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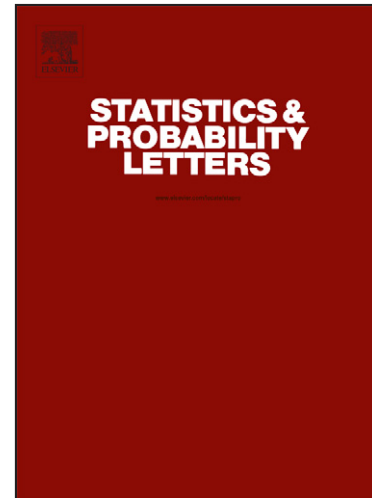
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Kolmogorov Inequalities for the Partial Sum of Independent Bernoulli Random Variables

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Abstract

In this paper two Kolmogorov inequalities are presented for the sample average of independent (but not necessarily identically distributed) Bernoulli random variables.

Keywords and phrases: Kolmogorov inequalities, Bernoulli random variables.

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1. Introduction

For a sequence of independent and identically distributed Bernoulli random variables X_1, X_2, \dots , with $E(X_1) = p$, Kolmogorov (1963) provided the following inequality

$$P(\sup_{k \geq n} |\bar{X}_k - p| > \epsilon) \leq 2e^{-2n\epsilon^2(1-\epsilon)}$$

where

$$\bar{X}_k = \frac{1}{k} \sum_{i=1}^k X_i, \quad \epsilon > 0.$$

Improvements, extensions and related results can be found in Hoeffding (1963), Young, Seaman and Marco (1987), Turner, Young and Seaman (1992), Young, Turner and Seaman (1988), Christofides (1991), Christofides (1994), Kambo and Kotz (1966) and Banjevic (1985). In this paper, two Kolmogorov inequalities are provided for the case of independent but not necessarily identically distributed Bernoulli random variables.

2. Preliminaries

The following results will be used:

Lemma 2.1

Let p_1, \dots, p_n be positive real numbers and $\bar{p} = \frac{1}{n} \sum_{i=1}^n p_i$. Then for $t > 0$

$$\prod_{i=1}^n (p_i e^t + 1 - p_i) \leq (\bar{p} e^t + 1 - \bar{p})^n.$$

Proof: From the arithmetic-geometric mean inequality, we have

$$\begin{aligned} \prod_{i=1}^n (p_i e^t + 1 - p_i) &\leq \left[\sum_{i=1}^n \frac{1}{n} (p_i e^t + 1 - p_i) \right]^n \\ &= (\bar{p} e^t + 1 - \bar{p})^n. \end{aligned}$$

The following result is due to Christofides (1994).

Lemma 2.2

Let $\epsilon < \frac{1}{2}$ and

$$g(\bar{p}, \epsilon) = (1 - \bar{p} - \epsilon) \ln\left(\frac{1 - \bar{p} - \epsilon}{1 - \bar{p}}\right) + (\bar{p} + \epsilon) \ln\left(\frac{\bar{p} + \epsilon}{\bar{p}}\right).$$

Then for $\bar{p} + \epsilon < \frac{1}{2}$ or $\frac{1}{2} + \frac{1}{3}\epsilon \leq \bar{p} \leq 1$

$$g(\bar{p}, \epsilon) \geq -\frac{1}{2} \ln(1 - 4\epsilon^2).$$

Lemma 2.3

Let $x = 2(\bar{p} + \epsilon) - 1$ and $y = 1 - 2\bar{p}$. Then,

$$\sum_{r=1}^{\infty} \frac{1}{2r(2r-1)} [x^{2r} + (2r-1)y^{2r} + 2rxy^{2r-1}] = g(\bar{p}, \epsilon)$$

where $g(\bar{p}, \epsilon)$ is the function defined in Lemma 2.2.

