

Adaptive Bayesian Estimation of a spectral density Judith Rousseau, Willem Kruijer

▶ To cite this version:

Judith Rousseau, Willem Kruijer. Adaptive Bayesian Estimation of a spectral density. 2011. haloo641485

HAL Id: hal-00641485 https://hal.science/hal-00641485

Preprint submitted on 16 Nov 2011

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers. L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

Adaptive Bayesian Estimation of a spectral density

Judith Rousseau^{a,b}, Kruijer Willem^c

^aCEREMADE Université Paris Dauphine, Place du Maréchal de Lattre de Tassigny 75016 Paris, FRANCE. ^bENSAE-CREST, 3 avenue Pierre Larousse, 92245 Malakoff Cedex, FRANCE . ^cWageningen University Droevendaalsesteeg 1 Biometris, Building 107 6708 PB Wageningen The Netherlands

Abstract

Rousseau et al. [8] recently studied the asymptotic behavior of Bayesian estimators in the FEXP-model for spectral densities of Gaussian time-series. For the L_2 -norm on the log-spectral densities, they proved that the convergence rate is at least $n^{-\frac{\beta}{2\beta+1}}(\log n)^{\frac{2\beta+2}{2\beta+1}}$, $\beta > \frac{1}{2}$ being the Sobolev-regularity of the true spectral density f_o . We will improve upon the logarithmic factor, and prove that given a prior only depending on $\beta_s > \frac{1}{2}$, we have adaptivity to any $\beta \geq \beta_s$.

Keywords: Bayesian non-parametric, rates of convergence, adaptive estimation, long-memory time-series, FEXP-model

1. Introduction

Let $X_t, t \in \mathbb{Z}$, be a stationary zero mean Gaussian time series with spectral density $f_o(\lambda), \lambda \in [-\pi, \pi]$ in the form

$$f_o(\lambda) = |1 - e^{i\lambda}|^{-2d_o} \exp\left\{\sum_{j=0}^{\infty} \theta_{o,j} \cos(j\lambda)\right\}, \quad \theta_o \in \Theta(\beta, L_o)$$
(1.1)

where $d_o \in (-\frac{1}{2}, \frac{1}{2})$, $\Theta(\beta, L_o) = \{\theta \in l_2(\mathbb{N}) : \sum_{j \ge 0} \theta_j^2 (1+j)^{2\beta} \le L_o\}$ is a Sobolev ball. The parameter d_o is called the long-memory parameter; we

Preprint submitted to Statistics and Probability Letters

October 10, 2011

will refer to $\exp\{\sum_{j=0}^{\infty} \theta_{o,j} \cos(j\lambda)\}\$ as the short-memory part of the spectral density. The parameter β controls the regularity of the short-memory part. It is then natural to use the fractionally exponential or FEXP-model (see Beran [2] and Moulines and Soulier [6] and references therein) $\mathcal{F} = \bigcup_{k\geq 0} \mathcal{F}_k$, where

$$\mathcal{F}_k = \left\{ f_{d,k,\theta}(\lambda) = |1 - e^{i\lambda}|^{-2d} \exp\left\{ \sum_{j=0}^k \theta_j \cos(j\lambda) \right\}, \ d \in \left(-\frac{1}{2}, \frac{1}{2} \right), \theta \in \mathbb{R}^{k+1} \right\}.$$

We study Bayesian estimation of f_o within this FEXP-model. Let $\pi(d, k, \theta)$ denote the prior on (d, k, θ) ; this induces a prior on \mathcal{F} which we also denote π . Let $T_n(f)$ denote the covariance matrix of the observations $X = (X_1, \ldots, X_n)$, and let l_n be the associated log-likelihood

$$l_n(d,k,\theta) = -\frac{k+1}{2}\log(2\pi) - \frac{1}{2}\log|T_n(f)| - \frac{1}{2}X'T_n^{-1}(f)X$$
(1.2)

Bayesian estimates of the spectral density f_o are based on the posterior

$$\pi(f \in A|X) = \frac{\int_A e^{l_n(d,k,\theta)} d\pi(f)}{\int_{\mathcal{F}} e^{l_n(d,k,\theta)} d\pi(f)}, \qquad A \subset \mathcal{F}.$$
(1.3)

For example the posterior mean or median could be taken as 'point'-estimators of f_o . In this work however we focus on the posterior itself, and study the rate of convergence at which the posterior concentrates at f_o . More precisely, we lower-bound the posterior mass on the sets

$$B(\epsilon_n) = \{ f \in \mathcal{F} : l(f, f_o) \le \epsilon_n^2 \},\$$

where ϵ_n is a sequence tending to zero and

$$l(f, f_o) = \frac{1}{2\pi} \int_{-\pi}^{\pi} (\log f_o(\lambda) - \log f(\lambda))^2 d\lambda$$

Whether $\pi(B(\epsilon_n)|X)$ tends to one for a certain sequence ϵ_n critically depends on the smoothness of f_o as well as the smoothness induced by the prior. In Theorem 4.2 of Rousseau et al. [8] (RCL hereafter) it is shown that when $\theta_o \in \Theta(\beta, L_o)$ and the prior on θ has support contained in a Sobolev ball $\Theta(\beta, L)$ with L large enough, then the rate is $\epsilon_0(L)n^{-\frac{2\beta}{2\beta+1}}(\log)^{\frac{4\beta+4}{2\beta+1}}$, for fixed $\beta > \frac{1}{2}$ and $\epsilon_0(L)$ large enough depending on L. In the present work we prove that such priors in fact lead to an adaptive concentration rate (in β) and we improve upon the constant $\epsilon_0(L)$ and the logarithmic factor. Adaptivity is of great interest since it is difficult to know the smoothness of the function f a priori. Improving on the constant ϵ_0 is crucial in Kruijer and Rousseau [4] but has also interest in its own. Indeed in Theorem 2.1 we prove that ϵ_0 depends only on L_o the radius of the Sobolev ball containing θ_o . In RCL however ϵ_0 depends on L, with the risk that would L be very large ϵ_0 might also be very large. Here we prove that this is not the case and that we can choose L as large as the application requires. This suggests that the result might actually hold without the constraint L in the prior on θ , but we have not been able to prove that.

