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**Assessing the Effect of Microfinance on Vulnerability and Poverty among  
Low Income Households**

**Abstract**

*We empirically investigate whether participation in Indian Self Help Group microfinance program (SHG) has helped reduced poverty and household vulnerability using cross-sectional SHG rural household survey data. The potential selection bias is eliminated by propensity score matching to estimate the average treatment on treated effect using nearest neighbour matching and local linear regression algorithm. We find that vulnerability in SHG members is not significantly higher than the non-SHG members, even though the SHG members have a high incidence of poverty. However, vulnerability declines significantly for those that have been SHG members for more than one year. These results are found to be robust using sensitivity analysis and Rosenbaum bounds method.*

**JEL Classifications:** D14, G21, I32.

**Keywords:** Microfinance, Vulnerability, Poverty, Self Help Groups.

## 1. Introduction

An extensive literature has examined the impact of microfinance in alleviating poverty (Morduch, 1999). While several studies have shown positive impact in reducing poverty, a number have challenged this view expounding that the results are more mixed (Morduch, 1999; Amin et al., 1999; Puhazhendi and Badatya, 2002; de Aghion and Morduch, 2006; Karlan, 2007).<sup>1</sup> Exploring beyond poverty, this paper investigates if microfinance reduces household vulnerability. In other words, do microfinance programs reduce the household exposure to future shocks and improve their ability to cope with them? Answering this question is crucial since the goal of poverty alleviation is not just about improving economic welfare via increased incomes and consumption. It is also about devising means for preventing households from falling into poverty and enabling them to meet their survival needs including food security, to make productive investments and to avoid selling their limited resources in times of income or expenditure shocks. Static poverty measures are helpful in assessing the current poverty status of households but tend to ignore poverty dynamics over time.<sup>2</sup> Thus even though average household incomes do not fall into poverty levels, their degree of vulnerability or the risk of being poor in the future, can still remain high. The cumulative impact of microfinance programs on the household's wellbeing may therefore not be captured by standard poverty measures alone. A limited literature on the impact of microfinance on vulnerability provides evidence that microfinance tends to strengthen crisis coping mechanisms, helps diversify income-earning sources, and enables asset creation. In fact, a few studies suggest that it has a more significant impact in reducing vulnerability than income-poverty (Hashemi et al., 1996; Morduch, 1999).

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3 Our objectives in this paper are two-fold. First, we estimate two important dimensions of  
4 well-being namely, poverty and ex-ante vulnerability of households in SHG and non-SHG  
5 groups using 2003 rural household survey data. Secondly, we empirically investigate whether  
6 microfinance programs like Self Help Group program (SHG) lead to a reduction in  
7 vulnerability or not. Vulnerability in our study is defined as a forward-looking, ex-ante  
8 measure of the household's ability to cope with future shocks and proneness to food  
9 insecurity that can undermine the household's survival and the development of its members'  
10 capabilities.  
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24 The empirical analysis is based on a 2003 household survey data collected on one of the  
25 largest microfinance programs in the developing world, the National Bank for Agriculture and  
26 Rural Development (NABARD) self-help group (SHG) program in ten rural districts in India.  
27 We estimate several poverty measures as well as an ex-ante vulnerability measure using  
28 Chauduri, Jalan and Suryahadi (2002) methodology, which allows for household vulnerability  
29 estimation using cross-sectional data. We also take into account any variation in the effect of  
30 SHG participation on vulnerability due to difference in the economic environment and the  
31 design of the SHG bank linkage. We correct for potential selection bias in the household  
32 sample using propensity score matching to obtain the average treatment on treated effect  
33 (impact) on vulnerability. Finally we test the sensitivity of the results to unobservables.  
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50 Some researchers suggest that the poor are likely to be more vulnerable (Prowse, 2003;  
51 Cannon et al., 2003; Feldbrügge and von Braun, 2002). If this is the case, then the SHG  
52 members, with a higher proportion of poor households, are likely to be more vulnerable.  
53 Controlling for selection bias, our results show that SHG member households are not more  
54 vulnerable than non-member households, even though a higher proportion of them are poor.  
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3 Among the more mature SHG members however, we find a significant reduction in  
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5 vulnerability compared to the non-SHG members. These results are found to be robust using  
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7 the sensitivity analysis and Rosenbaum bounds method.  
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12 The paper is organized as follows. Section 2 discusses the notion of vulnerability and the  
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14 conceptual framework used in the estimation of vulnerability. Section 3 explores the role of  
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16 microfinance SHGs in reducing vulnerability. Section 4 provides an overview of the sample  
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18 data used in our analysis and the methodologies used in addressing potential participation  
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20 bias, in estimating vulnerability, and in assessing the effect of SHG participation. Section 5  
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22 provides the results of the propensity score matching and the resulting poverty and  
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24 vulnerability estimates for SHG and non-SHG members. The results of sensitivity analyses  
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26 involving the use of affected treatment on treated (ATT) effect and Rosenbaum bounds  
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28 methods to test the robustness of the propensity score matching estimates are provided in  
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30 section 6. Concluding remarks are presented in the final section.  
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## 38 **2. Understanding Vulnerability**

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44 It should be noted that vulnerability as a notional concept, has been viewed differently by  
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46 researchers, thus leading to varied definitions and measures. Some see vulnerability as an  
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48 aspect which can cause poverty or hinder people from escaping out of poverty (Prowse, 2003:  
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50 9). This view that poor people are generally more vulnerable is shared by Cannon et al. (2003)  
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52 and Feldbrügge and von Braun (2002). Some have taken a different perspective of  
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54 vulnerability whereby poverty is viewed as one element, which may contribute to an  
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56 enhanced vulnerability (Cardona, 2004). Others such as Calvo (2008) treat vulnerability as a  
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58 dimension of poverty itself and define it as a threat of suffering any form of poverty in the  
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3 future.<sup>3</sup> In Calvo and Dercon's (2005) model, vulnerability is seen as a combination of  
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5 poverty (failure to reach a minimum outcome) and risk (dispersion over states of the world)  
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7 that translates into a threat of being poor in the next period. This notion of vulnerability builds  
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9 upon the probability of outcomes failing to reach the minimal standard as well as on the  
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11 uncertainty about how far households may fall below that threshold. This uncertainty is a  
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13 source of distress and impinges directly on well-being. Chauduri, Jalan and Suryahadi (2002)  
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15 in their study of Indonesian households, define vulnerability within the framework of poverty  
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17 eradication as the ex-ante risk that a household will, if currently non-poor, fall below the  
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19 poverty line, or if currently poor, will remain in poverty. (p. 4). On the other hand, Ligon and  
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21 Schechter (2003) take a utilitarian approach in defining vulnerability, arguing that it depends  
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23 not only on the mean of household consumption but also on variation in consumption in the  
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25 context of a risky environment. The risk faced by the household is decomposed into aggregate  
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27 and idiosyncratic risk. A growing number of empirical studies have proposed varied measures  
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29 and proxy indicators of vulnerability as well (Zimmerman and Carter, 2003; Calvo and  
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31 Dercon, 2005; Glewwe and Hall, 1999; Ligon and Schechter, 2003; Carter and Barrett, 2006;  
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33 Morduch, 2005). Some make use of household panel data, where available, to analyze the  
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35 extent of consumption fluctuations over time as households experience income fluctuations  
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37 (Morduch, 2005; Kamanou and Morduch, 2005). Other studies examine the impact of various  
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39 forms of shocks on households' consumption (Ligon and Schechter, 2003; Carter et al., 2007),  
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41 or other aspects of household well-being, for instance, health (Dercon and Hoddinott, 2005).  
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43 While there are efforts to address data issues, empirical analyses of vulnerability remain  
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45 severely constrained by the paucity of panel data in many developing countries and by limited  
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47 information on the idiosyncratic and covariate shocks experienced by households (Günther  
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49 and Harttgen, 2009:1222-23). Chauduri et al. (2002) propose a method for estimating  
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51 vulnerability that can be applied to cross-sectional household surveys such as the 2003 Indian  
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3 rural household survey, thus avoiding the data problems mentioned. It has been adopted in a  
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5 number of vulnerability studies including Zhang and Wan (2006), Günther and Harttgen  
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7 (2009) and Imai, Wang and Kang (2010).<sup>4</sup> A discussion of this vulnerability estimation  
8  
9 method is presented in section 4.  
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### 12 13 14 15 **3. Microfinance Self-help Groups and Household Vulnerability** 16 17

