



HAL
open science

Accounting for global-mean warming and scaling uncertainties in climate change impact studies: application to a regulated lake system

B. Hingray, N. Mouhous, A. Mezghani, K. Bogner, B. Schaefli, A. Musy

► To cite this version:

B. Hingray, N. Mouhous, A. Mezghani, K. Bogner, B. Schaefli, et al.. Accounting for global-mean warming and scaling uncertainties in climate change impact studies: application to a regulated lake system. *Hydrology and Earth System Sciences Discussions*, 2007, 11 (3), pp.1207-1226. hal-00305662

HAL Id: hal-00305662

<https://hal.science/hal-00305662>

Submitted on 18 Jun 2008

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

Accounting for global-mean warming and scaling uncertainties in climate change impact studies: application to a regulated lake system

B. Hingray, N. Mouhous, A. Mezghani, K. Bogner, B. Schaefli and A. Musy

Laboratory of Hydrology and Land Improvement, Ecole Polytechnique Fédérale de Lausanne, Switzerland

Email for corresponding author: benoit.hingray@gmail.com

Abstract

A probabilistic assessment of climate change and related impacts should consider a large range of potential future climate scenarios. State-of-the-art climate models, especially coupled atmosphere-ocean general circulation models and Regional Climate Models (RCMs) cannot, however, be used to simulate such a large number of scenarios. This paper presents a methodology for obtaining future climate scenarios through a simple scaling methodology. The projections of several key meteorological variables obtained from a few regional climate model runs are scaled, based on different global-mean warming projections drawn in a probability distribution of future global-mean warming. The resulting climate change scenarios are used to drive a hydrological and a water management model to analyse the potential climate change impacts on a water resources system. This methodology enables a joint quantification of the climate change impact uncertainty induced by the global-mean warming scenarios and the regional climate response. It is applied to a case study in Switzerland, a water resources system formed by three interconnected lakes located in the Jura Mountains. The system behaviour is simulated for a control period (1961–1990) and a future period (2070–2099). The potential climate change impacts are assessed through a set of impact indices related to different fields of interest (hydrology, agriculture and ecology). The results obtained show that future climate conditions will have a significant influence on the performance of the system and that the uncertainty induced by the inter-RCM variability will contribute to much of the uncertainty of the prediction of the total impact. This research has been conducted within the 2001–2004 EU funded project SWURVE.

Keywords: climate change impact, regional response pattern scaling, Monte Carlo simulation, hydrological modelling, uncertainty analysis, impact analysis, Switzerland.

Introduction

The assessment of climate change impacts on water resources is currently subject to intense research (see, e.g. Jasper *et al.*, 2004; Kim, 2005; Lettenmaier *et al.*, 1999; Loukas *et al.*, 2002). The quantification of these impacts is conditional on the availability of regional climate change predictions of different key meteorological variables such as precipitation and surface air temperature. Currently available regional climate changes projections are generally based on the results of coupled Atmosphere–Ocean General Circulation Models (AOGCMs) or on the results of Regional Climate Models (RCMs) that are driven by the outputs of the former and that are supposed to describe regional climatic variables better than AOGCMs because of their higher spatial resolution (0.5° latitude \times 0.5° longitude or higher).

Regional climate change projections based on climate model outputs are highly uncertain, mainly due to the unknown future greenhouse gas emissions but also due to the highly simplified representation of reality encoded in these models. As a consequence, different state-of-the-art AOGCMs generally simulate different climate evolutions for the same emission scenario (see, e.g. Arnell and Hulme, 2000; Räisänen, 2001, 2002). For a given AOGCM experiment, the corresponding RCM experiment is also subject to these modelling uncertainties. The results of several RCM experiments based on the same AOGCM outputs can therefore also differ significantly (see, e.g. Frei *et al.*, 2003; Räisänen *et al.*, 2004). Nevertheless, the uncertainty introduced by the RCM is generally considered to be substantially smaller than the one inherited by the driving AOGCM (Jenkins and Lowe, 2003). As part of the

2001–2004 EU project SWURVE (Sustainable Water, Uncertainty, Risk and Vulnerability in Europe, see Kilsby *et al.*, 2007), an analysis of the data of the EU project PRUDENCE (Prediction of Regional scenarios and Uncertainties for Defining European Climate change risks and Effects (Christensen *et al.*, 2002) suggests, however, that the inter-RCM variability cannot be neglected (Ekström *et al.*, 2007; Hingray *et al.*, 2007).

An evaluation of the full range of potential climate change impacts on a given water resources system would require taking into account the uncertainties due to different emission scenarios and due to the climate model structure and the corresponding parameterization (for both AOGCMs and RCMs). This would need a large number of AOGCM-RCM experiments with different emission scenarios. However, climate model simulations being highly time-consuming, only a few model experiments are currently available. Consequently, most impact studies are reduced to a scenario analysis: meteorological variables are obtained for a limited number of climate models and of probable greenhouse gas emission scenarios. Such approaches can neither simulate all possible system outcomes nor assign them a probability of occurrence.

Recent work addresses the assignment of such probabilities of occurrence through Monte Carlo simulation frameworks (New and Hulme, 2000; Jones, 2000; Prudhomme *et al.*, 2003), combining some of the major sources of uncertainty such as the emission scenarios, the global climate sensitivity and the inter-climate model variability. However, most of the studies currently available take no account of all sources of modelling uncertainty. Prudhomme *et al.* (2003), for example, include different Global Circulation Model (GCM) patterns and climate sensitivities but their analysis is based on only four emission scenarios defined by the Intergovernmental Panel on Climate Change (IPCC) in its Special Report on Emission Scenarios (SRES) (Nakienovi and Swart, 2000).

The scaling approach introduced by Santer *et al.* (1990) enables the generation of multiple climate change scenarios for a number of global-mean warming projections. The technique is currently widely used in climate scenario studies (New and Hulme, 2000; Jones, 2000; Hulme *et al.*, 2002) and is based on the assumption of a linear relationship between the annual global-mean warming (the so-called scaler) and the response pattern of regional climate changes obtained from any global or regional climate model. In the work of Jones (2000) and of New and Hulme (2000), the response pattern is expressed for each variable as a simple scaling ratio, i.e. a constant regional temperature or precipitation change per degree of global-mean warming. The response pattern used by Mitchell (2003) is the spatial

climate change pattern of a GCM. The advantage of this scaling technique is that the scaler can be obtained from so-called Simple Climate Models (SCMs) demanding fewer computing resources than GCMs (Wigley *et al.*, 2000). Based on such SCMs, thousands of Monte Carlo simulations can be produced (New and Hulme, 2000; Huntingford and Cox, 2000). Wigley and Raper (2001) present such an approach to determine the Probability Distribution Function (PDF) of global-mean warming between 1990 and different periods in the 21st century. A simple upwelling-diffusion energy balance model is calibrated against the results of seven state-of-the-art AOGCMs. The emission uncertainties as well as the uncertainties due to the climate sensitivity, the carbon cycle, the ocean vertical diffusivity and the aerosol forcing are characterised by appropriate PDFs, which are then used to drive the SCM resulting in PDFs of the global-mean warming. SCMs are, however, not sufficiently complex to represent with accuracy spatial or seasonal patterns of climate changes, especially if regional features are needed for small space and time resolutions. This problem can be overcome by a technique of pattern scaling that combines SCM and RCM (or GCM) outputs to produce a number of climate change scenarios (Mitchell, 2003).

The methodology presented in this paper is based on such a scaling approach. It was developed during the 2001–2004 EU project SWURVE. Future climate scenarios — in terms of seasonal changes of the mean value and of the variability of temperature and precipitation — are obtained for the period 2070–2099 by scaling the meteorological response pattern from a RCM by a global-mean warming projection randomly drawn from a PDF obtained from Wigley and Raper (2001). The uncertainty inherent in the regional climate response is taken into account through the use of different meteorological response patterns derived from a set of RCM experiments available through the PRUDENCE project (Christensen *et al.*, 2002).

The corresponding local scale time series of surface air temperature and precipitation are obtained through a classical perturbation procedure (Shabalova *et al.*, 2003). These series are used as an input to simulate the behaviour of a water resources system for future climate change conditions. The performance of the resulting system is compared with a control period based on an appropriate set of performance indices. This methodology is applied to a managed lake system located in the north-west of Switzerland. The control period analysed is 1961–1990 and the future period 2070–2099. These 30-year periods are assumed to cover a large part of the natural variability of the system and to be long enough for a climate change impact to be assessed.

In this paper, after introducing the models developed to simulate the behaviour of the system and indices of the impact of climate change, the climatic datasets and the climate scenarios are presented, followed by a discussion of the probabilistic framework for stochastic climate scenario generation. Finally, the main results and the relative importance of the different sources of uncertainty in climate modelling are discussed.

Case study: a regulated lake system

SYSTEM DESCRIPTION

The water resources system studied in the north-west of Switzerland (Fig. 1) comprises three interconnected lakes, formed 15 000 years ago during the retreat of the Rhone glacier. The system was modified significantly in the periods 1833-1884 and 1963-1972 by the construction of connection channels (Jaton, 1989). Different regulation waterworks have been built to reduce the flood risk and to transform the formerly unproductive areas into fertile arable land. The interconnection of the three lakes equilibrates their water levels and the mean daily level of all of them is roughly the

same. The catchment surface area of the three lakes is around 8300 km² (Fig. 1) and its elevation ranges from 420 m a.s.l. (at the outlet of lake Biemme) to around 4270 m a.s.l. with a mean elevation of 1150 m a.s.l. The three main inflows into the lakes are the rivers Aare, Broye and Orbe (Fig. 1, and see Table 1 for catchment characteristics).

HYDROLOGICAL REGIME

The hydrological regimes of the Aare and Broye rivers located in the Swiss Alps and Prealps are influenced significantly by snowmelt; less than 3% of the Aare catchment is covered by glaciers. The Orbe River and several other tributaries to Lakes Biemme and Neuchatel are in the Swiss Jura Mountains; consequently, their hydrological regimes are influenced by snowmelt and karst effects. The hydrological regime of all these rivers has a marked seasonality, with low flows in winter and high flows from spring to autumn due to snowmelt and rainfall. Artificial lakes (mainly used for hydropower production) and regulated natural lakes perturb the river regimes, especially that of the Aare River (Spreatico *et al.*, 1992).

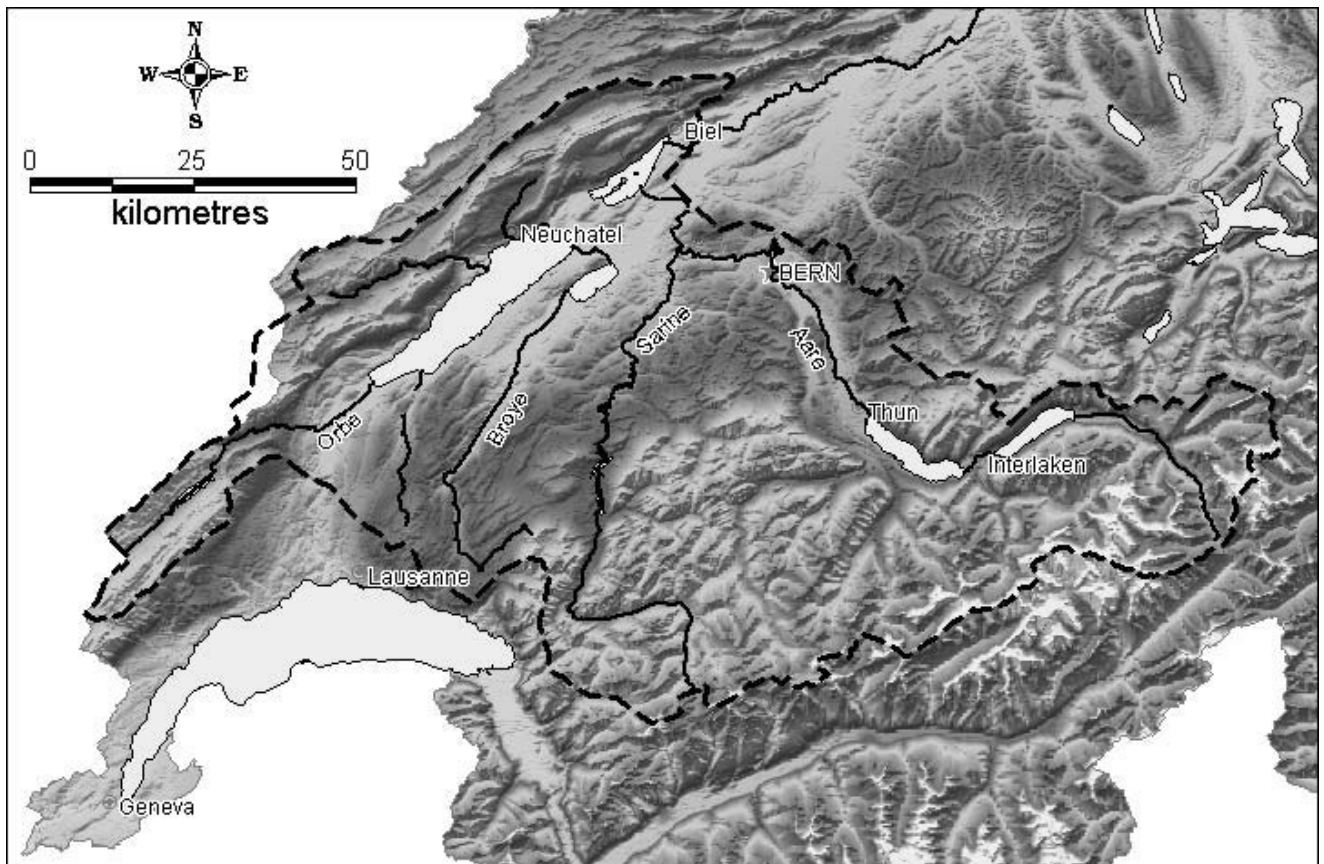


Fig. 1. Location of the lake system studied in the Swiss Alps (SwissTopo, 1997)