Notation:

The *m*-dimensional identity matrix is denoted I_m . For matrices A we write |A| for the Frobenius or Hilbert-Schmidt norm $|A| = \sqrt{\text{tr}AA^t}$, where A^t denotes the transpose of A. The operator or spectral norm is denoted $||A||^2 = \sup_{||x||=1} x'Ax$. We also use $||\cdot||$ for the Euclidean norm of finite dimensional vectors or sequences in $l^2(\mathbb{N})$, and for the L_2 -norm of functions. If $u \in l^1(\mathbb{N})$ we denote $||u||_1 = \sum_j |u_j|$. Given a sequence $\{u_j\}_{j\geq 0}$ and a nonnegative integer m, we write $u_{[m]}$ for the vector (u_0, \ldots, u_m) and $||u||_{>m}$ for the l^2 -norm of the sequence u_{m+1}, u_{m+2}, \ldots . When we write $\sum_{j\geq 0} (\theta_j - \theta_{o,j})^2$ or $\sum_{j\geq 0} |\theta_j - \theta_{o,j}|$ for a finite-dimensional vector θ and $\theta_o \in l_2(\mathbb{N})$, θ_j is understood to be zero when j > k. For any function $h \in L_1([-\pi, \pi])$, $T_n(h)$ is the matrix with entries $\int_{-\pi}^{\pi} e^{i|l-m|\lambda}h(\lambda)d\lambda$, $l, m = 1, \ldots, n$. For example, $T_n(f)$ is the covariance matrix of observations $X = (X_1, \ldots, X_n)$ from a time series with spectral density f. Let P_o denote the law associated with the true spectral density f_o and E_o expectations with respect to P_o .

2. Main results

Let $\beta_s > \frac{1}{2}$ be a fixed constant. We consider the following family of priors on (d, k, θ) . d is a priori independent of (k, θ) with density π_d with respect to Lebesgue measure. For some positive t < 1/2, the support of π_d is included in [-1/2 + t, 1/2 - t]. We consider two cases for the prior on k:

Deterministic sieve $\pi_k(k) = \delta_{k_{A,n}}(k)$, i.e. it is the Dirac mass at $k_{A,n} = |A(n/\log n)^{1/(2\beta_s+1)}|$, for some positive A.

Random sieve the support of π_k is \mathbb{N} and satisfies:

$$e^{-c_1k\log k} \le \pi_k(k) \le e^{-c_2k\log k},$$

for some positive c_1, c_2 and k large enough. $\pi_{\theta|k}$, the prior on θ given k, has a density with respect to the Lebesgue measure on \mathbb{R}^k . This density is also denoted $\pi_{\theta|k}$, and is such that, for some constants L > 0 and $\beta_s > 1/2$, $\pi_{\theta|k}$ is positive on $\Theta_k(\beta, L)$ and $\pi_{\theta|k} [\Theta_k(\beta_s, L)^c] = 0$. These priors have been considered in particular in RCL, in Holan et al. [3] and in Kruijer and Rousseau [4]. We now state the main result.

Theorem 2.1. Suppose we observe $X = (X_1, \ldots, X_n)$ from a stationary, zero mean Gaussian time-series whose spectral density f_o is as in (1.1), with $d_o \in [-\frac{1}{2} + t, \frac{1}{2} - t], \ \theta_o \in \Theta(\beta, L_o)$ and $\beta \ge \beta_s > \frac{1}{2}$. Consider a prior $\pi = \pi_d \pi_k \pi_{\theta|k}$ as described above such that there exists $c_0 > 0$ for which

$$\liminf_{n \to \infty} \min_{k \in \mathcal{K}_n} \inf_{\theta \in \Theta_k(\beta, L_o)} e^{c_0 k \log k} \pi_{\theta|k}(\theta) > 1.$$
(2.1)

where, for some B > 0 and $k_{B,n} = \lfloor B(n/\log n)^{1/(2\beta_s+1)} \rfloor$, $\mathcal{K}_n = \{0, \ldots, k_{B,n}\}$ in the case of the random sieve prior, and $\mathcal{K}_n = \{k_{A,n}\}$ in the case of the deterministic prior. Assume also that L is large enough.

• In the case of the random sieve prior, for any $\beta_2 > \beta_s$, we have the following uniform result:

$$\sup_{f_o \in \cup_{\beta_s \le \beta_2} \Theta(\beta, L_o)} E_o \pi((d, k, \theta) : l(f_{d,k,\theta}, f_o) \ge l_0^2 \epsilon_n^2(\beta) | X) \le n^{-3}, \quad (2.2)$$

where $\epsilon_n(\beta) = (n/\log n)^{-\frac{\beta}{2\beta+1}}$ and l_0 only depends on L_o . In particular, it is independent of L.

