

Risk of recurrent myocardial infarction with the concomitant use of clopidogrel and proton pump inhibitors.

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Risk of recurrent myocardial infarction with the concomitant use of clopidogrel and proton pump inhibitors.

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2 3	1	Risk of recurrent myocardial infarction with the concomitant use of clopidogrel
4 5 6	2	and proton pump inhibitors.
7 8	3	Vera E. Valkhoff ^{1,2} , Geert W. 't Jong ² , Eva M. van Soest ² , Ernst J. Kuipers ^{1,3} , and Miriam
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2 3	1	Authors' declaration of personal interests:
4 5	2	V.Valkhoff, E.van Soest, and G. 't Jong do not have any conflicts to declare.
6 7	3	As employee of Erasmus MC, M. Sturkenboom has been involved as project leader and in
, 8 9	4	analyses contracted by various pharmaceutical companies. She received unconditional
10	5	research grants from AstraZeneca, Pfizer, Merck, Johnson&Johnson, Amgen, Roche, GSK,
11 12	6	Boehringer, Yamanouchi, and Altana, none of which are related to the subject of this study.
13 14	7	M. Sturkenboom has been consultant to Pfizer, Servier and Lundbeck on issues not related to
15 16	8	this manuscript. E. Kuipers has served as a speaker and advisory board member for
17 18	9	this manuscript. E. Kuipers has served as a speaker and advisory board member for AstraZeneca.
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1 2 3	1	Key words:	
4 5 6	2	Proton pump inhibitors	
7 8 9	3	Clopidogrel	
10 11	4	Myocardial infarction	
12 13 14	5	Case-control study	
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Deleted: , but not when compared to past PPI use (OR:1.22, 95%CI:0.79-

1.88).

