

# An empirical comparison among aftershock decay models

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16	Abstract
17	We compare the ability of three aftershock decay models proposed in the literature to reproduce the
18	behavior of 24 real aftershock sequences of Southern California and Italy. In particular, we consider

der 19 the Modified Omori Model (MOM), the Modified Stretched Exponential model (MSE) and the 20 band Limited Power Law (LPL). We show that, if the background rate is modeled properly, the 21 MSE or the LPL reproduce the aftershock rate decay generally better than the MOM and are 22 preferable, on the basis of the Akaike and Bayesian information criteria, for about one half of the 23 sequences. In particular the LPL, which is usually preferable with respect to the MSE and fits well 24 the data of most sequences, might represent a valid alternative to the MOM in real-time forecasts of 25 aftershock probabilities. We also show that the LPL generally fits the data better than a purely 26 empirical formula equivalent to the aftershock rate equation predicted by the rate- and state-

- 27 dependent friction model. This indicates that the emergence of a negative exponential decay at long
- times is a general property of many aftershock sequences but also that the process of aftershock
- 29 generation is not fully described by current physical models.
- 30

#### 30 Introduction

31

The most commonly used formula to reproduce the decay of aftershock rate after a mainshock, also adopted in procedures for the real-time forecast of aftershock probabilities in California (Gerstenberger et al., 2007), is the Modified Omori Model (MOM, Utsu, 1961)

$$35 \qquad \qquad MOM \quad t \quad \frac{K}{t \quad c^{p}} \tag{1}$$

36 where  $_{MOM} t$  is the intensity (the rate) of a non stationary Poisson process, and p, c and K are free 37 parameters. The MOM is empirical in nature but it was found to be compatible with the rate- and 38 state-dependent friction model proposed by Dieterich (1994).

A characteristic of the MOM is to predict an infinite number of possible future aftershocks (that is 39 40 an infinite number of potential faults) if the power law exponent p is lower than or equal to 1. Since 41 such p values are often observed for real sequences, the MOM might appear physically unrealistic. 42 Few alternative formulations, proposed in the last decades, overcome this limitation of the MOM. 43 We mainly consider here two of them: the Modified Stretched Exponential Model (MSE, 44 Kisslinger, 1993; Gross and Kisslinger, 1994) and the Band Limited Power Law (LPL, Narteau et 45 al., 2002; 2003). Both MSE and LPL assume that a negative exponential decay emerges at long 46 times, hence they predict a finite number of aftershocks (and faults), independently of the value of 47 the power law exponent. The intensity of the MSE can be written as

48 
$$MSE t = 1 r N^* 0 \exp \left(\frac{d}{t_0}\right)^{1/r} \frac{t d}{t_0^{1/r}} \exp \left(\frac{t d}{t_0}\right)^{1/r}$$
[2]

49 where  $N^* 0$  is the total number of potential shocks at the time of the mainshock (*t*=0),  $t_0$  is the 50 relaxation time of the negative exponential decay process, *d* a delay time (corresponding to 51 parameter *c* of the MOM) and  $0 < r \le 1$  the power-law exponent.

52 The intensity of the LPL is given by

53 
$$\sum_{LPL} t \quad B \frac{q_{,b}t \quad q_{,a}t}{t^q}$$
 [3]

54 where q is the power-law exponent, B is a normalizing constant (similar to K of the MOM), 55  $_{a}$   $_{b}$  are two parameters (having the physical dimensions of rates) that controls the behavior at 56 long and short times respectively, and  $\gamma$  indicates the incomplete Gamma function

57 
$$q, x = \int_{0}^{x} z^{q-1} e^{-z} dz$$
 [4]

 $_{a}$  (as it may be assumed usually) the behavior of the LPL can be described as the 58 When 59 superposition of three regimes that control the rate at different times: an initial linear decay, which is followed by a power-law and, at large times, by a negative exponential. Narteau et al (2003) 60 61 suggested to considering two times  $t_1$  $t_{2}$ that correspond to the transition between the linear 62 and the power-law and between the power-law and the exponential decays respectively. They are 63 defined as the times at which the ratio between aftershock rates predicted by LPL and by a pure power-law is  $\chi$ . Narteau et al. (2003) report the values assumed by  $t_1$  and  $t_2$ , for values of 64 the ratio  $\chi$  ranging from 0.8 to 0.99 while Lolli et al. (2009) proposed to use  $t_b$   $t_1$  2  $^q$  and 65  $t_a = t_2 1/e$  (where e is the base of natural logarithms) as they corresponds approximately to c of 66 the MOM (or d of MSE) and  $t_0$  of the MSE respectively. We will adopt such derived parameters  $t_h$ 67 68 and  $t_a$  in the following references to the LPL.

69 Both MSE and LPL are based on reasonable physical assumptions but Lolli and Gasperini (2006) 70 showed that MSE and LPL are preferable with respect to the MOM for about one fourth of the real 71 aftershock sequences of Southern California and Italy only. They hypothesized that the limited 72 duration of the observing time interval they choose (one year) might penalize the MSE and the LPL 73 with respect to the MOM when the exponential decay emerges later than the end of such interval. In 74 this work we will test such hypothesis by considering a longer observing interval of four years. We 75 also consider here the possibility that the background rate (not modeled by Lolli and Gasperini, 76 2006) might play a role in assigning the preference to the MOM in some cases. The background 77 seismicity rate is accounted simply by a constant rate  $\mu$  (to be determined together with the other 78 parameters of the various decay models) added to the rate equations [1] [2] and [3].

79

#### 80 Data sources and sequence detection

81

82 We use essentially the same datasets analyzed by Lolli and Gasperini (2006) but we extend the 83 analysis to a longer time interval of four years after the mainshock and consider the catalogs of 84 southern California and Italy up to July 2008 and May 2008 respectively (instead of December 85 2004). For California we use the revised catalog from 1932 to 2008 available from the Southern 86 California Earthquake Center (SCEC) site (http://www.scecdc.scec.org/). For Italy, we merged 87 several catalogs of Italian instrumental earthquakes covering the time interval from 1960 to May 88 2008. From 1960 to 1980, we used the catalog of the Progetto Finalizzato Geodinamica (Postpischl, 89 1985) with magnitudes corrected according to Lolli and Gasperini (2003); from 1981 to 1996, we 90 used the Catalogo Strumentale dei Terremoti Italiani dal 1981 a 1996 Version 1.1 (CSTI Working Group, 2004); from 1997 to 2002, we used the Catalogo della Sismicità Italiana 1.1 (Castello et al., 91 92 2005); finally, from 2003 to 2008, the data are taken from the instrumental bulletin of the *Istituto* 93 Vulcanologia Nazionale di Geofisica e (INGV) available from site http://www.ingv.it/~roma/reti/rms/bollettino. Following Ouillon and Sornette (2005), we assumed 94 95 the completeness of the southern California catalog for  $M_L>3.0$  in 1932 and later years, for  $M_L>2.5$ 96 in 1975 and later years,  $M_I \ge 2.0$  in 1992 and later years, and  $M_I \ge 1.5$  in 1994 and later years. For 97 Italy we assumed the completeness for  $M_L > 2.5$  for 1984 and later (Lolli and Gasperini, 2003), and 98 for  $M_L > 3.0$  before 1984.

