

## Why climate sensitivity may not be so unpredictable?

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# Why climate sensitivity may not be so unpredictable?

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- Different explanations have been proposed as to why the range of climate
- sensitivity predicted by GCMs have not lessened substantially in the last decades,
- and subsequently if it can be reduced. One such study (Why is climate sen-
- sitivity so unpredictable?, Roe and Baker, 2007 [@]) addressed these questions
- using rather simple theoretical considerations and reached the conclusion that
- reducing uncertainties on climate feedbacks and underlying climate processes
- will not yield a large reduction in the envelope of climate sensitivity. In this
- letter, we revisit the premises of this conclusion. We show that it results from
- a mathematical artefact caused by a peculiar definition of uncertainty used
- by these authors. Applying standard concepts and definitions of descriptive
- statistics to the exact same framework of analysis as Roe and Baker, we show
- that within this simple framework, reducing inter-model spread on feedbacks
- does in fact induce a reduction of uncertainty on climate sensitivity, almost
- <sub>20</sub> proportionally. Therefore, following Roe and Baker assumptions, climate sen-
- 21 sitivity is actually not so unpredictable.

#### 1. Introduction

Uncertainties in projections of future climate change described in the last Assessment Report of the IPCC (IPCC, 2007 [@]) are high, as illustrated by the broad range of climate sensitivity – defined as the global mean temperature increase for a doubling of CO<sub>2</sub> – simulated by general circulation models (GCMs). Attempts to explain this fact have focused mainly on uncertainties in our understanding of the individual physical feedback processes (especially associated to clouds), difficulties to represent them faithfully in GCMs, nonlinearity of some processes and complex interactions among them giving rise to a chaotic behaviour of the climate system (Randall et al. [2007<sup>a</sup>]). A review of these explanations can be found in Bony et al., 2006 [@]. Nevertheless, in this letter, we leave aside all these considerations to focus our interest solely on the explanation proposed by Roe and Baker, 2007 [@] (RB07) which somewhat differ from the above-mentioned. This study uses the framework of feedback analysis, which has often been used to describe the relationship between physical processes involved in global warming and climate sensitivity (see for instance Lu and Cai, 2008 [@], Dufresne and Bony, 2008 [@], Soden and Held, 2006 [@]). The feedback analysis framework assumes a linear approximation of radiative feedbacks, resulting in a simple relationship between a global feedback gain f and climate 37 sensitivity  $\Delta T$ . In this classic setting, the main originality of RB07 approach consists in analyzing explicitly the way uncertainties on f, due to a limited understanding of their underlying physical processes, propagates into uncertainties on  $\Delta T$ : assuming f is a random variable with mean  $\bar{f}$  and standard deviation  $\sigma_f$ , RB07 uses this simple 41 probabilistic model to highlight several fundamental properties of uncertainty propagation

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from feedbacks to climate sensitivity. The most prominent conclusion of this analysis is
that reducing uncertainties on f does not yield a large reduction in the uncertainty of  $\Delta T$ , and thus that improvements in the understanding of physical processes will not yield

large reductions in the envelope of future climate projections. This conclusion, if true,

would clearly have crucial implications for climate research and policy.

In section 2, we revisit the premises of RB07 conclusion. We highlight that it is the result of a peculiar way of defining uncertainty. Moreover, we show in section 5 that this conclusion is a pure mathematical artefact with no connection whatsoever to climate. Since the basic question of uncertainty definition appears to be at stake, section 3 briefly recalls widely used definitions and elementary results on uncertainty and its propagation as they can be found in Descriptive Statistics textbooks. In section 4, we apply these standard concepts and definitions to the exact same framework of analysis as RB07. We show that within this simple framework, reducing inter-model spread on feedbacks does in fact induce a reduction of uncertainty on climate sensitivity, almost proportionally. Finally, section 6 concludes.