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19 Very few studies have explored the effect of microfinance in terms of reducing vulnerability.  
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21 Evidence on Bangladeshi microfinance institutions conclude that microfinance access has led  
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23 to consumption smoothing or a reduction in the variance in consumption by member  
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25 households across time periods (Khandker, 1998; Morduch, 1999; Zaman 2000). Puhazhendi  
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27 and Badatya (2002) study finds that microfinance provides loans for both production and  
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29 consumption purposes, thereby allowing consumption smoothing and enabling households to  
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31 mitigate the effects of negative shocks.  
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38 Building on these studies, we argue in this paper that microfinance SHG participation can  
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40 help member households in the face of liquidity constraints and multitude of risks, thereby  
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42 reducing their vulnerability. For instance, SHG program provide loans to those members who  
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44 face liquidity constraints in meeting investment needs as well as unexpected consumption  
45  
46 expenses. These production and consumption loans help ease the members' productivity and  
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48 earnings and help their households coping with contingencies and idiosyncratic shocks. The  
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50 training of members provided by the SHG program also can enhance their entrepreneurship  
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52 skills as well as their ability to perceive and process new information, evaluate and adjust to  
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54 changes, thus increasing both their productivity and self-confidence.  
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3 In addition, SHGs can promote or strengthen social networks that provide mutual support by  
4 facilitating the pooling of savings, regular meetings, etc that help empower their members,  
5 especially women. Group meetings are often used to discuss communal issues leading to  
6 improved ability of member households to manage risk and deal with shocks. These *non-*  
7 *pecuniary effect* of SHGs can reduce the vulnerability of the members and by association, that  
8 of their households in ways that may not be adequately captured by change in household  
9 earnings alone.

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12 While SHGs may help rural households deal with vulnerability to idiosyncratic shocks the  
13 protection afforded by them in dealing with covariate shocks such as epidemics, flooding or  
14 declining crop prices is likely to be weak (Zimmerman and Carter, 2003; Morduch, 2004;  
15 Dercon, 2005).<sup>5</sup> An enabling economic environment and the presence of services and  
16 infrastructural support such as health centers and flood control systems that reduce exposure  
17 to these aggregate shocks can help enhance the effectiveness of SHGs in reducing household  
18 vulnerability,

## 41 **4. Data and Methodology**

### 42 **4.1 Data Description**

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45 The NABARD SHG-bank linkage program in India is one of the largest and fastest-growing  
46 microfinance programs in the developing world. Initiated in 1996, the SHG program has  
47 grown to finance 687,000 SHGs in 2006-07 as compared to 198,000 SHGs in 2001-02.  
48 According to NABARD (2006), about 44,000 branches of 547 banks and 4,896 NGOs  
49 participate in the SHG bank linkage program. These microfinance SHGs typically include ten  
50 to twenty (primarily female) members in the village. In the initial months, the group members  
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3 save and lend amongst themselves to build group financial discipline. Once the group  
4 demonstrates stability for six months, it receives loans of up to four times the amount it has  
5 saved. The bank then disburses the loan and the group decides how to manage the loan. As  
6 savings increase through the group's life, the group accesses a larger amount of loans. The  
7 SHGs are linked to banks in several ways: SHGs that are formed and financed by banks  
8 (model 1), those formed by NGOs but directly financed by banks (model 2), and those that are  
9 formed by the NGOs and financed by the banks through the NGOs (model 3).  
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22 The data used for the empirical analyses in this paper was collected in 2003 as part of a larger  
23 study that investigates the NABARD SHG-bank linkage program.<sup>6</sup> The sample survey was  
24 conducted in two representative districts of the following five states: Orissa, Andhra Pradesh,  
25 Tamil Nadu, Uttar Pradesh, and Maharashtra.<sup>7</sup> NABARD's choice to expand the SHG  
26 program occurs at the district level without any specific policy to target certain villages (Bali  
27 Swain and Varghese, 2009). Thus, within the states, the study selected to sample at the district  
28 level, which is the basic administrative unit, avoiding those districts with over and under  
29 exposure of SHGs. The sampling strategy involved random selection of SHG member-  
30 households in each district. The control group (non-SHGs) was chosen to reflect a comparable  
31 socio-economic group as the SHG respondents. These households were selected from villages  
32 that were similar to the SHG villages in terms of the level of economic development, socio-  
33 cultural factors and infrastructural facilities, but did not have a SHG program. After refining  
34 the data further and dropping those with missing values, we are left with a sample of 840  
35 households.  
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58 Table 1 shows characteristics of SHG and non-SHG members and their households. In  
59 general, SHG members are younger, have higher level of education, and have less non-land  
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3 wealth compared to non-SHG respondents. They also have higher food consumption per  
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5 capita per month and bigger landholding size compared to non-SHG households, although  
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7 there is large variation in land quality. SHG households live in villages that are closer to  
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9 public transport and primary health care centres but further away from banks, compared to  
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11 non-SHG households. Using a subjective indicator based on the survey response as to whether  
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13 or not their household experienced severe shortage of food and/or cash in the past three years,  
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15 we find that 39% of the SHG households have experienced economic difficulties, compared  
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17 to 27% of non-SHG households. The t-test results confirm the significant difference between  
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19 the SHG members and non-members in terms of size of landholdings and their access to  
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21 market infrastructure and services, as well as incidence of food and/or cash shortage in the  
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<Table 1 about here>

