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Accurate Tests and Intervals Based on Non-Linear Cusum Statistics

by

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Abstract: Suppose that we are monitoring incoming observations for a change in mean via a cusum test statistic. The usual nonparametric methods give first and second order approximations for the one- and two-sided cases. Withers and Nadarajah [*Statistics and Probability Letters*, **79**, 2009, 689–697] showed how to improve the order of these approximations for linear statistics. Here, an extension is provided for non-linear statistics with non-zero asymptotic variances. This involves development of two calculi analogous to that of von Mises.

Keywords: Approximations; Cusums; Edgeworth-Cornish-Fisher; Non-linear statistics.

1 Introduction and Summary

Suppose that we are monitoring incoming observations for a change in mean via a cusum test statistic. The usual nonparametric asymptotic methods give a first order approximation in the one-sided case and a second order approximation in the two-sided case, where by *Ith order* we mean an error of magnitude $n^{-I/2}$ for *n* the sample size. Withers and Nadarajah (2009) showed how to improve on the order of these approximations for asymptotic covariance, asymptotic normality still holds and *I*th order approximations to their distributions and quantiles can be obtained by expanding their cumulants as power series in n^{-1} then applying Edgeworth-Cornish-Fisher technology. To do this two calculi are developed analogous but simpler to that of von Mises. In the case, where the mean is not changing, the corresponding tests can be viewed as one- and two-sided confidence intervals for the mean.

Set $X_0 = 0$ and let $X = X_1, X_2, \cdots$ be independent random variables in \mathcal{R}^p from some distribution F(x) say, with mean μ , and finite moments. (If p = 1 we denote the *r*th cumulant by κ_r and set $\sigma^2 = \kappa_2$.) These observations can be considered as a random process which may at some point go "out of control", changing their distribution. We define the average process of the observations as

$$M_n(t) = n^{-1} S_{[nt]} \tag{1.1}$$

for $S_0 = 0$, $S_i = \sum_{j=1}^{i} X_j$ and $0 \le t \le 1$, where [x] is the integral part of x. A change in mean can be tracked by monitoring the average process via some functional of it, say $T(M_n)$, referred to as a cusum statistic. We denote the mean of the average process (1.1) by

$$m_n(t) = E \ M_n(t) = \mu[nt]/n \to m(t) = \mu t$$

$$(1.2)$$

as $n \to \infty$ for $0 \le t \le 1$. Given a functional T, $\hat{\theta} = T(M_n)$ can be thought of as an estimate of $\theta = T(m)$. In this way we can if desired use $\hat{\theta}$ to provide an estimate of μ . We shall assume that $\hat{\theta}$ satisfies the standard cumulant expansion given by

$$\kappa_r(\widehat{\theta}) = \sum_{i=r-1}^{\infty} a_{ri} n^{-i} \tag{1.3}$$

for $r = 1, 2, \cdots$.

For univariate data the most common prospective and retrospective two-sided cusum statistics and functionals are

$$A_n = A(M_n) = \max_{k=1}^n |S_k - k\mu|/n,$$
(1.4)

$$B_n = B(M_n) = \max_{k=1}^n |S_k - k\bar{X}_n|/n$$
(1.5)

for $A(g) = \sup_{[0,1]} |g(t) - t\mu|$ and $B(g) = \sup_{[0,1]} |g(t) - tg(1)|$, where $\bar{X}_n = M_n(1) = S_n/n$. If $\hat{\sigma}$ is a consistent estimate of the standard deviation σ , then as $n \to \infty$

$$\begin{split} \widehat{\sigma}^{-1} n^{1/2} \{ M_n(t) - t\mu \} &\xrightarrow{\mathcal{L}} W(t), \\ \widehat{\sigma}^{-1} n^{1/2} \{ M_n(t) - tM_n(1) \} &\xrightarrow{\mathcal{L}} W_0(t) \end{split}$$

for W(t) a Wiener process and $W_0(t) = W(t) - tW(1)$ a Brownian Bridge. So,

$$\widehat{\sigma}^{-1} n^{1/2} A_n \xrightarrow{\mathcal{L}} \sup_{[0,1]} |W(t)|, \qquad (1.6)$$

$$\widehat{\sigma}^{-1} n^{1/2} B_n \xrightarrow{\mathcal{L}} \sup_{[0,1]} |W_0(t)|.$$
(1.7)

See Billingsley (1968) and Anderson and Darling (1952).

These asymptotic results are easily extended to functionals like $\sup\{|g(t) - t\mu_0| - b(t)\}$, that is to test $H_0: \mu = \mu_0$ versus $H_1: \mu \neq \mu_0$ by rejecting H_0 if $M_n(t) - t\mu_0$ crosses a given boundary $\pm b(t)$. An alternative is to use the L_1 norm. For example, the one-sided test of $H_0: \mu = \mu_0$ versus $H_1: \mu > \mu_0$ one can use $\int_0^1 \{g(t) - t\mu_0 - b(t)\} dt$, or equivalently $\int_0^1 g(t) dt$. This is an example of the statistics and functionals we consider here. They include $T(M_n)$ for

$$T(g) = \int_0^1 w(g(t), t) dt \text{ or } \int_0^1 w(g(t), t) dg(t)$$
(1.8)

