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# On the long time behavior of the TCP window size process

Djalil CHAFAÏ, Florent MALRIEU, Katy PAROUX

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## Abstract

The TCP window size process appears in the modeling of the famous Transmission Control Protocol used for data transmission over the Internet. This continuous time Markov process takes its values in  $[0, \infty)$ , is ergodic and irreversible. It belongs to the Additive Increase Multiplicative Decrease class of processes. The sample paths are piecewise linear deterministic and the whole randomness of the dynamics comes from the jump mechanism. Several aspects of this process have already been investigated in the literature. In the present paper, we mainly get quantitative estimates for the exponential convergence to equilibrium, in terms of the  $W_1$  Wasserstein coupling distance, for the process and also for its embedded chain.

**Keywords.** Network Protocols; Queueing Theory; Additive Increase Multiplicative Decrease Processes (AIMD); Piecewise Deterministic Markov Processes (PDMP); Exponential Ergodicity; Coupling.

**AMS-MS.** 68M12 ; 60K30 ; 60K25 ; 90B18

## 1 Introduction

The TCP protocol is one of the main data transmission protocols of the Internet. It has been designed to adapt to the various traffic conditions of the actual network. For a connection, the maximum number of packets that can be sent at each round is given by a variable  $W$ , called the *congestion window size*. If all the  $W$  packets are successfully transmitted, then  $W$  is increased by 1, otherwise it is multiplied by  $\delta \in [0, 1)$  (detection of a congestion). As shown in [4, 9, 19], a correct scaling of this process leads to a continuous time Markov process, called general TCP window size process. This process  $X = (X_t)_{t \geq 0}$  has  $[0, \infty)$  as state space and its infinitesimal generator is given, for any smooth function  $f : [0, \infty) \rightarrow \mathbb{R}$ , by

$$L(f)(x) = f'(x) + x \int_0^1 (f(hx) - f(x))H(dh) \quad (1)$$

for some probability measure  $H$  supported in  $[0, 1)$ . This window size  $(X_t)_t$  increases linearly (this is the  $f'$  part of  $L$ ) until the reception of a congestion signal which forces the reduction of the window size by a multiplicative factor of law  $H$  or equal to  $\delta$  in the simplest case (this is the jump part of  $L$ ). The sample paths of  $X$  are deterministic between jumps, the jumps are multiplicative, and the whole randomness of the dynamics relies on the jump mechanism. Of course, the randomness of  $X$  may also come from a random initial value. The process  $(X_t)_{t \geq 0}$  appears as an Additive Increase Multiplicative Decrease process (AIMD), but also as a very special Piecewise Deterministic Markov

Process (PDMP). In this direction, [15] gives a generalization of the scaling procedure to interpret various PDMPs as the limit of discrete time Markov chains. In the real world (Internet), the AIMD mechanism allows a good compromise between the minimization of network congestion time and the maximization of mean throughput.

Our aim in this paper is to get quantitative estimates for the convergence to equilibrium of this general TCP window size process. This process  $X$  is ergodic and admits a unique invariant law, as can be checked using a suitable Lyapunov function (for instance  $V(x) = 1 + x$ , see e.g. [2, 3, 17] for the Meyn-Tweedie-Foster-Lyapunov technique). Nevertheless, this process is irreversible since time reversed sample paths are not sample paths and it has infinite support. This makes Meyn-Tweedie-Foster-Lyapunov techniques inefficient for the derivation of quantitative exponential ergodicity.

The *embedded chain*  $\hat{X}$  of the process  $X$  is the sequence of its positions just after a jump. It is an homogeneous discrete time Markov chain with state space  $[0, \infty)$ . As already observed in [4], it is also the square root of a first order auto-regressive process with non-Gaussian innovations and random coefficients. We obtain the following results concerning  $\hat{X}$ . We show first that it admits a unique invariant probability measure  $\nu$ , and that it converges in law to  $\nu$  given any (random) initial value  $\hat{X}_0$ . More precisely, using a coupling technique on trajectories, we prove an ergodic theorem of geometric convergence to equilibrium with respect to any Wasserstein distance. Then we provide non asymptotic concentration bounds, thanks to Gross's logarithmic Sobolev inequalities.

Similarly, the continuous time process  $X$  admits a unique invariant probability measure  $\mu$ , and converges in law to  $\mu$ , for any (random) initial value  $X_0$ . The reader may find explicit series for the moments of  $\mu$  and  $\nu$  in [9, 15, 16]. We obtain three main results that describe the way  $X$  converges to equilibrium. We first show the existence of a coupling with exponential decay. We use this result to prove an exponential ergodic theorem in term of Wasserstein distance. Eventually, we provide a uniform bound over the starting law that implies strong ergodicity. This kind of uniform estimates, though classical for processes on a compact set, is rather unusual for real valued processes.

The remainder of the paper is organized as follows. In the next preliminary section, we introduce some notations and give the statements of the main results. In section 3, we focus on the embedded chain  $\hat{X}$  and establish its convergence to equilibrium. The last section is devoted to the study of the continuous time process  $X$  and contains the proof of the results announced in section 2.

## 2 Notations and main results

Let us first explain how the trajectories of the process  $X$  may be constructed. The *jump rate* (or *jump intensity*) of  $X$  is given by  $\lambda(x) = x$  for every  $x \in [0, \infty)$ . If  $X_0 = x$  then the process will experience its first jump at a random time  $T$  solution of

$$\int_0^T \lambda(X_s) ds = E,$$

where  $E$  is an exponential random variable of unit mean. Since the trajectories of  $X$  are piecewise deterministic with slope 1, this is nothing else but

$$\int_0^T \lambda(x + s) ds = E,$$

which leads to  $T = \sqrt{x^2 + 2E} - x$ . Then, the sample paths of the process  $X$  generated by (1) may be constructed recursively as follows. Let  $X_0$  be its non-negative random initial position,  $(E_n)_{n \geq 1}$  be a sequence of i.i.d. exponential random variables of unit mean, and  $(Q_n)_{n \geq 1}$  be a sequence of i.i.d. random variables of law  $H$ . Assume that  $X_0$ ,  $(E_n)_{n \geq 1}$  and  $(Q_n)_{n \geq 1}$  are independent. We define by induction the jump times  $(T_n)_{n \geq 1}$  and the positions just after the jumps  $(X_{T_n})_{n \geq 1}$  as

$$T_n = T_{n-1} + \sqrt{X_{T_{n-1}}^2 + 2E_n} - X_{T_{n-1}} \quad \text{and} \quad X_{T_n} = Q_n \sqrt{X_{T_{n-1}}^2 + 2E_n}. \quad (2)$$

If we set  $T_0 = 0$ , then for every  $n \geq 0$  and  $t \in [T_n, T_{n+1})$ , we have  $X_t = X_{T_n} + t - T_n$  and in particular,  $X_{T_n} = Q_n X_{T_n^-}$ . For every  $t \geq 0$ , one can also write the series representation

$$X_t = \sum_{n=0}^{\infty} (X_{T_n} + t - T_n) \mathbb{1}_{[T_n, T_{n+1})}(t).$$

The sequence  $\hat{X} = (X_{T_n})_{n \geq 0}$  is the *embedded chain* of  $X$ . According to (2), this discrete time Markov chain with state space  $[0, \infty)$  satisfies the recursion

$$\hat{X}_{n+1}^2 = Q_{n+1}^2 (\hat{X}_n^2 + 2E_{n+1}). \quad (3)$$

Thus, the embedded chain  $\hat{X}$  is the square root of a first order auto-regressive process with non-Gaussian innovations  $(2Q_n^2 E_n)_{n \geq 1}$  and random coefficients  $(Q_n^2)_{n \geq 1}$ , as already observed in [4]. The embedded chain  $\hat{X}$  is homogeneous, and its transition kernel  $K$  is given, for any  $x \geq 0$  and every bounded measurable  $f : [0, \infty) \rightarrow \mathbb{R}$ , by the formula

$$K(f)(x) = \int_0^{\infty} f(y) K(x, dy) = \mathbb{E} \left[ f \left( Q \sqrt{x^2 + 2E} \right) \right] \quad (4)$$

where  $E$  is an exponential random variable of unit mean and  $Q$  is a random variable of law  $H$  independent of  $E$ . We show in section 3 that the embedded Markov chain  $\hat{X}$  admits a unique invariant probability measure  $\nu$ , and converges in law to  $\nu$  given any (random) initial value  $\hat{X}_0$ . Similarly, the continuous time process  $X$  admits a unique invariant probability measure  $\mu$ , and converges in law to  $\mu$ , for any (random) initial value  $X_0$ . We recall that explicit series for the moments of  $\mu$  and  $\nu$  can be found in [9, 15, 16].

Despite the apparent simplicity of the dynamics (1), the quantitative study of the long time behavior of  $X$  is not easy, mainly because the jump rate depends on the position  $x$  of the process. Our strategy is to couple two trajectories starting at two different points in such a way that they get closer and closer. It seems difficult to stick the two trajectories in order to get total variation estimates since the sample paths are parallel between jump times. Thus, we provide quantitative bounds in terms of the Wasserstein coupling distance. Recall that for every  $p \geq 1$ , the  $W_p$  Wasserstein distance between two laws  $\mu$  and  $\nu$  on  $\mathbb{R}$  with finite  $p^{\text{th}}$  moment is defined by

$$W_p(\mu, \nu) = \left( \inf_{\Pi} \int_{\mathbb{R}^2} |x - y|^p \Pi(dx, dy) \right)^{p^{-1}} \quad (5)$$

where the infimum runs over all coupling of  $\mu$  and  $\nu$ . In other words,  $\Pi$  runs over the convex set of laws on  $\mathbb{R}^2$  with marginals  $\mu$  and  $\nu$ , see e.g. [20, 21]. It is well known that for

any  $p \geq 1$ , the convergence in  $W_p$  Wasserstein distance is equivalent to weak convergence together with convergence of all moments up to order  $p$ .