#### 33 34 35 36 **4.2 Propensity Score Matching Method** 37 38 39 40

41 The decision to participate in SHGs depends on the same attributes that determine the  
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43 vulnerability of the household. Self-selection bias could arise from the potentially  
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45 unobservable traits of the SHG members. For instance, higher entrepreneurship, ability to  
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47 recognize opportunity, and other critical aspects make the households more likely to  
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49 participate in the SHG program. However, the same characteristics could also affect their  
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51 vulnerability. A number of studies on microfinance have addressed the problem of selection,  
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53 reverse causality and other biases using different approaches.  
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3 To correct for selection bias created by program selection, we use the propensity score  
4 matching (PSM) method. This technique allows us to identify the program impact when a  
5 random experiment is not implemented, as long as there is counterfactual or control group. In  
6 contrast to other regression methods, the PSM does not depend on linearity and has a weaker  
7 assumption on the error term. The matching relies on the assumption of conditional  
8 independence of potential outcomes and treatment given observables. The data collection  
9 method meets the three conditions outlined in Heckman *et al.* (1997), thus allowing the use of  
10 the PSM method. First, the survey questionnaire is the same for participants and non-  
11 participants and therefore yields the same outcome measures. Second, both groups come  
12 from the same local environment or markets. Third, a rich set of observables for both outcome  
13 and participation variables are available for the performance of the PSM method.  
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31 As with any impact evaluation, the main problem with identifying SHG impact is that the  
32 outcome indicator for SHG member households with and without program is not observed  
33 because by definition, all the participants are SHG members in period 1. Since we only have  
34 information on the households once they participate in the program, there is need to identify a  
35 control group that allows us to infer what would have happened with the SHG participant  
36 household if the SHG program would not have been in place. The PSM uses the “Propensity  
37 Score” or the conditional probability of participation to identify a counterfactual group of non  
38 participants, given conditional independence.  
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53 The probability ( $P(X)$ ) of being selected is first determined by a logit equation and then this  
54 probability (the propensity score) is used to match the households.  $Y_1$  is the outcome indicator  
55 for the SHG program participants ( $T=1$ ), and  $Y_0$  is the outcome indicator for the SHG  
56 members ( $T=0$ ), then equation (1) denotes the mean impact:  
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$$\Delta = E[Y_1 | T = 1, P(X)] - E[Y_0 | T = 0, P(X)] \quad (1)$$

where the propensity score matching estimator is the mean difference in the outcomes over common support, weighted by the propensity score distribution of participants.

The literature proposes several propensity score matching methods to identify a comparison group.<sup>8</sup> Since the probability of two households being exactly matched is close to zero, distance measures are used to match households. Following Smith and Todd (2005), we first choose the neighbor to neighbor (NN) algorithm (with one person matching). This algorithm is the most straightforward and matches partners according to their propensity score. We further estimate the local linear regression (LLR) method (for bandwidths 1).<sup>9</sup> The LLR method uses the weighted average of nearly all individuals in the control group to construct the counterfactual outcome. Bootstrapped standard errors for the LLR procedures are used (Abadie and Imbens, 2007; Heckman et al., 1997).

### ***4.3 Estimating Poverty and Vulnerability***

We examine the poverty profile of the SHG and non-SHG households using standard measures of poverty such as the headcount ratio, poverty gap ratio and the squared poverty gap or Foster-Greer-Thorbecke (FGT). The head count ratio measures the proportion of population under the poverty line. The poverty gap ratio measures the depth of poverty and is the total amount that is needed to raise the poor from their present incomes to the poverty line as a proportion of the poverty line and averaged over the total population. The squared poverty gap or FGT index takes inequality among the poor into account and captures the severity of poverty.

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3 The poverty line used in our study is based on the official (consumption-based poverty) line  
4 for India, which assumes the minimum subsistence requirement of 2400 calories per capita  
5 per day for rural areas. The official poverty line estimate is derived from the household  
6 consumer expenditure data collected by National Sample Survey Organisation (NSSO) of the  
7 Ministry of Statistics and Programme Implementation, every fifth year. Since the poverty line  
8 estimate is drawn from the 61<sup>st</sup> round of the NSS which covers period July 2004 to June  
9 2005<sup>10</sup>, we adjust the official poverty line using the 2003 Consumer Price Index for  
10 agricultural workers in rural areas to correspond with the survey period. Hence our estimated  
11 2003 poverty line is Rs 356.3 per capita per month.  
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27 Next, we estimate the household's vulnerability using the Chauduri, Jayan and Suryahadi  
28 (2002) approach that allows the estimation of expected consumption and its variance with  
29 cross-section data. The Chauduri et al. approach is widely used in several studies on  
30 vulnerability (Jha and Dang, 2009; Zhang and Wan, 2006; Imai et al., 2010) and is considered  
31 to be one of the best estimators (Ligon and Schechter, 2004).<sup>11</sup> It is based on the notion of  
32 vulnerability as the probability of being poor and implies accounting for the expected (mean)  
33 consumption, as well as the volatility (variance) of its future consumption stream. The  
34 stochastic process generating the consumption of the household is dependent on the  
35 household characteristics and the error term (with mean zero). It captures the idiosyncratic  
36 shocks to consumption that are identically and independently distributed over time for each  
37 household. Hence, any unobservable sources of persistent or serially correlated shocks or  
38 unobserved household specific effects over time on household consumption are ruled out. It  
39 also assumes economic stability thereby ruling out the possibility of aggregate shocks. Thus  
40 the future consumption shocks are assumed to be idiosyncratic in nature. This does not mean  
41 however, that they are identically distributed across households. Furthermore, we assume that  
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3 the variance of the idiosyncratic factors (shocks) depend upon observable household  
4 characteristics.  
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10 Following Chauduri et al (2002) approach, we assume that the vulnerability level of a  
11 household  $h$  at time  $t$  is defined as the probability that the household finds itself to be  
12 consumption poor in period  $t + 1$ . The household's consumption level depends on several  
13 factors such as wealth, current income, expectation of future income (i.e. lifetime prospects),  
14 the uncertainty it faces regarding its future income and its ability to smooth consumption in  
15 the face of various income shocks. Each of these, in turn, depend on a number of household  
16 characteristics, both observed and unobserved, the socio-economic environment in which the  
17 household is situated, and the shocks that contribute to differential welfare outcomes for  
18 households that are otherwise observationally equivalent. Hence, the household's  
19 vulnerability level in terms of its future food consumption can be expressed as a reduced form  
20 for consumption determined by a set of variables  $X_{ht}$ :  
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$$\ln c_{ht} = \beta_0 + X_{ht} \beta_1 + \mu_{ht} \quad (2)$$

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43 where  $\ln c_{ht}$  represents log of consumption per capita on adult equivalence scale,  $X_{ht}$   
44 represents selected household and community level characteristics, and  $\mu_{ht}$  is the unexplained  
45 part of household consumption. Since the impact of shocks on household consumption is  
46 correlated with the observed characteristics, the variance of the unexplained part of  
47 consumption  $\mu_{ht}$  is:  
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$$\sigma_h^2 = \Phi_0 + \Phi_1 X_{ht} + \omega_{ht} \quad (3)$$

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3 which implies that the variance of the error term is not equal across households and depends  
4 upon  $X_{ht}$ . The latter include respondent's educational attainment, household composition,  
5 number of workers in the household, and household wealth proxy. We also take into account  
6 the environment characteristics such as access to paved roads, markets, health care services,  
7 and public transportation. Given data limitations, we cannot identify the particular stochastic  
8 process generating  $\beta$ . The expected mean and variance per capita household food  
9 consumption are estimated using a simple functional form by Amemiya's (1977) three-step  
10 feasible generalized least squares (FGLS).<sup>12</sup> Using the obtained  $\beta_1$  and  $\Phi_1$  estimates, we  
11 estimate the expected log consumption and the variance of log consumption for each  
12 household. These serve as vulnerability estimates.

