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Andrew Burke, Felix Fitzroy, Michael Nolan

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IS THERE A NORTH-SOUTH DIVIDE IN SELF-EMPLOYMENT IN ENGLAND?

by

Andrew E. Burke (Bettany Centre for Entrepreneurial Performance & Economics, Cranfield School of Management, UK and Max Planck Institute for Economics, Germany)

Felix R. FitzRoy (University of St. Andrews, UK and the IZA (Bonn), Germany)

and Michael A. Nolan (University of Hull, UK)

Professor Andrew Burke, Bettany Centre for Entrepreneurial Performance & Economics, Cranfield Management Research Institute (CMRI), Cranfield School of Management, Cranfield, Bedford, MK43 0AL. Email: andrew.burke@cranfield.ac.uk

Felix R FitzRoy, Department of Economics & Finance, University of St Andrews, Castlecliffe, St Andrews, Fife, KY16 9AL. Email: frf@st-andrews.ac.uk

Dr Michael A Nolan, Hull University Business School, University of Hull, HU6 7RX. Email: m.a.nolan@hull.ac.uk

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Abstract

Using decomposition analysis, the paper investigates why Northern England has fewer but higher performing self-employed individuals than the South. We find the causes are mainly structural differences rather than regional variation in individual characteristics. There are more self employed individuals in the South, but on average they create fewer jobs. Post compulsory education has a strong negative effect on the probability of self employment in the South, probably due to better employment opportunities there, but little influence in the North. Education has some positive effects on job creation by entrepreneurs in both regions. Aggregate studies may thus give misleading results.

Keywords: Self-employment, job creation, North-South divide, decomposition.

JEL classifications: R11, R23, J23

Est-il une fracture de Nord Sud en ce qui concerne les emplois indépendants en Angleterre?

Andrew E. Burke, Felix R. FitzRoy et Michael A. Nolan

Résumé

Se pliant sur une analyse par décomposition, cet article explique pourquoi le nord de l'Angleterre compte moins de travailleurs indépendants que le sud et pourquoi ils sont plus performants. Il apparaît que les causes sont principalement des différences structurelles plutôt que des variations régionales des caractéristiques individuelles. Il y a davantage de travailleurs indépendants dans le sud mais en moyenne ils créent moins d'emplois. Le post-enseignement obligatoire a un fort effet négatif sur la probabilité de l'emploi indépendant dans le sud ; cela est probablement dû aux meilleures chances d'emploi dans cette région mais elle a moins d'influence dans le nord. L'éducation à quelques effets positifs sur la création d'emplois par les entrepreneurs dans les deux régions. Des études complémentaires peuvent donc donner des résultats trompeurs.

Classifications JEL: R11, R23, J23

emplois indépendants, création d'emploi, fracture Nord Sud, décomposition

GIBT ES BEI DER FREIBERUFLICHEN TÄTIGKEIT UNTERSCHIEDE ZWISCHEN NORDEN UND SÜDEN? Andrew E. Burke, Felix R. FitzRoy and Michael A. Nolan Abstract

In diesem Beitrag untersuchen wir mit Hilfe einer Dekompositionsanalyse die Frage, warum es in Nordengland weniger, aber dafür leistungsfähigere Freiberufler gibt als im Süden. Wir stellen fest, dass die Ursachen weniger auf regionalen Schwankungen hinsichtlich individueller Merkmale beruhen, sondern vielmehr hauptsächlich auf strukturellen Unterschieden. Im Süden gibt es mehr Freiberufler, doch diese schaffen im Durchschnitt weniger Arbeitsplätze. Eine weiterführende Bildung hat im Süden - wahrscheinlich aufgrund der besseren Beschäftigungschancen - eine starke negative Auswirkung auf die Wahrscheinlichkeit einer freiberuflichen Tätigkeit, während sie im Norden nur wenig Einfluss ausübt. Die Bildung hat in beiden Regionen einige positive Auswirkungen auf die Schaffung von Arbeitsplätzen durch Unternehmer. Zusammengefasste Studien können daher irreführende Ergebnisse hervorbringen.

Keywords:
Freiberufliche Tätigkeit
Schaffung von Arbeitsplätzen
Nord-Süd-Unterschiede
Dekomposition

JEL classifications: R11, R23, J23

¿Existe una división norte-sur en cuanto al nivel de autónomos en Inglaterra? Andrew E. Burke, Felix R. FitzRoy and Michael A. Nolan

Abstract

Mediante un análisis de descomposición, en este artículo investigamos por qué en el norte de Inglaterra hay menos autónomos que en el sur pero con un mejor desempeño. Creemos que las causas radican en diferencias estructurales más que en variaciones regionales de las características individuales. Hay más autónomos en el sur pero de promedio crean menos trabajos. La educación postobligatoria tiene un efecto muy negativo en la probabilidad de convertirse en autónomo en la zona sur del país probablemente debido a mejores oportunidades laborales pero tiene poca influencia en el norte. La educación tiene algunos efectos positivos en la creación de trabajo por parte de empresarios en ambas regiones. Por tanto, los estudios acumulativos podrían dar resultados falsos.

Empleo autónomo
Creación de empleo
División norte-sur
Descomposición

JEL classifications: R11, R23, J23

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1. Introduction:

The regional dimension of the British economy has been well documented. Regional variation in economic performance is revealed in a North-South divide where the South has better economic performance than the North with lower unemployment and higher GDP per capita. This persistent divergence in performance has generated various economic policy responses (Lewis and Townsend (1989), Fothergill (2001) and Gudgin (1996)). In particular, a response to divergence in unemployment (Gray (2004)) or non-employment (Anyadike-Danes (2004)) has been to encourage job creation through self-employment (Shutt and Sutherland (2003)). Robson (1998), and Georgellis and Wall (2000) include a theoretical framework to underlie a relationship between the regional self-employment rate and a variety of characteristics of the regional economy, and estimate results for UK regional data over a short time series. Their models are based on Evans and Jovanovic (1989) and Khilstrom and Laffont's (1979) models of choice between wage work and self-employment. The probability of a person choosing self-employment is positively related to both the relative financial and non-pecuniary benefits compared to wage work, so econometric estimation requires variables affecting these components

To our knowledge, there has not been UK regional analysis at the level of the individual, which is one new contribution of this paper. We thus study regional variation in individual characteristics affecting the *ability* and *predisposition* for self-employment. These include skill, experience, education, psychological attitudes and culture. Variation in regional self-employment rates and job creation can therefore be attributed to both the structural disparities (well documented in research on the determinants of aggregate rates of self-employment) and differences in average individual characteristics which are difficult to measure and estimate at an aggregate level.

In this paper we thus include compositional as well as structural influences. With longitudinal data on individuals, we investigate both the probability of self-employment, and also job creation by the self-employed to explore the North-South divide in English self-employment. To the best of our knowledge, this approach is new.

Neglecting these issues, previous work has not identified the impact of certain individual characteristics on self-employment – if, for example, a significant effect in one region is cancelled by an opposing effect in another. Of particular interest to us is the role of post-secondary education which at a national GB level has been found to reduce the number of self-employed but increase job creation by entrepreneurs. Burke, FitzRoy and Nolan (henceforth BFN, 2000) and Cowling et al (2004) find that the net (aggregate) result of these opposing effects is positive so that education increases the total number of jobs created by the self-employed. Here we disaggregate this result and uncover substantial regional variation. We use data from the National Child Development Study (NCDS) – males and females being considered separately, as was justified previously by BFN (2002), following the 'self-employed female underperformance hypothesis' of Rosa et al (1996), and Du Rietz and Henrekson (2000).

In addition, we use decomposition analysis, to investigate how variation in self-employment between North and South results from differences in individual characteristics or from differential responses to given characteristics. Our work follows Reimers (1983), Cotton (1988), Neumark (1988), Oaxaca and Ransom (1994), and Fairlie (1999, 2003) – who extended the seminal analysis of decomposition by Blinder (1973) and Oaxaca (1973).

