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ON SOME GENERALIZED REINFORCED RANDOM WALK ON INTEGERS

OLIVIER RAIMOND AND BRUNO SCHAPIRA

ABSTRACT. We consider Reinforced Random Walks where transition probabilities are a function of the proportion of times the walk has traversed an edge. We give conditions for recurrence or transience. A phase transition is observed, similar to Pemantle [Pem1] on trees.

Key words: Reinforced random walk, urn processes.

A.M.S. classification: 60F20.

1. INTRODUCTION

Let $G = (V, E)$ be a graph with V the set of vertices and E the set of edges. The graph distance is denoted by d . Reinforced random walks $(X_n, n \geq 0)$ are walks on G (i.e. $X_n \in V$ and $(X_n, X_{n+1}) \in E$). To define their laws, denote by $(\mathcal{F}_n)_{n \geq 0}$ the natural filtration associated to $(X_n, n \geq 0)$ and for $(x, y) \in E$, set

$$a_n(x, y) = a_0(x, y) + \Delta \sum_{i=1}^n 1_{\{(X_{i-1}, X_i) = (x, y)\}}$$

with $a_0(x, y) > 0$ and $\Delta > 0$. The law of a directed reinforced random walk is such that for $(x, y) \in E$, on the event $\{X_n = x\}$,

$$\mathbb{P}(X_{n+1} = y \mid \mathcal{F}_n) = \frac{a_n(x, y)}{\sum_{\{z \mid (x, z) \in E\}} a_n(x, z)}.$$

When E is non-oriented (i.e. $(x, y) \in E$ implies $(y, x) \in E$), for $(x, y) \in E$, set

$$b_n(x, y) = b_0(x, y) + \Delta \sum_{i=1}^n 1_{\{(X_{i-1}, X_i) \in \{(x, y), (y, x)\}\}},$$

with $b_0(x, y) = b_0(y, x) > 0$ and $\Delta > 0$. The law of an undirected reinforced random walk is such that for $(x, y) \in E$, on the event $\{X_n = x\}$,

$$\mathbb{P}(X_{n+1} = y \mid \mathcal{F}_n) = \frac{b_n(x, y)}{\sum_{\{z \mid (x, z) \in E\}} b_n(x, z)}.$$

The law of a directed reinforced random walk is the same as the law of a random walk in a random environment. Indeed it is equivalent to attach independent Polya urns to all sites and then De Finetti theorem implies that it is equivalent to attach independent random probability vectors to each site. These probability vectors give the transition probabilities for the walk (when located at this site). When the graph is a non-oriented tree, the undirected reinforced random walk, with initial position ρ , has the same law as a directed reinforced random walk on G with $a_0(x, y) = b_0(x, y)$ if $d(\rho, y) = d(\rho, x) + 1$ and $a_0(x, y) = b_0(x, y) + \Delta$ if $d(\rho, y) = d(\rho, x) - 1$

and reinforced parameter 2Δ , or with $a_0(x, y) = b_0(x, y)/(2\Delta)$ if $d(\rho, y) = d(\rho, x) + 1$ and $a_0(x, y) = (b_0(x, y) + \Delta)/(2\Delta)$ if $d(\rho, y) = d(\rho, x) - 1$ and reinforced parameter 1. This representation was first observed by Coppersmith and Diaconis [CDia].

For this class of models, results on random walks in random environment can be applied. When the graph is \mathbb{Z} , Solomon's theorem shows that (directed and undirected) reinforced random walks are a.s. recurrent when $a_0(x, x+1) = a_0(x, x-1) = a_0 > 0$ and $b_0(x, x+1) = b_0 > 0$. When the graph is the binary tree, $b_0(x, y) = b_0 > 0$, the undirected reinforced random walk is transient for small Δ (or equivalently for large b_0) and recurrent for large Δ (or equivalently for small b_0). This last result was proved by R. Pemantle in [Pem1].

In this paper we address the question (posed by M. Benaïm to one of the authors) of what happens when the graph is \mathbb{Z} and on the event $\{X_n = x\}$,

$$\mathbb{P}(X_{n+1} = x + 1 \mid \mathcal{F}_n) = f \left(\frac{a_n(x, x + 1)}{a_n(x, x - 1) + a_n(x, x + 1)} \right),$$

where $f : [0, 1] \mapsto]0, 1[$ is a smooth function. For general functions f , these walks are no longer random walks in random environment. So we have to use different techniques. But one can still attach to each site independent urn processes (generalized Polya urns). Under the assumption that the number of fixed points of f is finite, if the walk is recurrent, stochastic algorithms techniques show that for all x , $a_n(x, x + 1)/(a_n(x, x - 1) + a_n(x, x + 1))$ converges a.s. towards a random variable α_x . This random variable takes its values in the set of fixed point of f . If $a_0(x, x + 1) = a_0(x, x - 1) = a_0 > 0$, the sequence $(\alpha_x, x \in \mathbb{Z})$ is i.i.d. Let us remark that Solomon's theorem states that the random walk in the random environment $(\alpha_x, x \in \mathbb{Z})$ is a.s. recurrent if and only if $\mathbb{E}[\ln(\alpha_x/(1 - \alpha_x))] = 0$.

We here focus on the case where all the fixed points are greater or equal to $1/2$ and $f(\alpha) \geq 1/2$ for $\alpha \geq 1/2$. We also assume $a_0(x, x + 1) = a_0(x, x - 1) = a_0 > 0$. We give criteria for recurrence and transience: when there exists one fixed point greater than $1/2$ the walk is transient and when $1/2$ is the unique fixed point, depending on the initial condition a_0 and on the shape of f around $1/2$, the walk can be either recurrent or transient. This last result shows that Solomon's criterion does not apply. The proofs of the theorems given here involve martingale techniques inspired by the work of Zerner on multi-excited random walks on integers [Zer, Zer2].

The paper is organized as follows: in section 2 reinforced random walks are defined and their representation with urn processes is given. In section 3 are given the results on urns that are needed to prove the theorems given in sections 5 and 6. A zero-one law is proved in section 4: recurrence occurs with probability 0 or 1. In section 5 and 6 the case where $f \geq 1/2$ and the case where there is a unique fixed point are studied. The last section develops some examples.

