



HAL
open science

A connection between extreme value theory and long time approximation of SDE's

Fabien Panloup

► **To cite this version:**

Fabien Panloup. A connection between extreme value theory and long time approximation of SDE's. Stochastic Processes and their Applications, 2009, 119 (10), pp.3583-3610. hal-00338401

HAL Id: hal-00338401

<https://hal.science/hal-00338401>

Submitted on 13 Nov 2008

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

A connection between extreme value theory and long time approximation of SDE's

Fabien Panloup*

November 13, 2008

Abstract

We consider a sequence $(\xi_n)_{n \geq 1}$ of *i.i.d.* random values living in the domain of attraction of an extreme value distribution. For such sequence, there exists (a_n) and (b_n) , with $a_n > 0$ and $b_n \in \mathbb{R}$ for every $n \geq 1$, such that the sequence (X_n) defined by $X_n = (\max(\xi_1, \dots, \xi_n) - b_n)/a_n$ converges in distribution to a non degenerated distribution.

In this paper, we show that (X_n) can be viewed as an Euler scheme with decreasing step of an ergodic Markov process solution to a SDE with jumps and we derive a functional limit theorem for the sequence (X_n) from some methods used in the long time numerical approximation of ergodic SDE's.

Keywords: stochastic differential equation ; jump process ; invariant distribution ; Euler scheme ; extreme value.

AMS classification (2000): 60G10, 60G70, 60J75, 65D15.

1 Introduction

Let $(\xi_n)_{n \geq 1}$ be a sequence of *i.i.d.* random values with common distribution function F . Set $M_n := \max(\xi_1, \dots, \xi_n)$ and let G denote one of the extreme values distribution functions:

$$\begin{aligned}\Lambda(x) &= \exp(-e^{-x}), & x \in \mathbb{R}, \\ \Phi_\alpha(x) &= \exp(-x^{-\alpha})1_{x>0} & \alpha > 0, \\ \Psi_\alpha(x) &= \exp(-(-x)^\alpha) \wedge 1, & \alpha > 0.\end{aligned}$$

One says that F is in the domain of attraction of one of the preceding extreme values distributions if there exist some sequences (a_n) and (b_n) , with $a_n > 0$ and $b_n \in \mathbb{R}$ for every $n \geq 1$, such that $((M_n - b_n)/a_n)_{n \geq 1}$ converges in distribution to G (see [10],

*Institut de Mathématiques de Toulouse, LSP, Université Paul Sabatier. E-mail: fpanloup@insa-toulouse.fr. Postal adress: Département GMM, INSA Toulouse, 135, Avenue de Rangueil, 31077 TOULOUSE Cedex 4.

[15] and Proposition 1 for background on extreme value theory).

In [9], Lamperti obtained a functional version of this result which is analogous to the Donsker theorem for sums of independent variables. More precisely, denoting by $(Y^{(n)})$ the sequence of càdlàg processes defined by

$$Y_t^{(n)} = \begin{cases} (M_{[nt]} - b_n)/a_n & \text{if } t \geq 1/n \\ (\xi_1 - b_n)/a_n & \text{if } 0 \leq t < 1/n, \end{cases}$$

he proved that $(Y^{(n)})_{n \geq 1}$ converges weakly on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$ (that denotes the space of càdlàg functions on \mathbb{R}_+ with values in \mathbb{R} endowed with the Skorokhod topology) to a càdlàg process Y called extremal process (see also [14]).

The aim of this paper is to obtain another functional limit theorem for the sequence $((M_n - b_n)/a_n)_{n \geq 1}$ by using that $((M_n - b_n)/a_n)_{n \geq 1}$ can be viewed as an approximation of an Euler scheme with decreasing step of an ergodic Markov process solution to a SDE with jumps.

The motivation is then twofold: on the one hand, we wish to connect the theory of long time discretization of SDE's and extreme value theory and on the other hand, we want to exhibit another functional asymptotic behavior of extremes of i.i.d random sequences.

Set $X_n := (M_n - b_n)/a_n$. Then, the sequence (X_n) can be recursively written as follows: $X_1 = \xi_1$ and for every $n \geq 1$,

$$X_{n+1} = \frac{a_n}{a_{n+1}} X_n + \frac{b_n - b_{n+1}}{a_{n+1}} + \frac{a_n}{a_{n+1}} \left(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} \right)_+.$$

For every $n \geq 1$, we set $\theta_n := \inf\{x, F(x) \geq 1 - 1/n\}$ and $\gamma_n := 1 - F(\theta_n)$. Note that $\gamma_n = 1/n$ if F is continuous and that $(\gamma_n)_{n \geq 1}$ is a nonincreasing sequence. Then, we denote by $(\rho_n)_{n \geq 1}$ and $(\beta_n)_{n \geq 1}$ the sequences defined by $\rho_n := \frac{a_{n-1} - a_n}{a_n \gamma_n}$ and $\beta_n := \frac{b_{n-1} - b_n}{a_n \gamma_n}$. With these notations, we have for every $n \geq 1$,

$$X_{n+1} = X_n + \gamma_{n+1}(\rho_{n+1} X_n + \beta_{n+1}) + (1 + \rho_{n+1} \gamma_{n+1}) \left(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} \right)_+. \quad (1)$$

Setting $\Gamma_n := \sum_{k=1}^n \gamma_k$, we denote by $(\mathcal{X}^{(n)})_{n \geq 1}$, the sequence of càdlàg processes defined by

$$\mathcal{X}_t^{(n)} = X_{N(n,t)} \quad \text{with} \quad N(n,t) = \inf\{k \geq n, \Gamma_{k+1} - \Gamma_n > t\}. \quad (2)$$

In particular, $\mathcal{X}_0^{(n)} = X_n$. Finally, we set $\mathcal{F}_n = \sigma(X_1, \dots, X_n)$.

In Equation (1), we try to write $X_{n+1} - X_n$ as the increment of an Euler scheme with step γ_n of a SDE (that we identify in the sequel). Then, the stepwise constant process $\mathcal{X}^{(n)}$ plays the role of a continuous-time version of the Euler scheme starting from time Γ_n .

At this stage, we can observe that two types of terms appear in the right-hand member of Equation (1): the first one is close to the time discretization of a drift term and the second one looks like a positive jump. This heuristic remark will be clarified in Theorem 1.

2 Main Result

The main result of this paper is Theorem 1. In this result, we show under some mild conditions that the sequence $(\mathcal{X}^{(n)})_{n \geq 1}$ converges weakly to a stationary Markov process for the Skorokhod topology on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$. In this way, we first need to recall the result by Gnedenko which characterizes the domain of attraction of each extreme value distribution (see [5]):

PROPOSITION 1. *Let (ξ_n) be a sequence of i.i.d. real-valued random variables with distribution function F and set $x_F := \sup\{x, F(x) < 1\}$. Then, there exists (a_n) and (b_n) such that $((M_n - b_n)/a_n)$ converges weakly to a random variable with a non-degenerated distribution function G if and only if one of the following conditions is fulfilled :*

- *Type 1 : There exists a positive function g such that*

$$\frac{1 - F(t + xg(t))}{1 - F(t)} \xrightarrow{t \nearrow x_F} -\ln(\Lambda(x)) \quad \forall x \in \mathbb{R}. \quad (3)$$

In this case, $G(x) = \Lambda(x)$ for every $x \in \mathbb{R}$ and the norming constants a_n and b_n may be chosen as $a_n = g(\theta_n)$ and $b_n = \theta_n$.

- *Type $(2, \alpha)$: $x_F = +\infty$ and for every $x > 0$*

$$\frac{1 - F(tx)}{1 - F(t)} \xrightarrow{t \rightarrow +\infty} -\ln(\Phi_\alpha(x)) \quad \text{where } \alpha > 0.$$

In this case, $G(x) = \Phi_\alpha(x)$ and the norming constants a_n and b_n may be chosen as $a_n = \theta_n$ and $b_n = 0$.

- *Type $(3, \alpha)$: $x_F < +\infty$ and for every $x < 0$,*

$$\frac{1 - F(x_F + xt)}{1 - F(x_F - t)} \xrightarrow{t \nearrow x_F} -\ln(\Psi_\alpha(x)) \quad \text{with } \alpha > 0.$$

In this case, $G(x) = \Psi_\alpha(x)$ and the norming constants a_n and b_n may be chosen as $a_n = x_F - \theta_n$ and $b_n = x_F$.

In the sequel, we will denote by ν_G the probability associated with G and by τ_G the limiting function that appears in Proposition 1:

$$\tau_G(x) = \begin{cases} \exp(-x) & \text{if } F \text{ is of type 1,} \\ x^{-\alpha} 1_{x>0} & \text{if } F \text{ is of type } (2, \alpha) \\ (-x)^\alpha 1_{x \leq 0} & \text{if } F \text{ is of type } (3, \alpha) \end{cases} \quad (4)$$

We will also denote by $\mathcal{C}_K^1(D_G)$ the set of \mathcal{C}_1 -functions with compact support in D_G with

$$D_G = \begin{cases} \mathbb{R} & \text{if } F \text{ is of type 1,} \\ (0, +\infty) & \text{if } F \text{ is of type } (2, \alpha), \\ (-\infty, 0] & \text{if } F \text{ is of type } (3, \alpha). \end{cases}$$

Finally, for a distribution function F of type 1, $(2, \alpha)$ or $(3, \alpha)$, $\mathcal{C}(F)$ will correspond to the set of sequences $(a_n, b_n)_{n \geq 1}$ (with $a_n > 0$ and $b_n \in \mathbb{R}$) such that, for every

sequence $(\xi_n)_{n \geq 1}$ of *i.i.d.* random variables with common distribution function F , $((\max(\xi_1, \dots, \xi_n) - b_n)/a_n)_{n \geq 1}$ converges weakly to ν_G . We recall that if $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$, then, $(\tilde{a}_n, \tilde{b}_n)_{n \geq 1} \in \mathcal{C}(F)$ if and only if,

$$a_n \underset{n \rightarrow +\infty}{\sim} \tilde{a}_n \quad \text{and} \quad \frac{b_n - \tilde{b}_n}{a_n} \underset{n \rightarrow +\infty}{\rightarrow} 0 \quad (\text{see e.g. [15], Proposition 0.2}). \quad (5)$$

Now, we introduce $\mathbf{H}_\Lambda^1(F)$ and $\mathbf{H}_\Lambda^2(F)$ which are some assumptions that are needed for Theorem 1 when F is of type 1:

$\mathbf{H}_\Lambda^1(F)$: There exists $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ such that $(\rho_n)_{n \geq 1}$ and $(\beta_n)_{n \geq 1}$ converge to some finite values.

