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Online Appendix A: Multi-event modeling of breeding experience 2

The multievent framework (Pradel 2005) distinguishes the events, coded in the capture 3 histories, from the states which are not directly observed. When studying experience, there 4 are several states to consider while there are only 2 events.

- The events are: 6
 - 0, 'not observed',
- 1, 'observed'. 8
- The states retained are: 9
 - *NB*₀, non-breeder with no previous experience at the onset of the season,
 - B_0 , breeder with no previous experience at the onset of the season,
 - NB_1 , non-breeder with one previous experience at the onset of the season,
 - B_1 , breeder with one previous experience at the onset of the season,
 - NB_2 , non-breeder with >1 previous experience at the onset of the season,
 - B_2 , breeder with >1 previous experience at the onset of the season.
- A multievent model has three kinds of parameters : the initial state probabilities, the 16
- transition probabilities, and the event probabilities (conditional on the underlying state). As 17
- every individual in our data set was ringed as a chick, all initial state probabilities are trivially 18

1 for the state NB_0 and this kind of parameters is not used here. 19

The transitions are best presented in stochastic matrix form with departure states in rows and arrival states in columns (see also Figure 1 in main text). ϕ is the survival probability, β the breeding probability. β is indexed according to the level of previous experience: 0 no previous experience, 1 one previous experience, and 2 two or more previous experiences.

	Ι	VB_0	B_0	NB_1	B_1	NB_2	B_2	dead
NB_0	(<i>φ</i>	$(1 - \beta_0)$	$\phi \beta_0$	0	0	0	0	$1-\phi$
B_0		0	0	$\phi(1-\beta_1)$	$\phi \beta_1$	0	0	$1-\phi$
NB_1		0	0	$\phi(1-\beta_1)$	$\phi \beta_1$	0	0	$1-\phi$
B_1		0	0	0	0	$\phi(1-\beta_2)$	$\phi\beta_2$	$1-\phi$
NB_2		0	0	0	0	$\phi(1-\beta_2)$	$\phi\beta_2$	$1-\phi$
B_2		0	0	0	0	$\phi(1-\beta_2)$	$\phi\beta_2$	$1-\phi$
dead		0	0	0	0	0	0	1)

²⁰ Note that most transitions are impossible, that the last state 'dead' can be reached from any

 $_{\mbox{\tiny 21}}$ $\,$ live state with probability $1-\phi$ and that once in this state an individual will remain in it

²² forever (transition from 'dead' to 'dead' is 1 : last row, last column). This basic model can be

²³ particularized, for instance to allow for a first-year survival lower than survival in later years

²⁴ or to introduce time-dependency on breeding probability.

The event probabilities relate the observations coded in the capture histories to the underlying states. Here, we assume that non-breeders, except chicks at the time of ringing (state NB_0), cannot be observed (event '0' with probability 1) and that breeders, whatever their experience, share the same probability p to be observed. A matrix form with the conditioning state in row and the event in column is again a particularly handy presentation.

$$\begin{array}{ccc} & & & & & & \\ & & & & & \\ NB_0 & & & & \\ B_0 & & & & \\ NB_1 & & & & \\ NB_1 & & & & \\ B_1 & & & & \\ NB_2 & & & & \\ B_2 & & & & \\ B_2 & & & & \\ dead & & & & \\ 1 & 0 \end{array} \right)$$

²⁵ Online Appendix B: Implementation of the generic model ²⁶ in E-SURGE

We present here how to implement the above model using program E-SURGE. The 27 specifics of the model have been kept as general as possible in order to be applicable to 28 data sets other than the one we used in the article. For instance, we have not included the 29 effect of one very cold year on survival. However, some details, which should be easy to 30 adapt, relate to the flamingo example. For instance, no breeding is possible until age 3. We 31 intend this as a template for other studies. 32 E-SURGE can be downloaded at http://www.cefe.cnrs.fr/biom/logiciels.htm. It can read 33 in data in a no-frill MARK format: no comments, no tabulations, no letters in the capture 34

histories (only digits); or alternatively in BIOMECO format (see the associated manual for a
 description).