2 3	1	Abstract
4 5	2	Background: The association between myocardial infarction (MI) and co-administration of
6 7 8	3	PPIs and clopidogrel remains controversial.
9 10	4	Aim: To quantify the association between concomitant use of PPIs and clopidogrel and
11 12	5	occurrence of recurrent MI.
13 14 15	6	Methods: We conducted a case-control study within a cohort of acute MI patients in
16 17	7	PHARMO Record Linkage System (1999-2008). Cases were patients readmitted for MI. PPI
18 19	8	exposure was categorized as current (3-1 days before MI), past (30-3 days before MI), or no
20 21	9	use (>30 days before MI). We used conditional logistic regression analyses.
22 23	10	Results: Among 23,655 patients hospitalized following MI, we identified 1,247 patients
24 25	11	readmitted for MI. Among clopidogrel users, current PPI use was associated with an increased
26 27	12	risk of recurrent MI (OR:1.62, 95%CI:1.15-2.27) when compared to no PPI use, but not when
28 29	13	compared to past PPI use (OR:0.95, 95%CI:0.38-2.41). Among clopidogrel-non-users, current
30 31	14	PPI use was associated with an increased risk of recurrent MI (OR:1.38, 95%CI:1.18-1.61)
32 33	15	when compared to no PPI use
34 35 36	16	Conclusions: The apparent association between recurrent MI and use of PPIs with clopidogrel
37 38	17	depends on the design and is affected by confounding by indication. The association is not
39 40	18	present when (un)measured confounding is addressed by design.
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1 2 3	1	Introduction
4 5	2	Clopidogrel is an oral antiplatelet agent commonly used in addition to aspirin to
6 7	3	reduce cardiovascular (CV) events. Clopidogrel is converted in the liver from an inactive pro-
8 9	4	drug to its active metabolite, that binds irreversible to P2Y12 adenosine diphosphate (ADP)
10	5	receptors on the platelet surface, thereby preventing platelet aggregation ¹ .
11 12	J	receptors on the platelet surface, thereby preventing platelet aggregation .
13 14	6	Clopidogrel conversion is catalyzed by several cytochrome P450 (CYP) enzymes, of
15 16	7	which CYP2C19 is the most important. Patients with loss-of-function polymorphisms in the
17 18	8	gene encoding for CYP2C19 have lower levels of the active metabolite and have reduced
19 20	9	platelet inhibition during clopidogrel treatment. This results in higher rates of acute
21 22	10	myocardial infarction (MI) ^{2,3} . In contrast, rapid metabolizers of clopidogrel (with CYP2C19
23	11	variants leading to increased enzyme activity) have a higher rate of clopidogrel activation and
24 25	12	more efficient platelet inhibition ⁴ .
26 27		
28 29	13	Proton pump inhibitors (PPIs) are routinely co-administered with clopidogrel to
30 31	14	prevent upper gastrointestinal (UGI) bleeding ⁵⁻⁷ , which is in line with expert consensus
32 33	15	guidelines ⁸ . PPIs are also metabolized by CYP2C19, and can competitively bind to its
34 35	16	catalytic site. Therefore, PPIs are potentially hindering the conversion of clopidogrel to its
36	17	active metabolite. Both the European Medicines Agency (EMA) and the US Food and Drug
37 38	18	Administration (FDA) recently advised against the concurrent use of PPIs and clopidogrel in
39 40	19	the absence of a strong indication ^{9, 10} , because several pharmacodynamic ¹¹⁻¹⁴ and clinical
41 42	20	outcome studies ¹⁵⁻²⁰ suggest a detrimental interaction between clopidogrel and PPIs.
43 44	24	Fridence on the second time between conditions and an educidate time of
45 46	21	Evidence on the association between cardiovascular events and co-administration of
47 48	22	PPIs with clopidogrel remains inconclusive. Observational studies show a small increased risk
49	23	of recurrent MI in patients using PPI – clopidogrel combination therapy compared to those
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2 3	1	using clopidogrel alone. Estimated relative risks vary from 0.92 to 1.93 ^{2,7,13,15-20} . These	
4 5	2	results are likely influenced by confounding by indication. Confounding by indication is	
6 7	3	introduced when more severely ill patients with a worse prognosis are more likely to receive	
8	4	PPIs than healthier patients. Furthermore, some published observational studies may have	
9 10	5	suffered from immortal time bias ^{15,21} . Both biases could distort the studied association in	Deleted: different
11 12 13	6	<u>either</u> directions.	
14 15	7	Other studies failed to show an interaction between the use of PPIs and clopidogrel ^{2, 6,}	
16 17	8	^{7, 13, 19, 22, 23} . Recently, results of a post hoc analysis of a randomized trial revealed no	
18 19	9	association between PPI use and the risk of the primary CV endpoint for patients treated with	
20 21	10	clopidogrel or the novel thienopyridine prasugrel (a prodrug also requiring metabolization	
22 23	11	through CYP enzymes) ¹³ . The main limitation of this randomized trial was that use of a PPI	
24 25	12	was not randomized. Preliminary analysis (pre-specified sample size and follow-up time not	
26 27	13	reached) of the unpublished Clopidogrel and the Optimization of Gastrointestinal Events	
28 29	14	(COGENT) trial, a randomized double-blind trial of omeprazole 20 milligram (mg) vs.	
30 31	15	placebo in patients taking dual therapy (clopidogrel and aspirin) demonstrated no significant	
32 33 34	16	difference in CV events between both study arms (hazard ratio = $1.02, 95\%$ CI: $0.70-1.51$) ⁶ .	
35 36	17	Because of the potential clinical consequences for a large patient group at risk for both	
37 38	18	recurrent CV as well as UGI events, residual uncertainty about this potential drug-drug	
39 40	19	interaction should be minimized. Therefore, we conducted a nested case-control study using	
40 41 42	20	data from the PHARMO Record Linkage System (RLS) (1999-2008) to quantify the	
43	21	association between use of PPIs and recurrent MI in the absence or presence of clopidogrel,	
44 45	22	while addressing the issues of both study design (avoidance of immortal time bias) and	
46 47	23	residual confounding (using past exposure to PPIs as the reference category and propensity	
48 49	24	score-based adjustments).	
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2	1	Patients and Methods
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8	3	Study design
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10	4	A population-based nested case-control study (1999-2008) was conducted within a
11 12		
12	5	cohort of patients admitted for acute MI during the study period.
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17	7	Setting
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20 21	8	The study was conducted using data from the PHARMO RLS. This system comprises
22	9	drug-dispensing records mostly from community pharmacies and hospital discharge records
23	9	drug-dispensing records mostry from community pharmacies and hospital discharge records
24	10	of more than three million inhabitants of 50 demographically defined areas in the
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26 27	11	Netherlands. For all participants, the computerized drug-dispensing histories contain data
27 28	12	concerning the name of the dispensed drug, dispensing date, dispensed amount, prescribed
29	12	concerning the name of the dispensed drug, dispensing date, dispensed amount, presented
30	13	dosing regimens, and the legend duration of use (prescription length). All drugs are coded
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32	14	according to the Anatomical Therapeutic Chemical (ATC) classification ²⁴ . The hospital
33 34	15	records include detailed information concerning discharge diagnosis, procedures, dates of
35	15	records merude detailed information concerning discharge diagnosis, procedures, dates of
36	16	hospital admission and discharge, discharge destination (or death in the hospital before
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38	17	discharge). Diagnoses are consistently classified according to the International Classification
39 40	18	of Diseases, Ninth Revision, Clinical Modification (ICD-9-CM) during the entire study
40 41	10	of Diseases, Whith Revision, Chinear Would and in (ICD-9-CW) during the entire study
42	19	period. For a detailed description of the database we refer to a previous publication ²⁵ .
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47 48	21	Study cohort
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1	The study cohort included all patients in the PHARMO RLS who were hospitalized
2	between January 1st, 1999 and December 31st, 2008 for a primary diagnosis of acute MI.
3	Cohort entry was date of discharge from hospital of an acute MI, registered as a primary
4	discharge diagnosis code 410 according to ICD-9-CM. Secondary diagnoses were not queried,
5	because it is likely that prevalent MI events (co-morbidity) are recoded as secondary
6	diagnosis. Because of the dynamic study population, patients were required to have had one
7	prescription filled at least one year preceding the date of cohort entry (i.e. one year of valid
8	database history) in order to allow uniform assessment of the presence of co-morbidity and
9	confounding factors. Patients were followed from cohort entry until diagnosis of recurrent MI,
10	last filled prescription, or the end of the study period (December 31st, 2008), whichever came
11	first.
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13	Case selection
14	The study outcome was hospitalization for a subsequent acute MI, registered as
15	primary diagnosis ICD-9-CM code 410. Only the first encountered recurrent acute MI during
16	follow-up was included in the analysis. The date of rehospitalization for the MI was used as
17	index date. To avoid misclassification of the exposure, we required a 30-day period between
18	discharge from baseline MI and recurrent MI, since outpatient time was needed to observe
19	filling of outpatient prescriptions. In a sensitivity analysis, we used the requirement of a 90-
20	day period between discharge from baseline MI and recurrent MI.
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22	Control selection
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1 For each case readmitted for acute MI, controls were randomly selected from the cohort, matched on gender, age (same year of birth), being at risk of a recurrent ML and 2 calendar time by means of incidence density sampling ²⁶. Only those controls were considered 3 to be at risk of developing a recurrent MI when they had a baseline acute $MI \ge 30$ days prior 4 5 to the index date.

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Definition of exposure 7

8 Drug exposure data were obtained from outpatient drug dispensing files of community 9 pharmacies, as recorded in the PHARMO RLS. The exposures of interest were clopidogrel and PPIs. PPIs included in the analyses were omeprazole, pantoprazole, lansoprazole, rabeprazole, and esomeprazole. Duration of use was obtained by dividing the total amount of dispensed units by the prescribed amount per day. We calculated the end of each prescription by adding the duration to the dispensing date.