for $w(x,t) : \mathbb{R}^p \times [0,1] \to \mathbb{R}^q$, as well as smooth functions of such functionals. These functionals have the advantage of being asymptotically normal, unlike (1.6) and (1.7), and of having distribution, density and quantiles given by their Edgeworth-Cornish-Fisher expansions. In contrast, expansions for the distribution, density and quantiles of the cusum statistics (1.4)–(1.7), are not available. The aim of this note is to extend the work of Withers and Nadarajah (2009) for nonlinear functionals. In this case, the cumulant expansions in (1.3) and hence the Edgeworth-Cornish-Fisher expansions (Section 2, Withers and Nadarajah, 2009) are obtained in terms of functional derivatives of T at m using the Euler-McLaurin expansion. Section 2 defines the functional derivatives needed. There are two types depending on whether integration is with respect to dt or dg(t). These are analogous to those of von Mises (1947) as extended in Withers (1983, 1988). We show that the chain rule applies to such derivatives. Examples include (1.8) and smooth functions of such functionals. Section 3 gives the cumulants of $T(M_n)$ needed to obtain its Edgeworth-Cornish-Fisher expansions (Section 2, Withers and Nadarajah, 2009). For example, as $n \to \infty$,

$$n \ var(T(M_n)) \to \sigma^2 \int T_m(t)^2 dt$$

for m(t) of (1.2) and $T_g(t)$ the first derivative of T(g) in the calculus of Section 2.

2 Functional Derivatives

Since $M_n(t)$ is not a distribution function, the calculus of von Mises (1947) does not apply to functionals of it. In this section we develop two differential calculi which can be applied to such functionals. The first calculus is for functionals like $\int_0^1 w(g(t), t)dt$ when g(0) = 0and for $f(\int wdg)$ without the condition g(0) = 0. The second calculus covers functionals like $\int_0^1 w(g(t), t)dt$ without the condition g(0) = 0.

2.1 First Calculus

Consider functions $g, h : I \to \mathcal{R}$, where $I \subset \mathcal{R}^p$, and a functional T(g) in \mathcal{R} . Suppose that for real ϵ ,

$$T(g+\epsilon h) = T(g) + \sum_{r=1}^{\infty} T_r(g,h)\epsilon^r/r!, \qquad (2.1)$$

where

$$T_r(g,h) = \int_I \cdots \int_I T_g(t_1,\ldots,t_r) dh(t_1)\ldots dh(t_r)$$

is assumed to exist. We call $T_g(t_1, \ldots, t_r)$ the rth order (functional) derivative of T(g) with respect to the first calculus, or simply the first rth order derivative. It is made unique by the condition

$$T_g(t_1, \dots, t_r)$$
 is symmetric in t_1, \dots, t_r . (2.2)

Restrictions on h may be required in order that (2.1) hold, as the examples show. That is, for a given T the expansion for $T(g_n) - T(g)$ may place restrictions on eligible $g_n - g$.

Note 2.1 Condition (2.1) can be weakened in two ways. Firstly, it need only hold in a neighbourhood of zero, say $|\epsilon| < \delta$. This is satisfied by $\epsilon = n^{-1/2}$ if $n > \delta^{-2}$. One can then take, for example, g(t) = m(t) and $h(t) = n^{1/2}(M_n(t) - m(t))$, so that the left hand side of

(2.1) is $T(M_n)$. Secondly, the infinite Taylor series (2.1) can be replaced by a partial series with a remainder term, analogous to the usual ordinary Taylor series for a real function of a real variable, given in Section 3.6.1 of Abramowitz and Stegun (1964).

Like von Mises' calculus, these calculi cannot be applied to sup functionals like that of (1.4) and (1.5).

 Set

$$\Delta_t(s) = I(t \le s) = \prod_{i=1}^p I(t_i \le s_i).$$
(2.3)

Taking $\lambda_1, \dots, \lambda_r$ in \mathcal{R} and $h(s) = \sum_{i=1}^r \lambda_i \Delta_{t_i}(s)$, where $\Delta_{t_i}(\cdot)$ is defined by (2.3), we see that $T_g(t_1, \dots, t_r)$ is the coefficient of $\lambda_1 \dots \lambda_r$ in $T_r(g, h)/r!$, so that $T_g(t_1, \dots, t_r)$ is the coefficient of $\lambda_1 \dots \lambda_r \epsilon^r$ in $T(g + \epsilon h)/r!$. So, the first derivative is

$$S(g) = T_g(t) = \lim_{\epsilon \to 0} \left\{ T(g + \epsilon \Delta_t) - T(g) \right\} / \epsilon,$$
(2.4)

where Δ_t is defined by (2.3). Similarly, by the method of proof of Theorem 2.1 of Withers (1983), it follows that

$$T_g(t_1, \dots, t_{r+1}) = S_g(t_{r+1}) \text{ for } S(g) = T_g(t_1, \dots, t_r).$$
 (2.5)

That is, the derivative of the *r*th derivative is the r + 1st derivative, just as for ordinary partial derivatives. The situation is analogous to, but simpler to the theory of derivatives of functionals of distribution functions developed by Von Mises (1947) and refined by Withers (1983). Note that in general $\int_0^1 T_g(t_1, \ldots, t_r) dt_1 \neq 0$, whereas for the von Mises derivative $\int T_F(x_1, \ldots, x_r) dF(x_1) = 0$: see (2.3) of Withers (1983).