#### 2. Overview of RB07 approach

RB07 uses the feedback analysis framework. Denoting  $\Delta T_0$  the Planck temperature response to the radiative perturbation and f the feedback gain (referred to as feedback factor in RB07), they obtain:

$$\Delta T = \frac{\Delta T_0}{1 - f} \tag{1}$$

RB07 then assumes uncertainty on Planck response to be neglictible so that the entire spread on  $\Delta T$  results from the uncertainty on the global feedback gain f. To model

- this uncertainty, RB07 assumes that f follows a gaussian distribution with mean f, standard deviation  $\sigma_f$  and implicit truncation for f > 1 (implications of this truncation are discussed in appendix 1). Then, they derive an exact mathematical expression of the distribution of  $\Delta T$  through equation (1). This simple probabilistic climatic model is then used by RB07 to analyze the way uncertainties on f, due to a limited understanding of their underlying physical processes, propagates into uncertainties on  $\Delta T$ . Their analysis highlights two fundamental properties of uncertainty propagation:
- Amplification: The term in  $\frac{1}{1-f}$  in equation (1) amplifies uncertainty on feedbacks, all the more intensely as  $\bar{f}$  is close to (though lower than) one. Small uncertainties on feedbacks are thus converted in large uncertainties on the rise of temperature.
- Insensitivity: Quoting RB07, "reducing uncertainty on f has little effect in reducing uncertainty on  $\Delta T$ ", also stated as "the breadth of the distribution of  $\Delta T$  is relatively insensitive to decreases in  $\sigma_f$ ."
- We fully subscribe to the first property and elaborate further on it in section 4. However,
  we are puzzled by the second property, that is, the claimed insensitivity of uncertainty
  on  $\Delta T$  to uncertainty on feedbacks. The reason why one may find this second assertion a
  priori puzzling, is that it intuitively seems to be at a contradiction with the first property
  highlighted. Indeed, if small uncertainties on f are amplified into large uncertainties
  on  $\Delta T$ , it suggests that a strong dependency exists between both uncertainties, rather
  than no or little dependency. We therefore dig into the details of RB07 argumentation
  regarding this assertion. To get to that conclusion, it appears that RB07 actually focus
  on the probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  that  $\Delta T$  lies in the interval  $[4.5^{\circ}C, 8^{\circ}C]$  in

response to a sustained doubling of the  $CO_2$  concentration. This interval is defined as immediately above the range obtained with the CMIP3/AR4 GCMs (IPCC, 2007 [@]). They study graphically how this probability fluctuates with the level of uncertainty on feedbacks, by plotting for several values of  $\sigma_f$  the obtained cumulative distribution of  $\Delta T$ . Doing this graphical analysis, they observe that the probability of large temperature increase  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  is insensitive to  $\sigma_f$ . This observation is easily verifiable: we replicated RB07 cumulative distribution chart in figure 1c, and we computed several values of  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  for  $\bar{f} = 0.65$  and  $\sigma_f$  ranging from 0.10 to 0.20, finding it to fluctuate between 0.18 and 0.20. Therefore, in agreement with RB07, it is fair to say that the probability of large temperature increase (i.e.  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$ ) is quite insensitive to  $\sigma_f$  in this domain. However, concluding from this observation that "the breadth of the distribution of  $\Delta T$  is relatively insensitive to decreases in  $\sigma_f$ " and that "reducing uncertainty on f has little effect in reducing uncertainty on  $\Delta T$ " implicitly assumes two very different definitions of uncertainty: while on the side of feedback the uncertainty is measured by standard deviation  $\sigma_f$ , on the side of sensitivity the probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  is used as a metric of uncertainty. As will be developed in section 3, standard deviation is a standard, consensual uncertainty metric but the probability to lie 102 in a fixed interval is not. While under this peculiar double definition of uncertainty RB07 conclusion holds, it is fair to ask whether it would still hold with a different uncertainty 104 metric for  $\Delta T$ ; second, whether the probability to lie in a fixed interval can be considered 105 an acceptable measure of distribution breadth; and third, what are the implications of using such an asymetric definition of uncertainty. The following sections attempt to answer these questions.

### 3. Standard measurement and propagation of distribution spread

To investigate the first question, which relates to the basic issue of uncertainty definition,
we briefly recall a few standard definitions and concepts, as they can be found almost
identically in most Descriptive Statistics textbooks. For details, the reader can refer for
instance to Barlow, 1989 [@], Van der Vaart, 2000 [@], Reinard, 2006 [@], James and
Eadie, 2006 [@] to mention but a few such textbooks.