14 Clopidogrel use was classified at the index date as either "no" use (no use or use at 15 least 30 days before index date), "past" use (end of prescription between seven and 30 days 16 before the index date), or "current" use (use at index date, or end of prescription less than seven days before index date). Because half-time values of PPIs are short (a few hours), the 17 18 current, past, and no use of PPIs was defined differently: "no" use (no use or use at least 30 days before index date), "past" use (end of prescription between three and 30 days before 19 20 index date), and "current" use (use at index date, or end of prescription less than three days before index date). To avoid protophatic bias, only prescriptions dispensed at least one day 21 22 before index date were considered "current" use. In sensitivity analyses, the definition of past 23 PPI use (which was the reference) was adjusted to use between three to 90 days or between 14

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and 90 days preceding the index date. In the latter case the definition of "current" PPI
 switched to use at index date, or final prescription date less than 14 days before index date.
 The study is not subject to immortal time bias as the outcome under study could be assessed
 equally for cases and controls after the 30- or 90-day period.

Covariates

At baseline, we considered the following factors as potential covariates: percutaneous transluminal coronary artery angiography (PTCA) within 30 days of baseline MI, prior hospitalization for cardiogenic shock, congestive heart failure, cancer, cerebrovascular disease, acute renal failure, chronic renal failure, cardiac dysrhythmia (including atrial fibrillation), or upper gastrointestinal (UGI) ulcer. The covariate diabetes mellitus was a combination of prior hospitalization for diabetes mellitus and anti-diabetic drug use. As a proxy for (the degree of) comorbidity, we aggregated all prescriptions per ATC-code in the one year preceding cohort entry and we refer to this 'score' as the number of different prescriptions. Additionally, we considered use of commonly prescribed cardiovascular medications (anti-hypertensive drugs, low-dose acetylsalicylic acid, β -adrenergic antagonists, digoxin, and lipid lowering drugs) or NSAIDs at baseline. We defined comorbidity and medication at baseline rather than at recurrent MI to avoid adjustment on intermediates in the pathway from exposure to outcome. Medications inducing or inhibiting cytochrome P450 2C19 or cytochrome P450 3A4 (directly affecting levels of clopidogrel and PPIs)²⁷, as well as standard secondary prevention medications such as lipid lowering drugs and anti-hypertensive agents were assessed at the index date. Furthermore, days of follow-up (number of days between date of discharge of baseline MI and index date) and length of hospital stay for the baseline MI were also considered as covariates.

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5 6	2	Statistical analysis
7 8	3	Descriptive analyses were performed to describe the cohort with regard to baseline
9 10	4	characteristics. Prescription rates for PPIs and clopidogrel for the cohort were calculated per
11 12	5	calendar month to assess prescription time trends. Survival rates were estimated using the
13 14	6	Kaplan-Meier method and evaluated the probability of recurrent MI free survival during
15 16	7	follow-up.
17 18	8	Baseline characteristics were compared between cases with a recurrent MI and
19 20	9	controls without recurrent MI in patients unexposed to PPIs and between all PPI users and PPI
21 22	10	non-users in patients without recurrent MI using univariate conditional logistic regression
23 24	11	analyses. We applied conditional logistic regression analysis to estimate matched and adjusted
25 26	12	odds ratios (OR) and 95% confidence intervals (95% CI) for the association between recurrent
27 28	13	MI and exposure to PPI. The analyses were performed in two subgroups: 1) current
29 30	14	clopidogrel users, and 2) clopidogrel non-users to examine whether there was a true drug-drug
31 32	15	interaction between clopidogrel and PPIs. The reference group were patients with no
33 34	16	prescription of PPIs in the previous month (non-PPI users). A second analysis that more
35 36	17	effectively deals with confounding by indication was performed applying past PPI users as the
37 38	18	reference category. In the adjusted model we included all factors that changed the unadjusted
39 40	19	odds ratio with more than 10 percent. To test the robustness of our finding, we examined the
41 42	20	effect of extension of the past PPI exposure window on the observed associations as a
43 44	21	sensitivity analysis. Finally, we estimated the OR for current use of omeprazole and
45 46	22	esomeprazole, using current pantoprazole as the reference group.
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Because the use of PPIs was not randomly assigned, we developed a propensity score (summary exposure risk score). The propensity score represents the likelihood of exposure to a PPI for each individual patient ²⁸. The propensity score for use of PPI was developed using logistic regression modeling. The log odds of the probability that a patient received a PPI at any point of time during follow-up was modeled as a function of all previously described baseline covariates and was included in our dataset as a separate confounder. The PPI propensity score was only included in the adjusted analysis when it changed the odds ratio with more than 10 percent, unless we adjusted for individual covariates, to avoid double adjustment for the same covariate.

Statistical significance was assumed for two-sided p-values <0.05. All statistical
analyses were performed using SPSS software version 15.0 (SPSS inc, Chicago, III).

