



HAL
open science

Least squares estimator for the parameter of the fractional Ornstein-Uhlenbeck sheet

Jorge Clarke de La Cerda, Ciprian A. Tudor

► **To cite this version:**

Jorge Clarke de La Cerda, Ciprian A. Tudor. Least squares estimator for the parameter of the fractional Ornstein-Uhlenbeck sheet. *Journal of the Korean Statistical Society*, 2012, 41, pp.341 - 350. hal-00619170

HAL Id: hal-00619170

<https://hal.science/hal-00619170>

Submitted on 5 Sep 2011

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

Least squares estimator for the parameter of the fractional Ornstein-Uhlenbeck sheet

Jorge Clarke De la Cerda ^{1, *} Ciprian A. Tudor ^{2, †}

¹ CI²MA, Departamento de Ingeniería Matemática, Universidad de Concepción, Casilla 160-C, Concepción, Chile.
jclarke@udec.cl

² Laboratoire Paul Painlevé, Université de Lille 1
F-59655 Villeneuve d'Ascq, France.
tudor@math.univ-lille1.fr

September 5, 2011

Abstract

We will study the least square estimator $\widehat{\theta}_{T,S}$ for the drift parameter θ of the fractional Ornstein-Uhlenbeck sheet which is defined as the solution of the Langevin equation

$$X_{t,s} = -\theta \int_0^t \int_0^s X_{v,u} dv du + B_{t,s}^{\alpha,\beta}, \quad (t, s) \in [0, T] \times [0, S].$$

driven by the fractional Brownian sheet $B^{\alpha,\beta}$ with Hurst parameters α, β in $(\frac{1}{2}, \frac{5}{8})$. Using the properties of multiple Wiener-Itô integrals we prove that the estimator is strongly consistent for the parameter θ . In contrast to the one-dimensional case, the estimator $\widehat{\theta}_{T,S}$ is not asymptotically normal.

2010 AMS Classification Numbers: 60G15, G0H07, 60G35, 62M40.

Key words: fractional Brownian sheet, parameter estimation, multiple Wiener-Itô integrals, strong consistency.

1 Introduction

We consider the two-parameter fractional Ornstein-Uhlenbeck process defined as the solution of the stochastic equation

*Partially supported by the MECESUP project UCO-0713 and the CONICYT-ECOS program C10E03.

†Associate member of the team Sann, Université de Paris 1 Panthéon-Sorbonne. Partially supported by the ANR grant "Mastérie" BLAN 012103.

$$X_{t,s} = -\theta \int_0^t \int_0^s X_{v,u} dv du + B_{t,s}^{\alpha,\beta}, \quad (t, s) \in [0, T] \times [0, S]. \quad (1)$$

Here $B^{\alpha,\beta}$ denotes a fractional Brownian sheet with Hurst parameters $\alpha, \beta \in (\frac{1}{2}, 1)$. We also suppose that $X_{0,0} = X_{t,0} = X_{0,s} = 0$ for every t, s . Our goal is to estimate the unknown parameter θ from the continuous time observation of the solution $(X_{t,s})_{(t,s) \in [0,T] \times [0,S]}$.

The development of the stochastic analysis for fractional Brownian motion (fBm) naturally led to the study of the statistical inference for stochastic equations driven by this process. There already exists an important literature related to these aspects. We refer, among others to [3], [4], [9], [13], [22]. Statistical analysis of the stochastic differential equations (SDE) driven by the fractional Brownian sheet has been less considered. We refer to the paper [18] for the study of the maximum likelihood estimator for a SDE with additive fractional Brownian sheet noise (see also [6] or [1] for the case when the noise is a standard Brownian sheet).

In this paper we propose a least square estimator for the unknown parameter θ following the approach in [9]. This estimator is obtained by formally minimizing with respect to θ the expression

$$\int_0^T \int_0^S \left| \frac{\partial^2}{\partial t \partial s} X_{t,s} + \theta X_{t,s} \right|^2 ds dt.$$

We obtain the following estimator

$$\widehat{\theta}_{T,S} = -\frac{\int_0^T \int_0^S X_{t,s} dX_{t,s}}{\int_0^T \int_0^S X_{t,s}^2 dt ds}. \quad (2)$$

The integral with respect to $dX_{t,s}$ is understood as the sum of the standard Lebesgue integral $-\theta \int_0^T \int_0^S dt ds X_{t,s}^2$ and of the stochastic integral $\int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta}$ which is a divergence type integral with respect to the fractional Brownian sheet $B^{\alpha,\beta}$ (it will be defined in Section 2, we also refer to [10], [11], [12], [20], [21] for the stochastic integration with respect to $B^{\alpha,\beta}$). Using (1) and (2) we can write

$$\widehat{\theta}_{T,S} - \theta = -\frac{\int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta}}{\int_0^T \int_0^S X_{t,s}^2 dt ds}. \quad (3)$$

We will study the asymptotic behavior of the least square estimator $\widehat{\theta}_{T,S}$ as $T, S \rightarrow \infty$. Our tools are the multiple stochastic integrals and the Malliavin calculus. Actually, the nominator and the denominator of the right hand side of (3) can be expressed as multiple integrals of order 2 with respect to the fractional Brownian sheet and from this we will obtain concrete estimates for their moments. We will prove that the estimator (2) is a strongly consistent estimator in the sense that it converges almost surely to the true value of the parameter θ . This result is similar to the one-dimensional case (see [9]), however the approach presented in [9] is not possible to be followed for the two-parameter case, instead

we use among other tools, the hypercontractivity of multiples integrals. By contrary, in the two-parameter case, the least square estimator does not preserve the asymptotic normality as in the one-parameter case. This will be noticed at the end of our work by using criteria for the asymptotic normality of sequences of multiple integrals in terms of Malliavin calculus.

Our paper is structured as follows. Section 2 contains some preliminaries on multiple integrals and fractional Brownian sheet. In Section 3 we discuss the relation between the solution to (1) and the Bessel function of order 0. Section 4 contains the proof of the consistency of the least square estimator while Section 5 is devoted to a discussion about the asymptotic normality of the estimator.

