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Multivariate shuffles and approximation of copulas

Fabrizio Durante^{a,*}, Juan Fernández-Sánchez^b

^aSchool of Economics and Management Free University of Bozen-Bolzano, Bolzano, Italy ^b Grupo de Investigación de Análisis Matemático Universidad de Almería, La Cañada de San Urbano, Almería, Spain

Abstract

We present and study a method for constructing multivariate copulas, which includes both the shuffles of Min and the ordinal sums. Such a method has been used in order to show that suitable transformations of a given copula constitute a dense set in the class of all copulas with respect to the L^{∞} norm.

Keywords: Copula, *d*–fold stochastic measure, ordinal sum, shuffle of Min. 2010 MSC: 62H05, 60E05

1. Introduction

Let I be the unit interval [0, 1] and let \mathbb{I}^d be the *d*-dimensional unit cube $(d \ge 2)$. A *d*-dimensional copula (copula, for brevity) is a distribution function on \mathbb{I}^d whose univariate margins are uniformly distributed on I. Copulas have been widely studied in Probability and Statistics due to the fact that they capture the scale-invariant dependence of continuous random vectors, as shown by Sklar's Theorem (Sklar, 1959). For more information about copulas and their applications, see, e.g., Jaworski et al. (2010); McNeil et al. (2005); Nelsen (2006); Salvadori et al. (2007).

A shuffle of Min (where Min indicates the comonotonicity copula $M_d(u_1, \ldots, u_d) = \min\{u_1, \ldots, u_d\}$) is a construction principle that generates new copulas by means of a suitable rearrangement of the mass distribution of M_d (Mikusiński et al., 1992, 2009). It has been used in several studies about approximations of copulas. In particular, it has been observed that shuffles of Min are dense in the class of bivariate copulas endowed with the L^{∞} norm (Mikusiński et al., 1992; Vitale, 1990).

Recently, Durante et al. (2009) have generalized the notion of shuffles of Min in order to allow that the probability mass of any bivariate copula, not only M_2 , can be used as a starting point for further rearrangements. Such a generalization is based on the use

^{*}Corresponding author. Tel. +39 0471013493. Fax +39 0471013009.

Email addresses: fabrizio.durante@unibz.it (Fabrizio Durante), juanfernandez@ual.es (Juan Fernández-Sánchez)

of special measure-preserving transformations, which conveniently describe the geometric approach commonly used in defining shuffles of Min.

In this paper, we aim at providing a further construction principle that is directly inspired by the idea of shuffle of Min (see section 2). Such a method encompasses both the multivariate shuffle of Min presented by Mikusiński et al. (2009) and the ordinal sum construction by Mesiar and Sempi (2010). Remarkably, copulas obtained by the introduced method have an interesting stochastic representation.

Moreover, starting with any copula C, it has been shown that the set of all transformations of C by means of such a construction is dense in the set of copulas endowed with the L^{∞} norm (see section 3). A feature of particular interest when one wants to approximate copulas by means of elements from a given basis set.

In order to reach our goal, we rely on measure-theoretic techniques that are grounded on the well-known one-to-one correspondence between copulas and special probability measures.

2. Definitions and basic properties

Let $d \in \mathbb{N}$, $d \geq 2$. Let λ be the Lebesgue measure on \mathbb{R} . Let Cop_d be the set of all *d*-dimensional copulas, i.e. the set of all distribution functions $C \colon \mathbb{R}^d \to \mathbb{I}$ whose univariate margins are uniformly distributed on \mathbb{I} . Notice that, since a copula C is uniquely determined by the values that it assumes on \mathbb{I}^d , we usually refrain from specifying the value of C outside \mathbb{I}^d . For basic results and notations about copulas, we refer to Durante and Sempi (2010).

We recall that a copula C induces a probability measure μ_C defined, for every nonempty orthotope $R \subseteq \mathbb{I}^d$, by $\mu_C(R) := V_C(R)$, where $V_C(R)$ denotes the C-volume of R, and extended by means of standard arguments to the Borel σ -algebra $\mathscr{B}(\mathbb{I}^d)$ (Nelsen, 2006). Such a μ_C is d-fold stochastic, i.e. the push-forward of μ_C under any projection p_i coincides with Lebesgue measure on \mathbb{I} , $\mu_C(p_i^{-1}(A)) = \lambda(A)$ for any Borel set $A \subseteq \mathbb{I}$. In particular, 2fold stochastic measures (better known as *doubly stochastic measures*) represent an infinitedimensional generalization of doubly stochastic matrices and originated from an idea by Birkhoff (1948, Problem 111). Notably, such measures also appear in several problems connected with optimal transportation (Ahmad et al., 2009; Gangbo and McCann, 2000; Rachev and Rüschendorf, 1998).

