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Adaptive Bayesian Estimation of a spectral density

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Abstract

Rousseau et al. [8] recently studied the asymptotic behavior of Bayesian estimators in the FEXP-model for spectral densities of Gaussian time-series. For the L_2 -norm on the log-spectral densities, they proved that the convergence rate is at least $n^{-\frac{\beta}{2\beta+1}}(\log n)^{\frac{2\beta+2}{2\beta+1}}$, $\beta > \frac{1}{2}$ being the Sobolev-regularity of the true spectral density f_o . We will improve upon the logarithmic factor, and prove that given a prior only depending on $\beta_s > \frac{1}{2}$, we have adaptivity to any $\beta \geq \beta_s$.

Keywords: Bayesian non-parametric, rates of convergence, adaptive estimation, long-memory time-series, FEXP-model

1. Introduction

Let X_t , $t \in \mathbb{Z}$, be a stationary zero mean Gaussian time series with spectral density $f_o(\lambda)$, $\lambda \in [-\pi, \pi]$ in the form

$$f_o(\lambda) = |1 - e^{i\lambda}|^{-2d_o} \exp\left\{\sum_{j=0}^{\infty} \theta_{o,j} \cos(j\lambda)\right\}, \quad \theta_o \in \Theta(\beta, L_o)$$
 (1.1)

where $d_o \in (-\frac{1}{2}, \frac{1}{2})$, $\Theta(\beta, L_o) = \{\theta \in l_2(\mathbb{N}) : \sum_{j \geq 0} \theta_j^2 (1+j)^{2\beta} \leq L_o\}$ is a Sobolev ball. The parameter d_o is called the long-memory parameter; we

will refer to $\exp\{\sum_{j=0}^{\infty} \theta_{o,j} \cos(j\lambda)\}$ as the short-memory part of the spectral density. The parameter β controls the regularity of the short-memory part. It is then natural to use the fractionally exponential or FEXP-model (see Beran [2] and Moulines and Soulier [6] and references therein) $\mathcal{F} = \bigcup_{k\geq 0} \mathcal{F}_k$, where

$$\mathcal{F}_k = \left\{ f_{d,k,\theta}(\lambda) = |1 - e^{i\lambda}|^{-2d} \exp\left\{ \sum_{j=0}^k \theta_j \cos(j\lambda) \right\}, \ d \in \left(-\frac{1}{2}, \frac{1}{2} \right), \theta \in \mathbb{R}^{k+1} \right\}.$$

We study Bayesian estimation of f_o within this FEXP-model. Let $\pi(d, k, \theta)$ denote the prior on (d, k, θ) ; this induces a prior on \mathcal{F} which we also denote π . Let $T_n(f)$ denote the covariance matrix of the observations $X = (X_1, \ldots, X_n)$, and let l_n be the associated log-likelihood

$$l_n(d, k, \theta) = -\frac{k+1}{2}\log(2\pi) - \frac{1}{2}\log|T_n(f)| - \frac{1}{2}X'T_n^{-1}(f)X$$
 (1.2)

Bayesian estimates of the spectral density f_o are based on the posterior

$$\pi(f \in A|X) = \frac{\int_A e^{l_n(d,k,\theta)} d\pi(f)}{\int_{\mathcal{F}} e^{l_n(d,k,\theta)} d\pi(f)}, \qquad A \subset \mathcal{F}.$$
 (1.3)

For example the posterior mean or median could be taken as 'point'-estimators of f_o . In this work however we focus on the posterior itself, and study the rate of convergence at which the posterior concentrates at f_o . More precisely, we lower-bound the posterior mass on the sets

$$B(\epsilon_n) = \{ f \in \mathcal{F} : l(f, f_o) < \epsilon_n^2 \},\$$

where ϵ_n is a sequence tending to zero and

$$l(f, f_o) = \frac{1}{2\pi} \int_{-\pi}^{\pi} (\log f_o(\lambda) - \log f(\lambda))^2 d\lambda.$$

Whether $\pi(B(\epsilon_n)|X)$ tends to one for a certain sequence ϵ_n critically depends on the smoothness of f_o as well as the smoothness induced by the prior. In Theorem 4.2 of Rousseau et al. [8] (RCL hereafter) it is shown that when $\theta_o \in \Theta(\beta, L_o)$ and the prior on θ has support contained in a Sobolev ball $\Theta(\beta, L)$ with L large enough, then the rate is $\epsilon_0(L)n^{-\frac{2\beta}{2\beta+1}}(\log)^{\frac{4\beta+4}{2\beta+1}}$, for fixed $\beta > \frac{1}{2}$ and $\epsilon_0(L)$ large enough depending on L. In the present work we prove

that such priors in fact lead to an adaptive concentration rate (in β) and we improve upon the constant $\epsilon_0(L)$ and the logarithmic factor. Adaptivity is of great interest since it is difficult to know the smoothness of the function f a priori. Improving on the constant ϵ_0 is crucial in Kruijer and Rousseau [4] but has also interest in its own. Indeed in Theorem 2.1 we prove that ϵ_0 depends only on L_o the radius of the Sobolev ball containing θ_o . In RCL however ϵ_0 depends on L, with the risk that would L be very large ϵ_0 might also be very large. Here we prove that this is not the case and that we can choose L as large as the application requires. This suggests that the result might actually hold without the constraint L in the prior on θ , but we have not been able to prove that.