Proof: We have $\bar{p} + \epsilon = \frac{x+1}{2}$ and $\bar{p} = \frac{1-y}{2}$. Then, $g(\bar{p}, \epsilon)$ from Lemma 2.2 is

$$g(\bar{p}, \epsilon) = \left(\frac{1-x}{2}\right) \ln\left(\frac{1-x}{1+y}\right) + \left(\frac{1+x}{2}\right) \ln\left(\frac{1+x}{1-y}\right).$$

Using the Taylor series expansions

$$\ln(1+x) = \sum_{r=1}^{\infty} (-1)^{r-1} \frac{x^r}{r} \quad \text{and} \quad \ln(1-x) = -\sum_{r=1}^{\infty} \frac{x^r}{r}$$

we have that

$$g(\bar{p}, \epsilon) = \frac{(1-x)}{2} \left[-\sum_{r=1}^{\infty} \frac{x^r}{r} - \sum_{r=1}^{\infty} (-1)^{r-1} \frac{y^r}{r} \right] + \frac{(1+x)}{2} \left[\sum_{r=1}^{\infty} (-1)^{r-1} \frac{x^r}{r} + \sum_{r=1}^{\infty} \frac{y^r}{r} \right]$$

and after algebraic manipulations we arrive at

$$g(\bar{p}, \epsilon) = \sum_{r=1}^{\infty} \frac{1}{2r(2r-1)} [x^{2r} + (2r-1)y^{2r} + 2rxy^{2r-1}].$$

Lemma 2.4

Let ν be a positive integer and $x > 1$. Then

$$2^{2\nu-2}(x^{2\nu-1} + 1) - (x+1)^{2\nu-1} > 0.$$

Proof: Let $F(x) = 2^{2\nu-2}(x^{2\nu-1} + 1) - (x+1)^{2\nu-1}$. Then

$$F'(x) = 2^{2\nu-2}(2\nu-1)x^{2\nu-2} - (2\nu-1)(x+1)^{2\nu-2}$$

$$= (2\nu-1)[2^{2\nu-2}x^{2\nu-2} - (x+1)^{2\nu-2}]$$

$$= (2\nu-1)[(2x)^{2\nu-2} - (x+1)^{2\nu-2}] > 0 \text{ since } x > 1.$$

Thus, $F(x)$ is an increasing function and $F(x) > F(1) = 0$.

Lemma 2.5

Let $y \geq 1$ and $\nu = 1, 2, \dots$. Then

$$H(y) = y^{2\nu} + 2\nu - 1 + 2\nu y - \frac{4}{2^{2\nu}}(y+1)^{2\nu} \geq 0.$$

Proof:

$$H'(y) = 2\nu \cdot y^{2\nu-1} + 2\nu - \frac{4}{2^{2\nu}} 2\nu (y+1)^{2\nu-1}$$

$$= 2\nu \frac{4}{2^{2\nu}} \left[\frac{2^{2\nu}}{4} y^{2\nu-1} + \frac{2^{2\nu}}{4} - (y+1)^{2\nu-1} \right]$$

$$\begin{aligned}
 &= 2^{3-2\nu} \cdot \nu [2^{2\nu-2} y^{2\nu-1} + 2^{2\nu-2} - (y+1)^{2\nu-1}] \\
 &= 2^{3-2\nu} \cdot \nu [2^{2\nu-2} (y^{2\nu-1} + 1) - (y+1)^{2\nu-1}].
 \end{aligned}$$

By Lemma 2.4 $H'(y) > 0$ implying that H is increasing and therefore $H(x) \geq H(1) = 0$.

Lemma 2.6

Let x, y be as in Lemma 2.3 and $r = 1, 2, \dots$. Then

$$x^{2r} + (2r-1)y^{2r} + 2rxy^{2r-1} \geq 4\left(\frac{x+y}{2}\right)^{2r}.$$

Proof: By lemma 2.5 for $c \geq 1$

$$c^{2r} + 2r - 1 + 2rc - \frac{4}{2^{2r}}(c+1)^{2r} \geq 0.$$

Take $c = \frac{x}{y}$. Then

$$\begin{aligned}
 \frac{x^{2r}}{y^{2r}} + 2r - 1 + 2r\frac{x}{y} &\geq \frac{4}{2^{2r}}\left(\frac{x}{y} + 1\right)^{2r} \\
 x^{2r} + (2r-1)y^{2r} + 2rxy^{2r-1} &\geq 4\left(\frac{x+y}{2}\right)^{2r}.
 \end{aligned}$$

3. Main Results

Theorem 3.1

Let Y_1, Y_2, \dots, Y_n be a sequence of independent Bernoulli random variables with $E(Y_i) = p_i$, $i = 1, \dots, n$ and $\epsilon < \frac{1}{2}$. Then, for $\bar{p} + \epsilon < \frac{1}{2}$ or $\bar{p} \geq \frac{1}{2} + \frac{1}{3}\epsilon$,

$$P(\bar{Y} - \bar{p} > \epsilon) \leq (1 - 4\epsilon^2)^{\frac{n}{2}}$$

where $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$ and $\bar{p} = \frac{1}{n} \sum_{i=1}^n p_i$.

