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Inmaculada Martínez-Zarzoso

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The Log of Gravity Revisited

By

Inmaculada Martínez-Zarzoso ^{a)}

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^{a)} Department of Economics and Ibero-America Institute for Economic Research, University of Goettingen, Germany, and Department of Economics and Instituto de Economía Internacional, Universidad Jaume I, Castellón, Spain. Financial support from the Spanish Ministry of Science and Innovation is gratefully acknowledged (ECO2010-15863). E-mail: martinei@eco.uji.es.

The Log of Gravity Revisited

Abstract

This paper evaluates the performance of alternative estimation methods for gravity models with heteroskedasticity and zero trade values. Both problematic issues, recently addressed by Santos Silva and Tenreyro in an influential paper, are re-examined here. We use Monte Carlo simulations to compare the Pseudo Poisson Maximum Likelihood (PPML) estimator recommended by Santos Silva and Tenreyro, a Gamma pseudo-maximum-likelihood (GPML), a Non-Linear Least Squares (NLS) estimator and a Feasible Generalized Least Squares (FGLS) estimator with more traditional techniques. Additionally, estimates of the gravity equation are obtained for three different data sets with the abovementioned methods. The results of the simulation study indicate that, although the PPML estimator is less affected by heteroskedasticity than others are, its performance is similar, in terms of bias and standard errors, to the FGLS estimator performance, in particular for small samples. GPML presents however the lowest bias and standard errors in the simulations without zero values. The results of the empirical estimations, using three different samples containing real data, indicate that the choice of estimator has to be made for each specific dataset. There is not a general “best” estimator and it is highly recommended to follow a model selection approach using a number of tests to select the more appropriate estimator for any application.

JEL classification: C5, F10

Keywords: Simulations; Poisson regression; constant-elasticity models; heteroskedasticity; Feasible Generalized Least Squares (FGLS); maximum likelihood

1. Introduction

Trade patterns can be extremely well described by the gravity model. The model, in its simplest form, states that bilateral trade between two countries is directly proportional to the product of the countries' income and negatively related to the distance between them. This model has been very successful empirically, mainly due to its high explanatory power. The most common practice in empirical applications has been to transform the multiplicative model by taking natural logarithms and to estimate the obtained log-linear model using Ordinary Least Squares (OLS)¹. Despite the extensive use of this practice, some econometric problems arise from the prevalence of zero values in trade flows and from heteroskedastic residuals.

Santos Silva and Tenreyro (2006, 2011) propose a Pseudo Poisson Maximum Likelihood (PPML) estimation technique that is consistent in the presence of heteroskedasticity and provides a way of dealing with zero values of the dependent variable. PPML is a special case of the generalized linear model (GLM) framework, in which the variance is assumed to be proportional to the mean. The authors recommend estimating constant-elasticity models using the PPML estimator, instead of applying traditional OLS estimation techniques to the log-linearized models. In their paper, the gravity equation for international trade is taken as an example of constant-elasticity models. Although the general practice has been to estimate the log-linearized version of the gravity model², in the related empirical literature, we can also find several attempts to deal separately with heteroskedasticity (e. g., Porojan, 2001) and zero values (e. g., Helpman, Melitz and Rubinstein, 2008).

Martin and Pham (2008) state that the PPML estimator is less subject to bias resulting from heteroskedasticity, but does not prove robust to the joint problems of heteroskedasticity and zero trade flows. Therefore, it could be appropriate for other typical applications of

¹ See Oguledo, V.I. and Macphee (1994) for a review of gravity studies using aggregated trade and Martínez-Zarzoso, Pérez-García and Suárez Burguet (2008) for an application using sectoral trade data.

²In Chapter 5, Feenstra (2004) presents a revision of articles applying the gravity model for international trade.

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3 multiplicative models with relatively few zero observations, such as the estimation of
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5 consumer-demand systems, Cobb-Douglas-type production functions and the STIRPAT
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7 (Stochastic Impacts by Regression on Population, Affluence and Technology) model in
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9 environmental economics.
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13 In this paper we consider cases with none or with a low percentage of zero values in the
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15 dependent variable and focus on the heteroskedasticity problem. Given the extended use of
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17 the Feasible Generalized Least Squares (FGLS) estimator to correct for heteroskedastic errors,
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19 it seems important to examine and compare its performance with the PPML estimator and
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21 with alternative GLM.
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25 The novelties of the present research are threefold. First, we argue that the performance of the
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27 PPML estimator should be also compared to FGLS techniques applied to the linearized model
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29 and not only to OLS techniques. Second, we consider two different ways of dealing with zero
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31 values of the dependent variable; we generate a dependent variable that becomes zero with a
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33 certain probability (p) and that remains continuous with probability $(1-p)$, and zeros are
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35 randomly created depending on a threshold value of 1 of the regressors to imitate some
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37 minimum conditions that are necessary for a positive trade value. Deviating from Santos Silva
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39 and Tenreyro (2006, 2011) and Martin and Pham (2008), we will evaluate the performance of
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41 the estimators not only by looking at the bias of the estimates, but also by computing their
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43 expected loss. Third, estimates of the gravity equation are obtained for three different data
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45 sets with two classes of estimators: traditional methods derived from least squares estimators
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47 for the log-linearized model and GLM for the unlogged model. For completeness we also
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49 estimate a Heckman sample-selection model. In the context of the gravity model of trade, the
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51 presence of zeros in the dependent variable is due primarily to the absence of trade, rather
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53 than to missing values. This presence raises a problem of selection bias. Heckman-type
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55 models are able to tackle the problem of sample selection bias (there are different motives for
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3 participating or not participating in a certain economic activity that have to be identified
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5 before running the regression for the agents participating).
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10 In line with Santos Silva and Tenreyro, we use Monte Carlo simulations to evaluate the
11 performance of alternative estimation methods (OLS, FGLS, Nonlinear Least Squares (NLS),
12 Gamma Pseudo Maximum Likelihood (GPML) and PPML) for multiplicative and log models
13 with heteroskedasticity. The results of the simulation study indicate that although the Pseudo
14 Poisson Maximum Likelihood estimator is less affected by heteroskedasticity than other
15 estimators, its performance is similar, in terms of bias and standard errors, to the FGLS
16 estimator performance.
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27 In addition, the results of the empirical estimations, using three different samples containing
28 real data indicate that the choice of estimator has to be made for each specific dataset. There
29 is not a general “best” estimator and it is highly recommended to follow a model selection
30 approach using a number of tests to select the more appropriate estimator for any application.
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38 The next section focuses on a discussion of the PPML and FGLS estimation techniques and
39 their consistency in the presence of heteroskedasticity. Section 3 presents the simulation
40 results in the presence of homoskedasticity and heteroskedasticity, and with a dependent
41 variable that can take on zero values. Section 4 applies the PPML and other alternative
42 estimation techniques to real data and discusses the estimation results. In Section 5, some
43 conclusions are drawn.
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51 **2. Alternative Estimators**

52 Let us assume that the original stochastic model is given by

$$53 y_i = \exp(x_i \beta) \varepsilon_i = \mu(x \beta) \varepsilon_i, \quad (1)$$

54 where $E(\varepsilon_i | x) = 1$.
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Assuming that y_i is positive, the model can be linearized by taking logs:

$$\ln y_i = x_i \beta_i + \ln \varepsilon_i, \quad (2)$$

where $\ln E(\varepsilon_i | x) = 0$; $E(\ln \varepsilon_i | x) \neq 0$.

The original multiplicative model given by Equation 1 can be directly estimated using NLS or maximum likelihood techniques, whereas Model 2 can be estimated simply by OLS. The inequality $E(\ln \varepsilon_i | x) \neq 0$ is known as Jensen's inequality. In the presence of heteroskedasticity, Least Squares (LS) estimation is no longer efficient. This problem can be resolved by controlling for heteroskedasticity.

The estimation technique proposed by Santos Silva and Tenreyro (2006) is a pseudo maximum likelihood estimator based on some assumptions about the functional form of the conditional variance. Under the assumption that the conditional variance is proportional to the conditional mean, β can be estimated by solving a set of first-order conditions:

$$\sum_{i=1}^n [y_i - \exp(x_i \tilde{\beta})] x_i = 0, \quad (3)$$

where y_i is the dependent variable, x_i are the explanatory variables, and β are the parameters to be estimated. The estimator based on Equation 3 gives the same weight to all observations, whereas the LS and NLS estimators give more weight to observations with large $\exp(x_i \beta)$.

The data do not have to follow the Poisson distribution and the dependent variable does not have to be an integer. Since the assumption of proportionality between the conditional variance and the conditional mean does not always hold, inference should instead be based on a robust covariance matrix estimator that specifically corrects for heteroskedasticity in the model.