The jump part of  $L$  ensures that the process will remain essentially in a compact set. The jumps act in a way like a confining potential. On the other hand, the jump rate is small when the process is close to the origin. This prevents the decay of the Wasserstein distance to be exponential for small times (see the remark below for more details).

In section 3 we first establish the following geometric convergence to equilibrium of the embedded Markov chain  $\hat{X}$  for any Wasserstein distance.

**Theorem 2.1** (Wasserstein exponential ergodicity for the generic embedded chain). *Let  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  be two processes generated by (1). Assume that  $\mathcal{L}(X_0)$  and  $\mathcal{L}(Y_0)$  have finite  $p^{\text{th}}$  moment for some real  $p \geq 1$ . Let  $\hat{X}$  and  $\hat{Y}$  be the embedded chains of  $X$  and  $Y$ . Then, for any  $n \geq 0$ , with a random variable  $Q \sim H$ ,*

$$W_p(\mathcal{L}(\hat{X}_n), \mathcal{L}(\hat{Y}_n)) \leq \mathbb{E}(Q^p)^{n/p} W_p(\mathcal{L}(X_0), \mathcal{L}(Y_0)).$$

In particular, if  $\nu$  is the invariant law of  $\hat{X}$  then

$$W_p(\mathcal{L}(\hat{X}_n), \nu) \leq \mathbb{E}(Q^p)^{n/p} W_p(\mathcal{L}(X_0), \nu).$$

We also establish in section 3 non asymptotic concentration bounds in the ergodic theorem by using Gross logarithmic Sobolev inequalities:

**Theorem 2.2** (Gaussian deviations for the ergodic theorem for the embedded chain). *Let  $\hat{X}$  be the embedded chain associated to (1) and starting from  $\hat{X}_0 = x \geq 0$ . Assume that  $H$  is the Dirac mass at point  $\delta \in (0, 1)$ . Then for any  $u \geq 0$  and any 1-Lipschitz function  $f : [0, \infty) \rightarrow \mathbb{R}$ ,*

$$\mathbb{P} \left( \left| \frac{1}{n} \sum_{k=1}^n f(\hat{X}_k) - \int f d\nu \right| \geq u + \frac{\delta}{1-\delta} W_1(\delta_x, \nu) \right) \leq 2 \exp \left( -\frac{n(1-\delta^2)u^2}{2\delta^2} \right).$$

The exponential convergence to equilibrium of the continuous time process  $X$  is addressed in section 4. The idea is to exhibit a *coupling*, i.e. a Markov process on  $[0, \infty)^2$  for which the marginal components are generated by (1), with prescribed initial laws. The infinitesimal generator  $\bar{L}$  of this coupling is defined for every smooth  $f : [0, \infty)^2 \rightarrow \mathbb{R}$  by

$$\bar{L}(f)(x, y) = \text{div}(f)(x, y) + x \int_0^1 (f(hx, hy)p(x, y) + f(hx, y)q(x, y) - f(x, y)) H(dh) \quad (6)$$

if  $x \geq y$  and

$$\bar{L}(f)(x, y) = \text{div}(f)(x, y) + y \int_0^1 (f(hx, hy)p(x, y) + f(x, hy)q(x, y) - f(x, y)) H(dh)$$

if  $x \leq y$ , where  $\text{div}(f) = \partial_1 f + \partial_2 f$  and where  $p(x, y)$  and  $q(x, y) = 1 - p(x, y)$  are chosen to recover the prescribed marginals (see section 4). This coupling is the only one such that the lower component never jump alone. Let us give the pathwise interpretation of this coupling. All the heuristic statements below are made more precise hereafter. The positions of both “components” increase linearly with slope 1. The jump rate is an increasing function of the position. Thus, “the higher a component is, the sooner it will jump”. The dynamics of the couple of components is as follows:

1. After an “appropriate” time which depends only on the initial position of the upper component, this one jumps.
2. Simultaneously, the other one “tosses an appropriate coin” whose probability of success depends on the positions on the two components to decide whether or not it jumps too.
3. In the case of joint jumps, both components use the same multiplicative factor.
4. Then, we repeat these three first steps again and again. . .

This coupling provides the following quantitative exponential upper bound.

**Theorem 2.3** (Existence of a coupling with exponential decay). *Given  $x, y \geq 0$ , there exists a Markov process  $(X, Y) = ((X_t, Y_t))_{t \geq 0}$  on  $[0, +\infty)^2$  such that  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  are generated by (1), with  $X_0 = x$  and  $Y_0 = y$ , and for every  $t \geq 0$ ,*

$$\mathbb{E}(|X_t - Y_t|) \leq \min(1, c|x - y|e^{-\rho t}).$$

In the general case, these constants can be chosen as

$$\rho = \frac{1}{2} \mathbb{E} \left( \frac{1 - Q}{1 + Q} \right) \quad \text{and} \quad c = 2 \tag{7}$$

where  $Q \sim H$ . In the particular case when  $H$  is the Dirac mass at point  $\delta \in [0, 1)$ , these constants can be improved as follows:

$$\rho = \frac{2 \sinh(1/2) \sqrt{1 - \delta}}{\sqrt{(e - 1)\delta + 1} + 1} \geq \frac{1 - \delta}{2} \tag{8}$$

and

$$c = \frac{e - 1}{\sqrt{(e - 1)\delta + 1} + 1} + 1 \leq \frac{e + 1}{2}. \tag{9}$$

In particular, if  $H = \delta_{1/2}$  then (8) and (9) become  $\rho \simeq 0.31$  and  $c \simeq 1.73$ .

If  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  are two processes generated by (1) and if  $\Pi$  is a coupling of  $\mathcal{L}(X_0)$  and  $\mathcal{L}(Y_0)$ , we deduce from Theorem 2.3 that for every  $t \geq 0$ ,

$$\begin{aligned} W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) &\leq \int_{[0, \infty)^2} \mathbb{E}(|X_t - Y_t| | X_0 = x, Y_0 = y) \Pi(dx, dy) \\ &\leq ce^{-\rho t} \int_{[0, \infty)^2} |x - y| \Pi(dx, dy). \end{aligned}$$

Taking the infimum over  $\Pi$  provides the following Theorem, which can be seen as a main result of the present paper:

**Theorem 2.4** (Wasserstein exponential ergodicity). *There exist explicit constants  $c \geq 1$  and  $\rho > 0$ , see (7) or (8-9), such that for any processes  $(X_t)_{t \geq 0}$  and  $(Y_t)_{t \geq 0}$  generated by (1) with finite first moment at initial time, and for any  $t > 0$ , we have*

$$W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq \min(1, ce^{-\rho t}) W_1(\mathcal{L}(X_0), \mathcal{L}(Y_0)).$$

In particular, when  $Y_0$  follows the invariant law  $\mu$  of (1), we get for every  $t \geq 0$

$$W_1(\mathcal{L}(X_t), \mu) \leq \min(1, ce^{-\rho t}) W_1(\mathcal{L}(X_0), \mu).$$

The jump rate  $\lambda$  goes to infinity at infinity. The following theorem, proved in section 4, shows that the convergence to equilibrium is in fact uniform over the starting laws, as it could be for a process living in a compact set.

**Theorem 2.5** (Strong ergodicity). *Let  $c$  and  $\rho$  be the constants given by (7) or (8-9), and let  $d = \sqrt{1 - \mathbb{E}(Q)}$  where  $Q$  is a random variable of law  $H$ . Then, for two processes  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  generated by (1) with arbitrary initial laws  $\mathcal{L}(X_0)$  and  $\mathcal{L}(Y_0)$  and for every  $t$  and  $s$  such that  $t > s > 0$ , one has*

$$W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq \frac{2ce^{\rho s}}{d \tanh(ds)} e^{-\rho t}.$$

Theorem 2.5 provides in particular a uniform bound in  $N \in (0, \infty)$  if  $X_0 = 0$  and  $Y_0 = N$ . This kind of uniform estimates are classical for processes on a compact set but rather unusual for real valued ones.

**Remark 2.6** (Role of the drift term). *Let us consider the continuous time Markov process  $Z = (Z_t)_{t \geq 0}$  with state space  $[0, \infty)$ , initial condition  $Z_0 = 1$ , and infinitesimal generator*

$$Af(z) = z(f(z/2) - f(z)).$$

*We denote by  $(T_n)_{n \geq 0}$  the sequence of jump times of  $Z$ , with the convention  $T_0 = 0$ . The holding times are given by  $\tau_n = T_n - T_{n-1}$  for every  $n \geq 1$ . The process  $Z$  admits a unique invariant law which is  $\delta_0$ . For any  $n \geq 1$ , we have  $Z_{T_n} = 2^{-n}$  and  $\tau_n$  follows an exponential law of mean  $2^{n-1}$ . Thus, for any  $n \geq 1$  and  $t > 0$ , we can write*

$$\mathbb{P}(Z_t \geq 2^{-n}) = \mathbb{P}(T_n > t) \geq \mathbb{P}(\tau_n \geq t) = e^{-2^{-(n-1)}t}.$$

*By choosing  $t = 2^{n-1}$  and by using Markov's inequality, we obtain*

$$\lim_{n \rightarrow \infty} 2^{n-1} \mathbb{E}(Z_{2^{n-1}}) \geq (2e)^{-1} > 0.$$

*As a consequence,  $t \mapsto t\mathbb{E}(Z_t)$  does not converge to 0 and  $Z_t$  converges to 0 as  $t$  tends to  $\infty$  with an at most linear speed. Of course, this dynamics without drift is very different from the TCP dynamics (1) since it is neither irreducible nor recurrent. This elementary example without drift shows that the drift term in (1) is crucial for the exponential decay phenomenon (after some "warmup"). The drift keeps the process away from zero, i.e. away from a slow zone (rare jumps).*

## Open questions and further remarks

The determination of the best constants in Theorem 2.3 remains an open problem. One may also ask for a version involving  $W_p$  for any  $p \geq 1$  or even the total variation distance. Beyond the TCP window size dynamics, one may ask about the speed of convergence of ergodic PDMPs, for which necessary and sufficient ergodicity criteria are already known, see e.g. [3]. One may also study the long time behavior of interacting processes associated to (1) or (13), for instance Mac Kean-Vlasov mean field interactions as in [7].

