

# Sensitivity analysis to identify key parameters influencing infection dynamics in a pig batch

Amandine Lurette, Suzanne Touzeau, Matieyendou Lamboni, Hervé Monod

# ▶ To cite this version:

Amandine Lurette, Suzanne Touzeau, Matieyendou Lamboni, Hervé Monod. Sensitivity analysis to identify key parameters influencing infection dynamics in a pig batch. Journal of Theoretical Biology, 2009, 258 (1), pp.43-52. 10.1016/j.jtbi.2009.01.026. hal-00554569

HAL Id: hal-00554569

https://hal.science/hal-00554569

Submitted on 11 Jan 2011

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

# Author's Accepted Manuscript

Sensitivity analysis to identify key parameters influencing *Salmonella* infection dynamics in a pig batch

Amandine Lurette, Suzanne Touzeau, Matieyendou Lamboni, Hervé Monod

PII: S0022-5193(09)00030-7

DOI: doi:10.1016/j.jtbi.2009.01.026

Reference: YJTBI 5441

To appear in: Journal of Theoretical Biology

Received date: 28 May 2008 Revised date: 21 January 2009 Accepted date: 21 January 2009



www.elsevier.com/locate/yjtbi

Cite this article as: Amandine Lurette, Suzanne Touzeau, Matieyendou Lamboni and Hervé Monod, Sensitivity analysis to identify key parameters influencing *Salmonella* infection dynamics in a pig batch, *Journal of Theoretical Biology* (2009), doi:10.1016/j.jtbi.2009.01.026

This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting galley proof before it is published in its final citable form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

# Sensitivity analysis to identify key parameters influencing Salmonella infection dynamics in a pig batch

Amandine Lurette <sup>a,b,\*</sup> Suzanne Touzeau <sup>c</sup>

Matievendou Lamboni <sup>c</sup> Hervé Monod <sup>c</sup>

<sup>a</sup>INRA, UMR1300 Bio-agression, Épidémiologie et Analyse de Risque, F-44300 Nantes, France

<sup>b</sup>ENVN, UMR1300 Bio-agression, Épidémiologie et Analyse de Risque, F-44300 Nantes, France

<sup>c</sup>INRA, UR341 Mathématiques et Informatique Appliquées, F-78350 Jouy-en-Josas, France

#### **Abstract**

In the context of managed herds, epidemiological models usually take into account relatively complex interactions involving a high number of parameters. Some parameters may be uncertain and/or highly variable, especially epidemiological parameters. Their impact on the model outputs then must then be assessed by a sensitivity analysis, allowing to identify key parameters. The prevalence over time is an output of particular interest in epidemiological models, so sensitivity analysis methods adapted to such dynamic output are needed.

In this paper, such a sensitivity analysis method, based on a principal component analysis and on analysis of variance, is presented. It allows to compute a

generalised sensitivity index for each parameter of a model representing Salmonella spread within a pig batch. The model is a stochastic discrete-time model describing the batch dynamics and movements between rearing rooms, from birth to slaughterhouse delivery. Four health states were introduced: Salmonella-free, seronegative shedder, seropositive shedder and seropositive carrier. The indirect transmission was modelled via an infection probability function depending on the quantity of Salmonella in the rearing room.

Simulations were run according to a fractional factorial design enabling the estimation of main effects and two-factor interactions. For each of the 18 epidemiological parameters, four values were chosen, leading to 4096 scenarios. For each scenario, 15 replications were performed, leading to 61 440 simulations. The sensitivity analysis was then conducted on the seroprevalence output.

The parameters governing the infection probability function and residual room contaminations were identified as key parameters. To control the *Salmonella* sero-prevalence, efficient measures should therefore aim at these parameters. Moreover, the shedding rate and maternal protective factor also had a major impact. Therefore, further investigation on the protective effect of maternal or post-infection antibodies would be needed.

Key words: epidemiological model, parameter uncertainty and variability, dynamic output, seroprevalence, analysis of variance

Suzanne.Touzeau@jouy.inra.fr (Suzanne Touzeau),

Matieyendou.Lamboni@jouy.inra.fr (Matieyendou Lamboni),

Herve.Monod@jouy.inra.fr (Hervé Monod).

<sup>\*</sup> Corresponding author. Tel.: +33(0)240687821; fax: +33(0)240687768. Email addresses: Amandine.Lurette@vet-nantes.fr (Amandine Lurette),

#### 1 1 Introduction

Epidemiological models are useful tools for representing infection processes. They are often used to predict the spread of a pathogen in a population or to assess the efficacy of control measures on that spread (Diekmann and Heesterbeek, 2000). In managed systems such as pig herds, they usually couple the population dynamics and the pathogen transmission, to take into account the relatively complex interactions between demography, management and infection. In general, parameters governing the herd population dynamics are well known, even though they may exhibit some stochasticity due the variability among individuals in the batch. Their values and range of variation can be assessed from various data sources (ITP, 2000, 2006). In contrast, hostpathogen interactions are not so well known. This lack of knowledge induces uncertainty on the parameters involved in the herd infection processes. The impact of these uncertain and variable parameters on the pathogen spread can generally be best assessed by a sensitivity analysis, as analytical tractability in such complex models is not ensured. A sensitivity analysis describes the effect parameter variations on model outputs. It identifies key parameters, i.e. parameters that influence the outputs most, which is useful for modelling and control purposes: on the one hand, improving knowledge on these parameters increases the model accuracy; on the other hand, control measures acting on these parameters will have a greater impact on the outputs.

The prevalence of seropositive animals over time (seroprevalence trajectory) is an output of particular interest in epidemiological models, as control measures often aim at reducing the prevalence and may not be applied uniformly over time. Most sensitivity analysis methods are adapted to scalar outputs (Saltelli

- et al., 2000) and cannot be applied to such dynamic outputs. However, some
- 27 methods have been developed to analyse time series outputs (Campbell et al.,
- 28 2006). They use classical statistical tools on variables that summarise the
- 29 trajectory variability. The method developed by Lamboni et al. (2008) is based
- on a principal component analysis applied to the trajectories. It maximises the
- information that is retained from the output variability. It has already been
- 32 used to analyse various agronomic models and can easily be applied to an
- 33 epidemiological model.
- The control of Salmonella prevalence in pig herds is a major step to reduce
- 35 the pork food chain contamination. Infected pigs are asymptomatic carriers
- that cannot be directly detected by observation. Serological and bacteriolog-
- ical detection methods do exist, but they cannot be systematically applied
- to all pigs sent to the slaughterhouse, for practical and economical reasons.
- Moreover, Salmonella are ubiquitous bacteria that survive well in the environ-
- 40 ment (Murray, 2002). Hence, eliminating the bacteria in a contaminated herd
- 41 is not a very realistic goal. Control measures should rather aim at reducing
- 42 the prevalence level. To achieve this goal, the parameters that influence the
- 43 Salmonella spread and prevalence most need to be identified. A modelling
- 44 approach coupled with a sensitivity analysis can be applied to identify these
- 45 key parameters.
- A simulation model representing the Salmonella spread within a farrow-to-
- 47 finish pig herd implementing batch-farrowing management was developed by
- Lurette et al. (2008a). This model details both the sow and pig dynamics
- <sup>49</sup> (Lurette et al., 2008b) and the Salmonella infection characteristics. In such a
- 50 herd, sows are distributed in groups of equal size called batches and their re-
- 51 productive cycles are synchronised. At regular intervals (usually three weeks),

- as a batch of sows is inseminated and their weaned pigs are managed together
  as a batch until their slaughterhouse delivery. Each batch of pigs is usually
  housed in a separate room, which is decontaminated between two batches.