2 Preliminaries

Let us introduce the elements from stochastic analysis that we will need in the paper. Consider \mathcal{H} a real separable Hilbert space and $(B(\varphi), \varphi \in \mathcal{H})$ an isonormal Gaussian process on a probability space (Ω, \mathcal{A}, P) , that is, a centered Gaussian family of random variables such that $\mathbf{E}(B(\varphi)B(\psi)) = \langle \varphi, \psi \rangle_{\mathcal{H}}$. Denote by I_n the multiple stochastic integral with respect to B (see [16]). This I_n is actually an isometry between the Hilbert space $\mathcal{H}^{\odot n}$ (symmetric tensor product) equipped with the scaled norm $\sqrt{n!} \|\cdot\|_{\mathcal{H}^{\otimes n}}$ and the Wiener chaos of order n which is defined as the closed linear span of the random variables $H_n(B(\varphi))$ where $\varphi \in \mathcal{H}$, $\|\varphi\|_{\mathcal{H}} = 1$ and H_n is the Hermite polynomial of degree $n \geq 1$

$$H_n(x) = \frac{(-1)^n}{n!} \exp\left(\frac{x^2}{2}\right) \frac{d^n}{dx^n} \left(\exp\left(-\frac{x^2}{2}\right) \right), \quad x \in \mathbb{R}.$$

The isometry of multiple integrals can be written as: for m, n positive integers,

$$\begin{aligned} \mathbf{E}(I_n(f)I_m(g)) &= n! \langle f, g \rangle_{\mathcal{H}^{\otimes n}} \quad \text{if } m = n, \\ \mathbf{E}(I_n(f)I_m(g)) &= 0 \quad \text{if } m \neq n. \end{aligned} \tag{4}$$

It also holds that

$$I_n(f) = I_n(\tilde{f})$$

where \tilde{f} denotes the symmetrization of f defined by

$$\tilde{f}(x_1, \dots, x_n) = \frac{1}{n!} \sum_{\sigma \in \mathcal{S}_n} f(x_{\sigma(1)}, \dots, x_{\sigma(n)}).$$

The Malliavin derivative acts on multiple integrals $F = I_n(f)$ in the following way: for every s

$$D_s I_n = n I_{n-1}(f(\cdot, s))$$

where " \cdot " above denotes $n-1$ variables. We recall the following hypercontractivity property for the L^p norm of a multiple stochastic integral (see [14, Theorem 4.1])

$$\mathbf{E} |I_m(f)|^{2m} \leq c_m (\mathbf{E} I_m(f)^2)^m \tag{5}$$

where c_m is an explicit positive constant and $f \in \mathcal{H}^{\otimes m}$.

In this work we use Malliavin calculus and multiple integrals with respect to the fractional Brownian sheet (fBs). Let us define this process and its associated Hilbert space. The fBs with Hurst parameters $\alpha, \beta \in (0, 1)$, $(B_{t,s}^{\alpha,\beta}, t, s \in [0, T] \times [0, S])$ is a zero mean Gaussian process with covariance

$$\begin{aligned} \mathbf{E} \left(B_{t,s}^{\alpha,\beta}, B_{u,v}^{\alpha,\beta} \right) &= \mathcal{R}^\alpha(t, u) \mathcal{R}^\beta(s, v) \\ &:= \frac{1}{2} (t^{2\alpha} + u^{2\alpha} - |t - u|^{2\alpha}) \frac{1}{2} (s^{2\beta} + v^{2\beta} - |s - v|^{2\beta}) \end{aligned} \quad (6)$$

given for all $t, u \in [0, T]^2$ and $s, v \in [0, S]^2$.

We assume that $B^{\alpha,\beta}$ is defined on a complete probability space $(\Omega, \mathcal{A}, \mathbb{P})$ such that \mathcal{A} is generated by $B^{\alpha,\beta}$. Fix a time interval $[0, T] \times [0, S]$, denote by ξ the set of real valued step functions on $[0, T] \times [0, S]$ and let $\mathcal{H}^{\alpha,\beta}$ be the Hilbert space defined as the closure of ξ with respect to the scalar product

$$\langle 1_{[0,t] \times [0,s]}, 1_{[0,u] \times [0,v]} \rangle_{\mathcal{H}^{\alpha,\beta}} = \mathcal{R}^\alpha(t, u) \mathcal{R}^\beta(s, v)$$

where $\mathcal{R}^\alpha(t, u) \mathcal{R}^\beta(s, v)$ is the covariance function of the fBs, given in (6). The mapping $1_{[0,t] \times [0,s]} \mapsto B_{t,s}^{\alpha,\beta}$ can be extended to a linear isometry between $\mathcal{H}^{\alpha,\beta}$ and the Gaussian space $\mathcal{H}_1^{\alpha,\beta}$ spanned by $B^{\alpha,\beta}$ which is a closed subspace of $L^2(\Omega, \mathcal{A}, \mathbb{P})$. We denote this isometry by $\varphi \mapsto B^{\alpha,\beta}(\varphi)$. Fix $\alpha, \beta > \frac{1}{2}$, in this case we have that for every $f, g \in \mathcal{H}^{\alpha,\beta}$ the scalar product has the form

$$\langle f, g \rangle_{\mathcal{H}^{\alpha,\beta}} = c(\alpha)c(\beta) \int_0^T \int_0^S \int_0^T \int_0^S f(a, b)g(m, n) |a - m|^{2\alpha-2} |b - n|^{2\beta-2} da db dm dn \quad (7)$$

and $c(\alpha) = \alpha(2\alpha - 1)$.

3 About the solution

The equation (1) has been studied in several papers (see [8], [15]). It has been showed that for $\theta > 0$ and $\alpha, \beta > \frac{1}{2}$ equation (1) admits an unique strong solution which can be expressed as

$$X_{t,s} = \int_0^T \int_0^S f(t, s, t_0, s_0) dB_{t_0, s_0}^{\alpha,\beta} \quad (8)$$

where

$$f(t, s, t_0, s_0) = 1_{[0,t]}(t_0) 1_{[0,s]}(s_0) \sum_{n \geq 0} (-1)^n \theta^n \frac{(t - t_0)^n (s - s_0)^n}{(n!)^2}. \quad (9)$$

We will call the process X solution to (1) as the fractional Ornstein-Uhlenbeck sheet. It is a Gaussian process since it is given by a multiple integral of order 1 (Wiener integral

actually) with respect to the Gaussian process $B^{\alpha,\beta}$. We mention that the solution to (1) behaves differently as its one-dimensional counterpart which is fractional Ornstein-Uhlenbeck process introduced in [5], this will make our analysis quite different from [9]. For example, we note that the solution of some stochastic differential equations driven by the Brownian sheet or fractional Brownian sheet, which are positive in the one-parameter case, can take negative values with strictly positive probability (see [15] or [16]).

