

**A 0 – 1 LAW FOR VERTEX-REINFORCED RANDOM WALKS ON
 \mathbb{Z} WITH WEIGHT OF ORDER k^α , $\alpha < 1/2$.**

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ABSTRACT. We prove that Vertex Reinforced Random Walk on \mathbb{Z} with weight of order k^α , with $\alpha \in [0, 1/2)$, is either almost surely recurrent or almost surely transient. This improves a previous result of Volkov who showed that the set of sites which are visited infinitely often was a.s. either empty or infinite.

1. INTRODUCTION

Linearly vertex reinforced random walks (VRRW for short), introduced by Pemantle in [P], were first studied on \mathbb{Z} by Pemantle and Volkov [PV], who showed that with positive probability these processes spend all large times on just five sites. Some times later, Tarrès [T1] managed to prove that this striking phenomenon, called localization, occurs in fact almost surely (and he recently gave a simplified proof in [T2]). Roughly in the mean-time Volkov [V1] proved that (linearly) VRRW localize as well on a large class of graphs with positive probability and almost surely on trees. Benaïm and Tarrès [BT] have recently generalized his result to a larger class of walks, with a completely different proof.

In the recent works [ETW1, ETW2], new models of self-interacting random walks are introduced, where the interaction is not restricted to nearest neighbors. Then the authors prove that localization can occur on sets of arbitrary size, depending on the parameters of the model.

What emerges from these remarkable results is the fact that, when studying self-interacting (or non Markovian) random walks on graphs, the first thing one should do is to determine the set of vertices which are visited infinitely often and see if this is empty, the whole graph or some nontrivial subgraph. According to Volkov's notation [V2] we shall denote this set by R' here. If it is the whole graph, we say that the walk is recurrent and if it is empty we say that the walk is transient. But as we just saw, it might be equal to something else and even have arbitrary size.

In 2006 Volkov started the study of VRRW on \mathbb{Z} with some weight ($w_k, k \geq 0$). Such process, say $(X_n, n \geq 0)$, is defined as follows. First $X_0 = 0$. Then for all $n \geq 0$, on the event $\{X_n = x\}$,

$$(1) \quad \mathbb{P}(X_{n+1} = X_n \pm 1 \mid \mathcal{F}_n) = \frac{w_{Z_n(x \pm 1)}}{w_{Z_n(x-1)} + w_{Z_n(x+1)}},$$

where $(\mathcal{F}_n, n \geq 0)$ is the natural filtration of X and for all $y \in \mathbb{Z}$,

$$Z_n(y) = \#\{m \leq n : X_m = y\},$$

is the local time in y at time n . Linearly VRRW correspond to the case when $w_k = k + 1$ for all $k \geq 0$. Volkov [V2] showed that when $\sum_k 1/w_k$ is finite, then X almost surely localizes on two sites, i.e. that R' has a.s. cardinality 2, and when

$w_k \asymp (k+1)^\alpha$,¹ for some $\alpha \in [0, 1)$, then a.s. R' cannot be nonempty and finite (actually he proved this result under slightly more general hypotheses, see [V2] for details). He conjectured also that R' should be in fact empty (Problem 3 in [V2]). It is even natural to believe that this should hold as soon as $w_k = \mathcal{O}(k^\alpha)$, for some $\alpha < 1$, and not only when w_k is exactly of order k^α , for some fixed α . However, to our knowledge, no progress on this conjecture has been made since then, even in the case $\alpha = 0$. Here we obtain the following result:

Theorem 1.1. *Assume that there exists some $\alpha \in [0, 1/2)$, such that $w_k \asymp (k+1)^\alpha$. Then the VRRW on \mathbb{Z} with weight $(w_k, k \geq 0)$ is either a.s. recurrent or a.s. transient.*

This result says that R' is either a.s. empty or a.s. equal to \mathbb{Z} . This first step toward Volkov's conjecture, called Problem 1 in [V2], gives strong evidence that the conjecture should be true, at least when $\alpha < 1/2$. Indeed since the process is "reinforced" it should be "more" recurrent than simple random walk and since it does not localize it should be recurrent. However, giving a rigorous proof to this kind of monotonicity argument (even formulating a correct statement) is still a real challenge.

For other results on VRRW, particularly on finite graphs, we refer the reader to [B, BT, LV, P]. We shall also mention that analogous results have been obtained in a continuous setting, for self-interacting diffusions, see [CLJ, HR, R].

