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EIGENVECTORS OF WIGNER MATRICES: UNIVERSALITY OF GLOBAL FLUCTUATIONS

FLORENT BENAYCH-GEORGES

ABSTRACT. We prove that for $[u_{i,j}]_{i,j=1}^n$ the eigenvectors matrix of a Wigner matrix, under some moments conditions, the bivariate random process

$$\left(\sum_{\substack{1 \leq i \leq ns, \\ 1 \leq j \leq nt}} (|u_{i,j}|^2 - 1/n) \right)_{(s,t) \in [0,1]^2}$$

converges in distribution to a bivariate Brownian bridge.

1. INTRODUCTION

It is well known that the matrix $U_n = [u_{i,j}]_{i,j=1}^n$ whose columns are the eigenvectors of a GOE or GUE matrix X_n can be chosen to be distributed according to the Haar measure on the orthogonal or unitary group. As a consequence, much can be said about the $u_{i,j}$'s: their joint moments can be computed via the so-called *Weingarten calculus* developed in [8, 9], any finite (or not too large) set of $u_{i,j}$'s can be approximated, as $n \rightarrow \infty$, by independent Gaussian variables (see [17, 7]) and the global asymptotic fluctuations of the $|u_{i,j}|$'s are governed by a theorem of Donati-Martin and Rouault, who proved in [10] that as $n \rightarrow \infty$, the bivariate càdlàg process

$$\left(B_{s,t}^n := \sqrt{\frac{\beta}{2}} \sum_{\substack{1 \leq i \leq ns, \\ 1 \leq j \leq nt}} (|u_{i,j}|^2 - 1/n) \right)_{(s,t) \in [0,1]^2}$$

(where $\beta = 1$ in the real case and $\beta = 2$ in the complex case) converges in distribution, for the Skorokhod topology, to the *bivariate Brownian bridge*, i.e. the centered continuous Gaussian process $(B_{s,t})_{(s,t) \in [0,1]^2}$ with covariance

$$(1) \quad \mathbb{E}[B_{s,t} B_{s',t'}] = (\min\{s, s'\} - ss')(\min\{t, t'\} - tt').$$

A natural question is the following:

What can be said beyond the Gaussian case, when the entries of the Wigner matrix X_n are general random variables ?

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For a general Wigner matrix¹, the exact distribution of the matrix U_n cannot be computed and few works had been devoted to this subject until quite recently. One of the reasons is that while the eigenvalues of an Hermitian matrix admit variational characterizations as *extremums* of certain functions, the eigenvectors can be characterized as the *argmax* of these functions, hence are more sensitive to perturbations of the entries of the matrix. However, in the last three years, the eigenvectors of general Wigner matrices have been the object of a growing interest, due in part to some relations with the universality conjecture for the eigenvalues. In several papers (see, among others, [13, 14, 15]), a *delocalization property* was shown for the eigenvectors of random matrices. More recently, Knowles and Yin in [19] and Tao and Vu in [25] proved that if the first four moments of the atom distributions of X_n coincide with the ones of a GOE or GUE matrix, then under some tail assumptions on these distributions, the $u_{i,j}$'s can be approximated by independent Gaussian variables as long as we only consider a finite (or not too large) set of $u_{i,j}$'s.

In this paper, we consider the global behavior of the $|u_{i,j}|$'s, and we prove (Theorem 2.3) that for Wigner matrices whose entries have moments of all orders, the process $(B_{s,t}^n)_{(s,t) \in [0,1]^2}$ has a limit in a weaker sense than for the Skorokhod topology and that this weak limit is the bivariate Brownian bridge if and only if the off-diagonal entries of the matrix have the same fourth moment as the GOE or GUE matrix (quite surprisingly, no hypothesis on the third moment is necessary). Under some additional hypotheses on the atom distributions (more coinciding moments and continuity), we prove the convergence for the Skorokhod topology (Theorem 2.6).

This result was conjectured by Djalil Chafaï, who also conjectures the same kind of universality for unitary matrices appearing in other standard decompositions, such as the singular values decomposition or the Housholder decomposition of matrices with no symmetry, as long as the matrix considered has i.i.d. entries with first moments agreeing with the ones of Gaussian variables. It would also be interesting to consider the same type of question in the context of *band matrices*, connecting this problem with the so-called *Anderson conjecture* (see *e.g.* the works of Erdős and Knowles [11, 12], of Schenker [20] or of Sodin [22], or, for a short introduction, the blog note by Chafaï [6]).

The paper is organized as follows. The main results are stated in Section 2, where we also make some comments on their hypotheses ; outlines of the proofs and the formal proofs are given in Section 3 ; and Section 4 is devoted to the definitions of the functional spaces and their topologies and to the proofs of several technical results needed in Section 3.

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2. MAIN RESULTS

For each n , let us consider a real symmetric or Hermitian random matrix

$$X_n := \frac{1}{\sqrt{n}} [x_{i,j}^{(n)}]_{i,j=1}^n.$$

¹A *Wigner matrix* is a real symmetric or Hermitian random matrix with independent, centered entries whose variance is one. Its *atom distributions* are the distributions of its entries.

For notational brevity, $x_{i,j}^{(n)}$ will be denoted by $x_{i,j}$.

Let us denote by $\lambda_1 \leq \dots \leq \lambda_n$ the eigenvalues of X_n and consider an orthogonal or unitary matrix $U_n = [u_{i,j}]_{i,j=1}^n$ such that

$$X_n = U_n \operatorname{diag}(\lambda_1, \dots, \lambda_n) U_n^*.$$

Note that U_n is not uniquely defined. However, one can choose it in any measurable way.

We define the bivariate càdlàg process

$$\left(B_{s,t}^n := \sqrt{\frac{\beta}{2}} \sum_{\substack{1 \leq i \leq ns, \\ 1 \leq j \leq nt}} (|u_{i,j}|^2 - 1/n) \right)_{(s,t) \in [0,1]^2},$$

where $\beta = 1$ in the real case and $\beta = 2$ in the complex case.

Assumption 2.1. *For each n , the random variables $x_{i,j}$'s are independent (up to the symmetry), with the same distribution on the diagonal and the same distribution above the diagonal.*

Assumption 2.2. *For each $k \geq 1$, $\sup_n \mathbb{E}[|x_{1,1}|^k + |x_{1,2}|^k] < \infty$. Moreover,*

$$(2) \quad \mathbb{E}[x_{1,1}] = \mathbb{E}[x_{1,2}] = 0, \quad \mathbb{E}[|x_{1,2}|^2] = 1$$

and $\mathbb{E}[|x_{1,2}|^4]$ has a limit as $n \rightarrow \infty$.

The bivariate Brownian bridge has been defined in the introduction and the definitions of the functional spaces and their topologies can be found in Section 4.1.

Theorem 2.3. *Suppose that Assumptions 2.1 and 2.2 are satisfied. Then the sequence*

$$(\text{distribution}(B^n))_{n \geq 1}$$

has a unique possible accumulation point supported by $C([0,1]^2)$. This accumulation point is the distribution of a centered Gaussian process which depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$, and which is the bivariate Brownian bridge if and only if $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4] = 4 - \beta$.

More precisions about the way the unique possible accumulation point depends on the fourth moment of the entries are given in Remark 2.8.

To get a stronger statement where the convergence in distribution to the bivariate Brownian bridge is actually stated, one needs stronger hypotheses.

Assumption 2.4. *The distributions of the entries of X_n are absolutely continuous with respect to the Lebesgue measure.*

The following hypothesis depends on an integer $m \geq 2$.