$\mathbf{H}_\Lambda^2(F)$: There exists a function g satisfying (3), a positive number λ and $\delta \in (-\infty, x_F)$ such that for sufficiently large n ,

$$\int_u^0 \frac{g(\theta_n)}{g(\theta_n + g(\theta_n)v)} dv \leq -\lambda u \quad \text{for every } u \in \left(\frac{\delta - \theta_n}{g(\theta_n)}, 0 \right).$$

REMARK 1. Assumption $\mathbf{H}_\Lambda^1(F)$ is only needed when F is of type 1 since it is always satisfied in the other cases. Actually, when F is of type $(2, \alpha)$ or $(3, \alpha)$, one can build $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ such that the pair of sequences $(\rho_n, \beta_n)_{n \geq 1}$ converges to $(-1/\alpha, 0)$ and $(1/\alpha, 0)$ respectively (see Section 5 for details). We are not able to obtain such a general result when F is of type 1 but one can check that $\mathbf{H}_\Lambda^1(F)$ is true in some standard cases:

- Exponential distribution: we can take $a_n = 1$ and $b_n = \log n$. Then, $\rho_n = 0$, and $\beta_n \underset{n \rightarrow +\infty}{\rightarrow} -1$.

- Normal distribution: $a_n = \frac{1}{\sqrt{2 \log n}}$ and $b_n = \sqrt{2 \log n} - \frac{1}{2\sqrt{\log n}}(\log \log n + \log 4\pi)$ belong to $\mathcal{C}(F)$. For these choices, $\rho_n \underset{n \rightarrow +\infty}{\rightarrow} 0$ and $\beta_n \underset{n \rightarrow +\infty}{\rightarrow} -1$.

Note that the limits of $(\rho_n)_{n \geq 1}$ and $(\beta_n)_{n \geq 1}$ correspond to ρ_G and β_G defined in Theorem 1. In particular, they depend only on the type of F in the sense that if $(\rho_n)_{n \geq 1}$ and (β_n) converge, then, the limits are systematically ρ_G and β_G respectively (see Section 5 for details).

Assumption $\mathbf{H}_\Lambda^2(F)$ is needed to control the evolution of $(X_k)_{k \geq 1}$ (see Lemma 2). First, using that for g and \tilde{g} satisfying (3), $g(t) \sim \tilde{g}(t)$ as $t \nearrow x_F$, one observes that Assumption $\mathbf{H}_\Lambda^2(F)$ does not depend on the choice of the function g . Second, $\mathbf{H}_\Lambda^2(F)$ is satisfied in the two preceding examples. Actually, for an exponential distribution, one can take $g(x) = 1$ and for a normal distribution, one can show that $g(x) \underset{x \rightarrow +\infty}{\sim} x^{-1}$ (see e.g. [10]). Hence, for a sufficiently large δ and n ,

$$\frac{g(\theta_n)}{g(\theta_n + g(\theta_n)v)} \leq C \frac{g(\theta_n)v + \theta_n}{\theta_n} \leq C \quad \forall v \in \left(\frac{\delta - \theta_n}{g(\theta_n)}, 0 \right).$$

More generally, one notices that $\mathbf{H}_\Lambda^2(F)$ is satisfied as soon as, g is bounded and cut by a positive number or, if there exists $\delta < x_F$ such that g is nonincreasing on $[\delta, x_F)$.

Let us now state our main result:

THEOREM 1. *Let F be a distribution function of type 1, $(2, \alpha)$ with $\alpha > 2$ or $(3, \alpha)$ with $\alpha > 0$ and let G denote one of the extreme values distribution functions. Assume $\mathbf{H}_\Lambda^1(F)$ and $\mathbf{H}_\Lambda^2(F)$ if F is of type 1. Then, for every $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$, the sequence of càdlàg processes $(\mathcal{X}^{(n)})_{n \geq 0}$ (defined in (2)) converges weakly, for the Skorokhod topology on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$, to a stationary Markov process with invariant distribution ν_G and infinitesimal generator \mathcal{A} defined for every $f \in \mathcal{C}_K^1(D_G)$ by*

$$\mathcal{A}f(x) = (\rho_G x + \beta_G) f'(x) + \int_0^{+\infty} (f(x+y) - f(x)) \varphi_G(x+y) dy$$

where $\varphi_G(x) = -\tau_G'(x)$ and (ρ_G, β_G) satisfies

$$(\rho_G, \beta_G) = \begin{cases} (0, -1) & \text{if } F \text{ is of type 1,} \\ (-1/\alpha, 0) & \text{if } F \text{ is of type } (2, \alpha), \\ (1/\alpha, 0) & \text{if } F \text{ is of type } (3, \alpha). \end{cases} \quad (6)$$

With the time discretization standpoint adopted in this paper, Theorem 1 is close to that obtained in [1] where the authors show a similar functional weak convergence result for the Euler scheme with decreasing step of ergodic Brownian diffusions. Then, if (X_n) was a “true” Euler scheme for the Markov process with infinitesimal generator \mathcal{A} , Theorem 1 would be only an adaptation of [1] to this type of SDE’s with jumps.

REMARK 2. In the literature about the numerical approximation of the stationary regime of Markovian SDE’s, another type of result could also be connected to extreme value theory. Actually, in [7, 8], the authors show the *a.s.* weak convergence of some weighted occupation measures of the Euler scheme of Brownian SDE’s (see also [11] and [13] for extensions). By adapting the approach of these papers to this context, it could be possible to retrieve the *a.s.* CLT for extreme values obtained in [4].

REMARK 3. The reader can observe that we assume $\alpha > 2$. This assumption can be viewed as a consequence of martingale methods in which the convergence needs some control of the moments. However, this assumption could be alleviated and it seems that at the price of technicalities, the result still holds if we only assume that $\alpha > 1$.

The proof of Theorem 1 is divided in three parts. First, in Section 3, we establish some stability properties for the sequence $(X_k)_{k \geq 1}$ and obtain the tightness of the sequence $(\mathcal{X}^{(n)})_{n \geq 1}$ on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$. Second, in Section 4, we identify the limit by showing that every weak limit of $(\mathcal{X}^{(n)})_{n \geq 1}$ is a solution to a martingale problem. Finally, Section 5 is devoted to some details, to a synthesis of the two preceding parts and to the uniqueness of the martingale problem.

Before going further, let us precise several notations of the proof. For a non-decreasing function g , we denote by g^\leftarrow , its left continuous inverse defined by $g^\leftarrow(x) = \inf\{y, g(y) \geq x\}$. Throughout the proof, C will denote a constant which may change from line to line.

3 Tightness of the sequence $(\mathcal{X}^{(n)})_{n \geq 1}$

The main result of this section is Proposition 2 where we obtain that the sequence $(\mathcal{X}^{(n)})_{n \geq 1}$ is tight on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$. Before stating it, we need to establish a series of technical lemmas.

In Lemma 1, we show that for every $\delta < x_F$, we can suppose that $F(x) = 0$ for every $x \in (-\infty, \delta)$. This assumption will be convenient for the sequel of the proof.

LEMMA 1. *Let F be a distribution function and δ be a real number such that $\delta < x_F$. Denote by F_δ the distribution function defined by $F_\delta(x) = F(x)1_{\{x \geq \delta\}}$. Let $(\mathcal{X}^{(n)})$ and $(\mathcal{X}^{(n,\delta)})$ denote some sequences of càdlàg processes built as in Equation 2 and corresponding to F and F_δ respectively. Then, for every bounded continuous functional $H : \mathbb{D}(\mathbb{R}_+, \mathbb{R}) \rightarrow \mathbb{R}$,*

$$\mathbb{E}[H(\mathcal{X}^{(n)}) - H(\mathcal{X}^{(n,\delta)})] \xrightarrow{n \rightarrow +\infty} 0. \quad (7)$$

Proof. Let $(\xi_n)_{n \geq 1}$ and $(\xi_n^\delta)_{n \geq 1}$ denote some sequences of *i.i.d.* random variables with distribution function F and F_δ respectively. Since this lemma only depends on the distribution of these sequences, we can assume that $(\xi_n)_{n \geq 1}$ and $(\xi_n^\delta)_{n \geq 1}$ are built as follows : $\xi_n = F^\leftarrow(U_n)$ and $\xi_n^\delta = F_\delta^\leftarrow(U_n)$ where $(U_n)_{n \geq 1}$ is a sequence of *i.i.d.* random variables such that $U_1 \sim \mathcal{U}_{[0,1]}$. In particular, we have $\xi_n = \xi_n^\delta$ on the event $\{U_n > F(\delta)\}$. It follows that $\mathcal{X}^{(n)} = \mathcal{X}^{(n,\delta)}$ on the event $A_n = \bigcup_{k=1}^n \{U_k > F(\delta)\}$. The sequence $(A_n)_{n \geq 1}$ is non-decreasing and $\mathbb{P}[\lim_{n \rightarrow +\infty} A_n] = 1$ since $\mathbb{P}[U_1 > F(\delta)] > 0$. Assertion (7) easily follows. \square

The aim of the following lemma is to obtain a recursive control of the conditional moments of order 1 and 2 for the last term of Equation (1). Keeping in mind Lemma 1, we show that this control is possible at the price of a potentially restriction of the support of F .