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30 To facilitate comparison of the vulnerability distribution among SHG and non-SHG  
31 households, we estimate additional measures using different thresholds in order to examine  
32 the sensitivity of our results to the choice of vulnerability threshold. The relative vulnerability  
33 threshold uses the observed poverty rate in the population, which is approximately equal to  
34 the mean vulnerability level within a group in the absence of aggregate shocks (Chauduri et  
35 al., 2002). Thus, vulnerability levels above the observed poverty rate threshold imply that the  
36 household's risk of poverty is greater than the average risk in the population, thus making it  
37 more vulnerable. We use the official rural poverty rate by the Planning Commission of India  
38 as the first vulnerability threshold.<sup>13</sup>

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54 Another vulnerability threshold is 0.50. Households with vulnerability levels between  
55 observed poverty rates and 0.50 threshold are termed *relatively vulnerable* whereas those  
56 above 0.50 are considered *highly vulnerable*. Finally, the vulnerability to poverty ratio,  
57 measures the fraction of the vulnerable population to the fraction that is poor. The higher the  
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3 vulnerability to poverty ratio the more spread is the distribution of vulnerability. Whereas a  
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5 lower vulnerability to poverty ratio implies greater concentration of vulnerability amongst a  
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7 few households. Admittedly, there is some arbitrariness involved in the selection of the  
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9 vulnerability thresholds so a comparison of the vulnerability estimates using additional  
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11 vulnerability thresholds shows the sensitivity of the results to the choice of vulnerability  
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13 threshold.  
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## 20 **5. Empirical Analysis**

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24 This section presents the logistic and the propensity score results of matching. This is  
25  
26 followed by a discussion on the poverty and the estimated vulnerability measures of SHG and  
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28 non-SHG member households. We then present the estimated average treatment on treated  
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30 (ATT) effect of SHG participation using different matching algorithms that takes potential  
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32 selection bias into account. The robustness of our results are then checked for sensitivity to  
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34 unobservables.  
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### 41 **5.1 Propensity Score Matching**

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45 We correct for potential selection bias using PSM method by first estimating a parsimonious  
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47 logistic equation in order to determine the probability of participating in the SHG program.<sup>14</sup>  
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49 The variables that likely affect both the participation in SHG and the outcome variable (real  
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51 food expenditure per capita per month) were chosen and these include age, age squared, sex,  
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53 education dummies, lack of cash or food 3 years ago, owned land 3 years ago, distance from  
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55 bank, health care centre, marketplace, and paved road.<sup>15</sup> We obtained very similar results with  
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57 both neighbour to neighbour algorithm (with one person matching) and log linear regression  
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3 method (for bandwidths 1). Appendix 1 shows the propensity score estimation using logistic  
4 regression. It indicates that landholding size in 2000, incidence of money or food shortage (in  
5 2000), and distance from the bank and market affect the probability of participating in the  
6 SHG. Other variables such as age, gender and education level of the respondent do not  
7 significantly explain SHG participation.  
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17 Using the derived propensity scores, we drop those SHG respondents with probabilities that  
18 cannot be matched to the propensity scores of the control group, leaving us with a sample of  
19 742 households comprised of 691 SHG and 51 non-SHG (control group) households. Of the  
20 691 SHG households, 532 have been members for more than one year (referred to as mature  
21 SHG members) and 159 belong to newly-formed groups. Only the households on the common  
22 support are retained to assure comparability. Prior to matching, the estimated mean propensity  
23 scores (standard error) for SHG members and non-SHG member were 0.94 (0.05) and 0.89  
24 (0.06) respectively. Appendix 2 provides the histograms of the estimated propensity scores  
25 for the two groups. After the matching, there was a negligible difference in the mean  
26 propensity scores of the two groups (0.93 (0.04) for SHG members and 0.89 (0.06) for non-  
27 SHG members).  
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## 46 **5.2 Poverty and Vulnerability Profile for SHG and non SHG members**

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50 We construct a poverty profile of the SHGs (treatment group) and the non-SHG member  
51 (control group) in 2003 using standard measures such as the headcount index, poverty gap  
52 index and the squared poverty gap index.<sup>16</sup> Table 2 presents the poverty profile of the SHG  
53 member and non-member households using standard poverty measures.<sup>17</sup> Our results show  
54 that a higher proportion of the SHG members are poor (72.5 per cent as compared to 60.8 per  
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3 cent for the non-members) although the depth of poverty is about the same between SHG and  
4 non-SHG households. Their aggregate poverty gap per household is Rs.123 compared to Rs  
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6 118 among non-SHGs. The FGT index shows that there is slightly greater inequality among  
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8 the non-SHG poor (0.24) compared to the SHG poor (0.22).  
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15 <Table 2 about here>  
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20 Following Chauduri et al. (2002) the vulnerability estimates are obtained from the FGLS  
21 estimates and are presented in Appendix 3. <sup>18</sup> The mean vulnerability level within the SHG  
22 member-household group is much lower (0.45) and statistically significant as compared to the  
23 SHG non-members (0.62). This implies that participation in SHGs may reduce the  
24 vulnerability of the households.  
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33 We also examine the mean vulnerability and sensitivity of the vulnerability estimate to the  
34 choice of a threshold. We use three different vulnerability thresholds in our study namely: a)  
35 the observed poverty rate; b) the vulnerability threshold of lying above the observed poverty  
36 rate but with a 50 percent probability of falling into poverty at least once in the next year; and  
37  
38 c) the highly vulnerable lying above the vulnerability threshold of 0.5 for a one year time  
39 period. We also report the ratio of the proportion of households that are vulnerable to the  
40 proportion that are poor. This is an indication of how dispersed vulnerability is in the  
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42 population.  
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55 The fraction of the population which is vulnerable with respect to these three thresholds is  
56 given in Table 2. Even though a higher proportion of SHG members are poor, they are  
57 relatively less vulnerable (0.55) as compared to the non-SHG (0.72). Not only are the non-  
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3 SHG members more vulnerable, a larger proportion of them (0.69) are highly vulnerable. The  
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5 non-members also have a higher vulnerability to poverty ratio (1.18) with a greater dispersion  
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7 in incidence of vulnerability. We further examine the subset of SHG participants that have  
8  
9 been members for more than one year. Their poverty and vulnerability profile is very similar  
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11 to that of the SHG members (see Appendix 4).  
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17 The above results indicate that there is a large proportion of currently poor SHG members,  
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19 whose vulnerability level is low enough for them to be classified as non-vulnerable. This  
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21 reflects the stochastic nature of the relationship between poverty and vulnerability. While  
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23 poverty and vulnerability are related concepts, the characteristics of those observed to be poor  
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25 at any given point in time may differ from the characteristics of those who are vulnerable to  
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27 poverty.  
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### 34 **5.3 Impact on Vulnerability Controlling for Selection Bias**