Assumption 2.5. *For each $k \geq 1$, $\sup_n \mathbb{E}[|x_{1,1}|^k + |x_{1,2}|^k] < \infty$. Moreover, for each $r, s \geq 0$,*

$$(3) \quad r + s \leq m - 2 \implies \mathbb{E}[\Re(x_{1,1})^r \Im(x_{1,1})^s] = \mathbb{E}[\Re(g_{1,1})^r \Im(g_{1,1})^s]$$

and

$$(4) \quad r + s \leq m \implies \mathbb{E}[\Re(x_{1,2})^r \Im(x_{1,2})^s] = \mathbb{E}[\Re(g_{1,2})^r \Im(g_{1,2})^s].$$

where the $g_{i,j}$'s are the entries of a standard GOE or GUE matrix.

Theorem 2.6. *Suppose that Assumptions 2.1 and 2.4 are satisfied, as well as Assumption 2.5 for $m = 12$. Then, as $n \rightarrow \infty$, the bivariate process B^n converges in distribution, for the Skorokhod topology in $D([0, 1]^2)$, to the bivariate Brownian bridge.*

Remark 2.7. Weakening of the assumptions. As explained above, the distributions of the entries $x_{i,j} = x_{i,j}^{(n)}$ are allowed to depend on n . In order to remove Assumption 2.4 below (which we did not manage to do yet), it might be useful to weaken Assumptions 2.2 and 2.5 in the following way: one can easily see that the proofs of Theorems 2.3 and 2.6 and Proposition 2.10 still work if, in Equations (2), (3) and (4), one replaces the identities by the same identities up to an error which is $O(n^{-\alpha})$ for all $\alpha > 0$.

Remark 2.8. Complements on Theorem 2.3. One can wonder how the unique accumulation point mentioned in Theorem 2.3 depends on the fourth moment of the entries of X_n . Let $G := (G_{s,t})_{(s,t) \in [0,1]^2}$ be distributed according to this distribution. We know that $(G_{s,t})_{(s,t) \in [0,1]^2}$ is the bivariate Brownian bridge only in the case where $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4] = 4 - \beta$. In the other cases, defining $F_{\text{semicircle}}$ as the cumulative distribution function of the semicircle law, the covariance of the centered Gaussian process

$$(5) \quad \left(\int_{u=-2}^2 u^k G_{s, F_{\text{semicircle}}(u)} du \right)_{s \in [0,1], k \geq 0}$$

can be computed thanks to Lemma 3.4 and Proposition 3.5. By Lemma 3.3, it determines completely the distribution of the process G . However, making the covariance of G explicit out of the covariance of the process of (5) is a very delicate problem, and we shall only stay at a quite vague level, saying that it can be deduced from the proof of Proposition 3.5 that the variances of the one-dimensional marginals of G are increasing functions of $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$. More specifically, one can deduce from Lemma 3.4 and Proposition 3.5 that for all $0 \leq s_1, s_2 \leq 1$,

$$\text{Cov} \left(\int_{u=-2}^2 u^2 G_{s_1, F_{\text{semicircle}}(u)} du, \int_{u=-2}^2 u^2 G_{s_2, F_{\text{semicircle}}(u)} du \right) = \frac{\mathbb{E}[|x_{1,2}|^4] - 1}{4} (\min\{s_1, s_2\} - s_1 s_2).$$

Remark 2.9. Comments on the hypotheses of Theorem 2.6 (1). In order to prove the convergence in the Skorokhod topology, we had to make several hypotheses on the atom distributions: absolute continuity, moments of all orders and coincidence of their 10 (on the diagonal) and 12 (above the diagonal) first moments with the ones of a GOE or GUE matrix. We needed these assumptions to control the discontinuities of the process B^n . Even though these hypotheses might not be optimal (especially the continuity one), a bound on the tails of the atom distributions seems to be necessary to avoid too large variations of the process B^n . Indeed, as illustrated by Figure 1, for a GOE matrix (*left picture*), $|u_{i,j}|^2$ is close to $1/n$ for all i, j with high probability, whereas when the atom distributions have not more than a second moment (*right picture*), the matrix X_n looks more like a *sparse matrix*, and so does U_n , which implies that for certain (i, j) 's, $|u_{i,j}|^2 - 1/n$ is not small enough. Since $|u_{i,j}|^2 - 1/n$ is the jump of the process B^n at $(s, t) = (i/n, j/n)$, this could be an obstruction to the existence of a continuous limit for the process B^n . That being said, we have hopes to prove the theorem under Assumption 2.5 for $m = 4$ instead of $m = 12$ (see Remark 2.11 below).

Note that it follows from the previous theorem that for all $0 \leq s < s' \leq 1$ and $0 \leq t < t' \leq 1$, the sequence of random variables

$$\frac{1}{\sqrt{(s' - s)(t' - t)}} \sum_{\substack{ns < i \leq ns' \\ nt < j \leq nt'}} (|u_{i,j}|^2 - 1/n)$$

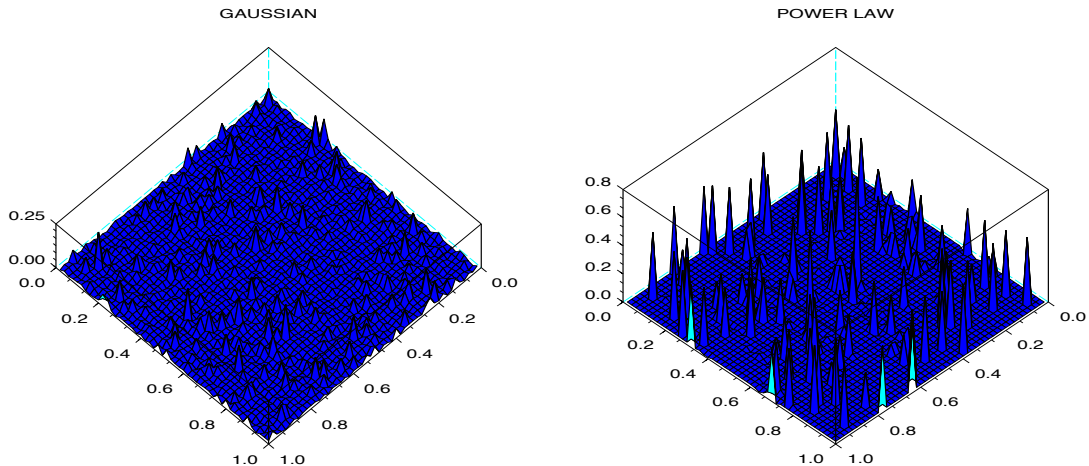


FIGURE 1. **Influence of the tails of the atom distributions of X_n on the $|u_{i,j}|^2$'s:** Plot of the map $(i/n, j/n) \mapsto ||u_{i,j}|^2 - 1/n|$ for two different choices of atom distributions. *Left:* GOE matrix. *Right:* Wigner matrix with atom distribution admitting moments only up to order $2 + \varepsilon$ for a small ε . For both pictures, the matrices are $n \times n$ with $n = 50$.

admits a limit in distribution as $n \rightarrow \infty$, hence is bounded in probability (in the sense of [23, Def. 1.1]). In the same way, it follows from [19] and [25] that the sequence $n|u_{i,j}|^2 - 1$ is bounded in probability. In the next proposition, we improve these assertions by making them uniform on s, s', t, t', i, j and upgrading them to the L^2 and L^4 levels. The proof of the proposition is postponed to Section 4.5.

Proposition 2.10. *Suppose that Assumptions 2.1, 2.4 and 2.5 for $m = 4$ are satisfied. Then as $n \rightarrow \infty$, the sequences*

$$(6) \quad n|u_{i,j}|^2 - 1 \quad \text{and} \quad \frac{1}{\sqrt{(s' - s)(t' - t)}} \sum_{\substack{ns < i \leq ns' \\ nt < j \leq nt'}} (|u_{i,j}|^2 - 1/n)$$

are bounded for the respective L^4 and L^2 norms, uniformly in $s < s', t < t', i, j$.