Actually, d-fold stochastic measures on $(\mathbb{I}^d, \mathscr{B}(\mathbb{I}^d))$ are in a one-to-one correspondence to the set of d-copula. In fact, from any d-fold stochastic measure μ a copula can be constructed by setting, for any $\mathbf{u} \in \mathbb{I}^d$,

$$C_{\mu}(\mathbf{u}) := \mu\left([0,\mathbf{u}]\right),$$

where $[0, \mathbf{u}] = [0, u_1] \times \cdots \times [0, u_d]$. In the following, when no confusion may arise, we may identify a copula and its corresponding measure.

Based on the isomorphism between copulas and measures, we propose the following construction principle for copulas.

Definition 2.1. Let $\mathscr{J}^1, \ldots, \mathscr{J}^d$ be systems of closed and non-empty intervals of \mathbb{I} , $\mathscr{J}^i = (J_n^i = [a_n^i, b_n^i])_{n \in \mathbb{N}}$ such that:

- (S1) N represents a finite or countable index set, i.e. $N = \{0, 1, \dots, \tilde{n}\}$ or $N = \mathbb{Z}_+$;
- (S2) for every $i \in \{1, 2, ..., d\}$ and $n, m \in N, n \neq m, J_n^i$ and J_m^i have at most one endpoint in common;
- (S3) for every $i \in \{1, 2, ..., d\}, \sum_{n \in N} \lambda(J_n^i) = 1;$
- (S4) for every $n \in N$, $\lambda(J_n^1) = \lambda(J_n^2) = \cdots = \lambda(J_n^d)$.

A system $S = (\mathcal{J}^i)_{i=1}^d$ satisfying the above properties is called *shuffling structure* (shortly, *S*-*structure*). The set of all S-structures based on an index set N is indicated by \mathscr{S}_N , while $\mathscr{S} = \bigcup \mathscr{S}_N$.

Definition 2.2. Let N be an index set as in (S1). Let $(\mu_n)_{n\in N}$ be a system of probability measures on $(\mathbb{I}^d, \mathscr{B}(\mathbb{I}^d))$ such that, for every $n \in N$, $\mu_n((0,1)^d) = 1$. Let $(\mathscr{J}^i)_{i=1}^d$ be an S-structure. Let $\mu \colon \mathscr{B}(\mathbb{I}^d) \to \mathbb{R}_+$ be the set-function defined, for every $A \in \mathscr{B}(\mathbb{I}^d)$, by

$$\mu(A) = \sum_{(n_1,\dots,n_d)\in N^d} \mu_{n_1,\dots,n_d} \left(A \cap \left(J_{n_1} \times \dots \times J_{n_d}\right)\right),\tag{1}$$

where, for all $(n_1, \ldots, n_d) \in N^d$, μ_{n_1,\ldots,n_d} is the set-function defined on the Borel sets of $J_{n_1}^1 \times \cdots \times J_{n_d}^d$ in the following way:

- (M1) $\mu_{n_1,\dots,n_d} = 0$ if $n_k \neq n_{k'}$ for some k, k';
- (M2) for every Borel set $A \subseteq J_n^1 \times \cdots \times J_n^d$

$$\mu_{n,\dots,n}(A) = \lambda(J_n^1)\mu_n\left(\varphi_n(A)\right),\tag{2}$$

where $\varphi_n \colon J_n^1 \times \cdots \times J_n^d \to \mathbb{I}^d$ is given by

$$\varphi_n(x_1,\ldots,x_d) = \left(\frac{x_1 - a_n^1}{\lambda(J_n^1)},\ldots,\frac{x_d - a_n^d}{\lambda(J_n^1)}\right).$$
(3)

The set-function μ is called *shuffling set-function* related to the S-structure $(\mathscr{J}^i)_{i=1}^d$ and to $(\mu_n)_{n\in N}$. It is indicated by the symbol $\mu = \left\langle (\mathscr{J}^i)_{i=1}^d, (\mu_n)_{n\in N}, \right\rangle$.

Notice that, by definition, a shuffling set-function μ is a measure, since any μ_{n_1,\dots,n_d} is a measure and $\bigcup_{(n_1,\dots,n_d)\in N^d} (J_{n_1}^1 \times \cdots \times J_{n_d}^d) = \mathbb{I}^d$. Moreover, the following result can be proved.

Proposition 2.1. Let $\mu = \left\langle (\mathscr{J}^i)_{i=1}^d, (\mu_n)_{n \in N}, \right\rangle$ be a shuffling set-function. Then μ is a probability measure. Moreover, if each μ_n is d-fold stochastic, then μ is d-fold stochastic.