2. NOTATION

2.1. The urn model. We consider an urn model where balls of the urn are only of two types or colors, let say Red and Blue. Given a function $f : [0, 1] \rightarrow [0, 1]$ we alter the draw by choosing at each step a Red ball with probability $f(\alpha)$, if α is the proportion of Red balls. Then we put back two Red (respectively Blue) balls in the urn if a Red (respectively Blue) ball was drawn. In other words an urn process associated to f is a Markov process $((\alpha_n, l_n), n \geq 0)$ on $[0, 1] \times]0, +\infty[$, where the transition probabilities are defined as follows: for all n , $l_{n+1} = l_n + 1$ and α_{n+1} is

equal to $(l_n \alpha_n + 1)/(l_n + 1)$ with probability $f(\alpha_n)$, or equal to $l_n \alpha_n/(l_n + 1)$ with probability $1 - f(\alpha_n)$. In fact α_n represents the proportion of Red balls and l_n the total number of balls in the urn at time n (at least if l_0 and $\alpha_0 l_0$ are integers). By abuse of notation we will sometime call the first coordinate $(\alpha_n, n \geq 0)$ an urn process associated to f (if there is no ambiguity on l_0). This model was introduced in [HLS], and then further studied in particular by Pemantle [Pem4], Duflo [D] and Benaim and Hirsch (see [BH] and [B]). We refer also the reader to the survey [Pem5] section 2.4, 2.5 and 3.2 for more details and references.

2.2. The reinforced random walk. Here we consider a particular model of reinforced random walk $(X_n, n \geq 0)$ on \mathbb{Z} where the evolution is driven by urns of the preceding type on each integer. For other models see [Pem5]. A first way to define it is as follows. Let $f : [0, 1] \rightarrow]0, 1[$ and $(\alpha_0^x, l_0^x)_{x \in \mathbb{Z}} \in ([0, 1] \times]0, +\infty[)^{\mathbb{Z}}$ be given. Then for $n \geq 0$ and $x \in \mathbb{Z}$, let

$$L_n^x := \sum_{k=0}^n 1_{\{X_k=x\}},$$

be the total time spent in x up to time n by the random walk. We will use also the convention $L_{-1}^x = 0$ for all x . Let then

$$\tilde{\alpha}_n^x := \frac{1}{l_0^x + L_{n-1}^x} \left\{ \alpha_0^x l_0^x + \sum_{k=0}^{n-1} 1_{\{X_k=x, X_{k+1}=x+1\}} \right\},$$

be the proportion of times it has moved to the right (up to some initial weights). Now if $X_n = x$, for some $x \in \mathbb{Z}$ and $n \geq 0$, then $X_{n+1} = x + 1$ with probability $f(\tilde{\alpha}_n^x)$ and $X_{n+1} = x - 1$ with probability $1 - f(\tilde{\alpha}_n^x)$. This defines recursively our random walk. Moreover for $n \geq 1$, define τ_n^x as the time of the n^{th} passage in x if it is finite or as τ_{n-1}^x if it is infinite (with the convention $\tau_0^x = 0$). Let also

$$\alpha_n^x := \tilde{\alpha}_{\tau_{n+1}^x - 1}^x,$$

with the convention $\tilde{\alpha}_{-1}^x = \alpha_0^x$. Then the processes $((\alpha_n^x, l_0^x + n), n \geq 0)$ form a family of independent urn processes of the type described above (possibly stopped at some random time).

There is another way to define our random walk which goes in the other direction. Assume first that we are given a family of independent urn processes $((\alpha_n^x, l_n^x), n \geq 0)_{x \in \mathbb{Z}}$ indexed by \mathbb{Z} . One can consider this as the full environment for instance. Then given the full environment our random walk evolves deterministically: it starts from X_0 . Next let $n \geq 0$ be given and assume that the random walk has been defined up to time n . Suppose that $X_n = x$ and $L_n^x = k$, for some $k \geq 1$. Then $X_{n+1} = X_n + 1$ if $\alpha_{k+1}^x > \alpha_k^x$, and $X_{n+1} = X_n - 1$ otherwise.

For $n \geq 0$, we define the environment $w_n \in ([0, 1] \times]0, +\infty[)^{\mathbb{Z}}$ at step n by

$$w_n := (\alpha_{L_n^x}^x, l_0^x + L_n^x)_{x \in \mathbb{Z}}.$$

We denote by \mathcal{F}_n the σ -algebra generated by (X_0, w_0, \dots, w_n) , or equivalently by (w_0, X_0, \dots, X_n) . For $x \in \mathbb{Z}$ and w some environment, we denote by $\mathbb{E}_{x,w}$ the law of the random walk starting from $X_0 = x$ and with initial environment $w_0 = w$. If no ambiguity on x or w is possible we will sometime forget them in the notation.

Observe that $((w_n, X_n), n \geq 0)$ is a Markov process (whereas $(X_n, n \geq 0)$ is not), and in particular:

$$\mathbb{E}_{x,w}[g(X_{n+1}) \mid \mathcal{F}_n] = \mathbb{E}_{X_n, w_n}[g(X_1)],$$

for any (x, w) and any measurable function g .

In order to include the framework of non-oriented reinforced random walks as discussed in the introduction we suppose the following: first we assume that $l_0^x = l_0 \in]0, +\infty[$ for all $x \neq 0$. Next we assume that $\alpha_0^x = \alpha_0 \in (0, 1)$, for $x > 0$, and $\alpha_0^x = 1 - \alpha_0$, for $x < 0$. The connection with the notation of the introduction is given by $a_0(x, x+1) = l_0 \alpha_0^x$.

2.3. Hypothesis on f and stable points. Throughout the paper $f : [0, 1] \rightarrow]0, 1[$ will be a regular function (C^3 is enough for our purpose). We say that p is a fixed point if $f(p) = p$. It is called stable if $f'(p) \leq 1$. We will assume that all fixed points of f are isolated.

3. PRELIMINARIES ON URNS

3.1. Convergence of urn processes. We recall here some known results about convergence of urn processes. In particular the next theorem is of fundamental importance in all this paper:

Theorem 3.1 ([HLS], [Pem4]). *Let $(\alpha_n, n \geq 0)$ be an urn process associated to some function f . Then almost surely α_n converges to a stable fixed point of f and for any stable fixed point, the probability that α_n converges to this point is positive.*

The convergence to a stable fixed point p with positive probability was first proved in [HLS], when $f(x) - x$ changes of sign near p , and in [Pem4] in the special case where the sign of $f(x) - x$ is constant near p . The non existence a.s. of other limiting points was also first proved in [HLS] (for extensions to more general settings see [D], [B], [Pem2], [Pem3]).

