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## THE LARGEST EIGENVALUES OF FINITE RANK DEFORMATION OF LARGE WIGNER MATRICES: CONVERGENCE AND NONUNIVERSALITY OF THE FLUCTUATIONS

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In this paper, we investigate the asymptotic spectrum of complex or real Deformed Wigner matrices  $(M_N)_N$  defined by  $M_N = W_N/\sqrt{N} + A_N$  where  $W_N$  is an  $N \times N$  Hermitian (resp., symmetric) Wigner matrix whose entries have a symmetric law satisfying a Poincaré inequality. The matrix  $A_N$  is Hermitian (resp., symmetric) and deterministic with all but finitely many eigenvalues equal to zero. We first show that, as soon as the first largest or last smallest eigenvalues of  $A_N$  are sufficiently far from zero, the corresponding eigenvalues of  $M_N$  almost surely exit the limiting semicircle compact support as the size  $N$  becomes large. The corresponding limits are universal in the sense that they only involve the variance of the entries of  $W_N$ . On the other hand, when  $A_N$  is diagonal with a sole simple nonnull eigenvalue large enough, we prove that the fluctuations of the largest eigenvalue are not universal and vary with the particular distribution of the entries of  $W_N$ .

**1. Introduction.** This paper lies in the lineage of recent works studying the influence of some perturbations on the asymptotic spectrum of classical random matrix models. Such questions come from statistics (cf. [20]) and appeared in the framework of empirical covariance matrices, also called non-white Wishart matrices or spiked population models, considered by Baik, Ben Arous and Pécché [8] and by Baik and Silverstein [9]. The work [8] deals with random sample covariance matrices  $(S_N)_N$  defined by

$$(1.1) \quad S_N = \frac{1}{N} Y_N^* Y_N,$$

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where  $Y_N$  is a  $p \times N$  complex matrix whose sample column vectors are i.i.d., centered, Gaussian and of covariance matrix a deterministic Hermitian matrix  $\Sigma_p$  having all but finitely many eigenvalues equal to 1. Besides, the size of the samples  $N$  and the size of the population  $p = p_N$  are assumed of the same order (as  $N \rightarrow \infty$ ). The authors of [8] first noticed that, as in the classical case (known as the Wishart model) where  $\Sigma_p = I_p$  is the identity matrix, the global limiting behavior of the spectrum of  $S_N$  is not affected by the matrix  $\Sigma_p$ . Thus, the limiting spectral measure is the well-known Marchenko–Pastur law. On the other hand, they pointed out a phase transition phenomenon for the fluctuations of the largest eigenvalue according to the value of the largest eigenvalue(s) of  $\Sigma_p$ . The approach of [8] does not extend to the real Gaussian setting and the whole analogue of their result is still an open question. Nevertheless, Paul was able to establish in [25] the Gaussian fluctuations of the largest eigenvalue of the real Gaussian matrix  $S_N$  when the largest eigenvalue of  $\Sigma_p$  is simple and sufficiently larger than 1. More recently, Baik and Silverstein investigated in [9] the almost sure limiting behavior of the extremal eigenvalues of complex or real nonnecessarily Gaussian matrices. Under assumptions on the first four moments of the entries of  $Y_N$ , they showed in particular that when exactly  $k$  eigenvalues of  $\Sigma_p$  are far from 1, the  $k$  first eigenvalues of  $S_N$  are almost surely outside the limiting Marchenko–Pastur support. Fluctuations of the eigenvalues that jump are universal and have been recently found by Bai and Yao in [6] (we refer the reader to [6] for the precise restrictions made on the definition of the covariance matrix  $\Sigma_p$ ). Note that the problem of the fluctuations in the very general setting of [9] is still open.

Our purpose here is to investigate the asymptotic behavior of the first extremal eigenvalues of some complex or real Deformed Wigner matrices. These models can be seen as the additive analogue of the spiked population models and are defined by a sequence  $(M_N)_N$  given by

$$(1.2) \quad M_N = \frac{1}{\sqrt{N}}W_N + A_N := X_N + A_N,$$

where  $W_N$  is a Wigner matrix such that the common distribution of its entries satisfies some technical conditions [given in (i) below] and  $A_N$  is a deterministic matrix of finite rank. We establish the analogue of the main result of [9], namely that, once  $A_N$  has exactly  $k$  (fixed) eigenvalues far enough from zero, the  $k$  first eigenvalues of  $M_N$  jump almost surely outside the limiting semicircle support. This result is universal (as the one of [9]) since the corresponding limits only involve the variance of the entries of  $W_N$ . On the other hand, at the level of the fluctuations, we exhibit a striking phenomenon in the particular case where  $A_N$  is diagonal with a sole simple nonnull eigenvalue large enough. Indeed, we find that in this case, the fluctuations of the largest eigenvalue of  $M_N$  are not universal and strongly depend

on the particular law of the entries of  $W_N$ . More precisely, we prove that the limiting distribution of the (properly rescaled) largest eigenvalue of  $M_N$  is the convolution of the distribution of the entries of  $W_N$  with a Gaussian law. In particular, if the entries of  $W_N$  are not Gaussian, the fluctuations of the largest eigenvalue of  $M_N$  are not Gaussian.

In the following section, we first give the precise definition of the Deformed Wigner matrices (1.2) considered in this paper and we recall the known results on their asymptotic spectrum. Then, we present our results and sketch the proofs. We also outline the organization of the paper.

**2. Model and results.** Throughout this paper, we consider complex or real Deformed Wigner matrices  $(M_N)_N$  of the form (1.2) where the matrices  $W_N$  and  $A_N$  are defined as follows:

- (i)  $W_N$  is an  $N \times N$  Wigner Hermitian (resp., symmetric) matrix such that the  $N^2$  random variables  $(W_N)_{ii}$ ,  $\sqrt{2}\Re((W_N)_{ij})_{i < j}$ ,  $\sqrt{2}\Im((W_N)_{ij})_{i < j}$  [resp., the  $N(N+1)/2$  random variables  $\frac{1}{\sqrt{2}}(W_N)_{ii}$ ,  $(W_N)_{ij}$ ,  $i < j$ ] are independent identically distributed with a symmetric distribution  $\mu$  of variance  $\sigma^2$  and satisfying a Poincaré inequality (see Section 3).
- (ii)  $A_N$  is a deterministic Hermitian (resp., symmetric) matrix of fixed finite rank  $r$  and built from a family of  $J$  fixed real numbers  $\theta_1 > \dots > \theta_J$  independent of  $N$  with some  $j_0$  such that  $\theta_{j_0} = 0$ . We assume that the non-null eigenvalues  $\theta_j$  of  $A_N$  are of fixed multiplicity  $k_j$  (with  $\sum_{j \neq j_0} k_j = r$ ), that is,  $A_N$  is similar to the diagonal matrix

$$D_N = \text{diag} \left( \underbrace{\theta_1, \dots, \theta_1}_{k_1}, \dots, \underbrace{\theta_{j_0-1}, \dots, \theta_{j_0-1}}_{k_{j_0-1}}, \underbrace{0, \dots, 0}_{N-r}, \underbrace{\theta_{j_0+1}, \dots, \theta_{j_0+1}}_{k_{j_0+1}}, \dots, \underbrace{\theta_J, \dots, \theta_J}_{k_J} \right). \tag{2.1}$$

Before going into the details of the results, we want to point out that the condition made on  $\mu$  (namely that  $\mu$  satisfies a Poincaré inequality) is just a technical condition: we conjecture that our results still hold under weaker assumptions (see Remark 2.1 below). Nevertheless, a lot of measures satisfy a Poincaré inequality (we refer the reader to [12] for a characterization of such measures on  $\mathbb{R}$ ; see also [1]). For instance, consider  $\mu(dx) = \exp(-|x|^\alpha) dx$  with  $\alpha \geq 1$ .

Furthermore, note that this condition implies that  $\mu$  has moments of any order (cf. Corollary 3.2 and Proposition 1.10 in [22]).

Let us now introduce some notations. When the entries of  $W_N$  are further assumed to be Gaussian, that is, in the complex (resp., real) setting when  $W_N$  is of the so-called GUE (resp., GOE), we will write  $W_N^G$  instead of  $W_N$ . Then  $X_N^G := W_N^G / \sqrt{N}$  will be said to be of the  $\text{GU}(\text{O})\text{E}(N, \frac{\sigma^2}{N})$  and we will let  $M_N^G = X_N^G + A_N$  be the corresponding Deformed  $\text{GU}(\text{O})\text{E}$  model.

In the following, given an arbitrary Hermitian matrix  $B$  of order  $N$ , we will denote by  $\lambda_1(B) \geq \dots \geq \lambda_N(B)$  its  $N$  ordered eigenvalues and by  $\mu_B = \frac{1}{N} \sum_{i=1}^N \delta_{\lambda_i(B)}$  its empirical measure.  $\text{Spect}(B)$  will denote the spectrum of  $B$ . For notational convenience, we will also set  $\lambda_0(B) = +\infty$  and  $\lambda_{N+1}(B) = -\infty$ .

The Deformed Wigner model is built in such a way that the Wigner theorem is still satisfied. Thus, as in the classical Wigner model ( $A_N \equiv 0$ ), the spectral measure  $(\mu_{M_N})$  converges a.s. toward the semicircle law  $\mu_{sc}$  whose density is given by

$$(2.2) \quad \frac{d\mu_{sc}}{dx}(x) = \frac{1}{2\pi\sigma^2} \sqrt{4\sigma^2 - x^2} 1_{[-2\sigma, 2\sigma]}(x).$$

This result follows from Lemma 2.2 of [2]. Note that it only relies on the two first moment assumptions on the entries of  $W_N$  and the fact that the  $A_N$ 's are of finite rank.

On the other hand, the asymptotic behavior of the extremal eigenvalues may be affected by the perturbation  $A_N$ . Recently, P ech e studied in [26] the Deformed GUE under a finite rank perturbation  $A_N$  defined by (ii). Following the method of [8], she highlighted the effects of the nonnull eigenvalues of  $A_N$  at the level of the fluctuations of the largest eigenvalue of  $M_N^G$ . To explain this in more detail, let us recall that when  $A_N \equiv 0$ , it was established in [33] that as  $N \rightarrow \infty$ ,

$$(2.3) \quad \sigma^{-1} N^{2/3} (\lambda_1(X_N^G) - 2\sigma) \xrightarrow{\mathcal{L}} F_2,$$

where  $F_2$  is the well-known GUE Tracy–Widom distribution (see [33] for the precise definition). Dealing with the Deformed GUE  $M_N^G$ , it appears that this result is modified as soon as the first largest eigenvalue(s) of  $A_N$  is (are) quite far from zero. In the particular case of a rank-1 perturbation  $A_N$  having a fixed nonnull eigenvalue  $\theta > 0$ , [26] proved that the fluctuations of the largest eigenvalue of  $M_N^G$  are still given by (2.3) when  $\theta$  is small enough and precisely when  $\theta < \sigma$ . The limiting law is changed when  $\theta = \sigma$ . As soon as  $\theta > \sigma$ , [26] established that the largest eigenvalue  $\lambda_1(M_N^G)$  fluctuates around

$$(2.4) \quad \rho_\theta = \theta + \frac{\sigma^2}{\theta}$$

(which is  $> 2\sigma$  since  $\theta > \sigma$ ) as

$$(2.5) \quad \sqrt{N} (\lambda_1(M_N^G) - \rho_\theta) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \sigma_\theta^2),$$

where

$$(2.6) \quad \sigma_\theta = (\sigma/\theta) \sqrt{\theta^2 - \sigma^2}.$$

Similar results are conjectured for the Deformed GOE but P ech e emphasized that her approach fails in the real framework. Indeed, it is based on the

explicit Fredholm determinantal representation for the distribution of the largest eigenvalue(s) that is specific to the complex setting. Nevertheless, Maïda [23] obtained a large deviation principle for the largest eigenvalue of the Deformed GOE  $M_N^G$  under a rank-1 deformation  $A_N$ ; from this result she could deduce the almost sure limit with respect to the nonnull eigenvalue of  $A_N$ . Thus, under a rank-1 perturbation  $A_N$  such that  $D_N = \text{diag}(\theta, 0, \dots, 0)$  where  $\theta > 0$ , [23] showed that

$$(2.7) \quad \lambda_1(M_N^G) \xrightarrow{\text{a.s.}} \rho_\theta \quad \text{if } \theta > \sigma$$

and

$$(2.8) \quad \lambda_1(M_N^G) \xrightarrow{\text{a.s.}} 2\sigma \quad \text{if } \theta \leq \sigma.$$

Note that the approach of [23] extends with minor modifications to the Deformed GUE. Following the investigations of [9] in the context of general spiked population models, one can conjecture that such a phenomenon holds in a more general and nonnecessarily Gaussian setting. The first result of our paper, namely the following Theorem 2.1, is related to this question. Before being more explicit, let us recall that when  $A_N \equiv 0$ , the whole spectrum of the rescaled complex or real Wigner matrix  $X_N = W_N/\sqrt{N}$  belongs almost surely to the semicircle support  $[-2\sigma, 2\sigma]$  as  $N$  goes to infinity and that (cf. [7] or Theorem 2.12 in [2])

$$(2.9) \quad \lambda_1(X_N) \xrightarrow{\text{a.s.}} 2\sigma \quad \text{and} \quad \lambda_N(X_N) \xrightarrow{\text{a.s.}} -2\sigma.$$

Note that this last result holds true in a more general setting than the one considered here (see [7] for details) and in particular only requires the finiteness of the fourth moment of the law  $\mu$ . Moreover, one can readily extend the previous limits to the first extremal eigenvalues of  $X_N$ , that is,

$$(2.10) \quad \text{for any fixed } k \geq 1, \quad \lambda_k(X_N) \xrightarrow{\text{a.s.}} 2\sigma \quad \text{and} \quad \lambda_{N-k}(X_N) \xrightarrow{\text{a.s.}} -2\sigma.$$

Here, we prove that, under the assumptions (i)–(ii), (2.10) fails when some of the  $\theta_j$ 's are sufficiently far from zero: as soon as some of the first largest (resp., last smallest) nonnull eigenvalues  $\theta_j$  of  $A_N$  are taken strictly larger than  $\sigma$  (resp., strictly smaller than  $-\sigma$ ), the same part of the spectrum of  $M_N$  almost surely exits the semicircle support  $[-2\sigma, 2\sigma]$  as  $N \rightarrow \infty$  and the new limits are the  $\rho_{\theta_j}$ 's defined by

$$(2.11) \quad \rho_{\theta_j} = \theta_j + \frac{\sigma^2}{\theta_j}.$$

Observe that  $\rho_{\theta_j}$  is  $> 2\sigma$  (resp.,  $< -2\sigma$ ) when  $\theta_j > \sigma$  (resp.,  $< -\sigma$ ) (and  $\rho_{\theta_j} = \pm 2\sigma$  if  $\theta_j = \pm\sigma$ ).

Here is the precise formulation of our result. For definiteness, we set  $k_1 + \dots + k_{j-1} := 0$  if  $j = 1$ .

**THEOREM 2.1.** *Let  $J_{+\sigma}$  (resp.,  $J_{-\sigma}$ ) be the number of  $j$ 's such that  $\theta_j > \sigma$  (resp.,  $\theta_j < -\sigma$ ).*

- (a)  $\forall 1 \leq j \leq J_{+\sigma}, \forall 1 \leq i \leq k_j, \lambda_{k_1+\dots+k_{j-1}+i}(M_N) \longrightarrow \rho_{\theta_j}$  a.s.
- (b)  $\lambda_{k_1+\dots+k_{J_{+\sigma}}+1}(M_N) \longrightarrow 2\sigma$  a.s.
- (c)  $\lambda_{k_1+\dots+k_{J_{-\sigma}}}(M_N) \longrightarrow -2\sigma$  a.s.
- (d)  $\forall j \geq J - J_{-\sigma} + 1, \forall 1 \leq i \leq k_j, \lambda_{k_1+\dots+k_{j-1}+i}(M_N) \longrightarrow \rho_{\theta_j}$  a.s.

**REMARK 2.1.** Following [9], one can expect that this theorem holds true in a more general setting than the one considered here, namely one that would only require four first moment conditions on the law  $\mu$  of the Wigner entries. As we will explain in the following, the assumption that  $\mu$  satisfies a Poincaré inequality is actually fundamental in our reasoning since we will need several variance estimates.

This theorem will be proved in Section 4. The second part of this work is devoted to the study of the particular rank-1 diagonal deformation  $A_N = \text{diag}(\theta, 0, \dots, 0)$  such that  $\theta > \sigma$ . We investigate the fluctuations of the largest eigenvalue of any real or complex Deformed model  $M_N$  satisfying (i) around its limit  $\rho_\theta$ . We obtain the following result.

**THEOREM 2.2.** *Let  $A_N = \text{diag}(\theta, 0, \dots, 0)$  with  $\theta > \sigma$ . Define*

$$(2.12) \quad v_\theta = \frac{t}{4} \left( \frac{m_4 - 3\sigma^4}{\theta^2} \right) + \frac{t}{2} \frac{\sigma^4}{\theta^2 - \sigma^2},$$

where  $t = 4$  (resp.,  $t = 2$ ) when  $W_N$  is real (resp., complex) and  $m_4 := \int x^4 d\mu(x)$ . Then

$$(2.13) \quad \sqrt{N} \left( 1 - \frac{\sigma^2}{\theta^2} \right)^{-1} (\lambda_1(M_N) - \rho_\theta) \xrightarrow{\mathcal{L}} \mu * \mathcal{N}(0, v_\theta).$$

Note that when  $m_4 = 3\sigma^4$  as in the Gaussian case, the variance of the limiting distribution of  $\sqrt{N}(\lambda_1(M_N) - \rho_\theta)$  is equal to  $\sigma_\theta^2$  (resp.,  $2\sigma_\theta^2$ ) in the complex (resp., real) setting [with  $\sigma_\theta$  given by (2.6)].

**REMARK 2.2.** Since  $\mu$  is symmetric, it readily follows from Theorem 2.2 that when  $A_N = \text{diag}(\theta, 0, \dots, 0)$  and  $\theta < -\sigma$ , the smallest eigenvalue of  $M_N$  fluctuates as  $\sqrt{N}(1 - \sigma^2/\theta^2)^{-1}(\lambda_N(M_N) - \rho_\theta) \xrightarrow{\mathcal{L}} \mu * \mathcal{N}(0, v_\theta)$ .

In particular, one derives the analogue of (2.5) for the Deformed GOE:

**THEOREM 2.3.** *Let  $A_N$  be an arbitrary deterministic symmetric matrix of rank 1 having a nonnull eigenvalue  $\theta$  such that  $\theta > \sigma$ . Then the largest eigenvalue of the Deformed GOE fluctuates as*

$$(2.14) \quad \sqrt{N}(\lambda_1(M_N^G) - \rho_\theta) \xrightarrow{\mathcal{L}} \mathcal{N}(0, 2\sigma_\theta^2).$$

Obviously, thanks to the orthogonal invariance of the GOE, this result is a direct consequence of Theorem 2.2.