Recent international research on new firm formation shows how it effects long term job creation (Audretsch and Fritsch, 2002, Fritsch and Mueller, 2004 and van Stel and Storey 2004). Fristch and Mueller (2004) claim that this effect evolves across three regimes. Initially, business start-ups have a direct positive effect on job creation which then turns negative as some of these firms grow and compete with incumbents. Later, a third inclusive phase occurs where these ventures have positive spillover effects on other businesses. Fritsch and Mueller's evidence is mainly based on German regional data (although their results have been replicated in other economies, see Fritsch and Schmunde (eds) 2006) and they argue that the positive effects (particularly, the direct effect) are more pronounced in high productivity regions. Mueller, Van Stel and Storey (2006), hereafter MVS (2006), find that in contrast the UK economy does not follow this productivity pattern and show that

the direct effect of job creation in the UK is higher in the less productive Northern English regions than in the South.

All of these studies use aggregate data, and our use of individual level data tests whether the UK is indeed an exception to the three regime regional model. Our data is a cross section snapshot of the self-employed and hence only provides a total net effect of self-employment on job creation in the self-employed sector. However, while this tempers the generality of how our findings relate to the displacement and spillover effects, it does not affect the measurement of direct positive effects – a key area where the UK evidence from MVS (2006) diverges from regional effects found in other countries.

Individual level data can also test the impact of education on labour productivity. In general, higher regional levels of education are associated with higher levels of regional productivity (see Lindsay, 2004 for UK evidence). Thus, one might expect that on average the self-employed in more highly educated regions will be more educated than those of less educated regions and therefore have higher productivity levels. While this is often true, and indeed explains the most common patterns observed in the aggregate analyses, our underlying model of self-employment choice (Evans and Jovanovic, 1989) can also generate the opposite effect. Namely, if qualified job opportunities are plentiful then better educated individuals may be drawn into wage work, resulting in lower education and productivity among the self-employed in a region with relatively high average education. By contrast, a less educated region may have higher levels of education among the self-employed if job opportunities are scarce and more of the educated are pushed into self-employment. If this is true and applies to the UK then it could explain why MVS (2006) find that the lower productivity Northern regions have higher positive direct employment effects than the South. undermining the three regime model, the English regions would then be consistent with it. However, the predictive power of the model in relation to productivity and job creation needs to include (dynamic) individual career choice following Evans and Jovanovic (1989), Blanchflower and

Oswald (1998) and BFN (2000, 2002). Our use of individual level data allows us to test this hypothesis and we find that it appears to hold.

To summarise, the novelty of our analytic approach is threefold. First, we complement aggregate regional analysis of self-employment in the UK by estimating, at the level of the individual, the probability of choosing self-employment and the determinants of job creation by the self-employed. Secondly, we introduce the regional dimension to existing studies of self-employment choice and performance at the level of the individual. Finally, we use decomposition analysis in order to separate out the impact of compositional from structural regional variation in determining entrepreneurial choice.

The paper proceeds with Section 2, which focuses on the data. Section 3 discusses methodology, concentrating on application of decomposition to this topic. Section 4 presents results. It focuses initially on analysis of the probability of self-employment, and decomposing differences between South and North into compositional and structural parts, and then proceeds to male self-employment job creation. There then follows a concluding section.

2. Data Description

The National Child Development Study (NCDS) surveys a cohort born in the week 3rd March 1958 to 9th March 1958 inclusive and living in Great Britain. These surveys (in 1958, 1965, 1969, 1974, 1981 and 1991) were used by Blanchflower and Oswald (1998) and BFN (2000, 2002). The first of these papers only considers the self-employment decision, while the second and third also analyse measures of entrepreneurial performance, and the third disaggregates by gender.

The precise extent of self-employment indicated by the fifth sweep NCDS data from 1991 (NCDS5 hereafter) depends upon the exact definition that is chosen. Blanchflower and Oswald (1998) choose to define as self-employed the 1,279 (out of 11,369) individuals who indicate self-employment to be their main economic activity – some of whom are only part-time self-employed (but they exclude those part-time self-employed whose main economic activity category is not self-

employment). We follow the broader definition used by BFN (2000, 2002) – including some people for whom part-time self-employment is not their main economic activity – by adding those with self-employment as their main economic activity to those reporting a self-employment income whose economic activity category is specified, and is not 'sick or disabled'. This yields a total of 1,558 self-employed. The maximum sample size, for our analysis of the probability of choosing self-employment rather than being solely in some other form of economic activity, is 11,113. We include the unemployed and, for example, housewives – many of whom are often considered to be economically inactive. In principle, these individuals could make a decision to become self-employed – provided appropriate incentives are offered so that self-employment provides them with greater utility than any feasible alternative. This indicates a self-employment probability of 0.140 across males and females. The self-employment proportions given for 1991 in Blanchflower and Oswald (1998), calculated on quite a different basis, are very similar at 14.2% of employment within the NCDS cohort, and 15% across Great Britain.

The 11,113 individuals in our sample include 5,432 males and 5,681 females. The self-employment probability for males is 0.195, while that for females is only 0.088. However, this puts the share of self-employment for women at about 32.0% – which is above the 24.8% indicated in OECD (2000) for the whole UK across 1990-96; while the same source indicates similar female shares of self-employment across 1990-97 for Italy (23.4%), Sweden (25.7%), France (26.0%) and Germany (28.3%), but a larger share in the USA (37.0%). Parker and Robson (2004) report a male self-employment rate of 17.7% of the workforce for the UK in 1990, and 7.4% for females. Corresponding figures for other countries include 12.7% and 4.7% for Sweden; and 10.4% and 6.2% for the USA (indicating that the greater female share of self-employment there exists in the context of a rather low overall rate of self-employment).

Of course, national self-employment rates mask significant regional variation. Table I (below) shows in more detail how the self-employment probability differs by region in the NCDS5 data. Although our paper does not discuss how the UK North-South divide has developed over time,

Georgellis and Wall (2000) show that while self-employment has changed across time, regional variation has not fluctuated much. This poses a particular challenge for policy makers because altering regional differences in self-employment is perceived as a key component of advancing regional economic development. Thus, understanding the causes of these differences in self-employment is a key input to policy aimed at alleviating the North-South divide in England. The last four rows of the table calculate self-employment probabilities for composite regions comprising at least two postcode-based Standard Regions (SRs). This paper focuses particularly on Southern England (a combination of Greater London, South East England and South West England) and Northern England (a combination of Yorkshire & The Humber, North West England and the North of England SR). The table shows clearly that, for the NCDS5 data, there is a noticeably higher self-employment rate in Southern England than in Northern England – 23% against 17% among males, and 11% against 8% among females. It is also clear – as expected – that self-employment rate is substantially higher for males than females (about 19½% against 9%).

The NCDS5 data does not offer detailed information about the industry and occupation categories of the self-employed. However, examination of Labour Force Survey (LFS) data for 1991 demonstrates only modest differences in the industry and occupation distributions of self-employment between Northern England and Southern England. Since the LFS is sample-based, and covers the entire age range, it is far from straightforward to perform a reliable and useful projection onto NCDS cohort data: improved comparability comes at the expense of higher relative sampling variation.

{Table I near here}

Turning from self-employment rates to performance, we use a measure of employment by the self-employed which is also provided by NCDS5 – where each self-employed cohort member indicates how many employees he/she has. Some summary statistics are shown in Table II below. Table II deals with 1526 self-employed individuals that report a value (nil in the majority of cases) for job creation. The higher job creation rate by the self-employed in Northern England is an

interesting feature to emerge from the table. This is true both for males – 3.529 jobs per self-employed individual on average, compared to 2.652 jobs for Southern England – and for females (3.477 jobs per self-employed individual, versus 3.079). However, while average male job creation in all northern SRs is above that for any southern SR, the highest female job creation average is for Greater London (7.820), well ahead of the highest northern SR (North West England, at 4.538). Another insight offered by Table II is that there is less of a job creation gap between self-employed men and self-employed women than there is a gap between the gender-specific self-employment probabilities. Indeed, there is almost no difference by gender in the job creation rate for the self-employed of Northern England.