  Batches of pigs may interact, mainly through batch mixing at the end of the
  finishing period. However, even if it occurs under normal pig growth conditions
  (i.e. adequate feed and no diseases affecting pig growth), it only concerns a
  low number of pigs (Lurette et al., 2008b). Salmonella transmission is mostly
  due to indirect contacts among animals via their rearing room. Infected pigs
  shed the bacteria intermittently (Kranker et al., 2003); it is then transmitted
  to susceptible pigs by the oro-fecal route. Due to the limited contacts between batches, Salmonella transmission is stronger at the within-batch level.
  So we chose to focus in this study on a single batch of pigs, from birth to
  slaughterhouse delivery.
- The aim of this paper is therefore to identify the parameters that contribute most to the variability of the *Salmonella* seroprevalence in a batch of pigs, using a sensitivity analysis based on a principal components analysis. We will first describe the pig batch epidemiological model, followed by the sensitivity analysis. Then we will present and discuss the results obtained with this method.

# 71 2 Model description

We developed a stochastic model to represent the *Salmonella* transmission within a batch of pigs, from their birth to their slaughterhouse delivery. The model has a discrete one-week time step, which is suited both to the representation of the farmer task planning and to the epidemiological time scale.

- As the carriage of Salmonella in pigs is generally subclinic, the infection does
- not affect the pig demography. Hence, the pig batch model and the epidemi-
- ological model are first presented separately. They are then coupled by dis-
- 79 tributing the animals of the batch into the infection states. Finally, the pa-
- 80 rameters and initial conditions are described.

#### 81 2.1 Pig batch model

- Pigs in a batch are born simultaneously from a batch of sows at time t=0.
- The litter size L is fixed. Pigs pass through three growing stages during their
- lifetime, denoted by X, corresponding to the occupancy of three rooms: (i)
- the suckling stage, a 4-week period from birth to weaning  $(X = S, D_S = 4)$ ,
- takes place with the sows in the farrowing room; (ii) the 8-week post-weaning
- period  $(X = PW, D_{PW} = 8)$  takes place in the post-weaning room; (iii) the
- 88 finishing stage takes place in the finishing room and lasts up to 14 weeks
- $_{\mbox{\tiny 89}} \ \ (X=F,D_F=14)$  . The number of pigs in the batch at time  $t\in\{0,1,\ldots,25\}$
- is given by P(t). At each time step, the number of pigs decreases because of
- 91 mortality and, at the end of the finishing stage, because of slaughterhouse
- 92 delivery.
- The mortality rate  $\mu_X$  is assumed to be constant in each growing stage X and
- 94 differs between stages.
- 95 At each delivery time (usually every two weeks), producers send groups of
- 66 finishers with a given slaughtering weight. As out-of-range pigs are sold at a
- of lower price, they are kept in the herd until they reach the slaughtering weight
- 98 or until their room needs to be emptied. Because of pig growth variability,

a batch leaves the herd at different times. To represent the proportion of the batch that has reached the slaughtering weight over time, a lognormal 100 distribution is used (with parameters  $a_{\alpha}$ , mean age at slaughtering weight 101 and  $a_{\sigma}$ , its standard deviation). This distribution represents the variability 102 among individual animals reported by technical data (ITP, 2006). Let  $\varphi$  be 103 the associated probability density function and  $\Delta t$  the time interval between 104 two deliveries (usually  $\Delta t = 2$  weeks), assuming they start at t = 1. At each delivery time  $t = k\Delta t (k \in \mathbb{N})$ , the proportion of pigs sent to the slaughterhouse to the initial number of pigs P(0) = L, corresponding to pigs that have reached the required weight since the last delivery, is given by:  $\int_{t-\Delta t}^{t} \varphi(s) ds$ . Moreover, delivery of finishing pigs only occurs between 22 and 25 weeks of age.

All remaining pigs are delivered at the last delivery time (25 weeks of age)  $t_m = \max_{k \in \mathbb{N}} \{k\Delta t : k\Delta t \leq 25\}$  because the room needs to be emptied. The proportion of pigs delivered to the slaughterhouse at time t, calculated from the pig population P(t-1), is given by:

$$d(t) = \begin{cases} 0 & \text{if } t < 22 \text{ or if } t \neq k\Delta t, k \in \mathbb{N} \text{ (not a delivery time)}, \\ \frac{\int_{t-\Delta t}^{t} \varphi(s) ds}{1-\int_{0}^{t-\Delta t} \varphi(s) ds} & \text{if } 22 \leqslant t = k\Delta t < t_m, k \in \mathbb{N} \text{ (delivery time)}, \\ 1 & \text{if } t = t_m \text{ (room is emptied)}. \end{cases}$$

The number of pigs at time t is hence obtained from the following equations:

$$P(t) = \begin{cases} L & \text{if } t = 0 \text{ (birth),} \\ (1 - \mu_S)P(t-1) & \text{if } 0 < t \leqslant 4 \text{ (suckling period),} \\ (1 - \mu_{PW})P(t-1) & \text{if } 4 < t \leqslant 12 \text{ (post-weaning period),} \\ (1 - \mu_F)(1 - d(t))P(t-1) & \text{if } 12 < t \leqslant 25, \text{ (finishing period).} \end{cases}$$

At each time step, mortality is assumed to take place first, so slaughterhouse deliveries are given by:  $D(t) = (1 - \mu_F)d(t)P(t-1)$ .

# 112 2.2 Salmonella epidemiological model

The epidemiological model was developed by Lurette et al. (2008a). Infected animals shed Salmonella in their feces and, hence, contaminate their local environment, i.e. their rearing room. The model represents the indirect fecal-oral transmission via encounters with the free-living Salmonella in contaminated feces on the room floor, the pen separations and the pig bodies. The free-living Salmonella in the room are modelled as a quantity of Salmonella infectious units denoted Q(t) at time t.

Four mutually exclusive health states (Y) were identified and selected from the literature and from expert opinion. A susceptible pig free of Salmonella 121 (Y = F) becomes a shedding pig (Y = Sh) after ingesting the bacteria. After 122 a seroconversion delay, it is called a seropositive shedding pig (Y = Shs). The shedding is intermittent. When it stops, the pig becomes a seropositive carrier 124 pig (Y = Cs). The shedding can be reactivated, especially under stressful 125 conditions. The number of pigs in each state Y at time t is given by P(t,Y), where  $P(t) = \sum_{Y} P(t, Y)$  represents the total population size. All transitions are stochastic: the number of pigs which undergo a transition between two states is drawn in a binomial law; the first parameter is the number of pigs in 129 the initial state, the second is the individual transition probability.

