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ON THE LINK BETWEEN DEPENDENCE AND INDEPENDENCE IN EXTREME VALUE THEORY FOR DYNAMICAL SYSTEMS

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ABSTRACT. We reformulate the standard conditions that allow reducing the study of extremes for dependent sequences to the classical extreme value theory. Essentially, we weaken the mixing type condition in such a way that, in the context of dynamical systems, it should follow from decay of correlations.

1. INTRODUCTION

Let (Ω, \mathcal{A}, P) be a probability space and $X : \Omega \rightarrow \mathbb{R}$ a random variable (r.v.) with distribution function (d.f.) given by $F(x) = P(X \leq x)$. Also, let $E(\cdot)$ denote expectation with respect to P so that $E(X) = \int X dP = \int x dF(x)$. Consider a stationary stochastic process X_0, X_1, \dots of r.v. defined on Ω with common d.f. F . For notational simplicity we assume that $X = X_0$. For every $i, j, n \in \mathbb{N}$, set

$$M_{i,j} = \max \{X_i, X_{i+1}, \dots, X_{i+j-1}\} \quad \text{and} \quad M_n = M_{0,n}. \quad (1.1)$$

The study of the limiting behavior for maxima of a stationary process can be reduced, under adequate conditions on the dependence structure, to the Classical Extreme Value Theory for sequences of i.i.d. r.v. Hence, to the stationary process X_0, X_1, \dots we associate an independent sequence of r.v. denoted by Z_0, Z_1, \dots with common d.f. F . We also set for each $n \in \mathbb{N}$

$$\hat{M}_n = \max \{Z_0, \dots, Z_{n-1}\}. \quad (1.2)$$

Let us focus on the conditions that allow us to relate the asymptotic distribution of M_n with that of \hat{M}_n . Following Leadbetter et al. (1983) we refer to these conditions as $D(u_n)$ and $D'(u_n)$, where u_n is a suitable sequence of thresholds converging to $\sup_{\omega \in \Omega} X(\omega) = 1$, as n goes to ∞ , that will be defined below. $D(u_n)$ imposes a certain type of distributional mixing property. Essentially, it says that the dependence between some special type of

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events fades away as they become more and more apart in the time line. $D'(u_n)$ restricts the appearance of clusters, that is, it makes the occurrence of consecutive ‘exceedances’ of the level u_n an unlikely event.

The purpose of this work is to verify that condition $D(u_n)$ can be stated in a weaker form and the result still prevails. The advantage of having this weaker requirement is that, in the context of Dynamical Systems, $D(u_n)$ should follow from decay of correlations.

We say that an *exceedance* of the level u_n occurs at time i if $X_i > u_n$. The probability of such an exceedance is $1 - F(u_n)$ and so the mean value of the number of exceedances occurring up to n is $n(1 - F(u_n))$. The sequences of levels u_n we consider are such that $n(1 - F(u_n)) \rightarrow \tau$ as $n \rightarrow \infty$, for some $\tau \geq 0$, which means that, in a time period of length n , the expected number of exceedances is approximately τ .

Condition $D(u_n)$ is a type of mixing requirement specially adapted to Extreme Value Theory. In this context, the events of interest are those of the form $\{X_i \leq u\}$ and their intersections. Observe that $\{M_n \leq u\}$ is just $\{X_0 \leq u, \dots, X_{n-1} \leq u\}$. Motivated by the work of Collet (2001), we propose:

Condition ($D(u_n)$). We say that $D(u_n)$ holds for the sequence X_0, X_1, X_2, \dots if for any integers ℓ, t and n

$$P(\{X_0 > u_n\} \cap \{M_{t,\ell} \leq u_n\}) - P(\{X_0 > u_n\})P(\{M_\ell \leq u_n\}) \leq \gamma(n, t),$$

where $\gamma(n, t)$ is nonincreasing in t for each n and $n\gamma(n, t_n) \rightarrow 0$ as $n \rightarrow \infty$ for some sequence $t_n = o(n)$, which means that $t_n/n \rightarrow 0$ as $n \rightarrow \infty$.

We remark that the actual definition of $D(u_n)$ appearing in Leadbetter et al. (1983), Section 3.2, is, in some sense, a stronger requirement.

The sequence u_n is such that the average number of exceedances in the time interval $\{0, \dots, [n/k]\}$ is approximately τ/k , which goes to zero as $k \rightarrow \infty$. However, the exceedances may have a tendency to be concentrated in the time period following the first exceedance at time 0. To avoid this we introduce

Condition ($D'(u_n)$). We say that $D'(u_n)$ holds for the sequence X_0, X_1, X_2, \dots if

$$\lim_{k \rightarrow \infty} \limsup_{n \rightarrow \infty} n \sum_{j=1}^{[n/k]} P(\{X_0 > u_n\} \cap \{X_j > u_n\}) = 0. \quad (1.3)$$

Notice that Condition 1.3 forbids the concentration of exceedances by bounding the probability of more than one exceedance in the time interval $\{0, \dots, [n/k]\}$. This guarantees that the exceedances should appear scattered through the time period $\{0, \dots, n-1\}$.