Our proof is very different from Volkov's proof, which was based on urns arguments and on Rubin's construction. We use instead a kind of domino principle, which works roughly as follows. Assume that some site $x \leq 0$ is visited infinitely often, but not $x - 1$, and let us fix some small constant $\epsilon > 0$. Then at k -th visit to x , with k large, the local time in $x + 1$ has to be at least of order $k^{1/\alpha - \epsilon}$. Otherwise, X would have jumped roughly $k^{\alpha\epsilon}$ times on $x - 1$, which is not allowed if k is large. Then we repeat this argument and show that before the $k^{1/\alpha - \epsilon}$ -th visit to $x + 1$, the local time in $x + 2$ has to be at least of order k^γ , with $\gamma = 1/\alpha + (1/\alpha - 1)^2 - \epsilon/\alpha$. Otherwise, during the $k^{1/\alpha - \epsilon}$ visits to $x + 1$, X would have jumped more than k times to x . By repeating this argument infinitely often, we get that the local time in $x + i$ has to be of order k^{γ_i} , with γ_i of order $(1/\alpha - 1)^i$, for all $i \geq 1$. This is of course not possible before the time of k -th visit to x , and we get a contradiction. However this argument only works when $\gamma_i \rightarrow \infty$, when $i \rightarrow \infty$, which explains why we need the hypothesis $\alpha < 1/2$. Then we deduce that a.s. R' is either empty or equal to \mathbb{Z} , see Sections 3 and 4 for more details. To really obtain a dichotomy, i.e. to prove that $\mathbb{P}[R' = \mathbb{Z}] \in \{0, 1\}$, we use in addition two general results, Lemma 2.1 and 2.2 below, which hold for any weight w . These lemmas assert that on one hand, if 0 is visited a.s. infinitely often, then X is a.s. recurrent. On the other hand, if there is positive probability that 0 is visited only finitely many times, then there is also positive probability that 0 is visited only once. By combining them with Borel-Cantelli like arguments, we can conclude.

2. TWO GENERAL RESULTS

Let us introduce some new notation. For any $w = (w_k(x))_{x \in \mathbb{Z}, k \geq 0}$, denote by \mathbb{P}_w the law of the VRRW in the "environnement" w . This process is defined as in (1) except that in the right hand side we replace $w_{Z_n(x \pm 1)}$ by $w_{Z_n(x \pm 1)}(x \pm 1)$.

¹we say that $f_k \asymp g_k$ when f_k/g_k is bounded from above and below by positive constants

Lemma 2.1. *If 0 has positive probability under \mathbb{P} to be visited only finitely many times, then for all $w \in (0, \infty)^{\mathbb{Z} \times \mathbb{N}}$, such that $w_k(x) = w_k$ for all $x \geq 0$ and $k \geq 0$, the probability under \mathbb{P}_w that 0 is visited only at time 0 is also positive. In particular $\mathbb{P}(X_n > 0 \forall n > 0) > 0$.*

Proof. By using sub-additivity and a symmetry argument we know that there exists some $M > 0$ such that $\mathbb{P}(X_n > 0 \text{ for all } n \geq M) > 0$. By conditioning now with respect to the first M steps, we see that there must exist some sequence (x_0, \dots, x_M) , with $x_M > 0$, such that conditionally on $E = \{(X_0, \dots, X_M) = (x_0, \dots, x_M)\}$, the probability that $X_n > 0$ for all $n > M$ is positive. But for any such sequence and any w as in the lemma, we have $\mathbb{P}_w(E) > 0$, and since $w_k(x) = w_k$ when $x \geq 0$, we have

$$\mathbb{P}_w(X_n > 0 \text{ for all } n > M \mid E) = \mathbb{P}(X_n > 0 \text{ for all } n > M \mid E) > 0.$$

Note that if X follows the path (x_0, \dots, x_M) during the first M steps and after stays on the right of 0, then certainly it always stays on the right of $-M$. Thus we also have

$$(2) \quad \mathbb{P}_w(X_n > -M \text{ for all } n \geq 0) > 0.$$

Now if $w' = (w'_k(x))_{k \geq 0, x \in \mathbb{Z}}$ is such that for all $k \geq 0$, $w'_k(x) = w_{k+1}$ if $x < 0$, and $w'_k(x) = w_k$ if $x \geq 0$, then by using the Markov property we get

$$\mathbb{P}_w(X_n > 0 \forall n > 0) \geq \mathbb{P}_w(X_1 = 1, \dots, X_M = M) \mathbb{P}_{w'}(X_n > -M \forall n \geq 0).$$

