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A NOTE ON CONDITIONED GALTON-WATSON TREES

ROMAIN ABRAHAM AND JEAN-FRANÇOIS DELMAS

ABSTRACT. We give a necessary and sufficient condition for the convergence in distribution of a conditioned Galton-Watson tree to Kesten's tree. This yields elementary proofs of Kesten's result as well as other known results on local limit of conditioned Galton-Watson trees. We then apply this condition to get new results, in the critical and sub-critical cases, on the limit in distribution of a Galton-Watson tree conditioned on having a large number of individuals with out-degree in a given set.

1. INTRODUCTION

Galton-Watson (GW) process is a very simple model of population growth where all individuals give birth independently from each others to a random number of children with the same reproduction law p . This population growth can be described by a genealogical tree τ that we call the GW tree. It is well-known that, in the sub-critical case (the mean number of children of a single individual is strictly less than 1) and in the critical case (the mean number of children of an individual is 1), the population becomes a.s. extinct. However, one can define in these two cases the tree τ^* , called Kesten's tree, conditioned on non-extinction as the local limit in distribution of the tree τ conditioned to reach height n , when n tends to infinity, see KESTEN [9]; this result is recalled here in Section 2.4.

It is also known that, at least in the critical case, other conditionings such as conditioning by the total progeny, see KENNEDY [8] and GEIGER and KAUFMANN [4], or by the number of leaves, see KORTCHEMSKI [10] or MYLLÄRI [12] for associated works, lead to the same local limit in distribution. See also the survey from JANSON [6].

For all those cases, the conditioning event can be written $\{\tau \in \mathbb{A}_n\}$ with \mathbb{A}_n of the form:

$$\mathbb{A}_n = \{\mathbf{t}, A(\mathbf{t}) \geq n\} \quad \text{or} \quad \mathbb{A}_n = \{\mathbf{t}, A(\mathbf{t}) = n\},$$

where $A : \mathbf{t} \mapsto A(\mathbf{t})$ is a functional defined on the set of trees and satisfying an additive property, see Equation (9). We prove our following main result, see Theorem 3.1 for a precise statement, which unifies all the previous conditioning and gives a necessary and sufficient condition to obtain Kesten's tree as a limit. In the non-degenerate critical case, if A satisfies the additive property (9) with some aperiodic condition, then the following two statements are equivalent:

- $\lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{A}_n) = 0$ and $\lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{A}_{n+1}) / \mathbb{P}(\tau \in \mathbb{A}_n) = 1$,
- The distribution of τ conditionally on $\{\tau \in \mathbb{A}_n\}$ converges to the distribution of Kesten's tree τ^* .

Using this result, we give elementary proofs for the convergence in distribution to Kesten's tree τ^* of the GW tree conditioned on:

- (i) Extinction after or at a large time (sub-critical and critical case), with $A(\mathbf{t}) = H(\mathbf{t})$ the height of the tree \mathbf{t} and conditioning event $\{H(\tau) = n\}$ or $\{H(\tau) \geq n\}$. See Sections 4.1 and 4.2.

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- (ii) Large total population size (critical case), with $A(\mathbf{t}) = \text{Card}(\mathbf{t})$ the total size of the tree and conditioning event $\{\text{Card}(\tau) = n\}$ or $\{\text{Card}(\tau) \geq n\}$. See Section 4.3.
- (iii) Large number of leaves (critical case), with $A(\mathbf{t}) = L_0(\mathbf{t})$ the total number of leaves of \mathbf{t} and conditioning event $\{L_0(\tau) = n\}$ or $\{L_0(\tau) \geq n\}$. See Section 4.4.

In fact the conditioning on the large total population size or on the large number of leaves are particular cases of tree conditioned on large number of individuals with a given number of children. This corresponds to the functional $A(t) = L_{\mathcal{A}}(\mathbf{t})$ which gives the total number of individuals of the tree \mathbf{t} whose number of children belongs to a given set \mathcal{A} . Such conditioning has already been studied by RIZZOLO [14], see also MIMAMI [11], but for global scaling limit and not local limit. We obtain the convergence in distribution to Kesten's tree τ^* of the GW tree conditioned on:

- (iv) Large number of individuals with number of children in a given set \mathcal{A} (critical case), with $A(\mathbf{t}) = L_{\mathcal{A}}(\mathbf{t})$ and conditioning event $\{L_{\mathcal{A}}(\tau) = n\}$ or $\{L_{\mathcal{A}}(\tau) \geq n\}$. One has to distinguish according to $0 \in \mathcal{A}$, Section 5.1, and $0 \notin \mathcal{A}$, Section 5.2.

Let us remark that the total progeny $\mathcal{A} = \mathbb{N}$, the number of leaves $\mathcal{A} = \{0\}$ and the number of internal nodes $\mathcal{A} = \mathbb{N} \setminus \{0\}$ are particular cases of this conditioning.

We then study the subcritical case and define a one-parameter family $(p_\theta, \theta \in I)$ of distributions on the set of integers such that the GW tree τ associated with the reproduction law p and the GW tree τ_θ associated with the reproduction law p_θ have the same conditional distributions given $L_{\mathcal{A}}$, see Proposition 5.6. This generalizes Kennedy's transformation [8] concerning the total progeny, and the pruning of ABRAHAM, DELMAS and HE [1] concerning the number of leaves. We then immediately deduce, see Corollary 5.8, that if there exists θ_0 such that p_{θ_0} is critical, then the distribution of τ conditionally on $\{L_{\mathcal{A}}(\tau) = n\}$ or on $\{L_{\mathcal{A}}(\tau) \geq n\}$ converges (in the aperiodic case) to the distribution of the Kesten's tree $\tau_{\theta_0}^*$ associated with the reproduction law p_{θ_0} . When there is no such θ_0 , then a condensation phenomenon may appear, see [7] or [6].

Finally, we consider another conditioning which does not enter into the framework of Theorem 3.1 : conditioning on the size on the n -th generation. However, we can adapt the proof of Theorem 3.1 to get an analogous result in that case, see Proposition 6.1. We apply this result to a geometric reproduction law where explicit computations can be performed to prove that the corresponding tree conditioned on the n -th generation being on size n converges in distribution to Kesten's tree.

The paper is organized as follows. In the Section 2, we recall the framework we use for discrete trees and define the GW tree τ and Kesten's tree τ^* associated with a reproduction law p . In Section 3, we state and prove the necessary and sufficient condition for convergence in distribution of the conditioned tree to Kesten's tree. We apply this result in Section 4 to recover the classical results on critical conditioned GW trees and we study in Section 5 the case of the number of individuals with out-degree in a given set for the critical and sub-critical case. Finally, we study in Section 6 the conditioning on the size of the n -th generation of the tree.

2. TECHNICAL BACKGROUND ON GW TREES

2.1. First notations. We denote by $\mathbb{N} = \{0, 1, 2, \dots\}$ the set of non-negative integers and by $\mathbb{N}^* = \{1, 2, \dots\}$ the set of positive integers.

If K is a subset of \mathbb{N}^* , we call the span of K the greatest common divisor of K . If X is an integer-valued random variable, we call the span of X the span of $\{n > 0, \mathbb{P}(X = n) > 0\}$ the restriction to \mathbb{N}^* of its support.

2.2. The set of discrete trees. We recall Neveu's formalism [13] for ordered rooted trees. We set

$$\mathcal{U} = \bigcup_{n \geq 0} (\mathbb{N}^*)^n$$

the set of finite sequences of positive integers with the convention $(\mathbb{N}^*)^0 = \{\emptyset\}$. For $u \in \mathcal{U}$ let $|u|$ be the length or generation of u defined as the integer n such that $u \in (\mathbb{N}^*)^n$. If u and v are two sequences of \mathcal{U} , we denote by uv the concatenation of the two sequences, with the convention that $uv = u$ if $v = \emptyset$ and $uv = v$ if $u = \emptyset$. The set of ancestors of u is the set:

$$(1) \quad A_u = \{v \in \mathcal{U}; \text{there exists } w \in \mathcal{U}, w \neq \emptyset, \text{ such that } u = vw\}.$$

The most recent common ancestor of a subset \mathbf{s} of \mathcal{U} , denoted by $\text{MRCA}(\mathbf{s})$, is the unique element u of $\bigcap_{u \in \mathbf{s}} A_u$ with maximal length $|u|$.

A tree \mathbf{t} is a subset of \mathcal{U} that satisfies:

- $\emptyset \in \mathbf{t}$,
- If $u \in \mathbf{t}$, then $A_u \subset \mathbf{t}$.
- For every $u \in \mathbf{t}$, there exists a non-negative integer $k_u(\mathbf{t})$ such that, for every positive integer i , $ui \in \mathbf{t}$ iff $1 \leq i \leq k_u(\mathbf{t})$.

The integer $k_u(\mathbf{t})$ represents the number of offsprings of the vertex $u \in \mathbf{t}$. The vertex $u \in \mathbf{t}$ is called a leaf if $k_u(\mathbf{t}) = 0$. The vertex \emptyset is called the root of \mathbf{t} . Let us remark that, for a tree \mathbf{t} , we have

$$(2) \quad \sum_{u \in \mathbf{t}} k_u = \text{Card}(\mathbf{t}) - 1.$$

Let \mathbf{t} be a tree. The set of its leaves is $\mathcal{L}_0(\mathbf{t}) = \{u \in \mathbf{t}; k_u(\mathbf{t}) = 0\}$, its height is defined by

$$H(\mathbf{t}) = \sup\{|u|, u \in \mathbf{t}\}$$

and can be infinite. For $u \in \mathbf{t}$, we define the sub-tree $\mathcal{S}_u(\mathbf{t})$ of \mathbf{t} "above" u as:

$$\mathcal{S}_u(\mathbf{t}) = \{v \in \mathcal{U}, uv \in \mathbf{t}\}.$$

Notice that by construction $\text{MRCA}(\mathcal{S}_u(\mathbf{t})) = u$.