It is worth noticing that, according to the Cramér–Lévy theorem (cf. [14], Theorem 1, page 525), the limiting distribution (2.13) is not Gaussian if  $\mu$  is not Gaussian. Thus, (2.13) depends on the particular law  $\mu$  of the entries of the Wigner matrix  $W_N$  which implies the nonuniversality of the fluctuations of the largest eigenvalue of rank-1 diagonal deformation of symmetric or Hermitian Wigner matrices (as conjectured in Remark 1.7 of [16]).

The latter also shows that in the non-Gaussian setting, the fluctuations of the largest eigenvalue depend, not only on the spectrum of the deformation  $A_N$ , but also on the particular definition of the matrix  $A_N$ . Indeed, in collaboration with S. Péché, the third author of the present article has recently stated in [16] the universality of the fluctuations of some Deformed Wigner models under a full deformation  $A_N$  defined by  $(A_N)_{ij} = \theta/N$  for all  $1 \leq i, j \leq N$  (see also [17]). Before giving some details on this work, we have to specify that [16] considered Deformed models such that the entries of the Wigner matrix  $W_N$  have sub-Gaussian moments. Nevertheless, thanks to the analysis made in [27], one can observe that the assumptions of [16] can be reduced and that it is, for example, sufficient to assume that the  $W_{i,j}$ 's have moments of any order. Thus, the conclusions of [16] apply to the setting considered in our paper. The main result of [16] establishes the universality of the fluctuations of the largest eigenvalue of the complex Deformed model  $M_N$  associated to a full deformation  $A_N$  and for any value of the parameter  $\theta$ . In particular, when  $\theta > \sigma$ , it is proved therein the universality of the Gaussian fluctuations (2.5). The approach of [16] is mainly based on a combinatorial method inspired by the work [31] (which handles the non-Deformed Wigner model) and some results of [26] on the Deformed GUE. The combinatorial arguments of [16] also work (with minor modifications) in the real framework and yield the universality of the fluctuations if  $\theta < \sigma$ . In the case where  $\theta > \sigma$  which is of particular interest here, the analysis made in [16] reduces the universality problem in the real setting to the knowledge of the particular Deformed GOE model (this remark is also valid in the case where  $\theta = \sigma$ ). Here, we will prove the needed results on the Deformed GOE which, thanks to the analysis of [16] and [27], allow us to claim the following universality.

**THEOREM 2.4.** *Let  $A_N$  be a full perturbation given by  $(A_N)_{ij} = \theta/N$  for all  $(i, j)$ . Assume that  $\theta > \sigma$ . Let  $W_N$  be an arbitrary real Wigner matrix*



with the underlying measure  $\mu$  being symmetric with a variance  $\sigma^2$  and such that  $\int |x|^q d\mu(x) < +\infty$  for any  $q$  in  $\mathbb{N}$ .

Then the largest eigenvalue of the Deformed model  $M_N$  has the Gaussian fluctuations (2.14).

REMARK 2.3. To be complete, let us notice that the previous result still holds when we allow the distribution  $\nu$  of the diagonal entries of  $W_N$  to be different from  $\mu$  provided that  $\nu$  is symmetric and has moments of any order.

The fundamental tool of this paper is the Stieltjes transform. For  $z \in \mathbb{C} \setminus \mathbb{R}$ , we denote the resolvent of the matrix  $M_N$  by

$$G_N(z) = (zI_N - M_N)^{-1}$$

and the Stieltjes transform of the expectation of the empirical measure of the eigenvalues of  $M_N$  by

$$g_N(z) = \mathbb{E}(\text{tr}_N(G_N(z))),$$

where  $\text{tr}_N$  is the normalized trace. We also denote by

$$g_\sigma(z) = \mathbb{E}((z - s)^{-1})$$

the Stieltjes transform<sup>1</sup> of a variable  $s$  with semicircular distribution  $\mu_{sc}$ .

Theorem 2.1 is the analogue of the main statement of [9] established in the context of general spiked population models. The conclusion of [9] requires numerous results obtained previously by Silverstein and co-authors in [30], [3] and [4] (a summary of all this literature can be found in [2], pages 671–675). From very clever and tedious manipulations of some Stieltjes transforms and the use of the matricial representation (1.1), these works highlight a very close link between the spectra of the Wishart matrices and the covariance matrix (for quite general covariance matrix which includes the spiked population model). Our approach mimics the one of [9]. Thus, using the fact that the Deformed Wigner model is the additive analogue of the spiked population model, several arguments can be quite easily adapted here (this point has been explained in Chapter 4 of the Ph.D. thesis [15]). Actually, the main point in the proof consists in establishing that for any  $\varepsilon > 0$ , almost surely,

$$(2.15) \quad \text{Spect}(M_N) \subset K_\sigma^\varepsilon(\theta_1, \dots, \theta_J)$$

for all  $N$  large, where we have defined

$$K_\sigma^\varepsilon(\theta_1, \dots, \theta_J) = K_\sigma(\theta_1, \dots, \theta_J) + (-\varepsilon, \varepsilon)$$

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<sup>1</sup>Note that in some papers to which we make reference, the Stieltjes transform is defined with the opposite sign.

and

$$K_\sigma(\theta_1, \dots, \theta_J) := \{\rho_{\theta_J}; \dots; \rho_{\theta_{J-J_{-\sigma}+1}}\} \cup [-2\sigma, 2\sigma] \cup \{\rho_{\theta_{J+\sigma}}; \dots; \rho_{\theta_1}\}.$$

This point is the analogue of the main result of [3]. The analysis of [3] is based on technical and numerous considerations of Stieltjes transforms strongly related to the Wishart context and that cannot be directly transposed here. Our approach to prove such an inclusion of the spectrum of  $M_N$  is very different from the one of [3]. Indeed, we use the methods developed by Haagerup and Thorbjørnsen in [18], by Schultz [29] and by the two first authors of the present article [13]. The key point of this approach is to obtain a precise estimation at any point  $z \in \mathbb{C} \setminus \mathbb{R}$  of the following type:

$$(2.16) \quad g_\sigma(z) - g_N(z) + \frac{1}{N}L_\sigma(z) = O\left(\frac{1}{N^2}\right),$$

where  $L_\sigma$  is the Stieltjes transform of a distribution  $\Lambda_\sigma$  with compact support in  $K_\sigma(\theta_1, \dots, \theta_J)$ . Indeed such an estimation allows us through the inverse Stieltjes transform and some variance estimates to deduce that a.s.,  $\text{tr}_N(\mathbb{1}_{cK_\sigma^\varepsilon(\theta_1, \dots, \theta_J)}(M_N)) = O(N^{-4/3})$ . Thus the number of eigenvalues of  $M_N$  in  ${}^cK_\sigma^\varepsilon(\theta_1, \dots, \theta_J)$  is almost surely an  $O(N^{-1/3})$  and since for each  $N$  this number has to be an integer, we deduce that it is actually equal to zero as  $N$  goes to infinity.

Dealing with the particular diagonal perturbation  $A_N = \text{diag}(\theta, 0, \dots, 0)$  such that  $\theta > \sigma$ , we obtain the fluctuations of the largest eigenvalue  $\lambda_1(M_N)$  (Theorem 2.2) by an approach close to the one of [25] and the ideas of [11]. The reasoning relies on the writing of the rescaled variable  $\sqrt{N}(\lambda_1(M_N) - \rho_\theta)$  in terms of the resolvent of a non-Deformed Wigner matrix. Then, to complete the analysis of [16] and justify Theorem 2.4, we focus on the particular Deformed GOE model and improve the previous convergence at the level of Laplace transform.

The paper is organized as follows. In Section 3, we introduce preliminary lemmas which will be of basic use later on. Section 4 is devoted to the proof of Theorem 2.1. We first establish an equation (called master equation or master inequality) satisfied by  $g_N$  up to some correction of order  $\frac{1}{N^2}$  (see Section 4.1). Then we explain how this master equation gives rise to an estimation of type (2.16) and thus to the inclusion (2.15) of the spectrum of  $M_N$  in  $K_\sigma^\varepsilon(\theta_1, \dots, \theta_J)$  (see Sections 4.2 and 4.3). In Section 4.4, we use this inclusion to relate the asymptotic spectra of  $A_N$  and  $M_N$  and then deduce Theorem 2.1. Section 5 deals with the fluctuations results. The proof of Theorem 2.2 is given in Section 5.2; Theorem 2.4 is justified in Section 5.3.

**3. Basic lemmas.** We assume that the distribution  $\mu$  of the entries of the Wigner matrix  $W_N$  satisfies a Poincaré inequality: there exists a positive constant  $C$  such that for any  $\mathcal{C}^1$  function  $f: \mathbb{R} \rightarrow \mathbb{C}$  such that  $f$  and  $f'$  are in  $L^2(\mu)$ ,

$$\mathbf{V}(f) \leq C \int |f'|^2 d\mu,$$

with  $\mathbf{V}(f) = \mathbb{E}(|f - \mathbb{E}(f)|^2)$ .

Let  $\text{Tr}$  denote the classical trace.

For any matrix  $M$ , define  $\|M\|_2 = (\text{Tr}(M^*M))^{1/2}$  the Hilbert–Schmidt norm. Let  $\Psi: (M_N(\mathbb{C})_{sa}) \rightarrow \mathbb{R}^{N^2}$  [resp.,  $\Psi: (M_N(\mathbb{C})_s) \rightarrow \mathbb{R}^{N(N+1)/2}$ ] be the canonical isomorphism which maps an Hermitian (resp., symmetric) matrix  $M$  to the real parts and the imaginary parts of its entries (resp., to the entries)  $(M)_{ij}, i \leq j$ .

**LEMMA 3.1.** *Let  $M_N$  be the complex (resp., real) Wigner Deformed matrix introduced in Section 2. For any  $\mathcal{C}^1$  function  $f: \mathbb{R}^{N^2}$  (resp.,  $\mathbb{R}^{N(N+1)/2}$ )  $\rightarrow \mathbb{C}$  such that  $f$  and the gradient  $\nabla(f)$  are both polynomially bounded,*

$$(3.1) \quad \mathbf{V}[f \circ \Psi(M_N)] \leq \frac{C}{N} \mathbb{E}\{\|\nabla[f \circ \Psi(M_N)]\|_2^2\}.$$

**PROOF.** According to Lemma 3.2 in [13],

$$(3.2) \quad \mathbf{V}[f \circ \Psi(X_N)] \leq \frac{C}{N} \mathbb{E}\{\|\nabla[f \circ \Psi(X_N)]\|_2^2\}.$$

Note that even if the result in [13] is stated in the Hermitian case, the proof is valid and the result still holds in the symmetric case. Now (3.1) follows putting  $g(x_{ij}; i \leq j) := f(x_{ij} + (A_N)_{ij}; i \leq j)$  in (3.2) and noticing that the  $(A_N)_{ij}$  are uniformly bounded in  $i, j, N$ .  $\square$

This lemma will be useful to estimate many variances. Now, we recall some useful properties of the resolvent (see [13, 21]).

**LEMMA 3.2.** *For an  $N \times N$  Hermitian or symmetric matrix  $M$ , for any  $z \in \mathbb{C} \setminus \text{Spect}(M)$ , we denote by  $G(z) := (zI_N - M)^{-1}$  the resolvent of  $M$ .*

*Let  $z \in \mathbb{C} \setminus \mathbb{R}$ .*

- (i)  $\|G(z)\| \leq |\Im m(z)|^{-1}$  where  $\|\cdot\|$  denotes the operator norm.
- (ii)  $|G(z)_{ij}| \leq |\Im m(z)|^{-1}$  for all  $i, j = 1, \dots, N$ .
- (iii) For  $p \geq 2$ ,

$$\frac{1}{N} \sum_{i,j=1}^N |G(z)_{ij}|^p \leq (|\Im m(z)|^{-1})^p.$$

(iv) *The derivative with respect to  $M$  of the resolvent  $G(z)$  satisfies*

$$G'_M(z) \cdot B = G(z)BG(z) \quad \text{for any matrix } B.$$

(v) *Let  $z \in \mathbb{C}$  such that  $|z| > \|M\|$ ; we have*

$$\|G(z)\| \leq \frac{1}{|z| - \|M\|}.$$

PROOF. We just mention that (v) comes readily noticing that the eigenvalues of the normal matrix  $G(z)$  are the  $\frac{1}{z - \lambda_i(M)}$ ,  $i = 1, \dots, N$ .  $\square$

We will also need the following estimations on the Stieltjes transform  $g_\sigma$  of the semicircular distribution  $\mu_{sc}$ .

LEMMA 3.3.  *$g_\sigma$  is analytic on  $\mathbb{C} \setminus [-2\sigma, 2\sigma]$  and*

(i)  $\forall z \in \{z \in \mathbb{C} : \Im m(z) \neq 0\}$ ,

$$(3.3) \quad \sigma^2 g_\sigma^2(z) - z g_\sigma(z) + 1 = 0,$$

$$(3.4) \quad |g_\sigma(z)| \leq |\Im m(z)|^{-1},$$

$$(3.5) \quad |g_\sigma(z)^{-1}| \leq |z| + \sigma^2 |\Im m(z)|^{-1},$$

$$(3.6) \quad |g'_\sigma(z)| = \left| \int \frac{1}{(z-t)^2} d\mu_\sigma(t) \right| \leq |\Im m(z)|^{-2},$$

$$(3.7) \quad \text{for } a > 0, \theta \in \mathbb{R}, \quad \left| \frac{1}{a g_\sigma(z) - z + \theta} \right| \leq |\Im m(z)|^{-1}.$$

(ii)  $\forall z \in \{z \in \mathbb{C} : |z| > 2\sigma\}$ ,

$$(3.8) \quad |g_\sigma(z)| \leq \frac{1}{|z| - 2\sigma}.$$

$$(3.9) \quad |g'_\sigma(z)| = \left| \int \frac{1}{(z-t)^2} d\mu_\sigma(t) \right| \leq \frac{1}{(|z| - 2\sigma)^2},$$

$$(3.10) \quad |g_\sigma(z)|^{-1} \leq |z| + \frac{\sigma^2}{|z| - 2\sigma}.$$

PROOF. For (3.3), we refer the reader to Section 3.1 of [2]. Equation (3.7) is a consequence of  $\Im m(g_\sigma(z))\Im m(z) < 0$ . Other inequalities derive from (3.3) and the definition of  $g_\sigma$ .  $\square$

**4. Almost sure convergence of the first extremal eigenvalues.** Sections 4.1, 4.2 and 4.3 below describe the different steps of the proof of the inclusion (2.15). We choose to develop the case of the complex Deformed Wigner model and just to point out some differences with the real model case (at the end of Section 4.3) since the approach would be basically the same. In these sections, we will often refer the reader to the paper [13] where the authors deal with several independent non-Deformed Wigner matrices. The reader needs to fix  $r = 1$ ,  $m = 1$ ,  $a_0 = 0$ ,  $a_1 = \sigma$  and to change the notation  $\lambda = z$ ,  $G_N = g_N$ ,  $G = g_\sigma$  in [13] in order to use the different proofs we refer to in the present framework. We shall denote by  $P_k$  any polynomial of degree  $k$  with positive coefficients and by  $C$ ,  $K$  any constants;  $P_k$ ,  $C$ ,  $K$  can depend on the fixed eigenvalues of  $A_N$  and may vary from line to line. We also adopt the following convention to simplify the writing: we sometimes state in the proofs below that a quantity  $\Delta_N(z)$ ,  $z \in \mathbb{C} \setminus \mathbb{R}$  is  $O(N^{-p})$ ,  $p = 1, 2$ . This means precisely that

$$|\Delta_N(z)| \leq (|z| + K)^l \frac{P_k(|\Im m(z)|^{-1})}{N^p}$$

for some  $k$  and some  $l$  and we give the precise majoration in the statements of the theorems or propositions.

Section 4.4 explains how to deduce Theorem 2.1 from the inclusion (2.15).

The goal of Sections 4.1 and 4.2 is to establish Proposition 4.4 below which is fundamental in the proof of the inclusion (2.15). Before describing rigorously the different ideas of these two sections, let us help the reader's intuition by a heuristic understanding of the approach. Assume that we can establish that  $g_N(z)$  satisfied the rough quadratic equation (also called master inequality):

$$\sigma^2 g_N^2(z) - z g_N(z) + 1 + \frac{1}{N} E_\sigma(z) = O\left(\frac{1}{N^2}\right).$$

Then, for any suitable  $z$ , divided by  $g_N(z)$  the last approximation would provide us an estimation of  $\Lambda_N(z) - z$  where  $\Lambda_N(z) = z_\sigma(g_N(z))$  with  $z_\sigma(g) = \frac{1}{g} + \sigma^2 g$  being the inverse function of  $g_\sigma$  (see Lemma 4.4 below). Then, intuitively, a Taylor expansion of  $g_\sigma$  between  $\Lambda_N(z)$  and  $z$  would lead to an estimation of the type

$$g_N(z) - g_\sigma(z) = -\frac{1}{N} (g_\sigma(z))^{-1} g'_\sigma(z) E_\sigma(z) + O\left(\frac{1}{N^2}\right).$$

This intuitive process may throw light on the expression (4.20) of  $L_\sigma(z)$  in Proposition 4.4 below.

4.1. *The master equation.*

4.1.1. *A first master inequality.* In order to obtain a master equation for  $g_N(z)$ , we first consider the Gaussian case, that is,  $X_N = X_N^G$  is distributed as the  $\text{GUE}(N, \sigma^2/N)$  distribution.<sup>2</sup>

Let us recall the integration by parts formula for the Gaussian distribution.

LEMMA 4.1. *Let  $\Phi$  be a complex-valued  $C^1$  function on  $(M_N(\mathbb{C}))_{sa}$  and  $X_N \sim \text{GUE}(N, \sigma^2/N)$ . Then,*

$$(4.1) \quad \mathbb{E}[\phi'(X_N) \cdot H] = \frac{N}{\sigma^2} \mathbb{E}[\phi(X_N) \text{Tr}(X_N H)]$$

for any Hermitian matrix  $H$ , or by linearity for  $H = E_{jk}$ ,  $1 \leq j, k \leq N$  where  $E_{jk}$ ,  $1 \leq j, k \leq N$  is the canonical basis of the complex space of  $N \times N$  matrices.