• In the case of the deterministic prior, for any $\beta_2 > \beta_s$, we have the following uniform result:

$$\sup_{f_o \in \bigcup_{\beta_s \le \beta \le \beta_2} \Theta(\beta, L_o)} E_o \pi((d, k, \theta) : l(f_{d,k,\theta}, f_o) \ge l_0^2 \epsilon_n^2(\beta_s) | X) \le n^{-3},$$
(2.3)

where l_0 only depends on L_o and is independent of L.

The constraint $\beta > 1/2$ is necessary to ensure that the short memory part $\exp(\sum_{j} \theta_{oj} \cos(jx))$ is bounded and continuous. As mentioned in the introduction, the fact that l_0 is independent of L is interesting since it allows us, in practice, to choose L arbitrarily high without penalizing the posterior concentration rate. It suggests that such results could hold with $L = \infty$, however we have no proof for it. The random sieve prior leads to an adaptive posterior concentration rate over the range $\beta \geq \beta_s$, since for all $\beta > 1/2$, $\epsilon_n(\beta)$ is the minimax (up to a log *n* term) rate over the class of FEXP spectral densities given by (1.1) and associated to $\theta \in \Theta(\beta, L_o)$. The deterministic sieve prior does not lead to an adaptive procedure since the posterior concentration rate is $\epsilon_n(\beta_s)$ in this case. Obtaining adaptation by putting a prior on the dimension of the model is a commonly used strategy in Bayesian non parametrics, see for instance Arbel [1] or Rivoirard and Rousseau [7].

3. Proof of Theorem 2.1

We first introduce some notions that are useful throughout the proof.

3.1. Notation and preliminary results

We first introduce various (pseudo)-distances. We denote the Kullback -Leibler divergence between the Gaussian distributions associated with spectral densities f_o and f by

$$KL_n(f_o; f) = \frac{1}{2n} \left\{ \operatorname{tr} \left[T_n(f_o) T_n^{-1}(f) - I_n \right] - \log \det(T_n(f_o) T_n^{-1}(f)) \right\},\$$

a symmetrized version of it by $h_n(f_o, f) = KL_n(f_o; f) + KL_n(f; f_o)$ and the variance of the log-likelihood ratio by

$$b_n(f_o, f) = \frac{1}{n} \operatorname{tr} \left\{ T_n^{-1}(f) (T_n(f_o - f) T_n^{-1}(f) T_n(f_o - f)) \right\}.$$

The limiting values of $b_n(f_o, f)$ and $h_n(f_o, f)$ are denoted

$$h(f_o, f) = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left[\frac{f_o(\lambda)}{f(\lambda)} + \frac{f(\lambda)}{f_o(\lambda)} - 2 \right] d\lambda, \quad b(f_o, f) = (2\pi)^{-1} \int_{-\pi}^{\pi} \left(\frac{f_o}{f}(\lambda) - 1 \right)^2 d\lambda$$

Then $h(f_o, f) \ge l(f_o, f)$ (RCL, p.6). Using Lemma 2 in RCL we find that for all $k \in \mathbb{N}$,

$$b_n(f_o, f_{d,k,\theta}) \leq \|T_n(f_o)^{1/2} T_n(f)^{-1/2}\|^2 h_n(f_o, f_{d,k,\theta}) \\ \leq C(\|\theta_o\|_1 + \|\theta\|_1) n^{2(d_o - d)_+} h_n(f_o, f),$$
(3.1)

where C is a universal constant. Similarly,

$$h_n(f_o, f) \le \|T_n^{-\frac{1}{2}}(f)T_n^{\frac{1}{2}}(f_o)\|^2 b_n(f_o, f).$$
(3.2)

In line with the notation of (1.2), let $\phi(x; d, k, \theta)$ denote the density of X, which is the Gaussian density with mean zero and covariance matrix $T_n(f_{d,k,\theta})$ and let $\phi(x; d_o, \theta_o)$ denote the Gaussian density associated with $T_n(f_o)$. We write $R_n(f_{d,k,\theta}) = \phi(X; d, k, \theta) / \phi(X; d_o, \theta_o)$ for the likelihood-ratio.

The proof of Theorem 2.1 contains two parts. First, it needs to be shown that the rate is $l_0^2 \epsilon_n^2$, for a constant l_0 that may depend on L and β_s . Then by re-insertion of the rate obtained in the first part, we improve upon the constant l_0 . In particular, it is shown to be independent of L for L large enough.

3.2. Proof of Theorem 2.1

Throughout the proof C denotes a universal constant. Let 0 < t < 1/2 and

$$\mathcal{G}_k(t,\beta_s,L) = \left\{ f_{d,k,\theta} : d \in \left[-\frac{1}{2} + t, \frac{1}{2} - t\right], \theta \in \Theta_k(\beta_s,L) \right\}, \quad \mathcal{G} = \bigcup_{k=0}^{\infty} \mathcal{G}_k(t,\beta_s,L)$$

