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# Random queues and risk averse users* 

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#### Abstract

We analyse Nash equilibrium in time of use of a congested facility. Users are risk averse with general concave utility. Queues are subject to varying degrees of random sorting, ranging from strict queue priority to a completely random queue. We define the key "no residual queue" property, which holds when there is no queue at the time the last user arrives at the queue, and prove that this property holds in equilibrium under all queueing regimes considered. The no residual queue property leads to simple results concerning the equilibrium utility of users and the timing of the queue.


Keywords: Congestion; Queuing; Risk aversion; Endogenous arrivals. JEL codes: D00; D80

[^0]
## 1 Introduction

We generalise the Vickrey (1969) analysis of bottleneck congestion to allow for random queue sorting as well as more general scheduling preferences. The paper shows that the fundamental insights of Vickrey remain valid in these circumstances. In spite of users being risk averse, random queue sorting turns out to play no role for the properties of equilibrium that are relevant for regulation of congestion.

Enormous amounts of time are lost queueing. Just for private transportation, the cost of congestion in Europe and the US is equivalent to more than 1 percent of GDP (International Transport Forum, 2007; Texas Transportation Institute, 2007) and unpriced congestion leads to excess urban sprawl (Arnott, 1979). Dynamic models of traffic congestion are reviewed in de Palma and Fosgerau (2011). Congestion arises not only on roads. Queues occur regularly also in supermarkets, banks, public offices, restaurants (Becker, 1991), movie theatres, concert ticket sales, at ski lifts (Barro and Romer, 1987) and toll road booths, in airports (Daniel, 1995), computer systems, communications systems, web services, call centers, and many other systems. Queueing is also relevant for understanding competitive markets, where queueing plays a role in allocating goods among consumers and trade from firms is congestible (Sattinger, 2002). So it is clearly important to understand queueing phenomena.

Economic analyses of congestion mostly assume strict first-in-first-out (FIFO) queue discipline, whereby the order of arrival at the queue is preserved. Many real queues, however, involve an element of random sorting. An extreme case is a pure random queue. ${ }^{1}$ An example is a (virtual) queue to get through on a busy telephone line (de Palma and Arnott, 1989), where every person present in the queue at a given time has the same probability of being served as any other person in the queue, regardless of how long each has been in the queue. Other queues also involve random queue sorting. There are random opportunities for overtaking on roads; in a supermarket, FIFO applies to individual checkout lines, but not to the supermarket checkout system as a whole (Blanc, 2009); also queueing for public transport is often not strictly FIFO (Yoshida, 2008). In general, we may think that strict FIFO rarely occurs. It is thus of interest to determine the properties of queues that are not strictly FIFO. ${ }^{2}$

The economic literature has previously paid attention to the properties of user equilibrium in queues with strict queue priority using the seminal Vickrey (1969)

[^1]bottleneck model. This model offers many insights that are central to the understanding of congested demand peaks. Arnott, de Palma and Lindsey (1993) summarise a number of these. In the Vickrey model, users arrive at a bottleneck where they wait in a FIFO queue until they are served by the bottleneck. The bottleneck serves users at a fixed rate. A continuum of users choose their time of arrival at bottleneck into the queue to minimise a scheduling cost, which is linear in time spent in the queue, time early and time late at the destination. The time-varying arrival rate at the bottleneck is then determined endogenously in response to the evolution of the queue. The model is closed by assuming Nash equilibrium. ${ }^{3}$

We extend the Vickrey model in two ways: first by allowing for random queue sorting, and second by allowing users to have general concave utility depending on duration in the queue as well as on time of exit from the queue. Random queue sorting causes randomness in outcomes and the concavity of utility implies that users are risk averse.

We then introduce the no residual queue (NRQ) property for a queue with a general random sorting mechanism. A residual queue is a queue that remains at the time of arrival at the bottleneck of the last user. The NRQ property is said to hold when the queue has vanished at the time of the last arrival. By definition, the equilibrium utilities of the first and the last user are equal. The NRQ property is then sufficient to establish the equilibrium time interval of arrivals. A number of useful results follow. In particular, we determine the equilibrium utility and the marginal utility of adding users under Nash equilibrium. This is the information that is needed in order to determine the optimal capacity provision as well the optimal constant toll.

The basic insight is then that it is the NRQ property that underlies the elegance of the Vickrey analysis of congestion. When the NRQ property holds, it does not matter that the queue is subject to random sorting. Remarkably, the optimal capacity, the optimal constant toll as well as the optimal time varying toll are unaffected by random queue sorting.

So it is of interest to establish when the NRQ property holds. We identify a condition on scheduling preferences that is sufficient for the NRQ property under any degree of random queue sorting. It turns out to be sufficient that users must be always willing to arrive one minute later in exchange for spending one minute less in the queue. This condition cannot be relaxed in general.

We also show that the optimal time varying toll is also not affected by random queue sorting, since there is no queue under the optimal time varying toll. This result holds regardless of whether the NRQ property holds in no toll equilibrium.

The paper is organised as follows. Section 2 presents the general framework,

[^2]introduces the NRQ property, and derives the results that follow from this property.

The remainder of the paper is devoted to establishing the NRQ property under various degrees of random queue sorting. First, Section 3 reviews and generalises the standard case of strict queue priority and establishes that the NRQ property holds here. Next, Section 4 considers the oppposite case of no queue priority where users to be served are chosen completely at random from the queue. We establish also the NRQ property for this case given the above condition on preferences.