We denote by \mathbb{T} the set of trees, by

$$\mathbb{T}_0 = \{\mathbf{t} \in \mathbb{T}; \text{Card}(\mathbf{t}) < +\infty\}$$

the subset of finite trees, by

$$\mathbb{T}^{(h)} = \{\mathbf{t} \in \mathbb{T}; H(\mathbf{t}) \leq h\}$$

the subset of trees with height at most $h \in \mathbb{N}$, and by

$$\mathbb{T}_1 = \{\mathbf{t} \in \mathbb{T}; \lim_{n \rightarrow +\infty} |\text{MRCA}(\{u \in \mathbf{t}; |u| = n\})| = +\infty\}$$

the subset of trees with a unique infinite spine. Notice that \mathbb{T}_0 and $\mathbb{T}^{(h)}$ are countable and \mathbb{T}_1 is uncountable. For $h \in \mathbb{N}$ the restriction function r_h from \mathbb{T} to \mathbb{T} is defined by:

$$r_h(\mathbf{t}) = \{u \in \mathbf{t}, |u| \leq h\}.$$

We endow the set \mathbb{T} with the ultrametric distance

$$d(\mathbf{t}, \mathbf{t}') = 2^{-\max\{h \in \mathbb{N}, r_h(\mathbf{t}) = r_h(\mathbf{t}')\}}.$$

A sequence $(\mathbf{t}_n, n \in \mathbb{N})$ of trees converges to a tree \mathbf{t} with respect to the distance d if and only if, for every $h \in \mathbb{N}$,

$$r_h(\mathbf{t}_n) = r_h(\mathbf{t}) \quad \text{for } n \text{ large enough.}$$

The Borel σ -field associated with the distance d is the smallest σ -field containing the singletons for which the restriction functions $(r_h, h \in \mathbb{N})$ are measurable. With this distance, the restriction functions are contractant. Since \mathbb{T}_0 is dense in \mathbb{T} and (\mathbb{T}, d) is complete, we get that (\mathbb{T}, d) is a Polish metric space.

Consider the closed ball $B(\mathbf{t}, 2^{-h}) = \{\mathbf{t}' \in \mathbb{T}; d(\mathbf{t}, \mathbf{t}') \leq 2^{-h}\}$ for some $\mathbf{t} \in \mathbb{T}$ and $h \in \mathbb{N}$ and notice that:

$$B(\mathbf{t}, 2^{-h}) = r_h^{-1}(\{r_h(\mathbf{t})\}).$$

Since the distance is ultrametric, the closed balls are open and the open balls are closed, and the intersection of two balls is either empty or one of them. We deduce that the family $((r_h^{-1}(\{\mathbf{t}\}), \mathbf{t} \in \mathbb{T}^{(h)}, h \in \mathbb{N})$ is a π -system, and Theorem 2.3 in [2] implies that this family is convergence determining for the convergence in distribution. Let $(T_n, n \in \mathbb{N}^*)$ and T be \mathbb{T} -valued random variables. We denote by $\text{dist}(T)$ the distribution on the random variable T (which is uniquely determined by the sequence of distributions of $r_h(T)$ for every $h \geq 0$), and we denote

$$\text{dist}(T_n) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(T)$$

for the convergence in distribution on the sequence $(T_n, n \in \mathbb{N}^*)$ to T .

We deduce from the Portmanteau theorem that the sequence $(T_n, n \in \mathbb{N}^*)$ converge in distribution to T if and only if for all $h \in \mathbb{N}$, $\mathbf{t} \in \mathbb{T}^{(h)}$:

$$\lim_{n \rightarrow +\infty} \mathbb{P}(r_h(T_n) = \mathbf{t}) = \mathbb{P}(r_h(T) = \mathbf{t}).$$

As we shall only consider \mathbb{T}_0 -valued random variables that converge in distribution to a \mathbb{T}_1 -valued random variable, we shall give an other characterization of convergence in distribution that holds for this restriction. To present this result, we introduce some notations. If $\mathbf{t}, \mathbf{s} \in \mathbb{T}$ and $x \in \mathcal{L}_0(\mathbf{t})$ we denote by:

$$\mathbf{t} \otimes (\mathbf{s}, x) = \{u \in \mathbf{t}\} \cup \{xv, v \in \mathbf{s}\}$$

the tree obtained by grafting the tree \mathbf{s} on the leaf x of the tree \mathbf{t} . For every $\mathbf{t} \in \mathbb{T}$ and every $x \in \mathcal{L}_0(\mathbf{t})$, we shall consider the set of trees obtained by grafting a tree on the leaf x of \mathbf{t} :

$$\mathbb{T}(\mathbf{t}, x) = \{\mathbf{t} \otimes (\mathbf{s}, x), \mathbf{s} \in \mathbb{T}\}.$$

It is easy to see that $\mathbb{T}(\mathbf{t}, x)$ is closed. It is also open, as for all $\mathbf{s} \in \mathbb{T}(\mathbf{t}, x)$ we have that $B(\mathbf{s}, 2^{-H(\mathbf{t})-1}) \subset \mathbb{T}(\mathbf{t}, x)$.

Moreover, notice that the set \mathbb{T}_1 is a Borel subset of the set \mathbb{T} .

Lemma 2.1. *Let $(T_n, n \in \mathbb{N}^*)$ and T be \mathbb{T} -valued random variables which belong a.s. to $\mathbb{T}_0 \cup \mathbb{T}_1$. The sequence $(T_n, n \in \mathbb{N}^*)$ converges in distribution to T if and only if for every $\mathbf{t} \in \mathbb{T}_0$ and every $x \in \mathcal{L}_0(\mathbf{t})$, we have:*

$$(3) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(T_n \in \mathbb{T}(\mathbf{t}, x)) = \mathbb{P}(T \in \mathbb{T}(\mathbf{t}, x)) \quad \text{and} \quad \lim_{n \rightarrow +\infty} \mathbb{P}(T_n = \mathbf{t}) = \mathbb{P}(T = \mathbf{t}).$$

Proof. The subclass $\mathcal{F} = \{\mathbb{T}(\mathbf{t}, x), \mathbf{t} \in \mathbb{T}_0, x \in \mathcal{L}_0(\mathbf{t})\} \cup \{\{\mathbf{t}\}, \mathbf{t} \in \mathbb{T}_0\}$ of the Borel sets on $\mathbb{T}_0 \cup \mathbb{T}_1$ forms a π -system since we have

$$\mathbb{T}(\mathbf{t}_1, x_1) \cap \mathbb{T}(\mathbf{t}_2, x_2) = \begin{cases} \mathbb{T}(\mathbf{t}_1, x_1) & \text{if } \mathbf{t}_1 \in \mathbb{T}(\mathbf{t}_2, x_2), \\ \mathbb{T}(\mathbf{t}_2, x_2) & \text{if } \mathbf{t}_2 \in \mathbb{T}(\mathbf{t}_1, x_1), \\ \{\mathbf{t}_1\} & \text{if } \mathbf{t}_1 = \mathbf{t}_2 \text{ and } x_1 \neq x_2, \\ \emptyset & \text{in the other cases.} \end{cases}$$

For every $h \in \mathbb{N}$ and every $\mathbf{t} \in \mathbb{T}^{(h)}$, we have that \mathbf{t}' belongs to $r_h^{-1}(\{\mathbf{t}\}) \cap \mathbb{T}_1$ if and only if \mathbf{t}' belongs to some $\mathbb{T}(\mathbf{s}, x)$ where x is a leaf of \mathbf{t} such that $|x| = h$ and \mathbf{s} belongs to $r_h^{-1}(\{\mathbf{t}\}) \cap \mathbb{T}_0$ such that x is also a leaf of \mathbf{s} . Since \mathbb{T}_0 is countable, we deduce that \mathcal{F} generates the Borel σ -field on $\mathbb{T}_0 \cup \mathbb{T}_1$. In particular \mathcal{F} is a separating class.

Since $A \in \mathcal{F}$ is closed and open as well, according to Theorem 2.3 of [2], to prove that the family \mathcal{F} is a convergence determining class, it is enough to check that for all $\mathbf{t} \in \mathbb{T}_0 \cup \mathbb{T}_1$ and $h \in \mathbb{N}$, there exists $A \in \mathcal{F}$ such that:

$$(4) \quad \mathbf{t} \in A \subset B(\mathbf{t}, 2^{-h}).$$

If $\mathbf{t} \in \mathbb{T}_0$, this is clear as $\{\mathbf{t}\} = B(\mathbf{t}, 2^{-h})$ for all $h > H(\mathbf{t})$. If $\mathbf{t} \in \mathbb{T}_1$, for all $\mathbf{s} \in \mathbb{T}_0$ and $x \in \mathcal{L}_0(\mathbf{s})$ such that $\mathbf{t} \in \mathbb{T}(\mathbf{s}, x)$, we have $\mathbf{t} \in \mathbb{T}(\mathbf{s}, x) \subset B(\mathbf{t}, 2^{-|x|})$. Since we can find such a \mathbf{s} and x such that $|x|$ is arbitrary large, we deduce that (4) is satisfied. This proves that the family \mathcal{F} is a convergence determining class.