Example 2.1 Suppose that p = 1 and $T(g) = \int_0^1 w(t)dg(t)$, where $w(t) : [0,1] \to \mathcal{R}$ and $g : [0,1] \to \mathcal{R}$ are chosen such that T(g) exists. Set $w_r = (\int w^r)^{1/r}$ for $0 < r < \infty$. Then $T_g(t) = w(t)$ and higher derivatives are zero. So, $\int T_m dm = \mu w_1 \neq 0$ in general.

Example 2.2 Suppose that p = 1 and $T(g) = \int_0^1 W(t)g(t)dt$. Then one needs to impose the condition h(0) = 0 giving $T_g(t) = \int_t^1 W(s)ds$ or the condition h(1) = 0 giving $T_g(t) = \int_0^t W(s)ds$. The higher derivatives are zero. The condition h(0) = 0 restricts application to $T(g_n) - T(g)$ with $g_n(0) - g(0) = 0$, a condition satisfied by $(g_n, g) = (M_n, m)$ of (1.1), (1.2). The condition h(1) = 0 restricts application to $T(g_n) - T(g)$ with $g_n(1) - g(1) = 0$, a condition not satisfied by $(g_n, g) = (M_n, m)$.

For the next examples it is helpful to use \wedge and \vee to denote min and max:

$$t_{r\wedge} = \min_{i=1}^{r} t_i, \ t_{r\vee} = \max_{i=1}^{r} t_i.$$
(2.6)

If p > 1 then min and max are interpreted element-wise.

Example 2.3 Take p = 1. Let s be any point in [0,1]. Consider functions $g : [0,1] \to \mathcal{R}$, and a smooth function $w : \mathcal{R} \to \mathcal{R}$. Denote its rth derivative by $w_{.r}$. If h(0) = 0 and T(g) = w(g(s)) then

$$T_g(t_1,\ldots,t_r) = w_{.r}(g(s))\Delta_{t_1}(s)\cdots\Delta_{t_r}(s) = w_{.r}(g(s))\Delta_{t_{r\vee}}(s),$$

where $\Delta_{t_1}(\cdot)$ is defined by (2.3). If h(s+) = h(s) and T(g) = w(g(s) - g(1)) then

$$T_g(t_1, \dots, t_r) = w_{r}(g(s) - g(1))(\Delta_{t_1}(s) - 1) \cdots (\Delta_{t_r}(s) - 1)$$

= $(-1)^r w_{r}(g(s) - g(1))I(s < t_{r\wedge}).$

Example 2.4 Let ν be a measure on some space L. For s in L let $T_s(g)$ be a functional with finite rth derivative $T_{sg}(t_1, \dots, t_r)$. For $S(g) = \int T_s(g) d\nu(s)$,

$$S_g(t_1,\ldots,t_r) = \int T_{sg}(t_1,\ldots,t_r) d\nu(s).$$

Putting these two examples together, we have

Example 2.5 Take p = 1. Let $w(x,t) : \mathcal{R} \times [0,1] \to \mathcal{R}$ have finite partial derivatives $w_{r0}(x,t) = (\partial/\partial_x)^r w(x,t)$. If g(0) = 0 and $S(g) = \int_0^1 w(g(s),s) ds$ then

$$S_g(t_1, \dots, t_r) = \int_{t_{r\vee}}^1 w_{.r0}(g(s), s) ds.$$
(2.7)

If g is right-continuous (for example, $g = M_n - m$) and

$$S(g) = \int_0^1 w(g(s) - g(1), s) ds$$
(2.8)

then

$$S_g(t_1, \dots, t_r) = (-1)^r \int_0^{t_{r\wedge}} w_{.r0}(g(s) - g(1), s) ds.$$
(2.9)

For example, alternatives to the statistics A_n and B_n of (1.4) and (1.5) are

$$C_n = C_n(M_n) = n^{-3} \sum_{j=1}^n (S_j - j\mu)^2 w(j/n)/2,$$
$$D_n = D_n(M_n) = n^{-3} \sum_{j=1}^n (S_j - jS_n/n)^2 w(j/n)/2$$

for

$$C_n(g) = \int_0^1 (g(s) - \mu[ns]/n)^2 w([ns]/n) ds/2,$$

$$D_n(g) = \int_0^1 (g(s) - g(1)[ns]/n)^2 w([ns]/n) ds/2.$$

Their non-zero derivatives are

$$C_{ng}(t_1) = \int_{t_1}^1 (g(s) - \mu[ns]/n)w([ns]/n)ds,$$

$$C_{ng}(t_1, t_2) = \int_{t_{m2}}^1 w([ns]/n)ds,$$

$$D_{ng}(t_1) = \int_0^1 (g(s) - g(1)[ns]/n)w([ns]/n)e_n(s, t_1)ds,$$

$$D_{ng}(t_1, t_2) = \int_0^1 e_n(s, t_1)e_n(s, t_2)ds,$$

where $e_n(s,t) = I(t \le s) - [ns]/n$. Now replace $C_n(g), D_n(g)$ by their limits

$$C(g) = \int_0^1 (g(s) - \mu s)^2 w(s) ds/2,$$

$$D(g) = \int_0^1 (g(s) - g(1)s)^2 w(s) ds/2$$

and set $C'_n = C(M_n), D'_n = D(M_n)$. If w(s) = 1 then one may show that $C_n - C'_n$ and $D_n - D'_n$ are $O_p(n^{-3/2})$ in probability.