Section 5 considers the intermediate case, which we refer to as loose queue priority. Under this regime, the probability of being served at time $t$, conditional on being in the queue at time $t$, increases with the time spent in the queue. We show that the above condition on marginal utilities is again sufficient to guarantee the NRQ property to hold in general when queue priority is loose. Some concluding remarks are provided in Section 6.

## 2 Model specification

Consider $N$ users treated as a continuum. They must all pass through a bottleneck which has a capacity of $\psi$ users per time unit. Users arrive at the bottleneck at the back of the queue at the locally bounded time dependent rate $\rho(a) \geq 0$ during the interval $\left[t_{0}, t_{1}\right]$, where $t_{0}$ and $t_{1}$ are the minimum and the maximum of the support of $\rho$. The cumulative arrival rate up to time $a$ is denoted by $R(a)=$ $\int_{t_{0}}^{a} \rho(s) d s$, and $R(\cdot)$ is continuous since $\rho(\cdot)$ is locally bounded. Furthermore, $R(\cdot)$ is differentiable at all points of continuity of $\rho(\cdot)$. Users enter a vertical queue of length $Q(a)$ at time $a$, which represents the number of users who have arrived at the entrance of the bottleneck but not yet exited. The queue length evolves according to ${ }^{4}$

$$
\begin{equation*}
Q(a)=R(a)-\int_{t_{0}}^{a}\left[\psi 1_{\{Q(s)>0\}}+\min (\psi, \rho(s)) 1_{\{Q(s)=0\}}\right] d s, \tag{1}
\end{equation*}
$$

so $Q(\cdot)$ is continuous and also differentiable at points of continuity of $\rho(\cdot)$. Denote the minimum and the maximum of the support of the queue length $Q(\cdot)$ as $\tau_{0}$ and $\tau_{1}$.

The last user exits the queue at time $\tau_{1}$. This implies that $\tau_{1} \geq t_{1}$. If $Q\left(t_{1}\right)=$ 0 , then $\tau_{1}=t_{1}$. If $Q\left(t_{1}\right)>0$, we say that there is a residual queue at time $t_{1}$. In this case, $\tau_{1}$ is given by $Q\left(t_{1}\right)=\psi\left(\tau_{1}-t_{1}\right)$, since the queue length at time $t \in\left[t_{1}, \tau_{1}\left[\right.\right.$ is strictly positive if $Q\left(t_{1}\right)>0$.

[^3]We shall consider various queueing regimes. At one extreme we have the strict queue priority case, considered by Vickrey (1969), where the queue obeys the first-in-first-out principle (FIFO). At the other extreme we have the no queue priority case, where the user to exit at each instant is chosen completely at random from the queue. Therefore the probability of exit from the queue at some instant is the same for all users present in the queue and does not depend on how much time each has spent in the queue. In between these two cases, we have the loose queue priority case. In this case, users who are in the queue in a given instant have a higher probability of exit if they have spent more time in the queue.

We formalise these cases below through the conditional density of exit times $f(t \mid a)$, which describes the probability of exit at time $t$ conditional on arrival at the bottleneck at time $a \leq t$. This conditional density depends on the arrival rate $\rho(\cdot)$, but it is exogenous from the perspective of a single atomistic user. In all cases, except the strict queue priority case that is treated separately, we assume that $f(t \mid a)$ is differentiable as a function of $a$.

A user arrives at the bottleneck at time $a$ and exits at time $t$ with $a \leq t$, such that his duration in the queue is $d=t-a$. The arrival time is chosen by the user while the exit time is determined by the queue. He has a preferred exit time $t^{*}$. Utility is associated with the duration in the queue and the deviation $t-t^{*}$ of the exit time from the preferred exit time. Assume homogenous users and write utility as $u\left(d, t-t^{*}\right)$. Utility is concave, has a unique maximum at $d=0$ for any $t-t^{*}$ and a unique maximum at $t=t^{*}$ for any duration in the queue. Given any exit time, users strictly prefer zero duration in the queue to anything else, and given any duration in the queue, users strictly prefer exiting at the preferred time to anything else. With these assumptions, utility is strictly decreasing in $d$, strictly increasing in $t$ for $t<t^{*}$ and strictly decreasing in $t$ for $t>t^{*}$. We normalise $t^{*}=0$ at no loss of generality.

Users choose their arrival time $a$ to maximise their expected utility given by

$$
\begin{equation*}
E(u \mid a)=\int_{a}^{\infty} u(t-a, t) f(t \mid a) d t . \tag{2}
\end{equation*}
$$

We specify the following assumptions concerning the utility function. Denote the partial derivatives of $u$ with respect to duration and exit time as $u_{1}$ and $u_{2}$, respectively. We require first and second derivatives to exist, except $u_{2}(d, 0)$ which is not required to exist. Clearly, users who exit late are always willing to exit one minute earlier in exchange for spending one minute less in the queue. We require that also users who exit early are always willing willing to exit one minute earlier in exchange for spending one minute less in the queue. This first condition is assumed throughout the paper.

Condition $1 u_{1}(d, t)+u_{2}(d, t)<0$ for all $t<0$.
We shall also have use for a second condition stating that users who exit late are always willing to exit one minute later in exchange for spending one minute less in the queue. For easy reference we shall call this the acceptable lateness condition. Clearly, users who exit early always satisfy the acceptable lateness condition. It is assumed where indicated.