{Table II near here}

We now turn to factors likely to determine self-employment choice and performance. These NCDS variables are used in our estimation of self-employment probability logits, and/or job creation tobits. They are motivated by previous papers on self-employment using this dataset, (Blanchflower and Oswald (1998) and BFN (2000, 2002)), which themselves use the self-employment versus wage work theoretical framework developed by Evans and Jovanovic (1989). Thus, the variables are those likely to affect self-employment income or non pecuniary satisfaction relative to wage work. They include the following:

1. Ability, education and training – more able individuals are likely to secure higher income in both self-employment and wage work. Therefore, relative impact will determine ability's role in influencing self-employment. In terms of job creation by the self-employed, the more able are likely to be more successful – but there is an income effect, which may allow such individuals to try less hard and settle for a given performance level. We use dummies to indicate whether the highest academic qualification achieved is O level (or equivalent), A level, first degree or higher degree; up to four pairs of dummies capture performance in reading and maths tests at age seven (NCDS2) and age sixteen (NCDS3). For each test, a dummy is used to indicate a score definitively (not tied) in the top quintile of the cohort and another indicates a score in the bottom quintile –

- leaving the middle 60% (plus ties) of each ability distribution as the base case. A dummy captures apprenticeship by 1981; another denotes receipt of a vocational qualification by 1991.
- 2. *Non-cognitive attributes* self-employment is often associated with unique psychological characteristics but empirical support for these propositions is limited (see Parker, 2004). Therefore, several psychological measures are included as discrete scores. Creativity comes from NCDS1 (1965) a zero value denoting no creativity, and other values rescaled to a maximum of 0.4; while unforthcomingness, withdrawal, depression, anxiety acceptance and hostility towards (other) children are taken from NCDS2 (1969) each with a zero minimum; and caution, flexibility, moodiness, timidity, sociability and laziness measures are derived from NCDS3 (1974) varying in the range [–2,+2]. There is a dummy for fear of new situations (1974). A number of dummies indicate what the cohort member regarded, in 1981 (NCDS4), as being most important when choosing a job. Included are promotion, being in charge, being one's own boss, lack of responsibility, job security and good pay. Cohort members responding with some other job characteristic form the base group. This set of dummies serves to capture the individual's primary motivation across pecuniary and non-pecuniary dimensions emphasised in BFN (2000).
- 3. Family background included to capture the impact of role model, mentor and network effects on entrepreneurial choice and performance. A dummy reflects family financial difficulties (NCDS1); another denotes use of the English Language at home in 1969 (NCDS2); a series of dummies are used to indicate occupation of the cohort member's father in 1969 including employee manager of small firm, employee manager of large firm, professional self-employed, professional employee, foreman (manual work), skilled manual, worker with own account (a type of self-employment), farmer employee-manager and farmer with own account; two grouped variables from NCDS3 indicate the age at which the cohort member's father and mother left full-time education; another grouped variable indicates, for the cohort member's 1974 school, the percentage of male parents in a non-manual job.

- 4. Current family a dummy variable captures having no children by 1991; we also investigate the interaction of this dummy with higher level qualification (at least A level). Being childless may be relevant because people with children face extra obligations and, thus, time constraints. BFN (2002) found that both males, and less qualified females, with children are more likely to be self-employed (perhaps due to self-employment's potential flexibility in working time). That paper also found that highly qualified males with children, once self-employed, hired more workers than otherwise similar childless entrepreneurs. To capture exogenous finance, three variables are also constructed (NCDS5) to capture the size (linearly and quadratically¹) and timing (year) of any inheritance to capture liquidity constraints.
- 5. Region similarly to BFN (2002), we construct four regions from the 11 SRs of Great Britain, of which two (Southern England and Northern England) are our focus here. Even within the composite regions, there may be some variation in costs (particularly housing) and demand conditions. In our logits of self-employment, we include two SR dummies with South-West England SR being the base part of our Southern England composite region, and North of England SR being the base within our broader definition of Northern England. In the self-employment tobit equations, we include the average SR unemployment rate as a control, rather than pairs of SR dummies.
- 6. Aspects of self-employment we include a control for the length in years by 1991 of a spell of self-employment ongoing at NCDS5. We also use a dummy for non-full-time self-employed.
- 7. *Missing value dummies* for some individual regressors, and some groups of regressors, an extra dummy is used to indicate missing data, and as a (rather limited) control for this fact. This approach is quite common.

3. Empirical methodology

The well-known logit model provides a straightforward method of estimating the individual's probability of self-employment. We perform basic decomposition analysis to determine

whether North-South differences in self-employment are primarily a result of 'compositional' or 'structural' differences. As in previous literature, compositional differences reflect between-region differences (on average) in the characteristics of individuals; and structural differences refer to between-region differential responses to given characteristics.

The initial work on decomposition by Oaxaca (1973) and Blinder (1973) pertained to linear regressions of the logarithm of wages. However, since probit and logit specifications also typically involve a linear specification, Gomulka and Stern (1990) and Fairlie (1999, 2003) have noted that decomposition is feasible for these models. When appropriately transformed, estimated coefficients from probit and logit models are typically similar – which is unsurprising, given the shapes of the normal and logistic distributions that underlie the respective models. For our decompositions, a useful feature of the logit model is that the predicted probability of a given outcome is identical to the actual probability, not only for the whole sample, but also for sub-samples where a given dummy variable takes a particular value.

The basic format of the decomposition is as follows:

$$\overline{Y}_{S} - \overline{Y}_{N} = \left[\sum_{i=1}^{n_{S}} \frac{F(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{N}}\right] + \left[\sum_{i=1}^{n_{N}} \frac{F(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{N}} - \sum_{i=1}^{n_{N}} \frac{F(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{N})}{n_{N}}\right],$$
(1)

where \overline{Y} indicates the mean of the dependent variable in a logit model, F(.) is the Cumulative Distribution Function of the logistic distribution, the S and N subscripts indicate (throughout) Southern and Northern England respectively, n denotes the number of individuals in a particular area, X_i' is an individual's vector of characteristics and $\hat{\beta}$ is the vector of estimated coefficients from the logit model. The estimated coefficients vectors have the subscript S or N attached because they are generated through separate estimations of the logit model for those from Southern and Northern England. The version of the decomposition shown in equation (1) is split into a compositional effect (the term in the first square bracket) and a structural effect. The compositional term looks at the average predicted probability of self-employment that would be generated if individuals from Northern England responded according to the logit estimates for Southern England.

relative to the average predicted probability of self-employment in Southern England. The structural term measures the influence on the self-employment probability of the difference between the response coefficients for Southern England and Northern England, together with the impact of unobserved regional differences in characteristics (which, by definition, cannot be captured within the logit estimation – and, in our case, include industry and occupation category).

The following expression for the decomposition is equivalent:

$$\overline{Y}_{S} - \overline{Y}_{N} = \left[\sum_{i=1}^{n_{S}} \frac{F(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{N})}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{N})}{n_{N}} \right] + \left[\sum_{i=1}^{n_{S}} \frac{F(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{N})}{n_{S}} \right].$$

$$(2)$$

The interpretation of equation (2) is the opposite to equation (1), where now individuals from the buoyant South react to Northern coefficients.