Since D(g) does not have the form (2.8), (2.9) does not apply. This can be fixed by applying the method not to $M_n(t)$ but to

$$N_n(t) = [nt]^{-1} \sum_{i=1}^{[nt]} X_i,$$

interpreted as 0 for $t \leq 1/n$. However, $\kappa(\hat{w}_1, \dots, \hat{w}_r)$ for $\hat{w}_i = N_n(t_i)$ is more complicated than for $\hat{w}_i = M_n(t_i)$, so we do not pursue this idea beyond the next example.

Example 2.6 Consider

$$D'_{n} = D'_{n}(N_{n}) = n^{-1} \sum_{j=1}^{n-1} (S_{j}/j - S_{n}/n)^{2} w(j/n)/2,$$

where

$$D'_{n}(g) = \int_{n^{-1}}^{1} (g(s) - g(1))^{2} w([ns]/n) ds/2.$$

The non-zero derivatives are

$$D'_{ng}(t_1) = -\int_{n^{-1}}^{1} (g(s) - g(1))w([ns]/n)I(s \le t)ds$$
$$D'_{ng}(t_1, t_2) = \int_{n^{-1}}^{1} w([ns]/n)I(s < t_{m2})ds.$$

Note that $D_n[w] = D_n$ and $D'_n[w] = D'_n$ are equivalent in that $D'_n[w'] = D_n[w]$ for $w'(t) = t^2w(t)$. Note that

$$2n^{3}D_{n}[1] = 2n^{3}D'_{n}[t^{2}] = \sum_{j=1}^{n} (S_{j} - j\bar{X}_{n})^{2}$$

provides a simple alternative to

$$nB_n = \max_{j=1}^n |S_j - j\bar{X}_n|,$$

since they provide L_2 and L_{∞} norms of $S_j - j\bar{X}_n$.

2.2 Chain Rule

Since $T_g(t_1, \ldots, t_r)$ is an exact analogue of the ordinary rth order partial derivative, it satisfies the chain rule. We now spell this out in detail. By Comtet (1974, page 137), Faa di Bruno's chain rule for the derivatives of a function of a function is most simply stated in terms of the exponential Bell polynomials tabled on page 307 of Comtet (1974): for $r \ge 1$, the rth derivative of a real function of a real function, say S(g) = f(T(g)), for g in \mathcal{R} , $T: \mathcal{R} \to \mathcal{R}, f: \mathcal{R} \to \mathcal{R}$ is given by

$$S^{(r)}(g) = \sum_{k=1}^{r} f_k B_{rk}(\mathbf{T})$$
(2.10)

at $f_k = f^{(k)}(T(g))$ and $\mathbf{T} = (T_1, T_2, \cdots)$, where $T_r = T^{(r)}(g)$, the *r*th derivative of T(g). This extends to g and T(g) multi-dimensional by a simple reinterpretation. For example, for r = 4, (2.10) gives

$$S^{(4)}(g) = f_1 T_4 + f_2 (4T_1 T_3 + 3T_2^2) + f_3 (6T_1^2 T_2) + f_4 T_1^4.$$

For g in p dimensions and T(g) in q dimensions this becomes

$$S_{.i_{1}i_{2}i_{3}i_{4}}(g) = f_{.j_{1}}T_{j_{1}.i_{1}i_{2}i_{3}i_{4}} + f_{.j_{1}j_{2}}(\sum_{i}^{4}T_{j_{1}.i_{1}}T_{j_{2}.i_{2}i_{3}i_{4}} + \sum_{i}^{3}T_{j_{1}.i_{1}i_{2}}T_{j_{2}.i_{3}i_{4}})$$

+ $f_{.j_{1}j_{2}j_{3}}\sum_{i}^{6}T_{j_{1}.i_{1}}T_{j_{2}.i_{2}}T_{j_{3}.i_{3}i_{4}} + f_{.j_{1}j_{2}j_{3}j_{4}}T_{j_{1}.i_{1}}T_{j_{2}.i_{2}}T_{j_{3}.i_{3}}T_{j_{4}.i_{4}}, (2.11)$

where, for i_1, i_2, \dots in $1, 2, \dots, p$ $S_{.i_1i_2\dots}(g) = \partial_{i_1}\partial_{i_2}\dots S(g)$ for $\partial_i = \partial/\partial_{g_i}$. Similarly, for the partial derivatives of $T_j(g)$ for $j = 1, \dots, q$; and for j_1, j_2, \dots in $1, 2, \dots, q$ $f_{.j_1j_2\dots}(T(g)) = \partial_{j_1}\partial_{j_2}\dots f(y)$ at y = T(g) for $\partial_j = \partial/\partial_{y_j}$. The convention is that j_1, j_2, \dots , the repeated pairs of suffixes in (2.11), are implicitly summed over their range $1, 2, \dots, q$. The expressions \sum_i^N in (2.11) mean summation over all N permutations of the *i*'s giving different terms. For example,

$$\sum_{i}^{4} T_{j_1.i_1} T_{j_2.i_2i_3i_4} = T_{j_1.i_1} T_{j_2.i_2i_3i_4} + T_{j_1.i_2} T_{j_2.i_1i_3i_4} + T_{j_1.i_3} T_{j_2.i_1i_2i_4} + T_{j_1.i_4} T_{j_2.i_1i_2i_3}.$$

