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## Replication of Restless Legs Syndrome Loci in Three European Populations

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#### Abstract

Background: Restless legs syndrome (RLS) is associated to common variants in three intronic and intergenic regions in MEIS1, BTBD9, and MAP2K5/LBXCOR1 on chromosomes $2 \mathrm{p}, 6 \mathrm{p}$ and 15 q .

Methods: Our study investigated these variants in 649 RLS patients and 1230 controls from Czech Republic (290 cases and 450 controls), Austria (269 cases and 611 controls) and Finland ( 90 cases and 169 controls). Ten SNPs within the three genomic regions were selected according to the results of previous genome-wide scans. Samples were genotyped using Sequenom platforms.

Results: We replicated associations for all loci in the combined samples set (rs2300478 in MEIS1, $\mathrm{P}=1.26 \times 10^{-5}$, odds ratio $(\mathrm{OR})=1.47$, rs3923809 in BTBD9, $\mathrm{P}=4.11 \times 10^{-5}, \mathrm{OR}=1.58$ and rs 6494696 in $M A P 2 K 5 / L B X C O R 1, \mathrm{P}=0.04764, \mathrm{OR}=$ 1.27). Analyzing only familial cases against all controls, all three loci were significantly associated. Using sporadic cases only, we would confirm the association only to BTBD9.

Conclusion: Our study shows that variants in these three loci confer consistent disease risks in patients of European descent. Among the known loci, BTBD9 seems to be the most consistent in its effect on RLS across populations and is also most independent of familial clustering.


## Introduction

Restless legs syndrome (RLS) is characterized by an urge to move the legs associated with unpleasant sensations in the lower limbs typically occurring at rest in the evening or at night.[1] Since the maximum of symptoms appears at bedtime, RLS can lead to disturbances of sleep leading to a decreased quality of life.[1] The diagnosis is further supported by the presence of periodic limb movements in sleep (PLMs) and positive response to dopaminergic treatement.[1]

A recent genome-wide association study (GWA) with German and Canadian RLS cases identified intronic or intergenic variants within three genomic regions: MEIS1 (myeloid ecotropic viral integration site homeobox 1) on chromosome 2 p, BTBD9 (BTB/POZ domain-containing protein 9) on chromosome 6 p , and a third region on chromosome 15q containing MAP2K5 (mitogen-activated protein kinase kinase 5) and LBXCOR1 (ladybird homeobox co-repressor 1).[2] A similar study conducted in Icelandic and US cases showed an association of $B T B D 9$ to PLMs.[3] MEISI belongs to the family of TALE homeobox genes involved in the limb development, the determination of the megacaryocytes and CNS structures such as the retina, cerebella granule cells, hindbrain and spinal motor neuron pool.[4,5,6] So far, very little is known about BTBD9. It consists of a BTB/POZ domain, a back domain and a coagulation factor domain. Known functions of similar proteins containing these domains include ubiquitin dependent protein degradation.[7] The variants located in the third genetic region are in strong linkage disequilibrium with two surrounding genes: MAP2K5 is critical at early stages of muscle cell differentiation [8] and LBXCOR1 which is a transcriptional corepressor of LBX and is highly expressed in spinal dorsal horn and midbrain-hindbrain border.[9] The involvement of these genes in the etiopathogenesis of RLS is still unknown.

The aim of our study was to investigate whether these variants are also relevant among other European (Czech, Austrian, and Finnish) was discussed that RLS has component of a developmental disorder. RLS cases and what is the difference of their impact between sporadic and familial cases.

## Patients and Methods

## Patients and Controls

The diagnosis of all RLS cases was made according to diagnostic criteria of the International RLS Study Group [1] by personal examination by a neurologist in the respective study center. The positive family history was defined as at least one firstdegree family member being affected by RLS (reported by the proband) in all three populations. The control samples originate from the general population and were not screened for presence of RLS.

Czech subjects - The patients were recruited in the center for Disorders of Sleep and Wakefulness, Department of Neurology of First Faculty of Medicine and the General Teaching Hospital, Prague. In total, 290 patients were included (107 males, mean age $55.7 \pm 15.3$ years $( \pm \mathrm{SD})$, mean age at onset of RLS $38.3 \pm 18.1$ years $)$. Positive family history was reported by 110 patients, in 175 cases it was negative and in 5 the data were not available. Altogether 450 sex matched controls were selected randomly from the Czech blood and bone marrow donors registry (166 males, mean age 45.3 $\pm 9.9$ ). Since maximum age for the controls was 63 years, 38 male and 51 female cases in the age group from 64 to 91 years could not be age matched.