By the results of RCL (Theorem 3.1, and Corollary 1 in the supplement) we have consistency for $h(f_o, f_{d,k,\theta})$ and $|d - d_o|$, i.e. for all $\delta, \epsilon > 0$, $\pi (f_{d,k,\theta} : h(f_{d,k,\theta}, f_o) < \epsilon^2, |d - d_o| < \delta |X)$ tends to one in probability. Hence it suffices to show that

$$\pi [W_n | X] = \frac{\int_{W_n} R_n(f) d\pi(f)}{\int R_n(f) d\pi(f)} := \frac{N_n}{D_n} \xrightarrow{P_o} 0,$$
(3.3)

where in the case of the random sieve prior,

$$W_n = \left\{ f_{d,k,\theta} \in \mathcal{G} : l(f_o, f_{d,k,\theta}) \ge l_0 \epsilon_n^2(\beta), h(f_o, f_{d,k,\theta}) \le \epsilon^2, \, |d - d_o| \le \delta \right\},$$

for a constant $l_0 > 0$ depending only on L_o , β_s and the prior on k. In the case of the deterministic sieve prior, we replace $\epsilon_n^2(\beta)$ in this definition by $\epsilon_n^2(\beta_s)$. We present the proof of (3.3) for the case of the random sieve prior; the proof for the case of the deterministic sieve prior can be deduced by replacing β by β_s . The proof consists of two parts: first we show that for some c > 0,

$$P_o\left[D_n < e^{-2nu_0\epsilon_n^2(\beta)}/2\right] \le e^{-cn\epsilon_n^2(\beta)},\tag{3.4}$$

for which we will establish a lower bound the prior mass on a Kullback-Leibler neighborhood of f_o . In the second part we show that under the event $D_n \geq \frac{1}{2}e^{-nu_0\epsilon_n^2(\beta)}$ we can control N_n/D_n . This will be done by giving a bound on the upper-bracketing entropy of the model.

For the proof of (3.4), note that RCL already found that if $\beta \geq \beta_s > 1/2$, there exists $u_0 \geq 0$ depending only on L_o such that

$$P_o\left[D_n < e^{-nu_0\epsilon_n^2(\beta)(\log n)^{1/(2\beta+1)}}/2\right] = o(n^{-1}).$$

To prove (3.4), we thus need to improve on the log *n* term in the preceding equation. Set

$$\bar{\mathcal{B}}_n = \{(d,k,\theta); KL_n(f_o, f_{d,k,\theta}) \le \frac{\epsilon_n^2(\beta)}{4}, b_n(f_o, f_{d,k,\theta}) \le \epsilon_n^2(\beta), d_o \le d \le d_o + \delta\},\$$

for some positive δ . Recall that

$$D_n = \sum_k \pi_k(k) \int e^{l_n(d,k,\theta) - l_n(f_o)} d\pi_{\theta|k}(\theta) d\pi_d(d),$$
$$P_o^n \left[D_n < e^{-2nu_0 \epsilon_n^2(\beta)} \right] \leq P_o^n \left[\int_{\bar{\mathcal{B}}_n} e^{l_n(f) - l_n(f_o)} d\pi(f) < e^{-2nu_0 \epsilon_n^2(\beta)} \right].$$

From the proof of Theorem 4.1 in RCL (section 5.2.1), it follows that

$$P_0^n \left(D_n \le e^{-nu_0 \epsilon_n^2(\beta)} \pi(\bar{\mathcal{B}}_n) / 2 \right) \le e^{-Cn \epsilon_n^2(\beta)}$$

for some constant C > 0 (independent of L). We now show that

$$\pi(\bar{\mathcal{B}}_n) \ge e^{-nu_0\epsilon_n^2(\beta)/4}.$$
(3.5)

Define

$$\tilde{\mathcal{B}}_n = \{ (d, k_{B,n}, \theta); d_o \le d \le d_o + \epsilon_n^2(\beta) n^{-a}, |\theta_j - \theta_{o,j}| \le (1+j)^{-\beta} \epsilon_n(\beta) n^{-a}, j = 0, \dots, k_{B,n} \}$$

We first prove that $\pi(\tilde{\mathcal{B}}_n) \geq e^{-nu_0\epsilon_n^2(\beta)/4}$ and then that $\tilde{\mathcal{B}}_n \subset \bar{\mathcal{B}}_n$. As in RCL (see the paragraph following equation (29) on p. 26), we find that

$$\sum_{j=1}^{\infty} (1+j)^{2\beta} \theta_j^2 \le 2L_o + \epsilon_n^2(\beta) n^{-2a} \le 3L_o, \forall \theta \in \tilde{\mathcal{B}}_n$$

for n large enough. Combined with condition (2.1) on $\pi_{\theta|k}$, this implies

$$\pi(\tilde{\mathcal{B}}_n) \ge \left(c\epsilon_n(\beta)k_{B,n}^{-\beta}n^{-a}\right)^{k+3} e^{-c_0k_{B,n}\log k_{B,n}} \ge e^{-c(\beta_s)k_0n\epsilon_n^2(\beta)}, \quad \forall \beta \ge \beta_s$$

This achieves the proof of (3.4) with $u_0 = c(\beta_s)k_0$.