There is also a central limit theorem, which can be extracted from the book of Duflou:

Theorem 3.2 ([D] Theorem 4.III.5). *Suppose that $p \in (0, 1)$ is a stable fixed point of f . Let $a = f'(p)$ and $v^2 = p(1-p)$. If $a < 1/2$, then conditionally on $\alpha_n \rightarrow p$, $\sqrt{n}(\alpha_n - p)$ converges in law, as n tends to $+\infty$, toward a normal variable with variance $v^2/(1-2a)$.*

3.2. Convergence of the drift. For $n \in \mathbb{N}$ we set

$$\delta_n = \sum_{k=0}^n (2f(\alpha_k) - 1).$$

If δ_n converges when $n \rightarrow +\infty$, we denote by δ_∞ its limit. We will need also to consider its negative and positive parts defined respectively by

$$\delta_n^- := \sum_{k=0}^n (2f(\alpha_k) - 1)_-,$$

and

$$\delta_n^+ := \sum_{k=0}^n (2f(\alpha_k) - 1)_+,$$

for all $n \geq 0$. In fact we can always define in the same way δ_∞^- and δ_∞^+ , even when δ_n does not converge. Moreover observe that if $\mathbb{E}[\delta_\infty^-]$ or $\mathbb{E}[\delta_\infty^+]$ are finite, then δ_n converges a.s. and $\mathbb{E}[\delta_\infty] = \mathbb{E}[\delta_\infty^+] - \mathbb{E}[\delta_\infty^-]$. It happens that for our purpose, such finiteness result will be needed.

The problem is that the convergence of the drift appears to be a non-trivial question. To be more precise, we were able to obtain a satisfying result essentially only in the case where f has a unique fixed point. When this fixed point is different from $1/2$, it is immediate to see that the drift converges a.s. toward $\pm\infty$. However to see that $\mathbb{E}[\delta_\infty^-]$ or $\mathbb{E}[\delta_\infty^+]$ is finite, some non-trivial argument is needed. Since it is the same than in the more difficult case where $1/2$ is the unique fixed point, we start by this case:

Proposition 3.1. *If $1/2$ is the unique fixed point of f and if $f'(1/2) = 0$, then either $\mathbb{E}[\delta_\infty^-]$ or $\mathbb{E}[\delta_\infty^+]$ is finite. In particular δ_n converges a.s. toward δ_∞ .*

Proof. For $x \in [-1/2, 1/2]$, let $h(x) := f(x + 1/2) - x - 1/2$, and for $n \geq 0$, let $x_n := \alpha_n - 1/2$. Let also ϵ_{n+1} be equal to 1 if $\alpha_{n+1} > \alpha_n$ and equal to -1 otherwise. By definition $(x_n, n \geq 0)$ satisfies the following stochastic algorithm:

$$(1) \quad x_{n+1} = x_n + \frac{h(x_n)}{l_0 + n + 1} + \frac{\xi_{n+1}}{l_0 + n + 1},$$

where $\xi_{n+1} = \epsilon_{n+1} - \mathbb{E}[\epsilon_{n+1} \mid \mathcal{F}_n]$. Consider first the case where the sign of $f - 1/2$ is constant in a neighborhood of $1/2$. To fix ideas let say that $f \geq 1/2$ in $[1/2 - \epsilon, 1/2 + \epsilon]$ for some constant $\epsilon > 0$. In this case we will prove that $\mathbb{E}[\delta_\infty^-] < +\infty$. For all $n \geq 0$,

$$\mathbb{E}[\delta_n^-] \leq \sum_{k=0}^n \mathbb{P}[x_k^2 \geq \epsilon^2].$$

Therefore it suffices to prove that this last series is convergent. This will be achieved by using Equation (1) and some ideas from the proof of Proposition 4.1.3 in [D]. First, since $1/2$ is the unique fixed point of f , there exists some constant $a > 0$ such that $xh(x) \leq -ax^2/2$ for all $x \in [-1/2, 1/2]$. Moreover we can always take $a < 1/2$. Next Equation (1) gives

$$x_{n+1}^2 \leq x_n^2 \left(1 - \frac{a}{l_0 + n + 1}\right) + \frac{2x_n \xi_{n+1}}{l_0 + n + 1} + \frac{u_{n+1}}{(l_0 + n + 1)^2},$$

where u_{n+1} is bounded, i.e. there exists $C > 0$ such that a.s. $|u_{n+1}| \leq C$ for all n . Let $s_n = \prod_{k=0}^n (1 - a/(l_0 + k + 1))$. By induction we get

$$x_n^2 \leq s_n x_0^2 + s_n \sum_{k=0}^n ((l_0 + k + 1)s_k)^{-1} x_k \xi_{k+1} + C s_n \sum_{k=0}^n ((l_0 + k + 1)^2 s_k)^{-1}.$$

Since $s_k \sim k^{-a}$, this gives

$$x_n^2 \leq C' s_n + s_n \sum_{k=0}^n ((l_0 + k + 1)s_k)^{-1} x_k \xi_{k+1},$$

for some constant $C' > 0$. Define the martingale $(M_n, n \geq 0)$ by

$$M_n = \sum_{k=0}^n ((l_0 + k + 1)s_k)^{-1} x_k \xi_{k+1} \quad \forall n \geq 0.$$

Then for n large enough and any integer $\alpha > 0$, we have

$$\mathbb{P}[x_n^2 \geq \epsilon^2] \leq \mathbb{P}[s_n | M_n| \geq \epsilon^2/2] \leq \frac{(2s_n)^{2\alpha}}{\epsilon^{4\alpha}} \mathbb{E}[M_n^{2\alpha}].$$

But Burkholder-Davis-Gundy inequality (see e.g. [W] p.151) implies that for some constant c_α ,

$$\mathbb{E}[M_n^{2\alpha}] \leq c_\alpha \mathbb{E}[\langle M \rangle_n^\alpha],$$

where $\langle M \rangle_n := \sum_{k=0}^n ((l_0 + k + 1)s_k)^{-2} x_k^2 \xi_{k+1}^2$. Moreover, since $a < 1/2$,

$$\langle M \rangle_n \leq \sum_{k=0}^n ((l_0 + k + 1)s_k)^{-2} \leq C,$$

for some constant $C > 0$. Thus

$$s_n^{2\alpha} \mathbb{E}[\langle M \rangle_n^\alpha] \leq C_\alpha n^{-2a\alpha},$$

with $C_\alpha > 0$ a constant. Taking now α large enough shows that

$$\sum_{k=0}^{+\infty} \mathbb{P}[x_n^2 \geq \epsilon^2] < +\infty,$$

as we wanted. It remains to consider the case where the sign of $f - 1/2$ changes near $1/2$. In this case $f''(1/2) = 0$. Thus there exists a constant $C > 0$ such that

$$|2f(\alpha_n) - 1| \leq C|x_n|^3 \quad \forall n \geq 0,$$

from which we get

$$\delta_n^- + \delta_n^+ \leq C \sum_{k=0}^n |x_k|^3.$$

Thus it suffices to prove that $\sum_{k=0}^{+\infty} \mathbb{E}[|x_k|^3]$ is finite. But since $f'(1/2) = 0$, there exists $\epsilon > 0$ such that $2xh(x) \leq -x^2$ when $|x| \leq \epsilon$. Therefore (1) gives in fact

$$\mathbb{E}[x_{n+1}^2] \leq \mathbb{E}[x_n^2] \left(1 - \frac{1}{n}\right) + 4 \frac{\mathbb{P}[|x_n| \geq \epsilon]}{n} + \frac{C}{n^2}.$$