Descriptive Statistics primary purpose is to provide metrics summarizing a sample of 114 observations and similarly, in probabilistic terms, metrics summarizing the probability 115 density function (pdf) underlying them. Technically, the correspondence between both 116 is simply that a sample summary is an estimator (a function of the data) which esti-117 mates a distribution summary estimand (a parameter). In the present case, we study 118 continuous random variables thus we are rather concerned about pdf metrics than sample metrics, even though these pdfs actually aim at fitting a sample of observations, in that case CMIP3/AR4 GCMs simulations (Meehl et al. [2007]). Descriptive Statistics usually group metrics under three categories: location, scale and shape parameters. The 122 so-called location parameters are meant to identify the center of a distribution. Most 123 common location measures are mean, mode and median. The so-called scale parameters, 124 also referred to as dispersion, variability, variation, scatter or spread measures, describe 125 how far from the above-defined center possible values covered by the distribution tend 126 to be. This second group of metrics is the one we are interested in for our discussion, 127

as it is concerned with the measurement of distribution spread. Most common measures are standard deviation, interquartile range (IQR), range or median absolute deviation (MAD), more rarely full width at half maximum (FWHM). Variance and coefficient of 130 dispersion should also be mentioned though they are not expressed in the same unit as the 131 variable. Above mentioned references give complete mathematical expressions, properties, 132 strengths and limitations of these. We underline a property of particular interest to our 133 discussion: above mentionned measures of spread are invariant in location and linear in 134 scale. In other words, denoting S any particular measure of spread amongst those listed 135 above, X a random variable and Y = aX + b then: 136

$$S_Y = |a| . S_X \tag{2}$$

Further, in the general case of a dependency of the type  $Y = \phi(X)$ :

$$S_Y \simeq |\phi'(M_X)| . S_X \tag{3}$$

where  $\phi'$  represents the first derivative of  $\phi$  and M is a location parameter. This linear approximation is commonly used to combine errors on measurements, though generally in its multivariate formulation, and is thus sometimes referred to as the error propagation framework. It may also be used to study the way uncertainty on some input variable(s) propagates into uncertainty on an output obtained from a determinist function, as in section 4.

#### 4. Standard uncertainty propagation in RB07 feedback model

We now analyse the dependency between uncertainty on feedbacks and uncertainty on climate sensitivity in RB07 model. Denoting  $S_{\Delta T}$  a measure of climate sensitivity spread,

 $S_f$  a measure of feedback spread and  $M_f$  a measure of feedback location, the uncertainty propagation recalled in equation (3) can be applied straightforward to equation (1), leading to:

$$S_{\Delta T} \simeq \frac{\Delta T_0}{\left(1 - M_f\right)^2} \cdot S_f \tag{4}$$

Note that Equation (4) holds for any choice of pdf for feedback factor f and thus applies more generally than in the particular case of a truncated gaussian pdf chosen by RB07. Equation (4) also provides a simple relationship between  $S_{\Delta T}$ ,  $S_f$  and  $M_f$  which translates into the following two properties:

- Amplification: In agreement with RB07 first above recalled result, for a fixed level of feedback uncertainty  $S_f$ , the level of sensitivity uncertainty  $S_{\Delta T}$  is amplified when feedback  $M_f$  approaches one. Since estimates of feedback parameters in CMIP3/AR4 models (Soden and Held, 2006 [@], Randall et al. [2007a]) suggest  $M_f$  is close enough to one ( $M_f \simeq 0.65$ ) and hence yields subtantial amplification, it seems that "the climate system is operating in a regime in which small uncertainties in feedbacks are amplified in the resulting climate sensitivity uncertainty", to quote RB07.
- Proportionality: In disagreement with RB07 second above recalled result, for a fixed level of average feedback  $M_f$ , the level of climate sensitivity uncertainty  $S_{\Delta T}$  is proportional to the level of feedback uncertainty  $S_f$  ( $S_{\Delta T} \simeq 9.8 \ S_f$  for  $M_f \simeq 0.65$ ). This simple relationship between both uncertainties is intuitive. Indeed, when  $S_f = 0$ , feedbacks are determinists and  $\Delta T$  also is, considering no other source of uncertainty in the climate system, hence  $S_{\Delta T} = 0$ . As values of f get increasingly scattered, resulting values of climate sensitivity also get more scattered proportionally (figure 1a and 1b).