We apply the above lemma to the function  $\Phi(X_N) = (G_N(z))_{ij} = ((zI_N - X_N - A_N)^{-1})_{ij}$ ,  $z \in \mathbb{C} \setminus \mathbb{R}$ ,  $1 \leq i, j \leq N$ . In order to simplify the notation, we write  $(G_N(z))_{ij} = G_{ij}$ . We obtain, for  $H = E_{ij}$ ,

$$\begin{aligned} \mathbb{E}((GHG)_{ij}) &= \frac{N}{\sigma^2} \mathbb{E}[G_{ij} \text{Tr}(X_N H)], \\ \mathbb{E}(G_{ii}G_{jj}) &= \frac{N}{\sigma^2} \mathbb{E}[G_{ij}(X_N)_{ji}]. \end{aligned}$$

Now, we consider the normalized sum  $\frac{1}{N^2} \sum_{ij}$  of the previous identities to obtain

$$\mathbb{E}((\text{tr}_N G)^2) = \frac{1}{\sigma^2} \mathbb{E}(\text{tr}_N(GX_N)).$$

Then, since

$GX_N = (z - X_N - A_N)^{-1}(X_N + A_N - zI_N - A_N + zI_N) = -I_N - GA_N + zG$ , we obtain the following master equation:

$$\mathbb{E}((\text{tr}_N G)^2) + \frac{1}{\sigma^2} (-z \mathbb{E}(\text{tr}_N G) + 1 + \mathbb{E}(\text{tr}_N GA_N)) = 0.$$

Now, it is well known (see [13, 18] and Lemma 3.1) that

$$\text{Var}(\text{tr}_N(G)) \leq \frac{C|\Im m(z)|^{-4}}{N^2}.$$

Thus, in the case where  $X_N = X_N^G$  we obtain:

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<sup>2</sup>Throughout this section, we will drop the subscript  $G$  in the interest of clarity.

PROPOSITION 4.1. *The Stieltjes transform  $g_N$  satisfies the following inequality:*

$$(4.2) \quad \left| \sigma^2 g_N^2(z) - z g_N(z) + 1 + \frac{1}{N} \mathbb{E}(\text{Tr}(G_N(z) A_N)) \right| \leq C \frac{|\Im m(z)|^{-4}}{N^2}.$$

Note that since  $A_N$  is of finite rank,  $\mathbb{E}(\text{Tr}(G_N(z) A_N)) \leq C$  where  $C$  is a constant independent of  $N$  (depending on the eigenvalues of  $A_N$  and  $z$ ).

We now explain how to obtain the corresponding (4.2) in the Wigner case. Since the computations are the same as in [13]<sup>3</sup> and [21],<sup>4</sup> we just give some hints of the proof.

STEP 1. The integration by parts formula for the Gaussian distribution is replaced by the following tool:

LEMMA 4.2. *Let  $\xi$  be a real-valued random variable such that  $\mathbb{E}(|\xi|^{p+2}) < \infty$ . Let  $\phi$  be a function from  $\mathbb{R}$  to  $\mathbb{C}$  such that the first  $p+1$  derivatives are continuous and bounded. Then,*

$$(4.3) \quad \mathbb{E}(\xi \phi(\xi)) = \sum_{a=0}^p \frac{\kappa_{a+1}}{a!} \mathbb{E}(\phi^{(a)}(\xi)) + \epsilon$$

where  $\kappa_a$  are the cumulants of  $\xi$ ,  $|\epsilon| \leq C \sup_t |\phi^{(p+1)}(t)| \mathbb{E}(|\xi|^{p+2})$ ,  $C$  depends on  $p$  only.

We apply this lemma with the function  $\phi(\xi)$  given, as before, by  $\phi(\xi) = G_{ij}$  and  $\xi$  is now one of the variables  $\Re e((X_N)_{kl})$ ,  $\Im m((X_N)_{kl})$ . Note that, since the above random variables are symmetric, only the odd derivatives in (4.3) give a nonnull term. Moreover, as we are concerned by estimation of order  $\frac{1}{N^2}$  of  $g_N$ , we only need to consider (4.3) up to the third derivative (see [13]). The computation of the first derivative will provide the same term as in the Gaussian case.

STEP 2. Study of the third derivative.

We refer to [13] or [21] for a detailed study of the third derivative. Using some bounds on  $G_N$  (see Lemma 3.2), we can prove that the only term arising from the third derivative in the master equation, giving a contribution of order  $\frac{1}{N}$ , is

$$\frac{1}{N} \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N G_{kk}^2 \right)^2 \right].$$

<sup>3</sup>This paper treats the case of several independent non-Deformed Wigner matrices.

<sup>4</sup>The authors considered a non-Deformed Wigner matrix in the symmetric real setting.

In conclusion, the first master equation in the Wigner case reads as follows:

THEOREM 4.1. *For  $z \in \mathbb{C} \setminus \mathbb{R}$ ,  $g_N(z)$  satisfies*

$$(4.4) \quad \left| \sigma^2 g_N(z)^2 - z g_N(z) + 1 + \frac{1}{N} \mathbb{E}[\text{Tr}(G_N(z)A_N)] \right. \\ \left. + \frac{1}{N} \frac{\kappa_4}{2} \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N (G_N(z))_{kk}^2 \right)^2 \right] \right| \leq \frac{P_6(|\Im m(z)|^{-1})}{N^2},$$

where  $\kappa_4$  is the fourth cumulant of the distribution  $\mu$ .

4.1.2. *Estimation of  $|g_N - g_\sigma|$ .* Since

$$|\mathbb{E}[\text{Tr}(G_N(z)A_N)]| \vee \left| \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N (G_N(z))_{kk}^2 \right)^2 \right] \right| \leq P_4(|\Im m(z)|^{-1}),$$

Theorem 4.1 implies that for any  $z \in \mathbb{C} \setminus \mathbb{R}$ ,

$$(4.5) \quad |\sigma^2 g_N(z)^2 - z g_N(z) + 1| \leq \frac{P_6(|\Im m(z)|^{-1})}{N}.$$

To estimate  $|g_N - g_\sigma|$  from (3.3) and (4.5), we follow the method initiated in [18] and [29]. We do not develop it here since it follows exactly the lines of Section 3.4 in [13] but we briefly recall the main arguments and results which will be useful later on. We define the open connected set

$$\mathcal{O}'_N = \left\{ z \in \mathbb{C}, \Im m(z) > 0, \frac{P_6(|\Im m(z)|^{-1})}{N} (\sigma^2 |\Im m(z)|^{-1} + |z|) < \frac{1}{4|\Im m(z)|^{-1}} \right\}.$$

For any  $z$  in  $\mathbb{C}$  such that  $\Im m(z) > 0$ , we set

$$(4.6) \quad \Lambda_N(z) := \sigma^2 g_N(z) + \frac{1}{g_N(z)}.$$

One can prove that for any  $z$  in  $\mathcal{O}'_N$ :

- $g_N(z) \neq 0$  and

$$(4.7) \quad \frac{1}{|g_N(z)|} \leq 2(\sigma^2 |\Im m(z)|^{-1} + |z|);$$

- from (4.5) and (4.7),

$$(4.8) \quad |\Lambda_N(z) - z| \leq \frac{P_6(|\Im m(z)|^{-1})}{N} 2(\sigma^2 |\Im m(z)|^{-1} + |z|)$$

and

$$(4.9) \quad \Im m(\Lambda_N(z)) \geq \frac{\Im m(z)}{2} > 0;$$



- writing (3.3) at the point  $\Lambda_N(z)$ , we easily get that

$$(4.10) \quad g_N(z) = g_\sigma(\Lambda_N(z))$$

on the nonempty open subset  $\mathcal{O}_N'' = \{z \in \mathcal{O}_N', \Im m(z) > \sqrt{2}\sigma\}$  and then on  $\mathcal{O}_N'$  by the principle of uniqueness of continuation.

Using

$$\begin{aligned} |g_N(z) - g_\sigma(z)| &= |\mathbb{E}[(z-s)^{-1}(\Lambda_N(z)-s)^{-1}(\Lambda_N(z)-z)]| \\ &\leq \Im m(z) \cdot \Im m(\Lambda_N(z)) \cdot |\Lambda_N(z) - z|, \end{aligned}$$

this allows us to get an estimation of  $|g_N(z) - g_\sigma(z)|$  on  $\mathcal{O}_N'$  and then to deduce:

PROPOSITION 4.2. *For any  $z \in \mathbb{C}$  such that  $\Im m(z) > 0$ ,*

$$(4.11) \quad |g_N(z) - g_\sigma(z)| \leq (|z| + K) \frac{P_9(|\Im m(z)|^{-1})}{N}.$$

4.1.3. *Study of the additional term  $\mathbb{E}[\text{Tr}(A_N G_N(z))]$ .* From now on and until the end of Section 4.1, we denote by  $\gamma_1, \dots, \gamma_r$  the nonnull eigenvalues of  $A_N$  ( $\gamma_i = \theta_j$  for some  $j \neq j_0$ ) in order to simplify the writing. Let  $U_N := U$  be a unitary matrix such that  $A_N = U^* \Delta U$  where  $\Delta$  is the diagonal matrix with entries  $\Delta_{ii} = \gamma_i$ ,  $i \leq r$ ;  $\Delta_{ii} = 0$ ,  $i > r$ . We set

$$(4.12) \quad h_N(z) := \mathbb{E}[\text{Tr}(A_N G_N(z))] = \sum_{k=1}^r \gamma_k \sum_{i,j=1}^N U_{ik}^* U_{kj} \mathbb{E}[G_{ji}].$$

Our aim is to express  $h_N(z)$  in terms of the Stieltjes transform  $g_N(z)$  for  $N$  large, using the integration by parts formula. Note that since we want an estimation of order  $O(N^{-2})$  in the master inequality (4.4), we only need an estimation of  $h_N(z)$  of order  $O(N^{-1})$ . As in the previous subsection, we first write the equation in the Gaussian case and then study the additional term (third derivative) in the Wigner case.

- (a) *Gaussian case.* Apply (4.1) to  $\Phi(X_N) = G_{jl}$  and  $H = E_{il}$  to get

$$\mathbb{E}[G_{ji} G_{ll}] = \frac{N}{\sigma^2} \mathbb{E}[G_{jl}(X_N)_{li}]$$

and

$$\frac{1}{N} \sum_{l=1}^N \mathbb{E}[G_{ji} G_{ll}] = \frac{1}{\sigma^2} \mathbb{E}[(GX_N)_{ji}].$$

Expressing  $GX_N$  in terms of  $GA_N$ , we obtain

$$(4.13) \quad I_{ji} := \sigma^2 \mathbb{E}[G_{ji} \text{tr}_N(G)] + \delta_{ij} - z \mathbb{E}[G_{ji}] + \mathbb{E}[(GA_N)_{ji}] = 0.$$

Now, we consider the sum  $\sum_{i,j} U_{ik}^* U_{kj} I_{ji}$ ,  $k = 1, \dots, r$  fixed and we denote  $\alpha_k = \sum_{i,j} U_{ik}^* U_{kj} G_{ji} = (UGU^*)_{kk}$ . Then, we have the following equality, using that  $U$  is unitary:

$$\sigma^2 \mathbb{E}[\alpha_k \text{tr}_N(G)] + 1 - z \mathbb{E}[\alpha_k] + \sum_{i,j} U_{ik}^* U_{kj} \mathbb{E}[(GA_N)_{ji}] = 0.$$

Now,

$$\begin{aligned} \sum_{i,j} U_{ik}^* U_{kj} \mathbb{E}[(GA_N)_{ji}] &= \mathbb{E}[(UGA_N U^*)_{kk}] \\ &= \mathbb{E}[(UGU^* \Delta U U^*)_{kk}] \\ &= \gamma_k \mathbb{E}[(UGU^*)_{kk}] = \gamma_k \mathbb{E}[\alpha_k]. \end{aligned}$$

Therefore,

$$\sigma^2 \mathbb{E}[\alpha_k \text{tr}_N(G)] + 1 + (\gamma_k - z) \mathbb{E}[\alpha_k] = 0.$$

Since  $\alpha_k$  is bounded and  $\text{Var}(\text{tr}_N(G)) = O(N^{-2})$ , we obtain

$$(4.14) \quad \mathbb{E}[\alpha_k] (\sigma^2 g_N(z) + \gamma_k - z) + 1 = O\left(\frac{1}{N}\right).$$

Then using (4.11) we deduce that  $\mathbb{E}[\alpha_k] (\sigma^2 g_\sigma(z) + \gamma_k - z) + 1 = O(N^{-1})$  and using (3.7):

$$(4.15) \quad h_N(z) = \sum_{k=1}^r \gamma_k \mathbb{E}[\alpha_k] = \sum_{k=1}^r \frac{\gamma_k}{z - \sigma^2 g_\sigma(z) - \gamma_k} + O\left(\frac{1}{N}\right).$$

(b) *The general Wigner case.* We shall prove that (4.14) still holds. We now rely on Lemma 4.2 to obtain the analogue of (4.13):

$$(4.16) \quad \begin{aligned} J_{ij} &:= \sigma^2 \mathbb{E}[G_{ji} \text{tr}_N G] + \delta_{ij} - z \mathbb{E}[G_{ji}] \\ &+ \mathbb{E}[(GA_N)_{ji}] + \frac{\kappa_4}{6N^2} \sum_{l=1}^N \mathbb{E}[A_{i,j,l}] \\ &= O\left(\frac{1}{N^2}\right). \end{aligned}$$

The term  $A_{i,j,l}$  is a fixed linear combination of the third derivative of  $\Phi := G_{jl}$  with respect to  $\text{Re}(X_N)_{il}$  (i.e., in the direction  $e_{il} = E_{il} + E_{li}$ ) and  $\text{Im}(X_N)_{il}$  [i.e., in the direction  $f_{il} := \sqrt{-1}(E_{il} - E_{li})$ ]. We do not need to write the exact form of this term since we just want to show that this term will give a contribution of order  $O(N^{-1})$  in the equation for  $h_N(z)$ . Let us write the derivative in the direction  $e_{il}$ :

$$\mathbb{E}[(Ge_{il} Ge_{il} Ge_{il} G)_{jl}]$$

which is the sum of eight terms of the form

$$(4.17) \quad \mathbb{E}[G_{ji} G_{i_2 i_3} G_{i_4 i_5} G_{i_6 l}],$$

where if  $i_{2q+1} = i$  (resp.,  $l$ ), then  $i_{2q+2} = l$  (resp.,  $i$ ),  $q = 0, 1, 2$ .

LEMMA 4.3. *Let  $1 \leq k \leq r$  fixed; then*

$$(4.18) \quad F(N) := \left| \sum_{i,j=1}^N U_{ik}^* U_{kj} \frac{1}{N} \sum_{l=1}^N \mathbb{E}[A_{i,j,l}] \right| \leq C |\Im m(z)|^{-4}$$

for a numerical constant  $C$ .

PROOF.  $F(N)$  is the sum of eight terms corresponding to (4.17). Let us write, for example, the term corresponding to  $i_1 = i$ ,  $i_3 = i$ ,  $i_5 = i$ :

$$\begin{aligned} \frac{1}{N} \sum_{i,j,l} U_{ik}^* U_{kj} \mathbb{E}[G_{ji} G_{li} G_{li} G_{ll}] &= \mathbb{E} \left[ \frac{1}{N} \sum_{i,l} U_{ik}^* (UG)_{ki} G_{li} G_{li} G_{ll} \right] \\ &= \mathbb{E} \left[ \frac{1}{N} \sum_i U_{ik}^* (UG)_{ki} (G^T G^D G^T)_{ii} \right], \end{aligned}$$

where the superscript  $T$  denotes the transpose of the matrix and  $G^D$  is the diagonal matrix with entries  $G_{ii}$ . From the bounds  $\|G_N(z)\| \leq |\Im m(z)|^{-1}$  and  $\|U\| = 1$ , we get the bound given in the lemma.

We give the majoration for the term corresponding to  $i_1 = l$ ,  $i_3 = l$ ,  $i_5 = l$ :

$$\frac{1}{N} \sum_{i,j,l} U_{ik}^* U_{kj} \mathbb{E}[G_{jl} G_{il}^3] = \mathbb{E} \left[ \frac{1}{N} \sum_{i,l} U_{ik}^* (UG)_{kl} G_{il}^3 \right].$$

Its absolute value is bounded by  $\mathbb{E}[\frac{1}{N} \sum_{i,l} |G_{il}|^3] |\Im m(z)|^{-1}$  and thanks to Lemma 3.2 by  $|\Im m(z)|^{-4}$ . The other terms are treated in the same way.  $\square$

As in the Gaussian case, we now consider the sum  $\sum_{i,j} U_{ik}^* U_{kj} J_{ji}$ . From Lemma 4.3 and the bound (using the Cauchy–Schwarz inequality)

$$\sum_{i,j=1}^N |U_{ik}^* U_{kj}| \leq N,$$

we still get (4.14) and thus (4.15). More precisely, we proved:

PROPOSITION 4.3. *For any  $z \in \mathbb{C}$  such that  $\Im m(z) > 0$ ,*

$$\left| \mathbb{E}[\mathrm{Tr}(A_N G_N(z))] - \sum_{k=1}^r \frac{\gamma_k}{z - \sigma^2 g_\sigma(z) - \gamma_k} \right| \leq \frac{P_{11}(|\Im m(z)|^{-1})}{N} (K + |z|).$$

4.1.4. *Convergence of  $\mathbb{E}[(\frac{1}{N} \sum_{k=1}^N G_{kk}^2)^2]$ .* We now study the last term in the master inequality of Theorem 4.1. For the non-Deformed Wigner matrices, it is shown in [21] that

$$R_N(z) := \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N G_{kk}^2 \right)^2 \right] \xrightarrow[N \rightarrow \infty]{} g_\sigma^4(z).$$

Moreover, Proposition 3.2 in [13], in the more general setting of several independent Wigner matrices, gives an estimate of  $|R_N(z) - g_\sigma^4(z)|$ . The above convergence holds true in the Deformed case. We just give some hints of the proof of the estimate of  $|R_N(z) - g_\sigma^4(z)|$  since the computations are almost the same as in the non-Deformed case. Let us set

$$d_N(z) = \frac{1}{N} \sum_{k=1}^N G_{kk}^2.$$

We start from the resolvent identity

$$\begin{aligned} zG_{kk} &= 1 + \sum_{l=1}^N (M_N)_{kl} G_{lk} \\ &= 1 + \sum_{l=1}^N (A_N)_{kl} G_{lk} + \sum_{l=1}^N (X_N)_{kl} G_{lk} \end{aligned}$$

and

$$zd_N(z) = \frac{1}{N} \sum_{k=1}^N G_{kk} + \frac{1}{N} \sum_{k=1}^N (A_N G)_{kk} G_{kk} + \frac{1}{N} \sum_{k,l=1}^N (X_N)_{kl} G_{lk} G_{kk}.$$

For the last term, we apply an integration by parts formula (Lemma 4.2) to obtain (see [13, 21])

$$\mathbb{E} \left[ \frac{1}{N} \sum_{k,l=1}^N (X_N)_{kl} G_{lk} G_{kk} \right] = \sigma^2 \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N G_{kk} \right) d_N(z) \right] + O\left(\frac{1}{N}\right).$$

It remains to see that the additional term due to  $A_N$  is of order  $O(N^{-1})$ :

$$\frac{1}{N} \sum_{k=1}^N (A_N G)_{kk} G_{kk} = \frac{1}{N} \sum_{p=1}^r \gamma_p (U G G^D U^*)_{pp}$$

and

$$\left| \frac{1}{N} \sum_{k=1}^N (A_N G)_{kk} G_{kk} \right| \leq \left( \sum_{p=1}^r |\gamma_p| \right) \frac{|\Im m(z)|^{-2}}{N}.$$

We thus obtain (again with the help of a variance estimate)

$$z\mathbb{E}[d_N(z)] = g_N(z) + \sigma^2 g_N(z)\mathbb{E}[d_N(z)] + O\left(\frac{1}{N}\right).$$

Then using (4.11) and since  $d_N(z)$  is bounded we deduce that

$$z\mathbb{E}[d_N(z)] = g_\sigma(z) + \sigma^2 g_\sigma(z)\mathbb{E}[d_N(z)] + O\left(\frac{1}{N}\right).$$

Thus [using (3.7)]

$$\mathbb{E}[d_N(z)] = \frac{g_\sigma(z)}{z - \sigma^2 g_\sigma(z)} + O\left(\frac{1}{N}\right) \xrightarrow{N \rightarrow \infty} \frac{g_\sigma(z)}{z - \sigma^2 g_\sigma(z)} = g_\sigma^2(z).$$

Now, using some variance estimate,

$$\mathbb{E}[d_N^2(z)] = (\mathbb{E}[d_N(z)])^2 + O\left(\frac{1}{N}\right) = g_\sigma^4(z) + O\left(\frac{1}{N}\right).$$

We can now give our final master inequality for  $g_N(z)$  following our previous estimates:

**THEOREM 4.2.** *For  $z \in \mathbb{C}$  such that  $\Im m(z) > 0$ ,  $g_N(z)$  satisfies*

$$\left| \sigma^2 g_N^2(z) - z g_N(z) + 1 + \frac{1}{N} E_\sigma(z) \right| \leq \frac{P_{14}(|\Im m(z)|^{-1})}{N^2} (|z| + K),$$

where  $E_\sigma(z) = \sum_{k=1}^r \frac{\gamma_k}{z - \sigma^2 g_\sigma(z) - \gamma_k} + \frac{\kappa_4}{2} g_\sigma^4(z)$ ,  $\kappa_4$  is the fourth cumulant of the distribution  $\mu$ .