Condition 2 (Acceptable lateness) $u_{1}(d, t)<u_{2}(d, t)$ for all $t>0$.
We shall refer to the special case of linear utility, which is the case investigated by Vickrey (1969) and Arnott et al. (1993). This will be important for results and also helps in facilitating interpretation of results. The linear utility formulation is ${ }^{5}$

$$
u(d, t)=-\alpha d-\beta t^{-}-\gamma t^{+}
$$

where the parameters $\alpha, \beta$ and $\gamma$ are strictly positive. For the linear case, condition 1 states that $\beta<\alpha$, while the acceptable lateness condition 2 states that $\gamma<\alpha$. Yoshida (2008) summarises empirical evidence and concludes that both cases $\gamma<$ $\alpha$ and $\gamma>\alpha$ are empirically relevant.

We consider Nash equilibrium in pure strategies as the benchmark for rational behavior. ${ }^{6}$ The Nash equilibrium is defined by the requirement that, conditional on the actions of other users, no user has incentive to change his own action. With identical users, this requirement turns into the condition that the expected utility is constant and maximal over the times at which users arrive, i.e. over the support of $\rho$.

Below we shall briefly touch the issue of optimal tolling. For this we need to specify how a toll payment enters utility and a social welfare function with respect to which optimality is defined. We take any toll payment to be simply subtracted from utility, which then must be in monetary units. When expected utility is constant over users, we define a social welfare function as $N$ times the equilibrium expected utility plus aggregate toll revenues.

In the strict queue priority case, the exit time is given deterministically as a function of the arrival time. We then require that utility is constant over all arrival times $a$ with $\rho(a)>0$.

In all other cases considered, exit time is random. The Nash condition implies that the expected utility is constant, i.e. $\frac{\partial E(u \mid a)}{\partial a}=0$, for all $a$ such that $\rho(a)>0$.

[^4]This leads to the equation

$$
-u(0, a) f(a \mid a)+\int_{a}^{\infty}\left[u(t-a, t) \frac{\partial f(t \mid a)}{\partial a}-u_{1}(t-a, t) f(t \mid a)\right] d t=0 .
$$

Recall that $t_{0}$ and $t_{1}$ are the times of the first and the last arrival. The following Lemma shows that in equilibrium the queue begins when the first user arrives at the bottleneck and that the queue ends at the earliest when the last user arrives.

Lemma 1 The support of $Q$ is a finite interval in Nash equilibrium, with $-\infty<$ $t_{0}=\tau_{0}<0$ and $0<t_{1} \leq \tau_{1}<\infty$.

All proofs are given in the appendix. We now introduce the no residual queue property.

Definition 1 The no residual queue (NRQ) property holds if $\tau_{1} \leq t_{1}$.
The NRQ property ensures that $\left[t_{0}, t_{1}\right]=\left[\tau_{0}, \tau_{1}\right]$ in Nash equilibrium by Lemma 1. This means that the first and last users experience no queue, and hence that $u\left(0, t_{0}\right)=u\left(0, t_{1}\right)$. Moreover, all users are able to pass the bottleneck during [ $\left.t_{0}, t_{1}\right]$, which implies that $t_{1}=t_{0}+N / \psi$. These two observations pin down the equilibrium utility as shown in the following Proposition.

Proposition 1 Consider Nash equilibrium where the NRQ property holds. Then the interval of arrival, $\left[t_{0}, t_{1}\right]$ with $t_{0}<0<t_{1}$, is uniquely determined by $t_{1}=$ $t_{0}+\frac{N}{\psi}$ and $u\left(0, t_{0}\right)=u\left(0, t_{0}+\frac{N}{\psi}\right)$. The expected utility of any user is $u\left(0, t_{0}\right)$. The marginal change in expected utility from additional users is

$$
\begin{equation*}
\frac{\partial E(u \mid a)}{\partial N}=\frac{1}{\psi} \frac{u_{2}\left(0, t_{0}\right) u_{2}\left(0, t_{1}\right)}{u_{2}\left(0, t_{1}\right)-u_{2}\left(0, t_{0}\right)}<0, \tag{3}
\end{equation*}
$$

which decreases in the number of users.
The preceding Proposition exhibits the central properties of the bottleneck model. In particular, the expected utility of any user is known as a function of the number of users, which makes it easy to derive the optimal capacity. If the number of users is allowed to be elastic, then Proposition 1 can be used to determine the optimal constant toll. Below we establish that the NRQ property holds under strict, loose and no queue priority and hence that Proposition 1 applies in all these regimes.

The optimal time varying toll eliminates queueing. Hence it is not affected by random queue sorting. This is formalised in the following Proposition, which is stated without proof.

Proposition 2 The optimal time varying toll is

$$
\left[u(0, a)-u\left(0, t_{0}\right)\right]^{+},
$$

where $t_{0}$ is the first arrival time in Nash equilibrium under strict queue priority.

## 3 Strict queue priority

This is the case considered by Vickrey (1969) and Arnott et al. (1993) in the context of transportation and telecommunication, except for our more general formulation of user preferences. Users exit strictly in the order in which they arrive, hence exit time is a deterministic function of arrival time. A user arriving at time $a$ is served at time $a+q(a)$, where $q(a)=Q(a) / \psi$. We have $q(a)=\frac{R(a)}{\psi}-\left(a-t_{0}\right)$, since there is always queue during $\left[t_{0}, t_{1}\right]$. Therefore

$$
\begin{equation*}
q^{\prime}(a)=\frac{\rho(a)}{\psi}-1 . \tag{4}
\end{equation*}
$$

The queue satisfies the NRQ property, since if the last user arrives at time $t_{1}$ when $Q\left(t_{1}\right)>0$, then his exit time will be $\tau_{1}>t_{1}$. This implies that he could postpone arrival until $\tau_{1}$ to obtain zero duration in the queue while leaving the exit time unchanged, in contradiction of Nash equilibrium. We highlight this in a Proposition.