This proportionality has general validity in the sense that it holds for any above-recalled standard spread measure and for any distribution of f. However, it is an approximation for small values of  $S_f$ . We therefore find it relevant to investigate how this linear dependency is affected when  $S_f$  increases. To perform this analysis, we exhibit more precise results on uncertainty propagation in RB07 model. First, when spread is measured by IQR, an exact relationship holds for any value of  $S_f$  and any distribution of f (appendix 2):

$$S_{\Delta T} = \frac{\Delta T_0}{\left(1 - M_f\right)^2} S_f \cdot \left\{ 1 - \frac{w_f}{1 - M_f} S_f - \frac{1 - w_f^2}{4(1 - M_f)^2} S_f^2 \right\}^{-1}$$
 (5)

where  $w_f$  measures the asymetry of f distribution. Hence, when  $S \equiv IQR$ , the dependency between  $S_{\Delta T}$  and  $S_f$  is always overlinear when  $w_f \geq 0$ , eg when f has a symetric or right skewed distribution. When it is left skewed, the dependency is sublinear for small values of  $S_f$  but eventually becomes overlinear when  $S_f$  is large enough. Second, when spread is measured by standard deviation, a second order Taylor expansion of equation (1) leads to a more accurate approximation (appendix 3):

$$S_{\Delta T} \simeq \frac{\Delta T_0}{\left(1 - M_f\right)^2} S_f \cdot \left\{ 1 + \frac{2w_f}{1 - M_f} S_f + \frac{k_f - 1}{(1 - M_f)^2} S_f^2 \right\}^{\frac{1}{2}}$$
 (6)

Again, overlinearity prevails when  $w_f \ge 0$  or  $S_f$  large enough, which is connected to the convexity of the dependency between  $\Delta T$  and f. Third, when S is standard deviation and f distribution is log-normal, an exact formula holds for any  $S_f$ :

$$S_{\Delta T} = \frac{\Delta T_0}{\left(1 - M_f\right)^2} \cdot S_f \cdot \left\{ 1 + \left[ \frac{S_f}{1 - M_f} \right]^2 \right\} \tag{7}$$

and is again overlinear. Finally, overlinear relationships can also be derived when the distribution of f is assumed to be gamma or beta (equations (12) and (14) in appendix 4).

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To summarize the above discussion, its main outcome is rather intuitive and has actually few to do with climate: if the spread of feedback factor values decreases, the resulting spread of climate sensitivity values also decreases. Secondly, the dependency is as follows: it is linear for small feedback spreads and tends to get overlinear for larger values. Last, the proportionality coefficient in the dependency sharply increases as feedback intensifies.

#### 5. Properties of the probability to lie in a fixed interval

We now focus on whether the probability to lie in a fixed interval can be considered 196 an acceptable measure of distribution breadth, as implicitly done by RB07 to reach their 197 main conclusion. We approach this question very generally: let X be a continuous random 198 variable with location  $M_X$ , spread  $S_X$  and pdf  $p_X$ . Let [a,b] be a fixed interval near but 199 above the center  $(M_X < a < b)$ . Then, when  $S_X \to 0$  the variable becomes determinist 200  $(X = M_X)$  and it results that  $\mathbb{P}(X \in [a,b])$  equals to zero since  $M_X \notin [a,b]$ . When 201  $S_X \to +\infty$  the distribution covers such a wide range of values that the probability to 202 exceed any given threshold slowly increases towards 0.5 (figure 2b). In particular  $\mathbb{P}(X > X)$  $(a) \rightarrow 0.5$  and  $\mathbb{P}(X > b) \rightarrow 0.5$ , hence  $\mathbb{P}(X \in [a,b]) = \mathbb{P}(X > a) - \mathbb{P}(X > b) \rightarrow 0$ (appendix 5). Hence the dependency between  $\mathbb{P}(X \in [a,b])$  and  $S_X$  is characterized by a non monotonous function that increases, flattens and then decreases to zero (figure 2a). In light of this non monotonous dependency, it is difficult to hold  $\mathbb{P}(X \in [a,b])$ 207 as a valid measure for the width of X distribution. Further, the observed insensitivity of  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  to feedback spread  $S_f$ , which lead authors to their conclusion, 209 happens to proceed directly from the above described dependency: this flattening of the dependency is a pure mathematical artefact which systematically manifests under these definitions, and has nothing to do with climate.