Within the GLM framework it is common to consider a more general form of the variance function given by

$$v(y | x) = k(\mu(x\beta))^\lambda \quad (4)$$

where λ is finite and non-negative and its value determines different GLM families. We obtain the NLS estimator when $\lambda=0$. We obtain the Poisson class when $\lambda=1$. We obtain the

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3 gamma, the homoskedastic log-normal, the Weibull and the chi square in the case $\lambda=2$ with
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5 the appropriate specification of a distribution. Finally, when $\lambda=3$, the Inverse-Gaussian family
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7 is obtained.
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10 As mentioned above, the traditional way to estimate constant-elasticity models is to linearize
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12 the original multiplicative model using a log-log transformation. The rationale for using a log-
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14 log model can come from a desire to generate an estimate that easily yields elasticities or from
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16 a need to deal with dependent variables that are skewed to the right. Log transformation also
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18 improves precision and reduces the influence of outliers on the estimates. However, since no
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20 one is interested in log model results in the log scale per se, those results have to be
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22 retransformed to the original scale in order to be able to quantify the average response to a
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24 covariate (Manning, 1998).
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29 Within the class of least squares estimators, the FGLS estimator is efficient in the presence of
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31 heteroskedasticity and can be applied to the linearized model. The FGLS estimator weighs the
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33 observations according to the square root of their variances³, and is given by
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35

$$\beta_{FGLS} = (x' \hat{\Omega}^{-1} x)^{-1} x' \hat{\Omega}^{-1} y, \quad (5)$$

36
37 where $\hat{\Omega}$ is the weighting matrix. $\hat{\Omega}$ is estimated using the OLS squares residuals (\hat{u}_i), to
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39 obtain an estimate for the variance of \hat{u}_i ($\widehat{\text{var}} \hat{u}_i$). Then, $\hat{\Omega} = 1/\sqrt{\widehat{\text{var}} \hat{u}_i}$.
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45 It can be shown that even if the weights used in FGLS estimation are biased (resulting in a
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47 biased estimation of the residuals variance), FGLS would still provide consistent estimates
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49 when the appropriate retransformation technique is used⁴. In case of an unknown form of
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51 heteroskedasticity, FGLS can be applied and the variance of the disturbances must be
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53 estimated. This method should be well suited to estimating regression coefficients in the
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³ Heteroskedasticity is assumed to be a linear function of the explanatory variables, as in Wooldridge (2009), pages 282-283.

⁴ With transformed dependent variables, the mean response is the mean of the retransformed estimate of y, which depends not only on the mean of the x's, but also on their distribution (Manning, 1998).

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3 presence of heteroskedasticity. Henceforth, the comparison should be made between FGLS-
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5 like estimators and GPML, NLS or PPML estimation.
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8 The main difficulty encountered when dealing with log dependent variables is that log scale
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10 results are of little interest. Most economic policy decisions are taken based on monetary
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12 values requiring therefore retransformation of the log scaled predictions. Retransformation
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14 does not present any problem when errors are normal and homoskedastic, however,
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16 retransformation biases arise when one of these do not hold. A common practice in health
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18 economics has been to switch to GLM from log-transformed models when errors are
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20 heteroskedastic. However, although GLM can deal with heteroskedastic errors and do not lead
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22 to retransformation bias, OLS/FGLS of log-transformed variables could be more efficient in
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24 some cases (Manning and Mullahy, 2001). GLM can suffer substantial precision losses if the
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26 log-scale residuals have high kurtosis or if the variance function is misspecified⁵. The key
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28 question that remains is how to determine which method is best for specific empirical
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30 applications. Manning and Mullahy (2001) developed a method for determining which
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32 estimation method to choose for any empirical application which is based on using a number
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34 of tests that are easy to implement.
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43 **3. Simulation Study**

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45 A data set with the same properties as described by Santos Silva and Tenreyro (2006) is
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47 generated to compare different estimation techniques in the presence of heteroskedasticity.
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49 We extend the procedures used by the authors in two ways. First, we add an alternative
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51 estimation technique, the FGLS estimator, which is applied to the linearized model to correct
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53 for heteroskedasticity. Second, instead of using a rounding approach to generate zeros in the
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55 dependent variable⁶, two alternative methods are proposed; one consists of setting
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⁵ However, estimates are still consistent.

⁶ As done in Santos Silva and Tenreyro (2006).

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3 observations to zero randomly and the other consists of setting observations to zero for
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5 particular groups.
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8 The multiplicative constant elasticity model considered is
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$$10 \quad y_i = \mu(x_i, \beta) \eta_i = (\tilde{\beta}_0 X_{1i}^{\beta_1} X_{2i}^{\beta_2}) \eta_i, \quad (6)$$

11
12 η_i is a log-normal random variable with mean 1 and variance σ_i^2 , in the absence of
13
14 heteroskedasticity. x_{1i} is log-normal and x_{2i} is dichotomous with values of e^0 and e^1 with a
15
16 probability of 0.6 for the first value and a probability of 0.4 for the second value.
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20 Equation 5 can also be expressed as
21

$$22 \quad \mu(x_i, \beta) := E[y_i | x] = \exp(\ln \tilde{\beta}_0 + \beta_1 \ln x_{1i} + \beta_2 \ln x_{2i}), \quad (7)$$

23
24 where $X_{1i} := \ln x_{1i}$; $X_{2i} := \ln x_{2i}$ and $\ln \tilde{\beta}_0 := \beta_0$.
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28 From distributional theory, it follows that X_{1i} is drawn from a standard normal distribution
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30 and X_{2i} is a binary variable that takes the value 1 with a probability of 0.4. The two covariates
31
32 are independent and the true values for the coefficients are: $\beta_0 = 0$, $\beta_1 = 1$ and $\beta_2 = 1$.
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36 In line with Santos Silva and Tenreiro (2006), we assess the performance of different
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38 estimators (NLS, GPML, PPML, OLS and FGLS)⁷ under homoskedasticity and
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40 heteroskedasticity. We can distinguish four cases:
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$$43 \quad \text{Case 1: } V(\eta_i) = \mu(x_i, \beta)^{-2}; V[y_i | x] = 1$$

$$44 \quad \text{Case 2: } V(\eta_i) = \mu(x_i, \beta)^{-1}; V[y_i | x] = \mu(x_i, \beta)$$

$$45 \quad \text{Case 3: } V(\eta_i) = 1; V[y_i | x] = \mu(x_i, \beta)^2$$

$$46 \quad \text{Case 4: } V(\eta_i) = \exp(x_{2i}) + \mu(x_i, \beta)^{-1}; V[y_i | x] = \mu(x_i, \beta) + \exp(x_{2i}) \mu(x_i, \beta)^2$$

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48 In Case 1, the NLS assumptions hold. In Case 2, the conditional variance equals its
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50 conditional mean, as in the PPML assumptions. OLS, FGLS and GPML conditions are
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52 fulfilled in Case 3 (homoskedasticity). And finally, in Case 4, the conditional variance does
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⁷ Tobit and truncated OLS simulation results are not presented since these methods show a very poor performance.

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3 not only depend on the mean, but is also a function of one of the explanatory variables
4 (heteroskedasticity). We argue that in Case 4 of heteroskedasticity, which was identified as a
5 severe problem in many applications (e.g., when analyzing trade flows in the framework of
6 the gravity model), FGLS, rather than OLS, should be used for comparison purposes.
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11 In a first set of simulations, we study the performance of the estimators for the different
12 models considered and for the four specifications of the disturbances outlined above. We
13 focus on the estimators with a good performance in the experiments of Santos Silva and
14 Tenreyro (2006), add the FGLS estimator and disregard those with a poor performance
15 (Tobit, OLS (y+1), truncated-OLS).
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24 The truncated and the censored OLS models (in logarithms) produce results that are inferior
25 to those obtained from the classical OLS model (that disregards the observation with zero
26 values) in terms of bias and standard errors. This result is consistent with Martin and Pham
27 (2008).
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34 Initially, we also considered the Tobit estimator which applies to situations in which
35 outcomes can only be observed over some range, either because actual outcomes cannot
36 reflect desired outcomes or because of measurement inaccuracy (rounding). However, the
37 gravity equation only predicts zero trade if incomes of one or both countries are zero;
38 therefore, desired outcomes are always positive. In addition, trade values are reported in
39 COMTRADE even for very small amounts (up to \$1). Therefore, zero flows seem to result
40 from a binary decision-making process, rather than from censoring. This is also in accordance
41 with the Helpman et al. (2008) model.
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53 Deviating from Santos Silva and Tenreyro (2006) and Martin and Pham (2008), we will
54 evaluate the performance of the estimators not only by looking at the bias of the estimates, but
55 also their expected loss. Most economic simulation studies consider unbiasedness or a small
56 bias to be the most desirable property of an estimator. Therefore, the bias is used as the main
57 criterion to compare the quality of different estimators. But this approach could be misleading
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3 in some cases since –due to the fact that over- and under-estimations cancel each other out–
4 unbiased estimators are not necessarily good estimators⁸. A feature of our results for some of
5 the considered estimators is the wide variation in the direction of the bias in parameter
6 estimates; for example, the bias in the estimate of the slope coefficients resulting from use of
7 the PPML estimator is not consistently negative. For this reason, although the bias is
8 considered to be very low in some of the four cases, the expected loss is not. By considering
9 the expected loss, the direction of the bias is also taken into account in evaluating the
10 performance of alternative estimators.
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13 In statistical decision theory, one therefore looks at the risk of an estimator, defined as its
14 expected loss:
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16