In a first step we use the same sequence detection algorithm adopted by Lolli and Gasperini (2006) that defines the influence zone of any shock as a circular area centered in the epicenter and assumes as mainshocks (triggering the sequences) all earthquakes with magnitude not lower than 5.0 that are not included in the influence zone of a larger shock. The time window is fixed to four years after the mainshock while the radius R of the influence zone is chosen as a function of magnitude as Log10(R)=0.1238M+0.983 (that closely corresponds to Table 1 of Gardner and Knopoff, 1974).

105 Only the shocks shallower than 40 km and with magnitude above completeness threshold are 106 included in sequences. To reduce the possible incompleteness in the first times after the mainshock 107 we only consider the aftershocks with magnitude not lower than mainshock magnitude  $M_m$  minus 108 3.5.

109 As the Gardner and Knopoff (1974) radius is likely to overestimate the size of the mainshock 110 influence zone, in a second step we performed an analysis of correlation between the shock rates 111 observed at different distances from the mainshock during a time interval of 200 days after the 112 mainshock. In particular, for distances r varying from 0 to R, we correlate the sequence of rates 113 observed (over 5 days bins) inside the circle with radius r and inside the circular ring with 114 minimum and maximum radius r and R respectively. For each sequence we assumed as influence 115 distance (reported in Table 1 as  $R_i$ ) the largest r for which the correlation between the sequences of 116 rates is significant at the 0.05 level.

To grant a reliable determination of model parameters we consider for the analysis only the sequences including 100 shocks at least within the four years time interval following the mainshock. Moreover, since all the simple decay models we consider are not suitable to reproduce complex sequences with strong secondary clustering we excluded from our dataset, by a visual analysis of the plot of the rate over 5 days bins, the sequences showing at later times one or more peaks of the shock rate with amplitude of the same order of magnitude of the peak following the mainshock.

The detected sequences are listed in Table 1. The longer time window (four years instead of one) and the different completeness thresholds and selection criteria here adopted reduces the number of sequences (from 37 to 18 for California and from 10 to 6 for Italy) with respect to those detected by Lolli and Gasperini (2006).

- 128
- 129 Analysis
- 130

131 We estimated the parameters of each decay model by the maximum likelihood method (Ogata, 132 1988). To maximize the likelihood we use an algorithm (Lolli et al., 2009) that combines a random 133 search over a reasonable interval of the parameters space and Newton-like optimizations (Dennis 134 and Schnabel, 1983) of the best random solutions. We estimate the parameters of our sequences 135 both with and without the inclusion of the constant background term and by considering different 136 lengths of the observing interval of 3, 6, 12, 24 and 48 months.

137 We compare the goodness-of-fit of the different decay models by three criteria: the corrected Akaike Information Criteria (AICc, Akaike, 1974; Hurvich and Tsai, 1989), the corrected Bayesian 138 139 Information Criteria (BIC, Schwartz, 1978; Draper, 1995) and the simple maximum log-likelihood 140 function  $l_{\text{max}}$ . For the AIC<sub>c</sub> and BIC we adopt (consistently with Lolli and Gasperini, 2006) the 141 following scores

- 142

× [5]

143

144 where *k* is the number of free parameters (3 for the MOM, 4 for MSE and LPL and one more for all 145 models when the background rate  $\prod$  is considered), and n is the number of data (the number of 146 aftershocks in each sequence). With these formulations, which differ from the usual ones for the 147 sign and for a factor of 2, the best model is the one giving the highest score.

×

148 In the following comparisons, we will also consider  $l_{max}$  because we might hypothesize that the 149 additional parameter of the MSE and LPL, which models the exponential decay, might not be able 150 improve significantly the fit (and increase correspondingly the log-likelihood function) when the 151 length of the observing time interval (the assumed duration of the sequence) is short with respect to 152 the relaxation time ( $t_0$  for the MSE and about  $t_a$  for the LPL). In these cases the penalty terms 153 assigned by  $AIC_c$  and BIC to the additional parameter might be oversized. Moreover, the model 154 with the highest log-likelihood, whatever the number of parameters, is the one that best reproduce

[6]

- 155 the behavior of the rate. Hence it is the most suitable for real-time forecast of aftershock 156 probabilities, when the peculiar properties of the active sequence are not known well.
- 157

#### 158 **Results and discussion**

159

160 The parameters of the various models and the relevant goodness-of-fit estimators without and with 161 background are listed in Table 2, for the longest time interval of 48 months (1460 days). For the 162 computations not considering the background rate (Table 2, left), the maximum likelihood estimates of parameter  $p_3$  (corresponding to the exponential decay characteristic time  $t_0$  for the MSE and  $t_a$  for 163 164 the LPL) are in most cases definitely larger than the duration of the observing time interval (highlighted with bold type) and often coincides with the upper limit (Up lim.) of  $10^7$  days we 165 imposed in likelihood maximization. Conversely, when the background rate is included in 166 167 computations (Table 2, right), the estimates of p<sub>3</sub> are in most cases shorter than the observing 168 interval (on the order of some weeks to some months). We can also note in Table 2 a general 169 increase of the estimated power law exponent (p1) for all of the models when the background term 170 is considered. For some sequences (e.g. cal08, cal11, cal12, cal16) such increase is particularly 171 relevant for the MOM (from about 1 to 1.5 and more). As shown by Gasperini and Lolli (2008) by 172 simulation of synthetic sequences, such high p values for the MOM might be the symptom of an 173 early startup of the exponential decay. Parameter p2 (the initial delay time) is not affected instead 174 very much by the inclusion of background. It only tends to slightly increase as a consequence of the 175 increase of the power law exponent, being the two parameters correlated to each other (Gasperini 176 and Lolli, 2006).

The inclusion of the background has also the effect to improve the fit of the three models as shown by the increase of maximum log-likelihoods ( $l_{max}$ ) for almost all sequences and models. The *AIC<sub>c</sub>* and *BIC* scores are also higher for most sequences. The sequences showing lower *AIC<sub>c</sub>* scores (cal02, cal03, cal06, cal07, cal10, cal17, ita01, ita06) and *BIC* (cal01, cal02, cal03, cal06, cal07,

181 cal10, cal17, ita01, ita06) for the best model (highlighted with bold type) are characterized by 182 relatively low background rates. In these cases the slight improvement of the likelihood function 183 induced by the additional parameter (the background rate) is not large enough to compensate the 184 penalty terms added by the information criteria.

In the following we will evaluate the relative efficiency of the different models by counting the number of sequences for which each model is the best among the alternatives, according to the three criteria. The results are represented as line plots of such counts as a function of the considered length (of 3, 6, 12, 24 and 48 months) of the time window.

189 Fig. 1 concerns sequences of Southern California when the background rate is neglected. We note 190 (in Fig. 1a) that the MOM is preferable with respect to the other models according to the  $AIC_c$  for 191 more than 2/3 of the sequences. Such prevalence is clearer for *BIC* (Fig. 1b), for which the MOM is 192 preferable with respect to MSE and LPL for more than 3/4 of the sequences. These results are very 193 similar to those obtained by Lolli and Gasperini (2006) on a different set of sequences. We can note 194 that the number of sequences that are better fitted by the MOM increases slightly with the 195 increasing the duration of the time window. Fig. 1c shows that, at the maximum duration of 48 196 months, the MOM has the highest maximum log-likelihood  $l_{max}$  for 7 of the 18 sequences.