To show that $\tilde{\mathcal{B}}_n$ is included in $\bar{\mathcal{B}}_n$, first note that equation (3.1) implies that it is enough to bound $h_n(f_o, f)$ on $\tilde{\mathcal{B}}_n$. To this end, we use the decomposition $f_o = f_{o,k_{B,n}} e^{\Delta_{d_o,k_{B,n}}}$, where $f_{o,k_{B,n}} = f_{d_o,k_{B,n},\theta_o}$ and

$$\Delta_{d_o,k_{B,n}}(\lambda) = \sum_{j=k_{B,n+1}}^{\infty} \theta_{o,j} \cos(jx), \quad \forall \lambda \in [-\pi,\pi].$$

Then we have the expansion

$$f_o = f_{o,k_{B,n}} (1 + \Delta_{d_o,k_{B,n}} + \Delta_{d_o,k_{B,n}}^2 / 2 + O(\Delta_{d_o,k_{B,n}}^3)), \quad |\Delta_{d_o,k_{B,n}}|_{\infty} = o(1)$$

and

$$h_n(f_o, f) \le 2[h_n(f_o, f_{o,k_{B,n}}) + h_n(f_{o,k_{B,n}}, f)].$$
 (3.6)

We first deal with the first term above. Let $b_{o,n} = e^{\Delta_{d_o,k_{B,n}}} - 1$ and without loss of generality we can assume that $b_{o,n}$ is positive in the expression of $h_n(f_o, f_{o,k_{B,n}})$ so that for all $\beta > 1/2$,

$$h_{n}(f_{o}, f_{o,k_{B,n}}) := \frac{1}{2n} tr \left[T_{n}^{-1}(f_{o})T_{n}(f_{o}b_{o,n})T_{n}^{-1}(f_{o})T_{n}(f_{o}b_{o,n}) \right] \\ \leq \frac{c}{n} tr \left[T_{n}^{-1}(g_{o})T_{n}(g_{o}b_{o,n})T_{n}^{-1}(g_{o})T_{n}(g_{o}b_{o,n}) \right] \\ = \frac{c}{n} tr [T_{n}^{2}(b_{o,n})] + c\gamma_{1} + c\gamma_{2}$$
(3.7)

where $g_o(\lambda) = |\lambda|^{-2d_o}$ and c depends only on $\sum_{j=0}^{\infty} |\theta_{o,j}| \leq L_o^{1/2} (2\beta - 1)^{-1/2}$, and

$$\gamma_{1} = \frac{1}{n} \left(tr \left[\left(T_{n} \left(\frac{1}{4\pi^{2} g_{o}} \right) T_{n} \left(g_{o} b_{o,n} \right) \right)^{2} \right] - tr \left[T_{n}^{2} \left(b_{o,n} \right) \right] \right)$$

$$\gamma_{2} = \frac{1}{n} \left(tr \left[\left(T_{n}^{-1} \left(g_{o} \right) T_{n} \left(g_{o} b_{o,n} \right) \right)^{2} \right] - tr \left[\left(T_{n} \left(\frac{1}{4\pi^{2} g_{o}} \right) T_{n} \left(g_{o} b_{o,n} \right) \right)^{2} \right] \right).$$

We first bound the first term of the right hand side of (3.7). Note that $b_{o,n}(\lambda) = \tilde{\Delta}_{d_o,k_{B,n},K_n} + R_0$, where

$$\tilde{\Delta}_{d_o,k_{B,n},K_n}(\lambda) = \sum_{j=k_{B,n}+1}^{K_n} \theta_{o,j} \cos(j\lambda), \quad K_n = \epsilon_n(\beta)^{-1/\beta} (\log n)^{1/\beta}$$

and $||R_0||^2 \le \epsilon_n^2(\beta)(\log n)^{-2}$ so that

$$\operatorname{tr}\left[T_{n}^{2}(b_{o,n})\right] = \operatorname{tr}\left[T_{n}^{2}(\tilde{\Delta}_{d_{o},k_{B,n},K_{n}})\right] + O(\log n[K_{n}^{-2\beta} + K_{n}^{-\beta}k_{B,n}^{-\beta}])$$

$$= \operatorname{tr}\left[T_{n}^{2}(\tilde{\Delta}_{d_{o},k_{B,n},K_{n}})\right] + O(\epsilon_{n}^{2}(\beta)), \qquad (3.8)$$

where the term $O(\log n[K_n^{-2\beta} + K_n^{-\beta}k_{B,n}^{-\beta}])$ comes from the fact that

$$\left|\operatorname{tr}\left[T_n^2(b_{o,n})\right] - \operatorname{tr}\left[T_n^2(\tilde{\Delta}_{d_o,k_{B,n},K_n})\right]\right| \le \operatorname{tr}\left[T_n^2(R_0)\right] + |T_n(R_0)||T_n(\tilde{b}_{o,n})|$$

and from the use of inequality (20) in Lemma 6 of RCL, with $f_1 = f_2 = 1$, $\delta = 0$ and b either equal to $\tilde{\Delta}_{d_o,k_{B,n},K_n}$ or R_0 . Note that the constant in the term $O(\epsilon_n^2(\beta))$ in (3.8) does not depend on L. Lemma 2.1 in Kruijer and Rousseau [5] together with the fact that

$$|\tilde{\Delta}_{d_o,k_{B,n},K_n}(\lambda) - \tilde{\Delta}_{d_o,k_{B,n},K_n}(y)| \le \sum_{j=k_{B,n}}^{K_n} j|\theta_{o,j}| \le C(\beta)L_0K_n^{-\beta+3/2} \lor k_{B,n}^{-\beta+3/2}$$

implies that for large enough n,

$$n^{-1} \left| \operatorname{tr} \left[T_n^2(\tilde{\Delta}_{d_o,k_{B,n},K_n}) \right] - 2\pi \operatorname{tr} \left[T_n(\tilde{\Delta}_{d_o,k_{B,n},K_n}^2) \right] \right| \\ \leq KC(\beta) L_o n^{-2\beta+1+\epsilon} \epsilon_n^2(\beta) = o(\epsilon_n^2(\beta)), \quad \forall \epsilon > 0,$$