$$17 \quad R(\beta, \hat{\beta}) = E_{\beta} L(\beta, \hat{\beta}). \quad (8)$$

18 The way the loss function $L(\beta, \hat{\beta})$ is defined depends on the individual needs in each
19 statistical analysis. As argued above, we only consider the class of loss functions where over-
20 and underestimations cannot cancel out, as in case of the bias. The simplest loss function
21 which is adequate for our purposes is the absolute error loss:
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$$24 \quad L(\beta, \hat{\beta}) = |\beta - \hat{\beta}|. \quad (9)$$

25 As an alternative loss function, the squared error loss could also be used.
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28 In a second set of experiments, we study how zero values in the dependent variable affect the
29 performance of the estimators, using a data-generation mechanism that produces a percentage
30 of observations with zero values in y_i , but that differs from those used by Santos Silva and
31 Tenreyro (2006)⁹ and Martin and Pham (2008). In Santos Silva and Tenreyro (2006), the
32 zeros in the dependent variable are generated by rounding to the nearest integer the values of
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58 ⁸ We give a simple example to illustrate this: let the true parameter be $\beta = 1$. We define an estimator $\hat{\beta} = 0$
59 half of the time and $\hat{\beta} = 2$ the other half of the time. Clearly, this is a very bad estimator, but it is unbiased.
60 The constant estimator $\hat{\beta} = 1.1$, in contrast, is biased but obviously better than the above estimator.

⁹ In Santos Silva and Teneyro (2006), the dependent variable was generated by rounding the values of y_i to the nearest integer.

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3 y_i obtained in the first set of simulations. The authors mention that zeros could be a result of
4 rounding errors or of missing observations that are wrongly recorded as zeros. As they
5 recognize, this procedure will bias the estimates and the purpose of the experiment is just to
6 measure the magnitude of the bias. To check the robustness of the results we suggest a
7 slightly different way to investigate cases of zero trade or related problems, namely setting a
8 certain percentage of the dependent variable to zero. We do this in two different ways: first,
9 we choose a certain percentage, say 15 percent¹⁰, of all observations and set it to zero, and
10 second, we only create zeros in a certain group, assuming that there is an underlying pattern
11 that is responsible for this, say half of the 30 percent poorest/smallest countries, according to
12 one or more of the independent variables. This pattern is reasonable for many relevant real
13 cases, including the gravity model, since zero trade mainly occurs among poor or small
14 countries.

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Tables 1 and 2 summarize simulation results for sample sizes of 1,000 and 10,000
replications. The expected loss, the bias and the standard error of the two parameters of
interest (β_1 and β_2) are presented. Table 1 shows the results for the first set of simulations.
The NLS shows a lower bias and expected loss than the other estimators only in Case 1. In
Cases 3 and 4, the expected loss and bias of the NLS estimator are the highest. The PPML is
the best in Case 2, in which the underlying error structure is consistent with the PPML
estimator. In Cases 3 and 4, the performance of FGLS and PPML are comparable. FGLS
estimates display lower expected loss, biases and standard errors than PPML in Case 3. The
bias in the estimate of β_1 is only -0.0003 for FGLS and -0.0027 for PPML. For β_2 the
corresponding biases are 0.0001 and -0.0008. In Case 4, FGLS estimates also present lower
expected loss and standard errors for β_1 and β_2 , but PPML shows slightly lower biases.

Table 1. Simulation results, regular case

¹⁰ Bilateral trade among developed countries presents this pattern. Around 10 to 15 percent of the bilateral flows are zeros.

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3 Table 2 shows the results for the second set of simulations, in which the dependent variable is
4 generated with a percentage of zero values (15 percent). The NLS estimator yields estimates
5 that are strongly biased in Cases 3 and 4. Traditional OLS estimates also present large biases
6 and expected losses in Cases 1, 2 and 4. However, correcting for heteroskedasticity using
7 FGLS considerably reduces the expected losses and biases in Cases 1, 2 and 4. When the zero
8 values are generated randomly, the Gamma and the FGLS perform quite well in Cases 3 and
9 4, and both present lower expected losses and standard errors than the PPML estimation.
10
11 However, when the zero values are generated with a given pattern, the Gamma estimator yield
12 estimates that are strongly biased in all cases, especially for β_1 , which bias is around 25
13 percent in all four cases. FGLS shows lower expected losses and standard errors than PPML
14 in Cases 2, 3 and 4. Also in terms of bias, the bias in the estimate of β_1 for FGLS is always
15 lower than for PPML and it is lower than 1 percent in Cases 2 and 3 and around 5 percent for
16 Cases 1 and 4.

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34 *Table 2. Simulation results with zeros*

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36 Additionally, we repeated the simulations with higher percentages of zero values (40 percent
37 and 60 percent). The results for FGLS and PPML are shown in the appendix. A similar
38 pattern persists for the FGLS estimator, with β_1 biases always lower than 2 percent for Cases
39 2 and 3, and around 7 percent and 5 percent, respectively, for Cases 1 and 4. The β_2 biases are
40 consistently lower than 1 percent for Cases 3 and 4, and around 5 percent and 8 percent,
41 respectively, for Cases 1 and 4. The biases of β_1 and β_2 for PPML increase with higher
42 percentages of zero values¹¹, and more rapidly than for FGLS. The expected losses are around
43 10 percent for β_1 and around 15 percent for β_2 in Cases 3 and 4, with 40 percent being zero
44 values.

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¹¹ Martin and Pham (2008), using different data-generating processes, also show that the PPML estimator is highly vulnerable to the presence of a high percentage of zero values in the dependent variable. However, these results have been challenged by Santos Silva and Tenreiro (2011).

As a robustness check, and as suggested by Santos Silva (personal communication), we increased the sample size to 10,000 observations, which amount is consistent with some of the cross-sectional analysis in the framework of the gravity equation. Whereas the bias of the PPML decreases, the FGLS bias remains almost constant. The standard errors of the PPML estimator is considerably lower, but still twice the FGLS standard errors. Therefore, using PPML when the sample is large and when there is measurement error in the dependent variable could be recommendable. However, for small samples, FGLS could be a perfect choice to deal with the problem of heteroskedastic disturbances.

4. The Gravity Model: Estimation Results

4.1 Original versus log-log version of the gravity model

According to the generalized gravity model of trade, the volume of exports between pairs of countries in year t , X_{ijt} is a function of their incomes (Y_{it} , Y_{jt}), their incomes per capita (YH_{it} , YH_{jt}), their geographical distance ($DIST_{ij}$) and a set of dummies that represents any other factors aiding or preventing trade between pairs of countries (F_{ij}):

$$X_{ijt} = \beta_0 Y_{it}^{\beta_1} Y_{jt}^{\beta_2} YH_{it}^{\beta_3} YH_{jt}^{\beta_4} DIST_{ij}^{\beta_5} F_{ij}^{\beta_6} u_{ijt}, \quad (10)$$

where u_{ijt} is the error term.

Equation (10) in its multiplicative form can be directly estimated using GLM. GLM assumes that there is a function that explains the relationship between the variance and the mean (see equation 4 above) and does not assume constant variance.