1 2 3	1	Results
4 5	2	The initial study population comprised 27,513 patients who were admitted to the
6 7 8 9 10 11	3	hospital with an acute MI during follow-up. After exclusion of patients of whom not one year
	4	of valid data before baseline MI was available, the final cohort comprised 23,655 patients.
	5	The mean age was 64.7 (± standard deviation (SD) 13.2) years and 15,897 (67.2%) were
12 13	6	male. The median follow-up time after discharge was 42.6 months (interquartile range (IQR)
14 15	7	16.8 to 71.7 months). Co-morbidity at baseline was substantial: 4.4% of patients had been
16 17	8	hospitalized for cancer, 3.3% for cardiac dysrhythmia, and 2.8% for cerebrovascular disease.
18 19	9	Prevalence of diabetes mellitus was 21.6%. The median number of distinct prescribed drugs
20 21	10	in the one year before cohort entry was 5 (IQR 2 to 9 prescriptions) (Table 1; column 2). At
22 23	11	baseline, 10.5% of the patients used proton pump inhibitors (PPIs) and 315 patients (1.3%)
24 25	12	used clopidogrel.
26 27 28	13	Figure 1 shows that prescription rates of both PPIs and clopidogrel changed over time
29 30	14	between January 1st, 1999 and December 31st, 2008, illustrating the increased use of the
31 32 33	15	study drugs over a decade.
34 35	16	Within the study cohort, 1,247 patients were readmitted to the hospital with an acute
36	17	MI. The risk of a recurrent MI at least 30 days after discharge was 2.8% (95% CI: 2.6% -
37 38 39	18	3.0%) in the first year, and 10.3% (95% CI: 9.1% - 11.5%) in 10 years.
40 41	19	Of the cases with a recurrent MI, 1,224 (98.2%) could be matched to at least one
42 43	20	control without a recurrent MI. Characteristics of cases and controls (in the unexposed group
44 45	21	to PPIs) are depicted in Table 1 (Columns 3 and 4). As expected, cases were more likely to
46 47	22	have been hospitalized for co-morbidities associated with an increased risk of recurrent MI,
48 49	23	such as congestive heart failure and diabetes mellitus. At baseline, cases used several types of
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1	cardiovascular drugs more frequently, for instance angiotensin-converting-enzyme (ACE)	
2	inhibitors and lipid lowering drugs. To understand whether selective prescribing (channeling)	
3	of PPIs to persons at increased risk of recurrent MI occurred, we assessed the presence of risk	
4	factors in PPI users and PPI non-users (in <u>patients who did not develop a recurrent MI</u>) (Table	Deleted: control group
5	1; columns 5 and 6). Patients using PPIs were older, were more frequently male, had been	
6	hospitalized more frequently prior to cohort entry for several co-morbid conditions, and more	
7	frequently used cardiovascular medications at baseline.	
8	Among current clopidogrel users, a significant association was observed between	
9	current PPI use and recurrent MI when compared to PPI non-use (OR: 1.62, 95% CI: 1.15-	
10	2.27). When applying past PPI use as the reference category rather than PPI non-use, we	
10	found no association between recurrent MI and current use of PPI (OR: 0.95, 95% CI: 0.38-	
12	2.41) (Table 2). Repeated analyses with variations in the definition of past PPI use (3 to 90	
13	days or 14 to 90 days) did not affect the study findings appreciably; OR: 1.15, 95% CI: 0.51-	
14	2.61, and OR: 1.34, 95% CI: 0.50-3.62, respectively.	
14	2.01, and O.A. 1.5 1, 55 /6 CH. 0.50 5.02, respectively.	
15	Among clopidogrel non-users, current use of PPI was associated with recurrent MI	
16	when compared to PPI non-use (OR:1.38, 95% CI: 1.18-1.61). When applying past PPI users	
17	as the reference category, a weak association between recurrent MI and current use of PPI	
18	was found (OR: 1.22, 95% CI: 0.79-1.88), although the latter's confidence interval cross the	
19	null, which was likely due to insufficient power (Table 3).	Deleted: Among clopidogrel non-users,
20	· · · · · · · · · · · · · · · · · · ·	current use of PPI was associated with recurrent MI when compared to PPI non- use (OR:1.38, 95% CI: 1.18-1.61). When
		applying past PPI users as the reference category, no
21	The analysis by type of PPI among current clopidogrel users showed that pantoprazole	Deleted: association between recurrent MI and current use of PPI was found (OR: 1.22, 95% CI: 0.79-1.88) (Table 3)
22	was most frequently used, followed by omeprazole and esomeprazole. Lansoprazole and	
23	rabeprazole were rarely used. Using current pantoprazole as the reference category, current	
	14	

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2 3	1	use of omeprazole and of esomeprazole were not associated with an increased risk of
4 5	2	recurrent MI (respectively, OR: 1.07, 95 % CI: 0.58-1.99 and OR: 0.83, 95 % CI: 0.40-1.69,
6	3	respectivily) (Table 4).
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1 Discussion

2	This population-based cohort study showed that the association between clopidogrel –
3	PPI co-therapy and risk of recurrent MI is highly affected by confounding by indication ²⁹ ,
4	which may explain the contrasting results in the literature. To illustrate the problem of
5	confounding we compared current PPI users to none or past PPI users in the absence of
6	clopidogrel. To deal with the problem of confounding by indication, we compared current use
7	of clopidogrel plus current use of PPI not only to current clopidogrel without PPIs, but also to
8	current clopidogrel use plus past use of PPIs, which reduces confounding by indication. When
9	current PPI use was compared to past PPI use the association disappeared, suggesting that the
10	observed association between current PPI use and recurrent MI when PPI non-use was the
11	reference, may have been the result of residual confounding.
12	Our findings are in line both with other observational studies showing an increased
13	risk among PPI users when compared to PPI non-users ¹⁵⁻²⁰ , and those that did not find such
14	an association if better control of confounding was applied ^{6, 13, 19, 30} . In observational studies
15	or post-hoc analyses in randomized controlled trials, the use of PPI is by definition not
16	randomly assigned, most likely leading to confounding by indication. It is arguable whether
17	adequate adjustment for this harmful type of confounding is possible since much of the
18	confounding will be subtle and unmeasurable. Design choices and progressive adjustment
19	techniques can be used to avoid residual confounding, as we have shown. By careful selection
20	of the study design one can match cases and controls on important covariates, including
21	calendar time. We illustrated the increased use of the study drugs over a decade (Figure 1),
22	and thereby the importance of matching on index date (thus calendar time) as potential
23	confounder. Additionally, one can use a so-called 'active comparator' as reference, such as
24	past users of PPIs instead of none users. For adjustment high-dimensional propensity scores
	16

