma 16. *Assume (H') is satisfied and $m \in Q \times \{0\}^q$. Then*

$$(15) \quad (m + F) \cap K^\circ \neq \emptyset$$

Moreover, for all $x \in K \cap \mathcal{G}$,

$$(16) \quad (x + m + F) \cap K^\circ \cap B_J(0, 1) \neq \emptyset.$$

Proof. We assume $(m + F) \cap K^\circ = \emptyset$, and infer the existence of some $u \in Q \times \{0\}^q$ such that $(m + F) \subset u^\perp$, thus contradicting the assumption (H').

Let us first consider the case where $F = H$ is a hyperplane. Then there exists $u \neq 0$ such that $H = u^\perp$. We shall prove that $m \in u^\perp$. Suppose on the contrary that $\langle m, u \rangle \neq 0$, then, possibly changing u to $-u$, we can assume that $\langle m, u \rangle > 0$. Now, using the homogeneity of H and K° , we see that K° does not intersect with

$$\bigcup_{\lambda > 0} (\lambda m + H) = \bigcup_{\lambda > 0} (\lambda u + u^\perp) = \{x : \langle x, u \rangle > 0\}.$$

Therefore, K° is included in u^- . But since $m \in K = \overline{(K^\circ)}$ (this equality holds for any convex set with non-empty interior), we obtain that $\langle m, u \rangle \leq 0$, which contradicts our hypothesis. Hence, m belongs to u^\perp and $m + H = u^\perp$. Finally, the non-intersecting hypothesis rewrites $u^\perp \cap K^\circ = \emptyset$ and is easily seen to be equivalent to $u \in \pm Q \times \{0\}^q$.

We now turn to the general case where F is any linear subspace. Since $m + F$ and K° are two disjoint convex sets, it follows from the Hyperplane separation theorem that there exists an affine hyperplane H_m that separates $m + F$ and K° . But, since m belongs to both $m + F$ and $K = \overline{K^\circ}$, it must belong to H_m , and therefore $H_m = m + H$, where H is a linear hyperplane. Now, F being a linear subspace, it can't be on one side of H unless it is contained in H . Therefore, we obtain that $m + F \subset m + H$ and $(m + H) \cap K^\circ = \emptyset$, and equation (15) follows by applying the first part of the proof to $m + H$.

Let us now show that (15) implies (16). Since $(m + F) \cap K^\circ$ is non-empty, there is some $f_0 \in F$ such that $m + f_0 \in K^\circ$. Therefore

$$m + \alpha f_0 \in K^\circ$$

for all $\alpha \in (0, 1]$ (since $m \in K$ and K is convex). Fix such an α so small that $\|\alpha f_0\|_J < 1$. For $x \in K \cap \mathcal{G}$, write $x = \lambda_1 m + f_1$ with $\lambda_1 \geq 0$ and f_1 in F , and set $f = \alpha f_0 - f_1$. Then,

$$x + m + f = \lambda_1 m + (m + \alpha f_0) \in K^\circ.$$

(since $\lambda_1 m \in K$, $m + \alpha f_0 \in K^\circ$ and $K + K^\circ \subset K^\circ$.) In addition,

$$\|x + m + f\|_J = \|\alpha f_0\|_J < 1,$$

thus proving that $x + m + f \in K^\circ \cap B_J(0, 1)$. \square

The second lemma is a technical tool. For any $x, y \in \mathbb{R}^d$, we write $y \leq x$ iff $x - y \in \mathbb{R}_+$.

Lemma 17. *Let (x_n) be a sequence in $K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$. There exists a bounded sequence (y_n) in $K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$, such that*

$$y_n \leq x_n \quad \text{and} \quad y_n^{(J)} = x_n^{(J)} \quad \text{for all } n.$$

Proof. Since $\mathcal{G} = \mathbb{R}_+ m + F$, where F is a linear subspace of \mathbb{R}^d , there exist $L \in M_d(\mathbb{R})$ and a linear form $\phi : \mathbb{R}^d \rightarrow \mathbb{R}$, such that

$$\mathcal{G} = \{x \in \mathbb{R}^d : L(x) = 0, \phi(x) \geq 0\}.$$

Recall that

$$\mathbb{R}^{(J)} = \{x \in \mathbb{R}^d : x_i = 0, \forall i \notin J\}.$$

Let $\mathbb{R}^{(I)}$ be its orthogonal complement, i.e.