Table 1. Characteristics of the main rivers feeding the studied lakes system

	<i>Aare River</i>	<i>Broye River</i>	<i>Orbe River</i>
Catchment area	5128 km ²	392 km ²	333 km ²
Gauge	Hageneck	Payerne	Orbe
Outlet	Lake of Biemme	Lake of Morat	Lake of Neuchâtel
Elevation range	450–4250 m a.s.l.	450–2000 m a.s.l.	450–1500 m a.s.l.
Hydrological regime influenced by	Snowmelt Lake regulation Hydropower	Snowmelt	Karsts Snowmelt Hydropower

LAKE REGULATION

The only discharge of the lake system, at Nidau-Port (Lake Biemme), is regulated by a weir and water is released according to ‘règlement 80/82’ approved by the Swiss Federal Government (VAW, 1996). Daily target releases are fixed as a function of the water level in the lakes and of the calendar day. These rules have been elaborated to take account of the sometimes conflicting interests of the different stakeholders involved in farming, fishery, hydropower production, flood protection and ecosystem conservation (including wetland protection). The resulting water level is low in winter and high in summer. The maximum annual water level fluctuation is, however, less than 1 m.

POTENTIAL FAILURES AND CLIMATE CHANGE IMPACTS

For most stakeholders, potential system failures are related directly to the water level of the lakes. The main failures occur when the lake level is above a critical level because of an unexpected or long-lasting high water inflow. Such situations can lead to water-logging and flooding of agricultural areas (particularly damaging during the sowing season) and riparian urban areas (Jaton, 1989). Other failure events are induced by persistent droughts. Such situations can lead to economic losses for agriculture (water restriction and irrigation targets unachievable) or for the hydropower plants located downstream of the lake. In addition, the welfare of many animal and vegetal species depends on the temporal progression of the water level throughout the year. Pikes, for example, need a sudden spring flood to access their spawning area and their alevins need a slow flood decrease to return to the deeper water (Zaugg *et al.*, 1994). Highly specialised birds build their nests in the reeds just above the maximum water level experienced during the days preceding the brooding. A significant water level increase

after nest construction floods some of them (Aebischer, 1994).

A potential climate change is expected to modify the hydrological regime of the different rivers feeding the lakes. The resulting modified temporal evolution of the lake levels potentially influences the frequency, duration and intensity of failures occurring.

System modelling

DATA COLLECTION

An integrated simulation tool, developed to simulate the performance of the system under different climates, includes hydrological models to simulate river discharges feeding the lakes with a management model for the simulation of the lake level evolution. The development and calibration of these models is based on observations. The observed precipitation and surface air temperature time series for the hydrological models are obtained from the Swiss Meteorological Institute at a daily time step. Potential evapotranspiration is estimated based on the Penman-Monteith model (Monteith and Unsworth, 1990) and direct evaporation from the lake surface is estimated by the Turc formula (Mermoud, 1995). Observations of river discharge and lake level are obtained from the Swiss Federal Office of Water and Geology. Based on these data, the hydrological and management models were calibrated for the years 1984–1992 and validated for the years 1993–2000.

HYDROLOGICAL MODELLING

The catchment feeding the lake system is split into its three main sub-catchments named Aare, Broye and Jura (the last includes several small rivers as well as the River Orbe). For each of these sub-catchments, a specific hydrological model has been developed. The present modelling context —

namely the simulation of a high number of scenarios and the limited available meteorological data for future scenarios — imposed the development of highly parsimonious discharge models.

All catchments are strongly influenced by snow accumulation and snowmelt processes. In the present simulation framework, these processes are taken into account by computing a rainfall equivalent to the sum of rainfall and simulated snowmelt. For a catchment glacier-covered in parts, this equivalent rainfall should include ice melt but the computation of the equivalent rainfall for the Aare catchment accounts only for snowmelt. This is because, for the control period, the area of the catchment on which ice melt processes occur is insignificant: only 3% of the catchment area is glacierised and the corresponding mean ablation area (the area where ice melt occurs) can be assumed to cover only around one-third of the total glacier area (see, e.g. Maisch *et al.*, 2000).

Hereafter, the snowmelt computation is presented, followed by a short description of the equivalent rainfall-discharge transformation that is specific to each of the three catchments. The daily snow accumulation and snowmelt depend strongly on the mean daily precipitation and temperature, both of which are spatially very variable and altitude-dependent. Given the large elevation ranges of the catchments studied, they were divided into several elevation bands to simulate the snow pack evolution.

For each elevation band, the daily precipitation is divided into rain- and snowfall according to a fuzzy-threshold established by Steel (1999) for the whole of Europe. The daily snowmelt is computed according to the modified degree-day method of Schreider *et al.* (1997). Instead of the constant degree-day factor as in the classical degree-day formulation (Martinec *et al.*, 1983), an albedo-related factor was calculated for each month of the year as a function of the mean monthly temperature. This function was established empirically by Whetton *et al.* (1996) and has been adapted to European conditions by Steel (1999).

The equivalent rainfall is transformed into discharge through an appropriate runoff model for each subcatchment. As the Broye catchment has a rather natural hydrological behaviour, a classical conceptual modelling approach is applied (a lumped rainfall–runoff model), combining a non-linear loss module based on the moisture deficit in the catchment with the routing module of the rainfall–runoff model IHACRES (Jakeman and Hornberger, 1993). The accounting scheme for the non-linear loss module converts equivalent rainfall into rainfall excess, while the routing module relates rainfall excess to total streamflow via two parallel linear reservoirs.

The Aare catchment is highly influenced by water

derivations, i.e. where water is collected from subcatchments and routed via artificial collectors to other subcatchments, and by several regulated natural and artificial lakes. The hydrological regime of the Jura catchment is influenced significantly by karsts. As it is unrealistic to simulate the discharge of these two catchments by simple conceptual models, the Data-Based Mechanistic (DBM) modelling approach proposed by Young and Beven (1994) has been tested. In this approach, the most parsimonious model structure is inferred statistically from the available time series, using the time variable parameter estimation method called Fixed Interval Smoothing (FIS) presented by Young (1984) and Young (1993).

For the Aare and Jura catchments, the DBM modelling approach highlighted the predominating effect of the lagged runoff, resulting in an autoregressive model where the rainfall as driving forces is negligible. Such an approach is, clearly, not applicable in climate change impact studies. The elimination of the autoregressive part of the DBM model led to a linear regression model between the daily discharge and the equivalent rainfall filtered by moving averages. The same filtering is applied to the temperature series in the regression model as a second explanatory variable. For the Aare catchment, the final model has seven parameters, the constants a_p , b_p , b_2 , b_3 , and the window sizes w_i for the moving average filtering:

$$Q_t = a_1 + b_1 \cdot \frac{1}{w_1} \sum_{k=1}^{w_1} P_{t-k}^{Aare} + b_2 \cdot \frac{1}{w_2} \sum_{k=1}^{w_2} T_{t-k} + b_3 \cdot \frac{1}{w_3} \sum_{k=1}^{w_3} P_{t-k}^{Sarine} \quad (1)$$

where Q_t is the simulated discharge on day t and T_{t-k} is the mean temperature for the entire Aare catchment for day $t-k$. P_{t-k}^{Aare} and P_{t-k}^{Sarine} are the area average equivalent rainfall for the two subcatchments upper-Aare and Sarine. A similar multiple regression model has been developed for the Jura catchment.

Using the coefficient of determination (R^2) as an objective-function, the values of R^2 for the calibration period are, respectively, 0.77, 0.68 and 0.89 for the Aare, Jura and Broye catchments; for the validation period the values are 0.67, 0.43 and 0.97. Bearing in mind the parsimony constraints and the highly disturbed hydrological regimes, the performance of the three models is rather good.

MANAGEMENT MODEL

The simulation of the temporal evolution of the water levels is based on a simple water balance model accounting for lake inflow, direct evaporation from the lake surfaces and water release at the unique system outlet. The water routing through the hydraulic system and especially between the

lakes is assumed to be negligible. The water release is simulated according to the regulation rules (VAW, 1996) that determine the release as functions of the observed lake level (estimated through appropriate storage-to-level functions) and of the calendar day. Except for a few extreme situations, the weir manager follows these regulation rules strictly and, consequently, the observed water levels are well simulated (Denk, 2002).

SYSTEM PERFORMANCE MEASUREMENT

Because of the high number of stakeholders directly concerned about the temporal evolution of the lake level, an integrated assessment of all potential impacts of climate change is impossible. Hence, attention has been directed at three particular fields of interest, namely, agriculture (focusing on spring crops), fishery (focusing on pikes), flood protection (focusing on flooding of riparian urban areas) and wetland protection (focusing on the reed population in the Grande Cariçaie ecosystem).

To judge the behaviour of water resources systems under different future climatic situations, a pertinent set of criteria has to be defined. These should be discriminant and easily calculable and they should not be redundant, i.e. there should be no or little correlation between the information encoded in respect of the different criteria. Because the main interest is concerned with the ability of the system to deal with future uncertain climate conditions without severe damage, the robustness of the system i.e. its capacity to absorb adverse conditions without lasting damage, must be one of the criteria. The system may fail but should be able to recover rapidly. Hashimoto *et al.* (1982) have defined the RRV criteria (Reliability, Resilience and Vulnerability); they measure how often a system failure occurs, how severe it is and how quickly the system returns to a satisfactory state afterwards. These criteria can be interpreted as probabilistic measures of the frequency, duration and intensity of occurring failures. They are quantitative risk criteria and, therefore, highly adapted to climate change impact analysis. Several applications appear in the literature (Lane *et al.*, 1999; Vogel *et al.*, 1999; Fowler *et al.*, 2003).

The use of RRV criteria for the lake system studied requires definition of the acceptability of outputs of the system under examination. This task may be straightforward — as for a drinking water supply system for example — or quite difficult — as for agriculture or flood protection in the present case study. For agriculture, high water levels during the sowing season will delay access to the land or may inundate the seeds by water logging after sowing. During the maturation season, flooding over several consecutive days could affect the harvest and reduce the

productivity significantly. For a given field, at a given altitude compared to the lake level, a critical lake water level can be identified. The different fields are however not at the same altitude compared to the lake level.

Similarly, the water level threshold for urban flood damage varies (430.35 m a.s.l., 430.70 m a.s.l., 430.25 m a.s.l. for the lakes of Neuchâtel, Morat and Bienne). The extent of damage can be modelled as a function of the number of buildings and infrastructure affected by the inundation. This depends on the intensity of the flood and on the altitude of the riparian urban zones. As was found for agriculture, a unique critical water threshold is difficult to define.