In Fig. 2 we report the same computations of Fig. 1 when the background rate is included into the 197 rate equations as an additive constant term  $\mu$ . It is evident how the MSE and the LPL definitely 198 199 improve their performances with respect to the MOM. For  $AIC_c$  (Fig. 1a), the LPL and the MSE are 200 preferred with respect to the MOM for more than one half of the sequences while for *BIC* (Fig. 1b) 201 the preference goes to MSE or to the LPL for 8 sequences over 18 (Table 3). For the longer time 202 window (48 months), the maximum log-likelihood (Fig. 1c) of the MOM is higher than those of the 203 two alternative models for only one sequence (cal12). We can also note that the numbers of 204 preferences for the various models are weakly dependent on the duration of the observing interval. 205 A similar behavior was shown by the Italian sequences. Even in this cases the values of p3

decreases while p1 and p2 increases when the background is included in computations (Table 2). In

Fig. 3 we report the number of preferences for the various model only for the  $AIC_c$ , which represent an intermediate weighting of the additional parameters between the simple maximum log-likelihood (zero weight) and the *BIC* (highest weight). We can see that the LPL definitely improves its performance with respect to both the MOM and the MSE when the background is considered. For the longer time window of 48 months, only two Italian sequences (ita02, ita03) show larger scores for the MOM.

These evidences indicate that the emergence of exponential decay is a general characteristic of most sequences both in Southern California and Italy and that the background rate (if not appropriately modeled) has the effect to hide such emergence in many cases.

In summary (Table 3), for the longer time window of 48 months the MSE or the LPL perform better than the MOM for 15 sequences over 24 for  $AIC_c$ , 11 for *BIC* and 23 for  $l_{max}$ , when the background is properly modeled. The MSE and the LPL individually have both a larger  $l_{max}$  than the MOM for 18 sequences over 24 but a larger  $AIC_c$  scores for 8 and 14 sequences respectively and larger *BIC* scores for 8 and 10 sequences respectively.

221 The direct comparison between LPL and MSE shows that the former is preferable with respect to 222 the latter for 16 sequences versus 8 for all of the scores (as the two models have the same number of 223 free parameters). We can also note from Table 2 that for the six sequences (cal03, cal07, cal09, 224 call2, cal 13 and call7) for which the MOM has a higher log-likelihood score than the LPL the log-225 likelihood difference is on the order of a few units at most, indicating that the two models show a 226 very similar fit. Hence, even though the exponential decay, on the basis of information criteria, 227 might be not necessary to reproduce some sequences, we can assert that the LPL represents the 228 most suitable model when the actual properties are not known well, as in real-time forecasting of 229 aftershock probabilities (Gerstenberger et al., 2007) of an active sequence.

230

#### 231 Comparing the LPL with the Dieterich (1995) rate equation

To better understand the physical implication of the preference given to the LPL we will attempt an empirical comparison between such model and the rate equation implied by the Dieterich (1994) rate- and state-dependent friction model

236 
$$\lambda(t) = \frac{\mu_r \dot{\tau} / \dot{\tau}_r}{\left[\frac{\dot{\tau}}{\dot{\tau}_r} \exp\left(\frac{-\Delta \tau}{A\sigma}\right) - 1\right] \exp\left[\frac{-t}{t_c}\right] + 1}$$
[7]

where  $_{r}$  is the background rate before the mainshock,  $\dot{\tau}_{r}$  and  $\dot{\tau}$  are the shear stress rates prior to and following the shear stress step induced by the mainshock, *A* is a fault constitutive parameter, in the normal stress and  $t_{c} = A\sigma/\dot{\tau}$  a characteristic relaxation time. When the stress after the mainshock is about constant ( $\dot{\tau} \approx 0$ ) eq. [7] becomes equivalent to the Omori's law [1] (with indication). For  $\dot{\tau} \neq 0$ , eq. [7] also gives the Omori's law at short times ( $t/t_{c}$  1) but merges to the steady state background rate at long times ( $t/t_{c}$  1).

The formulation of eq. [7] is not particularly suitable to empirically fitting real sequences, because the maximum likelihood method is not able to constrain independently the 6 unknown parameters  $(, \dot{r}, \dot{\tau}, \dot{\tau}, \dot{r}, ..., A, ...)$  from sequence data due to their mutual correlation. Although some assumptions could be made on the values of some parameters we will adopt here a purely empirical approach where the 6 free parameters of eq. [7] are combined into 3

248 
$$\lambda_{DRL}(t) = \frac{\mu}{\left[C-1\right] \exp\left[\frac{-t}{t_c}\right] + 1}$$
[8]

249 where  $\mu = \mu_r \dot{\tau} / \dot{\tau}_r$  (the steady state background rate after the mainshock),  $C = \frac{\dot{\tau}}{\dot{\tau}_r} \exp\left(\frac{-\Delta \tau}{A\sigma}\right)$  and  $t_c$ 

are empirical parameters to be determined by maximum likelihood estimation. For the sake of a comparison with the LPL,  $t_c$  has a meaning comparable  $t_a$ , and the product  $Ct_c$  roughly corresponds (see eq. [16] in Dieterich, 1994) to *c* of the MOM an then to  $t_b$  of the LPL. We must note that such form [8] of the Dieterich rate law (DRL) maximizes the ability to fit the data because it does not imply any physical constraint on the parameter values. So its performance might be

slightly better than those of the original formula (eq. [7]) when a physically consistent value isassigned, for example, to \_\_\_\_\_\_.

In Table 4 we compare the estimated values of parameters  $\mu$ ,  $Ct_c$  (p2) and  $t_c$  (p3) as well as the goodness-of-fit scores of the DRL model with the parameters and the scores of the LPL (including background), for the same set of sequences analyzed previously (Table1) and the maximum duration of 48 months. The values of  $\Box$  as well as of  $Ct_c$  with  $t_b$  (p2 in Table 4) appear reasonably consistent for most sequences. On the contrary  $t_c$  is usually larger than  $t_a$  (p3 in Table 4) in most cases.

For all but four sequences (cal03, cal07, cal09, cal17), the maximum log-likelihood  $l_{max}$  is larger for 263 264 the LPL than for the DRL. The  $AIC_c$  and BIC scores are slightly less favorable to the LPL, due to 265 the lower number of free parameters of DRL (k=3) with respect to the LPL (k=5 including the background rate). In Fig. 4 we plot the behavior of the number of preferences of  $AIC_c$ , BIC and  $l_{max}$ 266 267 scores for the two models as a function of the duration of the time window over which the 268 sequences are observed. For short durations (3 and 6 months) the, LPL appears preferable with 269 respect to the DRL for about 2/3 of the 24 sequences. The preferences for the LPL tend to increase 270 for increasing durations up to 24 months. Then the performance of the LPL worsens slightly and for 271 48 months (Table 3) the preferences for the LPL and DRL become respectively 17 versus 7 for 272 AIC<sub>c</sub>, 15 versus 8 for BIC, and 20 versus 4 for  $l_{max}$ .