Since, for $\mathbf{t} \in \mathbb{T}_0$ and $x \in \mathcal{L}_0(\mathbf{t})$ the sets $\mathbb{T}(\mathbf{t}, x)$ and $\{\mathbf{t}\}$ are open and closed, we deduce from the Portmanteau Theorem that if $(T_n, n \in \mathbb{N}^*)$ converges in distribution to T , then (3) holds for every $\mathbf{t} \in \mathbb{T}_0$ and every $x \in \mathcal{L}_0(\mathbf{t})$. \square

2.3. GW trees. Let $p = (p(n), n \in \mathbb{N})$ be a probability distribution on the set of the non-negative integers. We assume that

$$(5) \quad p(0) > 0, \quad p(0) + p(1) < 1, \quad \text{and} \quad m := \sum_{n=0}^{+\infty} np(n) < +\infty.$$

A \mathbb{T} -valued random variable τ is a Galton-Watson (GW) tree with reproduction law p if the distribution of $k_\emptyset(\tau)$ is p and for $n \in \mathbb{N}^*$, conditionally on $\{k_\emptyset(\tau) = n\}$, the sub-trees $(\mathcal{S}_1(\tau), \mathcal{S}_2(\tau), \dots, \mathcal{S}_n(\tau))$ are independent and distributed as the original tree τ . Equivalently, for every $h \in \mathbb{N}^*$ and every $\mathbf{t} \in \mathbb{T}^{(h)}$, we have

$$\mathbb{P}(r_h(\tau) = \mathbf{t}) = \prod_{u \in r_{h-1}(\mathbf{t})} p(k_u(\mathbf{t})).$$

In particular, the restriction of the distribution of τ on the set \mathbb{T}_0 is given by:

$$(6) \quad \forall \mathbf{t} \in \mathbb{T}_0, \quad \mathbb{P}(\tau = \mathbf{t}) = \prod_{u \in \mathbf{t}} p(k_u(\mathbf{t})).$$

The GW tree is called critical (resp. sub-critical, super-critical) if $m = 1$ (resp. $m < 1$, $m > 1$).

2.4. Conditioning on non-extinction. Let p be a reproduction law satisfying Assumption (5) with $m \leq 1$ (i.e. the associated Galton-Watson process is critical or sub-critical). We denote by $p^* = (p^*(n) = np(n)/m, n \in \mathbb{N})$ the corresponding size-biased distribution.

According to Kesten [9], we can define an infinite random tree τ^* which can be viewed as the tree τ conditioned on non-extinction in the following sense:

$$\forall h \in \mathbb{N}^*, \quad \forall \mathbf{t} \in \mathbb{T}^{(h)}, \quad \mathbb{P}(r_h(\tau^*) = \mathbf{t}) = \lim_{n \rightarrow +\infty} \mathbb{P}(r_h(\tau) = \mathbf{t} \mid H(\tau) \geq n).$$

The distribution of τ^* is as follows. Almost surely, τ^* belongs to \mathbb{T}_1 , that is there exists a unique infinite sequence $(V_k, k \in \mathbb{N}^*)$ of positive integers such that, for every $h \in \mathbb{N}$, $V_1 \cdots V_h \in \tau^*$, with the convention that $V_1 \cdots V_h = \emptyset$ if $h = 0$. The joint distribution of $(V_k, k \in \mathbb{N}^*)$ and τ^* is determined recursively as follows: for each $h \in \mathbb{N}$, conditionally given (V_1, \dots, V_h) and $r_h(\tau^*)$, we have:

- The number of children $(k_v(\tau^*), v \in \tau^*, |v| = h)$ are independent and distributed according to p if $v \neq V_1 \cdots V_h$ and according to p^* if $v = V_1 \cdots V_h$.
- Given also the numbers of children $(k_v(\tau^*), v \in \tau^*, |v| = h)$, the integer V_{h+1} is uniformly distributed on the set of integers $\{1, \dots, k_{V_1 \cdots V_h}(\tau^*)\}$.

As a direct consequence we get that for all $h \in \mathbb{N}$, $\mathbf{t} \in \mathbb{T}^{(h)}$, $u \in \mathbf{t}$ such that $|u| = h$:

$$\mathbb{P}(r_h(\tau^*) = \mathbf{t}, V_1 \cdots V_h = u) = m^{-h} \mathbb{P}(r_h(\tau) = \mathbf{t}),$$

and for all $\mathbf{t} \in \mathbb{T}_0$, $x \in \mathcal{L}_0(\mathbf{t})$:

$$(7) \quad \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) = m^{-|x|} \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x)).$$

Since, for $\mathbf{t} \in \mathbb{T}_0$ and $x \in \mathcal{L}_0(\mathbf{t})$, $\mathbb{P}(\tau = \mathbf{t}) = \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x), k_x(\tau) = 0) = \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x))p(0)$, we deduce that:

$$(8) \quad \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) = \frac{1}{m^{|x|}p(0)} \mathbb{P}(\tau = \mathbf{t}).$$

3. MAIN RESULT

Let A be an integer-valued function defined on \mathbb{T} satisfying the following additivity property: there exists an integer-valued function D defined on \mathbb{T} such that, for every $\mathbf{t} \in \mathbb{T}_0$, every $x \in \mathcal{L}_0(\mathbf{t})$ and for every $\tilde{\mathbf{t}}$ such that $A(\mathbf{t} \otimes (\tilde{\mathbf{t}}, x))$ is large enough,

$$(9) \quad A(\mathbf{t} \otimes (\tilde{\mathbf{t}}, x)) = A(\tilde{\mathbf{t}}) + D(\mathbf{t}, x).$$

Let $n_0 \in \mathbb{N} \cup \{+\infty\}$ be given. We define for all $n \in \mathbb{N}^*$, the subset of trees

$$\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; A(\mathbf{t}) \in [n, n + n_0]\}.$$

Typical values of n_0 are 1 and $+\infty$.

The following theorem states that the distribution of the GW tree τ conditioned to be in \mathbb{A}_∞ , the limit of \mathbb{A}_n , is distributed as τ^* as soon as the probability of \mathbb{A}_n satisfies some regularity. We denote by

$$\text{dist}(\tau | \tau \in \mathbb{A}_n)$$

the conditional law of τ given $\{\tau \in \mathbb{A}_n\}$.

Theorem 3.1. *Assume that Assumptions (5) and (9) hold as well as one of the two following conditions*

- $m = 1$ or
- $m < 1$ and $D(\mathbf{t}, x) = |x|$ for all $\mathbf{t} \in \mathbb{T}_0$, $x \in \mathcal{L}_0(\mathbf{t})$.

Then, if

$$(10) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{A}_n) = 0 \quad \text{and} \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau \in \mathbb{A}_{n+1})}{\mathbb{P}(\tau \in \mathbb{A}_n)} = m,$$

we have:

$$\text{dist}(\tau | \tau \in \mathbb{A}_n) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(\tau^*),$$

that is:

$$(11) \quad \forall h \in \mathbb{N}^*, \forall \mathbf{t} \in \mathbb{T}^{(h)}, \lim_{n \rightarrow +\infty} \mathbb{P}(r_h(\tau) = \mathbf{t} \mid \tau \in \mathbb{A}_n) = \mathbb{P}(r_h(\tau^*) = \mathbf{t}).$$

Conversely, if $\text{dist}(\tau | \tau \in \mathbb{A}_n) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(\tau^*)$ that is (11) holds and if the span of $\{D(\mathbf{t}, x); \mathbf{t} \in \mathbb{T}_0 \text{ and } x \in \mathcal{L}_0(\mathbf{t})\} \cap \mathbb{N}^*$ is one, then (10) holds.

Proof. Let us first remark that, as we supposed that $m \leq 1$, we have a.s. $\tau \in \mathbb{T}_0$ and thus we are in the setting of Lemma 2.1.