The analysis of impacts on animal or vegetal species is even more complex. The definition of a proper failure state is difficult mainly because the interaction between the physical system and the studied populations is difficult to measure and can evolve with time. In the present study, potential climate change impacts are quantified through so-called welfare indices. Two different types have been defined: one measures the success of reproduction of pikes and the other the viability of reed. For the pikes, the relationship between the reproduction success and the annual lake level evolution is known (Zaugg *et al.*, 1994). For reed, the long-term viability can be expressed as a function of the maximum fluctuation of the water level (Buttler *et al.*, 1995; Clerc, 1999). All welfare indices have been evaluated on a yearly time step. Based on these annual time series, an intensity–frequency–duration analysis of the occurring events is completed. Such an analysis is comparable to an RRV-analysis even if the events analysed are not necessarily failures.

The same approach has been adopted for analysing potential climate change impacts on all identified fields of interest. In the following, the events analysed have been summarised:

- Reed population: annual maximum water level fluctuation defined as the difference between the annual minimum daily water level and maximum daily water level.
- Pike population: the annual reproduction success score is SO, the product of four different scores S1 to S4, each of which has a value between 0 and 1 and is related to a different aspects of success in reproduction. To reach their spawning areas, pikes need a sudden spring flood, the magnitude of which gives score S1 and the minimum expected variation in level is 30 cm. Score S2 is determined by the arrival date of this spring flood and the best date is 15th February. The final water level, which determines the extent of accessible spawning areas, should reach at least 429.9 m a.s.l. and gives score

S3. Score S4 measures the speed of flood decrease. Pike leveins need a slow decrease in flood to return to deep water, ideally less than 10 cm per 100 days and this rate gives Score S4. If conditions are optimal in a given year, all the scores are set to 1; otherwise, they are decreased according to Zaugg *et al.* (1994).

- Flood protection of urban areas: annual maximum flood is defined as the annual maximum daily water level.
- Spring crop: seasonal failure events are defined for the sowing season (spring) and the maturation season (summer) if the minimum daily water level over n -consecutive days ($n = 2$ for sowing season, $n = 10$ for maturation season) is above a given threshold H_{s1} that is expected to produce water logging in the sowing period in spring and field inundation in summer when the crops mature.

Pattern scaling and time series generation

GENERATION OF FUTURE TIME SERIES FOR A GIVEN CLIMATE CHANGE SCENARIO

Based on a probability distribution of the future global-mean warming and on available AOGCM and RCM model outputs, the generated future time series must meet the following requirements. (i) They should reflect the seasonality of regional climate changes. For the catchments studied, the mean temperature increase is expected to be higher in summer than in winter (Hingray *et al.*, 2007). (ii) For the meteorological variables (surface air temperature and precipitation), they should reflect changes both in the mean and in variability. Potential climate related impacts are generally highly dependent on extreme events. As the occurrence of extreme events is highly sensitive to the width of statistical distributions, changes in variability are generally supposed to influence their occurrence more than changes in averages (Katz and Brown, 1992; Mearns *et al.*, 1996; Schär *et al.*, 2004). Nevertheless, changes in variability are rarely taken into account in the development of future time series scenarios. As they can be estimated easily from outputs of climate model experiments, they should be integrated in a time series generation framework.

The method of Shabalova *et al.* (2003) is well adapted for the development of such time series scenarios for daily mean precipitation and temperature. Local scale meteorological time series for a given future scenario (here 2070–2099) are generated by perturbing the observed series for a control period (here 1961–1990). The perturbation preserves the changes in mean and variability given by the climate change scenario. This scenario is defined by the

following seasonal statistics: The absolute change of the mean temperature (XMT) and the relative change of the mean precipitation (XMP), of the standard deviation of daily temperature ($XSDT$) and of the coefficient of variation of daily precipitation ($XCVP$). The perturbation equation for temperature is,

$$T_{scen,s}(t) = [T_{obs,s}(t) - MT_{obs,s}] * (XSDT_s + 1) + MT_{obs,s} + XMT_s \quad (2)$$

where $T_{scen,s}(t)$ ($^{\circ}\text{C}$) is the local scale scenario temperature on day t of season s ($s = 1$: DJF; $s = 2$: MAM; $s = 3$: JJA; $s = 4$: SON). $T_{obs,s}(t)$ ($^{\circ}\text{C}$) is the observed temperature on day t of season s and $MT_{obs,s}$ ($^{\circ}\text{C}$) is the observed mean temperature of season s . XMT_s is the absolute change in the mean temperature of season s (defined as the mean temperature of the RCM output for the future period minus the mean temperature of the RCM output for the control period) and $XSDT_s$ is the relative change of the standard deviation of the daily temperature of season s (defined as the difference between the values obtained from the RCM output for the future period and the control period divided by the value for the control period).

The perturbation of the daily rainfall series is described in Appendix 1. A perturbation method similar to that for temperature is not used because this would generate negative rainfall values. Shabalova *et al.* (2003) fit a Weibull distribution to the observed daily rainfall amounts and the Weibull parameters of the perturbed series are computed based on the parameters of the observed distribution according to the relative changes in precipitation mean and variability predicted by the climate scenario.

GENERATION OF CLIMATE CHANGES SCENARIOS

The objective of the generation of multiple climate change scenarios is twofold: (i) to enable the generation of regional climate change scenarios for a large range of possible global-mean warming predictions and (ii) to account for the uncertain regional response to a given global-mean warming scenario (i.e. for the uncertainty related to the inter-RCM variability). Additionally, the scenario generation methodology has to preserve the correlations between the seasonal changes in temperature and precipitation predicted by the RCM experiments.

The prediction uncertainty due to the global-mean warming has been highlighted in previous studies and has been characterised, for example, by Wigley and Raper (2001). The uncertainty due to inter-model variability has been described for the mean values of different key meteorological variables and for their variability (e.g.

Räisänen, 2001; Räisänen, 2002). The temporal correlation between some of the key variables might well be significant. New and Hulme (2000) have highlighted a significant positive correlation between winter precipitation and temperature changes for Scotland and East Anglia. For the case study presented in this paper, the correlation between the four mean seasonal temperature changes is not negligible, partly because these changes are strongly related to the global-mean warming. Another explanation arises from the high correlations between the seasonal scaling ratios obtained from some climate experiments, such as the correlation between seasonal DJF and MAM temperatures shown in Fig. 3). To take these correlations into account, the pattern scaling technique discussed earlier has been adopted.

PATTERN SCALING

The generation of temperature and precipitation time series according to the above perturbation methodology requires regional climate change statistics as input. These statistics correspond to a 16-values matrix $\mathbf{X} = [XMT_s, XSDT_s, XMP_s, XCVP_s]_{s:1..4}$. In the following, a different notation for \mathbf{X} is

$$\mathbf{X} = [X_{v,s}]_{v:1..4, s:1..4} \quad (3)$$

where $X_{v,s}$ is one key statistic v of the daily mean temperature or precipitation series ($X_{1,s} = XMT_s, X_{2,s} = XSDT_s, X_{3,s} = XMP_s, X_{4,s} = XCVP_s$) and where s refers to the season.

Given a RCM called r that has been run for the control period (1961–1990) and the future period (2070–2099), the response pattern is defined as the following matrix

$$\mathbf{Y}_r = [Y_{v,s,r}]_{v:1..4, s:1..4} \quad (4)$$

where $Y_{v,s,r}$ is the scaling ratio for one of the four key statistics used ($Y_{1,s,r} = YMT_{s,r}, Y_{2,s,r} = YSDT_{s,r}, Y_{3,s,r} = YMP_{s,r}, Y_{4,s,r} = YCVP_{s,r}$) for season s and RCM experiment r . For each variable, the scaling ratio is defined as its regional change (absolute or relative) per degree of global-mean warming. For a given RCM experiment r , the scaling ratios are estimated based on

$$Y_{v,s,r} = X_{v,s,r} / \Delta T_r \quad (5)$$

where $X_{v,s,r}$ is the change in variable v for season s predicted by RCM experiment r between the control and the future period and ΔT_r is the global-mean warming value obtained for the AOGCM used to drive RCM experiment r .

Based on the outputs of a RCM experiment, a range of regional climate change scenarios can be generated by

scaling their response patterns according to different global-mean warming predictions drawn in an appropriate global-mean warming PDF. If several RCM experiments are available, a range of scenarios can be generated for each of the corresponding response patterns and, accordingly, the uncertainty of the regional climate response to a given global-mean warming scenario can be quantified.

The pattern scaling technique is based on the critical assumption that there is a linear relationship between the scaler (annual global-mean warming) and the response pattern of regional climate changes obtained from the climate model. For a given climate model (in the present case for an AOGCM/RCM configuration), the validity of this assumption can be analysed if several climate experiments are available for different radiative forcing scenarios. Mitchell *et al.* (1999) and Mitchell (2003) have examined this assumption for spatial changes in mean temperature and precipitation from ten climate change experiments obtained with the HadCM2 GCM. Their results indicate that for the model studied, pattern scaling could be applicable to a wide range of variables. For the present work, not more than two climate change experiments are available for each AOGCM/RCM configuration. The reliability of the assumption of linearity is, thus, difficult to check. Figure 2 presents regional changes in temperature and precipitation versus global-mean warming for six AOGCM/RCM configurations having climate experiments for both A2 and B2 emission scenarios. From these experiments, the linearity assumption seems to hold well for changes in mean seasonal temperatures and precipitation but the relationship is weaker for changes in temperature and precipitation variability. In the absence of any further climate experiments, it was decided to retain the assumption of linearity for all the variables studied. Note that it would be difficult to assess the reliability of this linear scaling assumption from all climate experiments available in this work. Non-negligible deviations from a mean linear scaling relationship are actually expected (e.g. Fig. 2 in Hingray *et al.*, 2007). Such deviations do not, however, necessarily contradict the assumed linear scaling relationship between regional changes and global-mean warming. They could be also due to differences between the models or to the effects of the natural variability in the climate model response.

DATASET: REGIONAL RESPONSE PATTERNS OF 26 RCMS

The present study has used a set of 26 RCMs experiments, each comprising simulations for the control 1961–1990 and for the future period 2070–2099, carried out in the framework of the EU project PRUDENCE (Christensen *et*