### 3 Embedded chain

It is shown in [4, Proposition 8], by Laplace transform inversion, that if  $H$  is a Dirac mass at point  $\delta \in (0, 1)$ , the invariant measure of the embedded chain  $\nu = \nu_\delta$  has Lebesgue density

$$x \geq 0 \mapsto \frac{1}{\prod_{n=1}^{\infty} (1 - \delta^{2n})} \sum_{n=1}^{\infty} \frac{(-1)^{n-1} \delta^{-2n}}{\prod_{k=1}^{n-1} |1 - \delta^{-2k}|} x e^{-\delta^{-2n} x^2 / 2}. \quad (10)$$

It is unimodal, of order  $O(x \exp(-\delta^2 x^2 / 2))$  when  $x \rightarrow \infty$ , and all its derivatives vanish at  $x = 0$ .

If  $H$  is not a Dirac mass, the invariant measure  $\nu$  of the embedded Markov chain is no longer explicit. Nevertheless, the recursion formula (3) provides the following result, see [5, 6], which establish the existence of an invariant measure with sub-Gaussian tails.

**Theorem 3.1** (Convergence of the embedded chain, [5, 6]). *Given any  $\hat{X}_0$ , the embedded Markov chain  $\hat{X} = (\hat{X}_n)_{n \geq 0}$  associated to the dynamics (1) converges in distribution to the law of the random variable*

$$\left( 2 \sum_{n=1}^{\infty} Q_1^2 \cdots Q_n^2 E_n \right)^{1/2}$$

which is a.s. finite, where  $E_1, E_2, \dots$  and  $Q_1, Q_2, \dots$  are independent sequences of i.i.d. random variables following respectively the exponential law of unit mean and the law  $H$  which appear in (1). In particular,  $\nu$  is the unique invariant law of  $\hat{X}$  and

$$\int e^{sx^2} \nu(dx) = \mathbb{E} \left( \frac{1}{\prod_{n=1}^{\infty} (1 - 2sQ_1^2 \cdots Q_n^2)} \right),$$

which is finite if  $2sq^2 < 1$  and infinite if  $2sq^2 > 1$ , where  $q = \inf \{x, \mathbb{P}(Q > x) = 1\} \leq 1$ .

Let us now turn to our quantitative estimate for the convergence to equilibrium for the embedded chain.

*Proof of Theorem 2.1.* It is sufficient to provide a good coupling. Let  $x \geq 0$  and  $y \geq 0$  be two non-negative real numbers, and let  $(E_n)_{n \geq 1}$  and  $(Q_n)_{n \geq 1}$  be two independent sequences of i.i.d. random variables with respective laws the exponential law of unit mean and the law  $H$  which appears in (1). Let  $\hat{X}$  and  $\hat{Y}$  be the discrete time Markov chains on  $[0, \infty)$  defined by

$$\begin{aligned} \hat{X}_0 = x \quad \text{and} \quad \hat{X}_{n+1} &= Q_{n+1} \sqrt{\hat{X}_n^2 + 2E_{n+1}} \quad \text{for any } n \geq 0 \\ \hat{Y}_0 = y \quad \text{and} \quad \hat{Y}_{n+1} &= Q_{n+1} \sqrt{\hat{Y}_n^2 + 2E_{n+1}} \quad \text{for any } n \geq 0. \end{aligned}$$

By the analogue of (2) for (13), the law of  $\hat{X}$  (respectively  $\hat{Y}$ ) is the law of the embedded chain of a process generated by (1) and starting from  $x$  (respectively  $y$ ). Now, for any  $p \geq 1$ , since  $x \mapsto \sqrt{x^2 + a}$  is a 1-Lipschitz function on  $[0, \infty)$  for any  $a \geq 0$ , we get

$$\begin{aligned} \mathbb{E} \left( \left| \hat{X}_{n+1} - \hat{Y}_{n+1} \right|^p \right) &= \mathbb{E} \left( \left| Q_{n+1}^p \left| \sqrt{\hat{X}_n^2 + 2E_{n+1}} - \sqrt{\hat{Y}_n^2 + 2E_{n+1}} \right|^p \right. \right) \\ &\leq \mathbb{E} \left( Q_{n+1}^p \left| \hat{X}_n - \hat{Y}_n \right|^p \right) = \mathbb{E}(Q_{n+1}^p) \mathbb{E} \left( \left| \hat{X}_n - \hat{Y}_n \right|^p \right) \end{aligned}$$



A straightforward recurrence leads to

$$\mathbb{E}\left(\left|\hat{X}_n - \hat{Y}_n\right|^p\right) \leq \mathbb{E}(Q_1^p)^n |x - y|^p.$$

This gives the desired inequality when the initial laws are Dirac masses. The general case follows by integrating this inequality with respect to couplings of the initial laws.  $\square$

Let us now investigate some properties of the kernel  $K$  defined by (4) that will be used to provide concentration bounds for the ergodic theorem. The key point is that  $K^n$  and  $\nu$  satisfy a Gross (or logarithmic Sobolev) inequality.

**Definition 3.2** (Gross inequality). *A law  $\eta$  on  $\mathbb{R}^d$  satisfies a Gross (or logarithmic Sobolev [1, 8]) inequality with constant  $c > 0$  when for any smooth compactly supported  $f : \mathbb{R}^d \rightarrow \mathbb{R}$ ,*

$$\int f^2 \log(f^2) d\eta - \int f^2 d\eta \log \int f^2 d\eta \leq c \int |\nabla f|^2 d\eta.$$

We denote by  $\text{GROSS}(\eta) \in (0, \infty]$  the smallest constant for which this holds true.

If  $F \cdot \eta$  is the image of  $\eta$  by  $F$  then  $\text{GROSS}(F \cdot \eta) \leq \text{GROSS}(\eta) \|F\|_{\text{LIP}}^2$ . The Gross inequality contains an information on Gaussian concentration of measure: the function  $x \mapsto e^{ax^2}$  is  $\eta$ -integrable as soon as  $a < 1/\text{GROSS}(\eta)$ . Moreover, if  $\eta$  has covariance  $\Sigma$  with spectral radius  $\rho(\Sigma)$  then  $2\rho(\Sigma) \leq \text{GROSS}(\eta)$  and equality is achieved when  $\eta$  is Gaussian. Furthermore, for any  $\alpha$ -Lipschitz function  $f : \mathbb{R} \rightarrow \mathbb{R}$  and any  $\lambda > 0$ ,

$$\mathbb{E}_\eta\left(e^{\lambda f}\right) \leq e^{C\alpha^2\lambda^2/4} e^{\lambda \mathbb{E}_\eta f} \quad (11)$$

as soon as  $C \geq \text{GROSS}(\eta)$ . This means that  $\eta$  satisfies a *sub-Gaussian concentration of measure* for Lipschitz functions. For more details, see e.g. [11, 21] and references therein.

**Theorem 3.3** (Properties of the kernel of the embedded chain). *Let  $\hat{X}$  be the embedded chain associated to (1) with transition kernel (4). Assume that  $H$  is the Dirac mass at point  $\delta \in [0, 1)$ . If  $f$  is a 1-Lipschitz function from  $[0, +\infty)$  to  $\mathbb{R}$ , then  $x \mapsto K(f)(x)$  is a  $\delta$ -Lipschitz function from  $[0, +\infty)$  to  $\mathbb{R}$ . Moreover, for any  $x \geq 0$ , the law  $K(\cdot)(x)$  satisfies a Gross inequality with constant  $2\delta^2$ .*

*Proof.* If  $\delta = 0$ , then  $K$  is the Dirac mass at 0 and the result is trivial. For any smooth function  $f : [0, \infty) \rightarrow \mathbb{R}$ , we have from (4) that

$$|(Kf)'| = \delta \left| K\left(\frac{x}{\sqrt{x^2 + 2E}} f'\left(\delta\sqrt{x^2 + 2E}\right)\right) \right| \leq \delta K(|f'|). \quad (12)$$

Let us show now that for every  $x \geq 0$  the law  $K(x, \cdot) = \mathcal{L}(\hat{X}_1 | \hat{X}_0 = x)$  satisfies a Gross inequality with constant  $2\delta^2$ . Since  $E$  is exponential of mean 1, the law  $\eta$  of  $\sqrt{E/2}$  is a  $\chi$ -distribution with probability density and cumulative distribution functions given by

$$g : v \mapsto 4ve^{-2v^2} \mathbf{1}_{\{v>0\}} \quad \text{and} \quad G : v \mapsto (1 - e^{-2v^2}) \mathbf{1}_{\{v>0\}}.$$

On the other hand,  $2E \stackrel{d}{=} U_1^2 + U_2^2$  where  $U_1, U_2$  are i.i.d. standard Gaussians, and thus

$$\sqrt{E/2} = \frac{1}{2} \sqrt{2E} \stackrel{d}{=} \frac{1}{2} \sqrt{U_1^2 + U_2^2}.$$