Notation:

The m-dimensional identity matrix is denoted I_m . For matrices A we write |A| for the Frobenius or Hilbert-Schmidt norm $|A| = \sqrt{\operatorname{tr} A A^t}$, where A^t denotes the transpose of A. The operator or spectral norm is denoted $||A||^2 = \sup_{||x||=1} x'Ax$. We also use $||\cdot||$ for the Euclidean norm of finite dimensional vectors or sequences in $l^2(\mathbb{N})$, and for the L_2 -norm of functions. If $u \in l^1(\mathbb{N})$ we denote $||u||_1 = \sum_j |u_j|$. Given a sequence $\{u_j\}_{j\geq 0}$ and a nonnegative integer m, we write $u_{[m]}$ for the vector (u_0,\ldots,u_m) and $||u||_{>m}$ for the l^2 -norm of the sequence u_{m+1},u_{m+2},\ldots When we write $\sum_{j\geq 0} (\theta_j-\theta_{o,j})^2$ or $\sum_{j\geq 0} |\theta_j-\theta_{o,j}|$ for a finite-dimensional vector θ and $\theta_o \in l_2(\mathbb{N})$, θ_j is understood to be zero when j>k. For any function $h\in L_1([-\pi,\pi])$, $T_n(h)$ is the matrix with entries $\int_{-\pi}^{\pi} e^{i|l-m|\lambda}h(\lambda)d\lambda$, $l,m=1,\ldots,n$. For example, $T_n(f)$ is the covariance matrix of observations $X=(X_1,\ldots,X_n)$ from a time series with spectral density f. Let P_o denote the law associated with the true spectral density f_o and E_o expectations with respect to P_o .

2. Main results

Let $\beta_s > \frac{1}{2}$ be a fixed constant. We consider the following family of priors on (d,k,θ) . d is a priori independent of (k,θ) with density π_d with respect to Lebesgue measure. For some positive t < 1/2, the support of π_d is included in [-1/2+t,1/2-t]. We consider two cases for the prior on k:

Deterministic sieve $\pi_k(k) = \delta_{k_{A,n}}(k)$, i.e. it is the Dirac mass at $k_{A,n} = |A(n/\log n)^{1/(2\beta_s+1)}|$, for some positive A.

Random sieve the support of π_k is \mathbb{N} and satisfies:

$$e^{-c_1 k \log k} \le \pi_k(k) \le e^{-c_2 k \log k},$$

for some positive c_1, c_2 and k large enough. $\pi_{\theta|k}$, the prior on θ given k, has a density with respect to the Lebesgue measure on \mathbb{R}^k . This density is also denoted $\pi_{\theta|k}$, and is such that, for some constants L > 0 and $\beta_s > 1/2$, $\pi_{\theta|k}$ is positive on $\Theta_k(\beta, L)$ and $\pi_{\theta|k} [\Theta_k(\beta_s, L)^c] = 0$. These priors have been considered in particular in RCL, in Holan et al. [3] and in Kruijer and Rousseau [4]. We now state the main result.

Theorem 2.1. Suppose we observe $X = (X_1, ..., X_n)$ from a stationary, zero mean Gaussian time-series whose spectral density f_o is as in (1.1), with $d_o \in [-\frac{1}{2} + t, \frac{1}{2} - t]$, $\theta_o \in \Theta(\beta, L_o)$ and $\beta \geq \beta_s > \frac{1}{2}$. Consider a prior $\pi = \pi_d \pi_k \pi_{\theta|k}$ as described above such that there exists $c_0 > 0$ for which

$$\liminf_{n \to \infty} \min_{k \in \mathcal{K}_n} \inf_{\theta \in \Theta_k(\beta, L_o)} e^{c_0 k \log k} \pi_{\theta|k}(\theta) > 1.$$
 (2.1)

where, for some B > 0 and $k_{B,n} = \lfloor B(n/\log n)^{1/(2\beta_s+1)} \rfloor$, $\mathcal{K}_n = \{0, \ldots, k_{B,n}\}$ in the case of the random sieve prior, and $\mathcal{K}_n = \{k_{A,n}\}$ in the case of the deterministic prior. Assume also that L is large enough.

• In the case of the random sieve prior, for any $\beta_2 > \beta_s$, we have the following uniform result:

$$\sup_{f_o \in \cup_{\beta_s \le \beta \le \beta_2} \Theta(\beta, L_o)} E_o \pi((d, k, \theta)) : l(f_{d, k, \theta}, f_o) \ge l_0^2 \epsilon_n^2(\beta) | X) \le n^{-3}, \quad (2.2)$$

where $\epsilon_n(\beta) = (n/\log n)^{-\frac{\beta}{2\beta+1}}$ and l_0 only depends on L_o . In particular, it is independent of L.

• In the case of the deterministic prior, for any $\beta_2 > \beta_s$, we have the following uniform result:

$$\sup_{f_o \in \cup_{\beta_s \le \beta \le \beta_2} \Theta(\beta, L_o)} E_o \pi((d, k, \theta)) : l(f_{d,k,\theta}, f_o) \ge l_0^2 \epsilon_n^2(\beta_s) | X) \le n^{-3},$$
(2.3)

where l_0 only depends on L_o and is independent of L.