However, the most prevalent modelling approach is to use OLS or similar techniques to the log-linearized model. Taking natural logarithms from Equation 10 and replacing F_{ij} for a number of specific variables, the log-log model is given by

$$\begin{aligned} \ln X_{ijt} = & \alpha_{ij} + \beta_1 \ln Y_{it} + \beta_2 \ln Y_{jt} + \beta_3 \ln YH_{it} + \beta_4 \ln YH_{jt} + \\ & + \beta_5 \ln DIST_{ij} + \beta_6 \ln Border_{ij} + \beta_7 \ln Comlang_{ij} + \beta_8 \ln Z_{ijt} + \sum_k \lambda_k \ln FTA_{ij} + \mu_{ijt} \end{aligned} \quad (11)$$

where:

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3 L denotes variables in natural logs,
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5 X_{ijt} are the exports from country i to country j in period t in current US\$,
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7 Y_{it} , Y_{jt} indicate the GDP of countries i and j , respectively, in period t in current PPP US\$,
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10 YH_{it} , YH_{jt} denote incomes per capita of countries i and j , respectively, in current PPP US\$ per
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12 thousand inhabitants in period t ,
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15 $DIST_{ij}$ is the Great Circle distance between countries i and j .
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18 $Border$ is a dummy that takes the value of 1 when countries i and j share a border, and
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21 $Comlang$ is a dummy that takes the value of 1 when countries i and j speak the same
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23 language.
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26 Z are additional variables that could foster or prevent trade between pairs of countries and
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28 could be, in some cases, time variant, for example, a dummy that takes the value of 1 if the
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30 trading countries share a common currency, or a dummy that takes the value of 1 if they had a
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32 colonial relationship in the past, 0 otherwise.
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35 The model includes trading blocs' dummy variables, defined as FTA_{ij} , which evaluate the
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37 effects of free trading agreements (FTAs). Integration dummies are described in more detail
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39 below. α_{ij} are the specific effects associated with each bilateral trade flow. They are a control
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41 for all the omitted variables that are specific for each trade flow and that are time invariant.
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44 A high level of income in the exporting country indicates a high level of production, which
45
46 increases the availability of goods for export. Therefore, we expect β_1 to be positive. The
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48 coefficient of Y_j , β_2 , is also expected to be positive since a high level of income in the
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50 importing country attracts higher imports. The coefficient for income per capita of the
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52 exporters, β_3 , may be negatively or positively signed depending on the type of goods
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54 exported. The coefficient of the importer income per capita, β_4 , also has an ambiguous sign,
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56 for similar reasons. Another factor that may influence these coefficient estimates is the
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58 composition effect that influences supply and demand. Each country produces and exports a
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60 different mix of commodities (supply) and the mix of goods demanded is also different for

each country. The distance coefficient is expected to be negative since it is a proxy of all possible trade costs.

In some of the results presented below, specific trading bloc dummies are included to model trade-creation and trade-diversion effects of trade (see Soloaga and Winters (2001); Chen and Tsai (2005), Carrère (2005) and Martínez-Zarzoso, Nowak-Lehmann and Horsewood (2009)).

The specification of the Viner's trade creation and trade diversion for a single period is given by

$$LX_{ij} = EV_{ij} + \sum_k \gamma_k D_k + \sum_k \delta_k D_{ki} + \sum_k \rho_k D_{kj} \quad , \quad (12)$$

where LX_{ij} denotes exports from country i to country j in natural logarithms, and EV_{ij} is defined as the rest of explanatory variables of the gravity equation above. D_k is a dummy that takes the value 1 if both countries, i and j , belong to the same economic bloc, 0 otherwise. D_{ki} is a dummy that takes the value 1 if i is a member of bloc k and j belongs to the rest of the world, 0 otherwise. D_{kj} is a dummy that takes the value 1 if j is a member of bloc k and i belongs to the rest of the world, 0 otherwise. γ_k measures the extent to which trade exceeds normal levels if both countries, i and j , are members of the bloc, δ_k measures the extent to which members' exports levels to non-member countries are higher than normal, and ρ_k measures the extent to which members' import levels from non-member countries are higher than normal. δ_k and ρ_k could be interpreted as a measure for trade-diversion effects, but they might also combine trade-diversion and openness effects.

4.2 Main results

Different versions of Models 10 and 11 are estimated using two classes of estimators: least squares for models with logged dependent variables, and GLM with a logarithmic function.

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3 Within the first class we consider four estimators: OLS, FGLS, Harvey model¹² and the
4 Heckman selection model. Within the second class, we use tree estimators: PPML, GPML
5 and NLS. Observations with zero values are simply dropped when estimate the OLS, FGLS
6 and Harvey models. The FGLS estimator is computed two steps. The squared OLS residuals
7 are computed (e^2) and regressed upon the gravity-model explanatory variables and a constant.
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10 Then $1/\sqrt{\hat{e}^2}$ are the weights used in the second step weighted least squares regression.
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13 The gravity model is estimated for three different data sets. The first one is a sample of 180
14 countries over the period from 1980 to 2000 from Rose (2005). The second data set consists
15 of a sample of 65 countries from Márquez, Martínez and Suárez (2007), with data for every
16 five years over the period from 1980 to 1999. Finally, a third sample of 47 countries from
17 Martínez-Zarzoso et al. (2009) is used, covering the period from 1980 to 1999. The
18 percentage of zero export values in the three samples is 13, 15, and 23 percent, respectively.
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32 The traditional gravity model, as well as the theoretically justified gravity model (Anderson
33 and van Wincoop, 2003) with multilateral resistance terms specified as exporter and importer
34 dummy variables, are considered.
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38 The distribution of the dependent variable, total exports, in the three samples shows right-
39 skewness since there are no negative exports and some pair of countries do not trade to each
40 other. The summary statistics for exports in the first sample for the year 1990 are mean=
41 1.82e+08, variance= 3.29e+18, skewness= 28.59645, and kurtosis= 1155.022. An attempt to
42 make the distribution more normal consists on taking natural logarithms. The equivalent
43 summary statistics for log exports are mean= 14.17866, variance= 16.51199, skewness= -
44 .8381542, and kurtosis= 5.054987. Although transformation makes the distribution closer to a
45 normal, the usual test still rejects the normality assumption.
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¹² Harvey's model of multiplicative heteroskedasticity (Harvey, 1976) has also been estimated since it is a very flexible and model that includes most of the useful formulations as special cases. The general formulation is $\sigma_i^2 = \sigma^2 \exp(z_i' \alpha)$.

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3 In the other two samples with 65 and 47 countries respectively, whereas unlogged exports are
4 badly skewed to the right, the log of exports is approximately normal for 1990 with skewness
5 close to zero and kurtosis close to 3. According to Manning (1998), in such cases estimates
6 based on logged models are usually more precise and robust than direct analysis of the
7 unlogged original dependent variable. Retransformation presents no problem when errors are
8 linear, normal and homoskedastic. One of the main reasons to switching to GLM from log-
9 transformed models in applied research has been due to heteroskedasticity of unknown form.
10 But if errors are normally distributed, a transformation can still be applied that accounts for
11 any form of heteroskedasticity (Baser, 2007). We applied the method proposed by Baser
12 (2007) to calculate the retransformed expected value of exports.
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20 Table 3 shows the results for the traditional gravity model, as specified in Equations 9 and 10
21 for the first data set¹³. With comparative purposes, we show the results for the year 1990, as in
22 Santos Silva and Tenreyro (2006) (SST). However, the estimated parameters are not directly
23 comparable since the number of countries is different; as is the definition of the dependent
24 and some of the independent variables. Our purpose is only to present estimation results for
25 different data sets in order to evaluate/put into perspective the use of the PPML as the
26 proposed “new workhorse” for the estimation of gravity models¹⁴.
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Table 3. Estimation results: traditional gravity model

The estimated coefficients for the income variables are closer to the theoretical value of 1 when the equation is estimated using OLS and FGLS, Harvey and Heckman. The fact that the income elasticities for the exporter and importer countries are equal in magnitude, according to the PPML estimation, is not indicative of a correct specification. Asymmetries can also be theoretically justified. The standard errors are significantly lower for the least squares estimators and therefore the estimates are more accurate.

¹³ The theoretically modified gravity model was also estimated for this sample using OLS and FGLS techniques, but the PPML estimation presented some problems since the maximization algorithm could not find a solution.

¹⁴ SST page 649.