Even and Macpherson (1990, 1993) noted that the decomposition component attributable to differences in characteristics can be split on a variable by variable basis (or for a group of variables, where this is more appropriate). The numerator for the required ratio is given by the size of the difference in sample means across the two groups (regions in our case) for the single variable, weighted by its estimated coefficient. The denominator is the difference in sample means across the two groups for *all* variables (each difference being weighted by the corresponding estimated coefficient). Hence the contribution to the probability gap by regressor r is as follows:

$$\left[\sum_{i=1}^{n_{S}} \frac{F(\boldsymbol{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F(\boldsymbol{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{S})}{n_{N}}\right] \left[\frac{(\overline{X}_{rS} - \overline{X}_{rN})\hat{\boldsymbol{\beta}}_{rS}}{(\overline{\boldsymbol{X}}_{S}^{\prime} - \overline{\boldsymbol{X}}_{N}^{\prime})\hat{\boldsymbol{\beta}}_{S}}\right].$$
(3)

Before we proceed to our estimation results, we discuss the appropriate decomposition. In its applications to wages, the question of what constitutes the 'no discrimination' distribution is often discussed. In our case too, we might expect that norm coefficients should be something other than either those for the South (as in equation (1)) or those for the North (equation (2)). While Reimers (1983) suggested the use of the arithmetic mean of the two sets of coefficients, and Cotton (1988) proposed the use of a weighted mean (the relative weights being determined by relative sample sizes) it has become more usual to consider the coefficients resulting from pooled estimation across

the groups under examination – as in Neumark (1988) and Oaxaca and Ransom (1994), where it is demonstrated that the coefficients from a pooled regression can be written as a weighted sum of the regression coefficients for the two groups. In the case of the logit model, the decomposition can be written as follows:

$$\left[\sum_{i=1}^{n_{S}} \frac{F\left(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{P}\right)}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F\left(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{P}\right)}{n_{N}}\right] + \left[\sum_{i=1}^{n_{S}} \frac{F\left(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{S}\right)}{n_{S}} - \sum_{i=1}^{n_{N}} \frac{F\left(\mathbf{X}_{iS}^{\prime} \hat{\boldsymbol{\beta}}_{P}\right)}{n_{S}}\right] + \left[\sum_{i=1}^{n_{S}} \frac{F\left(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{P}\right)}{n_{N}} - \sum_{i=1}^{n_{N}} \frac{F\left(\mathbf{X}_{iN}^{\prime} \hat{\boldsymbol{\beta}}_{P}\right)}{n_{N}}\right], (4)$$

where the *P* subscript refers to the pooled sample of Southern plus Northern England. The first term reflects the compositional effect (viewing the pooled coefficients as applying to all individuals). The second and third terms each reflect how the regional coefficients depart from the pooled norm.

Job creation by the self-employed is estimated using a Tobit model. This affords easier comparison with the previous work of BFN (2000, 2002) – and broadly similar statistical significance results to the negative binomial model for count data. Ordinary Least Squares (OLS) estimation is rejected on the basis that, while job creation might not be formally censored at zero, it seems highly unlikely that all cases of zero job creation by the self-employed (more than half of the group) reflect similar propensities to create jobs. If that suspicion is correct, OLS estimates are biased. An intuitively unappealing feature of the OLS fitted values for job creation is that some (in this case, 25-30%) are negative.

4. Estimation results

We use Limdep (Greene (2002)). Firstly, we consider the factors which influence the probability of an individual being self-employed and Table III, below, shows logit maximum likelihood estimates for males – in Southern and Northern England separately. The regressors in Table III are those remaining after a general-to-specific process based principally on at least some weak evidence of statistical significance in one or both regions². A number of differences are apparent, beyond the higher mean self-employment in Southern England shown in Table I. Some of these differences are shown by the simple means of the regressors in the third and sixth columns of

numbers within Table III. For example, among NCDS cohort males, a higher proportion reach first degree level in Southern England (16%, versus 10% in Northern England). Nor is this North-South divide on education confined to the NCDS generation itself – since the regressor means for the (grouped) variables on parental education are both noticeably higher for the South than the North.

{Table III near here}

Comparing our results to previous work on this dataset (see Table IV), the negative link between post-compulsory qualifications and the probability of male self-employment (BFN, 2002) shows clearly for Southern England – particularly for those with children – but, for the North, it is notable by its absence (especially for those with children). This may be a result of the less buoyant economy having fewer employment opportunities for the relatively well educated (since these opportunities would normally tend to draw them away from self-employment). There is a similar result for vocational qualifications, but apprenticeship has a positive association with male self-employment probability in both regions. The significant positive coefficient (Southern England) on the dummy identifying low reading ability at age 16 might indicate that these individuals have relatively poor employment prospects given the generally higher levels of education in the South and are pushed into self-employment.

{Table IV near here}

Creativity, found by BFN (2000) to be positively – though weakly – linked to self-employment across both genders, is shown above to be almost significant for males in Southern England only. We also find that 'relatively depressed at age 11' Northerners are more likely to take up self-employment. The fact that 'lazy at 16' and 'sociable at 16' are each positively linked to self-employment only in the South may indicate a North-South gap in ways of working and self-employment activities. There is a negative relation between 'timid at 16' and self-employment at 33, for Northern England only, and a lower mean for timidity compared to the South (perhaps a timid individual is more unsuitable for self-employment in the North). A desire to be one's own boss at 23 is associated, as expected, with generally higher self-employment – and there is also the expected

negative link of self-employment with the desire for job security (although a higher proportion of males in the North rated job security as the most important job characteristic in 1981 – when unemployment was particularly high, especially in the North).

Having a male parent who was the manager of a small firm is less common in Northern England, but this only appears to have a significant positive impact on self-employment in the North (in contrast to Blanchflower and Oswald (1998) and BFN (2000)). Parental education lasts about 0.3 years longer on average for each parent in the South, and its links with self-employment also suggests a North-South divide – with a weak positive link from father's education (only) in Southern England, and a stronger positive link with mother's education (only) in the North.

The effect of inheritance on the self-employment probability found by Blanchflower and Oswald (1998) and BFN (2000, 2002) shows evidence of a non-linear component. Only in the last paper are males considered separately. The effect of inheritance on self-employment there – a statistically strong positive linear effect and a rather weak negative quadratic effect, plus a weak link with timing whereby recent inheritance comes with less chance of self-employment – is altered when the regional dimension is considered. This paper finds a positive linear effect of inheritance for Southern England only, and no evidence of a quadratic effect. There is also a statistically significant link between recent inheritance and lower self-employment probability in the South, but no evidence of any effect of the magnitude or timing of inheritance for Northern England.

Decomposition results:

To calculate our decompositions, we have dropped all dummies to capture Standard Region (SR) or a wider region – using the rather severe assumption that, under the 'no discrimination' distribution, there is no purely spatial aspect to variation in the self-employment probability. For the logit model, predicted probabilities reflect actual frequencies precisely. Decompositions are shown in Table V, below – for five alternative approaches. The compositional term is negative in three out of five instances, equation (2) yielding the most notable exception. The choice of 'no discrimination' distribution clearly does matter. In each case, however, the compositional term is dwarfed by the

structural part(s) of the overall gap in self-employment probability – so male self-employment appears to be subject to a substantial North-South divide unaccounted for by the characteristics of individuals.

{Table V near here}

Notwithstanding the small compositional term found above, we now investigate – using equation (3) following Even and Macpherson (1993) – the contributions of the various categories of regressor laid out previously³ in Section 2. The second term in equation (3) can be used to break down the coefficient-weighted difference in regressor sample means, in accordance with each reported method of decomposition in turn. The results are shown in Table VI, below. Each row has six columns of numbers. Each of the first five give the numerator for the second term in equation (3) for that particular category of regressor. The last column (which is the sum of the first five) is the denominator for the second term in equation (3). The second row illustrates an important problem in the use of equation (3) – where the denominator of the second term is very small relative to four of the numerators. A general point should be noted about the relationship between rows 1, 2 and 4 in Table VI – namely, that the definition of the second term of equation (3) combines with the Reimers (1983) definition of the 'no discrimination' coefficients (as the simple arithmetic mean of the coefficient vectors for Southern and Northern England) so that each element in row 4 is equal to the simple average of the corresponding elements from rows 1 and 2. Similarly, there is a relationship between rows 1, 2 and 5 – whereby each element in row 5 is equal to the weighted average of the corresponding elements from rows 1 and 2 (as in the Cotton (1988) definition of the 'no discrimination' coefficients).