Remark 2.11. Comments on the hypotheses of Theorem 2.6 (2). This proposition is almost sufficient to apply the tightness criterion that we use in this paper. Would the second term of (6) have been bounded for the $L^{2+\varepsilon}$ norm (instead of L^2), Assumption 2.5 for $m = 4$ would have been enough to prove that B^n converges in distribution, for the Skorokhod topology in $D([0, 1]^2)$, to the bivariate Brownian bridge.

3. PROOFS OF THEOREMS 2.3 AND 2.6

3.1. Outline of the proofs. Firstly, Theorem 2.6 can be deduced from Theorem 2.3 by proving that the sequence $(\text{distribution}(B^n))_{n \geq 1}$ is tight and only has $C([0, 1]^2)$ -supported accumulation points. This can be done via some upper bounds on the fourth moment of the increments of B^n and on its jumps (or discontinuities). These bounds are given in the proof of Lemma 3.1 below, and rely on the existing bounds in the case where X_n is a GOE or GUE matrix and on

the “one-by-one entries replacement method” developed by Terence Tao and Van Vu in recent papers, such as [24, 25].

Secondly, the proof of Theorem 2.3 relies on the following remark, inspired by some ideas of Jack Silverstein (see [2, Chap. 10] and [21]): even though we do not have any “direct access” to the eigenvectors of X_n , we have access to the process $(B_{s, F_{\mu_{X_n}}^n}^n)_{s \in [0,1], u \in \mathbb{R}}$, for $F_{\mu_{X_n}}(u) := \frac{1}{n} \#\{i; \lambda_i \leq u\}$. Indeed,

$$B_{s, F_{\mu_{X_n}}^n}^n(u) = \sqrt{\frac{\beta}{2}} \sum_{1 \leq i \leq ns} \sum_{\substack{1 \leq j \leq n \\ \text{s.t. } \lambda_j \leq u}} (|u_{i,j}|^2 - 1/n),$$

hence for all fixed $s \in [0, 1]$, the function $u \in \mathbb{R} \mapsto B_{s, F_{\mu_{X_n}}^n}^n(u)$ is the cumulative distribution function of the signed measure

$$(7) \quad \sqrt{\frac{\beta}{2}} \sum_{1 \leq i \leq ns} \sum_{j=1}^n (|u_{i,j}|^2 - 1/n) \delta_{\lambda_j},$$

which can be studied via its moments

$$\sum_{1 \leq i \leq ns} \left(e_i^* X_n^k e_i - \frac{1}{n} \text{Tr} X_n^k \right) \quad (k \geq 1),$$

the e_i 's being the vectors of the canonical basis. From the asymptotic behavior of the *moments* of the signed measure of (7), one can then find out the asymptotic behavior of its *cumulative distribution function*.

Once the asymptotic distribution of the process $(B_{s, F_{\mu_{X_n}}^n}^n)_{s \in [0,1], u \in \mathbb{R}}$ identified, one can obtain the asymptotic distribution of the process $(B_{s,t}^n)_{s \in [0,1], t \in [0,1]}$ because the function $F_{\mu_{X_n}}$ tends to the (non random) cumulative distribution function $F_{\text{semicircle}}$ of the semicircle law.

3.2. Formal proofs. By a standard *tightness + uniqueness of the accumulation point* argument, Theorem 2.6 will follow from the following lemma and Theorem 2.3. The proof of the lemma is postponed to Section 4.5.

Lemma 3.1. *Under Assumptions 2.1, 2.4 and 2.5 for $m = 12$, the sequence $(\text{distribution}(B^n))_{n \geq 1}$ is C -tight, i.e. is tight and has only $C([0, 1]^2)$ -supported accumulation points.*

So let us now prove Theorem 2.3 under Assumptions 2.1 and 2.2. Note that it suffices to prove that the sequence $(\text{distribution}(B^n))_{n \geq 1}$ has a unique possible accumulation point supported by $C([0, 1]^2)$ and that this accumulation point is the distribution of a centered Gaussian process which depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$ (and actually does depend on $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$). Indeed, in this case, by Theorem 1.1 of [10], where the case of GOE and GUE matrices is treated, this limit distribution is the bivariate Brownian bridge if and only if $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$ is the same as for X_n a GOE or GUE matrix, i.e. equal to $4 - \beta$.

The following proposition is the key of the proof, since it allows to transfer the problem from the eigenvectors to some more accessible objects: the weighted spectral distributions of X_n .

Proposition 3.2. *Let e_i denote the i th vector of the canonical basis. To prove Theorem 2.3, it suffices to prove that each finite-dimensional marginal distribution of the process*

$$\left(\sum_{1 \leq i \leq ns} \left(e_i^* X_n^k e_i - \frac{1}{n} \operatorname{Tr} X_n^k \right) \right)_{s \in [0,1], k \geq 1}$$

converges to a centered Gaussian measure and that the covariance of the limit process depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$ and actually does depend on $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$.

Proof. Let $\mu_{X_n} := \frac{1}{n} \sum_{i=1}^n \delta_{\lambda_i}$ be the empirical spectral law of X_n and $F_{\mu_{X_n}}(u) := \frac{1}{n} \#\{i; \lambda_i \leq u\}$ be its cumulative distribution function. It is a well known result [2, Th. 2.5] that $F_{\mu_{X_n}}$ converges in probability, as n tends to infinity, to the cumulative distribution function $F_{\text{semicircle}}$ of the semicircle law in the space $D_c(\mathbb{R}, [0, 1])$ (see Section 4.1 for the definitions of the functional spaces and their topologies). Moreover, the map

$$\begin{aligned} D_0([0, 1]^2) \times D_c(\mathbb{R}, [0, 1]) &\longrightarrow D_c([0, 1] \times \mathbb{R}) \\ ((G_{s,t})_{s,t \in [0,1]}, (g(u))_{u \in \mathbb{R}}) &\longmapsto (G_{s,g(u)})_{(s,u) \in [0,1] \times \mathbb{R}} \end{aligned}$$

is continuous at any pair of continuous functions. Hence for any continuous process $(B_{s,t})_{s,t \in [0,1]}$ whose distribution is an accumulation point of the sequence $(\text{distribution}(B^n))_{n \geq 1}$ for the Skorohod topology in $D([0, 1]^2)$, the process

$$(B_{s, F_{\mu_{X_n}}^n}^n)_{(s,u) \in [0,1] \times \mathbb{R}}$$

converges in distribution (up to the extraction of a subsequence) to the process

$$(B_{s, F_{\text{semicircle}}}^n)_{(s,u) \in [0,1] \times \mathbb{R}}$$

(for the reader who is not used to weak convergence of probability measures, this assertion relies on two results which can be found in [5]: Theorem 4.4 and Corollary 1 of Theorem 5.1 in Chapter 1). Now, note that $F_{\text{semicircle}} : \mathbb{R} \rightarrow [0, 1]$ admits a right inverse, so the distribution of the process $(B_{s,t})_{s,t \in [0,1]}$ is entirely determined by the one of the process $(B_{s, F_{\text{semicircle}}(u)})_{(s,u) \in [0,1] \times \mathbb{R}}$.

As a consequence, to prove Theorem 2.3, it suffices to prove that the sequence

$$(8) \quad \left(\text{distribution}((B_{s, F_{\mu_{X_n}}^n}^n)_{(s,u) \in [0,1] \times \mathbb{R}}) \right)_{n \geq 1}$$

has a unique possible $C_c([0, 1] \times \mathbb{R})$ -supported accumulation point as $n \rightarrow \infty$.

Now, note that any $f \in C_c([0, 1] \times \mathbb{R})$ is entirely determined by the collection of real numbers $(\int_{u \in \mathbb{R}} u^k f(s, u) du)_{s \in [0,1], k \geq 0}$. Let us prove that this fact remains true in the ‘‘distribution sense’’ in the case where f is random. More precisely, let us prove the following lemma.