A key element of our analysis is the fact that the solution X (more precisely the kernel f of the solution) can be expressed in terms of the Bessel function of the first kind. Let us consider the Bessel function of order 0 given, for every $x \in \mathbb{R}$, by

$$J_0(x) = \sum_{n \geq 0} \frac{(-1)^n}{n!^2} \left(\frac{x}{2}\right)^{2n}$$

This Bessel function admits the integral representation, for every $x \in \mathbb{R}$

$$J_0(x) = \frac{1}{\pi} \int_0^\pi \cos(x \sin \rho) d\rho.$$

The kernel f in (9) of the solution $(X_{t,s})_{t,s \in [0,T] \times [0,S]}$ can be expressed as

$$\begin{aligned} f(t, s, u, v) &= 1_{[0,t]}(u) 1_{[0,s]}(v) J_0\left(2\sqrt{\theta(t-u)(s-v)}\right) \\ &= 1_{[0,t]}(u) 1_{[0,s]}(v) \frac{1}{\pi} \int_0^\pi \cos\left(2\sqrt{\theta(t-u)(s-v)} \sin \rho\right) d\rho. \end{aligned} \quad (10)$$

Let us also recall the following property of the Bessel function (see e.g. [2]) which will play an important role for our estimates: for x large enough

$$J_0(x) \sim \sqrt{\frac{2}{\pi x}} \cos\left(x - \frac{\pi}{4}\right) \quad (11)$$

(the symbol \sim means that the two sides have the same limit as $x \rightarrow \infty$).

4 Asymptotic behavior of the least square estimator

In this section we study the asymptotic behavior of the estimator $\widehat{\theta}_{T,S}$ defined in (2). More precisely, we will show that this estimator is strongly consistent for the parameter θ , that is, $\widehat{\theta}_{T,S}$ converges to θ almost surely as $T, S \rightarrow \infty$. To this end we will analyze separately the nominator and the denominator appearing in the right hand side of the expression (3). Let us start with the study of the nominator. It can be written as the double stochastic integral

$$\int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta} := F_{T,S} := I_2(f(u, v, t, s)) \quad (12)$$

where the kernel f is given by (9) and the integral I_2 acts with respect to the variables $(u, v), (t, s)$.

We will estimate first the L^2 norm of $F_{T,S}$. We have the following result.

Proposition 1 For every $\varepsilon > 0$ and for $\alpha, \beta \in (\frac{1}{2}, \frac{5}{8})$,

$$\mathbf{E} \left(T^{-2\alpha + \frac{1}{4} - \varepsilon} S^{-2\beta + \frac{1}{4} - \varepsilon} \int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta} \right)^2 \rightarrow 0 \text{ when } T, S \rightarrow \infty. \quad (13)$$

Moreover for T, S large enough we have

$$\mathbf{E} \left(T^{-2\alpha + \frac{1}{4}} S^{-2\beta + \frac{1}{4}} \int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta} \right)^2 < C$$

where C is a strictly positive constant not depending on T, S .

Proof: We calculate the L^2 norm of the random variable $I_2(f(u, v, t, s))$. By the isometry property of multiple integrals (4) and since $\|\tilde{f}\|_{(\mathcal{H}^{\alpha,\beta})^{\otimes 2}} \leq \|f\|_{(\mathcal{H}^{\alpha,\beta})^{\otimes 2}}$ this norm can be handles as follows

$$\begin{aligned} I_{T,S} &\leq \int_{[0,T]^4} dt dt_0 du du_0 \int_{[0,S]^4} ds ds_0 dv dv_0 \\ &\quad \times f(t, s, u, v) f(t_0, s_0, u_0, v_0) |u - u_0|^{2\alpha-2} |v - v_0|^{2\beta-2} |t - t_0|^{2\alpha-2} |s - s_0|^{2\beta-2} \\ &= \int_0^T dt \int_0^t du \int_0^T dt_0 \int_0^{t_0} du_0 \int_0^S ds \int_0^s dv \int_0^S ds_0 \int_0^{s_0} dv_0 \\ &\quad \times J_0 \left(2\sqrt{\theta(t-u)(s-v)} \right) J_0 \left(2\sqrt{\theta(t_0-u_0)(s_0-v_0)} \right) \\ &\quad \times |t - t_0|^{2\alpha-2} |u - u_0|^{2\alpha-2} |s - s_0|^{2\beta-2} |v - v_0|^{2\beta-2}. \end{aligned}$$

By making the change of variables $\tilde{t} = \frac{t}{T}$, $\tilde{u} = \frac{u}{T}$ and similarly for the other variables, we obtain

$$\begin{aligned} I_{T,S} &= T^{4\alpha-4} T^4 S^{4\beta-4} S^4 \int_0^1 dt \int_0^t du \int_0^1 dt_0 \int_0^{t_0} du_0 \int_0^1 ds \int_0^s dv \int_0^1 ds_0 \int_0^{s_0} dv_0 \\ &\quad \times J_0 \left(2\sqrt{\theta(t-u)(s-v)TS} \right) J_0 \left(2\sqrt{\theta(t_0-u_0)(s_0-v_0)TS} \right) \\ &\quad \times |t - t_0|^{2\alpha-2} |u - u_0|^{2\alpha-2} |s - s_0|^{2\beta-2} |v - v_0|^{2\beta-2} \\ &:= T^{4\alpha} S^{4\beta} U_{T,S}. \end{aligned} \quad (14)$$

Using the asymptotic behavior of the Bessel function (11), we have that

$$\frac{J_0 \left(2\sqrt{\theta(t-u)(s-v)TS} \right)}{(TS)^{-\frac{1}{4} + \varepsilon}} \xrightarrow{T,S \rightarrow \infty} 0$$

for almost every $t, u, s, v \in (0, 1)$ and for every $\varepsilon > 0$. We will next apply the dominated convergence theorem. To this end, using again relation (11), it suffices to show that the

integral

$$\begin{aligned}
I &= \int_0^1 dt \int_0^t du \int_0^1 dt_0 \int_0^{t_0} du_0 \int_0^1 ds \int_0^s dv \int_0^1 ds_0 \int_0^{s_0} dv_0 \\
&\quad ((t-u)(s-v)(t_0-u_0)(s_0-v_0))^{-\frac{1}{4}} \\
&\quad |t-t_0|^{2\alpha-2} |u-u_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |v-v_0|^{2\beta-2}
\end{aligned} \tag{15}$$

is finite. This is proved in the following lemma. ■

Remark 1 *In the above statement we can replace the normalization $T^{-2\alpha+\frac{1}{4}-\varepsilon} S^{-2\beta+\frac{1}{4}-\varepsilon}$ by $T^{-2\alpha+\frac{1}{4}} S^{-2\beta+\frac{1}{4}} f(T, S)$ where $f(T, S)$ is a deterministic function which converges to zero as $T, S \rightarrow \infty$. This is a consequence of the proof below.*