Note that  $E_\sigma(z)$  can be written in terms of the distinct eigenvalues  $\theta_j$  of  $A_N$  as

$$(4.19) \quad E_\sigma(z) = \sum_{j=1, j \neq j_0}^J k_j \frac{\theta_j}{z - \sigma^2 g_\sigma(z) - \theta_j} + \frac{\kappa_4}{2} g_\sigma^4(z).$$

Let us set

$$(4.20) \quad L_\sigma(z) = g_\sigma(z)^{-1} \mathbb{E}((z - s)^{-2}) E_\sigma(z),$$

where  $s$  is a centered semicircular random variable with variance  $\sigma^2$ .

4.2. *Estimation of  $|g_\sigma(z) - g_N(z) + \frac{1}{N} L_\sigma(z)|$ .* The method is roughly the same as the one described in Section 3.6 in [13]. Nevertheless we choose to

develop it here for the reader's convenience. We have for any  $z$  in  $\mathcal{O}'_n$ , by using (4.6) and (4.10),

$$\begin{aligned}
 & \left| g_\sigma(z) - g_N(z) + \frac{1}{N}L_\sigma(z) \right| \\
 &= \left| g_\sigma(z) - g_\sigma(\Lambda_N(z)) + \frac{1}{N}L_\sigma(z) \right| \\
 &= \left| \mathbb{E} \left[ (z-s)^{-1}(\Lambda_N(z)-s)^{-1}(\Lambda_N(z)-z) + \frac{1}{N}g_\sigma(z)^{-1}(z-s)^{-2}E_\sigma(z) \right] \right| \\
 &\leq \left| \mathbb{E} \left[ (z-s)^{-1}(\Lambda_N(z)-s)^{-1} \left( \Lambda_N(z)-z + \frac{1}{N}g_\sigma(z)^{-1}E_\sigma(z) \right) \right] \right| \\
 &\quad + \mathbb{E} \left[ |(z-s)^{-1}\{(z-s)^{-1} - (\Lambda_N(z)-s)^{-1}\}| \right] \frac{1}{N} |g_\sigma(z)^{-1}E_\sigma(z)| \\
 &\leq 2|\Im m(z)|^{-2} \left| \Lambda_N(z) - z + \frac{1}{N}E_\sigma(z)g_\sigma(z)^{-1} \right| \\
 &\quad + \frac{P_8(|\Im m(z)|^{-1})}{N} |\Lambda_N(z) - z|(|z| + K),
 \end{aligned}$$

where we made use of the estimates (3.5), (4.9),  $\forall z \in \mathbb{C} \setminus \mathbb{R}$ ,

$$\begin{aligned}
 (4.21) \quad \forall x \in \mathbb{R} \quad & \left| \frac{1}{z-x} \right| \leq |\Im m(z)|^{-1}, \\
 & |E_\sigma(z)| \leq P_4(|\Im m(z)|^{-1}) \quad [\text{using (3.7)}].
 \end{aligned}$$

Let us write

$$\begin{aligned}
 & \left| \Lambda_n(z) - z + \frac{1}{N}g_\sigma(z)^{-1} \right| \\
 &= \frac{1}{g_N(z)} \left( \sigma^2 g_N^2(z) - z g_N(z) + 1 + \frac{E_\sigma(z)}{N} \right) \\
 &\quad + \frac{E_\sigma(z)/N}{g_N(z)g_\sigma(z)} (g_N(z) - g_\sigma(z)).
 \end{aligned}$$

We get from Theorem 4.2, (4.7), (4.11), (4.21), (3.5),

$$\left| \Lambda_N(z) - z + \frac{1}{N}E_\sigma(z)g_\sigma(z)^{-1} \right| \leq (|z| + K)^3 \frac{P_{15}(|\Im m(z)|^{-1})}{N^2}.$$

Finally, using also (4.8), we get for any  $z$  in  $\mathcal{O}'_n$ ,

$$\left| g_\sigma(z) - g_N(z) + \frac{1}{N}L_\sigma(z) \right| \leq (|z| + K)^3 \frac{P_{17}(|\Im m(z)|^{-1})}{N^2}.$$

Now, for  $z \notin \mathcal{O}'_n$ , such that  $\Im m(z) > 0$ ,

$$\begin{aligned} 1 &\leq 4 \frac{P_6(|(\Im m(z))^{-1}|)}{N} (|z| + \sigma^2 |\Im m(z)|^{-1}) |\Im m(z)|^{-1} \\ &\leq (|z| + K) \frac{P_8(|\Im m(z)|^{-1})}{N}. \end{aligned}$$

We get

$$\begin{aligned} \left| g_\sigma(z) - g_N(z) + \frac{1}{N} L_\sigma(z) \right| &\leq |g_\sigma(z) - g_N(z)| + \frac{1}{N} |L_\sigma(z)| \\ &\leq (|z| + K) \frac{P_8(|\Im m(z)|^{-1})}{N} \\ &\quad \times \left[ (|z| + K) \frac{P_9(|\Im m(z)|^{-1})}{N} + \frac{1}{N} P_7(|\Im m(z)|^{-1}) (|z| + K) \right] \\ &\leq (|z| + K)^2 \frac{P_{17}(|\Im m(z)|^{-1})}{N^2}. \end{aligned}$$

Thus, for any  $z$  such that  $\Im m(z) > 0$ ,

$$(4.22) \quad \left| g_\sigma(z) - g_N(z) + \frac{1}{N} L_\sigma(z) \right| \leq (|z| + K)^3 \frac{P_{17}(|\Im m(z)|^{-1})}{N^2}.$$

Let us denote for a while  $g_N = g_N^{A_N}$  and  $L_\sigma = L_\sigma^{A_N}$ . Note that we get exactly the same estimation (4.22) dealing with  $-A_N$  instead of  $A_N$ . Hence since  $g_\sigma(z) = -g_\sigma(-z)$ ,  $g_N^{-A_N}(z) = -g_N^{A_N}(-z)$  (using the symmetry assumption on  $\mu$ ) and  $L_\sigma^{-A_N}(z) = L_\sigma^{A_N}(-z)$ , it readily follows that (4.22) is also valid for any  $z$  such that  $\Im m(z) < 0$ . In conclusion:

PROPOSITION 4.4. *For any  $z \in \mathbb{C} \setminus \mathbb{R}$ ,*

$$(4.23) \quad \left| g_\sigma(z) - g_N(z) + \frac{1}{N} L_\sigma(z) \right| \leq (|z| + K)^3 \frac{P_{17}(|\Im m(z)|^{-1})}{N^2}.$$

4.3. *The spectrum of  $M_N$ .* The following step now consists of deducing Proposition 4.6 from Proposition 4.4 (from which we will easily deduce the appropriate inclusion of the spectrum of  $M_N$ ). Since this transition is based on the inverse Stieltjes transform, we start with establishing the fundamental Proposition 4.5 below concerning the nature of  $L_\sigma$ . To this aim, it will be relevant to rewrite  $L_\sigma$  as

$$(4.24) \quad L_\sigma(z) = g_\sigma(z)^{-1} \times g'_\sigma(z) \times \left( \sum_{j=1}^J k_j \frac{\theta_j}{(1/(g_\sigma(z))) - \theta_j} + \frac{\kappa_4}{2} g_\sigma^A(z) \right).$$

We recall that  $J_{+\sigma}$  (resp.,  $J_{-\sigma}$ ) denotes the number of  $j$ 's such that  $\theta_j > \sigma$  (resp.,  $\theta_j < -\sigma$ ). As in the [Introduction](#), we define

$$\rho_{\theta_j} = \theta_j + \frac{\sigma^2}{\theta_j}$$

which is  $> 2\sigma$  (resp.,  $< -2\sigma$ ) when  $\theta_j > \sigma$  (resp.,  $< -\sigma$ ).

PROPOSITION 4.5.  *$L_\sigma$  is the Stieltjes transform of a distribution  $\Lambda_\sigma$  with compact support*

$$K_\sigma(\theta_1, \dots, \theta_J) := \{\rho_{\theta_j}; \dots; \rho_{\theta_{J-J_{-\sigma}+1}}\} \cup [-2\sigma, 2\sigma] \cup \{\rho_{\theta_{J+\sigma}}; \dots; \rho_{\theta_1}\}.$$

The proof relies on the following characterization already used in [\[29\]](#).

THEOREM 4.3 [\[32\]](#).

- Let  $\Lambda$  be a distribution on  $\mathbb{R}$  with compact support. Define the Stieltjes transform of  $\Lambda$ ,  $l: \mathbb{C} \setminus \mathbb{R} \rightarrow \mathbb{C}$  by

$$l(z) = \Lambda\left(\frac{1}{z-x}\right).$$

Then  $l$  is analytic in  $\mathbb{C} \setminus \mathbb{R}$  and has an analytic continuation to  $\mathbb{C} \setminus \text{supp}(\Lambda)$ . Moreover:

- (c<sub>1</sub>)  $l(z) \rightarrow 0$  as  $|z| \rightarrow \infty$ ,
- (c<sub>2</sub>) there exist a constant  $C > 0$ , an  $n \in \mathbb{N}$  and a compact set  $K \subset \mathbb{R}$  containing  $\text{supp}(\Lambda)$  such that for any  $z \in \mathbb{C} \setminus \mathbb{R}$ ,

$$|l(z)| \leq C \max\{\text{dist}(z, K)^{-n}, 1\},$$

- (c<sub>3</sub>) for any  $\phi \in C^\infty(\mathbb{R}, \mathbb{R})$  with compact support

$$\Lambda(\phi) = -\frac{1}{\pi} \lim_{y \rightarrow 0^+} \Im \int_{\mathbb{R}} \phi(x) l(x+iy) dx.$$

- Conversely, if  $K$  is a compact subset of  $\mathbb{R}$  and if  $l: \mathbb{C} \setminus K \rightarrow \mathbb{C}$  is an analytic function satisfying (c<sub>1</sub>) and (c<sub>2</sub>) above, then  $l$  is the Stieltjes transform of a compactly supported distribution  $\Lambda$  on  $\mathbb{R}$ . Moreover,  $\text{supp}(\Lambda)$  is exactly the set of singular points of  $l$  in  $K$ .

The following properties of the Stieltjes transform  $g_\sigma$  will be useful for showing that  $L_\sigma$  fulfills the previous conditions.

LEMMA 4.4.  *$g_\sigma$  is analytic and invertible on  $\mathbb{C} \setminus [-2\sigma, 2\sigma]$  and its inverse  $z_\sigma$  satisfied*

$$z_\sigma(g) = \frac{1}{g} + \sigma^2 g \quad \forall g \in g_\sigma(\mathbb{C} \setminus [-2\sigma, 2\sigma]).$$



(a) *The complement of the support of  $\mu_\sigma$  is characterized as follows:*

$$x \in \mathbb{R} \setminus [-2\sigma, 2\sigma] \iff \exists g \in \mathbb{R}^* \text{ such that } \left| \frac{1}{g} \right| > \sigma \text{ and } x = z_\sigma(g).$$

(b) *Given  $x \in \mathbb{R} \setminus [-2\sigma, 2\sigma]$  and  $\theta \in \mathbb{R}$  such that  $|\theta| > \sigma$ , one has*

$$\frac{1}{g_\sigma(x)} = \theta \iff x = \theta + \frac{\sigma^2}{\theta} := \rho_\theta.$$

This lemma can be easily proved using, for example, the explicit expression of  $g_\sigma$  [derived from (3.3)], namely for all  $x \in \mathbb{R} \setminus [-2\sigma, 2\sigma]$ ,

$$g_\sigma(x) = \frac{x}{2\sigma^2} (1 - \sqrt{1 - 4\sigma^2/x^2}).$$

**PROOF OF PROPOSITION 4.5.** Using (4.24), one readily sees that the set of singular points of  $L_\sigma$  is  $[-2\sigma, 2\sigma] \cup \{x \in \mathbb{R} \setminus [-2\sigma, 2\sigma], \frac{1}{g_\sigma(x)} \in \text{Spect}(A_N)\}$ . Hence [using point (b) of Lemma 4.4] the set of singular points of  $L_\sigma$  is exactly  $K_\sigma(\theta_1, \dots, \theta_J)$ .

Now, we are going to show that  $L_\sigma$  satisfies (c<sub>1</sub>) and (c<sub>2</sub>) of Theorem 4.3. We have obviously that

$$|z - \sigma^2 g_\sigma(z) - \theta_j| \geq ||z - \theta_j| - |\sigma^2 g_\sigma(z)||.$$

Now, let  $\alpha > 0$  such that  $\alpha > 2\sigma$  and for any  $j = 1, \dots, J$ ,  $\alpha - |\theta_j| > \frac{\sigma^2}{\alpha - 2\sigma}$ . For any  $z \in \mathbb{C}$  such that  $|z| > \alpha$ ,

$$|z - \theta_j| \geq |z| - |\theta_j| > \frac{\sigma^2}{\alpha - 2\sigma}$$

and according to (3.8)

$$|\sigma^2 g_\sigma(z)| \leq \frac{\sigma^2}{|z| - 2\sigma} \leq \frac{\sigma^2}{\alpha - 2\sigma}.$$

Thus we get that for  $z \in \mathbb{C}$  such that  $|z| > \alpha$ ,

$$|z - \sigma^2 g_\sigma(z) - \theta_j| \geq |z| - |\theta_j| - \frac{\sigma^2}{\alpha - 2\sigma}.$$

Using also (3.8)–(3.10), we get readily that for  $|z| > \alpha$ ,

$$\begin{aligned} |L_\sigma(z)| &\leq \left( |z| + \frac{\sigma^2}{|z| - 2\sigma} \right) \frac{1}{(|z| - 2\sigma)^2} \\ &\times \left( \sum_{j=1}^J \frac{k_j |\theta_j|}{|z| - |\theta_j| - (\sigma^2/(\alpha - 2\sigma))} + \frac{|\kappa_4|}{2(|z| - 2\sigma)^4} \right). \end{aligned}$$

Then, it is clear that  $|L_\sigma(z)| \rightarrow 0$  when  $|z| \rightarrow +\infty$  and (c<sub>1</sub>) is satisfied.

Now we follow the approach of [29] (Lemma 5.5) to prove (c<sub>2</sub>). Denote by  $\mathcal{E}$  the convex envelope of  $K_\sigma(\theta_1, \dots, \theta_J)$  and define the interval

$$\begin{aligned} K &:= \{x \in \mathbb{R}; \text{dist}(x, \mathcal{E}) \leq 1\} \\ &= [\min\{x \in K_\sigma(\theta_1, \dots, \theta_J)\} - 1; \max\{x \in K_\sigma(\theta_1, \dots, \theta_J)\} + 1] \end{aligned}$$

and

$$D = \{z \in \mathbb{C}; 0 < \text{dist}(z, K) \leq 1\}.$$

- Let  $z \in D \cap \mathbb{C} \setminus \mathbb{R}$  with  $\Re(z) \in K$ . We have  $\text{dist}(z, K) = |\Im(z)| \leq 1$ . Using the upper bounds (3.4), (3.5), (3.6) and (3.7), we easily deduce that there exists some constant  $C_0$  such that for any  $z \in D \cap \mathbb{C} \setminus \mathbb{R}$  with  $\Re(z) \in K$ ,

$$|L_\sigma(z)| \leq C_0 |\Im(z)|^{-7} = C_0 \text{dist}(z, K)^{-7} = C_0 \max(\text{dist}(z, K)^{-7}; 1).$$

- Let  $z \in D \cap \mathbb{C} \setminus \mathbb{R}$  with  $\Re(z) \notin K$ . Then  $\text{dist}(z, K_\sigma(\theta_1, \dots, \theta_J)) \geq 1$ . Since  $L_\sigma$  is bounded on compact subsets of  $\mathbb{C} \setminus K_\sigma(\theta_1, \dots, \theta_J)$ , we easily deduce that there exists some constant  $C_1$  such that for any  $z \in D$  with  $\Re(z) \notin K$ ,

$$|L_\sigma(z)| \leq C_1 \leq C_1 \text{dist}(z, K)^{-7} = C_1 \max(\text{dist}(z, K)^{-7}; 1).$$

- Since  $|L_\sigma(z)| \rightarrow 0$  when  $|z| \rightarrow +\infty$ ,  $L_\sigma$  is bounded on  $\mathbb{C} \setminus D$ . Thus, there exists some constant  $C_2$  such that for any  $z \in \mathbb{C} \setminus D$ ,

$$|L_\sigma(z)| \leq C_2 = C_2 \max(\text{dist}(z, K)^{-7}; 1).$$

Hence (c<sub>2</sub>) is satisfied with  $C = \max(C_0, C_1, C_2)$  and  $n = 7$  and Proposition 4.5 follows from Theorem 4.3.  $\square$

We are now in position to deduce the following proposition from the estimate (4.23).