Assuming homogeneous mixing, each pig is exposed to Q(t)/P(t) Salmonella infectious units. Ingestion is considered to have a logarithmic effect, so the

effective quantity of Salmonella infectious units for each susceptible pig is:  $q(t) = \log(1 + Q(t)/P(t))$ . The probability of infection for a susceptible pig is an increasing function of q with two thresholds:

$$f(q) = \begin{cases} 0 & \text{if } q = 0, \\ b_1 & \text{if } 0 < q \leqslant q_1, \\ b_1 + \frac{b_2 - b_1}{q_2 - q_1} (q - q_1) & \text{if } q_1 < q \leqslant q_2, \\ b_2 & \text{if } q > q_2. \end{cases}$$

This semi-log sigmoid-like function is very similar to the classical dose-response relationships, very commonly used in pharmacokinetics. The number of susceptible pigs that become infected is calculated at each time step from the following binomial law: Bin(P(t,F),f(q(t))).

The probability that a shedder pig becomes seropositive derives from the seroconversion delay  $\lambda_1$  and is equal to:  $1 - \exp(-1/\lambda_1)$ . Similarly, the probability that a seropositive shedding pig stops shedding *Salmonella* depends on the shedding period duration  $\lambda_2(t)$ . The shedding duration follows a lognormal distribution (with parameters  $\lambda_{2\alpha}$  and  $\lambda_{2\sigma}$ , mean and standard deviation) and is recalculated at each time step t to account for environmental variability. The probability that a seropositive carrier pig starts shedding again is  $\beta$ . As previously, the pigs which undergo a state change at each time step are drawn from binomial laws using these three individual transition probabilities.

Hence, we obtain the following system of equations, that describes the popu-

lation evolution in the four health states:

$$\begin{cases} P(t,F) = P(t-1,F) - \text{Bin}(P(t-1,F), f(q(t-1))), \\ P(t,Sh) = P(t-1,Sh) + \text{Bin}(P(t-1,F), f(q(t-1))) \\ - \text{Bin}\left(P(t-1,Sh), 1 - \exp\left(\frac{-1}{\lambda_1}\right)\right), \\ P(t,Shs) = P(t-1,Shs) + \text{Bin}\left(P(t-1,Sh), 1 - \exp\left(\frac{-1}{\lambda_1}\right)\right) \\ - \text{Bin}\left(P(t-1,Shs), 1 - \exp\left(\frac{-1}{\lambda_2(t-1)}\right)\right) + \\ - \text{Bin}(P(t-1,Cs),\beta), \\ P(t,Cs) = P(t-1,Cs) + \text{Bin}\left(P(t-1,Shs), 1 - \exp\left(\frac{-1}{\lambda_2(t-1)}\right)\right) \\ - \text{Bin}(P(t-1,Cs),\beta). \end{cases}$$

Finally, the Salmonella infectious units are subject to a degradation process at each time step with a fixed rate  $\nu$ . They are incremented by the infectious units shed by seropositive and seronegative shedder pigs. The resulting dynamics of the Salmonella infectious units is the following:

$$Q(t) = (1 - \nu)Q(t - 1) + q_e P(t, Sh) + q_e \pi_+ P(t, Shs).$$

where  $q_e$  is the shedding rate, i.e. the quantity of infectious units shed by a seronegative shedder pig, and  $\pi_+$  is the shedding ratio between a seropositive and a seronegative shedder pig.

#### 2.3 Coupling the models

The pig batch model and the epidemiological models are coupled as follows.

At each time step, mortality and slaughterhouse delivery are first applied uniformly over the four epidemiological states in the batch. The number of

pigs in epidemiological state Y delivered to the slaughterhouse at time t is therefore :  $D(t,Y) = (1-\mu_F)d(t)P(t-1,Y)$ ;  $D(t) = \sum_Y D(t,Y)$  is the total number of pigs delivered. Then the epidemiological transitions are drawn and finally the number of Salmonella infectious units is updated.

Some epidemiological transitions and parameters depend on the pig growing stage. The susceptibility of piglets to Salmonella infection is considered lower during the suckling period than during the other growing stages, due to a maternal protective immunity. So, during the suckling period S, the probability of infection  $f^S(q) = \varepsilon f(q)$ ,  $\varepsilon$  being a maternal protective factor ( $\varepsilon < 1$ ); during all other stages  $X \in \{PW, F\}$ ,  $f^X(q) = f(q)$ . The shedding rate also depends on the growing stage. Let  $q_e$  be the shedding rate of a finisher pig and  $\pi_X$  be the shedding ratio between a pig in stage X = S, PW and a finishing pig; the shedding rate in stage X is  $q_e\pi_X$ . Finally, piglets undergo a supplementary stress at weaning, which increases the probability  $\beta$  that a carrier pig starts shedding again. Denoting by  $\gamma$  the value of this probability at weaning time t=5, at any other time, it is given by  $\beta=\pi_{\beta}\gamma$ , with  $\pi_{\beta}<1$ .

Between two batches, rooms are cleaned and disinfected, but residual Salmonella may remain in the room. When piglets are born, the room can also have been contaminated by the sows. So when the pig batch enters each rearing room, the number of Salmonella infectious units is set to a residual value denoted  $Q_X$ , X being the corresponding stage. The number of Salmonella infectious

units is then given by the following equations:

# 2.4 Model parameters and initial condition

- The demographic parameters used for the pig herd model are known from the
- published performance data (ITP, 2000, 2006) and are detailed in Table 1.
- The epidemiological parameters, however, are not so well known. The infor-
- mation available is presented in Table 2.
- At the beginning of the simulation, we assume that piglets are born Salmonella-
- free. So P(0,F)=L and P(0,Y)=0 for  $Y\neq F$ . Contamination of the batch
- is initiated by the residual Salmonella infectious units in the rooms  $Q_Y$ . Their

Table 1

Demographic parameters used in the batch model.

Name	Description	Value	
$D_S$	Duration of the suckling period	4 weeks*	
$D_{PW}$	Duration of the post-weaning period	8 weeks*	
$D_F$	Duration of the finishing period	14 weeks*	
$\mu_S$	Mortality rate for suckling pigs	$0.01~\mathrm{week^{-1}}^\dagger$	
$\mu_{PW}$	Mortality rate for post-weaning pigs	$0.005 \text{ week}^{-1\dagger}$	
$\mu_F$	Mortality rate for finishing pigs	$0.002 \text{ week}^{-1\dagger}$	
$a_{lpha}$	Mean age at slaughtering weight	25.5 weeks*	
$a_{\sigma}$	Standard deviation —	1.5 weeks*	
* from (ITP, 2000). $^{\dagger}$ from (ITP, 2006).			

values are given in Table 2. Slaughterhouse deliveries from the herd may occur every  $\Delta t = 2$  weeks, starting at t = 0; so the room is emptied at  $t_m = 25$ .