The constraint $\beta > 1/2$ is necessary to ensure that the short memory part $\exp(\sum_j \theta_{oj} \cos(jx))$ is bounded and continuous. As mentioned in the introduction, the fact that l_0 is independent of L is interesting since it allows us, in practice, to choose L arbitrarily high without penalizing the posterior concentration rate. It suggests that such results could hold with $L = \infty$,

however we have no proof for it. The random sieve prior leads to an adaptive posterior concentration rate over the range $\beta \geq \beta_s$, since for all $\beta > 1/2$, $\epsilon_n(\beta)$ is the minimax (up to a log n term) rate over the class of FEXP spectral densities given by (1.1) and associated to $\theta \in \Theta(\beta, L_o)$. The deterministic sieve prior does not lead to an adaptive procedure since the posterior concentration rate is $\epsilon_n(\beta_s)$ in this case. Obtaining adaptation by putting a prior on the dimension of the model is a commonly used strategy in Bayesian non parametrics, see for instance Arbel [1] or Rivoirard and Rousseau [7].

3. Proof of Theorem 2.1

We first introduce some notions that are useful throughout the proof.

3.1. Notation and preliminary results

We first introduce various (pseudo)-distances. We denote the Kullback -Leibler divergence between the Gaussian distributions associated with spectral densities f_o and f by

$$KL_n(f_o; f) = \frac{1}{2n} \left\{ \text{tr} \left[T_n(f_o) T_n^{-1}(f) - I_n \right] - \log \det(T_n(f_o) T_n^{-1}(f)) \right\},\,$$

a symmetrized version of it by $h_n(f_o, f) = KL_n(f_o; f) + KL_n(f; f_o)$ and the variance of the log-likelihood ratio by

$$b_n(f_o, f) = \frac{1}{n} \operatorname{tr} \left\{ T_n^{-1}(f) (T_n(f_o - f) T_n^{-1}(f) T_n(f_o - f)) \right\}.$$

The limiting values of $b_n(f_o, f)$ and $h_n(f_o, f)$ are denoted

$$h(f_o, f) = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left[\frac{f_o(\lambda)}{f(\lambda)} + \frac{f(\lambda)}{f_o(\lambda)} - 2 \right] d\lambda, \quad b(f_o, f) = (2\pi)^{-1} \int_{-\pi}^{\pi} \left(\frac{f_o}{f}(\lambda) - 1 \right)^2 d\lambda.$$

Then $h(f_o, f) \geq l(f_o, f)$ (RCL, p.6). Using Lemma 2 in RCL we find that for all $k \in \mathbb{N}$,

$$b_n(f_o, f_{d,k,\theta}) \le \|T_n(f_o)^{1/2} T_n(f)^{-1/2}\|^2 h_n(f_o, f_{d,k,\theta}) < C(\|\theta_o\|_1 + \|\theta\|_1) n^{2(d_o - d)_+} h_n(f_o, f),$$
(3.1)

where C is a universal constant. Similarly,

$$h_n(f_o, f) < \|T_n^{-\frac{1}{2}}(f)T_n^{\frac{1}{2}}(f_o)\|^2 b_n(f_o, f).$$
 (3.2)

In line with the notation of (1.2), let $\phi(x; d, k, \theta)$ denote the density of X, which is the Gaussian density with mean zero and covariance matrix $T_n(f_{d,k,\theta})$ and let $\phi(x; d_o, \theta_o)$ denote the Gaussian density associated with $T_n(f_o)$. We write $R_n(f_{d,k,\theta}) = \phi(X; d, k, \theta)/\phi(X; d_o, \theta_o)$ for the likelihood-ratio.

The proof of Theorem 2.1 contains two parts. First, it needs to be shown that the rate is $l_0^2 \epsilon_n^2$, for a constant l_0 that may depend on L and β_s . Then by re-insertion of the rate obtained in the first part, we improve upon the constant l_0 . In particular, it is shown to be independent of L for L large enough.

3.2. Proof of Theorem 2.1

Throughout the proof C denotes a universal constant. Let 0 < t < 1/2 and

$$\mathcal{G}_k(t,\beta_s,L) = \left\{ f_{d,k,\theta} : d \in \left[-\frac{1}{2} + t, \frac{1}{2} - t \right], \theta \in \Theta_k(\beta_s,L) \right\}, \quad \mathcal{G} = \bigcup_{k=0}^{\infty} \mathcal{G}_k(t,\beta_s,L).$$

By the results of RCL (Theorem 3.1, and Corollary 1 in the supplement) we have consistency for $h(f_o, f_{d,k,\theta})$ and $|d - d_o|$, i.e. for all $\delta, \epsilon > 0$, $\pi\left(f_{d,k,\theta}: h(f_{d,k,\theta}, f_o) < \epsilon^2, |d - d_o| < \delta |X|\right)$ tends to one in probability. Hence it suffices to show that

$$\pi [W_n | X] = \frac{\int_{W_n} R_n(f) d\pi(f)}{\int R_n(f) d\pi(f)} := \frac{N_n}{D_n} \stackrel{P_o}{\to} 0, \tag{3.3}$$

where in the case of the random sieve prior,

$$W_n = \left\{ f_{d,k,\theta} \in \mathcal{G} : l(f_o, f_{d,k,\theta}) \ge l_0 \epsilon_n^2(\beta), h(f_o, f_{d,k,\theta}) \le \epsilon^2, |d - d_o| \le \delta \right\},\,$$

for a constant $l_0 > 0$ depending only on L_o , β_s and the prior on k. In the case of the deterministic sieve prior, we replace $\epsilon_n^2(\beta)$ in this definition by $\epsilon_n^2(\beta_s)$. We present the proof of (3.3) for the case of the random sieve prior; the proof for the case of the deterministic sieve prior can be deduced by replacing β by β_s . The proof consists of two parts: first we show that for some c > 0,

$$P_o\left[D_n < e^{-2nu_0\epsilon_n^2(\beta)}/2\right] \le e^{-cn\epsilon_n^2(\beta)},\tag{3.4}$$

for which we will establish a lower bound the prior mass on a Kullback-Leibler neighborhood of f_o . In the second part we show that under the event