Lemma 1 *Let I be given by (15). Then for $\frac{1}{2} < \alpha, \beta < \frac{5}{8}$ the integral I is finite.*

Proof: Consider the integral

$$\begin{aligned}
&\int_0^{s_0} dv_0 ((t-u)(s-v))^{-\frac{1}{4}} ((t_0-u_0)(s_0-v_0))^{-\frac{1}{4}} \\
&\quad \times |t-t_0|^{2\alpha-2} |u-u_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |v-v_0|^{2\beta-2} \\
&= ((t-u)(s-v)(t_0-u_0))^{-\frac{1}{4}} |t-t_0|^{2\alpha-2} |u-u_0|^{2\alpha-2} |s-s_0|^{2\beta-2} \\
&\quad \times \int_0^{s_0} dv_0 |v-v_0|^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}}.
\end{aligned}$$

If $s_0 < v$ we have

$$\int_0^{s_0} dv_0 |v-v_0|^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}} = \int_0^{s_0} dv_0 (v-v_0)^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}}$$

making the change of variables $z = \frac{s_0-v_0}{v-v_0}$ we get

$$\begin{aligned}
&\int_0^{\frac{s_0}{v}} \left(\frac{v-s_0}{1-z} \right)^{2\beta-2} \left(\frac{z(v-s_0)}{1-z} \right)^{-\frac{1}{4}} \left(\frac{v-s_0}{(1-z)^2} \right) dz \\
&\leq (v-s_0)^{2\beta-\frac{5}{4}} \int_0^1 z^{-\frac{1}{4}} (1-z)^{\frac{1}{4}-2\beta} dz = (v-s_0)^{2\beta-\frac{5}{4}} \tilde{\beta} \left(\frac{3}{4}, \frac{5}{4} - 2\beta \right)
\end{aligned}$$

where $\tilde{\beta}$ is the *Beta* function. The expression above is finite for $\beta < \frac{5}{8}$.

Now, if $v \leq s_0$

$$\begin{aligned}
&\int_0^{s_0} dv_0 |v-v_0|^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}} \\
&= \int_0^v dv_0 (v-v_0)^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}} + \int_v^{s_0} dv_0 (v_0-v)^{2\beta-2} (s_0-v_0)^{-\frac{1}{4}}.
\end{aligned}$$

For the first integral in the right hand side we make the change of variables $z = \frac{v-v_0}{s_0-v_0}$, and for the second one we make $z = \frac{s_0-v_0}{v_0-v}$. Then we get

$$\begin{aligned} &\leq (s_0 - v)^{2\beta - \frac{5}{4}} \left(\int_0^1 z^{2\beta-2} (1-z)^{\frac{1}{4}-2\beta} dz + \int_0^\infty z^{-\frac{1}{4}} (1+z)^{\frac{1}{4}-2\beta} dz \right) \\ &= (s_0 - v)^{2\beta - \frac{5}{4}} \left(\tilde{\beta} \left(2\beta - 1, \frac{5}{4} - 2\beta \right) + \frac{\Gamma(\frac{3}{4})\Gamma(2\beta - 1)}{\Gamma(2\beta - \frac{1}{4})} {}_2F_1(0, 3/4, 2\beta - 1/4; 0) \right) \\ &= (s_0 - v)^{2\beta - \frac{5}{4}} \left(\tilde{\beta} \left(2\beta - 1, \frac{5}{4} - 2\beta \right) + \tilde{\beta} \left(\frac{3}{4}, 2\beta - 1 \right) {}_2F_1(0, 3/4, 2\beta - 1/4; 0) \right) \end{aligned}$$

wich is finite for $\frac{1}{2} < \alpha, \beta < \frac{5}{8}$. Here Γ is the *Gamma* function, ${}_2F_1$ is the *Hypergeometric* function, and we have make use of the property $\tilde{\beta}(x, y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}$ and the formula (see [19], formula 1.6.7)

$${}_2F_1(a, b, c; 1-x) = \frac{\Gamma(c)}{\Gamma(b)\Gamma(c-b)} \int_0^\infty w^{b-1} (1+w)^{a-c} (1+wx)^{-a} dw.$$

Proceeding in a similar way with the other seven integrals we conclude that I is finite. ■

Remark 2 *We may note that the prescence of the cos function in the asymptotic behavior of the Bessel function does not allow to obtain a renormalization for $F_{T,S}$ in terms of the powers of T and S like in [9].*

The following proposition is a consequence of the Proposition 1 and of the hypercontractivity property of multiple stochastic integrals.

Proposition 2 *For every $\varepsilon > 0$ and for $\frac{1}{2} < \alpha, \beta < \frac{5}{8}$ the sequence*

$$\frac{1}{\sqrt{T^{4\alpha - \frac{1}{2} + \varepsilon} S^{4\beta - \frac{1}{2} + \varepsilon}}} \int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta}$$

converges to zero a.s. as $T, S \rightarrow \infty$.

Proof: As in [9] we can replace the couple (T, S) by a discrete sequence (T_M, S_N) such that T_M, S_N converge to infinity as $M, N \rightarrow \infty$. This is possible since the nominator and denominator in (3) are continuous a.s. with respect to (T, S) . Indeed, the fact that the divergence integral in the nominator is continuous follows from [16], page 293 since the integrand X is regular enough, and the integral $dt ds$ in the nominator is clearly continuous a.s. with respect to the couple (T, S) . For simplicity, we will assume that $(T_M, S_N) = (M, N)$. Let us show that

$$A_{M,N} := M^{-2\alpha + \frac{1}{4} - \varepsilon} N^{-2\beta + \frac{1}{4} - \varepsilon} \int_0^M \int_0^N X_{t,s} dB_{t,s}^{\alpha,\beta}$$

converges to zero a.s. as M, N tend to infinity. We will use the Borel-Cantelli lemma. To do this, we will estimate $P(A_{M,N} > (MN)^{-\gamma})$ for some $\gamma > 0$. For every $p \geq 1$ we have

$$P(A_{M,N} > (MN)^{-\gamma}) \leq (MN)^{p\gamma} \mathbf{E}|A_{M,N}|^p$$

and since $A_{M,N}$ is a multiple integral in the second Wiener chaos, the inequality (5) and Proposition 1 implies that

$$\mathbf{E}|A_{M,N}|^p \leq c(p) (\mathbf{E}A_{M,N}^2)^{\frac{p}{2}} \leq c(p, \alpha, \beta)(MN)^{-\varepsilon p}.$$