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1	¹⁹ , summary disease risk scores (probability of short term mortality) ¹⁶ , or conventional	
2	propensity scores ^{13, 20} can be applied. Table 5 provides an overview of techniques used in	
3	other observational studies on this topic to deal with confounding by indication.	
4	Our study further confirms the relevance of confounding, illustrated by the fact that	
5	PPI users had more co-morbidity and used more co-medication than PPI non-users at baseline	
6	and also by the results of the analysis in clopidogrel non-users. The increased risk for	
7	recurrent MI with current PPI use (OR 1.38, 95% CI: 1.18-1.61) compared to no PPI use in	
8	the absence of clopidogrel therapy provides further support for this. An <u>other</u> explanation	Deleted: alternative
9	would be that PPIs have a harmful effect irrespective of clopidogrel status. We are aware of	Deleted: c
	/	Deleted: one
10	two studies that conducted similar analysis in clopidogrel non-users which also showed	Deleted: y
11	significantly elevated risks with PPI use (adjusted OR <u>1.29 and</u> 1.55) compared to PPI non-	Deleted: performed Deleted: a
12	use ^{23, 31} . In two other observational studies however, no significant increase in CV events was	
13	shown in patients prescribed PPIs without clopidogrel ^{15, 32} .	
14	We addressed the confounding issue by all possible design and adjustments measures,	
15	but still the comparison against none-use of PPI seemed confounded. Applying past PPI use	
16	as active comparator changed the association from increased to no effect and was the most	
17	powerful approach to deal with channeling of PPIs, more than adjustment for the propensity	
18	score. The propensity score did not have large explanatory power in this study due to the fact	
19	that relatively little clinical information was available to construct the score.	
20	Our study is internally valid since we used a population-based design, so selection bias	
21	was unlikely as all cases and controls came from the same source population. The same	
22	pertains to information bias, because data were gathered prospectively without knowledge of	
23	the hypothesis studied. The method to control for confounding was discussed previously. The	
24	database is proven valid for research as the diagnoses are labeled with ICD-9-CM codes in	
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	17	
	2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23	 propensity scores ^{13,20} can be applied. Table 5 provides an overview of techniques used in other observational studies on this topic to deal with confounding by indication. Our study further confirms the relevance of confounding, illustrated by the fact that PPI users had more co-morbidity and used more co-medication than PPI non-users at baseline and also by the results of the analysis in clopidogrel non-users. The increased risk for recurrent MI with current PPI use (OR 1.38, 95% CI: 1.18-1.61) compared to no PPI use in the absence of clopidogrel therapy provides further support for this. Another explanation would be that PPIs have a harmful effect irrespective of clopidogrel status. We are aware of two studies that conducted similar analysis in clopidogrel non-users which also showed significantly elevated risks with PPI use (adjusted OR 1.29 and 1.55) compared to PPI non- use ^{23,31}. In two other observational studies however, no significant increase in CV events was shown in patients prescribed PPIs without clopidogrel ^{15,32}. We addressed the confounding issue by all possible design and adjustments measures, but still the comparison against none-use of PPI seemed confounded. Applying past PPI use as active comparator changed the association from increased to no effect and was the most powerful approach to deal with channeling of PPIs, more than adjustment for the propensity score. The propensity score did not have large explanatory power in this study due to the fact that relatively little clinical information was available to construct the score. Our study is internally valid since we used a population-based design, so selection bias was unlikely as all cases and controls came from the same source population. The same pertains to information bias, because data were gathered prospectively without knowledge of <

each hospital by official coding personnel from the national registry of hospitalization discharge records and linked to pharmacies with complete information on outpatient dispensing. In the Netherlands patients usually go to one pharmacy because of billing purposes (invoices being directly sent to insurance without the need to pre-pay by patient) and medication surveillance. The exposure drugs, clopidrogel and PPIs, are dispensed through regular pharmacies. With regard to external validity, the individuals captured in the database are representative for the entire population of the Netherlands in terms of age, gender, socioeconomic background, morbidity and mortality, drug use, and geographic distribution. This study adds to the existing literature because we revealed the difficulties to deal

with confounding by indication in observational studies and we adjusted for confounding in the most optimal way by applying past PPI use as the reference category. Furthermore, few other studies could account for the use of low-dose aspirin as important risk factor for recurrent MI^{19, 20}. We avoided immortal time bias by assigning exposure to the category it belongs to and not using follow-up time to define exposure. Immortal time bias alludes to the fact that the outcome under study cannot be assessed when follow-up time is used to define the exposure status. We used a time varying exposure assessment whereas some other cohort studies applied a fixed exposure status ^{13, 20}, which leads to misclassification of exposure at the index date (e.g. patients were defined at baseline as PPI non-users irrespective of changes in PPI use during follow-up). In addition, we used an European database whereas many published studies so far were conducted in North-America. None of the PPIs were available over the counter during the study period and we have illustrated that there was no difference in the risk of recurrent MI between esomeprazole, omeprazole, and pantoprazole.