Finally, if one still wants to stick to this peculiar, asymetric definition of uncertainty, it
has to be noted that in RB07 model, even though the dependency is flat in the domain  $S_f \in [0.1, 0.2]$ , the dependency is strong for  $S_f < 0.1$  when  $M_f \approx 0.65$  and subsequently
leads to a steep decrease of  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  to zero (figure 1d). In fact, since
feedback current estimates suggest  $S_f \simeq 0.09$  and  $M_f \approx 0.65$  (Soden and Held, 2006 [@],
Randall et al. [2007a]), the domain of strong dependency may actually already be reached
to date.

#### 6. Conclusion

Developments in section 5 suggest that, while the probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$ may be of interest practically, this metric is irrelevant to describe "the breadth of the distribution of climate sensitivity" which was RB07 explicit intent. To address this question, any measure of distribution spread chosen amongst those clasically used in Descriptive Statistics and recalled in section 3, appear to us more appropriate. With such measures of spread, we showed in section 4 that in RB07 framework, when the spread of feedback parameter  $S_f$  decreases, the resulting spread of climate sensitivity  $S_{\Delta T}$  values also decreases. Further, we also highlighted that in this framework, the decrease is approximately linear for  $S_f$  small and tends to be overlinear (i.e. to be steeper) for larger values of  $S_f$  owing to the convexity of the dependency between  $\Delta T$  and f.

Other than the definition issue discussed here, the relevance of RB07 simplified model to
describe the dependency between climate sensitivity and feedbacks may also be discussed

but this was beyond the scope of this article. In any case, if one holds this model to be
accurate, a decrease of the spread on feedback will lead to a decrease of the uncertainty
on climate sensitivity and a narrowing of the enveloppe of future climate projections. If
enough studies are undertaken to better understand and assess the physical processes
involved in the different feedbacks, neither are doomed to remain at their current level.

### **Appendix**

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#### 1 – Implications of the truncation

Since the linear feedback model of RB07 implicitly assumes  $f \leq 1$ , the gaussian distribu-238 tion  $\mathcal{N}(\bar{f}, \sigma_f)$  proposed by RB07 is implicitly truncated for f > 1 — otherwise equation 239 (1) would produce negative values of  $\Delta T$ . This truncation has several implications. First,  $\sigma_f$  (resp.  $\bar{f}$ ) does not exactly match standard deviation (resp. mean) of the truncated 241 distribution. For instance, when  $(\bar{f}, \sigma_f) = (0.75, 0.25)$  the standard deviation of f equals 242 0.18 and its mean equals 0.67. Second, it introduces some negative skewness in the distribution of f (-0.39 in the same example) which becomes more and more asymetric as  $\sigma_f$  and  $\bar{f}$  increases. Finally, since the truncated gaussian pdf is finite and non zero in the vicinity of f = 1, the obtained pdf of climate sensitivity behave as a Pareto distribution in  $\mathcal{O}(\Delta T^{-2})$  for high values, and hence does not have a finite mean, nor a finite variance. Hence, the truncated gaussian model of RB07 forbids the use of standard deviation as a measure of climate sensitivity spread, which explains the use of IQR in figure 1. For the purpose of RB07 which is to study climate sensitivity spread, assuming a parametric 250 distribution of f – such as log-normal, gamma or beta – which leads to finite mean and 251 deviation for sensitivity and exact mathematical expressions of the dependency between

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the deviation of  $\Delta T$  and the deviation of f (appendix 3), would be in our view more convenient. However, the results on the dependency between  $S_{\Delta T}$  and  $S_f$  presented in section 4 are general and also hold under RB07 gaussian assumption. Therefore, RB07 truncated gaussian is in our view mathematically unconvenient, but it does not affect uncertainty propagation: for a gaussian distribution just as for any other, the spread dependency is approximately linear for small spreads and overlinear otherwise, as equation (4) and (5) demonstrate and as figure 1b illustrates.