LEMMA 2. *Let F be a distribution function of type 1, $(2, \alpha)$ with $\alpha > 2$ or $(3, \alpha)$ with $\alpha > 0$ and let $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$. Assume furthermore that Assumption $\mathbf{H}_\Lambda^2(F)$ holds (if F is of type 1). Then, for every $\varepsilon > 0$ and, $\varepsilon < \alpha - 2$ if F is of type 2, there exists $\delta_\varepsilon < x_F$ such that for every $\delta \in [\delta_\varepsilon, x_F[$, for every sequence $(\xi_k)_{k \geq 1}$ of *i.i.d.* random variables with distribution function F_δ (defined by $F_\delta(x) = F(x)1_{\{x \geq \delta\}}$), there exists $n_\varepsilon \in \mathbb{N}$ such that for every $k \geq n_\varepsilon$,*

(i)

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+ / \mathcal{F}_k\right] \leq C\gamma_{k+1} \begin{cases} (1 + e^{-\lambda X_k}) & (\lambda > 0) & \text{if } F \text{ is of type 1,} \\ (1 + |X_k|^{-(\alpha+\varepsilon)+1}) & & \text{if } F \text{ is of type } (2, \alpha), \\ (1 + |X_k|^{\alpha+\varepsilon+1}) & & \text{if } F \text{ is of type } (3, \alpha). \end{cases}$$

(ii)

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+^2 / \mathcal{F}_k\right] \leq C\gamma_{k+1} \begin{cases} (1 + e^{-\lambda X_k}) & (\lambda > 0) & \text{if } F \text{ is of type 1,} \\ (1 + |X_k|^{-(\alpha+\varepsilon)+2}) & & \text{if } F \text{ is of type } (2, \alpha), \\ (1 + |X_k|^{\alpha+\varepsilon+2}) & & \text{if } F \text{ is of type } (3, \alpha). \end{cases}$$

Proof. (i) Using (5), it is easy to check that (i) and (ii) hold for every sequence $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ if and only if they hold for a particular sequence $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$. Hence, we only prove this lemma with (a_n, b_n) chosen as in Proposition 1. We have:

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+ / \mathcal{F}_k\right] = \int_0^{+\infty} (1 - F(a_k(u + X_k) + b_k)) du. \quad (8)$$

Assume first that F is of type 1. We set $(a_n, b_n) = (g(\theta_n), \theta_n)$ where g is such that (3) holds. By Proposition 1.4 p.43 of [15], g can be chosen such that the following representation holds for every $x \in [z_0, x_F]$:

$$1 - F(x) = c(x) \exp\left(-\int_{z_0}^x \frac{1}{g(s)} ds\right) \quad \text{where} \quad c(x) \xrightarrow{x \nearrow x_F} c > 0. \quad (9)$$

Let δ be a real number such that $\delta \geq z_0$ and (ξ_n) be a sequence of *i.i.d.* random variables with distribution function F_δ . For every positive u , we have $a_k(u + X_k) + b_k > \delta$. Then, if $a_k(u + X_k) + b_k < x_F$, we have for sufficiently large k ,

$$\begin{aligned} \frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} &= \frac{1 - F(a_k(u + X_k) + b_k)}{1 - F(b_k)} \\ &= \frac{c(a_k(u + X_k) + b_k)}{c(b_k)} \exp\left(-\int_{b_k}^{b_k + a_k(u + X_k)} \frac{1}{g(s)} ds\right) \\ &= \frac{c(a_k(u + X_k) + b_k)}{c(b_k)} \exp\left(-\int_0^{u + X_k} \frac{a_k}{g(a_k s + b_k)} ds\right). \end{aligned}$$

First, there exists $\delta_1 \in [z_0, x_F[$ such that for every $x \in [\delta_1, x_F[$, $c/2 \leq c(x) \leq 2c$. Hence, for every $\delta \geq \delta_1$ and for sufficiently large k ,

$$\frac{c(a_k(u + X_k) + b_k)}{c(b_k)} \leq C.$$

Second, one derives from Assumption $\mathbf{H}_\Lambda^2(F)$, one can find $\delta_2 \in [z_0, x_F[$ such that for sufficiently large k ,

$$\int_{u + X_k}^0 \frac{a_k}{g(a_k s + b_k)} ds \leq -\lambda(u + X_k) \quad \text{when } u + X_k < 0.$$

Thus, for $\delta > \delta_0 = \delta_1 \vee \delta_2$ and sufficiently large k ,

$$\frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} \leq C \exp(-\lambda(u + X_k)) \quad \forall u \in [0, -X_k \vee 0). \quad (10)$$

Finally, if $u + X_k > 0$, we derive from Lemma 2.2 p.78 of [15] that, for every $\varepsilon > 0$, there exists $k_\varepsilon > 0$ such that

$$\frac{a_k}{g(a_k s + b_k)} \geq \frac{1}{1 + \varepsilon s} \quad \forall k \geq k_\varepsilon \quad \forall s > 0.$$

Hence, for sufficiently large k ,

$$\frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} \leq C_\varepsilon (1 + \varepsilon(u + X_k))^{-\frac{1}{\varepsilon}} \quad \forall u \in [0 \vee -X_k, +\infty). \quad (11)$$

Setting $\varepsilon = 1/2$ and $x^- = \max(-x, 0)$, we derive from (8), (10) and (11) that, with a sufficient restriction of the support of F , there exists $k_0 \in \mathbb{N}$ such that for every $k \geq k_0$,

$$\begin{aligned} \mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+ / \mathcal{F}_k\right] &\leq C\gamma_k \left(\int_0^{X_k^-} e^{-\lambda(u+X_k)} du + \int_{X_k^-}^{+\infty} \frac{1}{(1 + \frac{1}{2}(u + X_k))^2} du \right) \\ &\leq C\gamma_k(1 + e^{-\lambda X_k}). \end{aligned}$$

Assume now that F is of type $(2, \alpha)$. By the Karamata representation ([15], p. 58), for every $x \geq 1$, we have

$$1 - F(x) = c(x) \exp\left(-\int_1^x \frac{\alpha(t)}{t} dt\right) \quad (12)$$

with $c(x) \xrightarrow{x \rightarrow +\infty} c > 0$ and $\alpha(x) \xrightarrow{x \rightarrow +\infty} \alpha > 0$. Then, replacing F by F_1 defined by $F_1(x) = F(x)1_{\{x \geq 1\}}$, we have for every $u > 0$ and sufficiently large k ,

$$\begin{aligned} \frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} &= \frac{1 - F(a_k(u + X_k))}{1 - F(a_k)} \\ &= \frac{c(a_k(u + X_k))}{c(a_k)} \exp\left(-\int_{a_k}^{a_k(u+X_k)} \frac{\alpha(s)}{s} ds\right). \end{aligned}$$

Let $\varepsilon > 0$. Let $\delta > 1$ such that $c/2 \leq c(x) \leq 2c$ and $\alpha - \varepsilon \leq \alpha(x) \leq \alpha + \varepsilon$ for every $x \geq \delta$. Then, replacing F_1 by F_δ , we obtain that

$$\frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} \leq C \left((u + X_k)^{-\alpha+\varepsilon} 1_{\{u+X_k > 1\}} + (u + X_k)^{-\alpha-\varepsilon} 1_{\{u+X_k \leq 1\}} \right). \quad (13)$$

Hence, we derive from (8) that for $\varepsilon \in]0, \alpha - 1[$ and sufficiently large k ,

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+ / \mathcal{F}_k\right] \leq C\gamma_k \left(1 + \frac{1}{X_k^{\alpha+\varepsilon-1}}\right). \quad (14)$$

Finally, assume that F is of type $(3, \alpha)$. By Corollary 1.14 p.62 of [15], $1 - F$ can be written as follows:

$$1 - F(x) = c(x) \exp\left(-\int_{x_F-1}^x \frac{\alpha(s)}{x_F - s} ds\right) \quad \forall x < x_F \quad (15)$$

with $c(x) \xrightarrow{x \nearrow x_F} c > 0$ and $\alpha(x) \xrightarrow{x \nearrow x_F} \alpha > 0$. Let $\varepsilon \in]0, \alpha/2[$. There exists $\delta_\varepsilon < x_F$ such that for every $x > \delta_\varepsilon$, $c/2 \leq c(x) \leq 2c$ and $\alpha - \varepsilon \leq \alpha(x) \leq \alpha + \varepsilon$. Assume now that $(\xi_n)_{n \geq 1}$ is a sequence of *i.i.d.* random variables with distribution function F_δ . Setting $(a_k, b_k) = (x_F - \theta_k, x_F)$, we obtain that for every $\varepsilon > 0$, there exists $n_\varepsilon \in \mathbb{N}$

such that for every $k \geq n_\varepsilon$ and $u \in [0, -X_k[$,

$$\begin{aligned} \frac{1 - F(a_k(u + X_k) + b_k)}{\gamma_k} &= \frac{1 - F(a_k(u + X_k) + x_F)}{1 - F(x_F - a_k)} \\ &\leq \frac{c(a_k(u + X_k) + x_F)}{c(x_F - a_k)} \exp\left(-\int_{x_F - a_k}^{x_F + a_k(u + X_k)} \frac{\alpha(s)}{x_F - s} ds\right), \\ &\leq C \exp\left(\int_{a_k}^{-a_k(u + X_k)} \frac{\alpha(x_F - v)}{v} dv\right), \\ &\leq C\left((-u + X_k)^{\alpha - \varepsilon} 1_{\{u + X_k > -1\}} + (-u + X_k)^{\alpha + \varepsilon} 1_{\{u + X_k \leq -1\}}\right). \end{aligned}$$

Using that $1 - F(a_k(u + X_k) + b_k) = 0$ when $u + X_k > 0$, we finally derive from (8) that

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+ / \mathcal{F}_k\right] \leq C\gamma_k(1 + |X_k|^{1+\alpha+\varepsilon}).$$

(ii) We have

$$\mathbb{E}\left[\left(\frac{\xi_{k+1}}{a_k} - X_k - \frac{b_k}{a_k}\right)_+^2 / \mathcal{F}_k\right] = 2 \int_0^{+\infty} u(1 - F(a_k(u + X_k) + b_k)) du. \quad (16)$$

and the result follows easily from the controls established in (i). Details are left the reader. \square

In the next lemma, we obtain an uniform control of the moments of the sequence $(X_n)_{n \geq 1}$ in terms of the type of the distribution function F .

LEMMA 3. *Let F be a distribution function and assume that there exists $\delta < x_F$ such that $F(x) = 0$ for every $x < \delta$. Let $(a_n, b_n) \in \mathcal{C}(F)$. Then,*

(i)

$$\sup_{n \geq 1} \mathbb{E}[|X_n|^r] < +\infty \quad \begin{cases} \forall r \geq 0 & \text{if } F \text{ is of type 1 or } (3, \alpha) \\ \forall r \in [0, \alpha) & \text{if } F \text{ is of type } (2, \alpha). \end{cases}$$

(ii) *Assume that F is of type 1. Then, for every positive number λ , there exists $\delta_\lambda < x_F$ such that for every sequence $(\xi_n)_{n \geq 1}$ with distribution function F_δ with $\delta \in [\delta_\lambda, x_F)$,*

$$\sup_{n \geq 1} \mathbb{E}[\exp(-\lambda X_n)] < +\infty. \quad (17)$$

(iii) *Assume that F is of type $(2, \alpha)$ ($\alpha > 0$) such that $F(x) = 0$ when $x < 1$. Then,*

$$\sup_{n \geq 1} \mathbb{E}\left[\frac{1}{(X_n)^r}\right] < +\infty, \quad \forall r \geq 0.$$

Proof. As in Lemma 2, one derives easily from (5) that the assertions of Lemma 3 hold with every $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ if and only if they hold for a particular sequence $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$. Hence, we assume that $(a_n, b_n)_{n \geq 1}$ is as specified in Proposition

1.

(i) This is a consequence of Proposition 2.1 p.77 of [15].