 $^{177}\,\,$  All simulations were performed using Scilab 4.0.  $^{1}\,\,$ 

# <sup>178</sup> 3 Sensitivity analysis

The aim of a sensitivity analysis is to define which parameters have the greatest impact on given model outputs. We chose two outputs. The first one is the Salmonella seroprevalence in the batch over time, from the piglet birth until Scilab, a free scientific software for numerical computations: http://www.scilab.org/

 $\label{eq:continuous} \mbox{Table 2}$  Epidemiological parameters used in the batch model.

Name	Description	Values tested
$q_1$	Lower infection probability threshold	$1, 2, 3, 4 \log(S.i.u.)$
$q_2$	Saturation infection probability threshold	4, 4.7, 5, 6 log(S.i.u.)
$b_1$	Lower infection probability	$10^{-7}$ , $10^{-6}$ , $10^{-4}$ , $10^{-2}$
$b_2$	Saturation infection probability	0.05,0.1,0.3,0.5
$\lambda_1$	Seroconversion delay	$1, 2, 3, 4 \text{ weeks}^*$
$\lambda_{2_lpha}$	Shedding period duration: mean	$1, 3, 6, 10 \text{ weeks}^{\dagger}$
$\lambda_{2_{\sigma}}$	Shedding period duration:standard deviation	1, 2, 3, 4 weeks
$\gamma$	Probability of shedding reactivation due to stress	0.1,0.2,0.4,0.6
$\pi_{eta}$	Ratio: shedding reactivation / $\gamma$	0.1, 0.4, 0.7, 1
arepsilon	Maternal protective factor	0, 0.1, 0.3, 1
ν	Salmonella degradation rate	$0, 0.3, 0.5, 0.8 \text{ week}^{-1\ddagger}$
$q_e$	Shedding rate for seronegative finishing pigs	$10^2, 5.10^3, 10^5, 5.10^6$
		$S.i.u.(week)^{-1}$
$\pi_S$	Ratio: shedding rate for suckling pigs / $q_e$	0.1,  0.2,  0.4,  0.6
$\pi_{PW}$	Ratio: shedding rate for post-weaning pigs / $q_e$	0.1, 0.2, 0.4, 0.6
$\pi_+$	Ratio: shedding rate for seropositive pigs / $q_e$	0.1,0.4,0.7,1
$Q_S$	Residual contamination: farrowing room	$0, 10^2, 10^4, 10^6$ S.i.u.
$Q_{PW}$	Residual contamination: post-weaning room	$0, 10^2, 10^4, 10^6$ S.i.u.
$Q_F$	Residual contamination: finishing room	$0, 10^2, 10^4, 10^6 \text{ S.i.u.}$

S.i.u.: Salmonella infectious units.

<sup>\*</sup> from (van Winsen et al., 2001; Fravalo et al., 2003).

 $<sup>^{\</sup>dagger}$  (Davies et al., 1998; Nielsen et al., 1995; van Winsen et al., 2001; Kranker et al., 2003; Wong et al., 2002).

 $<sup>^{\</sup>ddagger}$  (Plachá et al., 2001; Boes et al., 2005; Nicholson et al., 2005).

their slaughterhouse delivery. The seroprevalence corresponds to the proportion of seropositive shedding and carrying pigs in the batch. It is calculated at each time step by:

$$\psi(t) = \frac{\sum_{Y \in Shs, Cs} P(t, Y)}{P(t)}.$$

The second output is the mean seroprevalence at slaughterhouse delivery, calculated as follows:

$$\Omega = \frac{\sum_{t=22}^{25} \sum_{Y \in Shs, Cs} D(t, Y)}{\sum_{t=22}^{25} D(t)}$$

As the demographic parameters are well known, we chose to perform the analysis on the epidemiological parameters given in Table 2, whose values are uncertain. We used a design of experiments to define which simulations to run.

Then the resulting outputs were analysed.

# 3.1 Design of experiments

The simulation model described in the previous section has 18 epidemiological parameters, also called "factors". To study the sensitivity of the outputs to these parameters, we used a factorial experimental design, with four parameter values, called "levels", per factor. As the model is stochastic, each combination of levels of factors, i.e. each parameter set or "scenario", was replicated 15 times. This number of replications was found to be largely sufficient to get robust results. It also allowed to obtain quite precise estimates of the variances between replications for each scenario.

A complete factorial design corresponds to  $4^{18}$  scenarios, that is  $15 \times 4^{18}$  simulations, which is not feasible. Therefore, we used a fractional factorial design (Kobilinsky, 1997; Box et al., 1978). A resolution V design is orthogonal and

it enables the estimation of main effects and two-factor interactions, assuming
that the higher order interactions are negligible. A resolution V design for
18 factors can be obtained with 4096 scenarios, which corresponds to 61 440
simulations with 15 replications. The fractional factorial design we used was
produced by the FACTEX procedure of the SAS 8.0/QC module. Alternative designs could have been used, in particular response surface designs
(Myers and Montgomery, 2002) or space-filling designs (Fang et al., 2006). In
the present study, we gave priority to orthogonal and balanced designs.

# 203 3.2 Sensitivity analysis on the seroprevalence over time

To analyse the seroprevalences over time  $\psi(t)$  obtained with the fractional factorial design described above, we used a method developed by Lamboni et al. (2008). We only give here a short description of this method. The seroprevalence outputs can be represented in a  $N \times p$  table denoted T, with  $N=61\,440$  the number of simulations and p=23 the number of seroprevalence output times  $(t=2,\ldots,24)$ . As the animals are born Salmonella-free at t=0 and need a week for seroconversion when they become infected,  $\psi(0)=\psi(1)=0$  for each simulation; moreover, their room is emptied at  $t=t_m=25$ , so  $\psi(25)=0$ . Hence, these times are not included in the analysis.

A principal component analysis (PCA) with N individuals and p variables is performed on this table T. It consists in an eigenvalue decomposition of the matrix of correlations between the columns of T,  $\tilde{T}'\tilde{T}$ , where  $\tilde{T}$  is obtained by centring and normalising each column of T. Let  $u_1$  denote the first eigenvector

<sup>&</sup>lt;sup>2</sup> SAS/QC 8 User's Guide, Introduction to the FACTEX procedure: http://v8doc.sas.com/sashtml/qc/chap14/index.htm

associated with the largest eigenvalue. The first component  $u_1$  thus obtained is the linear combination of the variables that explains the maximum of variability – or inertia – among the simulations. The second component is the second best combination to explain the variability, conditionally on being orthogonal to the first component, etc. Usually, four components are enough to explain most of the variability among the simulations. In this study, only three components were needed.

As a result of the PCA, each simulation corresponding to a line in table T is given a "score" on each component. For the first component, it corresponds to the scalar product of  $u_1$  and the corresponding line of the T matrix. The scores are similarly obtained for each subsequent component. Each simulation is hence described by a scalar value called score for each component.

An ANOVA was then performed on the simulation scores for each of the first three principal components, to compare the influence of the 18 epidemiological factors on the seroprevalence. The model included all main effects and all two-factor interactions. Sensitivity indices and total sensitivities were calculated for each component  $i \in \{1, 2, 3\}$ . The sensitivity index  $SI_w^i$  associated with each factorial term w (main effect or interaction) is defined as the ratio between the sum of squares associated with that term and the total sum of squares. The total sensitivity  $TS_f^i$  is given for each factor f by the sum of the sensitivity indices corresponding to that factor; it gives the contribution of the factor and its two-factor interactions (Saltelli et al., 2000).