Also,  $\eta$  is the image of the Gaussian law  $\mathcal{N}(0, I_2)$  on  $\mathbb{R}^2$  by a  $(1/2)$ -Lipschitz function, and this implies that  $\eta$  satisfies a Gross inequality with constant  $1/2$ . Moreover,

$$\begin{aligned} K(f)(x) &= \int_0^\infty f\left(\delta\sqrt{x^2 + 2u}\right) e^{-u} du \\ &= \int f(2\delta v) \frac{4ve^{-2v^2}}{e^{-x^2/2}} \mathbf{1}_{\{v > x/2\}} dv \\ &= \int f(2\delta v) \frac{g(v)}{1 - G(x/2)} \mathbf{1}_{\{v > x/2\}} dv. \end{aligned}$$

Thus,  $K(\cdot)(x)$  is just the image law by the Lipschitz map  $v \mapsto 2\delta v$  of the law  $\eta$  conditioned on  $(x/2, +\infty)$ . This conditional law is in turn the image of  $\eta$  by the function

$$t \mapsto G^{-1}(G(x) + (1 - G(x))G(t)) = G^{-1}(1 - \exp(-t^2 - x^2)) = \sqrt{x^2 + t^2}.$$

This function is 1-Lipschitz for any  $x \geq 0$ . Consequently, by using twice the stability of Gross inequalities by Lipschitz maps, we obtain that for every  $x > 0$ , the law  $K(x, \cdot)$  satisfies a Gross inequality with constant  $(2\delta)^2/2 = 2\delta^2$ .  $\square$

**Remark 3.4.** When  $\delta = 0$ , the embedded chain is the constant Markov chain equal to 0. Moreover, the chain  $(Z_n)_{n \geq 0}$  defined by  $Z_n = X_{T_n^-}$  is also quite simple to study. Indeed, the random variables  $(Z_n)_{n \geq 1}$  are i.i.d. and have the law  $\nu$  of  $\sqrt{2E}$ . The previous proof ensures that  $\nu$  satisfies a Gross inequality with constant 2. One of the most useful properties of Gross inequality is the tensorization property:  $\text{Gross}(\eta^{\otimes n}) \leq \text{Gross}(\eta)$  for every  $n \geq 1$ , see e.g. [1, Chapter 1]. Using now the concentration property, one has, for any 1-Lipschitz function and any  $u \geq 0$ ,

$$\mathbb{P}\left(\left|\frac{1}{n} \sum_{k=1}^n f(Z_k) - \int f d\nu\right| \geq u\right) \leq 2 \exp\left(-\frac{Nu^2}{2}\right).$$

In the more general case where  $\delta$  is positive,  $(\hat{X}_n)_{n \geq 1}$  is no longer i.i.d. Nevertheless, the Gross inequality holds true for the iterated kernels and for the invariant law  $\nu$ :

**Corollary 3.5** (Gross inequality for the embedded chain and its invariant law  $\nu$ ). *Let  $\hat{X}$  be the embedded chain associated to (1). Assume that  $H$  is the Dirac mass at point  $\delta \in (0, 1)$ . For every  $n \geq 0$ , let  $K^n$  be the iterated transition kernel of  $\hat{X}$ , defined recursively for every bounded measurable function  $f : [0, \infty) \rightarrow \mathbb{R}$  by*

$$K^0(f) = f \quad \text{and} \quad K^{n+1}(f) = K(K^n(f))$$

where  $K$  is the kernel of  $\hat{X}$  as in (4). Then for every integer  $n \geq 1$  and every real  $x \geq 0$ , the iterated kernel  $K^n(x, \cdot)$  of  $\hat{X}$  satisfies a Gross inequality and

$$\text{GROSS}(K^n(x, \cdot)) \leq 2\delta^2 \frac{1 - \delta^{2n}}{1 - \delta^2}.$$

Also, the invariant law  $\nu$  of  $\hat{X}$  (see theorem 3.1) satisfies a Gross inequality and

$$\text{GROSS}(\nu) \leq 2\delta^2(1 - \delta^2)^{-1}.$$

*Proof.* Recall that for every  $n \geq 0$ ,  $x \geq 0$ , and bounded measurable  $f : [0, \infty) \rightarrow \mathbb{R}$ ,

$$\mathbb{E}\left(f(\hat{X}_n) \mid \hat{X}_0 = x\right) = (K^n f)(x) = \int_0^\infty f(y) K^n(x, dy)$$

To show that  $K^n$  satisfies a Gross inequality, we use a semi-group decomposition technique borrowed from [14]. For any  $n \geq 1$  and any smooth function  $f : [0, \infty) \rightarrow \mathbb{R}$ , the quantity

$$E_n(f) := K^n(f^2 \log f^2) - K^n(f^2) \log K^n(f^2)$$

is equal to the telescopic sum

$$\sum_{i=1}^n \{K^i [K^{n-i}(f^2) \log K^{n-i}(f^2)] - K^{i-1} [K^{n-i+1}(f^2) \log K^{n-i+1}(f^2)]\}.$$

Since the measure  $K(\cdot)(x)$  satisfies a Gross inequality of constant  $2\delta^2$ , we get, with  $g_{n-i} = \sqrt{K^{n-i}(f^2)}$ ,

$$E_n(f) = \sum_{i=1}^n K^{i-1} [E_1(g_{n-i})] \leq 2\delta^2 \sum_{i=1}^n K^i (|\nabla g_{n-i}|^2),$$

Now, by using the commutation (12), we obtain, for all  $1 \leq i \leq n$ ,

$$|\nabla g_{n-i}|^2 = \frac{|\nabla K^{n-i}(f^2)|^2}{4K^{n-i}(f^2)} \leq \delta^2 \frac{(K|\nabla K^{n-i-1}(f^2)|)^2}{4KK^{n-i-1}(f^2)}$$

Next, the Cauchy–Schwarz inequality

$$\frac{(Kf)^2}{K(g)} \leq K\left(\frac{f^2}{g}\right)$$

gives

$$\frac{(K|\nabla K^{n-i-1}(f^2)|)^2}{4KK^{n-i-1}(f^2)} \leq K\left(\frac{|\nabla K^{n-i-1}(f^2)|^2}{4K^{n-i-1}(f^2)}\right) = K(|\nabla g_{n-i-1}|^2).$$

From these bounds, a straightforward induction gives

$$|\nabla g_{n-i}|^2 \leq \delta^{2(n-i)} K^{n-i} (|\nabla f|^2).$$

Consequently, by putting all together, we have

$$E_n(f^2) \leq 2\delta^2 \sum_{i=0}^{n-1} \delta^{2i} K^n (|\nabla f|^2) = 2\delta^2 \frac{1 - \delta^{2n}}{1 - \delta^2} K^n (|\nabla f|^2).$$

This gives

$$\text{GROSS}(K^n) \leq 2\delta^2 (1 - \delta^{2n})(1 - \delta^2)^{-1}.$$

Finally, from Theorem 3.1,  $K^n$  tends weakly to  $\nu$  as  $n$  tends to infinity and thus

$$\text{GROSS}(\nu) \leq \limsup_{n \rightarrow \infty} \text{GROSS}(K^n) \leq 2\delta^2 (1 - \delta^2)^{-1}.$$

□

The Gross inequality for  $K$  can also be used to derive Theorem 2.2.

*Proof of Theorem 2.2.* We shall establish that for any  $u \geq 0$  and any 1-Lipschitz function  $f : [0, \infty) \rightarrow \mathbb{R}$ ,

$$\mathbb{P}\left(\frac{1}{n} \sum_{k=1}^n f(\hat{X}_k) - \int f d\nu \geq u + \frac{\delta}{1-\delta} W_1(\delta_x, \nu)\right) \leq \exp\left(-\frac{n(1-\delta^2)u^2}{2\delta^2}\right)$$

and the desired result follows immediately from this bound used for  $f$  and  $-f$ . For any 1-Lipschitz function  $f$ , any  $r > 0$  and  $\lambda > 0$ , we have,

$$\mathbb{P}\left(\frac{1}{n} \sum_{k=1}^n f(\hat{X}_k) \geq r\right) \leq \mathbb{E}\left(e^{\lambda \sum_{k=1}^n f(\hat{X}_k)}\right) e^{-nr\lambda}.$$

Now the Markov property ensures that

$$\begin{aligned} \mathbb{E}\left(e^{\lambda \sum_{k=1}^n f(\hat{X}_k)}\right) &= \mathbb{E}\left(e^{\lambda \sum_{k=1}^{n-1} f(\hat{X}_k)} \mathbb{E}\left(e^{\lambda f(\hat{X}_n)} | X_{n-1}\right)\right) \\ &= \mathbb{E}\left(e^{\lambda \sum_{k=1}^{n-1} f(\hat{X}_k)} K\left(e^{\lambda f}\right)(X_{n-1})\right). \end{aligned}$$

From Theorem 3.3, the kernel  $K(x, \cdot)$  of  $\hat{X}$  satisfies a Gross inequality with constant  $2\delta^2$  for every  $x \geq 0$ . This inequality implies by (11) that for any  $c$ -Lipschitz function  $g$ ,

$$K\left(e^{\lambda g}\right) \leq \exp\left(\lambda K g + \frac{c^2 \delta^2 \lambda^2}{2}\right).$$

Consequently, the Laplace transform of the ergodic mean can be bounded as follows:

$$\mathbb{E}\left(e^{\lambda \sum_{k=1}^n f(\hat{X}_k)}\right) \leq \exp\left(\frac{\delta^2 \lambda^2}{2}\right) \mathbb{E}\left(e^{\lambda \sum_{k=1}^{n-2} f(\hat{X}_k)} \mathbb{E}\left(e^{\lambda(f+Kf)(\hat{X}_{n-1})} | \hat{X}_{n-2}\right)\right).$$