Proposition 3 The NRQ property holds in Nash equilibrium under strict queue priority.

Now $t_{1}=\tau_{1}$ so that Proposition 1 applies and $t_{1}=t_{0}+N / \psi$. We shall briefly review the analysis of the bottleneck model for the case of general concave scheduling preferences.

By concavity of $u, t_{0}$ is the unique solution to the equation

$$
u\left(0, t_{0}\right)=u\left(0, t_{0}+N / \psi\right) .
$$

The utility function is given by $u(q(a), a+q(a))$. We omit below the arguments of $u(\cdot)$ to economise on notation. The first-order condition for Nash equilibrium is $\frac{\partial u}{\partial a}=u_{1} \cdot q^{\prime}(a)+u_{2} \cdot\left[1+q^{\prime}(a)\right]=0, a \in\left[t_{0}, t_{1}\right]$. Using (4) leads to the equilibrium arrival rate

$$
\begin{equation*}
\rho(a)=\psi \frac{u_{1}}{u_{1}+u_{2}}>0 \tag{5}
\end{equation*}
$$

which is strictly positive on $\left[t_{0}, t_{1}\right]$ by Condition 1 . (Condition 2 is not necessary here.)


Figure 1: The evolution of the queue under strict queue priority with linear utility

By (5), $\rho(a)>\psi$ exactly when $u_{2}>0$, which occurs exactly when $a+q(a)<$ 0 . Thus the queue builds up until time $\tilde{a}<0$ defined by $\tilde{a}+q(\tilde{a})=0$, at which time the queue begins to diminish.

The arrival rate is decreasing. To see this for $a \neq \tilde{a}$, differentiate the equilibrium condition twice to find

$$
\left(q^{\prime}(a), 1+q^{\prime}(a)\right)\left(\begin{array}{ll}
u_{11} & u_{12} \\
u_{12} & u_{22}
\end{array}\right)\left(q^{\prime}(a), 1+q^{\prime}(a)\right)^{T}+\left(u_{1}+u_{2}\right) q^{\prime \prime}(a)=0
$$

The first term here is negative since $u(\cdot)$ is concave, and hence the second term is positive. Then $q^{\prime \prime}(a) \geq 0$ by Condition 1. Find from (4) that $\rho^{\prime}(a) / \psi=q^{\prime \prime}(a)$, such that $\rho^{\prime}(a) \geq 0$. The utility function is not required to be differentiable at the point $(q(\tilde{a}), \tilde{a}+q(\tilde{a}))$.

For any small $\varepsilon>0$, we have $u_{2}(q(\tilde{a}+\varepsilon), \tilde{a}+\varepsilon+q(\tilde{a}+\varepsilon))<0$ and $0<$ $u_{2}(q(\tilde{a}-\varepsilon), \tilde{a}-\varepsilon+q(\tilde{a}-\varepsilon))$, while $u_{1}(q(a), a+q(a))<0$. Hence $\rho(\cdot)$ can only jump down at $\tilde{a}$. Such a jump occurs in the linear case, where the arrival rate is $\rho(a)=\psi \frac{\alpha}{\alpha-\beta}$ for $a<\tilde{a}$, and $\rho(a)=\psi \frac{\alpha}{\alpha+\gamma}$ for $a>\tilde{a}$, which is piecewise constant with a downward jump at $\tilde{a}=-\frac{\beta}{\alpha} \frac{\gamma}{\beta+\gamma} \frac{N}{\psi}$.

Figure 1 shows the evolution of the queue under strict queue priority with linear utility. The curve $R(a)$ is the cumulative arrival rate, the kink occurs at the time where users exit at time $t^{*}=0$. The curve $\psi\left(t-t_{0}\right)$ represents the cumulative number of exits from the queue. The curve $q(a)$ shows the duration in the queue for users entering the queue at time $a$. It is maximal for users who exit at time $t^{*}$. The curve $a+q(a)$ indicates the exit time for users entering the queue at time $a$.

## 4 No queue priority

With no queue priority, users to exit at any time are chosen at random at the rate $\psi$ such that all users present in the queue have the same chance to exit. We first formalise this notion and show that if there is a residual queue at the time $t_{1}$ of the last arrival at the bottleneck, then the distribution of exit times conditional of being in the queue at time $t_{1}$ is uniform. Using this result, we then show that the acceptable lateness condition 2 is sufficient to guarantee the NRQ property in Nash equilibrium under no queue priority and that the equilibrium arrival rate is indeed positive. The acceptable lateness condition cannot be relaxed in general.

We formulate the no queue priority assumption by means of the hazard rate using concepts and results from duration analysis (Lancaster, 1990). The hazard rate does not depend on $a$ as all users present in the queue at time $t$ have the same probability to exit. Define the hazard rate of a user who is present in the queue at time $t$ as

$$
\begin{equation*}
\lambda(t)=\frac{f(t \mid a)}{1-F(t \mid a)}=\frac{\psi}{Q(t)}, \tag{6}
\end{equation*}
$$

where $f(t \mid a)$ and $F(t \mid a)$ are respectively the density and cumulative distribution of exit time $t$ conditional on being in the queue at time $a$. The survivor function $1-F(t \mid a)$ can be expressed in terms of the integrated hazard by

$$
\begin{equation*}
1-F(t \mid a)=e^{-\int_{a}^{t} \lambda(s) d s} \tag{7}
\end{equation*}
$$

The following technical Lemma concerns the conditional density of exit times when there is a residual queue after the last arrival. It states that when a pool of users exit with equal probability at a constant rate during some interval, then the exit time for each of them is uniformly distributed over this interval.