This behavior can be explained by the interplay between the log-likelihood differences and the penalty terms of the  $AIC_c$  and BIC scores: the likelihood difference between the LPL and the DRL tends increase (Fig 4c), for increasing durations from 3 to 12 months, while for larger durations the tendency reverses and even the penalty terms tend to favor more the DRL, due to the increasing number of data.

In summary, the LPL performs generally better than the DRL and is definitely the most suitable model when the aim is to reproduce well the behavior of the aftershock rate. However the DRL,

- 280 which with only three free parameters shows to explain well a significant portion of the sequences,
- appears to pick much of the physics of the process of aftershock generation (albeit not all).
- 282

#### 283 Visual comparison of rate decay models

284

285 To better describe the different performance of various models, we plot in Fig. 5 the behavior with 286 the time elapsed after the mainshock of the observed (symbols) and predicted (lines) rates for two 287 sequences (call6 and ita02) that show the clear emergence of the exponential decay at relatively 288 long times when the background is modeled. We can see how the LPL (blue) and the MSE (green) 289 models, are able to reproduce better than the MOM (red) the transition of the rate to the background level at times on the order of  $10^2$  days. Moreover they also seem both to describe better than the 290 291 MOM the rate evolution in the first times after the mainshock (< 0.1 days). In both cases the DRL 292 (black) appears to be too "rigid" to follow well the behavior of aftershock rate decay. It tends to 293 overestimate the rate at short and long times and to underestimate at intermediate times (0.1 to 10 294 days). We can argue that the assumed functional form of the transition to the background rate at 295 long times and the power law exponent fixed to 1 at intermediate times prevent a good fit even at 296 short times.

In Fig 6 we report, for the same two sequences, the differences between the observed  $N_{observed}(t)$  and predicted  $N_{model}(t)$  cumulative number of aftershocks with time. The predicted numbers are computed, for various models, as time integrals of the rate functions from the time  $t_1$  of the first aftershock to the time  $t_i$  of each *i*-th aftershock

301

×	[9]
	L -

while the observed cumulative number is simply  $N_{observed}(t_i)=i$ . These plots confirm that the MOM has generally larger differences than MSE and LPL both at short and long times, while the DRL largely underestimates (positive difference) the cumulative number of shocks at times between few

days and 200 days. For sequence call6 we can note a marked negative difference for all models at times between 1 and 3 years that rapidly converge to 0 at larger times. This might indicate that the decay of the main sequence is probably slightly faster than that predicted by all of the models but the occurrence of a burst of shock (maybe due a small sequence) at about 3 years after mainshock prevents a more accurate modeling of the behavior at long times. These late shocks do not affect much the fit at times shorter than three years because the maximum likelihood estimation of parameters is controlled by the more numerous shocks occurring in the first part of the sequence.

312 In Fig. 7 we show the behavior with time of the rate (a) and of the cumulative number difference (b) 313 for a sequence (cal08) characterized by a high power-law exponent (p=1.86 for the MOM and 314 q=1.54 for the LPL), when the background is included in computations. We can see how the MOM 315 (red) and the LPL (blue) curves are almost superimposed among each other. In fact, although the 316 LPL has higher  $l_{\text{max}}$  and AIC<sub>c</sub>, the MOM is preferable according to BIC (see Table 2). Such high 317 value of the power-law exponent prevents a good fit by the MSE (green), for which r is limited in 318 the interval ]0,1[ and by the DRL (black), which assumes a power law exponent equal to 1. This 319 suggests that Dieterich (1994) model might neglect some unknown physics properties of the 320 aftershock generation process.

321

#### 322 Control experiments

323

As the behavior of aftershock sequences in the first times after the mainshock might not be due to true physical processes but rather to the incompleteness of the seismic catalog (Narteau et al., 2002; Lolli and Gasperini, 2006), we tested the stability of our computations as a function of the starting time of the interval over which the analysis is performed. We rerun our computations, for the longest time interval of 48 months, but removing from the datasets the first 10 minutes (0.007 days), the first hour (0.042 days) or the first day. In the first two cases (Table 5), the results of the comparison between MOM, MSE and LPL are similar to those obtained using the entire sequences.

We can note only a slight increase of the preferences to the MOM with respect to the alternative models MSE and LPL (2 sequences more for AICc and 1 for BIC and  $l_{max}$ ) and an increase of preferences to LPL with respect to MSE (3 sequences more). The direct comparison between LPL and DRL also gives similar results with a slight increase of preferences to LPL.

335 When instead the first day is not considered in the analysis, the preferences to the MOM with 336 respect to MSE and LPL increase further by a couple of sequences for  $l_{max}$  but more clearly for the 337 other criteria so that about 20 of the 24 sequences give a preferences to the MOM according to 338 AICc and BIC. In this case we also have a dramatic increase of p above 1.5 (up to 4) and of c above 339 1 day (up to 31 days), for many sequences (Table 6). Such values of MOM parameters are rarely 340 reported in the literature and can be considered unrealistic and not justified physically. In fact, 341 many authors have argued that c should be zero in principle and that non-zero values are due to the 342 incompleteness of the catalog (Narteau et al., 2002; Lolli and Gasperini, 2006) or to physical 343 processes occurring at very short times after the mainshock (Nanjo et al., 2007). When the first part 344 of the sequence is not considered in computations, such incompleteness or such processes cannot 345 influence significantly the estimated parameters, hence the high values of p and c estimated when 346 *Ts*=1 should be explained otherwise.

347 Lolli et al. (2009) found that high p and c values are estimated when fitting (by a MOM) sequences 348 simulated according to a MSE or a LPL with an early onset of the exponential decay. We could 349 argue that a similar phenomenon occurs in this case. We might say that, as c is not useful to 350 reproduce the decay at short times because such times are excluded from computations, the MOM 351 'uses' c to reproduce the deviation of rate function from power law (due to the onset of the 352 exponential decay). The better AICc and BIC scores of the MOM with respect to the MSE and the 353 LPL can be explained as well by the fact that the latter models pay a penalty for a parameter (d for 354 MSE and  $t_b$  for LPL) that is not useful to improve the fit at short times while the MOM can 355 'recycle' parameter c to improve the fit at long times. Hence we might consider the behavior

observed when Ts=1 as a further evidence of the emergence of the exponential decay for many sequences and of the inadequacy of the MOM to reproduce consistently their behavior.

When Ts=1 even the direct comparison between LPL and DRL (Table 5) shows significant variations with a reduction of the preferences to LPL for both AICc (2 sequences less) and BIC (4 sequences less). Conversely, the maximum log likelihood  $l_{max}$  of the LPL becomes the largest for all of the sequences. This means that the DRL is able to describe as well as the LPL the behavior at long times while it is less appropriate at short times where the earthquake catalog might be incomplete. We can argue that the DRL captures much of the true physical properties but it is not particularly suited in general to reproduce the empirical behavior of real sequences.

