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▶ To cite this version:

Mohamed Darouach, Michel Zasadzinski, Driss Mehdi. State Estimation of stochastic singular linear systems. International Journal of Systems Science, 1993, 24 (2), pp.345-354. 10.1080/00207729308949493 . hal-00143947

HAL Id: hal-00143947 https://hal.science/hal-00143947

Submitted on 28 Apr 2007

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State Estimation of stochastic singular linear systems

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Abstract : In this paper we present a simple algorithm for the state estimation of stochastic singular linear systems based on the least squares method.

1. Introduction

The analysis and design of linear singular systems have received great attention in the last few years as can be seen in the survey of Lewis (1986), and also in Verghese et al. (1981) and Cobb (1984). The control of these systems requires, as in the standard case, the knowledge of the state vector. In the deterministic case one can use the observer theory to estimate the state vector (El-Tohami et al. (1987), Fahmy and O'Reilly (1989)). Unfortunately, in the stochastic case there are few works which treat the problem of control (Bender and Laub (1987 a, b)) and estimation (Dai (1989a)). The state estimation problem is considered under the assumption of regularity [det(sE - A) \neq 0] and causality where matrices E and A are square and singular.

In a recent paper (Darouach and Zasadzinski (1990)), an extension to the deterministic system with uncertain measurements, where matrices E and A are rectangular, was presented. In this paper we shall consider a generalization to the stochastic systems case.

The paper is organized as follows. First the concept of estimability is introduced and the uniqueness conditions of problem solution are given in section 2. Then in section 3 an algorithm for the state estimation is derived from the least squares method. A numerical example is used to illustrate this algorithm. Section 4 contains conclusion and remarks.

2. Formulation of the problem

Let us consider the discrete singular linear stochastic system described by :

$$E x_{k+1} = A x_k + w_k \tag{1}$$

$$\mathbf{z}_{\mathbf{k}} = \mathbf{H} \, \mathbf{x}_{\mathbf{k}} + \mathbf{v}_{\mathbf{k}} \tag{2}$$

where x_k is the n-dimensional state vector and z_k is the m-dimensional output vector. E and A are $p_x n$ constant matrices and H is an $m_x n$ constant matrix. w_k and v_k are $p_x 1$ and $m_x 1$ vectors of zero mean white sequences which covariance matrices are given by :

$$\begin{split} & E\{w_k w_i^T\} = \begin{cases} W > 0 \text{ if } i = k\\ 0 & \text{ if } i \neq k \end{cases} \\ & E\{v_k v_i^T\} = \begin{cases} V > 0 \text{ if } i = k\\ 0 & \text{ if } i \neq k \end{cases} \\ & E\{w_k v_i^T\} = 0 \text{ for all } k \text{ and } i. \end{split}$$

When E and A are square matrices (p = n), it is well known (Dai (1989b)) that the knowledge of the structure, especially the observability and controllability properties, is important in estimation and in optimal control. In this contribution, we introduce the notion of estimability in the general case where p may be different from n.

Definition :

System (1)-(2) is said to be estimable if for $w_k = v_k = 0$ and for some N > 0, the knowledge of output z_k where $k \in [0, N]$ and model equation (1) is sufficient to determine uniquely x_k ($k \in [0, N]$).

System (1)-(2) can be written as follows :

$$\Phi_{N} \begin{bmatrix} x_{0} \\ \vdots \\ \vdots \\ x_{N} \end{bmatrix} = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ z_{0} \\ \vdots \\ z_{N} \end{bmatrix}$$

r

where

$$\Phi_{\rm N} = \begin{bmatrix} -A & E & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \dots & -A & E \\ H & 0 & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \dots & 0 & H \end{bmatrix}$$

(3)

System (1)-(2) is estimable if and only if Φ_N has a full column rank.

Our aim is to find the estimate of x_k , denoted $\hat{x}_{k/k}$, based on the observations z over the time interval [0, k].