$$\mathbb{R}^{(I)} = \{x \in \mathbb{R}^d : x_i = 0, \forall i \in J\},$$

and write $x = x^{(I)} + x^{(J)}$ the orthogonal decomposition with respect to $\mathbb{R}^{(I)} \oplus \mathbb{R}^{(J)}$. Let $x_n = x_n^{(I)} + x_n^{(J)}$ be an element of $K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$ and define

$$P = \{z \in \mathbb{R}_+^{(I)} : L(z) = -L(x_n^{(J)}), \phi(z) \geq -\phi(x_n^{(J)})\}.$$

For all $z \in \mathbb{R}_+^{(I)}$, notice that $z \in P$ iff $z + x_n^{(J)} \in \mathcal{G}$. Therefore, $x_n^{(I)} \in P$. It follows from Corollary 23 that there exists $y_n^{(I)} \in P$ such that $y_n^{(I)} \leq x_n^{(I)}$ and

$$\|y_n^{(I)}\| \leq M(\|L(x_n^{(J)})\| + |\phi(x_n^{(J)})|),$$

where $M = M(L, \phi)$ only depends on L and ϕ (and not on x_n). Setting $y_n = y_n^{(I)} + x_n^{(J)}$ thus gives a bounded sequence in $K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$ that satisfies the conditions of the lemma. \square

3.2.2. *Proof of Proposition 14.* We begin with a lemma that asserts the existence of a time b and a radius R_0 such that the random walk started at $x \in K \cap \mathcal{G}$ with $\|x\|_J \leq R_0$ will be at a distance ≥ 1 from the boundary of K at time b and still located in $\overline{B_J(0, R_0)}$ with a probability that is bounded from below by some positive constant, uniformly in x .

Lemma 18. *There exist $b \geq 1$ and $R_0 > 0$ such that*

$$\inf_{x \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}} \mathbb{P}_\mu^x (S_b \in K_1, \|S_b\|_J \leq R_0) > 0.$$

Proof. Clearly, the lemma will follow from the existence of an integer $n \geq 1$ such that

$$\inf_{x \in K \cap \overline{B_J(0, 1)} \cap \mathcal{G}} \mathbb{P}_\mu^{x\sqrt{n}} (S_n \in K_1, \|S_n\|_J \leq \sqrt{n}) > 0.$$

So, let us assume that this assertion is false. Then, we can find a sequence of points $x_n \in K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$ such that

$$p_n := \mathbb{P}_\mu^{x_n\sqrt{n}} (S_n \in K_1, \|S_n\|_J \leq \sqrt{n}) \rightarrow 0.$$

Thanks to Lemma 17, we can assume that (x_n) is bounded, because for any sequence (y_n) with the same properties as in this lemma, the probability p_n where x_n is replaced by y_n is smaller than p_n , by inclusion of events. Furthermore, by extracting a subsequence, it can be assumed without loss of generality that (x_n) converges to some element x of the closed set $K \cap \overline{B_J(0, 1)} \cap \mathcal{G}$.

Now, let $\tilde{S}_n = S_n - nm$ denote the centered random walk associated with S_n . Since $\|\cdot\|_J$ is left invariant by a translation by m , the probability p_n can be written as

$$p_n = \mathbb{P}_\mu^0 (x_n\sqrt{n} + \tilde{S}_n \in K_1 - nm, \|x_n\sqrt{n} + \tilde{S}_n\|_J \leq \sqrt{n}).$$

Let $\eta > 0$ be fixed. For all $n \geq 1/\eta$, holds the inclusion

$$(K_1 - nm)/\sqrt{n} \subset K_\eta - m.$$

Therefore, the probability p_n is bounded from below by

$$\mathbb{P}_\mu^0 (x_n + \tilde{S}_n/\sqrt{n} \in K_\eta - m, \|x_n + \tilde{S}_n/\sqrt{n}\|_J \leq 1).$$

Since $x_n + \tilde{S}_n/\sqrt{n}$ converges in distribution to $x + X$, where X denotes a random variable with $\mathcal{N}(0, \Gamma)$ Gaussian distribution, we can use the Portmanteau theorem and let then $\eta \downarrow 0$ to get the lower bound

$$\liminf_{n \rightarrow \infty} p_n \geq \mathbb{P}(x + X \in (K^\circ - m) \cap B_J(0, 1)).$$