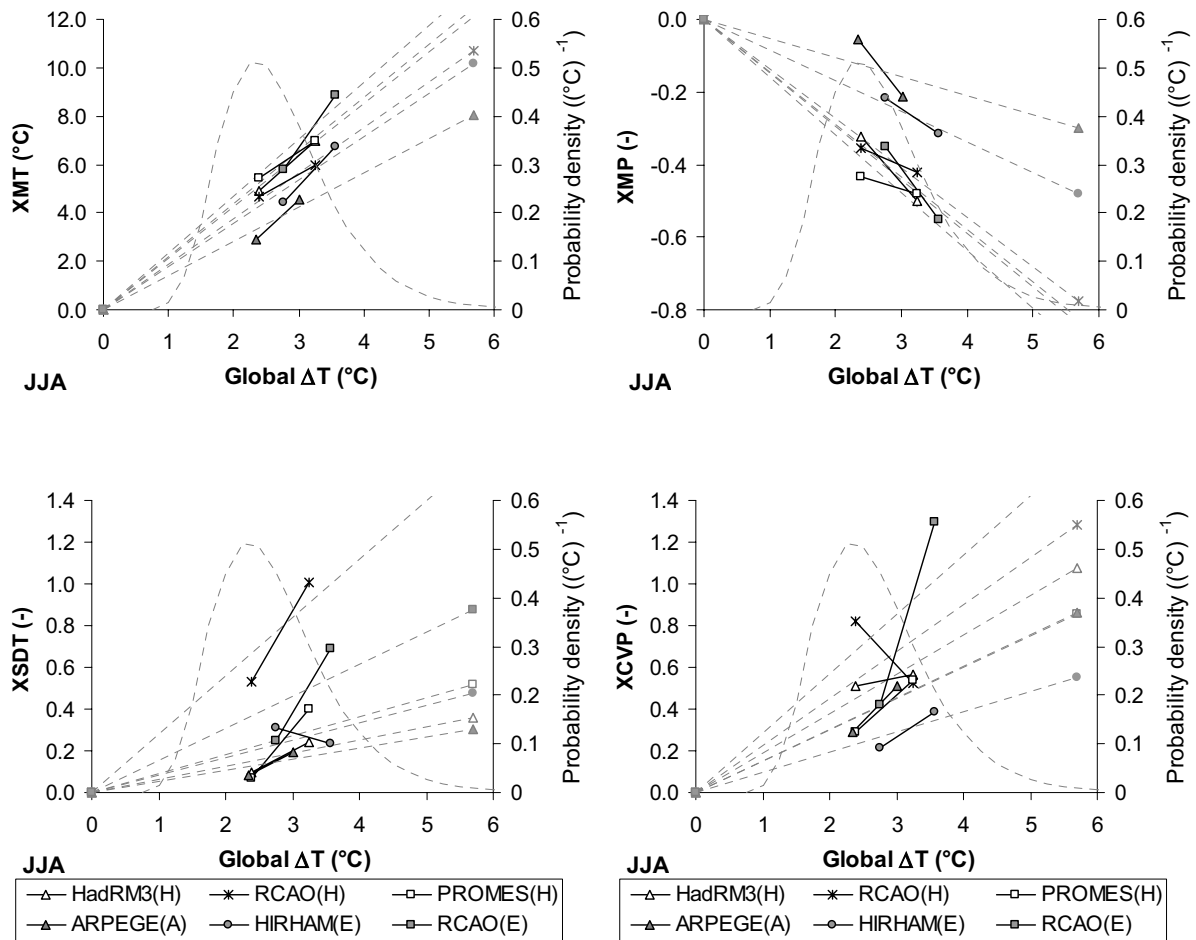


Fig. 2. Scaling relationships between regional climate changes and global-mean warming (ΔT) for the summer season. Left: regional changes in seasonal temperature (top: changes in mean, XMT; bottom: changes in variability, XSDT). Right: regional changes in seasonal precipitation (top: changes in mean, XMP; bottom: changes in variability, XCVP). Each symbol corresponds to one AOGCM/RCM configuration for which experiments are available for both A2 and B2 emission scenarios. If ensemble runs are available, only one run is plotted. For each configuration, the straight line is based on a least-squares fit to the two climate experiments. Linear regressions are forced to pass through the origin. Note that ΔT is not specific to the summer season. The grey dashed curve is the probability density function of global-mean warming from 1990 to the period 2070-2099 obtained from Wigley and Raper (2001).

al., 2002). The boundary conditions for these simulations have been obtained from the three AOGCMs used in the PRUDENCE project: ARPEGE/OPA (Royer *et al.*, 2002), HadCM3 (Gordon *et al.*, 2000; Pope *et al.*, 2000) and ECHAM4/OPYC3 (Roeckner *et al.*, 1999). In the case of HadCM3, use of a global model of the atmosphere alone (HadAM3H) between the global coupled model and the RCMs resulted in a much improved simulation of the present-day climate (Hulme *et al.*, 2002). All AOGCMs have been run for the SRES A2 and SRES B2 emission scenarios (Nakicenovic and Swart, 2000). The corresponding values of global-mean warming are given in Table 2.

For the present study, nine different RCMs were available (see Table 3 for a list of the models and the corresponding modelling institutions). However, one of these models,

ARPEGE, is not a regional but a global atmospheric model with variable horizontal resolution, from 50 km in the centre of the Mediterranean to 450 km in the southern Pacific Ocean (Gibelin and Déqué, 2003). In PRUDENCE, the low boundary conditions (sea surface temperatures, sea ice) have been obtained from ARPEGE/OPA and HadCM3 (Déqué, 2004, *pers. comm.*). For the lake system studied, the horizontal resolution of ARPEGE is comparable with that of the other RCMs and so ARPEGE is treated as an RCM in this paper.

Some RCMs have been run successively using the boundary conditions of different AOGCMs. All RCMs have been run at least for scenario A2 with one AOGCM. Some RCMs have also been run for scenario B2. In total, 19 different AOGCM/RCM experiments are available, 12 of

Table 2. Clusters of the RCM experiments of the PRUDENCE project according to the driving AOGCM and to the emission scenario (ΔT = global-mean warming, * = 2 ensemble runs are available, ** = 3 ensemble runs are available)

Scenario	AOGCM		RCM			
	No.	Name	ΔT	No.	Name	
A2	1	HADCM3/HADAM3H	3.25	1	CHRM(H)	
		HADCM3/HADAM3H	3.25	2	CLM(H)	
		HADCM3/HADAM3H	3.25	3**	HadRM3H(H)	
		HADCM3/HADAM3H	3.25	4*	HIRHAM(H)	
		HADCM3/HADAM3H	3.25	5	PROMES(H)	
		HADCM3/HADAM3H	3.25	6	RCAO(H)	
		HADCM3/HADAM3H	3.25	7	RegCM(H)	
		HADCM3/HADAM3H	3.25	8	REMO(H)	
		HADCM3/ARPEGE	3.25	9**	ARPEGE(H)	
	2	ARPEGE/ OPA	3.02	10	ARPEGE(A)	
3	ECHAM4/OPYC3	ECHAM4/OPYC3	3.56	11	HIRHAM(E)	
		ECHAM4/OPYC3	3.56	12	RCAO(E)	
B2	4	HADCM3/HADAM3H	2.39	13	HadRM3H(H)	
		HADCM3/HADAM3H	2.39	14	PROMES(H)	
		HADCM3/HADAM3H	2.39	15	RCAO(H)	
		HADCM3/ARPEGE	2.39	16	ARPEGE(H)	
	5	ARPEGE/ OPA	2.35	17**	ARPEGE(A)	
	6	ECHAM4/OPYC3	ECHAM4/OPYC3	2.76	18	HIRHAM(E)
			ECHAM4/OPYC3	2.76	19	RCAO(E)

Table 3. (a) the three AOGCM and (b) the nine RCM models used in the PRUDENCE project (Christensen *et al.*, 2002) and the corresponding modelling institutions

Acronym	Institution	AOGCM	Reference
(a)			
CNRM	Centre National de Recherches Météorologiques, Toulouse, France	ARPEGE/OPA	Royer <i>et al.</i> , 2002
HC	Hadley Centre for Climate Prediction and Research, Bracknell, United Kingdom	HadCM3	Gordon <i>et al.</i> , 2000 Pope <i>et al.</i> , 2000
MPI	Max-Planck-Institut für Meteorologie, Hamburg, Germany	ECHAM4/OPYC3	Roegner <i>et al.</i> , 1999
(b)			
CNRM	Centre National de Recherches Météorologiques, Toulouse, France	ARPEGE	Gibelin and Déqué, 2003
DMI	Danish Meteorological Institute, Copenhagen, Denmark	HIRHAM	Christensen <i>et al.</i> , 2001
ETHZ	Institute for Atmospheric and Climate Science, Zurich, Switzerland	CHRM	Vidale <i>et al.</i> , 2003
GKSS	Institute for Coastal Research, Geesthacht, Germany	CLM	Doms and Schättler, 1999
HC	Hadley Centre for Climate Prediction and Research, United Kingdom	HadRM3H	Hulme <i>et al.</i> , 2002
ICTP	International Centre for Theoretical Physics, Trieste, Italy	RegCM	Giorgi <i>et al.</i> , 1993ab
MPI	Max-Planck-Institut für Meteorologie, Hamburg, Germany	REMO	Jacob, 2001
SMHI	Swedish Meteorological and Hydrological Institute, Norrköping, Sweden	RCAO	Räisänen <i>et al.</i> , 2004
UCM	Universidad Complutense de Madrid, Toledo, Spain	PROMES	Arribas <i>et al.</i> , 2003

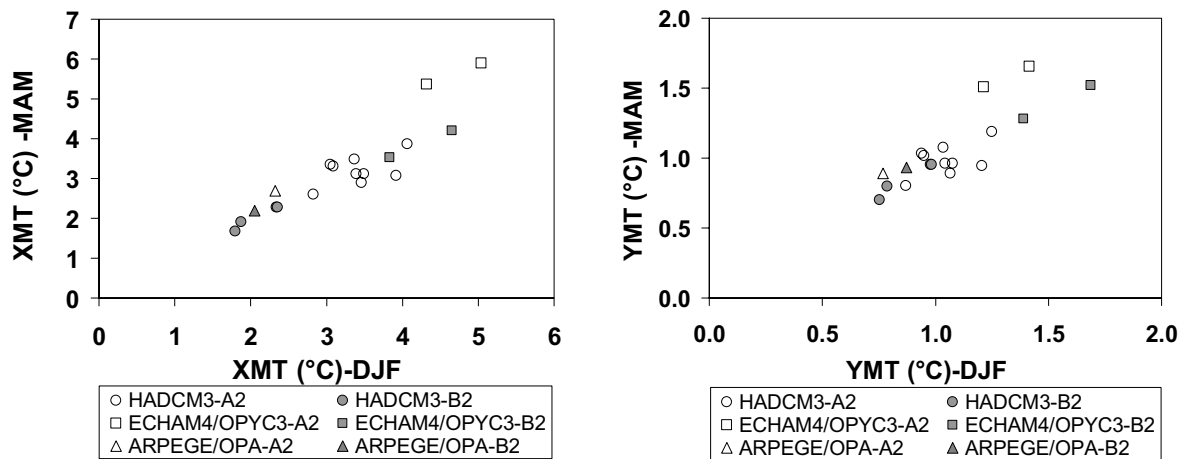


Fig. 3. Left: correlation between regional temperature changes during winter and spring ($XMT_{DJF,r}$ and $XMT_{MAM,r}$); right: correlation between corresponding scaling ratios ($YMT_{s,r} = XMT_{s,r} / \Delta T_r$); data from 19 RCMs obtained from Christensen et al., (2002), if ensemble runs are available, only one RCM is used.

them for A2 and 7 for B2 (Table 2). For some RCMs, ensemble runs exist (Table 2). This leads to a total of 26 response patterns used in the present work.

Within the PRUDENCE project, the seasonal means of surface air temperature (MT) and precipitation (MP) have been produced for each experiment for the whole of Europe on a 0.5° latitude by 0.5° longitude grid. In the absence of standard deviations of daily temperatures (SDT) and coefficients of variation of daily precipitations (CVP), PRUDENCE has produced grids of 90-days' temperature and precipitation variances for each season, obtained for each run (control and future) from the 30-values series (one series for each season) of 90-days mean temperatures (90-days precipitation amounts). These 90-days' statistics have been estimated directly from the RCM output series. For the 30 years from 2070–2099, a significant temperature trend has been simulated that could potentially influence the estimation of variances. For the seasonal variances of precipitation, this influence is probably small but, for the temperature variances, it could be significant. Consequently, the PRUDENCE data cannot be used directly for the perturbation of observed daily series but has to be pre-processed according to the method briefly explained in Appendix 2. For the lake system studied, the regional changes in mean and variability are extracted for the 35 grid boxes encompassing the entire catchment and are averaged over this area. The inter-RCM variability of regional climate changes and of related scaling ratios is given for all seasons in Fig. 4.

Note that the input data required for the hydrological models are daily series of precipitation, temperature, evaporation from free water surfaces and potential

evapotranspiration. The future evaporation and PET series are interpolated as a function of the future temperature based on the observed relationship for the control climate, assuming that this relationship remains constant in the future (strong linear regression relationships have been found between monthly PET and monthly temperature over the period 1961–1990). The pattern scaling approach could be extended to these variables but they were not available.

Simulation framework

Based on the pattern scaling technique, the uncertainty due to global-mean warming and to the regional response patterns can be quantified through Monte Carlo simulations. The global-mean warming is drawn randomly in a PDF for the global temperature change between 1990 and the period 2070–2099 derived from the results of Wigley and Raper (2001) for the years 2070, 2080, 2090 and 2100 (provided by Tom Wigley). The random draw of a response pattern among the 26 patterns available from PRUDENCE is based on a hierarchical resampling approach.