Lemma 3.3. *Let f be a random variable taking values in $C_c([0, 1] \times \mathbb{R})$ such that with probability one, $f(s, u) = 0$ when $|u| > 2$. Then the distribution of f is entirely determined by the finite dimensional marginals of the process*

$$(9) \quad \left(\int_{u \in \mathbb{R}} u^k f(s, u) du \right)_{s \in [0,1], k \geq 0}.$$

Moreover, in the case where the finite dimensional marginals of the process of (9) are Gaussian and centered, then so are the ones of f .

Proof. Let us fix $(s, u_0) \in [0, 1] \times [-2, 2]$ and let, for each $p \geq 1$, $(P_{p,q})_{q \geq 1}$ be a sequence of polynomials that is uniformly bounded on $[-3, 3]$ and that converges pointwise to $\mathbb{1}_{[u_0, u_0+1/p]}$ on $[-3, 3]$. Then one has, with probability one,

$$f(s, u_0) = \lim_{p \rightarrow \infty} p \int_{u_0}^{u_0+1/p} f(s, u) du = \lim_{p \rightarrow \infty} \lim_{q \rightarrow \infty} p \int_{u \in \mathbb{R}} P_{p,q}(u) f(s, u) du.$$

This proves the lemma, because any almost sure limit of a sequence of variables belonging to a space of centered Gaussian variables is Gaussian and centered. \square

Since the fourth moment of the entries of X_n is bounded, by e.g. [2, Th. 5.1], we know that the extreme eigenvalues of X_n converge to -2 and 2 . As a consequence, for any random variable f taking values in $C_c([0, 1] \times \mathbb{R})$ such that the distribution of f is a limit point of the sequence of (8), we know that with probability one, $f(s, u) = 0$ when $|u| > 2$.

As a consequence, it follows from the previous lemma and from what precedes that to prove Theorem 2.3, it suffices to prove that each finite dimensional marginal distribution of the process

$$\left(\int_{u \in \mathbb{R}} u^k B_{s, F_{\mu_{X_n}}^n}(u) du \right)_{s \in [0, 1], k \geq 0}$$

converges to a centered Gaussian measure and that the covariance of the limit process depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$ and actually does depend on $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$.

But for all $(s, u) \in [0, 1] \times \mathbb{R}$,

$$B_{s, F_{\mu_{X_n}}^n}(u) = \sqrt{\frac{\beta}{2}} \sum_{1 \leq i \leq ns} \sum_{\substack{1 \leq j \leq n \\ \text{s.t. } \lambda_j \leq u}} (|u_{i,j}|^2 - 1/n),$$

hence

$$(10) \quad B_{s, F_{\mu_{X_n}}^n}(u) = \sqrt{\frac{\beta}{2}} \sum_{1 \leq i \leq ns} F_{\mu_{X_n, e_i} - \mu_{X_n}}(u),$$

where:

- μ_{X_n} is still the empirical spectral law of X_n ,
- μ_{X_n, e_i} is the weighted spectral law of X_n , defined by $\mu_{X_n, e_i} := \sum_{j=1}^n |u_{i,j}|^2 \delta_{\lambda_j}$,
- $F_{\mu_{X_n, e_i} - \mu_{X_n}}$ is the cumulative distribution function of the null-mass signed measure $\mu_{X_n, e_i} - \mu_{X_n}$.

The following lemma will allow to conclude the proof of Proposition 3.2.

Lemma 3.4. *Let μ be a compactly supported null-mass signed measure and set $F_\mu(u) := \mu((-\infty, u])$. Then for all $k \geq 0$,*

$$\int_{u \in \mathbb{R}} u^k F_\mu(u) du = - \int_{x \in \mathbb{R}} \frac{x^{k+1}}{k+1} d\mu(x).$$

Proof. Let $a < b$ be such that the support of μ is contained in the open interval (a, b) . F_μ is null out of (a, b) and satisfies $F_\mu(u) = -\mu((u, b))$, so by Fubini's Theorem,

$$\int_{u \in \mathbb{R}} u^k F_\mu(u) du = \int_{x=a}^b \int_{u=a}^x -u^k du d\mu(x) = \int_{x=a}^b \frac{-x^{k+1}}{k+1} d\mu(x) + \frac{a^{k+1}}{k+1} \mu((a, b)) = \int_{x=a}^b \frac{-x^{k+1}}{k+1} d\mu(x).$$

□

It follows from this lemma and from (10) that for all $s \in [0, 1]$, $k \geq 0$,

$$\int_{u \in \mathbb{R}} u^k B_{s, F^\mu X_n}^n(u) du = \frac{-1}{k+1} \sqrt{\frac{\beta}{2}} \sum_{1 \leq i \leq ns} \left(e_i^* X_n^{k+1} e_i - \frac{1}{n} \text{Tr} X_n^{k+1} \right),$$

which proves Proposition 3.2. □

Now, Theorem 2.3 is a direct consequence of the following proposition, whose proof is postponed below to Section 4.2.

Proposition 3.5. *Under Assumptions 2.1 and 2.2, each finite-dimensional marginal distribution of the process*

$$\left(\sum_{1 \leq i \leq ns} (e_i^* X_n^k e_i - \frac{1}{n} \text{Tr}(X_n^k)) \right)_{s \in [0, 1], k \geq 1}$$

converges to a centered Gaussian measure. The covariance of the limit distribution, denoted by

$$(\text{Cov}_{s_1, s_2}(k_1, k_2))_{s_1, s_2 \in [0, 1], k_1, k_2 \geq 1},$$

depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$. Moreover, we have

$$\text{Cov}_{s_1, s_2}(2, 2) = \left(\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4] - 1 \right) (\min\{s_1, s_2\} - s_1 s_2).$$

4. PROOFS OF THE TECHNICAL RESULTS

4.1. Functional spaces and associated topologies. In this paper, we use several functional spaces:

- $C([0, 1]^2)$ (resp. $C_c([0, 1] \times \mathbb{R})$) is the set of continuous functions on $[0, 1]^2$ (resp. *compactly supported* continuous function on $[0, 1] \times \mathbb{R}$), endowed with the uniform convergence topology.
- $D_c(\mathbb{R}, [0, 1])$ is the set of *compactly supported* càdlàg functions on \mathbb{R} taking values in $[0, 1]$, endowed with the topology defined by the fact that $f_n \rightarrow f$ if and only if the bounds of the support of f_n tend to the ones of the support of f and for all $M > 0$, after restriction to $[-M, M]$, $f_n \rightarrow f$ (the topology of $D([-M, M])$ being deduced from the one of $D([0, 1])$ defined in [5]).
- $D([0, 1]^2)$ (resp. $D_c([0, 1] \times \mathbb{R})$) is the set of functions $f : [0, 1]^2 \rightarrow \mathbb{R}$ (resp. of *compactly supported* functions $f : [0, 1] \times \mathbb{R} \rightarrow \mathbb{R}$) admitting limits in all “orthants”, more precisely such that for each s_0, t_0 , for each pair of symbols $\diamond, \diamond' \in \{<, \geq\}$,

$$\lim_{\substack{s \diamond s_0 \\ t \diamond' t_0}} f(s, t)$$

exists, and is equal to $f(s_0, t_0)$ if both \diamond and \diamond' are \geq . The space $D([0, 1]^2)$ is endowed with the Skorokhod topology defined in [4] and the space $D_c([0, 1] \times \mathbb{R})$ is endowed with the topology defined by $f_n \rightarrow f$ if and only if for all $M > 0$, after restriction to $[0, 1] \times [-M, M]$, $f_n \rightarrow f$ (the topology of $D([0, 1] \times [-M, M])$ being deduced from the one of $D([0, 1]^2)$).

- $D_0([0, 1]^2)$ is the set of functions in $D([0, 1]^2)$ vanishing at the border of $[0, 1]^2$, endowed with the induced topology.