After opening a new session and reading in the data (2 first items of the bar menu), the number of states, 7, has to be specified through the 'Modify' button.

It is now possible to go through the 4 steps represented by the coloured buttons at the
 bottom left.

B.1 specifying the patterns (tool GEPAT)

This stage roughs out the model by specifying the probabilities that will not be used (impossible initial states, transitions, or events)– code '-' (minus sign) –and those that will be calculated from others (because they are the last of a set of exclusive complementary options)– code '*'.

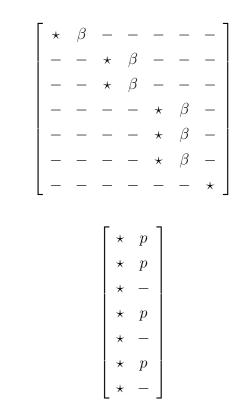
Initial States:

 $\left[\ \star \ - \ - \ - \ - \ - \
ight]$

⁴⁶ We decompose the transition in two steps: first survival, then the decision to breed. Survival Step:

-						_
ϕ	—	—	—	—	—	*
—	ϕ	—	_	—	—	*
—	—	ϕ	—	—	_	*
—	—	_	ϕ	—	_	*
—	—	_	—	ϕ	_	*
—	—	—	—	—	ϕ	* * * * * *
	—	_	—	—	_	*

Breeding Step:



Event probabilities:

- $_{
 m 47}$ The p on the first row (state NB_0) is needed because chicks are observed. We will clarify
- this point in the next step, GEMACO. (Note that GEPAT only displays greek letters; this is
- 49 not a problem.)

⁵⁰ B.2 specifying the effects (tool GEMACO, Choquet, 2008)

⁵¹ Initial States: there is nothing to do here Survival step:

- ⁵² This assumes that survival is constant over age past the first year of life. The last number
- ⁵³ (100 here) must exceed the number of time steps. Breeding step:

$$a(3_25).to(2) + a(4_25).to(4) + a(5_25).to(6) + others$$

- The breeding probabilities of inexperienced individuals (to(2)) are estimated from age 3 to
- ⁵⁵ age 25 (maximum age observable in the flamingo data set), those for individuals with one
- $_{\rm 56}~$ previous breeding experience (to(4)) from age 4 to 25, those for more experienced
- individuals (to(6)) from age 5 to 25. No individual breeds prior to age 3. The unspecified
- ⁵⁸ breeding probabilities, which should all be 0, are gathered together as a last breeding
- ⁵⁹ probability by the keyword '*others*'.

Event probabilities:

$$firste + nexte.[f(246)].t + others$$

60 Here, we distinguish the initial encounter probability (firste) from the subsequent

 $_{\rm 61}$ probabilities (*nexte*) of encountering breeders ($f(2\,4\,6)$) which are allowed to vary over years

62 (t). The remainder are the subsequent probabilities of encountering non-breeders (*others*).

B.3 setting some parameter values (tool IVFV)

⁶⁴ The last transition corresponds (among others) to the breeding probability prior to age 3

⁶⁵ and must be set to 0 (by entering the value 0 on the right and checking the box next to it).

 $_{66}$ The first event is the probability of encounter at the time of marking. It must be set to 1. The

⁶⁷ last event is the probability of encountering adult non-breeders. It has to be set to 0.

68 **B.4 RUN**

⁶⁹ The model can now be run.

70 Online Appendix C: Results of data analysis

71 C.1 Complete data set: years 1977-2001

72 C.1.1 Assessment of goodness-of-fit

There is no goodness-of-fit test available for the breeding probability models. We

examined the fit of the fully time- and age-dependent survival and capture model (Pollock,

- ⁷⁵ 1981). This test is obtained by running the test of the Cormack-Jolly-Seber model
- ⁷⁶ separately on each cohort and summing over cohorts (Table C1). Calculations were made
- vith program U-CARE (Choquet et al., 2009). This led to a variance inflation factor \hat{c} of
- ⁷⁸ 2.244 (1355.43/604) for the Pollock model. This value is conservative when applied to a
- ⁷⁹ model that performs better than the Pollock model. We used it for our analysis after
- ⁸⁰ introducing the Pollock model for comparison among the models fitted.