 $D_n \ge \frac{1}{2} e^{-nu_0 \epsilon_n^2(\beta)}$ we can control N_n/D_n . This will be done by giving a bound on the upper-bracketing entropy of the model.

For the proof of (3.4), note that RCL already found that if $\beta \geq \beta_s > 1/2$, there exists $u_0 \geq 0$ depending only on L_o such that

$$P_o\left[D_n < e^{-nu_0\epsilon_n^2(\beta)(\log n)^{1/(2\beta+1)}}/2\right] = o(n^{-1}).$$

To prove (3.4), we thus need to improve on the $\log n$ term in the preceding equation. Set

$$\bar{\mathcal{B}}_n = \{(d, k, \theta); KL_n(f_o, f_{d,k,\theta}) \le \frac{\epsilon_n^2(\beta)}{4}, b_n(f_o, f_{d,k,\theta}) \le \epsilon_n^2(\beta), d_o \le d \le d_o + \delta\},$$

for some positive δ . Recall that

$$D_{n} = \sum_{k} \pi_{k}(k) \int e^{l_{n}(d,k,\theta) - l_{n}(f_{o})} d\pi_{\theta|k}(\theta) d\pi_{d}(d),$$

$$P_{o}^{n} \left[D_{n} < e^{-2nu_{0}\epsilon_{n}^{2}(\beta)} \right] \leq P_{o}^{n} \left[\int_{\bar{\mathcal{B}}_{n}} e^{l_{n}(f) - l_{n}(f_{o})} d\pi(f) < e^{-2nu_{0}\epsilon_{n}^{2}(\beta)} \right].$$

From the proof of Theorem 4.1 in RCL (section 5.2.1), it follows that

$$P_0^n \left(D_n \le e^{-nu_0 \epsilon_n^2(\beta)} \pi(\bar{\mathcal{B}}_n) / 2 \right) \le e^{-Cn\epsilon_n^2(\beta)}$$

for some constant C > 0 (independent of L). We now show that

$$\pi(\bar{\mathcal{B}}_n) \ge e^{-nu_0\epsilon_n^2(\beta)/4}. (3.5)$$

Define

$$\tilde{\mathcal{B}}_n = \{ (d, k_{B,n}, \theta); d_o \le d \le d_o + \epsilon_n^2(\beta) n^{-a}, |\theta_j - \theta_{o,j}| \le (1+j)^{-\beta} \epsilon_n(\beta) n^{-a}, j = 0, \dots, k_{B,n} \}.$$

We first prove that $\pi(\tilde{\mathcal{B}}_n) \geq e^{-nu_0\epsilon_n^2(\beta)/4}$ and then that $\tilde{\mathcal{B}}_n \subset \bar{\mathcal{B}}_n$. As in RCL (see the paragraph following equation (29) on p. 26), we find that

$$\sum_{j=1}^{\infty} (1+j)^{2\beta} \theta_j^2 \le 2L_o + \epsilon_n^2(\beta) n^{-2a} \le 3L_o, \forall \theta \in \tilde{\mathcal{B}}_n$$

for n large enough. Combined with condition (2.1) on $\pi_{\theta|k}$, this implies

$$\pi(\tilde{\mathcal{B}}_n) \ge \left(c\epsilon_n(\beta)k_{B,n}^{-\beta}n^{-a}\right)^{k+3}e^{-c_0k_{B,n}\log k_{B,n}} \ge e^{-c(\beta_s)k_0n\epsilon_n^2(\beta)}, \quad \forall \beta \ge \beta_s.$$

This achieves the proof of (3.4) with $u_0 = c(\beta_s)k_0$.

To show that $\tilde{\mathcal{B}}_n$ is included in $\bar{\mathcal{B}}_n$, first note that equation (3.1) implies that it is enough to bound $h_n(f_o, f)$ on $\tilde{\mathcal{B}}_n$. To this end, we use the decomposition $f_o = f_{o,k_{B,n}} e^{\Delta_{d_o,k_{B,n}}}$, where $f_{o,k_{B,n}} = f_{d_o,k_{B,n},\theta_o}$ and

$$\Delta_{d_o, k_{B,n}}(\lambda) = \sum_{j=k_{B,n}+1}^{\infty} \theta_{o,j} \cos(jx), \quad \forall \lambda \in [-\pi, \pi].$$

Then we have the expansion

$$f_o = f_{o,k_{B,n}} (1 + \Delta_{d_o,k_{B,n}} + \Delta_{d_o,k_{B,n}}^2 / 2 + O(\Delta_{d_o,k_{B,n}}^3)), \quad |\Delta_{d_o,k_{B,n}}|_{\infty} = o(1)$$

and

$$h_n(f_o, f) \le 2[h_n(f_o, f_{o, k_{B,n}}) + h_n(f_{o, k_{B,n}}, f)].$$
 (3.6)