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39 We now estimate the impact on our outcome variables taking the selection bias from  
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41 participation into account. Heckman et al. (1997) suggest that in small samples the choice of  
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43 the matching algorithm can be important, due to trade-offs between bias and variances. Thus,  
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45 Caliendo and Kopeinig (2008) suggest that multiple algorithms should be tried and if they  
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47 give similar results, the choice may be unimportant.  
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53 Using two different algorithms for propensity score matching to identify the comparison  
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55 group, we estimate the ATT. Nearest Neighbour matching algorithm (NN) is the more  
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57 intuitive of the two as it matches each treated observation to a control observation with the  
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59 closest propensity score. We also employ the local linear regression (LLR) algorithm one to  
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3 one person matching (bandwidth 1), which is a generalised version of kernel matching that  
4 allows faster convergence at the boundary points.<sup>19</sup> Table 3 presents the Average Treatment  
5 on Treated estimates (ATT) of SHG participation impact on vulnerability and average food  
6 expenditure per capita per month.  
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20 The magnitude of the ATT estimates in Table 3, measures the impact of SHG participation on  
21 the outcome variables (vulnerability and food expenditure), controlling for the selection bias.  
22 Table 3, column 1 shows that the ATT point estimates (both NN and LLR) are positive but  
23 statistically insignificant for vulnerability. This indicates that after accounting for selection  
24 bias the SHG members are neither more nor less vulnerable as compared to the non  
25 members.<sup>20</sup> However, the SHG participants that have been members for more than a year,  
26 show a significantly lower level of vulnerability. This suggests that the impact of  
27 microfinance on vulnerability takes longer time. . By design, the SHG bank linkage program  
28 provides credit to those groups that have demonstrated financial maturity and stability during  
29 the first six months of their existence. Thus, the more mature (older than one year) groups are  
30 credit linked and have the possibility to use microfinance for reducing vulnerability whereas  
31 the newly formed SHGs are not. SHG participation on the other hand does lead to an increase  
32 in its average food expenditure per capita per month compared to that of non-SHG's using the  
33 LLR algorithm method (Table 3, column 2). A likely reason for this might be due to the  
34 provisioning of SHG loans that may be used for any purpose (including consumption) and  
35 thus helps the households cope with economic shocks. Taking the subset of the more mature  
36 SHGs however, the results do not show any significant increase in average food expenditure.  
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60 Our results show that even though the current poverty status of SHG member households has

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3 a very high proportion of poor with a higher aggregate poverty gap, their propensity to  
4 become poor in the next period (vulnerability) is not higher. The more mature SHG  
5 participants, however, have significantly lower level of vulnerability.  
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## 10 11 12 13 **6. Sensitivity Analysis - Robustness of Results** 14

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17 The propensity score matching hinges on the conditional independence or unconfoundedness  
18 assumption (CIA) and unobserved variables that affect the participation and the outcome  
19 variable simultaneously that may lead to a hidden bias due to which the matching estimators  
20 may not be robust. It is not possible to directly reject the unconfoundedness assumption  
21 however. Heckman and Hotz (1989) and Rosenbaum (1987) have developed indirect ways of  
22 assessing this assumption. These methods rely on estimating a causal effect that is known to  
23 be equal to zero. If the test suggests that this causal effect differs from zero, the  
24 unconfoundedness assumption is considered less plausible (Imbens, 2004).  
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38 Building on Rosenbaum and Rubin (1983) and Rosenbaum (1987), Ichino, Mealli and  
39 Nannicini (2008) propose a sensitivity analysis that we adopt in this paper. They suggest that  
40 if the CIA is not satisfied given observables but is satisfied if one could observe an additional  
41 binary variable (confounder), then this potential confounder could be simulated in the data  
42 and used as an additional covariate in combination with the preferred matching estimator. The  
43 comparison of the estimates obtained with and without matching on the simulated confounder  
44 shows to what extent the baseline results are robust to specific sources of failure of the CIA,  
45 since the distribution of the simulated variable can be constructed to capture different  
46 hypotheses on the nature of potential confounding factors.  
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3 To check the robustness of our ATT estimates, we use two covariates to simulate the  
4 confounder namely: young (respondents under the age of 26 years) and illiterate (with no  
5 education). These covariates are selected in order to capture the effect of ‘unobservables’ like  
6 ability, entrepreneurial skills, experience and risk aversion etc., which may have an impact on  
7 the member participation in the SHG program and on the vulnerability of the household. If the  
8 ATT estimate change dramatically with respect to the confounders, then it would imply that  
9 our results are not robust. We employ the Kernel matching algorithm with between-  
10 imputation standard error, in order to use only the variability of the simulated ATT across  
11 iterations. Since our outcome variable is continuous, the confounder is simulated on the basis  
12 of the binary transformation of the outcome along the 25<sup>th</sup> centile. The results of these two  
13 confounders<sup>21</sup> are presented in Table 4. For both the ‘young’ and ‘no education’ confounders  
14 the simulated ATT estimates are very close to the baseline estimate. The outcome and  
15 selection effect on vulnerability is positive but not very large. The results indicate a  
16 robustness of the matching estimates.  
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<Table 4 about here>

43 We further test the robustness of our results using Rosenbaum’s (2002) bounding approach  
44 and find our results to be robust (see Appendix 5, with discussion).  
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## 50 **7. Concluding Remarks**

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55 This paper explores an important dimension of household welfare that conventional measures  
56 of poverty do not address, namely vulnerability. We examine the likely effect of Self-Help  
57 microfinance groups (SHG) on the vulnerability of participating member households using an  
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3 Indian household sample survey data from 2003. We argue that a household's ability to  
4 mitigate risk and cope with shocks is enhanced through SHG participation by increasing  
5 household earnings through provision of microfinance and training, aiding the household in  
6 the face of shocks by providing consumption loans, and enhancing their resilience by  
7 strengthening social support and improving women's empowerment.  
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11 We use propensity score matching to extricate the potential selection bias that may arise due  
12 to unobservable attributes. Additionally, we empirically examine the current poverty status of  
13 households in SHG and non-SHG groups using several poverty measures and then make  
14 inferences about whether or not these households are currently vulnerable to future poverty  
15 using the Chauduri et al. approach. After matching the treated and comparison groups on the  
16 basis of their propensity scores, we estimate the average treatment on treated effect using  
17 nearest neighbour matching algorithm and local linear regression. The robustness is checked  
18 with help of sensitivity analysis and Rosenbaum bounds. Our main empirical results show that  
19 after we account for the selection bias, even though SHG-member households are found to be  
20 poorer than the non-SHG member (control group) households, they are not more vulnerable.  
21 Vulnerability is significantly lower for the more mature households as compared to the non-  
22 SHG members. These results are found to be robust using the sensitivity analysis and  
23 Rosenbaum bounds method.  
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50 The SHG bank linkage program is a joint liability microfinance program where the loan may  
51 be used for any purpose, be it production or consumption. Microfinance in this case provides  
52 an additional resource for consumption smoothing thus reducing the variability in food  
53 consumption levels and hence vulnerability. Finally, microfinance SHG can strengthen  
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3 mutual support networks that help reduce vulnerability of members and that of their  
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5 households in ways that may not be adequately captured by the change in household earnings.  
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## TABLES

Table 1

*Selected Characteristics of Survey Respondents and their Households  
(standard deviation in parentheses)*