Our goal is to show that M_n and \hat{M}_n have the same asymptotic distribution under $D(u_n)$ and $D'(u_n)$.

Theorem 1. *Let $(u_n)_{n \in \mathbb{N}}$ be such that $nP(X > u_n) = n(1 - F(u_n)) \rightarrow \tau$, as $n \rightarrow \infty$, for some $\tau \geq 0$. Assume that conditions $D(u_n)$ and $D'(u_n)$ hold. Then*

$$\lim_{n \rightarrow \infty} P(M_n \leq u_n) = \lim_{n \rightarrow \infty} P(\hat{M}_n \leq u_n).$$

2. EXTREMES ON DYNAMICAL SYSTEMS

In this section, we observe that the study of the partial maximum of deterministic processes arising from a certain dynamical system that possesses an invariant measure, with sufficient decay of correlations, can be reduced to prove $D'(u_n)$ and investigate the domain of attraction of maximum of the associated independent process.

Let $f : M \rightarrow M$ be a smooth map on the manifold M admitting an invariant Borel probability measure μ . Assume that there exists a nonincreasing $g : \mathbb{N} \rightarrow \mathbb{R}$ such that for all $\varphi, \psi : M \rightarrow \mathbb{R}$ with bounded variation, there is $C > 0$ independent of φ, ψ and n such that

$$\left| \int \varphi \cdot (\psi \circ f^n) d\mu - \int \varphi d\mu \int \psi d\mu \right| \leq C \text{Var}(\varphi) \|\psi\|_\infty g(n), \quad \forall n \geq 0, \quad (2.1)$$

where $\text{Var}(\varphi)$ denotes the total variation of φ and $ng(t_n) \rightarrow 0$, as $n \rightarrow \infty$ for some $t_n = o(n)$.

Consider now the r.v. $X : M \rightarrow \mathbb{R}$ of bounded variation on the probability space (M, \mathcal{B}, μ) , where \mathcal{B} is the Borel σ -algebra on M , with d.f. $F(x) = \mu(\{X^{-1}(-\infty, x]\})$. Define the stationary stochastic process X_0, X_1, X_2, \dots by $X_i = X \circ f^i$. Let

$$M_n = \max\{X_0, \dots, X_{n-1}\},$$

for each $n \in \mathbb{N}$. Denote by Z_0, Z_1, Z_2, \dots an independent sequence of r.v. with common d.f. F and set, for each $n \in \mathbb{N}$,

$$\hat{M}_n = \max\{Z_0, \dots, Z_{n-1}\}.$$

Taking $\varphi = \mathbf{1}_{\{X > u_n\}}$ and $\psi = \mathbf{1}_{\{M_\ell \leq u_n\}}$, then (2.1) implies that Condition $D(u_n)$ holds with $\gamma(n, t) = \gamma(t) = C \text{Var}(\mathbf{1}_{\{X > u_n\}}) \|\mathbf{1}_{\{M_\ell \leq u_n\}}\|_\infty g(t) \leq C(\text{Var}(X) + 1)g(t)$ and for the sequence t_n such that $t_n/n \rightarrow 0$ and $ng(t_n) \rightarrow 0$ as $n \rightarrow \infty$.

This means that if we are also able to prove $D'(u_n)$ for the sequence X_0, X_1, \dots then the study of the limiting behavior of M_n is reduced to study the domain of attraction for maximums of F using the usual tools of Classical Extreme Value Theory, available, for example, in Section 1.6 of Leadbetter et al. (1983). This type of study, namely the proof of condition $D'(u_n)$, was proved in the context of Benedicks-Carleson quadratic dynamical systems, in Freitas and Freitas (2007). We also mention the work of Haiman (2003) who obtained an extreme limiting law for the tent map using an approach relying on the linear property of the map.

Remark 2. Observe that often decay of correlations is known for Hölder continuous functions instead. In these situations one can use an Hölder continuous approximation of $\mathbf{1}_{\{X > u_n\}}$ and try to prove that the error introduced does not affect much, just as Collet (2001) did in Lemma 3.3.

3. PROOF OF THE RESULT

In this section, we prove Theorem 1 using the same type of approach used in Section 3 of Collet (2001). We begin by stating two abstract Lemmas.