The first probability on the right hand side is positive (since $w_0 > 0$ by hypothesis), and it follows from (2), with w' in place of w , that the second one is also positive. This finishes the proof of the lemma. \square

Lemma 2.2. *For any sequence $(w_k)_{k \geq 0}$ of positive reals, if a.s. 0 is visited infinitely often, then X is a.s. recurrent, i.e. all sites are a.s. visited infinitely often.*

Proof. If the result was not true, then for symmetry reasons, there would exist some $x_0 \leq 0$ such that x_0 is a.s. visited infinitely often and $x_0 - 1$ has positive probability to be visited only finitely many times. But then we claim that this would also be true under probability \mathbb{P}_w , for any w as in Lemma 2.1. So fix such w . We consider two cases.

Case $x_0 = 0$. First by using Lemma 2.1 we know that 0 is also \mathbb{P}_w -a.s. visited infinitely often. So we only need to prove that under \mathbb{P}_w , there is also positive probability for -1 to be visited only finitely many times. But if $x_0 = 0$, then as in the proof of Lemma 2.1 there exists $M > 0$ and a sequence (x_0, \dots, x_M) , with $x_M > 0$, such that if $E = \{(X_0, \dots, X_M) = (x_0, \dots, x_M)\}$, then

$$\mathbb{P}(X_n \geq 0 \text{ for all } n > M \mid E) > 0.$$

Now let z be the local time in -1 at time M on the event E , i.e. $z = \#\{i \leq M : x_i = -1\}$. For $k \geq 1$, denote by $Z(k)$ the random variable equal to the local time in 1 at k -th visit to 0 after time M . Then we have

$$\mathbb{P}\left(X_n \geq 0 \text{ for all } n \geq M \mid E, \{Z(k)\}_{k \geq 1}\right) = \prod_{k=1}^{\infty} \left(1 - \frac{w_z}{w_z + w_{Z(k)}}\right),$$

which is nonzero if, and only if, $\sum_k 1/w_{Z(k)}$ is finite. So we must have

$$\mathbb{P}(X_n \geq 0 \text{ for all } n \geq M \text{ and } \sum_{k \geq 1} 1/w_{Z(k)} < \infty \mid E) > 0.$$

We deduce that there exists some $T > 0$ and $K \geq 1$, such that

$$p := \mathbb{P} \left(X_n \geq 0 \text{ for all } n \geq M \text{ and } \frac{T}{2} \leq \sum_{k=1}^K \frac{1}{w_{Z(k)}} \leq \sum_{k \geq 1} \frac{1}{w_{Z(k)}} \leq T \mid E \right) > 0.$$

Now for $k \geq 1$, let T_k be the time of k -th visit to 0 after time M . We can write

$$(3) \quad \mathbb{P}_w(X_n \geq 0 \forall M \leq n \leq T_k \mid E) = \sum_{\underline{x}} \mathbb{P}_w((X_M, \dots, X_{T_k}) = \underline{x} \mid E),$$

where the sum is over all trajectories \underline{x} starting from x_M , visiting 0 at least k times, staying on the right of 0 up to k -th visit in 0 and stopped after this k -th visit in 0. But since the jump probabilities above 1 are the same under \mathbb{P} and \mathbb{P}_w , for all such trajectory, we have

$$\mathbb{P}_w((X_M, \dots, X_{T_k}) = \underline{x} \mid E) = \mathbb{P}((X_M, \dots, X_{T_k}) = \underline{x} \mid E) \frac{\prod_{i=1}^{k-1} \left(1 - \frac{w_z(-1)}{w_z(-1) + w_{z(i)}}\right)}{\prod_{i=1}^{k-1} \left(1 - \frac{w_z}{w_z + w_{z(i)}}\right)},$$

where for all $i \geq 1$, $z(i)$ is the local time in 1 at time T_i , when X follows the trajectory \underline{x} . Recall next that for u small enough, $e^{-u} \leq (1-u) \leq e^{-u/2}$. So there exists a constant $c > 0$ (depending only on T), such that for any $k \geq 2$ and any sequence $(z(i), i \geq 1)$, satisfying $T/2 \leq \sum_{i=1}^{k-1} 1/w_{z(i)} \leq T$, we have

$$\frac{\prod_{i=1}^{k-1} \left(1 - \frac{w_z(-1)}{w_z(-1) + w_{z(i)}}\right)}{\prod_{i=1}^{k-1} \left(1 - \frac{w_z}{w_z + w_{z(i)}}\right)} \geq c.$$