4.2. Proof of Proposition 3.5. Note that by invariance of the law of X_n under conjugation by any permutation matrix, the expectation of the random law μ_{X_n, e_i} does not depend on i . So for all $s \in [0, 1]$, $k \geq 1$,

$$(11) \quad \sum_{1 \leq i \leq ns} (e_i^* X_n^k e_i - \frac{1}{n} \text{Tr}(X_n^k)) = \sum_{1 \leq i \leq ns} (e_i^* X_n^k e_i - \mathbb{E}[e_i^* X_n^k e_i]) - \frac{\lfloor ns \rfloor}{n} \sum_{1 \leq i \leq n} (e_i^* X_n^k e_i - \mathbb{E}[e_i^* X_n^k e_i]).$$

Hence we are led to study the limit, as $n \rightarrow \infty$, of the finite-dimensional marginal distributions of the process

$$(12) \quad \left(\sum_{1 \leq i \leq ns} (e_i^* X_n^k e_i - \mathbb{E}[e_i^* X_n^k e_i]) \right)_{s \in [0, 1], k \geq 1}$$

The random variables of (12) are going to be studied via their joint moments. So let us fix $p \geq 1$, $s_1, \dots, s_p \in [0, 1]$ and $k_1, \dots, k_p \geq 1$. We shall study the limit, as n tends to infinity, of

$$(13) \quad \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right].$$

We introduce the set

$$(14) \quad \mathcal{E} := \{\bar{0}^1, \dots, \bar{k}_1^1\} \cup \dots \cup \{\bar{0}^p, \dots, \bar{k}_p^p\},$$

where the sets $\{\bar{0}^1, \bar{1}^1, \dots\}$, $\{\bar{0}^2, \bar{1}^2, \dots\}$, \dots , $\{\bar{0}^p, \bar{1}^p, \dots\}$ are p disjoint copies of the set of nonnegative integers. The set \mathcal{E} is ordered as presented in (14). For each partition π of \mathcal{E} , for each $x \in \mathcal{E}$, we denote by $\pi(x)$ the index of the class of x , after having ordered the classes according to the order of their first element (for example, $\pi(\bar{1}^1) = 1$; $\pi(\bar{2}^1) = 1$ if $\bar{1}^1 \approx \bar{2}^1$ and $\pi(\bar{2}^1) = 2$ if $\bar{1}^1 \not\approx \bar{2}^1$).

By Assumption 2.1, the expectation of (13) can be expanded and expressed as a sum on the set $\text{Part}(\mathcal{E})$ of partitions of the set \mathcal{E} introduced above. We get

$$(15) \quad \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = n^{-\frac{k_1 + \dots + k_p}{2}} \sum_{\pi \in \text{Part}(\mathcal{E})} A(n, \pi) \mathbb{E} \left[\prod_{\ell=1}^p \left(x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell - 1)^\ell, \pi(\bar{k}_\ell)^\ell} - \mathbb{E}[x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell - 1)^\ell, \pi(\bar{k}_\ell)^\ell}] \right) \right],$$

where for each $\pi \in \text{Part}(\mathcal{E})$, $A(n, \pi)$ is the number of families of indices of

$$(i_{\bar{0}^1}, \dots, i_{\bar{k}_1^1}, i_{\bar{0}^2}, \dots, i_{\bar{k}_2^2}, \dots, i_{\bar{0}^p}, \dots, i_{\bar{k}_p^p}) \in \{1, \dots, n\}^{\mathcal{E}}$$

whose level sets partition is π and who satisfy, for each $\ell = 1, \dots, p$,

$$(16) \quad 1 \leq i_{\bar{0}^\ell} = i_{\bar{k}_\ell^\ell} \leq ns_\ell.$$

For any $\pi \in \text{Part}(\mathcal{E})$, let us define G_π to be the graph with vertex set the set $\{\pi(x); x \in \mathcal{E}\}$ and edge set

$$E_\pi := \{ \{\pi(\bar{m} - \bar{1}^\ell), \pi(\bar{m}^\ell)\} ; 1 \leq \ell \leq p, m \in \{1, \dots, k_\ell\} \}.$$

For the term associated to π in (15) to be non zero, we need to have:

- (i) $\pi(\bar{0}^\ell) = \pi(\bar{k}_\ell^\ell)$ for each $\ell = 1, \dots, p$,
- (ii) each edge of G_π is visited at least twice by the union of the p paths $(\pi(\bar{0}^\ell), \dots, \pi(\bar{k}_\ell^\ell))$ ($\ell = 1, \dots, p$),
- (iii) for each $\ell = 1, \dots, p$, there exists $\ell' \neq \ell$ such that at least one edge of G_π is visited by both paths $(\pi(\bar{0}^\ell), \dots, \pi(\bar{k}_\ell^\ell))$ and $(\pi(\bar{0}^{\ell'}), \dots, \pi(\bar{k}_{\ell'}^{\ell'}))$.

Indeed, (i) is due to (16), (ii) is due to the fact that the $x_{i,j}$'s are centered and (iii) is due to the fact the the $x_{i,j}$'s are independent.

Let us define a function $s(\cdot)$ on the set \mathcal{E} in the following way: for each $\ell = 1, \dots, p$ and each $m = 0, \dots, k_\ell$, set

$$s(\bar{m}^\ell) = \begin{cases} s_\ell & \text{if } m = 0 \text{ or } m = k_\ell, \\ 1 & \text{otherwise,} \end{cases}$$

and

$$(17) \quad s_\pi := \prod_{B \text{ bloc of } \pi} \min_{x \in B} s(x).$$

Then one can easily see that, as $n \rightarrow \infty$,

$$(18) \quad A(n, \pi) \sim s_\pi n^{|\pi|}.$$

Thus for π to have a non zero asymptotic contribution to (15), we need the following condition, in addition to (i), (ii) and (iii):

$$(iv) \quad \frac{k_1 + \dots + k_p}{2} \leq |\pi|.$$

The following lemma is a generalization of [1, Lem. 2.1.34]. Its proof goes along the same lines as the proof of the former.

Lemma 4.1. *Let $\pi \in \text{Part}(\mathcal{E})$ satisfy (i), (ii), (iii). The the number c_π of connected components of G_π is such that $c_\pi \leq p/2$ and*

$$|\pi| \leq c_\pi - \frac{p}{2} + \frac{k_1 + \dots + k_p}{2}.$$

As a consequence, if π also satisfies (iv), we have

- (a) $c_\pi = p/2$,
- (b) p is even,
- (c) $|\pi| = \frac{k_1 + \dots + k_p}{2}$ (so that $k_1 + \dots + k_p$ is also even).

Note also that by (ii), we also have

$$(d) \quad |E_\pi| \leq \frac{k_1 + \dots + k_p}{2}.$$

To sum up, by (15) and (18), we have

$$(19) \quad \lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq n_{s_\ell}} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = \sum_{\pi} s_\pi \lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \left(x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_{\ell-1}^\ell), \pi(\bar{k}_\ell^\ell)} - \mathbb{E}[x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_{\ell-1}^\ell), \pi(\bar{k}_\ell^\ell)}] \right) \right],$$

where the sum is taken over the partitions π of \mathcal{E} which satisfy (i), (ii), (iii) and (iv) above, and such partitions also do satisfy (a), (b), (c) and (d) above.

Case where p is odd: By (b), we know that when p is odd, there is no partition π satisfying the above conditions. So by (19),

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = 0.$$

Case where $p = 2$: In this case, by (a), for each partition π satisfying (i), (ii), (iii) and (iv), G_π is connected, so that $|\pi| - 1 \leq |E_\pi|$. Hence by (c) and (d), $|E_\pi|$ is either equal to $\frac{k_1+k_2}{2}$ or to $\frac{k_1+k_2}{2} - 1$.