(ii) Let $\delta \in \mathbb{R}$ with $\delta < x_F$ and let $(\xi_n)_{n \geq 1}$ be a sequence of *i.i.d.* random variables with distribution function F_δ . First, since F_δ is of type 1 for every $\delta < x_F$, $(X_n)_{n \geq 1}$ converges in distribution. Second, $x \mapsto e^{-\lambda x}$ is bounded on $[-L, +\infty)$ for every $L > 0$. Hence, (17) holds if there exists $\delta_\lambda < x_F$ such that for every $\delta \in [\delta_\lambda, x_F]$,

$$\lim_{L \rightarrow +\infty} \limsup_{n \rightarrow +\infty} \mathbb{E} [\exp(-\lambda X_n) 1_{\{X_n < -L\}}] = 0. \quad (18)$$

Let us prove (18). We have:

$$\begin{aligned} \mathbb{E} [\exp(-\lambda X_n) 1_{\{X_n < -L\}}] &\leq \mathbb{E} [\exp(-\lambda X_n 1_{\{X_n < -L\}})], \\ &\leq 1 + \int_0^{+\infty} \lambda \exp(\lambda u) \mathbb{P}(X_n 1_{\{X_n < -L\}} < -u) du, \\ &\leq 1 + C \mathbb{P}(X_n < -L) + \int_L^{+\infty} \lambda \exp(\lambda u) \mathbb{P}(X_n < -u) du. \end{aligned}$$

Since ξ_1 has distribution function F_δ , $\mathbb{P}(X_n < -u) = 0$ if $u > a_n^{-1}(b_n - \delta)$ and $\mathbb{P}(X_n < -u) = F^n(-a_n u + b_n)$ if $u < a_n^{-1}(b_n - \delta)$. Hence, we derive from Lemma 2.2 p.78 of [15] that, for every $\varepsilon > 0$, there exists $\delta_\varepsilon < x_F$ such that for every $\delta \in [\delta_\varepsilon, x_F]$, for every $u > a_n^{-1}(\delta - b_n)$,

$$\mathbb{P}(X_n < -u) = F^n(-a_n u + b_n) \leq \exp\left(- (1 - \varepsilon)^2 (1 + \varepsilon |u|)^{\frac{1}{\varepsilon}}\right).$$

Setting $\varepsilon := 1/2$, we obtain that for such δ , $u \mapsto \exp(\lambda u) \mathbb{P}(X_n < -u)$ is dominated on \mathbb{R}_- by a Lebesgue-integrable function (uniformly in n). Hence, we derive from the dominated convergence theorem that

$$\lim_{L \rightarrow +\infty} \limsup_{n \rightarrow +\infty} \int_L^{+\infty} \lambda \exp(\lambda u) \mathbb{P}(X_n < -u) du = 0.$$

Finally, since

$$\lim_{L \rightarrow +\infty} \limsup_{n \rightarrow +\infty} \mathbb{P}(X_n < -L) = \lim_{L \rightarrow +\infty} \mathbb{P}(X_\infty < -L) = 0,$$

(18) follows.

(iii) Assume that $F(x) = 0$ for every $x < 1$. In particular, $X_n > 0$ *a.s.* for every $n \geq 1$. Then,

$$\frac{1}{X_n} = \frac{a_n}{\max(\xi_1, \dots, \xi_n)} = -a_n \max(\tilde{\xi}_1, \dots, \tilde{\xi}_n)$$

where $\tilde{\xi}_1 = -\frac{1}{\xi_1}$. Since ξ_1 is of type $(2, \alpha)$, it is easy to check that $\tilde{\xi}_1$ is of type $(3, \alpha)$ with $x_F = 0$. Set $\tilde{F}(x) := \mathbb{P}(\tilde{\xi}_1 \leq x)$ and $\tilde{a}_n := -\inf\{x, \tilde{F}(x) \geq 1 - \frac{1}{n}\}$. We have

$$-\tilde{a}_n = \inf\{x, F(-\frac{1}{x}) \geq 1 - \frac{1}{n}\} = -\frac{1}{a_n},$$

and the result follows from (i) (Control of the moments for F of type $(3, \alpha)$). \square

In Lemma 4, we state a simple criteria of C -tightness adapted to this problem.

LEMMA 4. *Let $(Z^{(n)})_{n \geq 1}$ be a sequence of càdlàg processes such that for every $n \geq 1$,*

$$Z_t^{(n)} = \sum_{k=n+1}^{N(n,t)} \gamma_k \phi_k(X_{k-1}) \quad (19)$$

where $(\phi_k)_{k \geq 1}$ is a sequence of real functions and assume that there exists $p > 1$ such that $\sup_{k \geq 1} \mathbb{E}[|\phi_k(X_{k-1})|^p] < +\infty$. Then, $(Z^{(n)})_{n \geq 1}$ is C -tight.

Proof. First, by (19) and the fact that $\sup_{k \geq 0} \mathbb{E}[|\phi_k(X_{k-1})|] < +\infty$, we have for every positive T and K ,

$$\begin{aligned} \mathbb{P}(\sup_{t \in [0, T]} |Z_t^{(n)}| > K) &\leq \frac{1}{K} \mathbb{E}[\sup_{t \in [0, T]} |Z_t^{(n)}|] \leq \frac{1}{K} \sum_{k=n+1}^{N(n, T)} \gamma_k \sup_{k \geq 0} \mathbb{E}[|\phi_k(X_{k-1})|] \\ &\leq \frac{C(\Gamma_{N(n, T)} - \Gamma_n)}{K} \leq \frac{CT}{K} \end{aligned}$$

thanks to the definition of $N(n, T)$. Hence, for every positive T ,

$$\mathbb{P}(\sup_{t \in [0, T]} |Z_t^{(n)}| > K) \xrightarrow{K \rightarrow +\infty} 0.$$

Then, according to Theorem VI.3.26 of [6], we have to show that for every positive T , ε and η , there exists $\delta > 0$ and $n_0 \in \mathbb{N}$ such that

$$\mathbb{P}(\sup_{|t-s| \leq \delta, 0 \leq s \leq t \leq T} |Z_t^{(n)} - Z_s^{(n)}| \geq \varepsilon) \leq \eta \quad \forall n \geq n_0.$$

In fact, using for instance proof of Theorem 8.3 of [2], it suffices to show that for every positive ε, η and T , there exists $\delta > 0$ and $n_0 \geq 1$ such that:

$$\frac{1}{\delta} \mathbb{P}(\sup_{t \leq s \leq t+\delta} |Z_t^{(n)} - Z_s^{(n)}| \geq \varepsilon) \leq \eta \quad \forall n \geq n_0 \quad \text{and} \quad 0 \leq t \leq T. \quad (20)$$

By (19), for every $t \in [0, T]$ and $s \in [t, t + \delta]$,

$$|Z_t^{(n)} - Z_s^{(n)}| \leq C \sum_{k=N(n,t)+1}^{N(n,s)} \gamma_k |\phi_k(X_{k-1})| \leq C \sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k |\phi_k(X_{k-1})|.$$

Let $p > 1$ such that $\sup_{k \geq 1} \mathbb{E}[|\phi_k(X_{k-1})|^p] < +\infty$ and set $\alpha_k = \gamma_k^{1-\frac{1}{p}}$ and $\mu_k = \gamma_k^{\frac{1}{p}} |\phi_k(X_{k-1})|$. We derive from the Holder inequality with $\bar{p} = \frac{p}{p-1}$ and $\bar{q} = p$ that

$$\sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k |\phi_k(X_{k-1})| \leq \left(\sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k \right)^{\frac{p-1}{p}} \left(\sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k |\phi_k(X_{k-1})|^p \right)^{\frac{1}{p}}.$$

It follows from the Markov inequality that

$$\mathbb{P}\left(\sup_{t \leq s \leq t+\delta} |Z_t^{(n)} - Z_s^{(n)}| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^p} \left(\sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k \right)^{\frac{p-1}{p}+1} \sup_{n \geq 1} \mathbb{E}[|\phi_n(X_{n-1})|^p].$$

Since $\sup_{n \geq 1} \mathbb{E}[|\phi_n(X_{n-1})|^p] < +\infty$, $\sum_{k=N(n,t)+1}^{N(n,t+\delta)} \gamma_k \leq 2\delta$ for sufficiently large n , we deduce that (20) holds for sufficiently small δ . \square

We can now state the main result of this section:

PROPOSITION 2. *Let F be a distribution function of type 1, $(2, \alpha)$ with $\alpha > 2$ or, $(3, \alpha)$ with $\alpha > 0$, and assume $\mathbf{H}_\Lambda^2(F)$ if F is of type 1. Let $(a_n, b_n) \in \mathcal{C}(F)$ such that $(\rho_n)_{n \geq 1}$ and $(\beta_n)_{n \geq 1}$ are bounded. Then, the sequence $(\mathcal{X}^{(n)})_{n \geq 1}$ is tight on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$.*

Proof. Using the convention $\sum_{\emptyset} = 0$, $\mathcal{X}_t^{(n)}$ can be written as follows:

$$\begin{aligned} \mathcal{X}_t^{(n)} &= X_n + D_t^{(n)} + Y_t^{(n)} \quad \text{where,} \\ D_t^{(n)} &= \sum_{k=n+1}^{N(n,t)} \gamma_k \left(\rho_k X_{k-1} + \beta_k + (1 + \rho_k \gamma_k) h_k(X_{k-1}) \right), \\ Y_t^{(n)} &= \sum_{k=n+1}^{N(n,t)} (1 + \rho_k \gamma_k) \Delta \bar{Y}_k \quad \text{with} \quad h_k(x) = \frac{1}{\gamma_k} \mathbb{E} \left[\left(\frac{\xi_1}{a_{k-1}} - x - \frac{b_{k-1}}{a_{k-1}} \right)_+ \right] \quad \text{and,} \\ \Delta \bar{Y}_k &= \left(\frac{\xi_k}{a_{k-1}} - X_{k-1} - \frac{b_{k-1}}{a_{k-1}} \right)_+ - \gamma_k h_k(X_{k-1}). \end{aligned}$$

By Lemma 1, we can assume that there exists $\delta < x_F$ such that $F(x) = 0$ for every $x < \delta$. Then, in this case, it follows from Lemma 3 and from the assumptions on α when F is of type $(2, \alpha)$ that $(Y_t^{(n)})_{n \geq 1}$ is a square-integrable martingale. We have

$$\begin{aligned} \langle Y^{(n)} \rangle_t &= \sum_{k=n+1}^{N(n,t)} (1 + \rho_k \gamma_k)^2 \mathbb{E} \left[(\Delta \bar{Y}_k)^2 / \mathcal{F}_{k-1} \right] \\ &= \sum_{k=n+1}^{N(n,t)} (1 + \rho_k \gamma_k)^2 \left(\mathbb{E} \left[\left(\frac{\xi_k}{a_{k-1}} - X_{k-1} - \frac{b_{k-1}}{a_{k-1}} \right)_+^2 / \mathcal{F}_{k-1} \right] - \gamma_k^2 h_k^2(X_{k-1}) \right). \end{aligned}$$