Table C1: Goodness-of-fit tests of the time-dependent survival and capture Cormack-Jolly-Seber model by cohort.

	2	
Cohort	χ^2	degrees of freedom
1977	165.95	64
1978	199.48	65
1979	169.96	67
1980	134.80	55
1981	174.96	60
1982	122.35	52
1983	78.69	48
1984	58.65	39
1985	80.05	39
1986	42.09	28
1987	40.08	23
1988	36.52	21
1989	14.53	12
1990	8.94	9
1991	7.51	9
1992	14.43	7
1993	6.43	5
1994	0	1
1995-1998	0	0
Total	1355.43	604
Total	1355.43	604

81 C.1.2 Additional models

We checked with two models that we could not rid ourselves of the cold spell effect on survival nor of the temporal variability of capture rates. Table C2: Some additional models in the analysis of the effects of age and breeding experience on the breeding probability β of greater flamingos marked as chicks from 1977 to 1997 and resighted as breeders in the Camargue, southern France, until 2001. Notations are as in Table 2 of the main document except for a^2 which means two age-classes, the first one corresponding to the first year of life. CS stands for 'cold spell'. In the models with a cold spell effect, first-year survival in normal years is fixed to 0.763, the sex-averaged value of Balkız (2006) estimates.

Model	Best or Reference model or	Deviance	k	Δ QAIC
	 hypothesis tested 			(ĉ=2.244)
$\beta_{a.e}, \phi_{CS.a2}, p_t$	Best model	53763.22	87	0
$\beta_{a.e}, \phi_{a2}, p_t$	 no effect of cold spell on survival 	53803.83	85	14.60
$eta_{a.e}$, $\phi_{CS.a2}$, p	 capture rate constant 	55295.91	66	641.50
$\phi_{a.t} p_{a.t}$	Reference Pollock model	53883.71	439	757.69

C.2 Detailed data set: years 1985-2001

C.2.1 Goodness-of-fit

⁸⁶ Same procedure as for the complete data set. Variance inflation factor \hat{c} estimated at

⁸⁷ 1.625 (247.07/152) for the Pollock model.

Table C3: Goodness-of-fit tests of the time-dependent survival and capture Cormack-Jolly-Seber model by cohort.

Cohort	χ^2	degrees of freedom
1985	78.04	38
1986	41.59	28
1987	38.74	22
1988	36.52	21
1989	14.53	12
1990	9.41	9
1991	7.51	9
1992	14.43	7
1993	6.30	5
1994	0	1
1995-1998	0	0
Total	247.07	152

88 C.2.2 estimates from best model

⁸⁹ With the reduced data set, more information was available about individuals. Also,

⁹⁰ during the restricted period, no major climatic events took place so that survival could be

considered constant within each class of individuals. We took advantage of this new context 91 to examine the influence of sex and of individual differences on breeding probabilities and to 92 refine the modelisation of survival. The age and sex survival structure was matched exactly 93 to that of Balkız (2006), who has found sex differences and a second-year survival still 94 somewhat lower than the adult survival. The cost of first reproduction in females (notation 95 1rf), first detected by Tavecchia et al. (2001), was introduced as a smoothed function of age 96 (logit(age) is a linear function of log(age) that cancels at 6, see Table C4 below). Individual 97 differences were introduced through the use of the residuals of the log-log regression of 98 weight on tarsus length at ringing as a covariate of breeding probability. The best model 99 retained is $\phi_{a3.s+1rf} \ \beta_{a_2.e+BC.s} \ p_{t+s}$ i.e. sex- and age-specific survival (with 3 age classes) 100 with an age-specific cost of first reproduction in females, capture rates varying over time in 101 parallel (on a logit scale) in the two sexes, and breeding probabilities age- and 102 experience-dependent and affected by the body condition differently in females and males. 103 A sex-specific effect of body condition on survival was also tested but not retained (see next 104 section). 105

Table C4: Survival estimates.