{Table VI near here}

One of the key features of Table VI is that the first category (ability, education and training) of regressors has a robust effect across four of the five forms of the decomposition – whereby the extra ability, education and training of an average individual in Southern England leads to a lower probability of self employment than in Northern England (probably through the extra opportunities

for employees with better education and training). However, the gap in family background characteristics (category 3) is shown in Table VI to have a robust effect in the opposite direction for the same four forms of decomposition – so that the type of family background enjoyed by the average individual in Southern England differs from that of his counterpart in Northern England in ways that, ceteris paribus, make self-employment more likely.

Female self-employment:

Table VII shows self-employment logit estimates for females – separately for Southern England and Northern England. In addition to the higher self-employment rate in Southern England already noted in Table I, the sample means shown in Table VII indicate some interesting regional differences – although the fact that there are fewer statistically significant regressors means that this table is shorter than Table III.

{Table VII near here}

As for males, females in the South are more highly educated on average than those in the North. For females, in contrast to males, there is little evidence of academic qualifications having an impact on self-employment. However, females in the South with A-levels as their highest qualification may be more likely to be in self-employment at age 33. No corresponding effect is evident for females in Northern England, although high reading ability at age 7 is associated with more likely self-employment in NCDS5.

We briefly summarise some other results on self-employment probability that differ from BFN (2000, 2002), or exhibit a North-South divide. Previous depression (positive) and timidity (negative) have impact only in the South. So too does the desire at age 23 to be one's own boss (positive) – this was not apparent in earlier work. The link between having a father working with his own account and subsequent self-employment of the child is now shown to stem from females in Northern England. The positive association between childlessness and female self-employment is significant only at the 10% level, and only disappears among highly qualified females for Northern

England. Magnitude of inheritance is only significant for females in the South – where the linear effect is statistically strong, but the quadratic (negative) effect is only significant at the 10% level.

Decomposition results:

The decompositions are again shown for the same five approaches. In contrast to the results for males, the compositional term for females is positive in every case – although the overall gap is of the same sign as it was for males. For equations (1) and (4), the compositional term accounts for about half the overall gap – and, for all five decomposition approaches, there is a greater relative importance for the regional differences in average characteristics in determining female probability of self-employment than was found for male self-employment probability. See Table VIII, below:

{Table VIII near here}

Viewing the respective balances of the compositional and structural elements in rows 1 and 2 as those for two opposite extremes in terms of the form of decomposition, the element balances for the other ('intermediate') forms of decomposition do fall in between. Although this ordering was widely assumed by previous authors, Oaxaca and Ransom (1994) gave a counter-example.

The breakdown of the coefficient-weighted difference in regressor sample means is shown in Table IX, below – although it should be noted that the effects of inheritance on female self-employment in Northern England (measured by statistically insignificant, but quite large coefficients) seem to distort the results for category 4 and the overall breakdown in rows 2, 4 and 5. One key difference between these results for females and the corresponding results for males in Table VI is the effect of regressors from category 1. Although, there is evidence of more education and training on average for females in Southern England, this has rather limited impact on the self-employment probability (and the effect is positive for most decompositions). There also seems to be more of a role for differences in non-cognitive characteristics than was evident for men. The effect of family background regressors from category 3, on the other hand, is rather smaller.

{*Table IX near here*}

Job creation by self-employed males:

Table X, below, shows estimates from censored (Tobit) regression for males – again, comparing Southern England and Northern England. Marginal effects can be readily calculated, via a scale factor (the Standard Normal Cumulative Distribution Function, evaluated - using the parameter estimates – at a chosen point, which is often the sample mean of the regressors). Among the male self-employed, job creation is positively associated with academic qualifications – but while having an A-level as highest is significant in the South, it is the more advanced first degree that is (weakly) significant for Northern England. Meanwhile, the positive effect of a professional qualification – found elsewhere in more aggregated samples – is preserved across the two separate regions. Unforthcomingness was found by BFN (2000) to have a statistically very weak negative link to job creation across all the self-employed. Table X illustrates that this relationship is statistically significant at the 1% level for males in Southern England. Timidity is negatively linked to job creation in the South – whereas, in the North, it is negatively linked to the self-employment probability. The classification of a job characteristic as most important does not appear to be a good indicator of self-employment job creation – although there is a very weak positive link for the promotion characteristic in the North. Having a father who was a professional employee is associated with greater job creation by self-employed males in Southern England, but not for the North (negative but insignificant here) – whereas BFN (2000) found a statistically weak positive effect aggregated across self-employed males and females. Another notable effect found in this family background category is the strong positive relationship between father's education and job creation in the South only.

The estimates on the inheritance regressors superficially appear different between the two regions – but the differences are not statistically significant. The positive linear coefficients is statistically significant at the 10% level for the South only. The negative quadratic estimate for the South is significant at the 10% level, while that for the North is insignificant. The other noticeable distinction is that the regressor means are quite different between the two regions: those for the

South are substantially higher, and given the greater gap for the quadratic regressor, this is an indication of some rather large inheritances having been received among the South's self-employed males. Unsurprisingly, for both regions, self-employed males that operate their business from home create fewer jobs than those based elsewhere.

{Table X near here}

5. Conclusions

This paper focuses on the North-South divide in England in self-employment and job creation by the self-employed, against the background of known differences in the regional economies of the two areas. The paper uses individual level data which so far has only been used to analyse UK self-employment at a national level. Males and females are studied separately, as this has been shown to be important previously. We also use decomposition analysis to clarify the distinction between regional structural effects and the effects of regional differences in individual characteristics. We have several new findings and implications.

For male self-employment, we find a negative effect of post-compulsory education only for Southern England – probably due to better employment opportunities there. Education levels are also higher in the south, so less educated individuals may be disadvantaged in the labour market and pushed into self-employment. We also find self-employment is associated with lower ability and motivation in the South. In the North, fewer job opportunities may push marginal well educated individuals into self-employment.

Blanchflower and Oswald (1998) note that inheritance is a good proxy for exogenous availability of finance – and should have a positive and significant effect on self-employment if finance constraints exist. Our results for Southern England include a positive effect of inheritance, and of the time since receipt – but neither of these effects appear for Northern England. This result is interesting because one would expect more finance availability in the South. The regional variation

we find may indicate that more business opportunities and competition among greater numbers of entrepreneurs for finance in the South means constraints are more likely there. Thus, the analysis uncovers some interesting regional differences compared to aggregate effects of inheritance on self-employment in the previous work of Blanchflower and Oswald (1998), and BFN (2000, 2002). Obviously, further research is needed but the results do raise issues of regional policy.

Decomposition analysis shows that higher male self-employment in Southern England occurs in spite of higher education, ability and training; and instead partly through an opposing effect resulting from differences in family background. These findings are inevitably missing in traditional aggregated studies of male self-employment.

Our results on male job creation indicate there are differences by region in which qualifications are associated with the creation of more jobs. Having a father who was a professional employee only seems to help job creation for those in the South; and job creation in this region (only) is also linked to paternal education.

For female self-employment, there is much less evidence of education having an impact – although A levels may influence self-employment in Southern England. Inheritance has no effect on self-employment for females in the North – another regional contrast. Higher female self-employment in Southern England is affected by both education and family background, but the situation is complicated by inheritance, which is only significant in the South.

The male results also give some new insights into why self-employment is higher in Southern England but the average entrepreneur creates more jobs in Northern England. Briefly, this seems to be due to differences in the regions and not the characteristics of the average individual in each region. Thus lower post-compulsory education in Northern England would actually generate higher levels of self-employment and similar job creation if the Northern economy was like the South.