From the proof of the preceding case, we deduce that

$$\mathbb{E}[x_{n+1}^2] \leq \mathbb{E}[x_n^2] \left(1 - \frac{1}{n}\right) + \frac{C'}{n^2}.$$

This by induction proves that $\mathbb{E}[x_n^2] \leq C/n$ for some constant $C > 0$. Let us now consider the moments of order 4. Since $4x^3h(x) \leq -3x^4$ in $[-\epsilon, \epsilon]$ for some $\epsilon > 0$, (1) gives similarly

$$\mathbb{E}[x_{n+1}^4] \leq \mathbb{E}[x_n^4] \left(1 - \frac{3}{n}\right) + \frac{C}{n^2},$$

for some constant $C > 0$. By induction, this gives $\mathbb{E}[x_n^4] \leq C'n^{-2}$, with $C' > 0$ another constant. Then Cauchy-Schwarz inequality gives (up to constants)

$$\mathbb{E}[|x_n|^3] \leq (\mathbb{E}[x_n^2] \mathbb{E}[x_n^4])^{1/2} \leq n^{-3/2},$$

which is summable. This finishes the proof of the proposition. \square

Observe that the argument given in the proof of the above proposition, applies as well when the unique fixed point of f is different from $1/2$. Thus we proved also the

Proposition 3.2. *If f has a unique fixed point $p > 1/2$, respectively $p < 1/2$, then $\mathbb{E}[\delta_\infty^-]$, resp. $\mathbb{E}[\delta_\infty^+]$, is finite. In particular δ_n converges a.s. toward $+\infty$, resp. $-\infty$.*

Our last result concerns the a.s. non-finiteness of δ_∞ . First if $p \neq 1/2$ is a stable fixed point of f , then conditionally on $\{\alpha_n \rightarrow p\}$, δ_n/n converges toward $2p - 1$ and thus $\delta_\infty = +\infty$. The next result investigates the case $p = 1/2$.

Corollary 3.1. *If $1/2$ is a stable fixed point of f , $f'(1/2) = 0$ and $f''(1/2) > 0$ (respectively $f''(1/2) < 0$), then conditionally on $\alpha_n \rightarrow 1/2$, almost surely $\delta_\infty = +\infty$ (respectively $\delta_\infty = -\infty$).*

Proof. To fix ideas assume that $f''(1/2) > 0$. The other case is analogue. A limited development of f near $1/2$ gives

$$(2) \quad |\delta_n - f''(1/2) \sum_{k=0}^n (\alpha_k - 1/2)^2| \leq C \sum_{k=0}^n |\alpha_k - 1/2|^3 \quad \forall n \geq 0,$$

with $C > 0$ some positive constant. For $n \geq 0$, we set $z_n := \sqrt{n}(\alpha_n - 1/2)$. We already saw in Theorem 3.2 that z_n converges in law toward a normal variable. In fact this holds in the sense of the trajectory. More precisely, an elementary calculus shows that $(z_n, n \geq 0)$ is solution of a stochastic algorithm of the form:

$$z_{n+1} = z_n - \frac{z_n/2 - r_{n+1}}{l_0 + n + 1} + \frac{\xi_{n+1}}{\sqrt{l_0 + n + 1}},$$

where $r_{n+1} = \mathcal{O}(\sqrt{n}(\alpha_n - 1/2)^2 + n^{-1})$. For $t \in [\log n, \log(n+1)]$, let

$$Y_t = z_{n+1} + (t - \log n)(-z_{n+1}/2 + r_{n+1}) + (t - \log n)^{1/2} \xi_{n+2}.$$

For $u \geq 0$, call $(Y_t^{(u)}, t \geq 0)$ the continuous time process defined by $Y_t^{(u)} = Y_{u+t}$ for $t \geq 0$. Then Theorem 4.II.4 in [D] says that the processes $(Y_t^{(u)}, t \geq 0)$ converge in law in the path space toward an Ornstein-Uhlenbeck process $(U_s, s \geq 0)$, when $u \rightarrow +\infty$ (the condition on r_n in the hypothesis of the theorem is not needed here, as one can see with Theorem 4.III.5 and its proof in [D]). Now we will deduce from this result that

$$\sum_{k=1}^{\infty} \frac{z_k^2}{k} = +\infty.$$

This together with (2) will achieve the proof. If we define z_t for all $t \geq 0$ by $z_t = z_{[t]}$, then one can check that $\sum_{k=0}^{\infty} z_k^2/k$ is comparable with $\int_1^{+\infty} z_t^2/t dt$. Moreover we have roughly $z_{e^t} = Y_t$. Therefore a change of variable shows that our series is comparable with $\int_1^{\infty} Y_t^2 dt$. So if the series is finite, then $\int_n^{n+1} Y_t^2 dt \rightarrow 0$ when $n \rightarrow +\infty$. But this a.s. does not hold since $\int_n^{n+1} Y_t^2 dt$ converges in law toward $\int_0^1 U_s^2 ds$, which is a.s. non-zero. Thus

$$\mathbb{P} \left[\sum_{k=0}^{\infty} z_k^2/k < +\infty \mid \alpha_n \rightarrow 1/2 \right] \leq \mathbb{P} \left[\int_n^{n+1} Y_t^2 dt \rightarrow 0 \mid \alpha_n \rightarrow 1/2 \right] = 0.$$

This finishes the proof of the corollary. \square

4. A ZERO-ONE LAW

We will now study reinforced random walks and try to relate their asymptotic behavior with urns characteristics. Our first result is general. It is a zero-one law for the property of recurrence. We say that the random walk is recurrent if all sites are visited infinitely often. We have

Lemma 4.1. *Let R be the property of recurrence. Then $\mathbb{P}[R] \in \{0, 1\}$.*

Proof. First Borel-Cantelli lemma implies that if a site is visited infinitely often, then the same holds for all sites. So there are only three alternatives. Either the random walk is recurrent, or it tends toward $+\infty$ or toward $-\infty$. Thus, if

$$T_n = \inf\{k \geq 0 \mid X_k = n\} \quad \forall n \geq 0,$$

then $1_{\{T_n < +\infty\}}$ converges toward $1_{R \cup \{X_n \rightarrow +\infty\}}$, when $n \rightarrow +\infty$. In the same way the event $\{X_n > 0 \quad \forall n > 0\}$ is included in $\{X_n \rightarrow +\infty\}$. In fact there is a stronger relation:

Lemma 4.2. *For any $k \geq 0$, $\mathbb{P}[X_n \rightarrow +\infty] > 0$ if, and only if, $\mathbb{P}_{k, w_0}[X_n > k \quad \forall n > 0] > 0$.*

Proof. We make the proof for $k = 0$. The other cases are identical. We reproduce Zerner's proof ([Zer] Lemma 8) for sake of completeness. We just have to prove the only if part. Call τ_2 the last time the random walk visits the integer 2. If \mathcal{C} is some path of length k starting from 0 and ending in 2 on \mathbb{Z} , call $E_{\mathcal{C}}$ the event that the random walk follows the path \mathcal{C} during the first k steps. Define also $w_{\mathcal{C}}$ as the state of all urns once the walker has performed the path \mathcal{C} . If $\mathbb{P}[X_n \rightarrow +\infty] > 0$, then for some path \mathcal{C} from 0 to 2, we have

$$0 < \mathbb{P}[E_{\mathcal{C}}, X_n > 2 \quad \forall n > k] = \mathbb{P}[E_{\mathcal{C}}] \times \mathbb{P}_{2, w_{\mathcal{C}}}[X_n > 2 \quad \forall n > 0].$$

Now construct \mathcal{C}' as follows: it starts by a jump from 0 to 1 and then we add (in chronological order) all the excursions of \mathcal{C} above level 1. Then clearly

$$\mathbb{P}_{2, w_{\mathcal{C}}}[X_n > 2 \quad \forall n > 0] = \mathbb{P}_{2, w_{\mathcal{C}'}}[X_n > 2 \quad \forall n > 0].$$

Moreover, since the range of f is in $]0, 1[$, it is elementary to see that $\mathbb{P}[E_{\mathcal{C}'}] > 0$. Thus

$$\mathbb{P}[X_n > 0 \quad \forall n > 0] \geq \mathbb{P}[E_{\mathcal{C}'}] \times \mathbb{P}_{2, w_{\mathcal{C}'}}[X_n > 2 \quad \forall n > 0] > 0,$$

which proves the lemma. \square

We can finish now the proof of Lemma 4.1. The martingale convergence theorem and the Markov property imply

$$\begin{aligned} 1_{\{X_n \rightarrow +\infty\}} &= \lim_{n \rightarrow +\infty} \mathbb{P}_{0, w_0}[X_n \rightarrow +\infty \mid \mathcal{F}_{T_n}] 1_{\{T_n < +\infty\}} \\ &= \lim_{n \rightarrow +\infty} \mathbb{P}_{n, w_{T_n}}[X_n \rightarrow +\infty] 1_{\{T_n < +\infty\}} \\ &\geq \limsup_{n \rightarrow +\infty} \mathbb{P}_{n, w_{T_n}}[X_m > n \quad \forall m > 0] 1_{\{T_n < +\infty\}} \\ &= \mathbb{P}_{1, w_0}[X_n > 1 \quad \forall n > 0] 1_{R \cup \{X_n \rightarrow +\infty\}}. \end{aligned}$$

Then multiply the left and right part of this inequality by 1_R and take expectation. This gives

$$\mathbb{P}_{1, w_0}[X_n > 1 \quad \forall n > 0] \mathbb{P}[R] = 0.$$

In the same way we have

$$\mathbb{P}_{-1, w_0}[X_n < -1 \quad \forall n > 0] \mathbb{P}[R] = 0.$$

These two equalities and Lemma 4.2 prove the lemma. \square

Remark 4.1. *Let T_0 be the first return time to 0. Then the usual equivalence $T_0 < +\infty$ a.s. if and only if 0 is a.s. visited infinitely often, is true. Indeed if $T_0 < +\infty$ a.s. then by Lemma 4.2, a.s. X_n does not converges toward $\pm\infty$. Then the 0 – 1 law says that R holds.*

5. THE CASE WITH ONLY NON-NEGATIVE DRIFT

Here we assume that $f \geq 1/2$. In this case we have a more precise zero-one law.

Lemma 5.1. *We have the alternative: either $(X_n, n \geq 0)$ is almost surely transient toward $+\infty$, or it is almost surely recurrent.*

Proof. Since $f \geq 1/2$, at each step the random walk has probability at least $1/2$ to jump to the right. Thus an elementary coupling argument (with the usual simple random walk on \mathbb{Z}) shows that a.s. the random walk does not converge toward $-\infty$. We conclude with Lemma 4.1. \square

For all $x \in \mathbb{Z}$, we set

$$\delta_n^x = \sum_{k=0}^n 2f(\alpha_k^x) - 1,$$

where $(\alpha_n^x, n \geq 0)$ is the urn process in x . The following theorem is an analogue of Zerner's criterion [Zer] for cookie random walks:

Theorem 5.1. *The random walk $(X_n, n \geq 0)$ is recurrent if, and only if, $\mathbb{E}[\delta_\infty^1] \leq 1$. In particular if $1/2$ is not the unique stable fixed point of f , or if $f''(1/2) > 0$, then $(X_n, n \geq 0)$ is transient.*

Proof. We follow essentially the proof of Zerner. Let us recall the main lines. First we introduce some notation. For $n \geq 0$, let

$$U_n = \sum_{k \leq n-1} 1_{\{X_k=0, X_{k+1}=-1\}},$$

and let

$$X_n^+ = \sum_{k \leq n-1} (X_{k+1} - X_k) 1_{\{X_k \geq 0\}}.$$

A straightforward computation gives the equation

$$X_n^+ = \max(X_n, 0) - U_n \quad \forall n.$$

We define the drift D_n^x accumulated in x up to time n by

$$D_n^x = \sum_{k=0}^{L_n^x-1} (2f(\alpha_k^x) - 1),$$

and the drift D_n^+ accumulated in the non-negative integers by

$$D_n^+ = \sum_{x \geq 0} D_n^x.$$

Let $(M_n^+, n \geq 0)$ be the process defined by

$$M_n^+ = X_n^+ - D_n^+ \quad \forall n.$$

It is a basic fact that $(M_n^+, n \geq 0)$ is a martingale. In particular for all $a \geq 0$ and all $n \geq 0$,

$$\mathbb{E}[\max(X_{T_a \wedge n}, 0)] = \mathbb{E}[U_{T_a \wedge n}] + \mathbb{E}[D_{T_a \wedge n}^+],$$

where

$$T_a = \inf\{k \geq 0 \mid X_k = a\}.$$

Now Lemma 5.1 implies that T_a is finite a.s. Moreover $(U_n, n \geq 0)$ and $(D_n^+, n \geq 0)$ are non-decreasing processes. Thus letting n go to $+\infty$ gives with the monotone convergence theorem

$$(3) \quad a = \mathbb{E}[U_{T_a}] + \mathbb{E}[D_{T_a}^+] \quad \forall a \geq 0.$$