There are several important limitations to this analysis. First, we were unable to study
actual drug utilization and adherence to dispensed drugs, because only dispensing data were

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1		
2 3	1	available in the database instead of more reliable proxies for drug use. Second, using the past
4 5	2	PPI group as reference ₂ , we may compare current PPI users to past PPI users who had not
6 7	3	been fully adherent to the prescribed PPI and therefore still had medication available and were
8 9	4	actually also current PPI users at the time of the recurrent MI. To address this point, we
10 11	5	conducted several sensitivity analyses varying the time window to define past PPI use and this
12 13	6	did not affect the estimates. Additionally, although the internal validity may be jeopardized by
14 15	7	non-compliance of the patient, this is unlikely to have affected the comparison between active
16 17	8	compounds (e.g. omeprazole/esomeprazole compared to pantoprazole). Third, since most
17 18 19	9	persons continue PPIs once started, applying past PPI use as reference reduces study power
20 21	10	substantially (resulting in wide confidence intervals). Finally, residual confounding remained
21 22 23	11	due to lack of data on important cardiac risk factors, such as smoking status, lipoproteins, or
24	12	type of coronary stents, which was likely providing an overestimation of the effect measure
25 26 27	13	and this was illustrated when applying PPI non-use as the reference.
28 29	14	In conclusion, this study provides an unique angle on the association between
30 31	15	clopidogrel -PPI co-therapy and risk of recurrent MI by using several techniques to deal with
32 33	16	confounding by indication and other methodological issues. Among clopidogrel users, current
34 35	17	use of PPIs was associated with an increased risk of recurrent MI when compared to PPI non-
36 37	18	use, but not when compared to past PPI use, which shows the magnitude of the confounding
38 39	19	by indication that is present. In clopidogrel-non-users, current PPI use was also associated
40 41	20	with an increased risk of recurrent MI when compared to PPI non-use. This study thus
42 43	21	demonstrates that previously reported associations between PPI-clopidogrel co-therapy and
44 45	22	risks of recurrent MI are influenced by confounding. It further demonstrates that such
46 47	23	confounding can be (at least partly) circumvented by careful selection of study design and
48 49	24	conventional adjustments techniques. Further work should concentrate on well-conducted
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Deleted: omeprazole and pantoprazole were available over the counter since respectively February 2004 and November 2007. We could not control for over the counter use of medication. However, since PPIs are only being reimbursed when on prescription, it is unlikely that we missed much (long-term) PPI use

Deleted: a drawback of Deleted: is that

- prospective randomized trials to dissolve the confusion around this controversial drug-drug
- interaction.

For Beer Berriew

1 2	1	Tables					
3							
4 5	2	Table 1: Characteristics of cohor	rt patients (co	olumn 1), of tho	se readmitted t	o the hospital wi	ith
6 7	3	MI (cases) versus their controls	(column 2 ar	nd 3), and of PP	l users versus n	on-PPI users in	
8 9	4	the case-control setting (column	4 and 5).				
10							
11- 12			Total	Cases with	Controles	PPI users	No PPI
13			n(%)	recurrent	without	n(%)	users
14 15				MI	recurrent		n(%)
16 17				n (%)	MI		II (<i>70</i>)
18				II (<i>%</i>)	IVII		
19 20					n (%)		
21- 22				In unexpos	ed to PPIs	In control.	s without
23						recurre	nt MI
24 25							
23 26							
27	Number	r	23,655	616 (100)	126,817	69,313 (100)	126,817
29			(100)		(100)		(100)
30 31 32	Mean a	ge (SD)	64.7	66.1 (13.2)	66.3 (10.4)	68.1 (10.3)	66.3 (10.4)
32 33	Wiedif a			00.1 (13.2)	00.5 (10.4)	00.1 (10.5)	00.5 (10.4)
34			(13.2)				
	Median	follow-up time in months (IQR)	42.6	11.4 (4.1-	29.5 (13.1-	34.5 (16.0-	29.5 (13.1-
37 38			(16.8-	29.4)	51.6)	58.4)	51.6)
39			71.7)				
40 41	Median	number of different	5 (2 - 9)	5 (2 - 8)	3 (1 - 6)	6 (3 - 10)	3 (1 - 6)
42			0(_))		U (1 U)		U (1 U)
43 44	prescrip	otions (IQR)*					
	Median	length of stay baseline MI in	7 (4 - 10)	7 (5 - 9)	7 (5 - 10)	7 (5 - 10)	7 (5 - 10)
	days (IC	QR)					
40 49							
50							

Male sex	15.007	40C (70 0)	72.020	<i></i>	70.000
Male sex	15,897	436 (70.8)	72,020	55,377	72,020
	(67.2)		(85.1)	(79.9)	(85.1)
PTCA during or within 30 days of	7,889	126 (20.5)	23,808	21,851	23,808
baseline MI	(33.4)		(28.1)	(31.5)	(28.1
Presence of diabetes mellitus	5,100	125 (20.3)	13,774	14,022	13,774
2	(21.6)		(16.3)	(20.2)	(16.3
⁴ Hospitalizations before baseline MI:					
5 Cardiogenic shock 7	12 (0.1)	1 (0.2)	31 (0.0)	37 (0.1)	31 (0.0
3 Congestive heart failure	551 (2.3)	14 (2.3)	858 (1.0)	961 (1.4)	858 (1.0
) Cancer	1,045	13 (2.1)	2,380 (2.8)	2,710 (3.9)	2,380 (2.8
1	(4.4)				
3 4 Cerebrovascular disease	651 (2.8)	11 (1.8)	1,378 (1.6)	1,836 (2.6)	1,378 (1.6
5 Acute renal failure	32 (0.1)	0 (0.0)	49 (0.1)	84 (0.1)	49 (0.1
7 3 Chronic renal failure	210 (0.9)	4 (0.6)	449 (0.5)	667 (1.0)	449 (0.5
) Cardiac dysrhythmia	775 (3.3)	18 (2.9)	1,898 (2.2)	1,960 (2.8)	1,898 (2.2
UGI ulcer and gastritis	153 (0.6)	1 (0.2)	101 (0.1)	630 (0.9)	101 (0.1
3 4 Medication use at the time of baseline					
5 5 MI:			7		
7 Angiotensin-converting-enzyme	3,141	84 (13.6)	9,122 (10.8)	8,601 (12.4)	9,122 (10.8
) j inhibitor	(13.3)		4		
Angiotensin-receptor antagonist	1,787	38 (6.2)	4,200 (5.0)	5,034 (7.3)	4,200 (5.0
3 4	(7.6)				
Acetylsalicylic acid	5,364	153 (24.8)	16,181	17,120	16,18
7 3	(22.7)		(19.1)	(24.7)	(19.1
5 9 0					