So in (3) if we restrict the sum on the right hand side to trajectories \underline{x} such that $T/2 \leq \sum_{i=1}^{k-1} 1/w_{z(i)} \leq T$, we get

$$\begin{aligned} & \mathbb{P}_w(X_n \geq 0 \forall M \leq n \leq T_k \text{ and } T/2 \leq \sum_{i=1}^{k-1} 1/w_{Z(i)} \leq T \mid E) \\ & \geq c \times \mathbb{P}(X_n \geq 0 \forall M \leq n \leq T_k \text{ and } T/2 \leq \sum_{i=1}^{k-1} 1/w_{Z(i)} \leq T \mid E) \\ & \geq c \times p > 0, \end{aligned}$$

for all $k \geq K + 1$. By letting now k go to infinity we get

$$\mathbb{P}_w(X_n \geq 0 \forall n \geq M) > 0,$$

which was the desired result.

Case $x_0 < 0$. First observe that -1 has to be visited infinitely often under probability \mathbb{P}_w as well. Indeed if this was not the case, then we could apply the above argument (from case $x_0 = 0$) to see that -1 would also have positive probability to be visited only finitely many times under \mathbb{P} , which contradicts $x_0 < 0$. But actually this argument shows as well by induction that under \mathbb{P}_w , x_0 is a.s. visited infinitely often and $x_0 - 1$ has positive probability to be visited finitely many times.

Let us conclude now the proof of the lemma. We just have proved that for all w as in the statement of Lemma 2.1,

$$\mathbb{P}_w(X_n \geq x_0 \text{ for all } n \geq 0) > 0.$$

But under \mathbb{P} , X has positive probability to jump to the right the first $|x_0| + 1$ steps. So by applying the Markov property and the above inequality with w equal to w' as defined in the proof of Lemma 2.1, we see that

$$\mathbb{P}(X_n \geq 1 \text{ for all } n \geq |x_0| + 1) > 0,$$

which contradicts the fact that 0 is a.s. visited infinitely often. This finishes the proof of the lemma. \square

3. AN INDUCTION ARGUMENT AND A NEW PROOF OF VOLKOV'S RESULT

We first present a kind of domino principle. In plain words it works as follows. Assume that there exists some $x \in \mathbb{Z}$, such that $\inf R' = x$. It means that x is visited infinitely often, but not $x - 1$. To simplify assume even that $x - 1$ has never been visited. Fix some large integer k and let n be the time of k -th visit to x . Then at each of the k first visits to x , the process has probability at least of order $1/Z_n(x + 1)^\alpha$ to jump to $x - 1$. Since it did not, this implies with high probability that $Z_n(x + 1)$ is at least of order $k^{1/\alpha}$. The idea is then to repeat the argument. More precisely the next lemma implies by induction that the local time in $x + i$ at time n is of order at least k^{γ_i} , with $\gamma_i = \sum_{j=0}^i (1/\alpha - 1)^j$, up to some error term and with probability going to 1 exponentially fast when $k \rightarrow \infty$. In particular when $\alpha < 1/2$ the error term is negligible and we get a contradiction, since the process X cannot visit an infinite number of sites before time n . See the next subsection for details. When α is larger than or equal to $1/2$, the error term becomes predominant when $i \rightarrow \infty$, and the argument blows up. However, it still implies that R' cannot be finite, which gives an alternative proof to Volkov's result, see Corollary 3.1 below.

Now for $x \in \mathbb{Z}$ and $k \geq 1$, set

$$T_x(k) = \inf\{n \geq 0 : Z_n(x) \geq k\}.$$

Denote also by $T_x := T_x(1)$ the hitting time of x . In the following each time we consider an event of the type $\{T < T'\}$, for two random times T and T' , we implicitly assume that it is contained in the set $\{T < \infty\}$.