Moreover the Markov property shows that for any $x \in [1, a]$,

$$\mathbb{E}_{0, w_0}[D_{T_a}^x] = \mathbb{E}[\mathbb{E}_{x, w_{T_x}}[D_{T_a}^x]] = E_{1, w_0}[D_{T_a - x + 1}^1].$$

Moreover $\mathbb{E}[D_{T_a}^0]$ and $E_{1, w_0}[D_{T_a - 1}^1]$ differ at most by $\mathbb{E}[N]$, where N is the number of visits to 0 before the first visit to 1. Since the probability to jump from 0 to 1 is bounded away from 0, $\mathbb{E}[N]$ is finite. Therefore

$$\lim_{a \rightarrow +\infty} \frac{1}{a} \mathbb{E}[D_{T_a}^+] = \mathbb{E}_{1, w_0}[D_\infty^1].$$

Then (3) gives the inequality

$$\mathbb{E}[D_\infty^1] \leq 1.$$

So if the random walk is recurrent, almost surely $D_\infty^1 = \delta_\infty^1$, and $\mathbb{E}[\delta_\infty^1] \leq 1$. This gives already the only if part of the theorem. Assume now that the random walk is transient. Then Lemma 4.2 shows that

$$\mathbb{E}[\delta_\infty^1 - D_\infty^1] \geq c \mathbb{E}[\delta_\infty^1 - \delta_0^1],$$

where $c = \mathbb{P}_{1, w_0}[X_n > 1 \quad \forall n > 0] > 0$. Now since the sequence $(\delta_n^1)_{n \geq 0}$ is non-decreasing, if $\mathbb{E}[\delta_\infty^1 - \delta_0^1]$ was equal to 0, this would mean that a.s. $\delta_n^1 = \delta_0^1$ for all n . In other words our walk would evolve like the simple random walk, which is recurrent. This is absurd. Thus $\mathbb{E}[\delta_\infty^1] > \mathbb{E}[D_\infty^1]$. It remains to prove that $\mathbb{E}[D_\infty^1] = 1$. From (3) we see that this is equivalent to prove the

Lemma 5.2. *If the random walk is a.s. transient, then $\lim_{a \rightarrow +\infty} \mathbb{E}[U_{T_a}]/a = 0$.*

Proof. This lemma can be proved by following the argument of Lemma 6 in [Zer], that we reproduce here. For $i \geq 1$, let

$$\sigma_i = \inf\{j \geq T_i \mid X_j = 0\}.$$

We have $\mathbb{E}[U_{T_a}] = \sum_{i=0}^{a-1} \mathbb{E}[U_{T_{i+1}} - U_{T_i}]$. Next $U_{T_{i+1}} - U_{T_i} \neq 0$ only on the set $A_i := \{\sigma_i < T_{i+1}\}$. Moreover (3) holds for any starting environment. Thus

$$\mathbb{E}[U_{T_{i+1}} - U_{T_i}] \leq \mathbb{E}[1_{A_i} \mathbb{E}_{0, w_{\sigma_i}}[U_{T_{i+1}}]] \leq (i+1) \mathbb{P}[A_i],$$

for all i . It remains to prove that

$$(4) \quad \frac{1}{a} \sum_{i=1}^a i \mathbb{P}[A_i] \rightarrow 0,$$

when $a \rightarrow +\infty$. Let $Y_i = \mathbb{P}[A_i \mid \mathcal{F}_{T_i}]$. Since the random walk is transient, the conditional Borel-Cantelli lemma implies¹

$$(5) \quad \sum_{i \geq 1} Y_i < +\infty \quad \text{a.s.}$$

Moreover a coupling argument with the simple random walk and standard results for this random walk show that a.s., $Y_i \leq 1/i$ for all i . Let $\epsilon > 0$. For all i ,

$$\mathbb{P}[A_i] = \mathbb{E}[Y_i 1_{\{Y_i < \epsilon/i\}}] + \mathbb{E}[Y_i 1_{\{Y_i \geq \epsilon/i\}}] \leq \frac{\epsilon}{i} + \frac{\mathbb{P}[Y_i \geq \epsilon/i]}{i}.$$

So we can divide the sum in (4) in two parts. One is lower than ϵ and the other one is equal to

$$\frac{1}{a} \mathbb{E}\left[\sum_{i=1}^a 1_{\{Y_i \geq \epsilon/i\}}\right].$$

But since (5) holds, a.s. the density of the $i \leq a$ such that $Y_i \geq \epsilon/i$ tends to 0 when a tends to $+\infty$. Thus the preceding sum converges to 0. This concludes the proof of the lemma. \square

The last assertion of Theorem 5.1 is a consequence of Corollary 3.1. \square

In section 7 we will see different examples of functions f which show in particular that, in the case where $1/2$ is the only stable fixed point and $f''(1/2) = 0$, both regimes (recurrence and transience) may appear.

6. THE CASE OF A UNIQUE FIXED POINT

Here we do not assume anymore that $f \geq 1/2$. But we assume that f has a unique fixed point. Then our main result is the

Theorem 6.1. *Assume that f has a unique fixed point p . If $p \neq 1/2$ then $\mathbb{P}[R] = 0$. If $p = 1/2$ and $f'(1/2) = 0$, then $\mathbb{E}[\delta_\infty^1] > 1$ or $\mathbb{E}[\delta_\infty^{-1}] < -1$ imply $\mathbb{P}[R] = 0$. In particular if $p = 1/2$, $f'(1/2) = 0$ and $f''(1/2) \neq 0$, then $\mathbb{P}[R] = 0$.*

The sufficient condition to get $\mathbb{P}[R] = 0$ in the case $p = 1/2$ has to be compared to the result of [KZer] in the context of cookie random walks, where they prove that it is also a necessary condition. Here we were not able to prove it.