1						
2 - 3	β-Adrenergic antagonist	5,531	148 (24.0)	16,915	16,891	16,915
4 5		(23.4)		(20.0)	(24.4)	(20.0)
6 7	Calcium-channel antagonist	3,150	92 (14.9)	9,257 (10.9)	10,031	9,257 (10.9)
8 9		(13.3)			(14.5)	
10 11	Digoxin	592 (2.5)	14 (2.3)	1,204 (1.4)	1,311 (1.9)	1,204 (1.4)
12 13	Spironolactone	336 (1.4)	5 (0.8)	455 (0.5)	559 (0.8)	455 (0.5)
13 14 15	Lipid lowering drugs	4,144	113 (18.3)	13,034	13,307	13,034
16 17		(17.5)		(15.4)	(19.2)	(15.4)
18	Thiazide diuretic	806 (3.4)	16 (2.6)	2,080 (2.5)	1,958 (2.8)	2,080 (2.5)
19 20	Other diuretic, excluding thiazide	2,713	63 (10.2)	5,299 (6.3)	6,094 (8.8)	5,299 (6.3)
21 22		(11.5)				
23 24	Nonsteroidal anti-inflammatory drug	1,720	30 (4.9)	3,910 (4.6)	6,384 (9.2)	3,910 (4.6)
25 26		(7.3)				
	fedication use at index date**:					
29 30	Angiotensin-converting-enzyme		213 (34.6)	29,328	2,4817	29,328
	hibitor			(34.6)	(35.8)	(34.6)
33 34	Angiotensin-receptor antagonist		50 (8.1)	9,301 (11.0)	9,535 (13.8)	9,301 (11.0)
35 36	Acetylsalicylic acid		408 (66.2)	56,621	45,511	56,621
37 38				(66.9)	(65.7)	(66.9)
39 40	β-Adrenergic antagonist		377 (61.2)	52,200	43,575	52,200
41 42				(61.7)	(62.9)	(61.7)
43 44	Calcium-channel antagonist		116 (18.8)	14,252	13,727	14,252
45 46				(16.8)	(19.8)	(16.8)
47 48	Lipid lowering drugs		356 (57.8)	54,475	45,776	54,475
49—						

1					
2 - 3			(64.4)	(66.0)	(64.4)
4 5	Thiazide diuretic	24 (3.9)	2,690 (3.2)	2,672 (3.9)	2,690 (3.2)
6 7	Other diuretic, excluding thiazide	130 (21.1)	12,606	14,510	12,606
8 9			(14.9)	(20.9)	(14.9)
	Aedication use at recurrent MI:				
12 13	Cytochrome P450 2C19 inhibitor(s)	7 (1.1)	679 (0.8)	676 (1.0)	679 (0.8)
14 15	Cytochrome P450 2C19 inducer(s)	13 (2.1)	1,043 (1.2)	2,973 (4.3)	1,043 (1.2)
16 16 17	Cytochrome P450 3A4 inhibitor(s)	56 (9.1)	5,725 (6.8)	5,854 (8.4)	5,725 (6.8)
18 19_	Cytochrome P450 3A4 inducer(s)	8 (1.3)	823 (1.0)	724 (1.0)	823 (1.0)

20 1 When cases differ statistically significant from controls or when PPI users differ statistically 21

2 significant from non-PPI users, the figures are depicted in bold. The numbers in columns
 23

three to six (reflecting the case-control setting) can be larger than the total number in column

26 4 two (reflecting the cohort setting) due to the control sampling method.

5 * Median number of different prescriptions on full ATC-level one year prior to cohort entry

6 ** Index date=date of recurrent MI. Only available for the case-control set.

 1 Table 2: Association between current clopidogrel plus current PPI and recurrent MI compared

2 to clopidogrel without PPI and to clopidogrel with past PPI.

	Cases	Controls	OR _{matched}	OR _{adjusted}
	n=1,224	n=153,967	(95% CI)	(95% CI)
Current clopidogrel / No PPI	90	11,147	1 (ref)	1 (ref)
Current clopidogrel / Current PPI	78	4,715	1.89 (1.37-2.63)	1.62 (1.15-2.27)
Other*	1,056	138,105		
Current clopidogrel / Past PPI	6	436	1 (ref)	1 (ref)
Current clopidogrel / Current PPI	78	4,715	1.15 (0.46-2.86)	0.95 (0.38-2.41):
Other**	1,140	148,816		
*No or past clopidogrel, or past PP **No or past clopidogrel, or no PPI				
# Adjusted for follow-up time in da	ays and tota	l number of pr	escriptions one yea	r prior to
baseline MI.				
± Adjusted for follow-up time in da	ays.			
				25
				23

1 Table 3: Association between current PPI use and recurrent MI compared to no PPI and to

2 past PPI use in the absence of clopidogrel.

	Cases	Controls	OR _{matched}	OR _{adjusted}
	n=1,224	n=153,967	(95% CI)	(95% CI)
No clopidogrel / No PPI	766	110,002	1 (ref)	1 (ref)
No clopidogrel / Current PPI	237	23,335	1.56 (1.34-1.81)	1.38 (1.18-1.61)
Other*	221	20,630		
No clopidogrel / Past PPI	25	2,941	1 (ref)	1 (ref)
No clopidogrel / Current PPI	237	23,335	1.22 (0.79-1.88)	1.22 (0.79-1.88)
Other**	962	127,691		
*Current or past clopidogrel, or p	oast PPI use.			
**Current or past clopidogrel, or	no PPI use.			
# Adjusted for total number of pr	escriptions or	ne year prior to	baseline MI.	
χ No adjustments.				
				26