Using (6), we have for every $\mathbf{t} \in \mathbb{T}_0$, $x \in \mathcal{L}_0(\mathbf{t})$ and $\tilde{\mathbf{t}} \in \mathbb{T}_0$:

$$\mathbb{P}(\tau = \mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) = \frac{1}{p(0)} \mathbb{P}(\tau = \mathbf{t}) \mathbb{P}(\tau = \tilde{\mathbf{t}}).$$

Let $\mathbf{t} \in \mathbb{T}_0$ and $x \in \mathcal{L}_0(\mathbf{t})$. We deduced that for n large enough, so that we can apply Equation (9):

$$\begin{aligned} \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x), \tau \in \mathbb{A}_n) &= \sum_{\tilde{\mathbf{t}} \in \mathbb{T}_0} \mathbb{P}(\tau = \mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) \mathbf{1}_{\{n \leq A(\mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) < n + n_0\}} \\ &= \frac{1}{p(0)} \sum_{\tilde{\mathbf{t}} \in \mathbb{T}_0} \mathbb{P}(\tau = \mathbf{t}) \mathbb{P}(\tau = \tilde{\mathbf{t}}) \mathbf{1}_{\{n \leq A(\tilde{\mathbf{t}}) + D(\mathbf{t}, x) < n + n_0\}} \\ &= \frac{1}{p(0)} \mathbb{P}(\tau = \mathbf{t}) \mathbb{P}(n - D(\mathbf{t}, x) \leq A(\tau) < n + n_0 - D(\mathbf{t}, x)) \\ &= m^{|x|} \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) \mathbb{P}(\tau \in \mathbb{A}_{n-D(\mathbf{t}, x)}), \end{aligned}$$

where we used (8) for the last equality. Therefore we have

$$(12) \quad \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x) \mid \tau \in \mathbb{A}_n) = \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) m^{|x|} \frac{\mathbb{P}(\tau \in \mathbb{A}_{n-D(\mathbf{t}, x)})}{\mathbb{P}(\tau \in \mathbb{A}_n)}.$$

Then, using (10) and that $D(\mathbf{t}, x) = |x|$ if $m < 1$, we obtain that:

$$(13) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x) \mid \tau \in \mathbb{A}_n) = \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)).$$

As $\lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{A}_n) = 0$, we also deduce that for all $\mathbf{t} \in \mathbb{T}_0$, $\mathbb{P}(\tau = \mathbf{t}, \tau \in \mathbb{A}_n) = 0$ for n large enough and thus:

$$(14) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(\tau = \mathbf{t} \mid \tau \in \mathbb{A}_n) = 0 = \mathbb{P}(\tau^* = \mathbf{t}).$$

We deduce from Lemma 2.1 that (11) holds.

Conversely, if (11) holds, then Lemma 2.1 implies that (13) and (14) hold. Then (14) readily implies that the first part of (10) holds. The fact that the span of $\{D(\mathbf{t}, x); \mathbf{t} \in \mathbb{T}_0 \text{ and } x \in \mathcal{L}_0(\mathbf{t})\} \cap \mathbb{N}^*$ is one and (12) imply, with Bezout theorem, that the second part of (10) holds. \square

4. EXAMPLES

4.1. Conditioning on extinction after large time. We give here a simple proof of Kesten's result for the convergence in distribution of a critical or sub-critical GW tree conditioned on non-extinction, see [10].

Proposition 4.1. *Let τ be a critical or sub-critical GW tree with reproduction law p satisfying Assumption (5). Then, we have*

$$(15) \quad \text{dist}(\tau \mid H(\tau) \geq n) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(\tau^*).$$

Proof. Consider $A(\mathbf{t}) = H(\mathbf{t})$ and $n_0 = +\infty$ that is $\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; H(\mathbf{t}) \geq n\}$. Notice that in this case for a tree $\tilde{\mathbf{t}}$ such that $H(\tilde{\mathbf{t}})$ is larger than $H(\mathbf{t})$, we have for every $x \in \mathcal{L}_0(\mathbf{t})$

$$(16) \quad A(\mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) = A(\tilde{\mathbf{t}}) + |x|.$$

Therefore, Condition (9) is satisfied by A .

According to Theorem 3.1, it suffices to prove

$$(17) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(H(\tau) \geq n) = 0, \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau \geq n+1)}{\mathbb{P}(\tau \geq n)} = m$$

to get (15). As we are in the critical or sub-critical case, $H(\tau)$ is finite a.s. and therefore, the first part of (17) holds.

We denote by φ the generating function of p and we define recursively $\varphi_1 = \varphi$ and for $n \geq 1$, $\varphi_{n+1} = \varphi_n \circ \varphi$. As φ_n is the generating function of the distribution of $\{u \in \tau; |u| = n\}$ the number of individuals at height n , we have $\mathbb{P}(\tau \in \mathbb{A}_n) = 1 - \varphi_n(0)$. We also have $\lim_{n \rightarrow +\infty} \varphi_n(0) = 1$ and

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau \in \mathbb{A}_{n+1})}{\mathbb{P}(\tau \in \mathbb{A}_n)} = \lim_{n \rightarrow +\infty} \frac{1 - \varphi(\varphi_n(0))}{1 - \varphi_n(0)} = \varphi'(1) = m$$

which is the second part of (17). \square

4.2. Conditioning on extinction at large time.

Proposition 4.2. *Let τ be a critical or sub-critical GW tree with reproduction law p satisfying Assumption (5). Then we have*

$$(18) \quad \text{dist}(\tau | H(\tau) = n) \longrightarrow \text{dist}(\tau^*).$$

Proof. We consider $A(t) = H(\mathbf{t})$ with $n_0 = 1$ that is $\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; H(\mathbf{t}) = n\}$. Since (16) is in force, we get that Condition (9) still holds. Again it suffices to prove

$$(19) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(H(\tau) = n) = 0, \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau = n+1)}{\mathbb{P}(\tau = n)} = m$$

to get (18). Since τ is a.s. finite, the first part of (19) is immediate.

Recall notation φ_n introduced in Section 4.1 and that $\lim_{n \rightarrow +\infty} \varphi_n(0) = 1$. We have $\mathbb{P}(\tau \in \mathbb{A}_n) = \varphi_{n+1}(0) - \varphi_n(0)$ and:

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau \in \mathbb{A}_{n+1})}{\mathbb{P}(\tau \in \mathbb{A}_n)} = \lim_{n \rightarrow +\infty} \frac{\frac{1 - \varphi(\varphi_n(0))}{1 - \varphi_n(0)} - \frac{1 - \varphi_2(\varphi_n(0))}{1 - \varphi_n(0)}}{1 - \frac{1 - \varphi(\varphi_n(0))}{1 - \varphi_n(0)}} = \frac{m - m^2}{1 - m} = m,$$

which is the second part of (19). \square

4.3. Conditioning on the total population size, critical case. We recover here results from [8] (under the hypothesis that $\text{Var}(X)$ is finite) or Theorem 7.1 in [6] (under no condition on the variance of X) on the convergence in distribution of a critical Galton-Watson tree conditioned on the size of its total progeny to Kesten's tree.

Proposition 4.3. *Let τ be a critical GW tree with reproduction law p satisfying Assumption (5). Let d be the span of $\text{Card}(\tau) - 1$ (that is the span of the set $\{k > 0, p(k) > 0\}$). Then we have*

$$(20) \quad \text{dist}(\tau | \text{Card}(\tau) = nd + 1) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(\tau^*). \quad \text{and}$$

and

$$(21) \quad \text{dist}(\tau | \text{Card}(\tau) \geq n) \xrightarrow[n \rightarrow +\infty]{} \text{dist}(\tau^*).$$

Remark 4.4. If we consider $A(\mathbf{t}) = \text{Card}(\mathbf{t})$ and $n_0 = +\infty$ that is $\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}, \text{Card}(\mathbf{t}) \geq n\}$, it is easy to adapt the end of the proof of Theorem 3.1 to get

$$(22) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\text{Card}(\tau) \geq n+1)}{\mathbb{P}(\text{Card}(\tau) \geq n)} = 1.$$

Indeed, notice that the span of $\{D(\mathbf{t}, x); \mathbf{t} \in \mathbb{T}_0 \text{ and } x \in \mathcal{L}_0(\mathbf{t})\} \cap \mathbb{N}^*$, with $D(\mathbf{t}, x) = A(\mathbf{t})$ given in (23), is d . Then use that $\mathbb{A}_{nd+1} = \dots = \mathbb{A}_{nd+d+1}$ to prove that (21) implies (22). Notice it is not clear how we can obtain (22) directly.

Proof of Proposition 4.3. Consider $A(\mathbf{t}) = \text{Card}(\mathbf{t})$ and $n_0 = d$. Then we have

$$\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; \text{Card}(\mathbf{t}) \in [n, n+d]\}.$$

We have for every $\mathbf{t} \in \mathbb{T}$, without any additional assumption,

$$(23) \quad A(\mathbf{t} \otimes (\tilde{\mathbf{t}}, x)) = A(\tilde{\mathbf{t}}) + A(\mathbf{t}),$$

so Condition (9) holds. Again, it therefore suffices to prove

$$(24) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(\text{Card}(\tau) \in [n, n+d]) = 0, \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\text{Card}(\tau) \in [n+1, n+1+d])}{\mathbb{P}(\text{Card}(\tau) \in [n, n+d])} = 1$$

to get (20). The first part of (24) is a direct consequence of $\text{Card}(\tau) < +\infty$ a.s.