PROPOSITION 4.6. *For any smooth function  $\varphi$  with compact support,*

$$(4.25) \quad \mathbb{E}[\text{tr}_N(\varphi(M_N))] = \int \varphi d\mu_{sc} + \frac{1}{N} \Lambda_\sigma(\varphi) + O\left(\frac{1}{N^2}\right).$$

*Consequently, for  $\varphi$  smooth, constant outside a compact set and such that  $\text{supp}(\varphi) \cap K_\sigma(\theta_1, \dots, \theta_J) = \emptyset$ ,*

$$(4.26) \quad \text{tr}_N(\varphi(M_N)) = O(N^{-4/3}) \quad a.s.$$

PROOF. Using the inverse Stieltjes transform, we get, respectively, that, for any  $\varphi$  in  $\mathcal{C}^\infty(\mathbb{R}, \mathbb{R})$  with compact support,

$$\mathbb{E}[\mathrm{tr}_N(\varphi(M_N))] - \int \varphi d\mu_{sc} - \frac{\Lambda_\sigma(\varphi)}{N} = -\frac{1}{\pi} \lim_{y \rightarrow 0^+} \Im m \int_{\mathbb{R}} \varphi(x) r_N(x + iy) dx,$$

where  $r_N = g_\sigma(z) - g_N(z) + \frac{1}{N} L_\sigma(z)$  satisfies, according to Proposition 4.4, for any  $z \in \mathbb{C} \setminus \mathbb{R}$ ,

$$|r_N(z)| \leq \frac{1}{N^2} (|z| + K)^\alpha P_k(|\Im m(z)^{-1}|),$$

where  $\alpha = 3$  and  $k = 17$ .

We refer the reader to the Appendix of [13] where it is proved using the ideas of [18] that

$$\limsup_{y \rightarrow 0^+} \left| \int_{\mathbb{R}} \varphi(x) h(x + iy) dx \right| \leq C < +\infty,$$

when  $h$  is an analytic function on  $\mathbb{C} \setminus \mathbb{R}$  which satisfies

$$|h(z)| \leq (|z| + K)^\alpha P_k(|\Im m(z)^{-1}|).$$

Dealing with  $h(z) = N^2 r_N(z)$ , we deduce that

$$\limsup_{y \rightarrow 0^+} \left| \int_{\mathbb{R}} \varphi(x) r_N(x + iy) dx \right| \leq \frac{C}{N^2}$$

and then (4.25).

Following the proof of Lemma 5.6 in [29], one can show that  $\Lambda_\sigma(1) = 0$ . Then, the rest of the proof of (4.26) sticks to the proof of Lemma 6.3 in [18] (using Lemma 3.1).  $\square$

Following [18] (Theorem 6.4), we set  $K = K_\sigma(\theta_1, \dots, \theta_J) + (-\frac{\varepsilon}{2}, \frac{\varepsilon}{2})$ ,  $F = \{t \in \mathbb{R}; \mathrm{dist}(t, K_\sigma(\theta_1, \dots, \theta_J)) \geq \varepsilon\}$  and take  $\varphi \in \mathcal{C}^\infty(\mathbb{R}, \mathbb{R})$  such that  $0 \leq \varphi \leq 1$ ,  $\varphi(t) = 0$  for  $t \in K$  and  $\varphi(t) = 1$  for  $t \in F$ . Then according to (4.26),  $\mathrm{tr}_N(\varphi(M_N)) = O(N^{-4/3})$  a.s. Since  $\varphi \geq 1_F$ , it follows that  $\mathrm{tr}_N(1_F(M_N)) = O(N^{-4/3})$  a.s. and thus the number of eigenvalues of  $M_N$  in  $F$  is almost surely an  $O(N^{-1/3})$  as  $N$  goes to infinity. Since for each  $N$  this number has to be an integer we deduce that the number of eigenvalues of  $M_N$  in  $F$  is zero almost surely as  $N$  goes to infinity. The fundamental inclusion (2.15) follows, namely, for any  $\varepsilon > 0$ , almost surely

$$\mathrm{Spect}(M_N) \subset K_\sigma(\theta_1, \dots, \theta_J) + (-\varepsilon, \varepsilon),$$

when  $N$  goes to infinity.

Such a method can be carried out in the case of Wigner real symmetric matrices; then the approximate master equation is the following [compare with (4.4)]:

$$\begin{aligned} & \sigma^2 g_N(z)^2 - z g_N(z) + 1 + \frac{1}{N} \frac{\kappa_4}{2} \mathbb{E} \left[ \left( \frac{1}{N} \sum_{k=1}^N G_{kk}(z)^2 \right)^2 \right] + \frac{\sigma^2}{N} \mathbb{E}(\text{tr}_N G_N(z)^2) \\ & + \mathbb{E}(\text{tr}_N [A_N G_N(z)]) = O\left(\frac{1}{N^2}\right). \end{aligned}$$

Note that the additional term  $\frac{\sigma^2}{N} \mathbb{E}(\text{tr}_N G_N(z)^2)$  already appears in the non-Deformed GOE case in [29]. One can establish in a similar way the analogue of (4.11) and then, following the proof of Corollary 3.3 in [29], deduce that

$$\mathbb{E}(\text{tr}_N G_N(z)^2) = \mathbb{E}((z-s)^{-2}) + O\left(\frac{1}{N}\right),$$

where  $s$  is a centered semicircular variable with variance  $\sigma^2$ . Hence by similar arguments as in the complex case, one gets the master equation

$$\sigma^2 g_N(z)^2 - z g_N(z) + 1 + \frac{1}{N} E_\sigma(z) = O\left(\frac{1}{N^2}\right),$$

where

$$E_\sigma(z) = \sum_{j=1, j \neq j_0}^J k_j \frac{\theta_j}{z - \sigma^2 g_\sigma(z) - \theta_j} + \frac{\kappa_4}{2} g_\sigma^4(z) + \mathbb{E}((z-s)^{-2}).$$

It can be proved that  $L_\sigma(z) := g_\sigma(z)^{-1} \mathbb{E}((z-s)^{-2}) E_\sigma(z)$  is the Stieltjes transform of a distribution  $\Lambda_\sigma$  with compact support  $K_\sigma(\theta_1, \dots, \theta_J)$ , too. The last arguments hold likewise in the real symmetric case.

Hence we have established:

**THEOREM 4.4.** *Let  $(M_N)_N$  be any real or complex Deformed model satisfying (i) and (ii) in Section 2. Let  $J_{+\sigma}$  (resp.,  $J_{-\sigma}$ ) be the number of  $j$ 's such that  $\theta_j > \sigma$  (resp.,  $\theta_j < -\sigma$ ). Then for any  $\varepsilon > 0$ , almost surely, there is no eigenvalue of  $M_N$  in*

$$(4.27) \quad \begin{aligned} & (-\infty, \rho_{\theta_J} - \varepsilon) \cup (\rho_{\theta_J} + \varepsilon, \rho_{\theta_{J-1}} - \varepsilon) \cup \dots \cup (\rho_{\theta_{J-J_{-\sigma}+1}} + \varepsilon, -2\sigma - \varepsilon) \\ & \cup (2\sigma + \varepsilon, \rho_{\theta_{J_{+\sigma}}} - \varepsilon) \cup \dots \cup (\rho_{\theta_2} + \varepsilon, \rho_{\theta_1} - \varepsilon) \cup (\rho_{\theta_1} + \varepsilon, +\infty), \end{aligned}$$

when  $N$  is large enough.

**REMARK 4.1.** As soon as  $\varepsilon > 0$  is small enough, the union (4.27) is made of nonempty disjoint intervals.

4.4. *The almost sure convergence result.* As announced in the [Introduction](#), Theorem 2.1 is the analogue of the main statement of [9] established for general spiked population models (1.1). The previous Theorem 4.4 is the main step of the proof since now, we can adapt the arguments needed for the conclusion of [9] viewing the Deformed Wigner model (1.2) as the additive analogue of the spiked population model (1.1).

Let us consider one of the positive eigenvalues  $\theta_j$  of the  $A_N$ 's. We recall that this implies that  $\lambda_{k_1+\dots+k_{j-1}+i}(A_N) = \theta_j$  for all  $1 \leq i \leq k_j$ . We want to show that if  $\theta_j > \sigma$  (i.e., with our notation, if  $j \in \{1, \dots, J_{+\sigma}\}$ ), the corresponding eigenvalues of  $M_N$  almost surely jump above the right endpoint  $2\sigma$  of the semicircle support as

$$\forall 1 \leq i \leq k_j \quad \lambda_{k_1+\dots+k_{j-1}+i}(M_N) \longrightarrow \rho_{\theta_j} \quad \text{a.s.},$$

whereas the rest of the asymptotic spectrum of  $M_N$  lies below  $2\sigma$  with

$$\lambda_{k_1+\dots+k_{J_{+\sigma}+1}}(M_N) \longrightarrow 2\sigma \quad \text{a.s.}$$

Analogous results hold for the negative eigenvalues  $\theta_j$  [see points (c) and (d) of Theorem 2.1]. To describe the phenomenon, one can say that, when  $N$  is large enough, the (first extremal) eigenvalues of  $M_N$  can be viewed as a “smoothed” deformation of the (first extremal) eigenvalues of  $A_N$ . According to the analysis made in the previous section [Lemma 4.4(b)], we already know that the limits  $\rho_{\theta_j}$  are related to the  $\theta_j$ 's through the Stieltjes transform  $g_\sigma$ . More precisely, one has

$$\text{for all } j \text{ such that } |\theta_j| > \sigma, \quad \frac{1}{g_\sigma(\rho_{\theta_j})} = \theta_j.$$

Our main purpose now is to establish the asymptotic link between the spectra of the matrices  $M_N = X_N + A_N$  and  $A_N$ .

Intuitively, this link seems rather natural when  $\sigma$  is close to zero. Indeed, when  $N$  goes to infinity, since the spectrum of  $X_N$  is concentrated in  $[-2\sigma, 2\sigma]$  [recall (2.9)], the spectrum of  $M_N$  should be close to the one of  $A_N$  as soon as  $\sigma$  will be close to zero (in other words, the spectrum of  $M_N$  is, viewed as a deformation of the one of  $A_N$ , continuous in  $\sigma$  in a neighborhood of zero). Thus given an interval  $[a, b] \subset {}^cK_\sigma(\theta_1, \dots, \theta_J)$ , the result of Theorem 4.4 saying that  $[a, b]$  does not contain eigenvalues of  $M_N$  should be improved: it should correspond to  $[a, b]$  some interval  $I$  close to  $[a, b]$ , lying outside the spectrum of  $A_N$  and such that the number of eigenvalues of  $M_N$  in one side of  $[a, b]$  is equal to the one of  $A_N$  in the corresponding side of  $I$ . Following [4], we will say that there is exact separation of eigenvalues of the matrices  $A_N$  and  $M_N$ .

In the following section, we justify that the exact separation phenomenon occurs regardless of the size of  $\sigma$ . The proof of Theorem 2.1 will then follow from some suitable choices of  $[a, b]$  (see Section 4.4.2).

4.4.1. *Exact separation of eigenvalues.* According to the previous discussion, we need to refine the analysis made on  $g_\sigma$  in order to identify and understand the link between intervals in  ${}^cK_\sigma(\theta_1, \dots, \theta_J)$  and the complement of the spectrum of the  $A_N$ 's. We also need to understand the dependence on  $\sigma$ . This is the aim of the following important Lemma 4.5.

As before, we denote (recall Lemma 4.4) by  $z_\sigma$  the inverse function of  $g_\sigma$  which is given by

$$z_\sigma(g) = \frac{1}{g} + \sigma^2 g.$$

Using Lemma 4.4, one readily sees that the set  ${}^cK_\sigma(\theta_1, \dots, \theta_J)$  can be characterized as follows:

$$(4.28) \quad x \in {}^cK_\sigma(\theta_1, \dots, \theta_J) \iff \exists g \in \mathcal{G}_\sigma \text{ such that } x = z_\sigma(g),$$

where

$$(4.29) \quad \mathcal{G}_\sigma := \left\{ g \in \mathbb{R}^*, \left| \frac{1}{g} \right| > \sigma \text{ and } \frac{1}{g} \notin \text{Spect}(A_N) \right\}.$$

Obviously, one has  $g = g_\sigma(x)$  if  $x \in {}^cK_\sigma(\theta_1, \dots, \theta_J)$ .

LEMMA 4.5. *Let  $[a, b]$  be a compact set contained in  ${}^cK_\sigma(\theta_1, \dots, \theta_J)$ . Then:*

- (i)  $[\frac{1}{g_\sigma(a)}, \frac{1}{g_\sigma(b)}] \subset (\text{Spect}(A_N))^c$ .
- (ii) *For all  $0 < \hat{\sigma} < \sigma$ , the interval  $[z_{\hat{\sigma}}(g_\sigma(a)), z_{\hat{\sigma}}(g_\sigma(b))]$  is contained in  ${}^cK_{\hat{\sigma}}(\theta_1, \dots, \theta_J)$  and  $z_{\hat{\sigma}}(g_\sigma(b)) - z_{\hat{\sigma}}(g_\sigma(a)) \geq b - a$ .*

PROOF. The function  $1/g_\sigma$  being increasing, (i) readily follows from (4.28).

Noticing that  $\mathcal{G}_\sigma \subset \mathcal{G}_{\hat{\sigma}}$  for all  $\hat{\sigma} < \sigma$  implies (recall also that  $g_\sigma$  decreases on  $[a, b]$ ) that  $[g_\sigma(b), g_\sigma(a)] \subset \mathcal{G}_{\hat{\sigma}}$ . Relation (4.28) combined with the fact that the function  $z_{\hat{\sigma}}$  is decreasing on  $[g_\sigma(b), g_\sigma(a)]$  leads to

$$[z_{\hat{\sigma}}(g_\sigma(a)), z_{\hat{\sigma}}(g_\sigma(b))] \subset {}^cK_{\hat{\sigma}}(\theta_1, \dots, \theta_J)$$

and the first part of (ii) is stated. Now, we have

$$\begin{aligned} l_\sigma(\hat{\sigma}) &:= z_{\hat{\sigma}}(g_\sigma(b)) - z_{\hat{\sigma}}(g_\sigma(a)) \\ &= \frac{1}{g_\sigma(b)} - \frac{1}{g_\sigma(a)} + \hat{\sigma}^2(g_\sigma(b) - g_\sigma(a)). \end{aligned}$$

Since  $g_\sigma$  decreases on  $[a, b]$ , we have  $g_\sigma(b) - g_\sigma(a) \leq 0$  and thus  $l_\sigma$  is decreasing on  $\mathbb{R}^+$ . Then the last part of (ii) follows since  $l_\sigma(\sigma) = b - a$ .  $\square$

The exact separation result can now be stated. Let  $[a, b]$  be an interval contained in  ${}^cK_\sigma(\theta_1, \dots, \theta_J)$ . By Theorem 4.4,  $[a, b]$  is outside the spectrum

of  $M_N$ . Moreover, from Lemma 4.5(i), there corresponds an interval  $I = [a', b']$  outside the spectrum of  $A_N$ , that is, there is  $i_N \in \{0, \dots, N\}$  such that

$$(4.30) \quad \lambda_{i_N+1}(A_N) < \frac{1}{g_\sigma(a)} := a' \quad \text{and} \quad \lambda_{i_N}(A_N) > \frac{1}{g_\sigma(b)} := b'.$$

$a$  and  $a'$  (resp.,  $b$  and  $b'$ ) are linked as follows:

$$a = \rho_{a'} := a' + \frac{\sigma^2}{a'} \quad (\text{resp., } b = \rho_{b'}).$$

We claim that  $[a, b]$  splits the eigenvalues of  $M_N$  exactly as  $I$  splits the spectrum of  $A_N$ . In other words:

**THEOREM 4.5.** *With  $i_N$  satisfying (4.30), one has*

$$(4.31) \quad \mathbb{P}[\lambda_{i_N+1}(M_N) < a \text{ and } \lambda_{i_N}(M_N) > b, \text{ for all large } N] = 1.$$

This result is the analogue of the main statement of [4] (cf. Theorem 1.2 of [4]) established in the spiked population setting (and in fact for quite general sample covariance matrices). Its proof is quite technical and is inspired by the work [4]. It mainly relies on results on eigenvalues of the rescaled Wigner matrix  $X_N$  combined with the following classical result (due to Weyl).

**LEMMA 4.6** (cf. Theorem 4.3.7 of [19]). *Let  $B$  and  $C$  be two  $N \times N$  Hermitian matrices. For any pair of integers  $j, k$  such that  $1 \leq j, k \leq N$  and  $j + k \leq N + 1$ , we have*

$$\lambda_{j+k-1}(B + C) \leq \lambda_j(B) + \lambda_k(C).$$

*For any pair of integers  $j, k$  such that  $1 \leq j, k \leq N$  and  $j + k \geq N + 1$ , we have*

$$\lambda_j(B) + \lambda_k(C) \leq \lambda_{j+k-N}(B + C).$$

**REMARK 4.2.** Note that this lemma is the additive analogue of Lemma 1.1 of [4] needed for the investigations of the spiked population model.

In particular, Lemma 4.6 gives that  $\lambda_{i_N+1}(M_N) \leq \lambda_{i_N+1}(A_N) + \lambda_1(X_N)$  and  $\lambda_{i_N}(M_N) \geq \lambda_{i_N}(A_N) + \lambda_N(X_N)$ . Besides, as both  $\lambda_1(X_N)$  and  $-\lambda_N(X_N)$  tend toward  $2\sigma$  as  $N \rightarrow \infty$  [this is (2.9)], the statement of Theorem 4.5 can be quite easily derived when  $\sigma$  is close enough to zero. To handle the general case, the key idea is that one can reduce to the previous situation by

introducing some parameters. More precisely, given  $L > 0$  and  $k \geq 0$ , we will introduce the Wigner matrix

$$W_N^{k,L} = W_N / \sqrt{1 + k/L}$$

and let

$$M_N^{k,L} = A_N + W_N^{k,L} / \sqrt{N}$$

be the Deformed Wigner matrix of parameter

$$\sigma_{k,L} = \sigma / \sqrt{1 + k/L}.$$

The proof will be organized as follows. On the one hand, as  $\sigma_{k,L} \rightarrow 0$  when  $k \rightarrow \infty$  (for any fixed  $L > 0$ ), we will readily prove that exact separation occurs for the matrices  $A_N$  and  $M_N^{K,L}$  as soon as  $K$  is large enough. On the other hand, we will show that exact separation also occurs for the eigenvalues of  $M_N = M_N^{0,L}$  and  $M_N^{K,L}$  choosing  $L$  large enough. This latter point will be established by induction on  $k$ ; the underlying idea is that when the parameter  $L$  is large, the matrices  $M_N^{k,L}$  and  $M_N^{k+1,L}$  are close to each other and hence split their spectrum in a similar way.