Lemma 2 Consider the no queue priority case. Let $t_{1}$ be the time of the last arrival and assume that $Q\left(t_{1}\right)>0$. Then the exit time conditional on being in the queue at time $a\left(t_{1} \leq a \leq \tau_{1}\right)$ is uniformly distributed over the interval $\left[a, \tau_{1}\right]$ with $f(t \mid a)=\lambda(a), t \in\left[a, \tau_{1}\right]$. Furthermore, $\lambda^{\prime}(a)=\lambda^{2}(a)$.

We shall now show that concave utility as defined above together with the acceptable lateness condition 2 is sufficient to establish the no residual queue property for the no queue priority case. The acceptable lateness condition states that the marginal disutility of lateness is smaller than the marginal disutility of duration in the queue. If the queue diminishes quickly enough as arrival time increases, users will then postpone arrival until the queue is no longer decreasing so quickly. The second half of the Proposition establishes that condition 2 is also necessary for the NRQ property under linear utility. Hence condition 2 cannot be relaxed in general.


Figure 2: The evolution of the queue under no queue priority with linear utility

Proposition 4 The acceptable lateness condition 2 is sufficient for the no residual queue property to hold. Under linear utility, condition 2 is also necessary.

Proposition 5 establishes that the equilibrium arrival rate is always strictly positive under the acceptable lateness condition 2 and that the condition cannot be relaxed in general.

Proposition 5 The acceptable lateness condition 2 is sufficient for the equilibrium arrival rate to be strictly positive over the interval $\left[t_{0}, t_{1}\right]$ defined by $u\left(0, t_{0}\right)=$ $u\left(0, t_{1}\right)$. Under linear utility, condition 2 is also necessary.

Figure 2 illustrates the evolution of the queue under no queue priority and linear utility. For comparison, the figure also shows the evolution of the queue under strict queue priority. The kinked curves are the cumulative arrival rates. Note that in the NQP case, the kink in the cumulative arrival rate occurs at time $t^{*}=0$. The straight curve represents the cumulative number of exits from the queue.

## 5 Loose queue priority

This section concerns the case of loose queue priority, which we shall define as an intermediate case between the cases examined so far of strict and no queue priority. We shall show that the acceptable lateness condition 2 is sufficient to establish the no residual queue property for the case of loose queue priority; hence Condition 2 implies that Proposition 1 holds.

Under strict queue priority, users exit strictly in the order in which they arrive. Under no queue priority, users present in the queue at any instant all have the same
probability of exit. The intermediate case of loose queue priority is defined by requiring that at any instant, users whose present duration in the queue is longer have a higher chance to exit than users whose present duration in the queue is shorter. So arrival time matters, even if queue priority is not strict. There are very many possibilities for explicitly defining processes that have this property. The example below provides one simple way to model loose priority.

Example 1 Introduce a variable $N(a, t)$ denoting the number of users in the quene at time $t$ who arrived at the queue after time $a$, $a \leq t$. We have $N(a, t) \leq$ $Q(t)$. Furthermore, $N(t, t)=0$ and $N\left(t_{0}, t\right)=Q(t)$. At time $t$, there are $Q(t)-N(a, t)$ users in the queue who arrived earlier than $a$. Users exit the queue at the rate $\psi$, but under loose queue priority the hazard is not the same for everybody, it depends on the time of arrival a. We want the hazard rate, denoted $\lambda(t \mid a)$, to increase with the duration of the stay in the queue. One possible way of achieving this is by specifying the hazard rate to be

$$
\lambda(t \mid a)=H\left(\frac{N(a, t)}{Q(t)}\right) \frac{\psi}{Q(t)},
$$

where $H(\cdot)$ is an increasing density on the unit interval with $H(0)<1$. This hazard rate increases with the duration in the queue. The definition encompasses strict and no queue priority as limiting cases as $H(\cdot)$ approaches either a point mass at 1 or a uniform density. The hazard for the last user has $\lambda\left(t \mid t_{1}\right)=H\left(\frac{N\left(t_{1}, t\right)}{Q(t)}\right) \frac{\psi}{Q(t)}=H(0) \frac{\psi}{Q(t)}<\frac{\psi}{Q(t)}\left(t_{1} \leq t\right)$.

Recall that $t_{1}$ is the time of the last arrival at the queue, while $\tau_{1}=t_{1}+$ $Q\left(t_{1}\right) / \psi$ is the time of the last exit from the queue. When there is a residual queue $Q\left(t_{1}\right)>0$ then $\tau_{1}>t_{1}$.

In the case of no queue priority we noted in Proposition 4 that the acceptable lateness condition 2 implies that $Q\left(t_{1}\right)>0 \Rightarrow E\left(u \mid \tau_{1}\right)>E\left(u \mid t_{1}\right)$, contradicting that we can have $Q\left(t_{1}\right)>0$ in Nash equilibrium. In this case the distribution of exit times conditional on entry at time $t_{1}$ is the uniform distribution over the interval $\left[t_{1}, \tau_{1}\right]$. We denoted this by $F\left(t \mid t_{1}\right)$.

In the case of strict queue priority we noted that $Q\left(t_{1}\right)>0 \Rightarrow u\left(\tau_{1}\right)>$ $u\left(t_{1}\right)$, which again contradicts that we can have $Q\left(t_{1}\right)>0$ in Nash equilibrium. This happens because the last user entering at time $t_{1}$ will exit at time $\tau_{1}$ with probability 1 .

In order to establish the no residual queue property for the case of loose priority, it is sufficient to give a condition on the distribution of exit times conditional on entry at time $t_{1}$. Denote this distribution by $\tilde{F}\left(\cdot \mid t_{1}\right)$. We require that loose queue priority satisfies the following condition.