Before proving the second part of (24), we first recall Dwass formula (see [3]). Let $(\tau_k, k \in \mathbb{N}^*)$ be independent GW trees distributed as τ . Set $W_k = \text{Card}(\tau_k)$. Let $(X_k, k \in \mathbb{N}^*)$ be independent integer-valued random variables distributed according to p . For $k \in \mathbb{N}^*$ and $n \geq k$, we have:

$$(25) \quad \mathbb{P}(W_1 + \dots + W_k = n) = \frac{k}{n} \mathbb{P}(X_1 + \dots + X_n = n - k).$$

We also recall some results on random walks. Let Y be an integrable random variable taking values in \mathbb{Z} , such that $\mathbb{E}[Y] = 0$, $\mathbb{P}(Y = 0) < 1$ and the span of $|Y|$ is 1. We consider the random walk $S = (S_n, n \in \mathbb{N})$ defined by:

$$(26) \quad S_0 = 0 \quad \text{and} \quad S_n = \sum_{k=1}^n Y_k \quad \text{for } n \in \mathbb{N}^*.$$

Then the random walk S is recurrent. If S is aperiodic, the strong ratio theorem for recurrent aperiodic random walks, see Theorem T1 p49 of [15], gives that, for $\ell \in \mathbb{Z}$:

$$(27) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(S_n = \ell)}{\mathbb{P}(S_n = 0)} = \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(S_n = 0)}{\mathbb{P}(S_{n+1} = 0)} = 1.$$

If S has period d , then for all $k \in \{1, \dots, d\}$, there exist $j_k \in \mathbb{Z}$ and $n_k \in \mathbb{N}^*$ such that

$$(28) \quad \forall n \geq n_k, \quad \mathbb{P}(S_{nd+k} = j_k) > 0.$$

The strong ratio theorem can then easily be adapted (using arguments similar to those used in the proof of Lemma 5.5) to get that, for $\ell \in \mathbb{Z}$, $k \in \{1, \dots, d\}$:

$$(29) \quad \lim_{m \rightarrow +\infty} \frac{\mathbb{P}(S_{md+k} = \ell d + j_k)}{\mathbb{P}(S_{md} = 0)} = 1.$$

By the definition of d , a.s. we have $A(\tau) \in d\mathbb{N} + 1$. We consider an integer valued random variable X distributed according to p and we set $Y = X - 1$ so that $\mathbb{E}[Y] = 0$ since we

supposed that $m = 1$. The random walk defined by (26) has period d and we can choose $j_1 = -1$ in (28) as $\mathbb{P}(Y = -1) > 0$. Dwass formula (25) implies that, for $m = \lfloor (n-1)/d \rfloor$:

$$\mathbb{P}(\tau \in \mathbb{A}_n) = \mathbb{P}(A(\tau) \in [n, n+d]) = \mathbb{P}(A(\tau) = md+1) = \frac{1}{md+1} \mathbb{P}(S_{md+1} = -1).$$

Using (29), we deduce that:

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\tau \in \mathbb{A}_{n+1})}{\mathbb{P}(\tau \in \mathbb{A}_n)} = \lim_{m \rightarrow +\infty} \frac{\mathbb{P}(S_{(m+1)d+1} = -1)}{\mathbb{P}(S_{md+1} = -1)} = 1$$

which implies readily the second part of (24).

Let's turn now to the proof of (21). As $\text{dist}(\tau | \tau \in \mathbb{A}_n)$ is a mixture of $\text{dist}(\tau | \text{Card}(\tau) = k)$ for $k \geq n$, we deduce from the first part of the proposition that (21) holds. \square

Remark 4.5. Notice that the local limit theorem gives asymptotics for $\mathbb{P}(S_n = -1)$ when the distribution of X belongs to the domain of attraction of a stable law, see Theorem 4.2.1 of [5] or Theorem 1.10 in [10]. This gives asymptotics for $\mathbb{P}(\tau \in \mathbb{A}_n)$ which in turns allow to recover Condition (10).

4.4. Conditioning on the number of leaves, critical case. For a finite tree $\mathbf{t} \in \mathbb{T}_0$, we denote by $L_0(\mathbf{t}) = \text{Card}(\mathcal{L}_0(\mathbf{t}))$ the number of leaves of \mathbf{t} . The next proposition (which seems to be a new result) is in fact a particular case of the proposition of the next section. However, we prove it separately for pedagogical purpose as its proof and in particular the construction of the Galton-Watson tree that codes $\mathcal{L}_0(\mathbf{t})$ of Remark 4.8 are much simpler in that particular case.

Proposition 4.6. *Let τ be a critical GW tree with reproduction law p satisfying Assumption (5). Let d_0 be the span of the random variable $L_0(\tau) - 1$. Then we have*

$$(30) \quad \text{dist}(\tau | L_0(\tau) = nd_0 + 1) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

and

$$(31) \quad \text{dist}(\tau | L_0(\tau) \geq n) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

Proof. We consider $A(\mathbf{t}) = L_0(\mathbf{t})$ and $n_0 = d_0$ which yields $\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; L_0(\mathbf{t}) \in [n, n+d_0]\}$. We have for every trees $\mathbf{t}, \tilde{\mathbf{t}} \in \mathbb{T}_0$ and every $x \in \mathcal{L}_0(\mathbf{t})$

$$(32) \quad A(\mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) = A(\tilde{\mathbf{t}}) + A(\mathbf{t}) - 1.$$

According to [11], see also Remark 4.8 below, $L_0(\tau)$ is distributed as the total size of a critical GW tree τ_0 with reproduction law given by the distribution of:

$$(33) \quad X_0 = \sum_{k=1}^{N-1} Z_k,$$

with $(Z_k, k \in \mathbb{N}^*)$ and N independent random variables such that $(Z_k, k \in \mathbb{N}^*)$ are independent and distributed as $X - 1$ conditionally on $\{X \geq 1\}$ (where X is a random variable distributed according to p) and N has a geometric distribution with parameter $p(0)$. As $\mathbb{E}[X_0] = 1$, we get that τ_0 is critical. Notice that d_0 is also the span of the random variable X_0 and that, if the span of X is d , then $d_0 = \max(1, d-1)$.

We deduce from Subsection 4.3 that Condition (24) is satisfied and thus:

$$(34) \quad \lim_{n \rightarrow +\infty} \mathbb{P}(L_0(\tau) \in [n, n+d_0]) = 0, \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_0(\tau) \in [n+1, n+1+d_0])}{\mathbb{P}(L_0(\tau) \in [n, n+d_0])} = 1.$$

Then use Theorem 3.1 to get that (30) holds.

If we consider $n_0 = +\infty$ that is:

$$\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}_0; L_0(\mathbf{t}) \geq n\},$$

arguing as in the proof of the second part of Proposition 4.3, we get (31). \square

Remark 4.7. We deduce from Remark 4.4 that (31) implies

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_0(\tau) \geq n + 1)}{\mathbb{P}(L_0(\tau) \geq n)} = 1.$$

Remark 4.8. We shall briefly recall how one can prove that $L_0(\tau)$ is distributed as the total size of a Galton-Watson process by mapping the set of leaves $\mathcal{L}_0(\tau)$ onto a GW tree, see [11].

Let \mathbf{t} be a tree. For $u \in \mathbf{t}$, we define the left branch starting from u as:

$$B_g^{\mathbf{t}}(u) = \{uv; |v| \geq 1 \text{ and } v = \{1\}^{|v|}\} \cap \mathbf{t}.$$

We also define the left leaf $G(u)$ of u and the left ancestors $A_g(v)$ of a leaf v as:

$$G^{\mathbf{t}}(u) = B_g^{\mathbf{t}}(u) \cap \mathcal{L}_0(\mathbf{t}) \text{ and } A_g^{\mathbf{t}}(v) = \{u \in A_v; G^{\mathbf{t}}(u) = v\}.$$

For a leaf $v \in \mathcal{L}_0(\mathbf{t})$, we define its leaf-children as:

$$C^{\mathbf{t}}(v) = \{G^{\mathbf{t}}(ui); u \in A_g^{\mathbf{t}}(v), 1 < i \leq k_u(\mathbf{t})\},$$

labeled according to the following order: $G^{\mathbf{t}}(ui) < G^{\mathbf{t}}(u'i')$ if $u < u'$ in the lexicographic order or if $u = u'$ and $i < i'$. This defines a tree, obtained from the leaves of \mathbf{t} , denoted by $\mathbf{t}_{\{0\}} = F_{\{0\}}(\mathbf{t})$. And we have $\text{Card}(\mathbf{t}_{\{0\}}) = L_0(\mathbf{t})$.

If τ is a GW tree then $\tau_{\{0\}} = F_{\{0\}}(\tau)$ is also a GW tree with reproduction law given by the distribution of X_0 in (33).

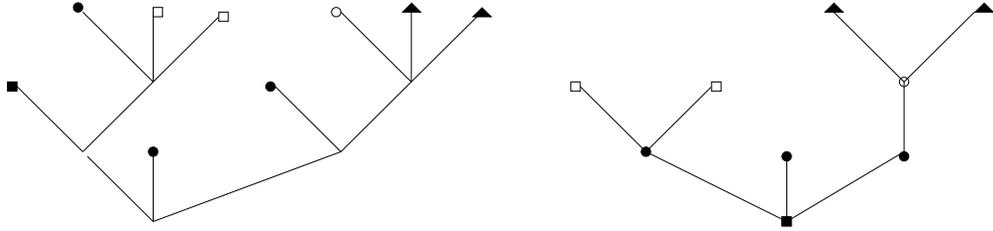


FIGURE 1. A tree \mathbf{t} on the left and the coding of $\mathcal{L}_0(\mathbf{t})$ by a tree $\mathbf{t}_0 = F(\mathbf{t})$ tree on the right.

5. CONDITIONING ON THE NUMBER OF INDIVIDUALS HAVING A GIVEN NUMBER OF CHILDREN

Let \mathcal{A} be a non-empty subset of \mathbb{N} . For a tree $\mathbf{t} \in \mathbb{T}$, we write $\mathcal{L}_{\mathcal{A}}(\mathbf{t}) = \{u \in \mathbf{t}; k_u(\mathbf{t}) \in \mathcal{A}\}$ the set individuals whose number of children belongs to \mathcal{A} and $L_{\mathcal{A}}(\mathbf{t}) = \text{Card}(\mathcal{L}_{\mathcal{A}}(\mathbf{t}))$ its cardinal. The case $\mathcal{A} = \{0\}$ represents the set of leaves of \mathbf{t} and has been treated in Section 4.4. We can also have $L_{\mathcal{A}}(\mathbf{t}) = \text{Card}(\mathbf{t})$ by taking $\mathcal{A} = \mathbb{N}$ or $L_{\mathcal{A}}$ can also be the number of internal nodes by taking $\mathcal{A} = \mathbb{N}^*$.