The following considerations motivated the use of this resampling approach: the RCM experiments and corresponding response patterns belong to different classes and sub-classes. For some RCMs, ensemble runs are available. The RCM experiments driven by the same AOGCM can be expected to give similar results and can, thus, be clustered according to the driving AOGCM. The response patterns obtained for the A2 and the B2 emission scenarios can differ significantly due to potentially significant non-linearities in the climate change response to the story lines (see Schneider and Thompson, 1981;

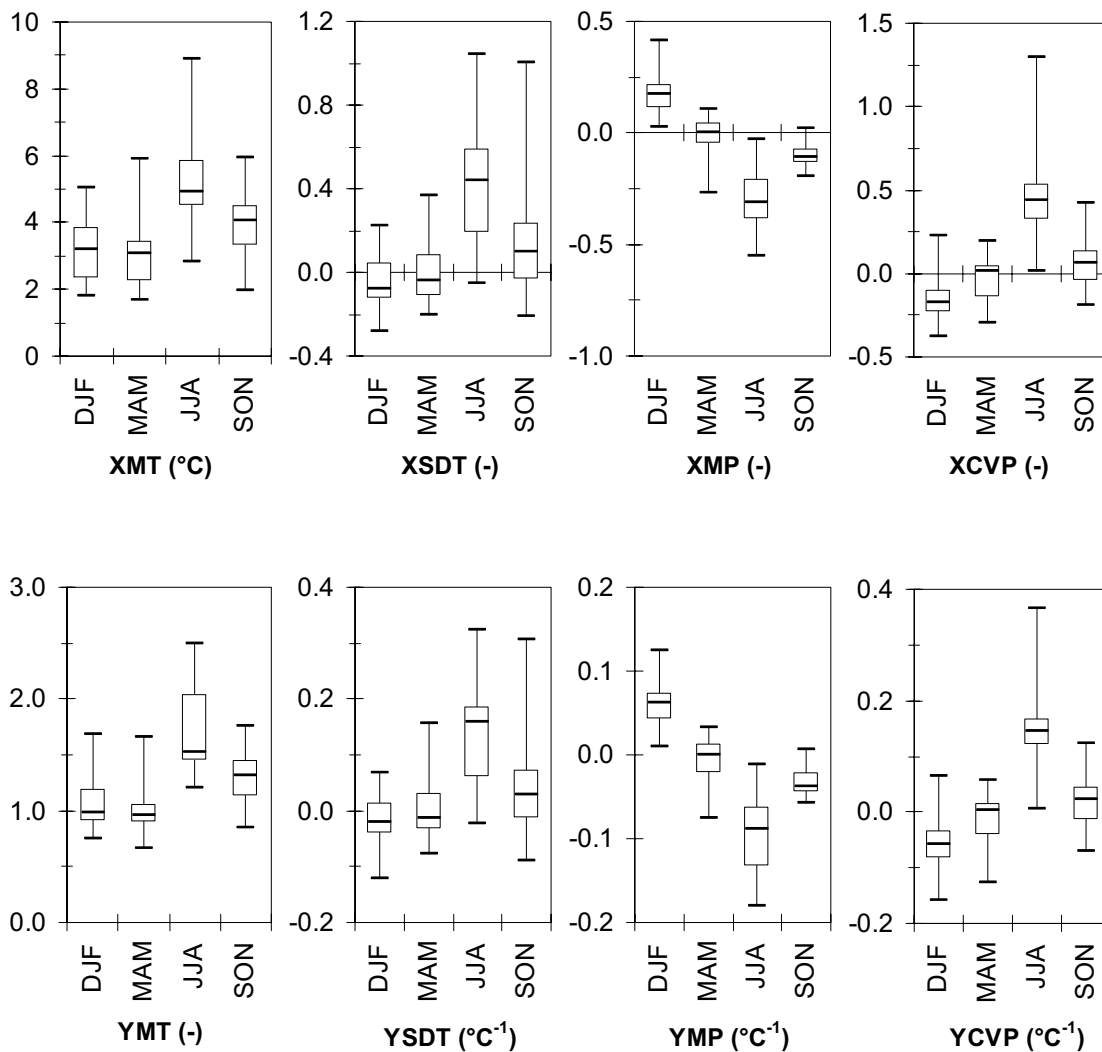


Fig. 4. Boxplots of regional climate changes X (top) and related seasonal scaling ratios Y (bottom) for the 4 key variables MT, SDT, MP and CVP. X and Y variables values are based on PRUDENCE outputs (from 90-days temperature and precipitation series); SDT and CVP values have been corrected before perturbation (see Appendix 2). Each boxplot displays the variation of the seasonal regional climate change and of the scaling ratio of the 26 RCM experiments of the PRUDENCE project (all available ensemble runs are included). The whiskers represent the smallest and the largest values, the box the quartiles and the horizontal bar the median.

Mitchell, 2003). Consequently, the AOGCM clusters can be clustered in turn according to the underlying emission scenario (see Table 2).

The A2 and B2 SRES-scenarios are assumed to have the same probability of occurrence. Similarly, all AOGCM and RCM models are assumed to perform equally well. Although this may not be the case, it is the simplest assumption to make. The number of experiments within each cluster is, however, variable (for example 14 experiments in the cluster HadCM3/HadAM3H-A2, 2 experiments in the cluster ECHAM4/OPYC3-A2; 17 and 9 experiments for A2 and B2 respectively). The assignment of a uniform occurrence probability to all experiments would therefore lead to an over-representation of some experiments. The ARPEGE/

OPA/ARPEGE-A2 experiment for example would be drawn 9 times more frequently than the HADCM3/HADAM3H/CHRM-A2 experiment. A detailed analysis of the data (not discussed here but some results are given in Hingray *et al.*, 2007) has shown that the inter-RCM variability and the inter-AOGCM variability have the same order of magnitude. Consequently, the same occurrence probability is assigned to all AOGCM/RCM experiments of an emission cluster.

Based on the above considerations, the following hierarchical resampling approach is adopted for this work:

- i. Randomly select an emission scenario (A2 or B2).
- ii. Randomly draw an AOGCM/RCM experiment within the selected emission scenario cluster.

iii. If ensemble runs are available, randomly draw the response pattern of one of the runs.

At each resampling step, all possible choices have the same occurrence probability. The uncertainties due to the differences between the emission scenarios and the uncertainties due to the system unpredictability under a given emission scenario are thus sampled uniformly.

The random generation of a response pattern could also be based on a stochastic approach using parametric models. Hingray *et al.* (2007) assumed that the scaling variables from the PRUDENCE dataset follow a normal distribution. The parameters of this distribution can be estimated from the observed scaling ratios using an appropriate unbalanced ANalysis Of VAriance framework (ANOVA). Following such an approach, the probability distributions of the regional temperature and precipitation change can be obtained through Monte Carlo simulation by combining a probability distribution of the global-mean temperature increase with the probability distributions of the appropriate scaling variables. Such PDFs are presented for five case study regions in Hingray *et al.* (2007). As mentioned before, some scaling variables are significantly correlated. A Principal Components Analysis (PCA) can be performed on the scaling variable dataset to produce a dataset of uncorrelated variables $[Z_{k,r}]_{k=1..16; r=1..26}$ where r refers to the RCM experiment and k to the k^{th} component (all components are used for the description of the response pattern Y_r). Each transformed $Z_{k,r}$ variable can be modelled by a normal distribution, the parameters of which are estimated with an ANOVA framework. Finally, a transformed $[Z_{k,r}]_{k=1..16; r=1..26}$ response pattern can be generated stochastically and transformed back to produce a response pattern $[YMT_s, YSDT_s, YMP_s, YCVVP]_{s=1..4}$ with correlated variables. Schaeffli *et al.* (2007) have applied this parametric approach to the generation of climate change scenarios. In the present study, the non-parametric resampling approach is used.

Results

The meteorological time series observed for the control period (1961–1990) are used to produce the reference discharge series for the Aare, Broye and Jura catchments and the corresponding time series of water levels, from which the reference statistics of the welfare indices are estimated. The same simulation procedure is applied for each climate change scenario. Note that the lake regulation rules for the control and the future period are those corresponding to the 1982–2000 period.

A main objective of the present study is to quantify, separately, the climate change prediction uncertainty due to

the global-mean warming and that induced by the regional response pattern. Consequently, climate change scenarios are generated according to the following three different simulation experiments:

- E1: random global-mean warming and random regional response pattern
- E2: random global-mean warming under median regional response pattern
- E3: random regional response pattern under median global-mean warming

For the experiments E1 and E2, 12 000 Monte Carlo simulations were carried out whereas for E3, only 26 simulations, corresponding to the 26 available RCM runs were completed. The E3 scenarios are generated based on the median global-mean warming value as predicted by the PDF adopted for the global-mean warming between 1990 and 2070–2099 ($\Delta T_{E3} = 2.6^\circ\text{C}$). For E2, the median response pattern is derived from the 26 PRUDENCE response patterns. Each scaling variable $Y_{v,s}$ of this median response pattern corresponds to the median of the 26 corresponding scaling variables $Y_{v,s,r}$.

RESULTS FOR E1 EXPERIMENT (OVERALL UNCERTAINTY)

Changes in regional meteorological variables

The simulations results show a higher regional mean warming than the global-mean warming: The median scaling ratio for the mean annual temperature is around $+1.3^\circ\text{C}$ per $^\circ\text{C}$ of global-mean warming. The 5th, 50th and 95th percentiles of the regional warming are $+1.8^\circ\text{C}$, $+3.3^\circ\text{C}$ and $+6.1^\circ\text{C}$. The regional warming in summer is higher, with a larger uncertainty range than the regional warming for the other seasons. The 99th percentile of the summer warming is as large as $+11^\circ\text{C}$ (only around $+7^\circ\text{C}$ for the other seasons). The 5th and 95th percentiles of the relative change of annual precipitation are -16% and $+2\%$. This suggests that the direction of the change is uncertain. Precipitation is, however, expected to increase in winter (up to 32% increase for the 99th percentile) and to decrease in summer and autumn (decrease of -24% in autumn and -68% in summer for the 1st percentile). No trend is predicted for spring precipitation: an increase is as likely as a decrease (-22% for the 5th percentile, $+11\%$ for the 95th percentile (Hingray *et al.*, 2007)).

The predicted range of daily temperature and precipitation variability change is large, especially for the summer season (Fig. 2). For most seasons and for both variables, a decrease in variability is as likely as an increase. A significant trend

is observed only for summer (the precipitation and temperature variability is expected to increase) and for winter (the precipitation variability is expected to decrease). These results are illustrated (Fig. 4) based on the regional changes predicted by the PRUDENCE RCMs (note that the predicted changes obtained by combining the global-mean warming and regional response uncertainties are larger than those of Fig. 4).

Changes in mean monthly water levels and potential impacts

An increase in the mean monthly lake level is as likely as a decrease (except for May and June, for which slight decreases of 5 and 10 cm, respectively, are expected, Fig. 5, top left). For all months, the difference between the 5th and 95th percentiles of these projections is small. The largest range of water level variation (around 20 cm) is observed in summer (May to August). In late autumn, in winter and early spring, this range is less than 5 cm. This suggests that the current regulation rules are also appropriate for the future climate scenarios.

For the majority of the future scenarios, the annual water level fluctuations simulated are less than those for the control period (Fig. 5 top right, Fig. 6 right). For the control period, the annual water level fluctuations are always higher than 50 cm; for 40% of the years, it is higher than 75 cm. For the median future conditions, these annual fluctuations (50 and 75 cm respectively) are exceeded for only 80% and 20% respectively of the years. The same results are obtained for the maximum water level fluctuation over a three- or five-year period. For the reed population, the future conditions are, thus, expected to be worse than the control conditions.

The simulated annual maximum water levels are also expected to be smaller for the future scenarios than for the control period (Fig. 6 left). The simulated future conditions are, thus, expected to be less critical for riparian urban areas than the control conditions. The same conclusion can be drawn about high water levels potentially damaging seeds and plants in maturation (Fig. 7, right).

The overall reproduction success of pikes is expected to decrease (Fig. 7 left). For most future scenarios, the number of zero global success scores increases significantly (+100% for median future conditions compared to the control period). These results are due mainly to a shift in the arrival date of the spring flood induced by earlier snowmelt following increases in temperature. For several years, this shift leads to a zero S2 score.