Condition 3 (Loose queue priority) Under loose queue priority, the distribution of exit times conditional on arriving last, $\tilde{F}\left(\cdot \mid t_{1}\right)$, first-order stochastically dominates $F\left(\cdot \mid t_{1}\right)$, where $F\left(\cdot \mid t_{1}\right)$ is the uniform distribution over $\left[t_{1}, \tau_{1}\right]$ with $\tau_{1}=t_{1}+Q\left(t_{1}\right) / \psi$.

The loose queue priority condition immediately implies that if there is a residual queue, then the last user to arrive is worse off under loose queue priority than under no queue priority (the utility function is decreasing in exit time, for any given arrival time). Hence Proposition 4 leads naturally to the following Proposition.

Proposition 6 Under loose queue priority, the acceptable lateness condition 2 implies the no residual queue property in Nash equilibrium.

Hence Condition 2 is sufficient to ensure that Proposition 1 applies, also in the case of loose queue priority.

## 6 Concluding remarks

This paper has considered a generalised version of the Vickrey bottleneck model of congestion users having general concave utility defined over the duration in the queue as well as the time of exit from the queue. The queue may be subject to varying degrees of random sorting, ranging from strict FIFO queue priority to no queue priority. The no residual queue (NRQ) property holds when the queue has vanished at the time of the last arrival. Proposition 1 shows that the NRQ property is sufficient to derive a number of results that are useful for designing policies to regulate congestion. In particular, the interval of arrival as well as the expected utility of users are independent of the queueing regime, provided the NRQ property holds. The remainder of the paper then establishes that the acceptable lateness condition 2 , restricting the relation between the marginal utilities of duration and exit time, is sufficient for the NRQ property to hold in Nash equilibrium under all queueing regimes considered and that this condition cannot be relaxed in general.

For simplicity, we have only considered the case where total usage is constant. The extension to endogenous total demand is however straightforward.

The paper leaves open the characterisation of Nash equilibrium when the NRQ property does not hold. In that case, the convenient results of Proposition 1 are not available. The paper also leaves open the question of what happens under random queue sorting when the acceptable lateness condition is not satisfied. It is possible that there are combinations of queueing regimes and strictly concave utility for which the NRQ property does hold.

We must acknowledge some further limitations of our analysis. A main simplification is that we assume homogenous users, whereas heterogeneity is likely in actual queueing situations. Lindsey (2004) presents an analysis of user heterogeneity for the bottleneck model with strict FIFO queue and scheduling utility which is separable in duration in the queue and time of exit from the queue. It may be possible to extend Lindsey's analysis to allow for random queue sorting. We leave this for the future.

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## A Proofs

## Proof of lemma 1.

Proof. All $N$ users can arrive and be served without queueing during an interval of length $N / \psi$, so $-\infty<-N / \psi \leq \tau_{0}, \tau_{1} \leq N / \psi<\infty$. There must be arrivals before the queue can start, so $t_{0} \leq \tau_{0}$. If $t_{0}<\tau_{0}$, some users can benefit from postponing arrival so $t_{0}=\tau_{0}$ in equilibrium. Similarly, $t_{1} \leq \tau_{1}$, since otherwise some users could benefit from arriving earlier. In equilibrium, there is always queue during $] \tau_{0}, \tau_{1}[$ since otherwise users could benefit from moving into the gap in the queue. The arrival rate is locally bounded so not all users can arrive at time 0 . The first arrival time occurs strictly before the preferred exit time 0 , since otherwise it would be possible to arrive at time 0 and be served immediately. Similarly, the last arrival time occurs strictly after time 0 .

## Proof of Proposition 1.

Proof. The NRQ property implies that $t_{1}=\tau_{1}$, which means that $Q\left(t_{1}\right)=0$. Hence the durations in the queue are zero at times $t_{0}$ and $t_{1}$ so that $u\left(0, t_{0}\right)=$ $u\left(0, t_{1}\right)$. By Lemma 1, the queue lasts from $t_{0}$ to $t_{1}$ such that $N=\psi\left(t_{1}-t_{0}\right)$. Consequently, $t_{0}$ and $t_{1}$ are unique due to concavity of $u(\cdot)$ and $t_{0}<0<t_{1}$. By the equilibrium condition, $E(u \mid a)=u\left(0, t_{0}\right)$ for all $a \in\left[t_{0}, t_{1}\right]$. Differentiating $N=\psi\left(t_{1}-t_{0}\right)$ leads to $1=\psi\left(\frac{\partial t_{1}}{\partial N}-\frac{\partial t_{0}}{\partial N}\right)$. Differentiating $u\left(0, t_{0}\right)=u\left(0, t_{1}\right)$ leads to $u_{2}\left(0, t_{0}\right) \frac{\partial t_{0}}{\partial N}=u_{2}\left(0, t_{1}\right) \frac{\partial t_{1}}{\partial N}$, so that

$$
\frac{\partial t_{0}}{\partial N}=\frac{1}{\psi} \frac{u_{2}\left(0, t_{1}\right)}{u_{2}\left(0, t_{0}\right)-u_{2}\left(0, t_{1}\right)}<0 .
$$

Then

$$
\frac{\partial u\left(0, t_{0}\right)}{\partial N}=\frac{1}{\psi} \frac{u_{2}\left(0, t_{0}\right) u_{2}\left(0, t_{1}\right)}{u_{2}\left(0, t_{0}\right)-u_{2}\left(0, t_{1}\right)}<0 .
$$

Straightforward computation establishes that when $u(\cdot)$ is concave, then the marginal utility decreases

$$
\frac{\partial^{2} u\left(0, t_{0}\right)}{\partial N^{2}}=\frac{1}{\psi^{2}} \frac{u_{2}\left(0, t_{0}\right)^{3} u_{22}\left(0, t_{1}\right)-u_{2}\left(0, t_{1}\right)^{3} u_{22}\left(0, t_{0}\right)}{\left(u_{2}\left(0, t_{0}\right)-u_{2}\left(0, t_{1}\right)\right)^{3}} \leq 0,
$$

with strict inequality when $u(\cdot)$ is strictly concave.
The following Lemma collects some relationships between the hazard rate and the corresponding conditional density and cumulative distribution function. We will use the results in the Lemma many times in the proofs below and will therefore omit references to the Lemma.