If the initial state x_0 is assumed to be gaussian with known mean \overline{x}_0 , and variance $P_0 > 0$, independent of v_k and w_k , the problem in the least squares sense can be formulated :

$$\min_{\hat{\mathbf{X}}_{i/k}} \mathbf{J}_{k} = \frac{1}{2} \left(\left\| \hat{\mathbf{x}}_{0/k} - \overline{\mathbf{x}}_{0} \right\|_{\mathbf{P}_{0}^{-1}}^{2} + \sum_{i=0}^{k-1} \left\| \mathbf{z}_{i+1} - \mathbf{H} \hat{\mathbf{x}}_{i+1/k} \right\|_{\mathbf{V}^{-1}}^{2} + \left\| \mathbf{E} \hat{\mathbf{x}}_{i+1/k} - \mathbf{A} \hat{\mathbf{x}}_{i/k} \right\|_{\mathbf{W}^{-1}}^{2} \right) (4)$$

The following theorem gives the conditions of the estimability and the uniqueness of the solution of the problem (4).

Theorem 1 :

System (1)-(2) is estimable, given the initial state x_0 , and problem (4) has a unique solution if and only if matrices $\begin{bmatrix} sE - A \\ H \end{bmatrix}$ and $\begin{bmatrix} E \\ H \end{bmatrix}$, where s is a complex variable, are of full column rank.

Proof :

From equation (3), the estimability condition is given by

$$rank(\Phi_{N}) = (N+1) n \tag{5}$$

Now assume that

$$\begin{bmatrix} sE - A \\ H \end{bmatrix} \text{ or } \begin{bmatrix} E \\ H \end{bmatrix}$$
(6)

is not a full column rank matrix, this is equivalent to the existence of a column vector $q \neq 0$ such that

$$\begin{bmatrix} sE - A \\ H \end{bmatrix} q = 0$$
(7)

o r

$$\operatorname{rank}\left(\begin{bmatrix} \mathbf{E}\\ \mathbf{H} \end{bmatrix}\right) < \mathbf{n} \tag{8}$$

where s is a complex variable and q is a finite polynomial vector in variable s given by

$$q(s) = q_0 + s q_1 + s^2 x_2 + \dots + s^k q_k$$
(9)

where k is the minimal index such that $q_k \neq 0$ (Gantmacher (1959)).

Substituting (9) into (7) gives

$$T_{k} \begin{bmatrix} q_{k} \\ q_{k-1} \\ \vdots \\ q_{0} \end{bmatrix} = 0$$
(10)

where

$$T_{k} = \begin{bmatrix} E0 \dots 0 & 0 \\ -A & E \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 \dots & -A & E \\ H & 0 \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 \dots & 0 & H \end{bmatrix}$$
(11)

is a $pmk_x(n+1k matrix)$. Equation (10) is equivalent to matrix T_k is not of full column rank. Now from (3) and (11), we can write

$$\Phi_{N} = \begin{bmatrix}
-A & 0 \\
0 & \vdots \\
\vdots & 0 \\
0 & E \\
H & T_{N-2} & 0 \\
0 & \vdots \\
\vdots & 0 \\
0 & H
\end{bmatrix}$$
(12)

and from (12) and the assumption (6), we can see that Φ_N is not a full column rank matrix, thus system (1)-(2) is not estimable, which proves the theorem.

Theorem 2 :

System (1)-(2) is estimable, given the initial state x_0 , and problem (4) has unique solution if and only if the matrix

$$\begin{bmatrix} E \\ H \end{bmatrix}$$
(13)

is of full column rank.