These indices describe the impact of the epidemiological parameters on the seroprevalence output. To assess the impact of the model stochasticity, which represents the variability among individuals in the batch, an extra sensitivity

- index  $RS^i$  was calculated for each component i as the ratio of the sum of squares within scenario summed over all scenarios to the total sum of squares.
- $^{244}$  This analysis was performed with the R software.  $^{3}$
- 245 3.3 Sensitivity analysis on the mean seroprevalence at slaughterhouse deliv-246 ery
- An ANOVA was also performed to compare the influence of the 18 epidemiological parameters on the mean seroprevalence of delivered pigs  $\Omega$ . It was calculated for each scenario defined by the fractional factorial design described above.

#### 251 4 Results

- We tested several sets of simulations with 10 replications per scenario. The sensitivity analysis was conducted on each set. Results were stable in terms of sensitivity index values and parameter order. Therefore, the 15 replications per scenario performed for this study ensure that the results are robust.
- 256 4.1 Influence of factors on the seroprevalence over time
- At each time step from birth to slaughterhouse delivery and for each scenario, the average seroprevalence was calculated over the 15 replications performed for the sensitivity analysis. The percentiles of the average seropreva-

<sup>&</sup>lt;sup>3</sup> The R Project for Statistical Computing: http://www.r-project.org/

lences across the 4096 scenarios are represented at each time step in Fig. 1. They show a global increase in seroprevalence over time, but they also ex-261 hibit a high variability. In 10% of the scenarios, no infection is found. In 10% of the scenarios, the batch is almost entirely infected at the slaughterhouse delivery. Box-plots at each time step and for each percentile are also drawn in Fig. 1. They show the seroprevalence variability between replicates due 265 to model stochasticity. Each box-plot is based on the 3×15 seroprevalences associated with the 3 scenarios closest to a given percentile at a given time step. As stated previously, no infection is found for the 10th percentile. For 268 the other percentiles, the variability due to model stochasticity appears to be quite lower than the variability due to the epidemiological factors, despite a few large outliers. 271

The inertia obtained for the first three principal components are 78.0%, 15.9% and 4.2%; the next component represented less than 1% of the total variability (Fig. 2). Results are described component by component, starting with the first component.

The sensitivity indices  $RS^i$  describing the impact of the model stochasticity for the first three components i are quite low:

$$RS^1 = 1.8\%, \quad RS^2 = 3.6\%, \quad RS^3 = 7.9\%.$$

The first principal component represents 78.0% of the total inertia. According to Fig. 3-A all time steps have comparable weights for this component, though somehow lower for the first time steps. So it corresponds more or less to an average seroprevalence effect. The most sensitive factors, which contributed to more than 20% of the variability for this component, are:  $q_1$ , the saturation

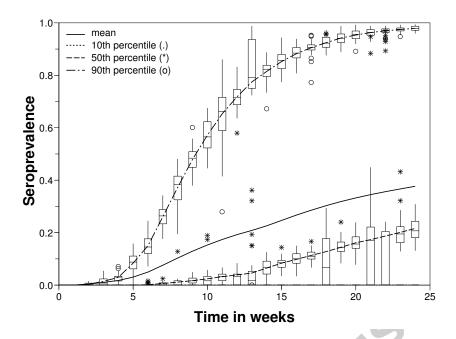


Fig. 1. Percentiles of the average seroprevalence at each time step (lines) and seroprevalence variability between replicates (boxplots). At each time step, percentiles
are calculated across the 4096 seroprevalences averaged over the 15 replications per
scenario: mean (plain line), 10<sup>th</sup> percentile (dotted line, superposed with the abscissa axis), 50<sup>th</sup> percentile (dashed line), and 90<sup>th</sup> percentile (dash-dotted line).
Each boxplot is calculated over all replications of the 3 scenarios that are most
closely associated with a given percentile. In the boxplots: outliers for the 10<sup>th</sup> percentile (.,none), 50<sup>th</sup> percentile (\*), and 90<sup>th</sup> percentile (o). Parameters are given in
Tables 1 and 2.

infection probability threshold,  $b_2$ , the saturation infection probability. The residual contamination in the farrowing room,  $Q_S$  and in the post-weaning room  $Q_{PW}$  explain more than 10% of the inertia;  $q_e$ ,  $b_1$  and  $\varepsilon$ , more than 5% each (Fig. 3-B). The cumulated sensitivity indices of the eight major main effects and two-factor interactions are higher than 70% (Fig. 3-C). For all epidemiological parameters, the contributions of their main effect and their

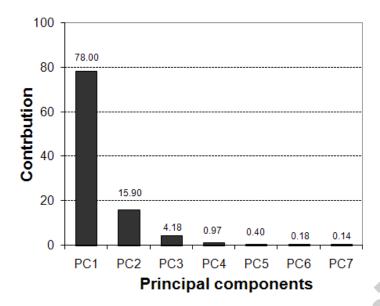


Fig. 2. Contribution to the total inertia of the seven first principal components: component contribution (black) and total inertia explained by the seven first components (grey).  $PC_X$  is the  $X^{\text{th}}$  component.

two-factor interactions are comparable. However, only one interaction appears in the eight most sensitive terms:  $b_2$ : $q_1$ .

The second component only corresponds to 15.9% of the total inertia. On 289 this component, the higher scores are attributed to simulations which exhibit 290 higher contrasts between the beginning and the end of the simulation (Fig. 4-A). The parameters explaining more variability differ from the first component 292 (Fig. 4-B). Two parameters contribute to ca. 40% of the inertia:  $\varepsilon$ , the maternal 293 protective factor, and  $Q_S$ , the residual contamination of the farrowing room. 294 Both are involved in the early infection of piglets.  $b_2$  and  $q_1$ , which govern the infection function, explain more than 10% of the inertia each. For all parameters, the contribution of the two-factor interactions is higher than the main effect contribution. The cumulated sensitivity indices of the eight major main factors and interactions is equal to 53 %; four interactions involving  $Q_S$ ,

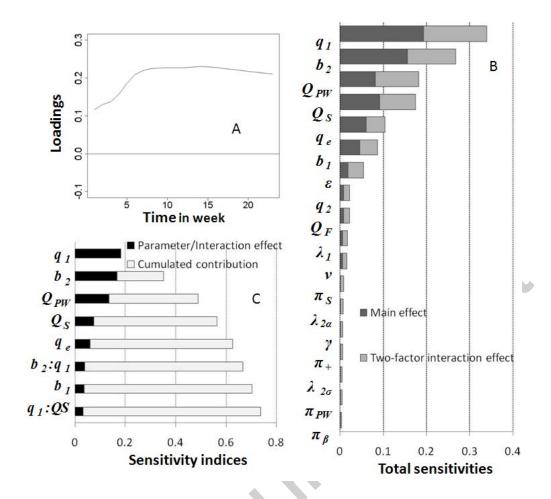


Fig. 3. Sensitivity analysis on the seroprevalence over time: results of the ANOVA performed on the simulation scores of the first component (inertia: 78.0%). – **A** Loadings defining the component for each time variable (in abscissa). **B** Total sensitivities  $TS_f^1$  for the 18 epidemiological parameters f ranked in descending order; sensitivities are split in main effect (black) and two-factor interactions (grey). **C** Sensitivity indices  $SI_w^1$  of the eight main factorial terms w (main effects or interactions) in descending order; individual (black) and cumulated (grey) contributions.