Lemma 3 Let the hazard rate $\lambda$ and the corresponding $f(t \mid a)$ and $F(t \mid a)$ be as defined above. Then the following relations hold.

$$
\begin{align*}
f(a \mid a) & =\lambda(a)  \tag{8}\\
\frac{\partial F(t \mid a)}{\partial a} & =-\frac{\lambda(a)}{\lambda(t)} f(t \mid a)  \tag{9}\\
\frac{\partial f(t \mid a)}{\partial a} & =\lambda(a) f(t \mid a) \tag{10}
\end{align*}
$$

Proof. The first assertion follows from (6), since $F(a \mid a)=0$. Differentiate (7) to find that

$$
\frac{\partial F(t \mid a)}{\partial a}=-\lambda(a) e^{-\int_{a}^{t} \lambda(s) d s}=-\lambda(a)(1-F(t \mid a)) .
$$

Then the second assertion follows by substitution from (6), while the third assertion follows by differentiation with respect to $t$.

## Proof of Lemma 2.

Proof. Evaluate first $1-F(t \mid a)$. Let $t_{1} \leq a \leq t \leq \tau_{1}$. Then by (7)

$$
1-F(t \mid a)=\exp \left(-\int_{a}^{t} \frac{\psi}{Q\left(t_{1}\right)-\psi\left(s-t_{1}\right)} d s\right)
$$

where we use that $Q(s)=Q\left(t_{1}\right)-\psi\left(s-t_{1}\right)$. Make the substitution $x=$ $Q\left(t_{1}\right) / \psi-\left(s-t_{1}\right)$ to find that

$$
\begin{aligned}
1-F(t \mid a) & =\exp \left(\int_{Q\left(t_{1}\right) / \psi-\left(a-t_{1}\right)}^{Q\left(t_{1}\right) / \psi-\left(t-t_{1}\right)} \frac{1}{x} d x\right) \\
& =\frac{Q\left(t_{1}\right) / \psi-\left(t-t_{1}\right)}{Q\left(t_{1}\right) / \psi-\left(a-t_{1}\right)}=\frac{\lambda(a)}{\lambda(t)} .
\end{aligned}
$$

Use (6) to see that $f(t \mid a)=\lambda(a)$. As the density of exit times conditional on $a$ is constant, the exit time is uniformly distributed. To verify the last statement of the Proposition, simply differentiate

$$
\frac{\partial \lambda(a)}{\partial a}=-\frac{\psi Q^{\prime}(a)}{Q^{2}(a)}=\frac{\psi^{2}}{Q^{2}(a)}=\lambda^{2}(a)
$$

## Proof of Proposition 4.

Proof. Assume a Nash equilibrium with a residual queue at time $t_{1}$ and consider $a>t_{1}$. The expected utility at time $a$, given by (2), is

$$
E(u \mid a)=\lambda(a) \int_{a}^{\tau_{1}} u(t-a, t) d t
$$

by Lemma 2. Using the last statement of Lemma 2, the derivative with respect to the arrival time $a$ is seen to be

$$
\begin{equation*}
\frac{1}{\lambda(a)} \frac{\partial E(u \mid a)}{\partial a}=E(u \mid a)-u(0, a)-\int_{a}^{\tau_{1}} u_{1}(t-a, t) d t . \tag{11}
\end{equation*}
$$

Considering the following identity

$$
u\left(\tau_{1}-a, \tau_{1}\right)-u(0, a)=\int_{a}^{\tau_{1}}\left[u_{1}(t-a, t)+u_{2}(t-a, t)\right] d t
$$

we may write

$$
\frac{1}{\lambda(a)} \frac{\partial E(u \mid a)}{\partial a}=E(u \mid a)-u\left(\tau_{1}-a, \tau_{1}\right)+\int_{a}^{\tau_{1}} u_{2}(t-a, t) d t .
$$

Add the two expressions for $\frac{\partial E(u \mid a)}{\partial a}$ to obtain

$$
\begin{aligned}
\frac{1}{\lambda(a)} \frac{\partial E(u \mid a)}{\partial a}= & {\left[E(u \mid a)-\frac{1}{2}\left(u(0, a)+u\left(\tau_{1}-a, \tau_{1}\right)\right)\right] } \\
+ & \frac{1}{2} \int_{a}^{\tau_{1}}\left[u_{2}(t-a, t)-u_{1}(t-a, t)\right] d t
\end{aligned}
$$

The first term on the RHS is positive by Jensen's inequality since $u(t-a, t)$ is concave as a function of $t$ and the second term is strictly positive by Condition 2. Thus, $E(u \mid a)$ is strictly increasing on $] t_{1}, \tau_{1}[$ so that

$$
\begin{equation*}
E\left(u \mid t_{1}\right)<E\left(u \mid \tau_{1}\right)=u\left(0, \tau_{1}\right), \tag{12}
\end{equation*}
$$

which contradicts Nash equilibrium.
To verify the second assertion of the Proposition, note that in the linear case,

$$
\begin{aligned}
\frac{1}{\lambda(a)} \frac{\partial E(u \mid a)}{\partial a} & =\frac{1}{2} \int_{a}^{\tau_{1}}\left[u_{2}(t-a, t)-u_{1}(t-a, t)\right] d t \\
& =\frac{1}{2}\left(\tau_{1}-a\right)(\alpha-\gamma) .
\end{aligned}
$$

Then $\frac{\partial E(u \mid a)}{\partial a}>0$ is equivalent to Condition 2 and so Condition 2 is also necessary.