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3 In contrast to SST (2006), the variable *common border* is also statistically significant when
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5 PPLM is used, compared to OLS and FGLS estimates. The OLS and FGLS estimates for
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7 *FTA*, indicates that the integration dummy is statistically significant, whereas using PPML, it
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9 is non-significant and much lower in magnitude than when using least squares methods.
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11 Instead, the GPML model produces a significant coefficient for the *FTA* dummy. Concerning
12
13 geographical distance, in line with SST the estimated elasticity is significantly lower in
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15 magnitude in the PPLM estimation, whereas in the FGLS and Harvey model, it is lower than
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17 in OLS but higher than in PPML.
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22 Table 4 shows the results for the theoretically justified gravity model for the second and third
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24 data sets. The first three columns show the estimated coefficients for the OLS, FGLS and
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26 Harvey models. Column 4 shows the Heckman selection model results that control for zero
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28 values for the dependent variable. The last three columns present the results obtained using
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30 GLM (PPML, GPLM and NLS).
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34 Comparing the results of the different specifications, we observe for both samples that the
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36 role of distance is significantly reduced under Poisson and NLS. Trading countries having a
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38 common language always have a positive effect on trade, but the effect is also substantially
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40 reduced (halved) when using Poisson in comparison to OLS, FGLS, Harvey or Heckman. A
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42 common border consistently has a positive effect on exports in the first sample (65 countries)
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44 under all estimation techniques. However, for the second sample (47 countries) OLS and
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46 FGLS and Harvey predict no significant effects and Poisson produces a positive and
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48 significant effect. The coefficient is not shown in the fourth column because the variable
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50 common border was only used in the selection equation.
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55 With respect to the free-trade agreement variables, for the first sample only PPML and NLS
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57 predict a significant and positive effect. For the second sample, PPML predicts always
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59 positive effects but the estimate is not significant for the EU, whereas the additional
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techniques produce negative effects for the EU agreement and positive and significant effects

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3 for the NAFTA, CACM and CARICOM agreements. However, as stated by some authors, the
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5 estimation of gravity models using cross-sectional data does not provide a good way to
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7 identify the effects of free trade agreements (Baier and Bergstrand, 2007).
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12 *Table 4. Estimation results: Theoretically justified gravity model*
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15 A panel data version of the gravity model is also estimated, adding to the traditional
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17 specification dyadic random effects and time effects. The results are shown in Table 5. The
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19 first column presents the result for the random effects model assuming homoskedasticity and
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21 uncorrelated errors. The second column allows for heteroskedastic errors and the third
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23 assumes autocorrelation of first order. Columns four to seven show the GLM estimates:
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25 PPLM, GPLM and NLS respectively. The log-linear models (columns 1 to 3) show more
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27 accurate coefficients than the GLM.
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32 *Table 5. Estimation results for panel data (1980-1999)*
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34 Since many other issues come into play when estimated gravity equations with panel data, as
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36 for example what is the correct way to control for multilateral resistance, we refer to the
37
38 results obtained in Martínez-Zarzoso et al. (2008) for the 47 country sample.
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41 **4.3 Robustness check**

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43 In this section, we first present the results obtained when testing for the pattern of
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45 heteroskedasticity. In addition, we investigate the out-of-sample forecast performance of the
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47 alternative estimates.
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50 To determine if the pattern of heteroskedasticity assumed by the different models is
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52 acceptable, a Park-type test is used to check for the adequacy of the log-linear model versus
53
54 alternative GLM. The Park-type test can be directly derived from Equation (4) in Section 2
55
56 above. The test consists of estimating the equation specified as
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$$59 \ln(y_i - \hat{y}_i)^2 = \ln(\lambda_0) + \lambda_1 \ln \hat{y}_i + \varepsilon_i. \quad (13)$$

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3 Acceptance of the null hypothesis: $H_0: \lambda_1=2$, based on a robust covariance estimator, would be
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5 in support of the log-linear model. The results of the test performed for the 65 countries
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7 sample are presented in the first part of Table 6. The point estimate for λ_1 in 1990 is 1.829 and
8
9 the estimates for other years are very similar. Below the point estimates, the 95 confidence
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11 intervals are also reported in Table 6. Although those intervals do not contain the value 2,
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13 they are closer to this value than to value 1, that will be in support of the PPML assumption
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15 (the variance is proportional to the mean).
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19 Since log-linearization of equation (4) is valid only under restrictive conditions of the
20
21 conditional distribution of y_i , we also used a robust alternative. As suggested by Manning and
22
23 Mullahy (2001) and Deb, Manning and Norton (2008) we have also estimated the modified
24
25 version of the Park test as a GLM with log link where the dependent variable is $(y_i - \hat{y}_i)^2$ and
26
27 the explanatory variable is $x\hat{\beta}$ from each one of the GLM estimators of y on x , using a robust
28
29 variance covariance matrix. The second part of Table 6 shows the results for several years
30
31 using Rose data. Among the GLM estimators we have chosen the Gamma family since it
32
33 presents the lowest AIC and BIC values. The results are presented in the second part of Table
34
35 6. The point estimate for λ_1 in 1990 is 1.586 and the estimates for other years are slightly
36
37 higher. Below the point estimates, the 95 confidence intervals are also reported. According to
38
39 those results, the variance always exceeds the mean and a Gamma family (within the GLM
40
41 models) would be more appropriate than a Poisson family. We also estimated $x\hat{\beta}$ using the
42
43 Poisson and Gaussian families, and the confidence intervals did not contain 1 and were
44
45 always closer to 2. For the 65 country sample the results for λ_1 lie on the 95% confidence
46
47 interval (1.66-1.75) and for the 47 country sample, the estimated confidence interval for λ_1 is
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49 (1.77-1.85). According to these results the more appropriate GLM family could be the
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51 gamma.
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3 SST used instead the results from a Gauss-Newton regression to test for the proportionality
4 assumption (the conditional variance is proportional to the conditional mean). The Gauss-
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Newton regression is given by

$$(y_i - \hat{y}_i)^2 / \sqrt{\hat{y}_i} = \lambda_0 \sqrt{\hat{y}_i} + \lambda_0 (\lambda_1 - 1) \ln(\hat{y}_i) \sqrt{\hat{y}_i} + \mu_i. \quad (14)$$

The proportionality hypothesis can be checked by testing for the significance of $\lambda_0 (\lambda_1 - 1)$ in Equation 14. The hypothesis will be accepted if the parameter is statistically insignificant. Equation 14 is derived using a Taylor series expansion around $\lambda_1 = 1$ (which is an approximation) and assuming that the variance is proportional to the mean to choose the appropriate weights in equation (14). We consider that the approach proposed by Manning and Mullahy (2001) and Deb, Manning and Norton (2008) is more straightforward and gives a confidence interval for λ_1 which is more useful to choose the more appropriate GLM estimator.

As pointed out by SST, the proportionality assumption is unlikely to hold; therefore, the PPLM does not fully account for the heteroskedasticity in the model. The authors mentioned that the inference should then be based on a robust covariance matrix estimator. However, another alternative is to specify the variance in a more accurate way: using a gamma family if λ_1 is close to 2, for example.

In this context (with λ_1 close to 2) FGLS could also be a good method. The shortcoming of standard FGLS techniques is that the relationship between the residuals and the independent variables is restricted to a linear model. In the presence of Jensen's inequality, as well as for other functional forms different from the linear model, this simplifying assumption lowers estimation quality. Nevertheless, for most practical purposes, this problem could be solved by using more sophisticated methods based on the results obtained in Carroll (1982), Robinson (1987) and Delgado (1992 and 1993). These authors propose a semi-parametric version of FGLS that consists of estimating the relationship between the residuals and the independent variables, allowing for an unknown functional form and using non-parametric techniques.

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3 The performance of the log-linearized models can be compared using the root mean squared
4 error (RMSE) reported in Tables 3 and 4 (in the first and second half of the table). Among
5 those models, the estimates of the FGLS model present the lowest RMSE. In order to use the
6 RMSE to compare log-log models with GLM models¹⁵, we have re-calculated the RMSE for
7 log-log models using the retransformed predicted values (retransformation proposed by Baser,
8 2007). The results presented in Tables 3 and 4 (RMSE retrans) indicate that NLS presents the
9 lowest value.

10
11 In order to compare different GLM models, the Deviance and Dispersion of the different
12 families can also be compared (See Tables 3 and 4, left hand side). The large Dispersion
13 shown by the Poisson estimates is another indicator that the Poisson distribution is not a good
14 choice. Since the distribution of exports presents signs of overdispersion (the variance is
15 larger than the mean), we scaled the standard errors using the square root of the Pearson chi-
16 square dispersion. The coefficients were identical to the previous analysis but the standard
17 errors were significantly higher in magnitude. The GPLM present (in all cases) the smaller
18 deviance and dispersion out of the three GML methods considered.

19
20 SST show the results obtained from a Ramsey-RESET test of specification error of functional
21 form that supported the PPML estimator. We also computed it, and the PPML estimator did
22 not pass the test. None of the models estimated for a single year pass the RESET test, apart
23 from the probit model estimated in the first step of the Heckman model. These results could
24 be due to the fact that relevant variables have been omitted. This problem is solved by using
25 panel data since the introduction of exporter-and-year and importer-and year fixed effects in
26 addition to dyadic fixed effects is an accepted way to control for the omission of relevant
27 variables in gravity models (Baldwin and Taglioni, 2006; Martínez-Zaroso et al., 2009).