Putting together the two above bounds, we get

$$\sum_{M>M_0, N>N_0} P(A_{M,N} > (MN)^{-\gamma}) \leq c(p, \alpha, \beta) \sum_{M>M_0, N>N_0} (MN)^{p(\gamma-\varepsilon)}$$

and this series is convergent when

$$(\varepsilon - \gamma)p > 1 \quad \text{or equivalently} \quad \gamma < \varepsilon - \frac{1}{p}.$$

For every given $\varepsilon > 0$ and for p large enough we can always chose a real number γ such that $0 < \gamma < \varepsilon - \frac{1}{p}$. This, together with the Borel-Cantelli lemma allows us to finish the proof. \blacksquare

The next step is to analyze the denominator in formula (3). We have the following estimate.

Proposition 3 *For any $\varepsilon > 0$ and for any $\alpha, \beta \in (\frac{1}{2}, \frac{5}{8})$*

$$\frac{1}{T^{2\alpha+\frac{1}{2}-\varepsilon} S^{2\beta+\frac{1}{2}-\varepsilon}} \mathbf{E} \int_0^T dt \int_0^S ds X_{t,s}^2 \rightarrow_{T,S \rightarrow \infty} \infty.$$

Proof: By the isometry of multiple integrals (4), the equation (7) and the expression of the solution to (1)

$$\begin{aligned} \mathbf{E}X_{t,s}^2 &= c(\alpha, \beta) \int_0^t du \int_0^t du_0 \int_0^s dv \int_0^s dv_0 \\ &\quad \times J_0(2\sqrt{\theta(t-u)(s-v)}) J_0(2\sqrt{\theta(t-u_0)(s-v_0)}) |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2}, \end{aligned}$$

making the change of variables $\tilde{t} = \frac{t}{T}, \tilde{s} = \frac{s}{S}$ and similar for the other variables we can write

$$\begin{aligned} \int_0^T \int_0^S ds dt \mathbf{E}X_{t,s}^2 &= c(\alpha, \beta) T^{2\alpha+1} S^{2\beta+1} \int_0^1 dt \int_0^1 ds \int_0^t du \int_0^t du_0 \int_0^s dv \int_0^s dv_0 \\ &\quad \times J_0(2\sqrt{\theta(t-u)(s-v)TS}) J_0(2\sqrt{\theta(t-u_0)(s-v_0)TS}) \\ &\quad \times |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \end{aligned}$$

and thus

$$\begin{aligned} \frac{\int_0^T \int_0^S ds dt \mathbf{E} X_{t,s}^2}{T^{2\alpha+\frac{1}{2}-\varepsilon} S^{2\beta+\frac{1}{2}-\varepsilon}} &= c(\alpha, \beta) \int_0^1 dt \int_0^1 ds \int_0^t du \int_0^t du_0 \int_0^s dv \int_0^s dv_0 \\ &\times \frac{J_0(2\sqrt{\theta(t-u)(s-v)TS})}{(TS)^{-\frac{1}{4}-\frac{\varepsilon}{2}}} \frac{J_0(2\sqrt{\theta(t-u_0)(s-v_0)TS})}{(TS)^{-\frac{1}{4}-\frac{\varepsilon}{2}}} \\ &\times |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2}. \end{aligned}$$

We will use the same idea as in the proof of Proposition 1. We notice first that

$$\frac{J_0(2\sqrt{\theta(t-u)(s-v)TS})}{(TS)^{-\frac{1}{4}-\frac{\varepsilon}{2}}}$$

converges to infinity as $T, S \rightarrow \infty$ by using (11); also for T, S large enough

$$\frac{J_0(2\sqrt{\theta(t-u)(s-v)TS})}{(TS)^{-\frac{1}{4}}}$$

is bounded by $c(|t-u||s-v|)^{-\frac{1}{4}}$ almost everywhere s, t, u, v . Also we note that the integral

$$\begin{aligned} &\int_0^1 dt \int_0^1 ds \int_0^t du \int_0^t du_0 \int_0^s dv \int_0^s dv_0 |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\ &(|t-u||s-v|)^{-\frac{1}{4}} (|t-u_0||s-v_0|)^{-\frac{1}{4}} \end{aligned}$$

is finite for $\alpha, \beta \in (\frac{1}{2}, \frac{5}{8})$ by using the same computations as in the proof of Lemma 1.

Using Fatou's lemma (we use the following version of the Fatou's lemma: if f_n is a sequence of functions such that $f_n \geq -g$ where g is positive and integrable, then $\underline{\lim} \int f_n \geq \int \underline{\lim} f_n$) this implies that

$$\begin{aligned} \underline{\lim}_{T,S \rightarrow \infty} \frac{\int_0^T \int_0^S ds dt \mathbf{E} X_{t,s}^2}{T^{2\alpha+\frac{1}{2}-\varepsilon} S^{2\beta+\frac{1}{2}-\varepsilon}} &\geq c(\alpha, \beta) \int_0^1 dt \int_0^1 ds \int_0^t du \int_0^t du_0 \int_0^s dv \int_0^s dv_0 \\ &\times \underline{\lim}_{T,S \rightarrow \infty} \frac{J_0(2\sqrt{\theta(t-u)(s-v)TS})}{(TS)^{-\frac{1}{4}-\frac{\varepsilon}{2}}} \frac{J_0(2\sqrt{\theta(t-u_0)(s-v_0)TS})}{(TS)^{-\frac{1}{4}-\frac{\varepsilon}{2}}} \\ &\times |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} = \infty. \end{aligned}$$

■

At this point we will need the following auxiliary lemma.