PROOF. Let $\mu = \left\langle (\mathscr{J}^i)_{i=1}^d, (\mu_n)_{n \in N}, \right\rangle$ be a shuffling set-function. Then

$$\iota\left(\mathbb{I}^{d}\right) = \sum_{\substack{(n_{1},\dots,n_{d})\in N^{d} \\ n_{1},\dots,n_{d}\in N^{d}}} \mu_{n_{1},\dots,n_{d}}\left(\cup_{(n_{1},\dots,n_{d})\in N^{d}}\left(J_{n_{1}}^{1}\times\cdots\times J_{n_{d}}^{d}\right)\right) \\
 = \sum_{n\in N}\mu_{n,\dots,n}\left(J_{n}^{1}\times\cdots\times J_{n}^{d}\right) = \sum_{n\in N}\lambda(J_{n}^{1})\mu_{n}\left(\varphi_{n}\left(J_{n}^{1}\times\cdots\times J_{n}^{d}\right)\right) = 1,$$

that is μ is a probability measure.

ŀ

Now, suppose that each μ_n is *d*-fold stochastic. In order to prove that μ is *d*-fold stochastic, we have to show that, for every every $i \in \{1, \ldots, d\}$ and for every $x_i \in \mathbb{I}$,

$$\mu\left(\mathbb{I}^{i-1}\times[0,x_i]\times\mathbb{I}^{d-i}\right)=\lambda([0,x_i]).$$

Suppose that i = 1 (the other cases can be treated analogously). Let $x_1 \in \mathbb{I}$. Since $\mu(\{x_1\} \times \mathbb{I}^{d-1}) = 0$, we can assume, without loss of generality, that $x_1 \notin \{a_n^1, b_n^1\}$ for every $n \in N$. Let N_1 be the following index set

$$N_1 = \{ n \in N \colon J_n^1 \cap [0, x_1] \neq \emptyset \text{ and } x_1 \notin J_n^1 \}.$$

Let $\hat{n} \in N$ such that $x_1 \in [a_{\hat{n}}^1, b_{\hat{n}}^1]$. It follows that

$$\mu \left([0, x_1] \times \mathbb{I}^{d-1} \right) = \left(\sum_{n \in N_1} \lambda(J_n^1) \mu_n \left(\varphi_n \left(J_n^1 \times \dots \times J_n^d \right) \right) \right) \\ + \lambda(J_{\widehat{n}}^1) \mu_{\widehat{n}} \left(\varphi_{\widehat{n}} \left(\left(J_{\widehat{n}}^1 \cap [0, x_1] \right) \times J_{\widehat{n}}^2 \times \dots \times J_{\widehat{n}}^d \right) \right) \\ = \sum_{n \in N_1} \lambda(J_n^1) + \lambda(J_{\widehat{n}}^1) \mu_{\widehat{n}} \left(\left[0, \frac{x_1 - a_{\widehat{n}}^1}{b_{\widehat{n}}^1 - a_{\widehat{n}}^1} \right] \times \mathbb{I}^{d-1} \right) \\ = \sum_{n \in N_1} \left(b_n^1 - a_n^1 \right) + \lambda(J_{\widehat{n}}^1) \frac{x_1 - a_{\widehat{n}}^1}{b_{\widehat{n}}^1 - a_{\widehat{n}}^1} = x_1,$$

which is the desired assertion.

A copula *C* is a *shuffling copula* (S–copula, for short) if its induced measure μ_C can be represented as a shuffling measure that is *d*–fold stochastic. Such a *C* is indicated by the symbol $\left\langle \left(\mathscr{J}^i \right)_{i=1}^d, \left(C_n \right)_{n \in \mathbb{N}} \right\rangle$, where C_n is the copula such that $\mu_n = \mu_{C_n}$.

Remark 2.1. Intuitively, an S-copula C is obtained by the following procedure:

- 1. define a suitable partition $\{J_{n_1}^1 \times \cdots \times J_{n_d}^d\}_{(n_1,\dots,n_d) \in N^d}$ of \mathbb{I}^d formed by *d*-dimensional orthotopes (the S-structure);
- 2. given a system of copulas $(C_n)_{n \in N}$, plug a transformation of the probability mass of C_n in $J_n^1 \times \cdots \times J_n^d$,
- 3. define C as the copula whose associated measure coincides in each $J_{n_1}^1 \times \cdots \times J_{n_d}^d$ with the previous transformation of μ_{C_n} .

An explicit expression of an S–copula is given by the following result.