Proof :

If x_0 is known, (3) becomes

$$\Phi'_{N} \begin{bmatrix} x_{1} \\ \vdots \\ x_{N} \end{bmatrix} = \begin{bmatrix} Ax_{0} \\ 0 \\ \vdots \\ 0 \\ z_{0} \\ \vdots \\ z_{N} \end{bmatrix}$$
(14)
where
$$\Phi'_{N} = \begin{bmatrix} E0 \dots 0 & 0 \\ -A & E \dots & 0 & 0 \\ -A & E \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 \dots & -A & E \\ H & 0 \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 \dots & 0 & H \end{bmatrix}$$
(15)

which can be written as

$$\Phi'_{N} = U \begin{bmatrix} E & 0 & 0 & .. & 0 \\ H & 0 & 0 & .. & 0 \\ -A & E & 0 & .. & 0 \\ 0 & -A & E & .. & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 0 & .. & 0 & -A & E \\ 0 & .. & 0 & H \end{bmatrix}$$
(16)

where U is a row permutation matrix.

From the echelon form (16) we can easily deduce that condition (13) is necessary and sufficient for the matrix Φ'_N to be of full column rank.

3. Problem solution

The solution to problem (4) is given by the following theorem.

Theorem 3: If matrices $\begin{bmatrix} sE - A \\ H \end{bmatrix}$ and $\begin{bmatrix} E \\ H \end{bmatrix}$ are of full column rank, then estimate $\hat{x}_{k/k}$ that minimizes criterion J_k is

$$\hat{\mathbf{x}}_{k/k} = \mathbf{P}_{k/k} \mathbf{E}^{\mathrm{T}} (\mathbf{W} + \mathbf{A} \mathbf{P}_{k-1/k-1} \mathbf{A}^{\mathrm{T}})^{-1} \mathbf{A} \hat{\mathbf{x}}_{k-1/k-1} + \mathbf{P}_{k/k} \mathbf{H}^{\mathrm{T}} \mathbf{V}^{-1} \mathbf{z}_{k} (17)$$

where

$$P_{k/k} = (E^T (W + A P_{k-1/k-1} A^T)^{-1} E + H^T V^{-1} H)^{-1}$$
(18)

with $P_{0/0} = P_0$ and $\hat{x}_{0/0} = \overline{x}_0$.

Proof :

Differentiating the cost function (4) with respect to $\hat{x}_{i/k}$ and equating it to zero, yields a two-point boundary-value problem :

$$D_{1} \hat{x}_{0/k} = P_{0} \overline{x}_{0}$$

$$- C \hat{x}_{i-1/k} + D \hat{x}_{i/k} - C^{T} \hat{x}_{i+1/k} = H^{T} V^{-1} z_{i}$$

$$- C \hat{x}_{k-1/k} + D_{f} \hat{x}_{k/k} = H^{T} V^{-1} z_{k}$$
(19)

where

and

$$\mathbf{C} = \mathbf{E}^{\mathrm{T}} \, \mathbf{W}^{-1} \, \mathbf{A} \tag{20}$$

From equations (19), after a few manipulations, we obtain the following recursive equation for $\hat{x}_{k/k}$:

$$\hat{\mathbf{x}}_{k/k} = (\mathbf{D}_{f} - \mathbf{C}\mathbf{K}_{k-1}\mathbf{C}^{T})^{-1}\mathbf{C}\mathbf{K}_{k-1}(\mathbf{D}_{f} - \mathbf{C}\mathbf{K}_{k-2}\mathbf{C}^{T})\hat{\mathbf{x}}_{k-1/k-1} + (\mathbf{D}_{f} - \mathbf{C}\mathbf{K}_{k-1}\mathbf{C}^{T})^{-1}\mathbf{H}^{T}\mathbf{V}^{-1}\mathbf{z}_{k}$$
(21)

where

$$K_{k} = (D - C K_{k-1} C^{T})^{-1}$$
(22)

and

$$K_0^{-1} = P_0^{-1} + A^T W^{-1} A$$
 (23)