Lemma 2 Consider a sequence of random variables $(A_N)_N$ such that $\sum_N P(A_N > cN^{-\gamma}) < \infty$ for some $\gamma > 0$ and for every $c > 0$ (which implies $A_N \rightarrow 0$ almost surely as $N \rightarrow \infty$.) Also consider a sequence of a.s. strictly positive random variables $(B_N)_N$ such that $\mathbf{E} B_N \rightarrow \infty$ as $N \rightarrow \infty$. Then

$$\frac{A_N}{B_N} \rightarrow_{N \rightarrow \infty} 0 \text{ almost surely}$$

Proof: We will use again the Borel-Cantelli lemma. Let $C > 0$ be arbitrary. Then

$$\begin{aligned} \sum_N P\left(\frac{A_N}{B_N} > N^{-\gamma}\right) &= \sum_N P\left(\frac{A_N}{B_N} > N^{-\gamma}, B_N > C\right) + \sum_N P\left(\frac{A_N}{B_N} > N^{-\gamma}, B_N < C\right) \\ &\leq \sum_N P\left(\frac{A_N}{C} > N^{-\gamma}, B_N > C\right) + \sum_N P\left(\frac{A_N}{B_N} > N^{-\gamma}, B_N < C\right). \end{aligned}$$

Using the fact that $\mathbf{E}B_N \rightarrow \infty$ as $N \rightarrow \infty$ we obtain that

$$P\left(\frac{A_N}{B_N} > N^{-\gamma}, B_N < C\right) = 0$$

for N large enough. By assumption, $\sum_N P\left(\frac{A_N}{C} > N^{-\gamma}, B_N > C\right) < \infty$. Therefore

$$\sum_N P\left(\frac{A_N}{B_N} > N^{-\gamma}\right) < \infty$$

and therefore the conclusion follows. ■

Remark 3 Lemma 2 can be extended without difficulty to two-parameter sequences. That is, if $A_{M,N}$ is a sequence of random variables such that $\sum_{M,N \geq 0} P(A_{M,N} > c(MN)^{-\gamma}) < \infty$ and $B_{M,N}$ is a sequence of positive random variables such that $\mathbf{E}B_{M,N} \rightarrow_{M,N \rightarrow \infty} \infty$ then

$$\frac{A_{M,N}}{B_{M,N}} \rightarrow_{M,N \rightarrow \infty} 0 \text{ almost surely.}$$

Let us state the main result of this section.

Theorem 1 Let $\theta_{T,S}$ be the estimator given by (2). Suppose that $\alpha, \beta \in (\frac{1}{2}, \frac{5}{8})$. Then $\theta_{T,S}$ is a strongly consistent estimator for the parameter θ , that is,

$$\widehat{\theta}_{T,S} \rightarrow \theta \text{ almost surely as } T, S \rightarrow \infty.$$

Proof: From relation (3), the difference between the estimator and the true parameter is

$$\begin{aligned} \widehat{\theta}_T - \theta &= \frac{T^{-2\alpha + \frac{1}{4} - \varepsilon} S^{-2\beta + \frac{1}{4} - \varepsilon} \int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta}}{T^{-2\alpha + \frac{1}{4} - \varepsilon} S^{-2\beta + \frac{1}{4} - \varepsilon} \int_0^T dt \int_0^S ds X_{t,s}^2} \\ &= \frac{T^{-2\alpha + \frac{1}{4} - \varepsilon} S^{-2\beta + \frac{1}{4} - \varepsilon} \int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta}}{T^{\frac{3}{4} - 2\varepsilon} S^{\frac{3}{4} - 2\varepsilon} T^{-2\alpha - \frac{1}{2} + \varepsilon} S^{-2\beta - \frac{1}{2} + \varepsilon} \int_0^T dt \int_0^S ds X_{t,s}^2}. \end{aligned}$$

The result is obtained by using Proposition 2, Proposition 3 and Lemma 2 (and the remark that follows after this lemma). ■

5 Asymptotic non-normality of the estimator

We proved in the previous section that the least square estimator $\widehat{\theta}_{T,S}$ is strongly consistent. This property has been proved in the one-dimensional case in ([9]). Nevertheless, we show in this paragraph that the limiting distribution of the estimator is not the same in the one-parameter and two-parameter cases. This different behavior is somehow expected since the fractional Ornstein-Uhlenbeck sheet does not keep the properties of the fractional Ornstein-Uhlenbeck process (for example, the kernel f given by (9) can take any real value while in the one parameter case it is positive since it is given by an exponential function). In order to notice the difference between the one-parameter and the two-parameter case, we will focus only on the nominator in the right hand side of (3). To check the asymptotic normality we use the following criterium: (see Theorem 4 in [17]).

Theorem 2 *Let $(F_k, k \geq 1)$, $F_k = I_n(f_k)$ (with $f_k \in \mathcal{H}^{\odot n}$ for every $k \geq 1$) be a sequence of square integrable random variables in the n th Wiener chaos such that $\mathbf{E}[F_k^2] \rightarrow 1$ as $k \rightarrow \infty$. Then the following are equivalent:*

- i) The sequence $(F_k)_{k \geq 0}$ converges in distribution to the normal law $\mathcal{N}(0, 1)$.*
- ii) $\|DF_k\|_{\mathcal{H}}^2$ converges to n in $L^2(\Omega)$ as $k \rightarrow \infty$.*

Denote by

$$\sigma_{T,S}^2 = \mathbf{E} \left(\int_0^T \int_0^S X_{t,s} dB_{t,s}^{\alpha,\beta} \right)^2 = \mathbf{E} (F_{T,S})^2.$$

Proposition 4 *Assume $\frac{1}{2} < \alpha, \beta < \frac{5}{8}$ and let $F_{T,S}$ be given by (12). Then when T, S tend to infinity,*

$$\frac{1}{\sigma_{T,S}} F_{T,S}$$

does not converges in distribution to the normal law $N(0, 1)$.

Proof: We will prove that $\|D_{\frac{1}{\sigma_{T,S}}} F_{T,S}\|_{\mathcal{H}^{\alpha,\beta}}^2$ does not converges to 2 in $L^2(\Omega)$ as $T, S \rightarrow \infty$.

Let us denote by \tilde{f} the symmetrization of f with respect to the variables $(t, s), (u, v)$

$$\tilde{f}((t, s), (u, v)) = \frac{1}{2} (f(t, s, u, v) + f(u, v, t, s)).$$

We have

$$D_{t,s} \frac{1}{\sigma_{T,S}} F_{T,S} = \frac{1}{\sigma_{T,S}} D_{t,s} I_2(f) = \frac{1}{\sigma_{T,S}} D_{t,s} I_2(\tilde{f}) = \frac{1}{\sigma_{T,S}} (I_1(f(\cdot, \cdot, t, s)) + I_1(t, s, \cdot, \cdot)).$$

Here " \cdot, \cdot " represents the variable with respect to which the integral I_1 acts.