Proposition 2.2. Let $C = \left\langle (\mathscr{J}^i)_{i=1}^d, (C_n)_{n \in N} \right\rangle$ be an *S*-copula. Then, for all $\mathbf{u} \in \mathbb{I}^d$,

$$C(\mathbf{u}) = \sum_{n \in N} \lambda(J_n^1) C_n\left(\frac{u_1 - a_n^1}{\lambda(J_n^1)}, \dots, \frac{u_d - a_n^d}{\lambda(J_n^1)}\right).$$
(4)

PROOF. Since Definition 2.2 it follows that, for every $\mathbf{u} \in \mathbb{I}^d$,

$$C(\mathbf{u}) = \mu_C([0,\mathbf{u}]) = \sum_{n \in N} \lambda(J_n^1) \mu_{C_n} \left(\varphi_n([0,\mathbf{u}] \cap \left(J_n^1 \times \dots \times J_n^d\right)) \right)$$
$$= \sum_{n \in N} \lambda(J_n^1) C_n \left(\frac{u_1 - a_n^1}{\lambda(J_n^1)}, \dots, \frac{u_d - a_n^d}{\lambda(J_n^1)} \right),$$

which is the desired assertion.

Thanks to eq. (4), an S-copula can be interpreted as a convex combination (with coefficients $\lambda(J_n^1)$) of the *d*-variate distribution functions $F_n(\mathbf{x}) = C_n\left(\frac{x_1-a_n^1}{\lambda(J_n^1)}, \ldots, \frac{x_d-a_n^d}{\lambda(J_n^1)}\right)$. As such, an S-copula $C = \left\langle \left(\mathscr{J}^i \right)_{i=1}^d, \left(C_n\right)_{n \in \mathbb{N}} \right\rangle$ has the following probabilistic interpretation, which is inspired by the methods of Mikusiński et al. (1991).

Let $(\Omega, \mathscr{A}, \mathbb{P})$ be a probability space. Assume that, for every $n \in N$, the random vector $\mathbf{U}^n = (U_1^n, \ldots, U_d^n)$ is distributed according to C_n . Let Z be a discrete random variable assuming values in N such that, for every $n \in N$, $P(Z = n) = \lambda(J_n^1)$. For every $n \in N$, consider the random vector

$$\mathbf{V}^{n} = (V_{1}^{n}, \dots, V_{d}^{n}) = \left(\lambda(J_{n}^{1})U_{1}^{n} + a_{n}^{1}, \dots, \lambda(J_{n}^{1})U_{d}^{n} + a_{n}^{d}\right).$$

Finally, let us consider the random vector \mathbf{W} given by

$$\mathbf{W} = \sum_{n \in N} \sigma_n(Z) \mathbf{V}^n,$$

where, for every $n \in N$, $\sigma_n(x) = 1$ if x = n, $\sigma_n(x) = 0$ otherwise. For every $\mathbf{u} \in \mathbb{I}^d$

$$\mathbb{P}(\mathbf{W} \le \mathbf{u}) = \sum_{n \in N} \mathbb{P}(\mathbf{W} \le \mathbf{u} \mid Z = n) \mathbb{P}(Z = n)$$
$$= \sum_{n \in N} \lambda(J_n^1) \mathbb{P}(\mathbf{V}^n \le \mathbf{u})$$
$$= \sum_{n \in N} \lambda(J_n^1) C_n \left(\frac{u_1 - a_n^1}{\lambda(J_n^1)}, \dots, \frac{u_d - a_n^d}{\lambda(J_n^1)}\right),$$

which coincides with (4).

Now, we show two relevant examples of S-copulas.

Example 2.1. In the bivariate case, a copula C is said to be *shuffle of Min* (Mikusiński et al., 1992, Definition 2.1) if there is a natural number n, two partitions

 $0 = s_0 < s_1 < \dots < s_{\tilde{n}} = 1$ and $0 = t_0 < t_1 < \dots < t_{\tilde{n}} = 1$

of I, and a permutation ς of $\{1, \ldots, \tilde{n}\}$ such that each $[s_{i-1}, s_i] \times [t_{\varsigma(i)-1}, t_{\varsigma(i)}]$ is a square in which C distributes a mass $s_i - s_{i-1}$ uniformly spread along one of the diagonals.

Now, it can be easily proved that every shuffle of Min can be represented as an S–copula of the form $\langle (\mathcal{J}^1, \mathcal{J}^2), (C_i)_{i \in N} \rangle$, where:

- $N = \{0, 1, \dots, \widetilde{n} 1\};$
- $\mathscr{J}^1 = (J_i^1)_{i \in N}$ is the partition given by $J_i^1 = [s_{i-1}, s_i];$
- $\mathcal{J}^2 = (J_i^2)_{i \in N}$ is the partition given by $J_i^2 = [t_{\varsigma(i-1)}, t_{\varsigma(i)}];$
- $(C_i)_{i \in N}$ is a system of copulas where $C_i = M_2$, if the mass is distributed along the main diagonal of $[s_{i-1}, s_i] \times [t_{\varsigma(i)-1}, t_{\varsigma(i)}], C_i = W_2$, otherwise.