Proof: Let $s > 0$. Then,

$$\begin{aligned}
 P(\bar{Y} - \bar{p} > \epsilon) &= P[s(\bar{Y} - \bar{p} - \epsilon) > 0] \\
 &\leq E[e^{s(\bar{Y} - \bar{p} - \epsilon)}] \\
 &= e^{-s(\bar{p} + \epsilon)} E[e^{s\bar{Y}}] \\
 &= e^{-s(\bar{p} + \epsilon)} E(e^{s \frac{1}{n} \sum_{i=1}^n Y_i}) \\
 &= e^{-s(\bar{p} + \epsilon)} \prod_{i=1}^n E(e^{\frac{s}{n} Y_i}) \\
 &= e^{-s(\bar{p} + \epsilon)} \prod_{i=1}^n (p_i e^{\frac{s}{n}} + 1 - p_i) \\
 &\leq e^{-s(\bar{p} + \epsilon)} (\bar{p} e^{\frac{s}{n}} + 1 - \bar{p})^n \tag{1} \\
 &= e^{-f(s)}
 \end{aligned}$$

where $f(s) = s(\bar{p} + \epsilon) - n \ln(\bar{p} e^{\frac{s}{n}} + 1 - \bar{p})$ and the last inequality follows from Lemma 2.1.

The function f is maximized at $s^{max} = n \ln[\frac{(1-\bar{p})(\bar{p}+\epsilon)}{\bar{p}(1-\bar{p}-\epsilon)}]$ and

$$\begin{aligned}
 f(s^{max}) &= n(\bar{p} + \epsilon) \ln \left[\frac{(\bar{p} + \epsilon)(1 - \bar{p})}{\bar{p}(1 - \bar{p} - \epsilon)} \right] - n \ln \left[\frac{1 - \bar{p}}{1 - \bar{p} - \epsilon} \right] \\
 &= n[(1 - \bar{p} - \epsilon) \ln(\frac{1 - \bar{p} - \epsilon}{1 - \bar{p}}) + (\bar{p} + \epsilon) \ln(\frac{\bar{p} + \epsilon}{\bar{p}})] \\
 &= n.g(\bar{p}, \epsilon).
 \end{aligned}$$

Thus, by Lemma 2.2

$$f(s^{max}) = n.g(\bar{p}, \epsilon) \geq -\frac{n}{2} \ln(1 - 4\epsilon^2)$$

and therefore

$$P(\bar{Y} - \bar{p} > \epsilon) \leq (1 - 4\epsilon^2)^{\frac{n}{2}}.$$

Notice that the restrictions on \bar{p} and ϵ arise from Lemma 2.2 which is due to Christofides (1994).

The following theorem gives an exponential bound under different conditions on \bar{p} and ϵ .

Theorem 3.2

Let Y_1, Y_2, \dots, Y_n be a sequence of independent Bernoulli random variables, with $E(Y_i) = p_i$, $i = 1, \dots, n$. Then for $\bar{p} + \epsilon > \frac{1}{2}$ or $\bar{p} < \frac{1}{2}$ and $\forall \epsilon < 1$,

$$P(\bar{Y} - \bar{p} > \epsilon) \leq e^{-n(2\epsilon^2 + \frac{1}{3}\epsilon^4 e^{\frac{\epsilon^2}{4}})}$$

where $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$ and $\bar{p} = \frac{1}{n} \sum_{i=1}^n p_i$.