Lemma 3.1. For any $\ell \in \mathbb{N}$ and $u \in \mathbb{R}$ we have

$$\sum_{j=0}^{\ell-1} P(X_j > u) \geq P(M_\ell > u) \geq \sum_{j=0}^{\ell-1} P(X_j > u) - \sum_{j=0}^{\ell-1} \sum_{i \neq j, i=0}^{\ell-1} P(\{X_j > u\} \cap \{X_i > u\})$$

Proof. This is a straightforward consequence of the formula for the probability of a multiple union on events. See for example first Theorem of Chapter 4 in Feller (1952). \square

Lemma 3.2. Assume that t, r, m, ℓ, s are nonnegative integers. Then, we have

$$0 \leq P(M_r \leq u) - P(M_{r+\ell} \leq u) \leq \ell \cdot P(X > u) \quad (3.1)$$

and

$$\begin{aligned} & \left| P(M_{s+t+m} \leq u) - P(M_m \leq u) + \sum_{j=0}^{s-1} P(\{X > u\} \cap \{M_{s+t-j,m} \leq u\}) \right| \\ & \leq 2s \sum_{j=1}^{s-1} P(\{X > u\} \cap \{X_j > u\}) + tP(X > u). \quad (3.2) \end{aligned}$$

Proof. Obviously,

$$0 \leq P(M_r \leq u) - P(M_{r+\ell} \leq u) \leq \sum_{j=0}^{\ell-1} (P(M_{r+j} \leq u) - P(M_{r+j+1} \leq u)) .$$

By the law of total probability we have, for any $i \geq 0$,

$$P(M_i \leq u) = P(M_{i+1} \leq u) + P(\{M_i \leq u\} \cap \{X_i > u\}) \leq P(M_{i+1} \leq u) + P(X_i > u)$$

and the first statement of the Lemma follows by stationarity.

For the second statement observe that

$$\{M_{s+t+m} \leq u\} = \{M_s \leq u\} \cap \{M_{s,t} \leq u\} \cap \{M_{s+t,m} \leq u\}.$$

Consequently,

$$(\{M_s \leq u\} \cap \{M_{s+t,m} \leq u\}) \setminus \{M_{s+t+m} \leq u\} \subset \{M_{s,t} > u\}.$$

Thus, using the first inequality of Lemma 3.1 we obtain

$$|P(\{M_s \leq u\} \cap \{M_{s+t,m} \leq u\}) - P(M_{s+t+m} \leq u)| \leq tP(X > u) .$$

Using stationarity and the first inequality in Lemma 3.1 we have

$$\begin{aligned} P(\{M_s \leq u\} \cap \{M_{s+t,m} \leq u\}) &= P(M_{s+t,m} \leq u) - P(\{M_s > u\} \cap \{M_{s+t,m} \leq u\}) \\ &= P(M_m \leq u) - P(\{M_s > u\} \cap \{M_{s+t,m} \leq u\}) \\ &\geq P(M_m \leq u) - \sum_{j=0}^{s-1} P(\{X_j > u\} \cap \{M_{s+t,m} \leq u\}). \end{aligned}$$

Now, by the second inequality in Lemma 3.1 and stationarity we have

$$\begin{aligned} P(\{M_s \leq u\} \cap \{M_{s+t,m} \leq u\}) &\leq P(M_m \leq u) - \sum_{j=0}^{s-1} P(\{X_j > u\} \cap \{M_{s+t,m} \leq u\}) \\ &\quad + \sum_{j=0}^{s-1} \sum_{j \neq \ell, \ell=0}^{s-1} P(\{X_j > u\} \cap \{X_\ell > u\} \cap \{M_{s+t,m} \leq u\}). \end{aligned}$$

Finally, stationarity and the last three inequalities give

$$\begin{aligned} \left| P(\{M_s \leq u\} \cap \{M_{s+t,m} \leq u\}) - P(M_m \leq u) + \sum_{j=0}^{s-1} P(\{X > u\} \cap \{M_{s+t-j,m} \leq u\}) \right| \\ \leq 2s \sum_{j=1}^{s-1} P(\{X > u\} \cap \{X_j > u\}), \end{aligned}$$

from which the result follows. \square

Proof of Theorem 1. Let $\ell = \ell_n = \lfloor \frac{n}{k} \rfloor$ where $\lfloor \frac{n}{k} \rfloor$ is the integer part of $\frac{n}{k}$. We begin by replacing $P(M_n \leq u_n)$ by $P(M_{k(\ell+t)} \leq u_n)$ for some $t > 1$. According to (3.1) of Lemma 3.2, we have

$$|P(M_n \leq u_n) - P(M_{k(\ell+t)} \leq u_n)| \leq ktP(X > u_n). \quad (3.3)$$

We now estimate recursively $P(M_{i(\ell+t)} \leq u_n)$ for $i = 0, \dots, k$. Using (3.2) of Lemma 3.2 and stationarity, we have for any $1 \leq i \leq k$

$$|P(M_{i(\ell+t)} \leq u_n) - (1 - \ell P(X > u_n))P(M_{(i-1)(\ell+t)} \leq u_n)| \leq \Gamma_{n,i},$$

where

$$\begin{aligned} \Gamma_{n,i} &= \left| \ell P(X > u_n)P(M_{(i-1)(\ell+t)} \leq u_n) - \sum_{j=0}^{\ell-1} P(\{X_j > u_n\} \cap \{M_{\ell+t, (i-1)(\ell+t)} \leq u_n\}) \right| \\ &\quad + tP(X > u_n) + 2\ell \sum_{j=1}^{\ell-1} P(\{X > u_n\} \cap \{X_j > u_n\}). \end{aligned}$$