Austrian subjects - 269 (104 males) patients were recruited in 2 centers: at the Department of Neurology, Medical University of Vienna and Department of Neurology, University Clinic Innsbruck, (mean age $59.0 \pm 14.3$, mean age at onset of RLS $37.14 \pm 19.5$ ). Positive family history was reported by 107 patients, in 108 cases it was negative and in 54 the data were not available. The patients were matched by sex to 611 controls from the German KORA project, which procedures were described elsewhere [10] (236 males, mean age $59.9 \pm 11.35$ ). KORA controls were
already used in previous GWA study, which showed only negligible effect of population stratification.[2]

Finnish subjects -90 (24 males) patients were recruited in Sleep Research Center in Turku (mean age $46.5 \pm 18.1$, mean age at onset of RLS $19.4 \pm 13.4$ ). Positive family history was reported by 81 patients and 9 patients had a negative family history. A random sample from the general Finnish population, comprising 169 sex matched individuals ( 45 males) was used as control. Data on age of controls were not available. Studies were performed according to the declaration of Helsinki and approved by the Ethical Committee of the respective study centers. Written informed consent was obtained from all RLS patients.

## Genotyping

Ten SNPs within the three genomic regions were selected according to the results of previous GWA scans.[2,3] Samples were genotyped on two Sequenom platforms in Munich and Helsinki (Sequenom MassArray system, Sequenom Inc, San Diego, CA, USA) with a genotype discordance rate of $1.3 \%$ in 158 comparisons, when analyzing repeatedly genotyped internal control samples. Automated genotype calling was done with SpectroTYPER 3.4 software and genotype clustering was visually checked by an experienced scientist. Assays were designed using AssayDesign 3.1.2.2 with iPLEX Gold chemistry default parameters. SNP quality control criteria leading to exclusion from analysis were a call rate $<90 \%, \mathrm{MAF}<1 \%$ and $\mathrm{P}<0.001$ for deviations from HWE in controls.

## Statistical analysis

Genotype data were analyzed using standard association tests (allelic, genotypic, dominant and recessive models) including Cochran-Armitage test for trend, Cochran-Mantel-Haenszel test for estimation of ORs in the stratified sample (including Breslow-Day test for homogeneity) and haplotype tests, as implemented in the PLINK statistical package v1.0.[11] The sample was stratified only according to the country of origin. Logistic regression implementing the Cochran-Armitage test for trend (using genotypes as ordinal values rather than categorical) in the combined sample using age, sex and country of origin as covariates was performed by generalized linear modeling routines incorporated in R package v.2.6.0 (http://www.r-project.org/). Bonferroni correction for multiple testing of ten markers was employed. All P-values given are one-sided, with the direction of the alternative hypothesis given by the original report.[3] Power calculations were performed using the Genetic Power Calculator (pngu.mgh.harvard.edu/~purcell/gpc/).[12] For input parameter we used a RLS prevalence of $8 \%$, an alpha-level of $0.05 \%$ and ORs and allele frequencies according to results from the GWA experiment.[2] Association tests were conducted in four different settings: 1. All patients (i.e. familial and sporadic) combined versus all controls, 2. familial cases versus controls, and 3 . sporadic cases versus controls.

## Results

All SNPs tested were in HWE ( $\mathrm{p}>0.01$ ) in both patients and controls. Under assumption of genetic homogeneity, the combined sample had good power to detect association using previously published parameters [2] (98\% for MEIS1 and BTBD9, $89 \%$ for MAP2K5/LBXCOR1). In the Czech sample alone the power was $82.5 \%$ for MEIS1 and BTBD9, $71.8 \%$ for MAP2K5/LBXCOR1, In the Austrian sample, the powers were $84.8 \%$ and $74.8 \%$, respectively, and in Finish sample separately 38.7 and $30.4 \%$.