The random variable $x + X$ admits a positive density with respect to Lebesgue measure on the affine space $x + F$ where $F = (\ker \Gamma)^\perp$, and

$$O := (x + F) \cap (K^\circ - m) \cap B_J(0, 1)$$

is an open subset of $x + F$. Thus it suffices to prove that O is non-empty to obtain a contradiction. But this is precisely what asserts Lemma 16: Indeed, since x belongs to $K \cap \mathcal{G}$, we obtain that

$$(x + m + F) \cap K^\circ \cap B_J(0, 1) \neq \emptyset.$$

Hence, subtracting m on both sides (recall that $\| * \|_J$ is left invariant by a translation by m) gives

$$O \neq \emptyset.$$

This implies

$$\mathbb{P}(x + X \in O) > 0,$$

thus contradicting our assumption that $\liminf p_n = 0$. Therefore, the lemma is proven. \square

By Lemma 18, there exist $b \geq 1$, $R_0 > 0$ and $\gamma > 0$, such that

$$(17) \quad \mathbb{P}_\mu^x(S_b \in K_1, \|S_b\|_J \leq R_0) \geq 2\gamma,$$

for all $x \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}$. Let us choose $\delta > 0$ and $R \geq R_0$ such that

$$\mathbb{P}_\mu^0 \left(\tau_{K_{-\delta}} > b, \max_{k \leq b} \|S_k\|_J \leq R - R_0 \right) \geq 1 - \gamma.$$

Then, by inclusion of events, we also have

$$(18) \quad \mathbb{P}_\mu^x \left(\tau_{K_{-\delta}} > b, \max_{k \leq b} \|S_k\|_J \leq R \right) \geq 1 - \gamma,$$

for all $x \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}$. Indeed, this follows from the relation $K + K_{-\delta} \subset K_{-\delta}$ and the triangle inequality for $\| * \|_J$. Now, combining (17) and (18), we obtain that

$$(19) \quad \mathbb{P}_\mu^x \left(\tau_{K_{-\delta}} > b, S_b \in K_1, \max_{k \leq b} \|S_k\|_J \leq R, \|S_b\|_J \leq R_0 \right) \geq \gamma,$$

for all $x \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}$. Set

$$p_\ell(x) = \mathbb{P}_\mu^x \left(\tau_{K_{-\delta}} > b\ell, S_{b\ell} \in K_\ell, \max_{k \leq b\ell} \|S_k\|_J \leq R \right).$$

Notice that $K_1 + K_\ell \subset K_{\ell+1}$. Hence, if we consider only trajectories such that $S_b \in K_1$ and $\|S_b\|_J \leq R_0$, and then use the Markov property at time b , we get the lower bound

$$p_{\ell+1}(x) \geq \gamma \times \inf_{y \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}} p_\ell(y),$$

for all $x \in K \cap \overline{B_J(0, R_0)} \cap \mathcal{G}$. This proves that $p_\ell(0) \geq \gamma^\ell$ for all $\ell \geq 1$, and Proposition 14 follows by inclusion of events since $K_\delta + K_{-\delta} \subset K$ and $K_\delta + K_\ell \subset K_\ell$.

3.3. On the exit time from a ball. This section is devoted to the proof of Proposition 15. In what follows, the abbreviation ‘‘f.s.’’ stands for ‘‘for some’’.

3.3.1. Preliminary estimate for Brownian motion. Let (B_t) denotes a true d -dimensional Brownian motion, i.e. the image of a standard d -dimensional Brownian motion (meaning a collection of d independent one-dimensional Brownian motions) by an invertible linear transformation.

Lemma 19. *For every $\epsilon > 0$, there exist $\delta > 0$ and $0 < \alpha < 1$ such that*

$$\mathbb{P}^x (\|B_t\| < 1 - \delta \text{ f.s. } t \in [\alpha, 1]) \geq 1 - \epsilon$$

for all $x \in \overline{B(0, 1)}$.

Proof. Suppose on the contrary that there is some $\epsilon_0 > 0$ for which we can pick $\alpha_n \downarrow 0$, $\delta_n \downarrow 0$ and $x_n \in \overline{B(0,1)}$ such that

$$p_n := \mathbb{P}^{x_n} (\|B_t\| < 1 - \delta_n \text{ f.s. } t \in [\alpha_n, 1]) < 1 - \epsilon_0$$

for all n . By compactness, it can also be assumed that $\|x_n - x\| \rightarrow 0$ for some $x \in \overline{B(0,1)}$. Now, for any $\eta > 0$, we have

$$p_n \geq \mathbb{P}^x (\|B_t\| < 1 - \eta \text{ f.s. } t \in [\alpha_n, 1])$$

as soon as $\|x_n - x\| + \delta_n \leq \eta$. Hence, taking the limit on both sides and letting $\eta \downarrow 0$ gives

$$\liminf_{n \rightarrow \infty} p_n \geq \mathbb{P}^x (\|B_t\| < 1 \text{ f.s. } t \in (0, 1]).$$