In the case where $|E_\pi| = \frac{k_1+k_2}{2} - 1$, the graph G_π has exactly one more vertex than edges, hence is a tree. As a consequence, the paths $(\pi(\bar{0}^1), \dots, \pi(\bar{k}_1^1))$ and $(\pi(\bar{0}^2), \dots, \pi(\bar{k}_2^2))$, which have the same beginning and ending vertices, have the property to visit an even number of times each edge they visit. By an obvious cardinality argument, only one edge is visited more than twice, and it is visited four times (twice in each sense). The other edges are visited once in each sense. It follows that the expectation associated to π in (19) is equal to $\mathbb{E}[|x_{1,2}|^4] - 1$.

In the case where $|E_\pi| = \frac{k_1+k_2}{2}$, by a cardinality argument again, we see that each edge of G_π is visited exactly twice (in fact, the configuration is the one described in [1, Sect. 2.1.7], where G_π is a *bracelet*). It follows that the expectation associated to π in (19) is equal to 1.

As a consequence, as n tends to infinity,

$$\mathbb{E} \left[\sum_{1 \leq i \leq ns_1} (e_i^* X_n^{k_1} e_i - \mathbb{E}[e_i^* X_n^{k_1} e_i]) \times \sum_{1 \leq i \leq ns_2} (e_i^* X_n^{k_2} e_i - \mathbb{E}[e_i^* X_n^{k_2} e_i]) \right]$$

converges to a number that we shall denote by

$$(20) \quad \text{Cov}_{s_1, s_2}^{\text{centered}}(k_1, k_2)$$

and which depends on the distributions of the $x_{i,j}$'s only through $\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4]$.

Case where p is > 2 and even: By (a) above, for each partition π satisfying (i), (ii), (iii) and (iv), G_π has exactly $p/2$ connected components. By (iii), each of them contains the support of exactly two of the p paths

$$(21) \quad (\pi(\bar{0}^\ell), \dots, \pi(\bar{k}_\ell^\ell)) \quad (\ell = 1, \dots, p).$$

Let us define σ_π to be the *matching* (i.e. a permutation all of whose cycles have length two) of $\{1, \dots, p\}$ such that for all $\ell = 1, \dots, p$, the paths with indices ℓ and $\sigma_\pi(\ell)$ of the collection (21) above are supported by the same connected component of G_π .

We shall now partition the sum of (19) according to the value of the matching σ_π defined by π . We get

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = \sum_{\sigma} \sum_{\pi} s_\pi \lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \left(x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell^\ell - 1), \pi(\bar{k}_\ell^\ell)} - \mathbb{E}[x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell^\ell - 1), \pi(\bar{k}_\ell^\ell)}] \right) \right],$$

where the first sum is over the matchings σ of $\{1, \dots, p\}$ and the second sum is over the partitions π satisfying (i), (ii), (iii) and (iv) such that $\sigma_\pi = \sigma$.

Note that for each matching σ of $\{1, \dots, p\}$, the set of partitions π of \mathcal{E} such that $\sigma_\pi = \sigma$ can be identified with the cartesian product, indexed by the set of cycles $\{\ell, \ell'\}$ of σ , of the set of

partitions τ of the set

$$\mathcal{E}_{\ell, \ell'} := \{\bar{0}^\ell, \dots, \bar{k}_\ell^\ell\} \cup \{\bar{0}^{\ell'}, \dots, \bar{k}_{\ell'}^{\ell'}\} \quad (\text{subset of } \mathcal{E})$$

satisfying the following conditions

- (i.2) $\tau(\bar{0}^\ell) = \tau(\bar{k}_\ell^\ell)$ and $\tau(\bar{0}^{\ell'}) = \tau(\bar{k}_{\ell'}^{\ell'})$,
- (ii.2) each edge of the graph G_τ is visited at least twice by the union of the 2 paths $(\tau(\bar{0}^\ell), \dots, \tau(\bar{k}_\ell^\ell))$ and $(\tau(\bar{0}^{\ell'}), \dots, \tau(\bar{k}_{\ell'}^{\ell'}))$,
- (iii.2) at least one edge of G_τ is visited by both previous paths,
- (iv.2) $\frac{k_\ell + k_{\ell'}}{2} \leq |\tau|$.

Moreover, by independence of the random variables $x_{i,j}$'s, the expectation

$$\mathbb{E} \left[\prod_{\ell=1}^p \left(x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell^\ell - 1), \pi(\bar{k}_\ell^\ell)} - \mathbb{E}[x_{\pi(\bar{0}^\ell), \pi(\bar{1}^\ell)} \cdots x_{\pi(\bar{k}_\ell^\ell - 1), \pi(\bar{k}_\ell^\ell)}] \right) \right]$$

factorizes along the connected components of G_π . The factor s_π , defined in (17), also factorizes along the connected components of G_π . It follows that we have

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = \sum_{\sigma} \prod_{\{\ell, \ell'\}} \lim_{n \rightarrow \infty} \mathbb{E} \left[\sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \times \sum_{1 \leq i \leq ns_{\ell'}} (e_i^* X_n^{k_{\ell'}} e_i - \mathbb{E}[e_i^* X_n^{k_{\ell'}} e_i]) \right],$$

where the sum is over the matchings σ of $\{1, \dots, p\}$ and for each such σ , the product is over the cycles $\{\ell, \ell'\}$ of σ .

By the definition of $\text{Cov}_{s_1, s_2}^{\text{centered}}(k_1, k_2)$ in (20), we get

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\prod_{\ell=1}^p \sum_{1 \leq i \leq ns_\ell} (e_i^* X_n^{k_\ell} e_i - \mathbb{E}[e_i^* X_n^{k_\ell} e_i]) \right] = \sum_{\sigma \text{ matching}} \prod_{\{\ell, \ell'\} \text{ cycle of } \sigma} \text{Cov}_{s_\ell, s_{\ell'}}^{\text{centered}}(k_\ell, k_{\ell'}).$$

By Vick's formula and Equation (11), we have proved the first part of Proposition 3.5.

Computation of $\text{Cov}_{s_1, s_2}(2, 2)$: We have, by the paragraph devoted to the case $p = 2$ above,

$$\text{Cov}_{s_1, s_2}^{\text{centered}}(2, 2) = \sum_{\pi, G_\pi \text{ is a tree}} s_\pi \times \left(\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4] - 1 \right) + \sum_{\pi, G_\pi \text{ is a bracelet}} s_\pi.$$

There is exactly one tree and zero bracelet with 2 vertices. We represent this tree in the following way: $\bullet - \circ$. There are two associated partitions π :

- the first one is defined by

$$(\pi(\bar{0}^1), \pi(\bar{1}^1), \pi(\bar{2}^1)) = (\pi(\bar{0}^2), \pi(\bar{1}^2), \pi(\bar{2}^2)) = (\bullet, \circ, \bullet),$$

and satisfies $s_\pi = \min\{s_1, s_2\}$,

- the second one is defined by

$$(\pi(\bar{0}^1), \pi(\bar{1}^1), \pi(\bar{2}^1)) = (\bullet, \circ, \bullet), \quad (\pi(\bar{0}^2), \pi(\bar{1}^2), \pi(\bar{2}^2)) = (\circ, \bullet, \circ),$$

and satisfies $s_\pi = s_1 s_2$.

As a consequence, $\text{Cov}_{s_1, s_2}^{\text{centered}}(2, 2) = (\lim_{n \rightarrow \infty} \mathbb{E}[|x_{1,2}|^4] - 1)(\min\{s_1, s_2\} + s_1 s_2)$.