uniformly over $\beta_s \leq \beta \leq \beta_2$ and $\theta_o \in \Theta(\beta, L_o)$. Consequently,

$$\frac{c}{n} tr[T_n^2(b_{o,n})] \le tr\left[T_n^2(\tilde{\Delta}_{d_o,k_{B,n},K_n})\right] + C(L_o,\beta)\epsilon_n^2(\beta),$$
(3.9)

for a constant $C(L_o, \beta)$ independent of L. Next we apply Lemma 2.4 in Kruijer and Rousseau [5] with $f = g_o$ and $b_1 = b_2 = b_{o,n}$; it then follows that

$$\gamma_1 \le \|b_{o,n}\|_{\infty}^2 n^{\delta-1} n^{\epsilon-1} = o(\epsilon_n^2(\beta)), \quad \forall \epsilon > 0.$$
 (3.10)

Finally Lemma 2.3 in Kruijer and Rousseau [5] implies that for all $\epsilon > 0$,

$$\gamma_2 \le \|b_{o,n}\|_{\infty}^2 n^{-1+\epsilon} = o(\epsilon_n^2(\beta)).$$
 (3.11)

Combining (3.9), (3.10) and (3.11), it follows that

$$h_n(f_{o,k_{B,n}}, f_o) = n^{-1} \operatorname{tr} \left[T_n(\tilde{\Delta}^2_{d_o,k_{B,n},K_n}) \right] + C(L_o,\beta)\epsilon_n^2(\beta) \le 2C'(L_o,\beta)\epsilon_n^2(\beta),$$

where also $C'(L_o, \beta)$ is independent of L. The last inequality follows from

$$n^{-1} \operatorname{tr} \left[T_n(\tilde{\Delta}_{d_o,k_{B,n},K_n}^2) \right] = \frac{1}{2\pi} \int_{-\pi}^{\pi} \tilde{\Delta}_{d_o,k_{B,n},K_n}^2(\lambda) d\lambda$$
$$= \sum_{j=k_{B,n}+1}^{K_n} \theta_{o,j}^2 \le C'(L_o,\beta)\epsilon_n^2(\beta).$$

We now bound the last term in (3.6), which we write as

$$h_n(f_{o,k_{B,n}},f) = \frac{1}{2n} \operatorname{tr} \left[T_n(f_{o,k_{B,n}})^{-1} T_n(fb) T_n(f)^{-1} T_n(fb) \right], \quad b = (f - f_{o,k_{B,n}}) / f.$$

Since $d \ge d_o$, $|b|_{\infty} < +\infty$ and applying Lemma 6 inequality (20) of RCL, we obtain if $d, \theta \in \tilde{\mathcal{B}}_n$ with a > 0,

$$h_n(f_{o,k_{B,n}}, f) \le C \log n \left(|b|_2^2 + |d - d_o| |b|_{\infty}^2 \right) \\\le C \log n \left(\sum_{j=1}^{k_{B,n}} (\theta_j - \theta_{o,j})^2 + n^{-a} \epsilon_n^2(\beta) \right) = o(\epsilon_n^2(\beta)),$$

which finally implies that $\pi(\bar{\mathcal{B}}_n) \geq \pi(\tilde{\mathcal{B}}_n)$ and that (3.4) is proved. We now find an upper bound on N_n . First write $\bar{W}_n = W_n \cap \mathcal{F}_n$ where $\mathcal{F}_n = \{f_{d,k,\theta}; k \leq k_{B_1,n}\}$ and $k_{B_1,n} = B_1(n/\log n)^{1/(2\beta+1)}$. Then, since the prior on k is Poisson,

$$\pi(\mathcal{F}_n^c) \le e^{-c_2 k_{B_1,n} \log k_{B_1,n}} \le e^{-2nu_0 \epsilon_n^2(\beta)}$$

if B_1 is large enough (depending on $L_o, c_2, \beta_s, \beta_2$) and W_n can be replaced by \overline{W}_n in the definition of N_n . Following the proof of RCL, we decompose $\overline{W}_n = \bigcup_{l=l_0}^{l_n} W_{n,l}$, where $l_0 \ge 2$, $l_n = \lceil \epsilon^2 / \epsilon_n^2(\beta) \rceil - 1$ and

$$W_{n,l} = \left\{ f_{d,k,\theta} \in \mathcal{G} : k \le k_{B_1,n}, h(f_{d,k,\theta}, f_o) \le \epsilon^2, |d - d_o| \le \delta, \\ \epsilon_n^2 l \le h_n(f_o, f_{d,k,\theta}) \le u_n(l+1) \right\}.$$

In addition let $N_{n,l} = \int_{W_{n,l}} R_n(f) d\pi(f)$; then $N_n = \sum_{l=l_0}^{l_n} N_{n,l}$, and we have

$$E_{o}\left[\frac{N_{n}}{D_{n}}\right] \leq P_{o}^{n}\left(D_{n} \leq e^{-nu_{0}\epsilon_{n}^{2}(\beta)}/2\right) + E_{o}\left[\sum_{l=l_{0}}^{l_{n}}\frac{N_{n,l}}{D_{n}}1_{\{D_{n}\geq e^{-nu_{0}\epsilon_{n}^{2}(\beta)}/2\}}\right].$$
(3.12)