NO

365

#### 366 **Conclusions**

367

368 We verified that if background rate is modeled properly, the most of a set of 24 real sequences in 369 Italy (6) and Southern California (18) show the emergence of a negative exponential decay of 370 aftershock rate after the initial time interval where the power-law dominates. In fact, two decay 371 models that predict such exponential decay – the Modified Stretched Exponential (MSE, Gross and 372 Kisslinger, 1994) and the band Limited Power Law (LPL, Narteau et al, 2002) - have a higher 373 maximum log-likelihood  $l_{max}$  than the Modified Omori Model (MOM, Utsu, 1971) for 23 sequences 374 over 24. The MSE and the LPL are to be preferred with respect to the MOM for about one half of 375 the sequences, on the basis of the corrected Akaike Information Criterion  $(AIC_c)$  and the Bayesian 376 Information Criterion (BIC). In particular, the LPL alone performs better than the MOM for 18 377 sequences over 24 according to  $l_{max}$ , 14 to AIC<sub>c</sub> and 10 to BIC. In most cases, the estimated 378 characteristic times of the exponential decay are on the order of some weeks to some months.

The inclusion of the background in the rate equation is necessary because, when neglected, the emergence of the exponential decay is somehow hidden for many sequences. The inclusion of the background term has also the effect to reduce the estimates of the characteristic time of the

exponential decay (for the MSE and LPL) and to increase the estimates of the power-law exponentand, to a minor extent, of the initial delay time for all of the models and particularly for the MOM.

As the LPL is generally preferable with respect to the MSE and is able to reproduce well the effective rate decay of real sequences in most cases, it is reasonable to adopt it (with background included) in future analyses of aftershock decay and particularly in real-time forecasts of aftershock probabilities (Gerstenberger et al., 2007) where the actual properties of the sequences are not known well.

We also found that an empirical rate formula equivalent to that predicted by the Dieterich (1994) rate- and state-dependent friction model, with only three free parameters, is able to explain quite well a significant portion of the sequences but performs generally worse than the LPL. This indicates that the Dieterich (1994) rate equation is able to describe well much of the physics of the process of aftershock generation but also that some further developments are needed to make it suitable for best reproducing the observed behavior of aftershock rates.

395

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452

#### **Table captions**

Table 1. List of sequences from Southern California (calxx) and Italy (itaxx) analyzed in this work. Dates and geographical coordinates refer to mainshock origin time and epicenter.  $M_m$  is the mainshock magnitude,  $M_{min}$  is the minimum magnitude of aftershocks, *Nev* is the number of aftershocks (above  $M_{min}$ ) within the influence zone of the mainshock, *D* is the effective duration of the sequence (the time difference between the mainshock and the last aftershock), and  $R_i$  the radius of the influence zone of the mainshock (see text).

459

460 Table 2. Parameters values and goodness-of-fit scores of the different decay models not including 461 (left) and including (right) the background rate ( $\mu$ , in shocks/day), for the maximum sequence 462 duration of 48 months.  $p_1$  is the power law exponent (p for the MOM, r for the MSE and q for the 463 LPL),  $p_2$  is the initial delay time (c for the MOM, d for the MSE,  $t_b$  for the LPL) in days, and  $p_3$  the 464 characteristic time of the negative exponential decay (not defined for the MOM,  $t_0$  for the MSE,  $t_b$ 465 for the LPL) in days. p<sub>3</sub> estimates definitely longer than the duration of the sequence (1460 days) 466 are highlighted with bold type and those at the lower (1 day) and upper (1  $10^7$  days) limit assumed 467 in maximum likelihood estimation are denoted by "Low lim." and "Up lim." respectively. AIC<sub>e</sub>, BIC, and  $l_{max}$  are the goodness-of-fits scores for each decay model (without and with background). 468 469 For each sequence, the higher score (highlighted with bold type) indicates the model preferred 470 according to each criterion.

471

Table 3. Counts of sequences for which the different criteria assign the preference to variousmodels, for the maximum sequence duration of 48 months.

474

Table 4. Parameters values and goodness-of-fit scores of the LPL (with background) and RDL
decay models for a sequence duration of 48 months. p<sub>1</sub> is the power law exponent for the LPL, p<sub>2</sub> is

the initial delay time ( $t_b$  for the LPL,  $Ct_c$  for the DRL) in days, and  $p_3$  the characteristic time of the negative exponential decay ( $t_b$  for the LPL and  $t_c$  for the DRL) in days.  $p_3$  estimates definitely longer than the duration of the sequence (1460 days) are highlighted with bold type and those at the lower (1 day) and upper (1 10<sup>7</sup> days) limit are denoted by "Low lim." and "Up lim." respectively. AIC<sub>c</sub>, BIC, and  $l_{max}$  are the goodness-of-fits scores for each decay model. For each sequence, the higher score (highlighted with bold type) indicates the model preferred according to each criterion. Table 5. Counts of sequences, from California and Italy, for which the different criteria assign the

485 preference to various models, for the maximum sequence duration of 48 months and different 486 starting times  $T_s$  (in days) of the observing time interval.

487

488 Table 6. Parameters values of the MOM for different starting times  $T_s$  (in days) of the observing 489 time interval. Values of p>1.5 and c>1 day are highlighted with bold type.

490

491

#### Table 1

### **Detected sequences**

Seq.	Year	Mo	Day	Lat	Lon	$M_m$	$M_{min}$	Nev	D	$R_i$
				(North)	(East)				(days)	(km)
cal01	1933	3	11	33.638	-117.973	6.4	3.0	269	1451	42
cal02	1946	3	15	35.702	-117.944	6.3	3.0	154	1364	30
cal03	1947	4	10	34.983	-116.531	6.5	3.0	124	1455	27
cal04	1954	3	19	33.298	-116.081	6.4	3.0	136	1452	-32
cal05	1968	4	9	33.167	-116.087	6.6	3.1	162	1445	52
cal06	1971	2	9	34.416	-118.370	6.6	3.1	291	1408	27
cal07	1979	3	15	34.327	-116.445	5.3	2.5	176	1451	13
cal08	1981	4	26	33.096	-115.624	5.8	2.5	186	1425	16
cal09	1986	7	8	33.999	-116.608	5.7	2.5	868	1452	44
cal10	1986	7	13	32.971	-117.874	5.5	2.5	1686	1457	22
cal11	1987	2	7	32.388	-115.305	5.4	2.5	225	1454	40
cal12	1987	11	24	33.015	-115.852	6.6	3.1	216	1446	41
cal13	1994	1	17	34.213	-118.537	6.7	3.2	344	1455	25
cal14	1999	10	16	34.594	-116.271	7.1	3.6	151	1367	52
cal15	2001	7	17	36.016	-117.874	5.2	1.7	2009	1457	28
cal16	2002	2	22	32.319	-115.322	5.7	2.2	785	1457	27
cal17	2003	12	22	35.709	-121.104	6.5	3.0	129	1444	32
cal18	2004	9	28	35.812	-120.379	6.0	2.5	147	1377	48
ita01	1980	11	23	40.800	15.367	6.5	3.0	105	1407	41
ita02	1984	4	29	43.204	12.585	5.2	2.5	130	1459	41
ita03	1990	5	5	40.650	15.882	5.6	2.5	109	1453	41
ita04	1997	9	26	43.015	12.854	5.8	2.5	774	1354	46
ita05	2002	9	6	38.381	13.654	5.6	2.5	181	1414	26
ita06	2002	10	31	41.717	14.893	5.4	2.5	163	1280	26