Lemma 3.1. *Assume that there exists some $\alpha \in [0, 1)$, such that $w_k \asymp (k + 1)^\alpha$. Then there exist constants $c > 0$ and $C > 0$, such that for all $x \in \mathbb{Z}$, all $\gamma > 1$, all $\epsilon \in (0, \alpha)$ and all $k \geq e^{C/\epsilon}$,*

$$\mathbb{P}\left[T_{x+1}(k^\gamma) < T_x(k) \wedge T_{x+2}(k^{\frac{\gamma-1}{\alpha}+1-\epsilon})\right] \leq \exp\left(-c \frac{k^{1-\alpha}}{|\ln \epsilon|^{1/(1-\alpha)}}\right).$$

Proof. Let $\epsilon \in (0, \alpha)$ and $\gamma > 1$ be given. Consider the event

$$A_0 := \left\{T_{x+1}(k^\gamma) < T_x(k) \wedge T_{x+2}(k^{\gamma'})\right\},$$

where $\gamma' := (\gamma - 1)/\alpha + 1 - \epsilon$. Set $K = \lceil 3 \ln \epsilon / \ln \alpha \rceil$. For $i = 1, \dots, K$, set

$$t_i := T_{x+1}\left(\frac{k^\gamma}{(K - i + 1)^2}\right),$$

and for $i \geq 2$,

$$N_i = \frac{k^\gamma}{(K-i+1)^2} - \frac{k^\gamma}{(K-i+2)^2}.$$

Set also $N_1 = k^\gamma/K^2$. Note that by hypothesis

$$(4) \quad N_i \geq k^\gamma/(K-i+2)^3 \geq 1,$$

for all $i \leq K$, and thus $t_i < t_{i+1}$. Moreover, since $w_k \asymp k^\alpha$, there exists some constant $c_0 \in (0, 1)$, such that for all $i_0 < j_0$, all $i \geq i_0$ and all $j \leq j_0$, $w_i/(w_i+w_j) \geq c_0 i_0^\alpha/j_0^\alpha$. In particular on the event A_0 , before time t_1 , at each visit to $x+1$, the probability to jump to x is larger than $p_1 := c_0/k^{\alpha\gamma'}$. Thus before time $1+t_1$, the number of jumps from $x+1$ to x stochastically dominates the sum of N_1 independent Bernoulli random variables with parameter p_1 . Therefore,

$$\mathbb{P}[A_0, Z_{1+t_1}(x) \leq N_1 p_1/2] \leq \exp(-c_1 N_1 p_1),$$

for some constant $c_1 > 0$. Define next inductively p_2, \dots, p_K , and A_1, \dots, A_{K+1} , by

$$p_i = c_0(N_{i-1}p_{i-1}/2)^{\alpha} k^{-\alpha\gamma'},$$

for $i \in \{2, \dots, K\}$, and

$$A_i := A_0 \cap \{Z_{1+t_i}(x) \geq N_i p_i/2\},$$

for $i \in \{1, \dots, K+1\}$. Now by using the same argument as above, we immediately get by induction that

$$(5) \quad \mathbb{P}[A_{i-1}, Z_{1+t_i}(x) \leq N_i p_i/2] \leq \exp(-c_1 N_i p_i),$$

for all $i \in \{2, \dots, K+1\}$. It is also straightforward to prove by induction, and by using (4), that

$$(6) \quad N_i p_i \geq \frac{2(c_0/2)^{1+\alpha+\dots+\alpha^i}}{\left(\prod_{j=1}^i (K-j+2)^{\alpha^{i-j}}\right)^3} k^{1-\alpha^i+\alpha\epsilon},$$

for all $i \leq K$. On the other hand it is immediate that

$$\sup_K \prod_{j=1}^K (K-j+2)^{\alpha^{K-j}} < \infty.$$

Thus there exists a constant $c' > 0$ such that

$$N_K p_K \geq c' k^{1-\alpha^K+\alpha\epsilon}.$$

By taking now $\epsilon \geq C/\ln k$, with C large enough, we deduce that $N_K p_K/2 > k+1$. Since on A_0 , $t_K = T_{x+1}(k^\gamma) \leq T_x(k)$, we get that A_{K+1} is empty. Finally note that for all $i \leq K$,

$$\prod_{j=1}^i (K-j+2)^{\alpha^{i-j}} \leq (K+1)^{1/(1-\alpha)},$$

so we also deduce from (5) and (6) that

$$\begin{aligned} \mathbb{P}[A_0] &\leq \sum_{j=1}^{K+1} \mathbb{P}[A_{i-1}, Z_{1+t_i}(x) \leq N_i p_i / 2] \\ &\leq (K+1) \exp(-c_2 k^{1-\alpha} / |\ln \epsilon|^{1/(1-\alpha)}) \\ &\leq \exp(-c k^{1-\alpha} / |\ln \epsilon|^{1/(1-\alpha)}), \end{aligned}$$

for some positive constants c_2 and c . This finishes the proof of the lemma. \square

We can now give an alternative proof to Volkov's result, in the case when w_k is of order k^α .