We construct tests $\bar{\phi}_l$ $(l = l_0, \ldots, l_n)$ and write

$$E_{o}\left[\sum_{l=l_{0}}^{l_{n}}\frac{N_{n,l}}{D_{n}}1_{\{D_{n}\geq e^{-nu_{n}}/2\}}(\bar{\phi}_{l}+1-\bar{\phi}_{l})\right]$$

$$\leq \sum_{l=l_{0}}^{l_{n}}E_{o}\left(\bar{\phi}_{l}\right)+2e^{nu_{n}}\sum_{l=l_{0}}^{l_{n}}E_{o}\left[N_{n,l}(1-\bar{\phi}_{l})\right].$$
(3.13)

The tests are based on a collection of spectral densities $H_{n,l} = \bigcup_{k=0}^{k_{B_1,n}} H_{n,l,k} \subset W_{n,l}$ defined as follows. Let D_l be a grid over $\{d : |d - d_o| \leq \delta\}$ with spacing $l\epsilon_n^2(\beta)/(\log n)$. Let $T_{l,k}$ denote the centers of hypercubes of radius $\frac{l\epsilon_n^2(\beta)}{k}$, covering $\Theta_k(\beta_s, L)$. We define $H_{n,l,k}$ as the collection of spectral densities $f_{l,i} = (2e)^{l\epsilon_n^2(\beta)} f_{d_{l,i},k,\theta^{l,i}}$, with $d_{l,i} \in D_l$ and $\theta^{l,i} \in T_{l,k}$. With every $f_{l,i}$ we associate a test

$$\phi_{l,i} = \mathbb{1}_{\{X'(T_n^{-1}(f_o) - T_n^{-1}(f_{l,i}))X \ge \operatorname{tr}\{I_n - T_n(f_o)T_n^{-1}(f_{l,i}) + \frac{n}{4}h_n(f_o, f_{l,i})\}\}},$$
(3.14)

and set $\bar{\phi}_l = \max_i \phi_{l,i}$.

The set $H_{n,l}$ can be seen as a collection of upper-bracket spectral densities, since for each $f_{d,k,\theta} \in W_{n,l}$ there exists a $f_{l,i} \in H_{n,l,k}$ such that $f_{l,i} \geq f_{d,k,\theta}$, $0 \leq d_{l,i} - d \leq l\epsilon_n^2(\beta)/(\log n)$ and

$$0 \le (2e)^{l\epsilon_n^2(\beta)} \exp\left\{\sum_{j=0}^k \theta_j^{l,i} \cos(jx)\right\} - \exp\left\{\sum_{j=0}^k \theta_j \cos(jx)\right\}$$
$$\le \frac{l\epsilon_n^2(\beta)}{32} (2e)^{l\epsilon_n^2(\beta)} \exp\left\{\sum_{j=0}^k \theta_j^{l,i} \cos(jx)\right\}.$$
(3.15)

The cardinality of $\cup_{k=0}^{k_{B_1,n}} H_{n,l,k}$ is at most

$$\left(l^{-1}k_{B_{1,n}}\epsilon_{n}(\beta)^{-4}\right)^{k_{B_{1,n}}}\frac{\delta\log n}{l\epsilon_{n}^{2}(\beta)} \le \exp\{2k_{B_{1,n}}\log n\} = C_{n,l}, \qquad (3.16)$$

for all $l \geq 2$. To bound the right hand side of (3.13), we use (3.16) in combination with the following error bounds for each of the tests $\phi_{l,i}$. Let $f \in W_{n,l}$ and let $f_{l,i} \in H_{n,l}$ be such that (3.15) holds, $\phi_{l,i}$ being the associated test-function. Then, from equation (4.4) in RCL together with the bound (3.1) on $b_n(f_o, f_{l,i})/h_n(f_o, f_{l,i})$ which again depends on $\|\theta^{l,i}\|_1$ and L_0 , we obtain that for all $0 < \alpha < 1$, there exists constants $d_1, d_2 > 0$ depending on L_o and $\|\theta^{l,i}\|_1$ such that

$$E_o \phi_{l,i} \le e^{-d_1 n l^{\alpha} \epsilon_n^2(\beta)}, \quad E_f^n(1 - \phi_{l,i}) \le e^{-d_2 n l^{\alpha} \epsilon_n^2(\beta)}.$$
 (3.17)

Using (3.17) we obtain the following bound on the term $\sum_{l=l_0}^{l_n} E_o(\bar{\phi}_l)$ in (3.13):

$$\sum_{l=l_0}^{l_n} E_o\left(\bar{\phi}_l\right) \le \sum_{l=l_0}^{l_n} C_{n,l} e^{-d_1 n l^\alpha \epsilon_n^2(\beta)} \le e^{2k_{B_1,n} \log n} \sum_{l=l_0}^{l_n} e^{-d_1 n l^\alpha \epsilon_n^2(\beta)} \to 0,$$

as soon as $l_0 \ge \left(\frac{2B_1}{d_1 u_0}\right)^2$, choosing $\alpha = 1/2$. Using (3.17) the last term in (3.13) is

$$\sum_{l=l_0}^{l_n} E_o \left[N_{n,l} (1-\bar{\phi}_l) \right] = \sum_{l=l_0}^{l_n} \int_{W_{n,l}} E_f^n (1-\bar{\phi}_l) d\pi(f) \le e^{-d_2 n l_0^{1/2} \epsilon_n^2(\beta)} \le e^{-2n\epsilon_n^2(\beta)}$$

as soon as $d_2 l_0^{1/2} \ge 2$, i.e. $l_0 \ge 4d_2^{-2}$. Note that the two lower bounds on l_0 depends on L_o and on $\|\theta^{l,i}\|_1$. Finally choosing, $l_0 = \max\left(4d_2^{-2}, \left(\frac{2B_1}{d_1u_0}\right)^2, 2, u_0\right)$ we obtain

$$P^{\pi}\left[h_n(f, f_o) \le l_0 \epsilon_n^2(\beta) | X^n\right] = o(n^{-1}).$$