But it follows from the classical cone condition (as found in [7, Proposition 3.3] for example) applied to the ball $B(0, 1)$ that x is regular for $B(0, 1)$, i.e. B_t immediately visits $B(0, 1)$ with full probability. Therefore, the last inequality reads

$$\liminf_{n \rightarrow \infty} p_n \geq 1$$

and contradicts our assumption. □

3.3.2. Application to random walks. In this subsection, $(S_n) \in \mathbb{R}^d$ is a square integrable random walk with increments distribution μ , mean $m = 0$ and any covariance matrix Γ .

The proof of Proposition 15 is based on the following basic idea. Given $\epsilon > 0$, find $R > 0$ and a time $n_0 \geq 1$ such that the random walk started at 0 returns to 0 at time n_0 without leaving the ball $B(0, R)$ with probability $\geq 1 - \epsilon$. If this can be done, then the result follows by concatenation (i.e Markov property). But this is asking for a property stronger than recurrence, thus we can not hope for such a simple argument. Instead of a return to 0, we can ask for a return in some ball $B(0, R_0)$, with $R_0 \leq R$. But then, in view of using concatenation, we need the previous probability to be greater than $1 - \epsilon$ uniformly for any starting point in the same ball $B(0, R_0)$. Lemma 20 below provides a result in this spirit that is sufficient for our purpose.

Lemma 20. *Suppose (here only) that the covariance matrix Γ is non-degenerate. Then, for every $\epsilon > 0$, there exist $0 < R_0 < R$ and $1 \leq \ell_0 \leq n_0$ such that*

$$\mathbb{P}_\mu^x \left(\max_{k \leq n_0} \|S_k\| \leq R \text{ and } \|S_k\| \leq R_0 \text{ f.s. } k \in [\ell_0, n_0] \right) \geq 1 - \epsilon$$

for all $x \in \overline{B(0, R_0)}$.

Proof. Let $\epsilon > 0$ be given and fix $\delta > 0$ and $0 < \alpha < 1$ so that the conclusion of Lemma 19 holds. Also, fix a parameter $\beta \in (0, \alpha)$.

Let (x_n) be a sequence in $\overline{B(0, 1)}$ that converges to some $x \in \overline{B(0, 1)}$, and set

$$p_n = \mathbb{P}_\mu^{x_n \sqrt{n}} (\|S_k\| \leq \sqrt{n} \text{ f.s. } k \in [\beta n, n])$$

We shall prove that

$$(20) \quad \liminf_{n \rightarrow \infty} p_n \geq 1 - \epsilon.$$

To do this, we consider the process with continuous path $(Z_n(t), t \in [0, 1])$ defined by

$$Z_n(t) = \frac{S_{[nt]}}{\sqrt{n}} + (nt - [nt]) \frac{\xi_{[nt]+1}}{\sqrt{n}},$$

where $\xi_k = S_k - S_{k-1}$ and $[a]$ denotes the integer part of a . The probability p_n can then be written in terms of the process Z_n as $\mathbb{P}_\mu^0(A_n)$, where

$$A_n = \{\|x_n + Z_n(t)\| \leq 1 \text{ f.s. } t = k/n \in [\beta, 1]\}.$$

We wish to use the Functional Central Limit Theorem [1, Theorem 10.1] together with the Portmanteau theorem [1, Theorem 2.1] in order to obtain a lower bound for the probability of this event, but the condition that t be rational ($t = k/n$) can not be handled directly, and must therefore be relaxed. To this end, we define

$$\widetilde{A}_n = \{\|x + Z_n(t)\| < 1 - \delta \text{ f.s. } t \in [\alpha, 1]\}$$

and

$$B_n = \left\{ \max_{k=0..n} \|\xi_k\| > (\delta/2)\sqrt{n} \right\}.$$

Since $\|Z_n(t) - Z_n(k/n)\| \leq \|\xi_{k+1}\|/\sqrt{n}$ for $k/n \leq t < (k+1)/n$, we have the inclusion of events

$$\widetilde{A}_n \cap \overline{B_n} \subset A_n$$

as soon as $\|x_n - x\| \leq \delta/2$ and $\alpha - \beta > 1/n$. Thus, for all sufficiently large n , we have

$$p_n \geq \mathbb{P}_\mu^0(\widetilde{A}_n) - \mathbb{P}_\mu^0(B_n).$$

Furthermore, it is a basic result in probability theory that $\mathbb{P}(B_n) \rightarrow 0$ for any i.i.d. sequence of zero-mean square integrable random variables (ξ_k) . Hence, we are left to prove that

$$\liminf_{n \rightarrow \infty} \mathbb{P}_\mu^0(\widetilde{A}_n) \geq 1 - \epsilon.$$