Table 2

Seq.	Model	el Without background						With background							
		p1	p2	p3	AICc	BIC		μ	p1	p2	р3	AICc	BIC		
cal01	MOM	1.090	0.061		210.778	208.188	213.823	0.012	1.160	0.082		211.467	208.029	215.542	
	MSE	0.886	0.029	Low Lim.	206.087	202.649	210.163	0.011	0.837	0.030	Low Lim.	205.811	201.533	210.926	
	LPL	1.080	0.055	Up. Lim.	211.975	208.538	216.051	0.011	1.130	0.073	9.1E+06	212.352	208.074	217.467	
cal02	MOM	1.180	1.360		-139.495	-141.213	-136.415	0.008	1.290	1.780		-139.746	-142.010	-135.611	
	MSE	0.847	0.861	Low Lim.	-142.022	-144.286	-137.888	0.010	0.795	1.030	Low Lim.	-142.147	-144.942	-136.944	
	LPL	1.130	0.868	Up. Lim.	-137.854	-140.118	-133.720	0.006	1.190	1.150	5.8E+06	-138.479	-141.274	-133.276	
cal03	MOM	0.974	0.008		-37.758	-39.132	-34.658	0.004	1.000	0.012		-38.578	-40.374	-34.409	
	MSE	0.948	0.008	Up. Lim.	-38.828	-40.624	-34.660	0.011	0.797	0.003	5	-39.249	-41.451	-33.995	
	LPL	0.955	0.005	8.7E6	-39.322	-41.119	-35.154	0.014	0.933	0.004	373	-40.007	-42.209	-34.753	
cal04	MOM	1.050	0.005		116.559	115.038	119.650	0.010	1.130	0.010		119.506	117.509	123.659	
	MSE	0.950	0.003	Low Lim.	114.436	112.439	118.588	0.014	0.742	0.001	Low Lim.	119.142	116.686	124.373	
	LPL	1.040	0.005	Up. Lim.	115.787	113.790	119.939	0.016	1.010	0.005	94	121.385	118.928	126.615	
cal05	MOM	0.850	0.008		-212.914	-214.713	-209.838	0.045	1.210	0.108		-196.063	-198.435	-191.935	
	MSE	0.831	0.007	Up. Lim.	-214.668	-217.040	-210.541	0.051	0.725	0.031	Low Lim.	-196.443	-199.375	-191.251	
	LPL	0.849	0.011	Up. Lim.	-213.667	-216.040	-209.540	0.052	0.979	0.036	59	-196.518	-199.450	-191.326	
cal06	MOM	1.050	0.002		650.169	647.457	653.210	0.003	1.060	0.002		649.272	645.671	653.342	
	MSE	0.872	0.001	Low Lim.	648.102	644.501	652.172	0.013	0.807	0.001	Low Lim.	648.499	644.015	653.604	
	LPL	1.030	0.001	Up. Lim.	648.319	644.718	652.389	0.022	0.992	0.001	190	650.093	645.609	655.198	
cal07	MOM	0.983	0.042		-79.017	-80.946	-75.947	0.002	0.993	0.046		-80.042	-82.590	-75.925	
	MSE	0.869	0.026	102	-79.962	-82.511	-75.845	0.017	0.697	0.002	12	-80.096	-83.251	-74.920	
	LPL	0.973	0.040	Up. Lim.	-80.996	-83.544	-76.879	0.025	0.799	3.4E-07	104	-81.636	-84.791	-76.459	
cal08	MOM	1.080	0.018		134.924	132.908	137.990	0.036	1.860	0.229		187.611	184.946	191.722	
	MSE	0.999	0.003	Low Lim.	130.356	127.690	134.466	0.036	0.501	1.0E-06	Low Lim.	175.349	172.046	180.516	
	LPL	1.090	0.028	Up. Lim.	136.149	133.483	140.259	0.036	1.540	0.139	46	187.865	184.562	193.031	
cal09	MOM	0.802	0.002		100.959	96.580	103.973	0.253	1.070	0.035		171.333	165.500	175.357	
	MSE	0.788	0.001	Up. Lim.	96.460	90.626	100.483	0.280	0.809	0.018	Low Lim.	170.929	163.643	175.964	
	LPL	0.803	0.003	Up. Lim.	100.924	95.090	104.947	0.274	1.030	0.024	773	168.214	160.928	173.249	

23

# Parameters values and goodness-of-fit scores

 Table 2 (continued)

Seq.	Model			Without	backgroun	d					With backg	round			
		p1	p2	p3	AICc	BIC		μ,Μ	p1	p2	р3	AICc	BIC		
cal10	MOM	0.725	0.031		183.112	177.731	186.119	1.2E-05	0.725	0.031		182.106	174.934	186.1	18
	MSE	0.598	1.0E-06	1260	193.267	186.095	197.279	1.2E-05	0.598	1.0E-06	1260	192.261	183.298	197.2	279
	LPL	0.630	4.1E-06	1790	194.930	187.757	198.942	1.2E-05	0.630	3.2E-06	1790	193.924	184.961	198.9	942
cal11	MOM	0.673	1.0E-06		-465.255	-467.568	-462.200	0.116	1.570	0.163		-431.516	-434.581	-427.4	125
	MSE	0.664	1.0E-06	Up. Lim.	-466.999	-470.064	-462.908	0.118	0.464	1.0E-06	Low Lim.	-432.487	-436.296	-427.3	\$50
	LPL	0.673	3.2E-05	Up. Lim.	-466.347	-469.413	-462.256	0.119	0.626	3.4E-06	2	-431.204	-435.012	-426.0	)67
cal12	MOM	0.962	0.009		15.087	12.837	18.143	0.055	1.450	0.090		44.980	42.000	49.0	)75
	MSE	0.945	0.010	Up. Lim.	13.480	10.500	17.575	0.060	0.552	0.001	Low Lim.	38.095	34.394	43.2	238
	LPL	0.961	0.011	Up. Lim.	15.495	12.515	19.590	0.053	1.340	0.064	4.0E+06	43.005	39.305	48.1	.48
cal13	MOM	1.150	0.066		529.989	527.020	533.024	0.016	1.260	0.110		533.566	529.620	537.6	525
	MSE	0.838	0.020	1	523.966	520.019	528.025	0.027	0.723	0.020	Low Lim.	535.597	530.679	540.6	686
	LPL	1.130	0.051	Up. Lim.	525.881	521.934	529.940	0.027	1.110	0.054	197	530.123	525.205	535.2	212
cal14	MOM	1.190	0.037		206.291	204.603	209.372	0.002	1.220	0.045		205.622	203.401	209.7	/59
	MSE	0.794	0.005	1	204.637	202.415	208.774	0.008	0.703	0.003	Low Lim.	209.143	206.401	214.3	349
	LPL	1.150	0.025	Up. Lim.	203.749	201.527	207.886	0.010	0.945	0.008	40	206.945	204.203	212.1	.52
cal15	MOM	0.951	0.165		2646.788	2641.143	2649.794	0.297	1.290	1.070		2702.098	2694.573	2706.1	.08
	MSE	0.922	0.160	Up. Lim.	2645.082	2637.557	2649.092	0.411	0.462	1.0E-06	12	2782.900	2773.496	2787.9	)15
	LPL	0.925	0.110	Up. Lim.	2634.159	2626.634	2638.169	0.423	0.547	0.002	26	2783.754	2774.350	2788.7	68
cal16	MOM	0.745	0.003		-365.589	-369.816	-362.574	0.308	1.740	1.150		-194.518	-200.148	-190.4	92
	MSE	0.732	0.002	Up. Lim.	-370.059	-375.689	-366.033	0.313	0.624	0.240	Low Lim.	-188.049	-195.080	-183.0	)10
	LPL	0.746	0.008	Up. Lim.	-366.548	-372.178	-362.523	0.317	0.489	8.5E-07	6	-184.342	-191.373	-179.3	303
cal17	MOM	0.994	0.025		-86.333	-87.770	-83.237	0.002	1.010	0.033		-87.325	-89.208	-83.1	.64
	MSE	0.895	0.014	55	-87.350	-89.233	-83.189	0.012	0.745	1.0E-06	7	-87.635	-89.946	-82.3	<b>391</b>
	LPL	0.976	0.018	Up. Lim.	-87.793	-89.676	-83.632	0.016	0.919	6.2E-06	253	-88.705	-91.016	-83.4	61
cal18	MOM	0.645	1.0E-06		-377.158	-378.803	-374.074	0.057	0.865	1.0E-06		-368.524	-370.688	-364.3	583
	MSE	0.639	1.0E-06	Up. Lim.	-378.495	-380.659	-374.354	0.057	0.850	1.0E-06	Up Lim.	-369.634	-372.302	-364.4	21
	LPL	0.645	1.4E-06	Up. Lim.	-378.249	-380.414	-374.109	0.065	0.856	2.8E-06	711	-369.560	-372.229	-364.3	347