From this formula and Equation (11), we easily deduce the formula of $\text{Cov}_{s_1, s_2}(2, 2)$. It concludes the proof of Proposition 3.5. \square

4.3. A preliminary result for the proofs of Proposition 2.10 and Lemma 3.1. Proposition 4.3 below will allow to prove Lemma 3.1. Let us now consider two independent Wigner matrices X_n, X'_n . Let us introduce a modified version of Assumption 2.5, also depending on an integer $m \geq 2$:

Assumption 4.2. For each $k \geq 1$, $\sup_n \mathbb{E}[|x_{1,1}|^k + |x_{1,2}|^k + |x'_{1,1}|^k + |x'_{1,2}|^k] < \infty$. Moreover, there exists $\varepsilon_0 > 0$ such that for each $r, s \geq 0$,

$$(22) \quad r + s \leq m - 2 \implies \mathbb{E}[\Re(x_{1,1})^r \Im(x_{1,1})^s] - \mathbb{E}[\Re(x'_{1,1})^r \Im(x'_{1,1})^s] = O(n^{-\varepsilon_0 + 1 + \frac{r+s-m}{2}})$$

and

$$(23) \quad r + s \leq m \implies \mathbb{E}[\Re(x_{1,2})^r \Im(x_{1,2})^s] - \mathbb{E}[\Re(x'_{1,2})^r \Im(x'_{1,2})^s] = O(n^{-\varepsilon_0 + \frac{r+s-m}{2}}).$$

Let $U'_n = [u'_{i,j}]_{i,j=1}^n$ be the associated eigenvectors matrix of a matrix X'_n (like $U_n = [u_{i,j}]_{i,j=1}^n$ for X_n).

Proposition 4.3. We suppose that X_n, X'_n both satisfy Assumptions 2.1 and 2.4, and satisfy Assumption 4.2. Let us fix a positive integer k and a polynomial function G on \mathbb{C}^k . For each n , let us consider a collection $(i_1, p_1, q_1), \dots, (i_k, p_k, q_k) \in \{1, \dots, n\}^3$ (this collection might depend on n). Then for a certain constant C independent of n and of the choices of $(i_1, p_1, q_1), \dots, (i_k, p_k, q_k)$ (but depending on k and on G)

$$(24) \quad |\mathbb{E}[G(n u_{p_1, i_1} \bar{u}_{q_1, i_1}, \dots, n u_{p_k, i_k} \bar{u}_{q_k, i_k})] - \mathbb{E}[G(n u'_{p_1, i_1} \bar{u}'_{q_1, i_1}, \dots, n u'_{p_k, i_k} \bar{u}'_{q_k, i_k})]| \leq C n^{2 - \frac{m}{2}}.$$

Proof. The proof follows the ideas developed by Terence Tao and Van Vu in recent papers, such as [24, 25]. More precisely, it makes an intensive use of the strategy and of some estimations given in [25].

An event $E = E^{(n)}$ depending on the parameter n will be said to be true with *overwhelming probability* if for all $c > 0$, $1 - \mathbb{P}(E) = O(n^{-c})$. One can neglect any such event in the proof of (24) because G has polynomial growth and the entries of unitary matrices are bounded by 1.

By Assumption 2.5 and Chebyshev's inequality, that for each k , as $t \rightarrow \infty$,

$$\mathbb{P}(|x_{i,j}| \geq t) = O(t^{-k}) \quad (\text{uniformly in } i, j \text{ and } n),$$

so that by the union bound, $\mathbb{P}(\max_{1 \leq i, j \leq n} |x_{i,j}| \geq n^\varepsilon) = O(n^{2-k\varepsilon})$ for any ε, k . Hence by what precedes, for a fixed ε_1 (a parameter fixed later), one can suppose that for each n , the $x_{i,j}$'s satisfy $|x_{i,j}| \leq n^{\varepsilon_1}$ for all i, j . Of course, one can also suppose that $|x'_{i,j}| \leq n^{\varepsilon_1}$ for all i, j .

To prove (24), it suffices to prove that we can replace the $x_{i,j}$'s by the $x'_{i,j}$'s one by one up to an error in the considered expectation which is $O(n^{1 - \frac{m}{2}})$ for each diagonal replacement and $O(n^{-\frac{m}{2}})$ for each off-diagonal replacement (and, of course, that these bounds on the errors are uniform in the $\frac{n(n+1)}{2}$ replacements).

So let us fix $1 \leq p \leq q \leq n$ and define, for each $z \in \mathbb{C}$ (or $z \in \mathbb{R}$ if $p = q$), $A_{p,q}(z)$ be the matrix obtained by "mixing" the (rescaled) matrices X_n and X'_n in the following way:

- for each $1 \leq i \leq j \leq n$ such that (i, j) is below (p, q) for the lexicographic order, the i, j entry of $A_{p,q}(z)$ and its symmetric one are the corresponding ones in the entries of nX'_n ,
- the p, q entry of $A_{p,q}(z)$ and its symmetric one are z and \bar{z} ,
- for each $1 \leq i \leq j \leq n$ such that (i, j) is above (p, q) for the lexicographic order, the i, j entry of $A_{p,q}(z)$ and its symmetric one are the corresponding ones in the entries of nX_n .

We set $U(z) = [u_{i,j}(z)]_{i,j=1}^n$ to be the eigenvectors matrix of $A_{p,q}(z)$ and

$$F(z) := G(n u_{p_1, i_1}(z) \bar{u}_{q_1, i_1}(z), \dots, n u_{p_k, i_k}(z) \bar{u}_{q_k, i_k}(z)).$$

Let $\mathbb{E}[\cdot | A_{p,q}(0)]$ denote the conditional expectation with respect to the σ -algebra generated by the entries of $A_{p,q}(0)$. We have to prove that with overwhelming probability, we have, uniformly on the probability space and on p, q ,

$$(25) \quad \mathbb{E}[F(\sqrt{n}x_{p,q}) | A_{p,q}(0)] = \mathbb{E}[F(\sqrt{n}x'_{p,q}) | A_{p,q}(0)] + O(n^{1-p-q-\frac{m}{2}}).$$

To prove it, we shall use the Taylor expansion of F around zero. Let us for example treat the case $p \neq q$. For $|x| \leq n^{\varepsilon_1}$, we have

$$F(\sqrt{n}x) = \sum_{r+s \leq m} \frac{n^{\frac{r+s}{2}}}{r!s!} \partial_{\Re(z)}^r \partial_{\Im(z)}^s F(0) \Re(x)^s \Im(x)^s + O \left(n^{(m+1)(\varepsilon_1+1/2)} \sup_{\substack{r+s=m+1 \\ |z| \leq n^{\varepsilon_1+1/2}}} |\partial_{\Re(z)}^r \partial_{\Im(z)}^s F(z)| \right).$$

Moreover, by [25, Propo. 20], we know that with overwhelming probability, $A_{p,q}(0)$ is a *good configuration* (in the sense of [25, Def. 18]). By Equation (27) of [25, Lem. 24], it implies that with overwhelming probability, for all $l = 1, \dots, k$, for all r, s such that $r + s \leq 10$,

$$\sup_{|z| \leq n^{1/2+\varepsilon_1}} |\partial_{\Re(z)}^r \partial_{\Im(z)}^s (n u_{p_l, i_l}(z) \bar{u}_{q_l, i_l}(z))| = O(n^{\varepsilon_1 - (r+s)})$$

(in the statement of [25, Lem. 24], m must be bounded above by 10, but the bound 10 can be replaced by any finite one). The function G is polynomial, let d be its degree. It follows that with overwhelming probability, for all r, s , we have

$$\sup_{|z| \leq n^{1/2+\varepsilon_1}} |\partial_{\Re(z)}^r \partial_{\Im(z)}^s F(z)| = O(n^{d\varepsilon_1 - (r+s)}).$$

As a consequence, by the bounds of (23),

$$\mathbb{E}[F(\sqrt{n}x_{p,q}) | A_{p,q}(0)] - \mathbb{E}[F(\sqrt{n}x'_{p,q}) | A_{p,q}(0)] = \sum_{r+s \leq m} O(n^{d\varepsilon_1 - \frac{m}{2} - \varepsilon_0}) + O(n^{(m+1+d)\varepsilon_1 - \frac{m+1}{2}})$$

Now, choosing ε_1 small enough so that $d\varepsilon_1 < \varepsilon_0$ and $(m+1+d)\varepsilon_1 < 1/2$, we get (25). \square

4.4. Proof of Proposition 2.10. Let us first prove that $\mathbb{E}[(n|u_{i,j}|^2 - 1)^4]$ is bounded independently of i, j and n . To do so, let us apply Proposition 4.3 for X'_n a GOE or GUE matrix and $m = 4$: it allows to reduce the problem to the case where the $u_{i,j}$'s are the entries of a Haar-distributed matrix. In this case, the distribution of $u_{i,j}$ does not depend on i, j and it is well known that the moments of $\sqrt{n}u_{1,1}$ converge to the ones of a standard real or complex Gaussian variable. We did not find any concise enough reference for this moments convergence, but in the unitary case, Proposition 3.4 of [3] allows to compute $\mathbb{E}[(n|u_{1,1}|^2 - 1)^4]$ and to verify its boundedness. To treat the orthogonal case, the most direct way to compute $\mathbb{E}[(n|u_{1,1}|^2 - 1)^4]$ is to use the fact that $u_{1,1}^2$ has the same distribution as $Z_1^2 / (Z_1^2 + \dots + Z_n^2)$ for Z_i independent standard Gaussian variables (see e.g. [18, Lem. 2.1]) and then to use [18, Lem. 2.4].