Now, since Γ is non-degenerate, the Functional Central Limit Theorem asserts that Z_n converges in distribution to a true d -dimensional Brownian motion (with covariance matrix Γ) on the space of continuous functions $w : [0, 1] \rightarrow \mathbb{R}^d$ equipped with the topology of uniform convergence. Since the set of continuous functions $w : [0, 1] \rightarrow \mathbb{R}^d$ such that $\|w(t)\| < 1 - \delta$ for some $t \in [\alpha, 1]$ is open with respect to the topology of uniform convergence, it follows from the Portmanteau theorem that

$$\liminf_{n \rightarrow \infty} \mathbb{P}_\mu^0(\widetilde{A}_n) \geq \mathbb{P}^0(\|x + B_t\| < 1 - \delta \text{ f.s. } t \in [\alpha, 1]) \geq 1 - \epsilon,$$

where the lower bound $1 - \epsilon$ comes from our choice of δ and α . This proves our first claim that $\liminf_n p_n \geq 1 - \epsilon$. By a standard compactness argument, this immediately implies that

$$\inf_{x \in B(0,1)} \mathbb{P}_\mu^{x\sqrt{n}}(\|S_k\| \leq \sqrt{n} \text{ f.s. } k \in [\beta n, n]) \geq 1 - 2\epsilon$$

for all sufficiently large n . Fix such $n_0 > 1/\beta$ and set $\ell_0 = [\beta n_0]$ and $R_0 = \sqrt{n_0}$. Then we have $0 < \ell_0 < n_0$, and the last inequality can be rewritten as

$$\inf_{x \in B(0, R_0)} \mathbb{P}_\mu^x(\|S_k\| \leq R_0 \text{ f.s. } k \in [\ell_0, n_0]) \geq 1 - 2\epsilon$$

In order to complete the proof, it suffices to notice that, for all $x \in \overline{B(0,1)}$ and $R \geq R_0$, we have

$$\mathbb{P}_\mu^x \left(\max_{k \leq n_0} \|S_k\| \leq R \right) \geq \mathbb{P}_\mu^0 \left(\max_{k \leq n_0} \|S_k\| \leq R - R_0 \right).$$

Since the last probability goes to 1 as $R \rightarrow \infty$, it is bounded from below by $1 - \epsilon$ for all sufficiently large R , and we conclude that for such a choice of R , we have

$$\mathbb{P}_\mu^x \left(\max_{k \leq n_0} \|S_k\| \leq R \text{ and } \|S_k\| \leq R_0 \text{ f.s. } k \in \llbracket \ell_0, n_0 \rrbracket \right) \geq 1 - 3\epsilon$$

for all $x \in \overline{B(0, R_0)}$. \square

3.3.3. Proof of Proposition 15. First of all, we notice that the variance-covariance matrix Γ can be assumed to be non-degenerate. Otherwise, the random walk lives on $(\ker \Gamma)^\perp$ where it has a non-degenerate variance-covariance matrix. Since the projection of a d -dimensional ball is still a ball in $(\ker \Gamma)^\perp$, the result will follow by application of the non-degenerate case to the projected random walk.

As explained earlier, the idea is now to concatenate the “high-probability path” given by Lemma 20. Let $\epsilon > 0$ be given. Thanks to Lemma 20, we can find $\delta > 0$, $0 < R_0 < R$ and $1 \leq \ell_0 \leq n_0$ such that

$$\inf_{x \in \overline{B(0, R_0)}} \mathbb{P}_\mu^x \left(\max_{k \leq n_0} \|S_n\| \leq R, H \leq n_0 \right) \geq 1 - \epsilon,$$

where H denotes the first hitting time of the ball $\overline{B(0, R_0)}$ after time ℓ_0 .