	<i>All</i>	<i>SHG members</i>	<i>Non-SHG<sup>††</sup></i>
N	840	789	51
Average Real food expenditure per capita per month	307 (442)	308 (453)	282 (194)
Average Age of Respondent	35 (8.41)	35 (8.44)	36 (8.08)
Proportion with some (in %)			
Primary education	18.	18	24 (0.43)
Secondary education	17	18	12
Post-Secondary education	3	3	2
Average number of children	1.5 (1.27)	1.5 (1.27)	1.4 (1.25)
Dependency ratio	0.66 (0.22)	0.66 (0.22)	0.62 (0.23)
Average number of workers in the household	2.48 (1.24)	2.46 (1.23)	2.70 (1.40)
Average number of workers engaged in primary activity	2.49 (1.37)	2.48 (1.37)	2.55 (1.30)
Mean size of owned land in 2000(in acres)	0.85 (1.43)	0.87 (1.45)	0.48** (1.12)
Mean value of non-land wealth years ago (in Rupees.) <sup>†</sup>	64,691 (90197)	63,708 (86775)	79,891 (132625)
Distance to Bank (in kms.)	7.33 (6.87)	7.48 (7.02)	4.96*** (3.16)
Distance to Health Care	3.55 (2.84)	3.46 (2.78)	4.95*** (3.30)
Distance to Market	5.39 (4.02)	5.38 (4.07)	5.46 (3.16)
Distance to Paved Road	3.06 (3.32)	3.03 (3.33)	3.59 (3.04)
Distance to Bus Stop	3.75 (3.55)	3.69 (3.59)	4.71** (2.76)
Lack of cash or food in 2000	0.38 (0.49)	0.39(0.49)	0.27* (0.45)

<sup>†</sup> Calculated with 2000 as the base year. <sup>††</sup> T- test results for equality of means of SHG members and non-SHG members are indicated by \*\*\* if significant at 1% level, \*\* if significant at 5% level, and \* if significant at 10% level.

Table 2

*Poverty and Vulnerability Estimates for SHG members and non members†*  
(Standard deviation in parentheses.)

	<i>SHG members</i>	<i>Non-SHG members††</i>
<i>All Households</i>		
<i>N</i>	691	51
<i>Poverty Profile for SHG members and non-members</i>		
Headcount ratio (per cent)	72.5	60.8
Aggregate poverty gap per observation	123	118
Poverty gap ratio (per cent)	35	34
Foster-Greer-Thorbecke (sqd poverty gap)	0.22	0.24
<i>Vulnerability Profile for SHG members and non-members</i>		
Mean	0.45 (0.39)	0.62*** (0.39)
Fraction vulnerability	0.55	0.72**
Fraction relatively vulnerable	0.08	0.03
Fraction highly vulnerable	0.47	0.69**
Vulnerability to poverty ratio	0.75	1.18

†The vulnerability estimates are based on the Chauduri et al (2002) method. †† T-test results for equality of means and proportion. \*\*\*, \*\* and \* indicate significance at 10%, 5% and 1% levels respectively.

Table 3

*Average Treatment on Treated Estimates of SHG Participation Impact on Vulnerability and Average Food Expenditure per Capita per Month*

<i>Matching Algorithm</i>	<i>(1) Vulnerability</i>	<i>(2) Av. food exp per capita per month</i>
<b>All SHG members</b>		
1 NN	0.09 (1.19)	29.04 (0.61)
LLR (bw 1)	0.11 (1.54)	68.35* (1.89)
<b>Mature SHG members</b>		
1NN	-0.15** (0.73)	39.33 (42.31)
LLR (bw1)	-0.11* (0.61)	66.80 (42.55)

*Notes:* \*\* Significant at the 5 % level. \* Significant at the 10 % level. NN = neighbor to neighbor, t-stats in parentheses. LLR= local linear regression, p-values in parentheses standard errors created by bootstrap replications of 200. Covariates of regression same as in Appendix 1.

Table 4

*Simulation-Based Sensitivity Analysis for Matching Estimators†  
Average Treatment on Treated Effect (ATT) Estimation on Vulnerability with Simulated  
Confounder General Multiple-Imputation Standard Errors ††*

<i>Confounder</i>	<i>(1) ATT</i>	<i>(2) Standard Error</i>	<i>(3) Outcome effect</i>	<i>(4) Selection effect</i>
<b>For All SHGs</b>				
Age	0.13	0.01	9.01	3.9
Education	0.14	0.01	5.2	1.1
<b>For Mature SHGs</b>				
Education	-0.17	0.008	6.830	1.009

† Based on the sensitivity analysis with kernel matching algorithm with between-imputation standard error. The binary transformation of the outcome is along the 25 centile. †† Age variable (=1 if age is less than 26 years; and = 0 otherwise) and education (=1 if no education; and zero otherwise).



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<sup>1</sup> The differences in the empirical findings arise from varying measures of poverty, different country contexts and types of microfinance organizations being analyzed, use of different theoretical models, survey designs and econometric techniques, and/or different time periods covered by the studies.

<sup>2</sup> See Glewwe and Hall, 1998; Calvo and Dercon, 2005; Carter and Ikegami, 2007; Ligon and Schechter, 2002; Dercon and Krishnan, 2000; Dercon, 2005.

<sup>3</sup> This concept is based on the notion that “future is uncertain, and the possibility of failing to reach some standard of minimal achievement in any well-being dimension is at least a disturbing background noise for some, and an ever-present, oppressing source of stress and dismay for many others” (Calvo, 2008: 1011).

<sup>4</sup> Chauduri et al (2002) measure of vulnerability is an unpublished working paper that has been adopted in several studies. Zhang and Wan (2006) explores the effect of livelihood diversification and education on household vulnerability in rural Chinese households. Günther and Harttgen (2009) examine the impact of idiosyncratic and covariate shocks in rural and urban households in Madagascar while the study by Imai, Wang and Kang (2010) analyses the impact of taxation policies on household welfare in China. We would like to thank the reviewer of this paper for bringing some of these studies to our attention.

<sup>5</sup> Rural livelihoods in developing countries like India often exhibit high correlations between risks faced by households in the same village or area. Hence, when farm prices decline, or there is a drought or flood in the area, all households are adversely affected simultaneously. Idiosyncratic shocks are, by definition, uncorrelated across households in a given community and therefore can be mutually insured within communities.

<sup>6</sup> The process involved discussion with statisticians, economists and practitioners at the stage of sampling design, preparing pre-coded questionnaires, translation and pilot testing with at least 20 households in each of the 5 states (100 households in total). The questionnaires were then revised, printed and the data collected by local surveyors that were trained and supervised by the supervisors. The standard checks were applied both on the field and during the data punching process.

<sup>7</sup> These districts (in parentheses) are Orissa (Koraput and Rayagada), Andhra Pradesh (Medak and Rangareddy), Tamil Nadu (Dharmapuri and Villupuram), Uttar Pradesh (Allahabad and Rae Bareli), and Maharashtra (Gadchiroli and Chandrapur).

<sup>8</sup> See Townsend, 1995; Dercon, 2005; Zimmerman and Carter, 2003; and Morduch, 2004.

<sup>9</sup> Bandwidths are smoothing parameters, which control the degree of smoothing for fitting the local linear regression.