We also find that total effects on self employment job creation differ in various ways between North and South. Thus, for example, first degrees have a negative total effect in the South by reducing self employment and not raising jobs per entrepreneur. In the North, however, first

degrees have a positive effect on job creation and no influence on self-employment, giving a positive total contribution. It follows that aggregate effects for the UK deduced in previous studies may be the net result of opposing tendencies in the different regions, and hence provide quite misleading guides for policy-makers.

This result can also explain why MVS (2006) find that UK regions do not match the pattern in Europe (Fritsch and Mueller, 2004), where direct job creation by new ventures is higher in high productivity/educated regions. Our results show that more highly educated regions do not necessarily have more highly educated self-employed sectors. Our use of individual level data and regional analysis show that since a lack of job opportunities can push well educated people into self-employment, then if push factors vary by region, less buoyant regions can have higher levels of education among the self employed than even regions with more highly educated workforces. Our results indicate that this appears to be the case in the UK, with the less educated North having higher direct job creation among the self-employed than the better educated South – as in MVS (2006).

In summary, the new results in this paper show that the explanations of self-employment prevalence and job creation are sufficiently different between North and South England as to require corresponding regional variation in enterprise policy – particularly in relation to education and finance. Our disaggregated analysis also indicates how unreliable predictions for the performance of regional self-employment are likely to be under the usual assumption of regional homogeneity. This is particularly important for education. For example, based on Southern estimates, lower post compulsory education in the North should boost the probability of Northern self-employment. Yet our regional analysis shows a different result – hypothetically higher post compulsory education in the North should not reduce self employment but could increase job creation by the self employed. The paper, therefore, highlights the importance of identifying differences across regional economies and provides new evidence that the North-South economic divide is not merely a traditional industrial phenomenon but also a hallmark of self-employment too.

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Table I: The self-employment probability by region – males and females separately

| | MALES | | |] | FEMALES | |
|-------------------------|-------|------|-------|-------|---------|-------|
| REGION | Total | S/E | Prob | Total | S/E | Prob |
| Greater London | 358 | 72 | 0.201 | 397 | 53 | 0.134 |
| South East England | 1302 | 274 | 0.210 | 1351 | 140 | 0.104 |
| South West England | 464 | 133 | 0.287 | 499 | 50 | 0.100 |
| East Anglia | 182 | 32 | 0.176 | 219 | 14 | 0.064 |
| East Midlands | 295 | 55 | 0.186 | 301 | 28 | 0.093 |
| West Midlands | 512 | 101 | 0.197 | 501 | 41 | 0.082 |
| Wales | 417 | 82 | 0.197 | 352 | 29 | 0.082 |
| Yorkshire & The Humber | 555 | 101 | 0.182 | 582 | 39 | 0.067 |
| North West England | 581 | 107 | 0.184 | 628 | 53 | 0.084 |
| North of England (NCDS) | 285 | 34 | 0.119 | 287 | 23 | 0.080 |
| Scotland | 462 | 64 | 0.139 | 529 | 25 | 0.047 |
| UNKNOWN | 19 | 4 | 0.211 | 35 | 4 | 0.114 |
| GREAT BRITAIN | 5432 | 1059 | 0.195 | 5681 | 499 | 0.088 |
| Southern England | 2124 | 479 | 0.226 | 2247 | 243 | 0.108 |
| Central England | 989 | 188 | 0.190 | 1021 | 83 | 0.081 |
| Northern England | 1421 | 242 | 0.170 | 1497 | 115 | 0.077 |
| Wales & Scotland | 879 | 146 | 0.166 | 881 | 54 | 0.061 |

Table II: Job creation by the self-employed by region – males and females separately

| | | MALES | |] | FEMALES | |
|-------------------------|------|-------|--------|------|---------|-------|
| REGION | Jobs | S/E | Mean | Jobs | S/E | Mean |
| Greater London | 164 | 67 | 2.448 | 391 | 50 | 7.820 |
| South East England | 762 | 265 | 2.875 | 265 | 139 | 1.906 |
| South West England | 302 | 131 | 2.305 | 80 | 50 | 1.600 |
| East Anglia | 119 | 32 | 3.719 | 23 | 13 | 1.769 |
| East Midlands | 315 | 55 | 5.727 | 61 | 27 | 2.259 |
| West Midlands | 305 | 100 | 3.050 | 60 | 40 | 1.500 |
| Wales | 179 | 82 | 2.183 | 155 | 29 | 5.345 |
| Yorkshire & The Humber | 398 | 101 | 3.941 | 109 | 37 | 2.946 |
| North West England | 349 | 106 | 3.292 | 236 | 52 | 4.538 |
| North of England (NCDS) | 100 | 33 | 3.030 | 41 | 22 | 1.864 |
| Scotland | 669 | 62 | 10.790 | 87 | 25 | 3.480 |
| UNKNOWN | 2 | 4 | 0.500 | 6 | 4 | 1.500 |
| GREAT BRITAIN | 3664 | 1038 | 3.530 | 1514 | 488 | 3.102 |
| Southern England | 1228 | 463 | 2.652 | 736 | 239 | 3.079 |
| Central England | 739 | 187 | 3.952 | 144 | 80 | 1.800 |
| Northern England | 847 | 240 | 3.529 | 386 | 111 | 3.477 |
| Wales & Scotland | 848 | 144 | 5.889 | 242 | 54 | 4.481 |

Table III: Male logits of the self-employment probability – South versus North

| | Sou | thern Englar | nd | Nor | thern Englai | nd |
|-----------------------------|-------------------------|--------------|--------|-------------------------|--------------|--------|
| VARIABLE | Estimate Est./S.E. Mean | | | Estimate Est./S.E. Mean | | |
| A level is highest | -0.912 | -3.71 | 0.121 | 0.152 | 0.47 | 0.080 |
| First degree is highest | -0.682 | -2.87 | 0.164 | -0.204 | -0.58 | 0.104 |
| Higher degree is highest | -1.385 | -2.60 | 0.028 | -1.113 | -1.39 | 0.019 |
| Professional qualification | 0.167 | 0.86 | 0.123 | 0.347 | 1.34 | 0.108 |
| Vocational qualification | -0.394 | -2.94 | 0.460 | -0.211 | -1.20 | 0.483 |
| Apprenticeship 1981 | 0.298 | 3.68 | 0.466 | 0.347 | 3.32 | 0.645 |
| Maths High Aged 7 | 0.047 | 0.30 | 0.201 | 0.292 | 1.41 | 0.182 |
| Maths Low Aged 7 | -0.300 | -1.53 | 0.107 | 0.139 | 0.59 | 0.142 |
| Reading High Aged 16 | -0.376 | -1.87 | 0.162 | 0.0003 | 0.001 | 0.121 |
| Reading Low Aged 16 | 0.505 | 2.67 | 0.097 | -0.387 | -1.52 | 0.141 |
| Creativity | 1.539 | 1.84 | 0.165 | 0.259 | 0.24 | 0.165 |
| Depression | -0.0003 | -0.01 | 0.877 | 0.145 | 3.04 | 0.986 |
| Caution | -0.139 | -1.55 | 0.178 | 0.062 | 0.54 | 0.173 |
| Laziness | 0.217 | 3.34 | -0.167 | 0.125 | 1.44 | -0.091 |
| Moodiness | 0.096 | 1.41 | -0.454 | 0.074 | 0.85 | -0.432 |
| Sociability | 0.175 | 2.49 | 0.449 | 0.058 | 0.66 | 0.404 |
| Timidity | -0.013 | -0.13 | 0.025 | -0.339 | -2.47 | 0.004 |
| Own boss important 1981 | 0.765 | 4.47 | 0.102 | 0.552 | 2.35 | 0.092 |
| Job security important 1981 | -0.406 | -2.70 | 0.217 | -0.376 | -1.94 | 0.289 |
| Dad manager of small firm | 0.258 | 1.46 | 0.114 | 0.652 | 2.35 | 0.067 |
| Dad professional employee | -0.158 | -0.55 | 0.056 | 0.678 | 1.62 | 0.033 |
| Dad worker own account | 0.268 | 1.02 | 0.038 | 0.606 | 1.44 | 0.025 |
| Dad farmer employee-manager | 0.732 | 1.50 | 0.010 | 1.866 | 2.69 | 0.007 |
| Dad farmer own account | 1.586 | 2.70 | 0.007 | 4.213 | 3.53 | 0.004 |
| Dad's years of education | 0.064 | 1.36 | 2.912 | -0.062 | -0.78 | 2.637 |
| Mum's years of education | 0.016 | 0.29 | 2.952 | 0.175 | 1.99 | 2.653 |
| No children | -0.349 | -2.43 | 0.353 | -0.046 | -0.23 | 0.274 |
| No children * higher quals | 0.486 | 1.77 | 0.142 | -0.368 | -0.85 | 0.072 |
| Inheritance | 0.082 | 3.15 | 0.687 | 0.179 | 1.42 | 0.243 |
| Inheritance squared | -0.023 | -0.61 | 0.184 | -0.480 | -1.01 | 0.018 |
| Year of inheritance | -0.423 | -2.62 | 0.268 | -0.190 | -0.74 | 0.192 |
| Sub-region 1 | -0.469 | -2.53 | 0.169 | 0.523 | 2.25 | 0.391 |
| Sub-region 2 | -0.438 | -3.28 | 0.613 | 0.551 | 2.37 | 0.409 |
| Constant | -1.294 | -4.12 | 1.000 | -3.066 | -6.98 | 1.000 |
| Log-likelihood | | -1023.262 | | -590.672 | | |
| Sample size | 2124 | | | 1421 | | |
| Mean of dependent variable | | 0.22552 | | | 0.17030 | |