The commutation relation (12) ensures that  $f + Kf$  is  $(1 + \delta)$ -Lipschitz and then

$$\mathbb{E}\left(e^{\lambda(f+Kf)(\hat{X}_{n-1})} | \hat{X}_{n-2}\right) \leq \exp\left(\frac{(1+\delta)^2 \delta^2 \lambda^2}{2}\right) e^{\lambda(Kf+K^2f)(\hat{X}_{n-2})}.$$

A straightforward recurrence ensures that

$$\mathbb{E}\left(e^{\lambda \sum_{k=1}^n f(\hat{X}_k)}\right) \leq \exp\left(\frac{n\delta^2 \lambda^2}{2(1-\delta^2)}\right) e^{\sum_{k=1}^n K^k f(x)}.$$

Choosing  $r = (1/n) \sum_{k=1}^n K^k f(x) + u$  leads to

$$\mathbb{P}\left(\frac{1}{n} \sum_{k=1}^n f(\hat{X}_k) - \frac{1}{n} \sum_{k=1}^n K^k f \geq u\right) \leq \exp\left(\frac{n\delta^2 \lambda^2}{2(1-\delta^2)} - n\lambda u\right).$$

The right hand side is minimum for  $\lambda = u(\delta^{-2} - 1)$ . At this point, we recall the dual formulation of  $W_1(\alpha, \beta)$  for every probability laws  $\alpha$  and  $\beta$ :

$$W_1(\alpha, \beta) = \sup_{\|f\|_{\text{Lip}} \leq 1} \left( \int f d\alpha - \int f d\beta \right) \quad \text{where} \quad \|f\|_{\text{Lip}} = \sup_{x \neq y} \frac{|f(x) - f(y)|}{|x - y|}.$$

Therefore, by using Theorem 2.1, one gets

$$\frac{1}{n} \sum_{k=1}^n K^k f(x) - \int f d\nu \leq \frac{1}{n} \sum_{k=1}^n W_1(K^k(\cdot)(x), \nu) \leq \frac{\delta}{1-\delta} W_1(\delta_x, \nu).$$

□

**Remark 3.6.** *A careful reading of the proof of Theorem 2.2 suggests that one may replace the initial law  $\delta_x$  by a more general initial law provided that it satisfies a sub-Gaussian concentration for Lipschitz functions (11).*

## 4 Continuous time process

As an introduction of our coupling method to prove Theorem 2.4, let us consider the following simpler dynamics, studied recently in [13, 18]. The window size is modeled by a Markov process  $X = (X_t)_{t \geq 0}$  that increases linearly with rate one. Congestion signals arrive according to a Poisson process with constant rate  $\lambda > 0$ , and upon receipt of the  $k^{\text{th}}$  signal, the window size is reduced by multiplication with a random variable  $Q_k$ . We assume that  $(Q_k)_{k \geq 0}$  is a sequence of i.i.d. random variables of law  $H$  with support in  $[0, 1)$ . In other words, the process  $X$  is generated by

$$L(f)(x) = f'(x) + \lambda \int_0^1 (f(hx) - f(x)) H(dh) \quad (13)$$

where  $\lambda$  is this time a positive real number. In this framework, one can compute explicitly the transient moments of  $X_t$  (see [13, 18]): for every  $n \geq 0$ , every  $x \geq 0$ , and every  $t \geq 0$ ,

$$\mathbb{E}((X_t)^n | X_0 = x) = \frac{n!}{\prod_{k=1}^n \theta_k} + n! \sum_{m=1}^n \left( \sum_{k=0}^m \frac{x^k}{k!} \prod_{\substack{j=k \\ j \neq m}}^n \frac{1}{\theta_j - \theta_m} \right) e^{-\theta_m t} \quad (14)$$

where for every real or integer  $p \geq 1$  the quantity  $\theta_p$  is as in our Theorem 4.1. In contrast with the original dynamics (1), the jump rate is constant and thus the jumps occur at Poissonian times. In this framework, we derive easily the following theorem, which states an exponential ergodicity in all Wasserstein distances.

**Theorem 4.1** (Wasserstein Exponential Ergodicity for constant jump rate). *Let  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  be two processes generated by (13). Assume that  $\mathcal{L}(X_0)$  and  $\mathcal{L}(Y_0)$  have finite  $p^{\text{th}}$  moment for some real  $p \geq 1$ . If one defines  $\theta_p = \lambda(1 - \mathbb{E}(Q^p))$  with  $Q \sim H$  then for every  $t \geq 0$ ,*

$$W_p(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq W_p(\mathcal{L}(X_0), \mathcal{L}(Y_0)) e^{-p^{-1}\theta_p t}.$$

We ignore if the exponential rate of convergence in Theorem 4.1 is optimal. One may try to get an upper bound from the moments formula (14).

*Proof of Theorem 4.1.* Let  $N = (N_t)_{t \geq 0}$  be a Poisson process with constant intensity  $\lambda$  and  $Q = (Q_k)_{k \geq 1}$  be i.i.d. random variables with law  $H$ , independent of  $N$ . For any  $x, y \geq 0$ , let us consider the processes  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  starting respectively

at  $x$  and  $y$  at time 0, that jump when  $N$  does, with a multiplicative factor  $Q_k$  for the  $k^{\text{th}}$  jump, and increase linearly with slope one between these jumps. It is quite clear that these processes are generated by (13). Moreover, between jumps,  $|X_t - Y_t|$  remains constant and at the  $k^{\text{th}}$  jump this quantity is multiplied by  $Q_k$ . Thus for every  $t \geq 0$  and  $p \geq 0$ ,

$$\begin{aligned} \mathbb{E}(|X_t - Y_t|^p) &= \sum_{k=0}^{\infty} \mathbb{E}(|X_t - Y_t|^p \mathbf{1}_{\{N_t=k\}}) \\ &= |x - y|^p \sum_{k=0}^{\infty} \mathbb{E}(Q^p)^k \mathbb{P}(N_t = k) \\ &= |x - y|^p e^{-\lambda t(1 - \mathbb{E}(Q^p))}. \end{aligned}$$

As a consequence, if  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  are now two processes generated by (13) with a constant jump intensity  $\lambda$  and arbitrary initial laws, we obtain that, for any coupling  $\Pi$  of their initial law  $\mathcal{L}(X_0)$  and  $\mathcal{L}(Y_0)$ , any  $t \geq 0$ , and any  $p \geq 1$ ,

$$W_p(\mathcal{L}(X_t), \mathcal{L}(Y_t))^p \leq e^{-\theta_p t} \int_{[0, \infty)^2} |x - y|^p \Pi(d(x, y)).$$

Taking the infimum over  $\Pi$  concludes the proof.  $\square$

Let us now turn to the TCP window size process generated by (1). Consider a two dimensional process where both components are generated by (1). Since the sample paths of both components have slope 1 between jumps, the distance between them remains constant except at jump times. If the components jump together with the same factor  $Q$ , then this distance is also multiplied by  $Q$ . Thus, our strategy will be to determine the greatest simultaneous jump probability  $p(x, y)$  in  $\bar{L}$ .

The section is organized as follows. We first establish elementary considerations about the laws of jump times as a function of the initial point. This allows us to construct rigorously our coupling. Then we provide the proof of Theorem 2.3. It relies on tricky controls of non-simultaneous jumps and on an original use of Gronwall's Lemma. Eventually, we give the rather simple proof of Theorem 2.5.

Let  $X = (X_t)_{t \geq 0}$  be a process generated by (1) starting at  $x \geq 0$ . As pointed out in the introduction, the law of the first jump time  $T_1$  is the one of  $\sqrt{x^2 + 2E} - x$  where  $E$  is an exponential random variable of unit mean. This law  $\nu_x$  admits the density  $f_x$  defined by

$$f_x(v) = (x + v)e^{-(x+v)^2/2} e^{x^2/2} \mathbf{1}_{\{v \geq 0\}} = (x + v)e^{-v^2/2 - xv} \mathbf{1}_{\{v \geq 0\}}. \quad (15)$$

Its cumulative distribution  $F_x$  is given by

$$F_x(v) = \left(1 - e^{-(x+v)^2/2 + x^2/2}\right) \mathbf{1}_{\{v \geq 0\}} = \left(1 - e^{-v^2/2 - xv}\right) \mathbf{1}_{\{v \geq 0\}}.$$

Note that the probability density function  $f_x$  satisfies the equation

$$f_x(v) = (x + v)(1 - F_x(v)) = \frac{f_0(x + v)}{1 - F_0(x)}.$$

Let us compare the laws of  $T_1$  as a function of the starting point (see Figure 1).