We first deal with the first term above. Let $b_{o,n} = e^{\Delta_{d_o,k_{B,n}}} - 1$ and without loss of generality we can assume that $b_{o,n}$ is positive in the expression of $h_n(f_o, f_{o,k_{B,n}})$ so that for all $\beta > 1/2$,

$$h_{n}(f_{o}, f_{o,k_{B,n}}) := \frac{1}{2n} tr \left[T_{n}^{-1}(f_{o}) T_{n}(f_{o}b_{o,n}) T_{n}^{-1}(f_{o}) T_{n}(f_{o}b_{o,n}) \right]$$

$$\leq \frac{c}{n} tr \left[T_{n}^{-1}(g_{o}) T_{n}(g_{o}b_{o,n}) T_{n}^{-1}(g_{o}) T_{n}(g_{o}b_{o,n}) \right]$$

$$= \frac{c}{n} tr \left[T_{n}^{2}(b_{o,n}) \right] + c\gamma_{1} + c\gamma_{2}$$

$$(3.7)$$

where $g_o(\lambda) = |\lambda|^{-2d_o}$ and c depends only on $\sum_{j=0}^{\infty} |\theta_{o,j}| \leq L_o^{1/2} (2\beta - 1)^{-1/2}$, and

$$\gamma_{1} = \frac{1}{n} \left(tr \left[\left(T_{n} \left(\frac{1}{4\pi^{2}g_{o}} \right) T_{n}(g_{o}b_{o,n}) \right)^{2} \right] - tr \left[T_{n}^{2}(b_{o,n}) \right] \right)$$

$$\gamma_{2} = \frac{1}{n} \left(tr \left[\left(T_{n}^{-1}(g_{o}) T_{n}(g_{o}b_{o,n}) \right)^{2} \right] - tr \left[\left(T_{n} \left(\frac{1}{4\pi^{2}g_{o}} \right) T_{n}(g_{o}b_{o,n}) \right)^{2} \right] \right).$$

We first bound the first term of the right hand side of (3.7). Note that $b_{o,n}(\lambda) = \tilde{\Delta}_{d_o,k_{B,n},K_n} + R_0$, where

$$\tilde{\Delta}_{d_o,k_{B,n},K_n}(\lambda) = \sum_{j=k_{B,n}+1}^{K_n} \theta_{o,j} \cos(j\lambda), \quad K_n = \epsilon_n(\beta)^{-1/\beta} (\log n)^{1/\beta}$$

and $||R_0||^2 \le \epsilon_n^2(\beta)(\log n)^{-2}$ so that

$$\operatorname{tr}\left[T_{n}^{2}(b_{o,n})\right] = \operatorname{tr}\left[T_{n}^{2}(\tilde{\Delta}_{d_{o},k_{B,n},K_{n}})\right] + O(\log n[K_{n}^{-2\beta} + K_{n}^{-\beta}k_{B,n}^{-\beta}])$$

$$= \operatorname{tr}\left[T_{n}^{2}(\tilde{\Delta}_{d_{o},k_{B,n},K_{n}})\right] + O(\epsilon_{n}^{2}(\beta)),$$
(3.8)

where the term $O(\log n[K_n^{-2\beta} + K_n^{-\beta}k_{B,n}^{-\beta}])$ comes from the fact that

$$\left| \operatorname{tr} \left[T_n^2(b_{o,n}) \right] - \operatorname{tr} \left[T_n^2(\tilde{\Delta}_{d_o,k_{B,n},K_n}) \right] \right| \le \operatorname{tr} \left[T_n^2(R_0) \right] + |T_n(R_0)| |T_n(\tilde{b}_{o,n})|$$

and from the use of inequality (20) in Lemma 6 of RCL, with $f_1 = f_2 = 1$, $\delta = 0$ and b either equal to $\tilde{\Delta}_{d_o,k_{B,n},K_n}$ or R_0 . Note that the constant in the term $O(\epsilon_n^2(\beta))$ in (3.8) does not depend on L. Lemma 2.1 in Kruijer and Rousseau [5] together with the fact that

$$|\tilde{\Delta}_{d_o,k_{B,n},K_n}(\lambda) - \tilde{\Delta}_{d_o,k_{B,n},K_n}(y)| \le \sum_{j=k_{B,n}}^{K_n} j|\theta_{o,j}| \le C(\beta)L_0K_n^{-\beta+3/2} \lor k_{B,n}^{-\beta+3/2}$$

implies that for large enough n,

$$n^{-1} \left| \operatorname{tr} \left[T_n^2 (\tilde{\Delta}_{d_o, k_{B,n}, K_n}) \right] - 2\pi \operatorname{tr} \left[T_n (\tilde{\Delta}_{d_o, k_{B,n}, K_n}^2) \right] \right|$$

$$\leq KC(\beta) L_o n^{-2\beta + 1 + \epsilon} \epsilon_n^2(\beta) = o(\epsilon_n^2(\beta)), \quad \forall \epsilon > 0,$$

uniformly over $\beta_s \leq \beta \leq \beta_2$ and $\theta_o \in \Theta(\beta, L_o)$. Consequently,

$$\frac{c}{n}tr[T_n^2(b_{o,n})] \le tr\left[T_n^2(\tilde{\Delta}_{d_o,k_{B,n},K_n})\right] + C(L_o,\beta)\epsilon_n^2(\beta), \tag{3.9}$$

for a constant $C(L_o, \beta)$ independent of L. Next we apply Lemma 2.4 in Kruijer and Rousseau [5] with $f = g_o$ and $b_1 = b_2 = b_{o,n}$; it then follows that

$$\gamma_1 \le \|b_{o,n}\|_{\infty}^2 n^{\delta - 1} n^{\epsilon - 1} = o(\epsilon_n^2(\beta)), \quad \forall \epsilon > 0.$$
(3.10)