We obtain

$$\begin{aligned}
\|D\frac{1}{\sigma_{T,S}}F_{T,S}\|_{\mathcal{H}^{\alpha,\beta}}^2 &= \frac{1}{\sigma_{T,S}^2}c(\alpha)c(\beta)\int_0^T dt\int_0^S ds\int_0^T dt_0\int_0^S ds_0 \\
&\quad \times [I_1(f(\cdot,\cdot,t,s)) + I_1(f(t,s,\cdot,\cdot))] [I_1(f(\cdot,\cdot,t_0,s_0)) + I_1(f(t_0,s_0,\cdot,\cdot))] \\
&\quad \times |t-t_0|^{2\alpha-2}|s-s_0|^{2\beta-2} \\
&= \frac{1}{\sigma_{T,S}^2}c(\alpha)c(\beta)\int_0^T dt\int_0^S ds\int_0^T dt_0\int_0^S ds_0 \\
&\quad \times [I_2(f(\cdot,\cdot,t,s)\otimes f(\cdot,\cdot,t_0,s_0)) + I_2(f(\cdot,\cdot,t,s)\otimes f(t_0,s_0,\cdot,\cdot))] \\
&\quad + [I_2(f(t,s,\cdot,\cdot)\otimes f(\cdot,\cdot,t_0,s_0)) + I_2(f(t,s,\cdot,\cdot)\otimes f(t_0,s_0,\cdot,\cdot))] \\
&\quad \times |t-t_0|^{2\alpha-2}|s-s_0|^{2\beta-2} + \mathbf{E}\|D\frac{1}{\sigma_{T,S}}F_{T,S}\|_{\mathcal{H}^{\alpha,\beta}}^2 \\
&:= A_{T,S}^{(1)} + A_{T,S}^{(2)} + A_{T,S}^{(3)} + A_{T,S}^{(4)} + \mathbf{E}\|D\frac{1}{\sigma_{T,S}}F_{T,S}\|_{\mathcal{H}^{\alpha,\beta}}^2
\end{aligned}$$

Let us note that that $\mathbf{E}\|D\frac{1}{\sigma_{T,S}}F_{T,S}\|_{\mathcal{H}^{\alpha,\beta}}^2$ is equal to 2. This follows from the fact that for any multiple integral of order n we have

$$\mathbf{E}\|DI_n(f)\|_{\mathcal{H}}^2 = n\mathbf{E}I_n(f)^2.$$

It remains to show that the terms containing multiple integrals of order 2 does not converges to zero in $L^2(\Omega)$ as $T, S \rightarrow \infty$. The first and the fourth summand are similar, as they are the second one and the third one. We will handle only the first summand denoted by $A_{T,S}^{(1)}$ (because the other three terms can be studied analogously). We can write, using the definition of the scalar product in the Hilbert space $(\mathcal{H}^{\alpha,\beta})^{\otimes 2}$ and the expression (10) of the kernel f in terms of the Bessel function J_0

$$\begin{aligned}
\mathbf{E} \left| A_{T,S}^{(1)} \right|^2 &= \frac{1}{\sigma_{T,S}^4} \mathbf{E} \left(\int_0^T dt \int_0^S ds \int_0^T dt_0 \int_0^S ds_0 |t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} \right. \\
&\quad \left. I_2(f(\cdot, \cdot, t, s) \otimes f(\cdot, \cdot, t_0, s_0)) \right)^2 \\
&\sim 2 \int_0^T dt \int_0^S ds \int_0^T dt_0 \int_0^S ds_0 \int_0^T du \int_0^S dv \int_0^T du_0 \int_0^S dv_0 \\
&\quad |t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\
&\quad \langle f(\cdot, \cdot, t, s) \otimes f(\cdot, \cdot, t_0, s_0), f(\cdot, \cdot, u, v) \otimes f(\cdot, \cdot, u_0, v_0) \rangle_{(\mathcal{H}^{\alpha,\beta})^{\otimes 2}} \\
&= C \frac{1}{\sigma_{T,S}^4} \int_0^T dt \int_0^S ds \int_0^T dt_0 \int_0^S ds_0 \int_0^T du \int_0^S dv \int_0^T du_0 \int_0^S dv_0 \\
&\quad \int_0^T dx \int_0^S dy \int_0^T dx_0 \int_0^S dy_0 \int_0^T da \int_0^S db \int_0^T da_0 \int_0^S db_0 \\
&\quad f(x, y, t, s) f(x_0, y_0, t_0, s_0) f(a, b, u, v) f(a_0, b_0, t_0, s_0) \\
&\quad |t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\
&\quad |x-a|^{2\alpha-2} |y-b|^{2\beta-2} |x_0-a_0|^{2\alpha-2} |y_0-b_0|^{2\beta-2} \\
&= C \frac{1}{\sigma_{T,S}^4} \int_0^T dx \int_0^S dy \int_0^T dx_0 \int_0^S dy_0 \int_0^T da \int_0^S db \int_0^T da_0 \int_0^S db_0 \\
&\quad \int_0^x dt \int_0^y ds \int_0^{x_0} dt_0 \int_0^{y_0} ds_0 \int_0^a du \int_0^b dv \int_0^{a_0} du_0 \int_0^{b_0} dv_0 \\
&\quad \overline{J_0(2\sqrt{\theta(x-t)(y-s)}) J_0(2\sqrt{\theta(x_0-t_0)(y_0-s_0)})} \\
&\quad \overline{J_0(2\sqrt{\theta(a-u)(b-v)}) J_0(2\sqrt{\theta(a_0-u_0)(b_0-v_0)})} \\
&\quad |t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\
&\quad |x-a|^{2\alpha-2} |y-b|^{2\beta-2} |x_0-a_0|^{2\alpha-2} |y_0-b_0|^{2\beta-2}
\end{aligned}$$

making $\tilde{t} = \frac{t}{T}$, $\tilde{s} = \frac{s}{S}$ and similar for the other variables we get

$$\begin{aligned}
\mathbf{E} \left| A_{T,S}^{(1)} \right|^2 &= C \frac{1}{\sigma_{T,S}^4} T^{8\alpha} S^{8\beta} \int_0^1 dx \int_0^1 dy \int_0^1 dx_0 \int_0^1 dy_0 \int_0^1 da \int_0^1 db \int_0^1 da_0 \int_0^1 db_0 \\
&\quad \int_0^x dt \int_0^y ds \int_0^{x_0} dt_0 \int_0^{y_0} ds_0 \int_0^a du \int_0^b dv \int_0^{a_0} du_0 \int_0^{b_0} dv_0 \\
&\quad \overline{J_0(2\sqrt{\theta(x-t)(y-s)TS}) J_0(2\sqrt{\theta(x_0-t_0)(y_0-s_0)TS})} \\
&\quad \overline{J_0(2\sqrt{\theta(a-u)(b-v)TS}) J_0(2\sqrt{\theta(a_0-u_0)(b_0-v_0)TS})} \\
&\quad |t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\
&\quad |x-a|^{2\alpha-2} |y-b|^{2\beta-2} |x_0-a_0|^{2\alpha-2} |y_0-b_0|^{2\beta-2}
\end{aligned}$$