#### 2 – Exact uncertainty propagation equation for IQR

If X is a continuous random variable X, we denote  $X_{\alpha}$  its  $\alpha$ -quantile,  $S_X = X_{0.75} - X_{0.25}$  its interquantile range,  $M_X = X_{0.50}$  its median and  $w_X = \frac{X_{0.75} + X_{0.25} - 2X_{0.50}}{X_{0.75} - X_{0.25}}$  a dimensionless, quantile-based metric of asymetry. We thus have  $X_{0.75} = M_X + \frac{1}{2}S_X(1 + w_X)$  and  $X_{0.25} = M_X - \frac{1}{2}S_X(1 - w_X)$ . Since when  $\Phi$  is a diffeomorphism, we also have  $[\Phi(X)]_{\alpha} = \Phi(X_{\alpha})$ , hence from (1):

$$\begin{split} S_{\Delta T} &= \Delta T_{0.75} - \Delta T_{0.25} = \frac{\Delta T_0}{(1 - f_{0.75})} - \frac{\Delta T_0}{(1 - f_{0.25})} = \frac{\Delta T_0}{(1 - f_{0.75})(1 - f_{0.25})} \ S_f \\ &= \frac{\Delta T_0}{\left(1 - M_f\right)^2} \ S_f \cdot \left\{ 1 - \frac{w_f}{1 - M_f} \ S_f - \frac{1 - w_f^2}{4(1 - M_f)^2} \ S_f^2 \right\}^{-1} \end{split}$$

#### 3 – Second order term in uncertainty propagation equation

Assuming  $Y = \phi(X)$ , we analyse the way the approximation of the relationship between both spread measures  $S_Y$  and  $S_X$  is modified when a second order term is introduced in the Taylor development of  $\phi$  about  $M_X$ :

$$Y \simeq \phi(M_X) + \phi'(M_X)(X - M_X) + \frac{1}{2}\phi''(M_X)(X - M_X)^2$$
 (8)

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When the chosen spread measure S is standard deviation, calculations can be performed explicitly:

$$S_Y \simeq |\phi'(M_X)| . S_X . \left\{ 1 + \left[ \frac{\phi''(M_X)}{\phi'(M_X)} w_X \right] S_X + \left[ \frac{\phi''(M_X)^2}{4\phi'(M_X)^2} (k_X - 1) \right] S_X^2 \right\}^{\frac{1}{2}}$$
 (9)

Equation (9) shows that non linear terms in the resulting relationship between  $S_Y$  and  $S_X$ 275 depends on the shape of the distribution p(x) through its skewness  $w_X$  (a dimensionless 276 measure of assymetry) and kurtosis  $k_X$  (a dimensionless measure of peakedness), and on 277 the shape of function  $\phi$  through the curvature factor  $\frac{\phi''(M_X)}{\phi'(M_X)}$  (the rate of increase of the 278 slope in  $M_X$ ). A remarkable consequence of equation (9) is that when X distribution is 279 symetric  $(w_X = 0)$  and since kurtosis always exceeds one (Jensen inequality) hence the 280 dependency of  $S_Y$  to  $S_X$  is always over linear. Actually, sublinearity would require quite 281 special conditions: a distribution p(x) with low kurtosis and high skewness, simultaneously 282 with a function  $\phi$  characterized by strong curvature with sign opposite to skewness. 283

Applying equation (9) to model (1), it follows:

$$S_{\Delta T} \simeq \frac{\Delta T_0}{\left(1 - M_f\right)^2} S_f \cdot \left\{ 1 + \frac{2w_f}{1 - M_f} S_f + \frac{k_f - 1}{(1 - M_f)^2} S_f^2 \right\}^{\frac{1}{2}}$$
 (10)

## 4 – Exact uncertainty propagation equations for standard deviation

Since the domain of value of f in RB07 model is  $]-\infty,1]$ , we assume single tailed distributions defined on this support to avoid a truncation and make mathematical developments more convenients. For several usual distributions, the relationship between  $S_{\Delta T}$  and  $S_f$  can thus be explicited. Assuming a log-normal distribution with pdf  $\frac{1}{(1-f)\sigma\sqrt{2\pi}}\exp\left[-\frac{(\ln(1-f)-\mu)^2}{2\sigma^2}\right]$ , mean  $M_f=1-e^{\mu+\frac{\sigma^2}{2}}$  and variance  $S_f^2=e^{2\mu+\sigma^2}(e^{\sigma^2}-1)$  we

obtain  $S_{\Delta T}^2=\Delta T_0^2$  .  $e^{-2\mu+\sigma^2}(e^{\sigma^2}-1)$  . Recombining :

$$S_{\Delta T} = \frac{\Delta T_0}{\left(1 - M_f\right)^2} \cdot S_f \cdot \left\{ 1 + \left[\frac{S_f}{1 - M_f}\right]^2 \right\}$$
 (11)