Using stationarity, $D(u_n)$ and, in particular, that $\gamma(n, t)$ is nonincreasing in t for each n we conclude

$$\begin{aligned} \Gamma_{n,i} &\leq \sum_{j=0}^{\ell-1} |P(X_0 > u_n)P(M_{(i-1)(\ell+t)} \leq u_n) - P(\{X_0 > u_n\} \cap \{M_{\ell+t-j, (i-1)(\ell+t)} \leq u_n\})| \\ &\quad + tP(X > u_n) + 2\ell \sum_{j=1}^{\ell-1} P(\{X > u_n\} \cap \{X_j > u_n\}) \\ &\leq \ell\gamma(n, t) + tP(X > u_n) + 2\ell \sum_{j=1}^{\ell-1} P(\{X > u_n\} \cap \{X_j > u_n\}). \end{aligned}$$

Define $\Upsilon_n = \ell\gamma(n, t) + tP(X > u_n) + 2\ell \sum_{j=1}^{\ell-1} P(\{X > u_n\} \cap \{X_j > u_n\})$. Then for every $1 < i \leq k$ we have

$$\left| P(M_{i(\ell+t)} \leq u_n) - (1 - \ell P(X > u_n))P(M_{(i-1)(\ell+t)} \leq u_n) \right| \leq \Upsilon_n$$

and for $i = 1$

$$\left| P(M_{(\ell+t)} \leq u_n) - (1 - \ell P(X > u_n)) \right| \leq \Upsilon_n.$$

Assume that k and n are large enough in order to have $\ell P(X > u_n) < 2$, which implies that $|1 - \ell P(X > u_n)| < 1$. A simple inductive argument allows to conclude

$$\left| P(M_{k(\ell+t)} \leq u_n) - (1 - \ell P(X > u_n))^k \right| \leq k\Upsilon_n.$$

Recalling (3.3), we have

$$\left| P(M_n \leq u_n) - (1 - \ell P(X > u_n))^k \right| \leq ktP(X > u_n) + k\Upsilon_n. \quad (3.4)$$

Since $nP(X > u_n) = n(1 - F(u_n)) \rightarrow \tau$, as $n \rightarrow \infty$, for some $\tau \geq 0$, we have

$$\lim_{k \rightarrow \infty} \lim_{n \rightarrow \infty} (1 - [\frac{n}{k}]P(X > u_n))^k = \lim_{k \rightarrow \infty} (1 - \frac{\tau}{k})^k = e^{-\tau}.$$

Now, observe that $nP(X > u_n) = n(1 - F(u_n)) \rightarrow \tau$ is equivalent to $P(\hat{M}_n \leq u_n) = (F(u_n))^n \rightarrow e^{-\tau}$, where the limits are taken when $n \rightarrow \infty$ and $\tau \geq 0$ (see Leadbetter et al. (1983), Theorem 1.5.1] for a proof of this fact). Hence,

$$\lim_{k \rightarrow \infty} \lim_{n \rightarrow \infty} (1 - [\frac{n}{k}]P(X > u_n))^k = \lim_{n \rightarrow \infty} P(\hat{M}_n \leq u_n). \quad (3.5)$$

It is now clear that, according to (3.4) and (3.5), M_n and \hat{M}_n share the same limiting distribution if

$$\lim_{k \rightarrow \infty} \lim_{n \rightarrow \infty} (ktP(X > u_n) + k\Upsilon_n) = 0,$$

that is

$$\lim_{k \rightarrow \infty} \lim_{n \rightarrow \infty} 2ktP(X > u_n) + n\gamma(n, t) + 2n \sum_{j=1}^{\ell} P(\{X > u_n\} \cap \{X_j > u_n\}) = 0. \quad (3.6)$$

Assume that $t = t_n$ where $t_n = o(n)$ is given by Condition $D(u_n)$. Then, for every $k \in \mathbb{N}$, we have $\lim_{n \rightarrow \infty} kt_n P(X > u_n) = 0$, since $nP(X > u_n) \rightarrow \tau \geq 0$. Finally, we use $D(u_n)$ and $D'(u_n)$ to obtain that the two remaining terms in (3.6) also go to 0. \square

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