Then, it is standard that if we want to show that $(\mathcal{X}^{(n)})_{n \geq 1}$ is tight on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$ it suffices to prove that $(D^{(n)})_{n \geq 1}$ and $(\langle Y^{(n)} \rangle)_{n \geq 1}$ are C -tight (see e.g. [6]). Let us show this last assertion. First, we observe that

$$\begin{aligned} D_t^{(n)} &= \sum_{k=n+1}^{N(n,t)} \gamma_k \phi_{k,1}(X_{k-1}) \quad \text{and} \quad \langle Y^{(n)} \rangle_t = \sum_{k=n+1}^{N(n,t)} \gamma_k \phi_{k,2}(X_{k-1}) \quad \text{with,} \\ \phi_{k,1}(x) &= \rho_k x + \beta_k + (1 + \rho_k \gamma_k) h_k(x) \quad \text{and,} \\ \phi_{k,2}(x) &= \frac{(1 + \rho_k \gamma_k)^2}{\gamma_k} \left(\mathbb{E} \left[\left(\frac{\xi_1}{a_{k-1}} - x - \frac{b_{k-1}}{a_{k-1}} \right)_+^2 \right] - \gamma_k^2 h_k^2(x) \right). \end{aligned}$$

Let $\varepsilon > 0$ and $\delta_\varepsilon < x_F$ such that Lemma 2 holds and assume that $F(x) = 0$ for every $x < \delta_\varepsilon$. Since (ρ_k) and (β_k) are bounded, we derive from Lemma 2 and 3 that

$$\phi_{k,1}(X_{k-1}) \leq C(1 + |X_{k-1}|) + C \begin{cases} e^{-\lambda X_{k-1}} & (\lambda > 0) & \text{if } F \text{ is of type 1,} \\ |X_{k-1}|^{-(\alpha+\varepsilon)+1} & & \text{if } F \text{ is of type (2, } \alpha), \\ |X_{k-1}|^{\alpha+\varepsilon+1} & & \text{if } F \text{ is of type (3, } \alpha), \end{cases}$$

and

$$\phi_{k,2}(X_{k-1}) \leq C \begin{cases} (1 + e^{-\lambda X_{k-1}}) & (\lambda > 0) & \text{if } F \text{ is of type 1,} \\ (1 + |X_{k-1}|^{-(\alpha+\varepsilon)+2}) & & \text{if } F \text{ is of type (2, } \alpha), \\ (1 + |X_{k-1}|^{\alpha+\varepsilon+2}) & & \text{if } F \text{ is of type (3, } \alpha). \end{cases}$$

Then, we derive from Lemma 3 that $\sup_{k \geq 1} \mathbb{E}[|\phi_{k,i}(X_{k-1})|^2] < +\infty$ for $i = 1, 2$ and it follows from Lemma 4 that $(D^{(n)})_{n \geq 1}$ and $(\langle Y^{(n)} \rangle)_{n \geq 1}$ are C -tight. \square

4 Characterization of the limit

Denote by \mathcal{A}_ρ^β the operator defined on $\mathcal{C}_K^1(D_G)$ by

$$\mathcal{A}_\rho^\beta f(x) = (\rho x + \beta) f'(x) + \int_0^{+\infty} (f(x+y) - f(x)) \varphi_G(x+y) dy. \quad (21)$$

The main objective of this section is to show that every weak limit \mathcal{X}^∞ of $(\mathcal{X}^{(n)})_{n \geq 1}$ solves the martingale problem $(\mathcal{A}_\rho^\beta, \nu_G, \mathcal{C}_K^1(D_G))$, *i.e.* we want to prove that if \mathcal{X}^∞ exists, the two following properties are satisfied: $\mathcal{L}(\mathcal{X}_0^\infty) = \nu_G$ and for every $f \in \mathcal{C}_K^1(D_G)$, (M_t^f) defined by

$$M_t^f = f(\mathcal{X}_t^\infty) - f(\mathcal{X}_0^\infty) - \int_0^t A_\rho^\beta f(\mathcal{X}_s^\infty) ds$$

is a martingale. These properties are obtained in Proposition 3. The main tool for this result is Lemma 6 where we show that the increments of X_n are ‘‘asymptotically equal’’ to those of the limit process. In this way, we need to establish the following lemma.

LEMMA 5. *Let F be of type 1, (2, α) ou (3, α) and let $(a_n, b_n) \in \mathcal{C}(F)$. Then,*

$$\sup_{x \in K} \left(\frac{1 - F(a_n x + b_n)}{1 - F(\theta_n)} - \tau_G(x) \right) \xrightarrow{n \rightarrow +\infty} 0 \quad \text{for any compact subset } K \text{ of } D_G.$$

Proof. Assume first that F is of type 1. Using that $a_k \sim g(\theta_k)$ and $b_k - \theta_k = o(g(\theta_k))$ when $k \rightarrow +\infty$, we derive from the Von Mises Representation (see (9)) that for sufficiently large k ,

$$\frac{1 - F(a_k x + b_k)}{1 - F(\theta_k)} = \frac{c(a_k x + b_k)}{c(\theta_k)} \exp \left(- \int_0^{x + \varepsilon_1(k,x)} \frac{g(\theta_k)}{g(\theta_k + s g(\theta_k))} ds \right) \quad (22)$$

where $c(x) \xrightarrow{x \nearrow x_F} c > 0$ and $(\varepsilon_1(k, \cdot))_{k \geq 1}$ is a sequence of functions which converges locally uniformly on \mathbb{R} to 0. According to Lemmas 1.2 p.40 and 1.3 p.41 of [15], the sequences of functions $(x \mapsto a_k x + b_k)_{k \geq 1}$ and $(x \mapsto \frac{g(\theta_k)}{g(\theta_k + xg(\theta_k))})_{k \geq 1}$ converge locally uniformly on \mathbb{R} , to x_F and 1 respectively. The result follows easily in this case.

Suppose now that F is of type $(2, \alpha)$ and consider a compact subset K of $(0, +\infty)$. Using that $a_k \sim \theta_k$ and $b_k = o(\theta_k)$ when $k \rightarrow +\infty$, the Karamata Representation (see (12)) yields for sufficiently large k ,

$$\frac{1 - F(a_k x + b_k)}{1 - F(\theta_k)} = \frac{c(a_k x + b_k)}{c(\theta_k)} \exp \left(- \int_1^{x + \varepsilon_2(k, x)} \frac{\alpha(\theta_k s)}{s} ds \right).$$

where $c(x) \xrightarrow{x \rightarrow +\infty} c > 0$, $\alpha(x) \xrightarrow{x \rightarrow +\infty} \alpha > 0$ and $(\varepsilon_2(k, \cdot))_{k \geq 1}$ is a sequence of functions which converges uniformly on K to 0. The result follows in this case from the fact that $\theta_n \xrightarrow{n \rightarrow +\infty} +\infty$.

Finally, if F is of type $(3, \alpha)$, one considers a compact subset K of \mathbb{R}_- . Using that $a_k \sim x_F - \theta_k$ and $b_k - x_F = o(a_k)$, one derives from Corollary 1.14 p.62 of [15] that for sufficiently large k and for every $x \in K$,

$$\frac{1 - F(a_k x + b_k)}{1 - F(\theta_k)} = \frac{c(a_k x + b_k)}{c(\theta_k)} \exp \left(- \int_{-1}^{x + \varepsilon_3(k, x)} \frac{\alpha((x_F - \theta_k)v + x_F)}{v} dv \right).$$

where $c(x) \xrightarrow{x \nearrow x_F} c > 0$, $\alpha(x) \xrightarrow{x \nearrow x_F} \alpha > 0$ and $(\varepsilon_3(k, \cdot))_{k \geq 1}$ is a sequence of functions which converges uniformly on K to 0. Therefore, the result follows from the fact that $x_F - \theta_n \xrightarrow{n \rightarrow +\infty} 0$. \square

LEMMA 6. *Let F be of type 1, $(2, \alpha)$ with $\alpha > 2$ or $(3, \alpha)$ and let $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ such that (ρ_n) and (β_n) converge to some finite numbers ρ and β respectively and such that the assertions of Lemmas 2 and 3 hold. Then, for every $f \in \mathcal{C}_K^1(D_G)$,*

$$\mathbb{E}[f(X_{n+1}) - f(X_n) / \mathcal{F}_n] = \gamma_{n+1} \mathcal{A}_\rho^\beta f(X_n) + \gamma_{n+1} R_n$$

where \mathcal{A}_ρ^β is defined by (21) and (R_n) is an (\mathcal{F}_n) -adapted sequence such that $R_n \rightarrow 0$ in L^1 .

Proof. First, we show that (R_n) defined for every $n \in \mathbb{N}$ by

$$R_n := \frac{1}{\gamma_{n+1}} \mathbb{E}[f(X_{n+1}) - f(X_n) / \mathcal{F}_n] - \mathcal{A}_\rho^\beta f(X_n)$$

is uniformly integrable. On the one hand, using that f is a Lipschitz function, it follows from Lemma 2 that

$$V_n := \frac{1}{\gamma_{n+1}} \left| \mathbb{E}[f(X_{n+1}) - f(X_n) / \mathcal{F}_n] \right| \leq C(|\rho_{n+1}| |X_n| + |\beta_{n+1}|) + \begin{cases} (1 + e^{-\lambda X_n}) & (\lambda > 0) & \text{if } F \text{ is of type 1,} \\ (1 + |X_n|^{-(\alpha+\varepsilon)+1}) & & \text{if } F \text{ is of type } (2, \alpha), \\ (1 + |X_n|^{\alpha+\varepsilon+1}) & & \text{if } F \text{ is of type } (3, \alpha). \end{cases}$$

Therefore, since (ρ_n) and (β_n) are bounded, we derive from Lemma 3 that there exists $\eta > 1$ such that

$$\sup_{n \geq 1} \mathbb{E}[|V_n|^\eta] < +\infty.$$

On the other hand, one checks that $A_\rho^\beta f$ is a bounded function. Actually, by an integration by parts, $A_\rho^\beta f$ can be written:

$$\mathcal{A}_\rho^\beta f(x) = (\rho x + \beta)f'(x) + \int_0^{+\infty} f'(x+y)\tau_G(x+y)dy. \quad (23)$$

Therefore, since $f \in \mathcal{C}_1^K(D_G)$,

$$|A_\rho^\beta f(x)| \leq \|f'\|_\infty (|\rho||x|1_{x \in D_G} + |\beta|) + \int_{D_G} \tau_G(u)du < C < +\infty.$$

It follows that there exists $\eta > 1$ such that

$$\sup_{n \geq 1} \mathbb{E}[|R_n|^\eta] < +\infty.$$

This implies in particular that (R_n) is uniformly integrable. For a uniformly integrable sequence, convergence in probability implies convergence in L^1 . One deduces that one only needs to prove that $R_n \rightarrow 0$ in probability.