Finally Lemma 2.3 in Kruijer and Rousseau [5] implies that for all $\epsilon > 0$,

$$\gamma_2 \le \|b_{o,n}\|_{\infty}^2 n^{-1+\epsilon} = o(\epsilon_n^2(\beta)).$$
 (3.11)

Combining (3.9), (3.10) and (3.11), it follows that

$$h_n(f_{o,k_{B,n}}, f_o) = n^{-1} \operatorname{tr} \left[T_n(\tilde{\Delta}_{d_o,k_{B,n},K_n}^2) \right] + C(L_o, \beta) \epsilon_n^2(\beta) \le 2C'(L_o, \beta) \epsilon_n^2(\beta),$$

where also $C'(L_o, \beta)$ is independent of L. The last inequality follows from

$$n^{-1}\operatorname{tr}\left[T_{n}(\tilde{\Delta}_{d_{o},k_{B,n},K_{n}}^{2})\right] = \frac{1}{2\pi} \int_{-\pi}^{\pi} \tilde{\Delta}_{d_{o},k_{B,n},K_{n}}^{2}(\lambda) d\lambda$$
$$= \sum_{j=k_{B,n}+1}^{K_{n}} \theta_{o,j}^{2} \leq C'(L_{o},\beta) \epsilon_{n}^{2}(\beta).$$

We now bound the last term in (3.6), which we write as

$$h_n(f_{o,k_{B,n}},f) = \frac{1}{2n} \operatorname{tr} \left[T_n(f_{o,k_{B,n}})^{-1} T_n(fb) T_n(f)^{-1} T_n(fb) \right], \quad b = (f - f_{o,k_{B,n}}) / f.$$

Since $d \ge d_o$, $|b|_{\infty} < +\infty$ and applying Lemma 6 inequality (20) of RCL, we obtain if $d, \theta \in \tilde{\mathcal{B}}_n$ with a > 0,

$$h_n(f_{o,k_{B,n}}, f) \le C \log n \left(|b|_2^2 + |d - d_o| |b|_{\infty}^2 \right)$$

$$\le C \log n \left(\sum_{j=1}^{k_{B,n}} (\theta_j - \theta_{o,j})^2 + n^{-a} \epsilon_n^2(\beta) \right) = o(\epsilon_n^2(\beta)),$$

which finally implies that $\pi(\bar{\mathcal{B}}_n) \geq \pi(\tilde{\mathcal{B}}_n)$ and that (3.4) is proved. We now find an upper bound on N_n . First write $\bar{W}_n = W_n \cap \mathcal{F}_n$ where $\mathcal{F}_n = \{f_{d,k,\theta}; k \leq k_{B_1,n}\}$ and $k_{B_1,n} = B_1(n/\log n)^{1/(2\beta+1)}$. Then, since the prior on k is Poisson,

$$\pi(\mathcal{F}_n^c) \le e^{-c_2 k_{B_1,n} \log k_{B_1,n}} \le e^{-2nu_0 \epsilon_n^2(\beta)}$$

if B_1 is large enough (depending on $L_o, c_2, \beta_s, \beta_2$) and W_n can be replaced by \bar{W}_n in the definition of N_n . Following the proof of RCL, we decompose $\bar{W}_n = \bigcup_{l=l_0}^{l_n} W_{n,l}$, where $l_0 \geq 2$, $l_n = \lceil \epsilon^2/\epsilon_n^2(\beta) \rceil - 1$ and

$$W_{n,l} = \left\{ f_{d,k,\theta} \in \mathcal{G} : k \le k_{B_1,n}, h(f_{d,k,\theta}, f_o) \le \epsilon^2, |d - d_o| \le \delta, \\ \epsilon_n^2 l \le h_n(f_o, f_{d,k,\theta}) \le u_n(l+1) \right\}.$$

In addition let $N_{n,l} = \int_{W_{n,l}} R_n(f) d\pi(f)$; then $N_n = \sum_{l=l_0}^{l_n} N_{n,l}$, and we have

$$E_o\left[\frac{N_n}{D_n}\right] \le P_o^n \left(D_n \le e^{-nu_0\epsilon_n^2(\beta)}/2\right) + E_o\left[\sum_{l=l_0}^{l_n} \frac{N_{n,l}}{D_n} 1_{\{D_n \ge e^{-nu_0\epsilon_n^2(\beta)}/2\}}\right]. \tag{3.12}$$