Let us first remark that for every $\mathbf{t} \in \mathbb{T}_0$, every $x \in \mathcal{L}_0(\mathbf{t})$ and every $\tilde{\mathbf{t}} \in \mathbb{T}$

$$L_{\mathcal{A}}(\mathbf{t} \otimes (\tilde{\mathbf{t}}, x)) = \begin{cases} L_{\mathcal{A}}(\mathbf{t}) + L_{\mathcal{A}}(\tilde{\mathbf{t}}) - 1 & \text{if } 0 \in \mathcal{A}, \\ L_{\mathcal{A}}(\mathbf{t}) + L_{\mathcal{A}}(\tilde{\mathbf{t}}) & \text{if } 0 \notin \mathcal{A}, \end{cases}$$

and hence $L_{\mathcal{A}}$ satisfies the additive property (9) with $D(\mathbf{t}, x) = L_{\mathcal{A}}(\mathbf{t}) - \mathbf{1}_{\{0 \in \mathcal{A}\}}$.

Theorem 5.1. *Let τ be a critical GW tree with reproduction law p satisfying Assumption (5). Let $d_{\mathcal{A}}$ be the span of the random variable $L_{\mathcal{A}}(\tau) - 1$. Then we have*

$$(35) \quad \text{dist}(\tau | L_{\mathcal{A}}(\tau) = nd_{\mathcal{A}} + 1) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

and

$$(36) \quad \text{dist}(\tau | L_{\mathcal{A}}(\tau) \geq n) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

Remark 5.2. If $0 \notin \mathcal{A}$, then $d_{\mathcal{A}} = 1$.

Remark 5.3. Arguing as in Remark 4.4, we get

$$(37) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) \geq n + 1)}{\mathbb{P}(L_{\mathcal{A}}(\tau) \geq n)} = 1.$$

In what follows, we denote by X a random variable distributed according to p . To avoid triviality, we assume that $0 < \mathbb{P}(X \in \mathcal{A}) \leq 1$. The proofs differ if $0 \in \mathcal{A}$ or not, the next two subsections are devoted each to one case. Then, we consider the same problem for the sub-critical case, $m = \mathbb{E}[X] < 1$, in a third subsection.

5.1. Proof of Theorem 5.1 when $0 \in \mathcal{A}$. We consider only $\mathbb{P}(X \in \mathcal{A}) < 1$, as the case $\mathbb{P}(X \in \mathcal{A}) = 1$ corresponds to the critical case with $\mathcal{A} = \mathbb{N}$ of Section 4.3. In this case, extending the decomposition of [11], it is easy to see that $L_{\mathcal{A}}(\tau)$ is distributed as the total size of a critical GW tree $\tau_{\mathcal{A}}$ whose reproduction law is given by the distribution of:

$$(38) \quad X_{\mathcal{A}} = \sum_{k=1}^{N-1} Y'_k + Y'',$$

where $(Y'_k, k \in \mathbb{N}^*)$, N and Y'' are independent random variables such that $(Y'_k, k \in \mathbb{N}^*)$ are independent and distributed as $X - 1$ conditionally on $\{X \notin \mathcal{A}\}$, N is geometric with parameter $\mathbb{P}(X \in \mathcal{A})$ and Y'' is distributed as X conditionally on $\{X \in \mathcal{A}\}$. See Remark 5.4 below for a construction of the corresponding GW tree.

As $\mathbb{E}[X_{\mathcal{A}}] = 1$, we get that $\tau_{\mathcal{A}}$ is critical. Remark that $d_{\mathcal{A}}$ is also the span of $X_{\mathcal{A}}$. Consider $n_0 = d_{\mathcal{A}}$ which gives

$$\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; L_{\mathcal{A}}(\mathbf{t}) \in [n, n + d_{\mathcal{A}}]\}.$$

As $L_{\mathcal{A}}(\tau)$ is distributed as the total size of a critical GW tree, we deduce from Subsection 4.3 that

$$(39) \quad \lim \mathbb{P}(L_{\mathcal{A}}(\tau) \in [n, n + d_{\mathcal{A}}]) = 0, \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) \in [n + 1, n + 1 + d_{\mathcal{A}}])}{\mathbb{P}(L_{\mathcal{A}}(\tau) \in [n, n + d_{\mathcal{A}}])} = 1$$

and thus by Theorem 3.1 that (35) holds. \square

Remark 5.4. We shall briefly indicate how one can prove that $L_{\mathcal{A}}(\tau)$ is distributed as the total size of a Galton-Watson process by mapping the set $\mathcal{L}_{\mathcal{A}}(\tau)$ onto a GW tree, generalizing the construction of Remark 4.8, see also [11] for $\mathcal{A} = [0, n]$.

We use notations from Remark 4.8. For a tree \mathbf{t} , $u \in \mathbf{t}$ and $v \in \mathcal{L}_{\mathcal{A}}(\mathbf{t})$, we define the left descendant of u in $\mathcal{L}_{\mathcal{A}}(\mathbf{t})$ and the left ancestors $A_g^{\mathbf{t}}(v)$ of v as:

$$G_{\mathcal{A}}^{\mathbf{t}}(u) = \operatorname{argmin} \{|w|; w \in B_g^{\mathbf{t}}(u) \cap \mathcal{L}_{\mathcal{A}}(\mathbf{t})\} \quad \text{and} \quad A_{g,\mathcal{A}}^{\mathbf{t}}(v) = \{u \in A_v; G_{\mathcal{A}}^{\mathbf{t}}(u) = v\}.$$

Remark that $G_{\mathcal{A}}^{\mathbf{t}}(u)$ exists since we supposed that $0 \in \mathcal{A}$.

For $v \in \mathcal{L}_{\mathcal{A}}(\mathbf{t})$, we define its children as:

$$C_{\mathcal{A}}^{\mathbf{t}}(v) = \begin{cases} \{G_{\mathcal{A}}^{\mathbf{t}}(ui); u \in A_{g,\mathcal{A}}^{\mathbf{t}}(v), 1 < i \leq k_u\} & \text{if } v \in \mathcal{L}(\mathbf{t}), \\ \{G_{\mathcal{A}}^{\mathbf{t}}(ui); u \in A_{g,\mathcal{A}}^{\mathbf{t}}(v), 1 < i \leq k_u\} \cup \{vi, 1 \leq i \leq k_v(\mathbf{t})\}, & \text{if } v \notin \mathcal{L}(\mathbf{t}), \end{cases}$$

labeled according to the following order: $G_{\mathcal{A}}^{\mathbf{t}}(ui) < G_{\mathcal{A}}^{\mathbf{t}}(u'i')$ if $u < u'$ in the lexicographic order or if $u = u'$ and $i < i'$; $G_{\mathcal{A}}^{\mathbf{t}}(ui) < vi'$ for $u \in A_{g,\mathcal{A}}^{\mathbf{t}}(v)$ and $1 \leq i' \leq k_v(\mathbf{t})$; and $vi < vi'$ for $1 \leq i < i' \leq k_v(\mathbf{t})$. This defines a tree, obtained from $\mathcal{L}_{\mathcal{A}}(\mathbf{t})$, denoted by $\mathbf{t}_{\mathcal{A}} = F_{\mathcal{A}}(\mathbf{t})$. And we have $\operatorname{Card}(\mathbf{t}_0) = L_{\mathcal{A}}(\mathbf{t})$.

If τ is a GW tree then $\tau_{\mathcal{A}} = F_{\mathcal{A}}(\tau)$ is also a GW tree with reproduction law given by the distribution of $X_{\mathcal{A}}$ in (38).

5.2. Proof of Theorem 5.1 when $0 \notin \mathcal{A}$. Notice that in that case, $\mathbb{P}(X \in \mathcal{A}) < 1$. For a tree \mathbf{t} and $u \in \mathcal{L}_{\mathcal{A}}(\mathbf{t})$ we define $\mathcal{S}_u^{\mathcal{A}}(\mathbf{t})$ the sub-tree rooted at u with no progeny in \mathcal{A} by:

$$\mathcal{S}_u^{\mathcal{A}}(\mathbf{t}) = \{w \in \mathcal{S}_u; A_w \cap A_u^c \cap \mathcal{L}_{\mathcal{A}}(\mathbf{t}) = \emptyset\}.$$

We set $p(\mathcal{A}) = \mathbb{P}(X \in \mathcal{A})$. We define $\tilde{p}(0) = p(0) + \sum_{k \in \mathcal{A}} p(k)$ and, for $k \in \mathbb{N}^*$, $\tilde{p}(k) = p(k) \mathbf{1}_{\{k \notin \mathcal{A}\}}$. Let \tilde{X} be a random variable with distribution \tilde{p} , so that $\mathbb{E}[\tilde{X}] = \mathbb{E}[X \mathbf{1}_{\{X \in \mathcal{A}^c\}}] < 1$. Notice that $\mathcal{S}_{\emptyset}^{\mathcal{A}}(\tau)$ is a GW tree with reproduction law \tilde{p} . We denote by \tilde{L} be the number of leaves of $\mathcal{S}_{\emptyset}^{\mathcal{A}}(\tau)$. Elementary computations yields:

$$(40) \quad \mathbb{E}[\tilde{L}] = \frac{\tilde{p}(0)}{1 - \mathbb{E}[\tilde{X}]} = \frac{p(0) + p(\mathcal{A})}{\mathbb{E}[X \mathbf{1}_{\{X \in \mathcal{A}\}}]},$$

where we used that $\mathbb{E}[X] = 1$ for the last equality.