28
29 We also computed the Akaike information criterion (AIC) and the Bayesian information
30 criterion (BIC). Although the results obtained for log-linear models and for GLM are not

¹⁵ Following Joao Santos Silva recommendation we retransform the predicted values obtained from log-.log models to make a valid comparison.

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3 comparable, we can use those two criteria to compare models within the first and second
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5 class. Whereas the FGLS and Harvey models always present the lowest AIC and BIC among
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7 the log-linear models. The Gamma always presents, for the three samples, the lowest AIC and
8
9 BIC¹⁶.

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12 Finally, we do some cross-validation. We use the estimates of the gravity model for the year
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14 1990 to calculate the predicted values in 1999. The fitted values are compared with the actual
15
16 values in 1999. In order to compare log-linearized models with GLM, we use for the log-
17
18 linear models the heteroskedasticity adjusted retransformation proposed by Baser (2007).
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20 Then the Root Mean Squared Predictor Error (RMSPE) is calculated. The results are
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22 summarized at the bottom of Tables 3 and 4 (first and second part) under RMSPE (1999). The
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24 RMSPE of different models can be compared directly, the model producing the smaller
25
26 RMSPE being the better one. We repeat the same experiment for the three samples. Our
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28 findings indicate that for the first dataset (Table 3) the Heckman and Harvey models present
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30 the smaller RMSPE, for the second dataset (first part of Table 4, last row) the PPLM present
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32 the smaller RMSPE and for the third (second part of Table 4, last row) the NLS present the
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34 smaller RMSPE.

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37 A matter of further research, outside the scope of the paper, is to check the accuracy and
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39 goodness of fit of two-part models. These models, the first part consisting of a probit or logit
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41 equation to distinguish between zero and positive outcomes and the second part using log-log
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43 models or GLM models on the positive values, allows different parameters and variables to
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45 determine the two parts of the data-generating process. Those models could be more
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47 appropriate than the Heckman model when the zeros represent outcomes of interest, as in the
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49 case of trade flows. This is recommendable when working with sectoral data, for which the
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¹⁶ Although AIC and BIC are not appropriate, strictly speaking, to evaluate models estimated by pseudo maximum likelihood, as stated by Santos Silva (personal communication), the criteria available based on the pseudo-likelihood function (PAIC, PBIC) add a penalty term that is very unstable and makes the PAIC and PBIC very inaccurate.

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3 amount of zero values in export flows is much higher than when working with aggregated
4 trade data.
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7 8 **5. Conclusions** 9

10 The extended use of gravity models to predict international trade flows has generated an
11 ongoing discussion concerning the estimation techniques to be applied. In particular, the
12 issues concerning zero trade values and heteroskedastic residuals are controversial. Both
13 issues also extend to gravity models used in social sciences to predict certain behaviour that
14 mimic gravitational interactions, namely migration, or capital flows. Immigration between
15 pairs of countries may be zero in a substantial percentage of observations, and omitting those
16 zero observations biases the regression results. Bilateral flows of international migrants also
17 exhibit tremendous variance both across destination countries and over time. Bilateral foreign
18 direct investment data contains also many zero values and the residuals of the related models
19 are in general heteroskedastic. This paper compares a number of estimation techniques that
20 have been proposed in the economic literature to deal with both problems. The results of the
21 simulations indicate that in these settings, PPML shows a similar performance to FGLS and
22 Gamma estimates when the bias and the expected loss are used as the criteria to evaluate the
23 performance of different estimators. The results also indicate that although the PPML is less
24 affected by heteroskedasticity than other estimators, its performance is similar, in terms of
25 bias and standard errors, to the FGLS estimator performance and in the case of no-zeros in the
26 dependent variable, the GPLM present the lowest biases.
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50 Nevertheless, it is worth noting that conclusions drawn from this type of simulation studies
51 should be made with care, and that the winning estimation method should not be labelled “the
52 new workhorse” for the estimation of constant-elasticity models. We have shown that small
53 changes of the simulation setting can lead to different outcomes. Hence, one should be
54 cautious and inspect each applied situation carefully in order to find the appropriate
55 estimation method.
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3 The results obtained when estimating the gravity model of trade using three different samples
4 indicate that finding the best estimator for a given dataset requires carrying out a large
5 number of tests and specification checks.
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10 The use of transformed least squares methods trades robustness for some improvement in
11 efficiency, whereas the use of GLM methods provides an easy way to deal with
12 heteroskedastic errors but the researcher should choose the best variance function and in some
13 cases the best GLM model cannot be estimated with adequate precision. These considerations
14 are not only valid for gravity models of trade; they can also be extended to gravity models
15 used to explain other types of inter-regional and international flows, including migration,
16 development aid, investment and transport.
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Table 1. Simulation results, regular case

	absolute error loss		bias		standard deviation	
	β_1	β_2	β_1	β_2	β_1	β_2
Case 1						
OLS ($y>0$)	0.3894	0.3560	0.3894	0.3560	0.0390	0.0529
NLS	0.0059	0.0137	0.0000	0.0002	0.0075	0.0173
Gamma	0.0535	0.0652	0.0139	0.0084	0.0672	0.0827
Poisson	0.0121	0.0215	-0.0001	0.0000	0.0161	0.0274
FGLS ($y>0$)	0.0943	0.0879	0.0795	0.0504	0.0737	0.0956
Case 2						
OLS ($y>0$)	0.2105	0.1992	0.2105	0.1992	0.0294	0.0497
NLS	0.0259	0.0452	0.0003	0.0012	0.0327	0.0569
Gamma	0.0336	0.0508	0.0043	0.0023	0.0431	0.0638
Poisson	0.0156	0.0321	0.0001	0.0000	0.0196	0.0401
FGLS ($y>0$)	0.0571	0.0931	0.0164	0.0008	0.0751	0.1285
Case 3						
OLS ($y>0$)	0.0213	0.0424	0.0001	-0.0001	0.0267	0.0532
NLS	0.2734	0.2331	0.1295	0.0255	3.4097	0.7309
Gamma	0.0253	0.0522	0.0003	-0.0006	0.0315	0.0650
Poisson	0.0551	0.0807	-0.0027	-0.0008	0.0721	0.1021
FGLS ($y>0$)	0.0219	0.0450	-0.0003	0.0001	0.0274	0.0564
Case 4						
OLS ($y>0$)	0.1327	0.1280	0.1326	-0.1247	0.0385	0.0754
NLS	0.9386	0.5972	0.7411	0.1663	20.5992	9.6913
Gamma	0.0463	0.0855	0.0030	-0.0022	0.0589	0.1073
Poisson	0.0777	0.1151	-0.0062	-0.0055	0.1030	0.1466
FGLS ($y>0$)	0.0591	0.1017	0.0492	-0.0881	0.0506	0.0861

Note: The absolute error loss is defined in Equations 7 and 8 in the main text.