We construct tests $\bar{\phi}_l$ $(l = l_0, \dots, l_n)$ and write

$$E_{o}\left[\sum_{l=l_{0}}^{l_{n}} \frac{N_{n,l}}{D_{n}} 1_{\{D_{n} \geq e^{-nu_{n}}/2\}} (\bar{\phi}_{l} + 1 - \bar{\phi}_{l})\right]$$

$$\leq \sum_{l=l_{0}}^{l_{n}} E_{o}(\bar{\phi}_{l}) + 2e^{nu_{n}} \sum_{l=l_{0}}^{l_{n}} E_{o}\left[N_{n,l} (1 - \bar{\phi}_{l})\right].$$
(3.13)

The tests are based on a collection of spectral densities $H_{n,l} = \bigcup_{k=0}^{k_{B_1,n}} H_{n,l,k} \subset W_{n,l}$ defined as follows. Let D_l be a grid over $\{d: |d-d_o| \leq \delta\}$ with spacing $l\epsilon_n^2(\beta)/(\log n)$. Let $T_{l,k}$ denote the centers of hypercubes of radius $\frac{l\epsilon_n^2(\beta)}{k}$, covering $\Theta_k(\beta_s, L)$. We define $H_{n,l,k}$ as the collection of spectral densities $f_{l,i} = (2e)^{l\epsilon_n^2(\beta)} f_{d_{l,i},k,\theta^{l,i}}$, with $d_{l,i} \in D_l$ and $\theta^{l,i} \in T_{l,k}$. With every $f_{l,i}$ we associate a test

$$\phi_{l,i} = 1_{\{X'(T_n^{-1}(f_o) - T_n^{-1}(f_{l,i}))X \ge \operatorname{tr}\{I_n - T_n(f_o)T_n^{-1}(f_{l,i}) + \frac{n}{4}h_n(f_o, f_{l,i})\}\}},$$
(3.14)

and set $\bar{\phi}_l = \max_i \phi_{l,i}$.

The set $H_{n,l}$ can be seen as a collection of upper-bracket spectral densities, since for each $f_{d,k,\theta} \in W_{n,l}$ there exists a $f_{l,i} \in H_{n,l,k}$ such that $f_{l,i} \geq f_{d,k,\theta}$, $0 \leq d_{l,i} - d \leq l\epsilon_n^2(\beta)/(\log n)$ and

$$0 \leq (2e)^{l\epsilon_n^2(\beta)} \exp\left\{\sum_{j=0}^k \theta_j^{l,i} \cos(jx)\right\} - \exp\left\{\sum_{j=0}^k \theta_j \cos(jx)\right\}$$

$$\leq \frac{l\epsilon_n^2(\beta)}{32} (2e)^{l\epsilon_n^2(\beta)} \exp\left\{\sum_{j=0}^k \theta_j^{l,i} \cos(jx)\right\}.$$

$$(3.15)$$

The cardinality of $\bigcup_{k=0}^{k_{B_1,n}} H_{n,l,k}$ is at most

$$(l^{-1}k_{B_1,n}\epsilon_n(\beta)^{-4})^{k_{B_1,n}} \frac{\delta \log n}{l\epsilon_n^2(\beta)} \le \exp\{2k_{B_1,n}\log n\} = C_{n,l},$$
 (3.16)

for all $l \geq 2$. To bound the right hand side of (3.13), we use (3.16) in combination with the following error bounds for each of the tests $\phi_{l,i}$. Let $f \in W_{n,l}$ and let $f_{l,i} \in H_{n,l}$ be such that (3.15) holds, $\phi_{l,i}$ being the associated test-function. Then, from equation (4.4) in RCL together with the bound (3.1) on $b_n(f_o, f_{l,i})/h_n(f_o, f_{l,i})$ which again depends on $\|\theta^{l,i}\|_1$ and L_0 , we

obtain that for all $0 < \alpha < 1$, there exists constants $d_1, d_2 > 0$ depending on L_o and $\|\theta^{l,i}\|_1$ such that

$$E_o \phi_{l,i} \le e^{-d_1 n l^{\alpha} \epsilon_n^2(\beta)}, \quad E_f^n(1 - \phi_{l,i}) \le e^{-d_2 n l^{\alpha} \epsilon_n^2(\beta)}.$$
 (3.17)

Using (3.17) we obtain the following bound on the term $\sum_{l=l_0}^{l_n} E_o(\bar{\phi}_l)$ in (3.13):

$$\sum_{l=l_0}^{l_n} E_o\left(\bar{\phi}_l\right) \le \sum_{l=l_0}^{l_n} C_{n,l} e^{-d_1 n l^{\alpha} \epsilon_n^2(\beta)} \le e^{2k_{B_1,n} \log n} \sum_{l=l_0}^{l_n} e^{-d_1 n l^{\alpha} \epsilon_n^2(\beta)} \to 0,$$

as soon as $l_0 \ge \left(\frac{2B_1}{d_1 u_0}\right)^2$, choosing $\alpha = 1/2$. Using (3.17) the last term in (3.13) is

$$\sum_{l=l_0}^{l_n} E_o\left[N_{n,l}(1-\bar{\phi}_l)\right] = \sum_{l=l_0}^{l_n} \int_{W_{n,l}} E_f^n(1-\bar{\phi}_l) d\pi(f) \le e^{-d_2 n l_0^{1/2} \epsilon_n^2(\beta)} \le e^{-2n\epsilon_n^2(\beta)}$$

as soon as $d_2 l_0^{1/2} \ge 2$, i.e. $l_0 \ge 4d_2^{-2}$. Note that the two lower bounds on l_0 depends on L_o and on $\|\theta^{l,i}\|_1$. Finally choosing, $l_0 = \max\left(4d_2^{-2}, \left(\frac{2B_1}{d_1u_0}\right)^2, 2, u_0\right)$ we obtain

$$P^{\pi}\left[h_n(f, f_o) \le l_0 \epsilon_n^2(\beta) | X^n\right] = o(n^{-1}).$$