Assuming a gamma distribution with pdf  $(1-f)^{k-1} \frac{\exp(-(1-f)/\theta)}{\Gamma(k)\theta^k}$ , mean  $M_f = 1 - \theta k$  and

variance  $S_f^2=\theta^2 k$ , we obtain  $S_{\Delta T}^2=\Delta T_0^2$ .  $[\theta^2(k-1)(k-2)]^{-1}$ . Recombining:

$$S_{\Delta T} = \frac{\Delta T_0}{(1 - M_f)^2} \cdot S_f \cdot \left\{ 1 - \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{-1} \cdot \left\{ 1 + \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{-\frac{1}{2}}$$
 (12)

Assuming a beta distribution with pdf  $\frac{\Gamma(2k)}{\theta\Gamma(k)^2} \left(1 - \frac{1-f}{\theta}\right)^{k-1} \left(\frac{1-f}{\theta}\right)^{k-1}$  on  $[1-\theta,1]$ , mean

 $M_f = 1 - \frac{\theta}{2}$  and variance  $S_f^2 = \theta^2 [8k+4]^{-1}$ , we obtain  $S_{\Delta T}^2 = \Delta T_0^2$ . [k(2k-1)].  $[\theta^2 (k-1)]$ 

 $(k-2)^{299}$  1)<sup>2</sup>(k-2)]<sup>-1</sup>. Recombining:

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$$S_{\Delta T} = \frac{\Delta T_0}{\left(1 - M_f\right)^2} \cdot S_f \cdot \left\{ 1 - \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{\frac{1}{2}} \left\{ 1 - 2 \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{\frac{1}{2}} \left\{ 1 - 3 \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{-1} \left\{ 1 - 5 \left[ \frac{S_f}{1 - M_f} \right]^2 \right\}^{-\frac{1}{2}}$$

$$\tag{13}$$

## 5 – Dependency between spread and probability weight of an interval

Assume  $X_1$  is a random real variable with pdf  $p_1(x)$ , cdf  $P_1(x)$ , center  $M_1$  and spread

 $S_1 > 0$ . Let [a, b] be a fixed interval near but above the center (eg  $M_1 < a$ ). For  $\lambda > 0$ ,

we introduce  $X_{\lambda} = \lambda(X_1 - M_1) + M_1$ , which has pdf  $\frac{1}{\lambda} p(\frac{x - M_1}{\lambda} + M_1)$ , cdf  $P(\frac{x - M_1}{\lambda} + M_1)$ ,

center  $M_1$  and spread  $\lambda S_1$ . To analyse the dependency between the probability of a real

variable to fall in [a, b] and the spread of its underlying distribution, we study  $F(\lambda; a, b) =$ 

 $\mathbb{P}(X_{\lambda} \in [a, b])$ . F can be expressed using the cdf of  $X_{\lambda}$ :

$$F(\lambda; a, b) = P(\frac{b - M_1}{\lambda} + M_1) - P(\frac{a - M_1}{\lambda} + M_1)$$

$$F(0; a, b) = P(-\infty) - P(-\infty) = 0 \quad \text{since } M_1 < a < b$$

$$F(+\infty; a, b) = P(M_1) - P(M_1) = 0$$
(14)

Since  $F(0;a,b)=F(+\infty;a,b)=0$ , and  $F\geq 0$ , then F reaches a maximum, and it

has the general pattern mentioned in the text. It is also straightforward to obtain that

F( $\lambda; a, b$ )  $\sim \frac{(b-a)p_1(M_1)}{\lambda^2}$  for large  $\lambda$ .