Table IV: Self-employment probability – statistical significance of regressors

| VARIABLE | S. England | N. England | BFN (00) | BFN(02) | BO(98) |
|------------------------------------|------------|------------|----------|---------|--------|
| A level is highest | **** | + | * | | N/A |
| First degree is highest | **** | * | **** | **** | N/A |
| Higher degree is highest | **** | * | * | | N/A |
| Professional qualification | + | + | ++ | + | N/A |
| Vocational qualification | **** | * | ** | *** | N/A |
| Apprenticeship 1981 | ++++ | ++++ | ++++ | ++++ | N/A |
| Maths High Aged 7 | + | + | N/A | N/A | N/A |
| Maths Low Aged 7 | * | + | N/A | N/A | N/A |
| Reading High Aged 16 | ** | + | N/A | N/A | N/A |
| Reading Low Aged 16 | ++++ | * | N/A | N/A | N/A |
| Unforthcomingness | (+) | (+) | * | N/A | * |
| Hostility towards (other) children | (+) | (+) | * | N/A | + |
| Anxiety acceptance | (+) | (+) | ++++ | N/A | *** |
| Creativity | ++ | + | + | N/A | N/A |
| Depression | * | ++++ | N/A | N/A | N/A |
| Caution | * | + | N/A | N/A | N/A |
| Laziness | ++++ | + | N/A | N/A | N/A |
| Moodiness | + | + | N/A | N/A | N/A |
| Sociability | +++ | + | N/A | N/A | N/A |
| Timidity | * | *** | N/A | N/A | N/A |
| Own boss important 1981 | ++++ | +++ | ++++ | ++++ | N/A |
| Job security important 1981 | **** | ** | **** | **** | N/A |
| Dad manager of small firm | + | +++ | ++++ | N/A | ++++ |
| Dad professional employee | * | + | ++ | N/A | N/A |
| Dad worker own account | + | + | ++++ | N/A | ++++ |
| Dad farmer employee-manager | + | ++++ | ++++ | N/A | N/A |
| Dad farmer own account | ++++ | ++++ | ++++ | N/A | ++++ |
| Dad's years of education | + | * | N/A | N/A | N/A |
| Mum's years of education | + | +++ | N/A | N/A | N/A |
| No children | *** | * | N/A | *** | N/A |
| No children * higher quals | ++ | * | N/A | + | N/A |
| Inheritance | ++++ | + | ++++ | ++++ | ++++ |
| Inheritance squared | * | * | **** | * | * |
| Year of inheritance | **** | * | *** | ** | N/A |

Key and points to note:

- 1. Asterisks indicate a negative sign one for statistical insignificance at the 10% level, and an extra asterisk for significance at each of the 10%, 5% and 1% levels.
- 2. Plus signs indicate a positive sign. The number of plus signs follows the same rules as above.
- 3. Of the five columns summarising signs, columns 1, 2 and 4 refer to results for males only. Columns 3 and 5 aggregate males and females.
- 4. For three regressors, a plus sign has been enclosed in brackets since these regressors were excluded from our preferred specification.
- 5. N/A indicates the absence of that regressor.

6. Blanchflower and Oswald (1998) did also include 12 parental social class dummies, and reported results for a variety of probit specifications. BFN (2000, 2002) also used a probit model, whereas this paper employs the logit model.

Table V: Male logit – decomposition of the self-employment probability gap

| VERSION | Compositional Term | Structural 1 | Structural 2 | Overall gap |
|--------------|---------------------------|--------------|--------------|-------------|
| Equation (1) | 0.2255 - 0.2366 = -0.0111 | NIL | 0.0663 | 0.0552 |
| Equation (2) | 0.1789 - 0.1703 = +0.0086 | 0.0467 | NIL | 0.0552 |
| Equation (4) | 0.2035 - 0.2032 = +0.0003 | 0.0220 | 0.0329 | 0.0552 |
| Reimers | 0.1956 - 0.1968 = -0.0012 | 0.0299 | 0.0265 | 0.0552 |
| Cotton | 0.2004 - 0.2037 = -0.0033 | 0.0251 | 0.0334 | 0.0552 |



Table VI: Male logit of the self-employment probability – breakdown of term 2 in (3)

| VERSION | 1 | 2 | 3 | 4 | 7 | Overall |
|--------------|---------|--------|--------|---------|---------|---------|
| Equation (1) | -0.1610 | 0.0237 | 0.0392 | 0.0047 | 0.0217 | -0.0718 |
| Equation (2) | -0.0508 | 0.0015 | 0.1083 | -0.0487 | -0.0055 | +0.0048 |
| Equation (4) | -0.1075 | 0.0183 | 0.0647 | 0.0022 | 0.0160 | -0.0063 |
| Reimers | -0.1059 | 0.0126 | 0.0737 | -0.0220 | 0.0081 | -0.0335 |
| Cotton | -0.1168 | 0.0148 | 0.0669 | -0.0167 | 0.0108 | -0.0411 |