## Proof of Proposition 5.

Proof. The expression for the expected utility conditional on arrival at time $a$ is (2). Using (10), we express the equilibrium condition for the no queue priority case as follows.

$$
\frac{\partial E(u \mid a)}{\partial a}=\lambda(a) E(u \mid a)-u(0, a) \lambda(a)-E\left(u_{1} \mid a\right)=0
$$

which can be solved using $\lambda(a)=\psi / Q(a)$ to yield

$$
\frac{Q(a)}{\psi}=\frac{E(u \mid a)-u(0, a)}{E\left(u_{1} \mid a\right)} .
$$

Differentiate again and use that (1) gives $Q^{\prime}(a)=\rho(a)-\psi$ to find

$$
\begin{equation*}
\frac{\rho(a)}{\psi}=1-\frac{u_{2}(0, a)}{E\left(u_{1} \mid a\right)}-\frac{\frac{\partial E\left(u_{1} \mid a\right)}{\partial a}}{\lambda(a) E\left(u_{1} \mid a\right)} . \tag{13}
\end{equation*}
$$

Multiply all terms in (13) by $-\lambda(a) E\left(u_{1} \mid a\right)>0$ to find that $\rho(a)>0$ iff

$$
\begin{equation*}
-\lambda(a) E\left(u_{1} \mid a\right)+\lambda(a) u_{2}(0, a)+\frac{\partial E\left(u_{1} \mid a\right)}{\partial a}>0 . \tag{14}
\end{equation*}
$$

Carry out the differentiation using Lemma 3 to find that

$$
\frac{\partial E\left(u_{1} \mid a\right)}{\partial a}=-\lambda(a) u_{1}(0, a)-E\left(u_{11} \mid a\right)+\lambda(a) E\left(u_{1} \mid a\right) .
$$

Insert this into the inequality (14) to find that it is equivalent to

$$
\begin{equation*}
\lambda(a)\left[u_{2}(0, a)-u_{1}(0, a)\right]-E\left(u_{11} \mid a\right)>0 . \tag{15}
\end{equation*}
$$

The second term is positive since $u$ is concave. Therefore Condition 2 implies that $\rho(a)>0$.

When utility is linear, (13) shows that the equilibrium arrival rate is

$$
\rho(a)=\left\{\begin{array}{c}
\psi \frac{\alpha+\beta}{\alpha}, a<0 \\
\psi \frac{\alpha-\gamma}{\alpha}, a>0 .
\end{array}\right.
$$

Then $\rho(a)>0$ implies Condition 2.
Proof of Proposition 6.

Proof. Assume that $Q\left(t_{1}\right)>0$. Then $E_{\tilde{F}}\left(u \mid t_{1}\right) \leq E_{F}\left(u \mid t_{1}\right)$, due to first-order stochastic dominance. But $E_{F}\left(u \mid t_{1}\right)<u\left(0, \tau_{1}\right)$ by (12) in the proof of Proposition 4. Then $E_{\tilde{F}}\left(u \mid t_{1}\right)<u\left(0, \tau_{1}\right)$ and the last user would prefer to arrive at $\tau_{1}$ rather than at $t_{1}$. This contradicts Nash equilibrium. Hence we must have $Q\left(t_{1}\right)=0$ in Nash equilibrium.


[^0]:    *This research is part of the SURPRICE project as well as of PREDIT-ADEME : TARIFICATION DES TRANSPORTS INDIVIDUELS ET COLLECTIFS A PARIS DYNAMIQUE DE L'ACCEPTABILITE and PREDIT: SCHEDULING, TRIP TIMING AND SCHEDULING PREFERENCES. We are grateful to Robin Lindsey, Katrine Hjorth, Hugo Harari-Kermadec, Søren Feodor Nielsen, Ken Small and seminar participants at the University of Copenhagen and at the Swedish Royal Institute of Technology for comments. Mogens Fosgerau is supported by the Danish Social Science Research Council. A special thanks is due to Richard Arnott, who gave as a number of very useful comments.
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[^1]:    ${ }^{1}$ It is also possible to conceive of queues with a queue manager. In this case, a last-in-first-out queue may be considered an opposite of a FIFO queue (Hassin, 1985).
    ${ }^{2}$ Arnott, de Palma and Lindsey (1996) and (Arnott, de Palma and Lindsey, 1999) analyze a situation in which capacity varies randomly from day to day, while the queue retains the FIFO property.

[^2]:    ${ }^{3}$ The operations research literature generally considers the arrival rate as exogenous, perhaps allowing the user to balk when he meets a long queue (Naor, 1969; Knudsen, 1972).

[^3]:    ${ }^{4} 1_{\{\cdot\}}$ is the indicator function for the event in curly brackets.

[^4]:    ${ }^{5} x^{+}=\max (x, 0)$, and $x=x^{+}-x^{-}$.
    ${ }^{6}$ The equilibrium concept is discussed by Arnott et al. (1993).