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Figure 1 – In all charts, f is truncated gaussian  $\mathcal{N}(M_f, \sigma_f)$  as in RB07. Upper left panel (a): 350 pdf of  $\Delta T$  with  $M_f = 0.65$  and  $\sigma_f = 0.20, 0.15, 0.10$ . Arrows represent the decreasing sensitivity 351 spread  $S_{\Delta T}$  obtained for decreasing values of  $\sigma_f$ . Upper right panel (b): climate sensitivity 352 spread  $S_{\Delta T}$  as a function of feedback spread  $S_f$ , for  $M_f = 0.60, 0.65, 0.70$ . Feedback spread  $S_f$  is 353 measured by standard deviation ( $\simeq \sigma_f$ ) but climate sensitivity spread  $S_{\Delta T}$  is measured by IQR 354 (see appendix 1 for explanation). Lower left panel (c): cdf of  $\Delta T$ . Arrows represent the stable 355 probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  obtained for decreasing values of  $\sigma_f = 0.20, 0.15, 0.10$ . Lower 356 right panel (d): probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  as a function of feedback spread  $S_f$ , spread 357 measured with IQR. 358

Figure 2 – X is centered gaussian with standard deviation  $S_X$ . Right panel: probability for X to exceed respectively 1 and 3, as functions of  $S_X$ . Left panel: probability for X to fall within interval [1,3] as a function of  $S_X$ .

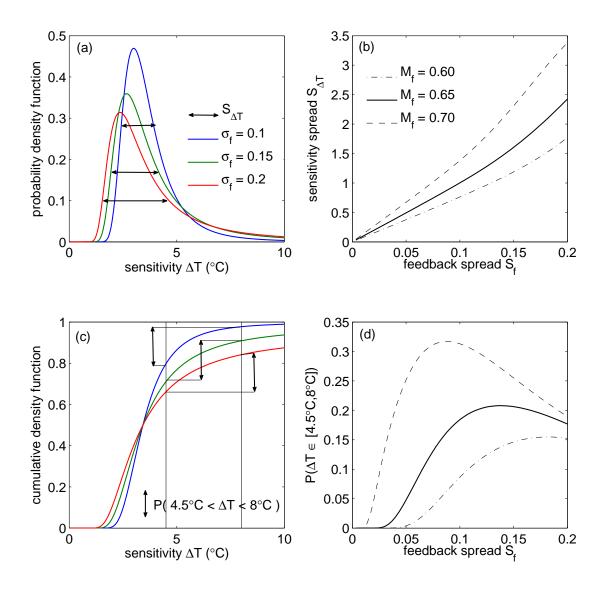
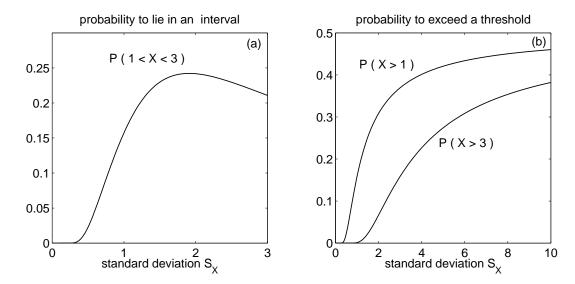


Figure 1. In all charts, f is truncated gaussian  $\mathcal{N}(M_f, \sigma_f)$  as in RB07. Upper left panel (a): pdf of  $\Delta T$  with  $M_f = 0.65$  and  $\sigma_f = 0.20$ , 0.15, 0.10. Arrows represent the decreasing sensitivity spread  $S_{\Delta T}$  obtained for decreasing values of  $\sigma_f$ . Upper right panel (b): climate sensitivity spread  $S_{\Delta T}$  as a function of feedback spread  $S_f$ , for  $M_f = 0.60$ , 0.65, 0.70. Feedback spread  $S_f$  is measured by standard deviation ( $\simeq \sigma_f$ ) but climate sensitivity spread  $S_{\Delta T}$  is measured by IQR (see appendix 1 for explanation). Lower left panel (c): cdf of  $\Delta T$ . Arrows represent the stable probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  obtained for decreasing values of  $\sigma_f = 0.20$ , 0.15, 0.10. Lower right panel (d): probability  $\mathbb{P}(\Delta T \in [4.5^{\circ}C, 8^{\circ}C])$  as a function of feedback spread  $S_f$ , spread measured with IQR.



**Figure 2.** X is centered gaussian with standard deviation  $S_X$ . Right panel: probability for X to exceed respectively 1 and 3, as functions of  $S_X$ . Left panel: probability for X to fall within interval [1,3] as a function of  $S_X$ .