For $x \in \overline{B(0, R_0)}$ and $n \geq n_0$, we use the strong Markov property at time H to get

$$\begin{aligned} & \mathbb{P}_\mu^x \left(\max_{k \leq n} \|S_k\| \leq R \right) \\ & \geq \mathbb{P}_\mu^x \left(\max_{k \leq n} \|S_k\| \leq R, H \leq n_0 \right) \\ & \geq \mathbb{E}_\mu^x \left(\max_{k \leq H} \|S_k\| \leq R, H \leq n_0, \mathbb{P}_\mu^{S_H} \left(\max_{k \leq n-j} \|S_k\| \leq R \right) \Big|_{j=H} \right) \\ & \geq \mathbb{P}_\mu^x \left(\max_{k \leq H} \|S_k\| \leq R, H \leq n_0 \right) \times \inf_{y \in \overline{B(0, R_0)}} \mathbb{P}_\mu^y \left(\max_{k \leq n-\ell_0} \|S_k\| \leq R \right). \end{aligned}$$

Thus

$$R(n) := \inf_{x \in \overline{B(0, R_0)}} \mathbb{P}_\mu^x \left(\max_{k \leq n} \|S_k\| \leq R \right)$$

satisfies the inequality

$$R(n) \geq (1 - \epsilon)R(n - \ell_0)$$

for all $n \geq n_0$. Since $R(n)$ is clearly decreasing, for $n_0 + k\ell_0 \leq n < n_0 + (k+1)\ell_0$, we obtain

$$R(n) \geq R(n_0 + (k+1)\ell_0) \geq (1 - \epsilon)^{k+1} R(n_0) \geq (1 - \epsilon)^{n+2}.$$

Hence,

$$\liminf_{n \rightarrow \infty} R(n)^{1/n} \geq 1 - \epsilon$$

and this proves the proposition.

APPENDIX A. MINIMAL POINTS OF A POLYHEDRON

This independent and (nearly) self-contained section provides the material we need for the proof of Lemma 17. Though I'm quite sure that the notion of minimality introduced here and the related results are not new, I was not able to find any reference for it. Thus, I have decided to present here my own proofs, not pretending it to be new.

The arguments developed here are highly inspired by the standard ideas of linear programming theory, as can be found in the book [2] for example.

Let $1 \leq m \leq n$. We consider here the polyhedron

$$P = \{x \in \mathbb{R}_+^n : \sum_{i=1}^n x_i C_i = b\},$$

where C_i, b are any vectors of \mathbb{R}^m , and x_i denote the coordinates of x in the canonical basis.

If P is not empty (a condition that we assume from now on), it is a closed convex set. Its extremal points are called the *vertices* of P . Note that 0 is a vertex iff $b = 0$. There exist a simple and well known characterization of vertices. To $x \in P$, let us associate the subset $I(x)$ of indices i such that $x_i > 0$. Then $x \neq 0$ is a vertex of P iff the vectors $\{C_i, i \in I(x)\}$ are linearly independent (see [2, Théorème 10.3-1]). This proves that there is only a finite number of vertices. The fact that there always exist (at least) a vertex is proved in [2, Théorème 10.3-3] for example, and will follow as a by-product of our analysis.

For $x, y \in \mathbb{R}^n$, we write $y \leq x$ iff $x - y \in \mathbb{R}_+^n$ and $y < x$ iff $y \leq x$ and $y \neq x$. Hence $x > 0$ means that $x \geq 0$ and one of its coordinates (at least) is > 0 . We shall say that $x \in P$ is *minimal* if P does not contain any $y < x$.

Let us begin with a useful characterization of minimality:

Lemma 21. *An element $x \neq 0$ of P is minimal iff the vectors $\{C_i, i \in I(x)\}$ are positively independent, that is:*

$$\sum_{i \in I(x)} u_i C_i = 0 \text{ and } u_i \geq 0 \text{ for all } i \in I(x) \Rightarrow u_i = 0 \text{ for all } i \in I(x).$$

Proof. Let x be an element of P . Assume first that x is not minimal. Then there exist $y \in P$ such that $y < x$. If $i \notin I(x)$, then $y_i = x_i = 0$ (since $0 \leq y_i \leq x_i = 0$). Thus, there is at least one $i \in I(x)$ such that $y_i < x_i$.

Set $u_i = x_i - y_i$ for $i \in I(x)$. Then, the u_i 's are ≥ 0 and at least one of them is > 0 . Furthermore,

$$\sum_{i \in I(x)} u_i C_i = \sum_{i \in I(x)} x_i C_i - \sum_{i \in I(x)} y_i C_i = \sum_{i=1}^n x_i C_i - \sum_{i=1}^n y_i C_i = b - b = 0.$$

Hence, the vectors $\{C_i, i \in I(x)\}$ are not positively independent.