<sup>10</sup> See Poverty Estimates for 2004-05, Government of India, Press Information Bureau, March 2007.

<sup>11</sup> In a comparative study of various vulnerability estimation strategies, Ligon and Schechter (2004) find that when the environment is stationary and consumption expenditures are measured without error, then the estimator proposed by Chauduri et al is the best estimator of vulnerability.

<sup>12</sup> For details on the statistical estimation refer to Chauduri et al., 2002.

<sup>13</sup> Planning Commission estimates, as accessed on 22 September 2010  
<http://www.planningcommission.gov.in/data/datatable/Data0910/tab%2019.pdf>

<sup>14</sup> Using saturated logit models as opposed to simple ones is debatable, as the purpose of logit equation is not only to predict SHG participation (as in selection models) but also for covariate balancing.

<sup>15</sup> The variables were chosen through ‘hit and miss’ method while keeping in mind the balance.

<sup>16</sup> The poverty gap is the average (over all individuals) gap between poor people’s living standards and the poverty line. It indicates the average extent to which individuals fall below the poverty line (if they do). It thus measures how much would have to be transferred to the poor to bring their income (or consumption) up to the poverty line. The poverty gap however does not capture the differences in the severity of poverty amongst the

poor and ignores “inequality among the poor”. To account for the inequality amongst the poor we calculate the squared poverty gap index which is defined as the average of the square relative poverty gap of the poor. The squared poverty gap index (Foster-Greer-Thorbecke Index) is a weighted sum of poverty gaps (as a proportion of the poverty line), where the weights are the proportionate poverty gaps themselves.

$$P_{\alpha} = \frac{1}{n} \sum_{i=1}^q \left( \frac{z - y_i}{z} \right)^{\alpha}$$
 The measures are defined for  $\alpha \geq 0$ , where  $\alpha$  is a measure of the sensitivity of the index to poverty. When  $\alpha=0$ , we have the headcount index (the proportion of the population for whom income (or other measures of living standard) is less than the poverty line),  $\alpha=1$  is the poverty gap index and  $\alpha=2$  is the squared poverty gap index.

<sup>17</sup> The poverty and vulnerability profile for the SHG and non-SHG member households is presented here for the sample on common support. Imposing common support condition in the estimation of the propensity score may improve the quality of the matches used to estimate ATT (Ichino and Becker, 2002).

<sup>18</sup> The three step feasible generalised least squares (FGLS) results are presented in Appendix 3. The results show that SHG membership leads to a statistically significant increment in the consumption. The coefficients of the control variables have the expected signs.

<sup>19</sup> We also employed NN (bandwidth 10) and LLR (bandwidth 4), both of which gave very similar results to those in table 3.

<sup>20</sup> The results in table 2 show that the SHG members have lower vulnerability as compared to the non-SHG; do not account for the selection bias and are hence biased.

<sup>21</sup> Both these confounders are “dangerous” confounders, since both the outcome and the selection effect are positive.

## APPENDIX

## Appendix 1

*Logistic regression for Participation in SHGs†*

	<i>Coefficient</i>	<i>T-Statistic††</i>
Age of Respondent	-0.13	1.05
Age square	0.001	0.85
Sex	-0.47	0.62
Primary education	-0.47	1.25
Secondary education	0.10	0.21
Post-Secondary education	0.27	0.25
Distance Bank (kms.)	0.18	2.52***
Distance Health Care	-0.26	3.93***
Distance Market	-0.001	0.02
Distance weekly market	-0.02	0.23
Land owned in 2000 (in acres)	0.31	2.02**
Lack of cash or food in 2000	0.69	2.05**

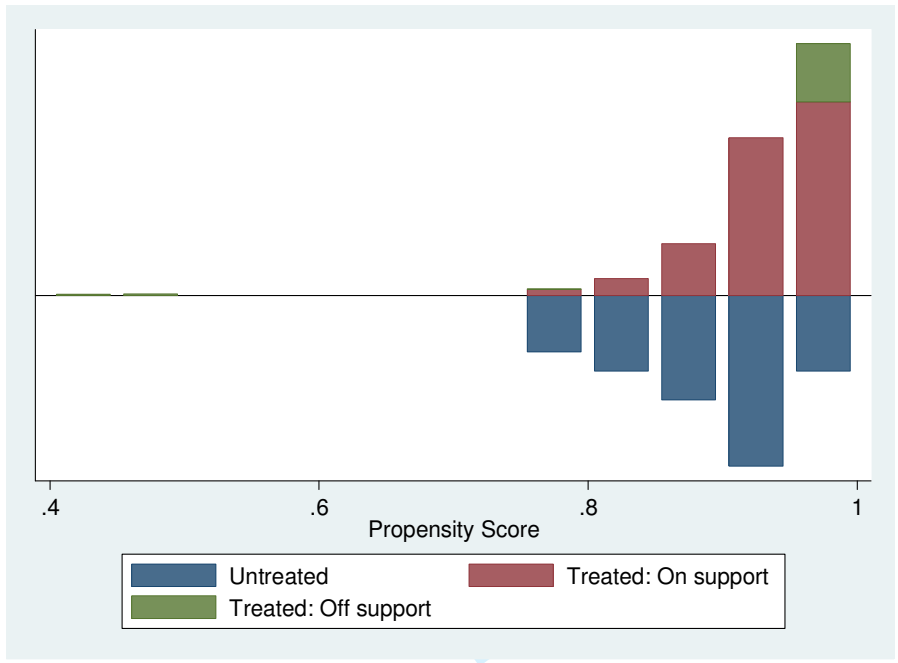
†Logistics Regression results with Nearest-Neighbor Matching algorithm.

†† Absolute t-ratios reported.\*\*\* Significant at 1% level. \*\* Significant at 5% level. \* Significant at 10% level.

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Appendix 2

Histograms of estimated propensity scores



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## Appendix 3

*FGLS regression of per capita consumption*

	<i>Coefficient</i>	<i>T-Statistic</i> <sup>†</sup>
Member	0.82	4.25***
Land owned in 2000 (in acres)	0.04	1.92**
Age of Respondent	0.34	18.64***
Age square	-0.004	16.93***
Primary education	-0.37	2.23**
Secondary education	0.34	3.42***
Post-Secondary education	0.59	3.59***
Number of children	0.02	0.72
Dependency ratio	-1.61	4.84***
Number of workers in the household	-0.27	5.46***
Number of workers engaged in primary activity	0.02	0.57
Real value of non-land wealth in 2000(Rs.)*	9.4e-8	0.14
Distance Bank (kms.)	-0.06	3.23***
Distance Health Care	-0.01	0.52
Distance Market	0.01	0.28
Distance Paved Road	-0.10	2.63***
Distance Bus Stop	0.16	7.26***
Lack of cash or food in 2000	0.59	8.29***

<sup>†</sup> Absolute t-ratios reported.\*\*\* Significant at 1% level. \*\* Significant at 5% level. \* Significant at 10% level.