Let us now prove that there is a constant C independent of n and of $0 \leq s < s' \leq 1$ and $0 \leq t < t' \leq 1$ such that

$$(26) \quad \mathbb{E}\left\{ \sum_{ns < i \leq ns'} \sum_{nt < j \leq nt'} (|u_{i,j}|^2 - 1/n)^2 \right\} \leq C(s' - s)(t' - t).$$

Set $I = \{i = 1, \dots, n; ns < i \leq ns'\}$ and $J = \{j = 1, \dots, n; nt < j \leq nt'\}$. We have

$$\mathbb{E}\left[\left\{\sum_{i \in I, j \in J} (|u_{i,j}|^2 - 1/n)\right\}^2\right] = n^{-2} \sum_{(i_\ell, j_\ell)_{\ell=1}^2 \in (I \times J)^2} \mathbb{E}\left[\prod_{\ell=1}^2 (n|u_{i_\ell, j_\ell}|^2 - 1)\right].$$

Let us now apply Proposition 4.3 for X'_n a GOE or GUE matrix and $m = 4$. Let $U'_n = [u'_{i,j}]_{i,j=1}^n$ be a Haar-distributed orthogonal or unitary matrix. By Proposition 4.3, we have

$$\mathbb{E}\left[\left\{\sum_{i \in I, j \in J} (|u_{i,j}|^2 - 1/n)\right\}^2\right] = n^{-2} O(n^{-2}(\#I \times J)^2) + n^{-2} \sum_{(i_\ell, j_\ell)_{\ell=1}^2 \in (I \times J)^2} \mathbb{E}\left[\prod_{\ell=1}^2 (n|u'_{i_\ell, j_\ell}|^2 - 1)\right],$$

where the term $n^{-2} O(n^{-2}(\#I \times J)^2)$ is uniform in s, s', t, t' , hence is bounded, for a certain constant C_1 , by $C_1(s' - s)^2(t' - t)^2 \leq C_1(s' - s)(t' - t)$. So it suffices to prove the result for U'_n instead of U_n . Since each $|u'_{i,j}|^2$ has expectation $1/n$,

$$\mathbb{E}\left[\prod_{\ell=1}^2 (n|u'_{i_\ell, j_\ell}|^2 - 1)\right] = n^2 \mathbb{E}[|u'_{i_1, j_1}|^2 |u'_{i_2, j_2}|^2] - 1.$$

Then, to compute $\mathbb{E}[|u'_{i_1, j_1}|^2 |u'_{i_2, j_2}|^2]$, one uses the *Weingarten calculus*, developed in [8, 9]. By the formulas of [9, Sec. 6] for the orthogonal group and [8, Sec. 5.2] for the unitary group,

$$\begin{aligned} \text{Wg}^{\text{ortho}}(Id_{\{1,2\}}) &= \frac{n+1}{n(n-1)(n+2)}, & \text{Wg}^{\text{ortho}}((12)) &= \frac{1}{n(n-1)(n+2)}, \\ \text{Wg}^{\text{unit}}(Id_{\{1,2\}}) &= \frac{1}{n^2-1}, & \text{Wg}^{\text{unit}}((12)) &= \frac{-1}{n(n^2-1)}, \end{aligned}$$

(where (12) denotes the transposition with support $\{1, 2\}$), one can easily verify that

$$n^2 \mathbb{E}[|u'_{i_1, j_1}|^2 |u'_{i_2, j_2}|^2] - 1 = \begin{cases} O(n^{-2}) & \text{if } i_1 \neq i_2, j_1 \neq j_2, \\ O(n^{-1}) & \text{if } (i_1 = i_2, j_1 \neq j_2) \text{ or } (i_1 \neq i_2, j_1 = j_2), \\ O(1) & \text{if } i_1 = i_2, j_1 = j_2, \end{cases}$$

which is enough to deduce (26).

4.5. Proof of Lemma 3.1. To prove Lemma 3.1, we shall use the following proposition, which is the obvious multidimensional generalization of Proposition 3.26 of Chapter VI of [16] (note also that the result of [16] is devoted to processes on $[0, \infty)$ and not on $[0, 1]$, so that the N of [16] means nothing to us).

For $f \in D([0, 1]^2)$ and $(s_0, t_0) \in [0, 1]^2$, we define $\Delta_{s_0, t_0} f$ to be the *maximal jump* of f at (s_0, t_0) , i.e.

$$\Delta_{s_0, t_0} f := \max_{\diamond, \diamond' \in \{<, \geq\}} \left| f(s_0, t_0) - \lim_{\substack{s \diamond s_0 \\ t \diamond' t_0}} f(s, t) \right|.$$

Proposition 4.4. *If the sequence $(\text{distribution}(B^n))_{n \geq 1}$ is tight and satisfies*

$$(27) \quad \forall \varepsilon > 0, \quad \mathbb{P}\left(\sup_{(s,t) \in [0,1]^2} \Delta_{s,t} B^n > \varepsilon\right) \xrightarrow{n \rightarrow \infty} 0,$$

then the sequence $(\text{distribution}(B^n))_{n \geq 1}$ is C -tight, i.e. is tight and can only have $C([0, 1]^2)$ -supported accumulation points.

So to prove Lemma 3.1, let us first prove that the sequence $(\text{distribution}(B^n))_{n \geq 1}$ is tight. Note that the process B^n vanishes at the border of $[0, 1]^2$. So according to [4, Th. 3] and to Cauchy-Schwartz inequality, it suffices to prove that there is $C < \infty$ such that for n large enough, for all $s < s', t < t' \in [0, 1]$,

$$\mathbb{E}[\{\sum_{ns < i \leq ns'} \sum_{nt < j \leq nt'} (|u_{i,j}|^2 - 1/n)\}^4] \leq C(s' - s)^2(t' - t)^2.$$

As in the proof of (26) above, one can suppose that the $u_{i,j}$'s are the entries of a Haar-distributed matrix. But in this case, the job has already been done in [10]: the unitary case is treated in Section 3.4.1 (see specifically Equation (3.25)) and the orthogonal case is treated, more elliptically, in Section 4.5 (to recover the details of the proof, join Equations (3.26), (4.5) and (4.17)).

Let us now prove (27). Note that $\sup_{(s,t) \in [0,1]^2} \Delta_{s,t} B^n = \max_{1 \leq i,j \leq n} ||u_{i,j}|^2 - 1/n|$. As a consequence, by the union bound, it suffices to prove that for each $\varepsilon > 0$, there is $C < \infty$ independent of i, j and n such that for all i, j ,

$$\mathbb{P}(|u_{i,j}|^2 - 1/n| > \varepsilon) \leq Cn^{-4},$$

which follows from Chebyshev's inequality and Proposition 2.10. \square

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