Recently, shuffles of Min have been also extended to the d-dimensional case (see (Mikusiński et al., 2009, section 6)). It is not difficult to show that also this extension can be represented in terms of S-copulas.

Example 2.2. Following Mesiar and Sempi (2010) (see also (Jaworski and Rychlik, 2008)), an ordinal sum of copulas can be introduced in the following way. Let L be a finite or countable set, let $([a_k, b_k])_{k \in L}$ be a system of sub-intervals of \mathbb{I} , and let $(C_k)_{k \in L}$ be a system in Cop_d . It is required that any two of the intervals $[a_k, b_k]$ ($k \in L$) have at most one endpoint in common. Then the *ordinal sum* C of $(C_k)_{k \in L}$ with respect to family of intervals $([a_k, b_k])_{k \in L}$ is the d-copula defined, for all $\mathbf{u} \in \mathbb{I}^d$ by

$$C(\mathbf{u}) = \begin{cases} a_k + (b_k - a_k) C_k \left(\frac{\min\{u_1, b_k\} - a_k}{b_k - a_k}, \dots, \frac{\min\{u_d, b_k\} - a_k}{b_k - a_k} \right), \\ \text{if } \min\{u_1, u_2, \dots, u_d\} \in]a_k, b_k[\text{ for some } k \in L, \\ \min\{u_1, u_2, \dots, u_d\}, \quad \text{elsewhere.} \end{cases}$$
(5)

Now, such a C can be represented as a suitable S–copula. In order to show this, without loss of generality, suppose that $([a_k, b_k])_{k \in L}$ is a partition of \mathbb{I} . Then, C given by (5) is an S–copula of type $\left\langle (\mathscr{J}^i)_{i=1,2,\dots,d}, (C_k)_{k \in L} \right\rangle$, where, for $i = 1,\dots,d$, $\mathscr{J}^i = (J^i_k)_{k \in L}$ is the partition given by $J^i_k = [a_k, b_k]$.

Finally, we end the section with a characterization of S-copulas in terms of measurepreserving transformations. We recall that it is known from Vitale (1996) (see also de Amo et al. (2010); Kolesárová et al. (2008)) that every copula can be represented in the form

$$C_{f_1,\dots,f_d}(\mathbf{u}) = \lambda \left(f_1^{-1}[0, u_1] \cap \dots \cap f_d^{-1}[0, u_d] \right)$$
(6)

for some measure-preserving transformations f_1, \ldots, f_d of \mathbb{I} .

Proposition 2.3. Let $C = \left\langle (\mathscr{J}^i)_{i=1}^d, (C_n)_{n \in N} \right\rangle$ be an *S*-copula. Suppose that every C_n can be represented in the form (6) by means of suitable f_1^n, \ldots, f_d^n . Then, *C* can be represented in the form (6) where f_i , for $i = 1, 2, \ldots, d$, is given by

$$f_i(t) = \begin{cases} f_i^n \left(\frac{t - a_n^i}{b_n^i - a_n^i} \right) (b_n^i - a_n^i) + a_n^i, & t \in]a_n^i, b_n^i[\\0, & otherwise. \end{cases}$$
(7)

PROOF. Let $C = \left\langle (\mathscr{J}^i)_{i=1}^d, (C_n)_{n \in N} \right\rangle$ be an S-copula. Let C^* be the copula that can be represented in the form (6) where, for $i = 1, 2, \ldots, d$, f_i is given by (7). In order to prove the assertion, it is enough to show that μ_C coincides with μ_{C^*} on every orthotope of type $J_n^1 \times \cdots \times J_n^d$. Now, let $\mathbf{u} \in J_n^1 \times \cdots \times J_n^d$. Then, it follows that:

$$\mu_{C^*} \left(\left(a_n^1, u_1 \right) \times \dots \times \left(a_n^d, u_d \right) \right) = \lambda \left(\left\{ t \colon f_1(t) \in \left(a_n^1, u_1 \right), \dots, f_d(t) \in \left(a_n^d, u_d \right) \right\} \right)$$

$$= \left(b_n^1 - a_n^1 \right) \lambda \left(\left\{ t \colon f_n^1(t) \in \left(0, \frac{u_1 - a_n^1}{\lambda(J_n^1)} \right), \dots, f_n^d(t) \in \left(0, \frac{u_d - a_n^d}{\lambda(J_n^1)} \right) \right\} \right)$$

$$= \left(b_n - a_n \right) \mu_{C_n} \left(\left(0, \frac{u_1 - a_n^1}{\lambda(J_n^1)} \right) \times \dots \times \left(0, \frac{u_d - a_n^d}{\lambda(J_n^1)} \right) \right)$$

$$= \mu_C \left(\left(a_n^1, u_1 \right) \times \dots \times \left(a_n^d, u_d \right) \right).$$

It follows that $\mu_{C^*}(A) = \mu_C(A)$ for every Borel set $A \subseteq J_n^1 \times \cdots \times J_n^d$, which is the desired assertion.