**Lemma 4.2** (Comparison of the jump time densities  $f_x$  and  $f_y$ ). *Assume that  $x > y \geq 0$ . Then there exists a unique value  $\tau(x, y)$  such that for every  $v \geq 0$ ,*

$$f_x(v) > f_y(v) \quad \text{if and only if} \quad v < \tau(x, y).$$

*Moreover, the function  $x \mapsto \tau(x, y)$  is increasing on  $[y, +\infty)$  whereas the function  $y \mapsto \tau(x, y)$  is decreasing on  $[0, x]$  and, for every  $0 \leq y < x$ , one has*

$$\frac{\sqrt{x^2 + 4} - x}{2} \leq \tau(x, y) \leq \frac{\sqrt{y^2 + 4} - y}{2} \leq 1.$$

*At last, if  $q(x, y) = 1 - p(x, y)$  is the area of the set (one may take a look at Figure 1)*

$$\{(u, v) \in [0, \infty) \times [0, \infty); f_y(u) \leq v \leq f_x(u)\},$$

*then, we have, shortening  $\tau(x, y)$  in  $\tau$ ,*

$$q(x, y) = e^{-\tau^2/2 - y\tau} - e^{-\tau^2/2 - x\tau}, \quad (16)$$

*and*

$$\partial_y q(x, y) = -\tau e^{-\tau y - \tau^2/2} < 0. \quad (17)$$

*Proof.* The function  $h_{x,y} : [0, \infty) \rightarrow \mathbb{R}$  defined by

$$h_{x,y}(v) = \log f_x(v) - \log f_y(v) = \log \left( \frac{v+x}{v+y} \right) - (x-y)v \quad (18)$$

is decreasing and convex since

$$h'_{x,y}(v) = -(x-y) \left( \frac{1}{(x+v)(y+v)} + 1 \right) < 0 \quad \text{and} \quad h''_{x,y}(v) = \frac{x-y}{(x+v)(y+v)} > 0.$$

Since  $h_{x,y}(0) > 0$  and  $h_{x,y}$  goes to  $-\infty$  as  $v \rightarrow +\infty$ , there is a unique  $\tau(x, y)$  such that  $h_{x,y}(\tau(x, y)) = 0$  or

$$\log \frac{\tau(x, y) + x}{\tau(x, y) + y} - (x-y)\tau(x, y) = 0. \quad (19)$$

For any  $x \geq 0$ , let us define

$$\tilde{\tau}(x) = (\sqrt{x^2 + 4} - x)/2.$$

The function  $g : [0, x] \rightarrow \mathbb{R}$  defined by

$$g(y) = \log \left( \frac{f_y(\tilde{\tau}(x))}{f_x(\tilde{\tau}(x))} \right)$$

is increasing and  $g(x) = 0$ . Then  $f_x(\tilde{\tau}(x)) > f_y(\tilde{\tau}(x))$  and then  $\tilde{\tau}(x) \leq \tau(x, y)$ . A similar argument leads to the upper bound. Let us differentiate (19) with respect to  $x$ :

$$\frac{\partial_x \tau(x, y) + 1}{\tau(x, y) + x} - \frac{\partial_x \tau(x, y)}{\tau(x, y) + y} - \tau(x, y) - (x-y)\partial_x \tau(x, y) = 0,$$

which is nothing else than

$$-\partial_x \tau(x, y)(x-y) \left( \frac{1}{(\tau(x, y) + x)(\tau(x, y) + y)} + 1 \right) = \frac{\tau(x, y)^2 + x\tau(x, y) - 1}{\tau(x, y) + x}.$$

Since  $\tau(x, y) \geq \tilde{\tau}$  and  $\tilde{\tau}$  is the largest root of the equation  $T^2 + xT - 1 = 0$ , we have  $\tau(x, y)^2 + x\tau(x, y) - 1 \geq 0$  and  $\partial_x \tau(x, y) \leq 0$ . A similar argument leads to  $\partial_y \tau(x, y) \leq 0$ .

At last, since  $f_y(u) \leq f_x(u)$  if and only if  $u \leq \tau$ , one has

$$q(x, y) = \int_0^\tau f_x(u) du - \int_0^\tau f_y(u) du = F_x(\tau) - F_y(\tau),$$

which is nothing else but (16). Moreover,

$$\partial_y q(x, y) = \partial_y \tau (f_x(\tau) - f_y(\tau)) - \partial_y F_y(\tau).$$

The fact that  $f_x(\tau) = f_y(\tau)$  provides (17). □

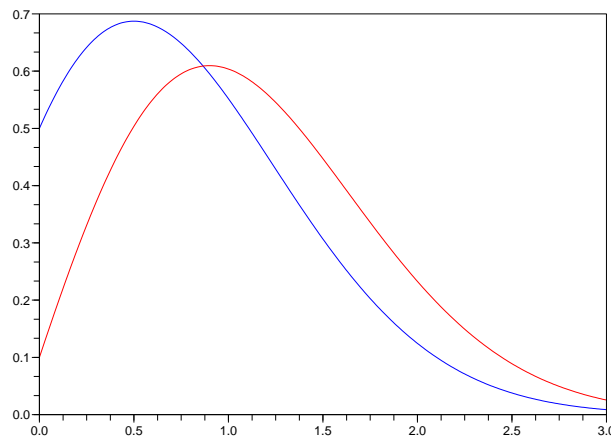


Figure 1: Plot of the densities  $f_{0.5}$  and  $f_{0.1}$ . The abscissa of the intersection is  $\tau(0.1, 0.5)$ .

Let us consider the functions  $g_1, g_2, g_3$  defined by

$$\begin{aligned} g_1 : v &\mapsto \frac{\min(f_x(v), f_y(v))}{p(x, y)} \mathbf{1}_{\{v > 0\}} \\ g_2 : v &\mapsto \frac{f_x(v) - f_y(v)}{q(x, y)} \mathbf{1}_{\{0 < v < \tau(x, y)\}} \\ g_3 : v &\mapsto \frac{f_y(v) - f_x(v)}{q(x, y)} \mathbf{1}_{\{v > \tau(x, y)\}}. \end{aligned}$$

From Lemma 4.2, there are three probability density functions and we can write  $f_x$  and  $f_y$  as two-component mixtures:

$$f_x(v) = p(x, y)g_1(v) + q(x, y)g_2(v) \quad \text{and} \quad f_y(v) = p(x, y)g_1(v) + q(x, y)g_3(v).$$

In other words, let us consider three random variables  $V_1, V_2$  and  $V_3$  with respective density functions  $g_1, g_2$ , and  $g_3$  independent of  $U$  with Bernoulli law  $\mathcal{B}(p(x, y))$ . Then,



$S = UV_1 + (1 - U)V_2$  and  $T = UV_1 + (1 - U)V_3$  have respective density functions  $f_x$  and  $f_y$  and  $\mathbb{P}(S = T) = p(x, y)$ . This procedure is a particular case of a well known maximal coupling Lemma (for the total variation distance, see e.g. [12] for a general formulation).

We are now able to construct the couple of processes  $X = (X_t)_{t \geq 0}$  and  $Y = (Y_t)_{t \geq 0}$  starting respectively from  $x$  and  $y$  that we shall use in the proof of Theorem 2.3. Let  $(E_n)_{n \geq 1}$  be a sequence of i.i.d. exponential random variables with unit mean, and let  $(Q_n)_{n \geq 1}$  be a sequence of i.i.d. random variables of law  $H$  where  $H$  is as in (1), independent of  $(E_n)_{n \geq 1}$ . Assume that  $x > y$  (since we will be interested in  $|X_t - Y_t|$ , this does not matter). For any  $t$  in the time interval  $[0, V_1)$  where  $V_1 = \sqrt{x^2 + 2E_1} - x$ , we have

$$X_t = x + t \quad \text{and} \quad Y_t = y + t.$$

At time  $V_1$ , the process  $X$  jumps to  $Q_1 X_{V_1^-}$  while  $Y$  jumps to  $Q_1 Y_{V_1^-}$  with probability  $p(X_{V_1^-}, Y_{V_1^-})$  and stays at  $Y_{V_1^-}$  with probability  $q(X_{V_1^-}, Y_{V_1^-})$ . Next we do the same by replacing  $x$  by  $\max(X_{V_1}, Y_{V_1})$  and  $y$  by  $\min(X_{V_1}, Y_{V_1})$  and  $E_1$  by  $E_2$  and  $Q_1$  by  $Q_2$ , etc. This recursive scheme produces two processes  $X$  and  $Y$  starting from  $x$  and  $y$ , and following the dynamics (1), and last but not least with maximal joint jump probability.

In other words, we have constructed a Markov process on  $[0, \infty)^2$  with infinitesimal generator given by  $\bar{L}$  (see (6)). We are now ready for the proof of Theorem 2.3.

*Proof of Theorem 2.3.* Let  $(X_t, Y_t)_{t \geq 0}$  be the couple of processes starting from  $(x, y)$  with  $x > y$  and constructed as explained before, by maximizing the joint jump probability. We assume for the moment that  $H$  is the Dirac mass at point  $\delta$ . Let  $\varphi : [0, \infty) \rightarrow [0, \infty)$  be a smooth nondecreasing and bounded function such that  $\varphi(0) = 1$ . The need of  $\varphi$  will be explained in Remark 4.4 below. For every  $t \geq 0$  we define

$$\gamma_{x,y}(t) = \mathbb{E} \left( \frac{|X_t - Y_t|}{\varphi(\max(X_t, Y_t))} \right).$$

The idea of the proof is to control the derivative of  $\gamma$  in terms of  $\gamma$  and then to use Gronwall's lemma. Let  $T$  be the time of the first jump. Recall that by the coupling construction, the law of  $T$  has density  $f_x$  defined by (15) and that, at time  $T$ ,  $X$  jumps automatically and  $Y$  jumps with probability  $p(X_T, Y_T)$ . We start with the identity

$$\gamma_{x,y}(t) = \mathbb{E} \left( \frac{|X_t - Y_t|}{\varphi(\max(X_t, Y_t))} (\mathbb{1}_{\{T \leq t, Y \text{ jumps at } T\}} + \mathbb{1}_{\{T \leq t, Y \text{ does not jump at } T\}} + \mathbb{1}_{\{T > t\}}) \right),$$

which gives

$$\begin{aligned} \gamma_{x,y}(t) &= \int_0^t (\gamma_{\delta x, \delta y}(t-s) p(x+s, y+s) + \gamma_{\delta x, y}(t-s) q(x+s, y+s)) f_x(s) ds \\ &\quad + \frac{|x-y|}{\varphi(x+t)} F_x(t) \\ &= \frac{\delta(x-y)}{\varphi(\delta x)} p(x, y) x t + \frac{|\delta x - y|}{\varphi(\delta x)} q(x, y) x t + \frac{|x-y|}{\varphi(x)} (1-xt) - \frac{(x-y)\varphi'(x)}{\varphi(x)^2} t + o(t). \end{aligned}$$

Consequently, we obtain

$$\gamma'_{x,y}(0) = \frac{x}{\varphi(\delta x)} (\delta|x-y|p(x, y) + |\delta x - y|q(x, y)) - \frac{x|x-y|}{\varphi(x)} - \frac{|x-y|\varphi'(x)}{\varphi^2(x)}. \quad (20)$$

Of course, this expression can be deduced from the expression (6) of the infinitesimal generator  $\bar{L}$  of  $(X, Y)$ . At this point, in order to use Gronwall's Lemma, we need to bound  $\gamma'_{x,y}$  in terms of  $\gamma_{x,y}$ . The following result makes this possible. Its proof can be skipped at first sight.