Conversely, suppose that there exist non-negative numbers $u_i, i \in I(x)$, not all zero, such that

$$\sum_{i \in I(x)} u_i C_i = 0.$$

For $i \notin I(x)$, set $u_i = 0$, and let $u \in \mathbb{R}_+^n$ denote the vector with coordinates u_i . Since one of the u_i 's is > 0 , we have $y = x - tu < x$ for any $t > 0$. But, if $t > 0$ is small enough, then y also belongs to \mathbb{R}_+^n (this follows from the fact that $x_i = 0 \Rightarrow u_i = 0$). Furthermore,

$$\sum_{i=1}^n y_i C_i = \sum_{i \in I(x)} (x - tu) = \sum_{i \in I(x)} x_i C_i + t \sum_{i \in I(x)} u_i C_i = b.$$

Hence, we find some $y \in P$ such that $y < x$, i.e. x is not minimal. \square

From this lemma, it should be clear that any vertex of P is minimal. Let us denote by P_m the set of minimal points of P . The main result of this section is the following:

Proposition 22. *Assume $P \neq \emptyset$. Then,*

- (1) *For any $x \in P$, there exist $y \in P_m$ such that $y \leq x$;*
- (2) *Every $y \in P_m$ is a convex combination of the vertices of P ;*
- (3) *There exists a number M , only depending on the C_i 's, such that P_m is bounded by $M\|b\|$.*

Proof. Let $x \in P$ and suppose that x is not minimal (otherwise there is nothing to prove). Then, by definition, there exist real numbers $u_i \geq 0, i \in I(x)$, such that

$$\sum_{i \in I(x)} u_i C_i = 0$$

and at least one of the u_i 's is > 0 . For $i \notin I(x)$, set $u_i = 0$, and let $u \in \mathbb{R}_+^n$ denote the vector with coordinates u_i . Now, we let $t \geq 0$ increase until the first time when $x - tu$ has a new 0 coordinate, that is we define

$$t_0 = \min\{x_i/u_i : u_i > 0\} > 0$$

and set

$$y = x - t_0 u.$$

Clearly, this new point y satisfies the inequality $0 \leq y < x$, and

$$\sum_{i=1}^n y_i C_i = \sum_{i \in I(x)} (x_i - t_0 u_i) C_i = b.$$

Thus $y < x$ also belongs to P and, furthermore, satisfies the strict inclusion relation $I(y) \subsetneq I(x)$. If y is not minimal, we repeat the argument (with y , and so on) until a minimal point is reached. The process indeed terminates since the set of positive indices $I(*)$ is finite and strictly decreasing along the process. This proves the first part of the proposition.

The proof of the second part is quite similar. Let $x \in P$ be minimal and suppose it is not a vertex of P . Then, the vectors $\{C_i, i \in I(x)\}$ are positively independent (by

minimality) but not independent. Therefore, there exist real numbers $\{u_i, i \in I(x)\}$, not all zero, such that

$$\sum_{i \in I(x)} u_i C_i = 0$$

and at least one of them is negative (< 0) and another is positive (> 0) (since else $-u$ would contradict the assumption). For $i \notin I(x)$, set $u_i = 0$, and let $u \in \mathbb{R}^n$ denote the vector with coordinates u_i . Now, we let $t \geq 0$ increase until the first time when $x + tu$ has a new 0 coordinate, that is we define

$$t^+ = \min\{-x_i/u_i : u_i < 0\} > 0 \quad \text{and} \quad x^+ = x + t^+u.$$

Similarly, we define

$$t^- = \min\{x_i/u_i : u_i > 0\} > 0 \quad \text{and} \quad x^- = x - t^-u.$$

These new points x^+ and x^- belong to P_m since they clearly are ≥ 0 , satisfy the equality

$$\sum_{i=1}^n x_i^\pm C_i = \sum_{i \in I(x)} (x_i \pm t^\pm u_i) C_i = b,$$

and the strict inclusion relation $I(x^\pm) \subsetneq I(x)$. Furthermore, x is a convex combination of x^+ and x^- since

$$(t^+ + t^-)x = t^+x^- + t^-x^+.$$

If x^+ and x^- are both vertices of P , then the proof is finished. Else, we repeat the argument until x is written as a convex combination of vertices. The process indeed terminates since the sets of positive indices $I(*)$ are finite and strictly decreasing along the process. This proves the second part of the proposition.