3. Approximation of copulas by means of shuffles

Here we show how the construction principle that we have introduced can be used in order to provide a way for approximating (in the L^{∞} norm) any copula. Firstly, we will consider some general aspects of the shuffling transformation.

Let N be an index set as in (S1). Let $\Psi \colon \mathscr{S}_N \times \operatorname{Cop}_d^N \to \operatorname{Cop}_d$ be the *shuffling mapping* defined by

$$\left(\left(\mathscr{J}^{i} \right)_{i=1}^{d}, (C_{n})_{n \in N} \right) \xrightarrow{\Psi} \left\langle \left(\mathscr{J}^{i} \right)_{i=1}^{d}, (C_{n})_{n \in N} \right\rangle, \tag{8}$$

which associates to each S–structure and to each system of copulas, the corresponding S–copula.

Now, for a fixed S-structure $S \in \mathscr{S}_N$, consider the function

$$\Psi_{(S,\cdot)} \colon \operatorname{Cop}_d^N \to \operatorname{Cop}_d, \quad (C_n)_{n \in N} \mapsto \left\langle S, (C_n)_{n \in N} \right\rangle.$$
(9)

Let d_{\sup} be the distance on Cop_d induced by the L^{∞} - norm. Let d_{\sup}^N be the distance defined on Cop_d^N by

$$d_{\sup}^{N}\left((C_{n})_{n\in N}, (C_{n}')_{n\in N}\right) = \sup_{n\in N} \sup_{\mathbf{u}\in\mathbb{I}^{d}} |C_{n}(\mathbf{u}) - C_{n}'(\mathbf{u})|$$
(10)

The following result holds.

Proposition 3.1. For any $S \in \mathscr{S}_N$, $\Psi_{(S,\cdot)}$ is a Lipschitz mapping between $(\operatorname{Cop}_d, d_{\sup}^N)$ and $(\operatorname{Cop}_d, d_{\sup})$.

PROOF. Let $S \in \mathscr{S}_N$. Let $(C_n)_{n \in N}$ and $(C'_n)_{n \in N}$ be in Cop_d^N . Set $\Psi_{(S,\cdot)}((C_n)_{n \in N}) = C$ and $\Psi_{(S,\cdot)}((C'_n)_{n \in N}) = C'$. We have

$$\sup_{\mathbf{u}\in\mathbb{I}^d} |C(\mathbf{u}) - C'(\mathbf{u})| = \sup_{n\in N} \left(\sup_{\mathbf{u}\in J_n^1\times\cdots\times J_n^d} |C(\mathbf{u}) - C'(\mathbf{u})| \right)$$

$$\leq \sup_{n\in N} \lambda(J_n^1) \sup_{\mathbf{u}\in\mathbb{I}^d} |C_n(\mathbf{u}) - C'_n(\mathbf{u})|$$

$$\leq \left(\sup_{n\in N} \lambda(J_n^1) \right) d_{\sup}^N \left((C_n)_{n\in N}, (C'_n)_{n\in N} \right),$$

which is the desired assertion.

Roughly speaking, for a given $S \in \mathscr{S}_N$, when two systems of copulas are close each other, so are the corresponding S-copulas.

Remark 3.1. Let $S \in \mathscr{S}_N$. Let $(C_n)_{n \in N}$ and $(C'_n)_{n \in N}$ be in Cop_d^N such that, for every $\mathbf{u} \in \mathbb{I}^d$, $C_n(\mathbf{u}) \leq C'_n(\mathbf{u})$ for every $n \in N$. Then it can be easily proved that $\Psi_{(S,\cdot)}(C_n)(\mathbf{u}) \leq \Psi_{(S,\cdot)}(C'_n)(\mathbf{u})$ for every $\mathbf{u} \in \mathbb{I}^d$. For more information about such a dependence order and its uses, see, for instance, (Joe, 1997; Nelsen, 2006).

A particular case of shuffling mapping is obtained when there exists $C \in \operatorname{Cop}_d$ such that $C_n = C$ for any $n \in N$. In such a case, for any $S \in \mathscr{S}_N$ the following mapping from Cop_d to Cop_d can be considered

$$\widetilde{\Psi}_{(S,\cdot)} \colon \operatorname{Cop}_d \to \operatorname{Cop}_d, \quad C \mapsto \left\langle S, (C_n = C)_{n \in N} \right\rangle,$$
(11)

A closer look at the proof of Proposition 3.1 allows to say that $\Psi_{(S,\cdot)}$ is a contraction and, hence, it has a unique fixed point. A possible way for for determining the fixed point of $\widetilde{\Psi}_{(S,\cdot)}$ is described below.