**Lemma 4.3** (Technical lemma). *For any  $x, y \geq 0$  and  $\delta \in [0, 1)$ , we have*

$$\beta(x, y) = -(1 - \delta)p(x, y) + \left( \frac{|\delta x - y|}{|x - y|} - 1 \right) q(x, y) \leq -a(1 - \delta)$$

where

$$a = 1 - e^{-1} \approx 0.63212.$$

Equivalently, for any  $x, y \geq 0$ , we have

$$\delta|x - y|p(x, y) + |\delta x - y|q(x, y) \leq b|x - y|$$

where

$$b = 1 - (1 - \delta)a < 1.$$

*Proof of Lemma 4.3.* If  $0 \leq y \leq \delta x$  then

$$\beta(x, y) = -(1 - \delta) \left( 1 - \frac{y}{x - y} q(x, y) \right) \leq -(1 - \delta).$$

Let us assume now that  $\delta x \leq y \leq x$ . We have then

$$\begin{aligned} \beta(x, y) &= -(1 - \delta)p(x, y) + \frac{2y - (\delta + 1)x}{x - y} q(x, y) \\ &= -(1 - \delta) + \frac{-2\delta x + (1 + \delta)y}{x - y} q(x, y). \end{aligned}$$

Let us study the function  $y \mapsto \beta(x, y)$  on  $[\delta x, x]$ .

$$\partial_y \beta(x, y) = (1 - \delta) \frac{x}{(x - y)^2} q(x, y) + \frac{2\delta x - (1 + \delta)y}{x - y} \partial_y p(x, y).$$

By using the identities (16) and (17) we get

$$\partial_y \beta(x, y) = \frac{\tau e^{-\tau^2/2 - y\tau}}{(x - y)} \left( (1 - \delta)x \frac{1 - e^{-(x-y)\tau}}{(x - y)\tau} + 2\delta x - (1 + \delta)y \right).$$

Since  $(1 - e^{-u})/u \geq 1 - u/2$  we obtain

$$\partial_y \beta(x, y) \geq \frac{\tau e^{-\tau^2/2 - y\tau}}{(x - y)} \left( (1 + \delta) - (1 - \delta) \frac{x\tau}{2} \right).$$

Since  $\tau \leq (\sqrt{y^2 + 4} - y)/2 = 2/(\sqrt{y^2 + 4} + y)$  and  $\delta x \leq y$ , we have

$$\begin{aligned} (1 + \delta) - (1 - \delta) \frac{x\tau}{2} &\geq \frac{(1 + \delta)\sqrt{y^2 + 4} + (1 + \delta)y - (1 - \delta)x}{\sqrt{y^2 + 4} + y} \\ &\geq \frac{(1 + \delta)\delta x - (1 - \delta)x}{\sqrt{y^2 + 4} + y} \geq 0. \end{aligned}$$

As a conclusion,  $\beta$  is increasing on the interval  $[\delta x, x]$ . Let us now compute the limit of  $\beta$  as  $y \rightarrow x$ . Since  $0 \leq \tau \leq 1$  and  $\tau \sim_{y \sim x} \tilde{\tau} = (\sqrt{x^2 + 4} - x)/2$ , we have, by using (17),

$$\lim_{y \rightarrow x} \beta(x, y) = -(1 - \delta) + (1 - \delta)x\tilde{\tau}e^{-\tilde{\tau}^2/2 - x\tilde{\tau}}.$$

Now, the identity  $\tilde{\tau}^2 + x\tilde{\tau} - 1 = 0$  yields

$$\lim_{y \rightarrow x} \beta(x, y) = -(1 - \delta) + (1 - \delta)x\tilde{\tau}e^{-x\tilde{\tau}/2}e^{-1/2}.$$

Next, by using

$$\frac{x\tilde{\tau}}{2} = \frac{1}{1 + \sqrt{1 + 4/x^2}} \leq \frac{1}{2}$$

and the elementary inequality  $ue^{-u} \leq (1/2)e^{-1/2}$  valid for  $u \in [0, 1/2]$ , we finally obtain

$$\lim_{y \rightarrow x} \beta(x, y) \leq -(1 - \delta)(1 - e^{-1}).$$

□

Thus from (20) and Lemma 4.3 we get

$$\gamma'_{x,y}(0) \leq - \left[ \left( 1 - b \frac{\varphi(x)}{\varphi(\delta x)} \right) x + \frac{\varphi'(x)}{\varphi(x)} \right] \gamma_{x,y}(0). \quad (21)$$

**Remark 4.4** (The need for  $\varphi$ ). *At this point, it is clear why we need  $\varphi$ . Indeed, if  $\varphi$  is constant and equal to 1, the previous computations lead to*

$$\frac{d}{dt} \mathbb{E}(|X_t - Y_t|) \leq -(1 - b) \mathbb{E}(\max(X_t, Y_t) |X_t - Y_t|) \leq 0$$

*which does not provide exponential convergence due to the presence of  $\max(X_t, Y_t)$  which can be arbitrarily small. However, this shows that  $t \mapsto \mathbb{E}(|X_t - Y_t|)$  is non-increasing. Note that the introduction of the  $\varphi$  function appears as a modification of the norm, a successful idea which is helpful in various situations, see for instance in [10, 22].*

It remains to find a good function  $\varphi$ . The ideal one should be the arg-maximum  $\varphi_0$  of the functional

$$\varphi \mapsto J(\varphi) = \inf_{x > 0} \left( \left( 1 - b \frac{\varphi(x)}{\varphi(\delta x)} \right) x + \frac{\varphi'(x)}{\varphi(x)} \right)$$

on the set of smooth positive bounded functions. . . We ignore how to determine such an optimal  $\varphi_0$ . However, one can seek for some  $\varphi$  which will give a relatively good bound. If for some function  $\varphi$ ,  $J(\varphi)$  is positive, equal to some  $\rho(\delta)$  (since  $b$  depends on  $\delta$ ), then, plugging this bound in Equation (21) leads to

$$\gamma'_{x,y}(0) \leq -\rho(\delta)\gamma_{x,y}(0).$$

Then, the Markov property ensures that, for any  $t \geq 0$ ,

$$\gamma'_{x,y}(t) \leq -\rho(\delta)\gamma_{x,y}(t),$$

which leads to

$$\mathbb{E}\left(\frac{|X_t - Y_t|}{\varphi(\max(X_t, Y_t))}\right) \leq \frac{|x - y|}{\varphi(\max(x, y))} e^{-\rho(\delta)t}.$$

Since  $1 \leq \varphi \leq \sup \varphi$ , we get

$$\mathbb{E}(|X_t - Y_t|) \leq \sup_{u>0} \varphi(u) |x - y| e^{-\rho(\delta)t}.$$

Let us now examine the problem of finding a *good* function  $\varphi$ . Let us define, for any  $u > 0$  and  $c > 1$ , the functions  $\varphi_{u,c}$  and  $\psi_{u,c}$  as

$$\varphi_{u,c}(x) = \begin{cases} \frac{c-1}{u}x + 1 & \text{for } 0 < x < u, \\ c & \text{for } x \geq u, \end{cases}$$

and

$$\psi_{u,c}(x) = \left(1 - b \frac{\varphi_{u,c}(x)}{\varphi_{u,c}(\delta x)}\right)x - \frac{\varphi'_{u,c}(x)}{\varphi_{u,c}(x)}.$$

Then

$$\psi_{u,c}(x) = \begin{cases} \left(1 - b \frac{(c-1)x + u}{(c-1)\delta x + u}\right)x + \frac{c-1}{(c-1)x + u} & \text{if } 0 < x \leq u, \\ \left(1 - \frac{ubc}{(c-1)\delta x + u}\right)x & \text{if } u < x < u/\delta, \\ (1-b)x & \text{if } x \geq u/\delta. \end{cases}$$

One can check that  $\psi_{u,c}$  is increasing on  $(u, +\infty)$ . Moreover

$$\psi_{u,c}(u^-) = \psi_{u,c}(u^+) + \frac{c-1}{cu}.$$

If one chooses

$$u = \sqrt{(c-1) \frac{(c-1)\delta + 1}{(c-1)\delta + 1 - bc}},$$

which requires that  $(c-1)\delta + 1 - bc > 0$  i.e.  $c \leq e$ , then

$$\psi_{u(c),c}(0) = \psi_{u(c),c}(u^+) = \frac{c-1}{u}.$$

Let us now explain why  $\psi_{u(c),c}$  is concave on  $[0, u)$  (which implies that the minimum of  $\psi_{u(c),c}$  is achieved at 0). For any  $x \in [0, u)$ , one can write  $\psi_{u(c),c}(x)$  as follows:

$$\psi_{u(c),c}(x) = Ax + B - \frac{b(1-\delta)v^2}{\delta^2} \frac{1}{\delta x + v} + \frac{1}{x + v},$$

where  $A$  and  $B$  are constants and  $v = u/(c-1)$ . As a consequence,  $\psi''_{u(c),c}(x)$  has the sign of

$$-\frac{b(1-\delta)v^2}{(\delta x + v)^3} + \frac{1}{(x + v)^3} = \frac{(\delta x + v)^3 - b(1-\delta)v^2(x + v)^3}{(\delta x + v)^3(x + v)^3}.$$

Then it has the sign of  $(\delta - w)x + v(1 - w)$  where  $w = (b(1 - \delta)v^2)^{1/3}$  since  $x$  and  $w$  are nonnegative. This sign is negative since  $w > 1$ . Indeed, using the expressions of  $u$  and  $b$ , we get

$$w^3 = be \frac{\delta(c-1) + c}{-(c-1)^2 + (e-1)(c-1)}.$$

Then,  $w > 1$  if and only if  $a \mapsto x^2 + (be\delta - e + 1)x + be$  is positive, which is equivalent to

$$g(\delta) = ((e-1)\delta^2 + \delta + 1 - e)^2 - 4(e - (e-1)(1-\delta)) > 0.$$

This function appears to be decreasing on  $[0, \delta_0]$  for some  $\delta_0 \in (0, 1)$  and increasing on  $(\delta_0, 1]$ . Since  $g(1) < g(0) < 0$ , the function  $g$  is negative on  $[0, 1]$ .