We decompose the increment $f(X_{n+1}) - f(X_n)$ as follows:

$$f(X_{n+1}) - f(X_n) = f(X_{n,1}) - f(X_n) + f(X_{n+1}) - f(X_{n,1}), \quad (24)$$

where $X_{n,1} = X_n + \gamma_{n+1}(\rho_{n+1}X_n + \beta_{n+1})$. First, by the Taylor formula,

$$f(X_{n,1}) - f(X_n) = \gamma_{n+1}f'(c_{n+1})(\rho_{n+1}X_n + \beta_{n+1}). \quad (25)$$

with $c_{n+1} \in [X_n, X_{n,1}]$. Denoting by \mathcal{L}_1 , the operator defined for every $f \in \mathcal{C}_K^1(\mathbb{R})$ by $\mathcal{L}_1 f(x) = f'(x)(\rho x + \beta)$, one deduces that

$$f(X_{n,1}) - f(X_n) = \gamma_{n+1}\mathcal{L}_1 f(X_n) + \gamma_{n+1}R_{n,1},$$

with

$$R_{n,1} = f'(X_n)((\rho_{n+1} - \rho)X_n + \beta_{n+1} - \beta) + (f'(c_{n+1}) - f'(X_n))(\rho_{n+1}X_n + \beta_{n+1}).$$

Using that f has compact support, that $(\rho_n)_{n \geq 1}$ and $(\beta_n)_{n \geq 1}$ are bounded and that $\gamma_n \xrightarrow{n \rightarrow +\infty} 0$, one checks that there exists $M > 0$ and $n_0 \geq 1$ such that for every x with $|x| \geq M$, for every $\theta \in [0, 1]$ and $n \geq n_0$,

$$f'(x + \theta\gamma_n(\rho_n x + \beta_n)) = f'(x) = 0. \quad (26)$$

Furthermore,

$$\sup_{|x| \in [-M, M]} (|\gamma_n(\rho_n x + \beta_n)|) \xrightarrow{n \rightarrow +\infty} 0.$$

Therefore, since f' is uniformly continuous with compact support, for every $\varepsilon > 0$, there exists $n_\varepsilon \in \mathbb{N}$ such that

$$|R_{n,1}| \leq C(|\rho_{n+1} - \rho| + |\beta_{n+1} - \beta| + \varepsilon).$$

Using that $\rho_n \rightarrow \rho$ and $\beta_n \rightarrow \beta$, it follows that

$$R_{n,1} \xrightarrow{n \rightarrow +\infty} 0 \quad a.s.$$

Let us now focus on the second term of the right-hand member of (24). We decompose it as follows:

$$\begin{aligned} f(X_{n+1}) - f(X_{n,1}) &= \Delta_{n+1,1} - \Delta_{n+1,2} + \Delta_{n+1,3} \quad \text{with,} \\ \Delta_{n+1,1} &= f\left(X_n + \left(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n}\right)_+\right) - f(X_n), \\ \Delta_{n+1,2} &= \left(f(X_{n,1}) - f(X_n)\right) 1_{\left\{\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} > 0\right\}} \quad \text{and,} \\ \Delta_{n+1,3} &= \left(f(X_{n+1}) - f\left(X_n + \left(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n}\right)_+\right)\right) 1_{\left\{\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} > 0\right\}}. \end{aligned}$$

Denote by \mathcal{L}_2 , the operator defined on $\mathcal{C}_K^1(D_G)$ by

$$\begin{aligned} \mathcal{L}_2 f(x) &= \int_0^{+\infty} \left(f(x+y) - f(x)\right) \varphi_G(x+y) dy \\ &= \int_0^{+\infty} f'(x+y) \tau_G(x+y) dy. \end{aligned} \quad (27)$$

In order to prove the lemma, it suffices now to show the three following points:

- (i) $\mathbb{E}[\Delta_{n+1,1}/\mathcal{F}_n] = \gamma_{n+1} \mathcal{L}_2 f(X_n) + \gamma_{n+1} R_{n,2}$ with $R_{n,2} \xrightarrow{\mathbb{P}} 0$.
- (ii) $R_{n,3} := \frac{1}{\gamma_{n+1}} \mathbb{E}[\Delta_{n+1,2}/\mathcal{F}_n] \xrightarrow{\mathbb{P}} 0$,
- (iii) $R_{n,4} := \frac{1}{\gamma_{n+1}} \mathbb{E}[\Delta_{n+1,3}/\mathcal{F}_n] \xrightarrow{\mathbb{P}} 0$.

(i) First, since ξ_{n+1} is independent of \mathcal{F}_n , we have

$$\mathbb{E}[\Delta_{n+1,1}/\mathcal{F}_n] = \psi_n(X_n) \quad \text{with} \quad \psi_n(x) = \mathbb{E}\left[f\left(x + \left(\frac{\xi_{n+1}}{a_n} - x - \frac{b_n}{a_n}\right)_+\right) - f(x)\right]$$

and a simple transformation yields

$$\begin{aligned} \psi_n(x) &= \int_0^{+\infty} f'(x+u) \mathbb{P}\left(\frac{\xi_{n+1}}{a_n} - x - \frac{b_n}{a_n} > u\right) du \\ &= \int_0^{+\infty} f'(x+u) (1 - F(a_n(x+u) + b_n)) du. \end{aligned}$$

Hence, it follows from (27) that

$$\mathbb{E}[\Delta_{n+1,1}/\mathcal{F}_n] = \gamma_{n+1} \mathcal{L}_2 f(X_n) + \gamma_{n+1} R_{n,2} \quad \text{where,} \quad R_{n,2} = \int_0^{+\infty} H_n(X_n, u) du,$$

$$\text{with} \quad H_n(x, u) = f'(x+u) \left(\frac{1 - F(a_n(x+u) + b_n)}{1 - F(\theta_n)} - \tau_G(x+u)\right).$$

Since $f \in \mathcal{C}_K^1(D_G)$, there exists a compact subset K_G of D_G such that $f'(x) = 0$ when $x \in K_G^c$. Then, $H_n(x, u) = 0$ as soon as $x + u \in K_G^c$.

Therefore, one derives from Lemma 5 that for every $\varepsilon > 0$, there exists $n_0 \in \mathbb{N}$ such that for every $n \geq n_0$, *a.s.*,

$$|H_n(X_n, u)| \leq \varepsilon 1_{\{(X_n+u) \in K_G\}} \|f'\|_\infty \implies |R_{n,2}| \leq \varepsilon \lambda(K_G) \|f'\|_\infty,$$

where $\|f'\|_\infty = \sup_{\{x \in K_G\}} |f'(x)|$ and λ denotes the Lebesgue measure. Hence, $R_{n,2} \xrightarrow{n \rightarrow +\infty} 0$ *a.s.*

(ii) First, since f has compact support, a similar argument as that used for (26) shows that there exists $M > 0$ and $n_0 \in \mathbb{N}$ such that for every x satisfying $|x| > M$, for every $\theta \in [0, 1]$ and $n \geq n_0$,

$$f'(x + \theta \gamma_n(\rho_n x + \beta_n)) = 0.$$

Therefore, since f is a Lipschitz continuous function and the sequences (ρ_n) and (β_n) are bounded, one obtains that

$$|f(X_{n,1}) - f(X_n)| \leq C \gamma_{n+1} 1_{\{|X_n| \leq M\}} \quad \forall n \geq n_0.$$

Hence,

$$\begin{aligned} |R_{n,3}| &\leq C \mathbb{P}\left(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} > 0 / \mathcal{F}_n\right) = C(1 - F(a_n X_n + b_n)) \\ &= C(1 - F(\max(\xi_1, \dots, \xi_n))) \xrightarrow{n \rightarrow +\infty} 0 \quad \textit{a.s.} \end{aligned} \quad (28)$$

(iii) We consider $R_{n,4}$. Since f is a bounded Lipschitz continuous function,

$$\begin{aligned} &|f(X_{n+1}) - f(X_n + (\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n})_+)| 1_{\{\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} > 0\}} \\ &\leq C \gamma_n |(\beta_n + \rho_n X_n) + \rho_n (\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n})_+| 1_{\{\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n} > 0\}}. \end{aligned}$$

It follows that

$$|R_{n,4}| \leq C(1 + |X_n|)(1 - F(a_n X_n + b_n)) + C \mathbb{E} \left[(\frac{\xi_{n+1}}{a_n} - X_n - \frac{b_n}{a_n})_+ / \mathcal{F}_n \right].$$

On the one hand, Lemmas 2 and 3 imply that the second term of the right-hand member converges to 0 in L_1 . On the other hand, for the first term, we use the Skorokhod Representation Theorem. By Proposition 1, $(X_n)_{n \geq 1}$ converges weakly to a random variable X_∞ with distribution ν_G . Hence, by the Skorokhod Representation Theorem, one can construct (\tilde{X}_n) and \tilde{X}_∞ on a probability space $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$, such that $\mathcal{L}(\tilde{X}_n) = \mathcal{L}(X_n)$, $\mathcal{L}(\tilde{X}_\infty) = \mathcal{L}(X_\infty)$ and $\tilde{X}_n \xrightarrow{n \rightarrow +\infty} \tilde{X}_\infty$ *a.s.* Since $\text{supp}(\nu_G) \subset D_G$, it implies that *a.s.*, there exists $n_0(\omega)$ such that $(\tilde{X}_n(\omega))_{n \geq n_0(\omega)}$ is contained in a compact set $K_G(\omega)$ of D_G . Hence, using Lemma 5 and the fact that $(\tilde{X}_n(\omega))$ is bounded, one obtains:

$$(1 + |\tilde{X}_n|)(1 - F(a_n \tilde{X}_n + b_n)) \xrightarrow{n \rightarrow +\infty} 0 \quad \textit{a.s.}$$

It follows that

$$(1 + |X_n|)(1 - F(a_n X_n + b_n)) \xrightarrow{\mathbb{P}} 0 \quad \text{as } n \rightarrow +\infty.$$

This completes the proof. \square

PROPOSITION 3. *Let F be of type 1, $(2, \alpha)$ with $\alpha > 2$ or $(3, \alpha)$ and let $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ such that (ρ_n) and (β_n) converge to some finite numbers ρ and β respectively. Assume $\mathbf{H}_\Lambda^2(F)$ if F is of type 1. Suppose that $(\mathcal{X}^{(n)})$ admits at least one weak limit denoted by \mathcal{X}^∞ . Then, \mathcal{X}^∞ solves the martingale problem $(\mathcal{A}_\rho^\beta, \nu_G, \mathcal{C}_K^1(D_G))$. Moreover, if existence and uniqueness hold for this martingale problem, \mathcal{X}^∞ is a stationary Markov process.*

Proof. First, by Lemma 1, we can choose $\delta < x_F$ such that $F(x) = 0$ for every $x \leq \delta$ and such that the assertions of Lemmas 2 and 3 hold. As a consequence, we can assume in the sequel that the conclusions of Lemma 6 hold.