 Table 2 (continued)

Seq.	Model			Without	backgroun	d			With background					
_		p1	p2	p3	AICc	BIC		μ	p1	p2	p3	AICc	BIC	
ita01	MOM	0.861	0.038		-202.562	-203.668	-199.444	7.5E-07	0.861	0.038		-203.644	-205.076	-199.444
	MSE	0.745	0.005	1260	-203.462	-204.895	-199.262	0.012	0.654	1.0E-06	60	-204.180	-205.917	-198.877
	LPL	0.829	0.026	6810	-203.817	-205.249	-199.617	0.019	0.730	7.5E-06	194	-204.178	-205.915	-198.875
ita02	MOM	0.908	0.005		-127.406	-128.856	-124.311	0.025	1.210	0.122		-120.859	-122.758	-116.699
	MSE	0.883	0.004	Up. Lim.	-128.684	-130.583	-124.524	0.032	0.559	1.0E-06	3	-115.481	-117.813	-110.239
	LPL	0.890	1.3E-07	Up. Lim.	-128.714	-130.614	-124.554	0.032	0.730	2.3E-06	19	-115.416	-117.748	-110.174
ita03	MOM	0.771	1.0E-06		-243.151	-244.317	-240.037	0.034	0.987	0.007		-236.613	-238.127	-232.420
	MSE	0.759	1.0E-06	Up. Lim.	-244.665	-246.180	-240.473	0.045	0.671	1.0E-06	2	-237.386	-239.228	-232.095
	LPL	0.771	8.1E-07	Up. Lim.	-244.237	-245.752	-240.045	0.047	0.753	4.6E-06	10	-237.106	-238.949	-231.815
ita04	MOM	1.340	4.050		497.048	492.843	500.064	5.8E-06	1.340	4.050		496.038	490.436	500.064
	MSE	0.572	0.048	42	511.769	506.167	515.795	0.046	0.455	1.0E-06	31	526.315	519.320	531.354
	LPL	0.750	0.083	330	491.742	486.141	495.768	0.050	0.551	0.008	71	523.765	516.770	528.804
ita05	MOM	0.960	0.061		-126.564	-128.537	-123.496	1.3E-06	0.960	0.061		-127.610	-130.218	-123.496
	MSE	0.720	0.007	70	-124.837	-127.444	-120.723	0.023	0.515	1.0E-06	15	-115.103	-118.333	-109.931
	LPL	0.899	0.026	3430	-128.744	-131.352	-124.631	0.025	0.623	1.2E-06	48	-112.481	-115.712	-107.310
ita06	MOM	0.974	0.331		-199.462	-201.270	-196.386	0.019	1.110	0.587		-199.809	-202.195	-195.683
	MSE	0.951	0.338	Up. Lim.	-200.571	-202.956	-196.444	0.007	0.998	0.356	2	-201.228	-204.177	-196.037
	LPL	0.974	0.358	Up. Lim.	-199.165	-201.551	-195.039	0.019	1.090	0.508	6.1E+06	-199.329	-202.278	-194.138

### Table 3

				-					
	Ca	liforni	a		Italy		Calife	ornia+I	Italy
Model	AICc	BIC		AICc	BIC		AICc	BIC	
MOM	7	10	1	2	3	0	9	13	1
MSE	2	2	7	1	1	1	3	3	8
LPL	9	6	10	3	2	5	12	8	15
MOM	8	11	6	2	3	0	10	14	6
LPL	10	7	12	4	3	6	14	10	18
MOM	13	13	5	3	3	1	16	16	6
MSE	5	5	13	3	3	5	8	8	18
MSE		7			1			8	
LPL		11			5			16	
LPL	14	13	14	3	3	6	17	16	20
DRL	4	5	4	3	3	0	7	8	4

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### Counts of preferred models (including background)