This chain rule in turn has a simple reinterpretation for functions of functionals. Consider the real functional S(g) = f(T(g)), where T(g) is a q-dimensional functional on functions $g: I \to \mathcal{R}$, where $I \subset \mathcal{R}^p$, and f(T) is an ordinary function from \mathcal{R}^q to \mathcal{R} . Then, for example, the fourth order functional derivatives of S(g) are given in terms of the functional derivatives of T(g) by (2.11) with $S_{.i_1i_2...}(g)$ replaced by $S_g(t_1t_2...)$ and $T_{j.i_1i_2...}$ replaced by $T_{jg}(t_1t_2...)$, the functional derivatives of $T_j(g)$:

$$S_{g}(t_{1}t_{2}t_{3}t_{4}) = f_{.j_{1}}T_{j_{1}g}(t_{1}t_{2}t_{3}t_{4}) + f_{.j_{1}j_{2}}\sum_{t}^{4}T_{j_{1}g}(t_{1})T_{j_{2}g}(t_{2}t_{3}t_{4}) + \sum_{t}^{3}T_{j_{1}g}(t_{1}t_{2})T_{j_{2}g}(t_{3}t_{4})) + f_{.j_{1}j_{2}j_{3}}\sum_{t}^{6}T_{j_{1}g}(t_{1})T_{j_{2}g}(t_{2})T_{j_{3}g}(t_{3}t_{4}) + f_{.j_{1}j_{2}j_{3}j_{4}}T_{j_{1}g}(t_{1})T_{j_{2}g}(t_{2})T_{j_{3}g}(t_{3})T_{j_{4}g}(t_{4}).$$

2.3 Extension to Vector q

We now extend the results of this section to g, h, ϵ s-vectors. Let ϵh denote the s-vector with *i*th component $\epsilon_i h_i$ for $i = 1, \dots, s$. For r in \mathcal{N}^s , where $\mathcal{N} = \{0, 1, 2, \dots\}$, set

$$|r| = \sum_{i=1}^{s} r_i, \ \epsilon^r / r! = \prod_{i=1}^{s} \epsilon^{r_i} / r_i!.$$

Now consider functionals T(g) such that for all s-vector functions g, h on $I \subset \mathcal{R}^p$,

$$T(g+\epsilon h) = \sum_{r\in\mathcal{N}^s} T_r(g,h)\epsilon^r/r!$$

for

$$T_r(g,h) = \int_I^{|r|} T_g(\{u_{ij}, i = 1, \cdots, s; j = 1, \cdots, r_i\}) \prod_{i=1}^s \prod_{j=1}^{r_i} dh_i(u_{ij}),$$

where $\int_{I}^{|r|}$ denotes |r| integrals over I. The rth derivative, the above integrand, is again made unique by the condition that it is symmetric in each of its s rows. Taking $h_i(u) = \sum_{j=1}^{r_i} \lambda_{ij} \Delta_{t_i}(u)$, where $\Delta_{t_i}(\cdot)$ is defined by (2.3), we see that this derivative is the coefficient of $\prod_{i=1}^{s} \prod_{j=1}^{r_i} \lambda_{ij}$ in $T_r(g, h)/r!$.

Let e_{is} be the *i*th unit vector in \mathcal{R}^s . Then as in (2.4), for *t* in \mathcal{R} the *s* first derivatives are

$$S_i(g) = \lim_{\epsilon \to 0} \left\{ T(g + \epsilon e_{is} \Delta_t) - T(g) \right\} / \epsilon,$$

where Δ_t is defined by (2.3). Similarly, (2.5) holds if t_1, \ldots, t_{r+1} and t_1, \ldots, t_r are interpreted as *i*th rows with the other rows implicit and the same on both sides, that is

Example 2.7 Take

$$I \subset \mathcal{R}^p, \ g: I \to \mathcal{R}, \ w: I \to \mathcal{R}^q, \ T(g) = \int w dg, S(g) = f(\int w dg).$$

Then

$$S_g(t_1,\cdots,t_r) = f_{j_1\cdots j_r}(\int w dg) w_{j_1}(t_1)\cdots w_{j_r}(t_r).$$

2.4 Second Calculus

Consider functions $g, h: I \to \mathcal{R}$, where $I \subset \mathcal{R}^p$, and a functional T(g) in \mathcal{R} . Suppose that for real ϵ , (2.1) holds, where now

$$T_r(g,h) = \int_I \cdots \int_I T_g(t_1,\ldots,t_r)h(t_1)dt_1\ldots h(t_r)dt_r$$

is assumed to exist. We call $T_g(t_1, \ldots, t_r)$ the rth order (functional) derivative of T(g) with respect to the second calculus, or simply the second rth order derivative. It is made unique by the condition (2.2).