Corollary 3.1. *Assume that there exists $\alpha \in [0, 1)$, such that $w_k \asymp (k+1)^\alpha$. Then a.s. R' is infinite.*

Proof. Fix some $x \in \mathbb{Z}$ and some integers $N \geq 1$ and $z_0 > 0$. We want to prove that the event $\{Z_\infty(x) = \infty\} \cap \{Z_\infty(x-1) < z_0\} \cap \{Z_\infty(x+N) \leq 1\}$ has probability zero, with the convention $Z_\infty(y) := \lim_{n \rightarrow \infty} Z_n(y)$, for all $y \in \mathbb{Z}$.

For this first observe that for any $\epsilon < 1/\alpha$, and any $m \geq 1$,

$$\mathbb{P}\left[T_x(m) < T_{x+1}(m^{1/\alpha-\epsilon}) \wedge T_{x-1}(z_0)\right] \leq \mathbb{P}\left[\sum_{j=1}^m \xi_j \leq z_0\right] \leq e^{-c(z_0) m^{\epsilon\alpha}},$$

where $c(z_0)$ is some constant and the ξ_j 's are i.i.d. Bernoulli random variables with parameter $c'm^{-1+\epsilon\alpha}$, for some other constant $c' > 0$.

Now define $\gamma_1, \dots, \gamma_N$, by $\gamma_1 = 1$, $\gamma_2 = 1/\alpha - \epsilon$, and for $i \geq 1$,

$$\gamma_{i+2} = \gamma_i(1 - \epsilon) + \frac{1}{\alpha}(\gamma_{i+1} - \gamma_i).$$

Note already, that if ϵ is small enough, then $\gamma_{i+1} > \gamma_i$, for all $i \leq N-1$. Thus, as soon as m is large enough, we can apply Lemma 3.1 with $k = m^{\gamma_i}$ and $\gamma = \gamma_{i+1}/\gamma_i$, for any $i \in \{1, \dots, N-1\}$, and we get

$$\mathbb{P}[T_{x+i}(m^{\gamma_{i+1}}) < T_{x+i-1}(m^{\gamma_i}) \wedge T_{x+i+1}(m^{\gamma_{i+2}})] \leq \exp\left(-c \frac{m^{\gamma_i(1-\alpha)}}{|\ln \epsilon|^{1/(1-\alpha)}}\right),$$

where c is the constant appearing in Lemma 3.1. Then,

$$\begin{aligned} &\mathbb{P}[\{Z_\infty(x) = \infty\} \cap \{Z_\infty(x-1) < z_0\} \cap \{Z_\infty(x+N) \leq 1\}] \\ &= \mathbb{P}[\cap_{m \rightarrow \infty} \{T_x(m) < T_{x-1}(z_0)\} \cap \{Z_\infty(x+N) \leq 1\}] \\ &= \lim_{m \rightarrow \infty} \mathbb{P}[\{T_x(m) < T_{x-1}(z_0)\} \cap \{Z_\infty(x+N) \leq 1\}] \\ &\leq \lim_{m \rightarrow \infty} \left\{ \mathbb{P}\left[T_x(m) < T_{x+1}(m^{1/\alpha-\epsilon}) \wedge T_{x-1}(z_0)\right] \right. \\ &\quad \left. + \sum_{i=1}^{N-1} \mathbb{P}[T_{x+i}(m^{\gamma_{i+1}}) < T_{x+i-1}(m^{\gamma_i}) \wedge T_{x+i+1}(m^{\gamma_{i+2}})] \right\} \\ &= 0, \end{aligned}$$

as wanted. Since this is true for any x , $N \geq 1$ and z_0 , this proves the corollary. \square

4. PROOF OF THEOREM 1.1

We assume now that $\alpha < 1/2$. First, according to Lemma 2.1 and 2.2, if $\mathbb{P}[X_n > 0 \forall n > 0] = 0$, then X is a.s. recurrent and there is nothing more to prove. So we assume that $\mathbb{P}[X_n > 0 \forall n > 0] = \mathbb{P}[X_n < 0 \forall n > 0] > 0$, and we want to prove that X is a.s. transient.