Proof: From the proof of Theorem 3.1

$$P(\bar{Y} - \bar{p} > \epsilon) \leq e^{-ng(\bar{p}, \epsilon)}$$

where

$$g(\bar{p}, \epsilon) = (1 - \bar{p} - \epsilon) \ln\left(\frac{1 - \bar{p} - \epsilon}{1 - \bar{p}}\right) + (\bar{p} + \epsilon) \ln\left(\frac{\bar{p} + \epsilon}{\bar{p}}\right).$$

By Lemma 2.3

$$g(\bar{p}, \epsilon) = \sum_{r=1}^{\infty} \frac{1}{2r(2r-1)} [x^{2r} + (2r-1)y^{2r} + 2rxy^{2r-1}].$$

x and y are positive values, imposing the restrictions on \bar{p} and ϵ .

Now, since $x + y = 2\epsilon$ and using Lemma 2.6 we have

$$\begin{aligned}
 g(\bar{p}, \epsilon) &\geq \sum_{r=1}^{\infty} \frac{1}{2r(2r-1)} \cdot 4 \left(\frac{2\epsilon}{2}\right)^{2r} \\
 &= \sum_{r=1}^{\infty} \frac{2\epsilon^{2r}}{r(2r-1)} \\
 &= 2\epsilon^2 + \epsilon^4 \sum_{r=2}^{\infty} \frac{2(\epsilon)^{2r-4}}{r(2r-1)} \\
 &\geq 2\epsilon^2 + 2\epsilon^4 \sum_{r=2}^{\infty} \frac{(\epsilon^2)^{r-2}}{6.4^{r-2}(r-2)!}
 \end{aligned}$$

since $r(2r-1) \leq 6.4^{r-2}(r-2)!$ for $r \geq 2$. The last inequality is clearly true for $r = 2$, and can be established (e.g., by induction) for $r > 2$. Then,

$$\begin{aligned}
 g(\bar{p}, \epsilon) &\geq 2\epsilon^2 + 2\epsilon^4 \frac{1}{6} \sum_{r=2}^{\infty} \left(\frac{\epsilon^2}{4}\right)^{r-2} \frac{1}{(r-2)!} \\
 &= 2\epsilon^2 + \frac{1}{3}\epsilon^4 \cdot e^{\frac{\epsilon^2}{4}}.
 \end{aligned}$$

Thus

$$P(\bar{Y} - \bar{p} > \epsilon) \leq e^{-n(2\epsilon^2 + \frac{1}{3}\epsilon^4 \cdot e^{\frac{\epsilon^2}{4}})}$$

and the proof is complete.

Theorem 3.3

Let Y_1, Y_2, \dots , be a sequence of independent Bernoulli random variables with $E(Y_i) = p_i$. Then,

(1) $\forall \epsilon < \frac{1}{2}$ and for $\bar{p}_k + \epsilon < \frac{1}{2}$ or $\bar{p}_k \geq \frac{1}{2} + \frac{1}{3}\epsilon$,

$$P[\sup_{k \geq n} (\bar{Y}_k - \bar{p}_k) > \epsilon] \leq (1 - 4\epsilon^2)^{\frac{n}{2}} \quad n = 0, 1, 2, \dots$$

while

(2) $\forall \epsilon < 1$ and for $\bar{p}_k + \epsilon > \frac{1}{2}$ or $\bar{p}_k < \frac{1}{2}$,

$$P[\sup_{k \geq n} (\bar{Y}_k - \bar{p}_k) > \epsilon] \leq e^{-n(2\epsilon^2 + \frac{1}{3}\epsilon^4 e^{\frac{\epsilon^2}{4}})} \quad n = 0, 1, 2, \dots$$

where $\bar{Y}_k = \frac{1}{k} \sum_{i=1}^k Y_i$ and $\bar{p}_k = \frac{1}{k} \sum_{i=1}^k p_i$.

Proof: Using Lemma 1 of Turner, Young and Seaman (1995) we can arrive at inequality (1) of Theorem 3.1 having the quantity $P[\sup_{k \geq n} (\bar{Y}_k - \bar{p}_k) > \epsilon]$ as our left-hand side. Then, one can follow the same steps of theorems 3.1 and 3.2 to reach the right-hand side, and the proof is complete.

Remarks

(1) Theorem 3.3 provides sharper bounds than that of the main result of Turner, Young and Seaman (1995), under of course restrictions on \bar{p} and ϵ .

(2) Theorem 3.1 is an extension of Corollary 3.2 of Christofides (1994).

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