Allele frequencies in the Czech and KORA control samples were not significantly different (lowest P in $\chi^{2}$ test $=0.2045$ for rs 4236060 ). Significant allele frequency differences were observed between the Finnish and the combined Czech and KORA control samples within BTBD9 ( $\mathrm{P}<7.67 \times 10^{-6}$ for all SNP markers within BTBD9). A similar, nominally significant, difference in allele frequencies in BTBD9 markers was also observed between Finnish cases and combined Czech and Austrian cases (in $\chi^{2}$ test lowest $\mathrm{P}=0.01063$ for rs9296249), but we did not observe a significant difference between allele frequencies of Czech and Austrian RLS patients (lowest P in $\chi^{2}$ test was 0.4608 for rs2300478). Logistic regression showed no significant interaction with country for any SNP tested, and Breslow-Day test showed homogeneous ORs in all samples.

Significant association after correction for multiple testing at significance level alpha $=5 \%$ was found in at least one SNP for all tested loci in the combined samples (Table 1), and in the Czech and Austrian samples separately. Analyzing the Finnish sample, we confirmed only the association to BTBD9. The association to rs2300478 in MEIS1 was only nominally significant and MAP2K5/LBXCOR1 showed no association (Table 2).

In the combined sample we observed a strong association with the haplotype formed by markers rs6710341 and rs12469063, both located within MEIS1. Carriers of the "AG" haplotype had ORs for developing RLS of $1.98\left(\mathrm{P}=9.1 \times 10^{-10}\right)$. Results for this haplotype were similar when testing the Czech $\left(\mathrm{P}=3.210^{-7}, \mathrm{OR}=2.38\right)$, Austrian $(\mathrm{P}=$ $8.3 \times 10^{-5}, \mathrm{OR}=1.82$ ), and Finnish samples $\left(\mathrm{P}=2.0 \times 10^{-4}, \mathrm{OR}=2.46\right)$ separately. No other common polymorphic phased haplotypes (MHF > 1\%) yielded significant results. An allele dosage model best described the association for MEIS1 and BTBD9 (Armitage trend test). In contrast, a recessive model for the risk allele fitted best for the MAP2K5/LBXCORI locus.

Analyzing only familial cases $(\mathrm{n}=217)$ and all controls, all three loci were significantly associated. Using sporadic cases only ( $\mathrm{n}=283$ ), we would confirm the association to BTBD9 but not to MEIS1 and MAP2K5/LBXCOR1. We omitted patients of Finnish origin from this sub-analysis due to very low proportion of sporadic cases and different allele frequencies in these samples. The Breslow-Day test did not show significant heterogeneity between sporadic and familial cases.

## Discussion

Our study showed an association of variants in MEIS1, BTBD9 and MAP2K5/LBXCOR1 with RLS in a combined sample of Czech, Austrian, and Finnish RLS cases. Similar findings were observed in the US population.[13] In accordance to the original report, the strongest effect was observed with the haplotype "AG" formed by markers rs6710341 and rs12469063 located in the $9^{\text {th }}$ intron of MEIS1, providing ORs of about 2.0 for this haplotype. However, the OR may be underestimated, because the controls samples were not screened to exclude RLS and therefore may contain approximately $10 \%$ of individuals actually affected by RLS. The best models observed for individual loci are in good agreement with previous findings in German and Canadian populations. The significance of these loci to RLS can therefore be regarded as well established.

The sub-analysis in Czech and Austrian populations show same trends for association as the combined sample, but in the Finnish sample, only association with $B T B D 9$ was confirmed and there was a trend for association to MEIS1. Moreover, the allele frequencies and proportions of familial cases in the Finnish sample were different from the other two, but the smaller size of this sample limits further implications.

In our sample set we have not observed significant differences between familial and sporadic cases concerning the BTBD9 locus. The 95\% confidence intervals of OR also overlapped between familial and sporadic cases for both MEIS1 (1.357-2.1 in familial and 1.019-1.534 in sporadic cases vs. all controls for rs12469063) and MAP2K5/LBXCOR1 (1.164-1.841 in familial and 0.951-1.408 in sporadic cases for rs6494696). There is a trend that MEIS1 and MAP2K5/LBXCOR1 possibly play a more important role in familial RLS, but due to limited number of patients, we were
not able to prove significant heterogeneity. Generally the risk alleles in these loci are common and exert only small to moderate effects. They do not explain the familial clustering of RLS.[2] Besides these association signals, six linkage regions for RLS on chromosomes $2 \mathrm{q}, 9 \mathrm{p}, 12 \mathrm{q}, 14 \mathrm{q}, 19 \mathrm{p}$ and $20 \mathrm{p}[14-18]$, under a recessive or autosomal dominant model of inheritance have been described. These variants must be of larger effects and less frequent, since only some have been successfully confirmed in independent populations or in single families.[19-24] Among the known loci, $B T B D 9$ seems to be the most consistent in its effect on RLS across populations and is also most independent of familial clustering. We conclude that the observed genetic determinants are risk factors for RLS in multiple populations.