Table VII: Female logits of the self-employment probability – South versus North

| | Sou | thern Englan | d | Nor | thern Englan | d |
|-------------------------------|----------|--------------|-------|----------|--------------|-------|
| VARIABLE | Estimate | Est./S.E. | Mean | Estimate | Est./S.E. | Mean |
| O level equivalent is highest | 0.304 | 1.59 | 0.409 | 0.164 | 0.64 | 0.444 |
| A level is highest | 0.556 | 2.12 | 0.123 | -0.148 | -0.32 | 0.081 |
| First degree is highest | 0.297 | 1.11 | 0.149 | -0.260 | -0.56 | 0.101 |
| Professional qualification | 0.150 | 0.61 | 0.081 | 0.275 | 0.73 | 0.066 |
| Vocational qualification | -0.072 | -0.46 | 0.338 | -0.013 | -0.06 | 0.304 |
| Apprenticeship 1981 | 0.415 | 2.60 | 0.083 | 0.452 | 2.00 | 0.068 |
| Reading High Aged 7 | 0.188 | 1.06 | 0.212 | 0.518 | 2.19 | 0.219 |
| Reading Low Aged 16 | -0.306 | -1.00 | 0.093 | 0.445 | 1.43 | 0.145 |
| Depression | 0.108 | 2.07 | 0.701 | -0.014 | -0.17 | 0.687 |
| Caution | -0.174 | -1.53 | 0.152 | -0.009 | -0.06 | 0.229 |
| Flexibility | 0.065 | 0.61 | 0.186 | -0.254 | -1.72 | 0.217 |
| Timidity | -0.296 | -2.15 | 0.069 | -0.138 | -0.79 | 0.100 |
| Promotion important 1981 | -0.434 | -1.18 | 0.054 | 0.275 | 0.62 | 0.041 |
| Own boss important 1981 | 1.627 | 6.16 | 0.033 | 0.842 | 1.67 | 0.021 |
| Job security important 1981 | -0.357 | -1.38 | 0.123 | -0.299 | -0.94 | 0.145 |
| Family financial difficulties | -0.028 | -0.07 | 0.045 | -0.743 | -1.39 | 0.068 |
| Dad manager of small firm | 0.289 | 1.35 | 0.107 | 0.313 | 0.95 | 0.084 |
| Dad worker own account | 0.111 | 0.28 | 0.034 | 1.099 | 2.20 | 0.023 |
| Dad farmer employee-mngr | 1.342 | 2.27 | 0.007 | 0.925 | 1.12 | 0.008 |
| No children | -0.482 | -1.90 | 0.249 | -0.631 | -1.75 | 0.206 |
| No children * higher quals | 0.461 | 1.33 | 0.112 | 1.028 | 1.83 | 0.078 |
| Inheritance | 0.096 | 3.00 | 0.748 | 0.435 | 1.28 | 0.585 |
| Inheritance squared | -0.076 | -1.98 | 0.410 | -7.865 | -1.07 | 2.351 |
| Sub-region 1 | 0.247 | 1.09 | 0.177 | -0.188 | -0.66 | 0.389 |
| Sub-region 2 | 0.086 | 0.47 | 0.601 | 0.131 | 0.48 | 0.420 |
| Constant | -2.794 | -11.19 | 1.000 | -2.566 | -7.61 | 1.000 |
| Log-likelihood | -710.91 | | | | -378.53 | |
| Sample size | | 2247 | | | 1496 | |
| Mean of dependent variable | | 0.10814 | | | 0.07687 | |

Table VIII: Female logit – decomposition of the self-employment probability gap

| VERSION | Compositional Term | Structural 1 | Structural 2 | Overall gap |
|--------------|---------------------------|--------------|--------------|-------------|
| Equation (1) | 0.1081 - 0.0931 = +0.0150 | NIL | 0.0163 | 0.0313 |
| Equation (2) | 0.0828 - 0.0769 = +0.0059 | 0.0254 | NIL | 0.0313 |
| Equation (4) | 0.1013 - 0.0871 = +0.0142 | 0.0068 | 0.0102 | 0.0313 |
| Reimers | 0.0891 - 0.0810 = +0.0081 | 0.0190 | 0.0042 | 0.0313 |
| Cotton | 0.0915 - 0.0827 = +0.0088 | 0.0167 | 0.0059 | 0.0313 |

Table IX: Female logit of the self-employment probability – breakdown of term 2 in (3)

| VERSION | 1 | 2 | 3 | 4 | 7 | Overall |
|--------------|---------|--------|--------|---------|--------|---------|
| Equation (1) | 0.0479 | 0.0447 | 0.0062 | 0.0197 | 0.0174 | +0.1358 |
| Equation (2) | -0.0400 | 0.0504 | 0.0174 | -2.7111 | 0.0328 | -2.6506 |
| Equation (4) | 0.0222 | 0.0502 | 0.0090 | 0.0229 | 0.0310 | +0.1352 |
| Reimers | 0.0039 | 0.0475 | 0.0118 | -1.3457 | 0.0251 | -1.2574 |
| Cotton | 0.0127 | 0.0470 | 0.0107 | -1.0717 | 0.0235 | -0.9778 |



Table X: Male tobits of self-employment job creation – South versus North

| | Sou | thern Englan | d | Nor | thern Englan | d |
|-------------------------------|----------|--------------|--------|----------|--------------|--------|
| VARIABLE | Estimate | Est./S.E. | Mean | Estimate | Est./S.E. | Mean |
| O level equivalent is highest | 2.306 | 1.06 | 0.380 | 1.731 | 1.01 | 0.375 |
| A level is highest | 10.662 | 2.71 | 0.076 | -0.340 | -0.11 | 0.079 |
| First degree is highest | -3.263 | -0.79 | 0.110 | 5.888 | 1.62 | 0.092 |
| Professional qualification | 10.753 | 3.24 | 0.093 | 6.415 | 2.36 | 0.117 |
| Maths High Aged 7 | -6.362 | -2.38 | 0.184 | 0.545 | 0.27 | 0.204 |
| Reading High Aged 16 | 5.927 | 1.76 | 0.102 | 8.102 | 2.69 | 0.117 |
| Unforthcomingness | -1.647 | -2.63 | 1.238 | 0.155 | 0.39 | 1.446 |
| Anxiety acceptance | -1.104 | -1.08 | 0.408 | 1.112 | 1.64 | 0.492 |
| Caution | 3.694 | 2.53 | 0.022 | -1.685 | -1.37 | 0.088 |
| Moodiness | -1.778 | -1.74 | -0.268 | 0.362 | 0.51 | -0.325 |
| Timidity | -3.046 | -1.71 | -0.076 | 0.817 | 0.54 | -0.133 |
| Promotion important 1981 | -0.980 | -0.29 | 0.076 | 4.842 | 1.41 | 0.050 |
| Dad professional employee | 13.951 | 2.77 | 0.039 | -0.550 | -0.13 | 0.046 |
| Dad professional self-emp | -12.317 | -1.44 | 0.011 | 36.118 | 3.00 | 0.004 |
| Dad worker own account | -9.589 | -2.02 | 0.058 | -0.353 | -0.09 | 0.038 |
| Dad's years of education | 1.725 | 2.59 | 2.765 | 0.472 | 0.66 | 2.692 |
| No children | -0.083 | -0.03 | 0.289 | -0.573 | -0.28 | 0.254 |
| No children * higher quals | -5.734 | -1.20 | 0.089 | -1.641 | -0.37 | 0.058 |
| Inheritance | 2.029 | 1.88 | 1.290 | 3.038 | 1.52 | 0.397 |
| Inheritance squared | -15.497 | -1.84 | 0.651 | -33.526 | -1.11 | 0.044 |
| NCDS region unemployment rate | 0.095 | 0.06 | 6.527 | 0.157 | 0.19 | 10.995 |
| Years self-employed | 0.420 | 1.80 | 4.708 | 0.093 | 0.47 | 4.375 |
| Not full time | 1.121 | 0.40 | 0.168 | -1.255 | -0.60 | 0.225 |
| Operated from home | -9.096 | -4.67 | 0.516 | -9.040 | -5.33 | 0.483 |
| Constant | -11.688 | -1.05 | 1.000 | -5.071 | -0.49 | 1.000 |
| Sigma | -15.614 | 18.50 | N/A | 9.363 | 14.18 | N/A |
| Log-likelihood | -890.94 | | | | -473.77 | |
| Sample size | 463 | | | | 240 | |
| Mean of dependent variable | | 2.65227 | | | 3.52917 | |

¹ These enter in present value (1991) form. The linear term is divided by 10000 (yielding a mean, across all cases with specified region and gender, of 0.5321) and the quadratic term is divided by 1.0×10^{10} (which gives a mean across all cases of 0.4996). The timing control indicates the year in which the inheritance was received (subtracting 1900 from the year in question, and then dividing by 100)

year in question, and then dividing by 100).

The two region dummies for which estimates are reported near the foot of Table III refer respectively to the Greater London and South-East England SRs in the case on Southern England (the South-West England SR forming the base); and to Yorkshire and the Humber and the North-West England SR in the case of Northern England (the North of England SR being the base here).

³ Categories 5 (regions) and 6 (characteristics of self-employment) do not apply to our decomposition of the probability of self-employment.