For $x \leq 0$ and $m \geq 1$, consider the event

$$E_x(m) := \{T_x(m) < T_{x-1}\}.$$

Then for $i \geq 1$, set $\epsilon_i = r/i^2$, with $r > 0$ some positive constant which will be fixed later. Consider the sequence $(\gamma_i, i \geq 1)$ defined inductively by $\gamma_1 = 1$, $\gamma_2 = (1/\alpha) - r$ and for $i \geq 1$,

$$\gamma_{i+2} = \gamma_i(1 - \epsilon_i) + \frac{1}{\alpha}(\gamma_{i+1} - \gamma_i).$$

Set

$$F_x(m) := \{T_{x+i}(m^{\gamma_{i+1}}) < T_{x+i-1}(m^{\gamma_i}) \text{ for all } i \geq 1\}.$$

Let us show that if r is small enough, then

$$(7) \quad \sup_{x \leq 0} \mathbb{P}[F_x(m)^c \cap E_x(m)] = \mathcal{O}\left(e^{-\kappa m^{r\alpha}}\right),$$

as $m \rightarrow \infty$, for some constant $\kappa > 0$. For this note that for all $i \geq 1$,

$$\gamma_{i+2} - \gamma_{i+1} = \left(\frac{1}{\alpha} - 1\right)(\gamma_{i+1} - \gamma_i) - \epsilon_i \gamma_i,$$

so by induction we get

$$(8) \quad \gamma_{i+2} - \gamma_{i+1} = (\gamma_2 - \gamma_1)\left(\frac{1}{\alpha} - 1\right)^i - \sum_{j=1}^i \epsilon_j \gamma_j \left(\frac{1}{\alpha} - 1\right)^{i-j}.$$

In particular $\gamma_{i+2} - \gamma_{i+1} \leq (1/\alpha - 1)^{i+1}$, for all $i \geq 1$, which implies $\gamma_i \leq C_0(1/\alpha - 1)^i$, for some constant $C_0 > 0$. Since $\sum 1/i^2 < \infty$, we see from (8) that if $r > 0$ is small enough, then there exists a constant $c_0 > 0$, such that

$$\gamma_{i+2} - \gamma_{i+1} \geq c_0 \left(\frac{1}{\alpha} - 1\right)^i.$$

Thus $\gamma_{i+2} \geq c_0(1/\alpha - 1)^i$, for all $i \geq 1$, and since $\alpha < 1/2$, γ_i grows exponentially fast with i . Therefore, as soon as m is large enough, we can apply Lemma 3.1 with $k = m^{\gamma_i}$, $\gamma = \gamma_{i+1}/\gamma_i$ and $\epsilon = \epsilon_i$, for all $i \geq 1$. Then we get

$$\begin{aligned} \mathbb{P}[F_x(m)^c \cap E_x(m)] &\leq \mathbb{P}[T_x(m) < T_{x+1}(m^{\gamma_2}) \wedge T_{x-1}(z_0)] \\ &\quad + \sum_{i \geq 1} \mathbb{P}[T_{x+i}(m^{\gamma_{i+1}}) < T_{x+i-1}(m^{\gamma_i}) \wedge T_{x+i+1}(m^{\gamma_{i+2}})] \\ &\leq e^{-\kappa m^{r\alpha}} + \sum_{i \geq 1} \exp\left(-c \frac{m^{\gamma_i(1-\alpha)}}{|\ln \epsilon_i|^{1/(1-\alpha)}}\right), \end{aligned}$$

where c is the constant appearing in Lemma 3.1, and κ some other constant, see the proof of Corollary 3.1. Since γ_i grows exponentially fast with i , (7) follows. But for any $x \leq 0$ and any $m \geq 1$, the event $F_x(m) \cap E_x(m)$ is empty since X cannot visit infinitely many sites in finite time. This proves that

$$(9) \quad \sup_{x \leq 0} \mathbb{P}[E_x(m)] = \mathcal{O}\left(e^{-\kappa m^{r\alpha}}\right),$$

as well, when $m \rightarrow \infty$. Similarly, if $E'_x := \{T_x(m) < T_{x+1}\}$, for $x \geq 0$, then

$$(10) \quad \sup_{x \geq 0} \mathbb{P}[E'_x(m)] = \mathcal{O}\left(e^{-\kappa m^{r\alpha}}\right),$$

as $m \rightarrow \infty$.