## Appendix 4

*Poverty and Vulnerability estimates by maturity of SHGs†*

	<i>SHGs older than 1 year</i>	<i>Newly formed SHGs</i>
<i>Poverty Profile</i>		
<i>N</i>	523	168
Headcount ratio (per cent)	72.7	71.7
Aggregate poverty gap per observation	126	113
Poverty gap ratio (per cent)	36	32
Foster-Greer-Thorbecke (squared poverty gap)	0.23	0.18
<i>Vulnerability</i>		
Mean	0.45 (0.40)	0.45 (0.40)
Fraction vulnerability	0.56	0.57
Fraction relatively vulnerable	0.09	0.08
Fraction highly vulnerable	0.47	0.49
Vulnerability to poverty ratio	1.30	1.26

†The vulnerability estimates are based on the Chauduri et al (2002) method, standard deviation in parentheses.

## Appendix 5

*Rosenbaum bounds – Vulnerability †*

$\Gamma$	<i>Hodges-Lehmann point estimates</i>				
	(1) <i>p-critical</i>	(2) <i>Upper bound</i>	(3) <i>lower bound</i>	(4) $CI_{max}$	(5) $CI_{min}$
1	7.3 (10 <sup>-13</sup> )	0.15	0.15	0.12	0.17
1.1	1.1 (10 <sup>-16</sup> )	0.14	0.16	0.09	0.19
1.2	0	0.12	0.17	0.07	0.20
1.3	0	0.10	0.19	0.05	0.21
1.4	0	0.08	0.19	0.04	0.22
1.5	0	0.06	0.21	0.02	0.23
1.6	0	0.05	0.21	-0.003	0.23
1.7	0	0.04	0.22	-0.03	0.24
1.8	0	0.02	0.23	-0.07	0.24
1.9	0	0.0004	0.23	-0.10	0.24
2	0	-0.02	0.23	-0.14	0.25

*Notes:* †  $\Gamma$ : log odds of differential assignment due to unobserved factors

*p-critical*: lower bound of significance level

$CI_{max}$ : upper bound confidence interval ( $\alpha=0.95$ )

$CI_{min}$ : lower bound confidence interval ( $\alpha=0.95$ )

We also check the robustness of our results using Rosenbaum's (2002) bounding approach in order to determine how strongly an unmeasured variable influences the selection process to undermine the implication of the matching analysis.<sup>1</sup> Rosenbaum bounds calculate the bounds for average treatment effects on the treated in the presence of unobserved heterogeneity (hidden bias) between treatment and control cases. It takes the difference in the response variable between treatment and control cases and then calculates Wilcoxon sign rank tests that give upper and lower bound estimates of significance levels at given levels of hidden bias.

<sup>1</sup> Instead of testing the unconfoundedness assumptions, the Rosenbaum's bounds provide evidence on the degree to which any significance result hinge on this assumption.



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3 The Hodges-Lehmann point estimates and confidence intervals for the average treatment  
4 effect on the treated are also provided.  
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9 Table 3 presents the results from the Rosenbaum bounds analysis for vulnerability using  
10 different Hodges-Lehmann point estimates. The analysis is conducted on the matching  
11 procedure using local linear regression (bandwidth 1) with a random draw and bootstrapped  
12 standard errors. The estimates illustrate the sensitivity of the results to potential hidden bias.  
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14 Our assumption about the potential endogeneity in assignment to treatment is given by  $\Gamma$   
15 which reflects the odds of participation in treatment. Matched units have the same probability  
16 of participation only if  $\Gamma=1$ . If the odds of participation differ from 1 then it must be due to  
17 hidden bias.  
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25 The Hodges-Lehmann point estimates reflect the uncertainty in the estimated Average  
26 Treatment Effect on the Treated at increasing levels of assumed hidden bias. At  $\Gamma = 1$ , there is  
27 no hidden bias and the estimates are equal (upperbound=lowerbound= 0.15). The confidence  
28 interval includes zero only when we cross  $\Gamma=1.5$ . This means that the unobserved effect  
29 would have to increase the odds of participation in SHG by more than 1.5 before one changes  
30 the conclusions about the effect of SHG participation on participants. Thus, the postulated  
31 effects of the SHG participation on mean vulnerability of the households due to unobservables  
32 would have to be quite large for us to doubt our results.  
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**Revisions based on Referee Comments for Manuscript FJDS-2009-Dec-0010.R1**

We are indebted to the editor and the referee for their support and suggestions that have improved our paper substantially. The manuscript has now been revised based on the 'few minor comments' by the referee and suggestions by the editor.

1. *Editor's suggestions:*

*Make the paper as succinct and short (maximum 9,000 words and as few tables as feasible). Place the detailed material in a separate Appendix file that will be available online. Ensure that the revised version follows our house style guide.*

Revision:

The paper has been revised as suggested and has been considerably shortened to 8691 words and contains a total of 4 tables. To the existing Appendix, we have added Table 5 (from the text) and the related discussion on Rosenbaum's bounding approach. These have been placed in a separate Appendix file for the interested reader. The manuscript has been revised once again to ensure that it follows the house style guide of JDS.

2. *Referee comment 1 was on revising the first sentence of the abstract; there was a reference missing (comment 2); and the Referee wanted us to add more references that used Chauduri et al. vulnerability measure (comment 3).*

Revision:

We have revised the abstract and added the missing reference. In addition to the two references suggested by the referee, we have also added Jha and Dang (2009). We have also added Ligon and Schechter (2004). In a comparative study of vulnerability estimation strategies, they find that when the environment is stationary and consumption expenditures are measured without error, then the estimator proposed by Chauduri is the best estimator of vulnerability.

3. *Referee comment 4*

*The referee wanted us to add another paragraph to the explanation of Table 3. There were three specific comments that the referee had:*

- (1) Explain in one or two sentences the meaning of the numbers in Table 3.*
- (2) The referee points that the LLR vulnerability SHG member is larger by 0.11 points at significance level 15%*
- (3) The referee suggests inserting more discussions on explaining how the results in table 2 and table 3 are consistent*

Revision

We have explained (1) in the first two sentences on page 15.

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Points (2) and (3) are related, as the confusion stems from the lack of explanation on Table 3. As suggested by the referee we have added another paragraph to explain table 3. We have also added the vulnerability results for the SHGs that are more than one year old and clearly show a reduction in vulnerability as compared to the non-members (Table 3). We have further clarified on page 18 and 19 that Heckman et al. (1997) suggest that in small samples the choice of the matching algorithm can be important, due to trade-offs between bias and variances. Thus, Caliendo and Kopeinig (2008) suggest that multiple algorithms should be used. This is why we have used NN (bandwidth 1 and 10) and LLR (bandwidth 1 and 4). LLR (bandwidth 1) result might be close to significance (for vulnerability) but it is *not* significant. ATT point estimates using other algorithm (bandwidths) confirm the non-significance. Furthermore, our sensitivity analyses (using two separate methods) confirm the reliability of our results. We have explicitly stated this under section 5.3 (pp. 20-21).

#### 4. Referee comment 5

*The referee wanted the statement, "Our main empirical results show ..." to be re-written.*

Revision

We have revised this statement, as suggested (page 22, second paragraph, last five lines).

The manuscript has been revised in strict accordance to the suggestions and comments by the editor and the referee. We hope that by making these revisions our manuscript will meet the standards of your journal and will be accepted for publication.

Sincerely,

Ranjula Bali Swain

Maria Floro