Then, one has to find  $c$  that maximizes  $\psi_{u(c),c}(0)$ . In other words, one has to maximize the function

$$c \in [1, e] \mapsto \sqrt{\frac{-(b-\delta)(c-1)^2 + (1-b)(c-1)}{\delta(c-1) + 1}}.$$

Since  $b - \delta = e^{-1}(1 - \delta)$  and  $1 - b = (1 - e^{-1})(1 - \delta)$ , this is equivalent to finding the maximum of the function

$$h(x) = \frac{-x^2 + (e-1)x}{\delta x + 1}$$

on  $[0, e-1]$ . The derivative of  $h$  has the sign of  $-\delta x^2 - 2x + e - 1$  then

$$\max_{x \in [0, e-1]} h(x) = h\left(\frac{\sqrt{1 + (e-1)\delta} - 1}{\delta}\right) = \frac{1}{\delta^2} \left(2 + (e-1)\delta - 2\sqrt{1 + (e-1)\delta}\right).$$

As a conclusion, one gets, for  $\delta \in (0, 1)$ ,

$$c(\delta) = \frac{\sqrt{1 + (e-1)\delta} - 1}{\delta} + 1 = \frac{e-1}{\sqrt{1 + (e-1)\delta} + 1} + 1$$

and

$$\begin{aligned} \psi_{u(c(\delta)),c(\delta)}(0) &= \sqrt{\frac{1-\delta}{e\delta^2} \left(2 + (e-1)\delta - 2\sqrt{1 + (e-1)\delta}\right)} \\ &= \frac{2 \sinh(1/2) \sqrt{1-\delta}}{\sqrt{(e-1)\delta + 1} + 1}. \end{aligned}$$

Notice that  $\delta \mapsto c(\delta)$  is decreasing. In the particular case of  $\delta = 1/2$ , one can choose

$$c = \sqrt{2e+2} - 1 \sim 1.73 \quad \text{and} \quad u \sim 2.34$$

to get

$$J(\varphi_{u(c),c}) = 3e^{-1} + 1 - 2\sqrt{2e+2} \sim 0.311798.$$

Some elementary but tedious computations (or a numerical graphical analysis) reveal that  $J(\varphi_{u(c(\delta)),c(\delta)}) \geq (1 - \delta)/2$  for any  $\delta \in (0, 1)$ . Let us now turn to the case when the probability measure  $H$  is not the Dirac mass at  $\delta$ . In that general model, we have to choose a function  $\varphi$  which does not depend on  $\delta$ . For sake of simplicity, let us choose

$u = c = 2$ . As above one can show that  $\varphi_{2,2}$  is concave on  $(0, 2)$  and increasing on  $(2, +\infty)$ . Then,

$$J(\varphi_{2,2}) = \min(\psi_{2,2}(0), \psi_{2,2}(2)) = \min\left(\frac{1}{2}, (2 - 4e^{-1})\frac{1 - \delta}{1 + \delta}\right) \geq \frac{1}{2} \frac{1 - \delta}{1 + \delta}.$$

Then the same proof (with an additional conditioning with respect to  $H$ ) provides

$$\gamma'_{x,y}(0) \leq -\rho_G \gamma_{x,y}(0) \quad \text{where} \quad \rho_G = \frac{1}{2} \mathbb{E}\left(\frac{1 - Q}{1 + Q}\right).$$

As in the deterministic case, we obtain, for every  $t \geq 0$ ,

$$\mathbb{E}(|X_t - Y_t|) \leq 2|x - y|e^{-\rho_G t}.$$

On the other hand, by Remark 4.4, we have

$$\mathbb{E}(|X_t - Y_t|) \leq |x - y|$$

for every  $t \geq 0$ . The desired result follows then immediately.  $\square$

**Remark 4.5.** *In the special but useful case where  $H$  is the Dirac mass at point  $\delta = 1/2$ , we have  $J(\varphi_{0.2,3.5}) \geq 0.46$  where  $\varphi_{u,c}(x) = c - (c - 1)e^{-ux}$  with  $c = 3.5$  and  $u = 0.2$ . This gives a better exponential rate  $\rho = 0.46$  but a greater constant in front of the exponential. Moreover, explicit controls for others values of  $\delta$  (and uniform ones) are not tractable.*

In the proof above, one has in all cases that  $\gamma'_{x,y}(t) \leq 0$ . This ensures that

$$W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq \mathbb{E}(|X_t - Y_t| | X_0 = x, Y_0 = y) \leq |x - y|.$$

By the coupling argument, it is still the case for any initial laws: for any  $t \geq 0$ ,

$$W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq W_1(\mathcal{L}(X_0), \mathcal{L}(Y_0)).$$

Let us now turn to the proof of Theorem 2.5.

*Proof of theorem 2.5.* The function  $f$  defined by  $f(x) = x$  for every  $x \geq 0$  satisfies to

$$Lf(x) = 1 - \kappa_1 x^2$$

where  $\kappa_1 = 1 - \int_0^1 h H(dh) \in (0, 1]$ . Now, for every  $x \geq 0$  and  $t \geq 0$ ,

$$\begin{aligned} \alpha_x(t) &:= \mathbb{E}(X_t | X_0 = x) \\ &= \alpha_x(0) + \int_0^t \alpha'_x(s) ds \\ &= x + \int_0^t \mathbb{E}((Lf)(X_s) | X_0 = x) ds \\ &= x + \int_0^t (1 - \kappa_1 \mathbb{E}(X_s^2 | X_0 = x)) ds. \end{aligned}$$

Also, since  $-\kappa_1$  is negative, we obtain, by using Jensen's inequality,

$$\alpha'_x(t) = 1 - \kappa_1 \mathbb{E}(X_t^2 | X_0 = x) \leq 1 - \kappa_1 \alpha_x(t)^2.$$

As a consequence,  $\alpha_x \leq \beta_x$  where  $\beta_x$  is the solution of the Riccati differential equation

$$\begin{cases} \beta_x(0) = x, \\ \beta'_x(t) = 1 - \kappa_1 \beta_x(t)^2 \text{ for } t > 0 \end{cases}$$

Denoting  $d = \sqrt{\kappa_1}$ , one gets, for  $x \geq 1/d$ ,

$$\beta_x(t) = \frac{1}{d} + \frac{2(x - 1/d)e^{-2dt}}{(dx + 1) - (dx - 1)e^{-2dt}} = \frac{1}{d} \frac{dx \cosh(dt) + \sinh(dt)}{dx \sinh(dt) + \cosh(dt)} \leq \frac{1}{d \tanh(dt)},$$

and therefore

$$\sup_{x \geq 1/d} \alpha_x(t) \leq \frac{1}{d \tanh(dt)}.$$

On the other hand, we have also  $\sup_{x \leq 1/d} \alpha_x(t) \leq 1/d$ , and thus for every  $t > 0$ ,

$$\sup_{x \geq 0} \alpha_x(t) \leq \frac{1}{d \tanh(dt)}.$$

Consider now two processes  $(X_t)_{t \geq 0}$  and  $(Y_t)_{t \geq 0}$  generated by (1) with arbitrary initial laws. For any  $s > 0$ ,  $\mathbb{E}(|X_s - Y_s|) \leq 2 \sup_x \alpha_x(s)$  and therefore the upper bound above gives

$$W_1(\mathcal{L}(X_s), \mathcal{L}(Y_s)) \leq \frac{2}{d \tanh(ds)}.$$

Together with Theorem 2.3, this gives the following uniform estimate, for every  $t \geq s > 0$ ,

$$\begin{aligned} W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) &\leq c W_1(\mathcal{L}(X_s), \mathcal{L}(Y_s)) e^{-\rho(t-s)} \\ &\leq \frac{2ce^{\rho s}}{d \tanh(ds)} e^{-\rho t}. \end{aligned}$$

As an example, when  $H$  is the Dirac measure at  $\delta$ , recall that one can choose  $\rho = (1 - \delta)/2$  and  $d = \sqrt{1 - \delta}$ . The function  $s \mapsto e^{\rho s} / \tanh(ds)$  is convex with a minimum at point

$$s_0 = \frac{1}{d} \operatorname{argth} \left( \frac{4}{\sqrt{1 - \delta} + \sqrt{17 - \delta}} \right).$$

In the case when  $\delta = 1/2$ , one gets, for any  $t > 1.72$

$$W_1(\mathcal{L}(X_t), \mathcal{L}(Y_t)) \leq 9e^{-.25t}.$$

□

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