Table 2. Simulation results with zeros in the dependent variable

zeros at random	absolute error loss		bias		standard deviation	
	β_1	β_2	β_1	β_2	β_1	β_2
Case 1						
OLS ($y>0$)	0.3900	0.3556	0.3900	0.3556	0.0429	0.0586
NLS	0.0694	0.0889	0.0043	0.0080	0.0907	0.1226
Gamma	0.0583	0.0721	0.0154	0.0072	0.0738	0.0920
Poisson	0.0290	0.0422	-0.0003	-0.0003	0.0379	0.0530
FGLS ($y>0$)	0.0979	0.0903	0.0826	0.0539	0.0760	0.0975
Case 2						
OLS ($y>0$)	0.2104	0.1993	0.2104	0.1993	0.0322	0.0529
NLS	0.0743	0.1038	0.0045	0.0073	0.0963	0.1371
Gamma	0.0372	0.0570	0.0045	0.0026	0.0473	0.0719
Poisson	0.0299	0.0486	-0.0011	0.0004	0.0383	0.0610
FGLS ($y>0$)	0.0575	0.0941	0.0195	0.0033	0.0749	0.1301
Case 3						
OLS ($y>0$)	0.0225	0.0462	-0.0003	0.0007	0.0282	0.0577
NLS	0.4154	0.3285	0.2529	0.0829	5.6568	2.4460
Gamma	0.0292	0.0596	-0.0002	0.0004	0.0366	0.0750
Poisson	0.0643	0.0937	-0.0036	-0.0028	0.0836	0.1185
FGLS ($y>0$)	0.0237	0.0483	0.0001	0.0008	0.0299	0.0607
Case 4						
OLS ($y>0$)	0.1322	0.1299	0.1321	-0.1254	0.0417	0.0830
NLS	0.8055	0.5406	0.5860	0.1639	7.8462	3.2434
Gamma	0.0495	0.0976	0.0046	0.0000	0.0626	0.1229
Poisson	0.0863	0.1281	-0.0084	-0.0045	0.1149	0.1649
FGLS ($y>0$)	0.0613	0.1053	0.0500	-0.0878	0.0536	0.0941
zeros with pattern						
Case 1						
OLS ($y>0$)	0.3408	0.2877	0.3408	0.2877	0.0448	0.0520
NLS	0.0116	0.0140	0.0112	0.0009	0.0082	0.0176
Gamma	0.2651	0.0801	0.2620	0.0097	0.0967	0.1014
Poisson	0.0669	0.0221	0.0668	0.0008	0.0166	0.0279
FGLS ($y>0$)	0.0760	0.0765	0.0571	0.0416	0.0714	0.0862
Case 2						
OLS ($y>0$)	0.1938	0.1776	0.1938	0.1776	0.0336	0.0488
NLS	0.0273	0.0468	0.0118	0.0021	0.0330	0.0586
Gamma	0.2490	0.0634	0.2487	0.0027	0.0644	0.0799
Poisson	0.0664	0.0335	0.0664	0.0010	0.0205	0.0420
FGLS ($y>0$)	0.0542	0.0889	0.0111	0.0054	0.0730	0.1227
Case 3						
OLS ($y>0$)	0.0248	0.0465	0.0002	-0.0007	0.0311	0.0583
NLS	0.3080	0.2827	0.1818	0.0921	3.6995	2.0455
Gamma	0.2405	0.0622	0.2405	0.0011	0.0408	0.0777
Poisson	0.0736	0.0834	0.0633	-0.0027	0.0716	0.1054
FGLS ($y>0$)	0.0258	0.0491	0.0001	0.0010	0.0325	0.0616
Case 4						
OLS ($y>0$)	0.1200	0.1545	0.1199	-0.1529	0.0448	0.0794
NLS	0.5178	0.4039	0.3373	0.1053	4.9428	3.1924
Gamma	0.2512	0.0999	0.2508	-0.0011	0.0773	0.1261
Poisson	0.0870	0.1194	0.0601	-0.0081	0.1047	0.1521
FGLS ($y>0$)	0.0584	0.1078	0.0465	0.0929	0.0540	0.0914

Note: The absolute error loss is defined in Equations 7 and 8 in the main text.

Table 3. Estimation results for 1990: traditional gravity model

Dep. Var.	LOG-LOG MODELS			TWO PART	GENERALISED LINEAR			
	OLS	FGLS	Harvey	HECKMAN	Poisson		Gamma	NLS
	log(exports)	log(exports)	log(exports)	log(exports)	exports	exports	exports	exports
lyi	1.08	0.932	1.077	1.069	0.38	0.799	0.739	0.86
	0.016	0.009	0.014	0.017	0.012	0.037	0.037	0.07
lyj	0.951	0.899	0.958	0.948	0.144	0.798	0.744	0.91
	0.016	0.008	0.014	0.016	0.011	0.03	0.032	0.079
lyhi	0.777	0.791	0.793	0.774	0.143	0.549	0.596	0.587
	0.028	0.019	0.025	0.027	0.019	0.08	0.043	0.129
lyhj	0.585	0.574	0.552	0.58	0.218	0.614	0.462	0.745
	0.029	0.016	0.025	0.027	0.019	0.054	0.055	0.183
ldist	-1.326	-0.903	-1.226	-1.316	-0.398	-0.713	-1.066	-0.797
	0.033	0.015	0.029	0.033	0.026	0.038	0.043	0.119
border	0.714	0.26	0.554	0.712	0.398	0.427	0.457	0.194
	0.136	0.04	0.132	0.154	0.162	0.113	0.151	0.124
comlang	0.329	0.335	0.396	0.325	0.085	0.585	0.211	0.551
	0.062	0.029	0.054	0.06	0.043	0.109	0.078	0.333
colony	1.635	0.679	1.377	1.627	8.598	0.101	1.219	-0.11
	0.112	0.042	0.124	0.165	8.14E+06	0.144	0.125	0.428
landl	-0.187	-0.3	-0.206	-0.19	0.059	-0.589	-0.268	-0.091
	0.051	0.027	0.046	0.05	0.034	0.108	0.092	0.474
island	0.008	0.354	0.048		0.126	0.391	-0.016	0.652
	0.05	0.024	0.046		0.035	0.077	0.069	0.18
landap	-0.047	-0.052	-0.058	-0.047	-0.027	-0.057	-0.016	-0.001
	0.011	0.005	0.009	0.01	0.007	0.024	0.019	0.066
regional	0.915	0.15	0.443	0.918	0.517	0.077	0.459	0.36
	0.145	0.037	0.155	0.171	0.187	0.093	0.151	0.156
custriect	1.714	0.698	1.851	1.691	0.529	-0.084	0.689	-0.015
	0.185	0.047	0.204	0.236	0.195	0.203	0.207	0.433
comcol	0.664	0.766	0.624	0.662	0.17	0.811	0.643	1.671
	0.101	0.076	0.085	0.085	0.053	0.296	0.121	1.124
_cons	-37.875	-36.377	-38.289	-37.5	-11.26	-28.51	-22.46	-36.549
	0.685	0.342	0.537	0.628	0.448	1.231	0.76	5.546
R-squared	0.619	0.862	0.679					
N	12134	12134	12134	13974		13974	13974	13974
ll	-28189.5	-18299.6	-26833.8	-32312.3		-	-234853	-302957
						7.90E+11		
rmse	2.4716	1.093957	2.210317	2.5114				
aic	56409.08	36629.1	53697.61	64686.49		1.13E+08	33.615	43.36218
bic	56520.13	36740.15	53808.67	64920.39		1.58E+12	-86841.5	5.54E+21
RMSE	1.81E+09	1.87E+09	1.70E+09	1.56E+09		7.39E+08	2.31E+09	6.30E+08
retrans								
RMSEP	9.94E+10	3.70E+09	3.40E+09	3.26E+09		4.22E+09	1.26E+10	5.61E+09

Note: Ly_i (Ly_j) and Ly_{hi} (Ly_{hj}) denote income and income per capita in country i (j), respectively. Border is a binary variable that is unity if countries i and j share a common border, 0 otherwise. *Comlang* is a binary variable that is unity if countries i and j share a common language, 0 otherwise. *Colony* is a dummy that takes the value of 1 if country i ever colonized j or vice versa, 0 otherwise. *Island* takes the value of 1 if country i or country j or both are islands. *Landl* is the number of landlocked countries in the country pair. *Landap* is the total area of both countries, i and j . *FTA* is a dummy that takes the value of 1 when the trading partners belong to the same trade agreement, 0 otherwise. *Custriect* is a dummy that takes the value of 1 when the trading partners share a common currency, 0 otherwise. *Comcol* is a dummy that takes the value of 1 if i and j were ever colonies after 1945 with the same colonizer, 0 otherwise. Standard errors reported below each estimate.