**PROOF OF THEOREM 4.5.** With our choice of  $[a, b]$  and the very definition of the spectrum of the  $A_N$ 's, one can consider  $\epsilon' > 0$  small enough such that, for all large  $N$ ,

$$\lambda_{i_{N+1}}(A_N) < \frac{1}{g_\sigma(a)} - \epsilon' \quad \text{and} \quad \lambda_{i_N}(A_N) > \frac{1}{g_\sigma(b)} + \epsilon'.$$

Given  $L > 0$  and  $k \geq 0$  (their size will be determined later), we define

$$a_{k,L} = z_{\sigma_{k,L}}(g_\sigma(a)) \quad \text{and} \quad b_{k,L} = z_{\sigma_{k,L}}(g_\sigma(b)),$$

where we recall that  $z_{\sigma_{k,L}}(g) = 1/g + \sigma_{k,L}^2 g$ . Note that for all  $L > 0$ , one has  $a_{0,L} = a$  and  $b_{0,L} = b$ .

We first choose the size of  $L$  as follows. We take  $L_0$  large enough such that for all  $L \geq L_0$ ,

$$(4.32) \quad \max((\sigma^2/L)(|g_\sigma(a)| + |g_\sigma(b)|); 3\sigma/L) < (b - a)/4.$$

From the very definition of the  $a_{k,L}$ 's and  $b_{k,L}$ 's, one can easily see that  $b_{k,L} - a_{k,L} \geq b - a$  [using the last point of (ii) in Lemma 4.5] and that this choice of  $L_0$  ensures that, for all  $L \geq L_0$  and for all  $k \geq 0$ ,

$$(4.33) \quad |a_{k+1,L} - a_{k,L}| < (b - a)/4 \quad \text{and} \quad |b_{k+1,L} - b_{k,L}| < (b - a)/4.$$

Now, we fix  $L$  such that  $L \geq L_0$  and we write  $a_k = a_{k,L}$ ,  $b_k = b_{k,L}$  and  $\sigma_k = \sigma_{k,L}$ .



We first show that there exists  $K$  large enough such that, for all  $k \geq K$ , there is exact separation of the eigenvalues of the matrices  $A_N$  and  $M_N^{k,L}$ , that is,

$$(4.34) \quad \mathbb{P}[\lambda_{i_N+1}(M_N^{k,L}) < a_k \text{ and } \lambda_{i_N}(M_N^{k,L}) > b_k \text{ for all large } N] = 1.$$

Lemma 4.6 first gives that

$$\lambda_{i_N+1}(M_N^{k,L}) \leq a_k - \epsilon' - \sigma_k^2 g_\sigma(a) + \frac{1}{\sqrt{1+k/L}} \lambda_1(X_N) \quad \text{if } i_N < N$$

and

$$\lambda_{i_N}(M_N^{k,L}) \geq b_k + \epsilon' - \sigma_k^2 g_\sigma(b) + \frac{1}{\sqrt{1+k/L}} \lambda_N(X_N) \quad \text{if } i_N > 0.$$

Furthermore, according to (2.9), the two first extremal eigenvalues of  $X_N$  are such that almost surely and for all  $N$  large enough,

$$0 < \max(-\lambda_N(X_N), \lambda_1(X_N)) < 3\sigma.$$

Thus for all  $k$ , almost surely, at least for  $N$  large enough ( $N$  does not depend on  $k$ ),

$$0 < \frac{1}{\sqrt{1+k/L}} \times \max(-\lambda_N(X_N), \lambda_1(X_N)) < 3\sigma_k.$$

As  $\sigma_k \rightarrow 0$  when  $k \rightarrow +\infty$ , there is  $K$  large enough such that for all  $k \geq K$ ,

$$\max(|3\sigma_k - \sigma_k^2 g_\sigma(a)|, |3\sigma_k + \sigma_k^2 g_\sigma(b)|) < \epsilon'$$

and then, almost surely, for all  $N$  large enough

$$(4.35) \quad \lambda_{i_N+1}(M_N^{k,L}) < a_k \quad \text{if } i_N < N$$

and

$$(4.36) \quad \lambda_{i_N}(M_N^{k,L}) > b_k \quad \text{if } i_N > 0.$$

Since  $\lambda_{N+1}(M_N^{k,L}) = -\lambda_0(M_N^{k,L}) = -\infty$ , (4.35) [resp., (4.36)] is obviously satisfied if  $i_N = N$  (resp.,  $i_N = 0$ ). Thus, we have established that for any  $i_N \in \{0, \dots, N\}$  satisfying (4.30), (4.34) holds for all  $k \geq K$ . In particular,

$$(4.37) \quad \mathbb{P}[\lambda_{i_N+1}(M_N^{K,L}) < a_K \text{ and } \lambda_{i_N}(M_N^{K,L}) > b_K \text{ for all large } N] = 1.$$

Now, we shall show that with probability 1: for  $N$  large,  $[a_K, b_K]$  and  $[a, b]$  split the eigenvalues of, respectively,  $M_N^{K,L}$  and  $M_N$  having equal amount of eigenvalues to the left sides of the intervals. To this aim, we will proceed by induction on  $k$  and establish that, for all  $k \geq 0$ ,  $[a_k, b_k]$  and  $[a, b]$  split the

eigenvalues of  $M_N^{k,L}$  and  $M_N$  (recall that  $M_N = M_N^{0,L}$ ) in exactly the same way. To begin, let us consider for all  $k \geq 0$ , the set

$$E_k = \{\text{no eigenvalues of } M_N^{k,L} \text{ in } [a_k, b_k], \text{ for all large } N\}.$$

By Lemma 4.5(ii) and Theorem 4.4, we know that  $\mathbb{P}(E_k) = 1$  for all  $k$ . In particular, from the fact that  $\mathbb{P}(E_0) = 1$ , one has for all  $\omega \in E_0$  and for all large  $N$ ,

$$(4.38) \quad \exists j_N(\omega) \in \{0, \dots, N\} \text{ such that } \lambda_{j_N(\omega)+1}(M_N) < a \text{ and } \lambda_{j_N(\omega)}(M_N) > b.$$

Extending the random variable  $j_N$  by setting, for instance,  $j_N := -1$  on  ${}^c E_0$ , we want to show that for all  $k$ ,

$$(4.39) \quad \mathbb{P}[\lambda_{j_N+1}(M_N^{k,L}) < a_k \text{ and } \lambda_{j_N}(M_N^k) > b_k, \text{ for all large } N] = 1.$$

This can be done by induction calling, one more time, on Lemma 4.6. By (4.38), this is true for  $k = 0$ . Now, let us assume that (4.39) holds true. One has

$$M_N^{k+1,L} = M_N^{k,L} + \left( \frac{1}{\sqrt{1+(k+1)/L}} - \frac{1}{\sqrt{1+k/L}} \right) X_N$$

so, by Lemma 4.6,

$$\lambda_{j_N+1}(M_N^{k+1,L}) \leq \lambda_{j_N+1}(M_N^{k,L}) + (-\lambda_N(X_N))/L.$$

But, for  $N$  large enough,  $0 < -\lambda_N(X_N) \leq 3\sigma$  a.s., so by the condition (4.32) on  $L$ ,

$$\lambda_{j_N+1}(M_N^{k+1,L}) < a_k + (b-a)/4 := \hat{a}_k.$$

Similarly, one can show that

$$\text{a.s.} \quad \lambda_{j_N}(M_N^{k+1,L}) > b_k - (b-a)/4 := \hat{b}_k.$$

By (4.33), one readily observes that  $\hat{a}_k - a_{k+1} = a_k - a_{k+1} + (b-a)/4 > 0$  and similarly that  $\hat{b}_k - b_{k+1} < 0$ . This implies that

$$[\hat{a}_k, \hat{b}_k] \subset [a_{k+1}, b_{k+1}].$$

As  $\mathbb{P}(E_{k+1}) = 1$ , we deduce that with probability 1,

$$\lambda_{j_N+1}(M_N^{k+1,L}) < a_{k+1} \quad \text{and} \quad \lambda_{j_N}(M_N^{k+1,L}) > b_{k+1} \quad \text{for all } N \text{ large.}$$

As a consequence, (4.39) holds for all  $k \geq 0$  and in particular for  $k = K$ . Comparing this with (4.37), we deduce that  $j_N = i_N$  a.s. and

$$\mathbb{P}[\lambda_{i_N+1}(M_N) < a \text{ and } \lambda_{i_N}(M_N) > b \text{ for all large } N] = 1.$$

This ends the proof of Theorem 4.5.  $\square$

Now, we are in position to prove the main Theorem 2.1.

4.4.2. *Proof of Theorem 2.1.* Our reasoning is close to the last Section 4 of [9]. It is enough to establish parts (a) and (b) since the assertions (c) and (d) can then be deduced by taking  $-M_N$  instead of  $M_N$ .

The proof of (a) is mainly based on successive applications of Theorem 4.5. Fix an integer  $1 \leq j \leq J_{+\sigma}$ , and let us consider for  $\epsilon > 0$ , the interval  $[a, b] = [\rho_{\theta_j} + \epsilon, \rho_{\theta_{j-1}} - \epsilon]$  which is included in the union (4.27) (at least for  $\epsilon$  small enough). We define  $K_{j(-1)} = k_1 + \dots + k_{j(-1)}$ . We also take  $\theta_0 := +\infty$  and recall the conventions that  $\lambda_0(M_N) = \lambda_0(A_N) = +\infty$  and  $K_0 = 0$ . Since  $1/g_\sigma(\rho_{\theta_k}) = \theta_k$  for  $k = j-1$  and  $j$  and since the function  $1/g_\sigma$  is continuous and increasing on  $[a, b]$ , the compact interval  $[a, b]$  satisfies (4.30) with  $i_N = K_{j-1}$ . Hence by Theorem 4.5, one has

$$\mathbb{P}[\lambda_{K_{j-1}}(M_N) \geq \rho_{\theta_{j-1}} - \epsilon \text{ and } \lambda_{K_{j-1}+1}(M_N) \leq \rho_{\theta_j} + \epsilon, \text{ for } N \text{ large}] = 1.$$

Similar arguments imply that for all  $j \in \{1, \dots, J_{+\sigma} - 1\}$ ,

$$\mathbb{P}[\lambda_{K_j}(M_N) \geq \rho_{\theta_j} - \epsilon \text{ and } \lambda_{K_{j+1}}(M_N) \leq \rho_{\theta_{j+1}} + \epsilon, \text{ for } N \text{ large}] = 1.$$

As a result, we deduce that for all  $1 \leq j \leq J_{+\sigma} - 1$ ,

$$(4.40) \quad \mathbb{P}[\rho_{\theta_j} - \epsilon \leq \lambda_{K_j}(M_N) \leq \dots \leq \lambda_{K_{j-1}+1}(M_N) \leq \rho_{\theta_j} + \epsilon \text{ for } N \text{ large}] = 1.$$

So, letting  $\epsilon$  go to zero, we obtain (a) for each integer  $j$  of  $\{1, \dots, J_{+\sigma} - 1\}$ .

Let us now quickly consider the case where  $j = J_{+\sigma}$ . Note first that, from the preceding discussion, we still have (for  $\epsilon$  small enough)

$$\mathbb{P}[\lambda_{K_{J_{+\sigma}-1}+1}(M_N) \leq \rho_{\theta_{J_{+\sigma}}} + \epsilon, \text{ for } N \text{ large}] = 1.$$

Then, using the fact that  $1/g_\sigma$  increases continuously on  $]2\sigma, +\infty[$  with  $1/g_\sigma(]2\sigma, +\infty[) = ]\sigma, +\infty[$ , one can show that once  $\epsilon > 0$  is small enough, the compact set  $[a, b] = [2\sigma + \epsilon, \rho_{\theta_{J_{+\sigma}}} - \epsilon]$  satisfies the assumptions of Theorem 4.5 with  $i_N = K_{J_{+\sigma}}$ . This leads to

$$\mathbb{P}[\lambda_{K_{J_{+\sigma}}}(M_N) \geq \rho_{\theta_{J_{+\sigma}}} - \epsilon \text{ and } \lambda_{K_{J_{+\sigma}+1}}(M_N) \leq 2\sigma + \epsilon, \text{ for } N \text{ large}] = 1.$$

Letting  $\epsilon \rightarrow 0$ , we deduce that (4.40) holds for  $j = J_{+\sigma}$  and the assertion (a) is established. For point (b), the preceding analysis gives that  $\limsup_N \lambda_{K_{J_{+\sigma}+1}}(M_N) \leq 2\sigma$  a.s. and it remains to prove that

$$\liminf_N \lambda_{K_{J_{+\sigma}+1}}(M_N) \geq 2\sigma \quad \text{a.s.}$$

This inequality follows from the fact that the spectral measure of  $M_N$  converges a.s. toward the semicircle law  $\mu_{sc}$  which is compactly supported in  $[-2\sigma, 2\sigma]$ . This completes the proof of Theorem 2.1.

**5. Fluctuations.** The (complex or real) Wigner matricial models under consideration are the same as previously [i.e., defined by (i) in Section 2] but now we assume that the perturbation  $A_N$  is diagonal:  $A_N = \text{diag}(\theta, 0, \dots, 0)$  with unique nonnull eigenvalue  $\theta > \sigma$ . According to the previous section, the a.s. convergence of  $\lambda_1(M_N)$  toward  $\rho_\theta$  is universal in the sense that it does not depend on  $\mu$ .

In the first part of this section, we will show that the fluctuations of  $\lambda_1(M_N)$  around this universal limit are not universal any more. Indeed, we are going to prove that  $\sqrt{N}(1 - \sigma^2/\theta^2)^{-1}(\lambda_1(M_N) - \rho_\theta)$  converges in distribution toward the convolution of  $\mu$  and a Gaussian distribution. Hence, the limiting distribution clearly varies with  $\mu$  and in particular cannot be Gaussian unless  $\mu$  is Gaussian.

In the second part of this section, we will sharpen the analysis of the particular Deformed GOE model and explain how this gives Theorem 2.4.

5.1. *Basic tools.* We start with the following results which will be of basic use later on. Note that in the following, a complex random variable  $x$  will be said to be *standardized* if  $\mathbb{E}(x) = 0$  and  $\mathbb{E}(|x|^2) = 1$ .

**THEOREM 5.1** (Lemma 2.7 [3]). *Let  $B = (b_{ij})$  be an  $N \times N$  Hermitian matrix and  $Y_N$  be a vector of size  $N$  which contains i.i.d. standardized entries with bounded fourth moment. Then there is a constant  $K > 0$  such that*

$$\mathbb{E}|Y_N^* B Y_N - \text{Tr } B|^2 \leq K \text{Tr}(B B^*).$$

**THEOREM 5.2** (cf. [6] or Appendix by J. Baik and J. Silverstein). *Let  $B = (b_{ij})$  be a  $N \times N$  random Hermitian matrix and  $Y_N = (y_1, \dots, y_N)$  be an independent vector of size  $N$  which contains i.i.d. standardized entries with bounded fourth moment and such that  $\mathbb{E}(y_1^2) = 0$  if  $y_1$  is complex. Assume that:*

- (i) *there exists a constant  $a > 0$  (not depending on  $N$ ) such that  $\|B\| \leq a$ ,*
- (ii)  *$\frac{1}{N} \text{Tr } B^2$  converges in probability to a number  $a_2$ ,*
- (iii)  *$\frac{1}{N} \sum_{i=1}^N b_{ii}^2$  converges in probability to a number  $a_1^2$ .*

*Then the random variable  $(1/\sqrt{N})(Y_N^* B Y_N - \text{Tr } B)$  converges in distribution to a Gaussian variable with mean zero and variance*

$$(\mathbb{E}|y_1|^4 - 1 - t/2)a_1^2 + (t/2)a_2,$$

*where  $t = 4$  when  $y_1$  is real and is 2 when  $y_1$  is complex.*

**PROOF.** This result is in fact a particular case of a more general result of [6] (Theorems 7.1 and 7.2) which follows from the method of moments.

We give an alternative elegant proof by J. Baik and J. Silverstein in the [Appendix](#) of the present paper.  $\square$

**THEOREM 5.3** (Theorem 1.1 in [5]). *Let  $f$  be an analytic function on an open set of the complex plane including  $[-2\sigma, 2\sigma]$ . If the entries of a general Wigner matrix  $W_N = ((W_N)_{ij})_{1 \leq i, j \leq N}$  satisfy the conditions:*

- for  $i \neq j$ ,  $\mathbb{E}(|(W_N)_{ij}|^4) = \text{const}$ ,
- for any  $\eta > 0$ ,  $\lim_{N \rightarrow +\infty} \frac{1}{\eta^4 N^2} \sum_{i, j} \mathbb{E}[|(W_N)_{ij}|^4 \mathbf{1}_{\{|(W_N)_{ij}| \geq \eta\sqrt{N}\}}] = 0$ ,

*then the random variable  $N(\text{tr}_N(f(\frac{1}{\sqrt{N}}W_N)) - \int f d\mu_{sc})$  converges in distribution toward a Gaussian variable.*

In our setting,  $\mu$  satisfies a Poincaré inequality and thus, as already noticed in Section 2,  $\mu$  satisfies  $\int |x|^q d\mu(x) < +\infty$  for any  $q$  in  $\mathbb{N}$ . Hence, the general Wigner matrices we consider obviously satisfy the conditions of Theorem 5.3. Nevertheless, in the following study of fluctuations, we do not use the Poincaré inequality; thus one can expect that Theorem 2.2 is still valid under assumptions on the only four first moments of  $\mu$  provided one can prove the a.s. convergence of  $\lambda_1(M_N)$  toward  $\rho_\theta$  under these weaker assumptions.