Let $\delta(t)$ denote the Dirac function on \mathcal{R} . For $t, s \in \mathcal{R}^q$, set $\delta(t) = \prod_{i=1}^q \delta(t_i)$ and $\delta_t(s) = \delta(t-s)$. Taking $\lambda_1, \dots, \lambda_r$ in \mathcal{R} and $h(s) = \sum_{i=1}^r \lambda_i \delta_{t_i}(s)$, we see that $T_g(t_1, \dots, t_r)$ is the coefficient of $\lambda_1 \dots \lambda_r$ in $T_r(g, h)/r!$. So, the first derivative is (2.4) with Δ_t replaced by δ_t . Similarly, (2.5) and the chain rule hold.

Example 2.8 Take p = 1. Suppose that $T(g) = \int_0^1 v(t)g(t)dt = \int vg \ say$. Then $T_g(t) = v(t)$ and higher derivatives are zero.

Example 2.9 Take p = 1. Let s be any point in [0,1]. Consider functions $g: [0,1] \to \mathcal{R}$, and $w: \mathcal{R} \to \mathcal{R}$. If T(g) = w(g(s)) then

$$T_g(t_1,\ldots,t_r) = w_{.r}(g(s))\delta_{t_1}(s)\cdots\delta_{t_r}(s)$$

If T(g) = w(g(s) - g(1)) then

$$T_g(t_1,\ldots,t_r) = w_{r}(g(s) - g(1))(\delta_{t_1}(s) - \delta_{t_1}(1))\cdots(\delta_{t_r}(s) - \delta_{t_r}(1)).$$

Since the analog of Example 3.4 in Withers and Nadarajah (2009) holds for the second calculus, we have in the notation of Example 2.5,

Example 2.10 Take

$$I \subset \mathcal{R}^p, g: I \to \mathcal{R}, v: I \to \mathcal{R}^q, T(g) = \int_I vg, S(g) = f(\int_I vg) dv$$

Then

$$S_g(t_1,\cdots,t_r) = f_{j_1\cdots j_r}(\int_I vg)v_{j_1}(t_1)\cdots v_{j_r}(t_r).$$

We note in passing that our method avoids the complexities of Gateaux, Fréchet and Hadamard derivatives. For more on these see, for example, Dudley (1992), Bednarski et al. (1991), Ren and Sen (1991, 1995) and Pons and de Turckheim (1991).

3 Cumulant Expansions

Here we apply the first calculus of the previous section with p = 1, I = [0, 1] in order to obtain the leading coefficients $\{a_{ri}\}$ in the cumulant expansion (1.3) for $\hat{\theta} = T(M_n)$. The higher order tests and confidence intervals can then be obtained as in Withers and Nadarajah (Section 2, 2009). By (2.1)

$$T(M_n) = T(m) + \sum_{r=1}^{\infty} \int_0^1 \cdots \int_0^1 T_m(t_1, \dots, t_r) \\ d(M_n(t_1) - m(t_1)) \dots d(M_n(t_r) - m(t_r))/r!.$$

Cumulant expansions for $T(M_n)$ could be obtained from this, following the method in Withers (1983), where cumulant expansions for $T(F_n)$ were obtained for F_n the empirical distribution of a random sample from F. However, in Withers (1988) we showed that these could be obtained much more easily by applying the cumulant expansions of Withers (1982) for $t(\hat{w})$ by identifying \hat{w}_i with $F_n(x_i), i = 1, \ldots, p$ and letting $p \to \infty$. Here we take the same short cut with $\hat{w}_i = M_n(t_i)$. We begin with some notation.

For $f: [0,1]^r \to \mathcal{R}$, set

$$(f)_r = \int_0^1 \cdots \int_0^1 f(t_1, \cdots, t_r) dt_1 \cdots dt_r.$$

For π_1, π_2, \cdots sets with elements in $\{1, 2, \cdots r\}$ set

$$[\pi_1, \pi_2, \cdots] = [\pi_1, \pi_2, \cdots]_T = (h)_r \text{ for } h(t_1, \cdots, t_r) = T_{\pi_1} T_{\pi_2} \cdots,$$

where, for $\pi = \{i, j, \dots\}, T_{\pi} = T_m(t_i, t_j, \dots)$. For example,

$$[1^{i}] = \int T_{m}(t_{1})^{i} dt_{1} = (T_{1}^{i})_{1},$$

$$[1^{i}, 11^{j}] = (T_{1}^{i}T_{11}^{j})_{1},$$

$$[1^{i}, 12^{j}, 2^{k}] = (T_{1}^{i}T_{12}^{j}T_{2}^{k})_{2}$$

for $(T_1^i)(t_1) = T_m(t_1)^i$, $(T_1^i T_{11}^j)(t_1) = T_m(t_1)^i T_m(t_1 t_1)^j$ and $(T_1^i T_{12}^j T_2^k)(t_1 t_2) = T_m(t_1)^i T_m(t_1 t_2)^j T_m(t_2)^k$. For $\hat{w}_i = M_n(t_i)$,