Proof. The idea of the proof is the same than for Theorem 5.1. However a priori we have to be careful when taking limits since the drift $(D_n^+)_{n \geq 0}$ is not anymore a non-decreasing function. But for any $x \geq 0$, Proposition 3.1 and Proposition 3.2 show that $\mathbb{E}[D_n^x]$ converges toward $\mathbb{E}[D_\infty^x]$. In fact since $\mathbb{E}[\delta_\infty^-] < +\infty$ or $\mathbb{E}[\delta_\infty^+] < +\infty$, if we replace n by any increasing sequence of stopping times τ_n converging toward τ_∞ , these propositions show that $\mathbb{E}[D_{\tau_n}^x]$ converges toward $\mathbb{E}[D_{\tau_\infty}^x]$. So in fact we get

$$\lim_{n \rightarrow +\infty} \mathbb{E}[D_{T_a \wedge n}^+] = \mathbb{E}[D_{T_a}^+].$$

¹since we were not able to find a reference we give here a short proof: let $(H_n, n \geq 0)$ be the $(\mathcal{F}_{T_n})_{n \geq 0}$ martingale defined by $H_n := \sum_{i=1}^n 1_{A_i} - Y_i$, for $n \geq 0$. Let $l \geq 1$ and let $T_l' = \inf\{k \mid H_k \geq l\}$. Then $H_{n \wedge T_l'}$ a.s. converges toward some limiting value $\alpha_l \in \mathbb{R}$, when $n \rightarrow +\infty$. If a.s. only a finite number of A_i 's occur, then a.s. T_l' is infinite for some $l \geq 1$. This implies the desired result.

Next observe that for all $a \geq 0$ and $n \geq 0$, $X_{T_a \wedge n}^+ \leq a$. Thus, using that $(M_n^+)_{n \geq 0}$ is a martingale, we have

$$\mathbb{E}[D_{T_a}^+] \leq a \quad \forall a \geq 0.$$

Now the Markov property implies here that if $1 \leq x \leq a$,

$$(6) \quad \mathbb{E}[D_{T_a}^x] = \mathbb{P}[T_x < +\infty] \times \mathbb{E}_{1, w_0}[D_{T_{a-x+1}}^1].$$

But $\mathbb{P}[T_x < +\infty]$ converges, when x tends to $+\infty$, toward $1 - \mathbb{P}[X_n \rightarrow -\infty]$. Now if $\mathbb{P}[X_n \rightarrow -\infty] = 1$, we are finished. If not, letting a tend to $+\infty$ in (6) gives

$$\mathbb{E}_{1, w_0}[D_\infty^1] \leq \frac{1}{1 - \mathbb{P}[X_n \rightarrow -\infty]}.$$

So if $\mathbb{P}[R] = 1$, then a.s. $D_\infty^1 = \delta_\infty^1$ and $\mathbb{E}[\delta_\infty^1] \leq 1$. The other inequality $\mathbb{E}[\delta_\infty^{-1}] \geq -1$ is similar. The different claims of the theorem now follow from Corollary 3.1. \square

Let us state now the following standard monotonicity argument.

Lemma 6.1. *Let $f \leq g$ be two functions. Then there exists a coupling of two urn processes $((\alpha_n, l_n), n \geq 0)$ and $((\beta_n, l'_n), n \geq 0)$ associated respectively to f and g , such that $l_0 = l'_0$, $\alpha_0 = \beta_0$, and almost surely $\alpha_n \leq \beta_n$ for all $n \geq 0$.*

Proof. The proof is standard. Let $(U_i, i \geq 0)$ be a sequence of i.i.d. random variables uniformly distributed on $[0, 1]$. We define two urn processes starting with initial conditions like in the lemma. Then at step n , $\alpha_{n+1} > \alpha_n$ if, and only if $f(\alpha_n) \geq U_n$. The same for β_{n+1} (with g in place of f). Assume now that for some n , $\alpha_n > \beta_n$. Assume also that n is the lowest index where such inequality occurs. This means that $\alpha_{n-1} = \beta_{n-1}$. But since $f \leq g$, by definition of our processes, we get an absurdity. \square

This lemma together with Theorem 6.1 allows to consider also the case where f has possibly more than one fixed point but under the condition $f \geq 1/2$ on $[1/2, 1]$. More precisely we have

Corollary 6.1. *Assume that $f \geq 1/2$ on $[1/2, 1]$ and that all fixed points of f are greater or equal to $1/2$. If $1/2$ is not a fixed point, then $\mathbb{P}[R] = 0$. If $1/2$ is a fixed point, but not the only fixed point, and $f'(1/2) = 0$, then $\mathbb{P}[R] = 0$.*

Proof. If any of the two hypothesis of the corollary is satisfied, then there exists a function g such that $g \leq f$, g has a unique fixed point equal to $1/2$, and $g'(1/2) = 0$. We can also assume that g is increasing on $[0, 1/2]$. Applying Lemma 6.1 we see that there exists an urn process $(\beta_n, n \geq 0)$ associated to g such that $\beta_n \leq \alpha_n$ for all n . Now the proof of Proposition 3.1 shows that

$$\sum_{n \geq 0} \mathbb{E}[(2g(\beta_n) - 1)_-] < +\infty.$$

Since g is increasing on $[0, 1/2]$ and $f \geq 1/2$ on $[1/2, 1]$, this implies that $\mathbb{E}[\delta_\infty^-]$ is finite. Moreover we know that $\delta_\infty = +\infty$ a.s. So we have everything to apply the proof of Theorem 6.1 and conclude. \square

7. SOME EXAMPLES

Our goal here is to give examples of functions f leading to interesting behavior for the random walk, in view of the previous results. In all this section we consider f , symmetric with respect to $1/2$, i.e. such that $f(1/2 - x) = f(1/2 + x)$ for all $x \in [0, 1/2]$, strictly decreasing on $[0, 1/2]$ and strictly increasing on $[1/2, 1]$. We assume also that f has a unique fixed point, equal to $1/2$, and that $f''(1/2) = 0$.

We start now by a comparison result. Let $u > 0$ be some positive real number. Define f_u by the equation $2f_u - 1 = u(2f - 1) \wedge 1$. One can see immediately that f_u has the same properties than f for all u , and moreover that $f_u \leq f_v$ if $u \leq v$. Call $(\alpha_n^u, n \geq 0)$ an urn process associated to f_u such that $\alpha_0^u = 1/2$, and set $\delta_\infty^u := \sum_{n \geq 0} 2f(\alpha_n^u) - 1$. Then

Lemma 7.1. *For all u , $\mathbb{E}[\delta_\infty^u] < +\infty$. The maps $u \mapsto \mathbb{E}[\delta_\infty^u]/u$ and $u \mapsto \mathbb{E}[\delta_\infty^u]$, are increasing respectively on $]0, 1]$ and $[0, +\infty[$. Moreover $\mathbb{E}[\delta_\infty^u] \rightarrow +\infty$, when $u \rightarrow +\infty$.*

Proof. The first claim results from the proof of Proposition 3.1. For the second part, consider first $0 < u < v \leq 1$. By symmetry, for any $k \geq 0$,

$$\mathbb{E}[2f_u(\alpha_k^u) - 1] = 2\mathbb{E}[(2f_u(\alpha_k^u) - 1)1_{\{\alpha_k^u \geq 1/2\}}].$$