Let $f \in \mathcal{C}_K^1(D_G)$. By Lemma 6, for every $n \geq 2$,

$$\sum_{k=1}^n \mathbb{E}[f(X_k) - f(X_{k-1})/\mathcal{F}_{k-1}] = \sum_{k=2}^n \gamma_k \mathcal{A}_\rho^\beta f(X_{k-1}) + \sum_{k=2}^n \gamma_k R_{k-1}.$$

where $(R_n)_{n \geq 1}$ is an (\mathcal{F}_n) -adapted sequence such that $R_n \xrightarrow{n \rightarrow +\infty} 0$ in L^1 . Denoting by $(M_n)_{n \geq 0}$ the martingale defined by

$$M_n = \sum_{k=1}^n (f(X_k) - f(X_{k-1}) - \mathbb{E}[f(X_k) - f(X_{k-1})/\mathcal{F}_{k-1}]),$$

one obtains that

$$f(X_n) = f(X_1) + \sum_{k=2}^n \gamma_k \mathcal{A}_\rho^\beta f(X_{k-1}) + \sum_{k=2}^n \gamma_k R_{k-1} + M_n.$$

It follows that for every $n \geq 2$,

$$f(\mathcal{X}_t^{(n)}) - f(\mathcal{X}_0^{(n)}) = M_t^{(n)} + \sum_{k=n+1}^{N(n,t)} \gamma_k (\mathcal{A}_\rho^\beta f(X_{k-1}) + R_{k-1}),$$

where $(M^{(n)})$ is the sequence of martingales defined by $M_t^{(n)} = M_{N(n,t)} - M_n$. One observes that

$$\sum_{k=n+1}^{N(n,t)} \gamma_k \mathcal{A}_\rho^\beta f(X_{k-1}) = \int_0^t \mathcal{A}_\rho^\beta f(\mathcal{X}_s^{(n)}) ds + \left(\sum_{k=n+1}^{N(n,t)} \gamma_k - t \right) \mathcal{A}_\rho^\beta f(\mathcal{X}_t^{(n)}).$$

Hence,

$$f(\mathcal{X}_t^{(n)}) - f(\mathcal{X}_0^{(n)}) - \int_0^t \mathcal{A}_\rho^\beta f(\mathcal{X}_s^{(n)}) ds = M_t^{(n)} + \tilde{R}_t^{(n)}, \quad (29)$$

where

$$\tilde{R}_t^{(n)} = \left(\sum_{k=n+1}^{N(n,t)} \gamma_k - t \right) \mathcal{A}_\rho^\beta f(\mathcal{X}_t^{(n)}) + \sum_{k=n+1}^{N(n,t)} \gamma_k R_{k-1}$$

First, using that f belongs to $\mathcal{C}_K^1(D_G)$ and (27), one checks that \mathcal{A}_ρ^β is a bounded uniformly continuous function. Then, using Theorem VI.1.14 of [6], it follows that $H^f : \mathbb{D}(\mathbb{R}_+, \mathbb{R}) \rightarrow \mathbb{D}(\mathbb{R}_+, \mathbb{R})$ defined by

$$(H^f(\alpha))_t = f(\alpha_t) - f(\alpha_0) - \int_0^t \mathcal{A}_\rho^\beta f(\alpha_s) ds$$

is continuous for the Skorokhod topology. Therefore, if $(\mathcal{X}^{(n_k)})$ denotes a convergent subsequence and \mathcal{X}^∞ its limit, $H^f(\mathcal{X}^{(n_k)})$ and $(M^{(n_k)} + \tilde{R}^{(n_k)})$ converge weakly to $H^f(\mathcal{X}^\infty)$.

Second, by the definition of $N(n, t)$ and the fact that (γ_n) decreases,

$$\left(\sum_{k=n+1}^{N(n,t)} \gamma_k - t \right) \leq \gamma_{N(n,T)+1} \leq \gamma_n.$$

Then, since $A_\rho^\beta f$ is bounded and $R_n \rightarrow 0$ in L^1 , we obtain that for every positive T ,

$$\mathbb{E} \left[\sup_{t \in [0, T]} |\tilde{R}_t^{(n)}| \right] \leq C\gamma_n + T \sup_{k \geq n} \mathbb{E}[|R_k|] \xrightarrow{n \rightarrow +\infty} 0.$$

Therefore, it follows from Lemma VI.3.31 of [6] that $(M^{(n_k)})$ converges weakly to $H^f(\mathcal{X}^\infty)$. Using that $(M^{(n_k)})$ is a sequence of martingales with uniformly bounded jumps (since f is bounded), we derive from Corollary IX.1.19 of [6] that $H^f(\mathcal{X}^\infty)$ is a local martingale for every $f \in \mathcal{C}_K^1(D_G)$. Furthermore, since f and $A_\rho^\beta f$ are bounded, H^f is a true martingale. Hence, \mathcal{X}^∞ solves the martingale problem $(\mathcal{A}_\rho^\beta, \mu, \mathcal{C}_K^1(D_G))$ where $\mu = \mathcal{L}(\mathcal{X}_0^\infty)$. Furthermore, since for every t , $(\mathcal{X}_t^{(n)})_{n \geq 1}$ is a subsequence of $(X_n)_{n \geq 1}$, one derives from Proposition 1 that \mathcal{X}_t^∞ has distribution ν_G for every t . This involves that, on the one hand, $\mathcal{L}(\mathcal{X}_0^\infty) = \nu_G$ and that, on the other hand, $\mathcal{L}(\mathcal{X}_t^\infty) = \mathcal{L}(\mathcal{X}_0^\infty)$ for every $t \geq 0$. Therefore, if existence and uniqueness hold for the martingale problem $(A_\rho^\beta, \nu_G, \mathcal{C}_K^1(D_G))$, then \mathcal{X}^∞ is a Markov process that is stationary since it satisfies $\mathcal{L}(\mathcal{X}_t^\infty) = \mathcal{L}(\mathcal{X}_0^\infty)$. \square

5 Proof of Theorem 1

By Proposition 2, we know that under the assumptions of Theorem 1, $(\mathcal{X}^{(n)})$ is tight and then admits at least one weak limit. Denoting by \mathcal{X}^∞ one of these weak limits, we derive from Proposition 3 that if (a_n, b_n) is a sequence that belongs to $\mathcal{C}(F)$ such that (ρ_n) and (β_n) converge to some finite numbers ρ and β , \mathcal{X}^∞ solves the martingale problem $(\mathcal{A}_\rho^\beta, \nu_G, \mathcal{C}_K^1(D_G))$. In particular, existence holds for this martingale problem under the preceding assumptions on the sequence (a_n, b_n) . Then, it also follows from Proposition 3 that if moreover, uniqueness holds for this martingale problem, $(\mathcal{X}^{(n)})$ converges weakly to \mathcal{X}^∞ for the Skorokhod topology on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$ and that \mathcal{X}^∞ is a stationary Markov process. Then, the reader can check that proving Theorem 1 comes now to show the four following assertions:

- (i) There exists $(a_n, b_n) \in \mathcal{C}(F)$ such that (ρ_n) and (β_n) converge to some finite numbers ρ and β .
- (ii) ρ and β satisfy (6).
- (iii) If Theorem 1 holds for a particular $(a_n, b_n) \in \mathcal{C}(F)$, then it holds for every $(a_n, b_n) \in \mathcal{C}(F)$.
- (iv) Uniqueness holds for the martingale problem $(\mathcal{A}, \nu_G, \mathcal{C}_K^1(D_G))$ where $\mathcal{A} = \mathcal{A}_{\rho_G}^{\beta_G}$.

(i) When F is of type 1, this assertion is exactly Assumption $\mathbf{H}_\Lambda^1(F)$. When F is of type $(2, \alpha)$, we know that $(\theta_n, 0) \in \mathcal{C}(F)$. With this choice, we have first that $\beta_n = 0$. Second, since $1 - F$ is a non-decreasing regularly varying function with index $-\alpha$, $(1/(1 - F))^\leftarrow$ is also regularly varying with index $1/\alpha$ (see *e.g.* [15], Proposition 0.8, p. 23). Then, the Karamata representation yields:

$$\theta_n = c(n) \exp \left(\int_1^n \frac{\rho(s)}{s} ds \right)$$

where $c(x) \xrightarrow{x \rightarrow +\infty} c > 0$ and $\rho(x) \xrightarrow{x \rightarrow +\infty} 1/\alpha$. Set $a_n := c \exp(\int_1^n (\rho(s)/s) ds)$. Since $a_n \sim \theta_n$, $(a_n, 0) \in \mathcal{C}(F)$. Furthermore,

$$\frac{a_{n-1}}{a_n} = \exp \left(- \int_{n-1}^n \frac{\rho(s)}{s} ds \right) = 1 - \frac{1}{\alpha n} + o\left(\frac{1}{n}\right) \quad \text{as } n \rightarrow +\infty.$$

This involves that $\rho_n \xrightarrow{n \rightarrow +\infty} -1/\alpha$. Finally, when F is of type $(3, \alpha)$, some very close arguments lead to the existence of a sequence $(a_n, b_n) \in \mathcal{C}(F)$ such that $\beta_n = 0$ for every $n \geq 1$ and $\rho_n \xrightarrow{n \rightarrow +\infty} 1/\alpha$.