		95%	6 C.I.	
parameter	estimate	lower boundary	higher boundary	s.e.
female 1st year	0.779	fixed pa	arameter	-
male 1st year	0.747	fixed pa	arameter	-
female 2nd year	0.944	fixed pa	arameter	-
male 2nd year	0.921	fixed pa	arameter	-
female 3+ year	0.970	0.951	0.981	0.007
male 2+ year	0.957	0.943	0.968	0.006
first breeding 3-year-old female	0.651	0.492	0.782	0.076
first breeding 4-year-old female	0.858	0.774	0.915	0.035
first breeding 5-year-old female	0.938	0.900	0.962	0.015

Figure C1: Breeding probability as a function of age and experience for greater flamingos breeding in the Camargue, south of France, estimated from the reduced data set (1985-2001). Plain circle: no previous breeding episode; empty circle: one previous breeding episode; triangle: 2 or more previous breeding episodes. These values are for an average individual in terms of body condition: value 0 of the covariate.

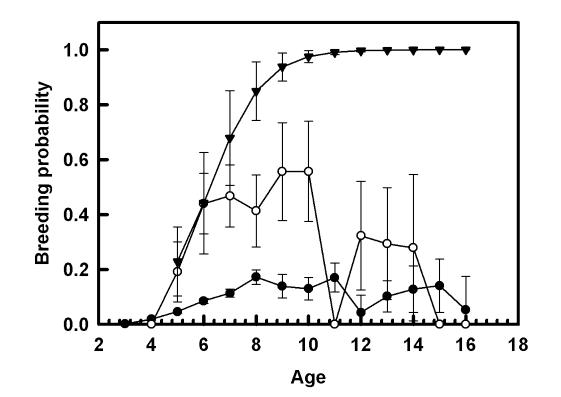
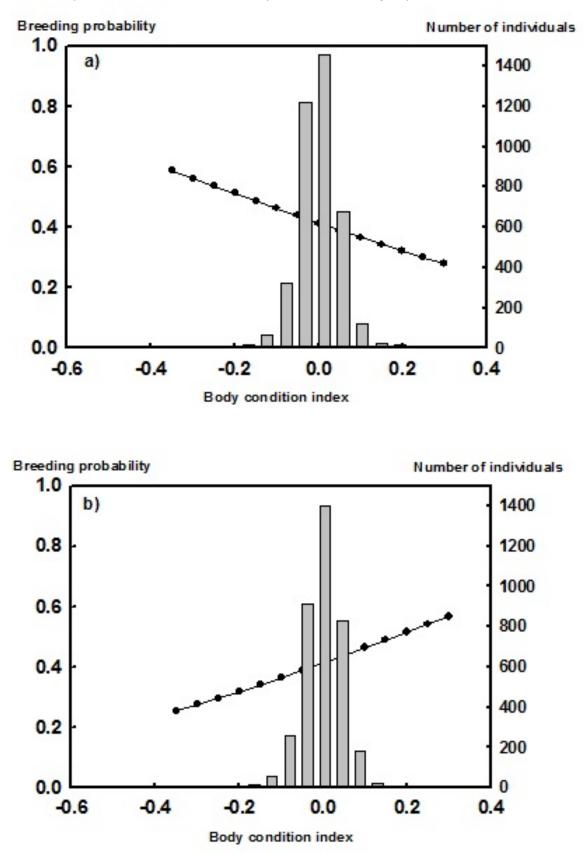


Figure C2: Influence of body condition on breeding probability: a) females, b) males. Positive values of the index correspond to larger than average individuals. The histogram shows the actual distribution of body condition in the population for each sex. The values of the curves are for 8-year-old individuals with one previous breeding experience.