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### Table 4

### Parameters values and goodness-of-fit scores

Seq.	Model	μ	p1	p2	р3	AICc	BIC	
cal01	LPL	0.011	1.130	0.073	9.1E+06	212.352	208.074	217.467
	DRL	1.0E-07		0.025	2.4E+08	205.206	202.616	208.251
cal02	LPL	0.006	1.190	1.150	5.8E+06	-138.479	-141.274	-133.276
	DRL	4.5E-09		0.545	4.4E+09	-144.569	-146.288	-141.489
cal03	LPL	0.014	0.933	0.004	373	-40.007	-42.209	-34.753
	DRL	0.007		0.011	1390	-37.499	-38.872	-34.399
cal04	LPL	0.016	1.010	0.005	94	121.385	118.928	126.615
	DRL	0.008		0.002	1340	116.246	114.725	119.337
cal05	LPL	0.052	0.979	0.036	59	-196.518	-199.450	-191.326
	DRL	0.047		0.029	223	-196.811	-198.610	-193.735
cal06	LPL	0.022	0.992	0.001	190	650.093	645.609	655.198
	DRL	2.0E-06		0.001	1.1E+07	646.748	644.036	649.789
cal07	LPL	0.025	0.799	3.4E-07	104	-81.636	-84.791	-76.459
	DRL	0.004		0.050	3890	-79.000	-80.929	-75.930
cal08	LPL	0.036	1.540	0.139	46	187.865	184.562	193.031
	DRL	0.030		4.5E-08	454	140.317	138.301	143.383
cal09	LPL	0.274	1.030	0.024	773	168.214	160.928	173.249
	DRL	0.290		0.021	173	170.673	166.294	173.687
cal10	LPL	1.2E-05	0.630	3.2E-06	1790	193.924	184.961	198.942
	DRL	0.346		1.760	597	115.124	109.743	118.131
cal11	LPL	0.119	0.626	3.4E-06	2	-431.204	-435.012	-426.067
	DRL	0.114		0.007	61	-434.675	-436.988	-431.621
cal12	LPL	0.053	1.340	0.064	4.0E+06	43.005	39.305	48.148
	DRL	0.043		0.009	346	28.513	26.263	31.569
cal13	LPL	0.027	1.110	0.054	197	530.123	525.205	535.212
	DRL	1.9E-07		0.016	1.6E+08	514.042	511.073	517.077
cal14	LPL	0.010	0.945	0.008	40	206.945	204.203	212.152
	DRL	3.1E-07		0.004	4.1E+07	194.681	192.993	197.762
cal15	LPL	0.423	0.547	0.002	26	2783.754	2774.350	2788.768
	DRL	0.265		0.200	761	2681.258	2675.612	2684.264
cal16	LPL	0.317	0.489	8.5E-07	6	-184.342	-191.373	-179.303
	DRL	0.307		0.077	147	-219.133	-223.360	-216.118
cal17	LPL	0.016	0.919	6.2E-06	253	-88.705	-91.016	-83.461
	DRL	0.003		0.026	4530	-86.282	-87.719	-83.186
cal18	LPL	0.065	0.856	2.8E-06	711	-369.560	-372.229	-364.347
	DRL	0.075		0.013	70	-371.459	-373.104	-368.375

### Table 4 (continued)

Seq.	Model	μ	p1	p2	р3	AICc	BIC	
ita01	LPL	0.019	0.730	7.5E-06	194	-204.178	-205.915	-198.875
	DRL	0.014		0.223	751	-203.952	-205.058	-200.833
ita02	LPL	0.032	0.730	2.3E-06	19	-115.416	-117.748	-110.174
	DRL	0.026		0.019	370	-121.550	-122.999	-118.455
ita03	LPL	0.047	0.753	4.6E-06	10	-237.106	-238.949	-231.815
	DRL	0.043		0.009	121	-236.343	-237.509	-233.229
ita04	LPL	0.050	0.551	0.008	71	523.765	516.770	528.804
	DRL	2.2E-08		0.517	4.4E+09	467.306	463.101	470.322
ita05	LPL	0.025	0.623	1.2E-06	48	-112.481	-115.712	-107.310
	DRL	0.007		0.097	2820	-126.766	-128.739	-123.698
ita06	LPL	0.019	1.090	0.508	6.1E+06	-199.329	-202.278	-194.138
	DRL	0.011		0.365	1760	-199.197	-201.005	-196.121
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Counts of preferred models using different starting times  $T_s$ 

Table 5

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		$T_s = t_1$		$T_s = 0.007$			$T_s$	=0.042	2	$T_s = 1$		
Model	AICc	BIC		AICc	BIC		AICc	BIC		AICc	BIC	
MOM	9	13	1	10	13	1	11	14	2	20	21	4
MSE	3	3	8	4	3	8	3	3	4	0	0	2
LPL	12	8	15	10	8	15	10	7	18	4	3	18
LPL	17	16	20	17	16	22	18	16	22	15	12	24
DRL	7	8	4	7	8	2	6	8	2	9	12	0

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#### Table 6

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MOM parameter estimates using different starting times  $T_s$ 

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		$T_s = t_1$			$T_s = 1$	
Seq.	μ	р	С	μ	р	С
cal01	0.012	1.160	0.082	0.003	1.030	0.000
cal02	0.008	1.290	1.780	0.005	1.170	0.100
cal03	0.004	1.000	0.012	0.005	0.988	0.006
cal04	0.010	1.130	0.010	0.015	2.130	9.890
cal05	0.045	1.210	0.108	0.049	1.450	1.140
cal06	0.003	1.060	0.002	0.013	1.340	2.350
cal07	0.002	0.993	0.046	0.009	1.140	1.190
cal08	0.036	1.860	0.229	0.036	1.750	0.000
cal09	0.253	1.070	0.035	0.273	1.220	0.830
cal10	0.000	0.725	0.031	0.000	0.817	2.790
cal11	0.116	1.570	0.163	0.119	4.000	5.120
cal12	0.055	1.450	0.090	0.056	1.500	0.000
cal13	0.016	1.260	0.110	0.020	1.330	0.000
cal14	0.002	1.220	0.045	0.007	1.780	2.000
cal15	0.297	1.290	1.070	0.395	2.180	12.600
cal16	0.308	1.740	1.150	0.315	3.130	9.170
cal17	0.002	1.010	0.033	0.011	1.290	1.450
cal18	0.057	0.865	0.000	0.055	0.906	1.680
ita01	0.000	0.861	0.038	0.004	0.942	0.370
ita02	0.025	1.210	0.122	0.032	2.400	6.090
ita03	0.034	0.987	0.007	0.043	1.330	0.000
ita04	0.000	1.340	4.050	0.032	1.870	18.600
ita05	0.000	0.960	0.061	0.024	2.790	31.000
ita06	0.019	1.110	0.587	0.029	1.210	0.000

514	Figure captions
515	
516	Figure 1. Number of sequences of Southern California for which each decay model is preferable
517	according to $AIC_c$ (a), $BIC$ (b) and maximum log-likelihood (c) as a function of the duration of the
518	sequence, when the background rate is neglected.
519	
520	Figure 2. Number of sequences of Southern California for which each decay model is preferable
521	according to $AIC_c$ (a), $BIC$ (b) and maximum log-likelihood (c) as a function of the duration of the
522	sequence, when the background rate is modeled.
523	
524	Figure 3. Number of sequences of Italy for which each decay model is preferable according to $AIC_c$
525	as a function of the duration of the sequence, when the background rate is neglected (a) or modeled
526	(b).
527	
528	Figure 4. Number of sequences of Italy and Souther California for which the LPL (with
529	background) or the DRL models are preferable according to $AIC_c$ (a) or $BIC$ (b) and maximum log-
530	likelihood (c) as a function of the duration of the sequence.
531	
532	Figure 5. Observed (symbols) and predicted (lines) aftershock rates as a function of time elapsed
533	after the mainshock when the background rate is modeled for two sequences showing the clear
534	emergence of the exponential decay at relatively long times.
535	
536	Figure 6. Differences between observed and predicted cumulative number of aftershocks as a
537	function of time elapsed after the mainshock for the same two sequences of Fig. 5.
538	

- 539 Figure 7. Observed (symbols) and predicted (lines) aftershock rates (a), and differences between
- 540 observed and predicted cumulative number of aftershocks (b) as a function of time elapsed after the
- 541 mainshock when the background rate is modeled, for a sequence characterized by an high power
- 542 law exponent (p=1.86 for the MOM and q=1.54 for the LPL).
- 543

#### 543

### Figure 1



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### Figure 2



Figure 3



Figure 4





### Figure 5













### Figure 7