The third component contribution is very low: 4.2%. It highlights the trajectories in which the seroprevalence during the intermediate time steps was very different from the beginning and end of the simulation (Fig. 5-A). As for

 $\varepsilon$ ,  $q_1$  or  $b_2$  are included (Fig. 4-C).

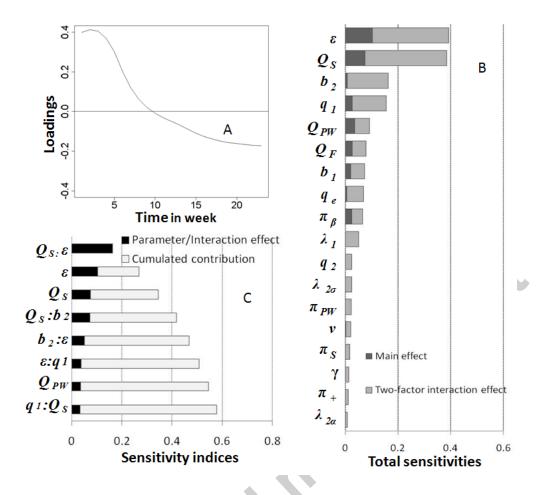


Fig. 4. Sensitivity analysis on the seroprevalence over time: results of the ANOVA performed on the simulation scores of the second component (inertia: 15.9%). – **A** Loadings defining the component for each time variable (in abscissa). **B** Total sensitivities  $TS_f^2$  for the 18 epidemiological parameters f ranked in descending order; sensitivities are split in main effect (black) and two-factor interactions (grey). **C** Sensitivity indices  $SI_w^2$  of the eight main factorial terms w (main effects or interactions) in descending order; individual (black) and cumulated (grey) contributions.

the second component, interactions contribute more to the inertia than major effects; they are comparable for the most influential parameter  $Q_F$ . The  $Q_F$ total sensitivity is higher than 20% (Fig. 5-B). Parameters  $Q_F$ ,  $Q_{PW}$ ,  $q_1$  and various two-factor interactions involving these three parameters as well as  $b_2$ ,

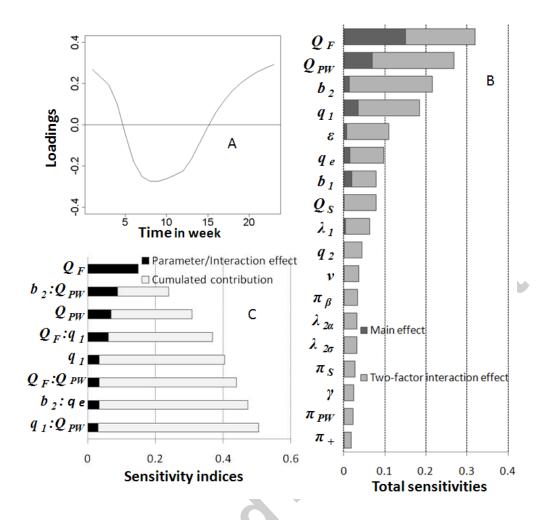


Fig. 5. Sensitivity analysis on the seroprevalence over time: results of the ANOVA performed on the simulation scores of the third component (inertia: 4.2%). – **A** Loadings defining the component for each time variable (in abscissa). **B** Total sensitivities  $TS_f^3$  for the 18 epidemiological parameters f ranked in descending order; sensitivities are split in main effect (black) and two-factor interactions (grey). **C** Sensitivity indices  $SI_w^3$  of the eight main factorial terms w (main effects or interactions) in descending order; individual (black) and cumulated (grey) contributions.

 $q_1$  and  $q_e$  explained more than 47% of the variability (Fig. 5-C).

# 309 4.2 Influence of factors on the seroprevalence at slaughterhouse delivery

The terms which had a higher than 5% effect on the seroprevalence at slaughterhouse delivery are, in descending order:  $q_1$  (16.1%),  $b_2$  (11.2%),  $b_1$  (5.3%), three of the four parameters governing the infection probability function, and the residual room contaminations  $Q_S$  (5.3%) and  $Q_{PW}$  (5.2%). The statistical model explains only 62% of the variability of the output.

#### 315 5 Discussion

The model developed in this study is a simplified version of a herd model, obtained by considering an independent pig batch that only interacts with the herd through residual room contaminations. It cannot take into account batch mixing. It corresponds to an all-in/all-out management system, which is the recommended and a frequent system under normal pig growth conditions, i.e. when pig growth duration is not affected by any disease or other external factor. This batch model has allowed us to assess the *Salmonella* seroprevalence in a batch of pigs, from birth to slaughterhouse delivery and in the groups delivered to the slaughterhouse.

Sensitivity analyses were conducted to identify the parameters which had the greatest influence on the variability of these model outputs. The first analysis, performed on the seroprevalence over time, was more informative than the analysis performed on the seroprevalence at slaughter age. Other sensitivity analysis methods were used in previously published model describing the Salmonella spread within pig herds. Some graphically assessed the effect of infection probability variations on the seroprevalence of pigs at slaughter age

(Hill et al., 2008). Others, in a meta-modelling approach, used a generalised linear mixed model to estimate the influence of epidemiological parameters on scalar outputs: the percentage of contaminated carcasses at the end of the slaughter process and their serological status (van der Gaag et al., 2004). In this study, we analysed the seroprevalence trajectories, estimating that the time variable would provide valuable information.

Sensitivity analyses rank sources of uncertainty and variability, according to
their influence on the model outputs. These sources need not be restricted to
model parameters. Here, they also included the variability among individuals
in a batch, because it induces notable variations on the batch seroprevalence.
Results eventually showed that this individual variability had a negligible impact on the Salmonella seroprevalence variability compared to the epidemiological parameters. The range of values tested for these parameters was chosen
especially broad to account for the uncertainty on these parameters and to emphasise their impact on the output variability. However, this result was not
obvious. It highlights in the present case an urgent need for more accurate
estimates of the epidemiological parameter values.

Several key epidemiological parameters were identified and can be grouped in three categories. The first category consists of  $q_1$ ,  $b_2$  and  $b_1$ , which govern the infection probability function. These parameters are the most influential ones, as together they explained more than 40% of the first component inertia (two-factor interaction not included), the first component representing more than 78% of the total inertia. In the second category, parameters related to the room contamination,  $Q_{PW}$ ,  $Q_S$ , and  $Q_F$  are found. Their contribution to the seroprevalence variability appeared in the top eight of the first three components. The third and last category contains the Salmonella shedding

rate for a seronegative finishing pig  $q_e$  and the maternal protective factor arepsilon.