We defined the children of u in \mathcal{A} as:

$$C^{\mathbf{t}}(u) = \mathcal{L}_0(\mathcal{S}_u^{\mathcal{A}}(\mathbf{t})) \cap \mathcal{L}_{\mathcal{A}}(\mathbf{t}).$$

The number of children in \mathcal{A} of the root \emptyset is then $\tilde{N}_{\emptyset} := \operatorname{Card}(C(\emptyset))$. Notice that conditionally given \tilde{L} , \tilde{N}_{\emptyset} has a binomial distribution with parameter $(\tilde{L}, p(\mathcal{A})/\tilde{p}(0))$. Let:

$$(41) \quad X_{\mathcal{A}} = \sum_{k=1}^{Z'} \tilde{N}_k,$$

where Z' is a random variable distributed as X conditionally on $\{X \in \mathcal{A}\}$, and $(\tilde{N}_k, k \in \mathbb{N}^*)$ are independent random variables distributed as \tilde{N}_{\emptyset} and independent of Z . Notice that for any $u \in \mathcal{L}_{\mathcal{A}}(\tau)$ with $u \neq \emptyset$, its number of children in \mathcal{A} , $C(u)$, is distributed as $X_{\mathcal{A}}$.

Let τ' be a GW tree with reproduction law given by the law of $X_{\mathcal{A}}$ and initial size 1. Let $(W'_k, k \in \mathbb{N}^*)$ be independent random variables distributed as $\operatorname{Card}(\tau')$ and independent of \tilde{N}_{\emptyset} . We deduce that $\mathcal{L}_{\mathcal{A}}(\tau)$ is distributed as $\sum_{k=1}^{\tilde{N}_{\emptyset}} W'_k$. Notice that $\mathcal{L}_{\mathcal{A}}(\tau)$ is empty if $\tilde{N}_{\emptyset} = 0$. By construction, notice the span of $X_{\mathcal{A}}$ is 1. The proof of the next lemma is postponed to the end of the section.

Lemma 5.5. *We have:*

$$(42) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(W'_1 = n)} = \mathbb{E}[\tilde{N}_\emptyset].$$

Then, using Subsection 4.3, we have $\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(W'_1 = n)}{\mathbb{P}(W'_1 = n+1)} = 1$, and thus we deduce that:

$$(43) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(L_{\mathcal{A}}(\tau) = n+1)} = 1.$$

As the span of $X_{\mathcal{A}}$ is 1, we consider $n_0 = 1$ that is:

$$\mathbb{A}_n = \{\mathbf{t} \in \mathbb{T}; L_{\mathcal{A}}(\mathbf{t}) = n\}.$$

Since τ is a.s. finite, the first part of Condition (39) is satisfied; and (43) readily implies the second part of Condition (39). As usual, we deduce that (35) holds. \square

Proof of Lemma 5.5. We have, using (40):

$$\begin{aligned} \mathbb{E}[X_{\mathcal{A}}] &= \mathbb{E}[Z'] \mathbb{E}[\tilde{N}_k] \\ &= \frac{1}{p(\mathcal{A})} \mathbb{E}[X \mathbf{1}_{\{X \in \mathcal{A}\}}] \mathbb{E}[\tilde{L}] \frac{p(\mathcal{A})}{p(0) + p(ca)} \\ &= 1. \end{aligned}$$

Let $S'_n = \sum_{k=1}^n (X'_k - 1)$, where $(X'_k, k \in \mathbb{N}^*)$ are independent random variables distributed as $X_{\mathcal{A}}$. By construction and using Dwass formula (25), we have:

$$\mathbb{P}(L_{\mathcal{A}}(\tau) = n) = \sum_{k=1}^{+\infty} \mathbb{P}(\tilde{N}_\emptyset = k) \mathbb{P}\left(\sum_{j=1}^k W'_j = n\right) = \frac{1}{n} \sum_{k=1}^{+\infty} k \mathbb{P}(\tilde{N}_\emptyset = k) \mathbb{P}(S'_n = -k).$$

We set:

$$\delta_n(k) = \frac{\mathbb{P}(S'_n = -k)}{\mathbb{P}(S'_n = -1)}.$$

The random walk $(S'_n, n \in \mathbb{N})$ is recurrent and aperiodic as the span of $X_{\mathcal{A}}$ is 1. We deduce from the strong ratio theorem (27) that:

$$(44) \quad \lim_{n \rightarrow +\infty} \delta_n(k) = 1.$$

Using Fatou's Lemma, we get:

$$\liminf_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(W'_1 = n)} \geq \sum_{k=1}^{+\infty} k \mathbb{P}(\tilde{N}_\emptyset = k) \liminf_{n \rightarrow +\infty} \delta_n(k) = \mathbb{E}[\tilde{N}_\emptyset].$$

As W'_1 is the total size of a Galton-Watson process with reproduction law given by $X_{\mathcal{A}}$, we have that W'_1 has the same distribution as:

$$1 + \sum_{k=1}^{X_{\mathcal{A}}} W'_k.$$

Using Dwass formula (25), we have:

$$\begin{aligned} \frac{\mathbb{P}(W'_1 = n+1)}{\mathbb{P}(W'_1 = n)} &= \sum_{k=1}^{+\infty} \mathbb{P}(Z_{\mathcal{A}} = k) \frac{\mathbb{P}\left(\sum_{j=1}^k W'_j = n\right)}{\mathbb{P}(W'_1 = n)} \\ &= \sum_{k=1}^{+\infty} k \mathbb{P}(Z_{\mathcal{A}} = k) \delta_n(k). \end{aligned}$$

Let $m_0 \geq 1$ such that $\mathbb{P}(Z' = m_0) > 0$. We deduce from (41) that:

$$\{\tilde{N}_1 = k, \tilde{N}_2 = 0, \dots, \tilde{N}_{m_0} = 0, Z' = m_0\} \subset \{X_{\mathcal{A}} = k\},$$

and thus, with $C^{-1} = \mathbb{P}(\tilde{N}_{\emptyset} = 0)^{m_0-1} \mathbb{P}(Z' = m_0) > 0$ we have for all $k \in \mathbb{N}$:

$$\mathbb{P}(\tilde{N}_{\emptyset} = k) \leq C \mathbb{P}(X_{\mathcal{A}} = k).$$

We get for any $K \in \mathbb{N}^*$:

$$\begin{aligned} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(W'_1 = n)} &= \sum_{k=1}^K k \mathbb{P}(\tilde{N}_{\emptyset} = k) \delta_n(k) + \sum_{k=K+1}^{+\infty} k \mathbb{P}(\tilde{N}_{\emptyset} = k) \delta_n(k) \\ &\leq \sum_{k=1}^K k \mathbb{P}(\tilde{N}_{\emptyset} = k) \delta_n(k) + C \sum_{k=K+1}^{+\infty} k \mathbb{P}(X_{\mathcal{A}} = k) \delta_n(k) \\ &= \sum_{k=1}^K k \mathbb{P}(\tilde{N}_{\emptyset} = k) \delta_n(k) - C \sum_{k=1}^K k \mathbb{P}(X_{\mathcal{A}} = k) \delta_n(k) + C \frac{\mathbb{P}(W'_1 = n+1)}{\mathbb{P}(W'_1 = n)}. \end{aligned}$$

Using (44), we get:

$$\limsup_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(W'_1 = n)} \leq \sum_{k=1}^K k \mathbb{P}(\tilde{N}_{\emptyset} = k) - C \sum_{k=1}^K k \mathbb{P}(X_{\mathcal{A}} = k) + C.$$

Letting K goes to infinity and using that $\mathbb{E}[X_{\mathcal{A}}] = 1$, we get:

$$\limsup_{n \rightarrow +\infty} \frac{\mathbb{P}(L_{\mathcal{A}}(\tau) = n)}{\mathbb{P}(W'_1 = n)} \leq \mathbb{E}[\tilde{N}_{\emptyset}].$$

This implies (42). □

5.3. The sub-critical case. Let p be a reproduction law. For every $\theta > 0$ such that $\sum_{k \in \mathbb{N}} \theta^k p(k)$ is finite, we define on \mathbb{N} the function p_{θ} by

$$\forall k \geq 0, p_{\theta}(k) = \begin{cases} c_{\mathcal{A}}(\theta) \theta^k p(k) & \text{if } k \in \mathcal{A}, \\ \theta^{k-1} p(k) & \text{if } k \notin \mathcal{A} \end{cases}$$

where the normalizing constant $c_{\mathcal{A}}(\theta)$ is given by:

$$c_{\mathcal{A}}(\theta) = \frac{1 - \sum_{k \notin \mathcal{A}} \theta^{k-1} p(k)}{\sum_{k \in \mathcal{A}} \theta^k p(k)}.$$

We denote by I the set of θ such that p_{θ} defines a probability distribution on \mathbb{N} . Notice that I contains $(0, 1]$.