From that we deduce a concentration rate in terms of the l norm, following RCL's argument in Appendix C. Let l_0 be an arbitrary constant and assume that $h_n(f, f_o) \leq l_0 \epsilon_n^2(\beta)$ and $f = f_{d,k,\theta}$. Then inequality (C.3) of Lemma 6 of RCL implies that

$$\frac{1}{n} \operatorname{tr} \left[T_n(f_o^{-1}) T_n(f_o - f) T_n(f^{-1}) T_n(f_o - f) \right] \le C_1 l_0 \epsilon_n,$$

where C_1 depends only on $\|\theta\|_1$ and on $\|\theta_o\|_1$. This implies that

$$\frac{1}{n} \operatorname{tr} \left[T_n (f_o^{-1} (f_o - f)) T_n (f^{-1} (f_o - f)) \right] \le 2C_1 l_0 \epsilon_n^2(\beta)$$

since the difference between the two terms is of order $O(n^{-1+2a}) \forall a > 0$, which also implies that $h(f_o, f) \leq 3C_1 l_0 \epsilon_n^2(\beta)$, for the same reason. Since $l(f_o, f) \leq h(f_o, f)$, we finally obtain that

$$P^{\pi}\left[l(f, f_o) \le 3C_1 l_0 \epsilon_n^2(\beta) | X^n\right] = o(n^{-1}).$$

To terminate the proof of Theorem 2.1, it only remains to prove that l_0 depends only on L_o, β_s . This is done using a simple re-insertion argument. Recall also that $k \leq k_{B_1,n} = B_1 n \epsilon_n^2(\beta)$, where B_1 is independent of the radius L of the sobolev-ball $\Theta(\beta_s, L)$ defining the support of the prior. We start with the following observation. From Kruijer and Rousseau [4] (equation (3.5)) it follows that for fixed d and k, the minimizer of $l(f_o, f_{d,k,\theta})$ over \mathbb{R}^{k+1} is

$$\bar{\theta}_{d,k} := \operatorname{argmin}_{\theta \in \mathbb{R}^{k+1}} l(f_o, f_{d,k,\theta}) = \theta_{o[k]} + (d_o - d)\eta_{[k]},$$

where η is defined by $\eta_j = -2/j$ $(j \ge 1)$ and $\eta_0 = 0$. Assuming that $l(f_o, f_{d,k,\theta}) \le 3C_1 l_0 \epsilon_n^2(\beta)$ and $k \le k_{B_1,n}$ leads to $l(f_o, f_{d,k,\bar{\theta}_{d,k}}) \le 3C_1 l_0 \epsilon_n^2$ and $\|\theta - \bar{\theta}_{d,k}\|^2 = l(f_{d,k,\theta}, f_{d,k,\bar{\theta}_{d,k}}) \le 12C_1 l_0 \epsilon_n^2$. Therefore

$$\sum_{j=0}^{k} |\theta_j| \le \sum_{j=0}^{k_{B_1,n}} |\theta_j - (\bar{\theta}_{d,k_n})_j| + \sum_{j=0}^{k_{B_1,n}} |(\bar{\theta}_{d,k_n})_j| \le \sqrt{12C}\sqrt{l_0}\epsilon_n k_{B_1,n}^{1/2} + 2|d - d_o|\log n + \sum_{j=0}^{k_{B_1,n}} j^{\rho}|\theta_{o,j}| \le 2(2\beta_s - 1)^{-1/2}\sqrt{L_o}$$

when n is large enough, where the second inequality comes from Lemma 3.1 of Kruijer and Rousseau [4]. This achieves the proof of Theorem 2.1.

Acknowledgement

This work was supported by the 800-2007–2010 grant ANR-07-BLAN-0237-01 "SP Bayes".

References

- [1] Arbel, J. (2010). Bayesian optimal adaptive estimation using a sieve prior, submitted.
- [2] Beran, J. (1993). Fitting long-memory models by generalized linear regression. *Biometrika*, 80(4):817–822.
- [3] Holan, S., McElroy, T., and Chakraborty, S. (2009). A Bayesian approach to estimating the long memory parameter. *Bayesian anlysis*, 4(1):159–190.
- [4] Kruijer, W. and Rousseau, J. (2011a). Bayesian semi-parametric estimation of the long-memory parameter under fexp-priors.

- [5] Kruijer, W. and Rousseau, J. (2011b). Bayesian semi-parametric estimation of the long-memory parameter under fexp-priors. (supplement).
- [6] Moulines, E. and Soulier, P. (2003). Semiparametric spectral estimation for fractional processes. In *Theory and applications of long-range dependence*, pages 251–301. Birkhäuser Boston, Boston, MA.
- [7] Rivoirard, V. and Rousseau, J. (2010). Bernstein-von mises theorem for linear functionals of the density.
- [8] Rousseau, J., Chopin, N., and Liseo, B. (2010). Bayesian nonparametric estimation of the spectral density of a long memory gaussian process.