Their contribution was lower, but their impact cannot be neglected.

**Infection probability function** The infection probability function seems to have a major influence on the Salmonella infection dynamics. However, its four parameters do not have the same impact on the seroprevalence variability. Actually,  $q_1$  and  $b_2$ , corresponding respectively to the lower threshold and the saturation infection probability, are the most sensitive parameters. This is consistent with a rapid growth of the quantity of Salmonella infectious units in a room. As the batch size does not vary much with time, this quantity Q(t)increases with the effective quantity ingested by a pig q(t). In low contaminated rooms  $(q(t) \leq q_1)$ , the transmission remains low. Once the room contamination exceeds a given threshold  $(q(t) > q_1)$ , the transmission increases rapidly and the infection probability reaches its saturation value ( $f(q(t) = b_2)$ ). Further investigation would be needed to define this function more accurately. The 371 major influence of the infection probability had already been shown by van der 372 Gaag et al. (2004): in this study, the rate of infection within a group of pigs was the most influential factor on the percentage of contaminated carcasses 374 at the end of the slaughter process. Control measures which would reduce the probability for a susceptible pig to become infected would be very efficient. 376 They would involve: (i) a reduction of the contact rate between susceptible pigs and all sources of infection, by implementing a strict all-in/all-out housing system for instance; and/or (ii) a reduction of the probability of infection given such a contact, by acid feed (Farzan et al., 2006) or vaccination (Roesler et al., 2006).

Residual room contaminations The residual room contaminations were, as expected, quite important for the prevalence, especially as they were used to 383 introduce the infection in the batch. Globally, the farrowing and post-weaning room contaminations had more impact on the seroprevalence over time. The batch was usually already infected when it entered the finishing room, so the contribution of shedding pigs to the room contamination overpowered the residual contamination. The second component brought out particularly strongly the residual farrowing room contamination. As this component emphasised trajectories exhibiting strong differences in seroprevalence between the beginning and the end of the simulation, it is quite consistent. The influence of the room contamination has been previously shown (Beloeil et al., 2003) and it highlights the need of cleaning and disinfecting the room between batches. This result is consistent with experimental results that showed that pigs removed at weaning or at 10 weeks of age from their herd and placed in Salmonella-free facilities became Salmonella-free delivered pigs (Dahl et al., 1997). Our study hence suggests that control measures should also concentrate on the residual room contamination, especially for the farrowing and the post-398 weaning rooms. This goal could be reached by: (i) a better cleaning-disinfection 399 process, (ii) reducing the quantity of Salmonella shed by infected pigs, and (iii) reducing the survival of these bacteria in the environment (quicker removal of contaminated feces).

Shedding Control measures reducing the shedding duration were shown to
be efficient by experimentation (e.g. with vaccination (Roesler et al., 2006),
or with acid feed (Farzan et al., 2006; van Winsen et al., 2001)). However, in
the range of variation tested here, the shedding duration did not have a great

impact on the seroprevalence variability. Therefore, further experiments should rather concentrate on the shedding quantity than on the shedding duration.

Maternal protection A partial maternal protection was represented by 409 lowering the probability of infection for piglets during the first four weeks 410 of life. The maternal protective factor had an intermediate influence in the 411 first component and a major effect in the second component. We used a large 412 range of variation for this factor, between 0 and 1, because not only is its value uncertain, its existence remains to be demonstrated. The transmission from sows to piglets seems to be a critical point for the initial contamination of pig 415 batches (Lurette et al., 2008a). However, little evidence was found in the lit-416 erature to support the existence of a maternal protective factor due to passive immunity. Indeed, even if passive immunity has been found in piglets (Proux et al., 2000), no experiments have yet shown a potential protection induced 419 by this immunity. Other authors modelling Salmonella spread within grower-420 finisher herds considered active immunity and represented either a complete immunity for a short period (10 days) following the infection (Hill et al., 2008), 422 or a partial immunity for the seropositive animals (van der Gaag et al., 2004). 423 We also attributed a partial immunity to the seropositive health states by reducing the shedding rate of seropositive animals  $(\pi_+)$ . Additional data concerning the potential protective effect of maternal antibodies or post-infection antibodies are however needed to investigate these matters further. 427

Hill et al. (2008) and van der Gaag et al. (2004) highlighted the importance of the infection prevalence in weaner pigs and then the importance of the transmission between sows and their piglets on the slaughterhouse delivery seroprevalence. Our study has shown that early infection, occurring between

birth and weaning, seems to be the critical point of the Salmonella spread within a pig batch, and possibly within a herd. Literature reported that the prevalence of Salmonella at slaughter age resulted from the late infection of pigs in cases of multi-site herds (Davies et al., 1998; Nielsen et al., 1995). Actually, the serotype identified at the slaughterhouse differed from those found in the breeding and in the multiplying herds. In this study, the birth and the growth of pigs take place at the same place. Moreover, other observed data exhibit a correlation between the sow herd Salmonella status and the finishing pigs status (Nollet et al., 2005), which tends to support our findings.

#### 441 6 Conclusion

This first work has allowed us to identify key parameters in Salmonella transmission from a model describing the Salmonella spread within a pig batch: first came parameters involved in the infection probability function, then parameters related to the room contamination, and finally the Salmonella shedding 445 rate and maternal protective factor. We deduced from these findings that efficient control measures should focus on reducing: (i) the probability of infection, like for instance acid feed, and vaccination; and (ii) the residual room contamination, especially for the farrowing room, with a better cleaning-disinfection process, or a quicker removal of contaminated feces. This study also high-450 lighted directions that should be further investigated to better understand the pathogen spread and control it better, such as the potential protective effect of 452 maternal antibodies or post-infection antibodies. Given that the residual contamination of the rearing rooms and the early infection of piglets were shown to have quite an impact on the pathogen spread, it would be interesting to

- 456 study the influence of these parameters within a farrow-to-finish herd where
- sows and pigs are housed together and where batch mixing can occur between
- batches (Hébert, 2006).

#### 459 References

- Beloeil, P. A., Chauvin, C., Proux, K., Rose, N., Queguiner, S., Eveno, E.,
- Houdayer, C., Rose, V., Fravalo, P., Madec, F., 2003. Longitudinal sero-
- logical responses to Salmonella enterica of growing pigs in a subclinically
- infected herd. Prev. Vet. Med. 60 (3), 207–226.
- Boes, J., Alban, L., Bager, J., Mogelmose, V., Baggesen, D. L., Olsen, J. E.,
- 465 2005. Survival of Escherichia coli and Salmonella Typhimurium in slurry
- applied to clay soil on a Danish swine farm. Prev. Vet. Med. 69 (3-4), 213-
- <sub>467</sub> 228.
- Box, G. E. P., Hunter, W. G., Hunter, J. S., 1978. Statistics for experimenters:
- an introduction to design, data analysis, and model building. Wiley series
- in probability and mathematical analysis. John Wiley & Sons, New York.
- Campbell, K., McKay, M. D., Williams, B. J., 2006. Sensitivity analysis when
- model output are functions. Reliab. Eng. Syst. Saf. 91 (10–11), 1468–1472.
- Dahl, J., Wingstrand, A., Nielsen, B., Baggesen, D. L., 1997. Elimination of
- Salmonella Typhimurium infection by the strategic movement of pigs. Vet.
- Rec. 140 (26), 679–681.
- Davies, P. R., Bovee, F. G., Funk, J. A., Morrow, W. E., Jones, F. T., Deen,
- J. R., 1998. Isolation of Salmonella serotypes from feces of pigs raised in a
- multiple-site production system. J. Am. Vet. Med. Assoc. 212 (12), 1925–
- 479 1929.
- Diekmann, O., Heesterbeek, J. A. P., 2000. Mathematical epidemiology of