**5.2. Proof of Theorem 2.2.** The approach is the same for the complex and real settings and is close to the one of [25] and the ideas of [11]. Let  $\widehat{M}_{N-1}$  be the  $(N-1) \times (N-1)$  matrix obtained from  $M_N$  removing the first row and the first column. Thus,  $\sqrt{N/(N-1)}\widehat{M}_{N-1}$  is a non-Deformed Wigner matrix associated with the measure  $\mu$ . We denote by  $\lambda_1(\widehat{M}_{N-1})$  [resp.,  $\lambda_{N-1}(\widehat{M}_{N-1})$ ] the largest (resp., lowest) eigenvalue of  $\widehat{M}_{N-1}$ .

Let  $0 < \delta < (\rho_\theta - 2\sigma)/4$ . Let us define the event

$$\Omega_N = \{\lambda_1(\widehat{M}_{N-1}) \leq 2\sigma + \delta; \lambda_{N-1}(\widehat{M}_{N-1}) \geq -2\sigma - \delta; \lambda_1(M_N) \geq \rho_\theta - \delta\}.$$

According to (2.9) and Theorem 2.1,  $\lim_{N \rightarrow +\infty} \mathbb{P}(\Omega_N) = 1$ . Thus, it is sufficient to restrict ourselves to the event  $\Omega_N$  in order to study the convergence in distribution of  $\sqrt{N}(1 - \sigma^2/\theta^2)^{-1}(\lambda_1(M_N) - \rho_\theta)$ .

Let  $V = {}^t(v_1, \dots, v_N)$  be an eigenvector corresponding to  $\lambda_1(M_N)$ . Define the following vectors in  $\mathbb{C}^{N-1}$ :

$$\widehat{V} = {}^t(v_2, \dots, v_N)$$

and

$$\check{M}_{.1} = {}^t((M_N)_{21}, \dots, (M_N)_{N1}) = \frac{1}{\sqrt{N}} {}^t((W_N)_{21}, \dots, (W_N)_{N1}).$$

Then,

$$M_N V = \lambda_1(M_N) V \iff \begin{cases} \theta v_1 + \frac{(W_N)_{11}}{\sqrt{N}} v_1 + \check{M}_{\cdot 1}^* \widehat{V} = \lambda_1(M_N) v_1, \\ \check{M}_{\cdot 1} v_1 + \widehat{M}_{N-1} \widehat{V} = \lambda_1(M_N) \widehat{V}. \end{cases}$$

On  $\Omega_N$ ,  $\lambda_1(M_N)$  is not an eigenvalue of  $\widehat{M}_{N-1}$  and one can write the eigen-equations using the resolvent  $\widehat{G}(\lambda_1(M_N)) := (\lambda_1(M_N) I_{N-1} - \widehat{M}_{N-1})^{-1}$  as follows:

$$(5.1) \quad \widehat{V} = v_1 \widehat{G}(\lambda_1(M_N)) \check{M}_{\cdot 1},$$

$$(5.2) \quad \lambda_1(M_N) v_1 = \theta v_1 + \frac{(W_N)_{11}}{\sqrt{N}} v_1 + v_1 \check{M}_{\cdot 1}^* \widehat{G}(\lambda_1(M_N)) \check{M}_{\cdot 1}.$$

Since  $v_1$  is obviously nonequal to zero, one gets from (5.2)

$$(5.3) \quad \lambda_1(M_N) = \theta + \frac{(W_N)_{11}}{\sqrt{N}} + \check{M}_{\cdot 1}^* \widehat{G}(\lambda_1(M_N)) \check{M}_{\cdot 1}.$$

Moreover, on  $\Omega_N$ ,  $\rho_\theta$  is not an eigenvalue of  $\widehat{M}_{N-1}$  (recall that  $\rho_\theta > 2\sigma$ ) and the resolvent  $\widehat{G}(\rho_\theta) := (\rho_\theta I_{N-1} - \widehat{M}_{N-1})^{-1}$  is well defined, too. Thus, (5.3) is equivalent to

$$\lambda_1(M_N) - \rho_\theta = \frac{(W_N)_{11}}{\sqrt{N}} + \check{M}_{\cdot 1}^* \widehat{G}(\rho_\theta) \check{M}_{\cdot 1} - \frac{\sigma^2}{\theta} + \check{M}_{\cdot 1}^* [\widehat{G}(\lambda_1(M_N)) - \widehat{G}(\rho_\theta)] \check{M}_{\cdot 1}.$$

Using  $\widehat{G}(\lambda_1(M_N)) - \widehat{G}(\rho_\theta) = -(\lambda_1(M_N) - \rho_\theta) \widehat{G}(\rho_\theta) \widehat{G}(\lambda_1(M_N))$  and  $g_\sigma(\rho_\theta) = 1/\theta$ , one gets (on  $\Omega_N$ )

$$\begin{aligned} & \lambda_1(M_N) - \rho_\theta \\ &= \frac{(W_N)_{11}}{\sqrt{N}} + \check{M}_{\cdot 1}^* \widehat{G}(\rho_\theta) \check{M}_{\cdot 1} - \sigma^2 g_\sigma(\rho_\theta) \\ & \quad - \check{M}_{\cdot 1}^* [(\lambda_1(M_N) - \rho_\theta) \widehat{G}(\rho_\theta) (\widehat{G}(\rho_\theta) - (\lambda_1(M_N) - \rho_\theta) \widehat{G}(\rho_\theta)) \\ & \quad \quad \quad \times \widehat{G}(\lambda_1(M_N)))] \check{M}_{\cdot 1}. \end{aligned}$$

Finally, defining  $f_\theta(z) := \frac{1}{\rho_\theta - z} \mathbf{1}_{|z| \leq 2\sigma + \delta}$ , we can easily deduce from the previous equality the following identity on  $\Omega_N$ :

$$(5.4) \quad \begin{aligned} & \{1 + c_N + \delta_1(N) + \delta_2(N)\} \sqrt{N} (\lambda_1(M_N) - \rho_\theta) \\ &= (W_N)_{11} + \sqrt{\frac{N}{N-1}} d_N + \sqrt{\frac{N}{N-1}} \delta_3(N), \end{aligned}$$

where

$$\begin{aligned} c_N &= \sigma^2 \operatorname{tr}_{N-1} [f_\theta^2(\widehat{M}_{N-1})], \\ d_N &= \sqrt{N-1} \{ \check{M}_{\cdot 1}^* \widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \check{M}_{\cdot 1} - \sigma^2 \operatorname{tr}_{N-1} \widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \}, \\ \delta_1(N) &= -(\lambda_1(M_N) - \rho_\theta) \check{M}_{\cdot 1}^* [\widehat{G}(\rho_\theta)]^2 \widehat{G}(\lambda_1(M_N)) \check{M}_{\cdot 1} 1_{\Omega_N}, \\ \delta_2(N) &= \check{M}_{\cdot 1}^* [\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}]^2 \check{M}_{\cdot 1} - \sigma^2 \operatorname{tr}_{N-1} [\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}]^2, \\ \delta_3(N) &= \sigma^2 \sqrt{N-1} \left\{ \operatorname{tr}_{N-1} (f_\theta(\widehat{M}_{N-1})) - \int f_\theta d\mu_{sc} \right\}. \end{aligned}$$

First

$$\begin{aligned} |\delta_1(N)| &\leq |\lambda_1(M_N) - \rho_\theta| \|\check{M}_{\cdot 1}\|^2 \|\widehat{G}(\rho_\theta)\|^2 \|\widehat{G}(\lambda_1(M_N))\| 1_{\Omega_N} \\ &\leq \frac{1}{(\rho_\theta - 2\sigma - 2\delta)(\rho_\theta - 2\sigma - \delta)^2} \frac{1}{N} \sum_{j=2}^N |(W_N)_{j1}|^2 \times |\lambda_1(M_N) - \rho_\theta|, \end{aligned}$$

[using Lemma 3.2(v)]. By the law of large numbers  $\frac{1}{N} \sum_{j=2}^N |(W_N)_{j1}|^2$  converges a.s. toward  $\sigma^2$  and according to Theorem 2.1,  $|\lambda_1(M_N) - \rho_\theta|$  converges a.s. to zero. Hence  $\delta_1(N)$  converges obviously in probability toward zero.

Now, since  $f_\theta$  is analytic on an open set including  $[-2\sigma, 2\sigma]$ , we deduce from Theorem 5.3 the convergence in probability of  $\delta_3(N)$  toward zero and of  $c_N$  toward  $\sigma^2 \int f_\theta^2 d\mu_{sc} = \frac{\sigma^2}{\theta^2 - \sigma^2}$ .

According to Theorem 5.1 and using Lemma 3.2(v),

$$\begin{aligned} \mathbb{E}(|\delta_2(N)|^2) &\leq \frac{K}{N-1} \mathbb{E}(\operatorname{tr}_N [\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}]^4) \\ &\leq \frac{K}{N-1} \mathbb{E}(\|\widehat{G}(\rho_\theta)\|^4 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}) \\ &\leq \frac{K}{N-1} \frac{1}{(\rho_\theta - 2\sigma - \delta)^4}. \end{aligned}$$

The convergence in probability of  $\delta_2(N)$  toward zero readily follows by Chebyshev inequality.

Let us check that  $\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}$  satisfies the conditions of Theorem 5.2.

- (i)  $\|\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}\| \leq \frac{1}{\rho_\theta - 2\sigma - \delta}$  by Lemma 3.2(v).
- (ii) As already noticed,  $\operatorname{tr}_{N-1} f_\theta^2(\widehat{M}_{N-1})$  converges in probability toward  $\int f_\theta^2 d\mu_{sc}$ . Since on the event  $\{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta\}$ , with limiting probability 1,  $\operatorname{tr}_{N-1} [\widehat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}]^2$  coincides with  $\operatorname{tr}_{N-1} f_\theta^2(\widehat{M}_{N-1})$ , it also converges in probability toward  $\int f_\theta^2 d\mu_{sc}$ .

(iii) It is proved in Proposition 3.1 in [13] that for any  $z \in \mathbb{C}$  such that  $\Im m(z) > 0$ ,  $\frac{1}{N-1} \sum_{i=1}^{N-1} ([\widehat{G}(z)]_{ii})^2$  converges in probability toward  $g_\sigma^2(z)$ . The same result holds for  $\frac{1}{N-1} \sum_{i=1}^{N-1} ([\widehat{G}(z)]_{ii})^2 \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}$ . For any  $\epsilon > 0$  and any  $\alpha > 0$ ,

$$\begin{aligned} & \mathbb{P} \left( \left| \frac{1}{N-1} \sum_{i=1}^{N-1} ([\widehat{G}(\rho_\theta)]_{ii})^2 \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} - g_\sigma^2(\rho_\theta) \right| > \epsilon \right) \\ & \leq \mathbb{P} \left( \left| \frac{1}{N-1} \sum_{i=1}^{N-1} \{([\widehat{G}(\rho_\theta)]_{ii})^2 - ([\widehat{G}(\rho_\theta + i\alpha)]_{ii})^2\} \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \right| > \frac{\epsilon}{3} \right) \\ & \quad + \mathbb{P} \left( \left| \frac{1}{N-1} \sum_{i=1}^{N-1} ([\widehat{G}(\rho_\theta + i\alpha)]_{ii})^2 \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} - g_\sigma^2(\rho_\theta + i\alpha) \right| > \frac{\epsilon}{3} \right) \\ & \quad + \mathbb{P} \left( |g_\sigma^2(\rho_\theta) - g_\sigma^2(\rho_\theta + i\alpha)| > \frac{\epsilon}{3} \right). \end{aligned}$$

Since

$$\begin{aligned} & \{([\widehat{G}(\rho_\theta)]_{ii})^2 - ([\widehat{G}(\rho_\theta + i\alpha)]_{ii})^2\} \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \\ & = [\widehat{G}(\rho_\theta) - \widehat{G}(\rho_\theta + i\alpha)]_{ii} [\widehat{G}(\rho_\theta) + \widehat{G}(\rho_\theta + i\alpha)]_{ii} \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \\ & = i\alpha [\widehat{G}(\rho_\theta) \widehat{G}(\rho_\theta + i\alpha)]_{ii} [\widehat{G}(\rho_\theta) + \widehat{G}(\rho_\theta + i\alpha)]_{ii} \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}, \end{aligned}$$

we get by using Lemma 3.2(v)

$$|([\widehat{G}(\rho_\theta)]_{ii})^2 - ([\widehat{G}(\rho_\theta + i\alpha)]_{ii})^2| \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta} \leq \frac{2\alpha}{(\rho_\theta - 2\sigma - \delta)^3}.$$

Similarly, we get that

$$|g_\sigma^2(\rho_\theta) - g_\sigma^2(\rho_\theta + i\alpha)| \leq \frac{2\alpha}{(\rho_\theta - 2\sigma)^3}.$$

Thus, choosing  $\alpha$  such that  $\frac{2\alpha}{(\rho_\theta - 2\sigma - \delta)^3} < \frac{\epsilon}{3}$ , we readily deduce the convergence in probability of

$$\frac{1}{N-1} \sum_{i=1}^{N-1} ([\widehat{G}(\rho_\theta)]_{ii})^2 \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}$$

toward  $g_\sigma^2(\rho_\theta)$ .

Since  $\widehat{G}(\rho_\theta) \mathbf{1}_{\|\widehat{M}_{N-1}\| \leq 2\sigma + \delta}$  and  $\check{M}_{\cdot 1}$  are independent, we can deduce from Theorem 5.2 that  $d_N$  converges in distribution toward a Gaussian law with mean zero and variance

$$v_\theta := \sigma^4 \left\{ \left( \mathbb{E} \left( \left| \frac{(W_N)_{12}}{\sigma} \right|^4 \right) - 1 - t/2 \right) \frac{1}{\theta^2} + \frac{t}{2} \frac{1}{\theta^2 - \sigma^2} \right\},$$



where  $t = 4$  in the real setting and  $t = 2$  in the complex one. Note that one readily verifies that  $v_\theta$  satisfies (2.12) in Section 2.

Let  $0 < \epsilon < 1$ . Since  $\delta_1(N) + \delta_2(N)$  converges in probability toward zero, the probability of the event

$$(5.5) \quad \tilde{\Omega}_N = \Omega_N \cap \{|\delta_1(N) + \delta_2(N)| \leq \epsilon\}$$

tends to 1. Now, since  $c_N \geq 0$  we have the following identity on  $\tilde{\Omega}_N$ :

$$(5.6) \quad \sqrt{N}(\lambda_1(M_N) - \rho_\theta) = \frac{1}{u_N} \left\{ (W_N)_{11} + \sqrt{\frac{N}{N-1}} d_N + \sqrt{\frac{N}{N-1}} \delta_3(N) \right\}$$

with  $u_N := 1 + c_N + \delta_1(N) + \delta_2(N)$  converging in distribution toward  $(1 - \sigma^2/\theta^2)^{-1}$ . Moreover, since  $(W_N)_{11}$  and  $d_N$  are independent,  $(W_N)_{11} + \sqrt{N/(N-1)}d_N + \sqrt{N/(N-1)}\delta_3(N)$  converges in distribution toward the convolution of  $\mu$  and a Gaussian distribution  $\mathcal{N}(0, v_\theta)$ .

Finally, we can conclude that  $\sqrt{N}(1 - \sigma^2/\theta^2)^{-1}(\lambda_1(M_N) - \rho_\theta)$  converges in distribution toward  $\mu * \mathcal{N}(0, v_\theta)$ .

*5.3. Proof of Theorem 2.4.* As before,  $\theta$  is assumed to be  $> \sigma$ . In Theorem 2.4, we consider the real Deformed models and claim that the full deformation  $A_N$  defined by  $(A_N)_{ij} = \theta/N$  exhibits universality of the Gaussian fluctuations of the largest eigenvalue around  $\rho_\theta$ . As already stated, the analogue of this result holds in the complex setting. This is one of the conclusions of the work [16] which also partly solves the real case (we recall to the reader that all the results of [16] readily extend to the framework of Theorem 2.4 calling on [27]). In order to explain this more precisely, let us summarize the main arguments developed by [16] in the complex setting. First, it is shown that the universality of the fluctuations follows from the universality of limits of expectations of traces of suitable high powers of any Deformed Wigner matrices (the powers are of the order of  $\sqrt{N}$ ). Second (this is the main part of the work [16]), to handle such expectations, the authors perform a combinatorial method inspired by [31] and then deduce that in the large limit  $N \rightarrow \infty$ , the previous expectations behave as in the Gaussian case. The last step of the analysis calls on the investigations of [26] on the Deformed GUE which allow to identify the value of these limits.

Actually, the combinatorial arguments also work in the real setting (see in particular Section 2 in [16]) and reduce the universality problem to the knowledge of the Deformed GOE. Thus, to get the result of Theorem 2.4, it suffices to prove (using the orthogonal invariance of the GOE) the following limit.

**PROPOSITION 5.1.** *Call  $L_\theta$  the Laplace transform of the law  $\mathcal{N}(0, 2\sigma_\theta^2)$ . Let  $M_N^G$  be the Deformed GOE with  $A_N = \text{diag}(\theta, 0, \dots, 0)$  and  $\theta > \sigma$ .*

For any  $t$  in  $[0, \rho_\theta[$ ,

$$(5.7) \quad \lim_N \mathbb{E}[\mathrm{Tr}(M_N^G / \rho_\theta)^{2[t\sqrt{N}]}] = L_\theta(2t / \rho_\theta).$$

The starting point of our computations is the following result which states that the previous expectation only involves (as  $N \rightarrow \infty$ ) the rescaled largest eigenvalue of the Deformed GOE

$$\xi_1^G = \sqrt{N}(\lambda_1(M_N^G) - \rho_\theta).$$

LEMMA 5.1. *For any  $t > 0$ ,*

$$(5.8) \quad \mathbb{E}[\mathrm{Tr}(M_N^G / \rho_\theta)^{2[t\sqrt{N}]}] = \mathbb{E}[\exp(2t\xi_1^G / \rho_\theta) 1_{|\xi_1^G| \leq N^{1/6}}](1 + o(1)).$$

This formula does not appear explicitly in [16] but all the arguments needed for its justification can be found in it (actually one can show that the formula holds for any Deformed Wigner model  $M_N$  satisfying the assumptions of Theorem 2.4). We will not give the proof and refer the reader to Section 2 in [16].

Hence, to derive Proposition 5.1, it remains to show the next lemma on  $\xi_1^G$ .