Example 3.1. This example is based on some results about self-similar measures. For more details see (Hutchinson, 1981, section 4).

Let \mathscr{X} be a complete metric space and let \mathscr{P} be the family of all probability measures defined on the Borel σ -algebra associated with \mathscr{X} . For every (finite or countable) family $\{F_t\}_{t\in T}$ of contraction maps on \mathscr{X} and every set of positive real numbers $\{\rho_t\}_{t\in T}$, we can define the mapping $F_{\rho} \colon \mathscr{P} \to \mathscr{P}$ such that $F_{\rho}(\mu)(A) = \sum_{t\in T} \rho_t \mu(F_t^{-1}(A))$ for every Borel set A. It follows from (Hutchinson, 1981, subsection 4.4) that there exists a unique probability measure μ such that $F_{\rho}(\mu) = \mu$.

Now, let $S \in \mathscr{S}_N$. It can be easily proved that the fixed point of $\Psi_{(S,\cdot)}$ is the copula C such that its induced measure μ_C is associated with the family of contraction maps $\{F_n\}_{n\in N}$, where $F_n: \mathbb{I}^d \to J_n^1 \times \cdots \times J_n^d$ is given by

$$F_n(x_1,\ldots,x_d) = \lambda(J_n^1)(x_1,\ldots,x_d) + (a_n^1,\ldots,a_n^d),$$

where, for $i \in \{1, 2, ..., d\}$, $J_n^i = [a_n^i, b_n^i]$ and the set of positive real numbers is given by $\{\lambda(J_n^1)\}_{n \in \mathbb{N}}$.

Now, given a copula C, let us consider the set $\{\widetilde{\Psi}_{(S,\cdot)}(C) \mid S \in \mathscr{S}\}$ formed by all Scopulas generated from C and any S-structure S by means of $\widetilde{\Psi}_{(S,\cdot)}$. The following result
holds.

Proposition 3.2. Let $C \in \operatorname{Cop}_d$. Then the set $\{\widetilde{\Psi}_{(S,\cdot)}(C) \mid S \in \mathscr{S}\}$ is dense in Cop_d with respect to the L^{∞} norm.

PROOF. Let $C \in \text{Cop}_d$. Let $p \in \mathbb{N}, p \geq 2$. Divide \mathbb{I}^d into p^d orthotopes of type

$$B_m^p = \left[\frac{a_m^1}{p}, \frac{a_m^1 + 1}{p}\right] \times \dots \times \left[\frac{a_m^d}{p}, \frac{a_m^d + 1}{p}\right]$$
(12)

where $m \in \mathcal{M} = \{0, 1, \dots, p^d - 1\}, a_m^i \in \{0, 1, \dots, p - 1\}$ for every $i \in \{1, \dots, d\}$, and m can be written in the form

$$m = a_m^1 + a_m^2 p + \dots + a_m^d p^{d-1}.$$
 (13)

We denote by $\mathscr{M}_{k,i}$ the subset of \mathscr{M} formed by all $m \in \mathscr{M}$ such that m can be written in the form (13) with $a_m^i = k$.

Notice that, since μ_C is *d*-fold stochastic, for every $m \in \mathcal{M}_{k,i}$

$$\sum_{m \in \mathscr{M}_{k,i}} \mu_C(B_m^p) = \lambda \left(\mathbb{I}^{i-1} \times \left[\frac{k}{p}, \frac{k+1}{p} \right] \times \mathbb{I}^{d-i} \right) = \frac{1}{p}.$$

Let $i \in \{1, 2, ..., d\}$. Along the *i*-th axis, divide each interval of type $\left[\frac{k}{p}, \frac{k+1}{p}\right]$ into a system of closed intervals $(J_m^i)_{m \in \mathcal{M}_{k,i}}$ such that:

- (a) $\lambda(J_m^i) = \mu_C(B_m^p)$ for every $m \in \mathscr{M}_{k,i}$;
- (b) if $a_{m_1}^i < a_{m_2}^i$, then $\sup (J_{m_1}^i) \le \inf (J_{m_2}^i)$;
- (c) if $a_{m_1}^i = a_{m_2}^i$ and $m_1 < m_2$, then $\sup (J_{m_1}^i) \le \inf (J_{m_2}^i)$.

Notice that, because of condition (a), some of the intervals J_m^i may be of measure zero. In such a case, they will not be considered.