RESULTS FOR THE E2 AND E3 EXPERIMENTS

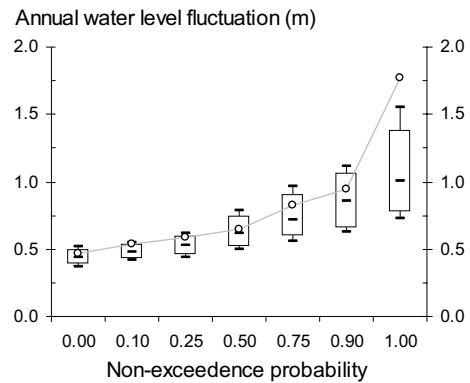
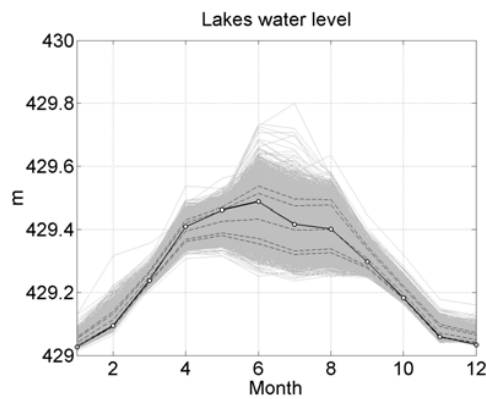
The previous results do not change significantly if one source of prediction uncertainty — either the global-mean warming

or the regional response — is considered separately. The trends identified (criterion increase or decrease) remain unchanged for all variables studied but the variability range decreases significantly and for some, this reduction may correspond to a decrease of 100% or more. This is clearly illustrated by the results for the mean monthly water levels (Fig. 5 left, top, centre and bottom) or for several percentiles of the maximum annual water level fluctuation (Fig. 5 right, top, centre and bottom). The contribution of the global-mean warming uncertainty to the variability of the system behaviour for future conditions is of the same order of magnitude as the contribution of the uncertainty in the regional response. Combining both sources of uncertainty does not just sum the respective variability ranges as the total variability is much higher. These results are in line with those obtained by Horton *et al.* (2006), Hingray *et al.* (2007) and by Schaeffli *et al.* (2007). For five case study regions, Hingray *et al.* (2007) show that the variance of the simulated regional climate changes due to RCM inter-model differences contributes substantially to the total variance. Schaeffli *et al.* (2007) analyse impacts of climate change on the performance of an alpine hydropower production system by successively integrating the uncertainties due to the management model, the hydrological model, the future glacier surface estimation, the regional response to global-mean warming and the global-mean warming itself. They also conclude that a significant part of the total prediction uncertainty can be attributed to the uncertainty in the regional response to climate change. In climate change studies currently available, the emission scenario and the driving climate model are generally assumed to determine a predominant part of the regional climate change uncertainty (e.g. Jenkins and Lowe, 2003). The results presented here suggest that this assumption may not always be valid and that more attention should be paid to the inter-RCM variability.

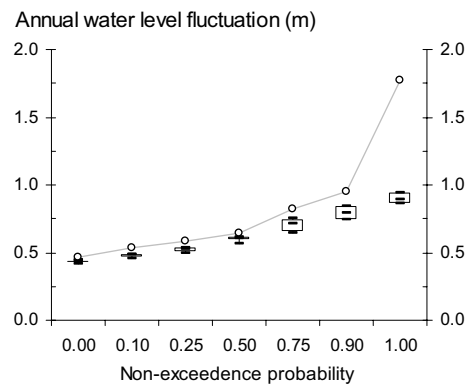
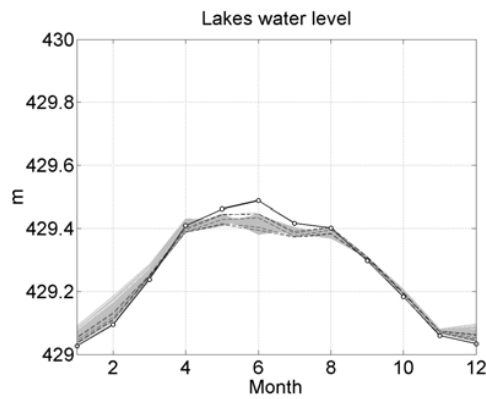
Conclusion

The methodology in this paper enables a probabilistic assessment of potential climate changes and their related impacts on a real-world water resources system. The climate changes predicted are highly uncertain but the results obtained show that they are likely to have a significant impact on the lake system under study.

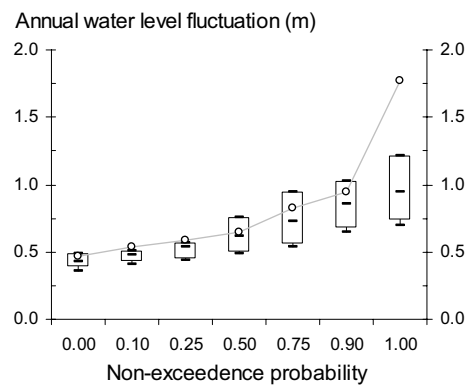
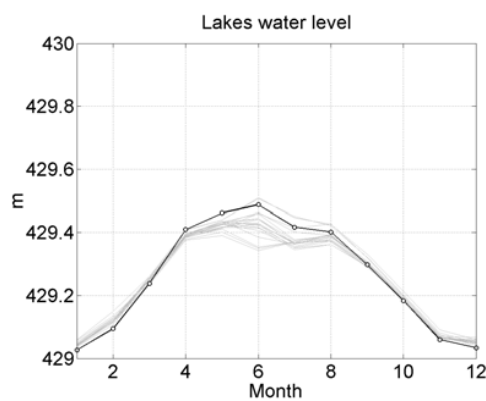
The regional climate change scenarios are obtained through scaling the regional response patterns resulting from RCM runs by global-mean warming projections from a simple climate model. This scenario generator allows global-mean warming uncertainty to be combined with the uncertainty inherent in the regional climate response. For



E1) Uncertain global-mean warming and uncertain scaling relationship (12000 simulations)



E2) Uncertain global-mean warming and median scaling relationship (12000 simulations)



E3) Median global-mean warming (2.6 °C) and uncertain scaling relationships (26 simulations)

Fig. 5. Variability of water level regimes obtained for the three experiments E1, E2, E3. Left: mean monthly lake levels; (circular marker = control situation; for E1 and E2, the dotted lines indicate the 5th, 10th, 50th, 90th and 95th percentiles obtained from the 12000 simulations for the future climate change scenarios; for E3, each grey line corresponds to the response to one of the 26 RCM response patterns). Right: boxplots of maximum annual water level fluctuation for different non-exceedence probabilities. For explanation of the boxplots, see Fig. 4.

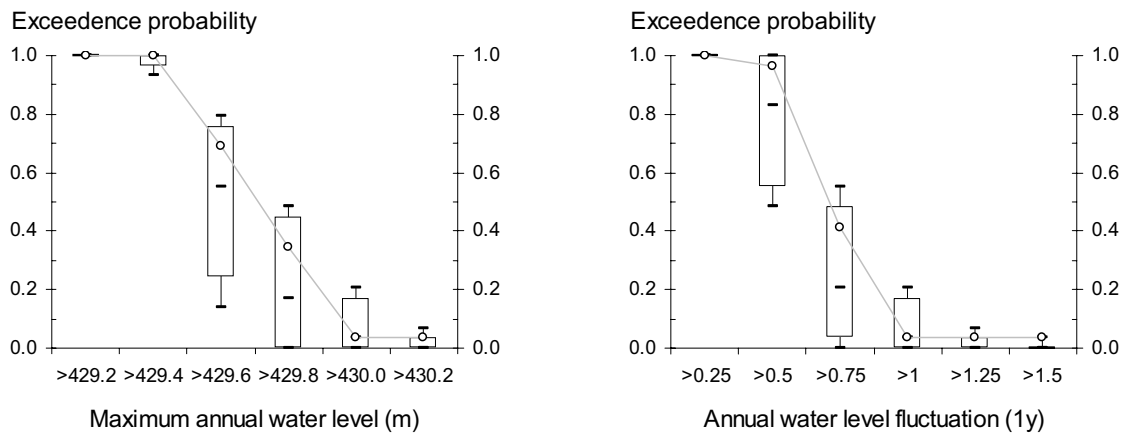


Fig. 6. Variability of the water level regimes obtained with experiment E1 (including global-mean warming uncertainty and regional response uncertainty). Left: exceedence probability of the maximum annual water level; right: exceedence probability of the annual water level fluctuation. The exceedence probability is obtained for each future scenario from the number of years within the 30-years simulation period for which a given threshold is passed. The boxplots refer to the variability of this variable within the set of 12000 future scenarios. For explanation of the boxplots, see Fig. 4.

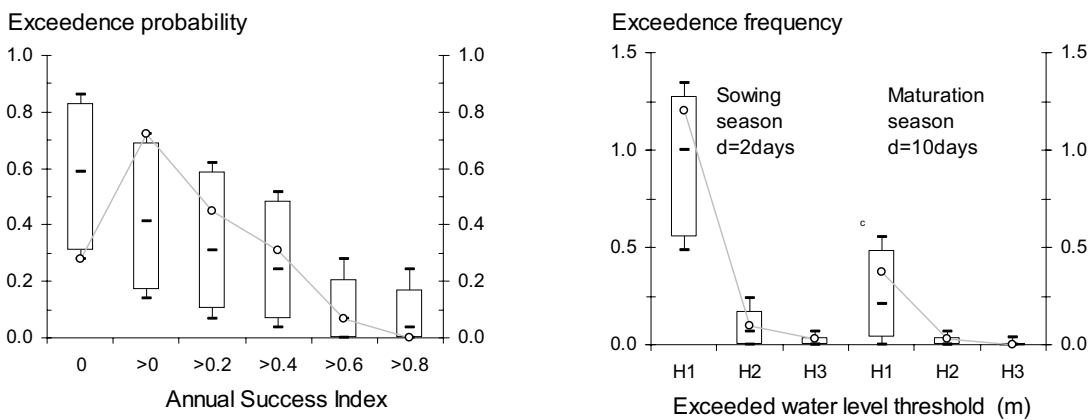


Fig. 7. Variability of performance indices obtained with experiment E1 (including global-mean warming uncertainty and regional response uncertainty); left: exceedence probability of the annual pike reproduction success; right: annual exceedence frequency of spring and summer flooding events related to agriculture (H1 = 429.5 m a.s.l., H2=429.75 m a.s.l. and H3=430 m a.s.l.). For explanation of the boxplots, see Fig. 4.

all fields studied, the uncertainty due to inter-RCM variability contributes much of the total uncertainty in predicting the impact of climatic change and, indeed, may be as large as the uncertainty induced by the global-mean warming. This shows that the uncertainty of the regional response to global-mean warming should not be neglected in studies of the impact of climatic change.

The results presented are conditional on the models, data and modelling assumptions used. The methodology for the generation of the local scale time series has the advantage of simplicity but has also some drawbacks as potential changes in weather frequency or persistence are not quantified (see Mearns *et al.*, 1996, for an illustration of possible related impacts). This is especially critical when extreme events are analysed (floods, droughts). Even on a seasonal time step, the risk of extreme events in a future

climate is not negligible (Palmer and Räisänen, 2002; Schär *et al.*, 2003). Potential long-term droughts such as that in 2003 experienced all over Europe or, possibly, *n*-days of extreme rainfall events, cannot be simulated with a simple scaling approach whatever the number of generated scenarios.

Using the RCM output time series directly as an input to local system models could help in overcoming these limits. As RCM experiments are very time-consuming, they cannot be used to simulate the number of scenarios required for a probabilistic impact assessment. Especially in mountain regions, the precipitation time series derived by these models are still too unreliable to be used directly as input to a hydrological model. Orographic precipitation is one of the most difficult variables to simulate in climate models (Giorgi and Mearns, 1991). Consequently, current research pays

increasing attention to downscaling models that connect the local meteorological variables to variables describing meteorological situations on a synoptic scale (for a review of such methods, see Xu, 1999; Zorita and von Storch, 1999; Prudhomme *et al.*, 2002). Such downscaling techniques could also be applied to simulate, for the control period, the natural climate variability, the characterisation of which would formally be required for a complete impact assessment (e.g., Hingray *et al.*, 2006).