For the proof of the third assertion, we first notice that the Manhattan norm $\|x\|_1 = \sum_{i=1}^n |x_i|$ coincides with the linear form $\sum_{i=1}^n x_i$ on \mathbb{R}_+^n . Hence, it easily follows from the second assertion of the proposition that

$$\sup_{x \in P_m} \|x\|_1 = \max\{\|v\|_1 : v \text{ vertex of } P\}.$$

(Recall that there is at least one vertex and only a finite number of vertices.) Let us denote by \mathcal{I} the family of subsets $I \subset \llbracket 1, n \rrbracket$ such that $\{C_i, i \in I\}$ is a family of linearly independent vectors. Set

$$\mathbb{R}^{(I)} = \{x \in \mathbb{R}^n : x_i = 0, \forall i \notin I\}.$$

For all $I \in \mathcal{I}$, the linear mapping

$$C_I : \mathbb{R}^{(I)} \rightarrow \mathbb{R}^m, u \mapsto \sum_{i \in I} u_i C_i$$

is a bijection onto its image. Let us denote C_I^{-1} its inverse.

If v is a vertex of P , then $I = I(v)$ belongs to \mathcal{I} , and

$$\sum_{i \in I} v_i C_i = b \Leftrightarrow (v_i)_{i \in I} = C_I^{-1}b$$

Therefore,

$$\|v\|_1 = \|(v_i)_{i \in I}\|_1 \leq \|C_I^{-1}\|_1 \|b\|_1,$$

and P_m is thus bounded by $M\|b\|_1$, where $M = \max_{I \in \mathcal{I}} \|C_I^{-1}\|_1$. \square

In view of application to the proof of Lemma 17, we need to extend some of the consequences of Proposition 22 to a larger class of polyhedra. So, let $L : \mathbb{R}^n \rightarrow \mathbb{R}^n$ be a linear mapping and $\phi : \mathbb{R}^n \rightarrow \mathbb{R}$ be a linear form. Fix $b \in \mathbb{R}^n$, $c \in \mathbb{R}$, and consider the polyhedron

$$P = \{x \in \mathbb{R}_+^n : L(x) = b, \phi(x) \geq c\}.$$

Corollary 23. *Assume $P \neq \emptyset$. There exists a number M , only depending on L and ϕ , such that, for all $x \in P$, there exists $y \in P$ with $y \leq x$ and $\|y\| \leq M(\|b\| + |c|)$.*

Proof. We use the very standard trick in linear programming problems that consists in increasing the dimension so that the new form of polyhedra fits with the previous one. Indeed, if we add the “ghost” equation $x_{n+1} = \phi(x) - c$ to the system of equalities and inequality that defines P , then the inequality $\phi(x) \geq c$ becomes $x_{n+1} \geq 0$ and the system of equations

$$\begin{cases} L(x) = b \\ x_{n+1} = \phi(x) - c \end{cases}$$

can be written as

$$L'(x') = b'$$

where $x' = (x, x_{n+1})$, $b' = (b, -c)$ and $L' \in M_{n+1}(\mathbb{R})$ only depends on L and ϕ . So, define

$$P' = \{x' \in \mathbb{R}_+^{n+1} : L'(x') = b'\}.$$

Then, the mapping $\Psi : x \mapsto x' = (x, \phi(x) - c)$ is a bijection from P onto P' . Applying Proposition 22 to P' , we obtain the existence of a number M , only depending on L' (and thus only depending on L and ϕ), such that any minimal point y' of P' satisfies

$$\|y'\|_1 \leq M\|b'\|_1 = M(\|b\|_1 + |c|).$$

We also know, from the same proposition, that for any $x' = \Psi(x) \in P'$, there exists $y' \in P'$ such that $y' \leq x'$ and y' is minimal. Taking $y = \Psi^{-1}(y')$ gives the desired result. \square

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REFERENCES

- [1] Billingsley, P. (1968) *Convergence of probability measures*. Wiley.
- [2] Ciarlet, P.G. (1982) *Introduction à l'analyse numérique matricielle et à l'optimisation*. Masson.
- [3] Denisov, D. and Wachtel, W. (2011). Random walks in cones. *Ann. Probab.*, to appear.
- [4] Duraj, J. (2014). Random walks in cones: the case of nonzero drift. *Stochastic Process. Appl.* **124** 1503–1518.
- [5] Garbit, R. (2007). Temps de sortie d'un cône pour une marche aléatoire centrée. *C. R. Math. Acad. Sci. Paris* **135** 587–591.
- [6] Garbit, R. and Raschel, K. (2013). On the exit time from a cone for random walks with drift. *Preprint arXiv:1306.6761*.

- [7] Port, Sidney C. and Stone, Charles J. (1978) *Brownian motion and classical potential theory*. Probability and Mathematical Statistics. Academic Press.

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