(ii) If F is of type $(2, \alpha)$ or $(3, \alpha)$, this assertion has been shown in (i). Then, let us suppose that F is of type 1. By Proposition 3, ν_Λ is an invariant distribution of \mathcal{A}_ρ^β . It follows that

$$\int \mathcal{A}_\rho^\beta f(x) \nu_\Lambda(dx) = 0 \quad \forall f \in \mathcal{C}_K^1(\mathbb{R}) \quad (\text{see } e.g. [3]).$$

Hence, for every $f \in \mathcal{C}_K^1(\mathbb{R})$, (ρ, β) is solution to

$$\begin{aligned} a_1(f, \nu_\Lambda) \rho + a_2(f, \nu_\Lambda) \beta + a_3(f, \nu_\Lambda, \Phi_\Lambda) &= 0 \quad \text{with} & (30) \\ a_1(f, \nu_\Lambda) &= \int x f'(x) \nu_\Lambda(dx), \quad a_2(f, \nu_\Lambda) = \int f'(x) \nu_\Lambda(dx) \quad \text{and,} \\ a_3(f, \nu_\Lambda, \Phi_\Lambda) &= \int \int_0^{+\infty} [f(x+y) - f(x)] \Phi_\Lambda(x+y) dy \nu_\Lambda(dx). \end{aligned}$$

By Remark 1, we deduce from some particular cases (exponential and normal distributions) that $(\rho, \beta) = (0, -1)$ is a solution to this system. Then, let us show that this is the only one: the fact that $(\rho, \beta) = (0, -1)$ is a solution to (30) involves in particular that $a_3(f, \nu_\Lambda, \Phi_\Lambda) = a_2(f, \nu_\Lambda)$. Hence, Equation (30) can be reduced to:

$$a_1(f, \nu_\Lambda) \rho + a_2(f, \nu_\Lambda) (\beta + 1) = 0 \quad \forall f \in \mathcal{C}_K^1(\mathbb{R}). \quad (31)$$

We now exhibit a function f such that $a_1(f, \nu_\Lambda) \neq 0$ and $a_2(f, \nu_\Lambda) = 0$. For every $t \geq 0$, denote by f_t the function defined by:

$$f_t(x) = \begin{cases} 1 + \cos(x) & \text{if } x \in [-\pi, 0], \\ 2 & \text{if } x \in [0, t], \\ 1 + \cos(x - t) & \text{if } x \in [t, t + \pi], \\ 0 & \text{otherwise.} \end{cases}$$

For every $t \geq 0$, f_t belongs to $\mathcal{C}_K^1(\mathbb{R})$. First, one observes that $x f_t'(x) \leq 0$ for every $x \in \mathbb{R}$ and that $x f_t'(x) < 0$ on $(-\pi, 0) \cup (t, t + \pi)$. Therefore, $a_1(f_t, \nu_\Lambda) < 0$

for every $t \geq 0$. Second, on the one hand, one can check that $\Lambda'(x) > \Lambda'(-x)$ for every $x > 0$. Then, using that $f'_0(x) = -f'_0(-x)$ for every $x \in \mathbb{R}$, it follows that $a_2(f_0, \nu_\Lambda) = \int f'_0(x)\Lambda'(x)dx < 0$. On the other hand, since $\Lambda'(x) \rightarrow 0$ when $x \rightarrow +\infty$, we derive that $a_2(f_t, \nu_\Lambda) > 0$ for sufficiently large t . It follows from the continuity of $t \rightarrow a_2(f_t, \nu_\Lambda)$ that there exists $t_0 > 0$ such that $a_2(f_{t_0}, \nu_\Lambda) = 0$ (and $a_1(f_{t_0}, \nu_\Lambda) < 0$). We deduce that $\rho = 0$ and then, that $\beta = -1$.

(iii) Let $(a_n, b_n)_{n \geq 1} \in \mathcal{C}(F)$ such that $(\mathcal{X}^{(n)})$ converges in distribution on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$ to \mathcal{X}^∞ and consider another sequence $(\tilde{a}_n, \tilde{b}_n)_{n \geq 0} \in \mathcal{C}(F)$. Set $\tilde{X}_n = (M_n - \tilde{b}_n)/\tilde{a}_n$. Since $a_n \sim \tilde{a}_n$ and $b_n - \tilde{b}_n = o(a_n)$ when $n \rightarrow +\infty$, we have:

$$X_n - \tilde{X}_n = \frac{M_n - b_n}{a_n} \left(1 - \frac{a_n}{\tilde{a}_n}\right) + \frac{\tilde{b}_n - b_n}{\tilde{a}_n} = \varepsilon_1(n)X_n + \varepsilon_2(n),$$

where $\varepsilon_1(n) \xrightarrow{n \rightarrow +\infty} 0$ and $\varepsilon_2(n) \xrightarrow{n \rightarrow +\infty} 0$. Then, we derive from the tightness of $(\mathcal{X}^{(n)})$ on $\mathbb{D}(\mathbb{R}_+, \mathbb{R})$ that, for every $\eta > 0$, for every $T > 0$,

$$\mathbb{P} \left(\sup_{k=n}^{N(n,T)} |\varepsilon_1(k)X_k| > \eta \right) \leq \mathbb{P} \left(\sup_{t \in [0, T]} |\mathcal{X}_t^{(n)}| > \frac{\eta}{\inf_{k \geq n} |\varepsilon_1(k)|} \right) \xrightarrow{n \rightarrow +\infty} 0.$$

Therefore, denoting by $\tilde{\mathcal{X}}^{(n)}$ the càdlàg process defined on \mathbb{R}_+ by $\tilde{\mathcal{X}}_t^{(n)} = \tilde{X}_{N(n,t)}$, it follows that for every $\eta > 0$, for every $T > 0$,

$$\mathbb{P} \left(\sup_{t \in [0, T]} |\mathcal{X}_t^{(n)} - \tilde{\mathcal{X}}_t^{(n)}| > \eta \right) \xrightarrow{n \rightarrow +\infty} 0.$$

Hence, $(\tilde{\mathcal{X}}^{(n)})$ converges weakly to \mathcal{X}^∞ (see *e.g.* [6], Lemma 3.31, p. 352).

(iv) First, by a change of variable, one obtains:

$$\mathcal{A}f(x) = (\rho_G x + \beta_G)f'(x) + \int_0^{+\infty} \left(f(x + (z - x)_+) - f(x) \right) \varphi_G(z) dz.$$

Then, \mathcal{A} can be viewed as the infinitesimal generator associated with the SDE:

$$dX_t = (\rho_G X_{t-} + \beta_G)dt + \int (z - X_{t-})_+ \varphi_G(z) N(dt, dz), \quad (32)$$

where N is a Poisson random measure with intensity measure $dt \otimes \lambda(dz)$. More precisely, if (X_t) is a solution to (32) with initial value X_0 , one can check that (X_t) is a solution to the martingale problem (\mathcal{A}, μ) where $\mu = \mathcal{L}(X_0)$. Now, by Lepeltier and Marchal ([12], Corollary II.10 and Theorem II.13), uniqueness for the martingale problem $(\mathcal{A}, \delta_{x_0})$ holds if pathwise uniqueness holds for SDE (32) with initial value x_0 . Therefore, if we want to prove uniqueness for the martingale problem (\mathcal{A}, ν_G) , we need only to show that for every $x_0 \in \text{supp}(\nu_G)$, pathwise uniqueness holds for the SDE with initial value x_0 .

Assume first that F is of type 1. Since $u \rightarrow (u)_+$ is a Lipschitz continuous function which is null on \mathbb{R}_- , one obtains that for every $N > 0$, for every $x, y \in [-N, N]$,

$$\int [(z - x)_+ - (z - y)_+]^2 (\varphi_G(z))^2 dz \leq C|x - y|^2 \int_{-N}^{+\infty} \exp(-2z) dz \leq C_N|x - y|^2.$$

Therefore, pathwise uniqueness for (32) follows for every $x_0 \in \mathbb{R}$ from [12] (Theorem III.3).

Second, suppose that F is of type $(2, \alpha)$. On the one hand, $x \rightarrow \rho_G x + \beta_G := -\alpha^{-1}x$ is a Lipschitz continuous function. On the other hand, for every positive ε and N such that $0 < \varepsilon < N$, for every $x, y \in [\varepsilon, N]$,

$$\int [(z-x)_+ - (z-y)_+]^2 (\varphi_G(z))^2 dz \leq C|x-y|^2 \int_\varepsilon^N \frac{\alpha}{z^{1+\alpha}} dz \leq C_{\varepsilon, N}|x-y|^2.$$

Now, since there are only positive jumps, any solution of (32) with initial value $x_0 \in \text{supp}(\nu_G) = (0, +\infty)$ is a positive process. Therefore, pathwise uniqueness follows again from Theorem III.3 of [12].

Finally, we assume that F is of type $(3, \alpha)$. Using that $X_{t^-} + (z - X_{t^-})_+ \leq z$ and that $\varphi_G(z) = 0$ when $z > 0$, a similar argument yields the result in this case.

References

- [1] Basak G.K., Hu I.C., Wei C.Z. (1997), Weak convergence of recursions, *Stochastic Process. Appl.*, **68**, pp. 65-82.
- [2] Billingsley, P. (1968) *Convergence of Probability Measures*, Wiley.
- [3] Ethier S., Kurtz T. (1986), *Markov processes, characterization and convergence*, Wiley series in probability and mathematical statistics, Wiley, New York.
- [4] Fahrner, I., Stadtmüller, U. (1998), On almost sure max-limit theorems, *Statist. Probab. Lett.*, **37**, pp. 229-236.
- [5] Gnedenko B.V. (1943), Sur la distribution limite du terme maximum d'une série aléatoire, *Ann. Math.*, **44**, pp. 423-453.
- [6] Jacod, J., Shiryaev, A. N. (1987), *Limit Theorems for Stochastic Processes*, Springer.
- [7] Lambertson D., Pagès G. (2002), Recursive computation of the invariant distribution of a diffusion, *Bernoulli*, **8**, pp. 367-405.
- [8] Lambertson D., Pagès G. (2003), Recursive computation of the invariant distribution of a diffusion: the case of a weakly mean reverting drift, *Stoch. Dynamics*, **4**, pp. 435-451.
- [9] Lamperti J. (1964), On extreme order statistics, *Ann. Math. Statist.*, **35**, pp. 1726-1737.
- [10] Leadbetter M., Lindgren G., Rootzen H. (1983), *Extremes and Related Properties of Random Sequences and Processes*, Springer, New York.
- [11] Lemaire V. (2006), An adaptative scheme for the approximation of dissipative systems, *Stochastic Process. Appl.*, 117, pp. 1491-1518.

- [12] Lepeltier J.P., Marchal B. (1976), Problème de martingales et équations différentielles stochastiques associées à un opérateur intégro-différentiel (French), *Annales de l'I.H.P.*, **12**, pp. 43-103.
- [13] Panloup F. (2008), Recursive computation of the invariant measure of a SDE driven by a Lévy process, *Annals Appl. Probab.*, **18**, pp. 379-426.
- [14] Resnick, S. (1975), Weak convergence to extremal processes, *Ann. Probab.*, **3**, pp. 951-960.
- [15] Resnick, S. (1987), *Extreme Values, Regular Variation and Point Processes*, Springer, New York.