It would also be interesting to quantify the uncertainties due to the hydrological modelling to show whether they increase the uncertainty ranges induced by the climate scenarios. In the present work, the rainfall-excess to runoff relationships are assumed to remain constant in the future. This assumption may be critical for the regression-based models applied to the Aare and Jura catchments. Schaeffli *et al.* (2007) showed for another case study in the Swiss Alps that the uncertainty in the estimation of the hydrological model parameters is small compared to those introduced by the climate change scenarios. The choice of the model structure could, however, have a larger influence (Schaeffli *et al.*, 2004; Schaeffli, 2005). It would also be interesting to compare the results of this study with results from a more detailed hydrological model including an explicit model of the large water reservoirs in the Aare catchment.

The impact assessment methodology is also worthy of further investigation. The evaluation of the impacts of climate change on highly different and sometimes competing water-dependent interests related to human activities or ecosystem conservation remains a very difficult task. The present study outlines a possible solution but does not pretend to approach this problem in its entire complexity. The welfare indices defined give at least a partial view of potential climate change impacts on the different fields of interests and enable a statistical analysis. Such a simple statistical analysis quantifying the intensity, frequency and duration of impacts is, however, insufficient for the analysis of natural systems. A meaningful analysis of the impacts of climate change on ecosystems should ideally be based on ecological simulation tools to model the long-term evolution of the populations studied.

Acknowledgements

This work is part of the 2001–2004 SWURVE (Sustainable Water: Uncertainty, Risk and Vulnerability in Europe) project, funded under the EU Environment and Sustainable Development programme, framework V, grant number EVK1-2000-00075. The Swiss part of this project was funded by the Federal Office for Education and Science, contract number 00.0117. The authors thank the Swiss

Federal Office of Hydrology and Geology for providing the discharge data and the national weather service MeteoSwiss for providing the meteorological time series.

The authors would like to express their gratitude to the members of the PRUDENCE (Prediction of Regional scenarios and Uncertainties for Defining European Climate change risks and Effects) project for having made available the seasonal RCM grids on which the present work is based. The PRUDENCE project is funded under the EU Environment and Sustainable Development programme, framework V, grant number EVK2-CT-2001-00132. Acknowledgement is also made to Tom Wigley for providing the global-mean warming PDFs, as well as to Bernard Büttiker (fishery inspector of Canton de Vaud), Catherine Strehler Perrin and Michel Antoniazza (study and management group of the Grande Caricaie) and Jean-Pierre Clément (Swiss Agency for the Environment, Forests and Landscape) for their help during the development of the impact indices.

References

- Aebischer, A., 1994. *La réussite de la reproduction des oiseaux nicheurs de la rive sud du lac de Neuchâtel*. Mandat du Laboratoire de recherches hydrauliques, Eidgenössische Technische Hochschule Zürich, Switzerland.
- Arnell, N.W. and Hulme, M., 2000. *Implications of climate changes for large dams and their management*. Thematic Review II.2 prepared as an input to the World Commission on Dams, Cape Town (<http://www.dams.org>).
- Arribas, A., Gallardo, C., Gaertner M.A., and Castro, M., 2003. Sensitivity of the Iberian Peninsula climate to a land degradation. *Clim. Dynam.*, **20**, 477–489.
- Buttler, A., Cornali, P. and Bueche, M., 1995. *Etude des effets de la régulation des lacs subjurassiens sur la végétation et le milieu*. Rapport final, Laboratoire d'écologie végétale et de phytosociologie de l'Université de Neuchâtel, Switzerland.
- Christensen, J.H., Carter, T.R. and Giorgi, F., 2002. PRUDENCE employs new methods to assess European Climate Change. *EOS Trans.*, AGU, **83**, 147.
- Christensen, J.H., Räisänen, J., Iversen, T., Bjorge, D., Christensen, O.B. and Rummukainen, M., 2001. A synthesis of regional climate change simulations - A Scandinavian perspective. *Geophys. Res. Lett.*, **28**, 1003–1006.
- Clerc, C., 1999. *Suivi scientifique des roselières lacustres. Résultats 1993-1998*. Champ-Pittet, Switzerland.
- Denk, M., 2002. *Modelling of a regulated lake system and sustainability evaluation of management rules. A case study on the three Jura lakes, Switzerland*. Diploma thesis, Ecole Polytechnique Fédérale de Lausanne, Switzerland.
- Doms, G. and Schättler, U., 1999. *The Nonhydrostatic Limited-Area Model LM (Lokal-Modell) of DWD, Part I: Scientific Documentation*, Deutscher Wetterdienst, Meteorologische Analyse und Modellierung, Offenbach, Germany.
- Ekström, M., Hingray, B., Mezghani, A. and Jones, P.D., 2007. Regional climate model data used within the SWURVE project 2: addressing uncertainty in regional climate model data for five European case study areas. *Hydrol. Earth Syst. Sci.*, **11**, 1085–1096.

- Fowler, H.J., Kilsby, C.G. and O'Connell, P.E., 2003. Modeling the impacts of climatic change and variability on the reliability, resilience, and vulnerability of a water resource system. *Water Resour. Res.*, **39**, 1222, doi:10.1029/2002WR001778.
- Frei, C., Christensen, J.H., Deque, M., Jacob, D., Jones, R.G. and Vidale, P.L., 2003. Daily precipitation statistics in regional climate models: Evaluation and intercomparison for the European Alps. *J. Geophys. Res.- Atmos.*, **108** (D3), 4124, doi: 10.1029/2002JD002287.
- Gibelin, A.L. and Déqué, M., 2003. Anthropogenic climate change over the Mediterranean region simulated by a global variable resolution model. *Clim. Dynam.*, **20**, 327–339.
- Giorgi, F. and Mearns, L.O., 1991. Approaches to the simulation of regional climate change: a review. *Rev. Geophys.*, **29**, 191–216.
- Giorgi, F., Marinucci, M.R. and Bates, G.T., 1993a. Development of a second generation regional climate model (REGCM2). Part I: Boundary layer and radiative transfer processes. *Mon. Weather Rev.*, **121**, 2794–2813.
- Giorgi, F., Marinucci, M. R., Bates, G. T. and DeCanio, G., 1993b. Development of a second generation regional climate model (REGCM2). Part II: convective processes and assimilation of lateral boundary conditions. *Mon. Weather Rev.*, **121**, 2814–2832.
- Gordon, C., Cooper, C., Senior, C.A., Banks, H., Gregory, J.M., Johns, T.C., Mitchell, J.F.B. and Wood, R.A., 2000. The simulation of SST, sea ice extents and ocean heat transports in a version of the Hadley Centre coupled model without flux adjustments. *Clim. Dynam.*, **16**, 147–168.
- Hashimoto, T., Stedinger, J.R. and Loucks, D.P., 1982. Reliability, Resiliency, and Vulnerability Criteria for Water Resource System Performance Evaluation. *Water Resour. Res.*, **18**, 14–20.
- Hingray, B., Mezghani, A., Schaeffli, B., Niggli, M., Faivre, G., Guex, F., Hamdi, Y., Musy, A. 2006. *Estimation des débits de crue du Rhône à Porte du Scex et autres points amont caractéristiques*. Projet CONSECRU 2. Rapport final. HYDRAM. Ecole Polytechnique Fédérale de Lausanne, Lausanne, Switzerland, 190p.
- Hingray, B., Mezghani, A. and Buishand, A., 2007. Development of probability distributions for regional climate change from uncertain global-mean warming and an uncertain scaling relationship. *Hydrol. Earth Syst. Sci.*, **11**, 1097–1114.
- Horton, P., Schaeffli, B., Mezghani, A., Hingray, B. and Musy, A., 2006. Assessment of climate change impacts on Alpine discharge regimes with climate model uncertainty. *Hydrol. Process.*, **20**, 2091–2109.
- Hulme, M., Jenkins, G.J., Lu, X., Turpenny, J.R., Mitchell, T.D., Jones, R.G., Lowe, J., Murphy, J.M., Hassell, D., Boorman, P., McDonald, R. and Hill, S., 2002. *Climate Change Scenarios for the United Kingdom: The UKCIP02 Scientific Report*, Tyndall Centre for Climate Change Research, School of Environmental Sciences, University of East Anglia, Norwich, UK.
- Huntingford, C. and Cox, P.M., 2000. An analogue model to derive additional climate change scenarios from existing GCM simulations. *Clim. Dynam.*, **16**, 575–586.
- Jacob, D., 2001. A note to the simulation of the annual and inter-annual variability of the water budget over the Baltic Sea drainage basin. *Meteorol. Atmos. Phys.*, **77**, 61–73.
- Jakeman, A.J. and Hornberger, G.M., 1993. How much complexity is warranted in a rainfall-runoff model. *Water Resour. Res.*, **29**, 2637–2649.
- Jasper, K., Calanca, P., Gyalistras, D. and Fuhrer, J., 2004. Differential impacts of climate change on the hydrology of two alpine river basins. *Climate Res.*, **26**, 113–129.
- Jaton, J.-F., 1989. *Ile Correction des eaux du Jura: effets du règlement 1980 - 82 de régulation des lacs en matière hydro-agricole - rapport final*, Laboratory of Hydrology and Land Improvement, Ecole Polytechnique Fédérale de Lausanne, Switzerland.
- Jenkins, G. and Lowe, J., 2003. Handling uncertainties in the UKCIP02 scenarios of climate change., *Hadley Centre, Technical note 44, Exeter, UK*.
- Jones, R.N., 2000. Analysing the risk of climate change using an irrigation demand model, *Climate Res.*, **14**, 89–100.
- Katz, R.W. and Brown, B.G., 1992. Extreme events in a changing climate: variability is more important than averages. *Climatic Change*, **21**, 289–302.
- Kim, J., 2005. A projection of the effects of the climate change induced by increased CO₂ on extreme hydrologic events in the western US. *Climatic Change*, **68**, 153–168.
- Lane, M.E., Kirshen, P.H. and Vogel, R.M., 1999. Indicators of impacts of global climate change on U. S. water resources. *J. Water Resour. Plan. Man.-ASCE*, **125**, 194–204.
- Lettenmaier, D.P., Wood, A.W., Palmer, R.N., Wood, E.F. and Stakhiv, E.Z., 1999. Water resources implications of global warming: An U.S. regional perspective. *Climatic Change*, **43**, 537–579.
- Loukas, A., Vasilades, L. and Dalezios, N.R., 2002. Potential climate change impacts on flood producing mechanisms in southern British Columbia, Canada using the CGCMA1 simulation results. *J. Hydrol.*, **259**, 163–188.
- Maisch, M., Wipf, A., Denzler, B., Battaglia, J. and Benz, C., 2000. *Die Gletscher der Schweizer Alpen: Gletscherhochstand 1850, aktuelle Vergletscherung, Gletscherschwund-Szenarien*. Schlussbericht NFP 31. Hochschul-Verlag an der ETH, Zürich, 373pp.
- Martinez, J., Rang, o A. and Major, E., 1983. *The snowmelt-runoff model (SRM) user's guide*. Nasa Reference Publ. 1100. Washington, D.C., USA.
- Mearns, L.O., Rosenzweig, C. and Goldberg, R., 1996. The effect of changes in daily and interannual climatic variability on cereals: A sensitivity study. *Climatic Change*, **32**, 257–292.
- Mermoud, A., 1995. *Méthodes et techniques d'irrigation*. Notes de cours, Ecole Polytechnique Fédérale de Lausanne, Switzerland.
- Mitchell, J.F.B., Johns, T.C., Eagles, M., Ingram, W.J. and Davis, R.A., 1999. Towards the construction of climate change scenarios. *Climatic Change*, **41**, 547–581.
- Mitchell, T.D., 2003. Pattern scaling - An examination of the accuracy of the technique for describing future climates. *Climatic Change*, **60**, 217–242.
- Monteith, J.L. and Unsworth, M., 1990. *Principles of environmental Physics*, Second edition, Editions Arnold, London, UK. 291pp.
- Nakicenovic, N. and Swart, R. (Eds.), 2000. *Special Report on Emissions Scenarios (Intergovernmental Panel on Climate Change)*. Cambridge University Press, Cambridge, UK. 570pp.
- New, M. and Hulme, M., 2000. Representing Uncertainty in Climate Change Scenarios: a Monte Carlo approach. *Integr. Assess.*, **1**, 203–213.
- Palmer, T.N. and Räisänen, J., 2002. Quantifying the risk of extreme seasonal precipitation events in a changing climate. *Nature*, **415**, 512–514.
- Pope, V.D., Gallani, M.L., Rowntree, P.R. and Stratton, R.A., 2000. The impact of new physical parametrizations in the Hadley Centre climate model - HadAM3. *Clim. Dynam.*, **16**, 123–146.
- Prudhomme, C., Jakob, D., and Svensson, C., 2003. Uncertainty and climate change impact on the flood regime of small UK catchments. *J. Hydrol.*, **277**, 1–23.