LEMMA 5.2. *For any  $t$  in  $[0, 2[$ ,*

$$(5.9) \quad \lim_N \mathbb{E}[\exp(t\xi_1^G) 1_{|\xi_1^G| \leq N^{1/6}}] = L_\theta(t).$$

PROOF. Observe first that it is enough to show that

$$(5.10) \quad \lim_N \mathbb{E}[\exp(t\xi_1^G) 1_{|\xi_1^G| \leq N^{1/6}} 1_{\tilde{\Omega}_N}] = L_\theta(t),$$

where the event  $\tilde{\Omega}_N$  was defined above by (5.5) choosing  $\delta > 0$  smaller than  $\min\{\frac{\rho_\theta - 2\sigma}{4}; \frac{1}{3} \int \frac{1}{\rho_\theta - x} d\mu_{sc}(x)\}$ . Indeed, by the Cauchy–Schwarz inequality,

$$(\mathbb{E}[\exp(t\xi_1^G) 1_{|\xi_1^G| \leq N^{1/6}} 1_{c\tilde{\Omega}_N}])^2 \leq \mathbb{E}[\exp(2t\xi_1^G) 1_{|\xi_1^G| \leq N^{1/6}}] \times \mathbb{P}(c\tilde{\Omega}_N).$$

The previous right-hand side is negligible as  $N \rightarrow \infty$  since the probability vanishes and the expectation is bounded since [16] proved that the left-hand side of (5.8) is bounded, too.

The occurrence of the event  $\tilde{\Omega}_N$  allows to make use of the relevant representation (5.6) of  $\xi_1^G$  obtained in the previous Section 5.2:

$$(5.11) \quad \xi_1^G = \frac{1}{u_N^G} \left\{ (W_N^G)_{11} + \sqrt{\frac{N}{N-1}} d_N^G + \sqrt{\frac{N}{N-1}} \delta_3^G(N) \right\}.$$

Second, by Fubini's theorem one can check that

$$(5.12) \quad \begin{aligned} & \mathbb{E}[\exp(t\xi_1^G)1_{|\xi_1^G| \leq N^{1/6}}1_{\tilde{\Omega}_N}] \\ &= \int_{\mathbb{R}} e^x \mathbb{P}[\{t\xi_1^G \geq x\} \cap \tilde{\Omega}_N \cap \{|\xi_1^G| \leq N^{1/6}\}] dx. \end{aligned}$$

By Theorem 2.2,

$$(5.13) \quad \mathbb{P}[\{t\xi_1^G \geq x\} \cap \tilde{\Omega}_N \cap \{|\xi_1^G| \leq N^{1/6}\}] \xrightarrow{N \rightarrow +\infty} \mathbb{P}[t\mathcal{N} \geq x],$$

where  $\mathcal{N}$  is a centered Gaussian variable with variance  $2\sigma_\theta^2$ . We want to deduce (5.10) from (5.12) and (5.13) by the dominated convergence theorem. Thus, we are going to prove that there exists a function  $h$  such that for  $N$  large enough and for any  $x$ ,

$$\mathbb{P}[\{t\xi_1^G \geq x\} \cap \tilde{\Omega}_N \cap \{|\xi_1^G| \leq N^{1/6}\}] \leq h(x)$$

with  $\int_{\mathbb{R}} e^x h(x) dx < +\infty$ . Note that for  $x \leq 0$ , the result is obvious setting  $h(x) = 1$ . Let  $x$  be nonnegative. We shall improve the general analysis made in the previous Section 5.2 thanks to the particular Gaussian setting considered here. For all  $N$  large enough,

$$\begin{aligned} & \mathbb{P}[\{t\xi_1^G \geq x\} \cap \tilde{\Omega}_N \cap \{|\xi_1^G| \leq N^{1/6}\}] \\ & \leq \mathbb{P}\left[(W_N^G)_{11} \geq \frac{x(1-\varepsilon)}{3t}\right] + \mathbb{P}\left[\sqrt{\frac{N}{N-1}}d_N^G \geq \frac{x(1-\varepsilon)}{3t}\right] \\ & \quad + \mathbb{P}\left[\sqrt{\frac{N}{N-1}}\delta_3^G(N) \geq \frac{x(1-\varepsilon)}{3t}\right] \\ & = J_N^{(1)}(x) + J_N^{(2)}(x) + J_N^{(3)}(x). \end{aligned}$$

$J_N^{(1)}(x) = J^{(1)}(x)$  does not depend on  $N$  and we have  $\int_{\mathbb{R}} e^x J^{(1)}(x) dx = \mathbb{E}[\exp(\frac{3t}{1-\varepsilon}(W_N^G)_{11})] < +\infty$ . Besides, one can easily see that the choice of  $\delta$  insures that  $J_N^{(3)}(x) = 0$ . By the Chebyshev inequality, we have

$$J_N^{(2)}(x) \leq \exp(-6x(1-\varepsilon)/3t)\mathbb{E}(\mathcal{E}),$$

where  $\mathcal{E} = \mathcal{E}'\mathcal{E}''$  with

$$\mathcal{E}' = \mathbb{E}[\exp(6\sqrt{N}^t \check{M}_{\cdot 1}^G \hat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}^G\| \leq 2\sigma + \delta} \check{M}_{\cdot 1}^G) | \widehat{M}_{N-1}^G],$$

$$\mathcal{E}'' = \exp[-6\sigma^2 \sqrt{N} \text{tr}_{N-1}(\hat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}^G\| \leq 2\sigma + \delta})].$$

Using the Gaussian assumptions (see [28], pages 90–91), one has

$$\mathcal{E}' = \det\left(I_{N-1} - 12\frac{\sigma^2}{\sqrt{N}}\hat{G}(\rho_\theta) 1_{\|\widehat{M}_{N-1}^G\| \leq 2\sigma + \delta}\right)^{-1/2}$$

$$= \prod_{i=1}^{N-1} (1 - 12\sigma^2\beta_i/\sqrt{N})^{-1/2}$$

for large enough  $N$ , where the  $\beta_i$ 's are the eigenvalues of  $\widehat{G}(\rho_\theta)1_{\|\widehat{M}_{N-1}^G\| \leq 2\sigma + \delta}$ .

Note that  $0 \leq \beta_i < \frac{1}{3\delta}$  so that the last identities make sense, for instance, for  $N > \frac{16\sigma^4}{\delta^2}$ . Hence,

$$\ln \mathcal{E}' \mathcal{E}'' \leq \frac{1}{2} \sum_{i=1}^{N-1} \{-\ln(1 - 12\sigma^2\beta_i/\sqrt{N}) - 12\sigma^2\beta_i/\sqrt{N}\}.$$

Let  $\alpha > \frac{1}{2}$ ; using that for any  $y$  in  $[0, 1 - \frac{1}{2\alpha}]$ , we have  $-\ln(1 - y) - y \leq \alpha y^2$ . So, as  $\beta_i < \frac{1}{3\delta}$ , we get that for  $N > 16\sigma^4\delta^{-2}(1 - \frac{1}{2\alpha})^{-2}$ ,

$$\ln \mathcal{E}' \mathcal{E}'' \leq \frac{\alpha 12^2 \sigma^4}{18\delta^2}.$$

Thus, there is some constant  $C_{\alpha, \sigma, \delta}$  such that  $\mathcal{E}' \mathcal{E}'' \leq C_{\alpha, \sigma, \delta}$  and

$$J_N^{(2)}(x) \leq C_{\alpha, \sigma, \delta} \exp(-2x(1 - \varepsilon)/t).$$

Now, for  $0 < t < 2(1 - \varepsilon)$ ,  $\int_0^{+\infty} \exp(x - 2x(1 - \varepsilon)/t) dx < \infty$ . The proof is complete.  $\square$

## APPENDIX: BY J. BAIK AND J. SILVERSTEIN

This [Appendix](#) presents the proof by J. Baik and J. Silverstein of the CLT (given by [Theorem 5.2](#)) needed in the previous section for the proof of [Theorem 2.2](#). Their proof is based on a writing of the expression

$$(A.1) \quad (1/\sqrt{N})(Y_N^* B Y_N - \text{Tr } B)$$

as a sum of martingale differences, and uses the following CLT.

**THEOREM A.1** (Theorem 35.12 of [\[10\]](#)). *For each  $N$ , let  $Z_{N1}, \dots, Z_{Nr_N}$  be a real martingale difference sequence with respect to the increasing  $\sigma$ -field  $\{\mathcal{F}_{N,j}\}$  having second moments. If, as  $N \rightarrow \infty$ ,*

$$(A.2) \quad \sum_{j=1}^{r_N} \mathbb{E}(Z_{Nj}^2 | \mathcal{F}_{N,j-1}) \xrightarrow{P} v^2,$$

where  $v^2$  is a positive constant, and for each  $\epsilon > 0$ ,

$$(A.3) \quad \sum_{j=1}^{r_N} \mathbb{E}(Z_{Nj}^2 1_{|Z_{Nj}| \geq \epsilon}) \rightarrow 0,$$

then

$$\sum_{j=1}^{r_N} Z_{Nj} \xrightarrow{\mathcal{L}} \mathcal{N}(0, v^2).$$

PROOF OF THEOREM 5.2. First, one can write (A.1) as a sum of martingale differences:

$$\begin{aligned} & (1/\sqrt{N})(Y_N^* B Y_N - \text{Tr } B) \\ &= (1/\sqrt{N}) \sum_{i=1}^N \left( (|y_i|^2 - 1)b_{ii} + \bar{y}_i \sum_{j<i} y_j b_{ij} + \bar{y}_i \sum_{j>i} y_j b_{ij} \right) \\ &= (1/\sqrt{N}) \sum_{i=1}^N \left( (|y_i|^2 - 1)b_{ii} + \bar{y}_i \sum_{j<i} y_j b_{ij} + y_i \sum_{j<i} \bar{y}_j b_{ji} \right) = \sum_{i=1}^N Z_i, \end{aligned}$$

where

$$Z_i = Z_{Ni} = (1/\sqrt{N}) \left( (|y_i|^2 - 1)b_{ii} + \bar{y}_i \sum_{j<i} y_j b_{ij} + y_i \sum_{j<i} \bar{y}_j b_{ij} \right).$$

Let  $\mathcal{F}_{N,i}$  (resp.,  $\mathcal{F}_{N,0}$ ) be the  $\sigma$ -field generated by  $y_1, \dots, y_i$  and  $B$  (resp., by  $B$ ). Let also  $\mathbb{E}_i(\cdot)$  denote conditional expectation with respect to  $\mathcal{F}_{N,i}$ . It is clear that  $Z_i$  is measurable with respect to  $\mathcal{F}_{N,i}$  and satisfies  $\mathbb{E}_{i-1}(Z_i) = 0$ .

We will show the conditions of Theorem A.1 are met.

To verify the Lindeberg condition (A.3), we need to show this property is closed under addition. This will follow from the following fact. For random variables  $X_1, X_2$ , and positive  $\epsilon$ ,

$$(A.4) \quad \mathbb{E}(|X_1 + X_2|^2 1_{|X_1 + X_2| \geq \epsilon}) \leq 4(\mathbb{E}(|X_1|^2 1_{|X_1| \geq \epsilon/2}) + \mathbb{E}(|X_2|^2 1_{|X_2| \geq \epsilon/2})).$$

Indeed, we have

$$\begin{aligned} \mathbb{E}(|X_1|^2 1_{|X_1 + X_2| \geq \epsilon}) &\leq \mathbb{E}(|X_1|^2 1_{(|X_1| \geq \epsilon/2)}) + \mathbb{E}(|X_1|^2 1_{(|X_1| < \epsilon/2, |X_2| \geq \epsilon/2)}) \\ &\leq \mathbb{E}(|X_1|^2 1_{(|X_1| \geq \epsilon/2)}) + (\epsilon^2/4) \mathbb{P}(|X_2| \geq \epsilon/2) \\ &\leq \mathbb{E}(|X_1|^2 1_{(|X_1| \geq \epsilon/2)}) + \mathbb{E}(|X_2|^2 1_{(|X_2| \geq \epsilon/2)}). \end{aligned}$$

The same bound starting with  $X_2$  leads to (A.4).

Write  $Z_i = X_1^i + X_2^i$ , with  $X_1^i = (1/\sqrt{N})(|y_i|^2 - 1)b_{ii}$ . Then for  $\epsilon > 0$ ,

$$(A.5) \quad \sum_{i=1}^N \mathbb{E}(|X_1^i|^2 1_{(|X_1^i| \geq \epsilon)}) \leq a^2 \mathbb{E}((|y_1|^2 - 1)^2 1_{(|x_1|^2 - 1| \geq \sqrt{N}\epsilon/a}) \rightarrow 0$$

as  $N \rightarrow \infty$ , by the dominated convergence theorem.

We have

$$\begin{aligned}
\mathbb{E} \left| \sum_{j<i} y_j b_{ij} \right|^4 &= \mathbb{E} \left( |y_1|^4 \sum_{j<i} |b_{ij}|^4 \right) + 2\mathbb{E} \left( \sum_{*} |b_{ij_1}|^2 |b_{ij_2}|^2 \right) \\
&\quad + \mathbb{E} \left( |y_1^2|^2 \sum_{*} b_{ij_1}^2 \bar{b}_{ij_2}^2 \right) \\
&\leq \mathbb{E} |y_1|^4 \mathbb{E} \left[ \max_j (B^2)_{jj} (B^2)_{ii} \right] + (2 + \mathbb{E} |y_1^2|^2) \mathbb{E} [(B^2)_{ii}^2] \\
&\leq a^4 [\mathbb{E} |y_1|^4 + 2 + \mathbb{E} |y_1^2|^2],
\end{aligned}$$

where the sum  $\sum_{*}$  is over  $\{j_1 < i, j_2 < i, j_1 \neq j_2\}$ . Therefore  $\mathbb{E} |X_2^i|^4 = o(N^{-1})$  so that for any  $\epsilon > 0$ ,

$$(A.6) \quad \sum_{i=1}^N \mathbb{E} (|X_2^i|^2 \mathbf{1}_{(|X_2^i| \geq \epsilon)}) \leq (1/\epsilon^2) \sum_{i=1}^N \mathbb{E} |X_2^i|^4 \rightarrow 0 \quad \text{as } N \rightarrow \infty.$$

Thus, by (A.5), (A.6) and (A.4),  $\{Z_i\}$  satisfies (A.3).

Now, we shall verify condition (A.2). We have

$$\begin{aligned}
&\sum_{i=1}^N \mathbb{E}_{i-1} Z_i^2 \\
&= (1/N) \sum_{i=1}^N \left\{ (\mathbb{E} |y_1|^4 - 1) b_{ii}^2 + \mathbb{E} \bar{y}_1^2 \left( \sum_{j<i} y_j b_{ij} \right)^2 \right. \\
&\quad \left. + \mathbb{E} y_1^2 \left( \sum_{j<i} \bar{y}_j \bar{b}_{ij} \right)^2 + 2\mathbb{E} (|y_1|^2 \bar{y}_1) b_{ii} \sum_{j<i} y_j b_{ij} \right. \\
&\quad \left. + 2\mathbb{E} (|y_1|^2 y_1) b_{ii} \sum_{j<i} \bar{y}_j \bar{b}_{ij} + 2 \left( \sum_{j<i} y_j b_{ij} \right) \sum_{j<i} \bar{y}_j \bar{b}_{ij} \right\}.
\end{aligned} \tag{A.7}$$

Let  $B_L$  denote the strictly lower triangular part of  $B$ . We have

$$\mathbb{E} \left[ (1/N) \sum_{i=1}^N b_{ii} \sum_{j<i} y_j b_{ij} \right] = 0$$

and using Cauchy–Schwarz,

$$\mathbb{E} \left| (1/N) \sum_{i=1}^N b_{ii} \sum_{j<i} y_j b_{ij} \right|^2 = \mathbb{E} \left| (1/N) \sum_{j=1}^{N-1} y_j \sum_{i>j} b_{ii} b_{ij} \right|^2$$

$$\begin{aligned}
&= (1/N^2) \mathbb{E} \left( \sum_{j=1}^{N-1} \sum_{i>j} b_{ii} b_{ij} \sum_{\underline{i}>j} b_{\underline{ii}} \bar{b}_{\underline{ij}} \right) \\
&= (1/N^2) \mathbb{E} \left( \sum_{\underline{ii}} b_{ii} b_{\underline{ii}} (B_L B_L^*)_{\underline{ii}} \right) \\
&\leq \mathbb{E} \left[ \left( \max_i b_{ii} \right)^2 (1/N) \left( \sum_{\underline{ii}} |(B_L B_L^*)_{\underline{ii}}|^2 \right)^{1/2} \right] \\
&= \mathbb{E} \left[ \left( \max_i b_{ii} \right)^2 (1/N) \operatorname{Tr}((B_L B_L^*)^2)^{1/2} \right] \\
&\leq \mathbb{E} \left[ \left( \max_i b_{ii} \right)^2 (1/\sqrt{N}) \|B_L\|^2 \right].
\end{aligned}$$

We apply the following bound (due to Mathias; see [24]):  $\|B_L\| \leq \gamma_N \|B\|$  where  $\gamma_N = O(\ln N)$ , and the bound  $\|B\| \leq a$  to conclude that

$$(1/N) \sum_{i=1}^N b_{ii} \sum_{j<i} y_j b_{ij} \xrightarrow{P} 0.$$

Then (recall that  $\mathbb{E}y_1^2 = 0$  when  $y_1$  is complex), (A.7) can be written as

$$\begin{aligned}
\sum_{i=1}^N \mathbb{E}_{i-1} Z_i^2 &= (1/N) \sum_{i=1}^N \left[ (\mathbb{E}|y_1|^4 - 1) b_{ii}^2 \right. \\
&\quad \left. + t \left( \sum_{j<i} y_j b_{ij} \right) \left( \sum_{j<i} \bar{y}_j \bar{b}_{ij} \right) \right] + o_P(1) \\
&= (1/N) \sum_{i=1}^N (\mathbb{E}|y_1|^4 - 1) b_{ii}^2 + t(1/N) Y_N^* B_L^* B_L Y_N + o_P(1),
\end{aligned} \tag{A.8}$$

where  $t = 4$  when  $y_1$  is real, and is 2 when  $y_1$  is complex.

Besides, from Lemma 2.7 in [3] (recalled in Theorem 5.1) we have

$$\begin{aligned}
\mathbb{E} |(1/N)(Y_N^* B_L^* B_L Y_N - \operatorname{Tr}(B_L^* B_L))|^2 &\leq (1/N^2) \mathbb{E} (\operatorname{Tr}(B_L^* B_L)^2) \\
&\leq K \mathbb{E} \|B\|^4 \frac{\ln^4 N}{N} \rightarrow 0
\end{aligned}$$

as  $N \rightarrow \infty$ . So, as

$$\operatorname{Tr} B_L^* B_L = \sum_{j<i} |b_{ij}|^2 = (1/2) \left( \operatorname{Tr} B^2 - \sum_i b_{ii}^2 \right),$$

(A.8) implies that condition (A.2) holds with

$$v^2 = (\mathbb{E}|y_1|^4 - 1 - t/2)a_1^2 + (t/2)a_2.$$

Thus, by Theorem A.1, we deduce that  $(1/\sqrt{N})(Y_N^*BY_N - \text{Tr } B)$  converges in distribution to a Gaussian variable with mean zero and variance  $v^2$ .  $\square$

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