Using the asymptotic behavior of the Bessel function when its variable is close to infinity (see (11)) we see that

$$\begin{aligned} \mathbf{E} \left| A_{T,S}^{(1)} \right|^2 &\approx C \frac{1}{\sigma_{T,S}^4} T^{8\alpha} S^{8\beta} \int_0^1 dx \int_0^1 dy \int_0^1 dx_0 \int_0^1 dy_0 \int_0^1 da \int_0^1 db \int_0^1 da_0 \int_0^1 db_0 \\ &\int_0^x dt \int_0^y ds \int_0^{x_0} dt_0 \int_0^{y_0} ds_0 \int_0^a du \int_0^b dv \int_0^{a_0} du_0 \int_0^{b_0} dv_0 \\ &((x-t)(y-s)(x_0-t_0)(y_0-s_0)(a-u)(b-v)(a_0-u_0)(b_0-v_0))^{-\frac{1}{4}} \\ &|t-t_0|^{2\alpha-2} |s-s_0|^{2\beta-2} |u-u_0|^{2\alpha-2} |v-v_0|^{2\beta-2} \\ &|x-a|^{2\alpha-2} |y-b|^{2\beta-2} |x_0-a_0|^{2\alpha-2} |y_0-b_0|^{2\beta-2}. \end{aligned}$$

and considering the fact that the last integral is finite for $\frac{1}{2} < \alpha, \beta < \frac{5}{8}$ (the proof of this fact is similar to the proof of Lemma 1) it is straightforward to see that, for T, S close to infinity, the quantity $\mathbf{E} \left| A_{T,S}^{(1)} \right|^2$ does not converges to zero. ■

Acknowledgement: The authors would like to thank Prof. Soledad Torres for interesting discussions.

References

- [1] M. Arató, G. Pap and M.C.A van Zuijlen (2001): *Asymptotic inference for spatial autoregression and orthogonality of Ornstein-Uhlenbeck sheet*. Comput. Math. Appl. 42 (1-2), 219-229.
- [2] George B. Arfken and Hans J. Weber (2005): *Mathematical Methods for Physicists*, 6th edition (Harcourt: San Diego, 2005). ISBN 0-12-059876-0.
- [3] K. Bertin, S. Torres and C. A. Tudor (2011): *Drift parameter estimation in fractional diffusions, martingales and random walks*. Statistics and Probability Letters. 81(2), 243-249.
- [4] K. Bertin, S. Torres and C. A. Tudor (2010): *Maximum likelihood estimators and random walks in long-memory models*. Statistics. 44(5), 1-14.
- [5] P. Cheridito, H. Kawaguchi and M. Maejima (2003): *Fractional Ornstein-Uhlenbeck processes*. Electronic Journal of Probability. 8, 1-14.
- [6] A. Dorogovcev, P. Knopov (1979): *An estimator of a two-dimensional signal from an observation with additive random noise*. Theor Probab Math Stat. 17, 67-86.
- [7] T. Duncan, Y. Hu and B. Pasik-Duncan (2000): *Stochastic calculus for fractional Brownian motion. I. Theory*. SIAM J. Control Optim. 38, 582-612.

- [8] M. Erraoui, D. Nualart, Y. Ouknine (2003): *Hyperbolic stochastic partial differential equations with additive fractional Brownian sheet*. Stoch. Dyn. 3(2),121-139.
- [9] Y. Hu and D. Nualart (2010): *Parameter estimation for fractional Ornstein-Uhlenbeck processes*. Statistics and Probability Letters. 80, 1030-1038.
- [10] Yoon Tae Kim (2009): *A note on the differentiation formula in the Stratonovich type for fractional Brownian sheet*. Journal of the Korean Statistical Society. 38, 259-265.
- [11] Yoon Tae Kim and Hyun Suk Park (2009): *Stratonovich calculus with respect to fractional Brownian sheet*. Stoch. Anal. Appl. 27(5), 962-983.
- [12] Yoon Tae Kim, Jong Woo Jeon and Hyun Suk Park (2008): *Various types of stochastic integrals with respect to fractional Brownian sheet and their applications*. J. Math. Anal. Appl. 341, no. 2, 1382-1398.
- [13] M. Kleptsyna, A. Le Breton (2002): *Statistical analysis of the fractional Ornstein-Uhlenbeck type process*. Stat. Infer. Stoch. Process. 5(3), 229-248.
- [14] P. Major (2005): *Tail behavior of multiple random integrals and U-statistics*. Probability Surveys. 2, 448-505.
- [15] I. Nourdin, C. A. Tudor (2006): *Some linear fractional stochastic equations*. Stochastics. 78(2), 51-65
- [16] D. Nualart (1995): *The Malliavin Calculus and Related Topics*. First edition, Springer.
- [17] D. Nualart and S. Ortiz-Latorre (2008): *Central limit theorems for multiple stochastic integrals and Malliavin calculus*. Stochastic Processes Appl. 118, 614-628.
- [18] T. Sottinen and C. A. Tudor (2008): *Parameter estimation for stochastic equations with additive fractional Brownian sheet*. Stat Infer Stoch Process. 11, 221-236.
- [19] L. Slater (1966): *Generalized Hypergeometric Functions*. First edition, Cambridge.
- [20] C. A. Tudor and F. Viens (2003): *Itô Formula and Local Time for the Fractional Brownian Sheet*. Electronic Journal of Probability. 8(14), 1-31.
- [21] C. A. Tudor and F. Viens (2006): *It formula for the two-parameter fractional Brownian motion using the extended divergence operator*. Stochastics 78 (2006), no. 6, 443-462.
- [22] C. A. Tudor and F. Viens (2007): *Statistical aspects of the fractional stochastic calculus*. The Annals of Statistics. 25(5), 1183-1212.