1 2 1 <i>Table 4:</i> Association betw	ween current clopic	dogrel plus specific o	current PPI use and r	ecurrent
2 MI compared to clopidog	_			
<u>;</u>	Cases	Controls	OR _{matched}	OR _{adjusted} **
3) O	n=1,224	n=153,967	(95% CI)	(95% CI)
1 Gurrent clopidogrel + pantoprazole	36	2,271	1 (ref)	1 (ref)
3 4 5 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9	26	1,339	1.03 (0.57-1.84)	1.07 (0.58-1.99)
6 Current clopidogrel+ esomeprazole 8	13	827	0.96 (0.48-1.92)	0.83 (0.40-1.69)
9 0 urrent clopidogrel + lanso- or	2	186	-	-
1 Pabeprazole 3				
24 Gther*	1,147	149,344		
6 7				
8 3 Combinations of PPIs we	ere excluded.			
9				
0 4 *No or past clopidogrel	or no or past PPI u	se		
1				
2				
3 5 **Adjusted for follow-up	p time in days and	propensity score.		
4 5				
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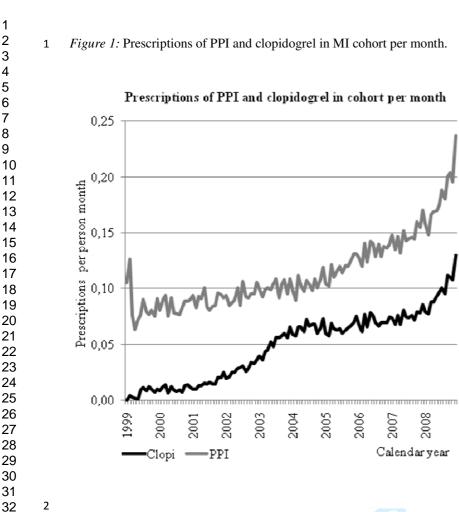
1	Table 5: Summary	v of used	methods to	deal with	confounding	in obser	vational	studies
-	<i>Tuble 5.</i> Summu	y or used	methous to	ucui wittii	comountaing	111 000501	valionai	studies

- 2 assessing the effect of PPI on CV events in clopidogrel users, excluding restrictions in study
- 3 design.

Study population Methods to deal with confounding Authors Study type Outcome Patients with - Multivariable adjustments Simon Prospective Death, MI, stroke et al. 2 cohort acute MI - Propensity analysis for CYP2C19 loss-of-function-alleles which was used to match 5 controls for each patient with 2 variant alleles Double-blind Bhatt et Patients with MI, stroke, - Randomization al. ⁶ CABG, PCI, CV randomized trial ACS undergoing coronary stent death placement Ray et Retrospective Patients serious CV - Multivariable adjustments al. 7 cohort study hospitalized for disease - Propensity score for the use of PPI MI, coronary ((non)fatal MI, was converted in deciles and included stroke, other CV artery as variable in model revascularization, death). or unstable AP on clopidogrel Patients with CV death, MI, O'Dono Retrospective - Multivariable adjustments cohort within ACS undergoing stroke ghue et - Propensity score for the use of PPI and

al. ¹³	RCT	PCI		strata matched on this score
II	Detres estima	Dationto mith	A11	
Ho et	Retrospective	Patients with	All-cause	- Multivariable adjustments
al. ¹⁵	cohort study +	ACS on	mortality,	- Matched on duration of follow-up
	case-control	clopidogrel	rehospitalization	
	study		for ACS	
Juurlink	Nested case-	Patients with	Recurrent MI,	- Multivariable adjustments
et al. ¹⁶	control	acute MI on	death	- Matched on age, PCI, date of hospita
		clopidogrel		discharge, predicted probability of
				short-term mortality
Pezalla	Retrospective	Patients adherent	Acute MI	- Subgroup analysis; restriction to
<i>et al</i> . ¹⁷	cohort study	to clopidogrel		patients who all had diagnosis of
				ischaemic heart disease, congestive
				heart failure, hypertension,
				hyperlipidemia, and diabetes before the
				start of clopidogrel therapy.
Stanek	Retrospective	Patients adherent	CV events (MI,	- Multivariable adjustments
et al. ¹⁸	cohort study	to clopidogrel	unstable AP,	
		following	TIA/stroke,	
		coronary stenting	coronary	
			vascularization,	
			CV death)	
			,	
				29

1						
2 3	Rassen		Retrospective	Patients	MI,	- Multivariable adjustments
4 5	et al. ¹⁹		cohort study	undergoing PCI	revascularization,	- High-dimensional propensity score
6 7				or hospitalized	all-cause	and strata matched on this score (1:1)
8 9				for ACS on	mortality	
10 11				clopidogrel		
12 13 14	Stockl		Retrospective	Patients with MI	Acute MI	- Propensity score and strata matched on
15	<i>et al.</i> ²⁰		cohort study	or coronary stent		this score (1:1)
16 17 18				on clopidogrel		
19 20	Collet		Prospective	Survivors of MI	CV death, non-	- Multivariable adjustments
21 22	<i>et al.</i> ²²		cohort	(< 45 yrs)	fatal MI, urgent	
23 24 25					revascularisation	
26	Dunn e	t	Retrospective	Patients	Death, MI,	- No information on adjustment to
27 28 29	al. ²³		cohort within	undergoing PCI	Stroke (1 year)	covariates
30 31			RCT			
32 33	1 1	MI	Myocardial infarc	tion, ACS: Acute Co	oronary Syndrome, C	CABG: Coronary artery bypass
34 35	2 ક	grai	ft, PCI: Percutaneo	us coronary interven	tion, CV: Cardiovas	cular, AP: Angina pectoris,
36 37	3 I	PPI	: Proton pump inhi	bitor, RCT: Random	ized controlled trial	, TIA: Transient ischemic
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Number of prescriptions per person month for PPIs and clopidogrel calculated per calendar

month during study period in cohort of acute MI patients (n=23,655).

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