Moreover, since f_v is increasing on $[1/2, 1]$ and since one may assume that $\alpha_k^u \leq \alpha_k^v$ a.s. by Lemma 6.1,

$$\begin{aligned} \mathbb{E}[(2f_u(\alpha_k^u) - 1)1_{\{\alpha_k^u \geq 1/2\}}] &= \frac{u}{v}\mathbb{E}[(2f_v(\alpha_k^u) - 1)1_{\{\alpha_k^u \geq 1/2\}}] \\ &\leq \frac{u}{v}\mathbb{E}[(2f_v(\alpha_k^v) - 1)1_{\{\alpha_k^v \geq 1/2\}}] \end{aligned}$$

The result follows by summation. The fact that $u \mapsto \mathbb{E}[\delta_\infty^u]$ is increasing on $[0, +\infty[$ is similar. It remains to find the limit when $u \rightarrow +\infty$. For this, fix some $n \geq 1$. Then one can observe that there exists $\epsilon > 0$, such that $|\alpha_{2k+1}^u - 1/2| \geq \epsilon$ for all $k \leq n$. This implies that for u large enough, $\mathbb{E}[\delta_\infty^u] \geq n/2$. Since this holds for all n , the result follows. \square

The preceding lemma and Theorem 5.1 show that there is a phase transition: there exists some $u_0 > 0$ such that for $u > u_0$, the random walk associated to f_u is transient, whereas for $u < u_0$ it is recurrent. In particular recurrence and transience may both appear. The question of what happens at u_0 is related to the continuity of $\mathbb{E}[\delta_\infty]$ with respect to f . But general results for Markov chains show that if $u \rightarrow u_0$, then for all n , $\mathbb{E}[\delta_n^u] \rightarrow \mathbb{E}[\delta_n^{u_0}]$. Together with the monotonicity of $\mathbb{E}[\delta_\infty^u]$ in u , this proves that $\mathbb{E}[\delta_\infty^u]$ is continuous in u . In particular for $u = u_0$ the random walk is recurrent.

Our second problem concerns what happens when initial conditions vary. Here also we will see that there is possibly a phase transition. For $(\alpha, l) \in [0, 1] \times]0, +\infty[$, we denote by $\mathbb{E}_{\alpha, l}$ the law of an urn process starting from (α, l) . First let us prove the

Lemma 7.2. *Let $\alpha \in [0, 1] \setminus \{1/2\}$ and $l \in]0, +\infty[$, be such that $2\alpha l - l \in \mathbb{N}$. Then $\mathbb{E}_{\alpha, 2l}[\delta_\infty] > \mathbb{E}_{1/2, 2l}[\delta_\infty]$.*

Proof. We use a standard coupling argument. Let $(U_i)_{i \geq 0}$ be a family of i.i.d. random variables, uniformly distributed in $[0, 1]$. Let start two urn processes $(\alpha_n, n \geq 0)$ and $(\beta_n, n \geq 0)$, respectively from $(1/2, 2l)$ and $(\alpha, 2l)$. They evolve

according to the following rule. If at step n , α_n or β_n is equal to $x \geq 1/2$, then we add one Red ball in the corresponding urn if $U_n \leq f(x)$. Now if $x < 1/2$, then we add a Red ball if $U_n \geq 1 - f(x)$. The condition $2\alpha l - l \in \mathbb{N}$ assures that the two urn processes (as well as their symmetric with respect to $1/2$) cannot cross each other without meeting them. Thus for all $n \geq 0$, $|\beta_n - 1/2| \geq |\alpha_n - 1/2|$. The lemma follows. \square

The preceding results show in particular that the property of recurrence or transience may depend on the initial condition of the urns (even if l_0 is fixed). Indeed it suffices to consider f such that $\mathbb{E}_{1/2, 2l_0}[\delta_\infty] = 1$, which is possible by Lemma 7.1 and the continuity of $\mathbb{E}[\delta_\infty^u]$ as explained above. Then the preceding lemma shows that for any $\alpha \neq 1/2$ satisfying the condition of the lemma, the random walk associated with urns starting from $(\alpha, 2l_0)$ is always transient, whereas it is recurrent if they start from $(1/2, 2l_0)$.

We arrive now to our last result.

Lemma 7.3. *The map $l \mapsto \mathbb{E}_{1/2, 2l}[\delta_\infty]$ is continuous on $]0, +\infty[$, decreasing, and converges toward 0 when $l \rightarrow +\infty$.*

Proof. The continuity of the map is similar to the continuity of $\mathbb{E}[\delta_\infty^u]$ in u , observed above. The fact that the map is decreasing can be seen by using a coupling argument like in the preceding lemma. Indeed let $l_0 < l_1$. Let $(\alpha_n, n \geq 0)$ and $(\beta_n, n \geq 0)$ be two urn processes starting respectively from $(1/2, 2l_1)$ and $(1/2, 2l_0)$. Define their joint law like in the previous lemma. Observe that each urn process cannot jump above $1/2$ without touching it. In the same way, if for some n , $1/2 \leq \beta_{n+1} < \beta_n$ and $1/2 \leq \alpha_{n+1} < \alpha_n < \beta_n$, then

$$\beta_n - \beta_{n+1} = \frac{\beta_n}{2l_0 + n + 1},$$

and

$$\alpha_n - \alpha_{n+1} = \frac{\alpha_{n+1}}{2l_1 + n}.$$

Thus

$$\beta_n - \beta_{n+1} \leq \beta_n - \alpha_{n+1}.$$

In other words the two urn processes cannot cross them without meeting them. Thus for all $n \geq 0$, $|\beta_n - 1/2| \geq |\alpha_n - 1/2|$, which proves the desired result. It remains to find the limit when $l \rightarrow +\infty$. But for each n , $\mathbb{E}_{1/2, 2l}[\delta_n]$ converges to 0 when $l \rightarrow +\infty$. Since moreover for fixed l it converges toward $\mathbb{E}_{1/2, 2l}[\delta_\infty]$, when $n \rightarrow +\infty$, the result follows. This finishes the proof of the lemma. \square

This lemma shows that for some functions f , there is a phase transition when l grows. Indeed if f and l_0 are such that $\mathbb{E}_{1/2, 2l_0}[\delta_\infty] > 1$, then there exists some l_1 such that for $l < l_1$, the random walk associated to urns starting from $(1/2, 2l)$ is transient, whereas when these urns start from $(1/2, 2l)$ with $l \geq l_1$, the random walk is recurrent. Observe that this phase transition is similar to the phenomenon observed by Pemantle on trees [Pem1]. Indeed in the non-oriented reinforced framework discussed in the introduction, starting with small l is equivalent to have a large reinforcement coefficient Δ .

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