- infectious diseases: model building, analysis and interpretation. Wiley Series
- in Mathematical and Computational Biology. Wiley.
- Fang, K.-T., Li, R., Sudjianto, A., 2006. Design and modeling for computer
- experiments. Computer Science and Data Analysis Series. Chapman &
- 485 Hall/CRC, Boca Raton, FL.
- Farzan, A., Friendship, R., Dewey, C. E., Warriner, K., Poppe, C., Klotins,
- 487 K., 2006. Prevalence of Salmonella spp. on Canadian pig farms using liquid
- or dry-feeding. Prev. Vet. Med. 73 (4), 241–254.
- Fravalo, P., Cariolet, R., Proux, K., Salvat, G., 2003. Le portage asymptoma-
- tique de Salmonella enterica par les porcs : résultats issus de la constitu-
- tion d'un modèle en conditions expérimentales. In: 35èmes Journées de la
- Recherche Porcine. ITP, INRA, Paris, France, pp. 393–400.
- 493 Hébert, H., 27 octobre 2006. Modalités de conduite en bandes en élevage
- porcin: effets sur les contacts entre animaux. Thèse vétérinaire, École Na-
- tionale Vétérinaire de Nantes, France.
- 496 Hill, A. A., Snary, E. L., Arnold, M. E., Alban, L., Cook, A. J. C., 2008.
- bynamics of Salmonella transmission on a British pig grower-finisher farm:
- a stochastic model. Epidemiol. Infect. 136 (3), 320–333.
- 499 ITP, 2000. Le mémento de l'éleveur de porcs. Institut Technique du Porc,
- Paris, France.
- 501 ITP, 2006. Le porc par les chiffres 2006. Institut Technique du Porc, Paris,
- 502 France.
- 503 Kobilinsky, A., 1997. Les plans factoriels. In: Droesbeke, J.-J., Fine, J.,
- Saporta, G. (Eds.), Plans d'expériences: applications à l'entreprise. Éditions
- <sup>505</sup> Technip, Paris, Ch. 3, pp. 69–209.
- 506 Kranker, S., Alban, L., Boes, J., Dahl, J., 2003. Longitudinal study of
- 507 Salmonella enterica serotype typhimurium infection in three danish farrow-

- to-finish swineherds. J. Clin. Microbiol. 41 (6), 2282–2288.
- Lamboni, M., Makowski, D., Monod, H., Apr. 2008. Multivariate global
- sensitivity analysis for discrete-time models, technical report available:
- http://www.inra.fr/miaj/public/nosdoc/rap2008-3.pdf.
- Lurette, A., Belloc, C., Touzeau, S., Hoch, T., Ezanno, P., Seegers, H., Fouri-
- chon, C., Apr. 2008a. Modelling salmonella spread within a farrow-to-finish
- pig herd, veterinary Research.
- Lurette, A., Belloc, C., Touzeau, S., Hoch, T., Seegers, H., Fourichon, C.,
- 2008b. Modelling batch farrowing management within a farrow-to-finish pig
- herd: influence of management on contact structure and pig delivery to the
- slaughterhouse. Animal 2 (1), 105–116.
- Murray, C., 2002. Mathematical epidemiology of infectious diseases: model
- building, analysis and interpretation. Wiley Series in Mathematical and
- 521 Computational Biology. Wiley.
- Myers, R. H., Montgomery, D. C., 2002. Response surface metholology: pro-
- cess and product optimization using designed experiments, 2nd Edition.
- Wiley Series in Probability and Statistics. Wiley-Interscience, New York.
- Nicholson, F., Groves, S. J., Chambers, B. J., 2005. Pathogen survival during
- livestock manure storage and following land application. Bioresour. Technol.
- 96 (2), 135–143.
- Nielsen, B., Baggesen, D., Bager, F., Haugegaard, J., Lind, P., 1995, The
- serological response to Salmonella serovars typhimurium and infantis in
- experimentally infected pigs. The time course followed with an indirect anti-
- LPS ELISA and bacteriological examinations. Vet. Microbiol. 47 (3-4), 205–
- 532 218.
- Nollet, N., Houf, K., DeWulf, J., Duchateau, L., Zutter, L. D., Kruif, A. D.,
- Maes, D., 2005. Distribution of Salmonella strains in farrow-to-finish pig

- berds: A longitudinal study. J. Food Prot. 68 (10), 2012–2021.
- 536 Plachá, I., Venglovský, J., Sasáková, N., Svoboda, I. F., 2001. The effect of
- summer and winter seasons on the survival of Salmonella typhimurium and
- indicator micro-organisms during the storage of solid fraction of pig slurry.
- J. Appl. Microbiol. 91 (6), 1036–1043.
- Proux, K., Houdayer, C., Humbert, F., Cariolet, R., Rose, V., Eveno, E.,
- Madec, F., 2000. Development of a complete ELISA using Salmonella
- lipopolysaccharides of various serogroups allowing to detect all infected pigs.
- Vet. Res. 31 (5), 481–490.
- Roesler, U., Heller, P., Waldmann, K., Truyen, U., Hensel, A., 2006. Immuniza-
- tion of sows in an integrated pig-breeding herd using homologous inactivated
- Salmonella vaccine decreases the prevalence of Salmonella typhimurium in-
- fection in the offspring. J. Vet. Med. B. 53 (5), 224–228.
- Saltelli, A., Chan, K., Scott, E. M. (Eds.), 2000. Sensitivity Analysis. Wiley
- series in Probability and Statistics. John Wiley & Sons.
- van der Gaag, M. A., Vos, F., Saatkamp, H. W., van Boven, M., van Beek, P.,
- Huirne, R. B. M., 2004. A state-transition simulation model for the spread
- of Salmonella in the pork supply chain. Eur. J. Oper. Res. 156 (3), 782–798.
- van Winsen, R. L., van Nes, A., Keuzenkamp, D., Urlings, H. A. P., Lip-
- man, L. J. A., Biesterveld, S., Snijders, J. M. A., Verheijden, J. H. M.,
- van Knapen, F., 2001. Monitoring of transmission of Salmonella enterica
- serovars in pigs using bacteriological and serological detection methods. Vet.
- <sup>557</sup> Microbiol. 80 (3), 267–274.
- 558 Wong, D. M. A. L. F., Hald, T., van der Wolf, P., Swanenburg, M., 2002.
- Epidemiology and control measures for Salmonella in pigs and pork. Liv.
- 560 Prod. Sci. 76 (3), 215–222.