			-		
# previous		point	95%	6 C.I.	
breeding episodes	age	estimate	lower boundary	higher boundary	s.e.
0	3	0.002	0.001	0.005	0.001
0	4	0.020	0.015	0.027	0.003
0	5	0.046	0.037	0.058	0.005
0	6	0.085	0.069	0.105	0.009
0	7	0.114	0.088	0.145	0.014
0	8	0.173	0.128	0.229	0.026
0	9	0.139	0.074	0.246	0.043
0	10	0.130	0.069	0.232	0.041
0	11	0.171	0.091	0.297	0.052
0	12	0.044	0.002	0.469	0.063
0	13	0.102	0.033	0.279	0.057
0	14	0.128	0.032	0.394	0.085
0	15	0.141	0.033	0.442	0.097
0	16	0.053	0.000	0.872	0.123
1	4	0.000	0.000	0.000	0.000
1	5	0.191	0.056	0.486	0.109
1	6	0.440	0.246	0.654	0.110
1	7	0.468	0.265	0.683	0.113
1	8	0.413	0.196	0.671	0.131
1	9	0.556	0.234	0.837	0.178
1	10	0.557	0.228	0.843	0.182
1	11	0.000	0.000	0.000	0.000
1	12	0.323	0.075	0.737	0.197
1	13	0.294	0.057	0.741	0.204
1	14	0.280	0.028	0.838	0.266
1	15	0.000	0.000	0.000	0.000
1	16	0.000	0.000	0.000	0.000
2+	5	0.229	0.069	0.543	0.125
2+	6	0.442	0.154	0.775	0.185
2+	7	0.679	0.310	0.909	0.172
2+	8	0.849	0.525	0.966	0.106
2+	9	0.938	0.731	0.988	0.051
2+	10	0.976	0.870	0.996	0.022
2+	11	0.991	0.943	0.999	0.009
2+	12	0.997	0.976	1.000	0.003
2+	13	0.999	0.990	1.000	0.001
2+	14	1.000	0.996	1.000	0.001
2+	15	1.000	0.998	1.000	0.000
2+	16	1.000	0.999	1.000	0.000

Table C5: Breeding estimates.

	95% C.I.						
year	estimate	lower boundary	higher boundary	s.e.			
		females					
1988	0.000	-	-	-			
1989	0.000	-	-	-			
1990	0.496	0.247	0.747	0.140			
1991	1.000	-	-	-			
1992	0.405	0.296	0.525	0.059			
1993	0.254	0.191	0.329	0.035			
1994	0.314	0.252	0.384	0.034			
1995	0.374	0.314	0.438	0.032			
1996	0.241	0.200	0.287	0.022			
1997	0.305	0.260	0.354	0.024			
1998	0.418	0.367	0.471	0.027			
1999	0.179	0.152	0.209	0.015			
2000	0.373	0.331	0.417	0.022			
2001	0.331	0.293	0.371	0.020			
males							
1988	0.000	-	-	-			
1989	0.000	-	-	-			
1990	0.569	0.304	0.800	0.138			
1991	1.000	-	-	-			
1992	0.477	0.360	0.597	0.062			
1993	0.313	0.239	0.398	0.041			
1994	0.381	0.310	0.457	0.038			
1995	0.445	0.379	0.512	0.034			
1996	0.298	0.250	0.352	0.026			
1997	0.370	0.318	0.425	0.027			
1998	0.490	0.434	0.547	0.029			
1999	0.226	0.193	0.263	0.018			
2000	0.443	0.397	0.491	0.024			
2001	0.399	0.355	0.443	0.023			

Table C6: Capture estimates. Confidence intervals are not available for the parameters estimated on the boundary. No individuals were at risk of capture prior to 1988.

106 C.2.3 effect of body condition on survival

To test for individual differences in survivorship after accounting for the cost of the first reproduction in females, body condition was introduced as a covariate of survival for males and females separately in the best model $\phi_{a3.s+1rf}$ $\beta_{a_2.e+BC.s}$ p_{t+s} (see above section). The effect was not significant. Table C7: Test of an effect of body condition on survival assessed from the best model $\phi_{a3.s+1rf} \beta_{a_2.e+BC.s} p_{t+s}$. Only the survival part is shown below as the remainder is unchanged.

Model	Best model or	Deviance	k	Δ QAIC
	- difference from best model			(ĉ=1.625)
$\phi_{a3.s+1rf}$	Best model	17575.27	49	0
$\phi_{a3.s+1rf+BC.s}$	- effect of body condition	17574.79	51	3.70
	on survival assessed sepa-			
	rately on males and females			

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