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## Table 1. Genotyped SNPs and Results of Association in Combined Samples

The Genetic positions in bp and gene alignments are derived from UCSC Genome browser (http://genome.ucsc.edu, assembly March 2006) [25], $\mathrm{r}^{2}$ - linkage disequilibrium relative to preceding marker, data were computed using genotypes observed in both cases and controls using Haploview 4.0 from HapMap project (http://www.hapmap.org, release 21a) [26], OR Odds ratio for the risk allele (Cochran-Mantel-Haenszel test) with 95\% confidence intervals, P nom - Logistic regression implementing Armitage trend test with country of origin, sex and age as covariates, P corr - adjusted P values for multiple testing, MAF - minor allele frequencies observed in combined Czech and Austrian sample, in sporadic and familial cases, Best model corresponds to model, under which lowest $P$ values were observed (TREND Armitage trend test, REC - recessive model), P corr Fam. - comparison of allele frequencies between familial cases and all controls, P corr Spor - comparison of allele frequencies between sporadic cases and all controls. $\dagger$ risk allele is the major allele

| Chr | Gene | SNP ID | Genome | $\mathbf{r}^{2}$ | $\begin{array}{\|ll} \hline \text { OR } \\ (95 \% ~ C o n f . ~ I n t) ~ \end{array}$ | P nom | P corr | MAF <br> Fam. | MAF <br> Spor. | Best <br> model | P corr <br> Fam. | P corr <br> Spor. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2p | MEIS1 | rs6710341 | 66611926 |  | 0.84 (0.64-1.11) | 0.30646 | 1 | 0.1270 | 0.1288 | TREND | 1 | 1 |
| 2p | MEIS1 | rs12469063 | 66617812 | 0.413 | 1.43 (1.16-1.78) | $4.15 \mathrm{E}-06$ | $4.15 \mathrm{E}-05$ | 0.3522 | 0.2727 | TREND | $2.24 \mathrm{E}-05$ | 0.3245 |
| 2p | MEIS1 | rs2300478 | 66634956 | 0.969 | 1.47 (1.18-1.82) | $1.26 \mathrm{E}-06$ | 1.26E-05 | 0.3575 | 0.2860 | TREND | $3.10 \mathrm{E}-05$ | 0.1520 |
| 6p | BTBD9 | rs9296249 | 38473818 |  | $1.59(1.26-2.01) \dagger$ | 0.00011 | 0.00107 | 0.1694 | 0.1553 | TREND | 0.0544 | 0.0012 |
| 6p | BTBD9 | rs3923809 | 38548947 | 0.512 | 1.58 (1.28-1.96) $\dagger$ | $4.11 \mathrm{E}-06$ | $4.11 \mathrm{E}-05$ | 0.2204 | 0.2330 | TREND | 0.0018 | 0.0022 |
| 6p | BTBD9 | rs4236060 | 38578315 | 0.829 | 1.49 (1.19-1.86) † | $1.93 \mathrm{E}-05$ | 0.00019 | 0.1882 | 0.2110 | TREND | 0.0008 | 0.0049 |
| 15q | MAP2K5 | rs11635424 | 65824631 |  | 1.26 (1.02-1.55) † | 0.00602 | 0.06023 | 0.2446 | 0.2992 | REC | 0.0203 | 1 |
| 15q | MAP2K5 | rs3784709 | 65859328 | 0.935 | $1.24(1.01-1.52) \dagger$ | 0.00530 | 0.05301 | 0.2392 | 0.2917 | REC | 0.0393 | 1 |
| 15q | MAP2K5 | rs1026732 | 65882138 | 0.966 | 1.27 (1.03-1.56) † | 0.00428 | 0.04278 | 0.2339 | 0.2936 | REC | 0.0116 | 1 |
| 15q | $\begin{aligned} & \text { MAP2K5/ } \\ & \text { LBXCOR1 } \end{aligned}$ | rs6494696 | 65890259 | 0.999 | 1.27 (1.03-1.56) † | 0.00476 | 0.04764 | 0.2339 | 0.2936 | REC | 0.0108 | 1 |