- Prudhomme, C., Reynard, N. and Crooks, S., 2002. Downscaling of global climate models for flood frequency analysis: where are we now? *Hydrol. Process.*, **16**, 1137–1150.
- Räisänen, J., 2001. CO₂-induced climate change in CMIP2 experiments: Quantification of agreement and role of internal variability. *J. Climate*, **14**, 2088–2104.
- Räisänen, J., 2002. CO₂-induced changes in interannual temperature and precipitation variability in 19 CMIP2 experiments. *J. Climate*, **15**, 2395–2411.
- Räisänen, J., Hansson, U., Ullerstig, A., Döscher, R., Graham, L. P., Jones, C., Meier, H.E.M., Samuelsson, P. and Willén, U., 2004. European climate in the late 21st century: regional simulations with two driving global models and two forcing scenarios. *Clim. Dynam.*, **22**, 13–31.
- Roeckner, E., Bengtsson, L., Feichter, J., Lelieveld, J. and Rodhe, H., 1999. Transient climate change simulations with a coupled atmosphere-ocean GCM including the tropospheric sulfur cycle. *J. Climate.*, **12**, 3004–3032.
- Royer, J.F., Cariolle, D., Chauvin, F., Déqué, M., Douville, H., Planton, S., Rascol, A., Ricard, J.L., Salas y Melia, D., Sevault, F., Simon, P., Somot, S., Tyteca, S., Terray, L. and Valcke, S., 2002. Simulation of climate changes during the 21st century including stratospheric ozone. *CR Geosci.*, **334**, 147–154.
- Santer, B.D., Wigley, T.M.L., Schlesinger, M.E. and Mitchell, J.F.B., 1990. *Developing climate scenarios for equilibrium GCM results*. MPI report 47, Hamburg, Germany.
- Schaefli, B., 2005. *Quantification of modelling uncertainties in climate change impact studies on water resources : Application to a glacier-fed hydropower production system in the Swiss Alps*. Doctoral Thesis, Ecole Polytechnique Fédérale de Lausanne, 209 pp, <http://library.epfl.ch/theses>, thesis no 3225.
- Schaefli, B., Hingray, B. and Musy, A., 2004. Improved calibration of hydrological models: use of a multi-objective evolutionary algorithm for parameter and model structure uncertainty estimation. In: *Hydrology: Science and Practice for the 21st Century*, B.W. Webb *et al.* (Eds.), British Hydrological Society, London, UK. 362–371.
- Schaefli, B., Hingray, B. and Musy, A., 2007. Climate change and hydropower production in the Swiss Alps: quantification of potential impacts and related modelling uncertainties. *Hydrol. Earth Syst. Sci.*, **11**, 1191–1205.
- Schär, C., Vidale, P.L., Lüthi, D., Frei, C., Häberli, C., Liniger, M.A. and Appenzeller, C., 2004. The role of increasing temperature variability in European summer heatwaves. *Nature*, doi: 10.1038/nature02300.
- Schneider, S.H. and Thompson, S.L., 1981. Atmospheric CO₂ and climate – importance of the transient response. *J. Geophys. Res. Oceans Atmos.* **86**, 3135–3147.
- Zorita, E. and von Storch, H., 1999. The analog method as a simple statistical downscaling technique: comparison with more complicated methods. *J. Climate*, **12**, 2474–2489.
- Schreider, S.Y., Whetton, P.H., Jakeman, A.J. and Pittock, A.B., 1997. Runoff modelling for snow affected catchments in the Australian alpine region, Eastern Victoria. *J. Hydrol.* **200**, 1–23.
- Shabalova, M.V., Van Deursen, W.P.A. and Buishand, T.A., 2003. Assessing future discharge of the river Rhine using regional climate model integrations and a hydrological model. *Climate Res.*, **23**, 223–246.
- Spreafico, M., Weingartner, R. and Leibundgut, C., 1992. *Atlas hydrologique de la Suisse*. Service Hydrologique et Géologique National, Bern, Switzerland.
- Steel, M., 1999. *Historic rainfall, climatic variability and flood risk estimation for Scotland*. PhD thesis, University of Dundee, UK.
- SwissTopo, 1997. *Digital National Maps of Switzerland - PM25*, Wabern, Switzerland.
- VAW, 1996. *Régularisation optimale de la lie correction des eaux du Jura (lie CEJ)*. 3941 F, Versuchsanstalt für Wasserbau, Hydrologie und Glaziologie der Eidgenössischen Technischen Hochschule Zürich, Switzerland.
- Vidale, P.L., Lüthi, D., Frei, C., Seneviratne, S. and Schär, C., 2003. Predictability and uncertainty in a Regional Climate Model. *J. Geophys. Res.-Atmos.*, **108** (D18), 4586, doi: 10.1029/2002JD002810, 2003.
- Vogel, R.M., Lane, M., Ravindran, R.S. and Kirshen, P.H., 1999. Storage reservoir behavior in the United States. *J. Water Resour. Plan. Man.-ASCE*, **125**, 245–254.
- Whetton, P., Haylock, M. and Galloway, R., 1996. Climate change and snow-cover duration in the Australian Alps. *Climatic Change*, **32**, 447–479.
- Wigley, T.M.L. and Raper, S.C.B., 2001. Interpretation of High Projections for Global-Mean Warming. *Science*, **293** (5529), 451–454.
- Wigley, T.M.L., Raper, S.C.B., Smith, S. and Hulme, M., 2000. *The Magic/ScenGen Climate Scenario Generator: version 2.4*: Technical Manual, CRU, UEA, Norwich, U.K.
- Xu, C.Y., 1999. From GCMs to river flow: a review of downscaling methods and hydrologic modelling approaches. *Prog. Phys. Geogr.*, **23**, 229–249.
- Young, P., 1984. *Recursive Estimation and Time-Series Analysis*. Springer, Berlin, Germany.
- Young, P., 1993. Time variable and state dependent modelling of non-stationary and non-linear time series. In: *Developments in Time Series Analysis*, T. Subba Rao (Ed.), Chapman and Hall, London, UK. 374–413.
- Young, P. and Beven, K., 1994. Data-based mechanistic modelling and the rainfall-flow non-linearity. *Environmetrics*, **5**, 335–363.
- Zaugg, B., Pedroli, J.C. and Zaugg, C., 1994. *Incidences des fluctuations du niveau du lac de Neuchâtel sur la faune piscicole*. Rapport du bureau Aquarius, Neuchâtel, Switzerland.

$$F(u) = \Pr(U \leq u) = 1 - \exp\left[-(u/\alpha)^c\right] \quad (\text{A1.1})$$

where $\Pr(U \leq u)$ is the probability that $U \leq u$ and α and c are the scale and the shape parameter. For such a distribution, the p th quantile has the expression:

$$u_p = \alpha[-\ln(1-p)]^{1/c} \quad (\text{A1.2})$$

For the observed data, both parameters α_{obs} and c_{obs} are estimated by the method of moments such that the distribution preserves the observed mean (MP_{obs}) and

Appendix 1 : Perturbation of daily rainfall series (according to Shabalova *et al.*, 2003)

The Weibull distribution is assumed to give a reasonable fit to the distribution of the daily rainfall amounts. For a random variable U (daily rainfall amount) with the distribution function $F(u)$ where u is a realisation of U , the Weibull distribution is defined as

coefficient of variation (CVP_{obs}). The theoretical values for the mean and the variance are:

$$E(u) = \alpha \Gamma\left(\frac{1}{C} + 1\right) \tag{A1.3}$$

$$Var(u) = \alpha^2 \left[\Gamma\left(\frac{2}{C} + 1\right) - \left[\Gamma\left(\frac{1}{C} + 1\right) \right]^2 \right] \tag{A1.4}$$

For the future time series scenario, the mean (MP_{scen}) and the variation coefficient (CVP_{scen}) of the daily rainfall amounts are estimated from the observed mean (MP_{obs}) and the observed variation coefficient (CVP_{obs}) and their relative changes given by the regional climate change statistics (XMP and $XCVP$). The Weibull parameters for the future scenario α_{scen} and c_{scen} are estimated by the method of moments (Eqns A1.3 and A1.4). It is assumed that if the daily rainfall amount $u_{obs}(t)$ observed on day t corresponds to the p th quantile in the observed series, the scenario value $u_{scen}(t)$ for the same date corresponds to the same quantile in the scenario series. From these assumptions and Eqn. A1.2 it follows:

$$u_{scen}(t) = \alpha_{scen} \cdot \left(\frac{u_{obs}(t)}{\alpha_{obs}} \right)^{\frac{C_{obs}}{C_{scen}}} \tag{A1.5}$$

Note that the Weibull distribution parameters are estimated independently for each season (for both the observed and the future scenario).

Appendix 2: Estimation of daily variances from PRUDENCE data

DETRENDING FUTURE TEMPERATURE SERIES

The variances used in PRUDENCE to describe the variability of the temperatures are estimated from the 30 values of 90-days temperature $T_{s,k}$, where k refers to the k^{th} year of the 30-years period and where $T_{s,k}$ is the mean temperature value for the s^{th} season of this year. The future temperature series can show considerable trends, especially the runs obtained for the A2 scenario, and this trend could have a non-negligible effect on the estimated variance and the related standard deviation. The time series $T_{s,k}$ should, therefore, be detrended. The effect of the temperature trend on the variance can be estimated without knowing the individual seasonal values. The expected temperature increase during the 21st century is assumed linear in time.

The difference in the mean seasonal temperature between the future and the control run will, therefore, give a reasonable estimate of the total trend Δ_s for the years 2070–2099 relative to season s (Eqn. A2.1):

$$\Delta_s = XMT_s * 30/110 \tag{A2.1}$$

where $XMT_s = MT_{future,s} - MT_{control,s}$ is the change in mean seasonal temperature over this 110-years period and where Δ_s is the trend for the years 2070–2099 for season s (in °C per 30 years). Note that for the period 1961–1990, the change in temperature is probably not large enough to have a marked influence on the standard deviations of the seasonal temperatures. Let $T_{s,1}, T_{s,2}, \dots, T_{s,n}$ be the mean 90-days temperature values as simulated by the RCM for the period 2070-2099 ($n = 30$) for season s and σ_s^2 the variance estimated based on this series:

$$\sigma_s^2 = \frac{1}{n-1} \sum_{i=1}^n (T_{s,i} - MT_{s,future})^2 \tag{A2.2}$$

If the trend is linear, it can be easily shown that the mean of σ_s^2 is given by

$$E(\sigma_s^2) = SDT_{90,s}^2 + \frac{1}{12} \Delta_s^2 (1 + 1/n) \approx SDT_{90,s}^2 + \frac{1}{12} \Delta_s^2 \tag{A2.3}$$

where $SDT_{90,s}^2$ is the true variance of the detrended series. If an estimate of σ_s and of Δ_s is known, an almost unbiased estimate of $SDT_{90,s}^2$ can be obtained.

PERTURBATION OF DAILY SERIES BASED ON 90-DAYS STATISTICS

The PRUDENCE data yields 90-day statistics that cannot be used directly to perturb the observed daily temperature and precipitation series. For the daily variation coefficients CVP , a linear regional relationship has been identified between this variable and the 90-days' coefficients of variation based on 15 observed rainfall time series located in the region under study. Such a linear relationship is identified for each of the four seasons s .

For temperatures, the change in the daily standard deviations $XSDT$ is set equal to the change in the 90-days standard deviations $XSDT_{90}$. It can be easily shown that, in this case, the perturbation of the observed daily series based on Eqn. 2 gives a future daily series that preserves the relative change in the 90-days standard deviations.