$$\kappa(\widehat{w}_{i_1},\cdots,\widehat{w}_{i_r}) = \kappa_r n^{1-r} k_n(t_{r\wedge})$$

for $t_{r\wedge}$ of (2.6), where $k_n(t) = [nt]/n = t + n^{-1}\ell_n(t)$ for $\ell_n(t) = [nt] - nt$. So, (1.3) holds with a_{ri} given by Withers (1982) with $k_{r-1}^{i_1\cdots i_r}$ replaced by $\kappa_r dt_{r\wedge}$, $k_r^{i_1\cdots i_r}$ replaced by $\kappa_r d\ell_n(t_{r\wedge})$, other $k_j^{i_1\cdots i_r}$ replaced by 0, Σ replaced by \int and $t_{ij\ldots}$ replaced by $T_m(t_1t_2\cdots)$. For example,

$$a_{21} = \sum_{ij} t_i k_1^{ij} t_j = \kappa_2 \int \int T_m(t_1) T_m(t_2) dt_{2\wedge} = \kappa_2 \int T_m(t)^2 dt = \kappa_2 \ [1^2]$$

and the second term in a_{11} is

$$\sum_{ij} t_{ij} k_1^{ij} / 2 = \kappa_2 \int \int T_m(t_1, t_2) dt_{2\wedge} / 2 = \kappa_2 \int \int T_m(t, t) dt / 2 = \kappa_2 \ [11] / 2.$$

The first term in a_{11} is

$$\begin{aligned} t_i k_1^i &= \kappa_1 \int T_m(t) d\ell_n(t) = \kappa_1 \alpha_1(T_1) + n^{-1} \kappa_1 \alpha_2(T_1) + O(n^{-2}) \\ &= \kappa_1 (T_m(1) - T_m(0))/2 + n^{-1} \kappa_1 (T_m(1) - T_m(0))/12 + O(n^{-2}), \end{aligned}$$

for $\alpha_i(g)$ of the Euler-McLaurin expansion (A.1) of Withers and Nadarajah (2009, Appendix A), since this implies that

$$\int_0^1 g(t)d\ell_n(t) = \sum_{k=0}^{2K-2} \alpha_{k+1}(g)n^{-k} + nr_{n,2K},$$

where $r_{nk} = O(n^{-k})$ if $g^{(k)}$ is bounded on [0, 1]. The first term in a_{22} is

$$t_i k_2^{ij} t_j = \kappa_2 \int \int T_1 T_2 d\ell_n(t_{2\wedge}) = \kappa_2 \int T_1^2 d\ell_n(t_1) = \kappa_2 \alpha_1(T_1^2) + O(n^{-1})$$

= $\kappa_2 (T_m(1)^2 - T_m(0)^2)/2 + O(n^{-1}).$

By the corrigendum to Withers (1982), the fifth term in a_{22} is

$$2t_i k_1^{ij} t_{jk} k_1^k = 2\kappa_1 \kappa_2 \int \int \int T_1 T_{23} dt_{2\wedge} d\ell_n(t_3) = 2\kappa_1 \kappa_2 \alpha_1(g_3) + O(n^{-1})$$
$$= \kappa_1 \kappa_2(g_3(1) - g_3(0)) + O(n^{-1})$$

for $g_3(t) = \int_0^1 T_m(s) T_m(s,t) ds$. In this way, discarding the terms in a_{rr} of magnitude $O(n^{-1})$ or less, one obtains

$$\begin{aligned} a_{10} &= T(m), \\ a_{21} &= \kappa_2 \ [1^2], \\ a_{11} &= \kappa_1 \alpha_1(T_m) + \kappa_2 \ [11]/2, \\ a_{32} &= \kappa_3 \ [1^3] + 3\kappa_2^2 [1, 12, 2], \\ a_{22} &= \kappa_2 \alpha_1(T_m^2) + \kappa_3 \ [1, 11] + \kappa_2^2 \ [12^2]/2 + \kappa_2^2 \ [1, 122] \\ &+ 2\kappa_1 \kappa_2 \alpha_1(g_3), \\ a_{43} &= \kappa_4 \ [1^4] + 12\kappa_2 \kappa_3 \ [1^2, 12, 2] + 12\kappa_2^3 \ [1, 12, 23, 3] \\ &+ 4\kappa_3^2 \ [1, 2, 3, 123]. \end{aligned}$$

$$\begin{aligned} (3.1) \\ (3.2) \\ (3.2) \\ (3.3) \\ (3.3) \end{aligned}$$

Note the formal equivalence, apart from the bias terms α_k , with the expressions for the cumulant coefficients of $T(F_n)$ given by Withers (1983, page 580). (The negative term in a_{43} does not appear here due to the different form there for the *r*th order cumulants of \hat{w} .) So, we obtain from Withers (1983) the coefficients needed to extend inference to $O(n^{-2})$:

$$\begin{aligned} a_{12} &= \kappa_1 \alpha_2(T_m) + \kappa_3 \ [111]/6 + \kappa_2^2 \ [1122]/8, \\ a_{33} &= 3\kappa_4 \ [1^2, 11]/2 + 3\kappa_2 \kappa_3 [1, 12, 22] + 3\kappa_2 \kappa_3 \ [1, 12^2] + 3\kappa_2 \kappa_3 \ [1, 2, 122] \\ &+ 3\kappa_2 \kappa_3 \ [1^2, 122]/2 + 3\kappa_2^3 \ [1, 2, 1233]/2 + \kappa_2^3 \ [12, 23, 31] \\ &+ 3\kappa_2^3 \ [1, 12, 233] + 3\kappa_2^3 \ [1, 23, 123] + d_3, \\ a_{54} &= \kappa_5 \ [1^5] + 20\kappa_2 \kappa_4 \ [1, 2^3, 12] + 60\kappa_2^2 \kappa_3 \ [1, 2, 3, 12, 23] \\ &+ 60\kappa_2^2 \kappa_3 \ [1, 2^2, 13, 23] + 30\kappa_2^2 \kappa_3 \ [1, 2, 3^2, 123] \\ &+ 5\kappa_2^4 \ [1, 2, 3, 4, 1234] + 60\kappa_2^4 \ [1, 2, 34, 13, 24] \\ &+ 60\kappa_4^4 \ [1, 2, 3, 14, 234] \end{aligned}$$

where d_3 of a_{33} is the bias term, obtainable from ${}_3\Delta^{abc}$ of Withers (1983, page 67) as $d_3 = \alpha_1(g) = (g(1) - g(0))/2$ for

$$g(s) = 3\kappa_1\kappa_3 \int_0^1 T_m(s,t)T_m(t)^2 dt + 3\kappa_1\kappa_2^2 \int_0^1 \int_0^1 \{T_m(s,t,u)T_m(t)T_m(u) + 2T_m(s,t)T_m(t,u)T_m(u)\} dt du.$$

Example 3.1 Suppose that $T(g) = \int_0^1 w(g(s), s) ds$. By (2.7) the rth derivative at g = m is $\int_{t_{r\vee}}^1 f_r(s) ds$, where $f_r(s) = w_{.r0}(m(s), s)$. So, a_{21}, a_{11}, a_{32} are given by (3.1)–(3.3) in terms of

$$[1^{r}] = \int \cdots \int s_{r\wedge} \prod_{i=1}^{r} \{f_{1}(s_{i})ds_{i}\},\$$

$$[11] = \int sf_{2}(s)ds,\$$

$$\alpha_{1}(T_{m}) = -\int f_{1}(s)ds/2,\$$

$$[1, 12, 2] = \int \int \int \int f_{2}(s_{0})ds_{0} \prod_{i=1}^{2} \{\min(s_{0}, s_{i})f_{1}(s_{i})ds_{i}\}.$$

In particular,

$$a_{21} = \kappa_2 \int ds_1 \int ds_2 f_1(s_1) f_1(s_2) s_{2\wedge},$$

$$2a_{11} = -\mu \int f_1(s) ds + \kappa_2 \int s f_2(s) ds.$$

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References

- Abramowitz, M. and Stegun, I. A. (1964). Handbook of Mathematical Functions. U.S. Department of Commerce, National Bureau of Standards, Applied Mathematics Series, Volume 55.
- [2] Anderson, T. W. and Darling, D. A. (1952). Asymptotic theory of certain 'goodness of fit' criteria based on stochastic processes. *Annals of Mathematical Statistics*, 23, 193–212.
- [3] Bednarski, T., Clarke, B. R. and Kolkiewicz, W. (1991). Statistical expansions and locally uniform Fréchet differentiability. *Journal of the Australian Mathematical Society*, A, 50, 88–97.
- [4] Billingsley, P. (1968). Convergence of Probability Measures. Wiley, New York.
- [5] Comtet, L. (1974). Advanced Combinatorics. Reidel, Dordrecht.
- [6] Dudley, R. M. (1992). Fréchet differentiability, p-variation and uniform Donsker classes. Annals of Probability, 20, 1968–1982.
- [7] Pons, O. and de Turckheim, E. (1991). Von Mises method, bootstrap and Hadamard differentiability for nonparametric general models. *Statistics*, 22, 205–214.
- [8] Ren, J. J. and Sen, P. K. (1991). On Hadamard differentiability of extended statistical functional. *Journal of Multivariate Analysis*, **39**, 30–43.
- [9] Ren, J. J. and Sen, P. K. (1995). Hadamard differentiability on D[0,1]^p. Journal of Multivariate Analysis, 55, 14–28.
- [10] Von Mises, R. (1947). On the asymptotic distribution of differentiable statistical functions. Annals of Mathematical Statistics, 18, 309–342.
- [11] Withers, C. S. (1982). Second order inference for asymptotically normal random variables. Sankhyā, B, 44, 1–9.
- [12] Withers, C. S. (1983). Expansions for the distribution and quantiles of a regular functional of the empirical distribution with applications to nonparametric confidence intervals. Annals of Statistics, 11, 577–587.

ACCEPTED MANUSCRIPT

- [13] Withers, C. S. (1988). Nonparametric confidence intervals for functions of several distributions. Annals of the Institute of Statistical Mathematics, 40, 727–746.
- [14] Withers, C. S. and Nadarajah, S. (2009). Accurate tests and intervals based on linear cusum statistics. *Statistics and Probability Letters*, **79**, 689–697.