By means of previous procedure, \mathbb{I} can be partitioned into intervals in the following way:

$$\mathbb{I} = \bigcup_{k=0}^{p-1} \left[\frac{k}{p}, \frac{k+1}{p} \right] = \bigcup_{k=0}^{p-1} \left(\bigcup_{m \in \mathscr{M}_{k,i}} J_m^i \right).$$
(14)

Let $\mathscr{J}^i = (J^i_m)_{m \in \mathscr{M}}$ be the system of intervals of eq. (14). Then it is easily proved that $(\mathscr{J}^i)^d_{i=1}$ forms an S-structure (which depends on p). Let C_p be the S-copula of type

 $\langle (\mathscr{J}^i)_{i=1}^d, (C_m = C)_{m \in \mathscr{M}} \rangle$. For every B_m^p of type (12), it can be proved that $\mu_{C_p}(B_m^p) = \mu_C(B_m^p)$.

Now, consider a continuous function f on \mathbb{I}^d . Then, for every $\varepsilon > 0$ and for every $m \in \mathcal{M}$, there exists $q \in \mathbb{N}$ such that, for every p > q and for every $\mathbf{x}, \mathbf{x}' \in B_m^p$, $|f(\mathbf{x}) - f(\mathbf{x}')| < \varepsilon$. Therefore

$$\begin{split} \sum_{m} \min_{\mathbf{x} \in B_{m}^{p}} \left(f(\mathbf{x}) \right) \mu_{C}(B_{m}^{p}) &= \sum_{m} \min_{\mathbf{x} \in B_{m}^{p}} \left(f(\mathbf{x}) \right) \mu_{C_{p}}(B_{m}^{p}) \\ &\leq \int_{\mathbb{I}^{d}} f(\mathbf{x}) d\mu_{C_{p}} \leq \sum_{m} \max_{\mathbf{x} \in B_{m}^{p}} \left(f(\mathbf{x}) \right) \mu_{C_{p}}(B_{m}^{p}) \\ &\leq \sum_{m} \max_{\mathbf{x} \in B_{m}^{p}} \left(f(\mathbf{x}) \right) \mu_{C}(B_{m}^{p}) \leq \sum_{m} \min_{\mathbf{x} \in B_{m}^{p}} \left(f(\mathbf{x}) \right) \mu_{C}(B_{m}^{p}) + \varepsilon, \end{split}$$

where the sums are taken over all possible indices m. It follows that

$$\left|\int_{\mathbb{I}^d} f(\mathbf{x}) d\mu_{C_p} - \int_{\mathbb{I}^d} f(\mathbf{x}) d\mu_C\right| < \varepsilon$$

when p > q. For the arbitrariness of f, we can conclude that C_p converges pointwise to C as p tends to $+\infty$.

Remark 3.2. By using the same idea of the proof of Proposition 3.2, it should be noticed that, any $C \in \text{Cop}_d$ could be approximated by means of S-copulas constructed from systems of copulas formed by elements not necessarily coinciding.

In the bivariate case, Proposition 3.2 was already known for the case $C = M_2$ (see (Mikusiński et al., 1992, Theorem 3.1)). In particular, since Proposition 3.2, it follows that the independence copula $\Pi_d(\mathbf{u}) = u_1 \cdots u_d$ can be approximated by means of elements of the set $\{\widetilde{\Psi}_{(S,\cdot)}(M_d) \mid S \in \mathscr{S}\}$. Intuitively speaking, taking into account the probabilistic interpretation of S-copulas and the fact that M_d describe comonotone dependence, it follows that any sequence of independent random variables can be approximated by a sequence of random variables that are completely dependent (in the sense that each random variable is a deterministic function of the others). This intriguing fact was noted for the first time by Kimeldorf and Sampson (1978).

However, we want to stress that Proposition 3.2 strongly depends on the topology that we are considering over Cop_d .

Consider, for example, the set of all finite signed measures on $(\mathbb{I}^d, \mathscr{B}(\mathbb{I}^d))$. Such a set is a real vector space that, equipped with the total variation norm, forms a Banach space. It is known that a sequence of finite signed measure $(\mu_n)_n$ converges to μ if, and only if, for every $B \in \mathscr{B}(\mathbb{I}^d), \ \mu_n(B) \to \mu(B)$ as $n \to +\infty$.

Now, we can consider Cop_d (precisely, its induced measure) as subset of the set of all signed measures. Then Cop_d can be equipped with the topology induced by the total variation norm. In such a topology, it holds that, given $C \in \operatorname{Cop}_d$, the set $\{\widetilde{\Psi}_{(S,\cdot)}(C) \mid S \in \mathscr{S}\}$ need not be dense in Cop_d . In fact, consider $C = M_2$. Then the elements of $\{\widetilde{\Psi}_{(S,\cdot)}(M_2) \mid S \in \mathscr{S}\}$ concentrates the probability mass on a set of 2-dimensional Lebesgue measure 0 and, thus, they cannot approximate in the total variation norm the copula Π_2 (or, in general, any copula with a non-trivial absolutely continuous component).

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