Now (9) and (10) show that

$$\sum_{x \geq 0} \mathbb{P}[\{T_x(x) < T_{x+1}\} \cup \{T_{-x}(x) < T_{-x-1}\}] < +\infty.$$

Thus according to Borel–Cantelli’s lemma, a.s. for x large enough, either $T_{x+1} < T_x(x)$ or $T_{-x-1} < T_{-x}(x)$. For $n \geq 1$, denote by x_n the n -th site visited by X , such that $T_{x_n} < T_{x_{n-1}}(x_n - 1)$ and $x_n > 0$, or $T_{x_n} < T_{x_{n+1}}(|x_n + 1|)$ and $x_n < 0$. Note now that for all n , if for instance $x_n > 0$, then X has probability of order at least $n^{-\alpha}$ to jump to $x_n + 1$ at time T_{x_n} , and similarly if $x_n < 0$. Hence,

$$\sum_{n \geq 1} \mathbb{P}[T_{x_{n+1}} = T_{x_n} + 1 \text{ or } T_{x_{n-1}} = T_{x_n} + 1 \mid \mathcal{F}_{T_{x_n}}] = \infty.$$

It then follows from Levy’s conditional Borel–Cantelli’s lemma (see for instance Lemma 5.1 in [T2]), that a.s. for infinitely many $n \geq 1$, either $T_{x_{n+1}} = T_{x_n} + 1$ (if $x_n > 0$) or $T_{x_{n-1}} = T_{x_n} + 1$ (if $x_n < 0$). But each time this happens, Lemma 2.1 shows that, independently of $\mathcal{F}_{T_{x_n}}$, X has some positive probability $p > 0$ to never come back to x_n after time T_{x_n} . It follows that a.s. this happens infinitely often, which proves well that X is a.s. transient, as wanted. \square

REFERENCES

- [B] **Benaïm M.:** *Vertex-reinforced random walks and a conjecture of Pemantle*, Ann. Probab. 25, (1997), 361–392.
- [BT] **Benaïm M., Tarrès P.:** *Dynamics of Vertex-Reinforced Random Walks*, to appear in Ann. Probab.
- [CLJ] **Cranston M., Le Jan Y.:** *Self-attracting diffusions: Two case studies*, Math. Ann. 303, (1995), 87–93.
- [ETW1] **Erschler A., Tóth B., Werner W.:** *Some locally self-interacting walks on the integers*, arXiv:1011.1102.
- [ETW2] **Erschler A., Tóth B., Werner W.:** *Stuck Walks*, to appear in Probab. Theory Related Fields.
- [HR] **Herrmann S., Roynette B.:** *Boundedness and convergence of some self-attracting diffusions*, Math. Ann. 325, (2003), 81–96.
- [LV] **Limic V., Volkov S.:** *VRRW on complete-like graphs: almost sure behavior*, Ann. Appl. Probab. 20, (2010), 2346–2388
- [P] **Pemantle R.:** *Vertex-reinforced random walk*, Probab. Theory Related Fields 92 (1992), no. 1, 117136.
- [PV] **Pemantle R., Volkov S.:** *Vertex-reinforced random walk on \mathbb{Z} has finite range*, Ann. Probab. 27, (1999), 1368–1388.
- [R] **Raimond O.:** *Self-attracting diffusions: Case of the constant interaction*, Probab. Theory Related Fields 107, (1997), 177–196.
- [T1] **Tarrès P.:** *Vertex-reinforced random walk on \mathbb{Z} eventually gets stuck on five points*, Ann. Probab. 32, (2004), 2650–2701.
- [T2] **Tarrès P.:** *Localization of reinforced random walks*, arXiv:1103.5536.
- [V1] **Volkov S.:** *Vertex reinforced random walk on arbitrary graphs*, Ann. Probab. 29, (2001), 66–91.
- [V2] **Volkov S.:** *Phase transition in vertex-reinforced random walks on \mathbb{Z} with non-linear reinforcement*, J. Theoret. Probab. 19, (2006), 691–700.

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