## Table 2 Analysis in individual populations

MAF - minor allele frequencies in each subsample in patients and healthy individuals, N numbers of successfully genotyped individuals bypassing quality control criteria, Best P corr - P values corrected for multiple testing according to the full association model in Table 1, OR - Odds ratio and corresponding 95\% confidence intervals.

|  | Czech Republic |  |  |  | Austria |  |  |  | Finland |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SNP ID | MAF <br> Cases $\mathrm{N}=276$ | MAF <br> Controls $\mathrm{N}=412$ | Best <br> P corr | $\begin{aligned} & \text { OR } \\ & (95 \% \text { Conf. Int }) \end{aligned}$ | $\begin{aligned} & \text { MAF } \\ & \text { Cases } \\ & \mathrm{N}=222 \end{aligned}$ | MAF <br> Controls $\mathrm{N}=570$ | Best <br> P corr | $\left\lvert\, \begin{aligned} & \text { OR } \\ & (95 \% \\ & \text { Conf. Int })\end{aligned}\right.$ | $\begin{aligned} & \text { MAF } \\ & \text { Cases } \\ & \mathrm{N}=88 \end{aligned}$ | MAF <br> Controls $\mathrm{N}=246$ | Best <br> P corr | OR $(95 \%$ Conf. Int $)$ |
| rs6710341 | 0.1309 | 0.1456 | 1 | 1.13 (1.55-0.83) | 0.1306 | 0.1412 | 1 | 0.91 (0.66-1.26) | 0.1207 | 0.1585 | 1 | 0.73 (0.43-1.22) |
| rs12469063 | 0.2971 | 0.2172 | 0.0492 | 1.52 (1.19-1.95) | 0.3108 | 0.2426 | 0.0064 | 1.41 (1.11-1.79) | 0.3161 | 0.2439 | 0.6093 | 1.43 (0.98-2.10) |
| rs2300478 | 0.3025 | 0.2209 | 0.0285 | 1.53 (1.20-1.96) | 0.3243 | 0.2487 | 0.0017 | 1.45 (1.14-1.84) | 0.3276 | 0.2459 | 0.3676 | 1.49 (1.02-2.18) |
| rs9296249 | 0.1649 | 0.2306 | 0.0252 | 1.52 (1.15-2.00) | 0.1644 | 0.2378 | 0.0116 | 1.59 (1.19-2.11) | 0.2414 | 0.3516 | 0.1081 | 1.70 (1.15-2.53) |
| rs3923809 | 0.2301 | 0.2998 | 0.0374 | 1.43 (1.12-1.84) | 0.223 | 0.3133 | 0.0049 | 1.59 (1.23-2.05) | 0.2651 | 0.4119 | 0.0124 | 1.94 (1.32-2.87) |
| rs4236060 | 0.2047 | 0.2662 | 0.1903 | 1.41 (1.09-1.83) | 0.1968 | 0.2891 | 0.0028 | 1.66 (1.27-2.17) | 0.2674 | 0.3921 | 0.0497 | 1.77 (1.20-2.60) |
| rs11635424 | 0.2772 | 0.3350 | 0.0135 | 1.31 (1.04-1.66) | 0.2793 | 0.3229 | 0.1014 | 1.23 (0.97-1.57) | 0.3046 | 0.2866 | 1 | 1.09 (0.75-1.59) |
| rs3784709 | 0.2754 | 0.3289 | 0.0124 | 1.29 (1.02-1.63) | 0.2725 | 0.3185 | 0.0522 | 1.25 (0.98-1.59) | 0.3046 | 0.2744 | 1 | 1.16 (0.79-1.69) |
| rs1026732 | 0.2717 | 0.3350 | 0.0050 | 1.35 (1.07-1.71) | 0.2748 | 0.322 | 0.0519 | 1.25 (0.98-1.60) | 0.3046 | 0.2764 | 1 | 1.15 (0.79-1.67) |
| rs6494696 | 0.2717 | 0.3350 | 0.0050 | 1.35 (1.07-1.71) | 0.2748 | 0.3229 | 0.0416 | 1.26 (0.99-1.60) | 0.3046 | 0.2764 | 1 | 1.15 (0.79-1.67) |