Proposition 5.6. *Let τ be a GW tree with reproduction law p satisfying $p(0) > 0$ and $p(0) + p(1) < 1$. For every $\theta \in I$, let τ_{θ} be a GW tree with reproduction law p_{θ} . Then the conditional distributions of τ given $\{\mathcal{L}_{\mathcal{A}}(\tau) = n\}$ and of τ_{θ} given $\{\mathcal{L}_{\mathcal{A}}(\tau_{\theta}) = n\}$ are the same.*

Remark 5.7. This proposition covers Kennedy's result [8] for $\mathcal{A} = \mathbb{N}$ and the pruning procedure of [1] for $\mathcal{A} = \{0\}$.

Proof. Let $\mathbf{t} \in \mathbb{T}_0$. Then we have, using the definition of p_θ and (2):

$$\begin{aligned} \mathbb{P}(\tau_\theta = \mathbf{t}) &= \prod_{v \in \mathbf{t}} p_\theta(k_v(\mathbf{t})) \\ &= \prod_{v \in \mathbf{t}, k_v(\mathbf{t}) \in \mathcal{A}} c_{\mathcal{A}}(\theta) \theta^{k_v(\mathbf{t})} p(k_v(\mathbf{t})) \prod_{v \in \mathbf{t}, k_v(\mathbf{t}) \notin \mathcal{A}} \theta^{k_v(\mathbf{t})-1} p(k_v(\mathbf{t})) \\ &= c_{\mathcal{A}}(\theta)^{L_{\mathcal{A}}(\mathbf{t})} \theta^{\sum_{v \in \mathbf{t}} k_v(\mathbf{t}) - L_{\mathcal{A}^c}(\mathbf{t})} \mathbb{P}(\tau = \mathbf{t}) \\ &= c_{\mathcal{A}}(\theta)^{L_{\mathcal{A}}(\mathbf{t})} \theta^{\text{Card}(\mathbf{t}) - 1 - L_{\mathcal{A}^c}(\mathbf{t})} \mathbb{P}(\tau = \mathbf{t}) \\ &= \theta^{-1} (\theta c_{\mathcal{A}}(\theta))^{L_{\mathcal{A}}(\mathbf{t})} \mathbb{P}(\tau = \mathbf{t}). \end{aligned}$$

We deduce that

$$\begin{aligned} \mathbb{P}(L_{\mathcal{A}}(\tau_\theta) = n) &= \sum_{\mathbf{t} \in \mathbb{T}_0, L_{\mathcal{A}}(\mathbf{t}) = n} \mathbb{P}(\tau_\theta = \mathbf{t}) \\ &= \theta^{-1} (\theta c_{\mathcal{A}}(\theta))^n \sum_{\mathbf{t} \in \mathbb{T}_0, L_{\mathcal{A}}(\mathbf{t}) = n} \mathbb{P}(\tau = \mathbf{t}) \\ &= \theta^{-1} (\theta c_{\mathcal{A}}(\theta))^n \mathbb{P}(L_{\mathcal{A}}(\tau) = n) \end{aligned}$$

and finally, for every $\mathbf{t} \in \mathbb{T}_0$ such that $L_{\mathcal{A}}(\mathbf{t}) = n$, we have

$$\begin{aligned} \mathbb{P}(\tau_\theta = \mathbf{t} \mid L_{\mathcal{A}}(\tau_\theta) = n) &= \frac{\mathbb{P}(\tau_\theta = \mathbf{t})}{\mathbb{P}(L_{\mathcal{A}}(\tau_\theta) = n)} \\ &= \frac{\theta^{-1} (\theta c_{\mathcal{A}}(\theta))^n \mathbb{P}(\tau = \mathbf{t})}{\theta^{-1} (\theta c_{\mathcal{A}}(\theta))^n \mathbb{P}(L_{\mathcal{A}}(\tau) = n)} = \mathbb{P}(\tau = \mathbf{t} \mid L_{\mathcal{A}}(\tau) = n). \end{aligned}$$

□

Corollary 5.8. *Let τ be a sub-critical GW tree with reproduction law p satisfying Assumption (5). For every $\theta \in I$, let τ_θ be a GW tree with reproduction law p_θ . If there exists $\theta_0 \in I$ such that p_{θ_0} is critical, then*

$$\text{dist}(\tau \mid L_{\mathcal{A}}(\tau) = nd_{\mathcal{A}} + 1) \longrightarrow \text{dist}(\tau_{\theta_0}^*).$$

and

$$\text{dist}(\tau \mid L_{\mathcal{A}}(\tau) \geq n) \longrightarrow \text{dist}(\tau_{\theta_0}^*).$$

6. CONDITIONING BY THE SIZE OF A HIGH GENERATION

We end this paper with a conditioning which does not enter into the framework of Theorem 3.1. However its proof can be easily adapted. For a tree \mathbf{t} , we denote by

$$\mathcal{G}_n(\mathbf{t}) = \text{Card}(\{u \in \mathbf{t}, |u| = n\})$$

the size of the n -th generation of \mathbf{t} . Then we have

Proposition 6.1. *Let τ be a critical GW tree with reproduction law p satisfying Assumption (5). Let $(\alpha_n, n \in \mathbb{N})$ be a sequence of positive integers. If*

$$(45) \quad \lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\mathcal{G}_{n+1}(\tau) = \alpha_n)}{\mathbb{P}(\mathcal{G}_n(\tau) = \alpha_n)} = 1,$$

then we have

$$(46) \quad \text{dist}(\tau | \mathcal{G}_n(\tau) = \alpha_n) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

Proof. For every tree $\mathbf{t} \in \mathbb{T}_0$, every $x \in \mathcal{L}_0(\mathbf{t})$ and every tree $\tilde{\mathbf{t}} \in \mathbb{T}$, we have

$$\mathcal{G}_n(\mathbf{t} \circledast (\tilde{\mathbf{t}}, x)) = \mathcal{G}_n(\mathbf{t}) + \mathcal{G}_{n-|x|}(\tilde{\mathbf{t}})$$

which generalizes Assumption (9).

The same computations as in the proof of Theorem 3.1 give for $\mathbf{t} \in \mathbb{T}_0$, $x \in \mathcal{L}_0(\mathbf{t})$ and $n \geq H(\mathbf{t})$:

$$\begin{aligned} \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x), \mathcal{G}_n(\tau) = \alpha_n) &= \frac{1}{p(0)} \mathbb{P}(\tau = \mathbf{t}) \mathbb{P}(\mathcal{G}_{n-|x|}(\tau) = \alpha_n - \mathcal{G}_n(\mathbf{t})) \\ &= \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) \mathbb{P}(\mathcal{G}_{n-|x|}(\tau) = \alpha_n). \end{aligned}$$

Therefore, we obtain by Assumption (45):

$$\begin{aligned} \lim_{n \rightarrow +\infty} \mathbb{P}(\tau \in \mathbb{T}(\mathbf{t}, x) | \mathcal{G}_n(\tau) = \alpha_n) &= \lim_{n \rightarrow +\infty} \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)) \frac{\mathbb{P}(\mathcal{G}_{n-|x|}(\tau) = \alpha_n)}{\mathbb{P}(\mathcal{G}_n(\tau) = \alpha_n)} \\ &= \mathbb{P}(\tau^* \in \mathbb{T}(\mathbf{t}, x)). \end{aligned}$$

The result follows from Lemma 2.1. \square

Corollary 6.2. *Let τ be a critical GW tree with reproduction law p given by a mixture of a geometric distribution with parameter $\rho \in (0, 1]$ and a Dirac mass at 0, i.e. $p(0) = 1 - \rho$ and $p(k) = \rho^2(1 - \rho)^{k-1}$ for $k \geq 1$. Let $(\alpha_n, n \in \mathbb{N})$ be a sequence of positive integers such that $\lim_{n \rightarrow +\infty} n^{-2}\alpha_n = 0$. Then we have:*

$$\text{dist}(\tau | \mathcal{G}_n(\tau) = \alpha_n) \xrightarrow{n \rightarrow +\infty} \text{dist}(\tau^*).$$

Proof. In that particular case, the generating function φ_n of $\mathcal{G}_n(\tau)$ is explicitly known and we have for every $s \in [0, 1]$

$$\varphi_n(s) = \frac{nc - (nc - 1)s}{(nc + 1) - ncs}$$

with $c = (1 - \rho)/\rho$. Expanding φ_n gives for every $k \geq 1$:

$$\mathbb{P}(\mathcal{G}_n(\tau) = k) = \frac{(nc)^{k-1}}{(nc + 1)^{k+1}},$$

and therefore

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{P}(\mathcal{G}_{n+1}(\tau) = \alpha_n)}{\mathbb{P}(\mathcal{G}_n(\tau) = \alpha_n)} = \lim_{n \rightarrow +\infty} \frac{n(nc + 1)}{(n + 1)((n + 1)c + 1)} \left(\frac{1 + \frac{1}{nc}}{1 + \frac{1}{(n+1)c}} \right)^{\alpha_n} = 1.$$

\square

Remark 6.3. As for Theorem 3.1, we can obtain the converse of Proposition 6.1. We deduce that, in the geometric case, the Galton-Watson tree τ conditioned on $\{\mathcal{G}_n(\tau) = j \lfloor n^a \rfloor\}$, with $j \in \mathbb{N}^*$, converges in distribution to Kesten's tree if and only if $a \in [0, 2)$.

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