Table 4. Estimation results for 1990: Theoretically justified gravity model

Countries'	LOG-LINEAR MODELS			TWO PART HECKMAN		GENERALIZED LINEAR MODELS		
	OLS	FGLS	Harvey	Heckman		Poisson	Gamma	NLS
Dep. Var.	log(exports)	log(exports)	log(exports)	log(exports)	exports	exports	exports	exports
Edist	-1.134	-1.121	-1.081	-1.095	-0.97	-0.539	-1.191	-0.454
10	0.045	0.042	0.04	0.039	0.078	0.044	0.057	0.085
Comlang	0.639	0.587	0.645	0.623	0.592	0.271	0.946	0.448
12	0.085	0.077	0.073	0.084	0.136	0.064	0.101	0.101
Border	0.393	0.113	0.245	0.402		0.658	0.533	0.588
14	0.151	0.136	0.137	0.148		0.089	0.168	0.187
FTA	-0.064	-0.106	-0.073		0.172	0.287	-0.017	0.540
17	0.045	0.041	0.037		0.120	0.050	0.056	0.072
Constant	12.824	15.036	14.236	12.388	6.429	11.761	16.464	10.335
19	0.55	0.573	0.542	0.505	0.771	0.559	0.693	0.863
R-squared	0.801	0.835	0.853		Deviance	4.96E+8	5168	2.01E+15
RMSE	1.395448	1.182354	1.14222		Dispersion	5.92e+8	13158	2.01E+15
22	3230	2938	3230	3220	3804	3804	3804	3804
23	-5593.633	-4595.019	-4946.855	-6211.515		-2.48E+08	-41675.05	-56731.24
24	11445.27	9448.038	10151.71	12939.03		130566	21.979	29.89497
AIC	12229.62	10220.17	10936.06	14549.93		4.97E+08	-25127.24	2.00E+15
BIC	12056872	4673505.5	3431688	8477757		978886.81	22653696	725514.25
RMSE retrans	15804256	8104227.5	8640781	7839773.5		3661013.5	20849120	4846363
28								
29 countries'	OLS	FGLS	Harvey	Heckman		Poisson	Gamma	NLS
Dep. Var.	log(exports)	log(exports)	log(exports)	log(exports)	exports	exports	exports	exports
Edist	-1.064	-1.04	-1.06	-0.841	-0.705	-0.564	-1.396	-0.294
32	0.058	0.044	0.05	0.09	0.055	0.052	0.069	0.078
Island	-0.37	-0.284	-0.279	-0.099	-0.161	-1.034	-1.024	-0.825
34	1.529	0.877	0.435	1.612	0.076	0.403	0.472	0.469
Comlang	1.416	1.184	1.259	1.23	0.51	0.775	1.74	0.745
36	0.122	0.089	0.103	0.141	0.119	0.069	0.147	0.07
Border	-0.354	-0.347	-0.349		-1.526	0.246	-0.218	0.213
38	0.215	0.162	0.176		0.258	0.079	0.207	0.077
EU	-0.369	-0.462	-0.461	-0.479	5.722	0.04	-1.729	-0.189
41	0.191	0.122	0.142	0.199	.	0.09	0.199	0.101
NAFTA	1.071	0.478	0.657	0.668	6.476	1.322	-0.037	2.571
43	0.642	0.426	0.598	0.671	.	0.184	0.456	0.363
CACM	1.805	1.765	1.841	1.884	-0.118	1.481	1.245	1.909
45	0.388	0.246	0.281	0.401	0.559	0.34	0.328	0.84
CARIC	1.393	1.44	1.447	1.763	-1.205	0.917	2.158	0.662
47	0.231	0.233	0.228	0.255	0.13	0.331	0.289	0.9
Constant	11.968	15.402	12.092	10.185	6.797	18.786	26.599	17.334
50	1.677	1.426	0.822	0.801	0.485	0.457	0.654	0.587
R-squared	0.817	0.903	0.878		Deviance	1.78E+8	3153	5.67E+14
RMSE	1.468021	1.005774	0.0003035		Dispersion	1.99E+8	7910	5.67E+14
53	1656	1656	1656	2162	2162	2162	2162	2162
54	-2933.42	-2307.191	11116.21			-8.91E+07	-21359.71	-31491.06
AIC	6068.839	4816.382	-22032.42	.		8.24E+04	19.85265	29.22484
BIC	6615.467	5363.011	-21491.21	.		1.78E+08	-12672.58	5.68E+14
RMSE retrans	13452503	6886919.5	9055090	12829089		666638.63	61030884	512352.28
RMSPPE (1999)	11174116	5939834	7465615.5	10453723		4320483	58862776	4307895.5

Note: The effects of income and income per capita variables cannot be estimated since exporter and importer effects are added as regressors. Border, Comlang, Island and FTA are defined as in Table 3. EU, NAFTA, CACM and CARICOM denote trade creation effects. Standard errors reported below each estimate.

Table 5. Estimation results for panel data (1980-1999)

47 countries' Sample	<u>LOG-LINEAR MODELS</u>			<u>GENERALIZED LINEAR MODELS</u>		
	OLS	FGLS, het	FGLS, het ar1	Poisson	Gamma	NLS
Dep. Var.	log(exports)	log(exports)	log(exports)	exports	exports	exports
Ly_i	0.941	0.889	0.907	0.782	0.848	0.759
	0.006	0.012	0.014	0.036	0.040	0.043
Ly_j	0.889	0.849	0.824	0.824	0.698	0.701
	0.005	0.01	0.01	0.034	0.034	0.039
Ly_{hi}	0.870	0.435	0.915	0.191	1.025	-0.033
	0.036	0.065	0.073	0.185	0.152	0.111
Ly_{hj}	0.217	-0.367	-0.300	-0.205	0.390	-0.256
	0.012	0.017	0.016	0.052	0.04	0.069
Ldist	-1.176	-1.276	-1.533	-0.618	-0.941	-0.470
	0.012	0.021	0.024	0.067	0.065	0.04
Border	-0.061	0.006	0.091	0.448	0.348	0.478
	0.028	0.049	0.071	0.115	0.211	0.108
Comlang	0.896	0.721	0.629	0.625	1.185	0.618
	0.023	0.049	0.057	0.113	0.279	0.086
Island	-0.048	-0.021	0.012	0.055	-0.075	0.12
	0.019	0.039	0.047	0.115	0.198	0.111
EU	0.952	0.521	0.561	0.339	0.889	-0.081
	0.026	0.031	0.032	0.123	0.151	0.095
EUX	0.125	-0.270	-0.160	-0.522	0.113	-0.510
	0.022	0.031	0.028	0.163	0.121	0.207
EUM	0.902	0.864	0.998	0.252	0.827	-0.052
	0.033	0.052	0.062	0.179	0.184	0.193
NAFTA	0.702	0.171	0.037	0.468	1.437	0.650
	0.182	0.127	0.123	0.257	0.354	0.153
NAFTAX	-0.288	-0.137	-0.140	-0.551	-0.397	-0.621
	0.03	0.031	0.03	0.132	0.134	0.113
NAFTAM	0.463	0.326	0.425	0.163	0.384	-0.206
	0.04	0.057	0.061	0.256	0.132	0.178
Constant	-35.748	-23.336	-26.421	-23.470	-32.502	-17.947
	0.37	0.663	0.721	2.05	1.851	2.154
Nobs	12507	12506	12506	13264	13264	13264
Deviance				6.10E+09	24015.77	4.62E+16
Chi2_dev				7.18E+09	50382.26	4.62E+16
Dispersion					3.798421	3.48E+12
Chi2_dis				541439.1	3.798421	3.48E+12

Note: Ly_i (Ly_j) and Ly_{hi} (Ly_{hj}) denote income and income per capita in country i (j), respectively. Ldist, Border, Comlang and Island are defined as in Table 3. Ldist, Border, Comlang and Island are defined as in Table 3. EU and NAFTA denote trade creation effects, EUM, NAFTA, CACMM and CARICOMM denote import-diversion effects; and EUX and NAFTAX denote export diversion effects. Standard errors reported below each estimate.

Table 6. Test for the implicit assumption in PPML and OLS/FGLS (Rose data, various years)

PARK TEST LOG-LOG	<u>1980</u>	<u>1990</u>	<u>1995</u>	<u>2000</u>
λ_1	1.803	1.829	1.834	1.856
95% Confidence Interval	1.786,1.820	1.816,1.842	1.822,1.846	1.846,1.865
St. error	0.011	0.008	0.007	0.007
Nobs	11852	13974	17616	15110
PARK TEST GAMMA	<u>1980</u>	<u>1990</u>	<u>1995</u>	<u>2000</u>
λ_1	1.68	1.586	1.639	1.646
95% Confidence Interval	1.534,1.826	1.451,1.722	1.571,1.708	1.472,1.819
St. error	0.074	0.069	0.035	0.088
Nobs	11852	13974	17616	15110
(Rose Data)				

Note: λ_1 determines the relationship between raw-scale mean and variance functions.

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Appendix

A.1 Simulations 40 percent zeros

40% Zeros		Absolute error loss		Bias		Standard deviation	
		β_1	β_2	β_1	β_2	β_1	β_2
Case 1	Poisson	0.0673	0.0221	0.0673	0.0010	0.0002	0.0003
	FGLS	0.0944	0.0873	0.0787	0.0507	0.0008	0.0009
Case 2	Poisson	0.0539	0.0788	-0.0018	-0.0014	0.0007	0.0010
	FGLS	0.0621	0.0989	0.0231	0.0054	0.0008	0.0014
Case 3	Poisson	0.0850	0.1237	-0.0068	-0.0028	0.0011	0.0016
	FGLS	0.0289	0.0574	-0.0003	0.0000	0.0004	0.0007
Case 4	Poisson	0.1066	0.1605	-0.0114	-0.0101	0.0014	0.0020
	FGLS	0.0667	0.1159	0.0504	-0.0884	0.0006	0.0011