From that we deduce a concentration rate in terms of the l norm, following RCL's argument in Appendix C. Let l_0 be an arbitrary constant and assume that $h_n(f, f_o) \leq l_0 \epsilon_n^2(\beta)$ and $f = f_{d,k,\theta}$. Then inequality (C.3) of Lemma 6 of RCL implies that

$$\frac{1}{n} \text{tr} \left[T_n(f_o^{-1}) T_n(f_o - f) T_n(f^{-1}) T_n(f_o - f) \right] \le C_1 l_0 \epsilon_n,$$

where C_1 depends only on $\|\theta\|_1$ and on $\|\theta_o\|_1$. This implies that

$$\frac{1}{n} \text{tr} \left[T_n(f_o^{-1}(f_o - f)) T_n(f^{-1}(f_o - f)) \right] \le 2C_1 l_0 \epsilon_n^2(\beta)$$

since the difference between the two terms is of order $O(n^{-1+2a}) \, \forall a > 0$, which also implies that $h(f_o, f) \leq 3C_1 l_0 \epsilon_n^2(\beta)$, for the same reason. Since $l(f_o, f) \leq h(f_o, f)$, we finally obtain that

$$P^{\pi} \left[l(f, f_o) \le 3C_1 l_0 \epsilon_n^2(\beta) | X^n \right] = o(n^{-1}).$$

To terminate the proof of Theorem 2.1, it only remains to prove that l_0 depends only on L_o , β_s . This is done using a simple re-insertion argument. Recall also that $k \leq k_{B_1,n} = B_1 n \epsilon_n^2(\beta)$, where B_1 is independent of the radius L of the sobolev-ball $\Theta(\beta_s, L)$ defining the support of the prior. We start with the following observation. From Kruijer and Rousseau [4] (equation (3.5)) it follows that for fixed d and k, the minimizer of $l(f_o, f_{d,k,\theta})$ over \mathbb{R}^{k+1} is

$$\bar{\theta}_{d,k} := \operatorname{argmin}_{\theta \in \mathbb{R}^{k+1}} l(f_o, f_{d,k,\theta}) = \theta_{o[k]} + (d_o - d)\eta_{[k]},$$

where η is defined by $\eta_j = -2/j$ $(j \geq 1)$ and $\eta_0 = 0$. Assuming that $l(f_o, f_{d,k,\theta}) \leq 3C_1l_0\epsilon_n^2(\beta)$ and $k \leq k_{B_1,n}$ leads to $l(f_o, f_{d,k,\bar{\theta}_{d,k}}) \leq 3C_1l_0\epsilon_n^2$ and $\|\theta - \bar{\theta}_{d,k}\|^2 = l(f_{d,k,\theta}, f_{d,k,\bar{\theta}_{d,k}}) \leq 12C_1l_0\epsilon_n^2$. Therefore

$$\sum_{j=0}^{k} |\theta_{j}| \leq \sum_{j=0}^{k_{B_{1},n}} |\theta_{j} - (\bar{\theta}_{d,k_{n}})_{j}| + \sum_{j=0}^{k_{B_{1},n}} |(\bar{\theta}_{d,k_{n}})_{j}|$$

$$\leq \sqrt{12C} \sqrt{l_{0}} \epsilon_{n} k_{B_{1},n}^{1/2} + 2|d - d_{o}| \log n + \sum_{j=0}^{k_{B_{1},n}} j^{\rho} |\theta_{o,j}| \leq 2(2\beta_{s} - 1)^{-1/2} \sqrt{L_{o}}$$

when n is large enough, where the second inequality comes from Lemma 3.1 of Kruijer and Rousseau [4]. This achieves the proof of Theorem 2.1.

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References

- [1] Arbel, J. (2010). Bayesian optimal adaptive estimation using a sieve prior, submitted.
- [2] Beran, J. (1993). Fitting long-memory models by generalized linear regression. *Biometrika*, 80(4):817–822.
- [3] Holan, S., McElroy, T., and Chakraborty, S. (2009). A Bayesian approach to estimating the long memory parameter. *Bayesian anlysis*, 4(1):159–190.
- [4] Kruijer, W. and Rousseau, J. (2011a). Bayesian semi-parametric estimation of the long-memory parameter under fexp-priors.

- [5] Kruijer, W. and Rousseau, J. (2011b). Bayesian semi-parametric estimation of the long-memory parameter under fexp-priors. (supplement).
- [6] Moulines, E. and Soulier, P. (2003). Semiparametric spectral estimation for fractional processes. In *Theory and applications of long-range dependence*, pages 251–301. Birkhäuser Boston, Boston, MA.
- [7] Rivoirard, V. and Rousseau, J. (2010). Bernstein-von mises theorem for linear functionals of the density.
- [8] Rousseau, J., Chopin, N., and Liseo, B. (2010). Bayesian nonparametric estimation of the spectral density of a long memory gaussian process.