If we define the estimation error $\boldsymbol{\epsilon}_k$ by :

$$\boldsymbol{\varepsilon}_{k} = \mathbf{x}_{k} - \hat{\mathbf{x}}_{k/k}$$

its covariance matrix $P_{k/k}$ is :

$$P_{k/k} = E(\varepsilon_k \ \varepsilon_k^T) \tag{24}$$

From equations (21), (23) and definition (24), we obtain after a few manipulations :

$$P_{1/1} = (D_f - C K_0 C^T)^{-1} = (K_1^{-1} - A^T W^{-1} A)^{-1}$$

 $K_1^{-1} = P_{1/1}^{-1} + A^T W^{-1} A$

and

$$K_1^{-1} = P_{1/1}^{-1} + A^T W^{-1} A$$

Now if we suppose that :

$$K_{k}^{-1} = P_{k/k}^{-1} + A^{T} W^{-1} A$$
(25)

from relations (21), (24) and (25), it can be shown that :

$$P_{k+1/k+1} = (D_f - C K_k C^T)^{-1}$$
(26)

and

$$K_{k+1}^{-1} = P_{k+1/k+1}^{-1} + A^{T} W^{-1} A$$
(27)

Substituting (22) and (26) into (21) gives (17). (18) can be obtained from (25) and (26).

Equation (18) represents a generalized Riccati difference equation.

Numerical example :

As an example, we consider the singular discrete-time system described by the following equations :

where

$$\mathbf{E} = \begin{bmatrix} 1 & 1 & 1 & 0 \\ 2 & 0 & -1 & 0 \\ 0 & 1 & 0 & 1 \end{bmatrix}, \mathbf{A} = \begin{bmatrix} 1 & 1 & 0 & 0.59 \\ 0 & -1 & 0 & 0.50 \\ 1 & 0 & 1 & 0.09 \end{bmatrix}, \mathbf{B} = \begin{bmatrix} 1 & 1 \\ 2 & 0 \\ 1 & 2 \end{bmatrix} \text{ and } \mathbf{H} = \begin{pmatrix} 1 & 0 & 0 & 1 \\ 0 & 1 & -0.5 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

The vectors x_k , z_k , w_k and v_k have the same definition as in equations (1) and (2) (n=4, m=3 and p=3). u_k is the q-dimensional input vector (q=2). The variance matrices of \overline{x}_0 , w_k and v_k are :

$$P_0 = \begin{bmatrix} 0.6 & 0 & 0 & 0 \\ 0 & 0.2 & 0 & 0 \\ 0 & 0 & 0.5 & 0 \\ 0 & 0 & 0 & 0.7 \end{bmatrix}, W = \begin{bmatrix} 0.6 & 0 & 0 \\ 0 & 0.8 & 0 \\ 0 & 0 & 0.7 \end{bmatrix} \text{ and } V = \begin{bmatrix} 0.3 & 0 & 0 \\ 0 & 0.3 & 0 \\ 0 & 0 & 0.6 \end{bmatrix}$$

The estimability conditions of theorem 1 are verified.

Inputs and outputs are plotted in figures 1 and 2. The true and estimated values of the state vector are shown in figures 3 to 6. The evolution of the norm $||P_{k/k}||$ (the largest singular value) is plotted in figure 7.

4. Conclusion

By using the notion of estimability for the general discrete-time singular systems (E $x_{k+1} = A x_k$), where matrices E and A are constant, not necessarily square, and by applying the least squares estimation method, we have established a simple algorithm for the state estimation of stochastic singular linear systems. A numerical example has been presented to illustrate the algorithm. The evolution of the norm $||P_{k/k}||$ was plotted. The

sequence $P_{k/k}$ is the solution of a generalized Riccati equation. The convergence conditions of this sequence are under study.

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figure 2 : outputs (---: z_1 , ----: z_2 , ----: z_3)



figure 3 : true and estimated values of x_1 (——: true values x_1 , ----- : $\hat{x}_{1,k/k}$)



figure 4 : true and estimated values of x_2 (---: true values x_2 , ---- : $\hat{x}_{2,k/k}$)



figure 5 : true and estimated values of x_3 (---: true values x_3 , ---- : $\hat{x}_{3,k/k}$)



figure 6 : true and estimated values of x_4 (---: true values x_4 , ---- : $\hat{x}_{4,k/k}$)



figure 7 : evolution of the norm $||\boldsymbol{\Sigma}_{k\,k}^k||$